Ch 8. Hypothesis testing: Part II

Likelihood Ratio Tests (Ch 6.3 & 8.3)

 $H_0: \theta \in \Omega_0$ versus $H_1: \theta \in \Omega \cap \Omega_0^c$

▶ Likelihood Ratio Test (LRT) of size α rejects H_0 if

$$\frac{\max_{\theta \in \Omega_0} L(\theta; \mathbf{X})}{\max_{\theta \in \Omega} L(\theta; \mathbf{X})} \le k,$$

with k satisfying

$$\max_{\theta \in \Omega_0} P\left(\frac{\max_{\theta \in \Omega_0} L(\theta; \mathbf{X})}{\max_{\theta \in \Omega} L(\theta; \mathbf{X})} \le k\right) = \alpha$$

 \to If $\Omega_0=\{\theta_0\}$ (i.e. $H_0:\theta=\theta_0$ versus $H_1:\theta\in\Omega\cap\Omega_0^c$), LRT of size α reject H_0 if

▶ Likelihood Ratio Test (LRT) of size α rejects H_0 if

$$\frac{L(\theta_0; \mathbf{X})}{L\left(\hat{\theta}^{MLE}; \mathbf{X}\right)} \le k,$$

with k satisfying

$$P_{\theta_0} \left(\frac{L(\theta_0; \mathbf{X})}{L(\hat{\theta}^{(MLE)}; \mathbf{X})} \le k \right) = \alpha$$

ightarrow LRT is not the "best" in general, but it is known that LRT is asymptotically "best".

- Equivalent form of LRT:

Reject H_0 if

$$2\left(\ell\left(\hat{\theta}^{\Omega_0};\mathbf{X}\right)-\ell\left(\hat{\theta}^{\Omega};\mathbf{X}\right)\right)\leq c$$

with

$$\max_{\boldsymbol{\theta} \in \Omega_0} P\left[2\left(\ell\left(\hat{\boldsymbol{\theta}}^{\Omega_0}; \mathbf{X}\right) - \ell\left(\hat{\boldsymbol{\theta}}^{\Omega}; \mathbf{X}\right)\right) \leq c\right] = \alpha$$

where

$$\hat{\theta}^{\Omega} = \underset{\theta \in \Omega}{\operatorname{arg\,max}} L(\theta; \mathbf{X})$$

$$\hat{\theta}^{\Omega_0} = \operatorname*{arg\,max}_{\theta \in \Omega_0} L(\theta; \mathbf{X})$$

- If $H_0: \theta=\theta_0$ versus $H_1: \theta \neq \theta_0$, the equivalent LRT is Reject H_0 if

$$2\left(\ell\left(\theta_{0};\mathbf{X}\right)-\ell\left(\hat{\theta}^{MLE};\mathbf{X}\right)\right)\leq c$$

with

$$P_{\theta_0}\left[2\left(\ell\left(\theta_0; \mathbf{X}\right) - \ell\left(\hat{\theta}^{MLE}; \mathbf{X}\right)\right) \le c\right] = \alpha$$

- Example 8.9 (two-sided test)

$$X_1,\ldots,X_n\stackrel{iid}{\sim} N(\mu,\sigma^2)$$
, σ^2 is known and $\mu\in\Omega=(-\infty,\infty)$ $H_0:\mu=\mu_0$ versus $H_1:\mu\neq\mu_0$

- Example 8.9 (One-sided test)

$$X_1,\ldots,X_n\stackrel{iid}{\sim} N(\mu,\sigma^2),\ \sigma^2$$
 is known and $\mu\in\Omega=(-\infty,\infty)$ $H_0:\mu\leq\mu_0$ versus $H_1:\mu>\mu_0$

 $X_1,\ldots,X_n\stackrel{iid}{\sim} Exp(heta)$, $H_0: heta= heta_0$ versus $H_1: heta
eq heta_0$

 $X_1,\dots,X_n\stackrel{iid}{\sim} N(\theta_1,\theta_2)$, θ_2 is unknown and $-\infty<\theta_1<\infty$, $\theta_2>0$

 $H_0: heta_1 = heta_0$ versus $H_1: heta_1
eq heta_0$

$$X_1, \ldots, X_n \stackrel{iid}{\sim} N(\mu, \theta), -\infty < \mu < \infty, \theta > 0$$

 $H_0: \theta = heta_0$ versus $H_1: heta
eq heta_0$

$$X_1,\ldots,X_n \overset{iid}{\sim} N(\mu_1,\sigma^2)$$
, $Y_1,\ldots,Y_m \overset{iid}{\sim} N(\mu_2,\sigma^2)$, $-\infty < \mu_1,\mu_2 < \infty$, $\sigma^2 > 0$ is known, $H_0: \mu_1 = \mu_2$ versus $H_1: \mu_1 \neq \mu_2$

 $X_1,\ldots,X_n\stackrel{iid}{\sim} N(\mu_1,\theta_1),\ Y_1,\ldots,Y_m\stackrel{iid}{\sim} N(\mu_2,\theta_2),\ H_0:\theta_1=\theta_2$ versus $H_1:\theta_1\neq\theta_2$

Other tests

Let $\hat{\theta}$ be the MLE of θ .

- Likelihood Ratio Test (LRT): Reject H_0 if $-2(\ell(\theta_0) - \ell(\hat{\theta})) > k$
- $lackbox{ Wald test:}$ Reject H_0 if $\left(\sqrt{nI(\hat{ heta})}(\hat{ heta}- heta_0)
 ight)^2>k$
- Score test: Reject H_0 if $\left(\frac{\ell'(\theta_0)}{\sqrt{nI(\theta_0)}}\right)^2 > k$

Theorem (Asymptotic likelihood ratio test)

Under regularity conditions (R0) \sim (R4), if $H_0: \theta = \theta_0 \in \mathbb{R}^p$ is true, we have

$$-2\log\Lambda \stackrel{d}{\to} \chi_p^2 \ where \ \Lambda = \frac{L(\theta_0)}{L(\hat{\theta})} \ and \ \hat{\theta} \ is \ the \ MLE \ of \ \theta$$
 That is, $-2(\ell(\theta_0)-\ell(\hat{\theta})) \stackrel{d}{\to} \chi_p^2$

- Example 8.10 (revisited)

$$X_1, \ldots, X_n \stackrel{iid}{\sim} Exp(\theta), H_0: \theta = \theta_0 \text{ versus } H_1: \theta \neq \theta_0$$

Exercises: 6.3.16, 6.3.18, 6.3.19