

1 **Anomalously warm temperatures are associated with increased injury deaths**

2

3 Robbie M Parks^{1,2}, James E Bennett^{1,2}, Helen Tamura-Wicks^{1,2}, Vasilis Kontis^{1,2}, Ralf Toumi³,

4 Goodarz Danaei⁴, Majid Ezzati^{1,2,5*}

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6 ¹MRC-PHE Centre for Environment and Health, Imperial College London, London, United
7 Kingdom

8 ²Department of Epidemiology and Biostatistics, School of Public Health, Imperial College
9 London, London, United Kingdom

10 ³Space and Atmospheric Physics, Imperial College London, London, United Kingdom

11 ⁴Harvard T.H. Chan School of Public Health, Boston, Massachusetts, USA

12 ⁵WHO Collaborating Centre on NCD Surveillance and Epidemiology, Imperial College
13 London, London, United Kingdom

14

15 * Correspondence to: majid.ezzati@imperial.ac.uk

16 Temperatures which deviate from long-term local norm affect human health, and are
17 projected to become more frequent as the global climate changes. There is limited data
18 on how such deviations affect deaths from different injuries. Here, we used data on
19 mortality and temperature over a 38-year period (1980-2017) in the entire contiguous
20 USA and formulated a Bayesian spatio-temporal model to quantify how anomalous
21 temperatures, defined as deviations of monthly temperature from the local average
22 monthly temperature over the entire analysis period, affect deaths from different
23 unintentional (transport, falls and drownings) and intentional (assault and suicide)
24 injuries, by age group and sex. We found that a 2°C anomalously warm year would be
25 associated with an estimated 2,135 (95% credible interval 1,906-2,368) additional injury
26 deaths in the contiguous USA. 84% of these additional deaths would occur in males,
27 concentrated mostly in adolescent to middle ages. These deaths would comprise of
28 increases in deaths from drownings, transport, assault and suicide, offset partly by an
29 overall decline in deaths from falls in older ages. The findings demonstrate the need for
30 targeted public health interventions against injuries during periods of anomalously high
31 temperatures, especially as these episodes increase with global climate change.

32

33 Anomalously warm and cold weather events are an important public health concern in today's
34 world, and one of the key drivers for seeking adaptation measures against anthropogenic
35 climate change.¹⁻³ Current assessments of the health effects of weather and climate, and by
36 extension of global climate change, largely focus on parasitic and infectious diseases and
37 cardiorespiratory and other chronic diseases.¹⁻⁶ Less research has been conducted on injuries,⁷⁻
38⁹ especially in a consistent way across injury types and demographic subgroups of the
39 population. There are two reasons to investigate a potential role for temperature anomalies on
40 injury mortality: First death rates from injuries vary seasonally and the seasonality varies by

41 age group,^{10,11} which motivates investigating whether temperature may play a role in their
42 pathogenesis. Second, there are plausible behavioural and physiological pathways for a
43 relationship between temperature and injury – for example changes in driving patterns and
44 performance,^{12–21} alcohol drinking,²² and levels of anger^{23–25} – which motivates testing whether
45 injury deaths are affected by temperature anomalies. Our aim was to evaluate how deaths from
46 various injuries in the USA may be affected by anomalously warm temperatures that occur
47 today and are expected to become increasingly common as a result of global climate change.

48

49 We used vital registration data on all injury deaths in the contiguous USA (i.e., excluding
50 Alaska and Hawaii) from 1980 to 2017, with information on sex, age at death, underlying cause
51 of death and county and state of residence. From 1980 to 2017, 4,145,963 boys and men and
52 1,825,817 girls and women died from an injury in the contiguous USA, accounting for 9.3%
53 and 4.2% of all male and female deaths respectively. 95.7% of male injury deaths and 94% of
54 female injury deaths were in those aged 15 years and older, and over half (52.3%) of male
55 injury deaths were in those aged 15–44 years (Figure 1). In contrast with males, there was less
56 of an age gradient in females after 15 years of age.

57

58 Injuries from transport, falls, drownings, assault, and suicide accounted for 78.6% of injury
59 deaths in males and 71.8% in females. The remainder were from a heterogeneous group of
60 “other injuries” (Figure 1), within which the composition of injuries that led to death varied by
61 sex and age group. Transport was the leading injury cause of death in women younger than 75
62 years and men younger than 35 years. Between 35 and 74 years of age, more men died of
63 suicide than any other injury. Above 75 years of age, falls were the largest cause of death in
64 both men and women.

65

66 There was a decline in age-standardised death rates of three out of five major injuries (transport,
67 drownings and assault) from 1980 to 2017, although assault deaths have more recently (since
68 2014) increased (Figure 2). In contrast, age-standardised death rates from falls increased over
69 time while those from suicide initially decreased followed by an increase to surpass 1980
70 levels. The largest overall decline over time was for transport deaths in both sexes and for
71 deaths from drownings in men, which declined by over 50% from 1980 to 2017. Age-
72 standardised death rates for transport injuries and drownings peaked in summer months but
73 deaths from other major injuries did not have clear seasonal patterns.

74

75 We defined a measure of anomalous temperature for each county and month, which represents
76 the deviation from the county's average temperature in that month over the entire analysis
77 period (Supplementary Figure 1). County-level anomalies were aggregated to state level with
78 use of population weights. Average size of anomaly over the study period (1980-2017), a
79 measure of how variable temperatures are around their state-month long-term average, ranged
80 from 0.4°C for Florida in September, to 3.4°C for North Dakota in February (Figure 3). The
81 average size of anomaly had a median value of 1.2°C across all states and months (Figure 3).
82 Temperature anomalies were largest in January and December and smallest in August and
83 September. They were larger in northern and central states than in southern and coastal ones.

84

85 We analysed the association of monthly injury death rates with anomalous temperature using
86 a Bayesian spatio-temporal model, described in detail in Methods. We used the resultant risk
87 estimates and the age-sex-specific death rates from each injury in 2017, to estimate additional
88 deaths if each month in each state were +2°C above its long-term average, consistent with the
89 upper bound of the Paris Climate Agreement. Based on this analysis, we estimated that there
90 would be an estimated 2,135 (95% credible interval 1,906-2,368) excess injury deaths,

91 equivalent to 1.0% of all injury deaths in 2017, in each year in which each month in each state
92 was +2°C warmer than its long-term average (Figure 4). Deaths from drowning, transport,
93 assault and suicide would increase, partly offset by a decline in deaths from falls in middle and
94 older ages and in winter months (Figure 4). Most excess deaths would be from transport injuries
95 (985; 866-1,085) followed closely by suicide (720; 593-841). 84% of the excess deaths would
96 occur in males and 16% in females. 92% of all male excess deaths would occur in those aged
97 15-64 years, who have higher rates of deaths from transport and suicide. In those aged 85 years
98 and older, there would be an estimated decline in injury deaths, because deaths from falls are
99 expected to decline in a warmer year.

100

101 Proportionally, deaths from drownings are estimated to increase more than those of other injury
102 types, by as much 18.3% (16.6, 20.2) for a +2°C anomaly in men aged 15-24 years (Figure 5).
103 The smallest proportional increase was that of assault and suicide (less than 4% in all age and
104 sex groups). There was a larger percent increase in transport deaths for males than for females,
105 especially in young and middle-ages (e.g., 2.7% (2.1, 3.4) for 25-34 year old men versus 0.7%
106 (-0.4, 1.9) for women of the same age) (Figure 5).

107

108 That anomalously warm temperature influences deaths from drowning, although not previously
109 quantified, is highly plausible because swimming is likely to be more common when
110 temperature is higher. The higher relative and absolute impacts on men compared with women
111 may reflect differences in their behaviours. For example, over half of swimming deaths for
112 males occur in natural water, compared to about one quarter for females.²⁶ The former may
113 rise more in warmer weather. Similarly, deaths from falls declined more in older ages because
114 falls in the elderly are more likely to be due to slipping on ice than those in younger people.²⁷⁻

115 ²⁹

116

117 The pathways from anomalous temperature to transport injury are more varied. Firstly, driving
118 performance deteriorates at higher temperatures.^{16–19} Further, alcohol consumption increases
119 in warm temperatures,²² which also provides an explanation for why teenagers, who are more
120 likely than other age groups to crash while intoxicated,³⁰ experience a larger proportional rise
121 in deaths from transport when temperatures are anomalously warm than older adults. Lastly,
122 warmer temperatures generally increase road traffic in North America;^{12–15,20,21} coupled with
123 more people outdoors in warmer weather,³¹ this increase could lead to more fatal collisions.

124

125 Pathways linking anomalously high temperatures and deaths from assault and suicide are less
126 established. One hypothesis is that, more time spent outdoors in anomalously warmer
127 temperatures leads to an increased number of face-to-face interactions, and hence arguments,
128 confrontations, and ultimately assaults.^{32,33} These effects could be compounded by the greater
129 anger levels linked to higher temperatures.^{23–25} However, further research on the association of
130 temperature and assault, and the factors mediating it, is needed.³⁴ Regarding suicide, it has been
131 hypothesised that higher temperature is associated with higher levels of distress in younger
132 people.³⁵ Nonetheless, the mechanisms for the links between temperature and mental health
133 requires further investigation, including whether the relationship varies by age and sex, as
134 indicated by our results. Future research should also investigate the extent to which the
135 increased risk of injury death as a result of anomalous temperature depends on community
136 characteristics such as poverty and deprivation, quality of roads and housing, emergency
137 response, and social services.

138

139 Our work highlights how deaths from injuries are currently susceptible to temperature
140 anomalies and could also be modified by rising temperatures resulting from climate change,

141 unless countered by social and health system interventions that mitigate these impacts. Though
142 absolute impacts on mortality are modest, some groups, especially men in young to middle-
143 ages, will experience larger impacts. Therefore, a combination of public health interventions
144 that broadly target injuries in these groups – for example targeted messaging for younger males
145 on the risks of transport injury and drowning – and those that trigger in relation to forecasted
146 high temperature periods – for example more targeted blood alcohol level checks – should be
147 a public health priority.

148

149 **Methods**

150 *Data sources*

151 We used data on deaths by sex, age, underlying cause of death and state of residence in the
152 contiguous USA from 1980 to 2017 through the National Center for Health Statistics (NCHS)
153 (https://www.cdc.gov/nchs/nvss/dvs_data_release.htm) and on population from the NCHS
154 bridged-race dataset for 1990 to 2017 (https://www.cdc.gov/nchs/bridged_race.htm) and
155 from the US Census Bureau prior to 1990 ([https://www.census.gov/data/tables/time-
156 series/demo/popest/1980s-county.html](https://www.census.gov/data/tables/time-series/demo/popest/1980s-county.html)). We did not include Alaska and Hawaii, (which
157 together made up 0.5% of the US population in 2017) because their climates and environment
158 are distinct from other states due to their substantial physical distance. We calculated monthly
159 population counts through linear interpolation, assigning each yearly count to July.

160

161 The underlying cause of death was coded according to the international classification of
162 diseases (ICD) system (9th revision from 1980 to 1998 and 10th revision thereafter). The 6
163 million injury deaths fell into six categories: transport, falls, drownings, assault, suicide and
164 an aggregate set of other injuries. We report the results of all of these categories except other

165 injuries (1,402,941 deaths or 23% of total injury deaths during 1980-2017), because the
166 composition of this aggregate group varies by sex, age group, state and time.

167

168 We obtained data on temperature from ERA5, which uses measurement data from global in-
169 situ and satellite measurements to generate a worldwide meteorological dataset, with full space
170 and time coverage over our analysis period.³⁶ We used gridded four-times-daily estimates at a
171 resolution of 30 km to generate monthly temperatures by county and following population-
172 weighting, monthly temperatures by state throughout the analysis period.

173

174 *Anomalous temperature metric*

175 With few exceptions,^{7,37} current climate change risk assessments extrapolate from associations
176 of daily mortality with daily temperature.^{5,6,38-40} Climate change, however, will fundamentally
177 modify weather, including seasonal weather patterns, compared to long-term averages, and
178 hence can disrupt existing forms of adaptation. To mimic the conditions that may arise with
179 global climate change, we developed methodology to examine how deviations from long-term
180 average temperature may impact injury death rates.

181

182 We first defined a measure of anomalous temperature for each county and month, which
183 represents the deviation from the average temperature of the county in that month over the
184 entire analysis period (Supplementary Figure 1). To calculate the magnitude of temperature
185 anomaly, we first calculated average temperatures for each month in each county over the entire
186 38 years of analysis. We subtracted these long-term average temperatures from respective
187 monthly temperature values to generate a temperature anomaly time series for each month and
188 year in each county (Supplementary Figure 1). The temperature anomaly metric measures the
189 extent that temperature experienced in a specific month, year and county is warmer or cooler

190 than the long-term average to which the population has acclimatised. These values can be
191 different for different months in the same county, and different counties in the same month.
192 Further, a county with higher, but more stable, temperature in a specific month has smaller
193 anomalies than one with lower but more inter-annually variable temperature. County-level
194 anomalies were aggregated to state level with use of population weights for analysing their
195 associations with mortality.

196

197 *Statistical methods*

198 We analysed the association of monthly injury death rates with anomalous temperature using
199 a Bayesian spatio-temporal model, which leveraged variations over space and time to infer
200 associations. We modelled the number of deaths in each month in each year as following a
201 Poisson distribution:

202
$$deaths_{state-time} \sim Poisson(death\ rate_{state-time} \cdot population_{state-time})$$

203 with log-transformed death rates modelled as a sum of components that depend on location
204 (state) of death, month of year, overall time (in months) and temperature anomaly:

205
$$\log(deaths_{state-time}) =$$

206
$$\alpha_0 + \beta_0 \cdot time +$$

207
$$\alpha_{state} + \beta_{state} \cdot time +$$

208
$$\alpha_{month} + \beta_{month} \cdot time +$$

209
$$\zeta_{state-month} +$$

210
$$\psi_{state-month} \cdot time +$$

211
$$\nu_{time} +$$

212
$$\gamma_{month} \cdot Anomaly_{state-time} +$$

213
$$\varepsilon_{state-time}$$

214
215 The model contained terms that represent the national level and trend in mortality, with α_0 as
216 the common intercept and β_0 the common slope with overall time. Death rates also vary by
217 month, which may be partly related to temperature and partly due to other monthly factors;
218 monthly variations tend to be smooth across adjacent months.¹⁰ Therefore, we allowed each
219 month of the year to systematically have a different mortality level and trend, with α_{month} the

220 month-specific intercept and β_{month} the month-specific slope with overall time. We used a
221 first-order random walk prior for the monthly random intercepts and slopes, widely used to
222 characterise smoothly varying trends.⁴¹ The random walk had a cyclic structure, so that
223 December was adjacent to January.

224

225 We also included state random intercepts and slopes for death rates, with α_{state} as the state-
226 specific intercept and β_{state} the state-specific slope with overall time. These terms measure
227 deviations of each state from national values, and allow variation in level and trend in mortality
228 by state. We modelled the state-level random intercepts and slopes using the Besag, York, and
229 Mollie (BYM) spatial model,⁴² which includes both spatially-structured random effects with
230 an intrinsic Conditional Autoregressive (ICAR) prior and spatially-unstructured independent
231 and identically distributed (IID) Gaussian random effects. The extent to which information is
232 shared between neighbouring states depends on the uncertainty of death rates in a state and the
233 empirical similarity of death rates in neighbouring states. We also included state-month
234 interactions for intercepts and slopes ($\zeta_{state-month}$ and $\psi_{state-month}$), to allow variation in
235 mortality levels and trends in a particular state for different months and vice-versa. These state-
236 month interactions were modelled as IID and therefore were of Type I space-time
237 interactions.⁴³ Non-linear change over overall time (in months) was captured by a first-order
238 random walk, v_{time} .⁴¹ In order to ensure identifiability each set of random walk terms or state
239 random effects was constrained to sum to zero.

240

241 Finally, we included a term that relates log-transformed death rate to the above-defined state-
242 month temperature anomaly, $\gamma_{month} \cdot Anomaly_{state-time}$. The coefficients of γ_{month} represent
243 the logarithm of the monthly death rate ratio per 1°C increase in anomaly. There was a separate
244 coefficient for each month which means that an anomaly of the same magnitude could have

245 different associations with injury mortality in different months. As with the month-specific
246 intercepts and trends, we used a cyclic first-order random walk to smooth the coefficient of the
247 temperature anomaly across months. An over-dispersion term ($\varepsilon_{state-time}$) captured the
248 variation unaccounted for by other terms in the model, modelled as $N(0, \sigma_\varepsilon^2)$. We used weakly
249 informative priors so that parameter estimation was driven by the data. As in previous
250 analyses,^{44,45} hyper-priors were defined on the logarithm of the precisions of the random
251 effects, in other words on $\log(1/\sigma^2)$. These were modelled as $\text{logGamma}(\theta, \delta)$ distributions
252 with shape $\theta = 1$ and rate $\delta = 0.001$. The same hyper-priors were used for all precision
253 parameters of the random effects in the model. For the common slope, we used $N(0, 1000)$ and
254 for the common intercept a flat prior.

255

256 In addition to representing the spatial (across states) and temporal (across months and years)
257 patterns of mortality, the intercept terms ($\alpha_{month}, \alpha_{state}, \zeta_{state-month}$) in our statistical model
258 implicitly adjust for unobserved factors that influence mortality at the state, month and state-
259 month level; the slope terms ($\beta_{month}, \beta_{state}, \psi_{state-month}$) do so for changes in these factors
260 over time.⁴⁴ This means that the only confounding factors would be those that have the same
261 state-month anomaly as temperature.

262

263 We fitted the models using integrated nested Laplace approximation (INLA), using the R-
264 INLA software, which offers orders of computational efficiency improvement in Bayesian
265 inference compared to traditional MCMC.⁴⁶ The uncertainty in our results were obtained from
266 5000 draws from the posterior marginal of each month's excess relative risk. The reported 95%
267 credible intervals are the 2.5th to 97.5th percentiles of the sampled values.

268

269 Analyses were done separately by injury type, because different injuries can have differing
270 associations with anomalously warm and cold temperature. Analyses were also done separately
271 by sex and age group (0-4 years, 10-year age groups from 5 to 84 years, and 85+ years) because
272 injury death rates vary by age group and sex, as might their associations with temperature. We
273 used the resultant risk estimates and the age-sex-specific death rates from each injury in 2017,
274 to calculate additional deaths if each month in each state were +2°C above its long-term
275 average, not only realistic in our lifetimes under current projections of global climate change
276 but an agreed upper bound chosen under the Paris Climate Agreement.⁴⁷ +2°C is also within
277 the range of anomaly size experienced by some states (Figure 3). For this calculation, we
278 multiplied the actual death counts for each month, sex, state and age group in 2017 by the
279 corresponding excess relative risk, which was calculated as the exponential of the coefficient
280 of the temperature anomaly term from the above analysis.

281

282 *Sensitivity analyses*

283 We conducted sensitivity analyses to assess how much our results might depend on the
284 temperature metric used to generate anomalous temperature. First, instead of building our
285 monthly temperature anomalies based on daily mean temperatures, we used daily maxima and
286 minima. These measures were strongly correlated to those generated from daily means
287 (Supplementary Table 3), and therefore we did not run models using these alternatives.

288

289 Second, together with temperature anomaly based on daily mean temperatures, we also
290 included a second measure of anomaly in the model. The additional measures were related to
291 more extreme anomalous situations which may be relevant if the impacts on injuries are related
292 to more extreme temperatures and how frequent they are in each month:

293

- 294 • temperature anomaly calculated based on 90th percentile (°C) of daily mean temperatures
295 within a month, compared to the average of 90th percentiles for each state and month
296 • number of days in a month above the long-term 90th percentile of average temperature for
297 each state and month (adjusted for length of month)
298 • number of 3+ day episodes above the long-term 90th percentile of average temperature for
299 each state and month (adjusted for length of month)

300

301 The correlations among these variables and anomaly based on mean were between 0.60 and
302 0.89 (Supplementary Table 4). The estimated rate ratios of temperature anomaly based on daily
303 means (i.e., the anomaly measure used in the main analysis) were robust to the addition of
304 alternative measures of anomaly, while the coefficients of the additional measures were
305 generally not significant and with large credible intervals. Therefore, we did not include the
306 alternative additional measures of extreme anomalous temperature in the main analysis.

307

308 *Comparison with prior studies*

309 While there are no previous studies of how deviations of monthly temperature from long-term
310 average are associated with injury mortality, our results are broadly in agreement with those
311 that have analysed associations with absolute temperature and for specific injury types. A study
312 of suicide in US counties over 37 years (1968-2004) estimated that 1°C higher monthly
313 temperature would lead to a 0.7% rise in suicides,⁷ compared to our findings of 0.9-2% in males
314 and 0.6-3.9% in females in different ages for a +2°C anomaly. A cross-sectional analysis in
315 100 US counties found that a 1°C higher temperature would lead to a 1.3% increase in death
316 rates from road traffic injuries,²¹ compared to our finding of 0.8-4.1% in males and 0.6-2.7%
317 in females for a +2°C anomaly. In a study of six French heatwaves during 1971-2003, mortality
318 from unintentional injuries rose by up to 4% during a heatwave period compared to a non-

319 heatwave baseline.⁸ A study of daily mortality from all injuries from Estonia found a 1.24%
320 increase in mortality when daily maximum temperature went from the 75th to 99th percentile of
321 long-term distribution.⁹

322

323 *Strengths and limitations*

324 The major strength of our study is that we have comprehensively modelled the association of
325 temperature anomaly with injury by type of injury, month, age group and sex. Our measure of
326 temperature anomaly internalises long-term historical experience of each state, and is closer to
327 what climate change may bring about than solely examining daily episodes, or average
328 temperature to which people have adapted. To utilise this metric, we integrated two large
329 disparate national datasets on mortality (US vital statistics) and meteorology (ERA5), and
330 developed a bespoke Bayesian spatio-temporal model. A limitation of our study is that, like all
331 observation studies, we cannot rule out confounding of results due to other factors. As
332 described above, our statistical model by design adjusts for factors related to month, state and
333 state-month that are either invariant over time or that change linearly. Rather, the confounding
334 factors would be those with anomalies that are similar to those of monthly temperature in each
335 state, such as air pollution. However, to our knowledge, there is currently no evidence of an
336 association between air pollution and injury mortality.

337

338 **Data availability**

339 ERA5 temperature data are downloadable from
340 <https://www.ecmwf.int/en/forecasts/datasets/reanalysis-datasets/era5>. Vital statistics files with
341 geographical information can be requested through submission of a proposal to NCHS
342 (<https://www.cdc.gov/nchs/nvss/nvss-restricted-data.htm>).

343

344 **Code availability**

345 The computer code for the Bayesian model used in this work will be available at

346 www.globalenvhealth.org/code-data-download upon publication of the paper.

347

348 **Supplementary information**

349 This file contains Supplementary Figure 1, Supplementary Table 1, Supplementary Table 2,

350 Supplementary Table 3 and Supplementary Table 4.

351

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362 studies.

363

364 **Author contributions**

365 All authors contributed to study concept and interpretation of results. RP, GD and ME collated

366 and organised temperature and mortality files. RP, JEB, VK and ME developed statistical

367 model, which was implemented by RP, JEB and VK. RP performed the analysis, with input

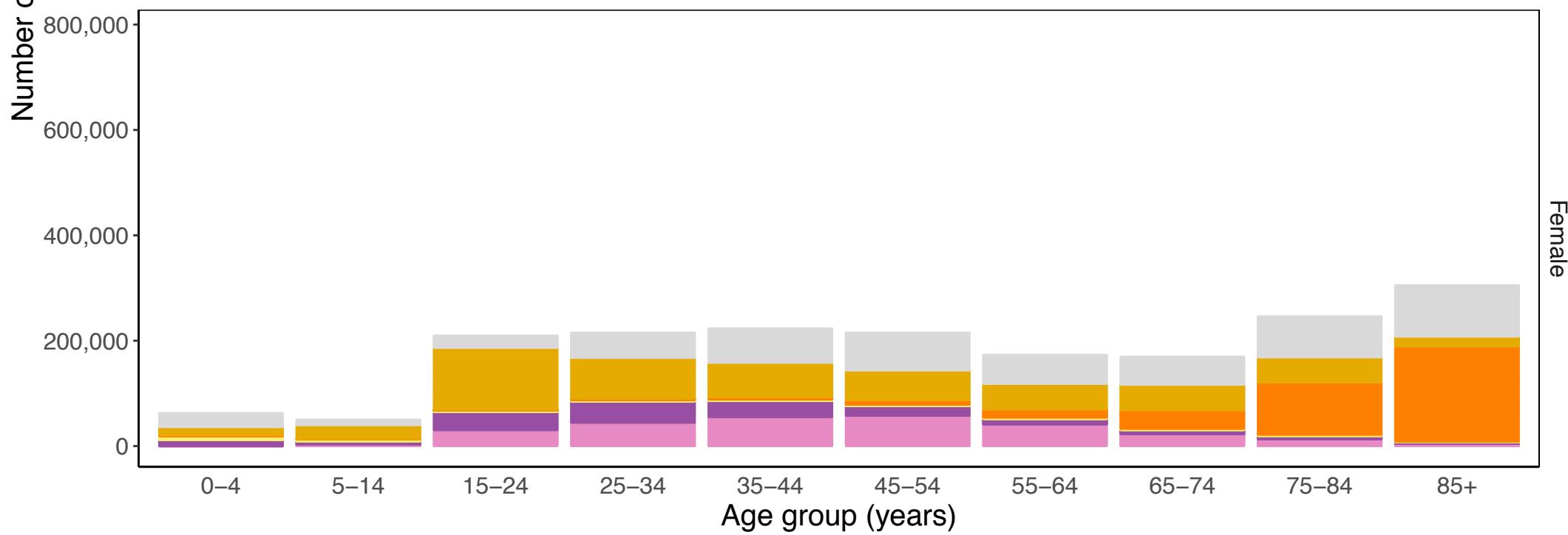
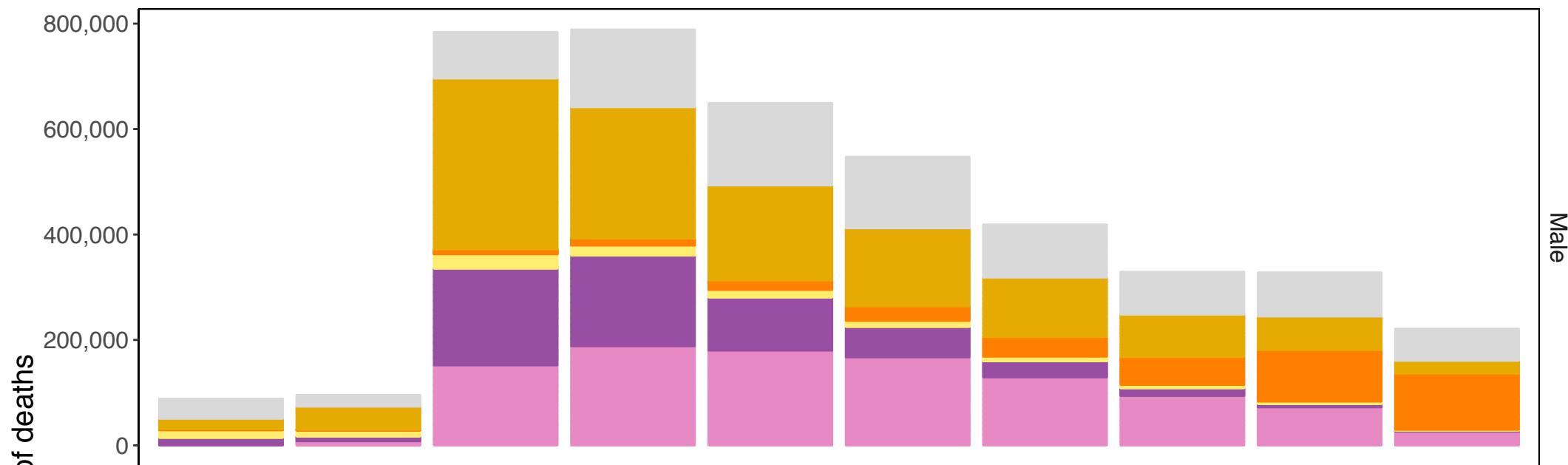
368 from other authors. RP and ME wrote the first draft of the paper; other authors contributed to
369 revising and finalizing the paper.

370

371 **Competing interests statement**

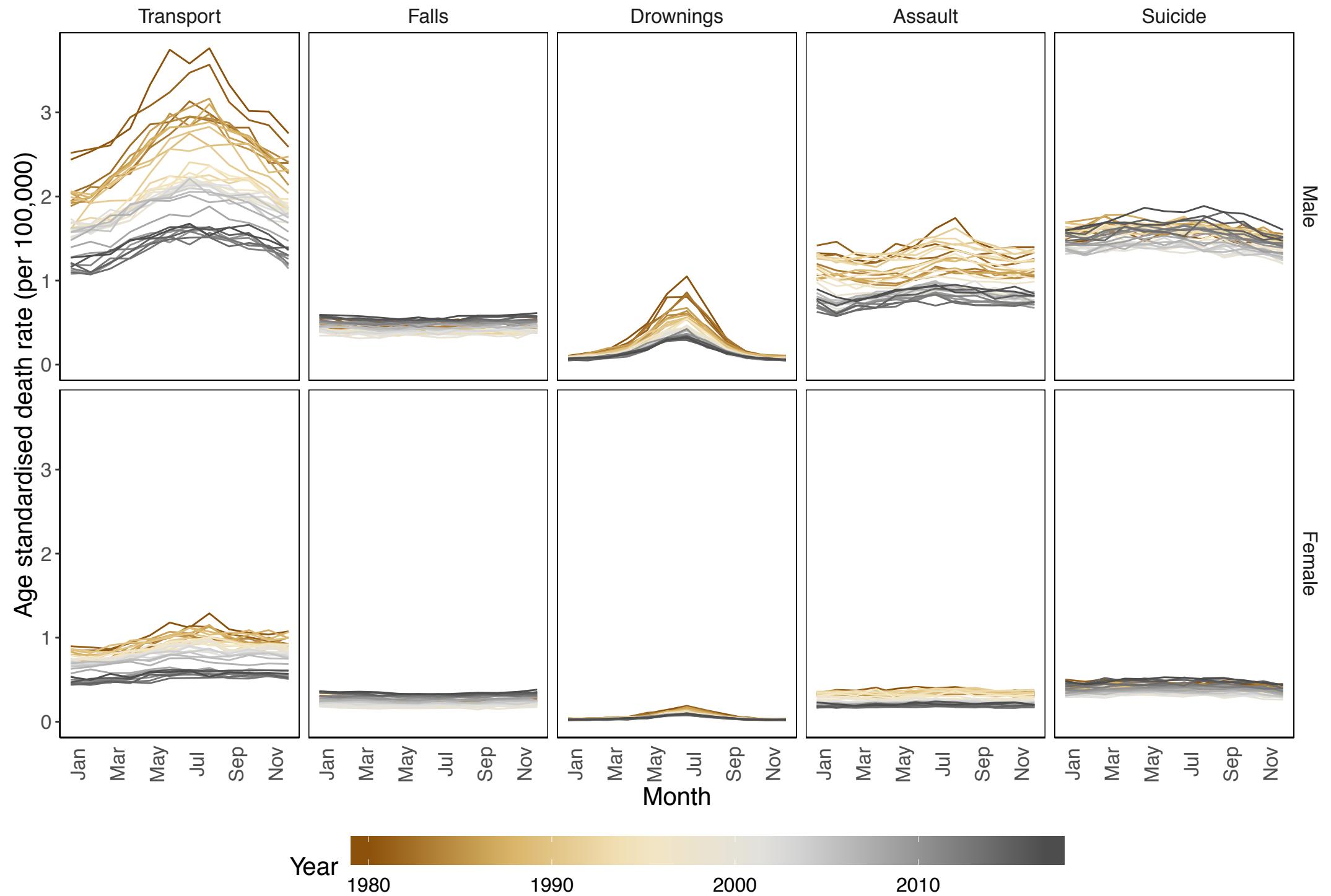
372 ME reports a charitable grant from AstraZeneca Young Health Programme, and personal fees
373 from Prudential, Scor, and Third Bridge, all outside the submitted work; all other authors
374 declare no competing interests.

375 **Figure 1.** Number of injury deaths, by type of unintentional (transport, falls, drownings, and
376 other) and intentional (assault and suicide) injury, by sex and age group in the contiguous USA
377 for 1980-2017.

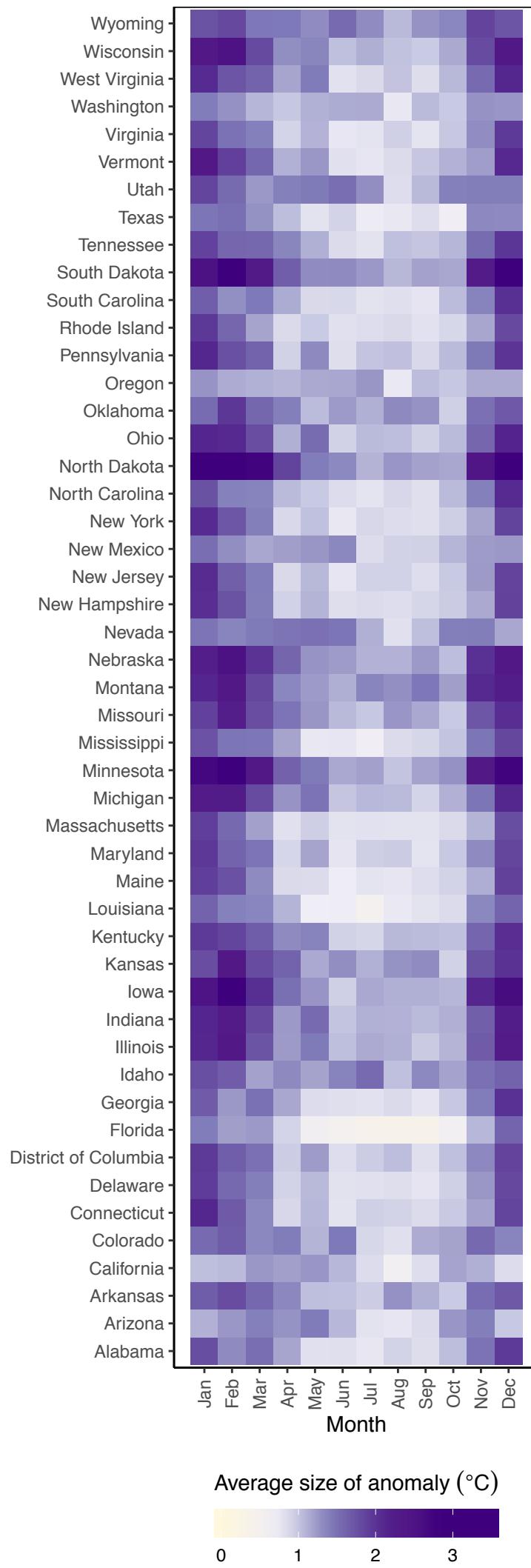


█ Transport
 █ Falls
 █ Drownings
 █ Assault
 █ Suicide
 █ Other injuries

378 **Figure 2.** National age-standardised death rates from 1980 to 2017, by type of injury, sex and
379 month.



380 **Figure 3.** Average size of temperature anomaly ($^{\circ}\text{C}$) from 1980 to 2017, by state and month.
381 The value for each state and month is the mean of the absolute size of anomaly, be it cold or
382 warm, and hence gives an indication of the scale of anomalies around the local average
383 temperatures.

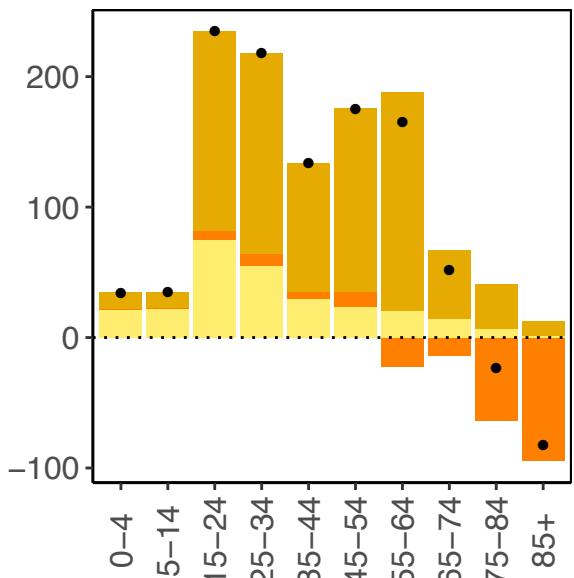


384 **Figure 4.** Additional annual injury deaths for the 2017 US population in year in which each
385 month was +2°C warmer compared with 1980-2017 average temperatures. The top row shows
386 breakdown by type of injury, sex and age group. The bottom row shows the break down by
387 type of injury, sex and month. Black dots represent net changes in deaths for each set of bars.

Additional deaths associated with a 2°C warmer year (based on 2017 population)

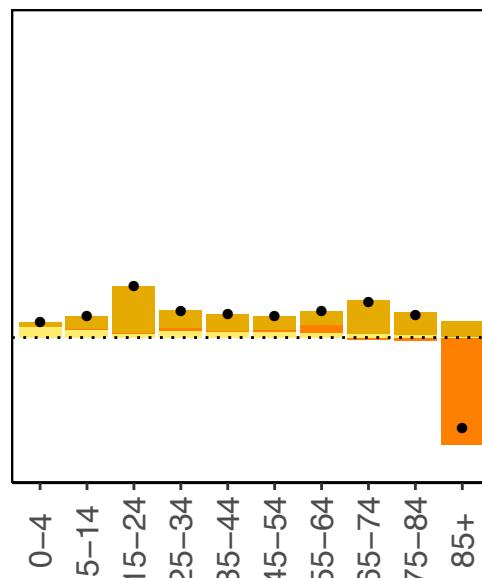
Unintentional

Male



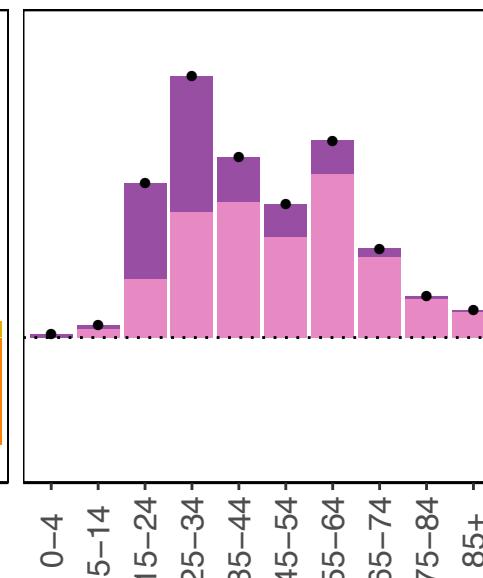
Unintentional

Female



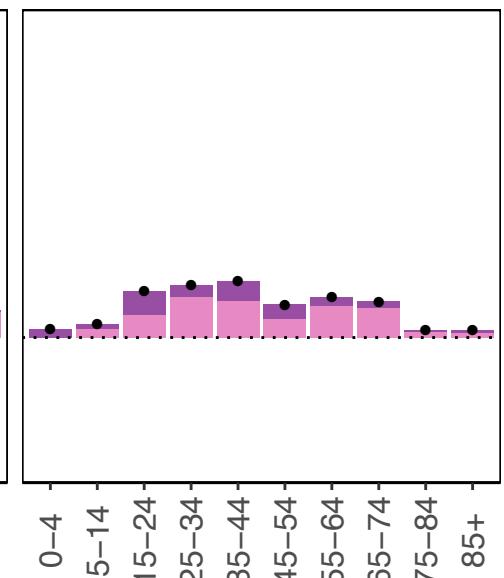
Intentional

Male



Intentional

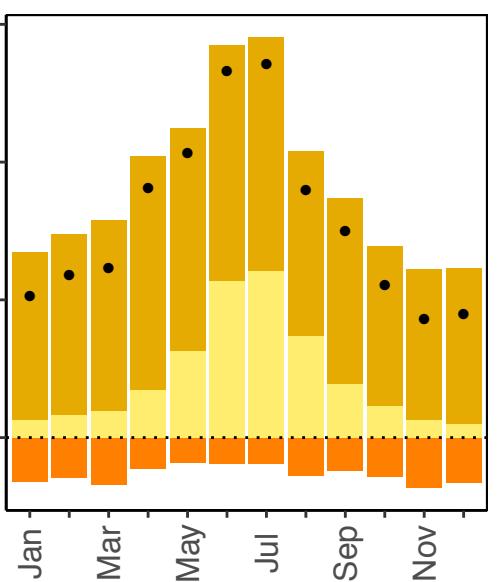
Female



Additional deaths associated with a 2°C warmer year (based on 2017 population)

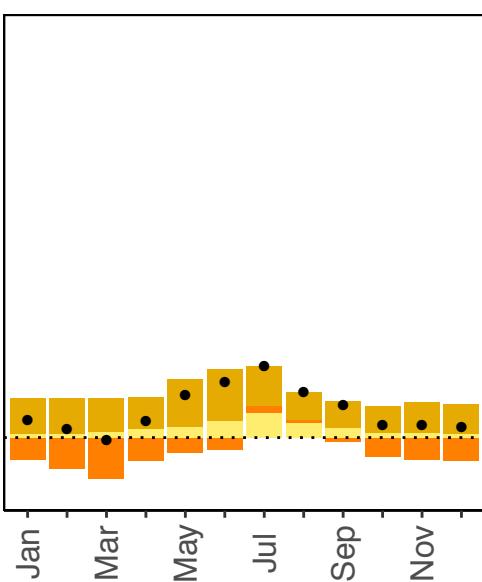
Unintentional

Male



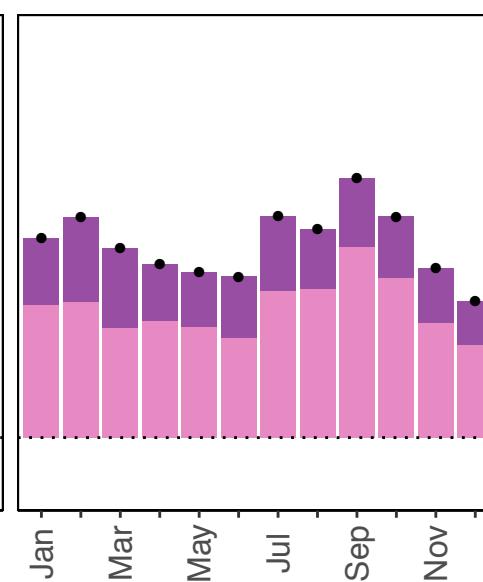
Unintentional

Female



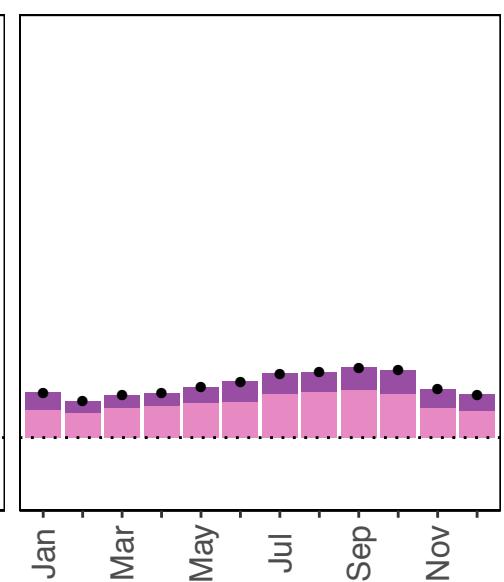
Intentional

Male



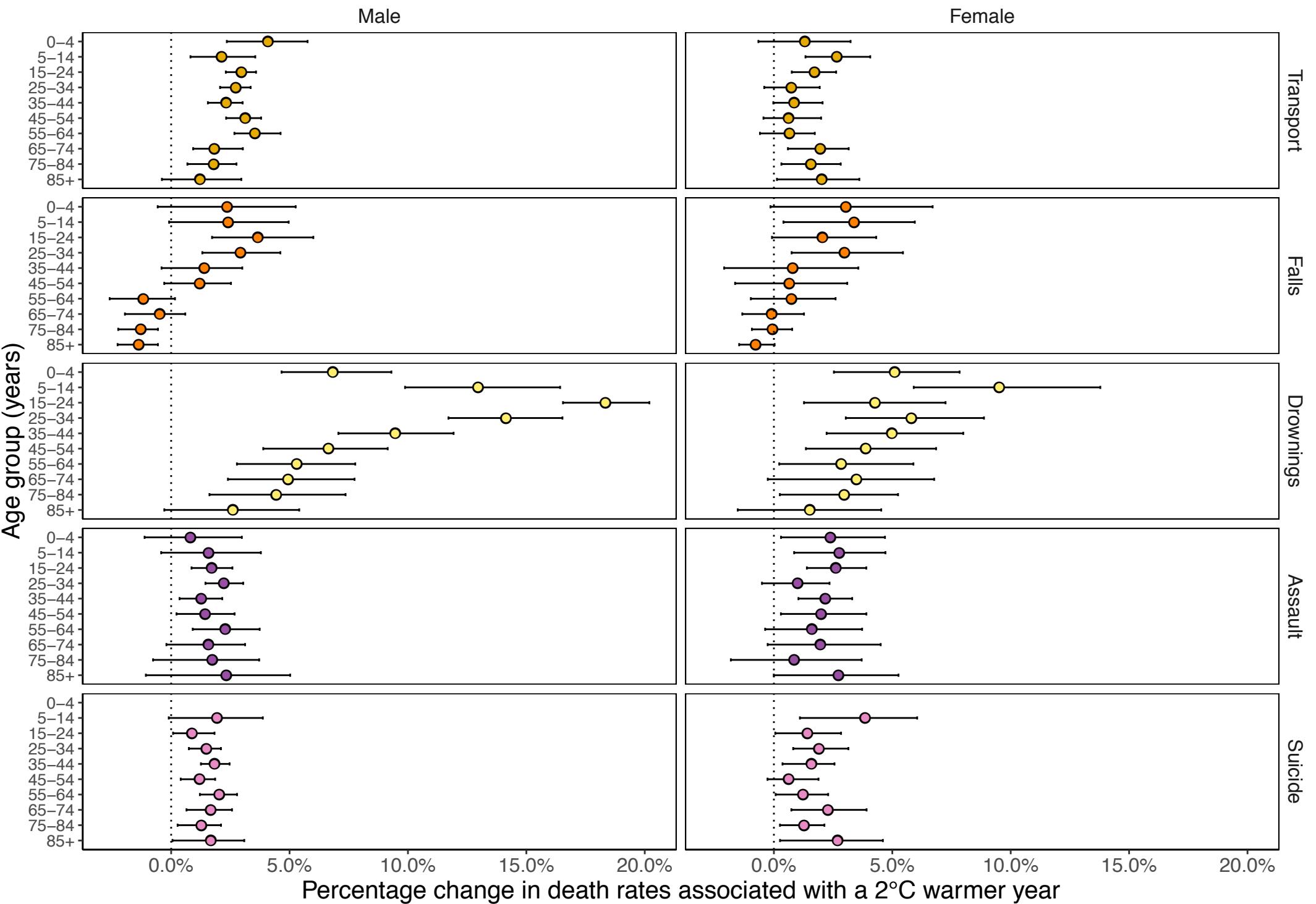
Intentional

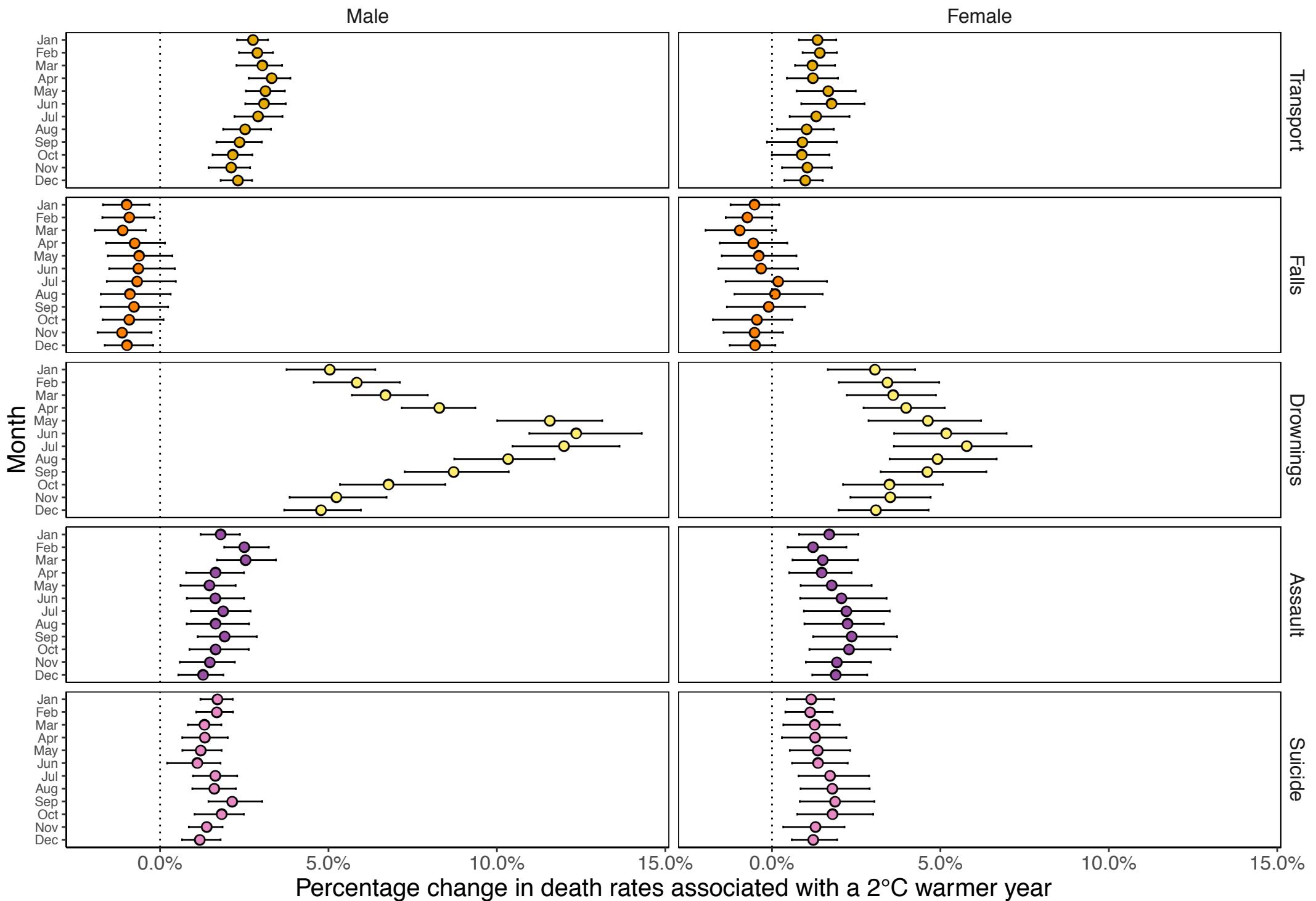
Female



Legend: Transport (Yellow), Falls (Orange), Drownings (Light Yellow), Assault (Purple), Suicide (Pink)

388 **Figure 5.** Percent change in death rates in year in which each month was +2°C compared with
389 1980-2017 average temperatures by type of injury, sex and (A) age group or (B) month.





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