COVER PAGE

STAT 608 Homework 02, Summer 2017

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Solution 1

$$\frac{S_{y}^{2} = \sum_{j=1}^{m} (y_{j} - \overline{y})^{2}}{\gamma - 1} = 100 = \frac{SST}{\eta - 1}$$

$$\frac{\hat{\nabla}^{2} = \sum_{j=1}^{m} (y_{j} - \hat{y})^{2}}{\eta - 2} = 10 = \frac{RSS}{\eta - 2}$$

SST = 100 X49 = 4900

=> RSS = 480

Since

SST= SSreg + RSS and all terms are positive. So Sy is related to total sample variability while T 2 is related to the variability unexplained by the model Clearly, Sy is not an Vertinator of o 2 (C)

Solution 2

(a) From the standardized residual us Distance plat, one value, lies outside the -2 to 2 interval Cook's distance for this point is found to be >4, (identified as City No 17) Clearly this is an outlier unexplained by the

model described Since the model is not valid, any claim based on the model are not useful

As far as the claim of 994/ variability explained by the model is concerned one of the points (City 13) lies far enough in the x-clirection to be classified as leverage point. This point has standardized residual inside -2 to 2 interval and is therefore a good leverage point. Good leverage points can increase the value of 12° and decrease the estimated regression coefficient.

(b) No (learly with current model there is a bad leverage point Ifiso there is a non-linear (non-random) battern in Standard residual plot which shows that the assumption of constant variance of residual errors is not satisfied to improve the model, first cheek using box-plot/normal Q-Q plot for Fare and distance variable if the samples resemble normal distribution. Accordingly, consider transforming any one or both followed by regression

Accordingly, consider transforming any one or both followed by regression Box-cox Method could be used for Distance vanish while Inverse Response plots could be used for Fare variable.

Solution 3 From Taylor series results on 1977 $Var(f(Y)) = [f'(E(Y))]^2 Var(Y)$

$$= \left[f'(\mu)\right]^2 \mu^2$$
Criven $Var\left(f(y)\right) = 1$

$$\Rightarrow \left[f'(y)\right]_{y=\mu}^2 = \frac{1}{\mu^2}$$

$$\Rightarrow f'(Y) = \frac{1}{y} \Rightarrow f'(Y) = \frac{1}{y}$$

$$\Rightarrow f(Y) = \log(Y)$$

Solution 4

$$Y_i = \beta x_i + e_i$$

Var
$$(e_{i}(x_{i}) = x_{i}^{2} - 2)$$

Comparing with residual definition, $W_{i} = \frac{1}{3x_{i}^{2}}$
 $e_{Wi} = \sqrt{W_{i}}(y_{i} - \hat{y}_{i})$
 $= x_{i}(y_{i} - \hat{y}_{i})$
 $= x_{i}(y_{i} - \hat{y}_{i})$
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For legat square extinate,

$$\frac{\partial WRSS}{\partial \beta} = 0$$

$$\frac{\partial WRSS}{\partial \gamma} = \frac{2}{2} \times \frac{1}{3} \times \frac{1}{3$$

Solution 5

- (a) Since Vi is the median of n_i observations so $Var(V_i) < \frac{1}{n_i}$ So weights $w_i = n_i$
- (b) From Yi vs XII and Yi vs XII plots the relationship doesnot seem like linear.
 From standardized residual plot vanance is not constent.
 Significant no of outliers from standardized residual plot.
- (C)(i) Find transformations $\Psi_{S_1}(x|i,\lambda_{X_1})$ and $\Psi_{S_2}(x_2i,\lambda_{X_2})$ that make XI $\times 2$ close to normal using Box-ox in find $\lambda_{X_1},\lambda_{X_2}$
 - (iv Consider regression model $V = g(\beta_1 + \beta_1 + \beta_2 + \beta_2)$ Using inverse response plut find g^{-1} and

 corresponding regression coefficients.