

STAT 675 Homework # 2

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5.1 Compute a Monte Carlo estimate of $\int_0^{\pi/3} \sin t dt$

```
set.seed(36)
m <- 10000
x <- runif(m, min = 0, max = pi/3)
theta.hat <- mean(sin(x)) * (pi/3)
theta <- cos(0) - cos(pi/3) #could just set this to .5 but shwowing the calc to be clear
theta.hat - theta

## [1] 0.001851

rm(theta.hat, theta, x)
```

Not impressively accurate with 10,000 samples.

5.3 Monte Carlo estimate of $\int_0^{0.5} e^{-x} dx$ with a uniform and exponential. Since the integrand is a pdf we can estimate directly by taking sample, Y , from this pdf and applying the indicator function $I(Y \leq x)$.

```
x <- runif(m, min = 0, max = 0.5)
y <- rexp(m) <= 0.5
theta.hat <- 0.5 * mean(exp(-x))
theta.star <- mean(y)
var.theta.hat <- (0.25/m) * var(exp(-x))
var.theta.star <- (theta.star * (1 - theta.star))/m
r <- matrix(c(theta.hat, theta.star, var.theta.hat, var.theta.star), 2, 2)
rownames(r) <- c("theta.hat", "theta.star")
colnames(r) <- c("Theta", "Var.Theta")
tbl1 <- xtable(r, caption = "The estimates and estimated variances from two estimates of theta.")
digits(tbl1) <- 10
print(tbl1)
```

	Theta	Var.Theta
theta.hat	0.3940505503	0.0000003238
theta.star	0.3926000000	0.0000238465

Table 1: The estimates and estimated variances from two estimates of theta.

```
rm(x, y, theta.hat, theta.star, var.theta.hat, var.theta.star, r, tbl1)
```

Since the integrand is close to .5 the variance of the binomial is maximized and the indicator estimate has a large variance.

5.4 Function for cdf of Beta(3,3) = $\int_0^y 30x^2(1-x)^2 dx$.

```
betacdf <- function(x, m) {
  theta <- rep(0, length(x))
  for (i in 1:length(x)) {
    u <- rbeta(m, 3, 3) <= x[i]
    theta[i] <- mean(u) # (30*u^2*(1-u)^2)*x
  }
  return(theta)
}
x <- seq(0, 1, 0.1)
r <- matrix(c(betacdf(x = x, m = 1e+06), pbeta(x, 3, 3)), 2, 11, byrow = TRUE)
colnames(r) <- x
rownames(r) <- c("My function", "R function")
tbl <- xtable(r, caption = "The estimates from a simple mote-carlo estimate and the R function.")
digits(tbl) <- 4
print(tbl)
```

	0	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9	1
My function	0.0000	0.0085	0.0579	0.1622	0.3179	0.4994	0.6833	0.8367	0.9423	0.9914	1.0000
R function	0.0000	0.0086	0.0579	0.1631	0.3174	0.5000	0.6826	0.8369	0.9421	0.9914	1.0000

Table 2: The estimates from a simple mote-carlo estimate and the R function.

```
rm(x, r, tbl)
```

My function does a pretty good job.

5.6 Improvement in antithetic estimation of $\int_0^1 e^x dx$.

$$\begin{aligned} Var\left(\frac{e^u + e^{1-u}}{2}\right) &= \frac{1}{4} [Var(e^u) + Var(e^{1-u}) + 2Cov(e^u, e^{1-u})] \\ &= \frac{1}{4} [2Var(e^u) + 2(Ee^u e^{1-u} - Ee^u Ee^{1-u})] \\ &\quad \text{Note : } Ee^u e^{1-u} = Ee = e \\ \text{Also : } Ee^{1-u} &= \int_0^1 e^{1-u} du = -e^{1-u} \Big|_0^1 = -e^0 + e^1 = e - 1 = Ee^u \end{aligned}$$

Therefore:

$$\begin{aligned} Var\left(\frac{e^u + e^{1-u}}{2}\right) &= \frac{1}{4} [2Var(e^u) + 2(e - (e - 1)^2)] \\ &= \frac{1}{2} \left[\left(\frac{e^2 - 1}{2} - (e - 1)^2\right) + (e - (e - 1)^2) \right] \end{aligned}$$

This is computed below

```
print(v.anti <- 0.5 * (((exp(1)^2 - 1)/2) - (exp(1) - 1)^2 + (exp(1) - ((exp(1) - 1)^2))))
```

```
## [1] 0.003912
```

From the book the variance of the simple Monte Carlo estimate is 0.1210178 so the variance reduction is:

```
100 * ((0.1210178 - v.anti)/0.1210178)
```

```
## [1] 96.77
```

```
5.7 anti <- function(m) {
  #function to comput the antithetic estimate
  u <- runif(m/2)
  v <- 1 - u
  theta <- mean((exp(u) + exp(v))/2)
  v <- var((exp(u) + exp(v))/2)
  return(c(theta, v))
}
theta.ant <- anti(m = m)
simple <- function(m) {
  #function to comput the simple estimate
  u <- runif(m/2)
  v <- runif(m/2)
  theta <- mean((exp(u) + exp(v))/2)
  v <- var((exp(u) + exp(v))/2)
  return(c(theta, v))
}
theta.simp <- simple(m = m)
theta.ant[1] #estimate

## [1] 1.718

theta.simp[1]

## [1] 1.719

theta.ant[2]

## [1] 0.00394

theta.simp[2]

## [1] 0.1202

reduction <- 100 * ((theta.simp[2] - theta.ant[2])/theta.simp[2]) #reduction in the variance of the
reduction

## [1] 96.72
```

5.8 Show that $\rho(X, X') = -1$.

$$\begin{aligned}
 \rho(X, X') &= \frac{\text{Cov}(X, X')}{\sqrt{\text{Var}(X)\text{Var}(X')}} = \frac{\text{Cov}(aU, a(1-U))}{\sqrt{\text{Var}(aU)\text{Var}(a(1-U))}} \\
 &= \frac{E[aU(a(1-U))] - E(aU)E(a(1-U))}{\sqrt{a^2\text{Var}(U)a^2\text{Var}(1-U)}} = \frac{a^2E(U - U^2) - aEU(a - aEU)}{a^2\text{Var}(U)} \\
 &= \frac{a^2EU - a^2EU^2 - [a^2EU - a^2(EU)^2]}{a^2\text{Var}(U)} = \frac{EU - EU^2 - EU - (EU)^2}{\text{Var}(U)} = \frac{-EU^2 + (EU)^2}{\text{Var}(U)} \\
 &= \frac{-1(\text{Var}(U))}{\text{Var}(U)} = -1
 \end{aligned}$$

The result would hold for any random variable with existing first and second moments.

5.14 Obtain a Monte Carlo estimate of $\int_1^\infty \frac{x^2}{\sqrt{2\pi}} e^{-x^2/2} dx$ by importance sampling.

I will use the pareto as my weighting distribution because it is the only distribution I could find in the back of Casella and Berger that could have the same support.

```

require(VGAM)
p <- rpareto(m, 1, 1)
g <- function(x) ((x^2)/sqrt(2 * pi)) * exp(-x^2/2)
fg <- g(p)/dpareto(p, 1, 1)
theta.hat <- mean(fg)
theta.hat

## [1] 0.4036

```