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Maximising the log-likelihood

$$\begin{aligned}\log p(X | \mu, \sigma^2) &= \log \prod_{i=1}^N \mathcal{N}(x_i | \mu, \sigma^2), \text{ to maximise} \\ &= \sum_{i=1}^N \log \left(\frac{1}{\sqrt{2\pi\sigma^2}} \exp -\frac{(x_i - \mu)^2}{2\sigma^2} \right) \\ &= -\left(\frac{N}{2}\right) \log 2\pi - \left(\frac{N}{2}\right) \log \sigma^2 - \frac{1}{2\sigma^2} \sum_{i=1}^N (x_i - \mu)^2\end{aligned}$$

How to maximise the log-likelihood?

Variational Calculus: $\frac{\partial \log\text{-likelihood}}{\partial \text{parameter}} = 0$

put it back in the log-likelihood, check for max/min

Better than 2nd deriv test, higher orders...

Parameters? μ, σ^2

$$\begin{aligned}1. \frac{\partial \log\text{-lh}}{\partial \mu} = 0 &\implies -\frac{1}{2\sigma^2} (2)(-1) \sum_{i=1}^N (x_i - \mu) = 0 \\ &\implies \sum_{i=1}^N x_i = \mu N \implies \mu_{ML} = \sum_{i=1}^N x_i / N \\ 2. \frac{\partial \log\text{-lh}}{\partial \sigma^2} = 0 &\implies -\frac{N}{2} \frac{1}{\sigma^2} + \frac{1}{2} \frac{1}{\sigma^4} \sum_{i=1}^N (x_i - \mu_{ML})^2 \implies \\ \frac{\sum_{i=1}^N (x_i - \mu_{ML})^2}{\sigma^2} &= N \implies \sigma^2 = \sum_{i=1}^N (x_i - \mu_{ML})^2 / N\end{aligned}$$

ML mean: sample mean, ML var: sample var

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Param Est: K D -D Gaussians

- **Given:** $\mathbf{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_N\}$, a set of N observations
- **Model:** K D -D Gaussians, means $\boldsymbol{\mu}_j$, Covs $\boldsymbol{\Sigma}_j$, Mixture coeffs π_j . 3 sets of parameters: $\boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}$
- **Assumptions:** Data points i.i.d. Independent: allows marginal prob multiplication without considering conditional dependence terms. Identically distributed: all from same model
- **Method:** $p(\mathbf{X} | \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})$: Likelihood, to maximise
- **Reasonable?** Find params which maximise the likelihood of getting these points, given our model
- $p(\mathbf{X} | \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \prod_{i=1}^N \sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)$, to max
- \equiv Maximise log-likelihood. Why? (John Napier)

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Maximising the log-likelihood

$$\log p(\mathbf{X} | \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{i=1}^N \log \sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)$$

$$\mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j) \triangleq \frac{1}{\sqrt{(2\pi)^D |\boldsymbol{\Sigma}_j|}} \exp -\frac{1}{2}(\mathbf{x}_i - \boldsymbol{\mu}_j)^T \boldsymbol{\Sigma}_j^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_j)$$

How to maximise the log-likelihood?

Variational Calculus: $\frac{\partial \log\text{-likelihood}}{\partial \text{parameter}} = 0$

$$\sum_{i=1}^N \frac{(+1)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} \frac{\partial \pi_j \mathcal{N}(\cdot)}{\partial \text{parameter}} = 0$$

- parameter #1: $\boldsymbol{\mu}_j$: $\frac{\partial \log\text{-lh}}{\partial \boldsymbol{\mu}_j} = 0 \implies$

$$\sum_{i=1}^N \frac{(+1) \pi_j \exp -\frac{1}{2}[(\mathbf{x}_i - \boldsymbol{\mu}_j)^T \boldsymbol{\Sigma}_j^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_j)]}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} \frac{2(-1)}{2} \boldsymbol{\Sigma}_j^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_j) = 0$$

$$\implies \frac{2}{2} \sum_{i=1}^N \left[\frac{\pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} \right] \boldsymbol{\Sigma}_j^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_j) = 0$$

$$\sum_{i=1}^N \gamma_j(\mathbf{x}_i) \cdot \mathbf{x}_i = \boldsymbol{\mu}_j \sum_{i=1}^N \gamma_j(\mathbf{x}_i) \implies \boldsymbol{\mu}_j = \frac{\sum_{i=1}^N \gamma_j(\mathbf{x}_i) \cdot \mathbf{x}_i}{\sum_{i=1}^N \gamma_j(\mathbf{x}_i)}$$

$$\implies \boldsymbol{\mu}_j = \frac{\sum_{i=1}^N \gamma_j(\mathbf{x}_i) \cdot \mathbf{x}_i}{N_j} \quad \text{prob (resp)-weighted mean}$$

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- parameter #2: $\Sigma_j = \frac{\sum_{i=1}^N \gamma_j(\mathbf{x}_i) \cdot (\mathbf{x}_i - \boldsymbol{\mu}_j)(\mathbf{x}_i - \boldsymbol{\mu}_j)^T}{N_j}$
probability (responsibility)-weighted Covariance

- parameter #3: $\pi_j: \frac{\partial \log \text{-lh}}{\partial \pi_j} = \sum_{i=1}^N \frac{(-1) \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \Sigma_j)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \Sigma_j)} = 0$
To modify the objective function, regularisation

$$E = \log p(\mathbf{X} | \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) + \lambda (\sum_{j=1}^K \pi_j - 1)$$

1. $\frac{\partial E}{\partial \boldsymbol{\mu}_j} = 0$: Doesn't affect the previous estimate
2. $\frac{\partial E}{\partial \Sigma_j} = 0$: Doesn't affect the previous estimate
3. $\frac{\partial E}{\partial \pi_j} = 0$: Hope to get some non-trivial solution
4. $\frac{\partial E}{\partial \lambda} = 0$: $\sum_{j=1}^K \pi_j - 1 = 0$, the constr to be imposed

For (3) above, $\sum_{i=1}^N \frac{(+1) \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \Sigma_j)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \Sigma_j)} + \lambda = 0$. Mult by π_j :

$$\begin{aligned} \sum_{i=1}^N \frac{\pi_j \mathcal{N}(\cdot)}{\sum_{j=1}^K \pi_j \mathcal{N}(\cdot)} &= -\lambda \pi_j \implies \sum_{i=1}^N \gamma_j(\mathbf{x}_i) = \lambda \pi_j \implies \\ -\sum_i (\sum_j \gamma_j(\mathbf{x}_i)) &= \lambda (\sum_j \pi_j) \implies \lambda (1) = -\sum_i (1) \implies \\ \lambda &= -N. \text{ Put back: } \pi_j \lambda = -N_j \implies \pi_j = N_j / N \end{aligned}$$



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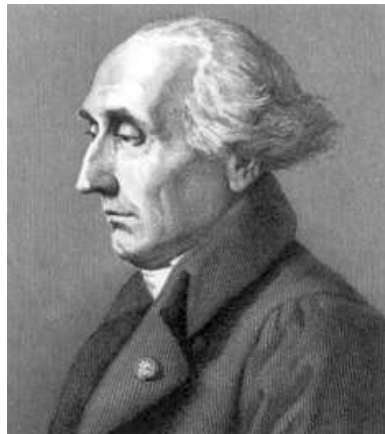
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Intertwined Histories



J.-L. Lagrange
[1736-1813]

https://upload.wikimedia.org/wikipedia/commons/1/19/Lagrange_portrait.jpg



A. Lavoisier
[1743-1794]

<https://upload.wikimedia.org/wikipedia/commons/4/44/Lavoisier-statue.jpg>



J.-B. J. Fourier
[1768-1830]

<https://upload.wikimedia.org/wikipedia/commons/a/aa/Fourier2.jpg>