Regression Discontinuity Designs

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- Sharp Regression Discontinuity
 - Identification
 - Estimation
 - Example
 - Diagnostics

- Puzzy Regression Discontinuity
 - Identification and Estimation

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Regression Discontinuity Design (RDD)

- RDD: A fairly old idea (Thistlethwaite and Campbell, 1960)
- Recently experienced a renaissance because of the formal theoretical foundation given by the causal inference framework
- Applicable when treatment is assigned according to a *rule* based on another variable (called the forcing or running variable)
- Often useful for analysis in a "rule-based" world (administrative programs, elections, etc.)
- High internal validity: One of the few observational designs that reproduced an experimental benchmark (Cook and Wong 2008)
- Limited external validity: Effect is only identified for a small subpopulation

Sharp RDD: Basic Setup

- $D_i \in \{0,1\}$: Treatment
- X_i: Forcing variable that perfectly determines the value of D_i with cutpoint c

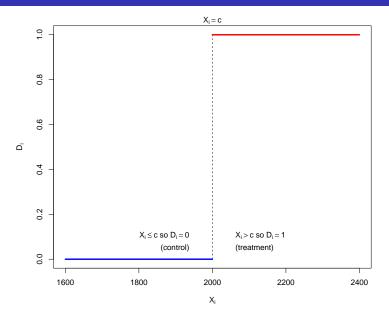
$$D_i = 1\{X_i > c\}$$
 or equivalently $D_i = \left\{ egin{array}{ll} 1 & ext{if } X_i > c \\ 0 & ext{if } X_i \leq c \end{array}
ight.$

- X_i may be correlated with $Y_i(0)$ and $Y_i(1)$, either directly or via other unobserved confounders
- Simply adjusting for X_i does not work because of lack of common support
- Basic intuition: Right at the cutpoint $X_i = c$, assignment to D_i may be as-if random

A Hypothetical Example: Effect of Scholarship

- Thistlethwaite and Campbell (1960) study the effects of college scholarships on later students' achievements
- Scholarships are given on the basis of whether or not a student's test score exceeds some threshold c
 - Treatment D_i is scholarship
 - Forcing variable X_i is SAT score with cutoff c
 - Outcome Y_i is subsequent earnings
 - $Y_i(0)$: potential earnings without the scholarship
 - $Y_i(1)$: potential earnings with the scholarship
- $Y_i(1)$ and $Y_i(0)$ are correlated with X_i : on average, students with higher SAT scores obtain higher earnings

Probability of Treatment in Sharp RDD



Identification of the Threshold Causal Effect

Key assumption: Continuity of average potential outcomes

$$\mathbb{E}[Y_i(d) \mid X_i = x]$$
 is continuous in x around $X_i = c$ for $d = 0, 1$

Causal estimand: Local ATE at the threshold

$$\tau_{SRD} = \mathbb{E}[Y_i(1) - Y_i(0) \mid X_i = c]$$

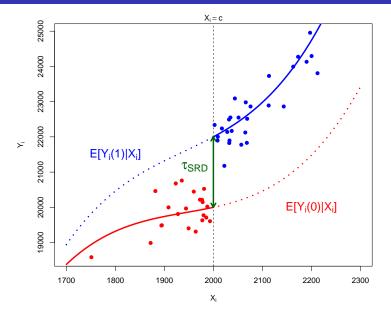
Identification result: If the continuity assumption holds, au_{SRD} is nonparametrically identified as

$$\tau_{SRD} = \lim_{x \downarrow c} \mathbb{E}[Y_i | X_i = x] - \lim_{x \uparrow c} \mathbb{E}[Y_i | X_i = x]$$

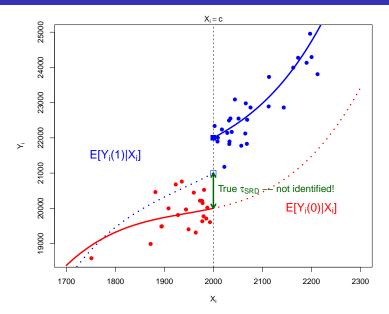
A "proof":

- D_i is wholly determined by X_i , so contional ignorability is trivially satisfied given X_i : $Y_i(1), Y_i(0) \perp \!\!\! \perp D_i \mid X_i$
- However, there is no common support, so conditioning on X_i in a usual way won't work.
- The continuity assumption allows us to do a tiny bit of extrapolation and compensate for the lack of common support at the threshold.

Graphical III ustration: Continuous $\mathbb{E}[Y_i(d) \mid X_i]$



Graphical Illustration: Discontinuous $\mathbb{E}[Y_i(d) \mid X_i]$



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Estimation of the LATE at the Threshold

- Trim the sample to a reasonable window around the threshold c (discontinuity sample)
 - $c h \le X_i \le c + h$, were h > 0 determines the width of the window
- **2** Recode forcing variable to deviations from threshold: $\tilde{X}_i = X_i c$
 - $\tilde{X}_i = 0$ if $X_i = c$
 - $\tilde{X}_i > 0$ if $X_i > c$ and thus $D_i = 1$
 - $\tilde{X}_i < 0$ if $X_i < c$ and thus $D_i = 0$
- **3** Decide on a model for $\mathbb{E}[Y_i|\tilde{X}_i]$:
 - linear, common slope for $\mathbb{E}[Y_i \mid \tilde{X}_i < 0]$ and $\mathbb{E}[Y_i | \tilde{X}_i > 0]$
 - linear, different slopes
 - non-linear
 - each model corresponds to a particular set of assumptions about the potential outcomes
 - always start with visual inspection (e.g. scatter plot with lowess) to check which model is plausible

Estimation with a Linear Model with a Common Slope

- Assumptions:

 - 2 Treatment effect, τ , does not depend on X_i

i.e.,

$$\mathbb{E}[Y_i(0)|X_i] = \alpha + \beta X_i$$
 and $\mathbb{E}[Y_i(1) - Y_i(0)|X_i] = \tau$

which implies

$$\mathbb{E}[Y_i(1)|X_i] = \tau + \mathbb{E}[Y_i(0)|X_i] = \tau + \alpha + \beta X_i$$

• Therefore, the model for the observed outcome should be:

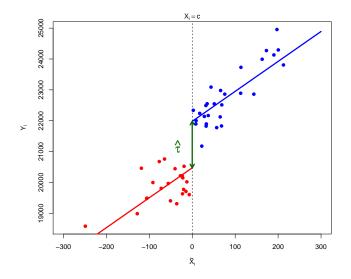
$$\mathbb{E}[Y_i|X_i, D_i] = D_i \cdot \mathbb{E}[Y_i(1)|X_i] + (1 - D_i) \cdot \mathbb{E}[Y_i(0)|X_i]$$

$$= \alpha + \tau D_i + \beta X_i$$

$$= \tilde{\alpha} + \tau D_i + \beta \tilde{X}_i \quad \text{(where } \tilde{\alpha} = \alpha + \beta c\text{)}$$

• So we just regress the observed outcome (Y_i) on D_i and \tilde{X}_i

Estimation with a Linear Model with a Common Slope



Estimation with a Linear Model with a Different Slope

- Assumptions:
 - ① $\mathbb{E}[Y_i(0)|X_i=x]$ and $\mathbb{E}[Y_i(1)|X_i=x]$ are both linear in x
 - ② But we now allow treatment effect to vary with X_i i.e.,

$$\mathbb{E}[Y_i(0)|X_i] = \alpha_0 + \beta_0 X_i \quad \text{and} \quad \mathbb{E}[Y_i(1)|X_i] = \alpha_1 + \beta_1 X_i$$

such that

$$\mathbb{E}[Y_i(1) - Y_i(0)|X_i] = (\alpha_1 - \alpha_0) + (\beta_1 - \beta_0)X_i$$

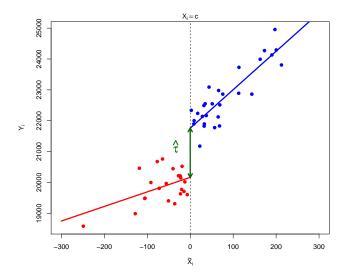
• The observed outcome model is therefore:

$$\begin{split} \mathbb{E}[Y_i|X_i,D_i] &= D_i \cdot \mathbb{E}[Y_i(1)|X_i] + (1-D_i) \cdot \mathbb{E}[Y_i(0)|X_i] \\ &= \alpha_0 + \beta_0 X_i + (\alpha_1 - \alpha_0)D_i + (\beta_1 - \beta_0)D_i X_i \\ &= (\alpha_0 + \beta_0 c) + \beta_0 \tilde{X}_i \\ &+ \{(\alpha_1 - \alpha_0) + (\beta_1 - \beta_0)c\}D_i + (\beta_1 - \beta_0)D_i \tilde{X}_i \\ &\equiv \tilde{\alpha} + \beta_0 \tilde{X}_i + \tau D_i + \tilde{\beta} D_i \tilde{X}_i \end{split}$$

Note that $\tau = \mathbb{E}[Y_i(1) - Y_i(0)|X_i = c]$, LATE at the threshold

• So, regress Y_i on \tilde{X}_i , D_i and the interaction $D_i\tilde{X}_i$

Estimation with a Linear Model with a Different Slope



Estimation with a Nonlinear Model

- Assumptions:
 - ① $\mathbb{E}[Y_i(0)|X_i=x]$ and $\mathbb{E}[Y_i(1)|X_i=x]$ are now allowed to be non-linear in X_i , but must be correctly specified
 - 2 Treatment effect is allowed to vary across X_i
- Include quadratic, cubic, etc. terms in \widetilde{X}_i and their interactions with D_i in the equation
- The specification with quadratic terms:

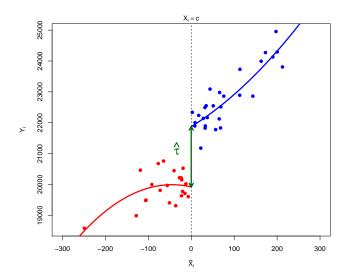
$$\mathbb{E}[Y_i|X_i,D_i] = \gamma_0 + \gamma_1 \tilde{X}_i + \gamma_2 \tilde{X}_i^2 + \tau D_i + \alpha_1 \tilde{X}_i D_i + \alpha_2 \tilde{X}_i^2 D_i$$

The specification with cubic terms is

$$\mathbb{E}[Y_i|X_i, D_i] = \gamma_0 + \gamma_1 \tilde{X}_i + \gamma_2 \tilde{X}_i^2 + \gamma_3 \tilde{X}_i^3 + \alpha_0 D_i + \tau \tilde{X}_i D_i + \alpha_2 \tilde{X}_i^2 D_i + \alpha_3 \tilde{X}_i^3 D_i$$

• In both cases, the coefficient on D_i corresponds to the LATE at the threshold: $\tau = \mathbb{E}[Y_i(1) - Y_i(0)|X_i = c]$

Estimation with a Nonlinear Model



Model Selection and Choice of Bandwidth in RDD

- How should we pick the "right" model and bandwidth?
- No ex-ante correct answer usually, but several data-driven procedures are available

Model choice:

- A tradeoff between bias and variance
- Standard practice: Use different specs and show robustness
- Local linear regression with a kernel smoother is a popular choice

Bandwidth selection:

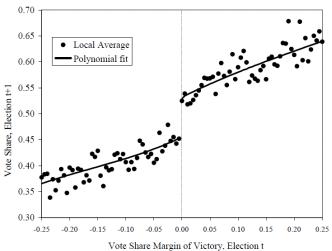
- Imbens-Kalyanaraman (IK) algorithm: Pick h that minimizes (a first-order approximation of) the MSE in $\hat{\tau}_{SRD}$
- Cross-validation: See Imbens and Lemieux (2008) for details

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Example: Party Incumbency Advantage

Figure IVa: Democrat Party's Vote Share in Election t+1, by Margin of Victory in Election t: local averages and parametric fit



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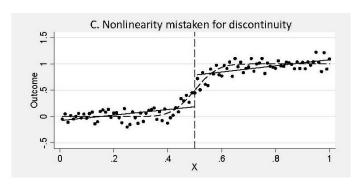
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Falsification Checks

Diagnoze the robustness of your results via falsification checks:

- Sensitivity: Are results sensitive to alternative specifications?
- ② Balance checks: Does any covariate Z_i jump at the threshold?
- Check if jumps occur at placebo thresholds c*?
- Sorting: Do units sort around the threshold?

Sensitivity to Specification



- RDD requires specification of the functional form and bandwidth
- Misspecification of either can lead to a spurious jump
- Take care not to confuse a nonlinear relation with a discontinuity!
- More flexibility (e.g. polynomials) reduces bias but increases variance
- Check sensitivity to size of bandwidth h

Balance Checks: Covariates as Placebo Outcomes

- Test for comparability of agents around the cutoff:
 - Visual tests: Plot $\mathbb{E}[Z_i|X_i,D_i]$ and look for jumps
 - Relation between covariates and treatment should be smooth around threshold
 - Use Z_i as a placebo outcome and see if there is inbalance:

$$\mathbb{E}[Z_i|X_i,D_i] = \beta_0 + \beta_1 \tilde{X}_i + \tau_z D_i + \beta_3 \tilde{X}_i D_i$$

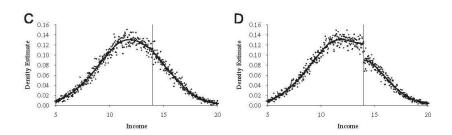
- $\tau_z = 0$ if Z_i is balanced at the threshold
- Discontinuity in Z_i indicates possible evidence of discontinuous $\mathbb{E}[Y_i(d) \mid X_i = x]$, violating the key assumption
- Inbalance can be addressed by incorporating Z_i in the analysis:
 - Use Z_i as an additional covariate in the model
 - Alternatively, regress Y_i on Z_i and use the residuals in the model, instead of Y_i itself
- Balance checks address only observables, not unobservables

Sorting Around the Threshold

- Agents' behavior can invalidate the continuity assumption:
 - Agents may exercise control over their values of X_i to fall on the beneficial side of the threshold
 - Administrators may strategically choose what X_i to use or which threshold to use
 - Such sorting of agents invalidate the continuity assumption
- When this occurs, distribution of X_i will discontinuously change at the threshold
- Diagnostics:
 - Visual inspection of Histograms (make sure no bin overlaps with the threshold!)
 - Formal tests (e.g. McCrary 2008)
- A related problem: Other treatments assigned by the exact same X_i and c (e.g. geographic boundary)
 - \longrightarrow Effect needs to be interpreted as a composite treatment effect

Example: Job Training Program

- ullet Beneficial job training program offered to agents with income < c
- Concern: People may withhold labor to lower their income just below the cutoff to gain access to the program



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Generalization to Imperfect Thresholds

- Threshold may not perfectly determine treatment exposure, but it
 may still create a discontinuity in the probability of treatment
- Hypothetical example: SAT and scholarship
 - Scholarship may not be wholly determined by SAT scores X_i , but the rule specifies a threshold $X_i = c$ to become eligible
 - Students above c get a chance of receiving the scholarship, but may not actually get it
 - Students below *c* is excluded from consideration, perhaps except ones who are otherwise remarkable
- That is, an encouragement to receive treatment is discontinuously determined by X_i at the threshold c
- Fuzzy RDD can therefore be thought of an instrumental variable version of an RDD

Fuzzy RDD: Setup and Identification

- $Z_i \in \{0,1\}$: Encouragement
- X_i : Forcing variable that perfectly determines the value of Z_i with cutpoint c

$$Z_i = 1\{X_i > c\}$$
 or equivalently $Z_i = \left\{ egin{array}{ll} 1 & ext{if } X_i > c \ 0 & ext{if } X_i \leq c \end{array}
ight.$

Identification assumptions:

- Both $\mathbb{E}[D_i(z) \mid X_i = x]$ (potential treatment) and $\mathbb{E}[Y_i(z) \mid X_i = x]$ (potential outcome) are continuous in x around $X_i = c$ for z = 0, 1
- ullet IV assumptions: Monotonicity, exclusion restriction, relevance of Z_i

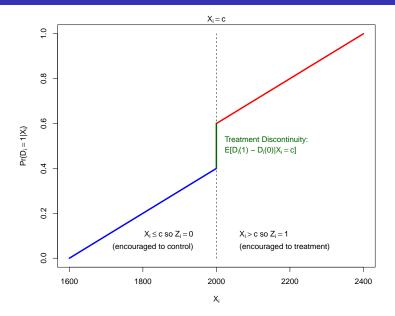
Causal estimand: Local ATE for compliers at the threshold

$$au_{FRD} = \mathbb{E}[Y_i(1) - Y_i(0) \mid \text{unit } i \text{ is a complier and } X_i = c]$$

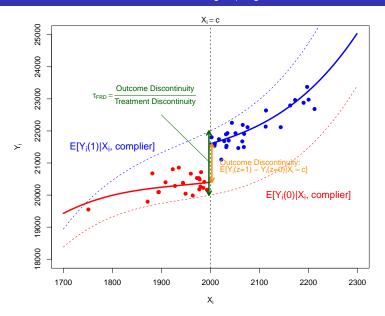
Identification result: Under the above assumptions, au_{FRD} is identified as

$$\tau_{FRD} = \frac{\lim_{x \downarrow c} \mathbb{E}[Y_i | X_i = x] - \lim_{x \uparrow c} \mathbb{E}[Y_i | X_i = x]}{\lim_{x \downarrow c} \mathbb{E}[D_i | X_i = x] - \lim_{x \uparrow c} \mathbb{E}[D_i | X_i = x]}$$

Probability of Treatment in Fuzzy RDD



Fuzzy RDD: Discontinuity in $\mathbb{E}[Y|X]$



Fuzzy RDD: Estimation

- lacktriangledown Trim the sample to a reasonable window above and below the threshold c (discontinuity sample)
- **②** Code the encouragement indicator: $Z_i = 1\{X_i > c\}$
- **3** Recode the forcing variable to deviation from c: $\tilde{X}_i = X_i c$
- Stimate the outcome model using two-stage least squares:

$$Y_i = \beta_0 + \beta_1 \tilde{X}_i + \beta_2 Z_i \tilde{X}_i + \tau D_i + \varepsilon_i,$$

where D_i is instrumented by Z_i

- ullet More flexible specifications can be used (e.g. polynomials of $ilde{X_i}$)
- **5** Then $\hat{ au}_{2SLS}$ consistently (but not unbiasedly) estimates au_{FRD}
 - In addition, it is also helpful to separately plot (and estimate) the outcome discontinuity and treatment discontinuity for interpretation
 - Usual diagnostics can be applied to check plausibility of assumptions