9.S918: Statistical Inference in Brain and Cognitive Sciences

Week 2 Day 1: Introduction to causal inference

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Introductory causal inference

- You have probably had previous exposure to both probability and statistics
- You are less likely to have had exposure to causal inference
- Causal inference uses probability and statistics, but it is something separate from the traditional construal of those two fields
- You can think of causal inference as being a framework extending more traditional statistics by:
 - Adding new probability-based mathematical constructs; and,
 - Developing a set of practice for statistical inference based on those constructs
- Two causal inference frameworks:
 - The potential outcomes framework
 - The causal graphical models framework

The potential-outcomes framework

- In epidemiology and many other areas of statistics, causal inference was developed out of the idea of potential outcomes (Neyman 1923, Rubin 1974)
- Consider an outcome, Y, and a potential treatment A

Example:

Y: an individual survives to the end of the year (0: no, 1: yes)

A: an individual with heart disease receives a heart transplant (0: no, 1: yes)

Potential-outcome random variables

- Suppose that A is discrete; for this case, $A \in \{0,1\}$
- The **potential outcomes**, or **counterfactual outcomes**, are random variables for Y for each potential value of A

$$Y^{a=0}$$
 The value that Y would take if A were 0 $Y^{a=1}$ The value that Y would take if A were 1

 Counterfactual risk is the expected value of each counterfactual-outcome random variable:

$$E[Y^{a=0}] E[Y^{a=1}]$$

Expected value, or expectation, is defined as follows:

$$E[X] = \sum x P(X = x)$$

• So we are interested in (and likewise for $Y^{a=1}$):

$$E[Y^{a=0}] = \sum_{x} yP(Y^{a=0} = y) = 0 \times P(Y^{a=0} = 0) + 1 \times P(Y^{a=0} = 1) = P(Y^{a=0} = 1)$$

Counterfactual data and causal effects

- Suppose we knew what would happen for each individual in the population under each value of the treatment
- Then we could compute the counterfactual risks:

$$E[Y^{a=0}] = 0.5$$
 $E[Y^{a=1}] = 0.5$

• The average causal effect of treatment A is defined as the difference of counterfactual risks:

$$E[Y^{a=1}] - E[Y^{a=0}] = 0$$

Here, treatment is ineffective

(Hernan & Robin		
	$Y^{a=0}$	$Y^{a=1}$
Rheia	0	1
Kronos	1	0
Demeter	0	0
Hades	0	0
Hestia	0	0
Poseidon	1	0
Hera	0	0
Zeus	0	1
Artemis	1	1
Apollo	1	0
Leto	0	1
Ares	1	1
Athena	1	1
Hephaestus	0	1
Aphrodite	0	1
Cyclope	0	1
Persephone	1	1
Hermes	1	0
Hebe	1	0
Dionysus	1	0
$P(Y^{a=*}) = 1$	0.5	

Remember, $E[Y^{a=i}] = P(Y^{a=i} = 1)$

Estimating causal effects

	L	A	Y	Y^0	Y^1
Rheia	0	0	0	0	?
Kronos	0	0	1	1	?
Demeter	0	0	0	0	?
Hades	0	0	0	0	?
Hestia	0	1	0	?	0
Poseidon	0	1	0	?	0
Hera	0	1	0	?	0
Zeus	0	1	1	?	1
Artemis	1	0	1	1	?
Apollo	1	0	1	1	?
Leto	1	0	0	0	?
Ares	1	1	1	?	1
Athena	1	1	1	?	1
Hephaestus	1	1	1	?	1
Aphrodite	1	1	1	?	1
Polyphemus	1	1	1	?	1
Persephone	1	1	1	?	1
Hermes	1	1	0	?	0
Hebe	1	1	0	?	0
D:	-	-	_	0	•

Dionysus

Naively, we might estimate the counterfactual risks $P(Y^{a=i}=1)$ directly from observed A and Y:

$$\hat{P}_{MLE}(Y=1 | A=0) = \frac{3}{7} \quad \hat{P}_{MLE}(Y=1 | A=1) = \frac{7}{13}$$

But under what circumstances

$$\hat{P}_{MLE}(Y|A=i) = \hat{P}_{MLE}(Y^{a=i}=1)$$
?

The following is certainly true:

$$\hat{P}_{MLE}(Y=1 \mid A=i) = \frac{\text{Count}(Y=1 \land A=i)}{\text{Count}(A=i)}$$

$$Consistency: \text{ when } A=i, Y=Y^{a=i}$$

$$= \frac{\text{Count}(Y^{a=1}=1 \land A=i)}{\text{Count}(A=i)}$$

$$= \hat{P}_{MLE}(Y^{a=i}=1 \mid A=i)$$

$$Crucial step; make sure you understand it!$$

So, the following condition suffices:

$$P(Y^{a=i} = 1 | A = i) = P(Y^{a=i} = 1)$$

• This is called EXCHANGEABILITY:

$$Y^a \perp A \mid \{\}$$

Exchangeability and randomization

Goal:
$$\hat{P}(Y^a = 1)$$

- Why is a randomized experiment so powerful?
- Recap of exchangeability criterion:

$$Y^a \perp A \mid \{\}$$

- If we ourselves determine A in a way that is *truly blind to* Y^a , it **imposes** exchangeability!
- We can now go ahead and estimate

$$\hat{P}(Y^{a=i} = 1) = \hat{P}(Y = 1 | A = i)$$

Hooray!!!

Does loss of randomization make things hopeless?

- In the real world, many datasets are not randomized this way
- **Example:** let's imagine some other variable that might affect whether treatment A is applied; e.g., L = whether the patient was in critical condition (1=yes, 0=no)
- In general, L will be related to Y^a
 - E.g., in this example, patients in critical condition are surely more likely to die overall!



	•
	L
Rheia	0
Kronos	0
Demeter	0
Hades	0
Hestia	0
Poseidon	0
Hera	0
Zeus	0
Artemis	1
Apollo	1
Leto	1
Ares	1
Athena	1
Hephaestus	1
Aphrodite	1
Polyphemus	1
Persephone	1
Hermes	1
Hebe	1
Dionysus	1

Conditional exchangeability

- But now suppose we have observed (i.e., it's in our dataset) the factor L that affected whether the treatment A was applied
- If the following condition holds, it can help us estimate the counterfactual risks $P(Y^a)$:

$A \perp Y^a \mid L$

- That is, L captures all the information available in A that is relevant to all Y^a
- This is called CONDITIONAL EXCHANGEABILITY

	L	A	$Y^{a=0}$	$Y^{a=1}$	Y
Rheia	0	0	0	1	0
Kronos	0	0	1	0	1
Demeter	0	0	0	0	0
Hades	0	0	0	0	0
Hestia	0	1	0	0	0
Poseidon	0	1	1	0	0
Hera	0	1	0	0	0
Zeus	0	1	0	1	1
Artemis	1	0	1	1	1
Apollo	1	0	1	0	1
Leto	1	0	0	1	0
Ares	1	1	1	1	1
Athena	1	1	1	1	1
Hephaestus	1	1	0	1	1
Aphrodite	1	1	0	1	1
Polyphemus	1	1	0	1	1
Persephone	1	1	1	1	1
Hermes	1	1	1	0	0
Hebe	1	1	1	0	0
Dionysus	1	1	1	0	0

Using conditional exchangeability

- Can we estimate $P(Y^{a=i} = 1 | L)$?
- It turns out we can!

$$P(Y^{a=i} = 1 \mid L) = P(Y^{a=i} = 1 \mid L, A)$$

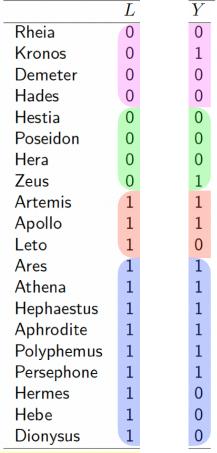
$$\hat{P}_{\mathsf{MLE}}(Y^{a=i} = 1 \mid L = j, A = k) = \frac{\mathsf{Count}(Y^{a=i} = 1, L = j, A = k)}{\mathsf{Count}(L = j, A = k)}$$

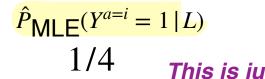
- In estimating this condprob:
 - when i = k we use CONSISTENCY

CONSISTENCY: when A = i, $Y = Y^{a=i}$

• WHEN $i \neq k$ we have "missing

		data", so	ignore those	e instances
L	\boldsymbol{A}	Count(L, A)	$Count(Y^{a=0}=1,\!L,A)$	$Count(Y^{a=1} = 1, L, A)$
0	0	4	1	?
0	1	4	?	1
1	0	3	2	?
1	1	9	?	6





1/4

2/3

2/3

This is just like estimating P(Y|L,A)!

Using conditional exchangeability

- We originally characterized our goal as estimating the counterfactual risks $E[Y^{a=i}] = P(Y^{a=i} = 1)$
- With conditional exchangeability, we estimated $P(Y^{a=i}=1\,|\,L)$ these are called **stratum-specific risks** (where each **stratum** is a value of L)
- Often, this may be all you need or want
 - If the causal effect of A depends on L, then "summarizing out" L discards information!
- But there are situations where the basic counterfactual risk $E[Y^{a=i}]$ may be of interest
 - e.g., "how many lives would it save if everyone who came to the hospital with heart disease received a heart transplant?
- But we can recover the basic counterfactual risks through standardization (or the mathematically equivalent inverse probability weighting)

Standardization

L = whether the patient was in critical condition (1=yes, 0=no)

L	\boldsymbol{A}	$\hat{P}_{MLE}(Y^{a=i} L)$
0	0	1/4
0	1	1/4
1	0	2/3
1	1	2/3

By the law of total probability,

We have just estimated this
$$P(Y^{a=i}) = \sum_{j} P(Y^{a=i} \mid L) P(L)$$
 We can estimate this from the data, too

Expanding the sum and plugging in our estimates we get:

$$P(Y^{a=i} = 1) = P(Y^{a=i} = 1 | L = 0)P(L = 0) + P(Y^{a=i} = 1 | L = 1)P(L = 1)$$

$$= \frac{1}{4} \times \frac{2}{5} + \frac{2}{3} \times \frac{3}{5}$$

$$=\frac{5}{10}=\frac{1}{2}$$

$= 0) + P(I^{w-1} = 1 L = 1)P(L = 1)$	$\hat{P}_{MLE}(L=1) = \frac{12}{20} = \frac{3}{5}$	5
(Because $\hat{P}_{MLE}(Y^{a=i} L)$ are the same for us the result for both counterfactual treatments		۷ gi

$$a=0$$
 and $a=1$, this work gives

IDENTIFIABILITY of causal effects

- IDENTIFIABILITY means, our assumptions allow the causal effect we are interested in to be uniquely estimated from the available data (set of observed/measured variables)
 - "Uniquely estimate": if we had an arbitrary large quantity of data, we could estimate the causal effect with arbitrarily high accuracy and precision
- Simple case of unidentifiability: Hernan & Robins's heart transplant example, if L (severity of disease) affects probability of a heart transplant and we don't measure it
 - Suppose that people with transplants have lower survival rates: $\hat{P}_{MLE}(Y|A=1) < \hat{P}_{MLE}(Y|A=0)$
 - Could be because heart transplants are dangerous
 - Or: sicker people are more likely to get transplants!

The three criteria for identifiability

- Consistency: $Y = Y^{a=i}$ whenever A = i
 - Consequence: different individuals' outcomes don't affect each other
 - Consequence: there can be no "multiple versions" of the same treatment A in terms of their influence on Y
- Conditional Exchangeability: for all i, $Y^{a=i} \perp A \mid Z$ for some set of observed variables Z
 - Consequence: there can be no "hidden common causes" or "hidden mediators" of A and $Y^{a=i}$
- **Positivity**: for all i and all values of Z, P(A = i | Z) > 0
 - e.g., in our example, it can't be the case that individuals with heart disease are *always* given transplants
- If all three criteria hold, we can estimate causal effects

Summary of intro to potential outcomes

- The potential outcomes framework formalizes causal effects (or risks) through counterfactual outcome (also called potential outcome) variables
- At most one counterfactual outcome is observable in each datum, so causal effects cannot in general be naively estimated from data

 However, if the three following conditions hold, the data can be viewed as a conditionally randomized experiment and causal effects can be estimated

Poseidon

Consistency	Conditional Exchangeability	Positivity
$Y = Y^{a=i}$ whenever $A = i$	$\exists Z. \forall i.Z$ is observed	$\forall i . \forall Z . P(A = i Z) > 0$
	$\wedge Y^{a=i} \perp A \mid Z$	

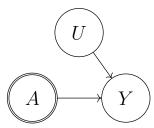
 This analysis also sheds light on the power of randomized experiments: they offer unconditional exchangeability 15

Why causal inference is challenging

- But, we don't get to see individualspecific counterfactual RV values!
- At most we see one actual outcome
- Let us bundle up everything we don't know about the process by which actual A and Y were determined into the RV U



 This is why randomized experiments are so powerful!



ernan & Robins, 2020, Table 1				
	$Y^{a=0}$)]	ra=1	1
Rheia	O		1	
Kronos	1		0	
Demeter	0		0	
Hades	0		0	
Hestia	0		0	
Poseidon	1		0	
Hera	0		0	
Zeus	0		1	
Artemis	1		1	
Apollo	$\lfloor 1 \rfloor$		0	
Leto	0		1	
Ares	1		1	
Athena	1		1	
Hephaestus	0		1	
Aphrodite	0		1	
Cyclope	0		1	
Persephone	1	1	1	
Hermes	1		0	
Hebe	1		0	
Dionysus	1		0	
		17		

Pearl's functional causal models

- Examples like these highlight the value of explicitly representing the unknowns – parts of the causal system that we will not in practice be able to observe
- For a system of RVs $V=\{X_i\}$, each gets an unseen "error term" $\{U_i\}$ (not necessarily independent)
- A functional causal model then consists of a deterministic system of equations of the form

$$x_i = f_i(pa_i, u_i)$$

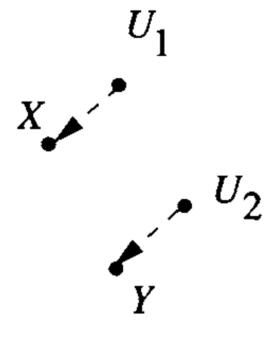
• where pa_i is the subset of V that are parents of X_i

Bayes Nets & functional causal models

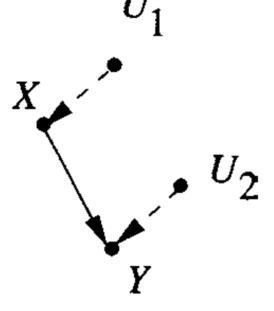
Bayes Net

X

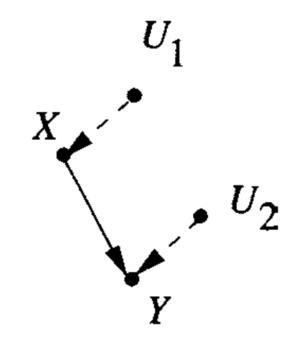
Corresponding functional causal model



Functional causal model with X causally upstream of Y



Example model & structural equations



$$x = f_1(u_1)$$

$$y = f_2(x, u_2)$$

An example structural equation:

$$y = a + bx + u_2$$

(Could be linear regression!)