Machine Learning Foundations

Chapter 6: Probability

Harish Guruprasad Ramaswamy
IIT Madras

Outline for Chapter 6: Probability

- 6.1 : Discrete Random Variables
- 6.2 : Continuous Random Variables
- 6.3 : Maximum Likelihood and other advanced topics

Outline for Chapter 6: Probability

- 6.1 : Discrete Random Variables
- 6.2 : Continuous Random Variables
- 6.3 : Advanced topics
 - 1. Bivariate and Multivariate normal
 - 2. Estimation of parameters using ML
 - 3. Gaussian Mixture Models and Expectation Maximisation
 - 4. Law of Large Numbers

Standard Normal Vector

$$Z, \Lambda N(0,1), Z_2 \Lambda N(0,1), \dots, Z_d \Lambda N(0,1)$$

$$Z := \begin{bmatrix} Z_1 \\ \vdots \\ Z_d \end{bmatrix}$$

$$f_{Z}(y): \prod_{i=1}^{J} \frac{1}{J_{ZT}} \exp\left(-\frac{1}{2}g_{i}^{2}\right)$$

$$f_{Z}(y): \frac{1}{J_{ZT}}dA \cdot \exp\left(-\frac{1}{2}\|g\|^{2}\right)$$

Let
$$PE[\Gamma',1] \cdot X_1 = Z_1 \quad j \quad X_2 = PZ_1 + \sqrt{1-P^2} Z_2$$

$$X = \begin{bmatrix} 1 & 0 & 0 \\ P & \sqrt{1-P^2} & 2 & j \end{bmatrix} \quad Z = \begin{bmatrix} 1 & 0 & 0 \\ P & \sqrt{1-P^2} & \sqrt{1-P^2} & 2 \\ Det(A) = \sqrt{1-P^2} & j \quad Det(A^{-1}) = \sqrt{1-P^2} \end{bmatrix}$$

$$Det(A) = \sqrt{1-P^2} \quad j \quad Det(A^{-1}) = \sqrt{1-P^2}$$

$$\begin{split} & \mathbb{E}\left[\mathbf{X}_{1}\right] \cdot \mathbb{E}\left[\mathbf{X}_{2}\right] \cdot O \\ & \mathbb{Cov}\left[\mathbf{X}_{1}, \mathbf{X}_{2}\right] \cdot \mathbb{E}\left[\mathbf{X}_{1}, \mathbf{X}_{2}\right] \cdot \mathbb{E}\left[\mathbf{X}_{1}, \mathbf{X}_{2}\right] \cdot \mathbb{E}\left[\mathbf{X}_{1}, \mathbf{X}_{2}\right] \\ & : \mathbb{E}\left[\left.\mathbf{P}\mathbf{Z}_{1}\right|^{2} + \left[\mathbf{I}-\mathbf{P}^{2}\mathbf{Z}_{1}\mathbf{Z}_{2}\right] \cdot \mathbf{P}\mathbf{E}\left[\mathbf{Z}_{1}^{2}\right] \cdot \mathbf{P} \end{split}$$

$$Var[X,] . I$$

$$Vav[X,] : E[(PZ + \sqrt{1-\rho'}Z)'] = \rho' + (1-\rho') . I$$

$$\therefore E[X] : \begin{bmatrix} 0 \\ 0 \end{bmatrix}$$

$$Cov[X] : \begin{bmatrix} 1 & P \\ P & I \end{bmatrix} = AA^{T} ; Det(E) . 1-\rho'$$

$$\begin{bmatrix} cov[X] \end{bmatrix}^{-1} : \frac{1}{1-\rho'} \begin{bmatrix} 1 & -P \\ -P & I \end{bmatrix} ; Det(E') . \frac{1}{1-\rho'}$$

$$f_{\chi}(\chi) = f_{Z}(A^{T}\chi) \cdot | Det(A^{T})|$$

$$= \frac{1}{2\pi \sqrt{1-\rho^{2}}} \cdot exp\left(-\frac{1}{2} || A^{T}\chi ||^{2}\right)$$

$$= \frac{1}{2\pi \sqrt{1-\rho^{2}}} \cdot exp\left(-\frac{1}{2} || A^{T}\chi ||^{2}\right)$$

$$= \frac{1}{2\pi \sqrt{1-\rho^{2}}} \cdot exp\left(-\frac{1}{2} || A^{T}\chi ||^{2}\right)$$

$$= \frac{1}{2\pi \sqrt{1-\rho^{2}}} \cdot exp\left(-\frac{1}{2(1-\rho^{2})}[x, x_{2}][1-\rho][x_{1}]\right)$$

f
$$(x)$$
 = $\frac{1}{2}$ exp $\left(-\frac{1}{2} \|A^{2}x\|^{2}\right)$

$$f_{\chi}(\chi) : \frac{1}{2\pi\sqrt{1-\rho^2}} \cdot \exp\left(-\frac{1}{2}\|A^{2}\chi\|^{2}\right)$$

$$= \frac{1}{2\pi} \int_{1-\rho^2}^{\infty} e^{2x\rho} \left[-\frac{1}{2} \left(x_1^2 + \left(\frac{-\rho}{I_{1-\rho^2}} x_1 + \frac{1}{I_{1-\rho^2}} x_2 \right)^2 \right) \right]$$

$$= \frac{1}{2\pi \sqrt{1-e^2}} e^{2ip\left(-\frac{1}{2}\chi_i^2\right)} \cdot e^{2ip\left(-\frac{1}{2(1-e^2)}(\chi_2-e^2\chi_i^2)\right)}$$

$$= \frac{1}{\sqrt{2\pi}} e^{2ip\left(-\frac{1}{2}\chi_i^2\right)} \cdot \frac{1}{\sqrt{2\pi}\sqrt{1-e^2}} e^{2ip\left(-\frac{1}{2(1-e^2)}(\chi_2-e^2\chi_i^2)\right)}$$

Simple Linear Transform of 2D-Normal
$$f(\alpha)$$
: $\frac{1}{2} \exp(\frac{-1}{2}x_1^2) \cdot \frac{1}{2} \exp(\frac{-1}{2}x_2^2) \cdot \frac$

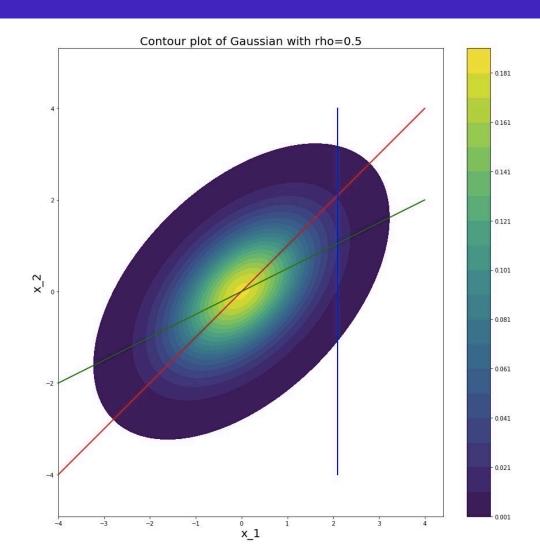
$$f_{\chi}(\chi) := \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}\chi_{1}^{2}\right) \cdot \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2(1-e^{2})}(\chi_{2}-e\chi_{1})^{2}\right)$$

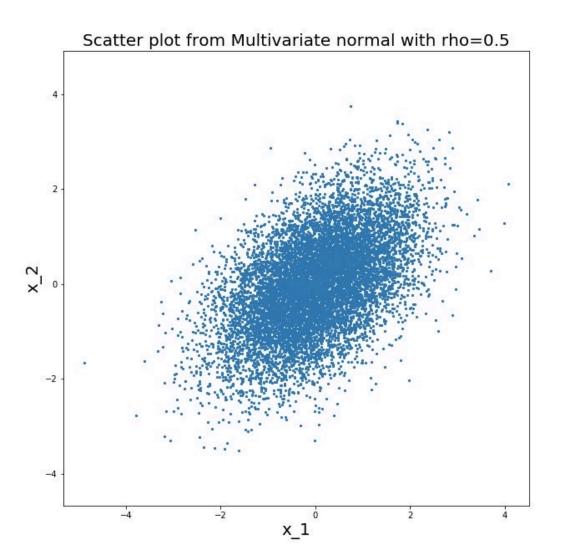
$$N(\chi_{1}|0,1) \qquad N(\chi_{2}|P\chi_{1},1-e^{2})$$

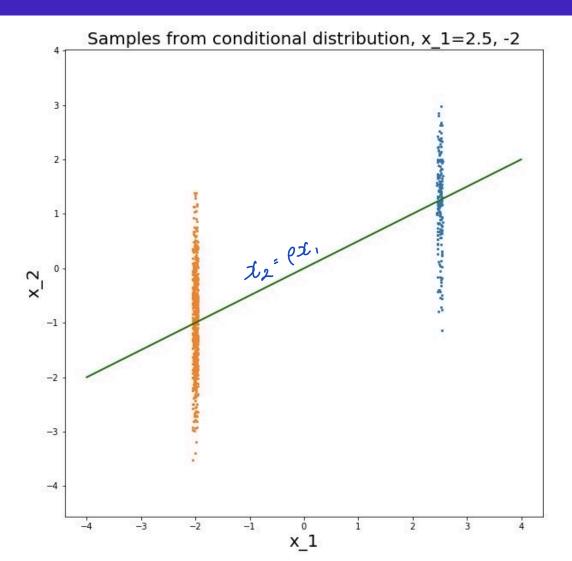
$$N(x, 10, 1)$$

$$N(x_2 | Px_1, 1-e^2)$$

$$f_{X_1}(x_1) \cdot f_{X_2}(x_2 | x_1)$$







The Bivariate Normal

$$x : AZ$$

$$f_{x}(x) = f_{z}(A^{-1}x) \cdot |Det(A^{-1})|$$
Let $\leq :AA^{T} = E[xx^{T}] = Cov[x]$
then $Det A^{T} : \int_{Det A} = \int_{Det E}$

$$f_{x}(x) = \frac{1}{2\pi \sqrt{Det(E)}} exp(-\frac{1}{2}x^{T}z^{T}x)$$

The Bivariate Normal

$$\begin{aligned}
\Xi &= A A^{T} &= \begin{bmatrix} a^{2} & \rho ab \\ \rho ab & b^{2} \end{bmatrix} \\
a, b &+ ve \\ \rho \in [-1, 1]
\end{aligned}$$

The Bivariate Normal

$$X : A = + \mu$$

$$Z = A^{-1}(X - \mu)$$

$$\Xi : A A^{T}$$

$$f_{X}(X) : \frac{1}{2\pi \sqrt{|\Xi|}} \exp(-\frac{1}{2}(X - \mu)^{T} \Xi'(X - \mu))$$

Multivariate Normal

$$X = AZ + \mu$$

$$Z = A^{-1}(X - \mu)$$

$$\Xi = AA^{T}$$

$$f_{X}(X) = \frac{1}{(2\pi)^{d_{X}}} \int_{|\Xi|} \exp(-\frac{1}{2}(X - \mu)^{T} \Xi'(X - \mu))$$

$$X = AZ + \mu$$

$$Z = A^{-1}(X - \mu)$$

$$Z = AA^{T}$$

$$(2\pi)^{d_{X}} \int_{|\Xi|} \exp(-\frac{1}{2}(X - \mu)^{T} \Xi'(X - \mu))$$

$$X = AZ + \mu$$

$$Z = A^{-1}(X - \mu)$$

$$Z = AA^{T}$$

Some Properties of the Multivariate Normal

Some Properties of the Multivariate Normal

Let
$$\times \times N(\mu, \leq)$$

i) $Y = \alpha^{T} \times \times N(\alpha^{T} \mu, \alpha^{T} \leq \alpha)$

(ii)
$$Y = A \times M (A \mu, A \not\equiv A^T)$$

$$C = C A^2 P a b T M = 0$$

(iii)

e.g.
$$Z = \begin{bmatrix} \alpha^2 & \rho \alpha b \\ \rho \alpha b & b^2 \end{bmatrix}$$
; $M = 0$; $A = \begin{bmatrix} \frac{1}{2} & 0 \\ 0 & \frac{1}{2} \end{bmatrix}$
Then $Y : AX \sim N(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} 1 & P \\ 0 & 1 \end{bmatrix})$

$$X_{i}, X_{j}$$
 are independent \iff $\leq_{i,j} : 0$