

Experimental Design

Note 7

2^k Factorial Design

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2^k Factorial Design

- Involving k factors to detect the important factors in the process with a minimum of experimental units.
- Each factor has two levels (often labeled + and -)
- Factor screening experiment (preliminary study) often used to at the early stage of experimentation to detect potential candidate factors for more detailed investigation.
- Identify important factors and their interactions
- Interaction (of any order) has **ONE** degree of freedom
- Factors need not be on numeric scale
- Ordinary regression model can be employed

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_{12} x_1 x_2 + \epsilon$$

where β_1 , β_2 , and β_{12} are related to main effects, interaction effects defined later.

Chemical Processes Example

| Factor | | Treatment | Replicate | | | |
|--------|---|----------------|-----------|----|-----|-------|
| A | B | Combination | I | II | III | Total |
| - | - | A low, B low | 28 | 25 | 27 | 80 |
| + | - | A high, B low | 36 | 32 | 32 | 100 |
| - | + | A low, B high | 18 | 19 | 23 | 60 |
| + | + | A high, B high | 31 | 30 | 29 | 90 |

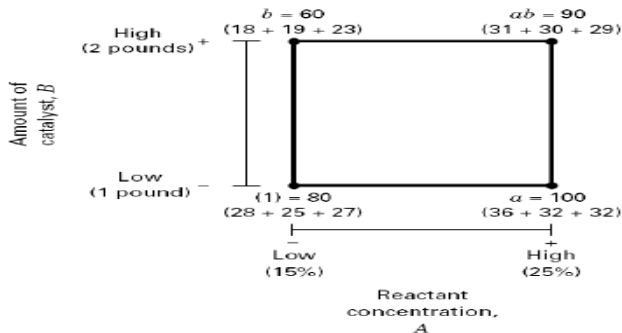
A =reactant concentration, B =catalyst amount, y =recovery

Analysis Procedure for a Factorial Design

- Estimate factor effects
- Formulate model
 - With replication, use full model
 - With an unreplicated design, use normal probability plots
- Statistical testing (ANOVA)
- Refine the model
- Analyze residuals (graphical)
- Interpret results

The Simplest Case: 2^2

Figure 6.1 Treatment combinations in the 2^2 design



- “-” and “+” denote the low and high levels of a factor, respectively.

Estimation of Factor Effects I

$$\begin{aligned}A &= \bar{y}_{A+} - \bar{y}_{A-} = \frac{ab + a}{2n} - \frac{b + (1)}{2n} = \frac{1}{2n} [ab + a - b - (1)] \\B &= \bar{y}_{B+} - \bar{y}_{B-} = \frac{ab + b}{2n} - \frac{a + (1)}{2n} = \frac{1}{2n} [ab + b - a - (1)] \\AB &= \frac{ab + (1)}{2n} - \frac{a + b}{2n} = \frac{1}{2n} [ab + (1) - a - b]\end{aligned}$$

The effect estimates are:

$$A = 8.33, B = -5.00, AB = 1.67$$

The quantities in brackets are **contrasts** in the treatment combinations.

Estimation of Factor Effects II

Effects and Contrasts

| factor | | | | effect (contrast) | | | |
|--------|---|-------|---------|-------------------|----|----|----|
| A | B | total | mean | 1 | A | B | AB |
| - | - | 80 | $80/3$ | 1 | -1 | -1 | 1 |
| + | - | 100 | $100/3$ | 1 | 1 | -1 | -1 |
| - | + | 60 | $60/3$ | 1 | -1 | 1 | -1 |
| + | + | 90 | $90/3$ | 1 | 1 | 1 | 1 |

- There is a one-to-one correspondence between effects and contrasts, and contrasts can be directly used to estimate the effects.

Estimation of Factor Effects III

- For a effect corresponding to contrast $c = (c_1, c_2, \dots)$ in 2^2 design

$$\text{effect} = \frac{1}{2} \sum_i c_i \bar{y}_i$$

where i is an index for treatments and the summation is over all treatments.
For example,

$$\begin{aligned} \text{effect}(A) &= \frac{1}{2n} \{ (y_{++\cdot} + y_{+-\cdot}) - (y_{-+\cdot} + y_{--\cdot}) \} \\ &= \frac{1}{2} \{ \bar{y}_{++} + \bar{y}_{+-} - \bar{y}_{-+} - \bar{y}_{--} \}, \\ \text{effect}(B) &= \frac{1}{2} \{ \bar{y}_{++} + \bar{y}_{-+} - \bar{y}_{+-} - \bar{y}_{--} \}, \\ \text{effect}(AB) &= \frac{1}{2} \{ \bar{y}_{++} + \bar{y}_{--} - \bar{y}_{+-} - \bar{y}_{-+} \}, \end{aligned}$$

Estimation of Factor Effects IV

Sum of Squares due to Effect

- Because effects are defined using contrasts, their sum of squares can also be calculated through contrasts.
- Recall for contrast $c = (c_1, c_2, \dots)$, its sum of squares is

$$SS_{\text{Contrast}} = \frac{(\sum_i c_i \bar{y}_i)^2}{\sum_i c_i^2 / n}$$

So

$$SS_A = \frac{(-\bar{y}(A_-B_-) + \bar{y}(A_+B_-) - \bar{y}(A_-B_+) + \bar{y}(A_+B_+))^2}{4/n} = 208.33$$

$$SS_B = \frac{(-\bar{y}(A_-B_-) - \bar{y}(A_+B_-) + \bar{y}(A_-B_+) + \bar{y}(A_+B_+))^2}{4/n} = 75.00$$

$$SS_{AB} = \frac{(\bar{y}(A_-B_-) - \bar{y}(A_+B_-) - \bar{y}(A_-B_+) + \bar{y}(A_+B_+))^2}{4/n} = 8.33$$

Estimation of Factor Effects V

Sum of Squares and ANOVA

| Source of Variation | Sum of Squares | Degrees of Freedom | Mean Square | F |
|---------------------|----------------|--------------------|-------------|---|
| A | SS_A | 1 | MS_A | |
| B | SS_B | 1 | MS_B | |
| AB | SS_{AB} | 1 | MS_{AB} | |
| Error | SSE | $N - 4$ | MSE | |
| Total | SS_T | $N - 1$ | | |

where $SS_T = \sum_{i,j,k} y_{ijk}^2 - y_{...}^2/N$, $SSE = SS_T - SS_A - SS_B - SS_{AB}$.

Revisit Chemical Process Example I

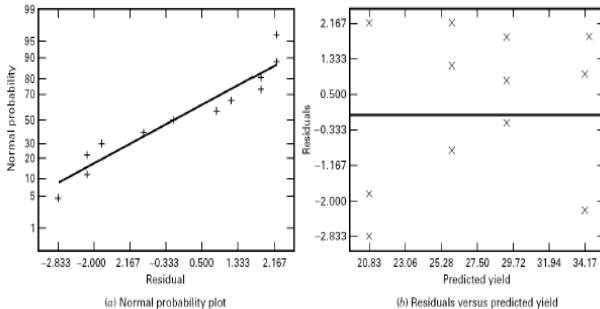
ANOVA Table

See Chemical-Process.SAS.

| Source of Variation | Sum of Squares | Degrees of Freedom | Mean Square | F | <i>P</i> -value |
|---------------------|----------------|--------------------|-------------|-------|-----------------|
| A | 208.33 | 1 | 208.33 | 53.15 | 0.0001 |
| B | 75.00 | 1 | 75.00 | 19.13 | 0.0024 |
| AB | 8.33 | 1 | 8.33 | 2.13 | 0.1826 |
| Error | 31.34 | 8 | 3.92 | | |
| Total | 323.00 | 11 | | | |

Revisit Chemical Process Example II

Residuals and Diagnostic Checking



■ FIGURE 6.2 Residual plots for the chemical process experiment

Analyzing 2^2 Experiment using Regression Model I

Because every effect in 2^2 design, or its sum of squares, has one degree of freedom, it can be equivalently represented by a numerical variable, and regression analysis can be directly used to analyze the data. The original factors are not necessarily continuous.

Code the levels of factor A and B as follow

| A | X_1 | B | X_2 |
|---|-------|---|-------|
| - | -1 | - | -1 |
| + | 1 | + | 1 |

Fit regression model

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_{12} x_1 x_2 + \epsilon$$

Analyzing 2^2 Experiment using Regression Model II

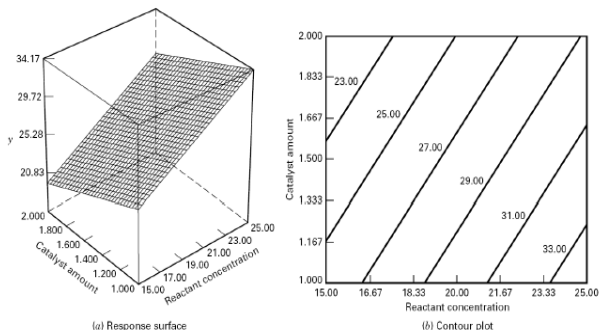
The fitted model should be

$$y = \bar{y}_{..} + \frac{A}{2}x_1 + \frac{B}{2}x_2 + \frac{AB}{2}x_1x_2$$

i.e., the estimated coefficients are half of the effects, respectively.
See Chemical-Process.SAS.

Analyzing 2^2 Experiment using Regression Model III

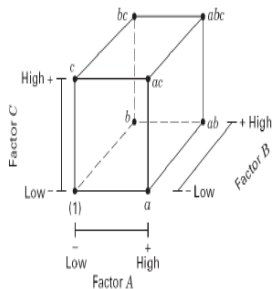
Response Surface



■ FIGURE 6.3 Response surface plot and contour plot of yield from the chemical process experiment

The 2^3 Factorial Design I

■ FIGURE 6.4
The 2^3 factorial
design



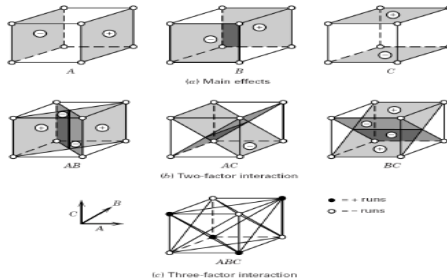
(a) Geometric view

| Run | Factor | | |
|-----|--------|---|---|
| | A | B | C |
| 1 | - | - | - |
| 2 | + | - | - |
| 3 | - | + | - |
| 4 | + | + | - |
| 5 | - | - | + |
| 6 | + | - | + |
| 7 | - | + | + |
| 8 | + | + | + |

(b) Design matrix

The 2^3 Factorial Design II

■ **FIGURE 6.5**
Geometric presentation
of contrasts
corresponding to the
main effects and
interactions in the
 2^3 design



$$A = \bar{y}_{A+} - \bar{y}_{A-}, \quad B = \bar{y}_{B+} - \bar{y}_{B-},$$

$$C = \bar{y}_{C+} - \bar{y}_{C-}, \quad \text{etc...}$$

Analysis is done via computer.

The 2^3 Factorial Design III

Table of - and + Signs for the 2^3 Factorial Design

■ TABLE 6.3

Algebraic Signs for Calculating Effects in the 2^3 Design

| Treatment Combination | Factorial Effect | | | | | | | |
|-----------------------|------------------|----------|----------|-----------|----------|-----------|-----------|------------|
| | <i>I</i> | <i>A</i> | <i>B</i> | <i>AB</i> | <i>C</i> | <i>AC</i> | <i>BC</i> | <i>ABC</i> |
| (1) | + | - | - | + | - | + | + | - |
| <i>a</i> | + | + | - | - | - | - | + | + |
| <i>b</i> | + | - | + | - | - | + | - | + |
| <i>ab</i> | + | + | + | + | - | - | - | - |
| <i>c</i> | + | - | - | + | + | - | - | + |
| <i>ac</i> | + | + | - | - | + | + | - | - |
| <i>bc</i> | + | - | + | - | + | - | + | - |
| <i>abc</i> | + | + | + | + | + | + | + | + |

The 2³ Factorial Design IV

Properties of the Table

- Except for column I , every column has an equal number of + and - signs
- The sum of the product of signs in any two columns is zero
- Multiplying any column by I leaves that column unchanged (identity element)
- The product of any two columns yields a column in the table:

$$A \times B = AB, \quad AB \times BC = AB^2C = AC$$

- Orthogonal design
- Orthogonality is an important property shared by all factorial designs

The 2^3 Factorial Design V

Contrasts for Calculating Effects in 2^3 Design

| | | | factorial effects | | | | | | | | |
|---|---|---|-------------------|----------|----------|----------|-----------|----------|-----------|-----------|------------|
| A | B | C | treatment | <i>I</i> | <i>A</i> | <i>B</i> | <i>AB</i> | <i>C</i> | <i>AC</i> | <i>BC</i> | <i>ABC</i> |
| - | - | - | (1) | 1 | -1 | -1 | 1 | -1 | 1 | 1 | -1 |
| + | - | - | a | 1 | 1 | -1 | -1 | -1 | -1 | 1 | 1 |
| - | + | - | b | 1 | -1 | 1 | -1 | -1 | 1 | -1 | 1 |
| + | + | - | ab | 1 | 1 | 1 | 1 | -1 | -1 | -1 | -1 |
| - | - | + | c | 1 | -1 | -1 | 1 | 1 | -1 | -1 | 1 |
| + | - | + | ac | 1 | 1 | -1 | -1 | 1 | 1 | -1 | -1 |
| - | + | + | bc | 1 | -1 | 1 | -1 | 1 | -1 | 1 | -1 |
| + | + | + | abc | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |

The 2^3 Factorial Design VI

Estimates:

$$\text{grand mean: } \frac{\sum_i \bar{y}_{i.}}{2^3}$$

$$\text{effect: } \frac{\sum_i c_i \bar{y}_{i.}}{2^{3-1}}$$

Contrast Sum of Squares:

$$SS_{\text{effect}} = \frac{(\sum_i c_i \bar{y}_{i.})^2}{2^3/n} = 2n(\text{effect})^2$$

Variance of Estimate:

$$\text{var}(\text{effect}) = \frac{\sigma^2}{n2^{3-2}}$$

The 2^3 Factorial Design VII

t -test for effects (confidence interval approach):

$$\text{effect} \pm t_{\alpha/2, 2^k(n-1)} SE(\text{effect})$$

Example I

Model Coefficients-Full Model

■ TABLE 6.4

The Plasma Etch Experiment, Example 6.1

| Run | Coded Factors | | | Etch Rate | | Total | Factor Levels | | |
|-----|---------------|----|----|-------------|-------------|-------------------|--|-----------|------|
| | A | B | C | Replicate 1 | Replicate 2 | | Low (-1) | High (+1) | |
| 1 | -1 | -1 | -1 | 550 | 604 | (1) = 1154 | A (Gap, cm) | 0.80 | 1.20 |
| 2 | 1 | -1 | -1 | 669 | 650 | <i>a</i> = 1319 | B (C ₂ F ₆ flow, SCCM) | 125 | 200 |
| 3 | -1 | 1 | -1 | 633 | 601 | <i>b</i> = 1234 | C (Power, W) | 275 | 325 |
| 4 | 1 | 1 | -1 | 642 | 635 | <i>ab</i> = 1277 | | | |
| 5 | -1 | -1 | 1 | 1037 | 1052 | <i>c</i> = 2089 | | | |
| 6 | 1 | -1 | 1 | 749 | 868 | <i>ac</i> = 1617 | | | |
| 7 | -1 | 1 | 1 | 1075 | 1063 | <i>bc</i> = 2138 | | | |
| 8 | 1 | 1 | 1 | 729 | 860 | <i>abc</i> = 1589 | | | |

A=Gap, B=Flow, C=Power, y=Etch Rate

Example II

Full Model

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + \beta_{12} x_1 x_2 + \beta_{13} x_1 x_3 + \beta_{23} x_2 x_3 + \beta_{123} x_1 x_2 x_3 + \epsilon$$

Model Coefficients for Full Model

| Factor | Coefficient Estimated | DF | Standard Error | 95% CI Low | 95% CI High |
|------------|--------------------------|----|-------------------|---------------|----------------|
| Intercept | 776.06 | 1 | 11.87 | 748.70 | 803.42 |
| A-Gap | -50.81 | 1 | 11.87 | -78.17 | -23.45 |
| B-Gas flow | 3.69 | 1 | 11.87 | -23.67 | 31.05 |
| C-Power | 153.06 | 1 | 11.87 | 125.70 | 180.42 |
| AB | -12.44 | 1 | 11.87 | -39.80 | 14.92 |
| AC | -76.81 | 1 | 11.87 | -104.17 | -49.45 |
| BC | -1.06 | 1 | 11.87 | -28.42 | 26.30 |
| ABC | 2.81 | 1 | 11.87 | -24.55 | 30.17 |

Example III

ANOVA Table

■ TABLE 6.6

Analysis of Variance for the Plasma Etching Experiment

| Source of Variation | Sum of Squares | Degrees of Freedom | Mean Square | F_0 | P -Value |
|---------------------|----------------|--------------------|--------------|--------|------------|
| Gap (A) | 41,310.5625 | 1 | 41,310.5625 | 18.34 | 0.0027 |
| Gas flow (B) | 217.5625 | 1 | 217.5625 | 0.10 | 0.7639 |
| Power (C) | 374,850.0625 | 1 | 374,850.0625 | 166.41 | 0.0001 |
| AB | 2475.0625 | 1 | 2475.0625 | 1.10 | 0.3252 |
| AC | 94,402.5625 | 1 | 94,402.5625 | 41.91 | 0.0002 |
| BC | 18.0625 | 1 | 18.0625 | 0.01 | 0.9308 |
| ABC | 126.5625 | 1 | 126.5625 | 0.06 | 0.8186 |
| Error | 18,020.5000 | 8 | 2252.5625 | | |
| Total | 531,420.9375 | 15 | | | |

Example IV

Reduced Model

$$y = \beta_0 + \beta_1 x_1 + \beta_3 x_3 + \beta_{13} x_1 x_3 + \epsilon$$

Model Coefficients for Reduced Model

| Factor | Coefficient Estimate | DF | Standard Error | 95% CI Low | 95% CI High |
|-----------|----------------------|----|----------------|------------|-------------|
| Intercept | 776.06 | 1 | 10.42 | 753.35 | 798.77 |
| A-Gap | -50.81 | 1 | 10.42 | -73.52 | 28.10 |
| C-Power | 153.06 | 1 | 10.42 | 130.35 | 175.77 |
| AC | -76.81 | 1 | 10.42 | -99.52 | -54.10 |

Example V

Model Summary Statistics

- R^2 and adjusted R^2 for reduced model

$$R^2 = \frac{SS_{Model}}{SS_T} = \frac{5.106 \times 10^5}{5.314 \times 10^5} = 0.9608$$

$$R^2_{Adj} = 1 - \frac{SSE/df_E}{SS_T/df_T} = 1 - \frac{20857.75/12}{5.314 \times 10^5/15} = 0.9509$$

- Standard error of full model coefficients

$$se(\hat{\beta}) = \sqrt{\frac{MSE}{n2^k}} = \sqrt{\frac{2252.56}{2 \times 8}} = 11.87$$

Example VI

- Confidence interval on model coefficients

$$\hat{\beta} - t_{\alpha/2, df_E} SE(\hat{\beta}) \leq \beta \leq \hat{\beta} + t_{\alpha/2, df_E} SE(\hat{\beta})$$