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STATISTICS 610-601: HW # 3



SEPTEMBER 28, 2017
DISTRIBUTION THEORY
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Homework 3 (due on Thursday September 28, 2017)**Problem 1.**

- a) Consider $f(x) = cx^{-3}1_{(1,\infty)}(x)$ with $c \in \mathbb{R}$. Determine c such that f is a probability density function and find the corresponding cumulative distribution function F .
- b) Consider $f(x) = c \sin(x)1_{(0,\pi)}(x)$ with $c \in \mathbb{R}$. Determine c such that f is a probability density function and find the corresponding cumulative distribution function F .

Problem 2. Consider a distribution P given by a pdf $f(\cdot)$ with $f(x) = 0$ for all $x \leq 0$. Assume F is a continuous antiderivative of f with $F'(x) = f(x)$ for all $x > 0$ and $F(x) = 0$ for all $x \leq 0$. Then, by definition,

$$P((a, b]) = \int_a^b f(x) dx = F(b) - F(a) \quad \text{for } 0 \leq a \leq b < \infty.$$

Further assume that there is a constant $\lambda > 0$ with

$$(*) \quad \lim_{0 < h \rightarrow 0} \frac{P((t, t+h] | (t, \infty))}{h} = \frac{1}{\lambda} \quad \text{for all } t > 0.$$

- a) Show that $f(x) = \frac{1}{\lambda}e^{-x/\lambda}$ for all $x > 0$ (exponential distribution).
- b) If P is the distribution of the lifetime of an individual or object, then $P((t, \infty))$ is the probability that the individual survives at least t time units. Given this, how would you interpret equation $(*)$? Is it reasonable to assume that humans' lifetimes are exponentially distributed?

Problem 3. Consider a random variable X with a Poisson(1) distribution and the transformation $Y = (X + 1)^{-1}$. Specify the probability mass function $f_Y(y)$ of Y . Check your answer by verifying that $f_Y(y)$ satisfies the criteria for a probability mass function.

Problems 4 - 6: numbers 1.53, 1.55, 2.1 from the textbook.

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Solution 1

$$(a) f(x) = cx^{-3} \mathbb{1}_{(1, \infty)}(x) \quad c \in \mathbb{R}$$

$$(i) f(x) \geq 0 \quad \forall x \Rightarrow c \geq 0$$

$$(ii) \int_{-\infty}^{\infty} f(t) dt = 1$$

$$\Rightarrow \int_1^{\infty} ct^{-3} dt = 1$$

$$\Rightarrow c \left[\frac{t^{-2}}{-2} \right]_1^{\infty} = 1$$

$$\Rightarrow -\frac{c}{2} [0 - 1] = 1 \Rightarrow c = 2$$

CDF

$$f(x) = 2x^{-3} \mathbb{1}_{(1, \infty)}(x)$$

 $x > 1$

$$\begin{aligned} F_x(x) &= \int_{-\infty}^x f(t) dt \\ &= \int_1^x 2t^{-3} dt \\ &= \left[\frac{2}{-2} [t^{-2}] \right]_1^x \\ &= -\left(\frac{1}{x^2} - 1 \right) = 1 - \frac{1}{x^2} \end{aligned}$$

$$F_x(x) = \begin{cases} 0 & x \leq 1 \\ 1 - \frac{1}{x^2} & x > 1 \end{cases}$$

$$(b) f(x) = c \sin(x) \mathbb{1}_{(0,\pi)}(x) \quad c \in \mathbb{R}$$

$$(i) f(x) \geq 0 \quad \forall x \Rightarrow c \geq 0 \quad [\sin(x) > 0 \text{ in } (0, \pi)]$$

$$(ii) \int_{-\infty}^{\infty} f(t) dt = 1$$

$$\Rightarrow \int_0^{\pi} c \sin(t) dt = 1$$

$$\Rightarrow -c [\cos(t)]_0^{\pi} = 1$$

$$\Rightarrow -c(-2) = 1 \Rightarrow c = \frac{1}{2}$$

CDF

$$f(x) = \frac{1}{2} \sin(x) \mathbb{1}_{(0,\pi)}(x)$$

$$x \geq \pi$$

$$F_x(x) = 1$$

$$x \leq 0$$

$$F_x(x) = 0$$

$$x \in (0, \pi)$$

$$F_x(x) = \int_{-\infty}^x f(t) dt$$

$$= \int_0^x \frac{1}{2} \sin(t) dt$$

$$= -\frac{1}{2} [\cos(t)]_0^x$$

$$= \frac{1 - \cos(x)}{2}$$

Solution 2

(a)

$$\lim_{0 < h \rightarrow 0}$$

$$\frac{P((t, t+h] | (t, \infty))}{h} = \frac{\frac{(t+h)^{\lambda}}{\lambda}}{\lambda} \quad \forall t > 0$$

$$\Rightarrow \lim_{0 < h \rightarrow 0}$$

$$\frac{P((t, t+h] \cap (t, \infty))}{P((t, \infty)) \cdot h} = \frac{\frac{(t)^{\lambda}}{\lambda}}{\lambda}$$

$$\Rightarrow \lim_{0 < h \rightarrow 0}$$

$$\frac{P((t, t+h])}{h} = \frac{P((t, \infty))}{\lambda}$$

$$\Rightarrow \lim_{0 < h \rightarrow 0}$$

$$\frac{\int_t^{t+h} f(x) dx}{h} = \frac{P((t, \infty))}{\lambda}$$

$$\Rightarrow \lim_{0 < h \rightarrow 0}$$

$$\frac{f(t) \cdot h}{h} = \frac{P((t, \infty))}{\lambda}$$

(As $h \rightarrow 0$, area under the curve $f(t)$ from t to $t+h$ becomes equal to area of rectangular strip with height $f(t)$ and width h).

$$\Rightarrow$$

$$f(t) = \frac{\int_t^{\infty} f(x) dx}{\lambda} = \frac{F(\infty) - F(t)}{\lambda} = \frac{1 - F(t)}{\lambda}$$

Taking derivative both sides

$$f'(t) = -\frac{f(t)}{\lambda}$$

$$\Rightarrow \frac{f'(t)}{f(t)} = -\frac{1}{\lambda} \quad \text{for } t > 0$$

$$\Rightarrow \int_0^t \frac{f'(t)}{f(t)} dt = -\frac{1}{\lambda} t + C$$

$$\Rightarrow \log f(t) = -\frac{1}{\lambda} t + C$$

$$\Rightarrow f(t) = e^{-t/\lambda} \cdot e^C$$

Now, since $f(t)$ is a pdf,

$$\int_{-\infty}^{\infty} f(t) dt = 1$$

$$\Rightarrow \int_0^{\infty} e^{-t/\lambda} e^C dt = 1$$

$$\Rightarrow \left[-\lambda e^{-t/\lambda} \right]_0^{\infty} e^C = 1$$

$$\Rightarrow \lambda e^C = 1 \Rightarrow e^C = \frac{1}{\lambda}$$

$$\therefore f(t) = \frac{1}{\lambda} e^{-t/\lambda}$$

(b) Given that the object/individual survives atleast t time units, it is the probability that it fails/dies at time unit t .

It is not reasonable to assume so because the probability of death is slightly higher during birth, then it reduces and starts increasing again after 65 or so years of age.

Solution 3

The Poisson distribution is given by the pmf

$$f(x|\lambda) = \frac{\lambda^x e^{-\lambda}}{x!} \mathbb{1}_{\{0,1,2,\dots\}}(x)$$

$$f(x) = \frac{e^{-1}}{x!} \mathbb{1}_{\{0,1,2,\dots\}}(x)$$

The transformation

$$Y = (X+1)^{-1}$$

\therefore the set $\{x : g(x) = y\}$ is a singleton.

The pdf of Y

$$f_Y(y) = \sum_{x \in g^{-1}(y)} f_X(x) = f_X(\frac{1}{y} - 1)$$

$$\Rightarrow f_y(y) = \frac{e^{-1}}{(\frac{1}{y}-1)!} \quad y = \frac{1}{1+x}, x = 0, 1, 2, \dots$$

$$\Rightarrow f_y(y) = \frac{e^{-1}}{(\frac{1}{y}-1)!} \quad y = 1, \frac{1}{2}, \frac{1}{3}, \dots$$

$$\text{Ansatz für } f_y(y) = \frac{e^{-1}}{(\frac{1}{y}-1)!} \quad y \in \{1, \frac{1}{2}, \frac{1}{3}, \dots\}$$

Verify

$$(i) f_y(y) \geq 0 \quad \forall y \text{ starts positive}$$

$$(ii) \sum_{y \in \{1, \frac{1}{2}, \dots\}} \frac{e^{-1}}{(\frac{1}{y}-1)!} = e^{-1} \sum_{x \in \{0, 1, 2, \dots\}} \frac{1}{x!} \quad \text{--- (1)}$$

To evaluate the summation, we use the hint provided in problem 6 of HW 2,

$$\lim_{n \rightarrow \infty} \left(1 + \frac{\lambda + \epsilon_n}{n}\right)^n = e^\lambda \quad \text{for } \lambda \in \mathbb{R}$$

$$\lim_{n \rightarrow \infty} \epsilon_n = 0$$

For $\lambda = 0, \epsilon_n = 0$ we have

$$\left(1 + \frac{1}{n}\right)^n \xrightarrow{n \rightarrow \infty} \left(1 + \frac{1}{n}\right)^n = e \quad \text{--- (2)}$$

Using binomial expansion, the k^{th} term on left side is given by

$${n \choose k} \frac{1}{n^k} = \frac{n(n-1) \cdots (n-k+1)}{n \cdot n \cdots n} \times \frac{1}{k!}$$

which converges to $\frac{1}{k!}$ in the $\lim_{n \rightarrow \infty}$.

left side of the equation ② can be expressed as

$$\lim_{n \rightarrow \infty} \left(1 + \frac{1}{n}\right)^n = \lim_{n \rightarrow \infty} \sum_{k=0}^n {n \choose k} \frac{1}{n^k}$$

$$= \sum_{k=0}^{\infty} \frac{1}{k!}$$

$$\Rightarrow \sum_{k=0}^{\infty} \frac{1}{k!} = e \quad (\text{using ②}) \quad \text{--- ③}$$

Using ③ in ①

$$e^x \sum_{x \in \{0, 1, 2, \dots\}} \frac{1}{x!} = e^x \cdot e = 1.$$

\therefore It satisfies the criteria of prf.

You can assume this $\sum_{n=0}^{\infty} \frac{1}{n!} = e$. No need to provide that proof.

Solution 4

$$F_Y(y) = P(Y \leq y) = 1 - \frac{1}{y^2} \quad 1 \leq y < \infty$$

(a) Verify that $F_Y(y)$ is a cdf.

(i) $\lim_{y \rightarrow \infty} F_Y(y) = \lim_{y \rightarrow \infty} 1 - \frac{1}{y^2} = 1 - 0 = 1$

$\lim_{y \rightarrow -\infty} F_Y(y) = 0 \quad (\because F_Y = 0 \forall y < 1)$

(ii) For $y_2 > y_1, \quad y_1, y_2 \geq 1$

$$F_Y(y_2) - F_Y(y_1) = -\frac{1}{y_2^2} + \frac{1}{y_1^2} = \frac{y_2^2 - y_1^2}{y_1^2 y_2^2} > 0$$

$\therefore F_Y(y)$ is non decreasing fn of y .

(iii) $\lim_{y_0 < y \rightarrow y_0} F_Y(y) = \lim_{y_0 < y \rightarrow y_0} 1 - \frac{1}{y^2}$

$$y_0 \geq 1$$

$$= \lim_{0 < h \rightarrow 0} 1 - \frac{1}{(y_0+h)^2} = 1 - \frac{1}{y_0^2} = F_Y(y_0)$$

$\therefore F_Y(y)$ is right continuous.

at least on $1 < y < \infty$ and $y_0 > 1$

$\Rightarrow F_Y(y)$ is a cdf. *justify both showing*

(b) Find $f_y(y)$

$$y > 1 \quad f_y(y) = \frac{d}{dy} F_y(y) = -\frac{d}{dy} (y^{-2}) = -(-2y^{-3}) \\ = \frac{2}{y^3}$$

✓

$$y \leq 1 \quad f_y(y) = 0$$

(c) $Z = 10(Y-1)$

Clearly $Z = g(y)$ is one to one increasing function
 and random variable Y is continuous (absolutely)
 with $F_Y^1(y) = f_y(y)$ for $y > 1$.

Using the change of variable formula for cdf

$$F_Z(z) = F_Y(g^{-1}(y))$$

$$= F_Y\left(\frac{z}{10} + 1\right)$$

$$= 1 - \frac{1}{\left(\frac{z}{10} + 1\right)^2}$$

$$z = g(y), y \geq 1$$

$$\Rightarrow \frac{z}{10} + 1 \geq 1 \\ \Rightarrow z \geq 0$$

$$= 1 - \frac{1}{\left(\frac{z}{10} + 1\right)^2}, z \geq 0$$

$$\text{and } F_Z(z) = 0, z < 0$$

Solution 5

$$T \leq 0 \Rightarrow V < 5$$

$$P(V \leq v) = 0 \quad v < 5$$

$$0 < T < 3 \Rightarrow V = 5$$

$$P(V \leq 5 | T \geq 3) P(T \geq 3)$$

$$\begin{aligned} P(V \leq v) &= P(V \leq 5 | 0 < T < 3) P(0 < T < 3) + \\ &= [P(V < 5 | 0 < T < 3) + P(V = 5 | 0 < T < 3)] P(0 < T < 3) \\ &= P(0 < T < 3) \\ &= \int_0^3 \frac{1}{1.5} e^{-t/1.5} dt \\ &= -\frac{1.5}{1.5} [e^{-2} - e^0] \\ &= 1 - e^{-2} \end{aligned}$$

$$T \geq 3 \Rightarrow V \geq 6$$

$$\begin{aligned} P(V \leq v) &= P(V \leq v | 0 < T < 3) P(0 < T < 3) + \\ &\quad P(V \leq v | T \geq 3) P(T \geq 3) \\ &= P(0 < T < 3) + P(T \leq \frac{v}{2} | T \geq 3) P(T \geq 3) \\ &= P(0 < T < 3) + \frac{P(3 \leq T \leq \frac{v}{2})}{P(T \geq 3)} P(T \geq 3) \\ &= P(0 < T < 3) + P(3 \leq T \leq \frac{v}{2}) \end{aligned}$$

$$= P(0 < T \leq \frac{v}{2}), \quad , \frac{v}{2} \geq 3$$

$$= \int_0^{\frac{v}{2}} \frac{1}{1.5} e^{-t/1.5} dt, \quad , v \geq 6$$

$$= -\frac{1.5}{1.5} [e^{-v/3} - e^0], \quad , v \geq 6$$

$$= 1 - e^{-v/3}, \quad , v \geq 6$$

$$5 < v < 6$$

$$P(V \leq v) = P(V \leq 5) \quad \therefore P(V = v) = 0 \quad \forall v > 5$$

$$F_V(v) = \begin{cases} 0 & v < 5 \\ 1 - e^{-2} & 5 \leq v < 6 \\ 1 - e^{-v/3} & v \geq 6 \end{cases}$$

Solution 6

$$(a) Y = X^3, f_X(x) = 42x^5(1-x) \quad 0 < x < 1$$

One to one transformation, continuous random variable X

$$g^{-1}(y) = y^{1/3} \quad 0 < y < 1$$

Using change of variable formula for pdf,

$$f_Y(y) = \begin{cases} f_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right| & 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$

$$\frac{d}{dy} g^{-1}(y) = \frac{1}{3} y^{-2/3}$$

$$f_Y(y) = \begin{cases} 42 y^{5/3} (1-y^{1/3}) \frac{1}{3} y^{-2/3} & 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$

$$= \begin{cases} 14y(1-y^{1/3}) & 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$

Verify

$$\int_{-\infty}^{\infty} f_Y(y) dy = \int_0^1 (14y - 14y^{4/3}) dy = \left[\frac{14}{2} y^2 - \frac{14}{7} y^{7/3} \right]_0^1 = 7 - 6 = 1$$

$$(b) Y = 4X + 3 \quad f_X(x) = 7e^{-7x} \quad x > 0$$

One to one transformation, continuous r.v. X

$$\Rightarrow g^{-1}(y) = \frac{y-3}{4} \quad \Rightarrow \frac{d}{dy} g^{-1}(y) = \frac{1}{4} \quad 3 < y < \infty$$

Using change of variable formula for pdf.

$$f_Y(y) = \begin{cases} f_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right| & 3 < y < \infty \\ 0 & \text{otherwise} \end{cases}$$

~~$$f_Y(y) = \begin{cases} \frac{7}{4} \exp\left(-\frac{7}{4}(y-3)\right) & 3 < y < \infty \\ 0 & \text{otherwise} \end{cases}$$~~

Verify

$$\int_3^{\infty} \frac{7}{4} \exp\left(-\frac{7}{4}(y-3)\right) dy = \frac{7}{4} \left[\exp\left(-\frac{7}{4}(y-3)\right) \right]_3^{\infty}$$

$$= -\frac{7}{4} \left[0 - 1 \right] = 1$$

$$(c) Y = X^2 \quad f_X(x) = 30x^2(1-x)^2 \quad 0 < x < 1 \quad (d)$$

For given range $0 < x < 1$, the transformation is one to one. Also X is continuous r.v.

$$g^{-1}(y) = \sqrt{y} \quad \frac{d}{dy} g^{-1}(y) = \frac{1}{2} y^{-1/2} \quad 0 < y < 1$$

Using change of variable formula for pdf.

$$f_Y(y) = \begin{cases} 30y(1-\sqrt{y})^2 \cdot \frac{1}{2} y^{-1/2} & 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$

$$= \begin{cases} 15 \cdot (1-\sqrt{y})^2 \sqrt{y} & 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$

Verify

$$\int_0^1 15(y^{1/2} + yy^{1/2} - 2y^{1/2}y^{1/2}) dy$$

$$= 15 \int_0^1 (y^{1/2} + y^{3/2} - 2y) dy = 15 \left[\frac{y^{3/2}}{\frac{3}{2}} + \frac{y^{5/2}}{\frac{5}{2}} - y^2 \right]_0^1$$

$$= 15 \left(\frac{2}{3} + \frac{2}{5} - 1 \right) = 10 + 6 - 15 = 1$$

Solution 6

MATLAB Live script for Solution 6 for verification using MATLAB Symbolic Toolbox.

```
% 1. isInc is 1 or -1 depending on whether the  
% transformation is an increasing or  
% decreasing function.  
% 2. l1y, l2y are lower and upper limits for  
% non-zero pdf of Y.  
% 3. check indicates whether pdf of Y  
% integrates to 1.  
x = sym('x');  
X = sym('X');  
% Part (a)  
disp("Solution 6(a)")
```

Solution 6(a)

$$f_X = 42x^5(1-x)$$

$$f_X = -42x^5(x-1)$$

$$Y = X^3$$

$$Y = X^3$$

```
isInc = 1;  
l1y = 0; l2y = 1;  
[fY,check] = find_pdf(fX,Y,X,isInc,l1y,l2y)
```

$$f_Y = -14y(y^{1/3} - 1)$$

$$\text{check} = 1$$

```
% Part (b)  
disp("Solution 6(b)")
```

Solution 6(b)

$$f_X = 7e^{-7x}$$

$$f_X = 7e^{-7x}$$

$$Y = 4X + 3$$

```
Y = 4X + 3  
  
isInc = 1;  
l1y = 3; l2y = inf;  
[fY,check] = find_pdf(fX,Y,X,isInc,l1y,l2y)
```

$$f_Y =$$

$\frac{7e^{\frac{21-7y}{4}}}{4}$
 check = 1

```
% Part (c)
disp("Solution 6(c)")
```

Solution 6(c)

$$f_X = 30x^2(1-x)^2$$

$$f_X = 30x^2 (x-1)^2$$

$$Y = X^2$$

$$Y = X^2$$

```
isInc = 1;
l1y = 0; l2y = 1;
[fY,check] = find_pdf(fX,Y,X,isInc,l1y,l2y)
```

$$f_Y = 15\sqrt{y} (\sqrt{y}-1)^2$$

check = 1

Following the function definition for finding pdf of Y:

```
function [fY,check] = find_pdf(fX,Y,X,isInc,l1y,l2y)
g = finverse(Y);
y = sym('y');
g = subs(g,X,y);
fY = compose(fX,g)*diff(g)*isInc;
check = int(fY,y,l1y,l2y);
end
```

Good Work !! Keep it up !!