Optimization of Advanced Marko Chain Monte Carlo Methods with Applications to Biological Systems

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Agenda

- Overview
 - Markov Chains & Stepsizes
 - Effectiveness & Expected Squared Jump Distance
 - Thompson Sampling
 - Once more: Why not to tune the acceptance rate!
- Results
- Improvements & More
 - Acceptance Rate Tuning with Thompson Sampling
 - Tune higher-order Autocorrelation Lags
 - Gaussian Process Time Costs
 - Monotonicity of Thinning

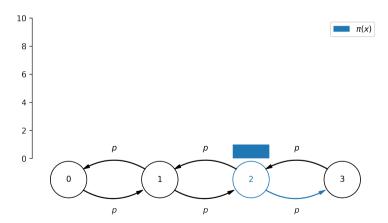
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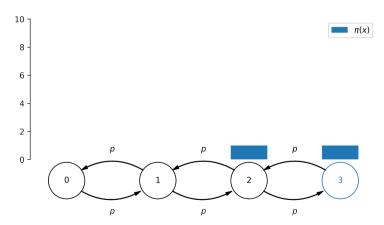
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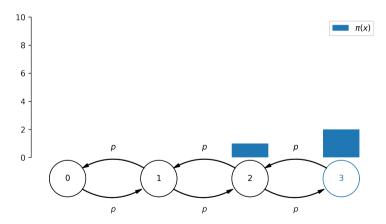
- A Markov chain is a stochastic process X_0, X_1, \ldots which takes values from some state space Θ and where the following properties hold:
 - $\mathbb{P}(X_0 = \theta_0) = \lambda(\theta_0)$
 - $\mathbb{P}(X_{t+1} = \theta_{t+1} | X_t = \theta_t, \dots, X_0 = \theta_0) = \mathbb{P}(X_{t+1} = \theta_{t+1} | X_t = \theta_t) =: P(\theta_t, \theta_{t+1})$

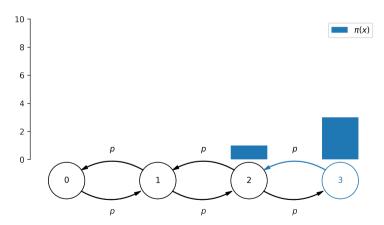
where λ is a distribution over Θ

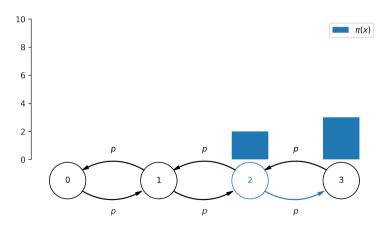
• An example:
$$\Theta = [0,3] \cap \mathbb{N}$$
 and $P = \begin{pmatrix} p & p \\ p & p \\ p & p \end{pmatrix}$ with $p = 0.5$

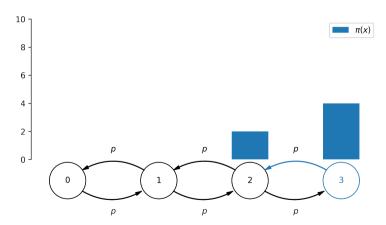


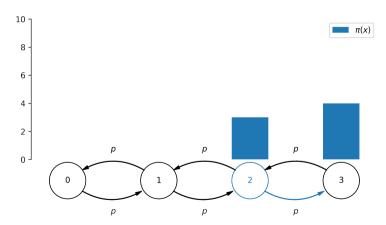


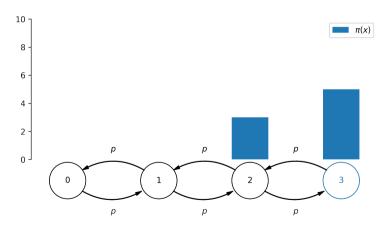


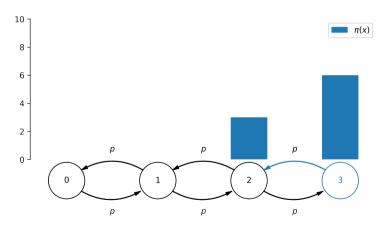


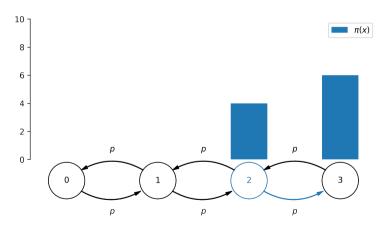


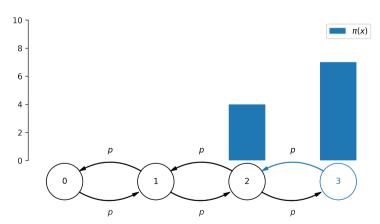


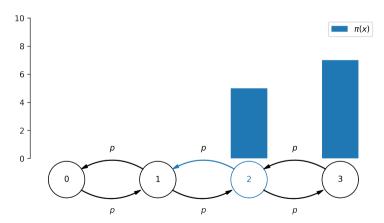


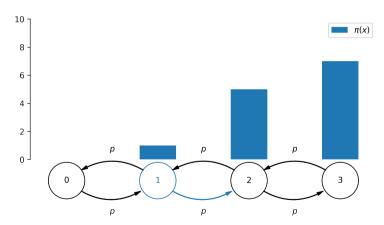


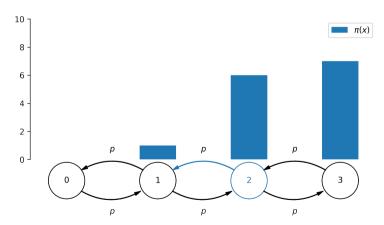


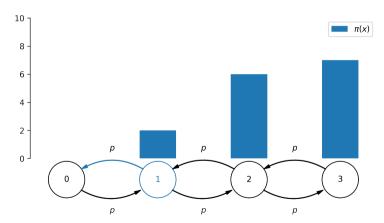




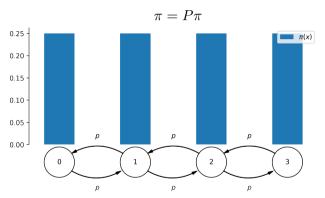








• At some point, we would reach the invariant or equilibrium distribution π



• Theory also tells us, that under suitable conditions (e.g finite Θ), we have that

$$\lim_{n \to \infty} (P^n)_i = \pi \quad \forall i \in \Theta$$

- We went from given transition matrix P to the **unique** invariant distribution π
- We can also go from given invariant distribution π (or measure!) to **non-unique** transition probabilities P!
- Given π , we have

$$\pi_i = \sum_j \pi_j P_{ji}$$

$$\sum_j P_{ij} \pi_i = \sum_j \pi_j P_{ji}$$

• Now,

$$\underbrace{\pi_i P_{ij} = \pi_j P_{ji}}_{detailed \ balance \ condition} \implies \sum_j P_{ij} \pi_i = \sum_j \pi_j P_{ji}$$

• First, rearrange ...

$$\pi_i P_{ij} = \pi_j P_{ji} \implies \frac{P_{ij}}{P_{ji}} = \frac{\pi_j}{\pi_i}$$

• ... and finally choose

$$P_{ij} = \min\left\{1, \frac{\pi_j}{\pi_i}\right\}$$

the famous Metropolis Filter

... & Stepsizes

• In continuous space, we split the transition probability $P(\theta, \theta')$ into proposal $q(\theta, \theta^*)$ and acceptance probability $\alpha(\theta, \theta^*)$ and use

$$\alpha(\theta, \theta^*) = \min \left\{ 1, \frac{\pi(\theta^*)q(\theta^*, \theta)}{\pi(\theta)q(\theta, \theta^*)} \right\}$$

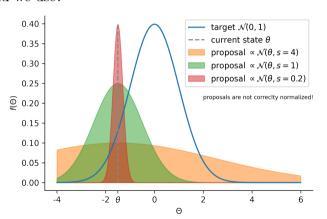
 \bullet q usually depends on some parameters, e.g. using

$$q(\theta, \theta^*) = q_s(\theta^*|\theta) \sim \mathcal{N}(\theta, s),$$

what *stepsize* s should we use?

... & Stepsizes

• What s should we use?



• Again: Transition kernels to a given invariant distribution are **not unique**

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Effectiveness ...

• Given samples $\theta_0, \ldots, \theta_n$, the sample average is $\bar{\theta} = \frac{1}{n} \sum_{i=0}^n \theta_i$ and the squared statistical error in \bar{x} is the variance

$$\begin{aligned} \operatorname{Var}[\bar{\theta}] &= \mathbb{E}[\bar{\theta}^2] - \mathbb{E}[\bar{\theta}]^2 \\ &= \frac{\sigma^2}{n^2} \left(n + \sum_{t=1}^{n-1} (n-t)\rho_t \right) \quad \text{with} \quad \rho_t = \frac{\mathbb{E}[\theta_0 \theta_t] - \mu^2}{\sigma^2} \\ &= \frac{\sigma^2}{n} \left(1 + \sum_{t=1}^{n-1} \left(1 - \frac{t}{n} \right) \rho_t \right) < \frac{\sigma^2}{n} \underbrace{\left(1 + \sum_{t=1}^{n-1} \rho_t \right)} \end{aligned}$$

• From this we derive the effective sample size as

$$\operatorname{Var}[\bar{x}] = \frac{\sigma^2}{n_{\text{eff}}}$$
 with $n_{\text{eff}} = \frac{n}{\tau}$

Effectiveness ...

- $\bullet\,$ The effective sample size is commonly used as a measure of quality for Markov chains
 - Note that $\tau = \tau((\theta_t)_{t=0}^n)$ and $(\theta_t)_{t=0}^n = ((\theta_t)_{t=0}^n)(s)$
- We now increase effectiveness by reducing τ , since $n_{\text{eff}} = \frac{n}{\tau}$
- Autocorrelations are usually assumed to be decreasing in time: $\rho_t \geq \rho_{t+k}$, for $k \geq 1$
- Thus, decrease autocorrelation for small lags!

... & Expected Squared Jump Distance

• We define

$$\operatorname{ESJD}(\theta) := \mathbb{E}[(\theta_{t+1} - \theta_t)^2] = \underbrace{\mathbb{E}[\theta_{t+1}^2]}_{=\mathbb{E}[\theta^2]} - 2\mathbb{E}[\theta_{t+1}\theta_t] + \underbrace{\mathbb{E}[\theta_t^2]}_{=\mathbb{E}[\theta^2]}$$

$$= 2(\mathbb{E}[\theta^2] - \mathbb{E}[\theta_{t+1}\theta_t] + \mu^2 - \mu^2)$$

$$= 2(\sigma^2 - (\mathbb{E}[\theta_{t+1}\theta_t] - \mu^2))$$

$$= 2\sigma^2 (1 - \frac{\mathbb{E}[\theta_{t+1}\theta_t] - \mu^2}{\sigma^2})$$

$$= 2\sigma^2 (1 - \rho_1)$$

• Since σ^2 is constant, maximizing the ESJD reduces the lag-1 autocorrelation, increasing efficiency.

... & Expected Squared Jump Distance

- However, the ESJD is a function of the **unknown** true target distribution, so we have to **estimate** it.
- We have a guess, how the function $\mathrm{ESJD}(s)$ looks like, but in general, it remains a black box function.

 \implies Thompson Sampling

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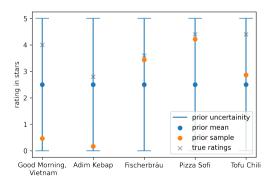
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Thompson Sampling

- A Bayesian optimization approach to optimize an unknown **black box** objective function:
 - We can evaluate it pointwise, but that's it!
- Thus, we have to explore the function and search for the optimum simultaneously
- Thompson Sampling:
 - Assume a distribution over the possible shapes of the black box function
 - Sample that distribution and evaluate the black box function, where the sampled function is maximized
 - Update the distribution over the possible function shapes based on the observed evaluation

Going to the restaurant!

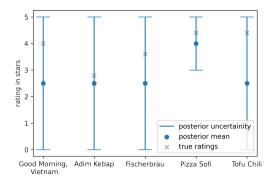
- You're coming to Vienna and there exist a total of 5 different restaurants, the internet has not yet been invented, I never went out in 8 years here
- Any restaurant might be good or bad, so we choose the first one uniformly ⇒ We already assumed a (uniform) distribution over the restaurants quality!



Going to the restaurant!

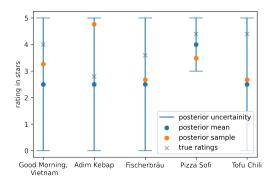
- We go to Pizza Sofi and rate it with 4 stars, but we only had a fraction of their menu
- Would we have rated the same, if we had eaten something else? Did the cook have an exceptional good day?

⇒ There is function noise and uncertainty left in our rating!



Going to the restaurant!

- Once more, we want to go out eating. What restaurant to choose?
- "Sample" from your posterior belief about the quality of the restaurants and go to the one, where your sample tells you, that you will have the best experience:
- Afterwards, update your beliefs based on your experience.



Dinner with William Thompson and Andrey Markov

- Our restaurants are the stepsizes s, that we test.
- Our friends, with whom we go to the restaurant, is the number of chains, that we use.
 - ⇒ Each friend rates the restaurant differently, each chain measures another value for the objective function
- \bullet The amount of money we spend, is the number of samples, we use to test the stepsize
- ⇒ We never take everything, that the restaurant has to offer. But taking more than just a salad is likely to reveal more about the kitchen's qualities, that is reduce the error in the estimate of the objective function

More formal Thompson Sampling

- We model our distribution over the possible black box functions as a Gaussian Process (GP)
- The algorithm then is:

```
for t \in 0, ..., m do
\begin{vmatrix}
\operatorname{sample} \hat{f}_t \sim GP(\mu_t, \Sigma_t) \\
\operatorname{let} x_t := \operatorname{arg} \max \hat{f}_t(x) \\
\operatorname{let} y_t := f(x_t) \\
\operatorname{let} D_{t+1} := D_t \cup \{(x_t, y_t)\} \\
\operatorname{update} \mu_{t+1} \operatorname{using} D_{t+1} \\
\operatorname{update} \Sigma_{t+1} \operatorname{using} D_{t+1}
\end{vmatrix}
```

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Once more: Why not to tune the acceptance rate!

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Results

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Acceptance Rate Tuning with Thompson Sampling

• Formally, the acceptance rate is

$$\alpha = \int_{\Theta^2} \underbrace{\pi(\theta)}_{!!} q(\theta, \theta^*) \alpha(\theta, \theta^*) d\theta d\theta^*$$

and thus it is a function of the **unknown** true target distribution

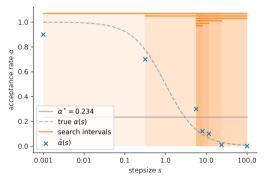
• Once more, we can only **estimate** it:

$$\hat{\alpha}(s) = \frac{1}{n} \sum_{i=0}^{n} \mathbb{1}_{\{\theta_{t+1} \neq \theta_t\}} = \frac{\text{\#accepted moves}}{n}$$

- As a function of the stepsize the acceptance rate is **strictly monotonous**
- Use binary search to find the intersection of $\alpha(s)$ with some desired target value α^* ?

Acceptance Rate Tuning with Thompson Sampling

• Use binary search to find the intersection of $\alpha(s)$ with some desired target value α^* ?



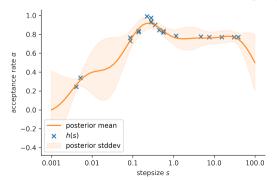
- No! This fails with noisy function evaluations which the estimator $\hat{\alpha}(s)$ gives us!
 - \rightarrow too deterministic! Instead, use Thompson Sampling!

Acceptance Rate Tuning with Thompson Sampling

• Given α^* and $\hat{\alpha}(s)$, define the score function:

$$h(s) := 1 - \Delta_{\alpha} := 1 - |\alpha^* - \hat{\alpha}(s)|$$

which is maximized when $\hat{\alpha}(s) = \alpha^*$ and takes values in [0, 1]



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Tune higher-order Autocorrelation Lags

• Recap the definition $\mathrm{ESJD}(\theta) = \mathbb{E}[(\theta_{t+1} - \theta_t)^2]$ and observe, that

$$\mathbb{E}[(\theta_{t+k} - \theta_t)^2] = \underbrace{\mathbb{E}[\theta_{t+k}^2]}_{=\mathbb{E}[\theta^2]} - 2\mathbb{E}[\theta_{t+k}\theta_t] + \underbrace{\mathbb{E}[\theta_t^2]}_{=\mathbb{E}[\theta^2]}$$

$$= 2(\mathbb{E}[\theta^2] - \mathbb{E}[\theta_{t+k}\theta_t] + \mu^2 - \mu^2)$$

$$= 2(\sigma^2 - (\mathbb{E}[\theta_{t+k}\theta_t] - \mu^2))$$

$$= 2\sigma^2 (1 - \frac{\mathbb{E}[\theta_{t+k}\theta_t] - \mu^2}{\sigma^2})$$

$$= 2\sigma^2 (1 - \rho_k)$$

Tune higher-order Autocorrelation Lags

• So we can minimize the lag-1 and lag-2 autocorrelation by maximizing the 1 and 2-jump ESJD:

$$\mathbb{E}[(\theta_{t+1} - \theta_t)^2] + \mathbb{E}[(\theta_{t+2} - \theta_t)^2] = 2\sigma^2(1 - \rho_1) + 2\sigma^2(1 - \rho_2)$$
$$= 2\sigma^2(2 - \rho_1 - \rho_2)$$

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Gaussian Process Time Costs

• Recap Thompson Sampling:

```
for t \in 0, \ldots, m do
```

end

with k = |X| the search grid, n test samples, c parallel chains and m iterations rounds.

• Overall, we have $O((mc)^4)^*$

^{*}oof

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Monotonicity of Thinning

Thanks!

53 / 53

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