

# Identifying Departures from the Fully Developed Speckle Hypothesis in Intensity SAR Data with Non-Parametric Estimation of the Entropy

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# Context and Problem Statement

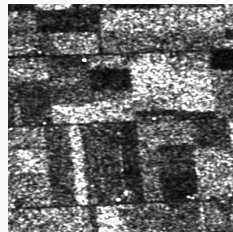
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## Fully Developed Speckle (FDS)

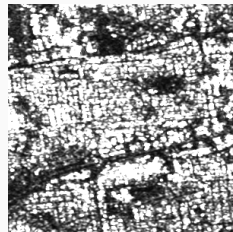
Homogeneous areas in SAR images are characterized by the lack of dominant scatterers, and the surface can be considered stationary. This is the case of the **fully developed hypothesis for the speckle**.

## Optimal model

The  $\mathcal{G}^0$  distribution is a suitable model for SAR intensity data. It includes the **Gamma law** as a limit case that results in the presence of fully developed speckle.



(a) FDS (agricultural area)



(b) Heterogeneous clutter (urban area)

# Problem and Proposal

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- **Applying**  $\mathcal{G}^0$  under fully developed speckle:
  - Tricky maximum likelihood estimation: Increased bias, flat likelihood, and numerical optimization may not converge.

## Our approach

We propose a test statistic to distinguish between fully developed speckle and heterogeneous clutter, based on the non-parametric entropy estimation.



## Model Setup

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## Intensity SAR Data

We denote  $Z \sim \Gamma_{\text{SAR}}(L, \mu)$  and  $Z \sim \mathcal{G}_I^0(\alpha, \gamma, L)$  to indicate that  $Z$  follows the distributions characterized by the respective probability density functions:

$$f_Z(z; L, \mu) = \frac{L^L}{\Gamma(L)\mu^L} z^{L-1} \exp\{-Lz/\mu\} \mathbb{1}_{\mathbb{R}_+}(z), \quad (1)$$

$$f_Z(z; \alpha, \gamma, L) = \frac{L^L \Gamma(L - \alpha)}{\gamma^\alpha \Gamma(-\alpha) \Gamma(L)} \cdot \frac{z^{L-1}}{(\gamma + Lz)^{L-\alpha}} \mathbb{1}_{\mathbb{R}_+}(z), \quad (2)$$

where  $L \geq 1$  is the number of looks,  $\Gamma(\cdot)$  is the gamma function, and  $\mathbb{1}_A(z)$  is the indicator function of the set  $A$ . In (1),  $\mu > 0$  is the mean; in (2)  $\gamma > 0$  is the scale,  $\alpha < -1$  measures the roughness.

## New Parametrization of $\mathcal{G}_I^0$

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Thus, the probability density function that characterize the  $G_I^0(\mu, \alpha, L)$  law is

$$f_Z(z; \mu, \alpha, L) = \frac{L^L \Gamma(L - \alpha)}{(-\mu(\alpha + 1))^\alpha \Gamma(-\alpha) \Gamma(L)} \cdot \frac{z^{L-1}}{(-\mu(\alpha + 1) + Lz)^{L-\alpha}} \mathbb{1}_{\mathbb{R}_+}(z).$$

## **Non-parametric Entropy Estimation Approach**

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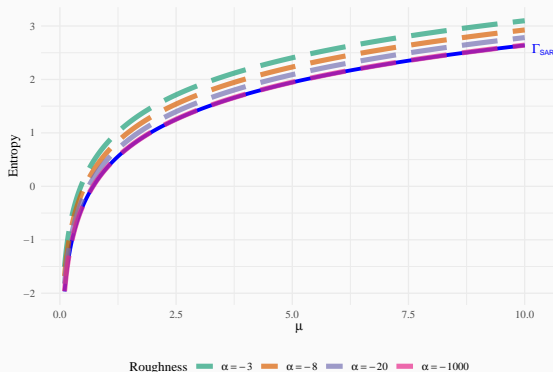
# Non-parametric Entropy Estimation Approach

## Shannon Entropy

$$H_{\Gamma_{\text{SAR}}}(L, \mu) = L - \ln L + \ln \Gamma(L) \\ + (1 - L)\psi^{(0)}(L) + \ln \mu,$$

$$H_{G_I^0}(\mu, \alpha, L) = H_{\Gamma_{\text{SAR}}} - \ln \Gamma(L - \alpha) \\ + (L - \alpha)\psi^{(0)}(L - \alpha) \\ - (1 - \alpha)\psi^{(0)}(-\alpha) + \ln(-1 - \alpha) \\ + \ln \Gamma(-\alpha) - L,$$

where  $\psi^{(0)}(\cdot)$  is the digamma function.



$H_{G_I^0}$  converges to the  $H_{\Gamma_{\text{SAR}}}$ ,  $L = 8$ .

# Non-parametric Entropy Estimation Approach

Vasicek introduced an non-parametric estimator in 1976:

$$\hat{H}_V(\mathbf{Z}) = \frac{1}{n} \sum_{i=1}^n \ln \left[ \frac{n}{2m} (Z_{(i+m)} - Z_{(i-m)}) \right],$$

where  $Z_{(i+m)} - Z_{(i-m)}$  is the  $m$ -spacing and  $Z_{(1)} \leq Z_{(2)} \leq \dots \leq Z_{(n)}$  are the **order statistics**.

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We consider superior adaptations:

- [Ebrahimi et al., 1994]:  $\hat{H}_E$ .
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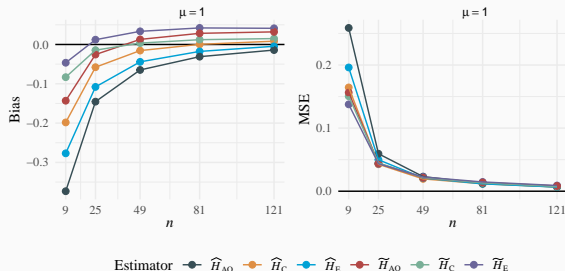
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## Enhanced Bootstrap Technique

$$\tilde{H} = 2\hat{\theta}(\mathbf{Z}) - \frac{1}{B} \sum_{b=1}^B \hat{\theta}_b(\mathbf{Z}^{(b)})$$



Bias and MSE of original and bootstrap versions,  $L = 5$  and  $\alpha = -10$ .

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We aim at testing the following hypotheses:

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$$S(\mathbf{Z}; L) = \tilde{H} - [H_{\Gamma_{\text{SAR}}}(L) + \ln \bar{\mathbf{Z}}] .$$

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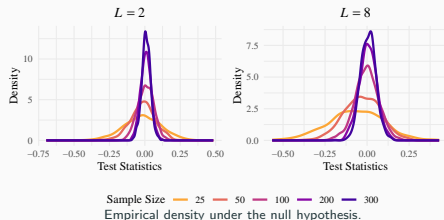
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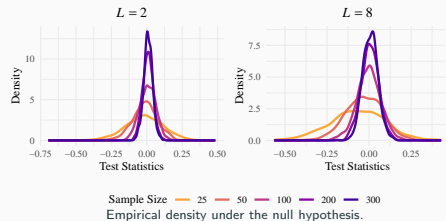
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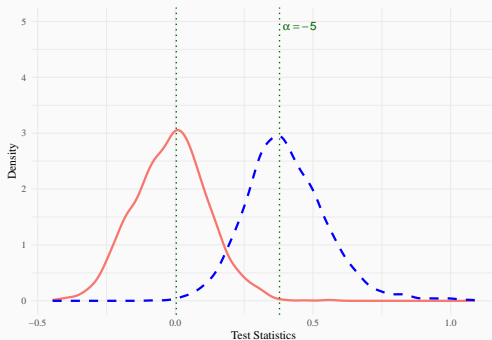
$L$	$n$	Mean	SD	Var	SK	EK	$p$ -value
2	25	-0.0199	0.1370	0.0188	-0.3492	1.2320	0.0586
	50	-0.0145	0.0896	0.0080	-0.1367	0.4826	0.0418
	100	0.0064	0.0562	0.0032	-0.0623	-0.1617	0.3938
	200	0.0067	0.0361	0.0013	0.0199	-0.0305	0.9273
	300	0.0093	0.0309	0.0010	0.0119	0.3522	0.1293
8	25	-0.0696	0.1623	0.0264	-0.1096	0.0158	0.1658
	50	-0.0301	0.1082	0.0117	-0.1741	0.0531	0.2961
	100	-0.0025	0.0711	0.0051	-0.0408	0.3446	0.2279
	200	0.0046	0.0520	0.0027	-0.0873	0.0466	0.5663
	300	0.0070	0.0448	0.0020	0.0818	0.0306	0.5431

Descriptive analysis of  $S(\mathbf{Z}; L)$  to verify the normality of the data with  $L \in \{2, 8\}$  and  $\mu = 1$ .



# Hypothesis Testing

## Empirical densities



Empirical distributions under the null (centered around zero) and alternative hypotheses (e.g.,  $\alpha = -5$ , showing a shift from zero), with  $\mu = 1$  and  $L = 8$ .

## Size and Power of the proposed test.

$L$	$n$	Size			Power		
		1%	5%	10%	1%	5%	10%
5	25	0.016	0.071	0.117	0.564	0.672	0.745
	50	0.014	0.052	0.113	0.640	0.754	0.806
	100	0.010	0.057	0.103	0.738	0.838	0.885
	200	0.010	0.054	0.092	0.807	0.937	0.963
	300	0.007	0.058	0.107	0.850	0.961	0.988
8	25	0.024	0.076	0.129	0.747	0.844	0.838
	50	0.015	0.055	0.104	0.857	0.919	0.939
	100	0.006	0.052	0.099	0.958	0.984	0.985
	200	0.014	0.043	0.109	0.989	0.999	1.000
	300	0.014	0.041	0.106	1.000	1.000	1.000

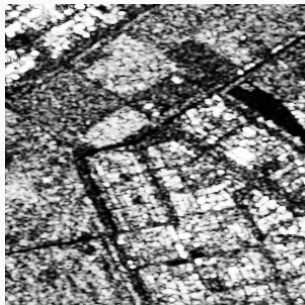
## Application to Actual Data

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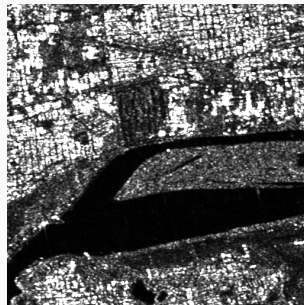
## Application to Actual Data

Images acquired by the Sentinel-1A satellite operating in the C band, with VV polarization, intensity format, and  $L = 5$  nominal looks.

Both images contain homogeneous and heterogeneous zones.



(a) Flevoland

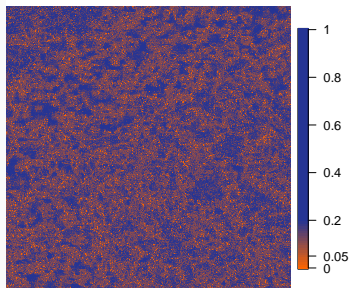


(b) Ottawa

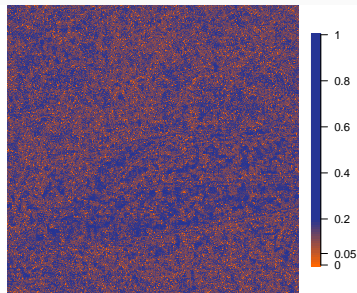
SAR images,  $300 \times 300$  and  $512 \times 512$  pixels, respectively.

## Application to Actual Data

The color table maps all above 0.05 into dark blue (no evidence to reject the fully-developed speckly hypothesis), and those below 0.05 into a gradient between blue and orange (evidence to reject the hypothesis).



(a) Flevoland



(b) Ottawa

Heatmap for a threshold of 0.05 of the  $p$ -values, generated after applying the proposed using  $7 \times 7$  local sliding windows.

## Conclusions & Future Perspectives

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## Conclusions & Future Perspectives

- We proposed a testing procedure to distinguish between fully-developed speckle and heterogeneous clutter.
- This tool is equipped with a test statistic based on a bootstrap-improved estimator of the Shannon entropy.
- Applying the test to actual SAR data yielded promising results, effectively distinguishing different regions in the image and demonstrating the test's potential to differentiate between homogeneous and heterogeneous zones.

Currently, we are working on detecting texturelessness in SAR images. To this end, we propose novel hypothesis tests based on classical and variants of the coefficient of variation.

*Thank you for your attention*

**Contact**

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*Scan to connect*

