# Experimental Design Note 7-2 Blocking and confounding $2^K$ factorial design

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### Recall: Blocking in Factorial Design-Example I

### **Battery Life Experiment:**

An engineer is studying the effective lifetime of some battery. Two factors, plate material and temperature, are involved. There are three types of plate materials (1, 2, 3) and three temperature levels (15, 70, 125). Four batteries are tested at each combination of plate material and temperature, and all 36 tests are run in random order. The experiment and the resulting observed battery life data are given below in Table 5.1.

If we assume further that four operators (1, 2, 3, 4) were hired to conduct the experiment. It is known that different operators can cause systematic difference in battery lifetime. Hence operators should be treated as blocks. The blocking scheme is every operator conduct a single replicate of the full factorial design.

### Recall: Blocking in Factorial Design-Example II

For each treatment (treatment combination), the observations were in the order of the operators 1, 2, 3, and 4. This is a **blocked factorial design**.

### Statistical Model for Blocked Factorial Experiment I

#### Model:

$$y_{ijk} = \mu + \tau_i + \beta_j + (\tau \beta)_{ij} + \delta_k + \epsilon_{ijk}$$

for  $i=1,2,\cdots,a, j=1,2,\cdots,b$  and  $k=1,2,\cdots,n, \delta_k$  is the effect of the kth block.

- randomization restriction is imposed. (complete block factorial design).
- interactions between blocks and treatment effects are assumed to be negligible.

### Statistical Model for Blocked Factorial Experiment II

The previous ANOVA table for the experiment should be modified as follows:

ADD: Block Sum of Squares

$$SS_{Blocks} = \frac{1}{ab} \sum_{k} y_{\cdot \cdot k}^2 - \frac{y_{\cdot \cdot \cdot}^2}{abn}, \quad df = n - 1.$$

MODIFY: Error Sum of Squares

$$(new)SSE = (old)SSE - SS_{Blocks}, \quad df = (ab - 1)(n - 1).$$

other inferences should be modified accordingly.



### Statistical Model for Blocked Factorial Experiment III

■ TABLE 5.20 Analysis of Variance for a Two-Factor Factorial in a Randomized Complete Block

Source of Variation	Sum of Squares	Degrees of Freedom	Expected Mean Square	$F_0$
Blocks	$\frac{1}{ab}\sum_{k}y_{.k}^{2}-\frac{y_{.k}^{2}}{abn}$	n-1	$\sigma^2 + ab\sigma_\delta^2$	
Α	$\frac{1}{bn}\sum_{i}y_{i\perp}^{2}-\frac{y_{\perp}^{2}}{abn}$	a-1	$\sigma^2 + \frac{bn\sum \tau_i^2}{a-1}$	$\frac{MS_A}{MS_E}$
В	$\frac{1}{an}\sum_{j}y_{\downarrow}^{2}-\frac{y_{\perp}^{2}}{abn}$	b-1	$\sigma^2 + \frac{an\sum \beta_j^2}{b-1}$	$\frac{MS_B}{MS_E}$
AB	$\frac{1}{n}\sum_{i}\sum_{j}y_{ij}^{2}-\frac{y_{}^{2}}{abn}-SS_{A}-SS_{B}$	(a-1)(b-1)	$\sigma^2 + \frac{n\sum\sum(\tau\beta)_{ij}^2}{(a-1)(b-1)}$	$\frac{MS_{AB}}{MS_E}$
Error	Subtraction	(ab-1)(n-1)	$\sigma^2$	
Total	$\sum_{i} \sum_{j} \sum_{k} y_{ijk}^{2} - \frac{y_{}^{2}}{abn}$	abn-1		

### Statistical Model for Blocked Factorial Experiment IV

- This is a blocking replicated design.
- If there are *n* replicates of the design, then each replicate is a block.
- Each replicate is run in one of the blocks (time periods, batches of raw material, etc.).
- Runs within the block are randomized.
- How about unreplicated case?

### Confounding in blocks I

- Now consider the unreplicated case.
- Clearly the previous discussion does not apply, since there is only one replicate.
- To illustrate, consider the situation of Example 6.2, the pilot plant filtration rate experiment.
- This is a  $2^4$ , n = 1 replicate.

### Confounding in blocks II

#### Filtration Rate Experiment (revisited)

factor							
$\boldsymbol{A}$	B	$\boldsymbol{C}$	D	original response			
-	_	_	_	45			
+	_	_	_	71			
_	+	_	_	48			
+	+	_	_	65			
_	_	+	_	68			
+	_	+	_	60			
_	+	+	_	80			
+	+	+	_	65			
_	_	_	+	43			
+	_	_	+	100			
_	+	_	+	45			
+	+	_	+	104			
_	_	+	+	75			
+	_	+	+	86			
_	+	+	+	70			
+	+	+	+	96			

### Confounding in blocks III

- Suppose there are two batches of raw material. Each batch can be used for only 8 runs. It is known these two batches are very different. Blocking should be employed to eliminate this variability.
- How to select 8 treatments (level combinations, or runs) for each block?

### Confounding the 2<sup>k</sup> Factorial Design with two blocks I

Suppose there are two factors (A, B) each with 2 levels, and two blocks  $(b_1, b_2)$  each containing two runs (treatments). Since  $b_1$  and  $b_2$  are interchangeable, there are three possible blocking scheme:

			Blocking scheme		
Α	В	Response	1	2	3
-	-	<i>y</i>	$b_1$	$b_1$	<i>b</i> <sub>2</sub>
+	-	$y_{+-}$	$b_1$	$b_2$	$b_1$
-	+	$y_{-+}$	$b_2$	$b_1$	$b_1$
+	+	<i>y</i> ++	$b_2$	$b_2$	$b_2$

### Confounding the 2<sup>k</sup> Factorial Design with two blocks II

### Comparing blocking schemes:

#### Scheme 1:

- block effect:  $b = \bar{y}_{b_2} \bar{y}_{b_1} = \frac{1}{2} (-y_{--} y_{+-} + y_{-+} + y_{++})$
- main effect:  $B = \frac{1}{2} (-y_{--} y_{+-} + y_{-+} + y_{++})$
- B and b are not distinguishable or confounded.

#### Scheme 2:

- block effect:  $b = \bar{y}_{b_2} \bar{y}_{b_1} = \frac{1}{2} \left( -y_{--} + y_{+-} y_{-+} + y_{++} \right)$
- main effect:  $A = \frac{1}{2} (-y_{--} + y_{+-} y_{-+} + y_{++})$
- A and b are not distinguishable or confounded.

#### Scheme 3:

- block effect:  $b = \bar{y}_{b_2} \bar{y}_{b_1} = \frac{1}{2} (y_{--} y_{+-} y_{-+} + y_{++})$
- main effect:  $AB = \frac{1}{2}(y_{--} y_{+-} y_{-+} + y_{++})$
- AB and b are not distinguishable or confounded.

## Confounding the 2<sup>k</sup> Factorial Design with two blocks III

- The reason for confounding: the block arrangement matches the contrast of some factorial effect.
- Confounding makes the effect Inestimable.
- Question: which scheme is the best (or causes the least damage)?

## Confounding the $2^k$ Factorial Design with two blocks IV

### 2<sup>k</sup> Design with Two Blocks via Confounding

- Confound blocks with the effect (contrast) of the highest order
- Block 1 consists of all treatments with the contrast coefficient equal to -1
- Block 2 consists of all treatments with the contrast coefficient equal to 1

## Confounding the 2<sup>k</sup> Factorial Design with two blocks V

Example 1. Block 2<sup>3</sup> Design

factorial effects (contrasts)							
- 1	Α	В	С	AB `	AC	BC	ABC
1	-1	-1	-1	1	1	1	-1
1	1	-1	-1	-1	-1	1	1
1	-1	1	-1	-1	1	-1	1
1	1	1	-1	1	-1	-1	-1
1	-1	-1	1	1	-1	-1	1
1	1	-1	1	-1	1	-1	-1
1	-1	1	1	-1	-1	1	-1
1	1	1	1	1	1	1	1

Defining relation: b = ABC

Block 1: 
$$(---),(++-),(+-+),(-++)$$

Block 2: 
$$(+--),(-+-),(-++)$$

### Confounding the 2<sup>k</sup> Factorial Design with two blocks VI

Example 2: For  $2^4$  design with factors, A, B, C, D, the defining contrast (optimal) for blocking factor (b) is

$$b = ABCD$$

In general, the optimal blocking scheme for  $2^k$  design with two blocks is given by  $b = AB \cdots K$ , where  $A, B, \cdots, K$  are the factors.

# Confounding the 2<sup>k</sup> Factorial Design with two blocks VII

### Analyze Unreplicated Block 2<sup>k</sup> Experiment

Filtration Experiment (four factors: A, B, C, D):

- Use defining relation: b = ABCD, i.e., if a treatment satisfies ABCD = -1, it is allocated to block 1  $(b_1)$ ; if ABCD = 1, it is allocated to block 2  $(b_2)$ .
- Assume that, all the observations in block 2 will be reduced by 20 because of the poor quality of the second batch of material, i.e., the true block effect = −20.

# Confounding the 2<sup>k</sup> Factorial Design with two blocks VIII

	fac	tor		blocks	
$\boldsymbol{A}$	B	C	$\boldsymbol{D}$	b = ABCD	response
_	_	_	_	1= <b>b</b> 2	45-20=25
+	_	_	_	$-1=b_1$	71
_	+	_	_	-1= <b>b</b> 1	48
+	+	_	_	$1=b_2$	65-20=45
_	_	+	_	$-1=b_1$	68
+	_	+	_	1=b2	60-20=40
_	+	+	_	$1=b_2$	80-20=60
+	+	+	_	$-1=b_1$	65
_	_	_	+	$-1=b_1$	43
+	_	_	+	$1=b_2$	100-20=80
_	+	_	+	1=b2	45-20=25
+	+	_	+	-1=b <sub>1</sub>	104
_	_	+	+	$1=b_2$	75-20=55
+	_	+	+	$-1=b_1$	86
_	+	+	+	-1=b <sub>1</sub>	70
+	+	+	+	$1=b_2$	96-20=76

See Unreplicated-Block.SAS.

### 2<sup>k</sup> Factorial Design with Four Blocks I

Need two 2-level blocking factors to generate 4 different blocks. Confound each blocking factors with a high order factorial effect. The interaction between these two blocking factors matters. The interaction will be confounded with another factorial effect. In another word, for four blocks, select two effects to confound, automatically confounding a third effect.

Optimal blocking scheme has least confounding severity.

### 2<sup>k</sup> Factorial Design with Four Blocks II

 $2^4$  design with four blocks: factors are A, B, C, D and the blocking factors are b1 and b2

### 2<sup>k</sup> Factorial Design with Four Blocks III

### Possible blocking schemes:

Scheme 1: Defining relationships:  $b_1 = ABC$ ,  $b_2 = ACD$  induce confounding

$$b_1b_2 = ABC * ACD = A^2BC^2D = BD$$

Scheme 2: Defining relationships:  $b_1 = ABCD$ ,  $b_2 = ABC$  induce confounding

$$b_1b_2 = ABCD * ABC = D$$

which is better?

### Confounding 2<sup>k</sup> factorial design in 2<sup>p</sup> blocks I

- k factors:  $A, B, \dots, K$ , and p is usually much less than k.
- *p* blocking factors:  $b_1, b_2, \dots, b_p$  with levels -1 and 1.
- confound blocking factors with k chosen high-order factorial effects, i.e.,  $b_1 = \text{effect } 1$ ,  $b_2 = \text{effect } 2$ , etc. (p defining relations).
- These p defining relations induce another  $2^p p 1$  confounding.
- treatment combinations with the same values of  $b_1, \dots, b_p$  are allocated to the dame block. Within each block.
- each block consists of  $2^{k-p}$  treatment combinations (runs).
- Given k and p, optimal schemes are tabulated, e.g. Table 7.9

# Confounding $2^k$ factorial design in $2^p$ blocks II

■ TABLE 7.9

Suggested Blocking Arrangements for the 2k Factorial Design

		Effects Chosen to Generate the Blocks	Interactions Confounded with Blocks	
3	2	4	ABC	ABC
	4	2	AB, AC	AB, AC, BC
4	2	8	ABCD	ABCD
	4	4	ABC, ACD	ABC, ACD, BD
	8	2	AB, BC, CD	AB, BC, CD, AC, BD, AD, ABCD
5	2	16	ABCDE	ABCDE
	4	8	ABC, CDE	ABC, CDE, ABDE
	8	4	ABE, BCE, CDE	ABE, BCE, CDE, AC, ABCD, BD, ADE
	16	2	AB, AC, CD, DE	All two- and four-factor interactions (15 effects)
6	2	32	ABCDEF	ABCDEF
	4	16	ABCF, CDEF	ABCF, CDEF, ABDE
	8	8	ABEF, ABCD, ACE	ABEF, ABCD, ACE, BCF, BDE, CDEF, ADF
	16	4	ABF, ACF, BDF, DEF	ABF, ACF, BDF, DEF, BC, ABCD, ABDE, AD, ACDE, CE, CDF, BCDEF, ABCEF, AEF, BE
	32	2	AB, BC, CD, DE, EF	All two-, four-, and six-factor interactions (31 effects)
7	2	64	ABCDEFG	ABCDEFG
	4	32	ABCFG, CDEFG	ABCFG, CDEFG, ABDE
	8	16	ABC, DEF, AFG	ABC, DEF, AFG, ABCDEF, BCFG, ADEG, BCDEG
	16	8	ABCD, EFG, CDE, ADG	ABCD, EFG, CDE, ADG, ABCDEFG, ABE, BCG, CDFG, ADEF, ACEG, ABFG, BCEF, BDEG, ACF, BDF
	32	4	ABG, BCG, CDG,	ABG, BCG, CDG, DEG, EFG, AC, BD, CE, DF, AE,
			DEG, EFG	BF, ABCD, ABDE, ABEF, BCDE, BCEF, CDEF, ABCDEFG, ADG, ACDEG, ACEFG, ABDFG, ABCEG BEG, BDEFG, CFG, ADEF, ACDF, ABCF, AFG, BCDFG
	64	2	AB, BC, CD, DE, EF, FG	All two-, four-, and six-factor interactions (63 effects

### General advice about blocking I

- When in doubt, block.
- Block out the nuisance variables you know about, randomize as much as possible and rely on randomization to help balance out unknown nuisance effects.
- Measure the nuisance factors you know about but cannot control (ANCOVA).
- It may be a good idea to conduct the experiment in blocks even if there is not an obvious nuisance factor, just to protect against the loss of data or situations where the complete experiment can't be finished