

Outline

- 1 ARIMA models
- 2 Lab Session 16
- 3 Seasonal ARIMA models
- 4 Lab Session 17
- 5 Forecast ensembles

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AR: autoregressive (lagged observations as inputs)

I: integrated (differencing to make series stationary)

MA: moving average (lagged errors as inputs)

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An ARIMA model is rarely interpretable in terms of visible data structures like trend and seasonality. But it can capture a huge range of time series patterns.

Stationarity

Definition

If $\{y_t\}$ is a stationary time series, then for all s, the distribution of (y_t, \ldots, y_{t+s}) does not depend on t.

Stationarity

Definition

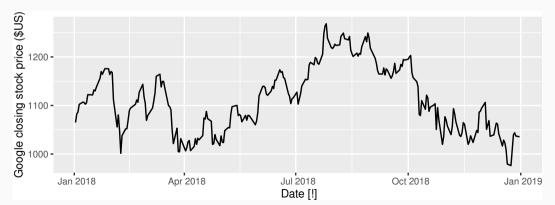
If $\{y_t\}$ is a stationary time series, then for all s, the distribution of (y_t, \ldots, y_{t+s}) does not depend on t.

A **stationary series** is:

- roughly horizontal
- constant variance
- no patterns predictable in the long-term

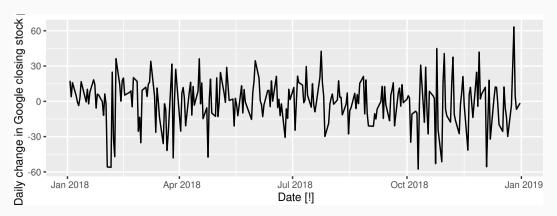
Stationary?

```
gafa_stock |>
  filter(Symbol == "G00G", year(Date) == 2018) |>
  autoplot(Close) +
  labs(y = "Google closing stock price ($US)")
```



Stationary?

```
gafa_stock |>
  filter(Symbol == "G00G", year(Date) == 2018) |>
  autoplot(difference(Close)) +
  labs(y = "Daily change in Google closing stock price")
```



Differencing

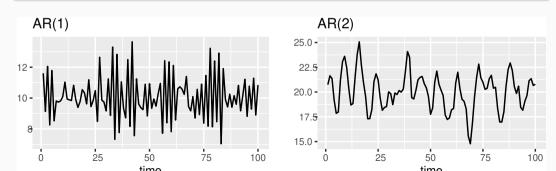
- Differencing helps to stabilize the mean.
- The differenced series is the *change* between each observation in the original series.
- Occasionally the differenced data will not appear stationary and it may be necessary to difference the data a second time.
- In practice, it is almost never necessary to go beyond second-order differences.

Autoregressive models

Autoregressive (AR) models:

$$y_t = c + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + \varepsilon_t$$

where ε_t is white noise. A multiple regression with **lagged** values of y_t as predictors.

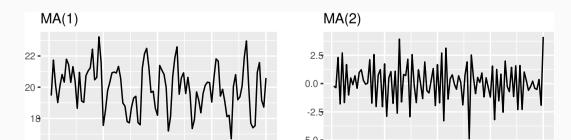


Moving Average (MA) models

Moving Average (MA) models:

$$y_t = c + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2} + \cdots + \theta_q \varepsilon_{t-q}$$

where ε_t is white noise. A multiple regression with **lagged errors** as predictors. Don't confuse with moving average smoothing!



Autoregressive Moving Average models:

$$y_{t} = c + \phi_{1}y_{t-1} + \dots + \phi_{p}y_{t-p}$$
$$+ \theta_{1}\varepsilon_{t-1} + \dots + \theta_{q}\varepsilon_{t-q} + \varepsilon_{t}.$$

Autoregressive Moving Average models:

$$y_{t} = c + \phi_{1}y_{t-1} + \dots + \phi_{p}y_{t-p}$$
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Predictors include both lagged values of y_t and lagged errors.

1

Autoregressive Moving Average models:

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■ Predictors include both lagged values of y_t and lagged errors.

Autoregressive Integrated Moving Average models

- Combine ARMA model with differencing.
- d-differenced series follows an ARMA model.
- Need to choose p, d, q and whether or not to include c.

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ARIMA(p, d, q) model

- AR: p = order of the autoregressive part
 - I: d = degree of first differencing involved
- MA: q = order of the moving average part.
 - White noise model: ARIMA(0,0,0)
 - Random walk: ARIMA(0,1,0) with no constant
 - Random walk with drift: ARIMA(0,1,0) with const.
 - \blacksquare AR(p): ARIMA(p,0,0)
 - \blacksquare MA(q): ARIMA(0,0,q)

```
fit <- global_economy |>
 model(arima = ARIMA(Population))
fit
# A mable: 263 x 2
# Key: Country [263]
                                           arima
   Country
   <fct>
                                         <model>
1 Afghanistan
                                 <ARIMA(4,2,1)>
2 Albania
                                 <ARIMA(0,2,2)>
3 Algeria
                                 <ARIMA(2,2,2)>
4 American Samoa
                                 < ARIMA(2,2,2) >
                        <ARIMA(2,1,2) w/ drift>
5 Andorra
6 Angola
                                 <ARIMA(4,2,1)>
 7 Antigua and Barbuda <ARIMA(2,1,2) w/ drift>
8 Arah World
                                 <ARIMA(0,2,1)>
```

```
fit |>
  filter(Country == "Australia") |>
  report()
Series: Population
Model: ARIMA(0,2,1)
Coefficients:
         ma1
      -0.661
s.e. 0.107
sigma^2 estimated as 4.063e+09: log likelihood=-699
ATC=1401 ATCc=1402
                       BTC=1405
```

```
fit |>
  filter(Country == "Australia") |>
  report()
Series: Population
Model: ARIMA(0,2,1)
Coefficients:
                                   V_t = 2V_{t-1} - V_{t-2} - 0.7\varepsilon_{t-1} + \varepsilon_t
           ma1
                                                  \varepsilon_t \sim \text{NID}(0.4 \times 10^9)
       -0.661
s.e. 0.107
sigma^2 estimated as 4.063e+09: log likelihood=-699
ATC=1401 ATCc=1402
                             BTC=1405
```

Understanding ARIMA models

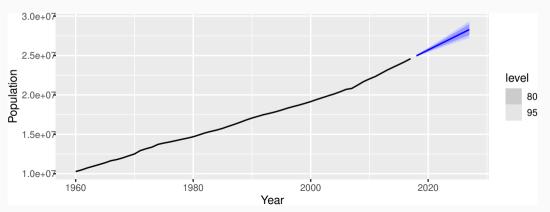
- If c = 0 and d = 0, the long-term forecasts will go to zero.
- If *c* = 0 and *d* = 1, the long-term forecasts will go to a non-zero constant.
- If c = 0 and d = 2, the long-term forecasts will follow a straight line.
- If $c \neq 0$ and d = 0, the long-term forecasts will go to the mean of the data.
- If $c \neq 0$ and d = 1, the long-term forecasts will follow a straight line.

Understanding ARIMA models

Forecast variance and d

- The higher the value of *d*, the more rapidly the prediction intervals increase in size.
- For d = 0, the long-term forecast standard deviation will go to the standard deviation of the historical data.

```
fit |>
  forecast(h = 10) |>
  filter(Country == "Australia") |>
  autoplot(global_economy)
```



Hyndman and Khandakar (JSS, 2008) algorithm:

- Select no. differences *d* via KPSS test.
- \blacksquare Select p, q and inclusion of c by minimising AICc.
- Use stepwise search to traverse model space.

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- Select no. differences d via KPSS test.
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AICc =
$$-2 \log(L) + 2(p+q+k+1) \left[1 + \frac{(p+q+k+2)}{T-p-q-k-2} \right]$$
.

where L is the maximised likelihood fitted to the differenced data, k = 1 if $c \ne 0$ and k = 0 otherwise.

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```
Step1: Select current model (with smallest AICc) from:
```

ARIMA(2, d, 2)

ARIMA(0, d, 0)

ARIMA(1, d, 0)

ARIMA(0, d, 1)

```
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```

- **Step 2:** Consider variations of current model:
 - vary one of p, q, from current model by ± 1 ;
 - ightharpoonup p, q both vary from current model by ± 1 ;
 - Include/exclude *c* from current model.

Model with lowest AICc becomes current model.

```
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Lab Session 16

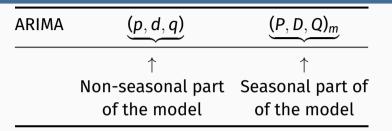
For the United States GDP data (from global_economy):

- Fit a suitable ARIMA model for the logged data.
- Produce forecasts of your fitted model. Do the forecasts look reasonable?

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Seasonal ARIMA models

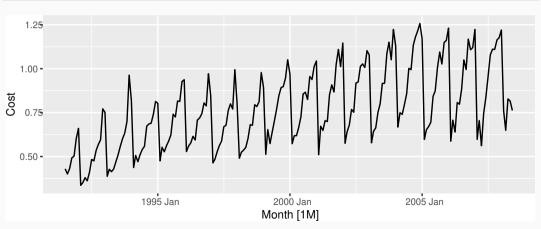


- \mathbf{m} = number of observations per year.
- d first differences, D seasonal differences
- p AR lags, q MA lags
- P seasonal AR lags, Q seasonal MA lags

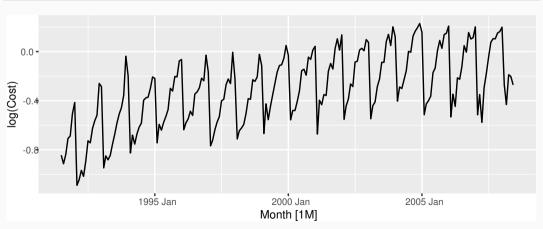
Seasonal and non-seasonal terms combine multiplicatively

```
h02 <- PBS |>
filter(ATC2 == "H02") |>
summarise(Cost = sum(Cost) / 1e6)
```

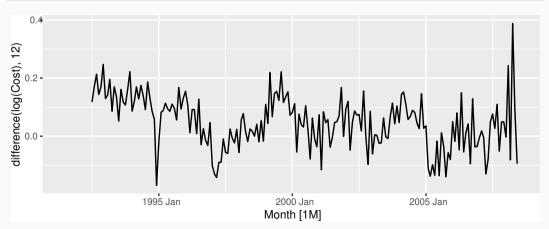
```
h02 |> autoplot(
   Cost
)
```



```
h02 |> autoplot(
   log(Cost)
)
```



```
h02 |> autoplot(
  log(Cost) |> difference(12)
)
```



Cortecosteroid drug sales

```
h02 |> autoplot(
   log(Cost) |> difference(12) |> difference(1)
difference(difference(log(Cost), 12),
     0.4
    0.2 -
    0.0 -
    -0.2 -
                                1995 Jan
                                                                                               2005 Jan
                                                                2000 Jan
```

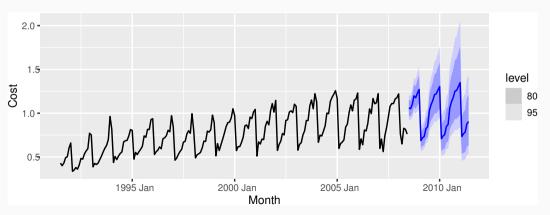
Month [1M]

Example: US electricity production

```
h02 |>
 model(arima = ARIMA(log(Cost))) |>
 report()
Series: Cost
Model: ARIMA(2,1,0)(0,1,1)[12]
Transformation: log(Cost)
Coefficients:
         ar1
             ar2
                          sma1
      -0.8491 -0.4207 -0.6401
s.e. 0.0712 0.0714 0.0694
sigma^2 estimated as 0.004387: log likelihood=245
ATC=-483 ATCc=-483
                      BTC=-470
```

Example: US electricity production

```
h02 |>
  model(arima = ARIMA(log(Cost))) |>
  forecast(h = "3 years") |>
  autoplot(h02)
```

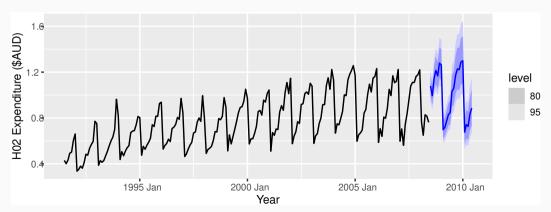


Cortecosteroid drug sales

```
fit <- h02 |>
 model(best = ARIMA(log(Cost),
   stepwise = FALSE.
   approximation = FALSE,
   order constraint = p + q + P + 0 \le 9
report(fit)
Series: Cost
Model: ARIMA(4,1,1)(2,1,2)[12]
Transformation: log(Cost)
Coefficients:
            ar2 ar3 ar4 ma1 sar1 sar2
                                                       smal sma2
        ar1
     -0.0425 0.210 0.202 -0.227 -0.742 0.621 -0.383 -1.202 0.496
s.e. 0.2167 0.181 0.114 0.081 0.207 0.242 0.118 0.249
                                                            0.213
sigma^2 estimated as 0.004049: log likelihood=254
ATC=-489 ATCc=-487
                    BTC=-456
```

Cortecosteroid drug sales

```
fit |>
  forecast() |>
  autoplot(h02) +
  labs(y = "H02 Expenditure ($AUD)", x = "Year")
```



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Lab Session 17

For the Australian tourism data (from tourism):

- Fit a suitable ARIMA model for all data.
- Produce forecasts of your fitted models.
- Check the forecasts for the "Snowy Mountains" and "Melbourne" regions. Do they look reasonable?

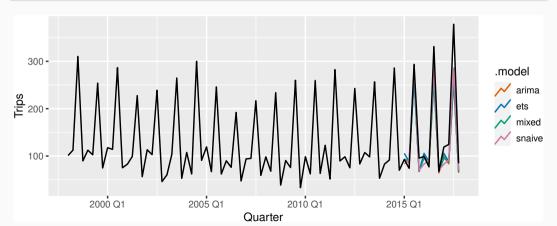
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```
train <- tourism |>
  filter(year(Quarter) <= 2014)
fit <- train |>
  model(
   ets = ETS(Trips),
   arima = ARIMA(Trips),
   snaive = SNAIVE(Trips)
) |>
  mutate(mixed = (ets + arima + snaive) / 3)
```

- Ensemble forecast mixed is a simple average of the three fitted models.
- forecast() will produce distributional forecasts taking into account the correlations between the forecast errors of the component models.

```
fc <- fit |> forecast(h = "3 years")
fc |>
  filter(Region == "Snowy Mountains", Purpose == "Holiday") |>
  autoplot(tourism, level = NULL)
```



```
accuracy(fc, tourism) |>
  group_by(.model) |>
  summarise(
    RMSE = mean(RMSE),
    MAE = mean(MAE),
    MASE = mean(MASE)
) |>
  arrange(RMSE)
```

```
# A tibble: 4 x 4
.model RMSE MAE MASE
<chr> <dbl> <dbl> <dbl> <dbl> 1 mixed 19.8 16.0 0.997
2 ets 20.2 16.4 1.00
3 snaive 21.5 17.3 1.17
4 arima 21.9 17.8 1.06
```

Can we do better than equal weights?

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- Hard to find weights that improve forecast accuracy.
- Known as the "forecast combination puzzle".
- Solution: FFORMA

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FFORMA (Feature-based FORecast Model Averaging)

- Vector of time series features used to predict best weights.
- A modification of xgboost is used.
- Method came 2nd in the 2018 M4 international forecasting competition.
- Main author: Pablo Montero-Manso (Monash U)
- Not (yet) available for fable.