



RAJAN KAPOOR

STATISTICS 610-601: HW # 9



NOVEMBER 16, 2017
DISTRIBUTION THEORY
Prof. Irina Gaynanova





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NOVEMBER 15, 2017 DISTRIBUTION THEORY

Homework 9 (due on Thursday November 16, 2017)

Problem 1: a random variable T that has a t distribution with n > 2 degrees of freedom can be represented as $T = \frac{Z}{\sqrt{Y/n}}$, where $Z \sim \text{normal}(0,1)$ and $Y \sim \text{chi-square}(n)$ are independent random variables. Use this to obtain E(T) and V

Problem 2: number 4.44 from the textbook

Problem 3: number 4.58 (a), (b), (c) from the textbook

Problem 4-5: numbers 5.21, 5.22 from the textbook

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Solution 1

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(1)
$$T = \frac{Z}{\sqrt{m}} \qquad Z \sim \mathcal{N}(0,1)$$
$$Y \sim \chi^{2}(n) \qquad n > 2$$

$$E(T) = E_{\gamma}(E(T|\gamma))$$

$$= E(0) = 0$$

$$Var(T) = Var(E(T|Y)) + E(Var(T|Y))$$

$$= Var(0) + E_{Y}(\frac{\eta}{Y}) = \eta E_{Y}(\frac{1}{Y})$$

$$E_{1}(\frac{1}{Y}) = \int_{0}^{\infty} \frac{1}{\Gamma(\frac{n}{2})} \frac{1}{2^{n/2}} \int_{0}^{\frac{n}{2}-1} e^{-\frac{n}{2}} dy$$

$$= \frac{1}{2} \frac{\Gamma(\frac{n}{2}-1)}{\Gamma(\frac{n}{2})} \int_{0}^{\infty} \frac{1}{\Gamma(\frac{n}{2}-1)} \frac{(\frac{n}{2}-1)^{-1}}{\Gamma(\frac{n}{2}-1)} e^{-\frac{n}{2}} dy$$

$$=\frac{1}{2}\frac{\Gamma(\frac{n}{2}-1)}{\Gamma(\frac{n}{2})}\times 1$$

$$= \underbrace{\frac{\Gamma(\frac{n}{2}-1)}{2}}_{1} \underbrace{\Gamma(\frac{n}{2}-1)}_{1} \underbrace{\Gamma(\frac{n}{2}-1)}_{1}$$

pdf of Chi-sq(h-2)

$$E_{Y}(\frac{1}{Y}) = \frac{1}{2} \frac{1}{n-2} = \frac{1}{n-2}$$

$$\Rightarrow Var(T) = \frac{n}{n-2}$$

Solution 2

(Proof by Induction)

Clearly base case is tone by Theorem 4.5.6.

$$Var(X_1+X_2) = VarX_1 + VarX_2 + 2Cov(X_1,X_2)$$

Assume it is true for n=k i.e.

$$Var(\overset{k}{\sum}X_i) = \overset{k}{\sum}VarX_i + 2\overset{k}{\sum}Cov(X_i,X_j)$$

$$\stackrel{|k|=|k|}{|k|=|k|} L_2$$

To Prove: It is true for h= K+1 i.e.

Using base case,

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$$= Var(X_{k+1}) + Var(\underbrace{\xi}_{i=1}X_i) + 2\underbrace{\xi}_{i=1}Cov(X_{k+1},X_i)$$

$$= \left(Var(X_{k+1}) + \underbrace{\xi}_{i=1}Var(X_i)\right) + 2\underbrace{\xi}_{i=i < j \leq k}Cov(X_i,X_j) + 2\underbrace{\xi}_{i=1}Cov(X_i,X_j) + 2\underbrace$$

Solution 3

(a)
$$Cov(X,Y) = Cov(X, E(Y|X))$$

 $Cov(X, E(Y|X)) = E(X \cdot E(Y|X)) - E(X) E(E(Y|X))$
 $= E(XY|X) - E(X) E(Y)$
 $= Cov(X,Y)$
 $= Cov(X,Y)$

(b)
$$Cov(X, Y - E(Y|X))$$

$$= cov(X, Y) - cov(X, E(Y|X))$$

$$= cov(X, Y) - cov(X, Y)$$
(Using result from (a))

(C)
$$Var(Y-E(Y|X)) = E(Var(Y|X))$$

Solution 4

$$P(max(X_1,X_2)>m)$$

$$= 1 - P(\max(x_1, x_2) < m)$$

$$= 1 - P(X_1 < m) P(X_2 < m)$$

$$= 1 - \left(\frac{1}{2}\right)\left(\frac{1}{2}\right)$$

$$=1-\frac{1}{4}=\frac{3}{4}$$

Solution 5

$$Z = min(X_iY)$$

$$F_{Z^2}(z) = P(Z^2 \leq z)$$

$$F_{Z^{2}}(z) = \frac{1}{2} \times 2 \times P(I \times I \leq \sqrt{z})$$

$$= P(-\sqrt{z} \leq X \leq \sqrt{z})$$

$$= P(X \leq \sqrt{z}) - P(X \leq -\sqrt{z})$$

$$= F_{X}(\sqrt{z}) - F_{X}(-\sqrt{z})$$

$$= f_{X}(\sqrt{z}) + f_{X}(\sqrt{z}) \frac{1}{2\sqrt{z}}$$

$$= f_{X}(\sqrt{z}) = \frac{1}{\sqrt{2}} e^{-2/2} \cdot z^{-1/2}$$

which is pof of $\chi^2(1)$ i.e. chi-square r.v. with I df.