# Lecture 5: GPs and Streaming regression

- Gaussian Processes
- Information gain
- Confidence intervals

## **Recall: Non-parametric regression**

- ullet Input space  $\mathcal{X}\subset\mathbb{R}^n$ , target space  $\mathcal{Y}=\mathbb{R}$
- Input matrix X of size  $m \times n$ , target vector y of size  $m \times 1$
- Feature mapping  $\phi: \mathcal{X} \mapsto \mathbb{R}^d$
- ullet Assumption:  $\mathbf{y} = \mathbf{\Phi} \mathbf{w}$
- Minimize

$$J_{\lambda}(\mathbf{w}) = \frac{1}{2}(\mathbf{\Phi}\mathbf{w} - \mathbf{y})^{\top}(\mathbf{\Phi}\mathbf{w} - \mathbf{y}) + \frac{\lambda}{2}\mathbf{w}^{\top}\mathbf{w} \qquad \lambda \ge 0$$

• The solution is

$$\mathbf{w} = (\mathbf{\Phi}^{\top} \mathbf{\Phi} + \lambda \mathbf{I}_n)^{-1} \mathbf{\Phi}^{\top} \mathbf{y} = \mathbf{\Phi}^{\top} (\mathbf{\Phi} \mathbf{\Phi}^{\top} + \lambda \mathbf{I}_m)^{-1} \mathbf{y}$$

## **Recall: Kernel regression**

• Let  $\mathbf{K} = \mathbf{\Phi} \mathbf{\Phi}^{\top}$  and  $\mathbf{k}(\mathbf{x}) = \phi(\mathbf{x})^{\top} \mathbf{\Phi}^{\top}$  be the kernel matrix/vector:

$$\mathbf{K} = \begin{bmatrix} k(\mathbf{x}_1, \mathbf{x}_1) & k(\mathbf{x}_1, \mathbf{x}_2) & \dots & k(\mathbf{x}_1, \mathbf{x}_m) \\ k(\mathbf{x}_2, \mathbf{x}_1) & k(\mathbf{x}_2, \mathbf{x}_2) & \dots & k(\mathbf{x}_2, \mathbf{x}_m) \\ \vdots & \vdots & & \vdots \\ k(\mathbf{x}_m, \mathbf{x}_1) & k(\mathbf{x}_m, \mathbf{x}_2) & \dots & k(\mathbf{x}_m, \mathbf{x}_m) \end{bmatrix} \qquad \mathbf{k}(\mathbf{x}) = \begin{bmatrix} k(\mathbf{x}, \mathbf{x}_1) \\ k(\mathbf{x}, \mathbf{x}_2) \\ \vdots \\ k(\mathbf{x}, \mathbf{x}_m) \end{bmatrix}$$

• The predictions for the input data are given by

$$\hat{\mathbf{y}} = \mathbf{\Phi} \mathbf{\Phi}^{\top} (\mathbf{\Phi} \mathbf{\Phi}^{\top} + \lambda \mathbf{I}_m)^{-1} \mathbf{y} = \mathbf{K} (\mathbf{K} + \lambda \mathbf{I}_m)^{-1} \mathbf{y}$$

ullet The prediction for a new input point x is given by

$$\hat{f}(\mathbf{x}) = \phi(\mathbf{x})^{\top} \mathbf{\Phi}^{\top} (\mathbf{K} + \lambda \mathbf{I}_m)^{-1} \mathbf{y} = \mathbf{k}(\mathbf{x}) (\mathbf{K} + \lambda \mathbf{I}_m)^{-1} \mathbf{y}$$

## Recall: Bayesian view of regression

- Consider noisy observations  $y = f(\mathbf{x}) + \epsilon = \phi(\mathbf{x})^{\top} \mathbf{w} + \epsilon$
- Recall Bayes' rule: posterior =  $\frac{\text{likelihood} \times \text{prior}}{\text{marginal likelihood}}$

$$P_{\phi}(\mathbf{w}|\mathbf{y}, \mathbf{X}) = \frac{P_{\phi}(\mathbf{y}|\mathbf{X}, \mathbf{w})P(\mathbf{w})}{P_{\phi}(\mathbf{y}|\mathbf{X})}$$

- $\Rightarrow$  Marginal likelihood is independent of weights  ${f w}$ 
  - With Gaussian noise  $\epsilon \sim \mathcal{N}(0, \sigma^2)$

$$P_{\phi}(\mathbf{y}|\mathbf{X}, \mathbf{w}) = \prod_{i=1}^{m} P_{\phi}(y_i|\mathbf{x}_i, \mathbf{w}) = \prod_{i=1}^{m} \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(y_i - \phi(\mathbf{x}_i)^{\top}\mathbf{w})^2}{2\sigma^2}\right)$$
$$= \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{\|\mathbf{y} - \mathbf{\Phi}\mathbf{w}\|^2}{2\sigma^2}\right) = \mathcal{N}_m\left(\mathbf{\Phi}\mathbf{w}, \sigma^2\mathbf{I}_m\right)$$

#### Posterior distribution on parameters

• With Gaussian prior on parameters  $\mathbf{w} \sim \mathcal{N}_d(0, \Sigma_{\mathbf{w}})$ 

$$P_{\phi}(\mathbf{w}|\mathbf{y}, \mathbf{X}) \propto \exp\left(-\frac{\|\mathbf{y} - \mathbf{\Phi}\mathbf{w}\|^{2}}{2\sigma^{2}}\right) \exp\left(-\frac{\mathbf{w}^{\top} \Sigma_{\mathbf{w}}^{-1} \mathbf{w}}{2}\right)$$

$$= \exp\left(-\frac{\mathbf{y}^{\top} \mathbf{y} - \mathbf{y}^{\top} \mathbf{\Phi} \mathbf{w} - \mathbf{w}^{\top} \mathbf{\Phi} \mathbf{y} + \mathbf{w}^{\top} \mathbf{\Phi}^{\top} \mathbf{\Phi} \mathbf{w} + \sigma^{2} \mathbf{w}^{\top} \Sigma_{\mathbf{w}}^{-1} \mathbf{w}}{2\sigma^{2}}\right)$$

$$= \exp\left(-\frac{\mathbf{y}^{\top} \mathbf{y} - \mathbf{y}^{\top} \mathbf{\Phi} \mathbf{w} - \mathbf{w}^{\top} \mathbf{\Phi} \mathbf{y} + \mathbf{w}^{\top} (\mathbf{\Phi}^{\top} \mathbf{\Phi} + \sigma^{2} \Sigma_{\mathbf{w}}^{-1}) \mathbf{w}}{2\sigma^{2}}\right)$$

$$\propto \exp\left((\mathbf{w} - \mathbf{b})^{\top} \mathbf{A}^{-1} (\mathbf{w} - \mathbf{b})\right)$$

where  $\mathbf{A}^{-1} = \sigma^{-2}(\mathbf{\Phi}^{\top}\mathbf{\Phi} + \sigma^{2}\Sigma_{\mathbf{w}}^{-1})$  and  $\mathbf{b} = (\mathbf{\Phi}^{\top}\mathbf{\Phi} + \sigma^{2}\Sigma_{\mathbf{w}}^{-1})^{-1}\mathbf{\Phi}^{\top}\mathbf{y}$  $\Rightarrow$  The posterior distribution is Gaussian!

#### **Predictive distribution**

The pointwise posterior predictive distribution is a normal distribution

$$\tilde{f}(\mathbf{x})|\mathbf{x}_1,\ldots,\mathbf{x}_m,y_1,\ldots,y_m \sim \mathcal{N}\left(\hat{f}(\mathbf{x}),s^2(\mathbf{x})\right)$$

of expectation

$$\hat{f}(\mathbf{x}) = \phi(\mathbf{x})^{\top} (\mathbf{\Phi}^{\top} \mathbf{\Phi} + \sigma^{2} \Sigma_{\mathbf{w}}^{-1})^{-1} \mathbf{\Phi}^{\top} \mathbf{y}$$
$$= \phi(\mathbf{x})^{\top} \Sigma_{\mathbf{w}} \mathbf{\Phi}^{\top} (\mathbf{\Phi} \Sigma_{\mathbf{w}} \mathbf{\Phi}^{\top} + \sigma^{2} \mathbf{I}_{m})^{-1} \mathbf{y}$$

and variance

$$\begin{split} s^2(\mathbf{x}) &= \sigma^2 \phi(\mathbf{x})^\top (\mathbf{\Phi}^\top \mathbf{\Phi} + \sigma^2 \Sigma_{\mathbf{w}}^{-1})^{-1} \phi(\mathbf{x}) \\ &= \phi(\mathbf{x})^\top \Sigma_{\mathbf{w}} \phi(\mathbf{x}) - \phi(\mathbf{x})^\top \Sigma_{\mathbf{w}} \mathbf{\Phi}^\top (\mathbf{\Phi}^\top \Sigma_{\mathbf{w}} \mathbf{\Phi} + \sigma^2 \mathbf{I}_m)^{-1} \mathbf{\Phi} \Sigma_{\mathbf{w}} \phi(\mathbf{x}) \\ &\to \text{using Sherman-Morrison} \end{split}$$

## Reinterpreting regularization

Recall kernel regression predictions:

$$\hat{f}(\mathbf{x}) = \mathbf{k}(\mathbf{x})^{\top} (\mathbf{K} + \lambda \mathbf{I}_m)^{-1} \mathbf{y}$$

• Using prior  $\Sigma_{\mathbf{w}} = \frac{\sigma^2}{\lambda} \mathbf{I}_d$ , the predictive mean rewrites as:

$$\hat{f}(\mathbf{x}) = \phi(\mathbf{x})^{\top} \Sigma_{\mathbf{w}} \mathbf{\Phi}^{\top} (\mathbf{\Phi} \Sigma_{\mathbf{w}} \mathbf{\Phi}^{\top} + \sigma^{2} \mathbf{I}_{m})^{-1} \mathbf{y}$$

$$= \phi(\mathbf{x})^{\top} \frac{\sigma^{2}}{\lambda} \mathbf{\Phi}^{\top} \left( \mathbf{\Phi} \frac{\sigma^{2}}{\lambda} \mathbf{\Phi}^{\top} + \sigma^{2} \mathbf{I}_{m} \right)^{-1} \mathbf{y}$$

$$= \mathbf{k}(\mathbf{x})^{\top} (\mathbf{K} + \lambda \mathbf{I}_{m})^{-1} \mathbf{y}$$

 $\Rightarrow \lambda$  encodes some prior on weights  ${f w}$ 

## Reinterpreting regularization (cont'd)

ullet Still using  $\Sigma_{\mathbf{w}}=rac{\sigma^2}{\lambda}\mathbf{I}_d$ , the predictive variance rewrites as:

$$s^{2}(\mathbf{x}) = \phi(\mathbf{x})^{\top} \Sigma_{\mathbf{w}} \phi(\mathbf{x}) - \phi(\mathbf{x})^{\top} \Sigma_{\mathbf{w}} \mathbf{\Phi}^{\top} (\mathbf{\Phi}^{\top} \Sigma_{\mathbf{w}} \mathbf{\Phi} + \sigma^{2} \mathbf{I}_{m})^{-1} \mathbf{\Phi} \Sigma_{\mathbf{w}} \phi(\mathbf{x})$$

$$= \phi(\mathbf{x})^{\top} \frac{\sigma^{2}}{\lambda} \phi(\mathbf{x}) - \phi(\mathbf{x})^{\top} \frac{\sigma^{2}}{\lambda} \mathbf{\Phi}^{\top} \left( \mathbf{\Phi}^{\top} \frac{\sigma^{2}}{\lambda} \mathbf{\Phi} + \sigma^{2} \mathbf{I}_{m} \right)^{-1} \mathbf{\Phi} \frac{\sigma^{2}}{\lambda} \phi(\mathbf{x})$$

$$= \frac{\sigma^{2}}{\lambda} k_{\lambda}(\mathbf{x}, \mathbf{x}) \quad \text{with}$$

$$k_{\lambda}(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}, \mathbf{x}') - \mathbf{k}(\mathbf{x})^{\top} (\mathbf{K} + \lambda \mathbf{I}_{m})^{-1} \mathbf{k}(\mathbf{x}')$$

## **Summary**

ullet Using prior  $\Sigma_{\mathbf{w}}=rac{\sigma^2}{\lambda}\mathbf{I}_d$ , we have

$$\hat{f}(\mathbf{x}) = \mathbf{k}(\mathbf{x})^{\top} (\mathbf{K} + \lambda \mathbf{I}_m)^{-1} \mathbf{y}$$

$$s^2(\mathbf{x}) = \frac{\sigma^2}{\lambda} k_{\lambda}(\mathbf{x}, \mathbf{x}) \quad \text{with}$$

$$k_{\lambda}(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}, \mathbf{x}') - \mathbf{k}(\mathbf{x})^{\top} (\mathbf{K} + \lambda \mathbf{I}_m)^{-1} \mathbf{k}(\mathbf{x}')$$

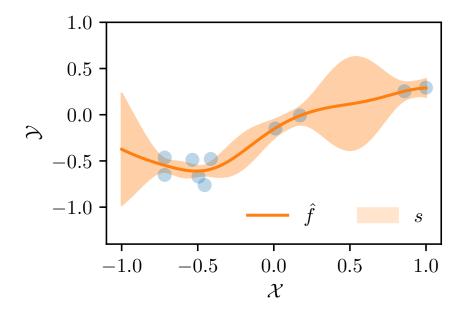
• Providing pointwise posterior prediction

$$\tilde{f}(\mathbf{x})|\mathbf{x}_1,\ldots,\mathbf{x}_m,y_1,\ldots,y_m \sim \mathcal{N}\left(\hat{f}(\mathbf{x}),s^2(\mathbf{x})\right)$$

 $\Rightarrow$  What does it mean to use  $\lambda \in \mathbb{R}_{>0}$ ?

## Pointwise posterior distribution

- ullet At each point  $\mathbf{x} \in \mathcal{X}$ , we have a distribution  $\mathcal{N}\left(\hat{f}(\mathbf{x}), s^2(\mathbf{x})\right)$
- $\bullet$  We can sample from these  $\tilde{f}(\mathbf{x}) \sim \mathcal{N}\left(\hat{f}(\mathbf{x}), s^2(\mathbf{x})\right)$



#### Joint distribution

- ullet Suppose you *query* your model at locations  ${f X}_*$
- Extend the prior to include query points:

$$egin{aligned} \begin{bmatrix} \mathbf{f} \ \mathbf{f}_* \end{bmatrix} | \mathbf{X}, \mathbf{X}_* & \mathcal{N}_{m+m_*} \left( \mathbf{0}, egin{bmatrix} \mathbf{K}_{\mathbf{X},\mathbf{X}} & \mathbf{K}_{\mathbf{X},\mathbf{X}_*} \ \mathbf{K}_{\mathbf{X}_*,\mathbf{X}_*} \end{bmatrix} 
ight) \ \mathbf{y} | \mathbf{f} &\sim \mathcal{N}_m(\mathbf{f}, \sigma^2 \mathbf{I}_m) \end{aligned}$$

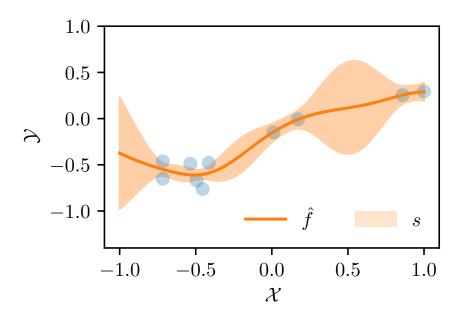
• Using joint normality of  $f_*$  and y:

$$egin{bmatrix} \mathbf{y} \ \mathbf{f}_* \end{bmatrix} \sim \mathcal{N}_{m+m_*} \left( \mathbf{0}, egin{bmatrix} \mathbf{K}_{\mathbf{X},\mathbf{X}} + \sigma^2 \mathbf{I}_m & \mathbf{K}_{\mathbf{X},\mathbf{X}_*} \ \mathbf{K}_{\mathbf{X}_*,\mathbf{X}} & \mathbf{K}_{\mathbf{X}_*,\mathbf{X}_*} \end{bmatrix} 
ight)$$

# Gaussian Process (GP)

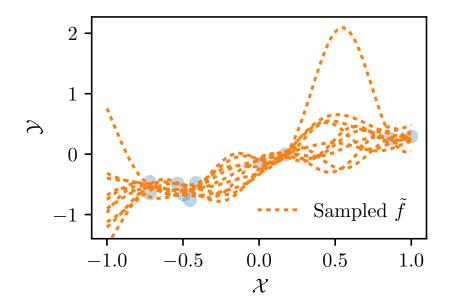
- By considering the covariance between *every points in the space*, we get a distribution over functions!
- Posterior distribution on *f*:

$$P[f|\mathbf{X}, \mathbf{y}] \sim \mathcal{N}_{|\mathcal{X}|} \left( \left[ \hat{f}(\mathbf{x}) \right]_{\mathbf{x} \in \mathcal{X}}, \frac{\sigma^2}{\lambda} \left[ k_{\lambda}(\mathbf{x}, \mathbf{x}') \right]_{\mathbf{x}, \mathbf{x}' \in \mathcal{X}} \right)$$



## **Sampling from a Gaussian Process**

- Generalization of normal probability distribution to the function space
  - From a normal distribution we sample variables
  - From a GP we sample *functions*!



## Sampling from a Gaussian Process: How to

Observe that

$$\mathcal{N}_{|\mathcal{X}|} \left( \left[ \hat{f}(\mathbf{x}) \right]_{\mathbf{x} \in \mathcal{X}}, \frac{\sigma^2}{\lambda} \left[ k_{\lambda}(\mathbf{x}, \mathbf{x}') \right]_{\mathbf{x}, \mathbf{x}' \in \mathcal{X}} \right)$$

defines a  $|\mathcal{X}|$ -dimensional multivariate Gaussian distribution

- $\rightarrow$  What if  $|\mathcal{X}| = \infty$  (e.g.  $\mathcal{X} = [-1, 1]$ )?
  - ullet We can consider a discrete, finite, set  $\mathbb{X}\subset\mathcal{X}$  and sample from

$$\mathcal{N}_{|\mathbb{X}|}\left(\left[\hat{f}(\mathbf{x})\right]_{\mathbf{x}\in\mathbb{X}}, \frac{\sigma^2}{\lambda}\left[k_{\lambda}(\mathbf{x}, \mathbf{x}')\right]_{\mathbf{x}, \mathbf{x}'\in\mathbb{X}}\right)$$

ullet This will result in a function  $\widetilde{f}$  evaluated at every  $\mathbf{x} \in \mathbb{X}$ 

## Learning the hyperparameters

- ullet If we assume that  $\Sigma_{\mathbf{w}} = \mathbf{I}_d$ , then we have  $\lambda = \sigma^2$
- Let  $\theta$  denote the kernel hyperparameters (e.g.  $\rho$ )
- In practice you may not know the noise and the *optimal* kernel hypers
- Recall: multivariate normal density

$$P(\mathbf{y}|\boldsymbol{\theta}) = \frac{\exp\left(-\frac{1}{2}\mathbf{y}^{\top}(\mathbf{K}_{\boldsymbol{\theta}} + \sigma^{2}\mathbf{I}_{m})^{-1}\mathbf{y}\right)}{\sqrt{(2\pi)^{D}|\mathbf{K}_{\boldsymbol{\theta}} + \sigma^{2}\mathbf{I}_{m}|}}$$

• Maximize the marginal likelihood  $\mathcal{L} = \log P(\mathbf{y}|\boldsymbol{\theta})$ :

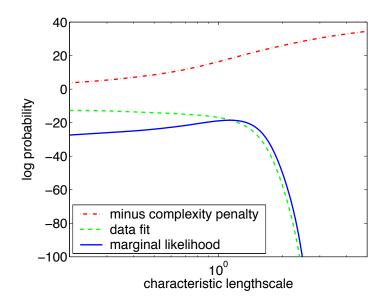
$$\mathcal{L} = -\frac{1}{2}\mathbf{y}^{\top}(\mathbf{K}_{\theta} + \sigma^{2}\mathbf{I}_{m})^{-1}\mathbf{y} - \frac{D}{2}\log(2\pi) - \frac{1}{2}\log|\mathbf{K}_{\theta} + \sigma^{2}\mathbf{I}_{m}|$$

## **Anatomy of marginal likelihood**

• Marginal likelihood:

$$\mathcal{L} = \log P(\mathbf{y}|\boldsymbol{\theta}) \propto -\frac{1}{2}\mathbf{y}^{\top}(\mathbf{K}_{\boldsymbol{\theta}} + \sigma^2 \mathbf{I}_m)^{-1}\mathbf{y} - \frac{1}{2}\log|\mathbf{K}_{\boldsymbol{\theta}} + \sigma^2 \mathbf{I}_m|$$

- 1st term: quality of predictions; 2nd term: model complexity
- Trade-off (from Rasmussen & Williams, 2006):



#### **Gradient-based optimization**

• Compute gradients:

$$\frac{\partial \mathcal{L}}{\partial \theta_i} = \frac{1}{2} \mathbf{y}^{\top} (\mathbf{K}_{\theta} + \sigma^2 \mathbf{I}_m)^{-1} \frac{\partial (\mathbf{K}_{\theta} + \sigma^2 \mathbf{I}_m)}{\partial \theta_i} (\mathbf{K}_{\theta} + \sigma^2 \mathbf{I}_m)^{-1} \mathbf{y}$$
$$- \frac{1}{2} \operatorname{Tr} \left( (\mathbf{K}_{\theta} + \sigma^2 \mathbf{I}_m)^{-1} \frac{\partial (\mathbf{K}_{\theta} + \sigma^2 \mathbf{I}_m)}{\partial \theta_i} \right)$$

- Minimize the negative
- Non-convex optimization task

#### **Summary**

- ullet Normal priors on the weights distribution o Gaussian Process
- ullet Regularization o prior on the weights covariance
- GP provides a posterior distribution on functions
  - Expectation: kernel regression model
  - Covariance → confidence intervals
- Sample discretized functions from a GP

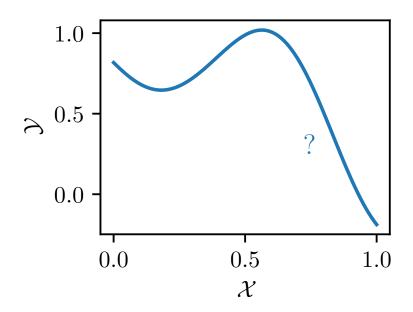
## **Typical supervised setting**

- ullet Have dataset  $(\mathbf{X}, \mathbf{y})$  of previously acquired data
- Learn model w on (X, y)
- Apply model w to provide predictions at new data points
- ullet Samples  $(\mathbf{X}, \mathbf{y})$  are often assumed to be i.i.d.

## **Streaming setting**

- Start with (possibly empty) dataset  $(\mathbf{X}_0, \mathbf{y}_0)$
- For each time step  $t = 1, 2, \ldots$ :
  - Fit model  $\mathbf{w}_t$  on  $\mathbf{X}_{t-1}$  and  $\mathbf{y}_{t-1}$
  - Acquire a new sample  $(\mathbf{x}_t, y_t)$  (possibly using  $\mathbf{w}_t$ )
  - Define  $X_t = [x_i]_{i=1...t}$ ,  $y_t = [y_i]_{i=1...t}$
- Samples may be dependent of model (not i.i.d.)
- $\rightarrow$  Samples influence model / Samples depend on model

## **Application: Online function optimization**



- ullet Unknown function  $f:\mathcal{X}\mapsto\mathbb{R}$
- ullet Sequentially select locations  $(\mathbf{x}_t)_{t\geq 1}$  at which to observe the function  $y_t$
- Try to maximize/minimize the observations

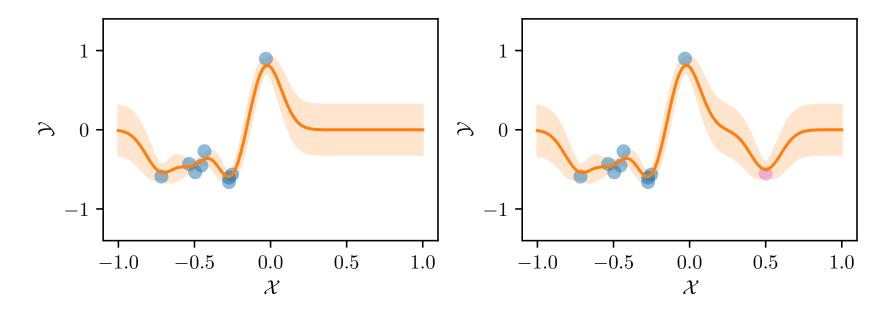
#### **E**xample

What is the optimal treatment dosage for a given disease?

- For each time step  $t = 1, 2, \ldots$ :
  - New patient t comes in
  - We decide on treatment dosage  $\mathbf{x}_t$
  - We observe the patient's response to treatement,  $y_t$
- Our goal is to cure patients as effectively as possible
- What you don't want: give bad dosages that were known to be bad
- What you want:
  - give good dosages
  - try informative dosages

#### **Information**

• An informative sample improves the model by reducing its uncertainty



How do we quantity the reduction of uncertainty after observing  $y_1, \ldots, y_t$  at locations  $\mathbf{x}_1, \ldots, \mathbf{x}_t$ ?

## A little bit of information theory

• Mutual information between underlying function f and observations  $y_1, \ldots, y_t$  at locations  $\mathbf{x}_1, \ldots, \mathbf{x}_t$ :

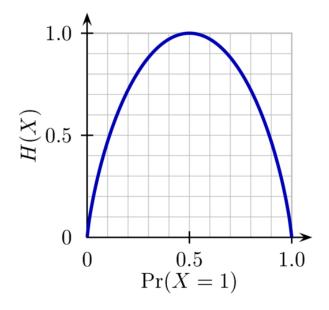
$$I(y_1,\ldots,y_t;f) = \underbrace{H(y_1,\ldots,y_t)}_{\text{Marginal entropy}} - \underbrace{H(y_1,\ldots,y_t|f)}_{\text{Conditional entropy}}$$

- It quantifies the "amount of information" obtained about one random variable, through the other random variable
- Entropy is the "amount of information" held by a random variable
- ullet H(Y|X)=0 if and only if the value of Y is completely determined by the value of X
- ullet H(Y|X) = H(Y) if and only if Y and X are independent random variables

#### **Amount of information**

$$H(X) = \mathbb{E}[-\ln P_X] = \sum_{i=1}^n p(x_i) \log_b (p(x_i))$$

Example: Tossing a coin



- A fair coin has maximal entropy: it is the less predictable!
- Every new sample that you get changes your *model* the most

#### **Mutual information**

• Entropy of  $X \sim \mathcal{N}_D(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ :

$$H(X) = \mathbb{E}\left[-\ln\frac{\exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^{\top}\boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right)}{\sqrt{(2\pi)^{D}}|\boldsymbol{\Sigma}|}\right] = \frac{D}{2} + \frac{D}{2}\ln 2\pi + \frac{1}{2}\ln|\boldsymbol{\Sigma}|$$

- $\rightarrow \ \mathsf{Using} \ \mathbb{E}\left[\mathbf{a}^{\top}\mathbf{M}^{-1}\mathbf{a}\right] = \mathbb{E}\left[\mathrm{Tr}\left(\mathbf{a}^{\top}\mathbf{M}^{-1}\mathbf{a}\right)\right] = \mathbb{E}\left[\mathrm{Tr}\left(\mathbf{a}\mathbf{a}^{\top}\mathbf{M}^{-1}\right)\right] = D$ 
  - Recall the joint distribution between m observations  $\mathbf{y}$  and  $m_*$  query points  $\mathbf{X}_*$ :

$$\begin{bmatrix} \mathbf{y} \\ \mathbf{f}_* \end{bmatrix} \sim \mathcal{N}_{m+m_*} \left( \mathbf{0}, \begin{bmatrix} \mathbf{K}_{\mathbf{X},\mathbf{X}} + \sigma^2 \mathbf{I}_m & \mathbf{K}_{\mathbf{X},\mathbf{X}_*} \\ \mathbf{K}_{\mathbf{X}_*,\mathbf{X}} & \mathbf{K}_{\mathbf{X}_*,\mathbf{X}_*} \end{bmatrix} \right)$$

• Then  $I(y_1,\ldots,y_t;f)=\frac{1}{2}\ln|\mathbf{K}_t+\sigma^2\mathbf{I}_t|$ 

## Another decomposition of mutual information

- Recall that  $y_1, \ldots, y_t | f(x_1), \ldots, f(x_t) \sim \mathcal{N}_t \left( \left( f(x_1), \ldots, f(x_t) \right), \sigma^2 \mathbf{I}_t \right)$
- $\rightarrow$  Using that  $y_i = f(x_i) + \epsilon$  with  $\epsilon \sim \mathcal{N}(0, \sigma^2)$ 
  - Pluging-in the conditional entropy of a multivariate normal distribution:

$$I(y_1, \dots, y_t; f) = H(y_1, \dots, y_t) - \frac{1}{2} \ln |2\pi e \sigma^2 \mathbf{I}_t|$$
  
=  $H(y_1, \dots, y_t) - \frac{t}{2} \ln 2\pi e - \frac{t}{2} \sigma^2$ 

• What is the marginal entropy  $H(y_1, \ldots, y_t)$ ?

## **Entropy of observations**

Recursively decompose

$$H(y_{1},...,y_{t}) = H(y_{1},...,y_{t-1}) + H(y_{t}|y_{1},...,y_{t-1})$$

$$= H(y_{1},...,y_{t-1}) + \frac{1}{2}\ln\left[2\pi e\left(\sigma^{2} + \frac{\sigma^{2}}{\lambda}k_{\lambda,t-1}(x_{t},x_{t})\right)\right]$$

$$\vdots$$

$$= H(y_{1}) + H(y_{2}|y_{1}) + \cdots + \frac{1}{2}\ln\left[2\pi e\left(\sigma^{2} + \frac{\sigma^{2}}{\lambda}k_{\lambda,t-1}(x_{s},x_{s})\right)\right]$$

$$= \sum_{s=1}^{t} \frac{1}{2}\ln\left[2\pi e\sigma^{2}\left(1 + \frac{1}{\lambda}k_{\lambda,s-1}(x_{s},x_{s})\right)\right]$$

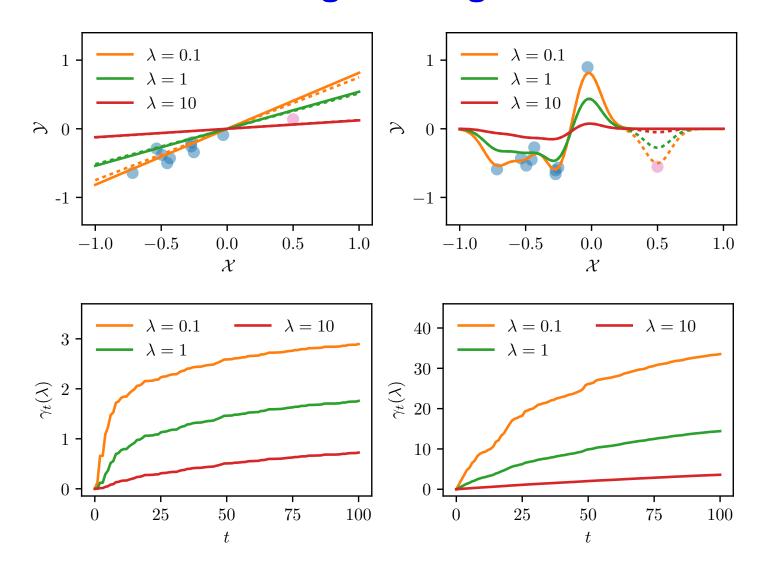
- $\rightarrow$  Still using that  $y_i = f(x_i) + \epsilon$  with  $\epsilon \sim \mathcal{N}(0, \sigma^2)$
- ightarrow Also using the uncertainty about the location of the *true* f

#### Information gain

$$I(y_1, \dots, y_t; f) = \sum_{s=1}^t \frac{1}{2} \ln \left[ 2\pi e \sigma^2 \left( 1 + \frac{1}{\lambda} k_{\lambda, s-1}(x_s, x_s) \right) \right] - \frac{t}{2} \ln(2\pi e) - \frac{t}{2} \sigma^2$$
$$= \sum_{s=1}^t \frac{1}{2} \ln \left[ 1 + \frac{1}{\lambda} k_{\lambda, s-1}(x_s, x_s) \right]$$

- Information gain  $\gamma_t(\lambda) = I(y_1, \dots, y_t; f)$ : reduction of uncertainty on f after observing  $y_1, \dots, y_t$
- ullet Information gain is inversely proportionnal to  $\lambda$
- ightarrow Limiting changes in function limits the contribution of samples

# Information gain vs regularization



#### **Summary**

- Streaming regression: sequentially gather (potentially non-i.i.d) samples
- Information gain measures the total information that could result from adding a new sample (observation) to the model
- Information gain is controlled by the information sharing capability of the kernel
- Information gain is controlled by the changes in model admitted by regularization
- Information gain will play a part in confidence intervals

#### **Confidence intervals**

- ullet Given that we have gathered t samples under the streaming setting, what kind of guarantees can we have on the resulting model?
- More specifically, could we guarantee that

$$|f(\mathbf{x}) - \hat{f}_t(\mathbf{x})| \leq \text{something}$$

simultaneously for all  $\mathbf{x} \in \mathcal{X}$  and for all  $t \geq 0$ ?

- Motivations:
  - Control bad behaviours in critical applications
  - Help selecting the next sample location
    - \* Maximize/minimize function?
    - \* Maximize model improvement?
  - Derive sound algorithms

## Result (Maillard, 2016)

• Under the assumption of subgaussian noise...

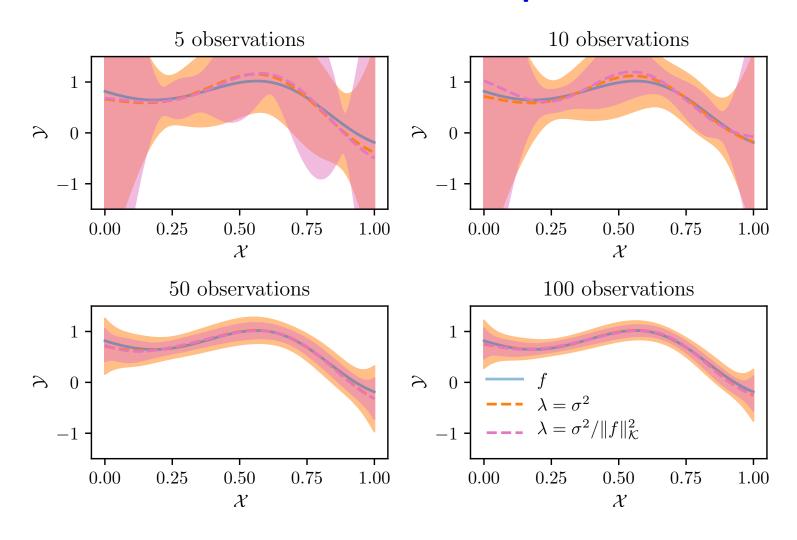
$$|f(\mathbf{x}) - \hat{f}_t(\mathbf{x})| \le \sqrt{\frac{k_{\lambda,t}(\mathbf{x},\mathbf{x})}{\lambda}} \left[ \sqrt{\lambda} ||f||_{\mathcal{K}} + \sigma \sqrt{2\ln(1/\delta) + 2\gamma_t(\lambda)} \right]$$

- ullet With probability higher than  $1-\delta$
- Simultaneously for all  $t \geq 0$ , for all  $\mathbf{x} \in \mathcal{X}$
- ⇒ Observe that the error bound scales with the information gain!

## **Subgaussian noise**

- ullet A real-valued random variable X is  $\sigma^2$ -subgaussian if  $\mathbb{E}\left[e^{\gamma X}\right] \leq e^{\gamma^2\sigma^2/2}$
- $\to$  The Laplace transform of X is dominated by the Laplace transform of a random variable sampled from  $\mathcal{N}(0,\sigma^2)$ 
  - Require that the tails of the noise distribution are dominated by the tails of a Gaussian distribution
  - For example, true for
    - Gaussian noise
    - Bounded noise

## **Confidence envelope**



## **Summary**

- It is possible to have guarantees even with non-i.i.d. data
- The predition error depends on
  - How well your model shares information across observations
  - How well your model is adapted to the function
  - How noisy the observations are
  - Regularization  $\rightarrow$  prior on  $\Sigma_{\mathbf{w}}$
- Confidence intervals/envelopes will be really useful for deriving algorithms (we will see that later)