

# Time Series Analysis

## chapter 7—Forecasting

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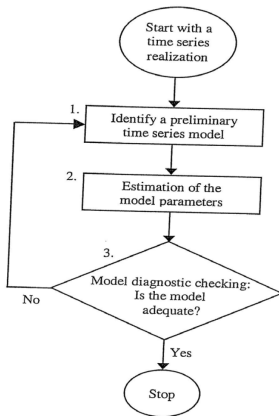
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- Model Steps
- Model Specification
- Model Estimation
- Model Diagnostic Checking
- Model Selection
- Forecasting

# Model Fitting (In-sample Fitting)

All models are wrong, but some are useful—George, E.P.Box



Chapter4\_strategy.eps

1919-2013

# Model Forecasting (Out-of-Sample Validation/Evaluation)

- Predict future values of a time series,  $y_{t+h}$ ,  $h = 1, 2, \dots$ , based on data to the present  $I_t = \{y_t, y_{t-1}, \dots\}$ .
- Let  $y_{t+h|t}$  denote a forecast of  $y_{t+h}$  made at time  $t$ , which has an associated *forecast error* or *prediction error*  $\varepsilon_{t+h|t}$ ,

$$\varepsilon_{t+h|t} = y_{t+h} - y_{t+h|t}.$$

Obviously, many different forecasts  $y_{t+h|t}$  exists.

- Find the *optimal forecast*  $y_{t+h|t}$  to *minimizes the mean squared error of the forecast*,

$$\text{MSE}(\varepsilon_{t+h|t}) \equiv E[\varepsilon_{t+h|t}^2] = E[(y_{t+h} - y_{t+h|t})^2].$$

- The *minimum MSE forecast* (best forecast) of  $y_{t+h}$  based on  $I_t$  is  $E[y_{t+h}|I_t]$ . (Why?)

Proof:

$$\begin{aligned} E[(y_{t+h} - y_{t+h|t})^2] &= E \{ [y_{t+h} - E(y_{t+h}|I_t) + E(y_{t+h}|I_t) - y_{t+h|t}]^2 \} \\ &= E \{ [y_{t+h} - E(y_{t+h}|I_t)]^2 \} + 2E \{ [y_{t+h} - E(y_{t+h}|I_t)][E(y_{t+h}|I_t) - y_{t+h|t}] \} \\ &\quad + E \{ [E(y_{t+h}|I_t) - y_{t+h|t}]^2 \} \\ &= E \{ [y_{t+h} - E(y_{t+h}|I_t)]^2 \} + E \{ [E(y_{t+h}|I_t) - y_{t+h|t}]^2 \} \end{aligned}$$

Taking the minimization with different values of  $y_{t+h|t}$ , we have

$$y_{t+h|t} = E(y_{t+h}|I_t).$$

Thus, the minimum MSE is  $E[(y_{t+h} - y_{t+h|t})^2] = E[(y_{t+h} - E(y_{t+h}|I_t))^2]$ .

# Forecasting AR( $p$ ) model

**1-Step-Ahead Forecast:** From the AR( $p$ ) model, we have

$$y_{t+1} = c + \phi_1 y_t + \cdots + \phi_p y_{t+1-p} + \varepsilon_{t+1}, \quad \varepsilon_t \sim \text{IID}(0, \sigma^2).$$

- Under the *MSE loss function*, the *point forecast* of  $y_{t+1}$  given  $I_t$  is

$$y_{t+1|t} = E[y_{t+1}|I_t] = c + \phi_1 y_t + \cdots + \phi_p y_{t+1-p},$$

- The *1-step-ahead forecast error* is

$$\varepsilon_{t+1|t} = y_{t+1} - y_{t+1|t} = \varepsilon_{t+1}.$$

- The variance of *1-step-ahead forecast error* is  $\text{Var}[\varepsilon_{t+1|t}] = \text{Var}(\varepsilon_{t+1}) = \sigma^2$ .
- If  $\varepsilon_t$  is *normally distributed*, then a *95% 1-step-ahead interval forecast* of  $y_{t+1}$  is  $y_{t+1|t} \pm 1.96\sigma$ .

# Forecasting AR( $p$ ) model

## 2-Step-Ahead Forecast:

From the AR( $p$ ) model, we have

$$y_{t+2} = c + \phi_1 y_{t+1} + \cdots + \phi_p y_{t+2-p} + \varepsilon_{t+2}. \quad (1)$$

- Taking conditional expectation, we have

$$y_{t+2|t} = E[y_{t+2}|I_t] = c + \phi_1 y_{t+1|t} + \phi_2 y_t + \cdots + \phi_p y_{t+2-p}, \quad (2)$$

- The associated forecast error

$$\varepsilon_{t+2|t} = y_{t+2} - y_{t+2|t} = \phi_1 (y_{t+1} - y_{t+1|t}) + \varepsilon_{t+2} = \varepsilon_{t+2} + \phi_1 \varepsilon_{t+1}.$$

- The variance of the forecast error is  $\text{Var}[\varepsilon_{t+2|t}] = (1 + \phi_1^2)\sigma^2$ .
- If  $\varepsilon_t$  is normally distributed, then a 95% 1-step-ahead interval forecast of  $y_{t+1}$  is  $y_{t+2|t} \pm 1.96\sigma\sqrt{1 + \phi_1^2}$ .
- It is interesting to see that  $\text{Var}[\varepsilon_{t+2|t}] \geq \text{Var}[\varepsilon_{t+1|t}]$ , meaning that as the forecast horizon increases the uncertainty in forecast also increases.

# Forecasting AR( $p$ ) model

**Multistep-Ahead Forecast:** In general, we have

$$y_{t+h} = c + \phi_1 y_{t+h-1} + \cdots + \phi_p y_{t+h-p} + \varepsilon_{t+h}. \quad (3)$$

- Taking conditional expectation, we have

$$y_{t+h|t} = E[y_{t+h}|I_t] = c + \sum_{i=1}^p \phi_i y_{t+h-i|t}, \quad (4)$$

where  $y_{t+\ell|t} = y_{t+\ell}$  if  $\ell \leq 0$ . This forecast can be computed recursively using forecasts  $y_{t+i|t}$  for  $i = 1, \dots, h-1$ .

- The  $h$ -step-ahead forecast error is

$$\varepsilon_{t+h|t} = y_{t+h} - y_{t+h|t}.$$

- How to obtain the forecasting interval for the AR( $p$ ) model?



# Forecasting MA( $q$ ) models

Consider the MA( $q$ ) model  $y_t = \sum_{i=0}^q \theta_i \varepsilon_{t-i}$ , with  $\theta_0 \equiv 1$ . Given the *IID* properties of  $\varepsilon_t$ , the *optimal forecast* is

$$y_{t+h|t} = \begin{cases} \sum_{i=h}^q \theta_i \varepsilon_{t+h-i}, & \text{for } h = 1, \dots, q, \\ 0, & \text{for } h > q, \end{cases}$$

whereas the corresponding forecast error follows as

$$\varepsilon_{t+h|t} = \begin{cases} \sum_{i=0}^{h-1} \theta_i \varepsilon_{t+h-i}, & \text{for } h = 1, \dots, q, \\ \sum_{i=0}^q \theta_i \varepsilon_{t+h-i}, & \text{for } h > q, \end{cases}$$

which can be simplified to  $\varepsilon_{t+h|t} = \sum_{i=0}^{h-1} \theta_i \varepsilon_{t+h-i}$  by defining  $\theta_i \equiv 0$  for  $h > q$ .

# Forecasting MA( $q$ ) models

Given the assumptions on  $\varepsilon_t$ , it follows that

$$E[\varepsilon_{t+h|t}] = 0,$$

and for *the mean squared error (MSE) of the forecast*

$$\text{MSE}(\varepsilon_{t+h|t}) = E[\varepsilon_{t+h|t}^2] = \sigma^2 \sum_{i=0}^{h-1} \theta_i^2.$$

Assuming normality, a 95% *forecasting interval* for  $y_{t+h}$  is bounded by

$$\left( y_{t+h|t} - 1.96 \cdot \text{RMSE}(\varepsilon_{t+h|t}), \quad y_{t+h|t} + 1.96 \cdot \text{RMSE}(\varepsilon_{t+h|t}) \right),$$

where  $\text{RMSE}(\varepsilon_{t+h|t})$  denotes the *square root* of  $\text{MSE}(\varepsilon_{t+h|t})$ .

# Forecasting ARMA( $p, q$ ) model:

The ARMA( $p, q$ ) model for  $y_t$  is

$$y_t = \phi_1 y_{t-1} + \cdots + \phi_p y_{t-p} + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \cdots + \theta_q \varepsilon_{t-q},$$

or, in lag operator form  $\phi_p(B)y_t = \theta_q(B)\varepsilon_t$ .

- The true forecast value is

$$y_{t+h} = \phi_1 y_{t+h-1} + \cdots + \phi_p y_{t+h-p} + \varepsilon_{t+h} + \theta_1 \varepsilon_{t+h-1} + \cdots + \theta_q \varepsilon_{t+h-q}.$$

- The minimum MSE forecast for  $y_{t+h}$  is

$$y_{t+h|t} = \phi_1 y_{t+h-1|t} + \cdots + \phi_p y_{t+h-p|t} + \theta_1 \varepsilon_{t+h-1|t} + \cdots + \theta_q \varepsilon_{t+h-q|t},$$

where  $y_{t+\ell|t} = y_{t+\ell}$  if  $\ell \leq 0$ ; and  $\varepsilon_{t+\ell|t} = 0$  if  $\ell > 0$ , otherwise  $\varepsilon_{t+\ell|t} = \varepsilon_{t+\ell}$ .

# Forecasting the ARMA( $p, q$ ) models by MA( $\infty$ ) representation

- It is convenient to rewrite the model as an MA( $\infty$ ) model, that is,  
 $y_t = \phi_p(L)^{-1} \theta_q(L) \varepsilon_t$  or

$$y_t = \varepsilon_t + \eta_1 \varepsilon_{t-1} + \eta_2 \varepsilon_{t-2} + \eta_3 \varepsilon_{t-3} + \cdots.$$

- The minimum MSE linear forecast (best linear predictor) of  $y_{t+h}$  based on  $I_t$  is

$$y_{t+h|t} = \eta_h \varepsilon_t + \eta_{h+1} \varepsilon_{t-1} + \cdots, \quad \text{with } \eta_0 \equiv 1.$$

[See Hamilton (1994) page 74].

- From which it follows that the *h-step-ahead prediction error* is given by

$$\varepsilon_{t+h|t} = \varepsilon_{t+h} + \eta_1 \varepsilon_{t+h-1} + \cdots + \eta_{h-1} \varepsilon_{t+1} = \sum_{i=0}^{h-1} \eta_i \varepsilon_{t+h-i},$$

and the MSE of the forecast error is

$$\text{MSE}(\varepsilon_{t+h|t}) = \sigma^2 \sum_{i=0}^{h-1} \eta_i^2,$$