

## Appendix B: Mathematical Details and Derivations

### B.1 Bernoulli Model and Likelihood

Each segment is represented as  $(x_i, y_i)$ , where  $x_i \in R^{38}$  is the feature vector and  $y_i \in \{0, 1\}$  is the case-level group label inherited at the segment level.

CourtShadow models the label as a Bernoulli random variable:

$$y_i \sim \text{Bernoulli}(p_i), \quad p_i = P(y_i = 1 | x_i).$$

Assuming conditional independence given  $x_i$  and parameter vector  $\theta$ , the likelihood of the entire dataset is:

$$L(\theta) = \prod_{i=1}^n P(y_i | x_i; \theta) = \prod_{i=1}^n p_i^{y_i} (1 - p_i)^{1-y_i}.$$

Taking logs, the log-likelihood is:

$$\ell(\theta) = \log L(\theta) = \sum_{i=1}^n [y_i \log p_i + (1 - y_i) \log(1 - p_i)].$$

The model uses a logistic link to parameterize  $p_i$ .

### B.2 Logistic Link and Negative Log-Likelihood

The logistic (sigmoid) function is:

$$\sigma(z) = \frac{1}{1 + e^{-z}}.$$

CourtShadow parameterizes the Bernoulli probability as:

$$p_i = \sigma(\theta^\top x_i) = \frac{1}{1 + e^{-\theta^\top x_i}}.$$

Substituting this into the log-likelihood, the *negative* log-likelihood (NLL) used as the loss is:

$$J(\theta) = -\ell(\theta) = -\sum_{i=1}^n [y_i \log \sigma(\theta^\top x_i) + (1 - y_i) \log(1 - \sigma(\theta^\top x_i))].$$

This is the standard cross-entropy loss for logistic regression.

### B.3 Gradient of the Logistic Loss

Define  $z_i = \theta^\top x_i$  and  $p_i = \sigma(z_i)$ . We compute the gradient of  $J(\theta)$  with respect to  $\theta$ .

First, observe:

$$\frac{\partial p_i}{\partial z_i} = \frac{\partial}{\partial z_i} \left( \frac{1}{1 + e^{-z_i}} \right) = p_i(1 - p_i).$$

By the chain rule,

$$\frac{\partial p_i}{\partial \theta} = \frac{\partial p_i}{\partial z_i} \frac{\partial z_i}{\partial \theta} = p_i(1 - p_i) x_i.$$

Now differentiate the loss:

$$J(\theta) = - \sum_{i=1}^n [y_i \log p_i + (1 - y_i) \log(1 - p_i)].$$

Taking  $\nabla_\theta$ ,

$$\nabla_\theta J(\theta) = - \sum_{i=1}^n \left[ \frac{y_i}{p_i} \frac{\partial p_i}{\partial \theta} - \frac{1 - y_i}{1 - p_i} \frac{\partial p_i}{\partial \theta} \right].$$

Substitute  $\frac{\partial p_i}{\partial \theta} = p_i(1 - p_i)x_i$ :

$$\nabla_\theta J(\theta) = - \sum_{i=1}^n \left[ \frac{y_i}{p_i} p_i(1 - p_i)x_i - \frac{1 - y_i}{1 - p_i} p_i(1 - p_i)x_i \right].$$

Simplify:

$$\nabla_\theta J(\theta) = - \sum_{i=1}^n [y_i(1 - p_i)x_i - (1 - y_i)p_i x_i].$$

Rearrange the terms inside:

$$y_i(1 - p_i) - (1 - y_i)p_i = y_i - y_i p_i - p_i + y_i p_i = y_i - p_i.$$

So:

$$\nabla_\theta J(\theta) = - \sum_{i=1}^n (y_i - p_i)x_i = \sum_{i=1}^n (p_i - y_i)x_i.$$

This is the gradient used by gradient-based optimizers.

## B.4 L2-Regularized Objective

To reduce overfitting on a small dataset and keep weights in a numerically stable regime, CourtShadow adds an L2 penalty term:

$$\Omega(\theta) = \lambda \sum_j \theta_j^2.$$

The regularized objective becomes:

$$J_{L2}(\theta) = J(\theta) + \Omega(\theta) = -\sum_{i=1}^n [y_i \log p_i + (1 - y_i) \log(1 - p_i)] + \lambda \sum_j \theta_j^2.$$

Differentiating the penalty term:

$$\nabla_\theta \Omega(\theta) = 2\lambda \theta.$$

Thus the gradient of the regularized loss is:

$$\nabla_\theta J_{L2}(\theta) = \sum_{i=1}^n (p_i - y_i)x_i + 2\lambda \theta.$$

In practice, many numerical packages absorb the factor of 2 into  $\lambda$ , but the idea is the same: larger weights incur larger penalties, shrinking coefficients and improving generalization on held-out data.

## B.5 Feature Scaling and Its Effect

Continuous features are standardized using

$$x' = \frac{x - \mu}{\sigma},$$

where  $\mu$  and  $\sigma$  are computed on the training set. Substituting  $x'$  into  $\theta^\top x$  yields:

$$\theta^\top x' = \sum_j \theta_j \frac{x_j - \mu_j}{\sigma_j}.$$

This has two benefits:

1. Features on different scales (e.g., token counts vs. rates) contribute comparably to the decision boundary.
2. The magnitude of  $\theta_j$  becomes more interpretable: it represents the effect of a one-standard-deviation change in feature  $j$ .

Binary topic indicators are left unscaled to preserve their direct “on/off” interpretation.

## B.6 Case-Level Aggregation

Segment-level probabilities  $p_j$  are aggregated to form a case-level *Linguistic Environment Score* (LES):

$$\bar{p}_{case} = \frac{1}{m} \sum_{j=1}^m p_j,$$

where  $m$  is the number of segments in a case.

From a statistical perspective,  $\bar{p}_{case}$  approximates the expected probability that a randomly sampled segment from that case is classified as Group 1 by the model. This aggregation reduces within-case noise (e.g., one unusually harsh turn) and focuses on the overall environment.

## B.7 Linear Contributions and Feature Families

Because logistic regression is linear in feature space, we can write the log-odds for a segment as:

$$\theta^\top x = \sum_{k=1}^d \theta_k x_k.$$

If we partition the indices  $1, \dots, d$  into disjoint families (structure, framing, pronouns, topics), the total log-odds decomposes as:

$$\theta^\top x = \underbrace{\sum_{k \in \text{structure}} \theta_k x_k}_{\text{structure}} + \underbrace{\sum_{k \in \text{framing}} \theta_k x_k}_{\text{framing}} + \underbrace{\sum_{k \in \text{pronouns}} \theta_k x_k}_{\text{pronouns}} + \underbrace{\sum_{k \in \text{topics}} \theta_k x_k}_{\text{topics}}.$$

This decomposition is used in the website's interpretability plots to show how each family pushes a segment or case toward Group A or Group B. Case-level family contributions are obtained by averaging these family sums across all segments in a case.

## B.8 ROC AUC and Calibration

The ROC area under the curve (AUC) is defined as:

$$AUC = P(s_{pos} > s_{neg}),$$

where  $s_{pos}$  and  $s_{neg}$  are scores for randomly chosen positive (Group 1) and negative (Group 0) examples. In practice, AUC is computed by ranking cases by  $\bar{p}_{case}$  and computing the fraction of correctly ordered positive-negative pairs.

Calibration is evaluated by binning predicted probabilities and comparing them to empirical frequencies:

$$\text{calibration}(\text{bin}) = E[Y \mid \hat{p} \in \text{bin}],$$

where  $\hat{p}$  is the model's predicted probability. If the model is well-calibrated, the reliability curve (empirical vs. predicted) lies near the diagonal. In CourtShadow, these diagnostics help confirm that the logistic probabilities are meaningful as *degrees of belief* about Group A environments, not merely ranking scores.