

Lindley and Smith (JRSS-B, 1972).

$$y_i | \theta_i, \mu \stackrel{\text{indep}}{\sim} N(\theta_i, \sigma^2), \quad \sigma^2 (> 0) \text{ known}$$

$$\theta_i | \mu \stackrel{\text{iid}}{\sim} N(\mu, \tau^2), \quad \tau^2 (> 0) \text{ known}$$

$$\mu \sim \text{uniform}(-\infty, \infty)$$

$$\Rightarrow P(\theta_1, \dots, \theta_n, \mu | y_1, \dots, y_n)$$

$$\begin{aligned} &\propto \exp\left\{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \theta_i)^2\right\} \exp\left\{-\frac{1}{2\tau^2} \sum_{i=1}^n (\theta_i - \mu)^2\right\} \\ &= \exp\left\{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \theta_i)^2\right\} \exp\left\{-\frac{1}{2\tau^2} \sum_{i=1}^n (\theta_i - \bar{\theta})^2\right\} \exp\left\{-\frac{n}{2\tau^2} (\mu - \bar{\theta})^2\right\} \underbrace{- \theta}_{\text{kernel of } N(\bar{\theta}, \frac{\tau^2}{n})} \end{aligned}$$

Then

$$P(\theta_1, \dots, \theta_n | y_1, \dots, y_n) = \int P(\theta_1, \dots, \theta_n, \mu | y_1, \dots, y_n) d\mu$$

$$\begin{aligned} &\propto \exp\left\{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \theta_i)^2\right\} \exp\left\{-\frac{1}{2\tau^2} \sum_{i=1}^n (\theta_i - \bar{\theta})^2\right\} \underbrace{\int \frac{1}{\sqrt{2\pi} \frac{\tau}{\sqrt{n}}} e^{-\frac{n}{2\tau^2} (\mu - \bar{\theta})^2} d\mu}_{=1} \\ &\quad \uparrow \\ &\text{by (*)} \quad \left. \begin{aligned} \sum_{i=1}^n y_i^2 - 2 \sum_{i=1}^n y_i \theta_i + \sum_{i=1}^n \theta_i^2 &= \sum_{i=1}^n \theta_i^2 - n \bar{\theta}^2 \\ &= \underline{\theta}^T \underline{\theta} - n \left(\frac{1}{n} \underline{1}^T \underline{\theta}\right)^2 \quad \text{where } \underline{\theta} = \begin{pmatrix} \theta_1 \\ \vdots \\ \theta_n \end{pmatrix} \\ &= \underline{\theta}^T \underline{\theta} - \frac{1}{n} (\underline{\theta}^T \underline{1}) (\underline{1}^T \underline{\theta}) \\ &\stackrel{\uparrow}{=} \underline{\theta}^T \underline{\theta} - \underline{\theta}^T \frac{1}{n} \underline{J} \underline{\theta} \end{aligned} \right\} \\ &\quad \underline{J} = \underline{1} \underline{1}^T \\ &\quad = \underline{\theta}^T \left( \underline{I} - \frac{1}{n} \underline{J} \right) \underline{\theta} \end{aligned}$$

$$\propto \exp\left\{-\frac{1}{2\sigma^2} (\underline{\theta}^T \underline{\theta} - 2 \underline{\theta}^T \underline{y}) - \frac{1}{2\tau^2} \underline{\theta}^T \left( \underline{I} - \frac{1}{n} \underline{J} \right) \underline{\theta} \right\}$$

$$= \exp \left[ -\frac{1}{2} \left\{ \underline{\theta}^T \left( \frac{1}{\sigma^2} \underline{I} + \frac{1}{\tau^2} \left( \underline{I} - \frac{1}{n} \underline{J} \right) \right) \underline{\theta} - 2 \frac{1}{\sigma^2} \underline{\theta}^T \underline{y} \right\} \right]$$

$$\left( \frac{1}{\sigma^2} + \frac{1}{\tau^2} \right) \underline{I} - \frac{1}{\tau^2 n} \underline{J}$$

$$= \exp \left[ -\frac{1}{2} \left\{ \underline{\theta}^T \left( \left( \frac{1}{\sigma^2} + \frac{1}{\tau^2} \right) \underline{I} - \frac{1}{\tau^2 n} \underline{J} \right) \underline{\theta} - 2 \frac{1}{\sigma^2} \underline{\theta}^T \underline{y} \right\} \right]$$

$$\propto \exp \left[ -\frac{1}{2} \left\{ \underline{Q} - \left( \left( \frac{1}{\sigma^2} + \frac{1}{\tau^2} \right) I - \frac{1}{n\tau^2} J \right)^{-1} \frac{1}{\sigma^2} \underline{y} \right\} \left( \left( \frac{1}{\sigma^2} + \frac{1}{\tau^2} \right) I - \frac{1}{\tau^2 n} J \right) \right. \\ \left. \right\{ \underline{Q} - \left( \left( \frac{1}{\sigma^2} + \frac{1}{\tau^2} \right) I - \frac{1}{n\tau^2} J \right)^{-1} \frac{1}{\sigma^2} \underline{y} \right\} \right]$$

∴ kernel of multivariate normal dist. with

mean  $\left( \left( \frac{1}{\sigma^2} + \frac{1}{\tau^2} \right) I - \frac{1}{\tau^2} \left( \frac{1}{n} J \right) \right)^{-1} \frac{1}{\sigma^2} \underline{y}$  and

Variance  $\left( \left( \frac{1}{\sigma^2} + \frac{1}{\tau^2} \right) I - \frac{1}{\tau^2 n} J \right)^{-1} = D^{-1}$

$$\text{Let } D = \left( \frac{1}{\sigma^2} + \frac{1}{\tau^2} \right) I - \frac{1}{\tau^2} \left( \frac{1}{n} J \right).$$

Then posterior variance is  $D^{-1}$

and posterior mean is  $\frac{1}{\sigma^2} D^{-1} \underline{y}$ .

Thus

$$\underline{Q} | \underline{y} \sim N \left( \frac{1}{\sigma^2} D^{-1} \underline{y}, D^{-1} \right).$$

$$\text{Note } D = \left( \frac{1}{\sigma^2} + \frac{1}{\tau^2} \right) \left( I - \frac{\sigma^2}{\sigma^2 + \tau^2} \left( \frac{1}{n} \underline{1} \underline{1}^T \right) \right) \\ = \frac{1}{\sigma^2 (1-B)} \left( I - B \left( \frac{1}{n} \underline{1} \underline{1}^T \right) \right)$$

$$B = \frac{\sigma^2}{\sigma^2 + \tau^2}$$

$$\left( \text{Remark If } A \text{ and } A + \underline{u} \underline{u}^T \text{ are both invertible, then} \right. \\ \left. (A + \underline{u} \underline{u}^T)^{-1} = A^{-1} - \frac{A^{-1} \underline{u} \underline{u}^T A^{-1}}{1 + \underline{u}^T A^{-1} \underline{u}} \right)$$

$$D^{-1} = \sigma^2 (I - B) (I - B (\frac{1}{n} \mathbf{1} \mathbf{1}^T))^{-1}$$

$$= \sigma^2 (I - B) (I + \underline{u} \underline{u}^T)$$

$$\uparrow$$

$$\underline{u} = \sqrt{B} \frac{1}{\sqrt{n}} \mathbf{1}$$

$$\underline{v} = -\underline{u}$$

$$= \sigma^2 (I - B) \left[ I + \frac{I^{-1} \sqrt{B} \frac{1}{\sqrt{n}} \mathbf{1} (-\underline{u})^T I^{-1}}{1 + (-\underline{u})^T I^{-1} \sqrt{B} \frac{1}{\sqrt{n}} \mathbf{1}} \right]$$

by Remark

$$= \sigma^2 (I - B) \left[ I + \frac{\sqrt{B} \frac{1}{\sqrt{n}} \mathbf{1} (\sqrt{B} \frac{1}{\sqrt{n}} \mathbf{1})^T}{1 - (\sqrt{B} \frac{1}{\sqrt{n}} \mathbf{1})^T \sqrt{B} \frac{1}{\sqrt{n}} \mathbf{1}} \right]$$

$$= \sigma^2 (I - B) \left[ I + \frac{\frac{B}{n} \mathbf{1} \mathbf{1}^T}{1 - \frac{B}{n} \underbrace{(\mathbf{1}^T \mathbf{1})}_{=n}} \right]$$

$$= \sigma^2 (I - B) \left( I + \frac{\frac{B}{n} J}{1 - B} \right)$$

$$= \sigma^2 [(I - B) I + B (\frac{1}{n} J)]$$

Thus,

$$\frac{1}{\sigma^2} D^{-1} \underline{y} = \frac{1}{\sigma^2} \left[ \sigma^2 \{ (I - B) I + B (\frac{1}{n} J) \} \right] \underline{y}$$

$$= (I - B) \underline{y} + B \left( \frac{1}{n} J \underline{y} \right)$$

$$\underbrace{\frac{1}{n} \mathbf{1}^T \underline{y}}_{= \sum y_i}$$

$$= (I - B) \underline{y} + B \frac{1}{n} \sum y_i \mathbf{1}$$

$$= (I - B) \underline{y} + B \bar{y} \mathbf{1}$$

Rewrite the posterior dist. of  $\underline{\theta}$  given  $\underline{y}$  is

$$\Rightarrow \underline{\theta} | \underline{y} \sim N \left( (I - B) \underline{y} + B \bar{y} \mathbf{1}, \sigma^2 [(I - B) I + B (\frac{1}{n} J)] \right)$$