The error in Bayesian phylogenetic reconstruction

when speciation is not instantaneous

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7 Abstract

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The tools for reconstructing phylogenetic relationships between taxonomic units (e.g. species) have become very advanced in the last three decades.

Among the most popular tools are Bayesian approaches, such as BEAST, MrBayes and RevBayes, that use efficient tree sampling routines to create a posterior probability distribution of the phylogenetic tree. A feature of these approaches is the possibility to incorporate known or hypothesized structure of the phylogenetic tree through the tree prior. It has been shown that the effect of the prior on the posterior distribution of trees can be substantial.

Currently implemented tree priors assume that speciation is instantaneous, where we know that speciation can be a gradual process.

Here we explore the effects of ignoring the protractedness of the speciation process with an extensive simulation study.

We compare the inferred tree to the simulated tree, and find that

Keywords: computational biology, evolution, phylogenetics, prior choice

$_{\scriptscriptstyle 24}$ 1 Introduction

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The computational tools that are currently available to the phylogeneticists go beyond the wildest imagination of those living four decades ago. Advances 26 in computational power allowed the first cladograms to be inferred from DNA alignments in 1981 (Felsenstein 1981), and the first Bayesian tools emerged in 1996 (Rannala & Yang 1996), providing unprecedented flexibility in the setup of a phylogenetic model. 30 Currently, the most popular Bayesian phylogenetics tools are BEAST (Drummond & Rambaut 2007) and its successor BEAST2 (Bouckaert et al. 2014), MrBayes (Huelsenbeck & Ronquist 2001) and RevBayes (Höhna et al. 2016). They allow to incorporate known or hypothesized structure of a phylogenetic tree-to-be-inferred through model priors. From these priors and an alignment 35 of DNA, RNA or protein sequences, they create a posterior distribution of parameter estimates (of the models used as a prior) and phylogenies, in which more probable combinations are represented more often. Each of these tools use efficient tree sampling routines to rapidly create an informative posterior. 39 The model priors in Bayesian phylogenetic reconstruction can be grouped into three categories: (1) site model, specifying nucleotide substitutions, (2) clock model, specifying the rate of mutation per lineage in time, and (3) tree model, constituting the speciation model underlying branching events (specia-43 tion) and branch termination (extinction). The choice of a wrong site model (Posada & Buckley 2004), clock model (Baele et al. 2012) or tree prior (Möller et al. 2018; Yang & Rannala 2005) is known to affect the posterior.

taneous, whilst we know that, speciation is often a gradual process (Schluter

Current phylogenetic tools use tree priors that assume speciation is instan-

⁴⁹ 2009). The (constant-rate) birth-death (BD) model is a commonly used tree ⁵⁰ prior, but it ignores this temporal aspect of speciation. The protracted birth-⁵¹ death (PBD) model, an extension of the BD model, does incorporate the idea ⁵² that speciation takes time. In this model, a branching event does not give rise ⁵³ to a new species, but to a new species-to-be, called an incipient species. Such an ⁵⁴ incipient species may go extinct, finish its speciation to become a good species, ⁵⁵ or give rise to new incipient species. Protracted speciation may explain observed ⁵⁶ declines in lineage accumulation (Etienne & Rosindell 2012).

Unfortunately, a tree prior according to this model, providing the probability of a species tree under the PBD model, is unavailable in current Bayesian phy-58 logenetic tools. Whilst an approximate formula for this probability has been derived (Lambert et al. 2015) and the approximation is very good (Simonet et al. 2018), it has not been implemented as tree prior yet. There are various reasons for this. First, the computation of this probability involves solving 62 set of non-linear differential equations, and while this computation is quite fast, it still takes much more time than the corresponding probability of the BD model which is a simple analytical formula. In a Bayesian MCMC chain, the tree prior probability must be calculated many times, and hence the total computation will take considerably longer with a PBD tree prior. Furthermore, 67 the approximate probability is a probability for the species tree assuming an underlying incipient species tree. It can be safely used as tree prior when only one individual per species is sampled, but if one has multiple samples per species -which is currently often the case- the methods to account for this such as the 71 multi-species coalescent (Heled & Drummond 2009) may not be compatible with the underlying incipient species tree. More precisely, the phylogeny under the PBD model may contain paraphylies, while the multi-species coalescent was developed exactly to avoid these by explaining them as arising from incomplete lineage sorting. Because of these paraphylies there is no such thing as a true species tree in the PBD model. To get a species-level tree one must sample one incipient species per species. Which incipient species is sampled may therefore have an impact on the species tree.

Here we aim to explore the effect of using the BD prior on PBD simulated phylogenies, taking into account possible sampling effects. In brief, we simulate protracted phylogenies using the PBD process, from which we sample a species tree in two very different ways. Given this species tree, we simulate a DNA sequence alignment. Then, we use BEAST2 on these alignments to infer a posterior of phylogenies, using a BD prior. We quantify the difference between the (BD) posterior phylogenies and the simulated (PBD) species tree.

Methods (but we are not allowed to keep this header)

The PBD model has five biological parameters (see 2), which we explore in a factorial fashion, excluding some combinations. We only simulate a PBD process for those combinations in which 95% of all simulated phylogenies are expected to have less than 1000 extant good species. [NOTE: use Rampals newest code. That new code assumes sirg = siri?]. We use 1000 good species as a threshold, to prevent overly taxon-poor and taxon-rich phylogenies respectively. The parameter values chosen are based on the parameter sets used by Etienne et al. 2014, as these parameters were shown to result in reasonably sized phylogenies and using the same set allows us to compare results. For the speciation initiation rates of good and incipient species, b_g and b_i respectively, we use 0.3 and 0.5 speciation initiation events per good/incipient species per time unit. The speciation completion rates we use are 0.1, 0.3, 1.0 and 10^9

speciation completion events per (incipient species) species per time unit. We use $10^9 \approx \infty$ to mimic the BD model, because the PBD model reduces to the BD model for $\lambda = \infty$. This allows us to measure the baseline error, which is the difference between inferred tree and true species tree that arises purely due to noise because the generating model and the model used in inference are identical in this case. The extinction rates of good and incipient species, μ_g and μ_i respectively, that we use are 0.0, 0.1 and 0.2 extinction events per good/incipient species per time unit.

From each biological parameter set, we simulate a protracted birth-death tree, using the PBD package (Etienne 2015) in the R programming language (R Core Team 2013), all with a crown age of 15 million years. Each protracted birth-death tree uses a different random number generatior seed, which makes all runs independent, resulting in a balanced data set. [NOTE: Rampal assumed runs with close seeds were related. I hope I have convinced him otherwise]

From each incipient species tree, we construct a species tree, by sampling one 116 incipient/good species per good species. For example, when an incipient species 117 branched off from its mother lineage, both of these subspecies are recognized 118 as representing the species, and hence both can be picked as an (equally good) 119 representative of the species. Here, we use three sampling scenarios, in which 120 we pick the representative randomly or in such a way that this results in either 121 the shortest or longest branch lengths. See the supplementary information for 122 visualization of these sampling methods. 123

Based on the sampled species tree, we simulate a DNA alignment that has
the same history as this species tree, using the phangorn package (Schliep 2011).
We assume that the nucleotides of the DNA alignment follow a Jukes-Cantor
(Cantor & Jukes 1969) nucleotide substitution model, in which all nucleotide-to-

nucleotide transitions are equally likely. In our Bayesian inference (see below) 128 we use the same site model as the (obviously correct) site model prior. One 129 could explore other substitution models in the simulations and in the Bayesian 130 inference, but we chose this simple model because we are primarily interested 131 in the effect of the choice of tree prior. If anything, our results are conservative: 132 with a more complex substitution model, there will be more noise and hence our 133 inference error will increase. We set the mutation rate in such a way to maximize 134 the information contained in the alignment. To do so, we set the mutation rate 135 such that we expect on average one (possibly silent) mutation per nucleotide between crown age and present, which equates to $\frac{1}{15}$ mutations per million years. 137 The DNA sequence length is chosen to provide a resolution of 10^3 years, that is, to have one expected nucleotide change per 10³ years per lineage on average. As 139 one nucleotide is expected to have on average one (possibly silent) mutation per 15 million years, $15 \cdot 10^3$ nucleotides result in 1 mutation per alignment per 10^3 141 years (which is coincidentally the same as Möller et al. 2018). The simulation 142 of these DNA alignments follows a strict clock model, which we will specify as 143 the known clock model prior in the Bayesian inference. 144

From an alignment, we run a Bayesian analysis and create a posterior dis-145 tribution of trees and parameters using the babette (?) package that sets the 146 input parameters similar to BEAUti 2 and then runs BEAST2. For our site and 147 clock model, we assume a Jukes-Cantor nucleotide substitution model and strict 148 clock model, as those are also used for simulating that alignment. We set the BD model as a tree prior, as gauging the effect of this incorrect assumption is 150 the goal of this study. We assume an MRCA prior with a tight normal distribu-15 tion around the crown age, by choosing the crown age as mean, and a standard 152 deviation of $0.5 \cdot 10^{-3}$ time units, resulting in 95% of the crown ages inferred have the same resolution (of 10^{-3} time units) as the alignment. We ran the 154

MCMC chain to generate 1111 states, of which we remove the first 10% (also called the 'burn-in'). Of the remaining 1000 MCMC states, the effective sample 156 size (ESS) of the posterior [NOTE: there is a parameter estimate called 157 'posterior'. I choose to pick that one, and I assume it is the wis-158 est choice of all BEAST2 parameter estimates, as displayed in table 159 4] must at least be 200 for a strong enough inference (Drummond & Bouck-160 aert 2015). An ESS can be increased by increasing the number of samples or 161 decreasing the autocorrelation between samples. If the ESS is less than 200, 162 we decrease autocorrelation by doubling the MCMC sampling interval of that simulation, until the ESS exceeds 200. 164

We compare each posterior phylogeny to the (sampled) species tree by the nLTT statistic (Janzen et al. 2015), using the nLTT package (Janzen 2015). The nLTT statistic equals the area between the normalized lineages-through-time-plots of two phylogenies, which has a range from zero (for identical phylogenies) to one. We use inference error and nLTT statistic interchangeably. Comparing the simulated species tree with each of the posterior species trees yields a distribution of nLTT statistics.

We produce two data sets as a comma-seperated file. We set the number of replicates for each parameter combination such, that this file and a possible copy can be handled in R's memory. Each row will then contain a parameter set and the generated nLTT statistics (see 3 for the exact data specification). The abovementioned memory constraints allows for $2 \cdot 10^3$ rows. With 48 [NOTE: recalculate] combinations of biological parameter, there will be 168 [NOTE: recalculate] replicates per parameter set.

For both data sets, we plot the nLTT statistics distribution per parameter set using a violin plot, as such a plot maintains information about the distribution. To simplify the interpretation of these plots, only nLTT statistics distribution

3 Results

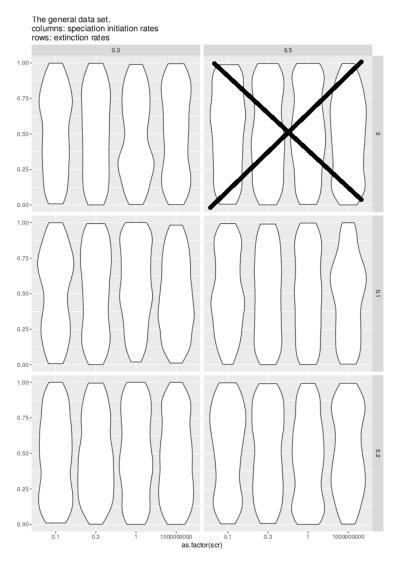


Figure 1: $\,$ nLTT statistic distribution per biological parameter set, using the balanced data set



Figure 2: nLTT statistic distribution per biological parameter set per sampling regime, using the data set conditioned on sampling regime having an effect

Term	Definition	
Phylogenetics	The inference of evolutionary relationships of groups of organ-	
	isms using genetics	
Model prior	Knowledge or assumptions about the ontogeny of evolutionary	
	histories	
Posterior	A collection of phylogenies and parameter estimates, in which	
	more probable combinations (determined by the data and the	
	model prior) are presented more frequently	
Protracted speciation	The process in which speciation takes two events to complete:	
	a speciation-initiation event and a speciation-completion event	
Speciation initiation	The start of a speciation event creating an incipient species	
Speciation completion	The end of a speciation event, in which an incipient species is	
	recognized as a good species	

Table 1: Glossary

4 Glossary

185 References

- Baele, G., Li, W.L.S., Drummond, A.J., Suchard, M.A. & Lemey, P. (2012) Ac-
- curate model selection of relaxed molecular clocks in bayesian phylogenetics.
- Molecular biology and evolution, **30**, 239–243.
- Bouckaert, R., Heled, J., Kühnert, D., Vaughan, T., Wu, C.H., Xie, D., Suchard,
- M.A., Rambaut, A. & Drummond, A.J. (2014) Beast 2: a software platform
- for bayesian evolutionary analysis. *PLoS computational biology*, **10**, e1003537.
- 192 Cantor, J. & Jukes, T. (1969) Mammalian protein metabolism. Evolution of
- protein molecules Academic Press, New York, NY, pp. 21–132.
- Drummond, A.J. & Bouckaert, R.R. (2015) Bayesian evolutionary analysis with
- 195 BEAST. Cambridge University Press.
- Drummond, A.J. & Rambaut, A. (2007) Beast: Bayesian evolutionary analysis
- by sampling trees. BMC evolutionary biology, 7, 214.

- Etienne, R.S. (2015) *PBD: Protracted Birth-Death Model of Diversification*. R
 package version 1.1.
- Etienne, R.S., Morlon, H. & Lambert, A. (2014) Estimating the duration of speciation from phylogenies. *Evolution*, **68**, 2430–2440.
- Etienne, R.S. & Rosindell, J. (2012) Prolonging the past counteracts the pull of
- $_{203}$ the present: protracted speciation can explain observed slowdowns in diver-
- sification. Systematic Biology, **61**, 204–213.
- $_{205}\,\,$ Felsenstein, J. (1981) Evolutionary trees from dna sequences: a maximum like-
- lihood approach. Journal of molecular evolution, 17, 368–376.
- Heled, J. & Drummond, A.J. (2009) Bayesian inference of species trees from
 multilocus data. *Molecular biology and evolution*, **27**, 570–580.
- Höhna, S., Landis, M.J., Heath, T.A., Boussau, B., Lartillot, N., Moore, B.R.,
- Huelsenbeck, J.P. & Ronquist, F. (2016) Revbayes: Bayesian phylogenetic
- inference using graphical models and an interactive model-specification lan-
- guage. Systematic biology, 65, 726–736.
- Huelsenbeck, J.P. & Ronquist, F. (2001) Mrbayes: Bayesian inference of phylo-
- genetic trees. *Bioinformatics*, **17**, 754–755.
- Janzen, T. (2015) nLTT: Calculate the NLTT Statistic. R package version 1.1.
- Janzen, T., Höhna, S. & Etienne, R.S. (2015) Approximate bayesian compu-
- tation of diversification rates from molecular phylogenies: introducing a new
- efficient summary statistic, the nltt. Methods in Ecology and Evolution, 6,
- 219 566-575.
- Lambert, A., Morlon, H. & Etienne, R.S. (2015) The reconstructed tree in
- the lineage-based model of protracted speciation. Journal of mathematical
- biology, **70**, 367–397.

- Möller, S., du Plessis, L. & Stadler, T. (2018) Impact of the tree prior on
- estimating clock rates during epidemic outbreaks. Proceedings of the National
- Academy of Sciences, p. 201713314.
- Posada, D. & Buckley, T.R. (2004) Model selection and model averaging in
- phylogenetics: advantages of akaike information criterion and bayesian ap-
- proaches over likelihood ratio tests. Systematic biology, **53**, 793–808.
- 229 R Core Team (2013) R: A Language and Environment for Statistical Computing.
- 230 R Foundation for Statistical Computing, Vienna, Austria.
- Rannala, B. & Yang, Z. (1996) Probability distribution of molecular evolution-
- ary trees: a new method of phylogenetic inference. Journal of molecular
- evolution, **43**, 304–311.
- Schliep, K. (2011) phangorn: phylogenetic analysis in r. Bioinformatics, 27,
- ₂₃₅ 592–593.
- ²³⁶ Schluter, D. (2009) Evidence for ecological speciation and its alternative. Sci-
- ence, **323**, 737–741.
- ²³⁸ Simonet, C., Scherrer, R., Rego-Costa, A. & Etienne, R. (2018) Robustness of
- the approximate likelihood of the protracted speciation model. Journal of
- evolutionary biology, **31**, 469–479.
- Yang, Z. & Rannala, B. (2005) Branch-length prior influences bayesian posterior
- probability of phylogeny. Systematic Biology, **54**, 455–470.

Parameter	Description	Values
b_g	Speciation initiation rate of a good species	0.3, 0.5
b_i^-	Speciation initiation rate of an incipient species	0.3, 0.5
λ	Speciation completion rate	$0.1, 0.3, 1.0, \infty$
μ_g	Extinction rate of a good species	0.0,0.1,0.2
μ_i^-	Extinction rate of an incipient species	0.0,0.1,0.2,0.4
t_c	Crown age	15
σ_c	Standard deviation around crown age	0.001
M	Sampling method	'shortest', 'longest' or random
r	Mutation rate	$\frac{1}{15}$
l_a	DNA alignment length	15K
f_i	MCMC sampling interval	1K or more
R_i	RNG seed incipient tree and randomly sampled species tree	1 to 20K
R_a	RNG seed alignment simulation	R_i
R_b	RNG seed BEAST2	R_i

Table 2: Overview of the 12 simulation parameters. Above the horizontal line is the biological parameter set. Sampling method M is random for the general data set. For the data set exploring the effect of sampling, we use 'shortest' for odd values of R_i , and 'longest' for even values of R_i . R_i is 1 for the first simulation, 2 for the next, etcetera.

\overline{n}	Description
12	simulation parameters, see table 2
1000	nLTT statistic values
11	ESSes of all parameters estimated by BEAST2 (see specs below)

Table 3: Specification of the data sets. Each row will contain one experiment, where the columns contain parameters, measurements and diagnostics. This table displays the content of the columns. n denotes the number of columns a certain item will occupy, resulting in a table of 1023 columns and 20K rows.

	D . 1.
#	Description
1	posterior
2	likelihood
3	prior
4	treeLikelihood
5	TreeHeight
6	BirthDeath
7	BDBirthRate
8	BDDeathRate
9	logP.mrca
10	mrcatime
11	clockRate

Table 4: Overview of the 11 BEAST2 estimated parameters