UGC/MATH 2018 (June set-a)-Q.106

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Question

Let X_i , $i \ge 1$, be a sequence of i.i.d. random variables with $E(X_i) = 0$ and $V(X_i) = 1$. Which of the following are true?

Definitions

Definition

A sequence of random variables Y, Y_1 , Y_2 ... converges in distribution to a random variable Y, if

$$\lim_{n\to\infty} F_{Y_n}(a) = F_Y(a) \ \forall a\in\mathbb{R}. \tag{1}$$

Definition

A sequence of random variables Y, Y_1 , Y_2 ... is said to converge in probability to Y, if

$$\lim_{n \to \infty} \Pr(Y_n - Y > \epsilon) = 0 \ \forall \epsilon > 0.$$
 (2)



Lemmas

Lemma

If $Y_n \to Y$ in probability, $Y_n \to Y$ in distribution.

Strong Law of Large Numbers

Let $X_1, X_2, ... X_n$ be i.i.d. random variables with expected value $E(X_i) = \mu < \infty$, then,

$$\lim_{n \to \infty} \Pr\left(\left| \frac{1}{n} \sum_{i=1}^{n} X_i - \mu \right| \ge \epsilon \right) = 0$$
 (3)

Or, $\frac{1}{n}\sum_{i=1}^{n} X_i$ converges in probability to μ .



Lemma

If X_i is a sequence of i.i.d. random variables, satisfying condition

$$F_{X_1}(x) = F_{X_2}(x) = \dots = F_{X_n}(x) = F_X(x)$$
 (4)

then,

$$F_{X_1^2}(x) = F_{X_2^2}(x) = \dots = F_{X_n^2}(x) = F_{X_n^2}(x)$$
 (5)

 $\forall x \in \mathbb{R}$ where $F_X(x)$ is the c.d.f. of X_i .

Proof

 X_i is a sequence of i.i.d. random variables, which means it satisfies the following condition.

$$F_{X_1}(x) = F_{X_2}(x) = \dots = F_{X_n}(x) = F_X(x)$$
 (6)

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Proof contd.

Let $Y_i = X_i^2$. For $y \ge 0$,

$$F_{Y_i}(y) = \Pr\left(Y_i \le y\right) = \Pr\left(X_i^2 \le y\right) \tag{7}$$

$$\implies F_{Y_i}(y) = \Pr\left(-\sqrt{y} \le X_i \le \sqrt{y}\right) \tag{8}$$

$$\implies F_{Y_i}(y) = \Pr\left(X_i \le \sqrt{y}\right) - \Pr\left(X_i \le -\sqrt{y}\right) = F_{X_i}(\sqrt{y}) - F_{X_i}(-\sqrt{y}) \tag{9}$$

Using (6),

$$F_{Y_i}(y) = F_X(\sqrt{y}) - F_X(-\sqrt{y}) \tag{10}$$

From (10),

$$F_{Y_1}(y) = F_{Y_2}(y) = \dots = F_{Y_n}(y) = F_Y(y)$$
 (11)

where $F_Y(y)$ is the c.d.f. of $Y_i = X_i^2$.

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Lemma

If X_i is a sequence of i.i.d. random variables, satisfying condition

$$F_{X_1,...X_n}(x_1...x_n) = F_X(x_1)F_X(x_2)...F_X(x_n)$$
 (12)

where $F_X(x)$ is the c.d.f. of X_i , then for $Y_i = X_i^2$

$$F_{Y_1,Y_2,...,Y_n}(y_1,y_2,...,y_n) = F_Y(y_1)F_Y(y_1)...F_Y(y_n)$$
 (13)

where $F_Y(y)$ is the c.d.f. of $Y_i = X_i^2$.

Proof

Let $Y_i = X_i^2$. Now, for $y_i \ge 0$, consider

$$F_{Y_1,Y_2,...,Y_n}(y_1,y_2,...,y_n) = \Pr(Y_1 \le y_1, Y_2 \le y_2,..., Y_n \le y_n)$$

$$= \Pr\left(X_1^2 \le y_1, \dots, X_n^2 \le y_n\right) \tag{14}$$

$$= \Pr\left(-\sqrt{y_1} \le X_1 \le \sqrt{y_1}, \dots, -\sqrt{y_n} \le X_n \le \sqrt{y_n}\right) \tag{15}$$

Proof contd.

Since X_1, X_2, \dots, X_n are independent,

$$F_{Y_{1},Y_{2},...,Y_{n}}(y_{1}, y_{2},...,y_{n}) = Pr(-\sqrt{y_{1}} \le X_{1} \le \sqrt{y_{1}}) Pr(-\sqrt{y_{2}} \le X_{2} \le \sqrt{y_{2}}) ... Pr(-\sqrt{y_{n}} \le X_{n} \le \sqrt{y_{n}})$$
(16)

From (8) and (11),

$$F_{Y_1,Y_2,...,Y_n}(y_1,y_2,...,y_n) = F_{Y_1}(y_1)F_{Y_2}(y_2)...F_{Y_n}(y_n)$$
 (17)

$$= F_Y(y_1)F_Y(y_2)\dots F_Y(y_n)$$
 (18)

So,

$$F_{Y_1,Y_2,...,Y_n}(y_1,y_2,...,y_n) = F_Y(y_1)F_Y(y_1)...F_Y(y_n)$$
 (19)

Some other lemmas

Lemma

If X_i is a sequence of i.i.d. random variables, it follows that X_i^2 is also a sequence of i.i.d. random variables.

Proof.

From Lemma 4 and Lemma 5, X_i^2 is also a sequence of i.i.d. random variables.

Lemma

If $X_1, X_2, \dots X_n$ are independent random variables,

$$V(\sum_{i=1}^{n} X_i) = \sum_{i=1}^{n} V(X_i)$$
 (20)

Chebyshev's Inequality

Let the random variable X have a finite mean μ and a finite variance σ^2 . For every $\epsilon>0$,

$$\Pr(X - \mu \ge \epsilon) \le \frac{\sigma^2}{\epsilon^2} \tag{21}$$

Central Limit Theorem

Let $X_1, X_2, \dots X_n$ be i.i.d.r.v with $E(X_i) = \mu$ and $V(X_i) = \sigma^2$. Then

$$Z_n = \frac{\bar{x} - \mu}{\frac{\sigma}{\sqrt{n}}} = \frac{\sum_{i=1}^n X_i - n\mu}{\sqrt{n}\sigma}$$
 (22)

converges in distribution to standard normal random variable as n $\to \infty$

$$\lim_{n\to\infty} \Pr(Z_n \le a) = \Phi(a) \ \forall a \in \mathbb{R}. \tag{23}$$

where $\Phi(a)$ is the standard normal CDF.

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From Lemma 6, $\{X_i^2\}$ is a sequence of i.i.d. random variables. Now, we know,

$$E(X_i^2) = V(X_i) + (E(X_i))^2$$
 (24)

Putting given values, we get,

$$E(X_i^2) = 1 \tag{25}$$

From Lemma 4,

$$\frac{1}{n} \sum_{i=1}^{n} X_i^2 \to 1 \tag{26}$$

Therefore, option 1 is incorrect.



Lemma

If $\{X_i\}$ is a sequence of i.i.d. random variables with $E(X_i)=0$ and $V(X_i)=1$, then the random variable $Y_n=\frac{1}{n^{3/4}}\sum_{i=1}^n X_i$. has $E(Y_n)=0$ and $V(Y_n)=\frac{1}{n^{1/2}}$

Proof

$$E(Y_n) = \frac{1}{n^{3/4}} E \sum_{i=1}^n X_i = 0$$
 (27)

Since
$$E(X_i) = 0$$
. Now, $V(Y_n) = V \frac{1}{n^{3/4}} \sum_{i=1}^n X_i = \frac{1}{n^{3/2}} V \sum_{i=1}^n X_i$

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Proof contd.

From Lemma 7, and since $V(X_i) = 1$,

$$\implies V(Y_n) = \frac{1}{n^{3/2}} \times n = \frac{1}{n^{1/2}}$$
 (28)

Now, from Lemma 10 and from Lemma 8

$$\lim_{n \to \infty} \Pr(|Y_n - 0| \ge \epsilon) \le \lim_{n \to \infty} \frac{1}{n^{1/2} \epsilon^2} = 0$$
 (29)

$$\implies \lim_{n \to \infty} \Pr\left(\left| \frac{1}{n^{3/4}} \sum_{i=1}^{n} X_i - 0 \right| \ge \epsilon \right) = 0$$
 (30)

From Definition 2,

$$\frac{1}{n^{3/4}} \sum_{i=1}^{n} X_i \to 0 \tag{31}$$

Thus, option 2 is correct.

Writing the random variable Z_n from Lemma 11 for $\{X_i\}$ where $\mu=0$ and $\sigma=1$,

$$Z_n = \frac{1}{n^{1/2}} \sum_{i=1}^n X_i \tag{32}$$

From Lemma 11

$$\frac{1}{n^{1/2}} \sum_{i=1}^{n} X_i \to N(0,1) \tag{33}$$

whereas the option states that

$$\frac{1}{n^{1/2}} \sum_{i=1}^{n} X_i \to 0 \tag{34}$$

from Lemma 3.

Therefore, option 3 is incorrect.

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From (26), option 4 is correct.

Therefore, options 2 and 4 are correct.