GEOG574/Math Introduction to Geostatistics

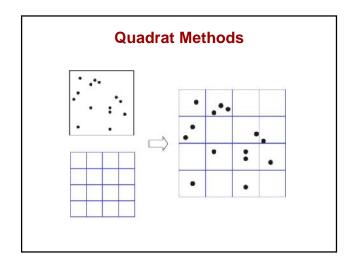
Point Pattern Analysis

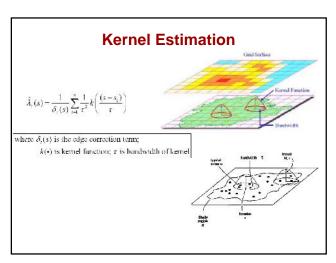
Daoqin Tong

School of Geography &Development 408 Harvill Building Email: daoqin@email.arizona.edu

Point Patterns

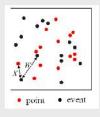
- A spatial point process is a spatial stochastic process (Y(s), s ∈ ℜ), where ℜ is random
- A spatial point pattern (y(s₁), y(s₂),..., y(s₂),..., y(s₂) is a realization of a spatial point process
- Exploring Spatial Point Patterns
 - First order effects
 - ⇒ Quadrat methods
 - \Rightarrow Kernel estimation
 - Second order effects
 - ⇒ Nearest neighbor distances
 - ⇒ K function





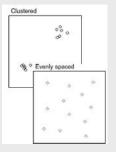
Nearest Neighbor Distances

- · The distribution of inter-event distances is closely related to the second order effects
- Two types of distances
 - W: event-event nearest neighbor distances (distances between a randomly chosen event and its nearest neighbor event)
 - X: point-event nearest neighbor distances (distances between a randomly chosen point in the study area and its nearest neighbor event)



Expected behavior of mean nearest neighbor distance

- For a *clustered* pattern
 - All nearest neighbor distances are short
 - So the mean is small
- · An evenly-spaced pattern
 - Minimum distances are longer
 - Mean is higher



Distance-based Measures

- W vs. X
 - To get W, a complete enumeration of all events is required
 - X can be used with random sampling
- Spatial dependence in point patterns can be explored through the observed distributions of W
- Two empirical cumulative probability distribution

$$\hat{G}(w) = \frac{\#(w_i \le w)}{n}; \qquad \qquad \hat{F}(x) = \frac{\#(x_i \le x)}{m}$$

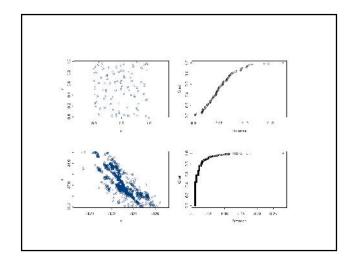
$$\hat{F}(x) = \frac{\#(x_i \le x)}{m}$$

n: the number of events;

m: the number of random points

Distance-based Exploratory Analysis

- Plot of *G(w)* against *w* or *F(x)* against *x*
 - If the empirical cumulative distribution function climbs very sharply in the early part before flattening out, then it indicates clustering
 - If it climbs very slowly at the beginning followed by steeply climbing, then it indicates a repulsion or regular pattern
- Plot of G(w) against F(x)
 - If no interaction, the two values should be very
- How to quantify these visual inspections?

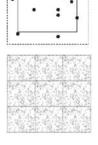


Distance-based Exploratory Analysis

- Edge effects
 - Guard area
 - Toroidal edge correction
 - Modified formula

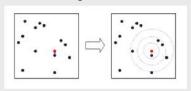
$$\hat{G}(w) = \frac{\#(b_i > w \ge w_i)}{\#(b_i > w)}$$

 b_i is the distance from event I to the nearest point on the boundary of R



The K Function

- Nearest neighbor distance analysis
 - Only consider nearest neighbor distances
- The K function
 - Reduced second order moment measure
 - Consider a range of distances



The K Function

- Need to be sure that it is valid to examine second order effects at which scale
- Need to make some assumption
 - Homogeneous or isotropic at the considered scale
 - Otherwise, it is not possible to estimate the second order effects directly from the observed patterns
 - Also, any second order effect can be due to the variation in the first order effect

Definition

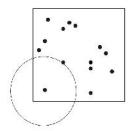
- •}K(h)=E(# of events within distance h of an arbitrary event)
- The expected # of events in n is } | n |
- The expected # of ordered pairs within distance h is ³ | M(h)
- The K can be estimated by

$$\tilde{K}(h) - \frac{1}{\lambda^2 \left|\Re\right|} \sum \sum_{i \neq i} I_k(d_y)$$

Edge Correction

$$\begin{split} \hat{K}(h) &= \frac{1}{\lambda^2 \left| \Re \right|} \sum_{i \neq j} \frac{I_k(d_g)}{w_g} \\ \hat{L}(h) &= \sqrt{\frac{\hat{K}(h)}{\pi}} - h \end{split}$$

 w_g is the proportion of circumference of circle centered on the *i*th event and passing through *j*th event



Estimation of Intensity

$$\begin{split} \lambda &= \frac{n}{|\Re|} \\ \hat{K}(h) &= \frac{|\Re|}{n^2} \sum_{i \neq j} \frac{I_h(d_{ij})}{w_{ij}} \end{split}$$

Modeling Spatial Point Patterns

- Exploratory analysis of spatial point patterns is rather informal and may not be sufficient
- Need more formal analysis
 - Hypothesis testing: comparing summary measures calculated from an observed point pattern with the expected observations under various hypothesized models
 - Statistical modeling: constructing specific models to explain observed patterns

Complete Spatial Randomness

- Complete Spatial Randomness (CSR)events follows a homogenous Poisson process over the study area
- Consider $\{Y(A), A \in \Re\}$
 - The probability distribution of Y(A) is a
 Poisson distribution with mean Y(A) , where Y
 is a constant

$$f_{Y(A)}(y) = \frac{(|A|)^y}{y!} e^{-|A|}$$

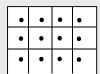
 $-Y(A_i),Y(A_j)$ are independent for any A_i and A_j

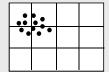
CSR

- CSR implies that conditional on n, events are independently and uniformly distributed over R
 - Any event has an equal probability of occurring at any location in R
 - The location of any event is independent of that of any other
 - CSR can be simulated by generating random events from uniform distribution over R
- CSR provides a baseline hypothesis for testing
 - Regular, clustered, or random

Quadrat count method

- Equally-spaced patterns will have most quadrats with similar counts
- Clustered patterns will have a few high count quadrats and many which are empty





Statistics of CSR

• The number of events that fall in any quadrat conforms to a binomial distribution

$$P(k \text{ events in quadrat}) = {n \choose k} \cdot \left(\frac{1}{m}\right)^k \left(1 - \frac{1}{m}\right)^{n-k}$$

where

k is the number of events in a quadrat, n is the total number of events, and m is the total number of quadrats

1/m is the fraction of the region occupied by a quadrat

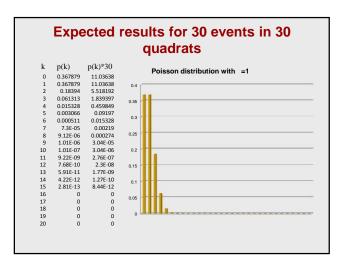
 $Exp(\text{the number of quadrats that contain } k \text{ events}) = P \times m$

Poisson distribution

• The binomial distribution is difficult to calculate because *n*! is often very large

• So in practice we use the *Poisson distribution* as an approximation:

$$P(k \text{ events in quadrat}) = \frac{\lambda^k e^{-\lambda}}{k!}$$
 $\} = \frac{n}{m}$



Quadrat Tests for CSR

- Divide the study region into m quadrats of equal size
- Find the mean number of points per quadrat $\bar{x} = \frac{1}{m} \sum_{i=1}^{m} x_i$
- Find the variance of the number of points per quadrat $s^2 = \frac{1}{m-1} \sum_{i=1}^{m} (xi \bar{x})^2$
- Calculate the variance-mean ratio (VMR) $_{VMR = \frac{S^2}{\pi}}$

Interpretation

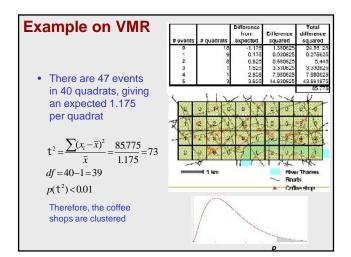
- If VMR 1 expected variation
 - Points approximate random distribution across the study region
- If VMR <1 less variation than expected
 - Points are spread out or uniform across the study region
- If VMR >1 more variation than expected
 - Point are more clustered distributed across the study region

Variance/mean Ratio (VMR) • Poisson distribution: mean = variance • variance/mean ratio (VMR) VMR<1 -> little variation (uniform) VMR=1 -> random VMR>1 -> good deal of variation (cluster) $\overline{x} = \frac{n}{m} = \frac{30}{30} = 1 \qquad n - \# \text{ of events} \\ m - \# \text{ of quadrats}$ $s^2 = \frac{\sum_{i=1}^{m} (x_i - \overline{x})^2}{m-1} \qquad x_i - \# \text{ of events in the } i^{th} \text{ quadrat}$ $= \frac{5 \times (0-1)^2 + 20 \times (1-1)^2 + 5 \times (2-1)^2}{30-1}$ $= 0.345 \qquad VMR = 0.345 < 1 \implies \text{ significantly uniform}$

Hypothesis Testing

- Hypotheses
 - H₀: Point pattern is random
 - H_A: Point pattern is not random (clustered, or regular)
- Statistical test
 - Chi-square test (df=m-1)
- Test statistic

$$\begin{split} \chi_{\text{tot}}^2 &= (m-1)VMR & \chi_{\text{tot}}^2 &< \chi_{\text{tot}}^2 < \chi_{\text{tot}}^2 < \chi_{\text{tot}}^2 = \text{prandom} \\ &- \frac{(m-1)s^2}{\overline{x}} - \frac{\sum_{i=1}^m (x_i - \overline{x})^2}{\overline{x}} & \chi_{\text{tot}}^2 < \chi_{\text{tot}}^2 < \chi_{\text{tot}}^2 = \text{prandom} \\ & \chi_{\text{tot}}^2 < \chi_{\text{tot}}^2 < \chi_{\text{tot}}^2 = \text{prandom} \end{split}$$



Indices

- Index of Dispersion $\frac{s^2}{\overline{x}}$
- Index of Cluster Size (ICS): $\frac{s^2}{\bar{x}}$ -1
 - Under CSR, E(ICS) =0
 - If ICS>0, the clustering is implied. ICS is the number of extra events
 - If ICS <0, the regularity is implied, ICS is the deficiency in events

Sampling

- Quadrat tests can be used in conjunction with the sampling of point patterns
 - Only need to count m quadrats which are randomly scattered over the study region
- Use sampled quadrats to estimate the intensity of the study region and its associated confidence interval

Estimated intensity 95% confide
$$\hat{\lambda} = \frac{\vec{x}}{|Q|}$$

$$\text{var}(\hat{\lambda}) = \frac{\hat{\lambda}}{m|Q|}$$

Limitation of Quadrat Tests

 The relative location of quadrats (if sampling is used) or the relative position of events within a quadrat is not take into account

Nearest Neighbor Tests

 Under CSR, events are independent and Y(A), the # of events in any area A is Poisson distributed

$$f_{Y(A)}(y) = \frac{(|A|)^y}{y!} e^{-|A|}$$

- Let x be the radius of a circle, the probability that no events fall within the circle is
- The distribution function F(x) of nearest neighbor point-event distances for CSR is $F(x) = P(X \le x) = 1 - e^{-1/fx^2}$

Nearest Neighbor Tests for CSR

Mean and variance of X

$$E(X) = \frac{1}{2\sqrt{f}}; Var(X) = \frac{4-f}{4f}$$

 Similarly, the distribution function G(w) of nearest neighbor event-event distances for CSR is

$$G(w) = P(W \le w) = 1 - e^{-fw^2}$$

$$E(W) = \frac{1}{2\sqrt{3}}; \quad Var(W) = \frac{4-f}{4f}$$

Nearest Neighbor Tests: Issues

- The theoretical distributions of X and W allow us to derive sampling distributions under CSR of summary statistics for observed nearest neighbor distances
 - The sampling distributions require independent sample point-event or event-event distances
 - Theoretical distributions do not consider edge effects
- Issues
 - Independence assumption may not be valid when large portions of events are used
 - Edge effects will introduce bias

- Tests are based on summary statistics of m randomly sampled nearest neighbor event-event distances (w₁,...,w_m) or point-event distances (x₁,...,x_m)
 - Sample size m can be chosen to be less then 0.1n
- Three common tests
 - Clark-Evans
 - Hopkins
 - Byth & Ripley

Clark-Evans

- Randomly sampled m nearest neighbor event-event distances (w₁,...,w_m)
- Calculate the sample mean $\overline{w} = \frac{\sum w_i}{m}$
- Compare with the sampling distribution of \overline{w} $N(\frac{1}{2\sqrt{1}}, \frac{4-f}{4)fm})$
 - Need to enumerate point pattern completely
 - Need to estimate intensity $\hat{y} = \frac{n}{|\Re|}$

Clark-Evens Test

- It is desirable to use all the nearest neighbor event-event distances (w₁,...,w_n)
- Correction to mean and variance

$$F(\overline{w}) = 0.5\sqrt{\frac{|R|}{n}} + 0.051\frac{P}{n} + 0.041\frac{P}{n^2}$$

$$Var(\overline{w}) = 0.07\frac{|R|}{n^2} - 0.037P\sqrt{\frac{|R|}{n^2}}$$

• Compare with the sampling distribution of $\overline{w} = \frac{\sum w_i}{m}$

 $N(E(\overline{w}), Var(\overline{w}))$

Hopkins

- Randomly sampled m nearest neighbor event-event distances (w₁,...,w_m) and point-event distances (x₁,...,x_m)
- Calculate the ratio of the two sample means \sum_{r}
- $H = \frac{\sum_{i=1}^{N_i} w_i^2}{\sum_{i=1}^{N_i} w_i^2}$ Compare with the sampling distribution

 $F_{2m,2m}$

of H

point-event distances (x₁,...,x_m)Calculate the statistic

$$BR = \frac{1}{m} \frac{\sum_{i} x_{i}^{2}}{\sum_{i} (x_{i}^{2} + w_{i}^{2})}$$

Byth & Ripley

• Randomly sampled m nearest neighbor

event-event distances (w₁,...,w_m) and

Compare with the sampling distribution of BR

$N(\frac{1}{2}, \frac{1}{12m})$

Nearest Neighbor Tests for CSR

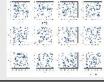
- The previous tests only utilize one single summary statistic
- With mapped point patterns, it is possible to compare the estimated distribution function \$\hat{\hat{\alpha}}(w)\$ or \$\hat{\epsilon}(x)\$ with their theoretic values \$\alpha(w)\$ or \$F(x)\$ under CSR
 - Need to make edge correction for estimated distribution functions
- However, it is hard to assess the significant of the deviation between the estimated from theoretic

Simulation

- Simulate m point patterns with n points under CSR
- Calculate Ĝ(w) for each point pattern without edge correction
- The theoretic value of G(w) without edge correction is estimated as

$$\overline{G}(w) = \sum \hat{G}_i(w)/m$$

• The observed G(w) $\overline{G}(w)$



Nearest Neighbor Tests: Simulation

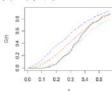
• Constructing the confidence envelopes

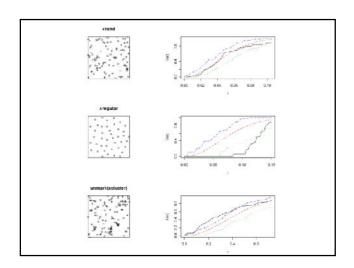
$$U(w) = \max_{l=1,\dots,m} \left\{ \hat{G}_l(w) \right\}$$

$$l(w) = \min_{l=1,\dots,m} \left\{ \hat{G}_l(w) \right\}$$

$$P(\hat{G}(w) > U(w)) = P(\hat{G}(w) < l(w)) = \frac{1}{m+1}$$

• Plot G(w), $\hat{G}(w)$, U(w), l(w)





K Function Tests for CSR

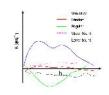
- Under CSR, the expected # of events within distance h of a randomly chosen event is)*fh²
- So under CSR, $K(h) = fh^2$
- Then the empirical K function $\frac{\tilde{K}(h) \frac{1}{\lambda^2 |\mathcal{H}|} \sum_{i,l} \sum_{i} I_{\lambda}(d_{\psi})}{\text{can be compared with } \kappa(h) = fh^2}$
- Or compared with the transformed form, L(h)=0 $\hat{L}(h)=\sqrt{\frac{\hat{K}(h)}{f}}-h$
- How to evaluate the significance?

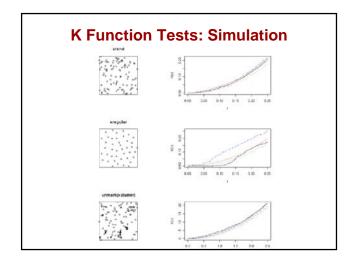
K Function Tests: Simulation

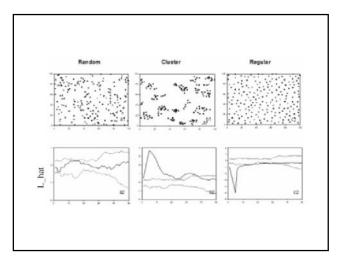
- Simulate m point patterns with n points under CSR
- Constructing the confidence envelopes $\frac{\mathcal{V}(b) = \max_{i \in \{L(b)\}} \{\hat{f}_i(b)\}}{|(b) \min_{i \in \{L(b)\}} \{L(b)\}}$

 $P(\hat{L}(h) > U(h)) = P(\hat{L}(h) < h(h)) = \frac{1}{m-1}$

• Plot $\hat{L}(h), U(h), L(h)$







Population Problem

- A serious difficulty with cluster detection is something we've been ignoring so far
 - The background population is not randomly distributed
 - In epidemiology the term used is the 'atrisk' population
- What other processes might we use?
 - Heterogeneous Poisson process
 - The Cox process