# UNIVERSITY OF TORONTO SCARBOROUGH

# Department of Computer & Mathematical Sciences

### STAD70H3 Statistics & Finance II

# April 2021 Final Examination

**Duration:** 2 hours

Aids allowed: Open book/notes, scientific calculator

**Instructor:** Sotirios Damouras

Last Name: _	
Student #: _	

#### **Instructions:**

- Read the questions carefully and answer only what is being asked.
- $\bullet\,$  Show your intermediate work, and write clearly and legibly.

Question:	1	2	3	4	Total
Points:	25	25	25	25	100
Score:					

1. (Risk Management & Betting Strategies)

Assume you start with an initial wealth of  $V_0$  and you can invest any amount on a sequence of investment opportunities with net returns  $\{R_i\}_{i\geq 1}$  which follow independent and identically distributed Uniform(-1,2) distributions. For example, if you invest \$1 in such an opportunity, your resulting wealth will be  $(1+R) \in (0,3)$ , i.e. anywhere between losing to tripling your wealth.

- (a) [2 points] Assume you bet your entire initial capital of  $V_0 = \$1$  on the first opportunity. Calculate the Value-at-Risk (VaR) of the investment at the  $\alpha = 5\%$  level.
- (b) [3 points] Calculate the Conditional VaR (CVaR) at the  $\alpha = 5\%$  level for the investment in part (a).
- (c) [5 points] Calculate the *Entropic* VaR (EVaR) at the  $\alpha = 5\%$  level for the investment in part (a). This is defined as  $\text{EVaR}_{\alpha} = \inf_{z>0} \{z^{-1} \ln(M_L(z)/\alpha)\}$ , where  $M_L(z)$  is the moment generating function of the loss (L). You are given that the infimum occurs at  $z \approx 18.122$ .
- (d) [5 points] Assume your strategy is to invest a fixed fraction (f) of your current wealth in each step. Show that your final wealth after n investments is

$$V_n = V_0 \prod_{i=1}^{n} (1 + fR_i)$$

- (e) [2 points] Assume you invest your entire wealth (f = 1) at each step. Find the probability of going bankrupt  $(\mathbb{P}(V_n = 0))$  after n steps?
- (f) [3 points] Assume you invest your entire wealth (f = 1) at each step. Find the expected value of your wealth  $(\mathbb{E}[V_n])$  after n steps.
- (g) [5 points] For general  $f \in (0,1)$ , show that the expected log-wealth after n investments is

$$\mathbb{E}[\ln(V_n)] = \ln(V_0) + \frac{n}{3} \left[ \left( 2 + \frac{1}{f} \right) \ln(1 + 2f) - \left( \frac{1}{f} - 1 \right) \ln(1 - f) - 1 \right]$$

(Hint: Use the fact that  $\int \ln(1+fx)dx = (x+1/f)\ln(1+fx) - x + c$ .)

#### **Solution:**

(a) The loss distribution is  $L=-R\sim \mathrm{Uniform}(-2,1),$  whose 5% quantile is

$$VaR_{.05} = 1 - .05 \times 3 = 0.85$$

(b) For the Uniform, the (conditional) expected value is the the midpoint of the range:

$$CVaR_{.05} = \frac{1 + 0.85}{2} = 9.25$$

(c) The MGF of the uniform distribution for the loss (L) is

$$M_L(z) = \mathbb{E}[e^{zL}] = \int_{-2}^1 e^{z\ell}/3d\ell$$
  
=  $\frac{1}{3z} \left[e^{z\ell}\right]_{\ell=-2}^1 = \frac{e^z - e^{-2z}}{3z}$ 

The  $\text{EVaR}_{\alpha}$  is

$$\begin{aligned}
\text{EVaR}_{\alpha} &= \ln \left( \frac{e^z - e^{-2z}}{3z(0.05)} \right) / z \Big|_{z=18.122} \\
&= \ln \left( \frac{e^{18.122} - e^{-2(18.122)}}{0.15(18.122)} \right) / 18.122 = 0.9448181
\end{aligned}$$

Note that:  $VaR_{.05} \le CVaR_{.05} \le EVaR_{.05}$ , as expected.

(d) Assuming you invest a fraction f, then your wealth after the first step will be:

$$V_1 = (1 - f)V_0 + fV_0(1 + R_1) = V_0 - fV_0 + fV_0 + fV_0R_1 = V_0(1 + fR_1)$$

Similarly for the next steps, we get

$$V_{n} = V_{n-1}(1 + fR_{n})$$

$$= V_{n-2}(1 + fR_{n-1})(1 + fR_{n})$$

$$= \vdots$$

$$= V_{0} \prod_{i=1}^{n} (1 + fR_{i})$$

- (e) If f = 1, you will go bankrupt if and only if any of the returns  $R_i$  is exactly equal -1. Since the return distribution is continuous, the probability of that happening is 0.
- (f) For f = 1

$$\mathbb{E}[V_n] = V_0 \mathbb{E}\left[\prod_{i=1}^n (1+R_i)\right]$$

$$= \prod_{i=1}^n \mathbb{E}[1+R_i] \quad \text{(by independence)}$$

$$= (\mathbb{E}[1+R_i])^n \qquad = (1.5)^n$$

$$\mathbb{E}[\ln(V_n)] = \mathbb{E}\left[\ln\left(V_0 \prod_{i=1}^n (1+fR_i)\right)\right]$$

$$= \mathbb{E}\left[\ln(V_0) + \sum_i i = 1^n \ln(1+fR_i)\right]$$

$$= \ln(V_0) + n \mathbb{E}[\ln(1+fR)]$$

$$= \ln(V_0) + n \int_{-1}^2 \ln(1+fx)/3dx$$

$$= \ln(V_0) + \frac{n}{3}\left[\left(x + \frac{1}{f}\right)\ln(1+fx) - x\right]_{x=-1}^2$$

$$= \ln(V_0) + \frac{n}{3}\left[\left(2 + \frac{1}{f}\right)\ln(1+f2) - 2 - \left(-1 + \frac{1}{f}\right)\ln(1-f) + 1\right]$$

$$= \ln(V_0) + \frac{n}{3}\left[\left(2 + \frac{1}{f}\right)\ln(1+2f) - \left(\frac{1}{f} - 1\right)\ln(1-f) - 1\right]$$

2. (Brownian Motion & Simulation)

Consider the standard Brownian motion  $\{W_t\}_{t\geq 0}$  and its running maximum  $M_t = \max_{0\leq s\leq t} \{W_s\}$ . For this question you will prove that the *conditional* distribution of the maximum given the final point of the Brownian motion (i.e.  $M_1|W_1$ ) follows a Rayleigh distribution, something we used in class.

(a) [7 points] Show that the joint PDF of  $M_1, W_1$  is

$$f_{W_1,M_1}(x,y) = \frac{2(2y-x)}{\sqrt{2\pi}} \exp\left\{-\frac{(2y-x)^2}{2}\right\}, \ \forall y > 0, x < y$$

(Hint: Differentiate the joint CDF  $F_{W_1,M_1}(x,y)$  w.r.t. x and y, using the fact that  $\mathbb{P}(W_1 \leq x, M_1 \leq y) = \mathbb{P}(W_1 \leq x) - \mathbb{P}(W_1 \leq x, M_1 > y)$ , where the latter probability is given by the reflection principle.)

(b) [7 points] Show that the conditional PDF of  $M_1|(W_1=x)$  is given by

$$f_{M_1|W_1}(y|x) = 2(2y-x)e^{-2y(y-x)}, \ \forall y > 0, x < y$$

(Hint: Use the fact that  $f_{M_1|W_1}(y|x) = f_{W_1,M_1}(x,y)/f_{W_1}(x)$ .)

(c) [4 points] Show that the conditional CDF of  $M_1|(W_1=x)$  is given by

$$F_{M_1|W_1}(y|x) = 1 - e^{-2y(y-x)}, \ \forall x \in \mathbb{R}, y > \min\{0, x\}$$

(d) [7 points] Let  $L_t = \min_{0 \le s \le t} \{W_t\}$  be the running minimum of the standard Brownian motion. Find a formula that simulates the conditional minimum  $L_1|(W_1 = x)$  of the standard Brownian motion given  $W_1 = x \in \mathbb{R}$ , based on a single Uniform(0,1) random number U.

(Hint: use the inverse conditional CDF of  $M_1|(W_1 = x)$  together with the symmetry of standard Brownian motion.)

#### Solution:

(a) Let  $\Phi()/\phi()$  be the standard normal CDF/PDF. The joint CDF of  $W_1, M_1$  is given by:

$$\begin{split} F_{W_1,M_1}(x,u) &= \mathbb{P}(W_1 \le x, M_1 \le y) \\ &= \mathbb{P}(W_1 \le x) - \mathbb{P}(W_1 \le x, M_1 > y) \\ &= \Phi(x) - \mathbb{P}(W_1 \ge 2y - x) \\ &= \Phi(x) - \Phi(-(2y - x)), \quad \forall y > 0, x < y \end{split}$$

The joint PDF is

$$f_{W_1,M_1}(x,u) = \frac{d^2}{dxdy} F_{W_1,M_1}(x,u)$$

$$= \frac{d^2}{dxdy} [\Phi(W_1 \le x) - \Phi(-(2y-x))]$$

$$= \frac{d}{dx} [-\phi(-(2y-x))(-2)]$$

$$= \frac{d}{dx} [2\phi(2y-x)]$$

$$= \frac{d}{dx} \left[ \frac{2}{\sqrt{2\pi}} \exp\left\{ -\frac{(2y-x)^2}{2} \right\} \right]$$

$$= \frac{2}{\sqrt{2\pi}} \exp\left\{ -\frac{(2y-x)^2}{2} \right\} \frac{d}{dx} \left( -\frac{(2y-x)^2}{2} \right)$$

$$= \frac{2(2y-x)}{\sqrt{2\pi}} \exp\left\{ -\frac{(2y-x)^2}{2} \right\}, \quad \forall y > 0, x < y$$

(b)

$$\begin{split} f_{M_1|W_1}(y|x) &= \frac{f_{W_1,M_1}(x,y)}{f_{W_1}(x)} \\ &= \frac{\frac{2(2y-x)}{\sqrt{2\pi}} \exp\left\{-\frac{(2y-x)^2}{2}\right\}}{\frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{x^2}{2}\right\}} \\ &= 2(2y-x) \exp\left\{-\frac{(2y-x)^2-x^2}{2}\right\} \\ &= 2(2y-x) \exp\left\{-\frac{4y^2-4yx+x^2-x^2}{2}\right\} \\ &= 2(2y-x) \exp\left\{-2y(y-x)\right\}, \quad \forall y > (0 \lor x) \end{split}$$

(c)

$$F_{M_1|W_1}(y|x) = \int_{0\vee x}^{y} 2(2u - x)e^{-2u(u - x)}du$$

$$= \int_{0\vee x}^{y} \left(-e^{-2u(u - x)}\right)' du$$

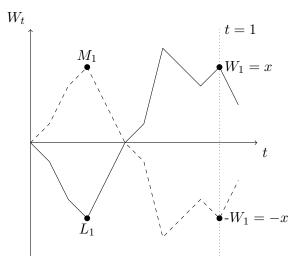
$$= \left[-e^{-2u(u - x)}\right]_{u=(0\vee x)}^{y}$$

$$= -e^{-2y(y - x)} + e^{0}, \quad (\text{because } (0\vee x)(x - (0\vee x)) = 0, \ \forall x \in \mathbb{R})$$

$$= 1 - e^{-2y(y - x)}, \quad y > (0\vee u)$$

(d) From the symmetry of standard Brownian motion, we have that the minimum conditional on  $W_1 = x$  has the same distribution as minus the maximum conditional on

 $W_1 = -x$ , i.e.  $L_1|(W_1 = x) \sim -M_1|(W_1 = -x)$ . The following plot illustrates the point:



Moreover, we can simulate the conditional maximum  $M_1$  given  $W_1 = x$  using the inverse CDF method. Letting  $M = M_1 | (W_1 = x)$ , we have

$$\begin{split} F_{M_1|(W_1=x)}(M) \sim U &\Leftrightarrow 1 - e^{-2M(M-x)} = U \\ &\Rightarrow -2M(M-x) = \ln(1-U) \Rightarrow 2M^2 - 2xM + \ln(1-U) = 0 \\ &\Rightarrow M = M_1|(W_1=x) = \frac{-(-2x) \pm \sqrt{(-2x)^2 - 4(2)\ln(1-U)}}{2(2)} \\ &= \frac{x + \sqrt{x^2 - 2\ln(1-U)}}{2} \quad \text{(rejecting the negative root)} \end{split}$$

Combining the two, we get the minimum as:

$$L_1|(W_1 = x) = -M_1|(W_1 = -x) = -\frac{-x + \sqrt{x^2 - 2\ln(1 - U)}}{2}$$
$$= \frac{x - \sqrt{x^2 - 2\ln(1 - U)}}{2}$$

Answer the following questions in an .R/.RMarkdown file.

### 3. (Simulation & Betting Strategies)

Consider the same setup as in question 1., i.e. a sequence of investment opportunities with i.i.d. Uniform (-1, 2) returns and an initial capital of  $V_0 = 1$ .

- (a) [5 points] Use the optimize() function to find the Kelly fraction  $f^*$  to invest, i.e. value of f which maximizes the expression in Q1.(g).
- (b) [10 points] Simulate 1,000 wealth paths consisting of 30 investments steps each using the Kelly fraction from the previous part (or, if you didn't get it, just f=.7). On the same axes, plot the *mean* simulated wealth, and the 5% and 95% quantiles of the simulated wealth at each step.
- (c) [10 points] In class, we mentioned that the Kelly criterion strategy reaches any wealth level faster than other strategies. You will test this claim with a simulation experiment. Fix a target wealth of 10,000 and simulate 1,000 paths with as many steps necessary to reach or exceed this level, under both the Kelly strategy  $(f^*)$  and the all-in strategy (f = 1). Report the average number of steps required to reach or exceed the target wealth under each strategy. Does it confirm the claim?

### 4. • (Industry Factor Models)

We have seen how to fit macroeconomic factor models, by regressing returns  $(\mathbf{R})$  on observed factors  $(\mathbf{F})$  one asset at a time, to estimate their loadings  $(\boldsymbol{\beta})$ . There is a related class of models called fundamental factor models, which are also fit by regression, but taking a different approach. The idea is to look at asset characteristics (i.e. fundamentals) to determine the asset loadings  $(\boldsymbol{\beta})$  based of their grouping with respect to common but unobserved factors. The simplest method is to group assets in non-overlapping industry categories, and assume all assets in a category are exposed to the same factor. In terms of the model, their loadings  $(\boldsymbol{\beta}$ 's) are either 1 for their industry's factor, or 0 for all other industry factors. The following model with 4 assets in 2 industries serves as an illustration:

$$\begin{bmatrix} R_{1,t} \\ R_{2,t} \\ R_{3,t} \\ R_{4,t} \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 1 & 0 \\ 0 & 1 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} F_{1,t} \\ F_{2,t} \end{bmatrix} + \begin{bmatrix} e_{1,t} \\ e_{2,t} \\ e_{3,t} \\ e_{4,t} \end{bmatrix} \Leftrightarrow \boldsymbol{R} = \boldsymbol{\beta}^{\top} \boldsymbol{F} + \boldsymbol{e}$$

In order to estimate the factor values  $(F_{1/2,t})$ , we use *cross-sectional* regression, i.e. we fix time t and look at returns as observations from the model

$$R_{i,t} = F_{1,t}I_{i,1} + F_{2,t}I_{i,2} + e_{i,t}, i = 1, \dots, n$$

where  $I_{i,1/2}$  are the indicator variables of asset i belonging to industry 1/2, serving as the explanatory variables, and the factors values  $F_{1/2,t}$  are the regression coefficients  $^1$ . The following code loads 250 daily net returns for n=121 assets (matrix R) belonging in 2 industries (factor sector). The loop fits 250 regressions to estimate the factor values (matrix  $\mathbf{f}$ ) and asset idiosyncratic errors (matrix  $\mathbf{e}$ ).

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<sup>&</sup>lt;sup>1</sup>In this simple case, this is effectively an ANOVA model, where factor values are simple averages of each industry's returns.

```
load("STAD70_W21_Final.RData")
b = model.matrix( object = ~ sectors - 1 )
f = matrix( 0, nrow = nrow(R), ncol = nlevels(sectors) )
e = matrix(0, nrow = nrow(R), ncol = ncol(R) )

for( i in 1:n_days){
   tmp = lm( R[i,] ~ b )
   f[i,] = tmp$coefficients
e[i,] = tmp$residuals
}
```

- (a) [5 points] Create a scatter-plot of the values of the two industry factors. Do the factors look independent? Explain your answer based on what you know about asset behaviour.
- (b) [3 points] Use the *sample covariance* of the returns (R) to calculate the minimum-variance-portfolio weights.
- (c) [10 points] Use the *industry factor model* to estimate the return covariance matrix (from  $\beta \Sigma_F \beta^\top + S_e$ ), and use that to calculate the minimum-variance-portfolio weights.
- (d) [3 points] Use the weights you found in the last two parts to create the returns of each portfolio, and calculate their Sharpe ratio assuming a risk-free rate of 0.
- (e) [4 points] Plot the *cumulative gross returns* of the two portfolios on the same axes.

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