

# **9.S918: Statistical Inference in Brain and Cognitive Sciences**

**Week 1 Day 2: Introduction to causal inference**

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# A tiny bit of statistics

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- On Tuesday we reviewed basics of **probability**: the logical calculus of uncertainty—a branch of mathematics
- The primary focus of this class is **statistics**: the mathematics, science, craft, and art of drawing inferences from data
- The two fields are fundamentally different
- But, probability is used extensively throughout statistics

# Perhaps the simplest probability **distribution**

- Consider a binary random variable  $Y$  with two possible outcomes: 0 and 1
- $Y$  is a **Bernoulli random variable** with **parameter**  $P(\text{heads}) = \pi$ , where  $0 \leq \pi \leq 1$
- Figuring out from observed data what the weighting is likely to be is **parameter estimation**
- In general, we will use  $\mathbf{y}$  to refer to observed-outcome **data** and  $\theta$  to refer to the model parameters to be estimated

# Statistical estimators

- **Estimator:** a procedure for guessing a quantity of interest within a population from a sample from that population
- For example, the **relative frequency estimator:** if we observe  $r$  instances of heads in  $n$  coin flips,

"this is an estimator" 

$$\hat{\pi} = \frac{r}{n}$$

- Data are stochastic, so estimators give random variables!

- **Bias** of an estimator is  $E[\hat{\theta}] - \theta$   Here we used **linearity of the expectation**

$$E[\hat{\pi}] = E\left[\frac{r}{n}\right] = \frac{1}{n}E[r] = \frac{1}{n} \sum_{i=1}^n E[Y_i] = \frac{1}{n}n\pi = \pi$$

...so  $\hat{\pi}$  is **unbiased**

- **Variance** of an estimator is ordinary variance

$$\text{Var}(X) \equiv E[(X - E[X])^2] \quad \text{Var}(\hat{\pi}) = \frac{\pi(1 - \pi)}{n} \quad (\text{see reading materials})$$

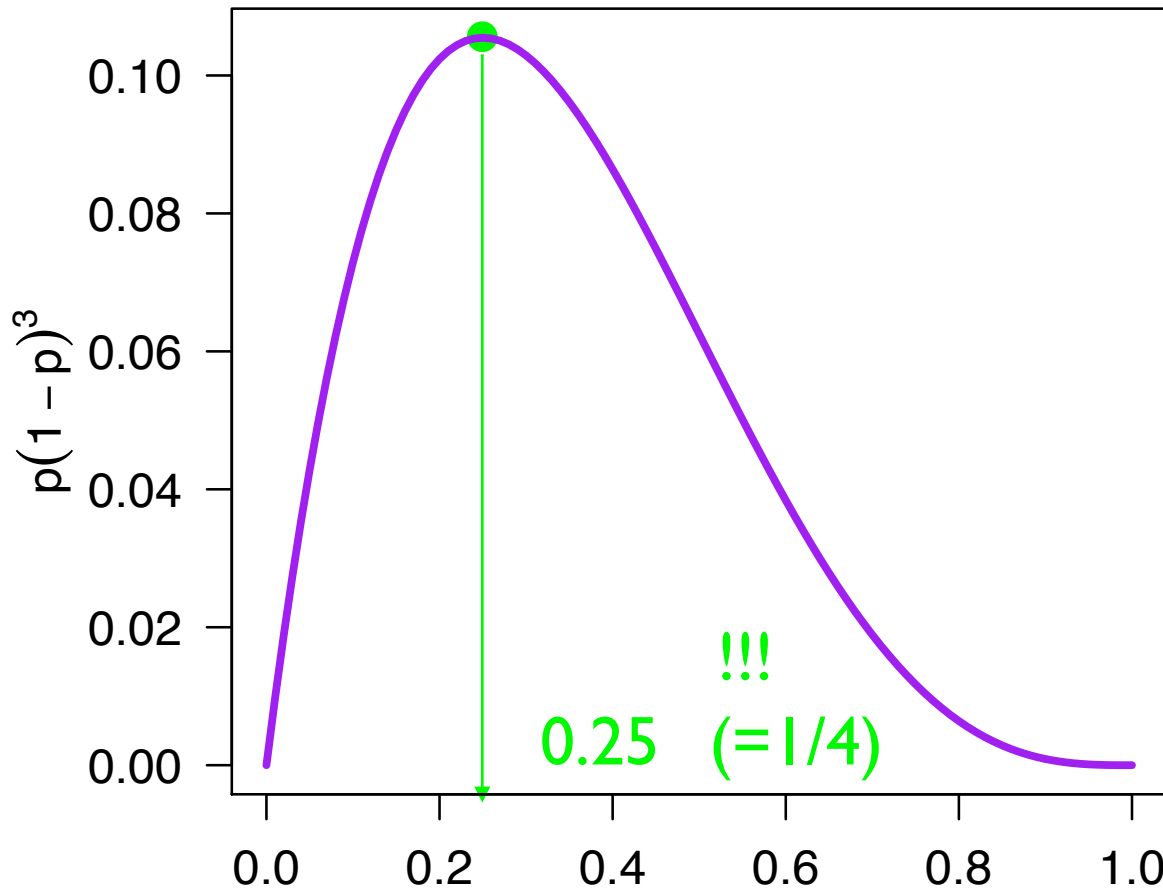
- Good estimators have favorable **bias–variance** tradeoff

# Maximum likelihood estimation

$$\text{Lik}(\boldsymbol{\theta}; \mathbf{y}) \equiv P(\mathbf{y}|\boldsymbol{\theta}) \quad \hat{\boldsymbol{\theta}}_{MLE} \stackrel{\text{def}}{=} \arg \max_{\boldsymbol{\theta}} \text{Lik}(\boldsymbol{\theta}; \mathbf{y})$$

$i$	$y_i$
1	T
2	T
3	H
4	T

- $p$  refers to the value of  $P(\text{coin toss}_i = \text{Heads})$
- Likelihood for the following dataset



This is choosing the *maximum likelihood estimate* (**MLE**)

The **MLE** also turns out to be the *relative frequency estimate* (**RFE**)

(repeat slide from lecture 3)

# Introductory causal inference

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- You have probably had previous exposure to both probability and statistics
- You are less likely to have had exposure to **causal inference**
- Causal inference uses probability and statistics, but it is something separate from the traditional construal of those two fields
- You can think of causal inference as being a framework extending more traditional statistics by:
  - Adding new probability-based mathematical constructs; and,
  - Developing a set of practice for statistical inference based on those constructs
- Two causal inference frameworks:
  - The **potential outcomes** framework
  - The **causal graphical models** framework

# The potential-outcomes framework

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- In epidemiology and many other areas of statistics, causal inference was developed out of the idea of **potential outcomes** (Neyman 1923, Rubin 1974)
- Consider an outcome,  $Y$ , and a potential **treatment**  $A$
- **Example:**  
 $Y$ : an individual survives to the end of the year (0: no, 1: yes)  
 $A$ : an individual with heart disease receives a heart transplant (0: no, 1: yes)

# Potential-outcome random variables

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- Suppose that  $A$  is discrete; for this case,  $A \in \{0,1\}$
- The **potential outcomes**, or **counterfactual outcomes**, are random variables for  $Y$  for each potential value of  $A$

$Y^{a=0}$                       The value that  $Y$  would take if  $A$  were 0

$Y^{a=1}$                       The value that  $Y$  would take if  $A$  were 1

- **Counterfactual risk** is the **expected value** of each counterfactual-outcome random variable:

$$E[Y^{a=0}]$$

$$E[Y^{a=1}]$$

- Expected value, or expectation, is defined as follows:

$$E[X] = \sum_x xP(X = x)$$

- So we are interested in (and likewise for  $Y^{a=1}$ ):

$$E[Y^{a=0}] = \sum_y yP(Y^{a=0} = y) = 0 \times P(Y^{a=0} = 0) + 1 \times P(Y^{a=0} = 1) = \boxed{P(Y^{a=0} = 1)}$$



# Counterfactual data and causal effects

- Suppose we knew **what would happen** for each individual in the population under each value of the treatment
- Then we could compute the counterfactual risks:

$$E[Y^{a=0}] = 0.5 \qquad E[Y^{a=1}] = 0.5$$

- The **average causal effect** of treatment  $A$  is defined as the difference of counterfactual risks:

$$E[Y^{a=1}] - E[Y^{a=0}] = 0$$

- Here, treatment is **ineffective**

(Hernan & Robins, 2020, Table 1.1)

	$Y^{a=0}$	$Y^{a=1}$
Rheia	0	1
Kronos	1	0
Demeter	0	0
Hades	0	0
Hestia	0	0
Poseidon	1	0
Hera	0	0
Zeus	0	1
Artemis	1	1
Apollo	1	0
Leto	0	1
Ares	1	1
Athena	1	1
Hephaestus	0	1
Aphrodite	0	1
Cyclope	0	1
Persephone	1	1
Hermes	1	0
Hebe	1	0
Dionysus	1	0
$P(Y^{a=*}) = 1$	0.5	0.5

# Estimating causal effects

Remember,  $E[Y^{a=i}] = P(Y^{a=i} = 1)$

	$L$	$A$	$Y$	$Y^0$	$Y^1$
Rheia	0	0	0	0	?
Kronos	0	0	1	1	?
Demeter	0	0	0	0	?
Hades	0	0	0	0	?
Hestia	0	1	0	?	0
Poseidon	0	1	0	?	0
Hera	0	1	0	?	0
Zeus	0	1	1	?	1
Artemis	1	0	1	1	?
Apollo	1	0	1	1	?
Leto	1	0	0	0	?
Ares	1	1	1	?	1
Athena	1	1	1	?	1
Hephaestus	1	1	1	?	1
Aphrodite	1	1	1	?	1
Polyphemos	1	1	1	?	1
Persephone	1	1	1	?	1
Hermes	1	1	0	?	0
Hebe	1	1	0	?	0
Dionysus	1	1	0	?	0

- Naively, we might estimate the counterfactual risks  $P(Y^{a=i} = 1)$  directly from observed  $A$  and  $Y$ :

$$\hat{P}_{MLE}(Y = 1 | A = 0) = \frac{3}{7} \quad \hat{P}_{MLE}(Y = 1 | A = 1) = \frac{7}{13}$$

- But under what circumstances  $\hat{P}_{MLE}(Y | A = i) = \hat{P}_{MLE}(Y^{a=i} = 1)$ ?
- The following is certainly true:

$$\hat{P}_{MLE}(Y = 1 | A = i) = \frac{\text{Count}(Y = 1 \wedge A = i)}{\text{Count}(A = i)}$$

**CONSISTENCY:** when  
 $A = i, Y = Y^{a=i}$

$$= \frac{\text{Count}(Y^{a=1} = 1 \wedge A = i)}{\text{Count}(A = i)}$$

$$= \hat{P}_{MLE}(Y^{a=i} = 1 | A = i)$$

*Crucial step;  
make sure you  
understand it!*

- So, the following condition suffices:
- This is called **EXCHANGEABILITY**:

$$Y^a \perp A | \{ \}$$

# Exchangeability and randomization

Goal:  $\hat{P}(Y^a = 1)$

- Why is a randomized experiment so powerful?
- Recap of exchangeability criterion:

$$Y^a \perp A \mid \{\}$$

- If we ourselves determine  $A$  in a way that is *truly blind* to  $Y^a$ , it **imposes** exchangeability!
- We can now go ahead and estimate

$$\hat{P}(Y^{a=i} = 1) = \hat{P}(Y = 1 \mid A = i)$$

- Hooray!!!

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Rheia  
Kronos  
Demeter  
Hades  
Hestia  
Poseidon  
Hera  
Zeus  
Artemis  
Apollo  
Leto  
Ares  
Athena  
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Aphrodite  
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Hermes  
Hebe  
Dionysus

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# Does loss of randomization make things hopeless?

- In the real world, many datasets are **not** randomized this way
- **Example:** let's imagine some other variable that might affect whether treatment  $A$  is applied; e.g.,  $L$  = whether the patient was in critical condition (1=yes, 0=no)
- In general,  $L$  will be related to  $Y^a$ 
  - E.g., in this example, patients in critical condition are surely more likely to die overall!

$A \perp Y \mid \{\}$



	$L$
Rheia	0
Kronos	0
Demeter	0
Hades	0
Hestia	0
Poseidon	0
Hera	0
Zeus	0
Artemis	1
Apollo	1
Leto	1
Ares	1
Athena	1
Hephaestus	1
Aphrodite	1
Polyphemus	1
Persephone	1
Hermes	1
Hebe	1
Dionysus	1