Logistic Regression with Polytomous & Ordinal Data

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A Look Back & A Look Ahead

- □ Last Week:
 - Binary Logistic Regression
 - □ DV = dichotomous
- □ This Week:
 - Polytomous Logistic Regression
 - □ aka Multinomial LR
 - □ DV = 3 or more nominal categories
 - Ordinal Logistic Regression
 - □ DV = 2 or more ordered categories



Review: Coin Flip Example



- □ Probability coin will be heads:
 - **50%**
- □ Odds coin will be heads:
 - probability coin will be heads = 50 = 1
 probability coin will be tails 50
- □ Logit coin will be heads:
 - Ln (odds) = $\ln (50/50) = 0$

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LR with Polytomous Data (aka Multinomial LR)

Running Example #1:

- □ Research Question: Do beliefs about impact of working mothers on children aid in predicting a women's work status?
- □ Data: GSS (http://sda.berkeley.edu/index.htm)
 - Only data from females used
- □ Independent Variable (1-4 scale):
 - A working mother can establish just as warm and secure a relationship with her children as a mother who does not work. (fechld)
 - □ Strongly Agree
 - □ Agree
 - □ Disagree
 - Strongly Disagree

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Running Example #1 (continued)

Binary LR:

- □ DV:
 - Working (interest group)
 - Not Working (reference group) (i.e., Unemployed)

Polytomous LR:

- □ DV:
 - Working (interest group)
 - Retired (interest group)
 - Student (interest group)
 - Not Working (reference group)

Why Not Run 4 Binary LRs???



- □ Could run 3 Binary LRs:
 - Working vs Not Working
 - Retired vs Not Working
 - Student vs Not Working
- □ Problem: Lose a lot of info!
 - Example: if we use Student vs Non-Student, we may NOT get a statistically significant effect for one IV (e.g., gender) if males:
 - □ Are <u>less</u> likely to be Unemployed
 - □ And more likely to be Working

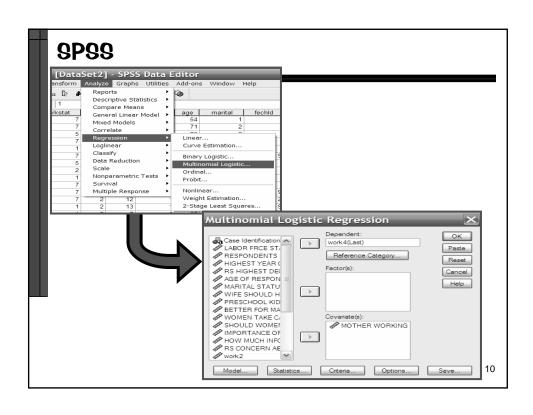
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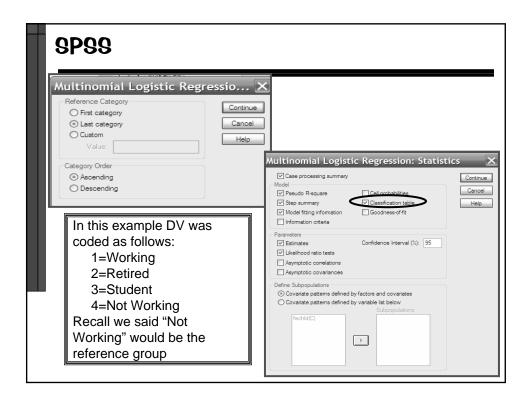
Polytomous LR

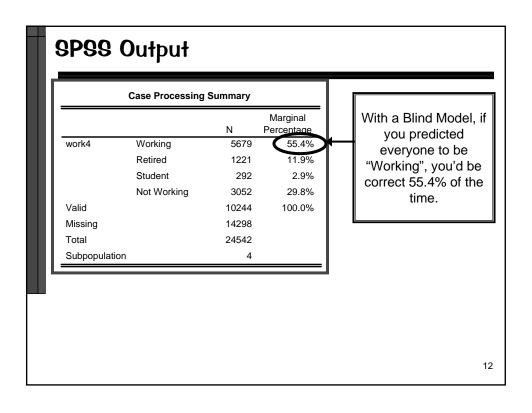
- □ Odds = interest group/reference group
- □ Need k-1 regression equations (i.e., one for each interest group):
 - Logit (working) = $a_w + b_w X$
 - Logit (retired) = $a_r + b_r X$
 - Logit (student) = $a_s + b_s X$
- Classify/predict people into one of 4 categories

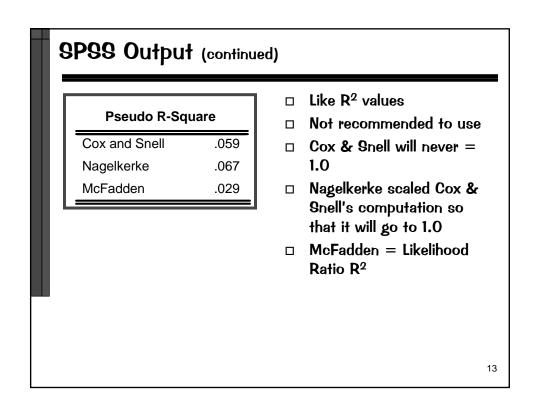
Note: For ease of reading the notes, for this lecture $exp[a + bX] = e^{a + bX}$

g









SPSS Output

The 2 models are nested, therefore: $\chi^2_{\text{difference}} = 744.048 - 123.544 = 620.536$ with 3 df, p <.001

Therefore, the Full model (with <u>fechld</u> as a predictor) is significantly better than the 'blind' model.

	Model Fi	tting Informati	on	
	Model Fitting Criteria	Likeliho	ood Ratio Te	ests
Model	-2 Log Likelihood	Chi-Square	df	Sig.
Intercept Only	744.048			
Final	123.544	620.504	3	.000

1 df for each regression equation. We have 4 groups, and k-1 regression equations. 4 – 1 = 3 df

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Full/Final Model vs Reduced Model

- □ Final Model includes the predictor:
 - 'Fechld'
- □ Reduced Model removes the predictor
 - Ex: 'Fechld' Reduced Model contains all predictors except 'Fechld'
 - □ In this case: Reduced model = Blind model, because only 1 predictor
 - □ ∴ redundant with "Model Fitting Information" table in SPSS
 - Likelihood Ratio Test:
 - $\quad \Box \quad \chi \textbf{2} \, = \, \textbf{-2LL}_{\text{REDUCED}} \, \textbf{-2LL}_{\text{FINAL}}$
 - allows you to test contribution of 'Fechld'

SPSS Output (continued)

The 2 models are nested, therefore:

 $\chi^2_{difference}$ = 744.048 - 123.544 = <u>620.536</u> with 3 df, p <.001

Therefore, the Full model (with <u>fechld</u> as a predictor) is significantly better than the 'blind' model.

Likelihood Ratio Tests

Model
Fitting
Criteria Likelihood Ratio Tests

-2 Log
Likelihood
of Reduced
Model Chi Square of Signature

 Effect
 Model Model
 Chi-Square
 df
 Sig.

 Intercept
 2037.841
 1914.297
 3
 .000

 fechld
 744.048
 620.504
 3
 .000

The chi-square statistic is the difference in -2 log-likelihoods between the final model and a reduced model. The reduced model is formed by omitting an effect from the final model. The null hypothesis is that all parameters of that effect are 0.

Likelihood Ratio Test: $\chi 2 = -2LL_{REDUCED} - 2LL_{FINAL}$

Tests if variable "fechld" contributes statistically significantly to the model.

(a little redundant w/ the 1 predictor model)

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SPSS Output (continued)

				Parame	ter Estii	nates			
	Vorking Intercept 1.821 .062 872.903 1 .000 fechId 566 .026 461.876 1 .000 .568 .539 .598 Retired Intercept 976 .094 108.062 1 .000 fechId .025 .037 .465 1 .495 1.026 .954 1.103 Student Intercept -1.065 .156 46.726 1 .000								
work4 ^a		В		Wald	df	Sig.	Exp(B)	Lower Bound	Upper Bound
Working	Intercept	1.821	.062	872.903	1	.000			
	fechld	566	.026	461.876	1	.000	.568	.539	.598
Retired	Intercept	976	.094	108.062	1	.000			
	fechld	.025	.037	.465	1	.495	1.026	.954	1.103
Student	Intercept	-1.065	.156	46.726	1	.000			
	fechld	609	.075	66.684	1	.000	.544	.470	.630

a. The reference category is: Not Working.

 \Box Logit_w = 1.821 - .566(fechld)

 \Box Logit_r = -0.976 + .025(fechld)

 \Box Logit_s = -1.065 - .609(fechld)

SPSS Output: Parameter Estimates (continued)

- Working: If IV changed 1 unit, the DV-odds would change by a multiplicative amount of .568
- Odds of being Student to Not Working are much lower than Working to Not Working
- □ Look at predictors for each regression equation separately!
 - Belief about working mothers is not a significant predictor of whether or not a person is retired or Not Working

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SPSS Output (continued) Classification Predicted Percent Retired Student Not Working Observed Working Correct Working 5415 264 95.4% Retired 1096 0 0 125 .0% Student 281 0 0 .0% 11 Not Working 2700 0 0 352 11.5% Overall Percentage 92.7% .0% .0% 7.3% 56.3%

- Only classifying 56.3% of women correctly
- Not classifying anyone as Retired or Student
 - •Hmmm, small percentage of respondents in these categories could \rightarrow problems
- Blind model predicted women work status correctly 55.4%

Classification Rule: Binary LR

- □ 2 equations:
 - $P(\text{Working}) = \frac{\exp[a + bX]}{1 + \exp[a + bX]}$
 - $P(\text{Not-Working}) = \frac{1}{1 + \exp[a + bX]}$

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Classification Rule: Polytomous LR

Recall:

Odds = odds of being in a particular group sum of odds of being in all groups

- □ (Relative to reference group)
- □ 4 equations:

$$P(\text{Group } j) = \frac{\exp[a_J + b_J X]}{1 + \sum \exp[a_J + b_J X]}$$

Classification Rule: Polytomous LR

□4 equations:

$$P(\text{Working}) = \frac{\exp[a_W + b_W X]}{1 + \exp[a_W + b_W X] + \exp[a_R + b_R X] + \exp[a_S + b_S X]}$$

$$\blacksquare P(Retired) = \frac{\exp[a_R + b_R X]}{1 + \exp[a_W + b_W X] + \exp[a_R + b_R X] + \exp[a_S + b_S X]}$$

$$\blacksquare P(\text{Student}) = \frac{\exp[a_S + b_S X]}{1 + \exp[a_W + b_W X] + \exp[a_R + b_R X] + \exp[a_S + b_S X]}$$

$$P(\text{Not-Working}) = \frac{1}{1 + \exp[a_W + b_W X] + \exp[a_R + b_R X] + \exp[a_S + b_S X]}$$

Regression Equations:

- ☐ If FechId = 4 (i.e., believe mother working negatively impacts children) then
 - $Logit_{working} = 1.821 .566(4) = -.443$
 - Logit_{retired} = -0.976 + .025(4) = -.876
 - $Logit_{student} = -1.065 .609(4) = -3.501$
 - exp(Logit_{working}) = exp(-.443) = .642
 odds(Working to Not Working)
 - exp(Logit_{retired}) = exp(-.876) = .416
 odds(Retired to Not Working)
 - exp(Logit_{student}) = exp(-3.501) = .030
 odds(9tudent to Not Working)

Regression Equations:

□ If Fechld = 4 (i.e., believe mother working negatively impacts children) then

■
$$P(\text{Working}) = \frac{.642}{1 + .642 + .416 + .030} = .31$$

■
$$P(\text{Retired}) = \frac{.416}{1 + .642 + .416 + .030} = .20$$

$$P(Student) = \frac{.030}{1 + .642 + .416 + .030} = .01$$

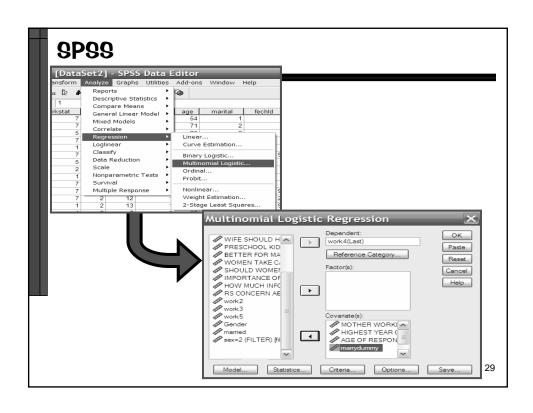
■
$$P(\text{Not Working}) = \frac{1}{1 + .642 + .416 + .030} = .48$$

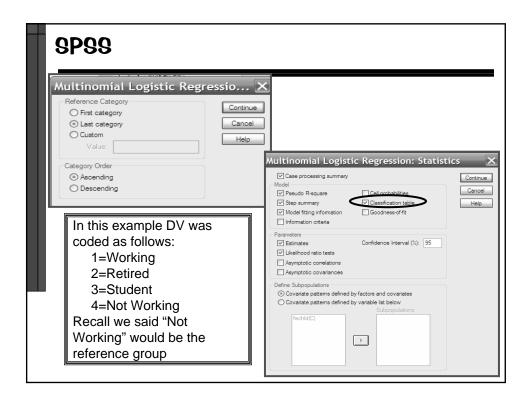
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Running Example #2

Polytomous LR w/ Multiple Predictors:

- □ DV:
 - Working (interest group)
 - Retired (interest group)
 - Student (interest group)
 - Not Working (reference group)
- □ IV's:
 - Fechld Belief about working mothers
 - Years of Education
 - Age
 - Marital Status (dummy coded 1=married)





SPSS Output

	Case Processing	Summary	
		N	Marginal Percentage
work4	Working	5645	55.5%
	Retired	1209	11.9%
	Student	290	2.8%
	Not Working	3035	29.8%
Valid		10179	100.0%
Missing		14363	
Total		24542	
Subpopulation		3734 ^a	
a. The dens	ndent variable has	only one va	الم

a. The dependent variable has only one value observed in 2500 (67.0%) subpopulations.

Pseudo R-Squ	uare
Cox and Snell	.423
Nagelkerke	.483
McFadden	.264

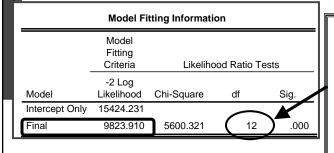
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SPSS Output

The 2 models are nested, therefore:

 $\chi^2_{difference}$ = 15424.231 - 9823.910 = **5600.321** with 12 df, p <.001

Therefore, the Full model (with <u>fechld</u>, <u>education</u>, <u>age</u>, and <u>marital</u> <u>status</u> as a predictors) is significantly better than the 'blind' model.



- 1 df for each predictor in each regression equation. We have 4 predictors, 4 groups, and k-1 regression equations.
- 4 1 = 3 regression equations * 4 predictors = 12

Full/Final Model vs Reduced Model

- □ Final Model has all 4 predictors:
 - 'Fechld'
 - Education
 - Age
 - Marital Status
- □ Reduced Model removes 1 of the predictors
 - Ex: 'Fechld' Reduced Model contains all predictors except 'Fechld'
 - Likelihood Ratio Test:
 - $\Box \quad \chi \mathbf{2} = -2\mathsf{LL}_{\mathsf{REDUCED}} 2\mathsf{LL}_{\mathsf{FINAL}}$
 - allows you to test contribution of 'Fechld'

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SPSS Output (continued)

Does the variable contribute statistically significantly to the model?

Likelihood Ratio Test = 10012.569 - 9823.910 = 188.659

	Likeliho	ood Ratio Tests	5	
	Model Fitting Criteria	Likeliho	od Ratio Tests	5
Effect	-2 Log Likelihood of Reduced Model	Chi Savara	ale.	Ci
	11079.234	Chi-Square 1255.323	df 3	Sig. .000
Intercept	11079.234	1200.020		
fechld	10012.569	188.659	3	.000
educ	10420.064	596.154	3	.000
age	13443.256	3619.346	3	.000
marrydummy	10031.433	207.523	3	.000
The chi-square	statistic is the	e difference in -2	2 log-likelihood	s

The chi-square statistic is the difference in -2 log-likelihoods between the final model and a reduced model. The reduced model is formed by omitting an effect from the final model. The null hypothesis is that all parameters of that effect are 0.

The Final model fits significantly better than a model WITHOUT 'fechld' in it (i.e., 'fechld' significantly contributes to the model).

All 4 predictors significantly contribute to the model

SPSS Output (continued)

			Std.						
work4 ^a		В	Error	Wald	df	Sig.	Exp(B)	Lower Bound	Upper Bound
Working	Intercept	850	.186	20.928	1	.000			
work4 ^a B Error Wald df Sig. Exp(B) Lower Bound Upper Bound Working Intercept 850 .186 20.928 1 .000 .687 .650 .72 educ .221 .010 492.222 1 .000 1.247 1.223 1.27 age 031 .002 336.890 1 .000 .970 .967 .97 marrydummy .563 .050 124.692 1 .000 1.755 1.590 1.93 Retired Intercept -10.897 .379 828.497 1 .000 .734 .667 .80 educ .161 .014 126.106 1 .000 .734 .667 .80 educ .161 .014 126.106 1 .000 1.175 1.142 1.20 age .125 .004 957.222 1 .000 1.133 1.124 1.14	.726								
	age 031 .002 336.890 1 .000 .970 .967 .973 marrydummy .563 .050 124.692 1 .000 1.755 1.590 1.938 ired Intercept -10.897 .379 828.497 1 .000								
	age	y .563 .050 124.692 1 .000 1.755 1.590 1.938							
	marrydummy	.563	.050	124.692	1	.000	1.755	1.590	1.938
Retired	Intercept	-10.897	.379	828.497	1	.000			
	fechld	309	.049	39.502	1	.000	.734	.667	.809
	.014	126.106	1	.000	1.175	1.142	1.209		
	age	.125	.004	957.222	1	.000	1.133	1.124	1.142
	marrydummy	.661	.090	54.274	1	.000	1.937	1.625	2.310
Student	Intercept	-2.625	.556	22.333	1	.000			
	fechld	339	.081	17.451	1	.000	.712	.608	.835
	educ	.275	.028	97.125	1	.000	1.317	1.246	1.391
	age	141	.009	263.167	1	.000	.868	.853	.883
	marrydummy	1.521	.157	93.666	1	.000	4.575	3.362	6.224

Parameter Estimates → Regression Equations

- Logit_w = -.850 .375(fechld)+.221(educ)-.031(age) +.563(marital)
- \Box Logit_R = -10.897 .309(fechld)+.161(educ)+.125(age) +.661(marital)
- Logit_S = -2.625 .339(fechld) + .275(educ) .141(age) + 1.521(marital)

☐ Working:

- If 'fechld' changed 1 unit, the log-odds of the DV would change by a multiplicative amount of .687, all else being held constant
- If 'educ' changed 1 unit, the log-odds of the DV would change by a multiplicative amount of 1.247, all else being held constant

SPSS Output (continued)

		Classificat	ion		
			Predicted		
Observed	Working	Retired	Student	Not Working	Percent Correct
Working	4975	68	0	602	88.1%
Retired	175	718) 0	316	59.4%
Student	284	\bigcup_{0}	$\overline{}$	6	.0%
Not Working	1690	411	0	934	30.8%
Overall Percentage	70.0%	11.8%	.0%	18.3%	65.1%

- Now classifying 65.1% of women correctly
- Not classifying anyone as Student
- Blind model predicted women work status correctly 55.4%
- 1 predictor model predicted women correctly 56.3%

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LR with Ordinal Data

Measurement Scales

□ Nominal:

■ Categories; no order

■ Ex: Goal/No Goal

□ Ordinal:

■ Rank order but cannot measure "by how much"

■ Ex: Degree (H9, Associate, Bachelor, Master, Doctorate)

□ Interval:

■ Rank order & equidistance

■ Ex: Years of Education

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My DV is Ordinal; What Should I Do??

- □ Pretend variable is on interval scale → OLS
- □ Treat as though it were *measured* on ordinal scale but is *really* interval/ratio underneath \rightarrow WLS
- □ **Treat as though it was measured on a true ordinal scale → Cumulative logit model
- □ Treat it as nominal → Polytomous LR

Slide Adapted from Roy Levy

Running Example #3:

- □ Research Question: Does father's education (faeduc) aid in predicting a person's highest degree obtained?
- □ Data: GSS (http://sda.berkeley.edu/index.htm)
- □ DV = Highest Degree Obtained:
 - Dropout (interest group)
 - High School (interest group)
 - Some College (interest group)
 - Bachelor (interest group)
 - Graduate (interest group)
- □ Independent Variable = <u>years</u> of father's education

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Now Predicting Cumulative Probabilities:

Drop	vs.	All other classes (HS, Some, Bachelor, Graduate)
Drop & H9	vs.	Some, Bachelor, Graduate
Drop, H9, & Some	vs.	Bachelor & Graduate
Drop, H9, Some, Bachelor	vs.	Graduate

 \square Need k-1 regression equations (5 – 1 = 4)

"Cumulative" Logits

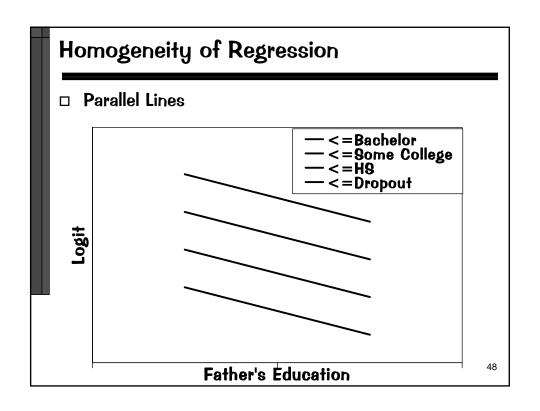
- □ Binary:
 - Logit expressions = probabilities of being in a given group
- □ Polytomous:
 - Logit expressions = probability of being in a given group compared to probability of being in reference group
- □ Ordinal:
 - Logit expressions = probabilities of being *in or below* a given group
 - □ i.e., "cumulative" logits
 - Logit (≤Drop) = logit for being in or below Dropout
 - Logit (≤HS) = logit for being in or below HS
 - Logit (\leq 9C) = logit for being *in or below* 9C
 - Logit (≤Bachelor) = logit for being in or below Bachelor⁴⁶

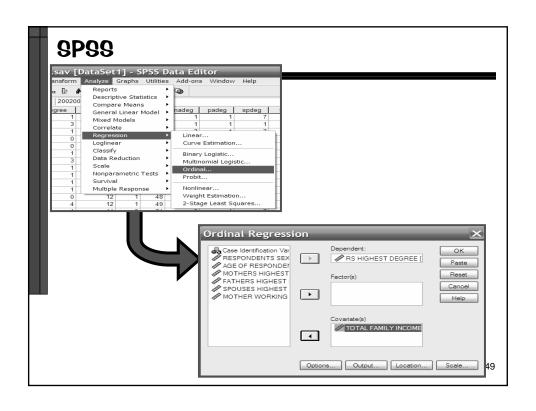
Homogeneity of Regression

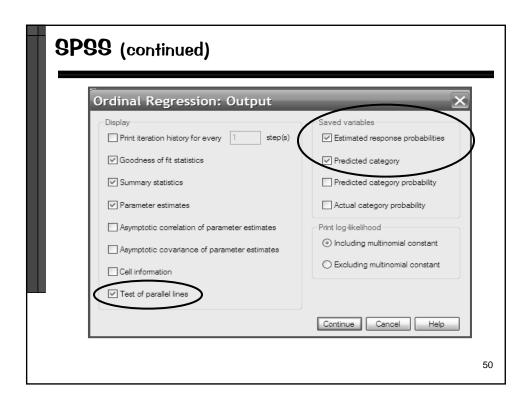
We know that:

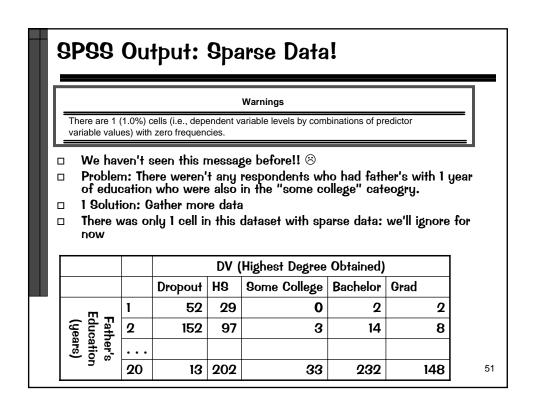
□ Logit(\leq Drop) < Logit (\leq H9) < Logit (\leq 9C) < Logit (\leq Bachelor)

- □ Logit (≤HS) = a_H b_1 X_1 b_2 X_2 ... b_P X_P □ Logit (≤SC) = a_C - b_1 X_1 - b_2 X_2 - ... - b_R X_P
- $\Box \quad \text{Logif ($\leq$ Bachelor)} = \begin{bmatrix} a_0 \\ a_B \end{bmatrix} \begin{bmatrix} b_1 \\ b_1 \end{bmatrix} X_1 \begin{bmatrix} b_2 \\ b_2 \end{bmatrix} X_2 \dots \begin{bmatrix} b_p \\ b_p \end{bmatrix} X_p$
- \Box Slopes (b₁, b₂ . . . b_p) are the SAME for each equation
- \Box Only the intercepts (a_D, a_H, a_C, a_B) differ









SPSS Output

	Case Processing Summary								
		N	Marginal Percentage						
RS HIGHEST	DROPOUT	5733	18.3%						
DEGREE	HIGH SCHOOL	16872	53.9%						
	SOME COLLEGE	1610	5.1%						
	BACHELOR	4842	15.5%						
	GRADUATE	2241	7.2%						
Valid		31298	100.0%						
Missing		12400							
Total		43698							

Pseudo R-Squa	are
Cox and Snell	.195
Nagelkerke	.211
McFadden	.085
Link function: Logit.	

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SPSS Output (continued)

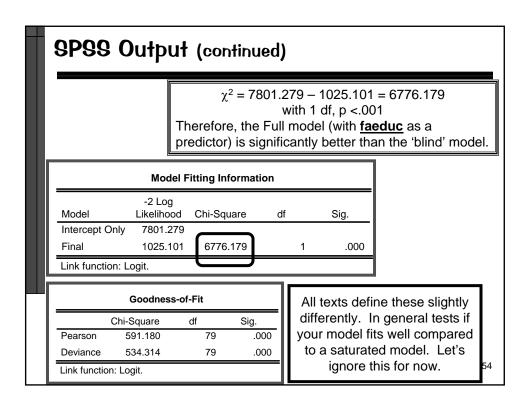
	Test of	Parallel Lines	ı	
Model	-2 Log Likelihood	Chi-Square	df	Sig.
Null Hypothesis	1025.101			
General	810.696	214.404	3	.000

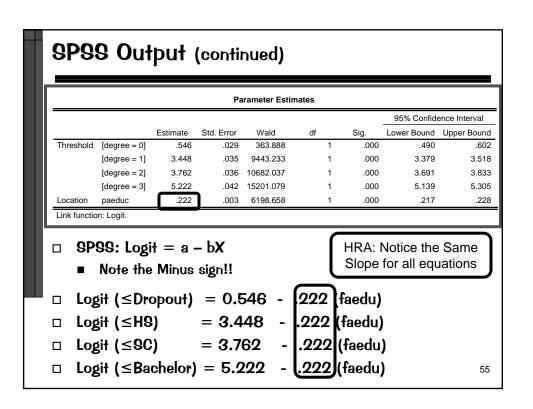
The null hypothesis states that the location parameters (slope coefficients) are the same across response categories.

a. Link function: Logit.

□ Tests Homogeneity of Regression Assumption

- lacktriangle Reject lacktriangle The slopes are different across logits lacktriangle
- Often anti-conservative i.e., reports p-values that are too low with large samples or continuous predictors
- We'll continue anyway for illustration purposes





'rol	babil	lities	ě čr (Cate	gorų	J Pr	rea	lict	ors	
∄*GSS	Degree Da	itaset.sav	[DataSet2	1 - SPSS D	ata Editor					_
	View Data			-						
	H + +									
			# III III III	: III		_				
1 : CASEI	_	20020001								
00050	EST1_1	EST2_1	EST3_1	EST4_1	EST5_1	PRE_1				
22953 22954	.11	.58	.06	.18	.07	1				
22954	.23	.62	.04	.09	.03	1				
22956	.13	.60	.04	.15	.06	1				
22957	.13	.47	.00	.13	.00	1				
22958	.41	/	.02	.00	.01					
22959	.31	.58	.03	.06	.02	1				
22960	.31	.58	.03	.06	.02	1				
22961	.31	.58	.03	.06	.02	- 1				
22962	.23	.62	.04	.09	.03	1				
22963	.11	.58	.06	.18	.07	- 1				
22964		-	-		-					
22965	.53	.43	.01	.03	.01	0				
22966	.53	.43	.01	.03	.01	0				
22967	.11	.58	.06	.18	.07	1				
22968	. 19	.62	.04	.11	.04	1				
22969	.11	.58	.06	.18	.07	1				
22970	.11	.58	.06	.18	.07	1				
22971	.13	.60	.06	.15	.06	1				
22972	.19	.62	.04	.11	.04	1				
22973	.23	.62	.04	.13	.03	1				
22974	.10	.02	.05	.13	.05					
22976	.13	.60	.06	.15	.06	1				
22977	.13	.62	.04	.09	.03	1				
22077	.23	.02	.04	.03	.00					

What are our Cut-Points?							
<u>years</u> faeduc	p(dropout)	p(hs)	p(sc)	p(bachelor)	p(grad)		
0	0.63	0.34	0.01	0.02	0.01		
1	0.58	0.38	0.01	0.02	0.01		
2	0.53	0.43	0.01	0.03	0.01		
3	0.469	0.472	0.02	0.03	0.01		
4	0.42	0.51	0.02	0.04	0.01		
5	0.36	0.55	0.02	0.05	0.02		
6	0.31	0.58	0.03	0.06	0.02		
7	0.27	0.60	0.03	0.07	0.02		
8	0.23	0.62	0.04	0.09	0.03		
9	0.19	0.62	0.04	0.11	0.04		
10	0.16	0.62	0.05	0.13	0.05 60		

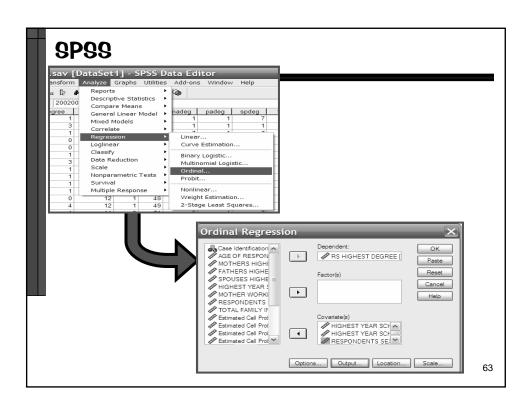
What are our Cut-Points?

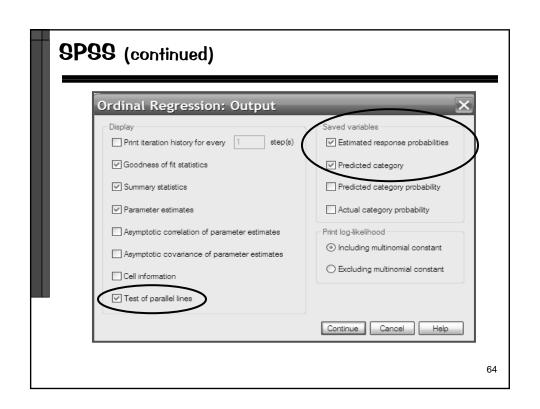
<u>years</u> faeduc	p(dropout)	p(hs)	p(sc)	p(bachelor)	p(grad)
11	0.13	0.60	0.06	0.15	0.06
12	0.11	0.58	0.06	0.18	0.07
13	0.09	0.55	0.07	0.21	0.09
14	0.07	0.51	0.07	0.23	0.11
15	0.06	0.47	0.08	0.26	0.13
16	0.05	0.43	0.08	0.29	0.16
17	0.04	0.38	0.08	0.31	0.19
18	0.03	0.335	0.07	0.332	0.22
19	0.02	0.29	0.07	0.34	0.27
20	0.02	0.25	0.07	0.35	0.32
					6

Running Example #4

Ordinal LR w/ Multiple Predictors:

- □ DV = Highest Degree Obtained:
 - Dropout (interest group)
 - High School (interest group)
 - Some College (interest group)
 - Bachelor (interest group)
 - Graduate (interest group)
- □ IV's:
 - Father's Education in years (faeduc)
 - Mother's Education in years (maeduc)
 - Gender





SPSS Output

Warnings

There are 1316 (38.4%) cells (i.e., dependent variable levels by combinations of predictor variable values) with zero frequencies.

Case Processing Summary						
		N	Marginal Percentage			
RS HIGHEST	DROPOUT	4932	16.9%			
DEGREE	HIGH SCHOOL	15787	54.2%			
	SOME COLLEGE	1540	5.3%			
	BACHELOR	4695	16.1%			
	GRADUATE	2181	7.5%			
Valid		29135	100.0%			
Missing		14563				
Total		43698				

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SPSS Output (continued)

Model Fitting Information							
Model	-2 Log Likelihood	Chi-Square	df	Sig.			
Intercept Only	15239.670						
Final	7685.279	7554.391	3	.000			
Link function: Logit.							

Goodness-of-Fit								
Chi-Square df Sig.								
Pearson	4569.821	2737	.000					
Deviance 3543.083 2737 .000								
Link function: Logit.								

Pseudo R-Square				
Cox and Snell	.228			
Nagelkerke	.248			
McFadden	.102			
Link function: Logit.				

SPSS Output (continued)

Parameter Estimates								
-	95% Confidence Interval							ence Interval
		Estimate	Std. Error	Wald	df	Sig.	Lower Bound	Upper Bound
Threshold	[degree = 0]	1.015	.052	379.292	1	.000	.913	1.117
	[degree = 1]	4.068	.058	4943.722	1	.000	3.955	4.181
	[degree = 2]	4.388	.059	5624.599	1	.000	4.273	4.503
	[degree = 3]	5.873	.063	8747.770	1	.000	5.750	5.996
Location	paeduc	.139	.004	1403.286	1	.000	.132	.146
	maeduc	.158	.004	1277.636	1	.000	.149	.167
	sex	127	.023	30.564	1	.000	172	082
Link function: Logit.								

- □ Hmmm, No likelihood-ratio tests 🕾
- □ Need to compute them yourself

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SPSS Output (continued)

Test of Parallel Lines ^a						
Model	-2 Log Likelihood	Chi-Square	df	Sig.		
Null Hypothesis	7685.279					
General	7286.928	398.351	9	.000		

The null hypothesis states that the location parameters (slope coefficients) are the same across response categories.

a. Link function: Logit.