

1) a)  $y|\theta=1 \sim N(1, \tau)$ ,  $y|\theta=2 \sim N(2, \tau)$

$$P(\theta=1) = 0.5, \quad P(\theta=2) = 0.5$$

$\Rightarrow$  let  $\theta$  be a random variable following a  $B(1, p)$ , where  $\theta \in \{1, 2\}$

$$\Rightarrow P(\theta) = \hat{p}^{\theta-1} (1-\hat{p})^{2-\theta}$$

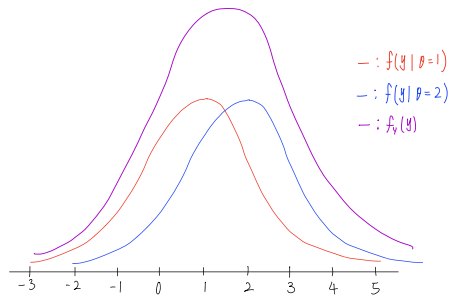
$$\Rightarrow f(y|\theta=1) = \frac{1}{\sqrt{2\pi\tau}} \exp\left(-\frac{(y-1)^2}{2\tau}\right)$$

$$f(y|\theta=2) = \frac{1}{\sqrt{2\pi\tau}} \exp\left(-\frac{(y-2)^2}{2\tau}\right), \text{ where } \tau=2$$

$$\Rightarrow f(y, \theta=1) = f(y|\theta=1)P(\theta=1) = \frac{1}{2\sqrt{2\pi}} \exp\left(-\frac{(y-1)^2}{4}\right), \text{ given } \tau=2$$

$$f(y, \theta=2) = f(y|\theta=2)P(\theta=2) = \frac{1}{2\sqrt{2\pi}} \exp\left(-\frac{(y-2)^2}{4}\right), \text{ given } \tau=2$$

$$\Rightarrow f_Y(y) = f(y, \theta=1) + f(y, \theta=2) = \frac{1}{2\sqrt{2\pi}} \exp\left(-\frac{(y-1)^2}{4}\right) + \frac{1}{2\sqrt{2\pi}} \exp\left(-\frac{(y-2)^2}{4}\right), \text{ mixture model}$$



b)  $P(y, \theta) = P(y|\theta)P(\theta)$ , where  $P(y|\theta) = \frac{1}{\sqrt{2\pi\tau}} \exp\left(-\frac{(y-1)^2}{2\tau}\right)I(\theta=1) + \frac{1}{\sqrt{2\pi\tau}} \exp\left(-\frac{(y-2)^2}{2\tau}\right)I(\theta=2)$  and  $P(\theta) = \hat{p}^{\theta-1} (1-\hat{p})^{2-\theta}$

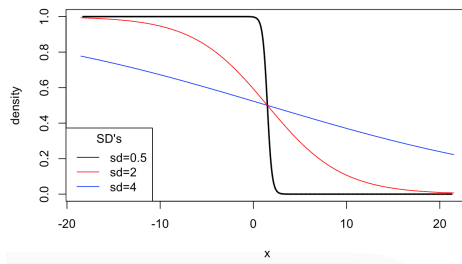
$$= \hat{p}^{2-\theta} (1-\hat{p})^{\theta-1} \left[ \frac{1}{\sqrt{2\pi\tau}} \exp\left(-\frac{(y-1)^2}{2\tau}\right)I(\theta=1) + \frac{1}{\sqrt{2\pi\tau}} \exp\left(-\frac{(y-2)^2}{2\tau}\right)I(\theta=2) \right]$$

$$\Rightarrow P(\theta|y) = \hat{p}^{2-\theta} (1-\hat{p})^{\theta-1} \left[ \frac{\frac{1}{\sqrt{2\pi\tau}} \exp\left(-\frac{(y-1)^2}{2\tau}\right)I(\theta=1) + \frac{1}{\sqrt{2\pi\tau}} \exp\left(-\frac{(y-2)^2}{2\tau}\right)I(\theta=2)}{\left[ \frac{1}{2\sqrt{2\pi\tau}} \exp\left(-\frac{(y-1)^2}{2\tau}\right) + \frac{1}{2\sqrt{2\pi\tau}} \exp\left(-\frac{(y-2)^2}{2\tau}\right) \right]} \right]$$

$$P(\theta=1|y=1) = P\left[\frac{1}{\sqrt{2\pi\tau}}\right] / \left[ \frac{1}{2\sqrt{2\pi\tau}} + \frac{1}{2\sqrt{2\pi\tau}} \exp\left(-\frac{1}{2}\right) \right], \quad p = \frac{1}{2}$$

$$= 0.53121$$

c)



$\Rightarrow$  As the standard deviation increases, the graph becomes flatter, which notes that the data has less information.

2) Find  $P(M)$  given  $P(G \cap G | M) = P$ ,  $P(B \cap B | M) = 1 - P$ ,  $P(G \cap B | M) = 0$ ,  $P(G \cap G | F) = P^2$ ,  $P(B \cap B | F) = (1 - P)^2$ ,  $P(G \cap B | F) = 2P(1 - P)$

$$P(G \cap G) = P(G \cap G \cap M) + P(G \cap G \cap F)$$

$$= P(G \cap G | M)P(M) + P(G \cap G | F)P(F)$$

$$\Rightarrow P(M) = \frac{P(G \cap G) - P(G \cap G | F)P(F)}{P(G \cap G | M)}$$

$\Rightarrow P(G \cap G)$  is unknown since we do not have any information about how associated they are.

$$\Rightarrow P(G \cap G | F)P(F) = P^2, \quad P(G \cap G | M) = P$$

$$\therefore P(M) = \frac{P(G \cap G) - P^2}{P}$$

3) a) Suppose  $\theta \sim \text{Beta}(\alpha, \beta)$ , then  $P(\theta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1}(1-\theta)^{\beta-1}$

$$\begin{aligned} \Rightarrow P(X, \theta) &= P(X|\theta)P(\theta) = \frac{(x+r-1)!}{x!(r-1)!} \theta^x(1-\theta)^{r-x} \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1}(1-\theta)^{\beta-1} \\ &= \frac{1}{P(x, r-1)} \cdot \frac{1}{\beta(\alpha, \beta)} \theta^{r+\alpha-1}(1-\theta)^{x+\beta-1} \propto \theta^{r+\alpha-1}(1-\theta)^{x+\beta-1} \end{aligned}$$

$\Rightarrow P(\theta|X) \sim \text{Gamma}(r+\alpha, x+\beta)$   $\therefore P(\theta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1}(1-\theta)^{\beta-1}$  is a conjugate prior

b) Let  $\beta \sim T(d, C)$ , then  $P(\beta) = \frac{C^d}{\Gamma(d)} \beta^{d-1} e^{-C\beta}$

$$\Rightarrow P(X, \beta) = P(X|\beta)P(\beta) = \frac{\beta^x}{\Gamma(x)} x^{\alpha-1} e^{-\beta x} \frac{C^d}{\Gamma(d)} \beta^{d-1} e^{-C\beta}$$

$$\begin{aligned} \Rightarrow P(X) &= \int_0^\infty P(X|\beta)P(\beta) d\beta \\ &= \int_0^\infty \frac{\beta^x}{\Gamma(x)} x^{\alpha-1} e^{-\beta x} \frac{C^d}{\Gamma(d)} \beta^{d-1} e^{-C\beta} d\beta \\ &= \frac{x^{\alpha-1} C^d}{\Gamma(x)\Gamma(d)} \frac{\Gamma(x+d)}{(x+d)^{x+d}} \int_0^\infty \frac{(x+d)^{x+d}}{\Gamma(x+d)} \beta^{x+d-1} e^{-\beta(x+d)} d\beta \\ &= \frac{\Gamma(x+d)}{\Gamma(x)\Gamma(d)} \frac{1}{(x+d)^{x+d}} x^{\alpha-1} C^d \\ \Rightarrow P(\beta|X) &= \frac{P(X, \beta)}{P(X)} = \frac{\frac{\beta^x}{\Gamma(x)} x^{\alpha-1} e^{-\beta x} \frac{C^d}{\Gamma(d)} \beta^{d-1} e^{-C\beta}}{\frac{\Gamma(x+d)}{\Gamma(x)\Gamma(d)} \frac{1}{(x+d)^{x+d}} x^{\alpha-1} C^d} \\ &= \frac{\beta^{x+d-1}}{\Gamma(x+d)} (x+d)^{x+d} e^{-(x+d)\beta} \propto T(x+d, x+d) \end{aligned}$$

$\therefore P(\beta) = \frac{C^d}{\Gamma(d)} \beta^{d-1} e^{-C\beta}$  is a conjugate prior

c) 포아송 분포, 지수 분포, 이항분포, 감마분포, 베타분포 등등 모두 해당되으나 나열지 않음.

4) a) Let  $X$  be the number of Californians supporting the death penalty, then  $P(X|\theta) = \binom{n}{x} \theta^x (1-\theta)^{n-x} I(0 \leq x \leq n)$ .

$\Rightarrow$  By the given conditions,  $\theta \sim \text{Beta}(\alpha, \beta)$ , so  $P(\theta) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1}$

Given that  $E(\theta) = 0.6$ ,  $\text{Var}(\theta) = 0.09$ , it is notable  $\frac{\alpha}{\alpha+\beta} = 0.6$  and  $\frac{\alpha\beta}{(\alpha+\beta+1)(\alpha+\beta)^2} = 0.09$

$$\Rightarrow \alpha = \frac{3}{2}\beta \Rightarrow \frac{\frac{3}{2}\beta^2}{(\frac{3}{2}\beta+\beta+1)(\frac{3}{2}\beta+\beta)^2} = \frac{\beta^2}{(\frac{5\beta+2}{2})(\frac{5\beta+1}{2})} = \frac{12\beta^2}{125\beta^2+50\beta} = \frac{9}{100}$$

$$1200\beta^2 = 1125\beta^2 + 450\beta$$

$$0 = 1125\beta^2 - 450\beta$$

$$= \beta^2(9\beta - 6) \quad \therefore \beta = \frac{2}{3}, \alpha = 1$$

b)  $P(X|\theta) = \binom{n}{x} \theta^x (1-\theta)^{n-x} I(0 \leq x \leq n)$ ,  $P(\theta) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1}$

$$\Rightarrow P(\theta|X) \propto \theta^{x+\alpha-1} (1-\theta)^{n-x+\beta-1}$$

$$\Rightarrow \theta|X \sim \text{Beta}(x+\alpha, n-x+\beta), E(\theta|X) = \frac{x+\alpha}{n+\alpha+\beta}, \text{Var}(\theta|X) = \frac{(x+\alpha)(n-x+\beta)}{(n+\alpha+\beta)^2(n+\alpha+\beta+1)}$$

It is given that  $n=1000$ , and  $\hat{\theta}=0.65$  implies  $x=650$ , and using the results of  $\alpha, \beta$  from (a),

$$E(\theta|X) = \frac{x+\alpha}{n+\alpha+\beta} = \frac{650+1}{1000+1+\frac{2}{3}} = 0.64992$$

$$\text{Var}(\theta|X) = \frac{(x+\alpha)(n-x+\beta)}{(n+\alpha+\beta)^2(n+\alpha+\beta+1)} = \frac{(650+1)(1000-650+\frac{2}{3})}{(1000+1+\frac{2}{3})^2(1000+1+\frac{2}{3}+1)} = 0.000227$$