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STATISTICS 610-601: HW # 5



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DISTRIBUTION THEORY
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Homework 5 (due on Thursday October 12, 2017)

Problem 1: Consider a r.v. X and the transformation $Y = e^X$ with $E(Y) = 1$. Can we determine the sign of $E(X)$? If no: why not? If yes: find it.

Problem 2: numbers 2.30 (a) - (c) from the textbook.

Problems 3 to 5: numbers 2.34, 3.6 and 3.18 from the textbook.

Problem 6: Suppose you toss a coin about which you know that, on average, in one third of all cases it shows “Tails” and in two thirds of all cases “Heads”. Consider the following situation: each time you observe “Heads” you roll a fair die and note the number of dots x . If the coin shows “Tails” you do *not* roll the die but, instead, note $x = 0$ for the observed number of dots. Find the moment generating function of X = “number of dots”.

Solution 9

(a) $f(x) = \frac{1}{c}$ $0 < x < c$

$$MGF_x(t) = \int_0^c e^{tx} \frac{1}{c} dx$$

$$= \left. \frac{e^{tx}}{t} \right|_0^c \frac{1}{c} = \left(\frac{e^{tc}}{t} - \frac{1}{t} \right) \frac{1}{c}$$

$$= \frac{1}{ct} (e^{tc} - 1)$$

(b) $f(x) = \frac{2x}{c^2}$ $0 < x < c$

$$MGF_x(t) = \int_0^c e^{tx} \frac{2x}{c^2} dx$$

$$= \left. \frac{2x}{c^2} \cdot \frac{e^{tx}}{t} \right|_0^c - \int_0^c \frac{2}{c^2} \cdot \frac{e^{tx}}{t} dx$$

$$= \left. \frac{2c}{c^2} \frac{e^{ct}}{t} - \frac{2}{tc^2} \left(\frac{e^{tx}}{t} \right) \right|_0^c$$

$$= \frac{2e^{ct}}{ct} - \frac{2e^{ct}}{t^2 c^2} + \frac{2}{tc^2}$$

$$= \frac{2}{c^2 t^2} (ct e^{ct} - e^{ct} + 1)$$

(c) $f(x) = \frac{1}{2\beta} e^{-|x-\alpha|/\beta}$ $-\infty < x < \infty$
 $-\infty < x < \infty$

$$= \begin{cases} \frac{1}{2\beta} e^{-(x-\alpha)/\beta} & x - \alpha > 0 \\ \frac{1}{2\beta} e^{(\alpha-x)/\beta} & x \leq \alpha \end{cases}$$

$$E(e^{tx}) = \int_{-\infty}^{\alpha} \frac{e^{tx}}{2\beta} e^{(\alpha-x)/\beta} dx + \int_{\alpha}^{\infty} \frac{e^{tx}}{2\beta} e^{-(x-\alpha)/\beta} dx$$

$$= \int_{-\infty}^{\alpha} \frac{\exp\left[t + \frac{1}{\beta}\right] \alpha}{2\beta} \cdot \exp\left(-\frac{\alpha}{\beta}\right) dx +$$

$$\int_{\alpha}^{\infty} \frac{\exp\left[t - \frac{1}{\beta}\right] x}{2\beta} \exp\left(\frac{x}{\beta}\right) dx$$

$$= \left. \exp\left(t + \frac{1}{\beta}\right) \alpha \right|_{-\infty}^{\alpha} \frac{e^{-\alpha/\beta}}{2\beta(t + \frac{1}{\beta})} + \left. \exp\left(t - \frac{1}{\beta}\right) \alpha \right|_{\alpha}^{\infty} \frac{e^{\alpha/\beta}}{2\beta(t - \frac{1}{\beta})}$$

$$= e^{tx} \cdot e^{\alpha/\beta} \cdot \frac{e^{-\alpha/\beta}}{2(\beta t + 1)} - e^{tx} \cdot \frac{e^{-\alpha/\beta} \cdot e^{\alpha/\beta}}{2(\beta t - 1)}$$

$$= \frac{e^{xt}}{2} \left[\frac{1}{\beta t + 1} - \frac{1}{\beta t - 1} \right]$$

$$= \frac{e^{xt}}{2} \cdot \frac{-2}{\beta^2 t^2 - 1} = \frac{e^{xt}}{1 - \beta^2 t^2}$$

for $-1 < t < \frac{1}{\beta}$

$$\Rightarrow MGF_x(t) = \begin{cases} \frac{e^{\frac{\alpha t}{\beta}}}{1-\frac{\alpha^2}{\beta^2}t^2} & -\frac{1}{\beta} < t < \frac{1}{\beta} \\ \infty & \text{else} \end{cases}$$

Solution 1

$$E(Y) = E(e^X) = 1$$

$$\Rightarrow \int_{-\infty}^{\infty} e^x f_x(x) dx = 1$$

$$\Rightarrow \int_{-\infty}^{\infty} \left(1+x + \sum_{k=2}^{\infty} \frac{x^k}{k!}\right) f_x(x) dx = 1$$

$$\Rightarrow \int_{-\infty}^{\infty} f_x(x) dx + \int_{-\infty}^{\infty} x f_x(x) dx + \int_{-\infty}^{\infty} \left(\sum_{k=2}^{\infty} \frac{x^k}{k!}\right) f_x(x) dx = 1$$

$$\Rightarrow 1 + E(X) = 1 - \int_{-\infty}^{\infty} g(x) f_x(x) dx$$

$$\Rightarrow E(X) = - \int_{-\infty}^{\infty} g(x) f_x(x) dx$$

$$g(x) = x^2 \left(\frac{1}{2!} + \frac{x}{3!} + \frac{x^2}{4!} + \dots \right)$$

$$= e^x - (1+x)$$

for $g(x) \geq 0$

$\Rightarrow e^x \geq 1+x$ which is true for all x

$$\Rightarrow g(x) \geq 0 \quad \forall x$$

$$\Rightarrow \int_{-\infty}^{\infty} g(x) f_x(x) dx \geq 0 \quad \forall x$$

$$\Rightarrow E(X) < 0$$

Yes. $E(X)$ is negative.

Solution 3

$$r=1$$

$$E(X) = 0$$

(mean of std normal)

$$E(Y) = \frac{\sqrt{3}}{6} - \frac{\sqrt{3}}{6} + 0 = 0$$

$$r=2$$

$$E(X^2) = \text{Var}(X) - E(X) = 1 - 0 = 1$$

$$E(Y^2) = 3 \times \frac{1}{6} + 3 \times \frac{1}{6} + 0 = \frac{1}{2} + \frac{1}{2} = 1$$

$$r=3$$

$$E(X^3) = 0$$

(odd moment of std normal)

$$E(Y^3) = \frac{3\sqrt{3}}{6} - \frac{3\sqrt{3}}{6} + 0 = 0$$

$$r=5$$

$$E(X^5) = 0$$

$$E(Y^5) = \frac{9\sqrt{3}}{6} - \frac{9\sqrt{3}}{6} + 0 = 0$$

$$r=4$$

$$MGF_X = e^{t^2/2}$$

2 marks

$$E(X^4) = \frac{d^4}{dt^4} e^{t^2/2} \Big|_{t=0}$$

$$= \frac{d^3}{dt^3} (t e^{t^2/2}) \Big|_{t=0}$$

$$= \frac{d^2}{dt^2} \left(e^{t^2/2} + t (t e^{t^2/2}) \right) \Big|_{t=0}$$

$$= \frac{d}{dt} \left(t e^{t^2/2} + 2t^2 e^{t^2/2} + t^2 (t e^{t^2/2}) \right) \Big|_{t=0}$$

$$= \frac{d}{dt} (3t e^{t^2/2} + t^3 e^{t^2/2}) \Big|_{t=0}$$

$$= 3e^{t^2/2} + 3t(t e^{t^2/2}) + \frac{d}{dt}(t^3 e^{t^2/2}) \Big|_{t=0}$$

$$= 3 + 0 + 0$$

2 marks

$$E(Y^4) = \frac{9}{6} + \frac{9}{6} + 0 = \frac{18}{6} = 3$$

Solution 6

X	0	1	2	3	4	5	6
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$$P(X=x) \quad \frac{1}{3} \quad \frac{1}{6} \times \frac{2}{3} \quad \frac{1}{6} \times \frac{1}{2} \quad \frac{1}{6} \times \frac{1}{2} \quad \frac{1}{6} \times \frac{1}{2} \quad \frac{1}{6} \times \frac{1}{2} \quad \frac{1}{6} \times \frac{1}{2}$$

$$MGF_X(t) = \sum_{x=0}^6 e^{tx} P(X=x)$$

$$P(\text{Head occurs}) = \frac{2}{3}$$

$$P(\text{Tail occurs}) = \frac{1}{3}$$

$$= \frac{e^0}{2} + \left(\frac{e^t}{12} + \frac{e^{2t}}{12} + \dots \right)$$

$$= \frac{1}{2} + \frac{1}{12} (e^t + e^t \cdot e^t + (e^t)^3 + \dots)$$

$$= \frac{1}{2} + \frac{1}{12} \frac{(e^t)^6 - 1}{e^t - 1}$$

$$= \frac{1}{2} + \frac{1}{12} \frac{e^{6t} - 1}{e^t - 1}$$

Solution 5MGF of Gamma($r, 1$)

$$M_X(t) = \left(\frac{1}{1-t} \right)^r$$

MGF of PY , $Y \sim \text{Negative Binomial}(r, p)$

$$MGF_{PY}(t) = M_Y(pt)$$

$$= \left(\frac{p}{1-(1-p)e^{pt}} \right)^r$$

Now,

$$\begin{aligned}
 \lim_{p \rightarrow 0} \frac{p}{1-(1-p)e^{pt}} &= \lim_{p \rightarrow 0} \frac{1}{-(1-p)e^{pt} \cdot t - (-1)e^{pt}} \\
 &= \lim_{p \rightarrow 0} \frac{1}{t(p-1)e^{pt} + e^{pt}} \\
 &= \lim_{p \rightarrow 0} \frac{1}{-te^{pt} + e^{pt} + tpe^{pt}} \\
 &= \frac{1}{1-t} \\
 \Rightarrow \lim_{p \rightarrow 0} MGF_{pY}(t) &\neq \left(\frac{1}{1-t}\right)^2 = MGF_X(t)
 \end{aligned}$$

Solution 4

(a) Binomial

success = "the insect survives"

$$p = 0.01 \quad n = 2000$$

$$(b) P(X < 100) = \sum_{k=0}^{99} \binom{2000}{k} p^k (1-p)^{2000-k}$$

$$= \sum_{k=0}^{99} \binom{2000}{k} (0.01)^k (0.99)^{2000-k}$$

(c) Approximate by Poisson distribution with $\lambda = np$

$$\begin{aligned}
 \lambda &= 2000 \times 0.01 \\
 &= 20
 \end{aligned}$$

(7)

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$$P(Y < 100) = \sum_{x=0}^{99} e^{-20} \frac{(20)^x}{x!}$$

$$= e^{-20} \sum_{x=0}^{99} \frac{(20)^x}{x!}$$