

Outline

- 1 Reconciliation via constraints
- The geometry of forecast reconciliation
- 3 Optimization and reconcilation
- 4 ML and regularization
- 5 In-built coherence

Outline

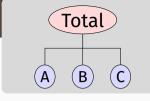
- 1 Reconciliation via constraints
- The geometry of forecast reconciliation
- 3 Optimization and reconcilation
- 4 ML and regularization
- 5 In-built coherence

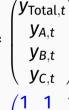
Notation reminder

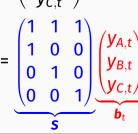
Every collection of time series with linear constraints can be written as

$$y_t = \mathbf{Sb_t}$$

- \mathbf{y}_t = vector of all series at time t
 - $y_{Total,t}$ = aggregate of all series at time t.
- $y_{X,t}$ = value of series X at time t.
- **\mathbf{b}_t** = vector of most disaggregated series at time t
- S = "summing matrix" containing the linear constraints.





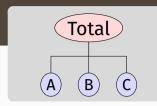


Notation reminder

Every collection of time series with linear constraints can be written as

$$y_t = \mathbf{Sb_t}$$

- \mathbf{v}_t = vector of all series at time t
- $y_{Total,t}$ = aggregate of all series at time t.
- $y_{X,t}$ = value of series X at time t.
- \mathbf{b}_t = vector of most disaggregated series at time t
- **S** = "summing matrix" containing the linear constraints.

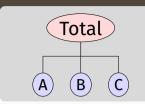


- Base forecasts: $\hat{\mathbf{y}}_{T+h|T}$
- Reconciled forecasts: $\tilde{\mathbf{y}}_{T+h|T} = \mathbf{S}\mathbf{G}\hat{\mathbf{y}}_{T+h|T}$
 - MinT:

$$G = (S'W_h^{-1}S)^{-1}S'W_h^{-1}$$

where W_h is
covariance matrix of
base forecast errors.

Notation



Aggregation matrix

$$y_t = \mathbf{Sb}_t$$

$$\begin{pmatrix} y_{\text{Total},t} \\ y_{A,t} \\ y_{B,t} \\ y_{C,t} \end{pmatrix} = \begin{pmatrix} 1 & 1 & 1 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} y_{A,t} \\ y_{B,t} \\ y_{C,t} \end{pmatrix}$$

$$\begin{pmatrix} \mathbf{a}_t \\ \mathbf{b}_t \end{pmatrix} = \begin{pmatrix} \mathbf{A} \\ \mathbf{I}_{n_b} \end{pmatrix} \mathbf{b}_t$$

Constraint matrix

where
$$Cy_t = 0$$

$$C = \begin{bmatrix} 1 & -1 & -1 & -1 \end{bmatrix}$$

$$= \begin{bmatrix} I_{n_a} & -A \end{bmatrix}$$

Aggregation matrix A

$$y_t = \begin{bmatrix} \boldsymbol{a}_t \\ \boldsymbol{b}_t \end{bmatrix} = \begin{bmatrix} \boldsymbol{A} \\ \boldsymbol{I}_{n_b} \end{bmatrix} \boldsymbol{b}_t = \boldsymbol{S} \boldsymbol{b}_t$$

Aggregation matrix A

$$y_t = \begin{bmatrix} a_t \\ b_t \end{bmatrix} = \begin{bmatrix} A \\ I_{n_b} \end{bmatrix} b_t = Sb_t$$

Constraint matrix C

$$Cy_t = 0$$

- Constraint matrix approach more general & more parsimonious.
- **C** = $[I_{n_a} -A]$.
- **S, A** and **C** may contain any real values (not just 0s and 1s).

Assuming **C** is full rank

$$\tilde{\mathbf{y}}_{T+h|T} = \mathbf{M}\hat{\mathbf{y}}_{T+h|T}$$

where $\mathbf{M} = \mathbf{I} - \mathbf{W}_h \mathbf{C}' (\mathbf{C} \mathbf{W}_h \mathbf{C}')^{-1} \mathbf{C}$

- Originally proved by Byron (1978) & Byron (1979) for reconciling data.
- Re-discovered by Wickramasuriya, Athanasopoulos, and Hyndman (2019) for reconciling forecasts.
- **M** = **SG** (the MinT solution)
- Leads to more efficient reconciliation than using G.

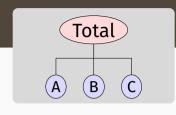
Suppose $W_h = I$. Then

$$\mathbf{M} = \mathbf{I} - \mathbf{W}_h \mathbf{C}' (\mathbf{C} \mathbf{W}_h \mathbf{C}')^{-1} \mathbf{C}$$

$$= \begin{pmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix} - \begin{pmatrix} 1 \\ -1 \\ -1 \\ -1 \end{pmatrix} \frac{1}{4} \begin{pmatrix} 1 & -1 & -1 & -1 \end{pmatrix}$$

$$= \begin{pmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix} - \begin{pmatrix} 1 & -1 & -1 & -1 \\ -1 & 1 & 1 & 1 \\ -1 & 1 & 1 & 1 \\ -1 & 1 & 1 & 1 \end{pmatrix}$$

$$= \begin{pmatrix} \frac{3}{4} & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{4} & \frac{3}{4} & -\frac{1}{4} & -\frac{1}{4} \\ \frac{1}{4} & -\frac{1}{4} & \frac{3}{4} & -\frac{1}{4} \\ \frac{1}{4} & -\frac{1}{4} & -\frac{1}{4} & \frac{3}{4} \end{pmatrix}$$

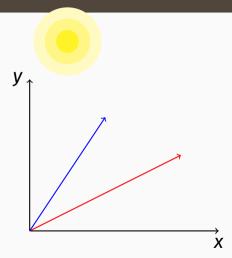


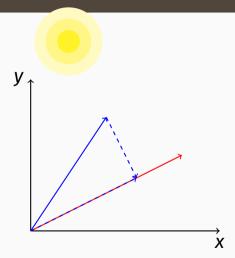
$$\mathbf{S} = \begin{pmatrix} \mathbf{A} \\ \mathbf{I}_{n_b} \end{pmatrix} = \begin{pmatrix} 1 & 1 & 1 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{pmatrix}$$
$$\mathbf{C} = (\mathbf{I}_{n_a} - \mathbf{A}) = (1 - 1 - 1 - 1)$$

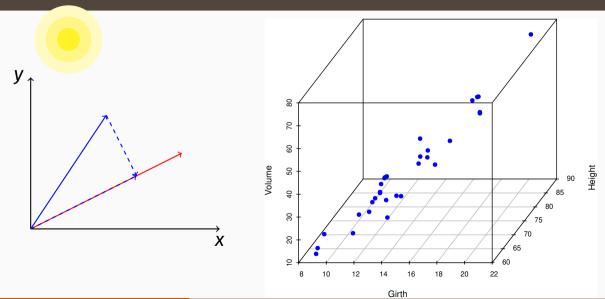
$$C = (I_{n_a} - A) = (1 - 1 - 1)$$

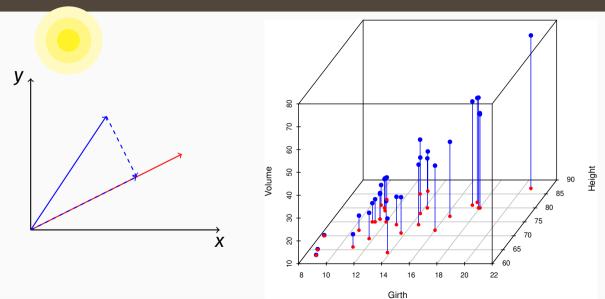
Outline

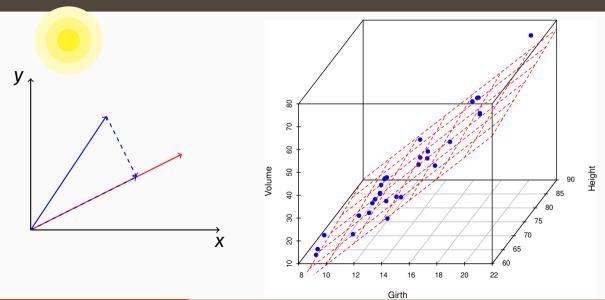
- 1 Reconciliation via constraints
- The geometry of forecast reconciliation
- 3 Optimization and reconcilation
- 4 ML and regularization
- 5 In-built coherence

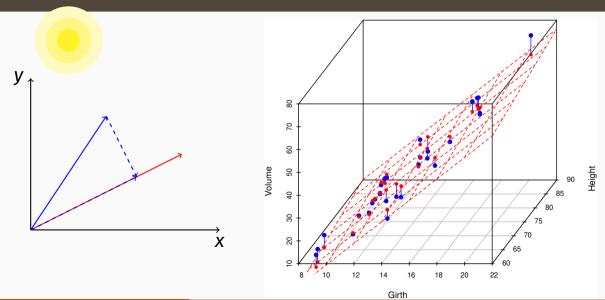












- A projection is a linear transformation M such that $M^2 = M$.
- i.e., *M* is idempotent: it leaves its image unchanged.
- **M** projects onto \mathfrak{s} if **My** = **y** for all $\mathbf{y} \in \mathfrak{s}$.
- All eigenvalues of **M** are either 0 or 1.
- All singular values of M are greater than or equal to 1 (with equality iff M is orthogonal).
- A projection is *orthogonal* if M' = M.
- If a projection is not orthogonal, it is called *oblique*.
- In regression, OLS is an orthogonal projection onto space spanned by predictors.

The coherent subspace

Coherent subspace

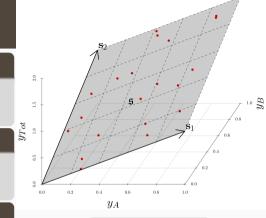
m-dimensional linear subspace $\mathfrak{s} \subset \mathbb{R}^n$ for which linear constraints hold for all $\mathbf{y} \in \mathfrak{s}$.

Hierarchical time series

An *n*-dimensional multivariate time series such that $\mathbf{v}_t \in \mathfrak{s} \quad \forall t$.

Coherent point forecasts

 $\tilde{\mathbf{y}}_{t+h|t}$ is coherent if $\tilde{\mathbf{y}}_{t+h|t} \in \mathfrak{s}$.



$$y_{Tot} = y_A + y_B$$

The coherent subspace

Coherent subspace

m-dimensional linear subspace $\mathfrak{s} \subset \mathbb{R}^n$ for which linear constraints hold for all $\mathbf{y} \in \mathfrak{s}$.

Hierarchical time series

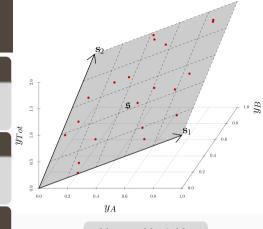
An *n*-dimensional multivariate time series such that $\mathbf{y}_t \in \mathfrak{s} \quad \forall t$.

Coherent point forecasts

 $\tilde{\mathbf{y}}_{t+h|t}$ is coherent if $\tilde{\mathbf{y}}_{t+h|t} \in \mathfrak{s}$.

Base forecasts

Let $\hat{\mathbf{y}}_{t+h|t}$ be vector of *incoherent* initial *h*-step forecasts.



 $y_{Tot} = y_A + y_B$

The coherent subspace

Coherent subspace

m-dimensional linear subspace $\mathfrak{s} \subset \mathbb{R}^n$ for which linear constraints hold for all $\mathbf{y} \in \mathfrak{s}$.

Hierarchical time series

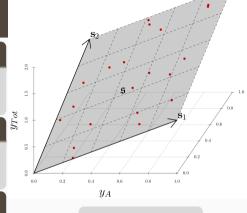
An *n*-dimensional multivariate time series such that $\mathbf{y}_t \in \mathfrak{s} \quad \forall t.$

Coherent point forecasts

 $\tilde{\mathbf{y}}_{t+h|t}$ is coherent if $\tilde{\mathbf{y}}_{t+h|t} \in \mathfrak{s}$.

Base forecasts

Let $\hat{\mathbf{y}}_{t+h|t}$ be vector of incoherent initial h-step forecasts.

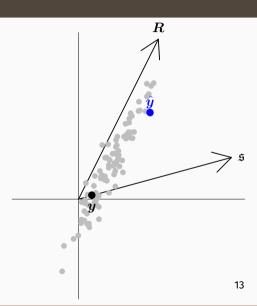


 $y_{Tot} = y_A + y_B$

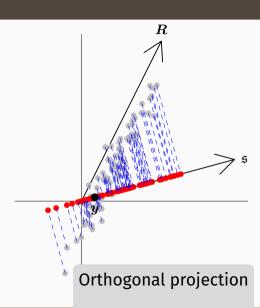
sciled forecasts

Reconciled forecasts Let ψ be a mapping, $\psi : \mathbb{R}^n \to \mathfrak{s}$. $\tilde{\mathbf{y}}_{t+h|t} = \psi(\hat{\mathbf{y}}_{t+h|t})$ "reconciles" $\hat{\mathbf{y}}_{t+h|t}$.

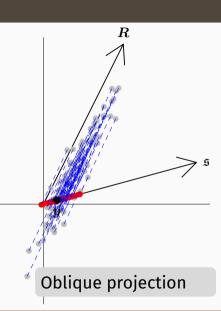
- \blacksquare *R* is the most likely direction of deviations from \mathfrak{s} .
- Grey: potential base forecasts



- R is the most likely direction of deviations from s.
- Grey: potential base forecasts
- Red: reconciled forecasts
- Orthogonal projections (i.e., OLS) lead to smallest possible adjustments of base forecasts.



- R is the most likely direction of deviations from s.
- Grey: potential base forecasts
- Red: reconciled forecasts
- Orthogonal projections (i.e., OLS) lead to smallest possible adjustments of base forecasts.
- Oblique projections (i.e., MinT) give reconciled forecasts with smallest variance.



$$\tilde{\mathbf{y}}_{t+h|t} = \psi(\hat{\mathbf{y}}_{t+h|t}) = \mathbf{M}\hat{\mathbf{y}}_{t+h|t}$$

- **M** is a projection onto $\mathfrak s$ if and only if My = y for all $y \in \mathfrak s$.
- Coherent base forecasts are unchanged since $M\hat{y} = \hat{y}$
- If \hat{y} is unbiased, then \tilde{y} is also unbiased since

$$\mathsf{E}(\tilde{\boldsymbol{y}}_{t+h|t}) = \mathsf{E}(\boldsymbol{M}\hat{\boldsymbol{y}}_{t+h|t}) = \boldsymbol{M}\mathsf{E}(\hat{\boldsymbol{y}}_{t+h|t}) = \mathsf{E}(\hat{\boldsymbol{y}}_{t+h|t}),$$

and unbiased estimates must lie on s.

- The projection is orthogonal if and only if M' = M.
- If **S** forms a basis set for \mathfrak{S} , then projections are of the form $\mathbf{M} = \mathbf{S}(\mathbf{S}'\Psi\mathbf{S})^{-1}\mathbf{S}'\Psi$ where Ψ is a positive definite matrix.

$$\tilde{\mathbf{y}}_{t+h|t} = \psi(\hat{\mathbf{y}}_{t+h|t}) = \mathbf{M}\hat{\mathbf{y}}_{t+h|t}, \quad \text{where} \quad \mathbf{M} = \mathbf{S}(\mathbf{S}'\Psi\mathbf{S})^{-1}\mathbf{S}'\Psi$$

OLS:
$$\Psi = I$$
 $M = S(S'S)^{-1}S'$ $= I - C'(CC')^{-1}C$
MinT: $\Psi = W_h$ $M = S(S'W_h^{-1}S)^{-1}S'W_h^{-1}$ $= I - W_hC'(CW_hC')^{-1}C$

- **M** is orthogonal iff Ψ = **I**.
- $\mathbf{W}_h = \text{Var}[\mathbf{y}_{T+h} \hat{\mathbf{y}}_{T+h|T} \mid \mathbf{y}_1, \dots, \mathbf{y}_T]$ is the covariance matrix of the base forecast errors.
- $V_h = \text{Var}[y_{T+h} \tilde{y}_{T+h|T} \mid y_1, ..., y_T] = MW_hM'$ is minimized when $\Psi = W_h$.

Mean square error bounds

Panagiotelis, Gamakumara, Athanasopoulos, and Hyndman (2021)

Distance reducing property

Let $\|\mathbf{u}\|_{\Psi}$ = $\mathbf{u}'\Psi\mathbf{u}$. Then

$$\| oldsymbol{y}_{t+h} - oldsymbol{ ilde{y}}_{t+h|t} \|_{\Psi} \leq \| oldsymbol{y}_{t+h} - oldsymbol{\hat{y}}_{t+h|t} \|_{\Psi}$$

- Ψ -projection is guaranteed to improve forecast accuracy over base forecasts using this distance measure.
- Distance reduction holds for any realisation and any forecast.
- OLS reconciliation minimizes Euclidean distance.
- Other measures of forecast accuracy may be worse.

Wickramasuriya (2021)

$$\|\mathbf{y}_{t+h} - \tilde{\mathbf{y}}_{t+h}\|_{2}^{2} = \|\mathbf{M}(\mathbf{y}_{t+h} - \hat{\mathbf{y}}_{t+h})\|_{2}^{2}$$

 $\leq \|\mathbf{M}\|_{2}^{2} \|\mathbf{y}_{t+h} - \hat{\mathbf{y}}_{t+h}\|_{2}^{2}$
 $= \sigma_{\max}^{2} \|\mathbf{y}_{t+h} - \hat{\mathbf{y}}_{t+h}\|_{2}^{2}$

- lacksquare σ_{\max} is the largest eigenvalue of $m{M}$
- lacksquare $\sigma_{\max} \geq$ 1 as **M** is a projection matrix.
- Every projection reconciliation is better than base forecasts using Euclidean distance.

Wickramasuriya (2021)

$$\begin{split} & \operatorname{\mathsf{tr}} \Big(\mathsf{E}[\boldsymbol{y}_{t+h} - \tilde{\boldsymbol{y}}_{t+h|t}^{\mathsf{MinT}}]' [\boldsymbol{y}_{t+h} - \tilde{\boldsymbol{y}}_{t+h|t}^{\mathsf{MinT}}] \Big) \\ & \leq & \operatorname{\mathsf{tr}} \Big(\mathsf{E}[\boldsymbol{y}_{t+h} - \tilde{\boldsymbol{y}}_{t+h|t}^{\mathsf{OLS}}]' [\boldsymbol{y}_{t+h} - \tilde{\boldsymbol{y}}_{t+h|t}^{\mathsf{OLS}}] \Big) \\ & \leq & \operatorname{\mathsf{tr}} \Big(\mathsf{E}[\boldsymbol{y}_{t+h} - \hat{\boldsymbol{y}}_{t+h|t}]' [\boldsymbol{y}_{t+h} - \hat{\boldsymbol{y}}_{t+h|t}] \Big) \end{split}$$

Using sums of variances:

- MinT reconciliation is better than OLS reconciliation
- OLS reconciliation is better than base forecasts

Outline

- 1 Reconciliation via constraints
- 2 The geometry of forecast reconciliation
- 3 Optimization and reconcilation
- 4 ML and regularization
- 5 In-built coherence

Minimum trace reconciliation

Minimum trace (MinT) reconciliation

If **SG** is a projection, then the trace of $\mathbf{V}_h = \text{Var}(\tilde{\mathbf{y}}_{t+h|t} - \mathbf{y}_{t+h})$ is **minimized** when

$$G = (S'W_h^{-1}S)^{-1}S'W_h^{-1}$$

$$\tilde{\mathbf{y}}_{T+h|T} = \mathbf{S}(\mathbf{S}'\mathbf{W}_h^{-1}\mathbf{S})^{-1}\mathbf{S}'\mathbf{W}_h^{-1}\hat{\mathbf{y}}_{T+h|T}$$

Reconciled forecasts

Base forecasts

- Trace of V_h is sum of forecast variances.
- MinT solution is L₂ optimal amongst linear unbiased forecasts.

Find the solution to the minimax problem

$$V = \min_{\tilde{\mathbf{y}} \in \mathfrak{s}} \max_{\mathbf{y} \in \mathfrak{s}} \left\{ \ell(\mathbf{y}, \tilde{\mathbf{y}}) - \ell(\mathbf{y}, \hat{\mathbf{y}}) \right\},$$

where ℓ is a loss function, and \mathfrak{s} is the coherent subspace.

- V < 0: reconciliation guaranteed to reduce loss.
- If $\ell(\mathbf{v}, \tilde{\mathbf{v}}) = \|\mathbf{v} \tilde{\mathbf{v}}\|_{\Psi} = (\mathbf{v} \tilde{\mathbf{v}})'\Psi(\mathbf{v} \tilde{\mathbf{v}})$, where Ψ is any symmetric pd matrix, then:
 - $\tilde{\mathbf{y}} = \mathbf{S}(\mathbf{S}'\Psi\mathbf{S})^{-1}\mathbf{S}'\Psi\hat{\mathbf{y}}$ will always improve upon the base forecasts;
 - The MinT solution $\tilde{\mathbf{y}} = \mathbf{S}(\mathbf{S}'\mathbf{W}_h^{-1}\mathbf{S})^{-1}\mathbf{S}'\mathbf{W}_h^{-1}\hat{\mathbf{y}}$ will optimise loss in expectation over any choice of Ψ .

Regularized empirical risk minimization problem:

$$\min_{\boldsymbol{G}} \frac{1}{Nn} \| \boldsymbol{Y} - \hat{\boldsymbol{Y}} \boldsymbol{G}' \boldsymbol{S}' \|_F + \lambda \| \text{vec} \boldsymbol{G} \|_1,$$

- \blacksquare N = T T₁ h + 1, T₁ is minimum training sample size
- $\|\cdot\|_F$ is the Frobenius norm
- $\mathbf{Y} = [\mathbf{y}_{T_1+h}, \dots, \mathbf{y}_T]'$
- lacksquare λ is a regularization parameter.

When
$$\lambda = 0$$
: $\hat{\boldsymbol{G}} = \boldsymbol{B}'\hat{\boldsymbol{Y}}(\hat{\boldsymbol{Y}}'\hat{\boldsymbol{Y}})^{-1}$ where $\boldsymbol{B} = [\boldsymbol{b}_{T_1+h}, \dots, \boldsymbol{b}_T]'$.

Non-negative forecasts

$$\min_{\boldsymbol{G}_h} \operatorname{tr} \left(\operatorname{E}[\boldsymbol{y}_{t+h} - \boldsymbol{S} \boldsymbol{G}_h \hat{\boldsymbol{y}}_{t+h|t}]' [\boldsymbol{y}_{t+h} - \boldsymbol{S} \boldsymbol{G}_h \hat{\boldsymbol{y}}_{t+h|t}] \right)$$
 such that $\boldsymbol{b}_{t+h|t} = \boldsymbol{G}_h \hat{\boldsymbol{y}}_{t+h|t} \geq 0$

Non-negative forecasts

$$\min_{\boldsymbol{G}_h} \operatorname{tr} \left(\operatorname{E}[\boldsymbol{y}_{t+h} - \boldsymbol{S}\boldsymbol{G}_h \hat{\boldsymbol{y}}_{t+h|t}]' [\boldsymbol{y}_{t+h} - \boldsymbol{S}\boldsymbol{G}_h \hat{\boldsymbol{y}}_{t+h|t}] \right)$$

such that $\boldsymbol{b}_{t+h|t} = \boldsymbol{G}_h \hat{\boldsymbol{y}}_{t+h|t} \geq 0$

Solve via quadratic programming:

$$\min_{\boldsymbol{b}} \frac{1}{2} \boldsymbol{b}' \boldsymbol{S}' \boldsymbol{W}_h^{-1} \boldsymbol{S} \boldsymbol{b} - \boldsymbol{b}' \boldsymbol{S}' \boldsymbol{W}_h^{-1} \hat{\boldsymbol{y}}_{T+h|T}$$
 s.t. $\boldsymbol{b} \geq 0$

(Wickramasuriya, Turlach, and Hyndman, 2020)

Non-negative forecasts

$$\mathsf{min}_{m{G}_h} \mathsf{tr} \Big(\mathsf{E}[m{y}_{t+h} - m{S}m{G}_h \hat{m{y}}_{t+h|t}]' [m{y}_{t+h} - m{S}m{G}_h \hat{m{y}}_{t+h|t}] \Big)$$
 such that $m{b}_{t+h|t} = m{G}_h \hat{m{y}}_{t+h|t} \geq 0$

Solve via quadratic programming:

$$\min_{\boldsymbol{b}} \frac{1}{2} \boldsymbol{b}' \boldsymbol{S}' \boldsymbol{W}_h^{-1} \boldsymbol{S} \boldsymbol{b} - \boldsymbol{b}' \boldsymbol{S}' \boldsymbol{W}_h^{-1} \hat{\boldsymbol{y}}_{T+h|T}$$
 s.t. $\boldsymbol{b} \geq 0$

(Wickramasuriya, Turlach, and Hyndman, 2020)

Set-negative-to-zero heuristic solution

- Negative reconciled forecasts at bottom level set to zero
- Remaining forecasts computed via aggregation
 (Di Fonzo and Girolimetto, 2023)

Immutable forecasts

Hollyman, Petropoulos, and Tipping (2021)

$$\min_{\boldsymbol{G}_h} \operatorname{tr} \left(\operatorname{E}[\boldsymbol{y}_{t+h} - \boldsymbol{S}\boldsymbol{G}_h \hat{\boldsymbol{y}}_{t+h|t}]' [\boldsymbol{y}_{t+h} - \boldsymbol{S}\boldsymbol{G}_h \hat{\boldsymbol{y}}_{t+h|t}] \right)$$
such that $\boldsymbol{C}\boldsymbol{S}\boldsymbol{G}_h \hat{\boldsymbol{y}}_{t+h|t} = \boldsymbol{d}$

 Differs from top-down approaches in that it can be done while also preserving the unbiasedness of base forecasts.

See also Di Fonzo and Girolimetto (2022). Zhang, Kang, Panagiotelis, and Li (2022).

Outline

- 1 Reconciliation via constraints
- 2 The geometry of forecast reconciliation
- 3 Optimization and reconcilation
- 4 ML and regularization
- 5 In-built coherence

ML and regularization

- Replace the linear regression formulation with a less restrictive method to obtain combinations of forecasts from the various hierarchical levels.
- Coherence is achieved via a bottom-up approach, or by embedding coherence in the ML training.

Gleason (2020) attempts to overcome the lack of focus on coherence by adjusting the objective function. Using neural network forecasts, he includes a regularisation term that penalises incoherences in the generated forecasts. This

Outline

- 1 Reconciliation via constraints
- The geometry of forecast reconciliation
- 3 Optimization and reconcilation
- 4 ML and regularization
- 5 In-built coherence

In-built coherence

Two-step approach: compute base forecasts \hat{y}_h , and then reconcile them to produce \tilde{y}_h .

One-step approaches: compute coherent $\tilde{\mathbf{y}}_h$ directly.

- Ashouri, Hyndman, and Shmueli (2022): linear regression models
- Pennings and Dalen (2017): state space models
- Villegas and Pedregal (2018): state space models

Suppose
$$\hat{y}_{t,i} = \hat{\beta}'_i \mathbf{x}_{t,i}$$
 with $\mathbf{x}_{t,i} = (1, x_{t,1,i}, \dots, x_{t,p,i}) \& \hat{\mathbf{y}}_i = (\hat{y}_{1,i}, \dots, \hat{y}_{T,i})$.

Suppose
$$\hat{y}_{t,i} = \hat{\beta}'_i \mathbf{x}_{t,i}$$
 with $\mathbf{x}_{t,i} = (1, x_{t,1,i}, \dots, x_{t,p,i}) & \hat{\mathbf{y}}_i = (\hat{y}_{1,i}, \dots, \hat{y}_{T,i}).$

$$\begin{pmatrix} \hat{\mathbf{y}}_1 \\ \hat{\mathbf{y}}_2 \\ \vdots \\ \hat{\mathbf{y}}_n \end{pmatrix} = \begin{pmatrix} \mathbf{X}_1 & 0 & \dots & 0 \\ 0 & \mathbf{X}_2 & \ddots & \vdots \\ \vdots & \ddots & \ddots & 0 \\ 0 & \dots & 0 & \mathbf{X}_n \end{pmatrix} \begin{pmatrix} \hat{\boldsymbol{\beta}}_1 \\ \hat{\boldsymbol{\beta}}_2 \\ \vdots \\ \hat{\boldsymbol{\beta}}_n \end{pmatrix}, \qquad \mathbf{X}_i = \begin{pmatrix} 1 & x_{1,i,1} & x_{1,i,2} & \dots & x_{1,i,p} \\ 1 & x_{2,i,1} & x_{2,i,2} & \dots & x_{2,i,p} \\ \vdots & \vdots & \vdots & & \vdots \\ 1 & x_{T,i,1} & x_{T,i,2} & \dots & x_{T,i,p} \end{pmatrix}$$

Suppose
$$\hat{y}_{t,i} = \hat{\beta}'_i \mathbf{x}_{t,i}$$
 with $\mathbf{x}_{t,i} = (1, x_{t,1,i}, \dots, x_{t,p,i}) \& \hat{\mathbf{y}}_i = (\hat{y}_{1,i}, \dots, \hat{y}_{T,i})$.

$$\begin{pmatrix}
\hat{\mathbf{y}}_1 \\
\hat{\mathbf{y}}_2 \\
\vdots \\
\hat{\mathbf{y}}_n
\end{pmatrix} = \begin{pmatrix}
\mathbf{X}_1 & 0 & \dots & 0 \\
0 & \mathbf{X}_2 & \ddots & \vdots \\
\vdots & \ddots & \ddots & 0 \\
0 & \dots & 0 & \mathbf{X}_n
\end{pmatrix} \begin{pmatrix}
\hat{\boldsymbol{\beta}}_1 \\
\hat{\boldsymbol{\beta}}_2 \\
\vdots \\
\hat{\boldsymbol{\beta}}_n
\end{pmatrix}, \qquad \mathbf{X}_i = \begin{pmatrix}
1 & X_{1,i,1} & X_{1,i,2} & \dots & X_{1,i,p} \\
1 & X_{2,i,1} & X_{2,i,2} & \dots & X_{2,i,p} \\
\vdots & \vdots & \vdots & & \vdots \\
1 & X_{T,i,1} & X_{T,i,2} & \dots & X_{T,i,p}
\end{pmatrix}$$

$$\hat{\mathbf{B}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$$

Suppose
$$\hat{y}_{t,i} = \hat{\beta}'_i \mathbf{x}_{t,i}$$
 with $\mathbf{x}_{t,i} = (1, x_{t,1,i}, \dots, x_{t,p,i}) & \hat{\mathbf{y}}_i = (\hat{y}_{1,i}, \dots, \hat{y}_{T,i}).$

$$\begin{pmatrix} \hat{\mathbf{y}}_{1} \\ \hat{\mathbf{y}}_{2} \\ \vdots \\ \hat{\mathbf{y}}_{n} \end{pmatrix} = \begin{pmatrix} \mathbf{X}_{1} & 0 & \dots & 0 \\ 0 & \mathbf{X}_{2} & \ddots & \vdots \\ \vdots & \ddots & \ddots & 0 \\ 0 & \dots & 0 & \mathbf{X}_{n} \end{pmatrix} \begin{pmatrix} \hat{\boldsymbol{\beta}}_{1} \\ \hat{\boldsymbol{\beta}}_{2} \\ \vdots \\ \hat{\boldsymbol{\beta}}_{n} \end{pmatrix}, \qquad \mathbf{X}_{i} = \begin{pmatrix} 1 & X_{1,i,1} & X_{1,i,2} & \dots & X_{1,i,p} \\ 1 & X_{2,i,1} & X_{2,i,2} & \dots & X_{2,i,p} \\ \vdots & \vdots & \vdots & & \vdots \\ 1 & X_{T,i,1} & X_{T,i,2} & \dots & X_{T,i,p} \end{pmatrix}$$

$$\hat{\boldsymbol{B}} = (\boldsymbol{X}'\boldsymbol{X})^{-1}\boldsymbol{X}'\boldsymbol{Y} \qquad \hat{\boldsymbol{y}}_{t+h} = \boldsymbol{X}_{t+h}^* \hat{\boldsymbol{B}}$$

Suppose
$$\hat{y}_{t,i} = \hat{\beta}'_i \mathbf{x}_{t,i}$$
 with $\mathbf{x}_{t,i} = (1, x_{t,1,i}, \dots, x_{t,p,i}) & \hat{\mathbf{y}}_i = (\hat{y}_{1,i}, \dots, \hat{y}_{T,i}).$

$$\begin{pmatrix} \hat{\mathbf{y}}_{1} \\ \hat{\mathbf{y}}_{2} \\ \vdots \\ \hat{\mathbf{y}}_{n} \end{pmatrix} = \begin{pmatrix} \mathbf{X}_{1} & 0 & \dots & 0 \\ 0 & \mathbf{X}_{2} & \ddots & \vdots \\ \vdots & \ddots & \ddots & 0 \\ 0 & \dots & 0 & \mathbf{X}_{n} \end{pmatrix} \begin{pmatrix} \hat{\boldsymbol{\beta}}_{1} \\ \hat{\boldsymbol{\beta}}_{2} \\ \vdots \\ \hat{\boldsymbol{\beta}}_{n} \end{pmatrix}, \qquad \mathbf{X}_{i} = \begin{pmatrix} 1 & X_{1,i,1} & X_{1,i,2} & \dots & X_{1,i,p} \\ 1 & X_{2,i,1} & X_{2,i,2} & \dots & X_{2,i,p} \\ \vdots & \vdots & \vdots & & \vdots \\ 1 & X_{T,i,1} & X_{T,i,2} & \dots & X_{T,i,p} \end{pmatrix}$$

$$\hat{\boldsymbol{B}} = (\boldsymbol{X}'\boldsymbol{X})^{-1}\boldsymbol{X}'\boldsymbol{Y} \qquad \hat{\boldsymbol{y}}_{t+h} = \boldsymbol{X}_{t+h}^*\hat{\boldsymbol{B}} \qquad \boldsymbol{X}_{t+h}^* = \operatorname{diag}(\boldsymbol{x}_{t+h,i}', \dots, \boldsymbol{x}_{t+h,n}')$$

Suppose
$$\hat{y}_{t,i} = \hat{\beta}'_i \mathbf{x}_{t,i}$$
 with $\mathbf{x}_{t,i} = (1, x_{t,1,i}, \dots, x_{t,p,i}) \& \hat{\mathbf{y}}_i = (\hat{y}_{1,i}, \dots, \hat{y}_{T,i})$.

$$\begin{pmatrix} \hat{\mathbf{y}}_{1} \\ \hat{\mathbf{y}}_{2} \\ \vdots \\ \hat{\mathbf{y}}_{n} \end{pmatrix} = \begin{pmatrix} \mathbf{X}_{1} & 0 & \dots & 0 \\ 0 & \mathbf{X}_{2} & \ddots & \vdots \\ \vdots & \ddots & \ddots & 0 \\ 0 & \dots & 0 & \mathbf{X}_{n} \end{pmatrix} \begin{pmatrix} \hat{\boldsymbol{\beta}}_{1} \\ \hat{\boldsymbol{\beta}}_{2} \\ \vdots \\ \hat{\boldsymbol{\beta}}_{n} \end{pmatrix}, \qquad \mathbf{X}_{i} = \begin{pmatrix} 1 & x_{1,i,1} & x_{1,i,2} & \dots & x_{1,i,p} \\ 1 & x_{2,i,1} & x_{2,i,2} & \dots & x_{2,i,p} \\ \vdots & \vdots & \vdots & & \vdots \\ 1 & x_{T,i,1} & x_{T,i,2} & \dots & x_{T,i,p} \end{pmatrix}$$

$$\hat{\mathbf{B}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$$
 $\hat{\mathbf{y}}_{t+h} = \mathbf{X}_{t+h}^*\hat{\mathbf{B}}$ $\mathbf{X}_{t+h}^* = \text{diag}(\mathbf{x}_{t+h,i}', \dots, \mathbf{x}_{t+h,n}')$

$$\tilde{\mathbf{y}}_{t+h} = \mathbf{S}(\mathbf{S}'\mathbf{W}_{h}\mathbf{S})^{-1}\mathbf{S}'\mathbf{W}_{h}\hat{\mathbf{y}}_{t+h} = \mathbf{S}(\mathbf{S}'\mathbf{W}_{h}\mathbf{S})^{-1}\mathbf{S}'\mathbf{W}_{h}\mathbf{X}_{t+h}^{*}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$$

$$\mathbf{V}_{h} = \sigma^{2}\mathbf{S}(\mathbf{S}'\mathbf{W}_{h}\mathbf{S})^{-1}\mathbf{S}'\mathbf{W}_{h}\left[1 + \mathbf{X}_{T+h}^{*}(\mathbf{X}'\mathbf{X})^{-1}(\mathbf{X}_{T+h}^{*})'\right]\mathbf{W}_{h}\mathbf{S}'(\mathbf{S}'\mathbf{W}_{h}\mathbf{S})^{-1}\mathbf{S}'$$

Reference: Ashouri, Hyndman, and Shmueli (2022)

In-built coherence

Pennings and Dalen (2017) propose the state space model

$$\mathbf{y}_t = \mathbf{S}\mu_t + \mathbf{Z}_t \boldsymbol{\beta} + \boldsymbol{\varepsilon}_t, \qquad \qquad \boldsymbol{\varepsilon}_t \sim N(\mathbf{0}, \boldsymbol{\Sigma}_{\varepsilon}),$$
 (1)

$$\mu_t = \mu_{t-1} + \eta_t, \qquad \qquad \eta_t \sim N(\mathbf{0}, \Sigma_{\eta}).$$
 (2)

- Coherent forecasts arise naturally using the Kalman filter
- Covariance matrices difficult to estimate except for small hierarchies.

In-build coherence

A related state space approach was proposed by Villegas and Pedregal (2018), who show that their formulation subsumes bottom-up, top-down, and some forms of forecast reconciliation and combination forecasting.

- Ashouri, M, RJ Hyndman, and G Shmueli (2022). Fast forecast reconciliation using linear models. J Computational & Graphical Statistics 31(1), 263–282.
- Ben Taieb, S and B Koo (2019). Regularized regression for hierarchical forecasting without unbiasedness conditions. In: *Proceedings of the 25th ACM SIGKDD International Conference on Knowledge Discovery & Data Mining.* KDD '19. Anchorage, AK, USA: Association for Computing Machinery, pp.1337–1347.
- Byron, RP (1978). The estimation of large social account matrices. *Journal of the Royal Statistical Society, Series A* **141**(3), 359–367.
- Byron, RP (1979). Corrigenda: The estimation of large social account matrices. Journal of the Royal Statistical Society. Series A 142(3), 405.

- Di Fonzo, T and D Girolimetto (2022). Forecast combination-based forecast reconciliation: Insights and extensions. *International Journal of Forecasting* **forthcoming**.
- Di Fonzo, T and D Girolimetto (2023). Spatio-temporal reconciliation of solar forecasts. *Solar Energy* **251**, 13–29.
- Gleason, JL (2020). Forecasting hierarchical time series with a regularized embedding space. In: MileTS '20: 6th KDD Workshop on Mining and Learning from Time Series 2020. San Diego, California, USA, pp.883–894.
- Hollyman, R, F Petropoulos, and ME Tipping (2021). Understanding forecast reconciliation. European Journal of Operational Research **294**(1), 149–160.
- Panagiotelis, A, P Gamakumara, G Athanasopoulos, and RJ Hyndman (2021). Forecast reconciliation: A geometric view with new insights on bias correction. *International J Forecasting* **37**(1), 343–359.

- Pennings, CL and J van Dalen (2017). Integrated hierarchical forecasting. European Journal of Operational Research **263**(2), 412–418.
- van Erven, T and J Cugliari (2015). "Game-theoretically optimal reconciliation of contemporaneous hierarchical time series forecasts". In: Modeling and Stochastic Learning for Forecasting in High Dimension. Ed. by A Antoniadis, JM Poggi, and X Brossat. Cham: Springer International Publishing, pp.297–317.
- Villegas, MA and DJ Pedregal (2018). Supply chain decision support systems based on a novel hierarchical forecasting approach. *Decision Support Systems* 114, 29–36.
- Wickramasuriya, SL (2021). Properties of point forecast reconciliation approaches. arXiv preprint arXiv:2103.11129.

- Wickramasuriya, SL, G Athanasopoulos, and RJ Hyndman (2019). Optimal forecast reconciliation for hierarchical and grouped time series through trace minimization. *J American Statistical Association* **114**(526), 804–819.
- Wickramasuriya, SL, BA Turlach, and RJ Hyndman (2020). Optimal non-negative forecast reconciliation. *Statistics & Computing* **30**(5), 1167–1182.
- Zhang, B, Y Kang, A Panagiotelis, and F Li (2022). Optimal reconciliation with immutable forecasts. European Journal of Operational Research forthcoming.