## Methods 3: Multilevel Statistical Modeling and Machine Learning

Week 6: *Mid-way evaluation and Machine Learning Intro*November 2, 2021

by: Lau Møller Andersen

#### These slides are distributed according to the CC BY 4.0 licence:

https://creativecommons.org/licenses/by/4.0/



Attribution 4.0 International (CC BY 4.0)

## Study Café how is it going?

# Reminder: Office hours (If assignments are hard, why is no one coming?)

## Python book in Stakbogladen

https://www.stakbogladen.dk/soegning.asp?phrase=

9781783555130

Also available online at the Royal Library (thanks, Emil!)



BOG

Python machine learning: unlock deeper insights into machine learning with this vital guide to cutting-edge predictive analytics

Sebastian Raschka author; Randal S Olson author of foreword 2015; 1st edition

#### Practical exercise tomorrow

To make sure that *Python* runs within *R Markdown*, make sure you have the *reticulate* package installed install.packages('reticulate')

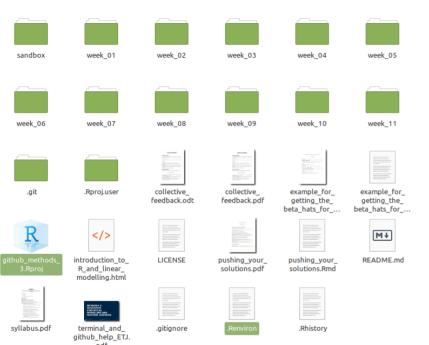
Also create a text file that is called *.Renviron* (remember the dot) placed in the folder where your *RProj* file is. It should have a single line: RETICULATE\_PYTHON=PATH where PATH is the path to your *methods3* conda environment. Use the commands below to find the paths:

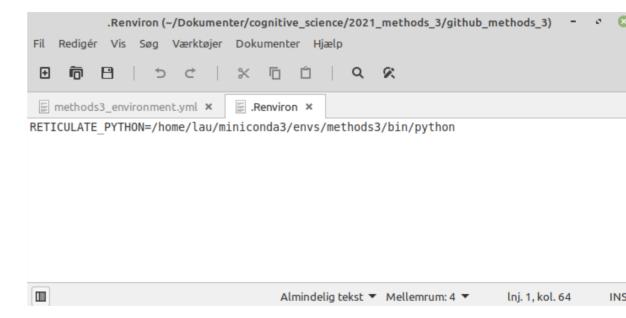
```
library(reticulate)
print(conda_list())

## name python
```

```
miniconda3
## 1
                                     /home/lau/miniconda3/bin/python
                       /home/lau/miniconda3/envs/methods3/bin/python
         methods3
## 2
                            /home/lau/miniconda3/envs/mne/bin/pvthon
## 3
              mne
## 4
         mne 0.17
                       /home/lau/miniconda3/envs/mne 0.17/bin/python
                   /home/lau/miniconda3/envs/mne func sig/bin/python
## 5 mne func sig
## 6
           mnedev
                         /home/lau/miniconda3/envs/mnedev/bin/python
## 7
         psychopy
                       /home/lau/miniconda3/envs/psychopy/bin/python
## 8
        fslpython
                                 /usr/local/fsl/fslpython/bin/python
        fslpython /usr/local/fsl/fslpython/envs/fslpython/bin/python
## 9
```

#### Practical exercise tomorrow





**NB!** No spaces around equals sign!

#### Practical exercise tomorrow

Update environment

conda env create --force -f methods3\_environment.yml

Overwrite old environment

Update environment (new packages have been added). Run this command from the folder week\_06

### Mid-way evaluation

#### Mid-way evaluation

- 1) Write something you liked about the course so far
- 2) Write something you did not like about the course so far
- 3) What would you change?

## Next time: I'll summarise the feedback on the three points and what I'll change

#### Learning goals

Evaluating and comparing models

- 1) Learning tools for comparing models
  - 1) Variance explained
  - 2) Likelihood ratio tests
  - 3) Information criteria
- 2) Bridging to out-of-sample

#### Why are we modelling?

Remember Emil's slides from week 03

- To be able to understand the world
- To be able to predict and manipulate the world

$$F = G \frac{m_1 m_2}{r^2}$$

**EXPLANATION** 



NASA/Bill Ingalls

**PREDICTION** 

#### What constitutes a good model?

Remember Emil's slides from week 03

- Accurate estimation
   of the underlying
   parameters of the
   population distribution
- Generalisation to new data

**EXPLANATION** 

**PREDICTION** 

Within an **explanatory** framework, how can we assess whether we have done a good job?

#### Variance explained

- Pros
  - R<sup>2</sup> is intuitive
- Cons
  - More complex models will always explain more variance
  - Hard to interpret in the case of collinearity
  - R<sup>2</sup> doesn't give us what we want

#### Likelihood ratio

#### Pros

- Models can be compared in a principled way by reference to a theoretical distribution,  $\chi^2$ . (In the single level case, F can be calculated)

#### Cons

- Models have to be nested in one another
- Maximum likelihood fitting may be biased for complex models
- Requires large sample sizes
- Be careful if collinearity is high

#### Information criteria

#### Pros

 Models can be compared even though one is not nested within the other (response data has to be the same though)

#### Cons

- Number of effective parameters not well defined for multilevel models
- Maximum likelihood fitting may be biased for complex models

#### Did you learn? (it's not easy)

Evaluating and comparing models

- 1) Learning tools for comparing models
  - 1) Variance explained
  - 2) Likelihood ratio tests
  - 3) Information criteria
- 2) Bridging to out-of-sample

#### Learning goals

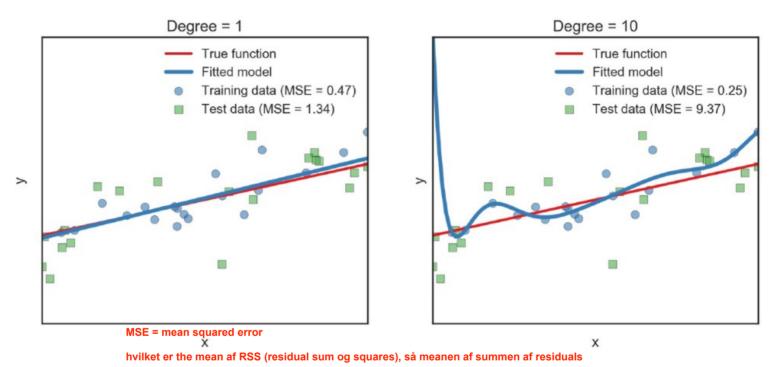
Explanation and prediction

- 1) Understanding that fitting (explaining) often leads to overfitting
- 2) Learning methods to prevent overfitting by introducing *bias*
- 3) Understanding how the error can be decomposed into *bias* and *variance*

#### To fit is to overfit

(Yarkoni and Westfall, 2017)

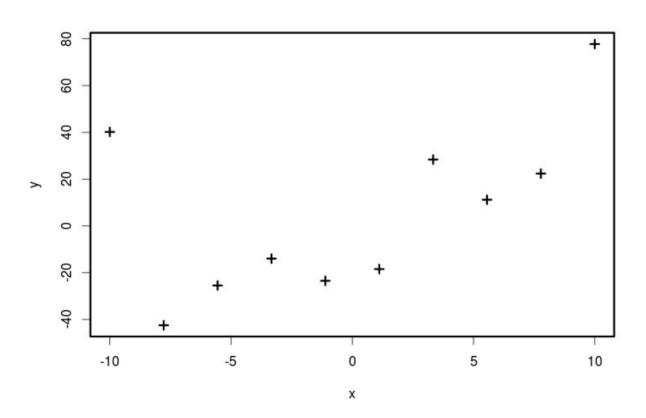
**Overfitting**: fitting sample-specific noise, which is thus not representative of the population



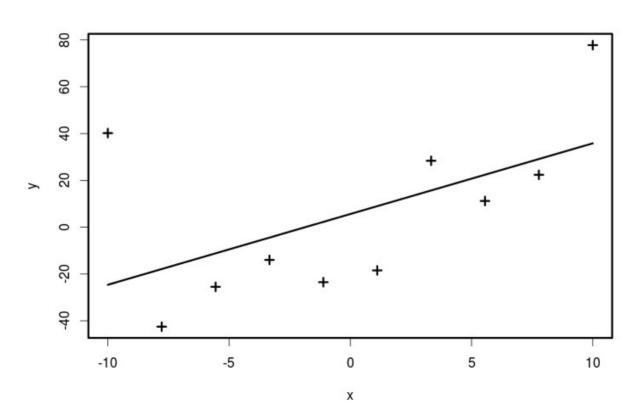
**Fig. 1.** Training and test error produced by fitting either a linear regression (left) or a 10th-order polynomial regression (right) when the true relationship in the population (red line) is linear. In both cases, the test data (green) deviate more from the model's predictions (blue line) than the training data (blue). However, the flexibility of the 10th-order polynomial model facilitates much greater overfitting, resulting in lower training error but much higher test error than the linear model. MSE = mean squared error.

(Yarkoni and Westfall, 2017)

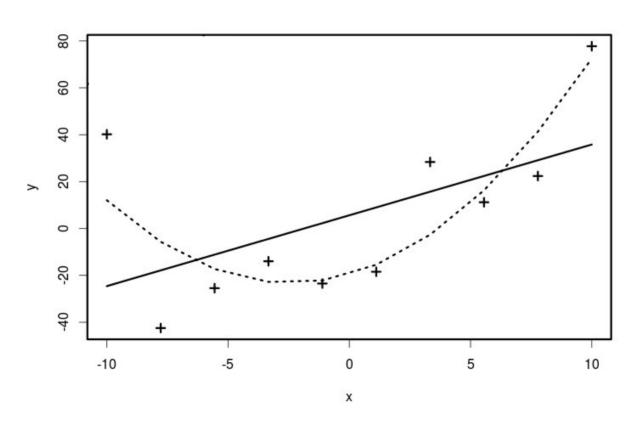
## A sample of 10 linear or quadratic?



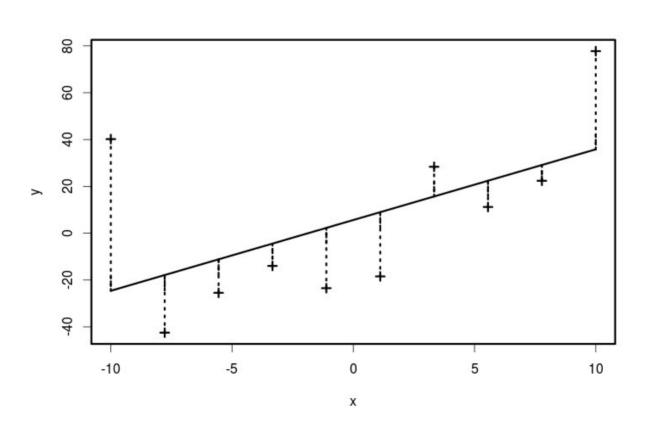
## A sample of 10 linear or quadratic?



## A sample of 10 linear or quadratic?

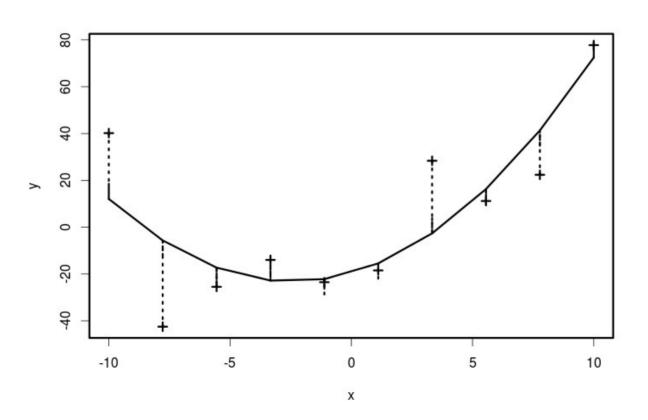


#### Residuals (linear)



#### Residuals (quadratic)

når den er quadratic har den en extra predictor (ectra predictor variable)



#### Quadratic:

$$ax^2 + bx + c$$

Linear:

$$bx + c$$

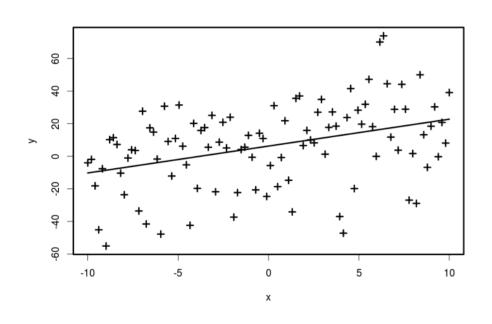
#### **Estimates**

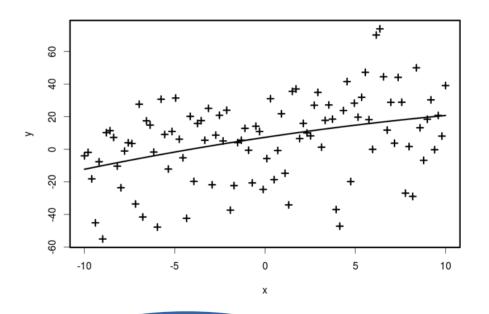
quadratic er bedre

$$a = 0,6184$$
;  $b = 3,0201$ 

$$b = 3,020$$

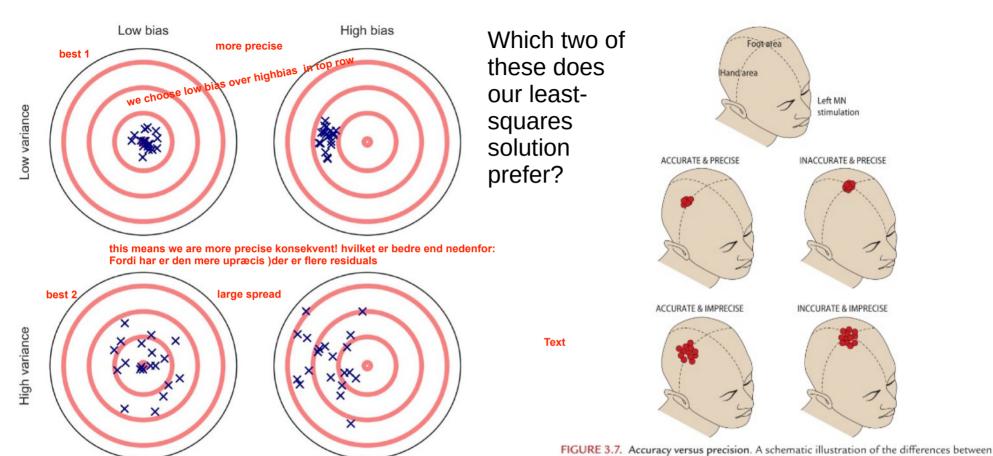
#### Now a sample of 100





$$b = 1,650$$

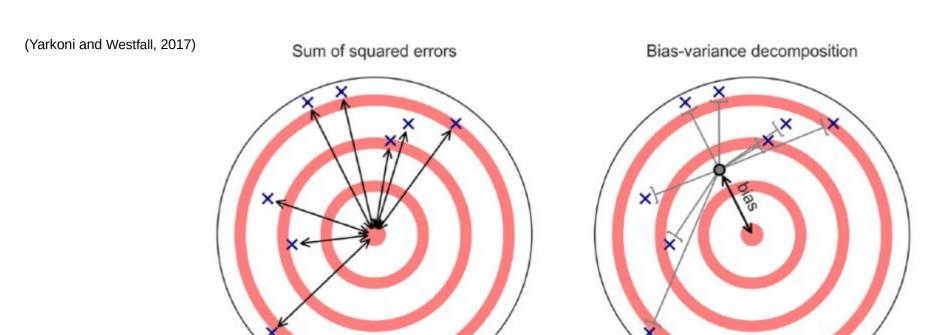
$$a = -0.03074 \approx 0.5b = 1.650$$



accuracy and precision of source localization. After left median-nerve stimulation, activations is expected in the right-hemisphere hand region of the primary somatosensory cortex. The foot area is shown at the top of the head. See text for further explanation.

(Yarkoni and Westfall, 2017)

of variance, or imprecision (compare bottom panels with top panels).



its impossible to know the bias, because it comes form the true function (true estimates) which we never really know..

**Fig. 3.** Schematic illustration of the bias-variance decomposition. (Left) Under the classical error model, prediction error is defined as the sum of squared differences between true scores and observed scores (black lines). (Right) The bias-variance decomposition partitions the total sum of squared errors into two separate components: a bias term that captures a model's systematic tendency to deviate from the true scores in a predictable way (black line) and a variance term that represents the deviations of the individual observations from the model's expected prediction (gray lines).

#### Bias variance decomposition

expectation of the error between true value and predicted value

$$E[(y_0 - \hat{f}(x_0))^2] = bias(\hat{f}(x_0))^2 + o^2$$

- 2) we dont know the bias, but we can estimate it
- 3) sample size 100
- 4) mx0 is one specific obsercation



this slide is not totally correct, lau will put new slides up for correction

bias (fhat )x0)) = E[ fhat(x0) ] - f(x0)  $\longrightarrow$  the difference between the fitted values and the how much we are off from the true function f(x0)

The only ting we know here is the true function, and with the true function we can calculate the bias (but in reality we never know the true function, but we have estimated paramters of the true function)..

is the expected squared remove between the true value  $y_0$  and its estimates based on fits  $\hat{f}(x_0)$  this is still stue tho

#### We'll look more into this during tomorrow's exercise

## Multilevel modelling as a *bias* introducer

"For example, some readers may be surprised to learn that multilevel modeling approaches to analyzing clustered data—which have recently seen a dramatic increase in adoption in psychology—improve on ordinary least squares (OLS) approaches to estimating individual cluster effects by deliberately biasing (through "shrinking" or "pooling") the cluster estimates toward the estimated population average"

(Yarkoni and Westfall, 2017)

#### Introducing bias

"In a widely used form of penalized regression called lasso regression (Tibshirani, 1996, 2011), this leastsquares criterion is retained, but the overall cost function that the estimation seeks to minimize now includes an additional penalty term that is proportional to the sum of the absolute values of the coefficients."

(Yarkoni and Westfall, 2017)

#### Penalised regression

 $RSS = \sum_{i} (y_i - \hat{y}_i)^2 (minimise to obtain least squares solution)$ 

lasso regression: RSS+
$$\lambda \sum_{j=1}^{p} |\beta_j|$$
 (minimise this sum) betas will be sat to 0 = will be shurnken

... setting a parameter (predictor) to 0 is just saying that it does not expalin anything

in a model with multiple predictors, when the vairables are correlated (e.g wt and mpg), then we can set some of them to 0

ridge regression: RSS+
$$\lambda \sum_{j=1}^{p} (\beta_j^2)$$
 (minimise this sum)

*i*:observations

*p*: predictor variables

 $\lambda$ : a constant

#### Penalised regression

RSS = 
$$\sum (y_i - \hat{y}_i)^2$$
 (minimise to obtain least squares solution)

lasso regression : RSS+
$$\lambda \sum_{j=1}^{p} |\beta_{j}|$$
 (minimise this sum)

ridge regression : RSS + 
$$\lambda \sum_{j=1}^{p} (\beta_{j}^{2})$$
 (minimise this sum)

*i*:observations

*p*: predictor variables

 $\lambda$ : a constant

#### **Group discussion**

In each case: what happens when?

- 1.  $\lambda$  increases?
- 2.  $\lambda$  decreases?
- 3.  $\lambda$  is 0?
- 4.  $\lambda$  goes towards infinity?

## Penalised regression

 $RSS = \sum_{i} (y_i - \hat{y}_i)^2 (minimise to obtain least squares solution)$ 

$$\underset{\lambda}{\operatorname{argmin}} = \operatorname{RSS} + \lambda \sum_{j=1}^{p} |\beta_j|$$

$$\underset{\lambda}{\operatorname{argmin}} = \operatorname{RSS} + \lambda \sum_{j=1}^{p} \left( \beta_{j}^{2} \right)$$

*i*:observations

*p*: predictor variables

 $\lambda$ : a constant

## How to choose $\lambda$ ?

```
##
## Call:
## lm(formula = hp ~ mpg + wt + drat + qsec, data = mtcars)
##
## Coefficients:
## (Intercept) mpg wt drat qsec
## 473.779 -2.877 26.037 4.819 -20.751
```

#### What is $\lambda$ equal to here?

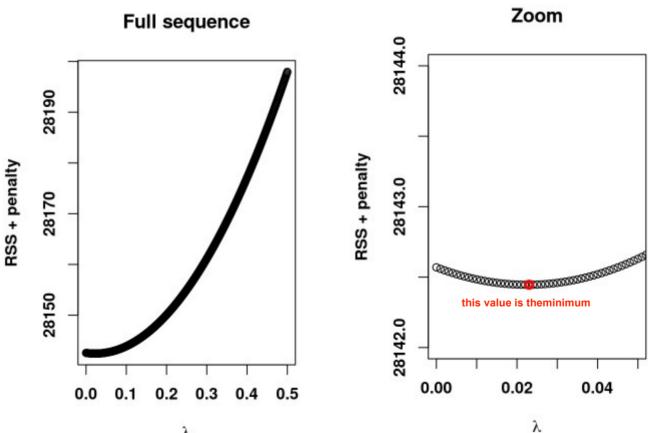
```
sum.of.squares.total <- sum((y - mean(y))^2)
sum.of.squared.errors.lm <- sum(residuals(linear_model)^2)
print(r.squared.lm <- 1 - sum.of.squared.errors.lm/sum.of.squares.total)</pre>
```

```
## [1] 0.8072553
```

## How to choose $\lambda$ (lasso)?

```
##
          glmnet(x = x, y = y, alpha = 1, lambda = c(0, 0.2, 2, 4, 20,
100))
##
                             lambda
                                             RSS
                                                      penalty
                                                                       sum
##
     Df
          %Dev Lambda
                                                                 145726.9
                              100.0 145726.9
        0.0000
                 100.0
##
      2 0.6567
                  20.0
                                                      14.05496
                           20.00000 50025.64218
                                                               50039.69714
##
  3
      3 0.8004
                   4.0
                            4.00000 29082.92003
                                                      41.34363 29124.26366
                            2.00000 28408.50683
                                                     44.99057 28453.49740
##
                   2.0
      3 0.8051
                   0.2
                            0.20000 28097.60764
                                                     52.47741 28150.08505
      4 0.8072
                            0.00000 28088.09951
                                                      54.46997 28142.56948
## 6
      4 0.8073
                   0.0
                                                    something is wrong here too
```

## How to choose $\lambda$ ?



## What does \(\lambda\) do (ridge)?

$$\hat{\beta}_{ridge} = (X^T X + \lambda I)^{-1} X^T y$$

I: an identity matrix with p predictor variables

 $\lambda$ : a constant (small) lambda can only be positive

lambda is timed with our identity matrix

## What is X<sup>T</sup>X?

head (X) this is our X matrix

print(cov.X) then we look at the covariated of the X matrix

$$X_{COV} = X^T X$$

## mpg 14042.310 1909.7528 2380.2770 11614.745
## wt 1909.753 360.9011 358.7190 1828.095
## drat 2380.277 358.7190 422.7907 2056.914
## qsec 11614.745 1828.0946 2056.9140 10293.480

graphically

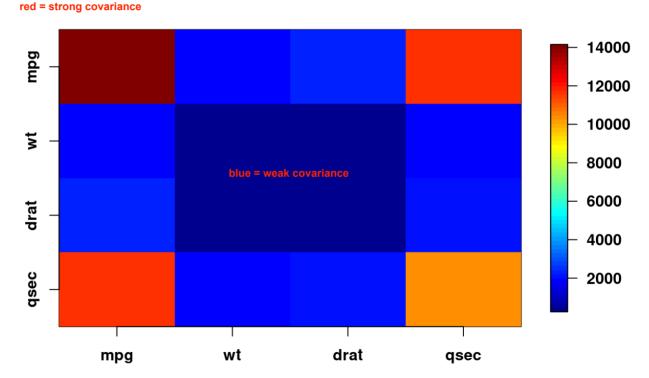
$$X_{COV} = X^T X$$

in the ideal world, a covariance matrix would look like this

X cov = its a part of our least squares solution

**Covariance matrix** 

The fact that the off-diagonal > 0, indicates that there is collinearity



```
Collinearity can be bad

remember to give the correct order of predictors, cuz they have agency to impact in the way they are ordered

##

## Call:

## lm(formula = hp ~ mpg + wt + drat + qsec, data = mtcars)

##

## Coefficients:

with colinarity being present, when we cant interpret the paramters below like we ususally do:
```

mpg

-2.877

Assuming no collinearity, what is the interpretation of the coefficients? With collinearity, is that interpretation possible?

26.037

qsec

-20.751

drat

4.819

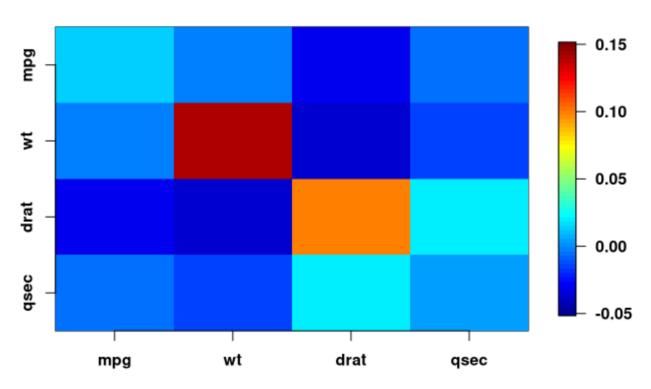
##

(Intercept)

473.779

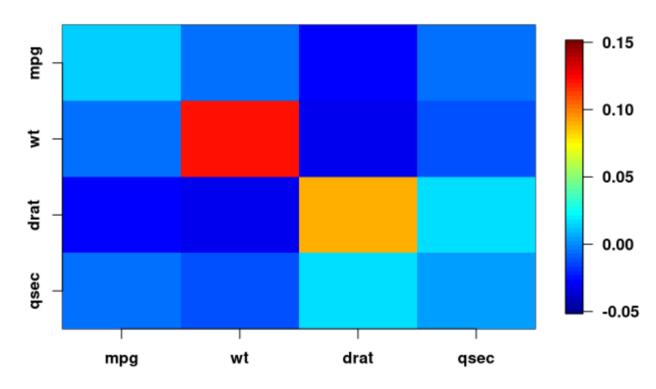
$$\hat{\beta_{ridge}} = (X^T X + \lambda I)^{-1} X^T y$$

#### Inverted Covariance matrix, regularized:(lambda= 0)



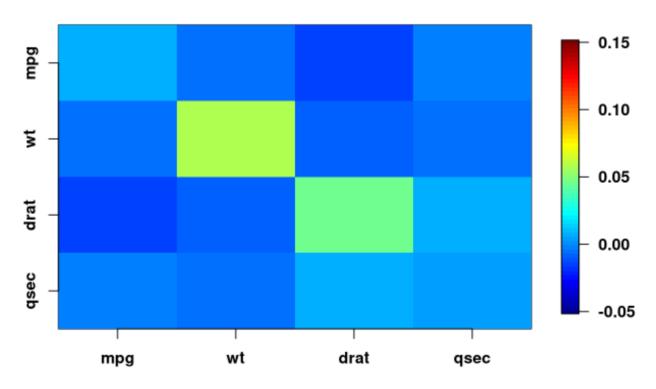
$$\hat{\beta_{ridge}} = (X^T X + \lambda I)^{-1} X^T y$$

#### Inverted Covariance matrix, regularized:(lambda= 1)



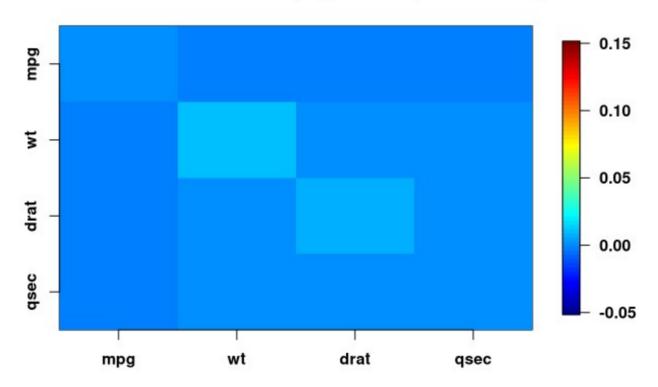
$$\hat{\beta_{ridge}} = (X^T X + \lambda I)^{-1} X^T y$$

#### Inverted Covariance matrix, regularized:(lambda= 10)



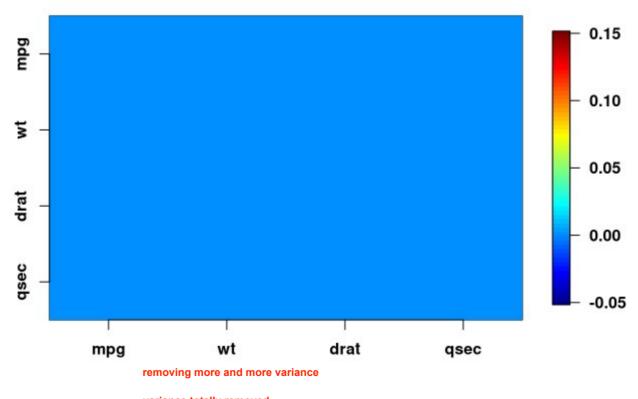
$$\hat{\beta_{ridge}} = (X^T X + \lambda I)^{-1} X^T y$$

#### Inverted Covariance matrix, regularized:(lambda= 100)



$$\hat{\beta_{ridge}} = (X^T X + \lambda I)^{-1} X^T y$$

#### Inverted Covariance matrix, regularized:(lambda= 1000)



variance totally removed 53

# So why is it called regularisation?

## Two notes about the inverted matrix:

with increase of  $\lambda$ 

we want low bias

- 1. Diagonal shrinks (bias is added)
- 2. Off-diagonal shrinks (collinearity is reduced, which improves the stability of the model)

its a trade offer= we give more bias but less variance



#### In a stable model:

# Feeding new data or adding new predictor variables will not change the parameter estimates a lot

so adding another parameter just exptains more noise, and if we keep on doing this, we expalain more variance BUT we also start to fit noise (so we are moving torwards overfitting)

and that is why adding another predictor or adding data is not optimizing the model (it does to a certain extend) but we cant just keep on adding predictos, then we will start to overfit:

ENTER LAMBDA!!!!!!!!!!!!

We have succeeded in finding a  $\lambda$  making our model more stable (improved **in-sample** validity), but we haven't found a  $\lambda$  that optimises predictive power – (**out-of-sample**)

## Out-of-sample as validity check



mtcars.1 <- mtcars[1:10, ]</pre>



## Out-of-sample as validity check

```
Call:
lm(formula = hp \sim mpg + wt + drat + qsec, data = mtcars.1)
Coefficients:
(Intercept)
                                                 drat
                                     wt
                                                                qsec
                      mpg
    414.541
                  -13.638
                                 12.753
                                               11.263
                                                             -5.042
Call:
 lm(formula = hp \sim mpq + wt + drat + qsec, data = mtcars)
Coefficients:
                                                  drat
 (Intercept)
                                      wt
                       mpg
                                                                qsec
     473.779
                    -2.877
                                  26.037
                                                 4.819
                                                             -20.751
```

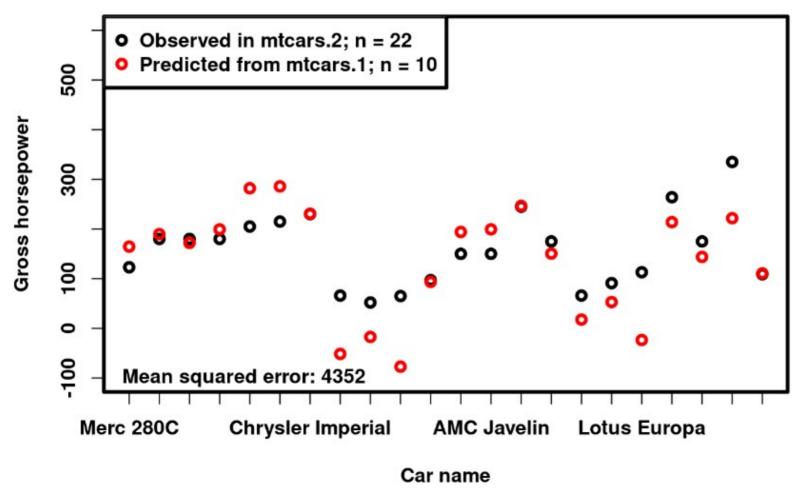
### Suddenly, someone shows up with



## Let's check our model

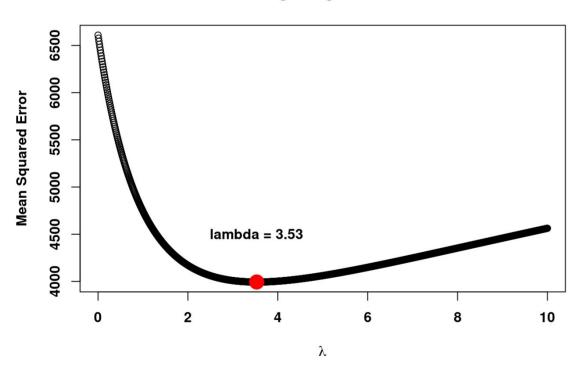
mtcars.2 <- mtcars[11:32, ]</pre>

#### Predictions based on mtcars.1



## Finding optimal lambda

#### **Ridge Regression**



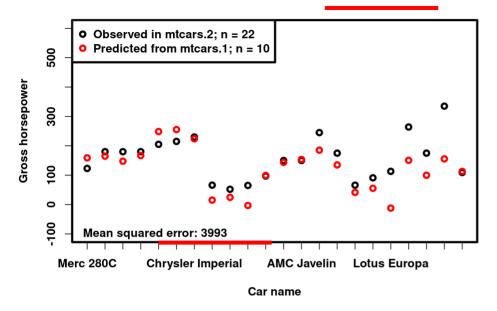
$$MSE = mean((y - \hat{y})^2)$$

Which dataset is the MSE calculated on?

by introducing bias with lambda, then we minimize the MSE the most! we introduce bias to get better predictions

```
Call:
lm(formula = hp \sim mpg + wt + drat + qsec + 0, data = mtcars.1)
Coefficients:
   mpg wt drat qsec
-8.379 76.588 45.705 -5.850
print(beta.hat.ridge <- ridge.regression(X, y, mtcars.1, min.lambda))</pre>
##
             mpg wt drat gsec
## [1,] -7.421887 33.637 22.0378 4.70691
```

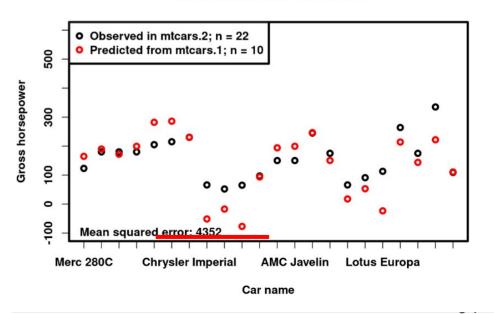
#### What has happened to the coefficients?



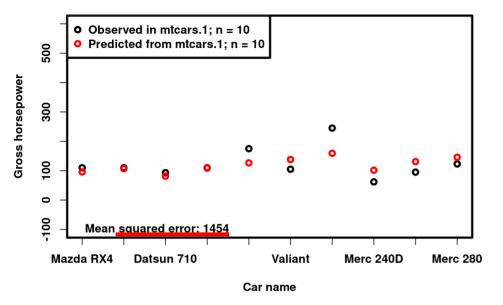
#### **Prediction on mtcars.2**

$$\lambda = 0$$





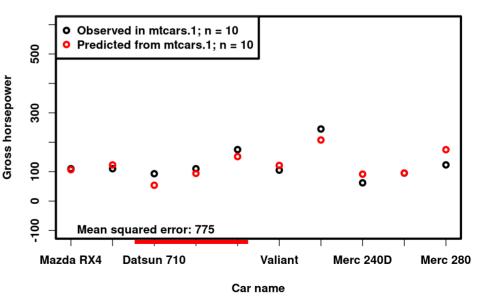
#### Predictions based on mtcars.1; lambda=3,53



#### "Prediction" on mtcars.1

$$\lambda = 0$$

#### Predictions based on mtcars.1; lambda=0



### Nomenclature

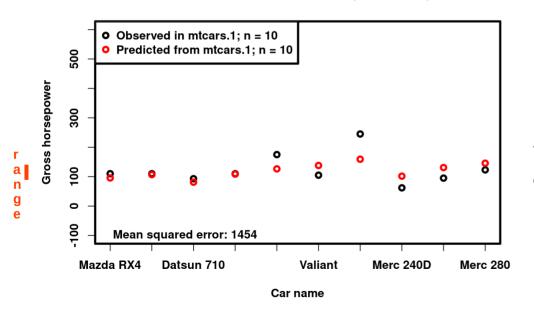
- mtcars.1 -> training set
- mtcars.2 -> test set
  - NB! Normally, we prefer that out training set is bigger than our test set
- By introducing bias in our training set, we at the same time reduce the variance of our training set, increasing the reliability of our predictions on a test set

## "Testing" on training set

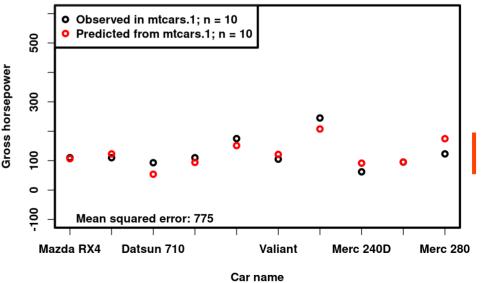
greater bias lesser variance

smallest bias greater variance (of  $\hat{y}$ )

Predictions based on mtcars.1; lambda=3,53

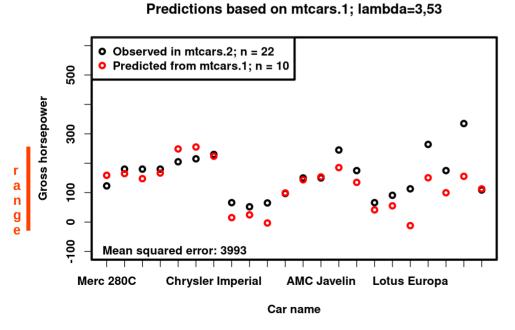


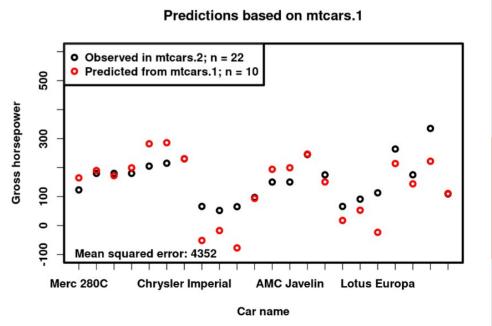
Predictions based on mtcars.1; lambda=0



## Optimal $\lambda$

lesser bias lesser variance more bias greater variance (of  $\hat{y}$ )





### Did you learn?

Explanation and prediction

- 1) Understanding that fitting (explaining) often leads to overfitting
- 2) Learning methods to prevent overfitting by introducing *bias*
- 3) Understanding that the error can be decomposed into *bias* and *variance*

### Next time

- The Perceptron
- Adaline
- Linear regression

#### References

- Bolker, B.M., Brooks, M.E., Clark, C.J., Geange, S.W., Poulsen, J.R., Stevens, M.H.H., White, J.-S.S., 2009. Generalized linear mixed models: a practical guide for ecology and evolution. Trends in Ecology & Evolution 24, 127–135. https://doi.org/10.1016/j.tree.2008.10.008
- Gelman, A., Hill, J., 2006. Data Analysis Using Regression and Multilevel/Hierarchical Models. Cambridge University Press.
- Hari, R., Puce, A., 2017. MEG-EEG Primer. Oxford University Press, New York, NY, US.
- Yarkoni, T., Westfall, J., 2017. Choosing Prediction Over Explanation in Psychology: Lessons From Machine Learning. Perspect Psychol Sci 12, 1100–1122. https://doi.org/10.1177/1745691617693393