Lecture 1

Distribution-free inference: aims and algorithms

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Why do we want "distribution-free" guarantees?

When we analyze data, we...

- Run a model/algorithm that is valid under certain assumptions
 (parametric model / smoothness conditions / sparsity assumption / ...)
- But if the assumptions don't hold, can we trust the output?
 (parameter estimate / predicted value / error bound / hypothesis test / ...)

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The goal of distribution-free inference is to provide guarantees that are valid universally over all data distributions.

What are inference questions we might want to ask, distribution-free?

- Prediction: the unobserved response Y_{n+1} will lie in [some range]
- Effect size: the dependence between X and Y lies in [some range]
- Independence: test if X & Y independent given [some confounders]
- Regression: the distribution of Y given X satisfies [some property]

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Lecture 1 — algorithms for prediction Lecture 2 — challenges with other inference problems

If we fit a regression model $\widehat{\mu}: \mathcal{X} \to \mathbb{R}$ on training data, how to build a prediction interval $\widehat{\mu}(X_{n+1}) \pm [...]$ for Y_{n+1} ?

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- If we use the training residuals $R_i = |Y_i \widehat{\mu}(X_i)|$, these are too small due to overfitting
- Holdout method: fit $\widehat{\mu}$ on training set $i=1,\ldots,\frac{n}{2}$, then use residuals $\{R_i=|Y_i-\widehat{\mu}(X_i)|\}_{i=\frac{n}{2}+1,\ldots,n}$ to choose width

Prediction interval: fitted on training data $i = 1, ..., \frac{n}{2}$

$$\widehat{C}_n(X_{n+1}) = \widehat{\mu}(X_{n+1}) \quad \pm \quad \underbrace{Q_{1-\alpha}(R_{\frac{n}{2}+1}, \dots, R_n)}_{\text{the } (1-\alpha)(\frac{n}{2}+1)\text{-th smallest resid.}}$$
in the holdout set $i = \frac{n}{2} + 1, \dots, n$

Theorem: if data points are i.i.d., $\mathbb{P}\left\{Y_{n+1} \in \widehat{C}_n(X_{n+1})\right\} \geq 1 - \alpha$

¹Vovk, Gammerman, Shafer 2005, Algorithmic Learning in a Random World

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Proof: $\{(X_i, Y_i)\}_{i=\frac{n}{2}+1,...,n+1}$ are i.i.d. conditional on training data \Rightarrow residuals $\{R_i\}_{i=\frac{n}{2}+1,...,n+1}$ are i.i.d. cond. on training data

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• In the above construction,

$$\widehat{C}_n(X_{n+1}) = \widehat{\mu}(X_{n+1}) \pm [\ldots] = \{ \text{ all } y \text{ values with } |y - \widehat{\mu}(X_{n+1})| \leq [\ldots] \}$$

• We can generalize to any score function:²

$$\widehat{C}_n(X_{n+1}) = \{ \text{ all } y \text{ values with } S(X_{n+1}, y) \leq [...] \}$$

where S(x,y) measures "nonconformity" of the data point (x,y)

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The method:

- Using data i = 1, ..., n/2, fit "nonconformity score" S(x, y)
- Compute $S_i = S(X_i, Y_i)$ for i = n/2 + 1, ..., n, & $Q_{1-\alpha}(S_{n/2} + 1, ..., S_n)$
- Prediction interval:

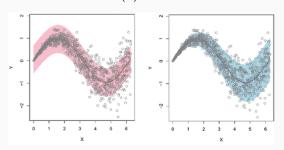
$$\widehat{C}_n(X_{n+1}) = \{ y : S(X_{n+1}, y) \le Q_{1-\alpha}(S_{n/2} + 1, \dots, S_n) \}$$

²Vovk, Gammerman, Shafer 2005, Algorithmic Learning in a Random World

Why any score function?

• If noise level varies with X, may want varying interval width:³

$$S(x,y) = \frac{|y - \widehat{\mu}(x)|}{\widehat{\sigma}(x)} \quad \Rightarrow \quad \widehat{C}_n(X_{n+1}) = \widehat{\mu}(X_{n+1}) \pm \widehat{\sigma}(X_{n+1}) \cdot Q_{1-\alpha}(...)$$



(figure from Lei et al 2018)

³Lei et al 2018, Distribution-Free Predictive Inference for Regression

Why any score function?

- If Y is discrete (or even categorical), 4 not natural to take $\widehat{C}_n(X_{n+1}) = [\text{some continuous interval}]$
- ullet Instead, take $\widehat{C}_n(X_{n+1})=\{ ext{all categories }y ext{ that are likely given }X_{n+1}\}$

⁴Sadinle et al 2019, Least ambiguous set-valued classifiers with bounded error levels

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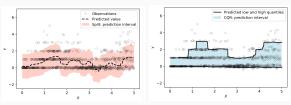
- If Y is discrete (or even categorical), ⁴ not natural to take $\widehat{C}_n(X_{n+1}) = [\text{some continuous interval}]$
- Instead, take $\widehat{C}_n(X_{n+1}) = \{ \text{all categories } y \text{ that are likely given } X_{n+1} \}$
- Compute estimate $\widehat{p}(y|x) \approx \mathbb{P}\left\{Y = y|X = x\right\}$, then use $S(x,y) = 1/\widehat{p}(y|x) \Rightarrow \widehat{C}_n(X_{n+1}) = \{y : \widehat{p}(y|X_{n+1}) \geq [...]\}$

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Why any score function?

- If the shape of distrib. of Y|X varies with X, $\widehat{C}_n(X_{n+1}) = \widehat{\mu}(X_{n+1}) \pm [...]$ is not an optimal form for the interval
- Instead, can estimate conditional quantiles directly:⁵

$$S(x,y) = \max\{\widehat{q}_{\alpha/2}(x) - y, \ y - \widehat{q}_{1-\alpha/2}(x)\} \quad \text{or} \ \max\left\{\frac{\widehat{q}_{0.5}(x) - y}{\widehat{q}_{0.5}(x) - \widehat{q}_{\alpha/2}}, \ \frac{y - \widehat{q}_{0.5}(x)}{\widehat{q}_{1-\alpha/2}(x) - \widehat{q}_{0.5}(x)}\right\}$$



(figure from Romano et al 2019)

⁵Romano et al 2019, Conformalized quantile regression Kivaranovic et al 2019 Adaptive, distribution-free prediction intervals for deep neural networks

All methods so far rely on data splitting:

- Training: use n/2 data points to develop a score function S(x, y)
- Validation: use n/2 data points to learn the distrib. of S(X,Y)
- Then we can predict $S(X_{n+1}, Y_{n+1}) \rightsquigarrow$ can predict Y_{n+1}

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The drawback: sample splitting means that we only use n/2 data points to fit the model / the score function

Conformal prediction: use all data points for training & validation⁶

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Conformal prediction: use all data points for training & validation⁶

 $^{^6\}mathrm{Vovk}$, Gammerman, Shafer 2005, Algorithmic Learning in a Random World

• Suppose we observe training + test data:

$$(X_1, Y_1), \ldots, (X_n, Y_n), (X_{n+1}, Y_{n+1})$$

• Fit model $\widehat{\mu}$ to all n+1 data points, & get residuals

$$R_i = Y_i - \widehat{\mu}(X_i), \ i = 1, \dots, n, \quad R_{n+1} = Y_{n+1} - \widehat{\mu}(X_{n+1})$$

• Check if $|R_{n+1}| \leq [(1-\alpha)$ quantile of $|R_1|, \ldots, |R_n|, |R_{n+1}|]$



By exchangeability of R_1, \ldots, R_{n+1} , this event has $\geq 1 - \alpha$ probability

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By exchangeability of R_1, \ldots, R_{n+1} ,

this event has $\geq 1-lpha$ probability if we plug in ${\it y}={\it Y}_{\it n+1}$

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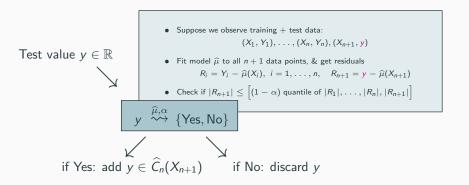
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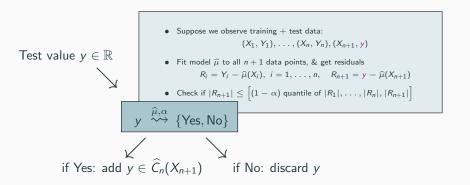
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ceil$







$$\mathbb{P}\left\{Y\in\widehat{C}_n(X_{n+1})
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"Full" conformal can be run with any score function on data sets:

$$((x_1,y_1),\ldots,(x_{n+1},y_{n+1}))\mapsto (S_1,\ldots,S_{n+1})$$

 S_i = "nonconformity score" of data point i, relative to rest of the data

- Regression: $S_i = |y_i \widehat{\mu}(x_i)|$ where $\widehat{\mu}$ is fitted on all data
- Quantile regression: S_i measures gap between y_i and $\widehat{q}(x_i)$
- Classification: $S_i = 1/\widehat{p}(y_i|x_i)$
- & many more

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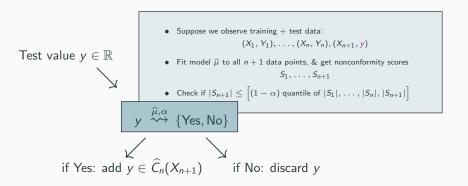
- Fit model $\widehat{\mu}$ to all n+1 data points, & get nonconformity scores S_1,\ldots,S_{n+1}
- $\bullet \quad \mathsf{Check} \; \mathsf{if} \; |S_{n+1}| \leq \Big\lceil (1-\alpha) \; \mathsf{quantile} \; \mathsf{of} \; |S_1|, \ldots, |S_n|, |S_{n+1}| \Big\rceil$

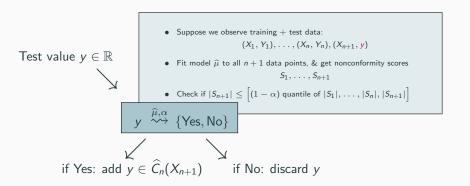
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Conformal prediction (or holdout method) assumes: training & test data are from the same distribution

One possible violation: covariate shift

- Marginal distribution of X is different in training vs. test data (e.g., some subpopulations are over- or under-represented in the training data)
- But, distribution of Y|X is the same

Assuming we know the shift (i.e., $dP_X^{\text{test}}(x) \propto w(x) \cdot dP_X^{\text{train}}(x)$), conformal can adjust for the shift by using weighted exchangeability⁷

⁷Tibshirani, B., Ramdas, & Candès 2019, Conformal prediction under covariate shift Hu & Lei 2020, A distribution-free test of covariate shift using conformal prediction

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known fn.

Given n + 1 data points...

- With (unweighted) exchangeability, each one is equally likely to be the test point $\rightsquigarrow R_{n+1} \leq Q_{1-\alpha}(R_1,\ldots,R_{n+1}) \text{ with prob. } 1-\alpha$
- With weighted exchangeability, the distribution is nonuniform:

$$\mathbb{P}\left\{(x,y) \text{ is the test point}\right\} \propto w(x)$$

$$ightharpoonup$$
 need to compute a weighted quantile: $Q_{1-\alpha}\left(\sum_{i} \frac{w(X_i)}{\sum_{i'} w(X_{i'})} \delta_{R_i}\right)$

⁷Tibshirani, B., Ramdas, & Candès 2019, Conformal prediction under covariate shift Hu & Lei 2020, A distribution-free test of covariate shift using conformal prediction

Application: survival analysis & censored data⁸

- "Clean" data $(X_i, Y_i) = (features, survival time)$
- Censored observations (X_i, \tilde{Y}_i) where $\tilde{Y}_i = \min\{C_i, Y_i\}$
- Main idea: choose a cutoff c_0 so that "usually" $Y_i \leq c_0$, & keep only data with $C_i \geq c_0$ (i.e., most Y's are not censored)

⁸Candès, Lei, Ren 2021, Conformalized survival analysis

Weighted conformal prediction

Application: survival analysis & censored data⁸

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- Main idea: choose a cutoff c_0 so that "usually" $Y_i \leq c_0$, & keep only data with $C_i \geq c_0$ (i.e., most Y's are not censored)
 - On this data set, can use CP to predict survival time Y
 - But, this may be a different distribution (population with $C_i \geq c_0 \neq \text{general population}$)
 - If distrib. of C|X known, can use weighted CP to correct for distribution shift

⁸Candès, Lei, Ren 2021, Conformalized survival analysis

Weighted conformal prediction

Application: estimating individual treatment effects⁹

- Data $(X_i, T_i, Y_i) =$ (features, treatment group = 0 or 1, outcome)
- ITE_i = (value of Y_i , if $T_i = 1$) (value of Y_i , if $T_i = 0$)
- ullet Challenge: treatment assignment may depend on X
- Main idea: if propensity score $\mathbb{P}\left\{T=1\mid X=x\right\}$ is known, can use weighted CP to adjust for X|T=1 versus X|T=0

⁹Lei & Candès 2020, Conformal inference of counterfactuals and individual treatment effects

Weighted conformal prediction

A related problem — label shift (for categorical Y / classification)¹⁰

- Marginal distribution of Y is different in training vs. test data (e.g., some subpopulations are over- or under-represented in the training data)
- But, distribution of X|Y is the same

If the label shift is known (i.e., the weight function $w(y) = \frac{P_{Y,\text{test}}(y)}{P_{Y,\text{train}}(y)}$), can use weighted exchangeability to guarantee coverage

 $^{^{10}}$ Podkopaev & Ramdas 2021, Distribution-free uncertainty quantification for classification under label shift

"Full" conformal prediction requires that the model fitting algorithm is re-run:

- For each test value X_{n+1} of interest
- For every possible value $y \in \mathbb{R}$

Approaches:

- In practice use a grid of *y* values (but no theory)
- Specialized methods for specific algorithms e.g. Lasso¹¹
- Discretized CP —
 "discretize the data" or "discretize the model" 12

 $^{^{11}}$ Lei 2017, Fast Exact Conformalization of Lasso using Piecewise Linear Homotopy

¹²Chen, Chun, & B. 2017, Discretized conformal prediction for efficient distribution-free inference

A preliminary observation—

It is valid to run CP on the interval $[\min_{1 \le i \le n} Y_i, \max_{1 \le i \le n} Y_i]$

- With prob. $\geq 1 \frac{2}{n+1}$, including Y_{n+1} doesn't change the endpoints
- So, coverage is $\geq 1 \alpha \frac{2}{n+1}$

Conformal prediction:



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Conformal prediction with rounding (informal version):

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Conformal prediction:



Conformal prediction with rounding (informal version):

Conformal prediction:

$$\longleftrightarrow y \in \mathbb{R}$$
 $\widehat{C}_n(X_{n+1})$

Conformal prediction with rounding (informal version):

Problems to solve:

- Theory to guarantee coverage rate 1α ?
- Avoid wider intervals due to discretized grid?

Why do we lose the coverage guarantee?

- If only fit $\widehat{\mu}$ on $(X_1, Y_1), \dots, (X_n, Y_n), (X_{n+1}, y)$ for y in a grid...
- Equivalent to: fit $\widehat{\mu}$ on $(X_1,Y_1),\dots,(X_n,Y_n),(X_{n+1},[Y_{n+1}])$ $Y_{n+1} \text{ rounded to grid}$

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To maintain exchangeability:

- Fit $\widehat{\mu}$ on $(X_1, [Y_1]), \dots, (X_n, [Y_n]), (X_{n+1}, y)$
- Residuals $R_i = [Y_i] \widehat{\mu}(X_i), R_{n+1} = y \widehat{\mu}(X_{n+1})$

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- Residuals $R_i = [Y_i] \widehat{\mu}(X_i), R_{n+1} = \mathbf{y} \widehat{\mu}(X_{n+1})$
- Prediction set on the grid:

$$[\widehat{C}_n](X_{n+1}) = \{y \text{ values on the grid such that } |R_{n+1}| \leq ...\}$$

• Final prediction set:

$$\widehat{C}_n(X_{n+1}) = \{ y \in \mathbb{R} : y \text{ rounds to any value in } [\widehat{C}_n](X_{n+1}) \}$$

Alternate approach—discretize the model, not the data

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• Run a discretized algorithm for model fitting:

$$(X_1,Y_1),\ldots,(X,y) \stackrel{[\]}{\longrightarrow} (X_1,[Y_1]),\ldots,(X,[y]) \stackrel{\text{fit model }\widehat{\mu}}{\longrightarrow} [\widehat{\mu}]$$

Calculate residuals

$$R_i = Y_i - [\widehat{\mu}](X_i), \ i = 1, \dots, n, \quad R_{n+1} = y - [\widehat{\mu}](X)$$

• Check if $|R_{n+1}| \leq \left[(1-\alpha) \text{ quantile of } |R_1|, \ldots, |R_n|, |R_{n+1}| \right]$

- Model is fitted once for each possible value of $[y] \in grid$
- Residual is calculated for each possible value of $y \in \mathbb{R}$

If we run conformal prediction w/ discretized data:

- ✓ Theory to guarantee coverage rate 1α ?
- ? Avoid wider intervals due to discretized grid?

If we run conformal prediction w/ discretized model:

- ✓ Theory to guarantee coverage rate 1α ?
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Holdout methods vs full conformal (lose sample size) (high computational cost)
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Holdout methods vs full conformal (lose sample size) (high computational cost)

To avoid this tradeoff, can we use cross-validation?

Split data into $S_1 \cup \cdots \cup S_K$

For each $i \in S_k$, $R_i^{CV} = |Y_i - \widehat{\mu}_{-S_k}(X_i)|$

$$C(X_{n+1}) = \widehat{\mu}(X_{n+1}) \pm Q_{1-\alpha}(R_i^{CV})$$

- Computational cost: K + 1 regressions
- Problem: theory from holdout setting no longer holds

Jackknife a.k.a. leave-one-out cross-validation
$$(K=n)$$
 Residuals $R_i^{\mathsf{LOO}} = |Y_i - \widehat{\mu}_{-i}(X_i)|$
$$C(X_{n+1}) = \widehat{\mu}(X_{n+1}) \pm \mathsf{Q}_{1-\alpha}\Big(R_i^{\mathsf{LOO}}\Big)$$

 $^{^{13}}$ Steinberger & Leeb 2018, Conditional predictive inference for high-dimensional stable algorithms

Jackknife a.k.a. leave-one-out cross-validation (K = n) Residuals $R_i^{\text{LOO}} = |Y_i - \widehat{\mu}_{-i}(X_i)|$

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Predictive coverage under algorithmic stability assumption:¹³

$$\mathbb{P}\left\{|\widehat{\mu}(X_{n+1}) - \widehat{\mu}_{-i}(X_{n+1})| \le \epsilon\right\} \ge 1 - \nu$$

 $^{^{13}}$ Steinberger & Leeb 2018, Conditional predictive inference for high-dimensional stable algorithms

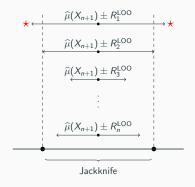
Jackknife+:14

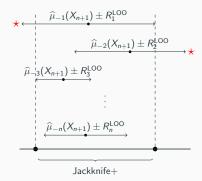
$$C(X_{n+1}) = \left[Q_{\alpha} \left(\widehat{\mu}_{-i}(X_{n+1}) - R_i^{\mathsf{LOO}} \right), \ Q_{1-\alpha} \left(\widehat{\mu}_{-i}(X_{n+1}) + R_i^{\mathsf{LOO}} \right) \right]$$

Compare to jackknife:

$$C(X_{n+1}) = \left[Q_{\alpha} \left(\widehat{\mu}(X_{n+1}) - R_i^{\mathsf{LOO}} \right), \ Q_{1-\alpha} \left(\widehat{\mu}(X_{n+1}) + R_i^{\mathsf{LOO}} \right) \right]$$

¹⁴B., Candès, Ramdas, Tibshirani 2019





Extension to CV+:15

$$C(X_{n+1}) = \left[Q_{\alpha} \Big(\widehat{\mu}_{-S_k}(X_{n+1}) - R_i^{\text{CV}} \Big), \ Q_{1-\alpha} \Big(\widehat{\mu}_{-S_k}(X_{n+1}) + R_i^{\text{CV}} \Big) \right]$$

Closely related to the cross-conformal prediction method¹⁶

¹⁵B., Candès, Ramdas, Tibshirani 2019

 $^{^{16}}$ Vovk 2015, Vovk et al 2018

Theorem: For any distrib. P and any A, jackknife+ satisfies

$$\mathbb{P}\{Y_{n+1} \in C(X_{n+1})\} \ge 1 - 2\alpha.$$

Theorem: For any distrib. P and any A, jackknife+ satisfies

$$\mathbb{P}\left\{Y_{n+1}\in C(X_{n+1})\right\}\geq 1-2\alpha.$$

Theorem: For any distrib. P and any A, K-fold CV+ satisfies

$$\mathbb{P}\left\{Y_{n+1} \in C(X_{n+1})\right\} \geq \begin{cases} 1 - 2\alpha - 1/K & (\text{new}) \\ 1 - 2\alpha - 2K/n & (\text{Vovk et al}) \end{cases}$$

$$\Rightarrow \quad \geq 1 - 2\alpha - \sqrt{2/n}.$$

Proof idea: embed jackknife+ into a larger exchangeable problem

Proof idea: embed jackknife+ into a larger exchangeable problem

- Exchangeable data $\{(X_1, Y_1), \dots, (X_{n+1}, Y_{n+1})\}$
- ullet $\binom{n+1}{2}$ leave-two-out regressions: $\widetilde{\mu}_{-\{i,j\}}$ for $1 \leq i,j \leq n+1$
- We can observe n of these, i.e., $\widetilde{\mu}_{-\{i,n+1\}} = \widehat{\mu}_{-i}$ for $1 \leq i \leq n$

Online / streaming / time series data

Conformal prediction can also be applied to an online setting...

• If data points are iid, conformal p-values are valid (and \bot) at each time t \Rightarrow can use conformal to predict / to test for changepoints¹⁷

¹⁷Vovk, Gammerman, Shafer 2005, Algorithmic Learning in a Random World

¹⁸Gibbs & Candès 2021, Adaptive conformal inference under distribution shift

¹⁹Xu & Xie 2021, Conformal prediction interval for dynamic time-series

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- If data points are iid,
 conformal p-values are valid (and ⊥) at each time t
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- If data points are independent but there is distribution drift,
 can bound cumulative error¹⁸

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- If data points are independent but there is distribution drift,
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- If data points form a time series,
 some standard assumptions ensure asymptotic coverage¹⁹

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Looking ahead...

Lecture 2 will discuss:

- Limitations of the predictive validity guarantee
- Challenges for inference on problems other than prediction
- Is "distribution free" the right framework?