

# Stochastic Processes (STA3021)

## HW7 Solution

### 1. Chapter 5 #4

The problem is asking

$$P(A > B + C, A > B)$$

under the following service time assumptions.

(a) For  $A = B = C = 10$ ,

$$P(A > B + C, A > B) = 0$$

since it is impossible.

(b) Desired event can be expressed as

$$\begin{aligned} P(A > B + C, A > B) &= P(A > 1 + C, A > 1 | B = 1) P(B = 1) \\ &+ P(A > 2 + C, A > 1 | B = 2) P(B = 2) + P(A > 3 + C, A > 3 | B = 3) P(B = 3) \\ &= P(A = 3, C = 1) P(B = 1) + 0 + 0 = \frac{1}{27}. \end{aligned}$$

(c)

$$\begin{aligned} P(A > B + C, A > B) &= \int_0^\infty P(A > b + C, A > b | B = b) f_B(b) db \\ &= \int_0^\infty \int_0^\infty P(A > b + c, A > b | B = b, C = c) f_C(c) dc f_B(b) db \\ &= \int_0^\infty \int_0^\infty P(A > b + c | B = b, C = c) f_C(c) dc f_B(b) db \\ &= \int_0^\infty \int_0^\infty e^{-\mu(b+c)} \mu^2 e^{-(b+c)\mu} db dc = \mu^2 \int_0^\infty e^{-2b\mu} db \int_0^\infty e^{-2c\mu} dc = \frac{1}{4}. \end{aligned}$$

### 2. Chapter 5 #18

(a)(b)

$$\begin{aligned} P(X_{(1)} \leq x) &= 1 - P(X_{(1)} > x) = 1 - P(X_1 > x, X_2 > x) \\ &= 1 - P(X_1 > x) P(X_2 > x) = 1 - e^{-2\mu x}. \end{aligned}$$

Thus,  $X_{(1)}$  follows Exponential distribution with rate  $2\mu$ . It gives that

$$EX_{(1)} = \frac{1}{2\mu}, \quad \text{Var}X_{(1)} = \frac{1}{4\mu^2}.$$

(c)(d) For  $X_{(2)}$ , which is the maximum of  $(X_1, X_2)$ , note that

$$P(X_{(2)} \leq x) = P(X_1 < x, X_2 < x) = (1 - e^{-\mu x})^2.$$

Thus, the pdf is given by

$$f(x) = \frac{d}{dx}(1 - e^{-\mu x})^2 = 2\mu e^{-\mu x}(1 - e^{-\mu x}).$$

$$\begin{aligned} EX_{(2)} &= \int_0^\infty x 2\mu e^{-\mu x}(1 - e^{-\mu x}) dx = 2\mu \left( \int_0^\infty x e^{-\mu x} dx - \int_0^\infty x e^{-2\mu x} dx \right) \\ &= 2\mu \left( \frac{1}{\mu^2} - \frac{1}{(2\mu)^2} \right) = \frac{3}{2\mu}, \end{aligned}$$

where we use the property of Exponential distribution

$$\int_0^\infty x \mu e^{-\mu x} dx = \frac{1}{\mu}.$$

Similarly,

$$\begin{aligned} \text{Var} X_{(2)} &= \int_0^\infty x^2 2\mu e^{-\mu x}(1 - e^{-\mu x}) dx - (EX_{(2)})^2 \\ &= 2\mu \left\{ \int_0^\infty x^2 e^{-\mu x} dx - \int_0^\infty x^2 e^{-2\mu x} dx \right\} - \left( \frac{3}{2\mu} \right)^2 \\ &= 2\mu \left\{ \frac{\Gamma(3)}{\mu^3} - \frac{\Gamma(3)}{(2\mu)^3} \right\} - \left( \frac{3}{2\mu} \right)^2 = \frac{5}{4\mu^2}, \end{aligned}$$

where the Gamma distribution integrates to 1 implies that

$$\int_0^\infty x^{\alpha-1} e^{-\mu x} dx = \frac{\Gamma(\alpha)}{\mu^\alpha}.$$

Also remark that the general pdf formula for  $r$ -th order statistics from the IID random sample  $X_1, \dots, X_n$  is given by

$$\frac{n!}{(r-1)!1!(n-r)!} F(x)^{r-1} f(x) (1 - F(x))^{n-r}.$$

### 3. Chapter 5 #30

Let  $T_d$  and  $T_c$  denote the life time of dog and cat which follows Exponential distribution with  $\lambda_d$  and  $\lambda_c$  respectively. Denote  $A$  be the additional life time, then

$$EA = E(A|T_d > T_c)P(T_d > T_c) + E(A|T_d < T_c)P(T_d < T_c).$$

As done in class,

$$P(T_d > T_c) = \frac{\lambda_c}{\lambda_c + \lambda_d}, \quad P(T_d < T_c) = \frac{\lambda_d}{\lambda_c + \lambda_d}$$

Hence,

$$EA = \frac{1}{\lambda_d} \frac{\lambda_c}{\lambda_c + \lambda_d} + \frac{1}{\lambda_c} \frac{\lambda_d}{\lambda_c + \lambda_d}.$$

4. Chapter 5 #31

Let  $T_i$  denote the amount of time that  $i$ -th appointment takes at the doctor's office. Then, the expected time can be expressed as

$$ET_2 = E(T_2|T_1 < 30)P(T_1 < 30) + E(T_2|T_1 > 30)P(T_1)$$

Since, the first patient see a doctor with Exponential distribution with mean 30,

$$P(T_1 < 30) = 1 - e^{-1}, \quad P(T_1 > 30) = e^{-1}.$$

When the second patient came to doctor's office and found that the first patient was still seeing a doctor (which is the latter case), then the second patient should wait till the first patient finish the appointment and her own amount of time seeing a doctor. This is on average 60 min due to memoryless property of Exponential distribution. All is all,

$$ET_2 = 30(1 - e^{-1}) + 60(e^{-1}) = 30 + 30e^{-1}.$$

5. Chapter 5 #42

(a) Straightforward calculation gives

$$ES_4 = E\left(\sum_{i=1}^4 T_i\right) = \frac{4}{\lambda}.$$

(b) Note that since  $N(1) = 2$ , the arrivals are given by

$$S_1 < S_2 \leq 1 < S_3 < S_4.$$

Now consider the counting process start at time 1, that is,

$$N'(t) = \sum_{i=3}^{\infty} 1_{\{S_i \leq t\}},$$

then  $N'(t)$  is a Poisson process with rate  $\lambda$ . Denote the inter-arrivals of  $N'(t)$  as  $T'_i$ , then

$$\begin{aligned} E(S_4|N(1) = 2) &= E(T_1 + T_2 + T_3 + T_4|N(1) = 2) \\ &= E(T_1 + T_2 + (1 - T_1 - T_2) + (T_1 + T_2 + T_3 - 1) + T_4|N(1) = 2) \\ &= 1 + E((T_1 + T_2 + T_3 - 1) + T_4|N(1) = 2) = 1 + E(T'_1 + T'_2) = 1 + \frac{2}{\lambda}. \end{aligned}$$

(c) Since the PP has independent increments

$$E(N(4) - N(2)|N(1) = 3) = E(N(4) - N(2)) = 2\lambda.$$

6. Chapter 5 #50

Let  $T$  and  $X$  denote the number of hours between successive train arrivals and the number of people who get on the next train respectively. Since the train just leaved, we should wait  $T$  amount of time to get on the next train. During the  $T$  time, the passengers arrive with Poisson process process with rate 7, it is deduced that  $X|T$  follows poisson distribution with rate  $7T$ .

(a) By conditioning on  $T$ , it follows that

$$EX = E(E(X|T)) = E(7T) = \frac{7}{2}.$$

(b) Similarly, by conditioning on  $T$ ,

$$\text{Var}X = E(\text{Var}(X|T)) + \text{Var}(E(X|T)) = E(7T) + \text{Var}(7T) = \frac{7}{2} + \frac{49}{12} = \frac{91}{12}.$$