Principal Components Analysis for Two Variables

There is software that computes principal components (PCs) very efficiently. No one wants to compute PCs by hand. But you may gain insight by seeing how it works in the simple case of two X's. So suppose you have data $(x_{11}, x_{12}, x_{13}, ..., x_{1n})$ and $(x_{21}, x_{22}, x_{23}, ..., x_{2n})$ on two variables X_1 and X_2 . I will derive the two PCs and their eigenvalues manually for the case of standardized X's.

For the case of standardized X's, the PCs are the eigenvectors of the correlation matrix. So the derivation begins with the correlation matrix $R = \begin{bmatrix} corr(X_1, X_1) & corr(X_1, X_2) \\ corr(X_1, X_2) & corr(X_1, X_2) \end{bmatrix} = \begin{bmatrix} 1 & r \\ r & 1 \end{bmatrix}$, where $r = corr(X_1, X_2)$.

First, compute the eigenvalues λ_1, λ_2 . By definition, they are the solutions to the determinant equation $|\mathbf{R} - \lambda \mathbf{I}| = 0$, i.e., $\begin{bmatrix} 1 & r \\ r & 1 \end{bmatrix} - \lambda \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = 0$, or $\begin{bmatrix} 1 - \lambda & r \\ r & 1 - \lambda \end{bmatrix} = 0$. Expand the determinant to get $(1 - \lambda)^2 - r^2 = 0$ and solve for λ : $\lambda_1 = 1 + r$ and $\lambda_2 = 1 - r$ are the two solutions. [Note: The sum of the eigenvalues is $\lambda_1 + \lambda_2 = (1 + r) + (1 - r) = 2$, as it should be.]

<u>Second</u>, for each eigenvalue, compute the eigenvectors (v_1, v_2) . These will be the principal component coefficients. By definition, the eigenvectors are solutions to the set of equations $[\mathbf{R} - \lambda_i \mathbf{I}] \mathbf{v} = \mathbf{0}$ or $\begin{bmatrix} 1 & r \\ r & 1 \end{bmatrix} - \lambda_i \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} v_1 \\ v_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}$. (There is one eigenvector for each eigenvalue,

so you must solve this pair of equations for $\lambda_i = \lambda_1 = 1 + r$ and then again for $\lambda_i = \lambda_2 = 1 - r$.) The

equations are
$$\begin{bmatrix} 1-\lambda_i & r \\ r & 1-\lambda_i \end{bmatrix} \begin{bmatrix} v_1 \\ v_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}. \text{ Hence } \begin{bmatrix} (1-\lambda_i)v_1 + rv_2 \\ rv_1 + (1-\lambda_i)v_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}. \text{ For } \lambda_i = \lambda_1 = 1+r \text{ , this }$$

becomes $\begin{bmatrix} -rv_1 + rv_2 \\ rv_1 - rv_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}$. The equations are linearly dependent, so there are infinitely many

solutions. Any vector is a solution if $v_1 = v_2$. For $\lambda_i = \lambda_2 = 1 - r$, the equations become

$$\begin{bmatrix} rv_1 + rv_2 \\ rv_1 + rv_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}$$
. So any vector is a solution if $v_1 = -v_2$. Let (a, a) denote a generic solution

when $\lambda_i=\lambda_1=1+r$. Let (b,-b) denote a generic solution when $\lambda_i=\lambda_2=1-r$. Then the equations that transform (x_1,x_2) into their principal components (ξ_1,ξ_2) have the form

$$\begin{bmatrix} \xi_1 \\ \xi_2 \end{bmatrix} = \begin{bmatrix} a & a \\ b & -b \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}$$
. In order for this rotation matrix to qualify as the PC rotation, it must be

¹ By convention, λ_1 is the larger and λ_2 is the smaller. So if $r \ge 0$, then $\lambda_1 = 1 + r$ and $\lambda_2 = 1 - r$ are the two solutions; if $r \le 0$, then $\lambda_1 = 1 - r$ and $\lambda_2 = 1 + r$.

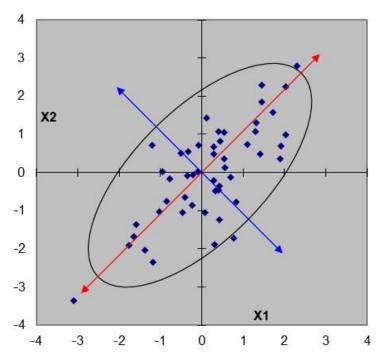
orthonormal. Let us check that out. Now, $(x_1, x_2) = (1,0)$ is transformed into (a,b); and $(x_1, x_2) = (0,1)$ is transformed into (a,-b). For orthonormality, we must have the transformed lengths = 1; i.e., $\sqrt{a^2 + b^2} = 1$. And the angle between them must have cosine = 0; i.e. $(a,b) \cdot (a,-b) = a^2 - b^2 = 0$. Substituting the latter into the length constraint, we have $\sqrt{a^2 + a^2} = 1$. Hence $a = 1/\sqrt{2} = b$. Thus, the PC transformation is

$$\begin{bmatrix} \xi_1 \\ \xi_2 \end{bmatrix} = \begin{bmatrix} 1/\sqrt{2} & 1/\sqrt{2} \\ 1/\sqrt{2} & -1/\sqrt{2} \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = \begin{bmatrix} \frac{1}{\sqrt{2}} x_1 + \frac{1}{\sqrt{2}} x_2 \\ \frac{1}{\sqrt{2}} x_1 - \frac{1}{\sqrt{2}} x_2 \end{bmatrix}$$

We see that the first PC $\xi_1 = \frac{1}{\sqrt{2}}x_1 + \frac{1}{\sqrt{2}}x_2$ is essentially the sum or the mean of the two X's, the

second PC $\xi_2 = \frac{1}{\sqrt{2}}x_1 - \frac{1}{\sqrt{2}}x_2$ is essentially their difference. Curiously, the solution does not

depend upon the value of the correlation between the two X's. These are the PCs for any two X's! However, the variances of the two PCs do depend upon the correlation. The variances are $Var(\xi_1) = \lambda_1 = 1 + r$ and $Var(\xi_2) = \lambda_2 = 1 - r$.



The figure above illustrates the situation. The two PCs are the two perpendicular arrows. PC1 (red) has slope +1 and PC2 (blue) has slope -1. They cross at the origin. The larger the correlation, the more the data stretch out in the direction of PC1, the larger the variance of PC1, and the more elliptical the data plot appears. The closer the correlation is to 0, the less the data spread out, the smaller the variance of PC1, and the more circular the data plot appears. If the

correlation is negative, then the orientation of the ellipse will rotate to upper left to lower right; PC1 will remain the long axis, but it will have slope of -1; and PC2 will remain the short axis, but with slope of +1, as footnote 1 suggests. In the original coordinates, the standardized X_1 and X_2 are correlated – indicated by the angled orientation of the data ellipse to the X_1 and X_2 axes. The principal components transformation replaces the original coordinate axes by the major and minor axes of the ellipse (the arrowed lines). In that new coordinate system, the data ellipse is flat – the data orientation is no longer angled; the correlation is zero.

It is worth noting that neither PC is the regression line for either variable. The regression of X_2 on X_1 has slope r, and the regression of X_1 on X_2 has slope 1/r. The regression lines rotate toward the original axes – illustrating the phenomenon of regression to the mean. The regression line for X_2 on X_1 minimizes the total squared *vertical up-and-down* distance from the points to the regression line; PC1 minimizes the total squared *perpendicular* distance from the points to the PC1 line.