Chapter 7: Forecasting

Time Series Analysis WISE, XMU

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§7.1 What is forecasting?

- ▶ There are two components of forecasting:
 - 1. One of the primary objectives of building a model for a time series is to be able to **forecast** the values for that series at future times.
 - Of equal importance is the assessment of the precision of those forecasts, measured by the **prediction limits** or confidence intervals.
- For a time series $\{Z_t\}$, we assume that the model is known **exactly**, including specific values for all parameters. Although this is not true in practice, the use of estimated parameters for large sample sizes does not seriously affect the results.

- ▶ Based on the available history of the series up to time n, namely $Z_1, ..., Z_n$, we want to forecast the value of Z_{n+1} , where n is called the **forecast origin** and I the **lead time**, and denote the forecasting value by $\widehat{Z}_n(I)$.
 - 1. By the criterion of the **minimum mean square error**, the forecasting value $\widehat{Z}_n(I)$ is just the conditional expectation,

$$\widehat{Z}_n(I) = E(Z_{n+I}|Z_n,...,Z_1).$$

2. The **forecast error** is defined as $a_n(I) = Z_{n+I} - \widehat{Z}_n(I)$, and the **forecast error variance** is $\text{var}[a_n(I)]$, which is also a conditional variance.

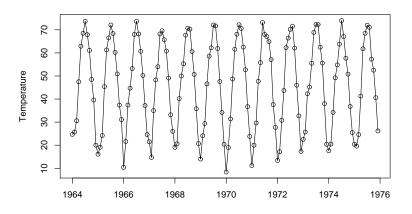
3. If the normality Z_t or a_t is also verified, then, conditional on $Z_n, ..., Z_1$, the random variable Z_{n+1} will follow normal

distribution with mean $\widehat{Z}_n(I)$ and variance $\text{var}[a_n(I)]$. Then the **prediction limit** with significance level of $1-\alpha$ will be

$$\widehat{Z}_n(I) \pm z_{1-\alpha/2} \sqrt{\text{var}[a_n(I)]},$$
 where $z_{1-\alpha/2}$ is $1-\alpha/2$ percentile of the standard normal

where $z_{1-\alpha/2}$ is $1-\alpha/2$ percentile of the standard normal distribution. Hence, there are **only two** important quantities, $\widehat{Z}_n(I)$ and $\text{var}[a_n(I)]$ (or $a_n(I)$), in time series forecasting.

► For example, consider the average monthly temperatures (in degrees Fahrenheit) in Dubuque, Iowa from January, 1964 to December 1975.



Suppose we have the cosine model, $Z_t = \mu_t + a_t$, to explain the data, where $\sigma_2^2 = 3.706^2$ and

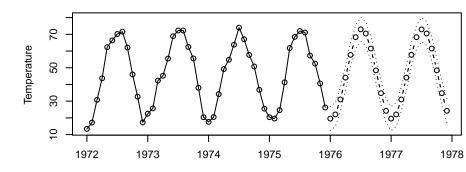
$$\mu_t = 46.2660 + (-26.7079)\cos(2\pi \frac{t-1}{12}) + (-2.1697)\sin(2\pi \frac{t-1}{12}).$$

Note that the observed sequence is $Z_1, ..., Z_{144}$. We want to make forecasting for the value at June, 1976, which is corresponding to the time point 150. Then,

$$Z_{144}(6) = E(Z_{150}|Z_{144},...,Z_1) = \mu_{150} = 68.3 \, {}^{\circ}F,$$

 $a_{144}(6) = a_{150}$, $var[a_{144}(6)] = 3.706^2$, and the 95% confidence interval will be

$$[68.3 - 1.96 \times 3.706, 68.3 + 1.96 \times 3.706] = [61.0, 75.6].$$



Year

§7.2 Forecasting with ARIMA models

► Consider the AR(1) model,

$$Z_t = \phi Z_{t-1} + a_t.$$

1. $Z_{n+1} = \phi Z_n + a_t$. Hence the **one-step-ahead forecast** is

$$\widehat{Z}_n(1) = E(Z_{n+1}|Z_n,...,Z_1) = E(\phi Z_n + a_{n+1}|Z_n,...,Z_1)$$

= $\phi Z_n + E(a_{n+1}|Z_n,...,Z_1) = \phi Z_n.$

$$var(a_n(1)) = var(a_t) = \sigma^2$$
.

2. For the *I*-step forecast,

 $Z_{n+l} = \phi Z_{n+l-1} + a_{n+l} = \phi(\phi Z_{n+l-2} + a_{n+l-1}) + a_{n+1} = \phi^2 Z_{n+l-1}$

 $= \phi^{l} Z_{n} + \phi^{l-1} a_{n+1} + \cdots + \phi a_{n+l-1} + a_{n+l}$

Remark: Since $|\phi| < 1$, for large I, we have that

 $var(a_n(I)) = var(\phi^{l-1}a_{n+1} + \cdots + a_{n+l}) = [\phi^{2(l-1)} + \cdots + 1]\sigma^2.$

 $\widehat{Z}_n(I) \approx 0 = E(Z_t)$ and $var(a_n(I)) \approx \sigma^2/(1-\phi^2) = var(Z_t)$.

- 3. In general, the forecast could be conducted as below.
 - (a) Rewrite Z_{n+l} recursively until you express it using Z_t with $t \le n$.
 - (b) Then calculate \hat{Z}_{n+l} and $var(a_n(l))$. Note that to forecast \hat{Z}_{n+l} , the steps above is equivalent to

Note that to forecast \hat{Z}_{n+1} , the steps above is equivalent to that we first predict Z_{n+1} , and use it to predict Z_{n+2} until we forecast Z_{n+1} .

To derive $var(a_n(I))$, it is equivalent to using the MA representation. Say we have $Z_{n+I} = \sum_{j=0}^{\infty} \psi_j a_{n+I-j}$ Then the forecast error is

$$a_n(I) = a_{n+1} + \psi_1 a_{n+l-1} + \cdots + \psi_{l-1} a_{n+1}.$$

The forecast error variance is

$$var[a_n(I)] = \sigma_a^2(1 + \psi_1^2 + \dots + \psi_{I-1}^2).$$

- **Note that:** The forecast error only depends on the form of MA representations, i.e. the results will be same for general ARMA models.
- 4. The **one-step-ahead forecast error** is $a_n(1) = a_{n+1}$. Hence, the white noise process $\{a_t\}$ can now be reinterpreted as a sequence of one-step-ahead forecast errors.

► Consider the MA(1) model,

$$Z_t = a_t - \theta a_{t-1}$$
.

1. $Z_{n+1} = a_{n+1} - \theta a_n$. The one-step-ahead forecast is

$$Z_n(1) = E(a_{n+1} - \theta a_n | Z_n, ..., Z_1) = -\theta a_n,$$

the forecast error is $a_n(1) = Z_{n+1} - Z_n(1) = a_{n+1}$ and the forecast error variance is $var[a_n(1)] = \sigma_a^2$.

2. The *I*-step (I > 1) forecast is

$$Z_n(I) = E(a_{n+1} - \theta a_{n+l-1} | Z_n, ..., Z_1) = 0,$$

the forecast error is $a_n(I) = Z_{n+l} - Z_n(I) = a_{n+l} - \theta a_{n+l-1} = Z_{n+l}$ and the forecast error variance is $var[a_n(I)] = \sigma_a^2(1 + \theta^2) = var(Z_{n+l})$.

- 3. How to calculate a_n since it appears in the one-step-ahead forecast?
- ► We can calculate them by

$$a_t = \sum_{i=0}^{\infty} \pi_j a_{t-j} = \sum_{i=0}^{t-1} \pi_j Z_{t-j}, \quad 1 \le t \le n,$$

where the initial values $Z_s = 0$ for $s \le 0$. Note that the invertibility is needed here.

- We may alternatively calculate them by the iterative method.
 - Suppose that $a_0 = 0$, and then

$$a_1 = Z_1 + \theta a_0 = Z_1, \ a_2 = Z_2 + \theta a_1, ..., \ a_n = Z_n + \theta a_{n-1}.$$

4. The method to forecast MA(q) models is similar.

▶ Consider the ARMA(1,1) model, $Z_t = \phi Z_{t-1} + a_t + \theta_1 a_{t-1}$.

1.
$$Z_{n+1} = \phi Z_n + a_{n+1} + \theta_1 a_n$$
. For one-step forecasting,

$$\widehat{Z}_n(1) = E(Z_{n+1}|Z_n,...,Z_1)$$

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 $= \phi Z_n + \theta_1 a_n$ $a_n(1) = Z_{n+1} - \widehat{Z}_n(1) = a_{n+1},$

 $\operatorname{var}[a_n(1)] = \sigma_n^2$

 $= \phi Z_n + \theta_1 a_n + E(a_{n+1}|Z_n,...,Z_1)$

$$\widehat{Z}_n(1) = E(Z_{n+1}|Z_n,...,Z_1)$$

= $E(\phi Z_n + a_{n+1} + \theta_1 a_n | Z_n,...,Z_1)$

2. For two-step forecasting, you could directly calculate $\widehat{Z}_n(2) = \phi \widehat{Z}_n(1) = \phi(\phi Z_n + \theta_1 a_n)$. Or we first find out that

$$Z_n(2)=\phi Z_n(1)=\phi(\phi Z_n+ heta_1 a_n).$$
 Or we first find out that
$$Z_n(2)=\phi Z_{n+1}+a_{n+2}+ heta_1 a_{n+1}$$

 $= \phi(\phi Z_n + \theta_1 a_n) + \phi a_{n+1} + a_{n+2} + \theta_1 a_{n+1}$

$$Z_{n}(2) = \phi Z_{n+1} + a_{n+2} + \theta_{1} a_{n+1}$$

$$= \phi (\phi Z_{n} + a_{n+1} + \theta_{1} a_{n}) + a_{n+2} + \theta_{1} a_{n+1}$$

$$= \phi (\phi Z_{n} + a_{n+1} + \theta_{1} a_{n}) + a_{n+2} + \theta_{1} a_{n+1}$$

 $\widehat{Z}_n(2) = \phi(\phi Z_n + \theta_1 a_n)$

 $var[a_n(2)] = [1 + (\phi + \theta_1)^2]\sigma_2^2$

3. For three-step, four-step, · · · .

 $a_n(2) = a_{n+2} + (\phi + \theta_1)a_{n+1}$

Example For each of the following models,

$$\begin{array}{ll} (i) & (1-0.9B)Z_t=a_t & --\mathsf{AR}(1) \text{ model with mean } 0 \\ (ii) & Z_t=10+a_t-0.9a_{t-1} & --\mathsf{MA}(1) \text{ model with mean } 10 \end{array}$$

where $\sigma_a^2=2$. Given $Z_1=1.2$ and $Z_2=0.1$, find the I-step ahead forecast values and forecast variances for I=1,2,3 (Suppose $a_1=0.3, a_2=-1.2$).

Solution:

Reference

Please read Chapter 9 of Cryer & Chan (2008).