1.

1.1

a.

X and Y are independent, so $E[XY^T] = E(X)E(Y^T)$

$$cov(X,Y) = E\left[\left(X - E(X)\right)\left(Y - E(Y)\right)^T\right] = E[XY^T] - E[X]E[Y^T] = E[X]E[Y^T] - E[X]E[Y^T] = 0$$

b.

For each value x_i, y_i in X and Y

$$E(x_i + y_i) = E(x_i) + E(y_i)$$
 for all integer $i \in [1, m]$

We can get E(X + Y) = E(X) + E(Y)

For a mxm matrix A, j-th number for AX is $\sum_{i=1}^m A_{ii} x_i$

$$E\left(\sum_{i=1}^{m} A_{ii} x_i\right) = \sum_{i=1}^{m} A_{ii} E(x_i)$$

We can get E(AX) = AE(X)

Let
$$W = AX(AX)^T$$

$$W_{ij} = \sum_{k=1}^{m} \sum_{l=1}^{m} A_{il} x_{l} x_{k} A_{kj}$$

$$E(W) = AE(x_l x_k) A^T$$

We can conclude that E(X + AY) = E(x) + AE(Y)

$$Var(X + AY) = Var(X) + Var(AY) + 2Cov(X, AY) = Var(X) + E(AYY^{T}A^{T}) - E(AY)E(AY)^{T}$$
$$= Var(X) + AE(YY^{T})A^{T} - AE(Y)E(Y)^{T}A^{T} = Var(X) + AVar(Y)A^{T}$$

We can conclude that $Var(X + AY) = Var(X) + AVar(Y)A^{T}$

c.

AX is normally distributed since X is normally distributed

Then we have E(AX) = AE(X)

$$Var(AX) = AVar(X)A$$

$$X \sim N(\mu, \Sigma)$$
 implies $AX \sim N(A\mu, A\Sigma A^{T})$

1.2

a.

Yes, the uniform distribution on the interval [0,0.5] has probability density f(x)=2 for $0 \le x \le 0.5$ and f(x)=0 elsewhere.

h

$$f(x) = \frac{1}{\sqrt{\frac{\pi}{50}}} e^{-50x^2}$$

C.

At 0 , the value of this pdf is $\frac{10}{\sqrt{2\pi}}$

d.

$$P(x = 0) = \int_0^0 f(x) dx = 0$$

1.3
a.
$$\frac{\partial}{\partial x} = x^T y = y^T$$
b.
$$\frac{\partial}{\partial x} = 2x^T$$
c.
$$\frac{\partial}{\partial x} = x^T (A^T + A)$$
d.
$$\frac{\partial}{\partial x} = A$$

2.

a.

$$\hat{\beta} = (X^T X)^{-1} X^T Y$$

$$Y|X, \beta \sim N(X\beta, \sigma^2 I)$$

 $\hat{\beta}$ is normally distributed

$$E(\hat{\beta}) = E((X^T X)^{-1} X^T Y) = (X^T X)^{-1} X^T E(Y) = (X^T X)^{-1} X^T X \beta = \beta$$

$$Var(\hat{\beta}) = Var((X^T X)^{-1} X^T Y) = (X^T X)^{-1} X^T Var(Y) ((X^T X)^{-1} X^T)^T = \sigma^2 (X^T X)^{-1}$$

b.

$$\begin{split} \mathsf{L}(\mathsf{Y}|\mathsf{X},\beta) &= \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi\sigma^{2}}} e^{-\frac{(y_{i}-x_{i}\beta)^{2}}{2\sigma^{2}}} = (2\pi\sigma^{2})^{-\frac{n}{2}} e^{\sum_{i=1}^{n} -\frac{(y_{i}-x_{i}\beta)^{2}}{2\sigma^{2}}} \\ \mathsf{l}(\mathsf{Y}|\mathsf{X},\beta) &= \mathsf{log}(\mathsf{L}) = -\frac{n}{2} \mathsf{log}(2\pi\sigma^{2}) - \sum_{i=1}^{n} \frac{(y_{i}-x_{i}\beta)^{2}}{2\sigma^{2}} \\ \frac{\partial}{\partial \beta} \mathsf{l}(\mathsf{Y}|\mathsf{X},\beta) &= \frac{\sum_{i=1}^{n} x_{i}^{T} y_{i} - x_{i}^{T} x_{i}\beta}{\sigma^{2}} = \frac{\mathsf{X}^{T} \mathsf{Y} - \mathsf{X}^{T} \mathsf{X}\beta}{\sigma^{2}} \end{split}$$

$$P(|\widehat{\beta}_i - \beta_i| \le \varepsilon) = P(\beta_i - \varepsilon \le \widehat{\beta}_i \le \varepsilon + \beta_i) = F(\varepsilon + \beta_i) - F(\beta_i - \varepsilon) \text{ where F(t) is the cdf of } \widehat{\beta}$$

2.2

a.

$$P(\beta|Y) = \frac{P(Y|\beta)P(\beta)}{P(Y)}$$
 based on bayes rule

$$P(Y|\beta) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y_i-x_i\beta)^2}{2\sigma^2}}$$
 where x_I is the i-th row of x

$$P(\beta) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(\beta)^2}{2\tau^2}}$$

Since $P(\beta|Y) \propto P(Y|\beta)P(\beta)$

We have $argmax(P(\beta|Y)) = argmax(P(Y|\beta)P(\beta))$

By expanding $P(Y|\beta)P(\beta)$ we have

$$(2\pi\sigma^2)^{-\frac{n}{2}} * (2\pi\sigma^2)^{-\frac{1}{2}} e^{-\sum_{i=1}^{n} \frac{(y_i - x_i\beta)^2}{2\sigma^2} - \frac{(\beta)^2}{2\tau^2}}$$

$$\operatorname{argmax} \left(P(Y|\beta) P(\beta) \right) = \operatorname{argmin} \left(\sum_{i=1}^{n} \frac{(y_i - x_i \beta)^2}{2\sigma^2} + \frac{(\beta)^2}{2\tau^2} \right)$$

To get result of argmin, we set derivative to 0

$$\frac{\partial}{\partial \beta} = \sum_{i=1}^{n} -x_i^T \frac{y_i - x_i \beta}{\sigma^2} + \frac{\beta}{\tau^2} = 0$$

$$\sum_{i=1}^{n} -x_i^T \frac{y_i - x_i \beta}{\sigma^2} = -\frac{\beta}{\tau^2}$$

$$\sum_{i=1}^{n} x_i^T (y_i - x_i \beta) = \frac{\sigma^2 \beta}{\tau^2}$$

$$\sum_{i=1}^{n} x_i^T (y_i - x_i \beta) = \lambda \beta$$

$$X^T Y - X^T X \beta = \lambda \beta$$

$$(X^T X + \lambda I) \beta = X^T Y$$

$$\beta = (X^T X + \lambda I)^{-1} X^T Y$$

b.

We have
$$X^* = \begin{bmatrix} & X \\ \sqrt{\lambda} & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \sqrt{\lambda} \end{bmatrix} \quad Y^* = \begin{bmatrix} & Y \\ 0 & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & 0 \end{bmatrix}$$

$$X^{*T}X^* = X^TX + \lambda I$$

$$X^{*T}Y^* = X^TY$$

MLE of β

$$P(Y|\beta) = \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y_i - x_i \beta)^2}{2\sigma^2}}$$

Same as part a, we can try to minimize $\sum_{i=1}^{n}(y_i-x_i\beta)^2$

$$\frac{\partial}{\partial \beta} = \sum_{i=1}^{n} -2x_i^T (y_i - x_i \beta) = 0$$

$$-X^TY + X^TX\beta = 0$$

$$\beta = (X^T X)^{-1} X^T Y$$

Substitute *X*, *Y* with X^* , Y^* , we have $\beta = (X^*TX^*)^{-1}X^*TY^* = (X^TX + \lambda I)^{-1}X^TY$

which is same as $\hat{\beta}$ in ridge regression.

2.3

Code:

import scipy.io as sio import random import numpy as np import matplotlib.pyplot as plt

part a load in mat data
dataset = sio.loadmat('./dataset.mat')
data_train_X = dataset['data_train_X']
data_train_y = dataset['data_train_y'][0]
data_test_X = dataset['data_test_X']
data_test_y = dataset['data_test_y'][0]

use random.shuffle to get random permutation of data def shuffle_data(data):

temp = data[:]
random.shuffle(temp)
return temp

def split data(data, num folds, fold):

```
# get each split's size
  size = int(len(data) / num_folds)
  total = []
  index = 0
  for i in range(num_folds):
    total.append(data[index:(index+size)])
    index += size
  # get specified data fold based on index fold
  data_fold = total[fold-1]
  total.pop(fold-1)
  data_rest = []
  for i in total:
    data rest.extend(i)
  return data_fold, data_rest
def train model(data, lambd):
  # create correct size matrices for X and Y
  y, x = np.empty([1, len(data)]), np.empty([len(data), 400])
  for i in range(len(data)):
    y[0][i] = data[i][0]
    x[i] = data[i][1]
  I = np.identity(400)
  beta = np.dot(np.dot(np.linalg.inv((np.dot(x.transpose(), x) + np.dot(lambd, I))), x.transpose()), y.transpose())
  return beta
def predict_model(data, model):
  res = []
  for d in data:
    res.append(np.dot(d[1].reshape([1, 400]), model.reshape([400, 1])))
  return res
# get summation of all differences then divided by number of value
def loss(data, model):
  prediction = predict_model(data, model)
  real, diff = [], []
  for i in data:
    real.append(i[0])
  for i in range(len(real)):
    diff.append((real[i] - prediction[i]) ** 2)
  return sum(diff)/len(diff)
def cross_validation(data, num_folds, lambd_seq):
  data = shuffle_data(data)
  cv_error = []
  for i in range(50):
    lambd = lambd seq[i]
    cv_loss_lmd = 0
    for fold in range(1, num_folds+1):
      val_cv, train_cv = split_data(data, num_folds, fold)
      model = train_model(train_cv, lambd)
```

```
cv loss lmd += loss(val cv, model)
    cv_error.append(cv_loss_lmd / num_folds)
  return cv_error
# part b
total_train = []
for i in range(len(data_train_X)):
  total_train.append((data_train_y[i], data_train_X[i]))
total_test = []
for i in range(len(data_test_X)):
  total_test.append((data_test_y[i], data_test_X[i]))
lambd seq = np.linspace(0.02, 1.5, 50)
cv_5_err = cross_validation(total_train, 5, lambd_seq)
cv_10_err = cross_validation(total_train, 10, lambd_seq)
# part c
def compute_loss(train_data, test_data, lambd_seq):
  train_loss, test_loss = [], []
  for i in lambd_seq:
    model = train_model(train_data, i)
    train_loss.append(loss(train_data, model))
    test_loss.append(loss(test_data, model))
  return train_loss, test_loss
train_error, test_error = compute_loss(total_train, total_test, lambd_seq)
# part d
def clip_data(data):
  res = []
  for i in data:
    res.append(i[0][0])
  return res
def plot graph():
  x = np.linspace(0.02, 1.5, 50)
  plt.plot(x, clip data(train error), label="train error")
  plt.plot(x, clip_data(test_error), label="test error")
  plt.plot(x, clip_data(cv_5_err), label="5 fold")
  plt.plot(x, clip_data(cv_10_err), label="10 fold")
  plt.xlabel('lambda')
  plt.ylabel('loss')
  plt.legend()
  plt.show()
plot_graph()
Graph:
```

