# Probability Theory, Mutual Information, and Taxonomy of Probabilistic models

Robert Bamler · 12 May 2022

This lecture is a part of the Course "Data Compression With and Without Deep Probabilistic Models" at University of Tübingen.

More course materials (lecture notes, problem sets, solutions, and videos) are available at:

https://robamler.github.io/teaching/compress22/



#### Recap From Last Week (1 of 3): Two probability distributions

- $\triangleright$   $p_{\text{data}}$ : true probability distribution of the data generative process
  - $\triangleright$  typically unknown, i.e., we can't *evaluate* the true probability  $p_{data}(\mathbf{x})$  of a given message  $\mathbf{x}$ ;
  - but we may have access to a data set  $\mathcal{D}$  of empirical samples from  $p_{\text{data}}$ .
    - $\Rightarrow$  then we can *estimate* expectations under  $p_{data}$  as follows:

$$\mathbb{E}_{\mathbf{x} \sim p_{\text{data}}} \big[ f(\mathbf{x}) \big] = \lim_{|\mathcal{D}| \to \infty} \frac{1}{|\mathcal{D}|} \sum_{\mathbf{x} \in \mathcal{D}} f(\mathbf{x}) \qquad \text{(assuming i.i.d. samples and expectation exists)}$$

- p<sub>model</sub>: probabilistic model of the data source
  - approximates p<sub>data</sub>;
  - let's assume, for now, that we can evaluate  $p_{data}(\mathbf{x}) \in [0, 1]$  for any given message  $\mathbf{x}$ .

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# Recap From Last Week (2 of 3): Entropy vs. Cross Entropy

Consider the expected bitrate  $\mathbb{E}_{\mathbf{x} \sim p_{\text{data}}}[R_C(\mathbf{x})]$  of a lossless compression code C:

fundamental lower bound: true entropy of the data source:

Entropy: 
$$H[p_{data}] = \mathbb{E}_{\mathbf{x} \sim p_{data}}[-\log_2 p_{data}(\mathbf{x})]$$

- (i.e., independent of our model).
- $\odot$  We can't reach this bound because we can't optimize C for  $p_{\text{data}}$ .
- $\odot$  We can't even calculate  $H[p_{data}]$  because we can't evaluate  $p_{data}(\mathbf{x})$ .
- **practically achievable expected bit rate:** cross entropy between  $p_{\text{data}} \& p_{\text{model}}$ :

Cross entropy: 
$$H(p_{\text{data}}, p_{\text{model}}) = \mathbb{E}_{\mathbf{x} \sim p_{\text{data}}}[-\log_2 p_{\text{model}}(\mathbf{x})]$$

- $\odot$  Assumes that the code *C* is optimal for  $p_{\text{model}}$ , which is more realistic.



#### Recap From Last Week (3 of 3): Kullback-Leibler (KL) Divergence

We need accurate probabilistic models to achieve good compression performance.

▶ Modeling error: How many additional bits do we need to transmit (in expectation) due to an inaccurate model?

Kullback-Leibler Divergence:  $D_{\mathsf{KL}}(p_{\mathsf{data}} \mid\mid p_{\mathsf{model}}) := H(p_{\mathsf{data}}, p_{\mathsf{model}}) - H[p_{\mathsf{data}}]$ 

- ▶ Problem 3.1: explicit proof that  $D_{KL}(p,q) \ge 0$  for any p and q ("Gibb's theorem")
- ▶ Problem 3.2: fit  $p_{\text{model}}$  to a data set by minimizing  $D_{\text{KL}}(p_{\text{data}}, p_{\text{model}})$  numerically
  - ▶ To reach low  $D_{KL}(p_{data}, p_{model})$ , we need an expressive model architecture.

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## Interlude: Probability Theory & Random Variables

What makes up a probabilistic model:

- $\triangleright$  sample space  $\Omega$  (abstract space of "all states of the world")
  - event  $E \subset \Omega$ : "event E occurs" := "the world is in some state  $\omega \in E$ ".
- ▶ probability measure: a function  $P : \Sigma \rightarrow [0, 1]$  where
  - ightharpoonup Σ is a so-called  $\sigma$ -algebra on Ω. (a set of all "expressible" events  $E \subset \Omega$ )
  - $ightharpoonup P(\emptyset) = 0$  and  $P(\Omega) = 1$ .
  - ► countable additivity:  $P\left(\bigcup_{i=1}^{\infty} E_i\right) = \sum_{i=1}^{\infty} P(E_i)$  if all  $E_i$  are pairwise disjoint.
  - ▶ therefore:  $P\left(\bigcup_{i=1}^{k}\right) = \sum_{i=1}^{k} P(E_i)$  if all  $E_i$  are pairwise disjoint. (proof: set  $E_i = \emptyset \ \forall i > k$ )
  - ▶ therefore:  $P(E) + P(\Omega \setminus E) = P(\Omega) = 1 \quad \forall E \in \Sigma$ .

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## **Examples of Probability Measures**

- 1. Simplified Game of Monopoly: (throw two fair three-sided dice)
  - $\blacktriangleright \ \text{ sample space: } \Omega = \{1,2,3\}^2 = \big\{(1,1),(1,2),(1,3),(2,1),(2,2),(2,3),(3,1),(3,2),(3,3)\big\}$
  - $\blacktriangleright \ \ \text{sigma algebra:} \ \Sigma = 2^{\Omega} := \big\{ \text{all subsets of} \ \Omega \ (\text{including} \ \emptyset \ \text{and} \ \Omega) \big\}$
  - ▶ probability measure *P*: for all  $E \subset \Sigma$ , let  $P(E) := |E|/|\Omega| = |E|/|9$



#### **Examples of Probability Measures (cont'd)**

- 2. Departure times of the next three buses from "Sternwarte":
  - sample space (in a simple model):  $\Omega = \mathbb{R}^3$
  - ▶ sigma algebra: all "measurable subsets" of  $\mathbb{R}^3$  (think of this as all subsets of  $\mathbb{R}^3$  except for extremely pathological exceptions)
  - probability measure P: complicated function, but we know it satisfies certain relations, e.g., P("next bus departs before 1.15 pm") = P("next bus departs before 1.10 pm") + P("next bus departs between 1.10 and 1.15 pm").
  - **Question:** what is the probability that the next bus departs *exactly* at 1.10 pm? I.e., what is  $P(\{1.10 \text{ pm}\} \times \mathbb{R}^2)$ ?
  - **Question:** what is the probability that the next bus departs between 1.10 pm and 1.15 pm?  $P\left(\underbrace{[1.10 \text{ pm}, 1.15 \text{ pm}]}_{=:\mathcal{I}} \times \mathbb{R}^2\right) = P\left(\bigcup_{x_1 \in \mathcal{I}} \{x_1\} \times \mathbb{R}^2\right) \stackrel{?}{=} \sum_{x_1 \in \mathcal{I}} P(\{x_1\} \times \mathbb{R}^2) = 0$

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#### **Random Variables**

- ▶ Often, we we're not interested in the *full* description of the state  $\omega \in \Omega$ , but only in certain properties of it.
- ▶ *Def. random variable:* function  $X : \Omega \to \mathbb{R}$  (not necessarily injective)

#### **Examples:**

- 1. Simplified Game of Monopoly;  $\Omega = \{(a, b) \text{ where } a, b \in \{1, 2, 3\}\}$ 
  - ► total value:  $X_{sum}((a,b)) = a + b \in \{2,3,4,5,6\}$
  - value of the red die:  $X_{red}((a,b)) = a$
  - value of the blue die:  $X_{\text{blue}}((a,b)) = b$
- 2. In our bus schedule model from before;  $\Omega = \mathbb{R}^3$ 
  - ▶ Time between the next bus and the one after it:  $X_{gap}((x_1, x_2, x_3)) = x_2 x_1$

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# **Properties of Individual Random Variables**

▶ "Probability that a random variable *X* has some given value *x*":

$$P(X = X) := P(X^{-1}(X)) = P(\{\omega \in \Omega : X(\omega) = X\})$$

- ▶ Example 1 (Simplified Game of Monopoly):  $P(X_{sum} = 3) =$
- **Example 2 (bus schedule):**  $P(X_{gap} = 20 \text{ minutes}) =$
- ▶ When we write just P(X), then we mean the *function* that maps  $x \mapsto P(X = x)$ .
- ► Expectation value of a random variable X under a model P
  - ▶ discrete case:  $\mathbb{E}_{P}[X] := \sum_{\omega \in \Omega} P(\{\omega\}) X(\omega) = \sum_{x \in X(\Omega)} P(X = x) x$ 
    - examples:  $\mathbb{E}_P[X_{\mathsf{red}}] =$ ;  $\mathbb{E}_P[X_{\mathsf{blue}}] =$ ;  $\mathbb{E}_P[X_{\mathsf{sum}}] =$
  - **•** continuous case:  $\mathbb{E}_P[X] := \int_{\Omega} X(\omega) \, dP(\omega)$  (see next slide)



#### Properties of Individual Random Variables (cont'd)

- ▶ Cumulative Density Function (CDF):  $P(X \le x) := P(\{\omega \in \Omega : X(\omega) \le x\})$ 
  - ▶ Example 1 (Simplified Game of Monopoly):  $P(X_{sum} \le 3) =$
  - **Example 2 (bus schedule):**  $P(X_{gap} \le 20 \text{ minutes}) \in [0, 1]$  (can be nonzero)
- ▶ Analogous definitions for: P(X < x),  $P(X \ge x)$ , P(X > x),  $P(X \in \text{some set})$ , . . .
- ▶ *Probability Density Function (PDF)* of a real-valued random variable *X*:

$$p(x) := \frac{d}{dx}P(X \le x)$$
 (if derivative exists)

 $\rightarrow$  expectation value:  $\mathbb{E}_P[X] = \int X(\omega) dP(\omega) = \int_{-\infty}^{\infty} x p(x) dx$  (if a density p(x) exists)

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## Multiple Random Variables: Joint & Marginal Probability Distributions

▶ Def. *joint probability distribution* of two random variables *X* and *Y*:

$$P(X=x, Y=y) = P(\{\omega \in \Omega : X(\omega) = x \land Y(\omega) = y\})$$

- ▶ (notation: "P(X, Y)": function that maps  $(x, y) \mapsto P(X = x, Y = x)$ )
- ▶ If we know P(X, Y), then we can calculate  $P(X) = \sum_{y} P(X, Y = y)$  (for discrete Y)
  - ► This process is called "marginalization".
  - for continuous random variables:  $P(X) = \int P(X, Y = y) dy$

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# Multiple Random Variables: Statistical Independence

- ▶ Def.: X and Y are (statistically) independent if and only if: P(X, Y) = P(X) P(Y)(i.e., if  $P(X=x, Y=y) = P(X=x) P(Y=y) \forall x, y$ )
- ► Examples (Simplified Game of Monopoly):
  - $\triangleright$   $X_{\text{red}}$  and  $X_{\text{blue}}$  are statistically independent.
  - $ightharpoonup X_{red}$  and  $X_{sum}$  are *not* statistically independent. (proof: Problem 3.1)
- ▶ Def.: conditional independence of X and Y given Z: see later



## **Conditional Probability Distributions**

 $|"X \& Y \text{ are } not \text{ statistically independent"}| \Leftrightarrow |"knowing X reveals something about Y"$ 

, ,			<u> </u>				
<b>Examples:</b> (Simplified Game of Monopoly; $P(E) = \frac{ E }{9}$ )	X =	1	2	3	4	5	6
What are the (marginal) probability distributions $P(X_{red})$		-		1 3 2	0	0	0
and $P(X_{sum})$ of the red die and the sum, respectively?	$P(X_{sum} = x) =$	0	9	<u>2</u> 9	3	2 9	<u>1</u> 9
Assume that you only accept throws where the red die comes up with value 1, and you keep rethrowing both dice until this condition is satisfied. What is the probability distribution of $X_{\text{sum}}$ in your first accepted throw? We call this the <i>conditional</i> probability distribution $P(X_{\text{sum}} \mid X_{\text{red}} = 1)$ .	$P(X_{\text{sum}} = x \mid X_{\text{red}} = 1) =$						
Now you only accept throws where the sum of both dies is 3. What is the conditional probability distribution of $X_{\text{red}}$ ?	$P(X_{\text{red}} = x \mid X_{\text{sum}} = 3) =$						
Finally, assume you only accept throws where $X_{\text{blue}} = 1$ . What is the conditional probability distribution of $X_{\text{red}}$ ?	$P(X_{\text{red}} = x \mid X_{\text{blue}} = 1) =$						

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## Conditional Probability Distributions (cont'd)

- ▶ Def. "conditional probability of event  $E_2$  given event  $E_1$ ":  $P(E_2 \mid E_1) := \frac{P(E_1 \cap E_2)}{P(E_1)}$ 
  - ▶ Thus,  $P(E_2 | E_1)$  is a (properly normalized) probability distribution with respect to the first parameter, i.e.,  $P(E_2 \mid E_1) + P(\Omega \setminus E_2 \mid E_1) = \frac{P(E_2 \cap E_1) + P((\Omega \setminus E_2) \cap E_1)}{P(E_1)} = \frac{P(E_1)}{P(E_1)} = 1$ .
- ▶ Def. "conditional probability distribution of random variable Y given random variable X":  $P(Y | X) := \frac{P(X,Y)}{P(X)}$  i.e.,  $P(Y = y | X = x) := \frac{P(X = x,Y = y)}{P(X = x)}$   $\forall x, y$ 
  - ▶ Thus, if *X* and *Y* are statistically independent (but only then!):  $P(Y | X) = \frac{P(X,Y)}{P(X)} = \frac{P(X)P(Y)}{P(X)} = P(Y)$ ("knowing X reveals no new information about Y")
  - ▶ In the general case: "chain rule" of probability theory: (follows directly from above def.)  $P(X_1, X_2, X_3, ...) =$

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## Warning: Conditionality $\neq$ Causation

- ▶ We'll often specify a joint problty. distribution as, e.g.,  $P(X, Y) = P(X) P(Y \mid X)$ .
- ▶ But just because we write down "P(Y | X)", this does not necessarily mean that X is the cause for Y.
- Example: (Simplified Game of Monopoly):
  - $ightharpoonup X_{red}$  and  $X_{blue}$  can be considered the cause for  $X_{sum}$ .
  - ▶ But, in the examples two slides ago, we were still able to calculate, e.g.,  $P(X_{red} | X_{sum})$ . (i.e., the probability of the cause  $X_{red}$  given its effect  $X_{sum}$ )
    - → This is called "posterior inference". (more in Lectures 6 and 7)



# Next Step:

# tying it back to information theory

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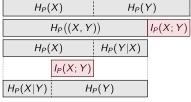
Definitions as you'd expect:

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- information content of the statement "X = x":
- entropy of a random variable X under a model P:  $H_P(X) :=$
- joint and conditional information content and entropy: see Problems 4.2 and 4.3.
- Subadditivity of entropies: ∀ random vars X and Y:

$$H_P((X, Y)) \le H_P(X) + H_P(Y)$$
 (proof: Problem 4.4)

- equality holds if X and Y are statistically independent (proof: Problem 2.3 (b))
- ► Thus, wrongfully assuming independence (to simplify the model) leads to a compression overhead of  $I_P(X; Y)$  bits:



(figure adapted from MacKay book)

Def. mutual information:  $|I_P(X; Y) := H_P(X) + H_P(Y) - H_P((X, Y)) \ge 0$ (see Problem 4.4)





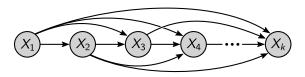
## Deep Probabilistic Models: Overview and Taxonomy

- Assume that the message is a sequence of symbols:  $\mathbf{X} = (X_1, X_2, \dots, X_k)$
- ▶ Subadditivity of entropies:  $H(\mathbf{X}) \leq \sum_{i=1}^{n} H(X_i)$
- $\triangleright$  Thus: instead of modeling each symbol  $X_i$  independently, we should model the message **X** as a whole (without completely sacrificing computational efficiency).
  - autoregressive models (e.g., Problem 3.2)
  - latent variable models (planned for Problem Set 6; also: basis for variational autoencoders)



#### **Probabilistic Models at Scale**

▶ All probabilistic models P over messages  $\mathbf{X} = (X_1, X_2, \dots, X_k)$  satisfy chain rule:  $P(\mathbf{X}) = P(X_1) P(X_2 \mid X_1) P(X_3 \mid X_1, X_2) P(X_4 \mid X_1, X_2, X_3) \cdots P(X_k \mid X_1, X_2, \dots, X_{k-1})$ 



- ▶ Assume each symbol is from alphabet  $\mathfrak{X} = \{1, 2, 3\}$ .
  - $\blacktriangleright$  How many model parameters do we need to specify an arbitrary distribution  $P(X_1)$ ?
  - ▶ How many parameters for an arbitrary conditional distribution  $P(X_2 \mid X_1)$ ?
  - ▶ How many parameters for an arbitrary conditional distribution  $P(X_k | X_1, X_2, ..., X_k)$ ?

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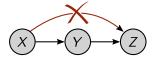


## **Expressive Yet Efficient Probabilistic Models**

- ▶ **Goal:** Find approximation to arbitrary models P(X) that
  - captures relevant correlations
  - but is still *computationally efficient*:
    - → reasonably compact representation of the model in memory
    - ightarrow reasonably efficient evaluation of probabilities  $P(\mathbf{X} = \mathbf{x})$
- ► **General Strategy:** enforce *conditional independence*:

X & Z are conditionally independent given  $Y : \iff P(X, Z \mid Y) = P(X \mid Y) P(Z \mid Y)$ 

 $\iff$  P(X, Y, Z) = P(X) P(Y | X) P(Z | Y) (proof: Problem Set 5)



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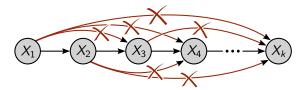
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## **Four Approximation Schemes**

- (a) Markov Process: assume symbols  $X_i$  are generated by a memoryless process
  - ► Each symbol  $X_{i+1}$  is conditioned on the immediately preceding symbol  $X_i$  but not on any earlier symbols:  $P(\mathbf{X}) = P(X_1) P(X_2 \mid X_1) P(X_3 \mid X_2) P(X_4 \mid X_3) \cdots P(X_k \mid X_{k-1})$



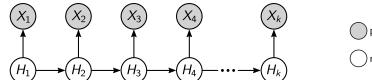
- $\blacktriangleright$  i.e., for all j < i, the symbols  $X_{i+1}$  and  $X_i$  are conditionally independent given  $X_i$ .
- $\odot$  only  $O(k |\mathfrak{X}|^2)$  (or even  $O(|\mathfrak{X}|^2)$ ) model parameters
- $\odot$  simplistic assumption; e.g., in English text, the string "the" is very frequent.  $\Rightarrow P(X_{i+1} = \text{`e'} \mid X_i = \text{`h'}, X_{i-1} = \text{`t'}) > P(X_{i+1} = \text{`e'} \mid X_i = \text{`h'}) \text{ (i.e., not cond. indep.)}$



#### Four Approximation Schemes (cont'd)

#### (b) Hidden Markov Model:

▶ Morkov Process of "hidden" states *H<sub>i</sub>* that are only indirectly observed



- part of the message ("observed")
- not part of the message ("latent")

$$P(\mathbf{X}) = \int P(\mathbf{X}, \mathbf{H}) d\mathbf{H}$$
 with  $P(\mathbf{X}, \mathbf{H}) = P(H_1) P(X_1 \mid H_1) \prod_{i=2}^{k} P(H_i \mid H_{i-1}) P(X_i \mid H_i)$ 

- © can model long-range correlations (exercise)
- $\odot$  overhead: in order to model  $P(X_i | H_i)$ , decoder has to decode  $H_i$ , even though it's not part of the message (solution: "bits-back coding", Lecture 6)

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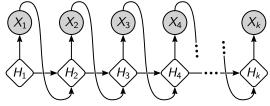




## Four Approximation Schemes (cont'd)

#### (c) Autoregressive Model:

 $\blacktriangleright$  similar to a hidden Markov model, but the hidden process is *deterministic*:  $H_{i+1} = f(H_i, X_i)$ 



- part of the message ("observed")
- deterministic function of its inputs

$$P(\mathbf{X}) = P(X_1 \mid H_1) \prod_{i=0}^{k} P(X_i \mid H_i)$$
 where  $H_1 = \text{fixed}; H_{i+1} = f(H_i, X_i)$ 

- © no compression overhead for reconstructing **H** (see Problem 3.2)
- © encoding & decoding are not parallelizable (⇒ slow on modern hardware)

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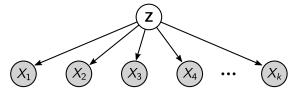




## Four Approximation Schemes (cont'd)

# (d) Latent Variable Models: "explain" message X by some unobserved Z

 Intuition: Z captures message at a higher level of abstraction (e.g., the "topic" of a piece of text, or basic shapes in an image)



- part of the message ("observed")
- not part of the message ("latent")

$$P(\mathbf{X}) = \int P(\mathbf{X}, \mathbf{Z}) d\mathbf{Z}$$
 with  $P(\mathbf{X}, \mathbf{Z}) = P(\mathbf{Z}) \prod_{i=1}^{k} P(X_i | \mathbf{Z})$ 

- © can model correlations (see Problem Set 6) and is parallelizable
- © compression overhead for encoding Z (solution: "bits-back coding", Lecture 6)

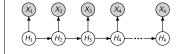


#### **Recap: Four Approximation Schemes**

#### **Markov Process**

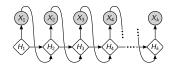
## $(X_1) \longrightarrow (X_2) \longrightarrow (X_3) \longrightarrow (X_4) - \cdots \longrightarrow (X_k)$

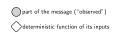
#### **Hidden Markov Model**



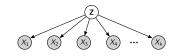


#### **Autoregressive Model**





#### **Latent Variable Model**



part of the message ("observed")

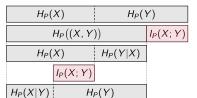
not part of the message ("latent")

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# Outlook



▶ Problem Set:

- ▶ Next ~4 weeks: lossless compression with deep probabilistic models
  - → Different model architectures require different compression algorithms.
    - ► Problem Set 3 (discussed tomorrow): compressing English text with a learnt autoregressive model (using recurrent neural networks)
    - Lecture 5 (next week): stream codes with first-in-first out vs. last-in-first-out
    - Lecture 6: (net-)optimal lossless compression with latent variable models
    - ▶ Lectures 7 and 8: deep-learning based latent variable models
- Afterwards: Lossy compression

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# Stream Codes

# outperforming Huffman coding by amortizing over symbols



#### Stream Codes vs. Symbol Codes

- ▶ Reminder: Huffman coding [Huffman, 1952] creates an optimal symbol code; but:
  - Symbol codes are restrictive: each symbol contributes an *integer* number of bits.
  - Modern machine-learning based (lossy) compression methods typically use models with very low entropy *per symbol* (e.g.,  $H_P[X_i] \approx 0.3$  bits).
    - ⇒ even if most code words are only 1 bit long: about 200 % overhead
- ▶ Naive idea (Problem 2.4): apply Huffman coding to the *entire message*  $\mathbf{x} \in \mathfrak{X}^k$  rather than to individual symbols
  - ▶ Problem: exponential cost  $O(|\mathfrak{X}|^k)$
- ▶ Better idea: stream codes amortize efficiently over multiple symbols
  - ► Arithmetic Coding and Range Coding [Rissanen and Langdon, 1979; Pasco, 1976]
  - Asymmetric Numeral Systems (ANS) [Duda et al., 2015]

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## **Amortizing Compressed Bits Over Symbols**

- ► Intuitively: "pack" information content as closely as possible
- We can no longer associate each bit in the compressed representation with any specific symbol