

Problem 1.1: Linear Regression

We can have two approaches for solving this problem. Both of them are correct.

Matrix way

Part (a) [5 points]

Since the noise terms are independent, we can write the covariance matrix of them Σ as a diagonal matrix with elements $(\sigma_1^2, \dots, \sigma_N^2)$. The inverse of this matrix is another diagonal matrix with elements $(\sigma_1^{-2}, \dots, \sigma_N^{-2})$. Thus, using the pdf of multivariate normal distribution, we can write the likelihood of the data as follows:

$$P(D) = (2\pi)^{-\frac{N}{2}} |\Sigma|^{-\frac{1}{2}} \exp \left(-\frac{1}{2} (\mathbf{y} - X\boldsymbol{\beta})^\top \Sigma^{-1} (\mathbf{y} - X\boldsymbol{\beta}) \right).$$

The negative log likelihood can be written as:

$$\begin{aligned} -\log P(D) &= \frac{1}{2} (\mathbf{y} - X\boldsymbol{\beta})^\top \Sigma^{-1} (\mathbf{y} - X\boldsymbol{\beta}) + \text{const.}, \\ &= \frac{1}{2} \|H(\mathbf{y} - X\boldsymbol{\beta})\|_2^2, \end{aligned}$$

where in the last step H is a diagonal matrix with $(\sigma_1^{-1}, \dots, \sigma_N^{-1})$ on its diagonal.

Part (b) [5 points]

Now, observing the fact that by defining $\tilde{\mathbf{y}} = H\mathbf{y}$ and $\tilde{X} = HX$, we have a ordinary regression problem, we can use its solution and write:

$$\hat{\boldsymbol{\beta}} = (\tilde{X}^\top \tilde{X})^{-1} \tilde{X}^\top \tilde{\mathbf{y}} = (X^\top H^\top H X)^{-1} X^\top H^\top H \mathbf{y} = (X^\top \Sigma^{-1} X)^{-1} X^\top \Sigma^{-1} \mathbf{y}.$$

Summation way

Part (a)

$$y_n = \mathbf{x}_n^\top \boldsymbol{\beta} + \varepsilon_n$$

Here, $\varepsilon \sim \mathcal{N}(0, \sigma_n)$, σ_n not given to be equal for all n . First, we note that each y_n is taken from the distribution $\mathcal{N}(\mathbf{x}_n^\top \boldsymbol{\beta}, \sigma_n)$. Thus:

$$\begin{aligned} P(y_n | \boldsymbol{\beta}, \mathbf{x}_n) &= (2\pi\sigma_n^2)^{-\frac{1}{2}} \exp \left\{ -\frac{1}{2\sigma_n^2} (y_n - \mathbf{x}_n^\top \boldsymbol{\beta})^2 \right\} \\ P(D) &= \prod_n P(y_n | \boldsymbol{\beta}, \mathbf{x}_n) = \prod_n (2\pi\sigma_n^2)^{-\frac{1}{2}} \exp \left\{ -\frac{1}{2\sigma_n^2} (y_n - \mathbf{x}_n^\top \boldsymbol{\beta})^2 \right\} \\ \log P(D) &= \sum_n \log(P(y_n | \boldsymbol{\beta}, \mathbf{x}_n)) = \sum_n \left[-\frac{1}{2} \log(2\pi\sigma_n^2) - \frac{1}{2\sigma_n^2} (y_n - \mathbf{x}_n^\top \boldsymbol{\beta})^2 \right] \end{aligned}$$

Part (b)

$$\frac{\partial}{\partial \boldsymbol{\beta}} \log P(D) = \sum_n \frac{1}{\sigma_n^2} (y_n - \mathbf{x}_n^\top \boldsymbol{\beta}) \mathbf{x}_n^\top$$

Set equal to zero and rearrange.

$$\sum_n \frac{1}{\sigma_n^2} \mathbf{x}_n^\top \boldsymbol{\beta} \mathbf{x}_n^\top = \sum_n \frac{1}{\sigma_n^2} y_n \mathbf{x}_n^\top$$

$$\beta^\top \sum_n \frac{1}{\sigma_n^2} \mathbf{x}_n \mathbf{x}_n^\top = \sum_n \frac{1}{\sigma_n^2} \mathbf{x}_n^\top y_n$$

$$\hat{\beta}^\top = \left[\sum_n \frac{1}{\sigma_n^2} \mathbf{x}_n \mathbf{x}_n^\top \right]^{-1} \sum_n \frac{1}{\sigma_n^2} \mathbf{x}_n^\top y_n$$

Problem 1.2: Smooth Coefficients

Part (a)

The regularizer representing $(\beta_i - \beta_{i+1})^2$ is given by: $\sum_i^{p-1} (\beta_i - \beta_{i+1})^2$. This can be rearranged into vector form. Define matrix $D \in \mathbb{R}^p$ as the following

$$D = \begin{bmatrix} 1 & -1 & 0 & \cdots & 0 \\ 0 & \ddots & \ddots & 0 & \vdots \\ 0 & \cdots & 1 & -1 & 0 \\ 0 & \cdots & 0 & 1 & -1 \\ 0 & \cdots & \cdots & 0 & 0 \end{bmatrix}$$

The regularizer is then *[5 points]* :

$$\beta^\top D^\top D \beta = \|D\beta\|_2^2$$

The full optimization problem is *[5 points]* :

$$L(\beta) = \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \lambda \|\beta\|_2^2 + \mu \|D\beta\|_2^2$$

$$L(\beta) = (\mathbf{y}^\top \mathbf{y} - 2\mathbf{y}^\top \mathbf{X}\beta + \beta^\top \mathbf{X}^\top \mathbf{X}\beta) + \lambda \beta^\top \beta + \mu \beta^\top D^\top D \beta$$

for some λ and μ hyper-parameters.

Part (b) *[5 points]*

$$\nabla L(\beta) = 2\mathbf{X}^\top \mathbf{X}\beta - 2\mathbf{X}^\top \mathbf{y} + 2\lambda\beta + 2\mu D^\top D \beta$$

Set equal to zero.

$$(\mathbf{X}^\top \mathbf{X} + \lambda I_p + \mu D^\top D) \hat{\beta} = \mathbf{X}^\top \mathbf{y}$$

$$\hat{\beta} = (\mathbf{X}^\top \mathbf{X} + \lambda I_p + \mu D^\top D)^{-1} \mathbf{X}^\top \mathbf{y}$$

Problem 1.3: Constrained Linear Regression

Here we take the immediate jump to the L_2 minimization problem, as presented in class.

$$\min_{\beta} \|\mathbf{y} - \mathbf{X}\beta\|_2 \quad \text{s.t.} \quad A\beta = \mathbf{b}$$

Write the Lagrangian.

$$\mathcal{L}(\beta, \lambda) = \|\mathbf{y} - \mathbf{X}\beta\|_2 + \lambda^\top (A\beta - \mathbf{b})$$

Take the derivative with respect to β .

$$\frac{\partial}{\partial \beta} \mathcal{L}(\beta, \lambda) = 2\mathbf{X}^\top \mathbf{X} \beta - 2\mathbf{X}^\top \mathbf{y} + A^\top \lambda$$

Set equal to zero and solve for β :

$$\beta = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y} - \frac{1}{2} (\mathbf{X}^\top \mathbf{X})^{-1} A^\top \lambda$$

Use the constraint by applying A to both sides.

$$A\beta = \mathbf{b} = A(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y} - \frac{1}{2} A(\mathbf{X}^\top \mathbf{X})^{-1} A^\top \lambda$$

Solve for λ .

$$A(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y} - \mathbf{b} = \frac{1}{2} A(\mathbf{X}^\top \mathbf{X})^{-1} A^\top \lambda$$

$$\lambda = 2(A(\mathbf{X}^\top \mathbf{X})^{-1} A^\top)^{-1} (A(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y} - \mathbf{b})$$

Plug λ into the derivative.

$$\hat{\beta} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y} - (\mathbf{X}^\top \mathbf{X})^{-1} A^\top (A(\mathbf{X}^\top \mathbf{X})^{-1} A^\top)^{-1} (A(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y} - \mathbf{b})$$

Problem 2: Perceptron

Conditions: From any current step parameters \mathbf{w}_i we want to update the classifier such that $\text{sign}(\mathbf{w}_{i+1}^\top \mathbf{x}_{i+1}) = y_{i+1}$. However, we also would like $\|\mathbf{w}_{i+1} - \mathbf{w}_i\|_2$ to be small.

Solution: If $y_{i+1} = \text{sign}(\mathbf{w}_{i+1}^\top \mathbf{x}_{i+1})$, then let $\mathbf{w}_{i+1} = \mathbf{w}_i$ (do nothing). Otherwise we need the smallest amount of movement such that then point \mathbf{x}_{i+1} is on the correct side of the plane. Do the following:

$$\mathbf{w}_{i+1} = \arg \min_{\mathbf{w}} \frac{1}{2} \|\mathbf{w} - \mathbf{w}_i\|_2^2 \quad \text{s.t.} \quad \mathbf{w}^\top \mathbf{x}_{i+1} y_{i+1} = 0$$

Writing the Lagrangian, yields

$$\mathcal{L}(\mathbf{w}, \lambda) = \frac{1}{2} (\mathbf{w} - \mathbf{w}_i)^\top (\mathbf{w} - \mathbf{w}_i) + \lambda \mathbf{w}^\top \mathbf{x}_{i+1} y_{i+1}$$

Take a derivative w.r.t. \mathbf{w}

$$\frac{\partial}{\partial \mathbf{w}} \mathcal{L}(\mathbf{w}, \lambda) = (\mathbf{w} - \mathbf{w}_i) - \lambda \mathbf{x}_{i+1} y_{i+1} = 0$$

$$\mathbf{w} = \lambda \mathbf{x}_{i+1} y_{i+1} + \mathbf{w}_i$$

Transpose and multiply by $\mathbf{x}_{i+1} y_{i+1}$ on both sides, then apply the equality $\mathbf{w}^\top \mathbf{x}_{i+1} y_{i+1} = 0$:

$$\mathbf{w}^\top \mathbf{x}_{i+1} y_{i+1} = 0 = \lambda (\mathbf{x}_{i+1} y_{i+1})^\top (\mathbf{x}_{i+1} y_{i+1}) + \mathbf{w}_i^\top (\mathbf{x}_{i+1} y_{i+1})$$

$$\lambda = -\frac{\mathbf{w}_i^\top (\mathbf{x}_{i+1} y_{i+1})}{\|\mathbf{x}_{i+1}\|_2^2}$$

Plug back in, and let this be the update rule.

$$\mathbf{w}_{i+1} = \mathbf{w}_i - \frac{\mathbf{w}_i^\top \mathbf{x}_{i+1}}{\|\mathbf{x}_{i+1}\|_2^2} \mathbf{x}_{i+1}$$

Geometrically this is the same as finding some vector that is perpendicular to \mathbf{x}_{i+1} and projecting \mathbf{w}_i onto it, taking the projection as the new normal vector.

Problem 3

Part (a) [5 points] : Given:

$$K_3 = a_1 K_1 + a_2 K_2$$

where $a_1, a_2 \geq 0$ and K_1, K_2 positive semi-definite. For any $\mathbf{x} \in \mathbb{R}^N$:

$$\mathbf{x}^\top K_3 \mathbf{x} = \mathbf{x}^\top (a_1 K_1 + a_2 K_2) \mathbf{x} = a_1 \mathbf{x}^\top K_1 \mathbf{x} + a_2 \mathbf{x}^\top K_2 \mathbf{x}$$

By assumption for any $\mathbf{x} \in \mathbb{R}^N$ both $\mathbf{x}^\top K_1 \mathbf{x} \geq 0$ and $\mathbf{x}^\top K_2 \mathbf{x} \geq 0$ (definition of positive semi-definite). Thus, the non-negative combination of the two is also ≥ 0 . So $\mathbf{x}^\top K_3 \mathbf{x} \geq 0$.

Part (b) [5 points] : Given:

$$K_4 : k_4(\mathbf{x}, \mathbf{x}') = f(\mathbf{x})f(\mathbf{x}')$$

for any real valued function f .

Let $\mathbf{f} = (f(\mathbf{x}_1), \dots, f(\mathbf{x}_1))$. We can write $K_4 = \mathbf{f} \mathbf{f}^\top$, thus $\mathbf{x}^\top K_4 \mathbf{x} = (\mathbf{x}^\top \mathbf{f})^2 \geq 0$ for any vector $\mathbf{x} \in \mathbb{R}^N$.

Part (c) [5 points] : Given:

$$K_5 : k_5(\mathbf{x}, \mathbf{x}') = k_1(\mathbf{x}, \mathbf{x}')k_2(\mathbf{x}, \mathbf{x}')$$

We can also see that

$$K_5 = K_1 \circ K_2$$

where \circ is the elementwise product. K_5 follows from the following proof. We also use an identity of the elementwise product, namely $\mathbf{x}^\top (A \circ B) \mathbf{y} = \text{tr}(A \text{diag}(\mathbf{x}) B \text{diag}(\mathbf{y}))$.

$$\mathbf{x}^\top K_1 \circ K_2 \mathbf{x} = \text{tr}(K_1 \text{diag}(\mathbf{x}) K_2 \text{diag}(\mathbf{x}))$$

K_1 and K_2 are assumed to be positive semi-definite, so they admit a root. $K_1 = (K_1^{\frac{1}{2}})(K_1^{\frac{1}{2}})$, $K_2 = (K_2^{\frac{1}{2}})(K_2^{\frac{1}{2}})$.

$$\text{tr}((K_1^{\frac{1}{2}})(K_1^{\frac{1}{2}}) \text{diag}(\mathbf{x})(K_2^{\frac{1}{2}})(K_2^{\frac{1}{2}}) \text{diag}(\mathbf{x})) = \text{tr}((K_1^{\frac{1}{2}}) \text{diag}(\mathbf{x})(K_2^{\frac{1}{2}})(K_2^{\frac{1}{2}}) \text{diag}(\mathbf{x})(K_1^{\frac{1}{2}}))$$

This is equivalent to $\text{tr}(A^\top A)$ for some matrix A , which is the trace of a gram matrix, which is always greater than zero. Thus $K_1 \circ K_2$ is positive semi-definite.

—Solutions based on interpretation of this kernel as the covariance of a random vector which is obtained by elementwise product of two independent random vector with covariances K_1 and K_2 is also acceptable. See the Wikipedia page for Schur product theorem.