SNITCH: Seeking a simple, comparative star formation history inference tool

R. J. Smethurst, 1 et al.

School of Physics and Astronomy, The University of Nottingham, University Park, Nottingham, NG7 2RD, UK

30 July 2018

ABSTRACT

We present snitch, an open source code written in Python, developed to infer a simple toy model of the star formation history of a galaxy spectrum from the emission and absorption features. Snitch utilises FSPS, the Manga DAP and emcee in order to infer the $[Z,t_q,\log\tau]$ parameters which best describe an exponentially declining quenching history to fit the input spectral features. This code was written for use on the Manga spectral data cubes but can be used for any scenario where a galaxy spectrum has been obtained, or spectral feature measurements have been provided. This code is not intended for an accurate inference of a single spectrum SFH but as a comparison of SFHs across many different spectra, either for IFU data or across a large galaxy sample.

Key words: software – description

1 INTRODUCTION

Whilst there are many publicly available codes which provide a full spectral fit to a galaxy spectrum in order to determine its star formation history (SFH), there are few providing a simpler, quicker inference of a toy model of SFH for comparative studies.

Here we present an open source *Python* software which uses Bayesian statistics and MCMC methods to infer a simple toy model of SFH from the absorption and emission features of a galaxy (or IFU) spectrum.

Herein we describe our SFH inference code, SNITCH (bayeSian iNference, given emIssion and absorpTion features, of quenChing Histories), in Section 2, the expected output of the code in Section 3, along with the rigorous testing procedures applied to SNITCH in Section 4. Where necessary we adopt the Planck 2015 (Planck Collaboration et al. 2016) cosmological parameters with $(\Omega_m, \Omega_\lambda, h) = (0.31, 0.69, 0.68)$.

2 DESCRIPTION OF CODE

As a black box, SNITCH takes absorption and emission spectral features and their associated errors as inputs, assumes a simple SFH model and convolves it with a SPS to generate a synthetic spectra. These synthetic spectra are then fitted to give resulting predicted absorption and emission spectral features which the code uses to infer the best fit model using Bayesian statistics and MCMC methods.

We describe this process below, first defining our toy model of SFH, describing how we convolve this with SPS

models to produce synthetic spectra (Section 2.2), how these spectra are then fitted to provide predicted model spectral features (Section 2.3) and then how these are used to infer the best fit SFH given the input parameters (Section 2.5).

2.1 Star Formation History Model

The SFH used in this code was first described in Smethurst et al. (2015). We have reproduced its description here. The quenched star formation history of a galaxy can be simply modelled as an exponentially declining star formation rate (SFR) across cosmic time as:

$$SFR = \begin{cases} I_{sfr}(t_q) & \text{if } t \leqslant t_q \\ I_{sfr}(t_q) \times exp\left(\frac{-(t-t_q)}{\tau}\right) & \text{if } t > t_q \end{cases}$$
 (1)

where t_q is the onset time of quenching, τ is the timescale over which the quenching occurs and I_{sfr} is an initial constant star formation rate dependent on t_q . A smaller τ value corresponds to a rapid quench, whereas a larger τ value corresponds to a slower quench. This model is clearly not a fully hydrodynamical simulation, it is a deliberately simple model built in order to test our understanding of the evolution of galaxy, or stellar, populations. This SFH model has previously been shown to appropriately characterise quenching galaxies (Weiner et al. 2006; Martin et al. 2007; Noeske & et al. 2007; Schawinski et al. 2014). For galaxies which are still star forming, this model results in a constant SFR.

Here, we assume that all galaxies formed at a time t = 0 Gyr with an initial burst of star formation, $I_{sfr}(t_q)$. This initial constant star formation rate must be defined in order

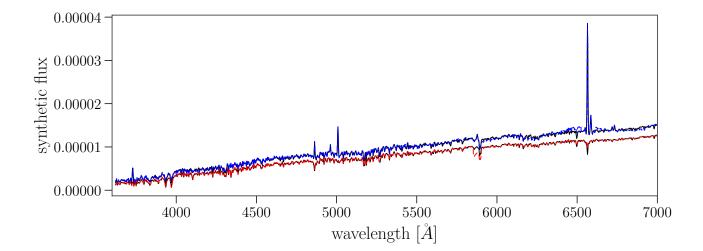


Figure 1. Example synthetic spectra constructed using FSPS, shown by the black solid lines, along with the fit returned by the MaNGA DAP (i.e. ppxf, emission lines and absorption features) shown by the red dashed line for a spectra which has already quenched with $[Z, t_q, \tau] = [1 Z_{\odot}, 11.5 \text{ Gyr}, 0.1 \text{ Gyr}]$ and the blue dashed line for a spectra which still has some residual star formation $[Z, t_q, \tau] = [1 Z_{\odot}, 10.0 \text{ Gyr}, 1.0 \text{ Gyr}]$.

to ensure the 'model' galaxy has a reasonable stellar mass by $z \sim 0$. This value will be dependent on the epoch at which quenching is modelled to occur, hence the dependence of this initial star formation rate on quenching time in Equation 1.

Peng & et al. (2010, Equation 1) define a relation between the average specific SFR (sSFR = SFR/ $\rm M_{*}$) and redshift by fitting to measurements of the mean sSFR of blue star forming galaxies from SDSS, zCOSMOS and literature values from Elbaz et al. (2007) and Daddi et al. (2007) measured at increasing redshifts with data from the GOODS survey:

$$sSFR(m,t) = 2.5 \left(\frac{m}{10^{10} M_{\odot}}\right)^{-0.1} \left(\frac{t}{3.5 \text{ Gyr}}\right)^{-2.2} \text{Gyr}^{-1}.$$
(2)

Beyond $z\sim 2$ the characteristic sSFR flattens and is roughly constant back to $z\sim 6$. This flattening can be seen across similar observational data (Peng & et al. 2010; González et al. 2010; Béthermin et al. 2012); the cause is poorly understood but may reflect a physical limit to the sSFR of a galaxy.

Motivated by these observations, the relation defined in Peng & et al. (2010) is taken up to a cosmic time of t=3 Gyr ($z\sim2.3$) and prior to this the value of the sSFR at t=3 Gyr is used. At the point of quenching, t_q , the SFH models are therefore defined to have an $I_{sfr}(t_q)$ which lies on this relationship for the sSFR, for a galaxy with mass, $m=10^{10.27}M_{\odot}$. This choice of $I_{sfr}(t_q)$ is an important one to produce realistic synthetic galaxy properties, however does not impact on the predicted spectral features output by the model as it is merely a normalisation factor on the SFH. The SFH parameters $[t_q, \log \tau]$, which set the shape of the SFH, and the metallicity, Z, which can affect the strength of a spectral feature, are the crucial parameters.

2.2 Synthetic spectra generation

We then employ stellar population synthesis models (SPS) in order to construct synthetic spectra with the SFHs defined in Section 2.1. These synthetic spectra will be measured the same as an observed spectra (see Section 2.3) in order to make a direct comparison using Bayesian statistics (see Section 2.5) to determine the 'best fit' SFH model for a given spectral features input.

In order to derive a realistic synthetic spectra with our defined SFHs we utilised the Flexible Stellar Population Synthesis (FSPS) Model¹ code of Conroy et al. (2009) and Conroy & Gunn (2010), which is written in FORTRAN, in conjunction with an existing Python wrapper² by Foreman-Mackey et al. (2014). SPS methods rely on stellar evolution calculations during all stages of stellar life, stellar spectral libraries, dust models and initial mass functions (IMFs) to translate the evolution of a hypothetical number of stars of varying ages and metallicities into a predicted integrated spectrum. FSPS also integrates CLOUDY (Ferland et al. 2013) into its output so that stellar emission lines can be synthesised along with the stellar continuum.

The FSPS Python wrapper makes it possible to generate spectra (or magnitudes) for any arbitrary stellar population in just two lines of *Python* code.

In SNITCH we set up the FSPS models to produce spectra using the Padova isochrones (Girardi et al. 2002) and MILES spectral library (Vazdekis et al. 2016) with nebular emission, emission from dust Draine & Li (2007), a Chabrier (2003) IMF and a Calzetti et al. (2000) dust extinction curve. We also smooth the generated synthetic spectra to have the minimum velocity dispersion measurable by MaNGA, 77 km s⁻¹ (Bundy et al. 2015). Spectra are gener-

¹ https://github.com/cconroy20/fsps

http://dfm.io/python-fsps/current/

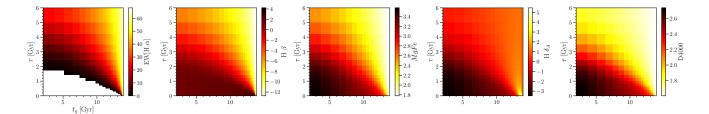


Figure 2. The variation of model spectral features across the logarithmically binned two dimensional $[t_q, \log \tau]$ parameter space measured at $t_{obs} = 13.8 \, \mathrm{Gyr}$ and solar metallicity, $Z = Z_{\odot}$. The features shown from left to right are the equivalent width of the $H\alpha$ emission line and the spectral absorption indices D4000, H β , H $\delta_{\rm A}$ and MgFe'. Note that when a model is no longer star forming, the fitting code cannot measure an equivalent width of $H\alpha$ therefore these values are masked out in the bottom left corner of the left most panel. This figure shows how each feature is sensitive to the changing SFH and how they can be used to break the degeneracies that plague photometric studies of SFH.

ated for the 22 metalicities provided in the MILES models, ranging from $0.011~Z_{\odot}$ to $1.579~Z_{\odot}$ across a logarithmic age range spanning the Universe's history. FSPS does not allow for chemical enrichment of stellar birth material with time, i.e. the metallicity is kept constant for a given SFH at all epochs.³ These spectra are generated across a logarithmically spaced 4-dimensional array in $[t_{obs},~Z,~t_q,~\tau]$ in order to facilitate faster run time during inference (see Section 2.5).

Two example synthetic spectra generated with FSPS for solar metallicity are shown by the solid black line in Figure 1. The fits are shown by the red dashed line for a spectra which has already quenched with $[Z,\ t_q,\ \tau]=[1\ Z_\odot,\ 11.5\ {\rm Gyr},\ 0.1\ {\rm Gyr}]$ and by the blue dashed line for a spectra which still has some residual star formation $[Z,\ t_q,\ \tau]=[1\ Z_\odot,\ 10.0\ {\rm Gyr},\ 1.0\ {\rm Gyr}]$ both observed at a redshift, z=0.1 (i.e. $t_{obs}=12.1\ {\rm Gyr}$).

2.3 Measuring the synthetic spectral features

This code was originally developed for a specific science case using MaNGA IFU survey data (an integral-field spectroscopic survey of 10,000 galaxies undertaken by the fourth phase of the Sloan Digital Sky Survey, SDSS-IV; Bundy et al. 2015). We therefore wished to measure our synthetic spectra generated using FSPS (see Section 2.2) in the same way as the MaNGA data. It is for that reason that we use the functions defined in the MaNGA Data Analysis Pipeline (DAP; Westfall prep) version 2.0.2 in order to measure the features in our synthetic spectra. If the user has a predefined method for measuring emission and absorption features in their spectra, the measure_spec function in SNITCH can simply be adapted.

The MaNGA DAP is described in full detail in Westfall (prep), however here we lay out the functions used in SNITCH to fit our synthetic spectra. Firstly we used pPXF (Cappellari & Emsellem 2004) to extract a fit of the stellar continuum of our synthetic spectra. Here we use the version

of pPXF coded into the MaNGA DAP using the PPXFFit object and the MILES template spectral libraries. Secondly, we used the obtained stellar continuum to fit the emission line features in the spectra using the Elric object and the "ELPFULL" emission line database of 26 lines provided in the Manga Dap. This procedure provides emission line fluxes, equivalent widths, and kinematics from single component Gaussian fits. All strong lines are fit, as well as the Balmer series up to $H\epsilon$ and other weaker lines. Finally, we then fit the absorption indices in our synthetic spectra using the SpectralIndices object and the "EXTINDX" index database of 42 indices provided in the MaNGA DAP. Spectral-index measurements including the 4000Å break, TiO bandhead features and the full Lick system. All indices are measured at the MaNGA resolution (specified for each index) and corrections are provided to a nominal, $\sigma_v = 0$ measurement. The measurements of the Lick indices are provided by convolving the MaNGA data to the Lick resolution.

2.4 Choosing which spectral features to use

Whilst there are many star formation sensitive spectral features used previously in the literature (?) here we adopted a from first principles approach. We observed how each of the 26 emission and 42 absorption features measured by the MaNGA DAP (see Section 2.3), changed across the model parameter space $[Z, t_q, \log \tau]$ with time of observation to determine which spectral features were most sensitive to star formation rate, metallicity, time of observation etc.

We looked at plots similar to those shown in Figure 2 for all 26 emission features and 42 absorption features and selected the following features with which to infer the SFH:

- (i) EW[H α] as it is the most sensitive to current star formation rate
- ii) D4000 as it is most sensitive to the age of the stellar population
- iii) H β as it is most sensitive to the metallicity of the stellar population irrespective of the population age
- iv) Hδ_A as it is the most sensitive to A-stars and therefore star formation that has been cut off abruptly in the recent past
 v) MgFe' as it is most sensitive to the metallicity of the stellar population irrespective of the current star formation rate

Each of these features is measured in the synthetic spectra in order for the input values to be compared to model

 $^{^3}$ Whilst we could attempt to provide a feature to implement (v) chemical enrichment into these models this would firstly be full of uncertainty (the propagation of which would be unquantifiable) and secondly move us out of the regime of a toy SFH model and into hydrodynamic simulation territory.

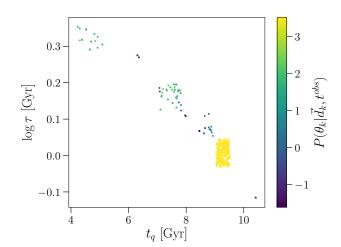


Figure 3. This figure shows the walker positions marginalized over the Z dimension into the two dimensional model $[t_q, \log \tau]$ space and coloured by their log probability value. The higher the value of their log probability, the more likely the model is. The lower values of log probability for some groups of walkers suggests that these are indeed stuck in local minima. These clusters of walkers in local minima can be 'pruned' (see Section 2.6) away to leave only the global minimum in the final output.

values to find the best fit SFH. Less than 5 spectral features can be provided to SNITCH, although not providing one of the five does restrict the accuracy to which a SFH can be inferred (see Section 4.2). An estimate of the error on these measured values is also needed for SNITCH to run. THe more precise the measurement of the spectral feature, the most precise the inferred SFH.

2.5 Bayesian inference of SFH parameters

For the SFH problem at hand, using a Bayesian approach requires consideration of all possible combinations of the model parameters $\theta \equiv [Z,t_q,\log\tau]$ (the hypothesis in this instance). Assuming that all galaxies formed at t=0 Gyr, we can assume that the 'age' of a spectrum is equivalent to an observed time, t_k^{obs} . We used this 'age' to calculate the predicted, p, spectral features at this cosmic time for a given combination of θ : $\vec{d}_{s,p}(\theta_k,t_k^{obs})$ for each of the five spectral features: EW[H α], D4000, H β , H δ A and MgFe'. The predicted spectral features can now directly be compared with input observed spectral features, so that for a single spectrum, k, with features $\vec{d}_k = \{s_k\}$, where S is the number of features, the likelihood of a given model $P(d_k|\theta_k,t_k^{obs})$ can be written as:

$$P(\vec{d}_k | \theta_k, t_k^{obs}) = \prod_{s=1}^{S} \frac{1}{\sqrt{2\pi\sigma_{s,k}^2}} \exp\left[-\frac{(d_{s,k} - d_{s,p}(\theta_k, t_k^{obs}))^2}{\sigma_{s,k}^2}\right]$$
(3)

Here we have assumed that $P(d_{s,k}|\theta_k, t_k^{obs})$ are all independent of each other and that the errors on the observed features, $\sigma_{s,k}$, are also independent (a simplifying

assumption but difficult to otherwise constrain). To obtain the probability of a combination of θ values given the data: $P(\theta_k|\vec{d}_k,t^{obs})$, i.e. how likely a single SFH model is given the observed spectral features of a spectrum, we utilise Bayes' theorem as:

$$P(\theta_k | \vec{d}_k, t^{obs}) = \frac{P(\vec{d}_k | \theta_k, t^{obs}) P(\theta_k)}{\int P(\vec{d}_k | \theta_k, t^{obs}) P(\theta_k) d\theta_k}.$$
 (4)

We assume the following prior on the model parameters so that the probability drops off at the edges of the parameter space: $P(\theta_k) = 1$ if $0 < Z[Z_{\odot}] \le 1.5$ and $0 < t_q$ [Gyr] ≤ 13.8 and $0 < \tau$ [Gyr] ≤ 5.9 and $P(\theta_k) = 2 \times \exp(\log_{10}[5.9]) - \exp(\log_{10}[\tau])$ otherwise.

As the denominator of Equation 4 is a normalisation factor, comparison between likelihoods for two different SFH models (i.e., two different combinations of $\theta_k = [Z, t_q, \log \tau]$) is equivalent to a comparison of the numerators. Markov Chain Monte Carlo (MCMC; Mackay 2003; Foreman-Mackey et al. 2013; Goodman & Weare 2010) analysis provides a robust comparison of the likelihoods between θ values.

MCMC allows for a more efficient exploration of the parameter space than a simple χ^2 analysis by avoiding those areas with low likelihood. A large number of 'walkers' are started at an initial position (i.e. an initial hypothesis, θ), where the likelihood is calculated; from there they individually 'jump' a randomised distance to a new area of parameter space. If the likelihood in this new position is greater than the original position then the 'walkers' accept this change in position. Any new position then influences the direction of the 'jumps' of other walkers (this is the case in ensemble MCMC as used in this investigation but not for simple MCMC, which is much slower at converging). This is repeated for the defined number of steps after an initial 'burnin' phase. The length of this burn-in phase is determined after sufficient experimentation to ensure that the 'walkers' have converged on a region of parameter space. Here we use emcee, a Python module which implements an affine invariant ensemble sampler to explore the parameter space, written by Foreman-Mackey et al. (2013). emcee outputs the positions of these 'walkers' in the parameter space, which are analogous to the regions of high posterior probability.

With each 'walker' jump to a new place in parameter space, a synthetic spectra must be created which must then be measured as described in Section 2.3 to produce predicted spectral parameters. Since this is very computationally expensive, a 4-dimensional look up table of spectral parameters was generated across a logarithmically spaced grid in $[t_{obs}, Z, t_q, \tau]$. We initialised our look up table over a non-regular grid, this was to optimise the number of useful t_q values for each t_{obs} value, i.e. quenched SFHs with $t_q \leq t_{obs}$. Those SFHs with $t_q > t_{obs}$ had constant SFR and so returned the same values for the spectral parameters regardless of the t_q , $\log \tau$ values. This allowed us to construct a finer array in t_q for each value of t_{obs} to pinpoint recent changes in the SFH more precisely.

The look up table is interpolated over (using a nearest neighbour approach to speed up run time over the irregular

⁴ dan.iel.fm/emcee/

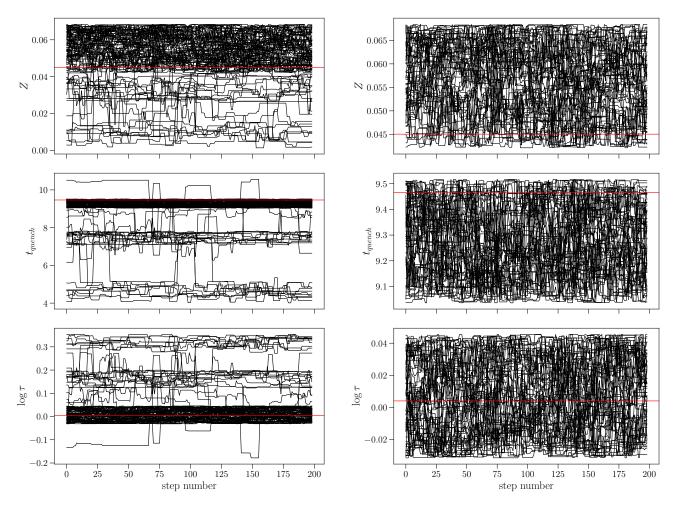


Figure 4. The positions traced by the *emcee* walkers with step number (i.e. time) in each of the $[Z, t_q, \log \tau]$ dimensions in the post burn-in phase before pruning (left) and after pruning (right). The red lines show the known true values in each panel. Walkers have got stuck in local minima (see Figure 3) but some have managed to find the global minimum which are the walkers which remain after pruning in the panel on the right.

grid) to find spectral parameters for each 'walker' jump to any new position in $[t_{obs},\ Z,\ t_q,\ \log\tau]$ parameter space.

For each run of SNITCH, the inference run is initialised with 100 walkers with a burn-in phase of 1000 steps before a main run of 200 steps. Acceptance fractions for each walker are difficult to estimate due to the fact that walkers often get stuck in local minima during a run (see Section 2.6 for more information).

2.6 Pruning walkers stuck in local minima

After running SNITCH and upon inspection of the walker positions it became apparent that the walkers of *emcee* would often get stuck in local minima. We therefore implemented a pruning method, as described in Hou et al. (2012), in order to remove those walkers in local minima leaving only the global minima from which to derive inferred SFH parameters. The method outlined in Hou et al. (2012) is a simple one dimensional clustering method wherein the average negative log-likelihood for each walker is collected. This results

in L numbers:

$$\bar{l}_k = \frac{1}{T} \sum_{t=1}^{T} l(\vec{\theta}_k(t)|D),$$
(5)

where T is the total number of steps each walker, k, takes. These L numbers, \bar{l}_k are therefore characteristic of the well which walker k is in, so that walkers in the same well will have similar \bar{l}_k (see Figure 3 in which walkers are coloured by their characteristic $\bar{l}_{(k)}$ value).

The walkers are all then ranked in order of decreasing average log likelihood, $\bar{l}_{(k)}$, or increasing $-\log \bar{l}_{(k)}$. If there are big jumps in the $-\log \bar{l}_{(k)}$, these are easy to spot and are indicative of areas where walkers have got stuck in local minima. The difference in $-\log \bar{l}_{(k)}$ for every adjacent pair of walkers is then calculated. The first pair whose difference is a certain amount times bigger than the average difference previously is then identified like so:

$$-\log \bar{l}_{(j+1)} + \log \bar{l}_{(j)} > Const \frac{\log \bar{l}_{(j)} + log\bar{l}_{(1)}}{j-1}.$$
 (6)

All the walkers with with k > j are thrown away and

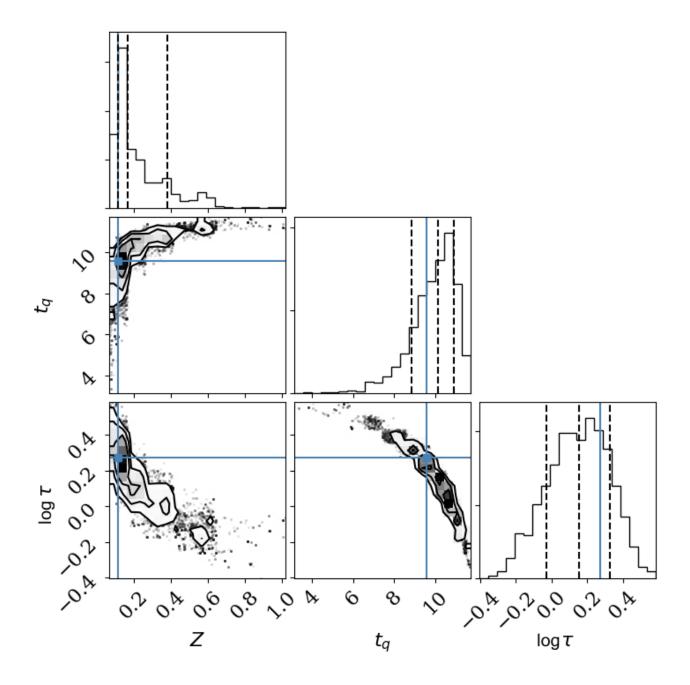


Figure 5. Example output from SNITCH showing the posterior probability function traced by the MCMC walkers across the three dimensional parameter space $[Z, t_q, \log \tau]$. Dashed lines show the 18th, 50th and 64th percentile of each distribution function which can be interpreted as the 'best fit' with 1σ . The blue lines show the known true values which SNITCH has managed to recover.

only the ones with $k \leq j$ are kept after being identified as part of the global minimum. This can be seen in Figure 4 wherein the walker positions at each step before pruning in the burn-in phase are shown in comparison to those after pruning in the main run stage. In the cases where the 'walkers' did not get stuck in local minima, this pruning routine leaves the walker chains untouched.

3 OUTPUT OF CODE

The burn-in and main run walker positions and posterior probabilities at each step are written to disc by snitch. From this three dimensional MCMC chain charting the $[Z,t_q,\log\tau]$ positions of the walkers around parameter space, the 'best fit' $[Z,t_q,\log\tau]$ values along with their uncertainties can be determined from the 16th, 50th and 84th percentile values of the walker positions. An example output from snitch for a single synthetic spectrum constructed

Table 1. The mean uncertainties $(\pm 1\sigma)$ on the best fit and difference in known and best fit values $(\Delta[Z, t_q, \tau])$ for the 10 synthesised galaxy spectra returned when each spectral feature is omitted in turn.

Spectral feature omitted	None	$_{ m Hlpha}$	D4000	$_{\mathrm{H}eta}$	${ m H}\delta_{ m A}$	MgFe'	
Average uncertainty, Z 1σ	0.1	0.1	0.1	0.1	0.1	0.2	_
Average uncertainty, $t_q 1\sigma$	0.2	0.6	0.5	0.2	0.3	0.1	
Average uncertainty, τ 1σ	0.1	0.4	0.2	0.1	0.1	0.1	
$\Delta Z [\mathrm{Gyr}]$	0.1	0.1	0.1	0.1	0.2	0.2	
$\Delta t_q [\mathrm{Gyr}]$	0.3	1.3	1.6	0.8	2.3	2.5	
$\Delta \tau [\mathrm{Gyr}]$	0.2	0.3	0.5	0.4	0.9	0.7	

with the FSPS models (see Section 2.2) is shown in Figure 5. This figure is also written to disc by SNITCH upon completion of a run on a single spectrum.

The routine outlined in Section 2 has been released as an open source package named snitch⁵. The required inputs for snitch to run on a single spectrum are at least one, if not all, of EW[H α], D4000, H β , H δ _A and MgFe' and their associated errors and the spectrum redshift, z.

4 TESTING

4.1 Testing precision

In order to test that SNITCH can find the correct quenched SFH model for a given set of spectral features, 25 synthesised galaxy spectra were created with known SFHs (i.e. known randomised values of $\theta = [Z, t_q, \log \tau]$) from which spectral features were generated using the FSPS models (see Section 2.2). These were input into SNITCH to test whether the known values of θ were reproduced, within error, for each of the 25 synthesised galaxies. In all cases the true values reside within the parameter space explored by the walkers left over after pruning, which trace the global minimum of the posterior probability. SNITCH therefore succeeds in locating the true parameter values within the degeneracies of the SFH model for known values. However, the spread in the walker positions gets broader as the inferred τ value gets larger (i.e. longer quench) and the inferred t_q value gets smaller (i.e. earlier quench). This test therefore demonstrates how snitch is precise in recovering known SFHs, however that precision varies across the parameter space. The median difference between known and inferred paramter values for 25 random SFHs is $[\Delta Z, \Delta t_q, \Delta \tau] = [0.1 \text{ Z}_{\odot}, 0.3 \text{ Gyr}, 0.2 \text{ Gyr}]$ and the maximum difference between the inferred and true values are $[\Delta Z, \Delta t_q, \Delta \tau] = [0.7 \text{ Z}_{\odot}, 3.7 \text{ Gyr}, 0.7 \text{ Gyr}].$

4.2 Testing precision when less spectral information provided

SNITCH is designed so that not all of the spectral features have to be provided for the code to return an inferred quenching history. This is a particularly useful feature if the user is unable to obtain or measure a certain spectral feature, if for example, measurements are being obtained from archival data or a feature lies outside of the wavelength range of their spectrum.

Users should note that quenching histories inferred given less inputs results in a larger uncertainty on the quoted

best fit parameters returned by snitch. To quantify this we generated 10 random $[Z,t_q,\log\tau]$ values and used them to generate synthetic spectra the measurements for which we ran through snitch, each time omitting one of the spectral features from the list of inputs. The mean uncertainties on the best fit and difference between known and best fit values returned when each spectral feature is omitted are quoted in Table 1.

For further combinations of missing parameters, we suggest the user completes their own tests to determine how the quoted uncertainty will change with the omission of more than one spectral feature. However, we do not recommend using SNITCH if the number of available spectral features is less than 4. If this is the case, the number of inputs given to the code will be equal to or less than the number of parameters to be inferred and the inferred SFH will be unreliable.

4.3 Population testing

A further test of SNITCH is to determine whether the inferred SFH parameters of $[Z,\ t_q,\ \log\tau]$ can reproduce the distribution of spectral parameters of a population of galaxy spectra. We randomly selected a sample of 100 SDSS DR14 galaxies with $z<0.15^6$ and used the measured spectral parameters provided by SDSS as inputs to SNITCH. We then used the inferred SFHs returned by SNITCH to estimate inferred spectral parameters for each of the 100 selected SDSS galaxy spectra. We also added a random multiple, drawn from a Gaussian distribution with mean of 0 and standard deviation of 0.4 (i.e. roughly distributed between -1 and 1), of the error on each measured spectral parameter to the inferred value.

The inferred values for each spectral parameter lie in a similar range to those measured in the SDSS DR14 spectra. Figure 6 shows the fractional difference in the inferred and measured spectral parameters and highlights how for the majority of spectra the inferred values are very close to the original measured values. However, in this test we can see that SNITCH has overestimated the value of MgFe' and underestimated the value of D4000 in the majority of cases. A two-dimensional Anderson Darling test (Anderson & Darling 1952) between the distributions of inferred and measured spectral parameters also confirms that for MgFe' and D4000 we must accept the null hypothesis that the two distributions are statistically distinguishable.

Upon inspection of the stellar mass - star formation

⁵ http://www.github.com/rjsmethurst/snitch/

⁶ This redshift limit was imposed as it is the maximum redshift of the MaNGA sample and therefore the maximum redshift at which the look up table values were calculated (see Section 2.5).

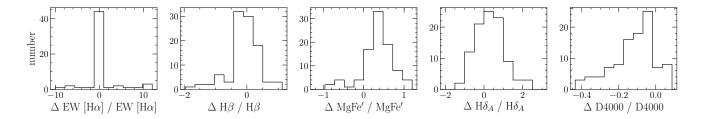


Figure 6. The fractional difference in the inferred and measured spectral parameters for 100 randomly selected SDSS DR14 galaxy spectra. From left to right the difference in EW[H α], $H\beta$, MgFe', H δ _A and D4000.

rate plane⁷ of the 100 measured spectra, those with the largest offsets in inferred and measured spectral parameters occurred for spectra with the highest star formation rates. This is to be expected since SNITCH fits an exponentially declining SFH to each spectra which is a poor fit to a star forming galaxy SFH. The results in Figure 6 suggest that SNITCH can accurately, within error, reproduce the spectral parameters characterising a population of galaxies (or IFU spectra within a single galaxy) which do not lie above the star forming main sequence.

4.4 Testing accuracy

We have shown in the previous section that SNITCH can return precise known values for SFHs, however now we must test its accuracy. In order to quantify this we have run SNITCH on spectra which have previously derived star formation histories. Firstly, on those which have had toy models derived (Section 4.4.1) and then on spectra with star formation histories from hydrodynamic simulations (Section 4.4.2).

4.4.1 Comparing with other toy model SFH codes

In the case of the previously fitted toy models, we have compared the results of snitch with the toy star formation histories derived by Tojeiro et al. (2013) for 6 stacked SDSS spectra of 13959 red ellipticals, 381 blue ellipticals, 5139 blue late-type spirals, 294 red late-type spirals, 1144 blue early-type spirals and 1265 red early-type spirals. We measured the spectral features of each of the 6 stacked spectra using the method outlined in Section 2.3 and input them into snitch. Since Tojeiro et al. (2013) quoted their results in terms of the fraction of stars formed (SFF) in a given time period, we have followed the same method. In Table 2 we have listed the SFF for the six samples used by Tojeiro et al. and the SFF for the best fit parameters returned by snitch along with the median 1σ error. These results are also plotted in Figure 7.

We can see from these results that SNITCH broadly agrees with the results of Tojeiro et al. (2013) for the stacked

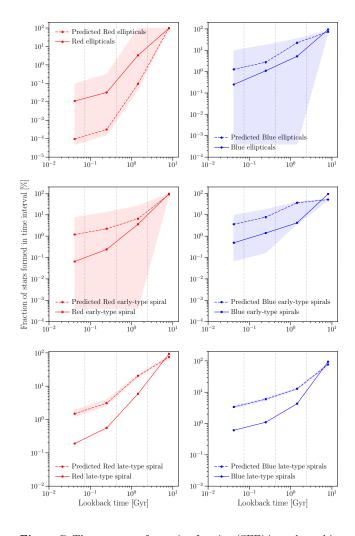


Figure 7. The mean star formation fraction (SFF) in each age bin for the six galaxy samples analysed by (Tojeiro et al. 2013, solid lines) and returned by SNITCH (dashed lines). The shaded region shows the 1σ error on the predicted SFF inferred by SNITCH.

spectra of ellipticals and early-type galaxies, i.e. those galaxies which typically have lower star formation rates. This suggests that SNITCH does indeed return an accurate toy model of star formation history for those galaxies which are quenching or quenched.

However, we do not see agreement between the SNITCH derived SFFs and those of Tojeiro et al. (2013)

⁷ Stellar masses and star formation rates were derived from the MPA-JHU catalogue (Brinchmann et al. 2004, which are corrected for aperture bias) and are available at http://wwwmpa. mpa-garching.mpg.de/SDSS/DR7/

for their late-type galaxies, i.e. those galaxies which are typically star forming. This is to be expected since SNITCH fits a quenching SFH model to a galaxy spectrum and so returns an inaccurate SFH for star forming spectra (see Section 4.5.1).

4.4.2 Comparing with 'realistic' known SFHs from hydrodynamic simulations

We generated 8238 simulated galaxy SFHs using the LGalaxies suite of hydrodynamic simulations (Henriques et al. 2015) at a redshift of z = 0.043 (the mean redshfit of the MaNGA DR14 sample) with a range of star formation rates, $0 < SFR [M_{\odot} yr^{-1}] < 1$, and stellar masses $10^9 < M_*[M_{\odot}] < 10^{11}$. Of these 8238 simulated galaxies we selected all of those flagged by LGalaxies to have a quasar accretion rate above zero. This resulted in 104 simulated galaxies SFHs. We used the FSPS models of Conroy & Gunn (2010) to generate synthetic spectra for each of these 104 simulated SFH (as described in Section 2.2) and then measured their spectral features using the MaNGA DAP functions outlined in Section 2.3. We then input these measurements into SNITCH to derive the best fit $[Z, t_q, \log \tau]$ parameters for our toy model of SFH to compare with the known SFH output by the hydrodynamic simulation. This test is therefore very similar to our tests with different known SFHs that we generated in Section 4.5, however the SFHs generated by the hydrodynamic simulation can be classed as both more varied and more "realistic" in this case.

Figure 8 shows the normalised star formation fractions as generated by Lgalaxies and inferred for their spectra by SNITCH for 10 randomly selected simulated SFHs. We can see that the output from SNITCH largely agrees with the known SFHs of Lgalaxies, largely by identifying the most recent epoch of a dramatic change in the star formation rate.

We can also generalise the SFHs generated by Lgalaxies and returned by SNITCH into two parameters, the time of maximum SFR, $\log t_{\rm max}$, and the time for the SFR to drop to half of the maximum value (log $t_{1/2}$). These generalised parameters roughly trace the exponential SFH parameters of t_q and τ , but allow for a comparison to the SFHs generated by Lgalaxies which are not constrained to a specific form. Figure 9 shows the difference between the generated and inferred values of $\log t_{\text{max}} \& \log t_{1/2}$. We can see that for the majority of synthetic spectra the inferred generalised SFH parameters are comparable to those generated by Lgalaxies. However there is a much larger spread in $\Delta \log t_{1/2}$ (shown in the right panel of Figure 9) than in $\Delta \log t_{\text{max}}$ (shown in the left panel), suggesting that for galaxies with more realistic SFHs, SNITCH will return a more accurate value for the time of quenching, t_q , than for the rate that quenching occurs, $\log \tau$.

4.5 Testing performance with different SFH definitions

4.5.1 Star Forming SFHs

We must also understand how SNITCH behaves when spectral parameters derived from a star forming spectra are input. For example, Figure 10 shows the example output from SNITCH across the three dimensional parameter space

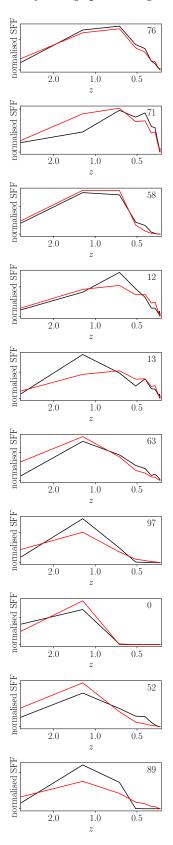


Figure 8. Comparison of the SFFs generated by Lgalaxies (black) and those inferred by SNITCH (red) for 10 randomly selected synthetic spectra with SFRs in the range $0 < {\rm SFR} \ [{\rm M}_{\odot} \ {\rm yr}^{-1}] < 1$, and stellar masses $10^9 < M_* \ [{M}_{\odot}] < 10^{11}$. Note how SNITCH is sensitive to the most recent change in the SFF.

Table 2. The mean star formation fraction (SFF) in each age bin for the six galaxy samples quoted by (Tojeiro et al. 2013, TSFF) and returned by SNITCH. Each value is quoted with a 1σ uncertainty, for the Tojeiro et al. (2013) values this is quoted as the standard error on the mean for each bin . The SFF and 1σ errors are given in units of 10^{-3} .

	$0.01 - 0.074 \; \mathrm{Gyr}$		0.074 - 0.425 Gyr		0.425 - 2.44 Gyr		2.44 - 13.7 Gyr	
	TSFF	SNITCH SFF	TSFF	SNITCH SFF	TSFF	SNITCH SFF	TSFF	SNITCH SFF
Red ellipticals	0.11	0.005 ± 0.001	0.32	0.02 ± 0.01	33	2.5 ± 1	966	997 ± 4
Red early-type spirals	0.65	11 ± 34	2.4	20 ± 62	36	74 ± 142	960	891 ± 245
Red late-type spirals	1.9	19 ± 4	5.6	34 ± 7	59	86 ± 14	933	856 ± 26
Blue ellipticals	2.5	12 ± 51	11	24 ± 92	52	163 ± 173	934	797 ± 327
Blue early-type spirals	4.9	37 ± 49	14	78 ± 90	42	364 ± 173	938	511 ± 310
Blue late-type spirals	6.1	31 ± 35	11	68 ± 62	43	358 ± 142	939	535 ± 255

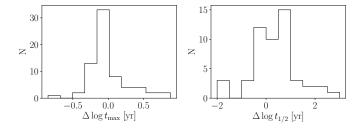


Figure 9. Comparison of the difference between the calculated and inferred time of maximum SFR ($\Delta \log t_{\rm max}$; left) and time for the SFR to drop to half of the maximum value ($\Delta \log t_{1/2}$; right) for the 104 synthetic SFHs with a non zero quasar accretion rate generated by LGalaxies.

 $[Z,t_q,\log au]$ for a synthetic galaxy spectrum which is still star forming at a constant rate by the time of observation. Note that the walkers have explored only the parameter space where $t_q > t_{obs}$, i.e. the observed redshift of the galaxy (see Section 2.2), and all possible values of $\log au$ since the synthetetic galaxy has not yet quenched and therefore all quenching rates are equally likely.

4.5.2 Different Forms of Quenching SFHs

Obviously, not all galaxies will be accurately described by an exponentially quenching SFH. In special use cases (for example studying post starburst galaxies) a different SFH may be defined by the user by replacing the expsfh function with their own.

However, we have also tested how SNITCH behaves when spectra with known SFHs of different forms are input. We tested spectra with constant, burst, many burst, normal and log-normal models of SFH, all of which are often used in the literature to model simple SFHs.

We found that SNITCH was always sensitive to the most recent epoch of star formation or quenching. For the constant SFR model, snitch returned a very recent t_q and a very large τ , i.e. a galaxy which has had constant SFR up until very recently at which point it started to decline very slowly. For the burst and many-burst models, snitch returns a constant SFR up until the peak of the last burst at which point quenching happens very rapidly. Similarly for the log normal and normal SFHs, snitch returns a best fit SFH with constant SFR until the peak of the normal and which point it declines at a rate comparable to the drop off of the Gaussian SFH. All of these tests suggest that snitch is most sensitive to the most recent epoch of star formation

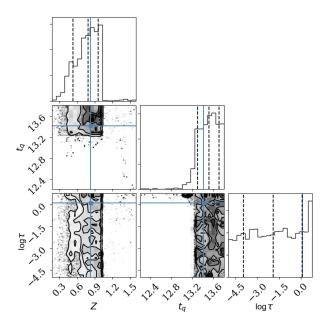


Figure 10. Example output from SNITCH showing the posterior probability function traced by the MCMC walkers across the three dimensional parameter space $[Z,t_q,\log\tau]$, for a synthetic galaxy spectrum which has a SFH with constant SFR. Dashed lines show the 18th, 50th and 64th percentile of each distribution function which can be interpreted as the 'best fit' with 1σ . The blue lines show the known true values which snitch has managed to recover, within error.

but can also roughly trace the quenching of star formation even if the true decline does not occur at an exponential rate.

5 CONCLUSIONS

We have described our quenching history inference code SNITCH (bayeSian iNference, given emIssion and absorpTion spectral features, of quenChing Histories) which uses absorption and emission features measured from a spectra to infer a simple exponentially declining toy model of star formation history. We have demonstrated that it is both precise and accurate at deriving a toy model of star formation history. We have completed a rigorous set of tests the results of which suggest that SNITCH is most sensitive to the most recent epoch of star formation but can also trace the quench-

ing of star formation even if the true decline does not occur at an exponential rate. We advocate the use of this code as a comparative tool across a large sample of spectra, rather than to derive a detailed star formation history of a single spectra.

REFERENCES

- Anderson T. W., Darling D. A., 1952, Ann. Math. Statist., 23, 193
- Béthermin M., Daddi E., Magdis G., Sargent M. T., Hezaveh Y., Elbaz D., Le Borgne D., Mullaney J., Pannella M., Buat V., Charmandaris V., Lagache G., Scott D., 2012, ApJ, 757, L23
- Brinchmann J., Charlot S., White S. D. M., Tremonti C., Kauffmann G., Heckman T., Brinkmann J., 2004, MN-RAS, 351, 1151
- Bundy K., Bershady M. A., Law D. R., Yan R., Drory N., MacDonald N., Wake D. A., Cherinka B., Sánchez-Gallego J. R., Weijmans A.-M., Thomas D., et a., 2015, ApJ, 798, 7
- Calzetti D., Armus L., Bohlin R. C., Kinney A. L., Koornneef J., Storchi-Bergmann T., 2000, ApJ, 533, 682 Cappellari M., Emsellem E., 2004, PASP, 116, 138

Chabrier G., 2003, PASP, 115, 763

Conroy C., Gunn J. E., 2010, ApJ, 712, 833

- Conroy C., Gunn J. E., White M., 2009, ApJ, 699, 486
- Daddi E., Dickinson M., Morrison G., Chary R., Cimatti A., Elbaz D., Frayer D., Renzini A., Pope A., Alexander D. M., Bauer F. E., Giavalisco M., Huynh M., Kurk J., Mignoli M., 2007, ApJ, 670, 156
- Draine B. T., Li A., 2007, ApJ, 657, 810
- Elbaz D., Daddi E., Le Borgne D., Dickinson M., Alexander D. M., Chary R.-R., Starck J.-L., Brandt W. N., Kitzbichler M., MacDonald E., Nonino M., Popesso P., Stern D., Vanzella E., 2007, A&A, 468, 33
- Ferland G. J., Porter R. L., van Hoof P. A. M., Williams R. J. R., Abel N. P., Lykins M. L., Shaw G., Henney W. J., Stancil P. C., 2013, Rev. Mexicana Astron. Astrofis., 49, 137
- Foreman-Mackey D., Hogg D. W., Lang D., Goodman J., 2013, PASP, 125, 306
- Foreman-Mackey D., Sick J., Johnson B., 2014, 10.5281/zenodo.12157
- Girardi L., Bertelli G., Bressan A., Chiosi C., Groenewegen M. A. T., Marigo P., Salasnich B., Weiss A., 2002, A&A, 391, 195
- González V., Labbé I., Bouwens R. J., Illingworth G., Franx M., Kriek M., Brammer G. B., 2010, ApJ, 713, 115

Goodman J., Weare J., 2010, CAMCS, 5, 65

- Henriques B. M. B., White S. D. M., Thomas P. A., Angulo R., Guo Q., Lemson G., Springel V., Overzier R., 2015, MNRAS, 451, 2663
- Hou F., Goodman J., Hogg D. W., Weare J., Schwab C., 2012, ApJ, 745, 198
- Mackay D. J. C., 2003, Information Theory, Inference and Learning Algorithms. Cambridge University Press
- Martin D. C., Wyder T. K., Schiminovich D., Barlow T. A., Forster K., Friedman P. G., Morrissey P., Neff S. G., Seibert M., Small T., Welsh B. Y., Bianchi L., Donas J., Heck-

man T. M., Lee Y.-W., Madore B. F., Milliard B., Rich R. M., Szalay A. S., Yi S. K., 2007, ApJS, 173, 342

Noeske K. G., et al. 2007, ApJ, 660, L43

Peng Y.-j., et al. 2010, ApJ, 721, 193

- Planck Collaboration Ade P. A. R., Aghanim N., Arnaud M., Ashdown M., Aumont J., Baccigalupi C., Banday A. J., Barreiro R. B., Bartlett J. G., et al. 2016, A&A, 594, A13
- Schawinski K., Urry C. M., Simmons B. D., Fortson L., Kaviraj S., Keel W. C., Lintott C. J., Masters K. L., Nichol R. C., Sarzi M., Skibba R., Treister E., Willett K. W., Wong O. I., Yi S. K., 2014, MNRAS, 440, 889
- Smethurst R. J., Lintott C. J., Simmons B. D., Schawinski K., Marshall P. J., Bamford S., Fortson L., Kaviraj S., Masters K. L., Melvin T., Nichol R. C., Skibba R. A., Willett K. W., 2015, MNRAS, 450, 435
- Tojeiro R., Masters K. L., Richards J., Percival W. J., Bamford S. P., Maraston C., Nichol R. C., Skibba R., Thomas D., 2013, MNRAS, 432, 359
- Vazdekis A., Koleva M., Ricciardelli E., Röck B., Falcón-Barroso J., 2016, MNRAS, 463, 3409
- Weiner B. J., Willmer C. N. A., Faber S. M., Harker J., Kassin S. A., Phillips A. C., Melbourne J., Metevier A. J., Vogt N. P., Koo D. C., 2006, ApJ, 653, 1049 Westfall K., in prep., ApJ