STAT 135 15. Multiple linear regression and inference

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Inference for matrix formulation of linear regression

Matrix notation

$$Y = X\beta + \epsilon$$

$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ \vdots \\ y_n \end{bmatrix} = \begin{bmatrix} 1 & x_{1,1} & x_{1,2} & \dots & x_{1,p-1} \\ 1 & x_{2,1} & x_{2,2} & \dots & x_{2,p-1} \\ 1 & x_{3,1} & x_{3,2} & \dots & x_{3,p-1} \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ 1 & x_{n,1} & x_{n,2} & \dots & x_{n,p-1} \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_{p-1} \\ (p \times 1) \end{bmatrix} + \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \vdots \\ \epsilon_n \end{bmatrix}$$

$$(n \times 1)$$

The LS prediction is

$$|\widehat{\boldsymbol{\beta}} = (X^T X)^{-1} X^T Y|$$

 $sale_price_i = \beta_0 + \beta_1 area_i + \beta_2 quality_i + \beta_3 year_i + \beta_4 beds_i + \epsilon_i$

$$Y = X\beta + \epsilon$$

$$\begin{bmatrix} \frac{177000}{130000} \\ \frac{130000}{130000} \\ \frac{88000}{140000} \\ \frac{125000}{125000} \\ \frac{125000}{127000} \\ \frac{155000}{155000} \\ \frac{105000}{98000} \end{bmatrix} = \begin{bmatrix} 1 & 1913 & 6 & 1928 & 3 \\ 1 & 1709 & 6 & 2004 & 3 \\ 1 & 858 & 4 & 1971 & 3 \\ 1 & 1164 & 7 & 1969 & 3 \\ 1 & 960 & 5 & 1958 & 3 \\ 1 & 1376 & 6 & 1920 & 3 \\ 1 & 1484 & 8 & 2006 & 2 \\ 1 & 1484 & 8 & 2006 & 2 \\ 1 & 1200 & 7 & 2004 & 2 \\ 1 & 1436 & 4 & 1910 & 3 \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_{p-1} \\ (p \times 1) \end{bmatrix} + \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \vdots \\ \epsilon_n \end{bmatrix}$$

The LS prediction is

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$$\widehat{\boldsymbol{\beta}} = \left(X^T X \right)^{-1} X^T Y$$

$$Y = \begin{bmatrix} \frac{177000}{130000} \\ \frac{130000}{88000} \\ \frac{140000}{140000} \\ \frac{125000}{125000} \\ \frac{165000}{265900} \\ \frac{127000}{155000} \\ \frac{155000}{98000} \end{bmatrix}$$

$$X = \begin{bmatrix} 1 & 1913 & 6 & 1928 & 3 \\ 1 & 1709 & 6 & 2004 & 3 \\ 1 & 858 & 4 & 1971 & 3 \\ 1 & 1164 & 7 & 1969 & 3 \\ 1 & 960 & 5 & 1958 & 3 \\ 1 & 1376 & 6 & 1920 & 3 \\ 1 & 1484 & 8 & 2006 & 2 \\ 1 & 864 & 5 & 1972 & 3 \\ 1 & 1200 & 7 & 2004 & 2 \\ 1 & 1436 & 4 & 1910 & 3 \end{bmatrix}$$

$$(n \times 1)$$

$$X^{T} = \begin{bmatrix} 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 \\ 1913 & 1709 & 858 & 1164 & 960 & 1376 & 1484 & 864 & 1200 & 1436 \\ 6 & 6 & 4 & 7 & 5 & 6 & 8 & 5 & 7 & 4 \\ 1928 & 2004 & 1971 & 1969 & 1958 & 1920 & 2006 & 1972 & 2004 & 1910 \\ 3 & 3 & 3 & 3 & 3 & 3 & 2 & 3 & 2 & 3 & 2 \end{bmatrix}$$

The LS prediction is

$$\widehat{\boldsymbol{\beta}} = \left(X^T X \right)^{-1} X^T Y$$

$$Y = \begin{bmatrix} \frac{177000}{130000} \\ \frac{180000}{88000} \\ \frac{140000}{140000} \\ \frac{125000}{125000} \\ \frac{165000}{265900} \\ \frac{127000}{155000} \\ \frac{155000}{98000} \end{bmatrix} \qquad X = \begin{bmatrix} \frac{1}{1} & 1913 & 6 & 1928 & 3 \\ 1 & 1709 & 6 & 2004 & 3 \\ 1 & 858 & 4 & 1971 & 3 \\ 1 & 1164 & 7 & 1969 & 3 \\ 1 & 960 & 5 & 1958 & 3 \\ 1 & 1376 & 6 & 1920 & 3 \\ 1 & 1484 & 8 & 2006 & 2 \\ 1 & 864 & 5 & 1972 & 3 \\ 1 & 1200 & 7 & 2004 & 2 \\ 1 & 1436 & 4 & 1910 & 3 \end{bmatrix}$$

$$(n \times p)$$

The LS prediction is
$$\left| \widehat{\boldsymbol{\beta}} = \left(X^T X \right)^{-1} X^T Y \right|$$

$$Y = \begin{bmatrix} 177000 \\ 130000 \\ 88000 \\ 140000 \\ 125000 \\ 165000 \\ 265900 \\ 127000 \\ 155000 \\ 98000 \end{bmatrix}$$
 $(n \times 1)$

$$\hat{\beta} = (X^T X)^{-1} X^T Y =$$
int
living_area
quality_score
year_built
bedrooms

[731418.02374]
18.97778
26094.62901
-330.71265
-39533.52530

$\hat{\beta}$ is unbiased

$\hat{\beta}$ is unbiased

Assume that the errors, ϵ_i , are IID with $E(\epsilon_i) = 0$ and $Var(\epsilon_i) = \sigma^2$

This means that $Cov(\epsilon) = \sigma^2 I_{n\times n}$ and $E(\epsilon) = 0$

The LS estimator is:

$$\hat{\beta} = (X^T X)^{-1} X^T Y$$

The LS estimate, $\hat{\beta}$ is unbiased:

$$E[\hat{\beta}] = \beta$$

Proof:

$$E(\widehat{\boldsymbol{\beta}}) = E((X^T X)^{-1} X^T Y)$$

$$= (X^T X)^{-1} X^T E(Y)$$

$$= (X^T X)^{-1} X^T E(X \boldsymbol{\beta} + \boldsymbol{\epsilon})$$

$$= (X^T X)^{-1} X^T X \boldsymbol{\beta}$$

$$= \boldsymbol{\beta}$$

Cov(B)

\hat{eta} variance-covariance matrix

Assume that the errors, ϵ_i , are IID with $E(\epsilon_i) = 0$, $Var(\epsilon_i) = \sigma^2$ This means that $Cov(\epsilon) = \sigma^2 I_{n \times n}$ and $E(\epsilon) = 0$ Note if A is a constant matrix and Y is a random vector, then:

$$Cov(AY) = ACov(Y)A^{T}$$

 $Cov(A + Y) = Cov(Y)$

The LS estimator is:

$$\hat{\beta} = (X^T X)^{-1} X^T Y$$

The LS estimate, $\hat{\beta}$ is has variance-covariance matrix:

$$Cov(\hat{\beta}) = \sigma^2(X^TX)^{-1}$$

$$(p \times p)$$

Proof:

$$Cov\left(\widehat{\boldsymbol{\beta}}\right) = Cov\left(\left(X^{T}X\right)^{-1}X^{T}Y\right)$$

$$= \left(X^{T}X\right)^{-1}X^{T}Cov(Y)X\left(X^{T}X\right)^{-1}$$

$$= \left(X^{T}X\right)^{-1}X^{T}Cov(X\boldsymbol{\beta} + \boldsymbol{\epsilon})X\left(X^{T}X\right)^{-1}$$

$$= \left(X^{T}X\right)^{-1}X^{T}Cov(\boldsymbol{\epsilon})X\left(X^{T}X\right)^{-1}$$

$$= \left(X^{T}X\right)^{-1}X^{T}\sigma^{2}I_{n\times n}X\left(X^{T}X\right)^{-1}$$

$$= \sigma^{2}\left(X^{T}X\right)^{-1}$$

The distribution of $\hat{\beta}$

The theoretical distribution of the \hat{eta} vector

Assume that the errors, $\epsilon_i \stackrel{\mathit{IID}}{\sim} N(0,\sigma^2)$ i.e., $\epsilon \stackrel{\mathit{IID}}{\sim} N(0,\sigma^2 I_{n\times n})$

The LS estimator is:

$$\hat{\beta} = (X^T X)^{-1} X^T Y$$

And we have shown that:

$$E[\hat{\beta}] = \beta \qquad Cov(\hat{\beta}) = \sigma^2(X^TX)^{-1}$$

Because the components of $\hat{\beta}$ are linear combinations of independent normal RVs, they are themselves normal, and so

The LS estimate, $\hat{\beta}$ is has distribution:

$$\hat{\beta} \sim N\left(\beta, \sigma^2(X^TX)^{-1}\right)$$

This means that each element

$$\hat{\beta}_j \sim N(\beta_j, \sigma^2 c_{jj})$$

(Where
$$C = (X^T X)^{-1}$$
)

But this assumes we know $\hat{\sigma}^2$

Distribution of $\hat{\beta_j}$ when estimating variance

Assume that the errors, $\epsilon_i \stackrel{\mathit{IID}}{\sim} N(0,\sigma^2)$ i.e., $\epsilon \stackrel{\mathit{IID}}{\sim} N(0,\sigma^2 I_{n\times n})$

We have that $\hat{\beta}_j \sim N(\beta_j, \sigma^2 c_{jj})$, where $C = (X^T X)^{-1}$ (c_{jj} is the jth diagonal entry)

If we estimate σ^2 with some $\hat{\sigma}^2$, then we can estimate the variance of $\hat{\beta}_j$ as:

$$\hat{\sigma}_{\hat{\beta}_i}^2 = \hat{\sigma}^2 c_{jj}$$

Then, we have

$$\frac{\hat{\beta}_{j} - \beta_{j}}{\hat{\sigma}_{\hat{\beta}_{i}}} \sim t_{n-p}$$

But what should $\hat{\sigma}^2$ be?

Estimating σ^2 Assume that the errors, $\epsilon_i \stackrel{IID}{\sim} N(0, \sigma^2)$ i.e., $\epsilon \stackrel{IID}{\sim} N(0, \sigma^2 I_{n \times n})$

The variance-covariance matrix is: $Cov(\hat{\beta}) = \sigma^2(X^TX)^{-1}$

Similarly to the 1-dim case, we can estimate σ^2 using the RSS Let $r = Y - X\hat{\beta}$ be the vector of residuals

Then, an unbiased estimator for σ^2 is:

$$\hat{\sigma}^2 = \frac{RSS}{n-p} = \frac{\|r\|_2^2}{n-p} = \frac{\sum_i r_i^2}{n-p}$$

The proof of unbiasedness requires some understanding of matrix properties such as the *trace*. If you're curious, check out Section 14.4.3 of Rice.

$$\hat{Y} = \hat{X}\hat{\beta} = \hat{X}\hat{\beta} = \begin{bmatrix} \frac{1}{1} & \frac{1913}{1769} & 6 & \frac{1928}{2004} & \frac{3}{3} \\ \frac{1}{1} & \frac{1709}{1709} & 6 & \frac{2004}{2004} & \frac{3}{3} \\ \frac{1}{1} & \frac{15858}{1709} & 4 & \frac{1971}{3} & \frac{3}{3} \\ \frac{1}{1} & \frac{15858}{1709} & 4 & \frac{1971}{3} & \frac{3}{3} \\ \frac{1}{1} & \frac{1564}{1709} & 7 & \frac{1969}{3} & \frac{3}{3} \\ \frac{1}{1} & \frac{1376}{1800} & 6 & \frac{1920}{3} & \frac{3}{3} \\ \frac{1}{1} & \frac{1376}{1800} & 6 & \frac{1920}{3} & \frac{3}{3} \\ \frac{1}{1} & \frac{15800}{1800} & \frac{1972}{3} & \frac{3}{3} \\ \frac{1}{1} & \frac{1200}{1436} & \frac{7}{1969} & \frac{3}{3} \\ \frac{1}{1} & \frac{1960}{1436} & \frac{7}{1969} & \frac{3}{1969} & \frac{3}{1969} & \frac{3}{1969} \\ \frac{1}{1} & \frac{1960}{1436} & \frac{7}{1969} & \frac{3}{1969} & \frac{3}{1969} & \frac{3}{1969} & \frac{3}{19$$

<u>177</u>000

130000

<u>88</u>000

<u>98</u>000

The residual vector is:

$$Y = Y - \hat{Y} = \begin{bmatrix} \frac{65000}{140000} \\ \frac{125000}{165000} \\ \frac{165000}{127000} \\ \frac{127000}{155000} \\ \frac{155000}{00000} \end{bmatrix}$$

$$Y = Y - \hat{Y} = \begin{pmatrix} \frac{177000}{130000} \\ \frac{130}{30000} \\ \frac{88000}{88000} \\ \frac{140000}{125000} \\ \frac{165000}{265900} \\ \frac{127000}{127000} \\ \frac{155000}{98000} \end{pmatrix} = \begin{pmatrix} \frac{168075.72}{139070.09} \\ \frac{8924.283}{-9070.088} \\ \frac{6355.741}{-26396.771} \\ \frac{113973.89}{160530.35} \\ \frac{225861.45}{107522.04} \\ \frac{19477.958}{195038.554} \\ \frac{-14786.886}{-14786.886} \end{pmatrix}$$

We can estimate σ^2 using:

$$\hat{\sigma}^2 = \frac{\|r\|_2^2}{n-p} = \frac{\sum_i r_i^2}{n-p} = \frac{8924.3^2 + 9070.1^2 + \dots + 14786.9^2}{10-5} = 968,972,338$$

Which we can plug into:

$$\hat{\sigma}_{\hat{\beta}_{j}}^{2} = \hat{\sigma}^{2} c_{jj} \quad \cdots \quad C = (X^{T}X)^{-1} = \begin{bmatrix} \text{int} \\ \text{living_area} \\ \text{quality_score} \\ \text{year_built} \\ \text{bedrooms} \end{bmatrix} \begin{bmatrix} 722.3664 & -0.0109 & 2.2965 & -0.3457 & -15.1243 \\ -0.0109 & 0.0000 & -0.0002 & 0.0000 & -0.0001 \\ 2.2965 & -0.0002 & 0.1667 & -0.0019 & 0.2462 \\ -0.3457 & 0.0000 & -0.0019 & 0.0002 & 0.0051 \\ -0.3457 & -0.0001 & 0.2462 & 0.0051 & 1.3988 \end{bmatrix}$$

int living_area quality_score year_built bedrooms

So the SD estimate for $\hat{\beta}_{quality}$ is:

$$\hat{\sigma}_{\hat{\beta}_{quality}}^{2} = \hat{\sigma}^{2} c_{quality,quality} = 968,972,338 \times 0.1667 = 161,527,689$$

$$\implies \hat{\sigma}_{\hat{\beta}_{quality}} = 12,709$$

Hypothesis testing and confidence intervals for $\hat{\beta}$

Hypothesis testing for β_j

Assume that the errors, $\epsilon_i \stackrel{\mathit{IID}}{\sim} N(0,\sigma^2)$ i.e., $\epsilon \stackrel{\mathit{IID}}{\sim} N(0,\sigma^2 I_{n\times n})$

$$H_0: \beta_j = 0 \text{ against } H_1: \beta_j \neq 0$$

Test statistic:
$$t = \frac{\hat{\beta}_j}{\hat{\sigma}_{\hat{\beta}_i}}$$
 P-value: $P(|T| \ge |t|)$

Where $T \sim t_{n-p}$ (p is the number of parameters in the model)

These standardized coefficients are comparable to one another

Where

$$\hat{\sigma}_{\hat{\beta}_{j}}^{2} = \hat{\sigma}^{2} c_{jj} \qquad \text{And} \qquad \hat{\sigma}^{2} = \frac{RSS}{n-p} = \frac{\|r\|_{2}^{2}}{n-p} = \frac{\|Y - X\hat{\beta}\|_{2}^{2}}{n-p}$$

Confidence interval for β_j

Assume that the errors, $\epsilon_i \stackrel{\mathit{IID}}{\sim} N(0,\sigma^2)$ i.e., $\epsilon \stackrel{\mathit{IID}}{\sim} N(0,\sigma^2 I_{n\times n})$

A
$$(1 - \alpha)$$
 % confidence interval is:

$$[\hat{\beta}_{j} - t_{n-p,\alpha/2}\hat{\sigma}_{\hat{\beta}_{j}}, \hat{\beta}_{j} + t_{n-p,\alpha/2}\hat{\sigma}_{\hat{\beta}_{j}}]$$

Where

$$\hat{\sigma}_{\hat{\beta}_{j}}^{2} = \hat{\sigma}^{2} c_{jj}$$
 And $\hat{\sigma}^{2} = \frac{RSS}{n-p} = \frac{\|r\|_{2}^{2}}{n-p} = \frac{\|Y - X\beta\|_{2}^{2}}{n-p}$

$$C = (X^T X)^{-1} = \begin{bmatrix} \text{int living_area quality_score year_built bedrooms} \\ 722.3664 & -0.0109 & 2.2965 & -0.3457 & -15.1243 \\ -0.0109 & 0.0000 & -0.0002 & 0.0000 & -0.0001 \\ 2.2965 & -0.0002 & 0.1667 & -0.0019 & 0.2462 \\ -0.3457 & 0.0000 & -0.0019 & 0.0002 & 0.0051 \\ -15.1243 & -0.0001 & 0.2462 & 0.0051 & 1.3988 \end{bmatrix}$$

$$\hat{\sigma}^2 = \frac{\|r\|_2^2}{n-p} = 968,972,338$$

$$\hat{\beta}_{j} \qquad (\hat{\sigma}_{\hat{\beta}_{j}}^{2} = \hat{\sigma}^{2}c_{jj}) \qquad (t_{\hat{\beta}_{j}} = \hat{\beta}_{j}/\hat{\sigma}_{\hat{\beta}_{j}}) \qquad P(|T| \ge |t|)$$

$$\hat{\beta}_{int} = 731,418.02 \qquad \hat{\sigma}_{\hat{\beta}_{int}} = 836,632.00 \qquad t_{\hat{\beta}_{int}} = 0.874 \qquad p_{\hat{\beta}_{int}} = 0.4220$$

$$\hat{\beta}_{area} = 18.98 \qquad \hat{\sigma}_{\hat{\beta}_{area}} = 35.12 \qquad t_{\hat{\beta}_{area}} = 0.540 \qquad p_{\hat{\beta}_{area}} = 0.6122$$

$$\hat{\beta}_{quality} = 26,094.63 \qquad \hat{\sigma}_{\hat{\beta}_{quality}} = 12,708.79 \qquad t_{\hat{\beta}_{quality}} = 2.053 \qquad p_{\hat{\beta}_{quality}} = 0.0952$$

$$\hat{\beta}_{year} = -330.71 \qquad \hat{\sigma}_{\hat{\beta}_{year}} = 406.57 \qquad t_{\hat{\beta}_{year}} = -0.813 \qquad p_{\hat{\beta}_{year}} = 0.4530$$

$$\hat{\beta}_{beds} = -39,533.53 \qquad \hat{\sigma}_{\hat{\beta}_{beds}} = 36,816.08 \qquad t_{\hat{\beta}_{beds}} = -1.074 \qquad p_{\hat{\beta}_{beds}} = 0.3320$$

```
> ls_fit <- lm(sale_price ~ ., ames_train)</pre>
> summary(ls_fit)
Call:
lm(formula = sale_price \sim ., data = ames_train)
Residuals:
                                                19478 -40039 -14787
  8924 -9070
               6356 -26397 11026
                                  4470 40039
Coefficients:
                  Estimate Std. Error t value Pr(>|t|)
                 731418.02
                           836632.00
                                       0.874
                                               0.4220
(Intercept)
total_living_area
                     18.98
                                35.12
                                       0.540
                                               0.6122
                                               0.0952 .
quality_score
                  26094.63
                             12708.79 2.053
year_built
                -330.71
                               406.57
                                      -0.813
                                               0.4530
bedrooms
                 -39533.53 36816.08 -1.074
                                               0.3320
Signif. codes:
               0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 31130 on 5 degrees of freedom
Multiple R-squared: 0.785, Adjusted R-squared: 0.6129
F-statistic: 4.563 on 4 and 5 DF, p-value: 0.06353
```

Im_inference_multi.R