

Ch 8. Hypothesis testing: Part I

Motivating example

- Suppose that we have a coin and we want to test whether or not this coin is fair. To get sample, we flip a coin six times and get results.
 - ▶ Suppose we have three Heads and three Tails. Can we say a given coin is fair?
 - ▶ Suppose we have six Heads. Can we say a given coin is fair?

Terminologies

- ▶ Hypothesis: A statement about parameter
- ▶ Null hypothesis (H_0): A hypothesis that the parameter takes a particular value.
- ▶ Alternative hypothesis (H_1): A hypothesis that the parameter falls in some alternative range of values.
- ▶ Type I error: H_0 is rejected when H_0 is true
- ▶ Type II error: H_0 is not rejected when H_0 is false
- ▶ Significance level (α): Probability that Type I error occurs
- ▶ β : Probability that Type II error occurs
- ▶ Power($1 - \beta$): Probability that H_0 is rejected when H_0 is false
- ▶ Critical region(\mathcal{C}): A subset of sample space that results in rejection of H_0
- ▶ $\gamma_{\mathcal{C}}(\theta)$: Power function defined on a range of θ

Type I error and Type II error

- ▶ Type I error probability (α): The probability that we reject H_0 when H_0 is true.
- ▶ Type II error probability (β): The probability that we accept H_0 when H_0 is not true.

		True	
		H_0	H_1
Decision	H_0	No error	Type II error
	H_1	Type I error	No error

- Example 8.1

► $X_1, X_2, X_3, X_4 \stackrel{iid}{\sim} b(1, \theta), \theta \in \Omega = \{1/3, 2/3\}$

► $H_0 : \theta = 1/3$ versus $H_1 : \theta = 2/3$

► Critical region:

$$\mathcal{C} = \{(X_1, X_2, X_3, X_4) : X_1 + X_2 + X_3 + X_4 \geq 2\}$$

► Type I error probability(α):

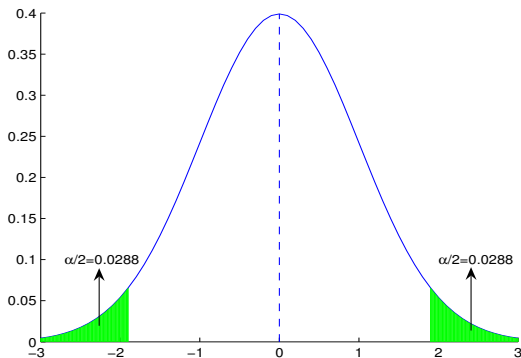
► Type II error probability(β):

- Example 8.2

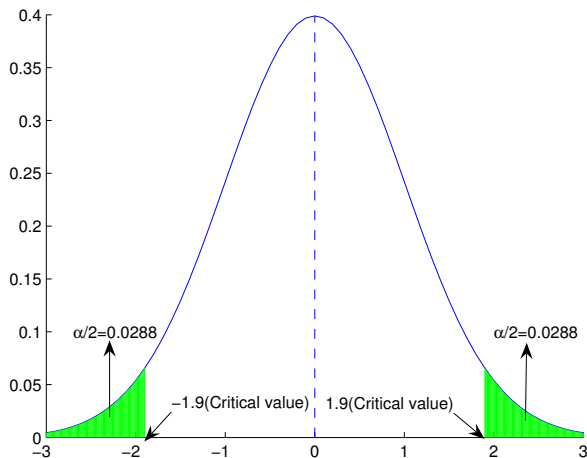
- Suppose we have 10 random samples from $N(\mu, 2.5^2)$.
- $H_0 : \mu = 50$ versus $H_1 : \mu \neq 50$
- Critical (or rejection) region:

$\mathcal{C} = \{X_1, \dots, X_n | Z_0 < -1.9 \text{ or } Z_0 > 1.9\}$, where

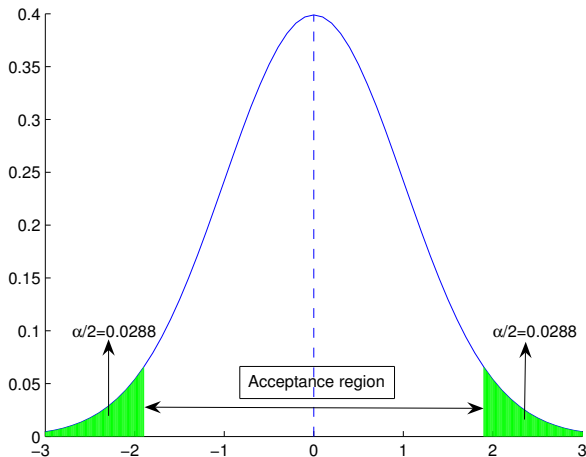
$$Z_0 = \frac{\bar{X}_n - \mu_0}{\sigma/\sqrt{n}} = \frac{\bar{X}_n - 50}{2.5/\sqrt{10}}$$



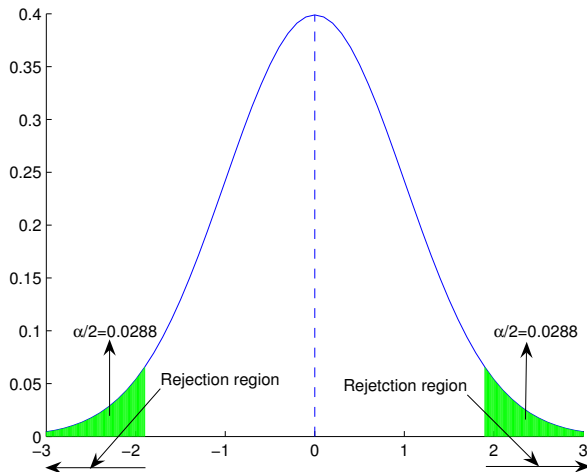
Critical value



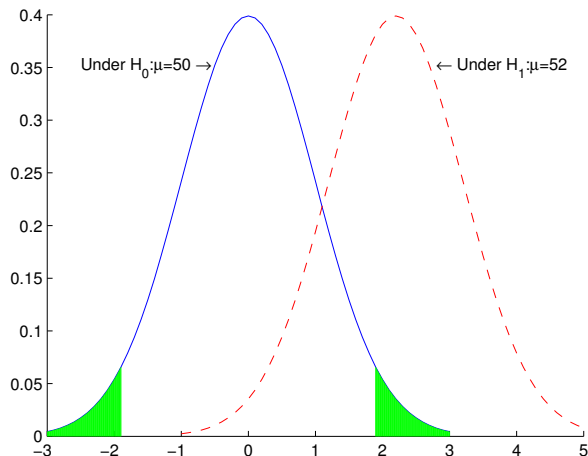
Acceptance region



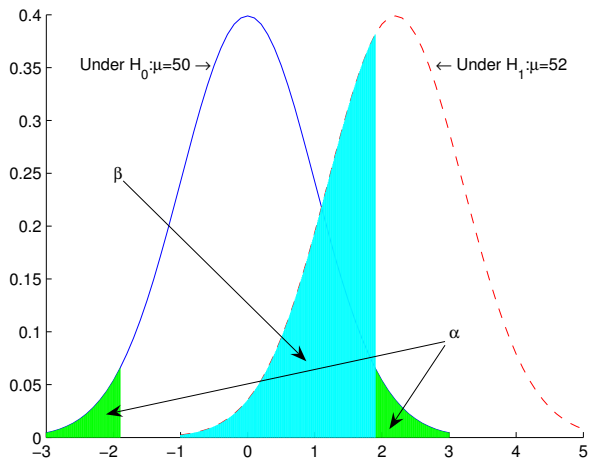
Critical (Rejection) region



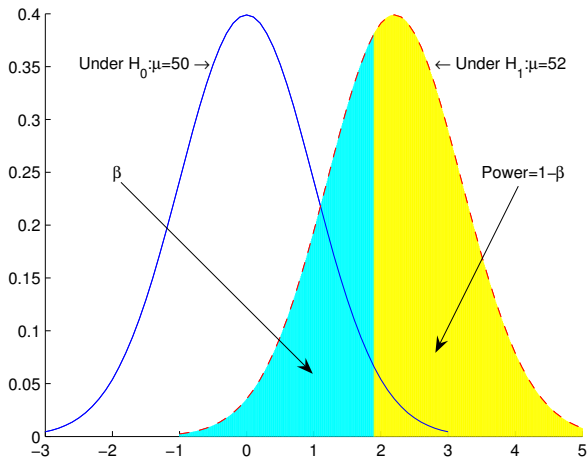
Under alternative hypothesis



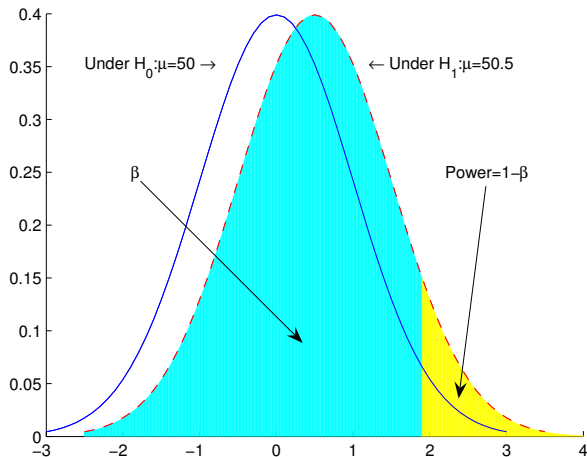
Type I error and Type II error



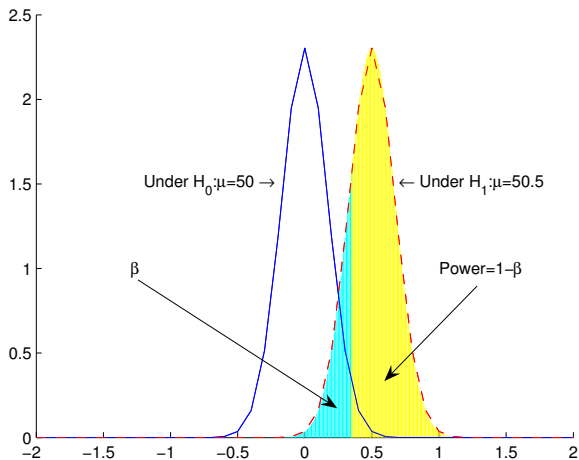
Power



What if H_0 and H_1 are too close



Sample size and power



- Exercises: 4.5.3, 4.5.4, 4.5.8

How to find the critical region \mathcal{C} ?

- ▶ Question 1: Can you construct a test with $\alpha = 0$ or $\beta = 0$?
 - ▶ Question 2: Can you construct a “best” test?
 - ▶ Question 3: What is “best”?
- In general, we cannot control both type I and type II errors for a given data. Our strategy to find a test is
- (i) Fix α
 - (ii) Find a test that makes β as small as possible

Definition (p. 484)

For hypothesis testing $H_0 : \theta = \theta'$ versus $H_1 : \theta = \theta''$, we will say that \mathcal{C} is a best critical region of size α if

- (a) $P(\mathbf{X} \in \mathcal{C} | \theta = \theta') = \alpha$
- (b) For any critical region \mathcal{A} that satisfies $P(\mathbf{X} \in \mathcal{A} | \theta = \theta') = \alpha$, we have $P(\mathbf{X} \in \mathcal{C} | \theta = \theta'') \geq P(\mathbf{X} \in \mathcal{A} | \theta = \theta'')$

- Example 8.3: $X \sim b(5, \theta)$, $H_0 : \theta = \frac{1}{2}$ versus $H_1 : \theta = \frac{3}{4}$

x	0	1	2	3	4	5
$P(X = x \theta = \frac{1}{2})$	$\frac{1}{32}$	$\frac{5}{32}$	$\frac{10}{32}$	$\frac{10}{32}$	$\frac{5}{32}$	$\frac{1}{32}$
$P(X = x \theta = \frac{3}{4})$	$\frac{1}{1024}$	$\frac{15}{1024}$	$\frac{90}{1024}$	$\frac{270}{1024}$	$\frac{405}{1024}$	$\frac{243}{1024}$

► What is the best critical region of size $\alpha = \frac{1}{32}$?

- Example 8.3: $X \sim b(5, \theta)$, $H_0 : \theta = \frac{1}{2}$ versus $H_1 : \theta = \frac{3}{4}$

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► What is the best critical region of size $\alpha = \frac{6}{32}$?

Neyman-Pearson Theorem

Theorem (Neyman-Pearson Theorem, p.486)

$$X_1, \dots, X_n \stackrel{iid}{\sim} f(x; \theta), \Omega = \{\theta', \theta''\}$$

For hypothesis testing $H_0 : \theta = \theta'$ versus $H_1 : \theta = \theta''$, \mathcal{C} is a best critical region of size α if

$$(a) \frac{L(\theta'; \mathbf{X})}{L(\theta''; \mathbf{X})} \leq k \text{ for each } \mathbf{X} \in \mathcal{C}.$$

$$(b) \frac{L(\theta'; \mathbf{X})}{L(\theta''; \mathbf{X})} \geq k \text{ for each } \mathbf{X} \in \mathcal{C}^c.$$

$$(c) P(\mathbf{X} \in \mathcal{C} | \theta = \theta') = \alpha.$$

- Example 8.4 (p. 489)

$$X_1, \dots, X_n \stackrel{iid}{\sim} N(\theta, 1), \theta \in \Omega = \{0, 1\}$$

$H_0 : \theta = 0$ versus $H_1 : \theta = 1$

- Example 8.4 (continued)

$$X_1, \dots, X_n \stackrel{iid}{\sim} N(\theta, 1), \theta \in \Omega = [0, \infty)$$

$H_0 : \theta = 0$ versus $H_1 : \theta > 0$

Exercises: 8.1.2, 8.1.5, 8.1.9

Definition (p. 493)

For hypothesis testing $H_0 : \theta = \theta'$ versus $H_1 : \theta \in \Omega_1$, where $\Omega = \Omega_1 \cup \{\theta'\}$, we will say that C^{UMP} is a **uniformly most powerful (UMP)** critical region of size α if

- (a) $P(\mathbf{X} \in C^{UMP} | \theta = \theta') = \alpha$
- (b) $P(\mathbf{X} \in C^{UMP} | \theta = \theta'') \geq P(\mathbf{X} \in \mathcal{C} | \theta = \theta'')$ for all $\theta'' \in \Omega_1$
and all size α critical region \mathcal{C} .

- Example 8.4 (continued)

$$X_1, \dots, X_n \stackrel{iid}{\sim} N(\theta, 1), \theta \in \Omega = (-\infty, \infty)$$

$$H_0 : \theta = 0 \text{ versus } H_1 : \theta \neq 0$$

Definition

$L(\theta; \mathbf{X})$ is said to have monotone likelihood ratio (MLR) property in the statistic $Y = u(\mathbf{X})$ if $\frac{L(\theta_1; \mathbf{X})}{L(\theta_2; \mathbf{X})}$ is a monotone function of $Y = u(\mathbf{X})$ for $\theta_1 < \theta_2$.

- Example 8.4 (continued)

$$X_1, \dots, X_n \stackrel{iid}{\sim} N(\theta, 1), \theta \in \Omega = [0, \infty)$$

$$H_0 : \theta = 0 \text{ versus } H_1 : \theta > 0$$

- Example 8.5

$X_1, \dots, X_n \stackrel{iid}{\sim} N(0, \theta), H_0 : \theta = \theta' \text{ versus } H_1 : \theta > \theta'$

- UMP test for $H_0 : \theta = \theta'$ versus $H_1 : \theta > \theta'$ with a regular exponential family

- $X_1, \dots, X_n \stackrel{iid}{\sim} f(x; \theta) = \exp\{p(\theta)K(x) + H(x) + q(\theta)\}$

► If $p(\theta)$ is monotone increasing,

$$\mathcal{C}^{UMP} = \left\{ \sum_{i=1}^n K(X_i) \geq c \right\} \text{ for some } c \text{ satisfying } P(\mathbf{X} \in \mathcal{C}^{UMP}) = \alpha$$

► If $p(\theta)$ is monotone decreasing,

$$\mathcal{C}^{UMP} = \left\{ \sum_{i=1}^n K(X_i) \leq c \right\} \text{ for some } c \text{ satisfying } P(\mathbf{X} \in \mathcal{C}^{UMP}) = \alpha$$

- Example 8.4 (revisited)

$$X_1, \dots, X_n \stackrel{iid}{\sim} N(\theta, 1), \theta \in \Omega = [0, \infty)$$

$H_0 : \theta = 0$ versus $H_1 : \theta > 0$

- Example 8.5 (revisited)

$X_1, \dots, X_n \stackrel{iid}{\sim} N(0, \theta), H_0 : \theta = \theta' \text{ versus } H_1 : \theta > \theta'$

- In fact, if X_1, \dots, X_n is a random sample from a regular exponential family $f(x; \theta) = \exp\{p(\theta)K(x) + H(x) + q(\theta)\}$, then

- ▶ The MLE is a function of $\sum K(X_i)$
- ▶ $\sum K(X_i)$ is a complete sufficient statistic for θ
- ▶ If $p(\theta)$ is monotone function, it has MLR property
- ▶ If $p(\theta)$ is monotone function, \mathcal{C}^{UMP} for one-sided test should be either $\{\sum_{i=1}^n K(X_i) \geq c\}$ or $\{\sum_{i=1}^n K(X_i) \leq c\}$

- Example 8.6

$X_1, \dots, X_n \stackrel{iid}{\sim} \text{Exp}(\theta)$, $H_0 : \theta = 2$ versus $H_1 : \theta > 2$

- Example 8.7

$X_1, \dots, X_n \stackrel{iid}{\sim} \text{Gamma}(3, \theta)$, $H_0 : \theta = \theta_0$ versus $H_1 : \theta < \theta_0$

- Example 8.8

$X_1, \dots, X_n \stackrel{iid}{\sim} U[0, \theta]$, $H_0 : \theta = \theta_0$ versus $H_1 : \theta < \theta_0$

Exercises: 8.2.5, 8.2.6, 8.2.11, 8.2.12 except (d)