

ASM Homework 1

Linear Model for IDMB data

Maria Gkotsopoulou & Ricard Monge Calvo & Amalia Vradi

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```
#####  
#####          DATA LOAD          #####  
#####
```

```
imdb <- read.csv("IMDB.csv", stringsAsFactors = F, sep=";")  
  
data.frame(variable = names(imdb),  
            class = sapply(imdb, class),  
            first_values = sapply(imdb, function(x) paste0(head(x), collapse = ", ")),  
            row.names = NULL)
```

```
##          variable      class  
## 1      movietitle character  
## 2          gross    integer  
## 3         budget    integer  
## 4      duration    integer  
## 5      titleyear    integer  
## 6    directorfl    integer  
## 7      actor1fl    integer  
## 8      actor2fl    integer  
## 9      actor3fl    integer  
## 10         castfl    integer  
## 11 facenumber_in_poster integer  
## 12          genre character  
##  
##                                     first_values  
## 1 10 Days in a Madhouse, 12 Years a Slave, 13 Going on 30, 21 & Over, 21 Grams, 25th Hour  
## 2                14616, 56667870, 56044241, 25675765, 16248701, 13060843  
## 3                12000000, 20000000, 37000000, 13000000, 20000000, 15000000  
## 4                111, 134, 98, 93, 124, 108  
## 5                2015, 2013, 2004, 2013, 2003, 2002  
## 6                0, 0, 56, 24, 0, 0  
## 7                1000, 2000, 3000, 552, 6000, 22000  
## 8                445, 660, 2000, 528, 979, 3000  
## 9                247, 500, 533, 499, 430, 346  
## 10               2059, 4251, 6742, 2730, 7567, 26050  
## 11                1, 0, 1, 0, 0, 0  
## 12               Drama, Drama, Comedy, Comedy, Drama, Drama
```

We first check for any missing values and see that there are no NAs.

```
summary(imdb)
```

```
##   movietitle      gross      budget  
## Length:940      Min.   :    3330   Min.   :   400000  
## Class :character 1st Qu.: 11816543 1st Qu.: 10000000
```

```
## Mode :character Median : 33428175 Median : 24000000
## Mean : 57813237 Mean : 40484550
## 3rd Qu.: 70756664 3rd Qu.: 48000000
## Max. :760505847 Max. :300000000
## duration titleyear directorfl actor1fl
## Min. : 74.0 Min. :2000 Min. : 0.0 Min. : 0.0
## 1st Qu.: 95.0 1st Qu.:2004 1st Qu.: 11.0 1st Qu.: 831.5
## Median :104.0 Median :2008 Median : 56.0 Median : 2000.0
## Mean :108.9 Mean :2008 Mean : 757.2 Mean : 9006.8
## 3rd Qu.:119.0 3rd Qu.:2012 3rd Qu.: 189.8 3rd Qu.: 13000.0
## Max. :280.0 Max. :2016 Max. :22000.0 Max. :640000.0
## actor2fl actor3fl castfl
## Min. : 0.0 Min. : 0.0 Min. : 0
## 1st Qu.: 462.5 1st Qu.: 255.0 1st Qu.: 2422
## Median : 756.0 Median : 501.0 Median : 4868
## Mean : 2391.7 Mean : 891.1 Mean : 13466
## 3rd Qu.: 1000.0 3rd Qu.: 748.2 3rd Qu.: 17659
## Max. :137000.0 Max. :19000.0 Max. :656730
## facenumber_in_poster genre
## Min. : 0.000 Length:940
## 1st Qu.: 0.000 Class :character
## Median : 1.000 Mode :character
## Mean : 1.624
## 3rd Qu.: 2.000
## Max. :31.000
```

Given the range of gross and budget we can switch to working in unit numbers by dividing by a million.

```
imdb<- imdb%>%
  mutate(gross = gross/1000000,
         budget = budget/1000000)
```

Exploratory Data Analysis

We are interested in predicting the gross of a movie basic on its characteristics. First let's analyze the target variable.

```
basicStats(imdb%>%dplyr::select(gross))
```

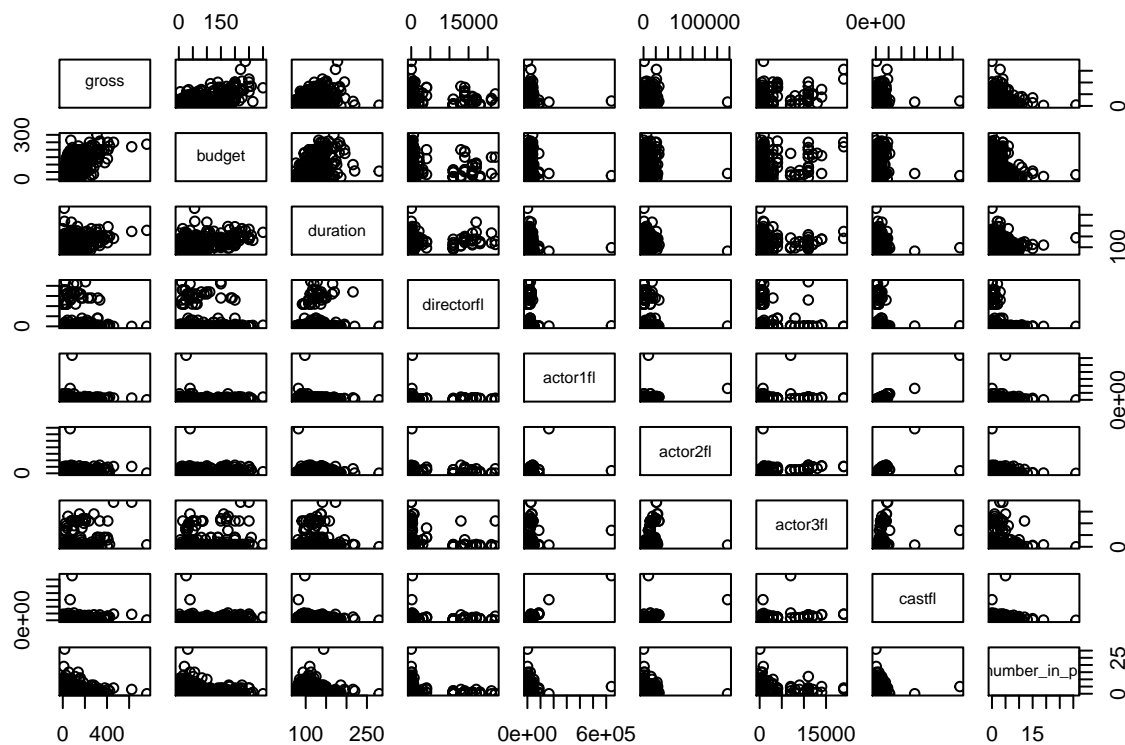
```
## gross
## nobs 940.000000
## NAs 0.000000
## Minimum 0.003330
## Maximum 760.505847
## 1. Quartile 11.816543
## 3. Quartile 70.756664
## Mean 57.813237
## Median 33.428175
## Sum 54344.442575
## SE Mean 2.515068
## LCL Mean 52.877432
## UCL Mean 62.749041
## Variance 5946.031921
## Stdev 77.110518
```

```
## Skewness      3.099129
## Kurtosis      14.530608
```

Using the `basicStats` we obtain the excess kurtosis, $K(X) - 3$ and we see that we have a considerable positive one and that it has a right skewed distribution. So, it is not normal. We should consider that the skewness and kurtosis could be due to outliers.

We look at an overview of the relationship between all variables in our dataset:

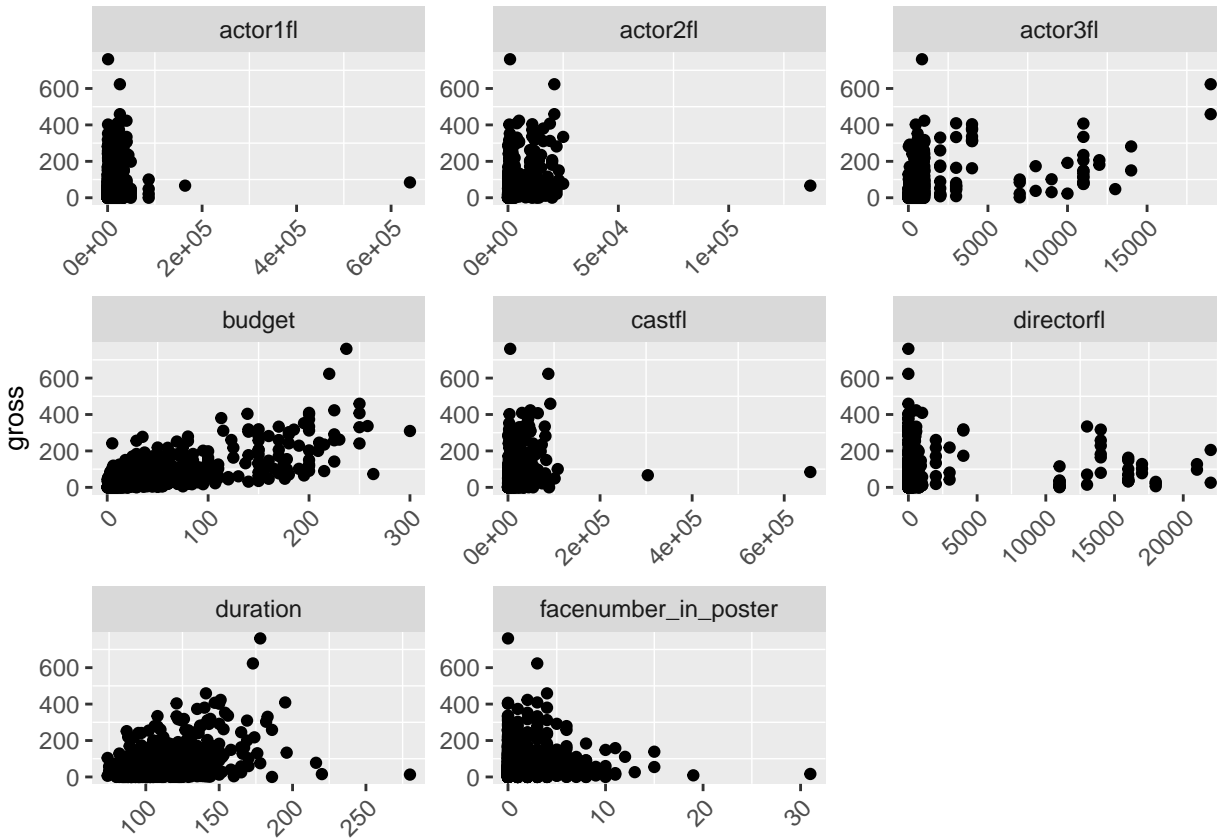
```
pairs(~.,imdb %>% select(-c(movietitle,genre, titleyear)))
```



In this plot we observe that some variables seem to be correlated, such as *actor1fl* with *castfl*, as well as, *budget* with *duration*. However, this correlation would present a problem, in the form of multicollinearity, in the case that both variables were to be included in the final model.

We now look closer into the relation between *gross* and all the numerical variables.

```
imdb %>%
  select(-c(movietitle,genre, titleyear)) %>%
  gather(-gross, key = "some_var_name", value = "some_value_name") %>%
  ggplot(aes(x = some_var_name, y = gross)) +
  geom_point() +
  facet_wrap(~ some_var_name, scales = "free")+
  ggstyleFonts +
  theme(panel.grid.major = element_blank(),
        axis.title.x = element_blank())
```



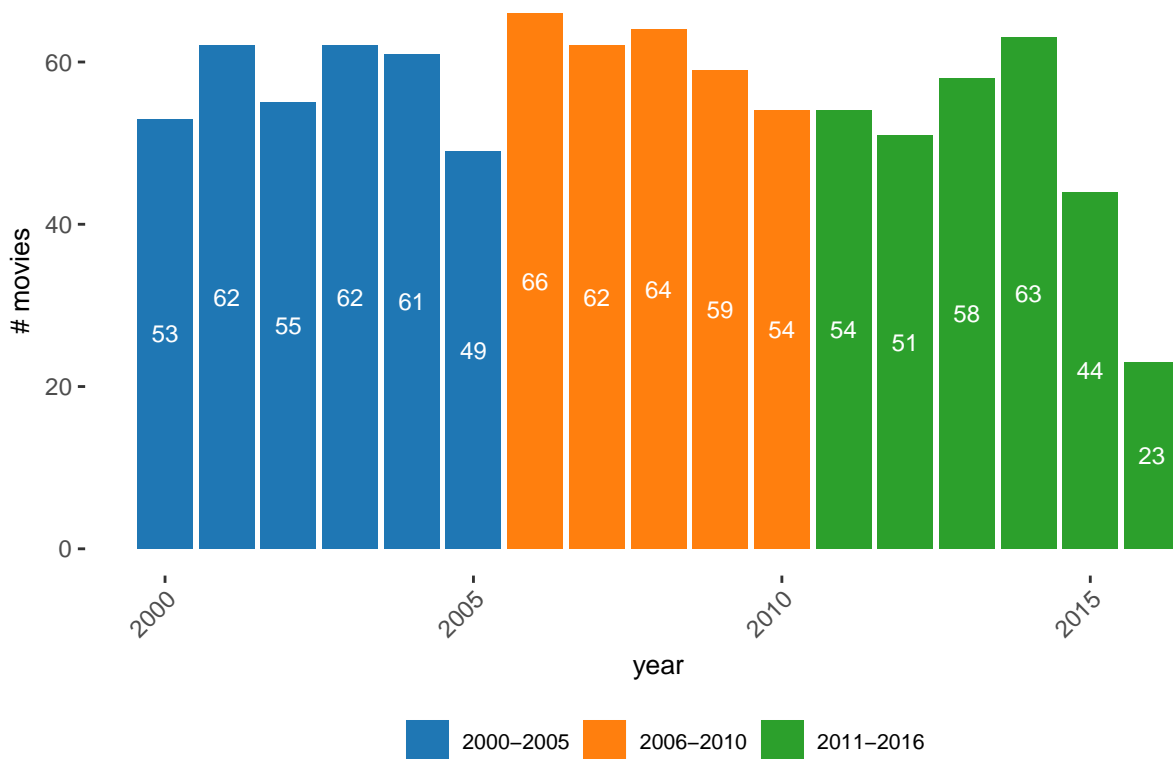
We observe more of a linear relation between the pairs of gross and budget, as well as with, duration. We can't discern any pattern between the pairs of gross and the Facebook variables: *directorfl*, *actor1fl*, *actor2fl*, *actor3fl*, *castfl*. The *actor3fl*, *directorfl* could be separated in 2 clusters at the cutoff point of 5000 likes and for the latter at the cutoff point of 10000 likes.

We create a categorial variable (*yearcat*) with 3 levels: 2000-2005, 2006-2010 and 2011-2016 based on the *titleyear* of the movie.

```
imdb <- imdb%>%
  mutate(yearcat = ifelse(titleyear < 2006, "2000-2005",
    ifelse(titleyear < 2011, "2006-2010", "2011-2016" )))

ggplot(imdb%>%
  group_by(titleyear, yearcat)%>%
  summarise(movies = n()) ,
  aes(x=titleyear, y=movies, fill= yearcat))+
  geom_bar(stat="identity") +
  geom_text(aes(label= movies),
    position=position_stack(vjust=0.5), colour="white" ,size=3) +
  scale_fill_d3(name="") +
  labs(y = "# movies", x = "year" , title = "Cluster movies into 3 categories by year") +
  ggstyle+
  theme(legend.position="bottom")
```

Cluster movies into 3 categories by year



```
imdb%>%
  group_by(titleyear, yearcat)%>%
  summarise(movies = n()) %>%
  group_by(yearcat) %>%
  mutate(avgMovies = mean(movies)) %>%
  summarise(movies = sum(movies),
            avgMovies = max(avgMovies)) %>%
  mutate(pcn = movies/sum(movies))
```

```
## # A tibble: 3 x 4
##   yearcat  movies avgMovies  pcn
##   <chr>    <int>    <dbl> <dbl>
## 1 2000-2005    342      57  0.364
## 2 2006-2010    305      61  0.324
## 3 2011-2016    293     48.8 0.312
```

The movies are roughly uniformly distributed between the three categories. However, on average more movies were released between the years 2006 and 2010. In addition, based on the significant difference between 2016 and all the previous years it is highly probable that we don't have data for the whole year. So, we have two categorical variables: the year category and the genre. Let's see how the economical variables relates to *genre*.

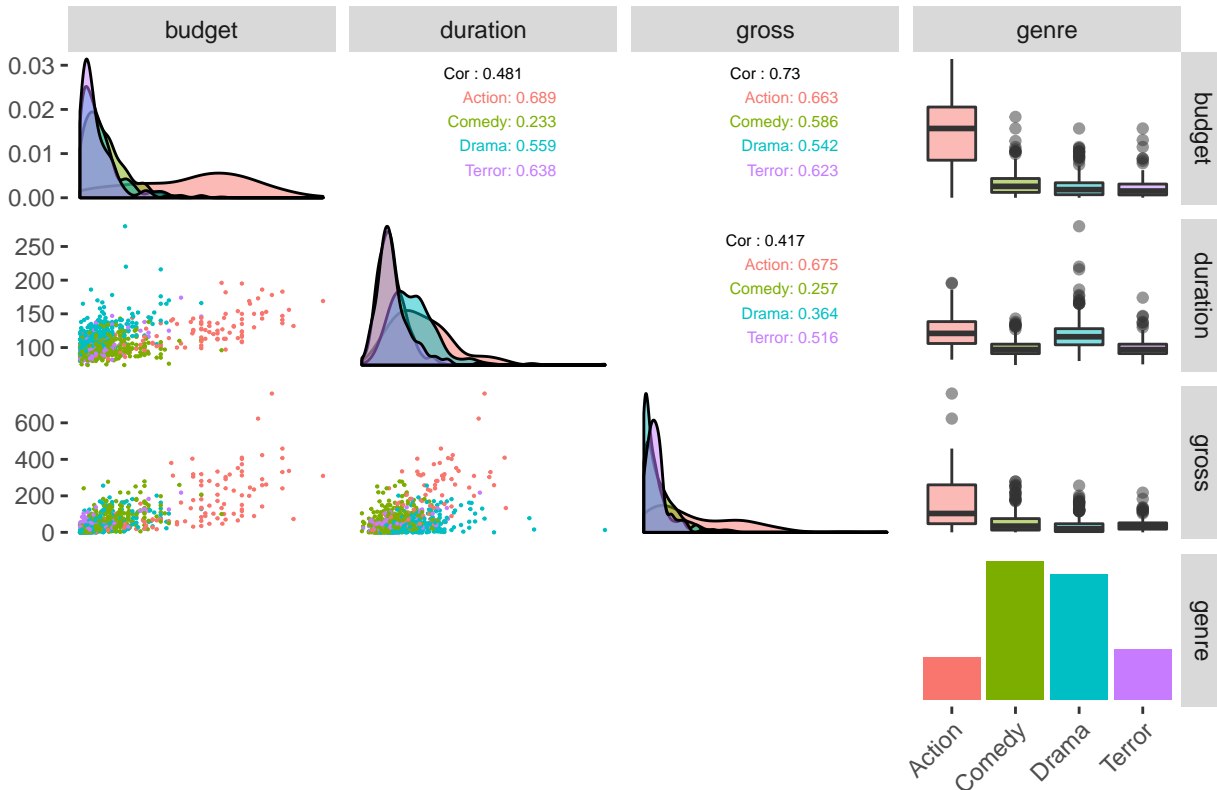
```
imdb %>%
  select(c(budget,duration,gross,genre)) %>%
  ggpairs(.,
    title = "Imdb economical variables relation by genre",
    mapping = ggplot2::aes(colour=genre),
    lower = list(#continuous = wrap("smooth", alpha = 0.3, size=0.1),
                 continuous = wrap("points", size=0.1),
```

```

discrete = "blank", combo="blank"),
diag = list(discrete="barDiag",
            continuous = wrap("densityDiag", alpha=0.5 )),
upper = list(combo = wrap("box_no_facet", alpha=0.5),
            continuous = wrap("cor", size=2, alignPercent=0.8))) +
ggstyle+
theme(panel.grid.major = element_blank())

```

Imdb economical variables relation by genre



We observe two outliers in the Action genre based on their *gross* value, which turn out to be blockbusters.

```
imdb %>% dplyr::filter(gross > 600) %>% pull(movietitle)
```

```
## [1] "Avatar" "The Avengers"
```

The distribution of *gross* for the Action genre is skewed to the right and has a higher IQR than the rest of the genres. However, it is also the genre with the smallest number of movies. Similarly, *budget* has excess kurtosis with more heavier tails than *gross* especially for the action movies. In the linear relation that we observed before between *gross* and *budget* we add now the genre which confirms this relation, particularly more for the Action movies.

```

imdb %>%
  select(c(budget,duration,gross,yearcat)) %>%
  ggpairs(.,
    title = "Imdb economical variables relation by Year group",
    mapping = ggplot2::aes(colour=yearcat),
    lower = list(#continuous = wrap("smooth", alpha = 0.3, size=0.1),
      continuous = wrap("points", size=0.1),
      discrete = "blank", combo="blank"),
    diag = list(discrete="barDiag",

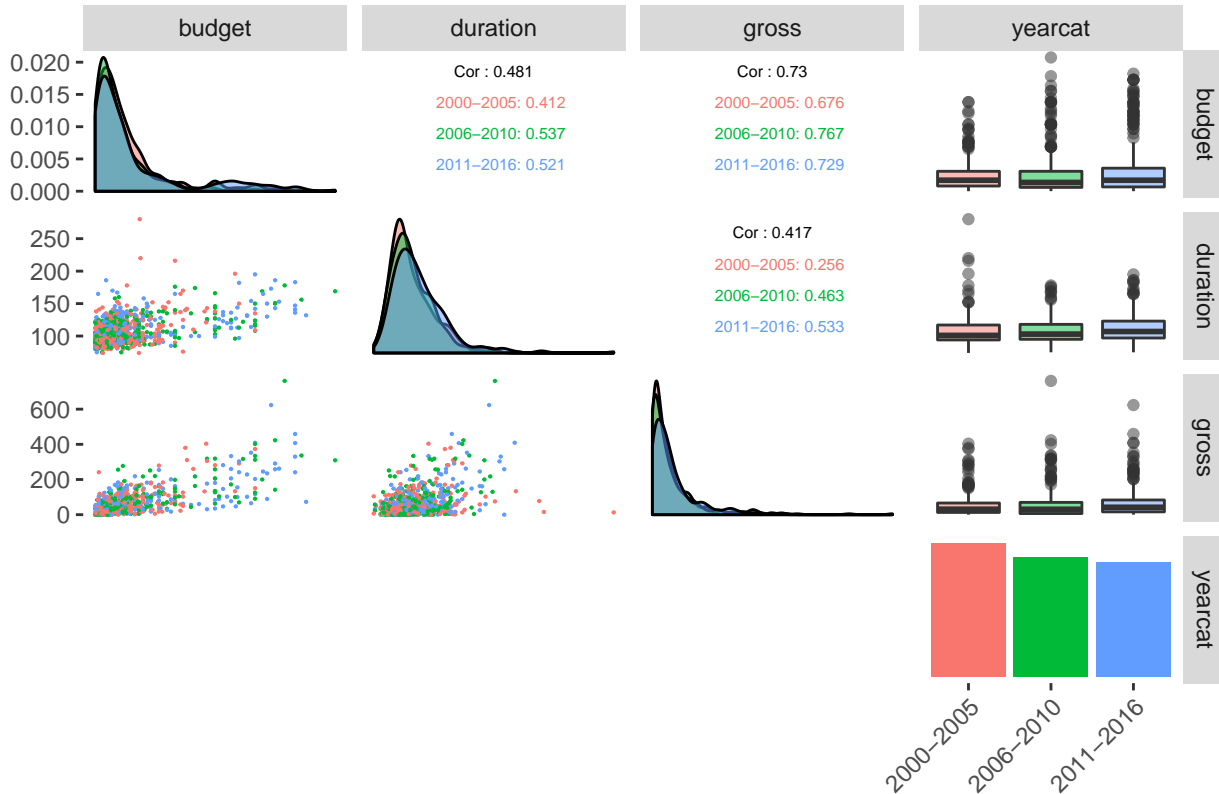
```

```

continuous = wrap("densityDiag", alpha=0.5 ),
upper = list(combo = wrap("box_no_facet", alpha=0.5),
             continuous = wrap("cor", size=2, alignPercent=0.8))) +
ggstyle+
theme(panel.grid.major = element_blank())

```

Imdb economical variables relation by Year group



On the other hand, we don't observe any differences between the different years.

Fit complete model

We first fit the complete model including as predictors, all the numerical variables, the two categorical variables, the categorical-categorical interactions and the interaction between numerical-categorical.

```

rownames(imdb) <- imdb$movietitle
imdb <- imdb %>% dplyr::select(-c(movietitle,titleyear))
mc<-lm(gross~(.-genre-yearcat)*(genre+yearcat)+genre:yearcat, imdb)

glance(mc)

```

```

## # A tibble: 1 x 11
##   r.squared adj.r.squared sigma statistic  p.value    df logLik   AIC
##   <dbl>      <dbl> <dbl>      <dbl>    <dbl> <int>  <dbl> <dbl>
## 1     0.659        0.636  46.5      28.8 5.50e-166    60 -4912. 9946.
## # ... with 3 more variables: BIC <dbl>, deviance <dbl>, df.residual <int>

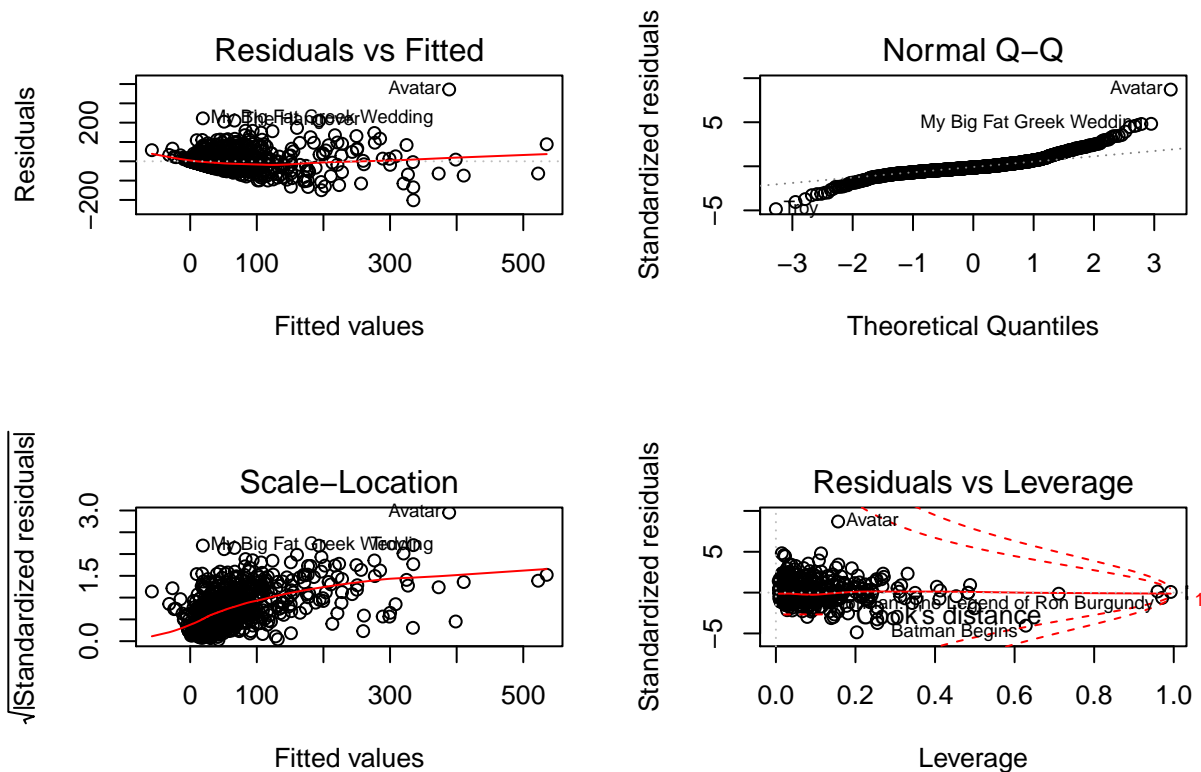
```

Roughly 64% of the variance found in the response variable (*gross*) can be explained by the predictor variables. The obtained p-value (*omnibustest*) indicates that the overall model is significant.

```

op<-par(mfrow=c(2,2))
plot(mc)

```



```
par(op)
```

From the *Normal Q-Q* plot we see that there is asymmetry in the distribution and we can conclude that normality of the residuals is not met. From the *Residuals vs Fitted* plot, we seek to validate the assumption of homoskedasticity, which does not seem to hold in our case. What's more, we observe, a non random distribution of the points along the $y - axis$. All in all, we can't validate this model. We look into this with more detail with the final model.

Select significant variables

We use the stepwise procedure, by using the *BIC* criterion, to select the significant variables. Since our objective is the prediction of *gross revenue* per movie, we choose as starting point the complete model, in contrast to starting from the null which would result in a simpler model, albeit loss in predictability.

```
summary(m1<-step(mc,direction="both",k=log(nrow(imdb)), trace = 0))
```

```
##
## Call:
## lm(formula = gross ~ budget + duration + actor1fl + actor2fl +
##      castfl + genre + yearcat + budget:yearcat + duration:genre +
##      actor1fl:genre + castfl:genre, data = imdb)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -220.59  -22.63   -7.31   14.33   364.11
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -2.257e+02  2.367e+01  -9.538  < 2e-16 ***
```



```
## budget          9.984e-01  9.146e-02  10.916 < 2e-16 ***
## duration        2.031e+00  2.122e-01   9.568 < 2e-16 ***
## actor1fl       -7.284e-03  8.989e-04  -8.103 1.70e-15 ***
## actor2fl       -3.135e-03  1.037e-03  -3.022 0.00258 **
## castfl         5.718e-03  6.329e-04   9.036 < 2e-16 ***
## genreComedy     1.745e+02  3.146e+01   5.546 3.82e-08 ***
## genreDrama      2.218e+02  2.712e+01   8.178 9.58e-16 ***
## genreTerror     2.087e+02  3.618e+01   5.768 1.10e-08 ***
## yearcat2006-2010 -2.321e+00  5.005e+00  -0.464 0.64289
## yearcat2011-2016 1.460e+01  5.144e+00   2.838 0.00464 **
## budget:yearcat2006-2010 3.773e-02  9.337e-02   0.404 0.68624
## budget:yearcat2011-2016 -4.085e-01  9.043e-02  -4.518 7.07e-06 ***
## duration:genreComedy -1.336e+00  2.942e-01  -4.541 6.33e-06 ***
## duration:genreDrama -1.970e+00  2.330e-01  -8.452 < 2e-16 ***
## duration:genreTerror -1.717e+00  3.369e-01  -5.099 4.16e-07 ***
## actor1fl:genreComedy 4.942e-03  1.074e-03   4.602 4.77e-06 ***
## actor1fl:genreDrama 3.686e-03  1.133e-03   3.254 0.00118 **
## actor1fl:genreTerror 3.622e-03  1.384e-03   2.616 0.00903 **
## castfl:genreComedy -3.340e-03  7.326e-04  -4.559 5.84e-06 ***
## castfl:genreDrama -2.189e-03  7.194e-04  -3.043 0.00241 **
## castfl:genreTerror -2.311e-03  8.920e-04  -2.590 0.00974 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 46.94 on 918 degrees of freedom
## Multiple R-squared:  0.6378, Adjusted R-squared:  0.6295
## F-statistic: 76.96 on 21 and 918 DF,  p-value: < 2.2e-16
```

Similarly to the complete model, we see that roughly 63% of the variance can be explained and the obtained p-value indicates that the overall model is significant.

In addition, we see that neither *actor3fl* nor *directorfl* are included in the model and hence, we could consider categorizing these variables as it was mentioned in the exploratory analysis and see whether we obtain a better model.

When dealing with categorical variables we should use the *Anova* method. The *p – value* obtained will allow us to say if the interaction variables are significant.

```
car::Anova(m1)
```

```
## Anova Table (Type II tests)
##
## Response: gross
##
```

	Sum Sq	Df	F value	Pr(>F)	
## budget	451563	1	204.9587	< 2.2e-16	***
## duration	57495	1	26.0962	3.951e-07	***
## actor1fl	105109	1	47.7078	9.236e-12	***
## actor2fl	20121	1	9.1325	0.0025808	**
## castfl	157997	1	71.7127	< 2.2e-16	***
## genre	57406	3	8.6853	1.099e-05	***
## yearcat	1002	2	0.2275	0.7965685	
## budget:yearcat	100097	2	22.7163	2.349e-10	***
## duration:genre	161349	3	24.4114	3.339e-15	***
## actor1fl:genre	46983	3	7.1083	0.0001007	***

```
## castfl:genre      47668    3    7.2119 8.709e-05 ***
## Residuals        2022528 918
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

We see that the interaction variables *budget:yearcat*, *duration:genre*, *actor1fl:genre* and *castfl:genre* are significant, so we keep them in our model. Nevertheless, the variable *yearcat* seems to not be significant, so we could remove it from our model. **PREGUNTAR SI TIENE SENTIDO QUEDARSE CON YEARCAT Y/O SUS INTERACCIONES → IF THE INTERACTION OF THE VARIABLE WITH OTHERS IS SIGNIFICANT BUT THE VARIABLE ITSELF IS NOT, WE KEEP BOTH VARIABLE AND INTERACTION.**

Check for multicollinearity

Strong associations between predictors will increase standard errors, and therefore increase the probability of a type-II error. The diagnostic that we will use is the variance-inflation factor.

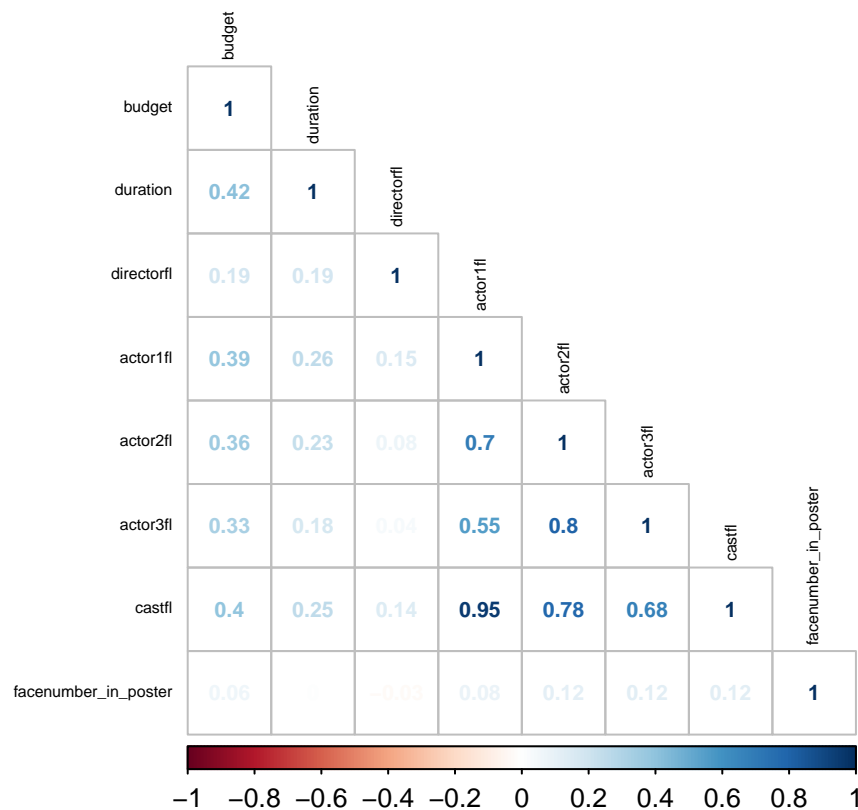
```
car::vif(m1)
```

```
##              GVIF Df GVIF^(1/(2*Df))
## budget          8.800183  1      2.966510
## duration        8.298531  1      2.880717
## actor1fl       198.636593  1     14.093849
## actor2fl       17.143583  1      4.140481
## castfl         132.845959  1     11.525882
## genre          72188.876411  3      6.452753
## yearcat         3.227704  2      1.340366
## budget:yearcat  13.103015  2      1.902579
## duration:genre 81351.402009  3      6.582550
## actor1fl:genre 94415.699560  3      6.747984
## castfl:genre   67158.366057  3      6.375536
```

COMO SE LEE ESTE RESULTADO???? → LEER COLUMNA “NORMALIZADA” CUANDO TENEMOS CATEGORICAL VARIABLES. SI SALE > 5 ES POSIBLE QUE TENGA COLLINIARITY. TRY WITHOUT ONE THE VARIABLES WITH HIGH VIF AND SEE IF THE VIF IMPROVES. IF WE WANT TO BETTER INTERPRET THE RELATIONS WITH THE RESPONSE VARIABLES, WE WANT TO REMOVE THE COLLINIARITY TO BETTER INTERPRET THE COEFFICIENTS, EVEN WITH WORSE R².

castfl and *actor1fl*, *actor2fl* which are included in the model are correlated, as well as *actor1fl* and *actor2fl*. Consequently, the coefficients can't be directly interpreted. The following correlations confirm this idea:

```
corr_mat=cor(imdb %>%
  select_if(is.numeric) %>%
  select(-c(gross)),method="s")
corrplot::corrplot(corr_mat,type = "lower", method = "number",
  tl.col = "black", tl.cex = 0.5,number.cex = 0.7)
```



To proceed we would need to select which of these three correlated variables would result in a better model.

```
lm.actor1 <- update(m1, .~.(castfl+actor2fl))
lm.actor2 <- update(m1, .~.(castfl+actor1fl))
lm.cast <- update(m1, .~.(actor1fl+actor2fl))
```

```
glance(lm.actor1)
```

```
## # A tibble: 1 x 11
##   r.squared adj.r.squared sigma statistic    p.value    df logLik   AIC
##   <dbl>      <dbl> <dbl>      <dbl>      <dbl> <int> <dbl> <dbl>
## 1     0.634        0.626  47.1        79.6 1.01e-184    21 -4945. 9934.
## # ... with 3 more variables: BIC <dbl>, deviance <dbl>, df.residual <int>
```

```
glance(lm.actor2)
```

```
## # A tibble: 1 x 11
##   r.squared adj.r.squared sigma statistic    p.value    df logLik   AIC
##   <dbl>      <dbl> <dbl>      <dbl>      <dbl> <int> <dbl> <dbl>
## 1     0.638        0.629  46.9        77.0 1.03e-185    22 -4941. 9927.
## # ... with 3 more variables: BIC <dbl>, deviance <dbl>, df.residual <int>
```

```
glance(lm.cast)
```

```
## # A tibble: 1 x 11
##   r.squared adj.r.squared sigma statistic    p.value    df logLik   AIC
##   <dbl>      <dbl> <dbl>      <dbl>      <dbl> <int> <dbl> <dbl>
## 1     0.634        0.626  47.1        79.6 1.01e-184    21 -4945. 9934.
## # ... with 3 more variables: BIC <dbl>, deviance <dbl>, df.residual <int>
```

Since we do not obtain a significantly better model with any of the variables, we conclude we need other information, such as domain expert knowledge, to decide which predictor to keep. In our case, we decide to keep *castfl* as its and added variable of the likes of the whole movie cast (it includes the other measures).

```
m1 <- update(m1, ~.-actor1fl-actor2fl)
```

We check the *Anova* of the new model to see how it has changed.

```
car::Anova(m1)

## Anova Table (Type II tests)
##
## Response: gross
##
```

	Sum Sq	Df	F value	Pr(>F)
budget	458768	1	206.403	< 2.2e-16 ***
duration	60467	1	27.204	2.262e-07 ***
castfl	127264	1	57.257	9.280e-14 ***
genre	74532	3	11.177	3.295e-07 ***
yearcat	827	2	0.186	0.8303
budget:yearcat	93148	2	20.954	1.262e-09 ***
duration:genre	161420	3	24.208	4.407e-15 ***
genre:actor1fl	139159	4	15.652	2.135e-12 ***
castfl:genre	100795	3	15.116	1.310e-09 ***
Residuals	2042648	919		

```
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

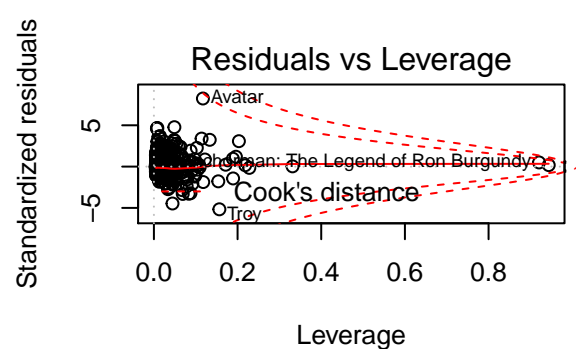
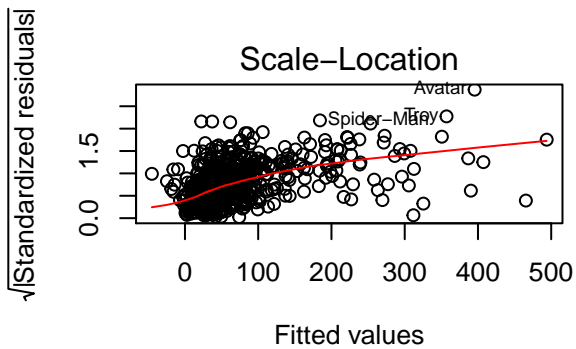
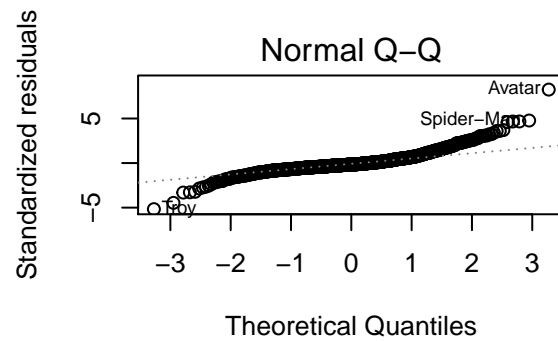
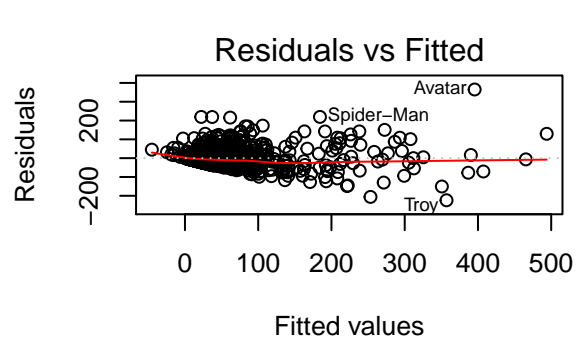
```
car::vif(m1)

##
```

	GVIF	Df	GVIF ^{1/(2*Df)}
budget	8.799477e+00	1	2.966391
duration	8.293560e+00	1	2.879854
castfl	6.684936e+01	1	8.176146
genre	7.207881e+04	3	6.451112
yearcat	3.223282e+00	2	1.339907
budget:yearcat	1.299691e+01	2	1.898716
duration:genre	8.127410e+04	3	6.581507
genre:actor1fl	1.385808e+05	4	4.392511
castfl:genre	4.405568e+04	3	5.942926

Validate model's assumptions

```
op<-par(mfrow=c(2,2))
plot(m1)
```



```
par(op)
```

Model interpretation

```
# plot(allEffects(m1))
```

TO INTERPRET VARIABLES WITH INTERACTIONS, WE PLOT THE EFFECT OF THE RESPONSE TO THE VARIABLE. WHEN INTERACTION IS WITH A CATEGORICAL VARIABLE THE EFFECTS ARE PLOTS STRATIFIED BY CATEGORICAL VARIABLE LEVEL.