

# Advanced Forecasting Models

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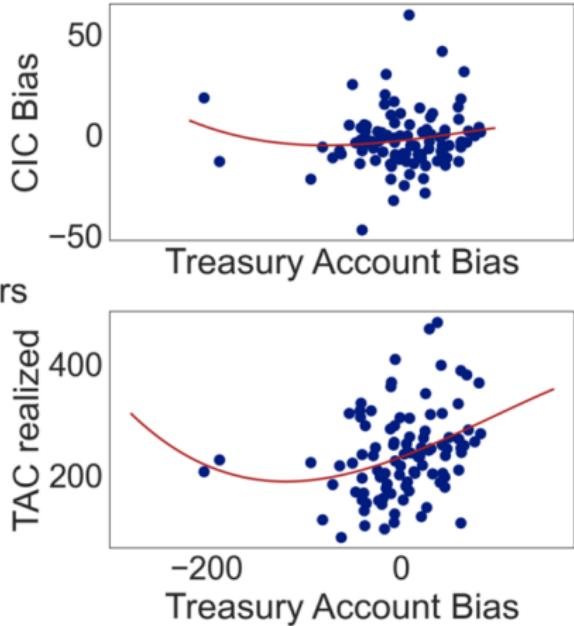
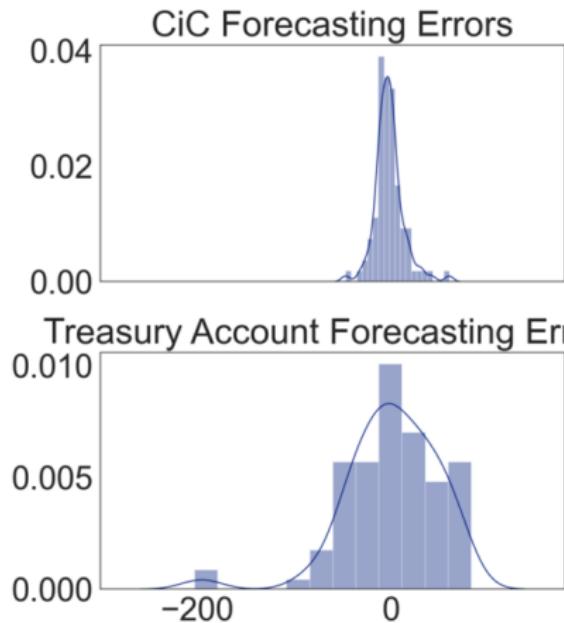
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- ① Advanced Seasonal Models
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# Bias Identification

## Autonomous Factors Forecasting Analysis



Credit: Author

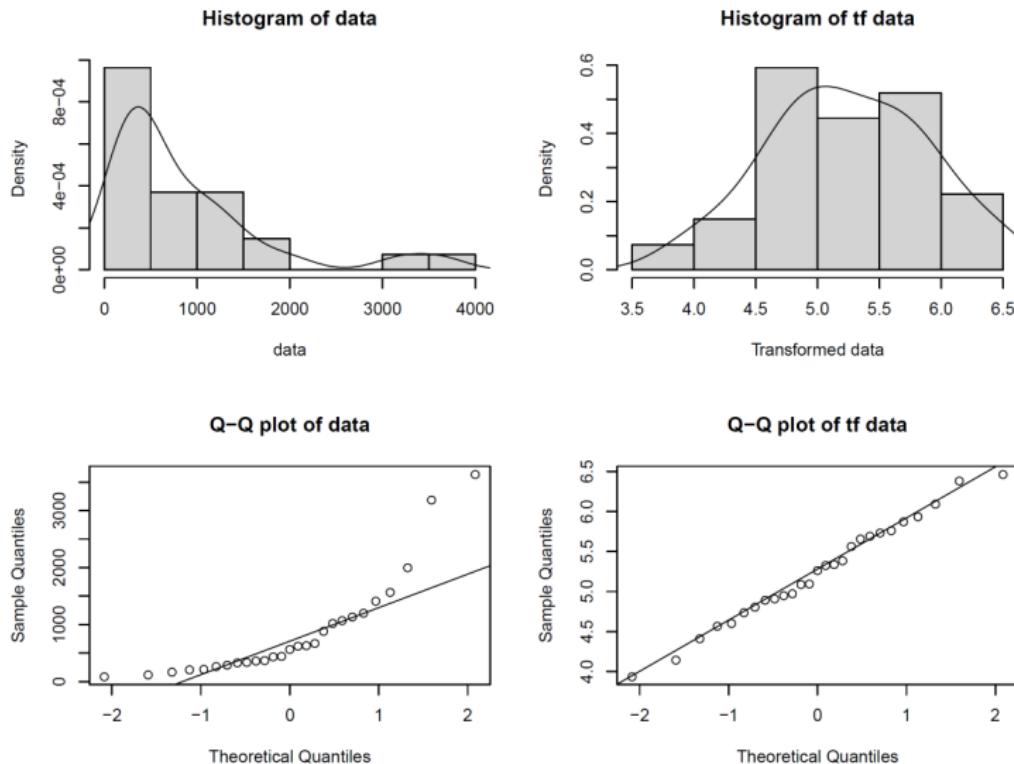
# Advanced Models Typology

- **ETS with trigonometric** multiple seasonalities and indicators for events
- **ARIMA with trigonometric** multiple seasonalities and indicators for events
- **ARIMA with stochastic** multiple seasonalities and indicators for events
- **TBATS** – a state space model with exponential smoothing, stochastic trigonometric patterns, Cox-Box transformation and ARIMA errors. TBATS cannot incorporate events
- **Composite TBATS** – where the local level and events are modelled with ETS and the seasonality/trend with TBATS

# Box Cox Transformation

- The Box-Cox transformation is used to transform non-Gaussian data  $y_t$  into Gaussian data
- Working with Gaussian data makes it easier to apply the standard model in econometrics
- The Box Cox transformation works with a parameters  $\lambda$  that takes values between  $[-5, 5]$
- For positive  $y_t$  data:
  - $y(\lambda) = \frac{y^\lambda - 1}{\lambda}$  if  $\lambda \neq 0$
  - $\log(y)$  if  $\lambda = 0$
- For negative  $y_t$  data, use an extra parameter  $\lambda_2$ 
  - $y(\lambda) = \frac{(y + \lambda_2)^\lambda - 1}{\lambda}$  if  $\lambda \neq 0$
  - $\log(y + \lambda_2)$  if  $\lambda = 0$

# Transforming Data with the Box Cox Transformation

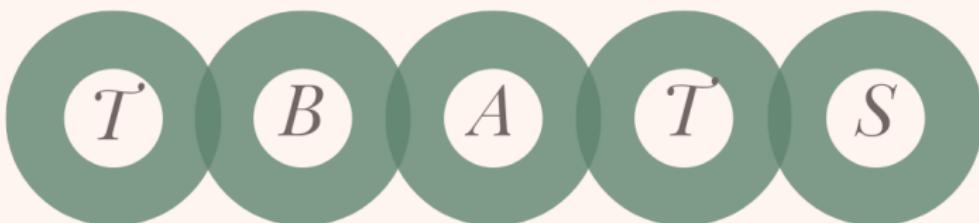


Credit: <https://universeofdatascience.com/>

# The TBATS Model: Overview

## TBATS

Forecasting Model



TBATS: Trigonometric seasonality, Box-Cox transformation, ARMA errors, Trend and Seasonal components.

@nadeemoffl

Credit: <https://medium.com/analytics-vidhya/>

# The TBATS Model: Overview

- Combines in a **single state-space** model:
  - Trigonometric terms to account for seasonality
  - Box-Cox transformation to account for non-normality and heterogeneity
  - ARIMA errors to model the residuals short-term dynamic
  - Trend (potentially damped)
  - Seasonal (including multiple and non-integer period)
- Main specifications:
  - Handles non-integer seasonality, multiple seasonal periods
  - Entirely automated
  - Predictions interval often wide
  - Very slow on long series

# The TBATS Model: Specification

## Model:

$$y_t^{(\lambda)} = l_{t-1} + \phi b_{t-1} + \sum_{i=1}^T s_{t-m_i}^{(i)} + d_t$$

$$l_t = l_{t-1} + \phi b_{t-1} + \alpha d_t$$

$$b_t = \phi b_{t-1} + \beta d_t$$

$$d_t = \sum_{i=1}^p \varphi_i d_{t-i} + \sum_{i=1}^q \theta_i e_{t-i} + e_t$$

## Seasonal part:

$$s_t^{(i)} = \sum_{j=1}^{k_i} s_{j,t}^{(i)}$$

$$s_{j,t}^{(i)} = s_{j,t-1}^{(i)} \cos(\omega_i) + s_{j,t-1}^{*(i)} \sin(\omega_i) + \gamma_1^{(i)} d_t$$

$$s_{j,t}^{*(i)} = -s_{j,t-1}^{(i)} \sin(\omega_i) + s_{j,t-1}^{*(i)} \cos(\omega_i) + \gamma_2^{(i)} d_t$$

$$\omega_i = 2\pi j/m_i$$

## Where:

$y_t^{(\lambda)}$  - time series at moment  $t$  (Box-Cox transformed)

$s_t^{(i)}$  -  $i$ th seasonal component

$l_t$  - local level

$b_t$  - trend with damping

$d_t$  - ARMA( $p,q$ ) process for residuals

$e_t$  - Gaussian white noise

## Model parameters:

$T$  - Amount of seasonalities

$m_i$  - Length of  $i$ th seasonal period

$k_i$  - Amount of harmonics for  $i$ th seasonal period

$\lambda$  - Box-Cox transformation

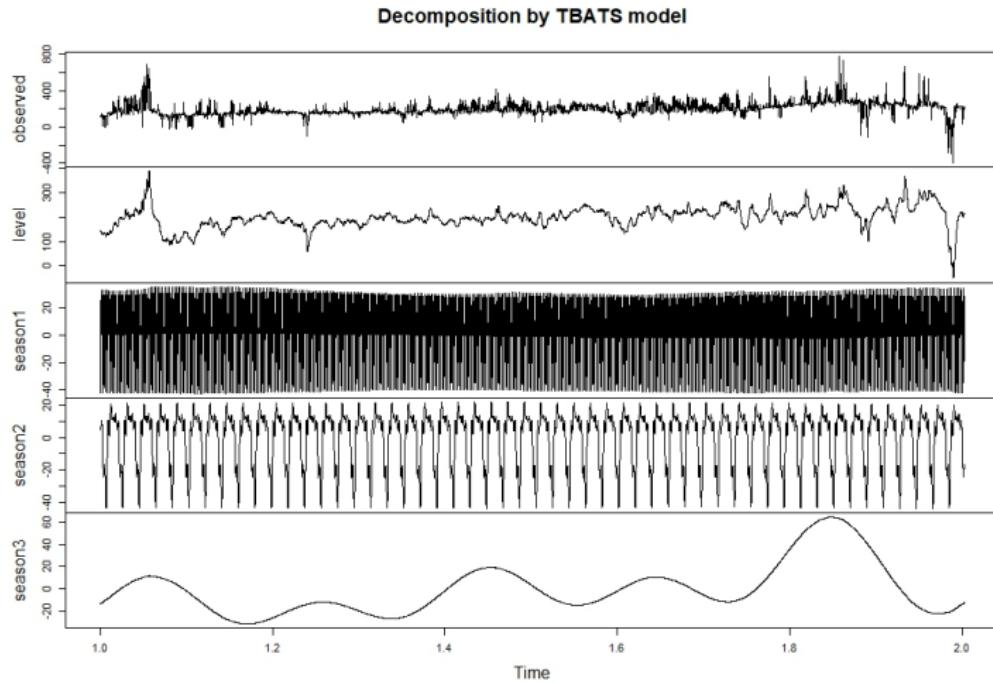
$\alpha, \beta$  - Smoothing

$\phi$  - Trend damping

$\varphi_i, \theta_i$  - ARMA( $p, q$ ) coefficients

$\gamma_1^{(i)}, \gamma_2^{(i)}$  - Seasonal smoothing (two for each period)

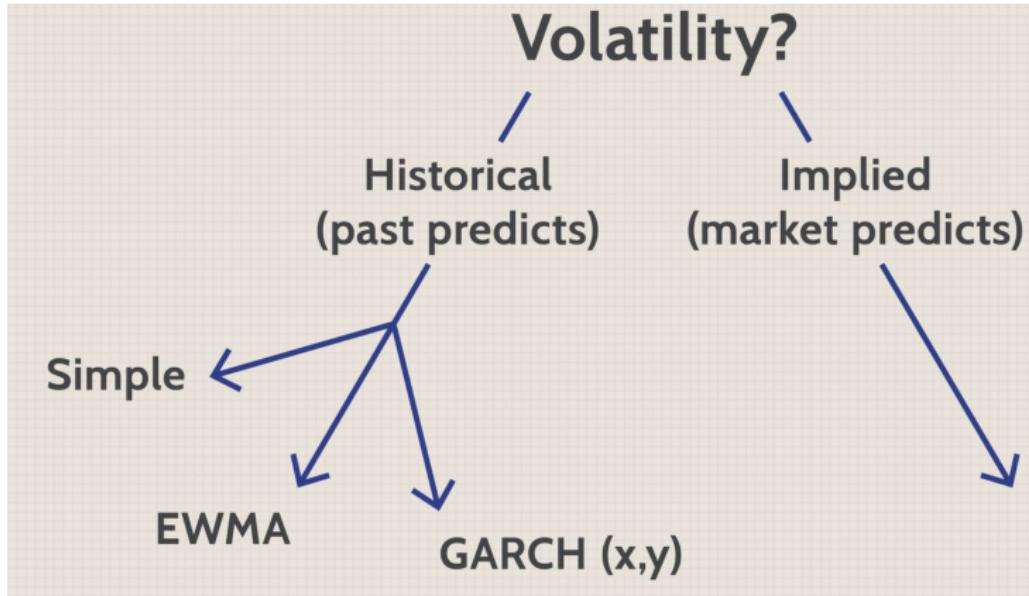
# The TBATS Model: Decomposition



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# Volatility Estimation



Credit: *Sabrina Jiang*

# EWMA Model

Squared Periodic Return	$U_i^2$	$U_{i-1}^2$	$U_{i-2}^2$	$U_{i-3}^2$
Weight:	$(1-\lambda)\lambda^0$	$(1-\lambda)\lambda^1$	$(1-\lambda)\lambda^2$	$(1-\lambda)\lambda^3$

Credit: *Sabrina Jiang*

$$\sigma_n^2(\text{EWMA}) = \lambda\sigma_{n-1}^2 + (1 - \lambda)u_{n-1}^2$$

**where:**

EWMA = Exponentially weighted moving average

$\sigma_n^2$  = Variance today

$\lambda$  = Degree of weighting

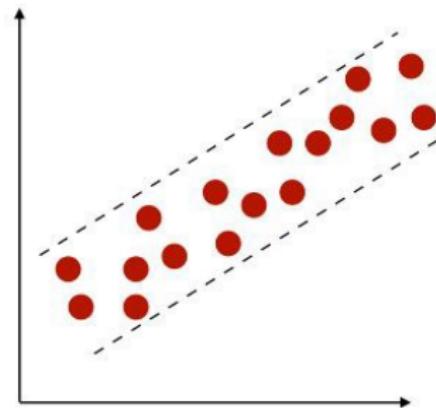
$\sigma_{n-1}^2$  = Variance yesterday

$u_{n-1}^2$  = Squared return yesterday

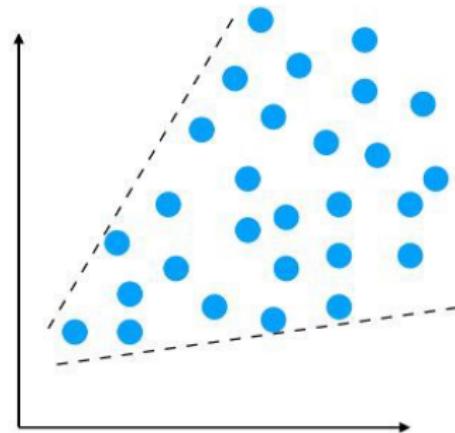
## EWMA Weights

Simple ↓	EWMA ↓	
Equal Weight	Weight ( $\alpha$ )	$\alpha \text{Return}^2$ ( $\alpha \mu^2$ )
0.196%	6.00%	0.000010%
0.196%	5.64%	0.000004%
0.196%	5.30%	0.000886%
0.196%	4.98%	0.000034%
0.196%	4.68%	0.001189%
0.196%	0.00%	0.000000%
100%	100%	

# Homoscedasticity and Heteroscedasticity



Homoscedasticity



Heteroscedasticity

Credit: *Medium.com*

# GARCH Models

## Meaning

**GARCH:** Generalized AutoRegressive Conditional Heteroscedasticity

*Reference: Bollerslev, T. (1986), Generalized Autoregressive Conditional Heteroskedasticity. Journal of Econometrics, 31, 307-327*

## Definition: GARCH Model

The stochastic process  $\{\epsilon_t, t \in \mathbb{Z}\}$  is said to be a **GARCH(p,q)** process if:

$$\epsilon_t = Z_t \sigma_t$$

where  $Z_t$  is a sequence of i.i.d variables with  $\mathbb{E}(Z_t) = 0$  and  $\mathbb{V}(Z_t) = 1$ , and  $\sigma_t$  is a non-negative process such that:

$$\sigma_t^2 = \omega + \sum_{i=1}^p \alpha_i \epsilon_{t-i}^2 + \sum_{i=1}^p \beta_i \sigma_{t-i}^2$$

with  $\omega > 0$ ,  $\forall i, (\alpha_i, \beta_i) \in \mathbb{R}^{+,2}$  and  $\sum_{i=1}^p \alpha_i + \sum_{i=1}^p \beta_i < 1$

# GARCH: Intuition

- The conditional variance of a GARCH(p, q) depends on:
  - The first p lag of the  $\epsilon_t^2$  (e.g. the squared error terms)
  - The first q lag of the conditional variance  $\sigma^2$

$$\sigma_t^2 = \omega + \underbrace{\sum_{i=1}^p \alpha_i \epsilon_{t-i}^2}_{\text{ARCH Components}} + \underbrace{\sum_{i=1}^q \beta_i \sigma_{t-i}^2}_{\text{GARCH components}}$$

- The parameters  $\alpha_i$  are often called the **ARCH parameters**
- The parameters  $\beta_i$  are often called the **GARCH parameters**

# GARCH(1, 1)

## Tip

GARCH(1,1) specifications are generally sufficient to capture the dynamics of the conditional variance

## Special Case: GARCH(1, 1)

The stochastic process  $\{\epsilon_t, t \in \mathbb{Z}\}$  is said to be a **GARCH(1,2)** process if:

$$\epsilon_t = Z_t \sigma_t$$

where  $Z_t$  is a sequence of i.i.d variables with  $\mathbb{E}(Z_t) = 0$  and  $\mathbb{V}(Z_t) = 1$ , and  $\sigma_t$  is a non-negative process such that:

$$\sigma_t^2 = \omega + \alpha \epsilon_{t-1}^2 + \beta \sigma_{t-1}^2$$

with  $\omega > 0$ ,  $\alpha \geq 0$ ,  $\beta \geq 0$  and  $\alpha + \beta < 1$

# Conditional Variance Persistences: Intrinsic and Extrinsic

- The conditional variance  $\sigma_t^2 = \omega + \alpha\epsilon_{t-1}^2 + \beta\sigma_{t-1}^2$  depends on two effects:
  - An **intrinsic persistence** effect through the first lag of the conditional variance
  - An **extrinsic persistence** effect
- Following a positive (or negative) shock at time  $t-1$ , the conditional variance at time  $t$  increases (impact effect) and thus it has an impact on  $\epsilon_t = Z_t\sigma_t$

$$\text{shock } z_{t-1} > 0 \Rightarrow \epsilon_{t-1} \uparrow \Rightarrow \sigma_t \uparrow \dots$$

- Starting from the next period (at time  $t$ ), the effect of the shock at  $t-1$  on the conditional variance at  $t+1$  passes through the conditional variance at time  $t$  (intrinsic persistence)

$$\text{shock } z_{t-1} \Rightarrow \sigma_t \uparrow \Rightarrow \sigma_{t+1}^2 \uparrow$$

- The overall effect of a shock can be decomposed into a

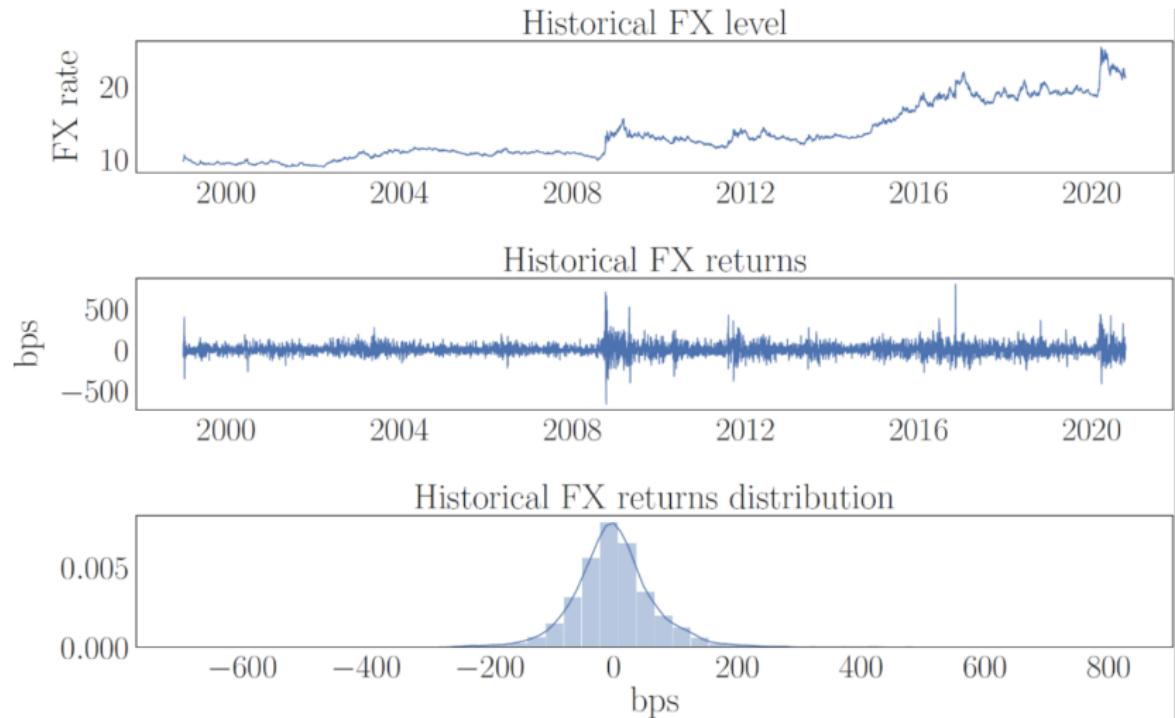
## Remarks

- It is often the case that:
  - ➊ The sum of the estimates of  $\alpha$  and  $\beta$  are generally close (but below 1)
  - ➋ The estimate of  $\beta$  is generally greater than the one of  $\alpha$
  - ➌ The estimate of  $\beta$  is generally larger than 0.90 for daily returns and the estimate of  $\alpha$  is below 0.1
- Be careful: it is not a general rule, just an observation.

# GARCH Extensions

- I won't cover them during the course but many extensions have been proposed to extend the GARCH model, including model for dealing with **asymmetric responses**, **persistence** and **long-memory**
  - **IGARCH model:** Integrated GARCH (model persistence)
  - **GARCH-M model:** GARCH mean, when the mean depends on the volatility
  - **GJR-GARCH model:** Glosten, Jagannathan and Runkle (1993): asymmetric with non-linearities
  - **TGARCH model:** Threshold GARCH captures regimes change
  - **EGARCH model:** Exponential GARCH captures asymmetric responses and big shocks

# Application: FX Modeling



Credit: Lafarguette and Veyrune (2021)

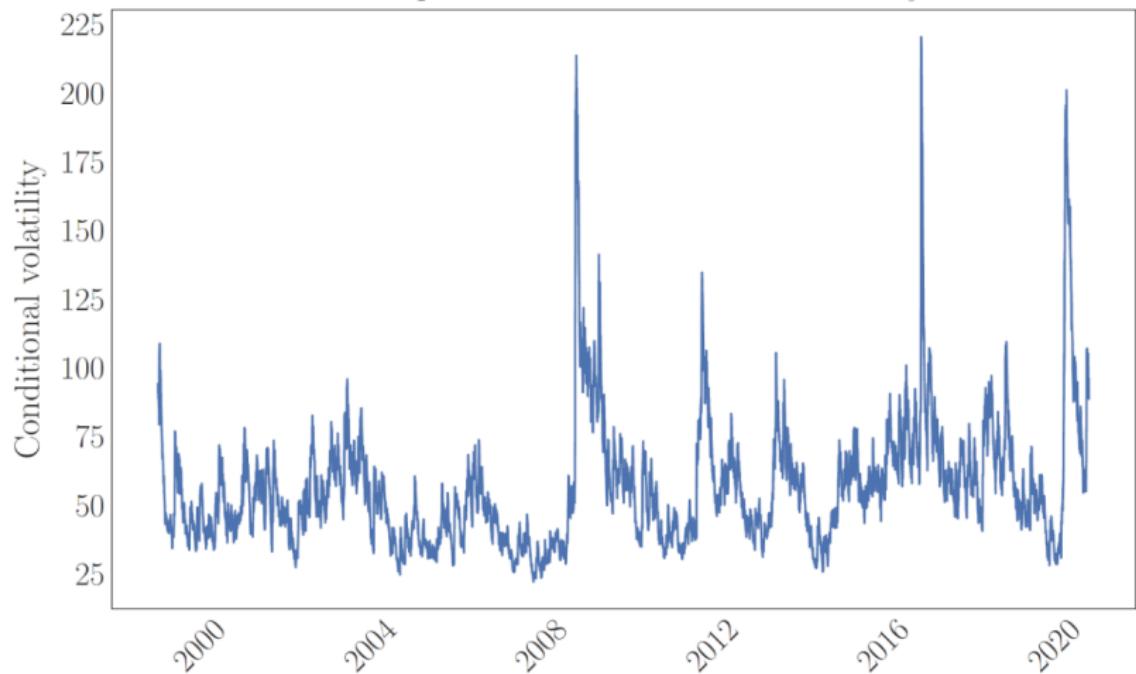
# Application: GARCH Estimation

	Microstructure	CIP	Dollar move	Risk appetite	Baseline
<b>Regressors</b>					
Intercept	-2.33***	-2.29	-1.84	-2.55	-1.63
Lag FX log returns	-0.07***	-0.08	-0.08***	-0.08***	-0.08***
Bid-ask abs	5.71	24.39	-35.66	-2.42	3.23
Min-max abs	35.56***	34.63	34.32	34.55*	26.21
Forward points first difference	23.29***	17.79***	26.44***	19.8***	19.44***
Interbank rate vs Libor		33.61***	39.32***	34.75***	33.86***
EUR/USD log returns			-0.14***	-0.17***	-0.16***
VIX first diff				15.66***	15.37***
Oil prices log returns					-0.02***
FX intervention dummy lag					2.23
<b>GARCH Parameters</b>					
Omega	0.13***	0.13	0.12***	0.11***	0.12***
Alpha	0.17***	0.17*	0.16***	0.16***	0.15***
Gamma	0.07***	0.06***	0.06***	0.05***	0.05***
Beta	0.98***	0.99***	0.99***	0.99***	0.99***
Nu	8.33***	8.66***	8.92***	8.71***	8.54***
Lambda	0.08***	0.07	0.09	0.07*	0.08***
R2	0.058	0.067	0.104	0.273	0.276
R2 adjusted	0.058	0.066	0.104	0.272	0.275
Number of observations	5986	5986	5682	5682	5680
<i>Significance thresholds</i>					
*10%, **5%, ***1%					

Credit: Lafarguette and Veyrune (2021)

# Application: MXN/USD Conditional Volatility

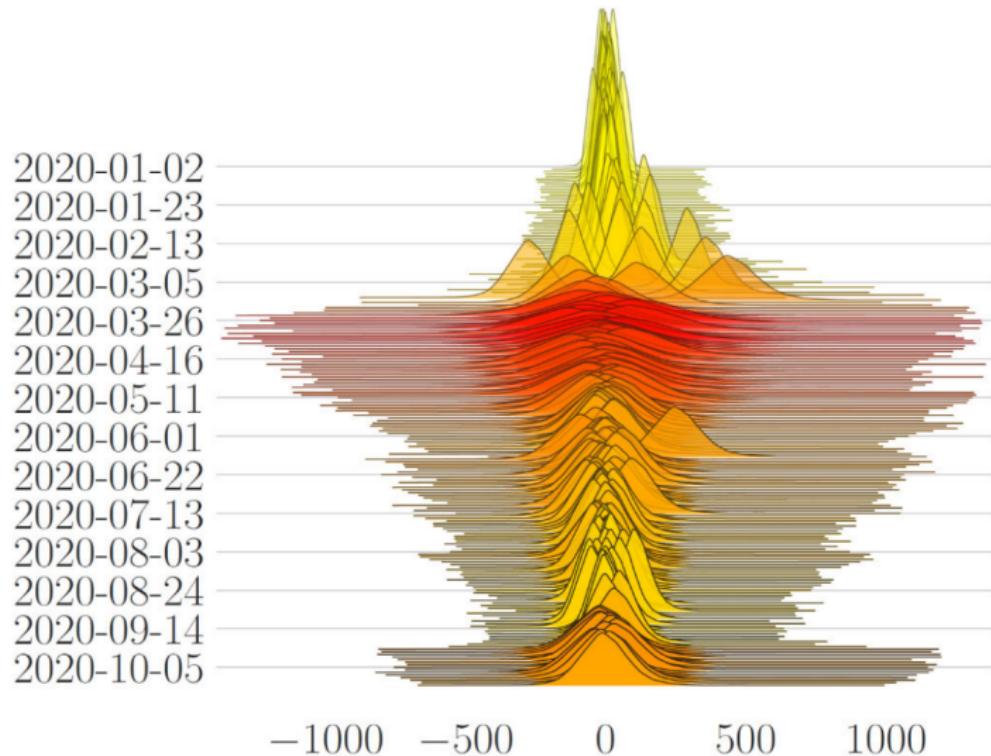
**Figure 3. Conditional FX Volatility over Time**



Credit: Lafarguette and Veyrune (2021)

# Application: MXN/USD Conditional Density

Figure 4. Out-of-Sample Conditional Density

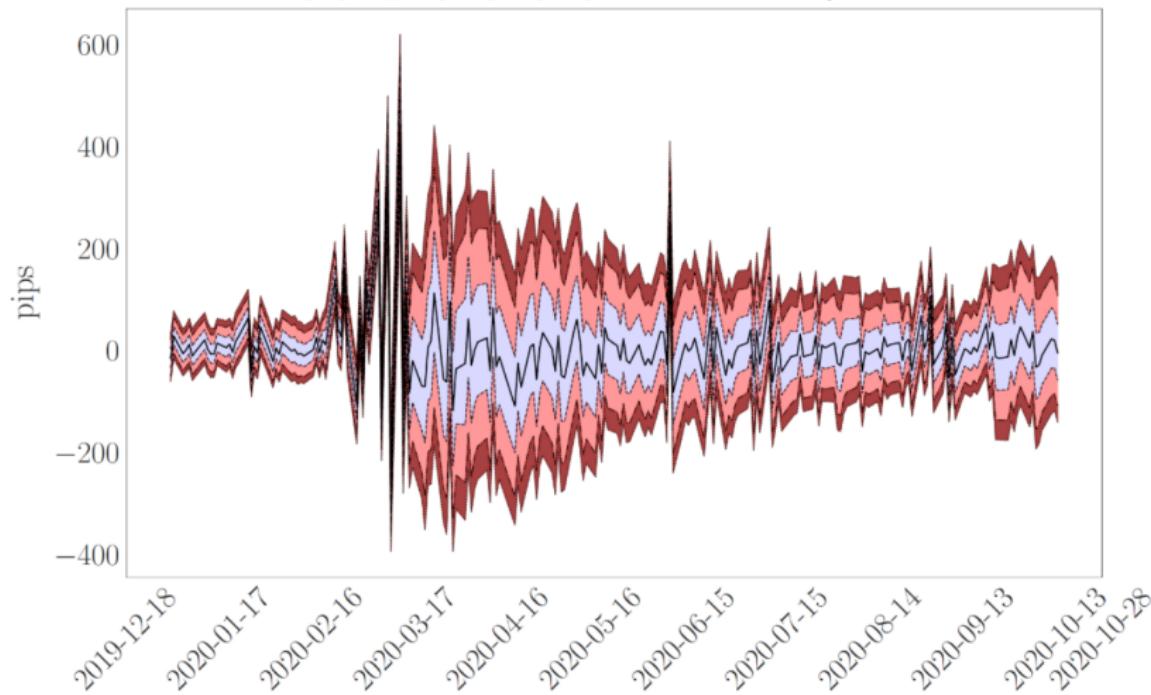


Credit: Lafarguette and Veyrune (2021)

# Application: MXN/USD Fan Chart

Figure 5. Out-of-Sample Fan Chart

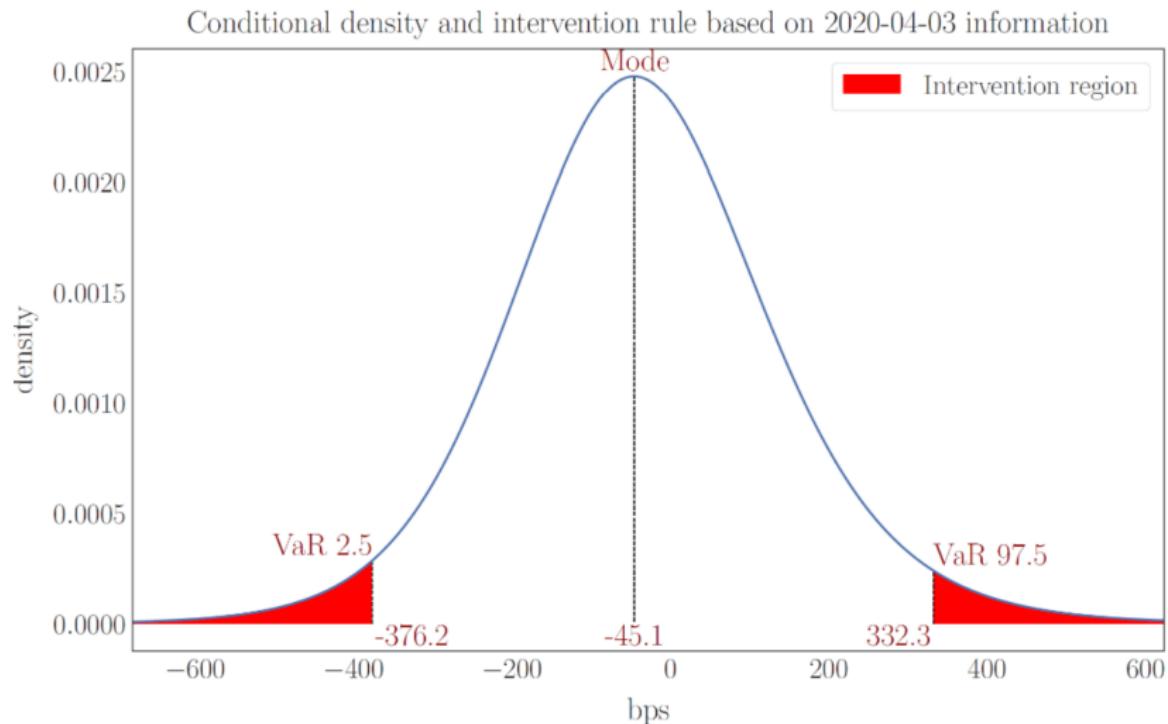
Fan chart of predictive FX log returns  
1, 5, 10, 25, 50, 75, 90, 95 Conditional Quantiles



Credit: Lafarguette and Veyrune (2021)

# Application: MXN/USD Intervention Rule

**Figure 7. VaR FX Intervention Rule Based on a Given Information Set**



Credit: Lafarguette and Veyrune (2021)

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# From Point Forecasts to Distributional Forecasts

- A forecast  $\hat{y}_{T+h|T}$  is (usually) the mean of the conditional distribution:  $Y_{T+h}|Y_1, \dots, Y_T$
- Most models produce Gaussian distributed forecasts
- Because models assume Gaussian residuals: remember that the residuals are the stochastic components that are determining the distribution of both the estimators and the forecasts
- The forecast distribution describes the probability to forecast any future value

# Distributional Forecasts

Assuming the **residuals are uncorrelated** with variance  $\sigma^2$ , and with an estimate  $\hat{\sigma}^2$

## Model Typology

- **Mean:**  $Y_{T+h|T} = \mathcal{N}(\bar{Y}, (1 + \frac{1}{T})\hat{\sigma}^2)$
  - **Naive:**  $Y_{T+h|T} = \mathcal{N}(Y_T, h\hat{\sigma}^2)$
  - **Seasonal Naive:**  $Y_{T+h|T} = \mathcal{N}(Y_{T+h-m(k+1)}, (k+1)\hat{\sigma}^2)$
  - **Drift:**  $Y_{T+h|T} = \mathcal{N}(Y_T + \frac{h}{T-1}(Y_T - Y_1), h\frac{T+h}{T}\hat{\sigma}^2)$
- 
- where  $k$  is the integer part of  $\frac{h-1}{m}$
  - Note that when  $h = 1$  and  $T$  is large, these all give the same approximate variance  $\hat{\sigma}^2$

# Prediction Intervals

## Definition

A prediction interval gives **a region** within which we expect  $Y_{T+h}$  to lie with **a specified probability**

- Assuming that the forecasting errors are normally distributed, then a 95% prediction interval is:

$$Y_{T+h|T} \pm 1.96\hat{\sigma}_h$$

- where  $\hat{\sigma}_h$  is the standard deviation of the h-step distribution
- when  $h = 1$ ,  $\hat{\sigma}_h$  can be estimated from the residuals

# Prediction Interval Fit

```
brick_fc %>% hilo(level = 95)

## # A tsibble: 80 x 5 [1Q]
## # Key:     .model [4]
##   .model      Quarter    Bricks .mean     `95%
##   <chr>       <qtr>     <dist> <dbl>     <hilo>
## 1 Seasonal_naive 2005 Q3 N(428, 2336)  428 [333, 523]95
## 2 Seasonal_naive 2005 Q4 N(397, 2336)  397 [302, 492]95
## 3 Seasonal_naive 2006 Q1 N(355, 2336)  355 [260, 450]95
## 4 Seasonal_naive 2006 Q2 N(435, 2336)  435 [340, 530]95
## 5 Seasonal_naive 2006 Q3 N(428, 4672)  428 [294, 562]95
## 6 Seasonal_naive 2006 Q4 N(397, 4672)  397 [263, 531]95
## 7 Seasonal_naive 2007 Q1 N(355, 4672)  355 [221, 489]95
## 8 Seasonal_naive 2007 Q2 N(435, 4672)  435 [301, 569]95
## 9 Seasonal_naive 2007 Q3 N(428, 7008)  428 [264, 592]95
```

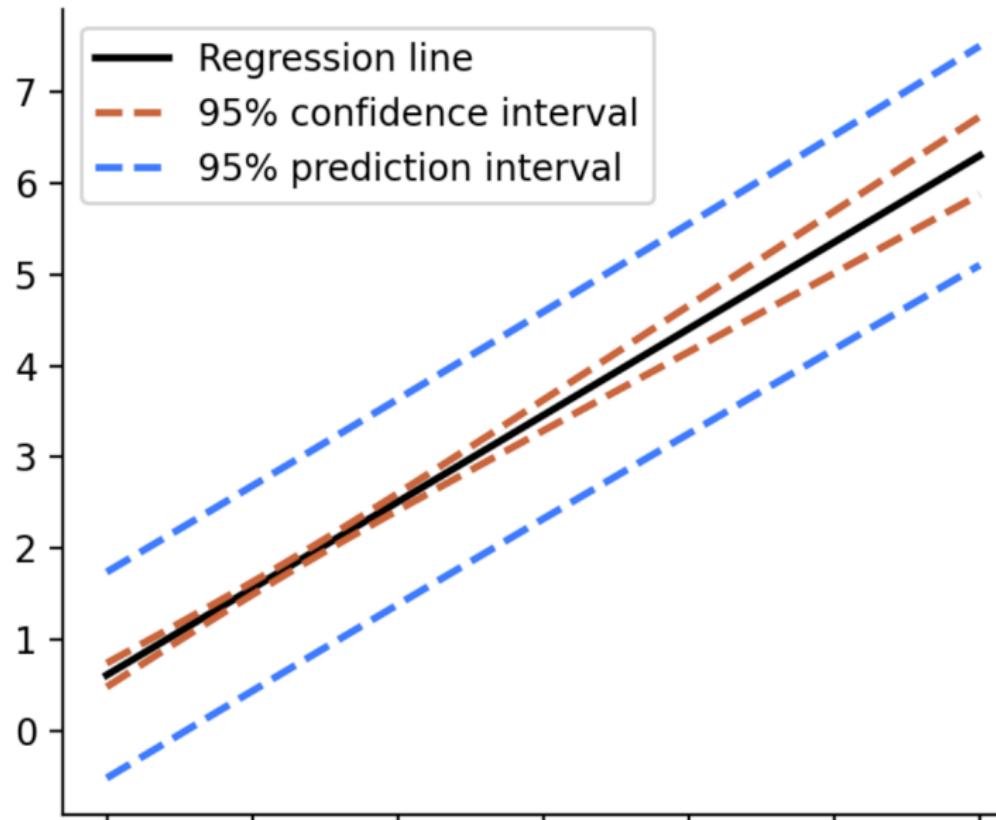
# Why Prediction Intervals Matter

- Point forecasts are often useless without a measure of uncertainty (such as prediction intervals)
- Prediction intervals require a **stochastic model** (with random errors, etc.)
- For most models, prediction intervals get wider as the forecast horizon increases
- The degree of confidence (the probability) impacts the width of the prediction interval
- Usually too narrow due to unaccounted uncertainty: pay attention

# Difference between Confidence Interval and Prediction Interval

- A confidence interval informs about where the true parameter of a model can be
  - **Confidence interval quantifies the uncertainty about the model**, or the distance between the model and reality
  - Confidence interval are associated with a wide range of parameters, values, etc.
  - They inform about how the model represents well the reality
  - Wide confidence intervals are associated with a less accurate model, and/or a very volatile model
- The prediction interval predicts in what range a future individual observation will fall
  - **Prediction interval quantifies the uncertainty about the future**, or the distance between today and the future
  - Prediction interval are not about the parameters of the model, but about the dependent variable ( $y_t$ )
  - The problem is that prediction intervals tend to neglect the uncertainty about the parameters used to generate the forecasts...

# Difference between Confidence Interval and Prediction Interval



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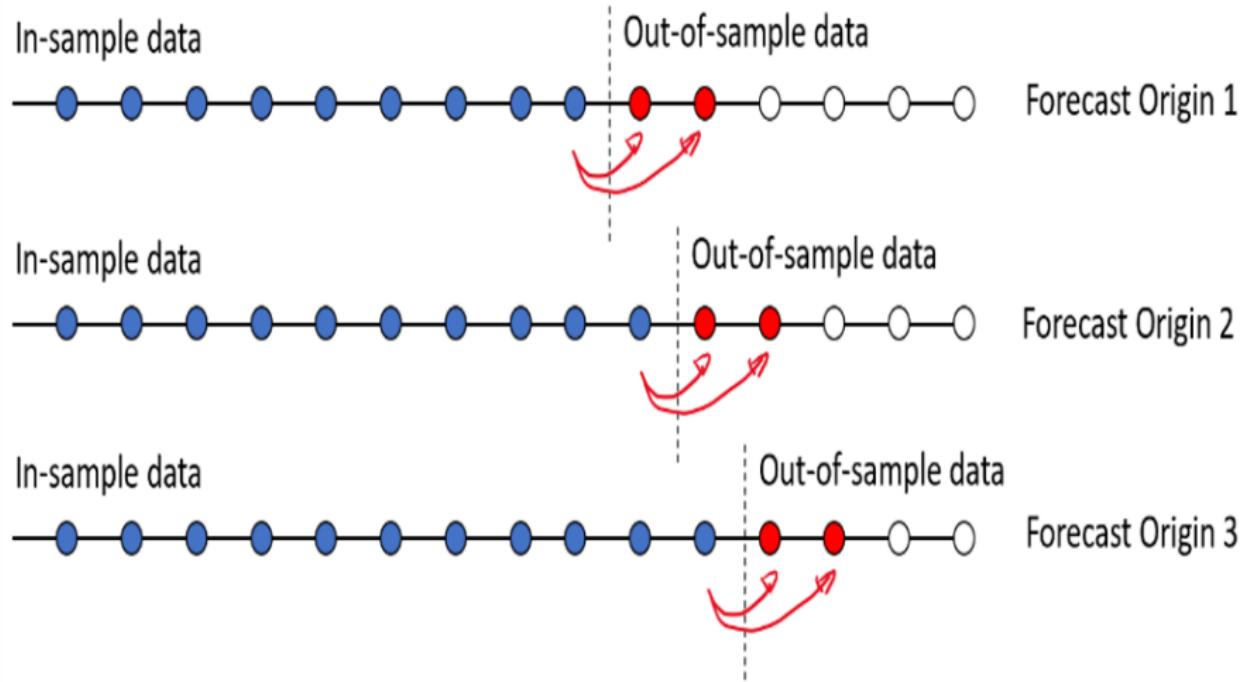
# Fitting and Forecasting

## Be careful

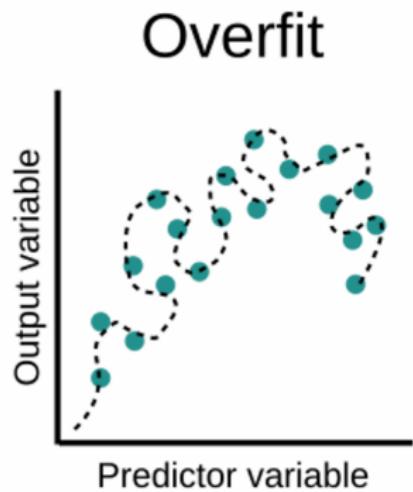
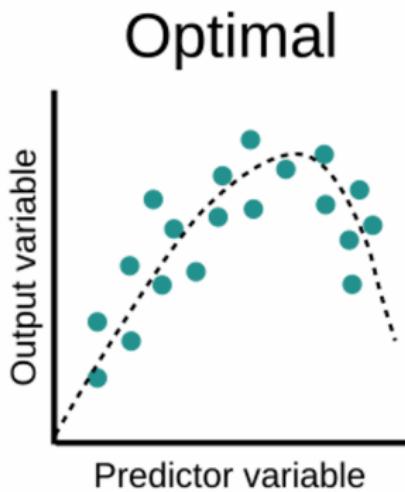
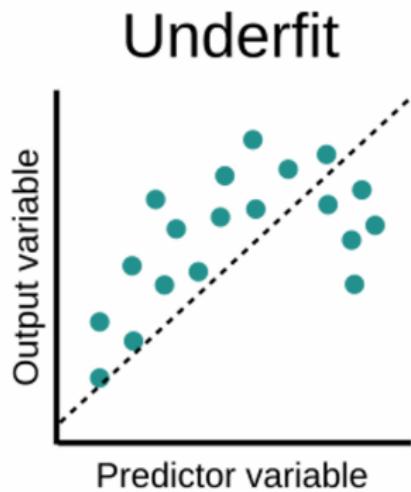
A model that fits the data well (in sample) might not necessarily forecast well

- A perfect in-sample fit can always be obtained by using a model with enough parameters
- Over-fitting a model to data is just as bad as failing to identify a systematic pattern in the data
- Need to split the model between
- The test set must not be used to *any* aspect of model development or calculation of forecasts
- Forecast accuracy is only based on the test set

## Train and Test Set



# Underfit, Optimal, Overfit



# Forecast Errors

## Definition: Forecast Errors

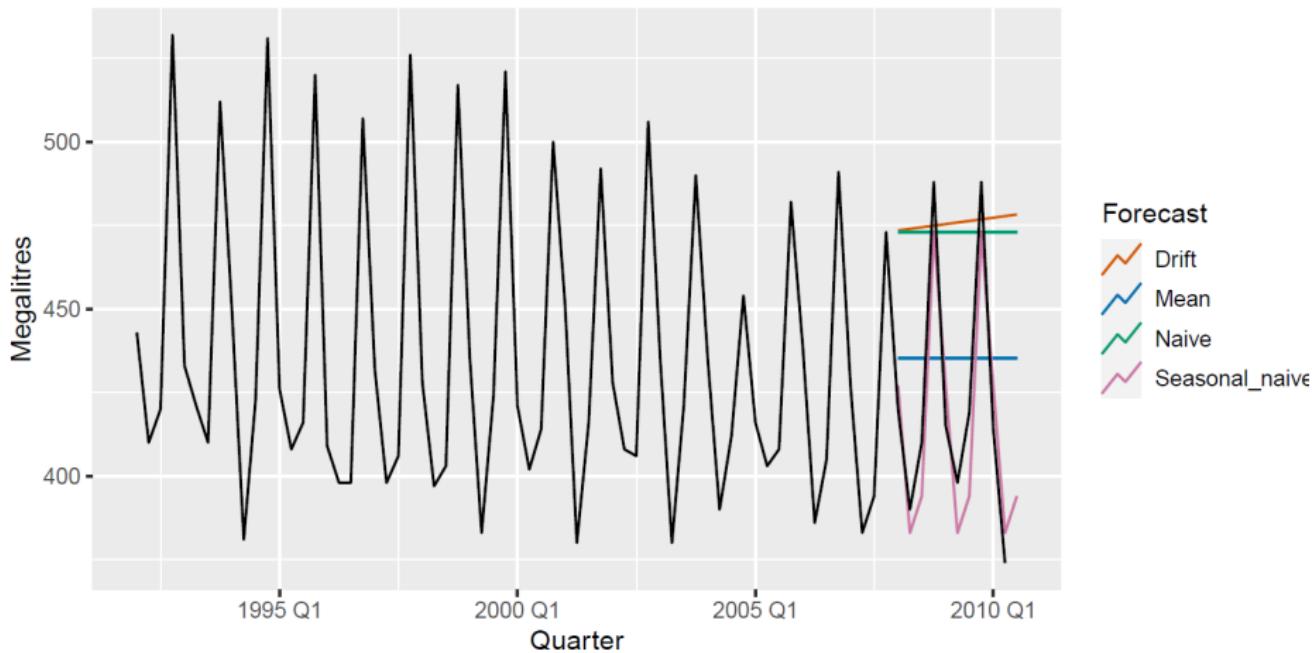
A forecast error is the difference between an observed value and its forecast

$$e_{T+h} = y_{T+h} - \hat{y}_{T+h}|Y_T, \dots, Y_1$$

- The conditional set  $Y_T, \dots, Y_1$  should only be taken from the training dataset
- The true value  $y_{T+h}$  is taken from the test set
- Unlike residuals, forecast errors on the test involve multi-step forecasts
- These are the **true** forecast error, as the test data is not used to compute  $\hat{y}_{T+h}$

# Example: Forecasting Beer Production

Forecasts for quarterly beer production



# Measures of Forecast Accuracy

## Main Metrics

- **MAE:** mean absolute errors  $\frac{1}{S} \sum_{s \in S} |e_{s,T+h}|$
- **MSE:** mean squared errors  $\frac{1}{S} \sum_{s \in S} (e_{s,T+h})^2$
- **MAPE:** mean absolute percentage errors  $\frac{1}{S} 100 * \sum_{s \in S} \frac{|e_{s,T+h}|}{|y_{s,t+h}|}$
- **RMSE:** root mean squared errors:  $\sqrt{\frac{1}{S} \sum_{s \in S} (e_{s,T+h})^2}$

With:

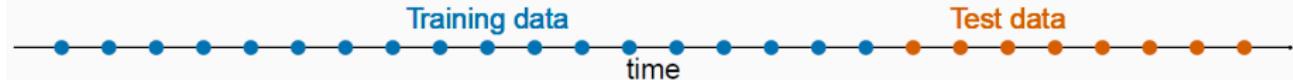
- $y_{T+h}$ : T+h observation, h being the horizon ( $h = 1, 2, \dots, H$ )
- $\hat{y}_{T+h|T}$ : the forecast based on data up to time  $T$
- $e_{T+h} = y_{T+h} - \hat{y}_{T+h|T}$ : The forecast errors
- $S$  is the testing sample

# Scaling

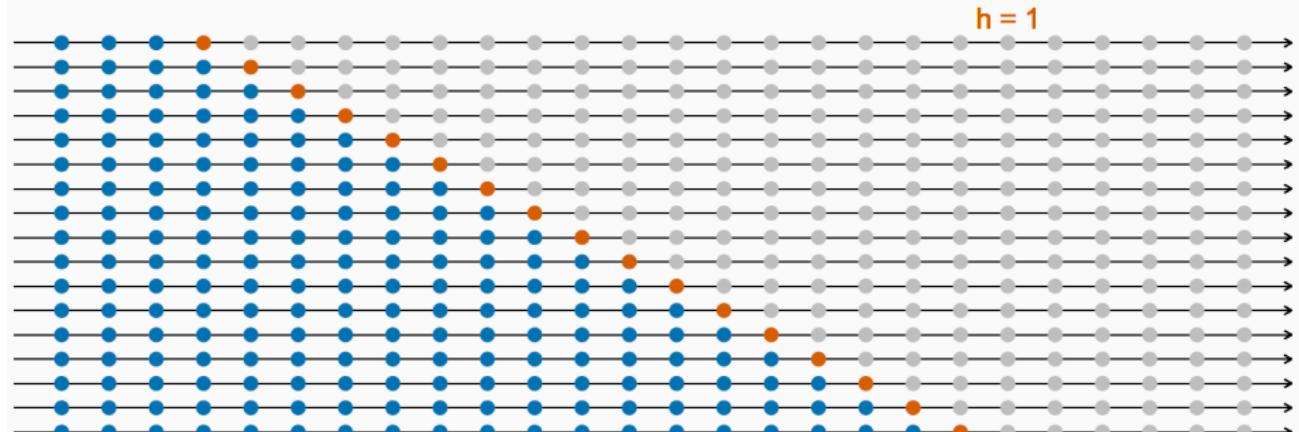
- MAE, MSE and RMSE are all **scale dependent**
- MAPE is scale independent but is only sensible if  $y_t >> 0 \quad \forall t$
- **Most commonly used: Time Cross-Validation with the lowest RMSE**

# Time Series Cross-Validation

## Traditional evaluation

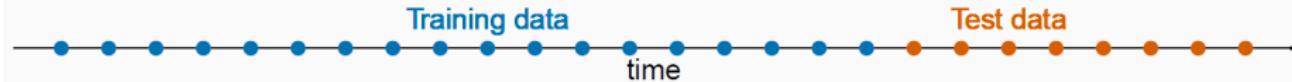


## Time series cross-validation

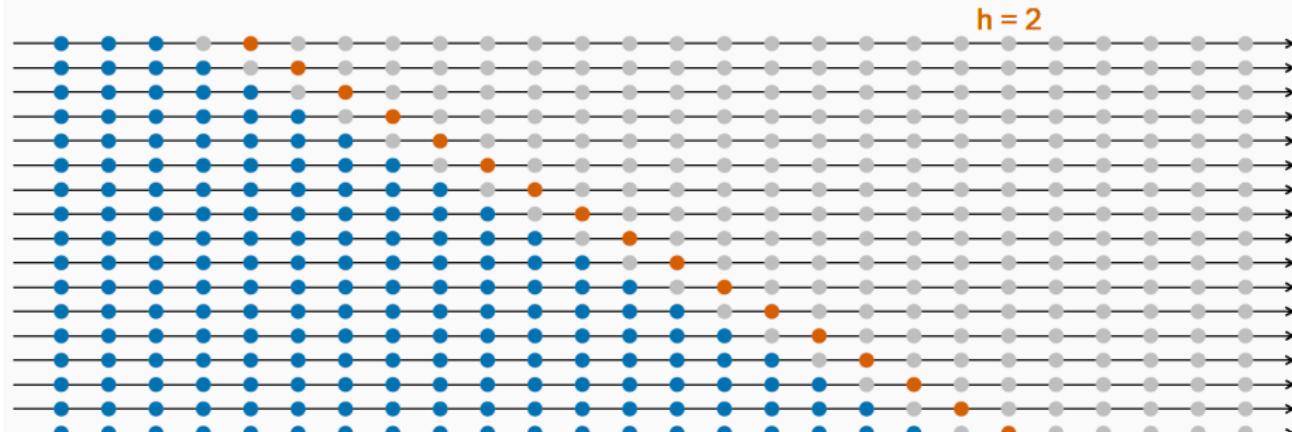


# Time Series Cross-Validation

## Traditional evaluation



## Time series cross-validation

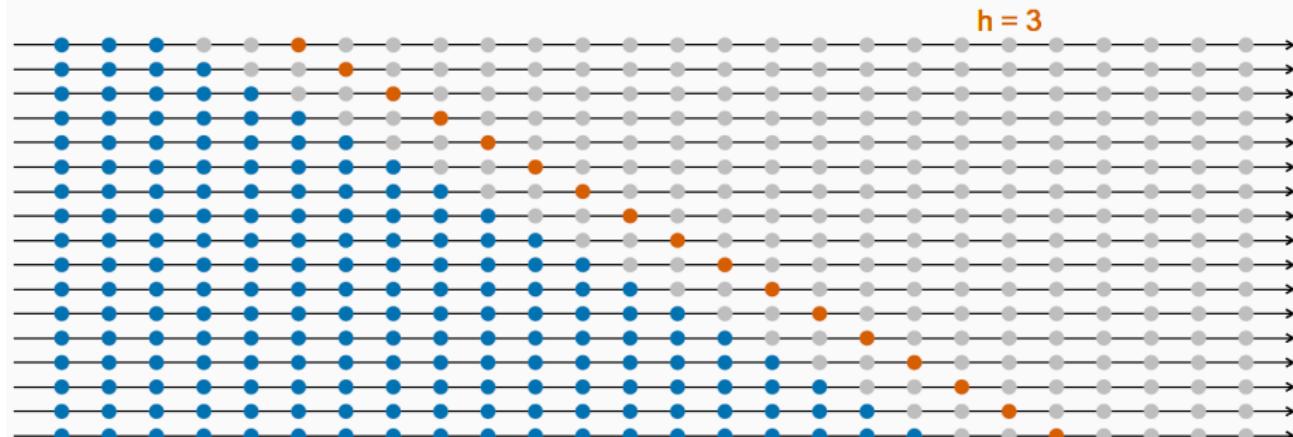


# Time Series Cross-Validation

## Traditional evaluation

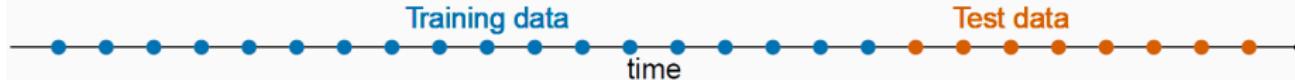


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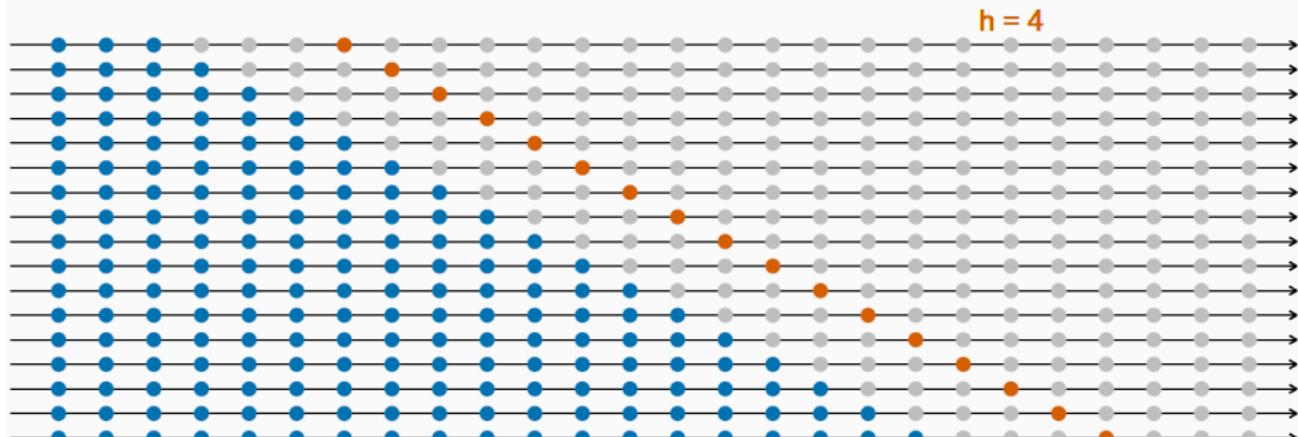


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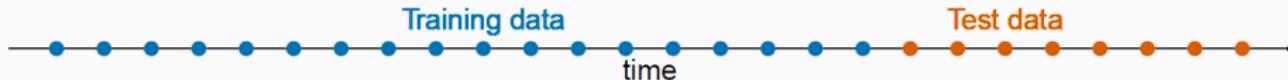


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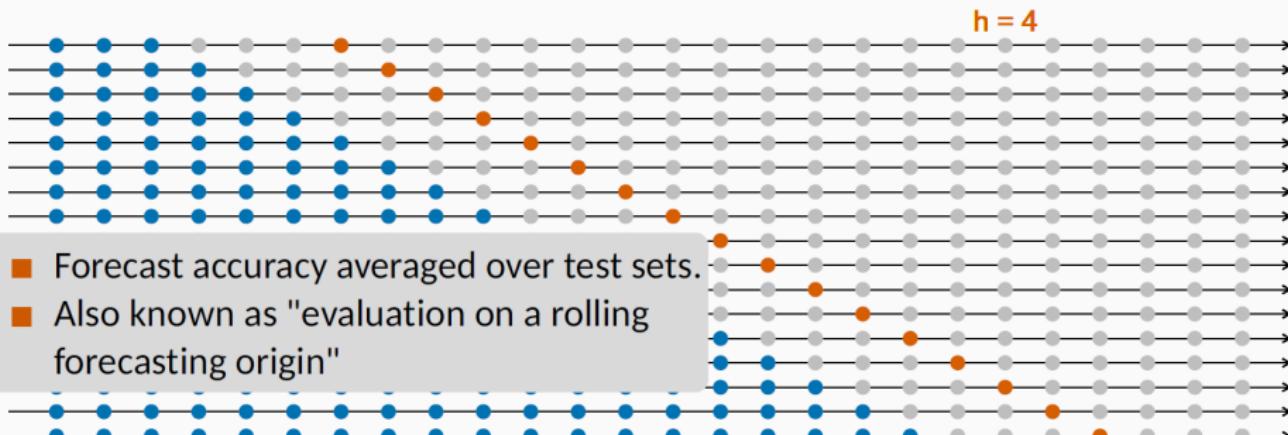


# Time Series Cross-Validation

## Traditional evaluation



## Time series cross-validation



# Framework RMSE Output

## RMSE

The RMSE is defined as

$$RMSE = \sqrt{\sum_{t \in T_{real}} (y_t - \hat{y}_t(h))^2}$$

where  $y_t$  is the observed value and  $\hat{y}_t(h)$  is the h-step ahead forecast of  $y_t$ . The RMSE for each model and HORIZON is shown below. Lower values indicate better forecasts, the best individual model is in bold, while combinations are colored.

		Show 10 entries Search: <input type="text"/>														
	Method	h=1	h=2	h=3	h=4	h=5	h=6	h=7	h=8	h=9	h=10	h=11	h=12	h=13	h=14	mean
1	Dynamic Model Selection	211.6	457.5	1,097.4	1,599.8	2,048.9	2,315.7	2,400.4	2,501.1	2,622.4	3,204.7	3,451.4	3,672.2	3,573.2	3,604.6	2,340.1
2	Model Average (best 2)	174.8	374.8	1,308.2	1,661.9	2,028.1	2,405.4	2,554.2	2,719.8	2,899.5	3,241.8	3,392.5	3,525.4	3,582.3	3,653.2	2,394.4
3	ARIMA	210.0	491.7	1,435.5	1,785.8	2,146.9	2,468.8	2,608.2	2,775.3	2,965.3	3,245.6	3,371.7	3,472.3	3,484.7	3,549.0	2,429.4
4	ETS	179.7	355.0	1,250.8	1,624.1	2,000.8	2,405.1	2,558.2	2,726.0	2,904.0	3,312.7	3,493.8	3,661.4	3,744.5	3,815.3	2,430.8
5	ETS (non-seasonal)	206.8	424.1	1,389.8	1,781.4	2,092.3	2,414.7	2,548.7	2,709.6	2,893.1	3,337.7	3,510.2	3,649.8	3,712.3	3,783.3	2,461.0
6	Model Average (all models)	175.3	371.9	1,342.3	1,721.6	2,085.8	2,471.4	2,622.7	2,795.4	2,965.0	3,401.1	3,576.6	3,729.1	3,819.1	3,896.2	2,499.5
7	Seasonal ARIMA	161.6	351.6	1,428.1	1,863.5	2,289.5	2,742.0	2,911.0	3,106.6	3,320.9	3,883.0	4,120.1	4,331.3	4,503.8	4,594.2	2,829.1

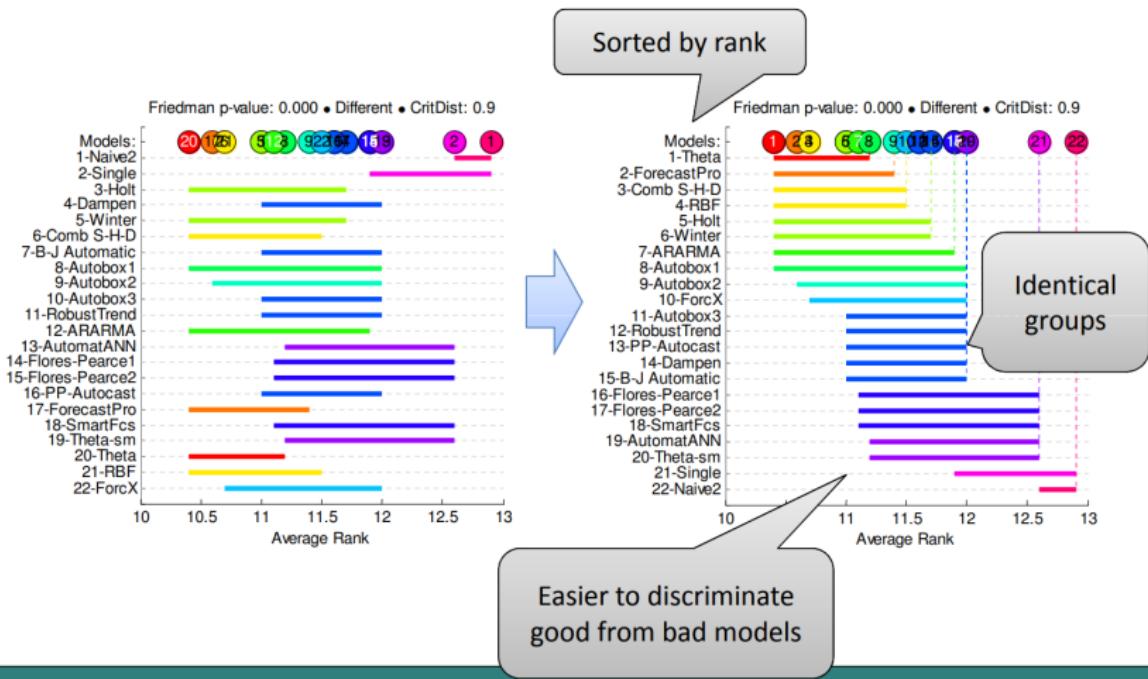
Credit: IMF Framework

# Nemenyi Test

- We can rank the model by RMSE (or another metric), but are the RMSE significantly different?
- Maybe Model 1 can have a lower RMSE than Model 2, but the difference in RMSE is non-significant
- In which case, we could pool the two models together
- Use a non-parametric test to test the hypothesis of equal RMSE, with the test statistic:

$$r_{\alpha,K,N} \approx \frac{q_{\alpha,K}}{\sqrt{2}} \sqrt{\frac{K(K+1)}{6N}}$$

# Nemenyi Test in Practice



Credit: Nikolaos Kourentzes

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- ② Volatility Models: EWMA and GARCH Models
- ③ Distributional Forecasts and Prediction Intervals
- ④ Evaluating Forecast Accuracy
- ⑤ Model Combination and Selection
- ⑥ Forecasts Reconciliation

## Combination of Models

- Combining model provides a more reliable forecast, overcoming the risk of relying on a single model
- Estimate all the models, determine the cross-validated errors
- Define a pool of "good" models and reject the other models, based on their out of sample performance. Define a **rejection threshold** to determine the "bad" models, based on the distribution of the RMSE

$$\text{Threshold} = Q(0.75) + 1.5 * \text{IQR}$$

- Where  $Q(0.75)$  is the 75th quantile of the RMSE distribution and IQR the interquantile range ( $Q(0.75) - Q(0.25)$ )
- Then, weight each forecast of the eligible model by the relative performance of their model, the best models having the highest weight

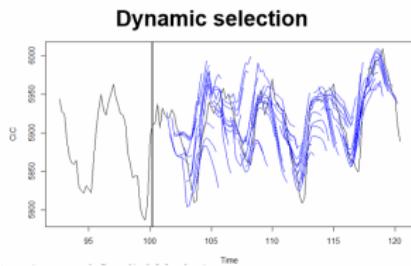
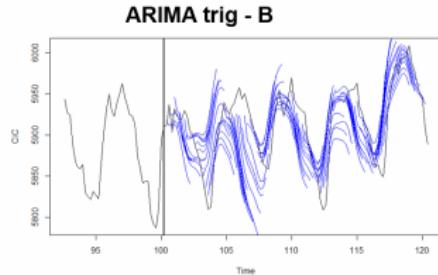
# Dynamic Model Selection: Principles

- Some models are better for modelling long-term dynamics, while others are better for short-term dynamics
- Time-varying volatility matters. As volatility increases, models focusing on the short term have an advantage
- Dynamically account for changes in the behaviour of the time series:
  - Combining high performing forecasts
  - Selecting a well performing forecast based on local performance

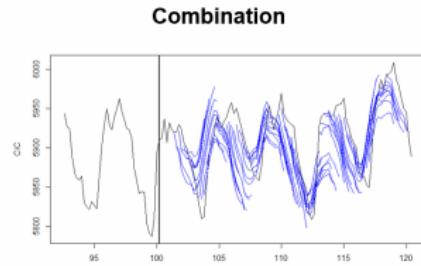
# Dynamic Model Selection: Implementation

- **Combining high-performance forecasts:**
  - Compute unweighted average of forecasts. Although it seems better to estimate “optimal” weights, weights estimation introduces additional uncertainty
- Selecting a well performing forecast based on **recent performance**
  - **Rolling performance of the previous week** to select the forecast for the next origin.
- Consider ETStrig, ARIMAtrig, ARIMA stoch, and TBATS comp.  
All these models capture information differently
  - **Model diversity is advantageous** both for selection and combination between forecasts

# Out-of-Sample Performances via Rolling Origins



- Both selection and combination have fewer extreme forecasts.
  - ▶ Combination provides smoother forecasts (combining models acts as a shrinkage estimator)
  - ▶ while the selection is able to pick up details



Credit: Author

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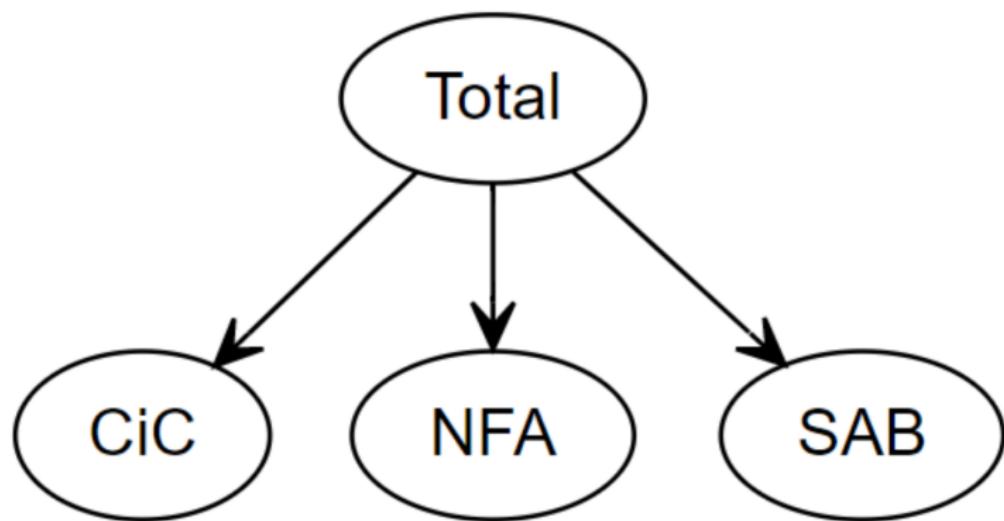
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# Forecasts Reconciliation of the Autonomous Factors

Forecast the autonomous factors using a hierarchical modeling approach, with different models for each autonomous factor:

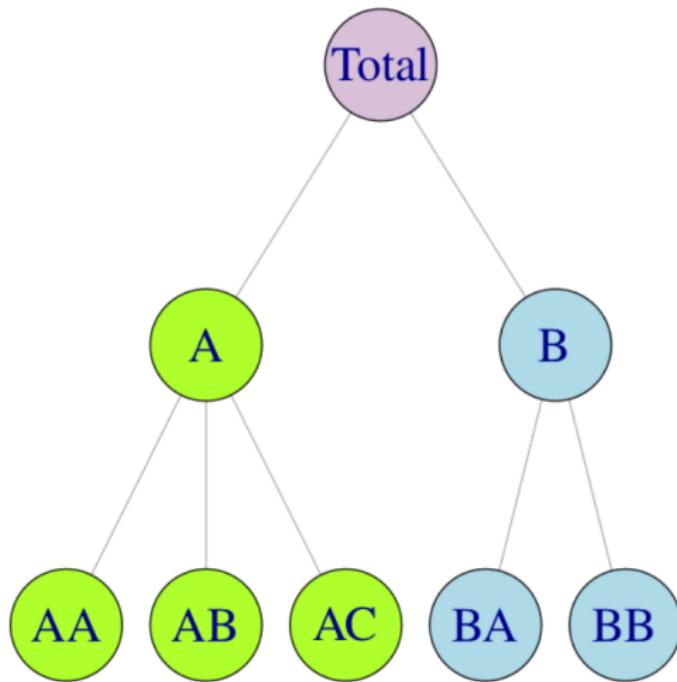
- Forecasting the sum of the autonomous factors separately
- Reconciling the forecasts using aggregation technics as described for instance in Hyndman, Rob J., et al. "Optimal combination forecasts for hierarchical time series." (2011)
- Use either OLS or the “minT” approach: minimize the trace of the VarCov forecast matrix. Exploit the covariance among regressors to maximize the accuracy of the hierarchical forecasts

# Reconciliation via Hierarchical Forecasts



Credit: *Author*

# General Approach



Credit: <https://otexts.com/fpp3/reconciliation.html>

# Formalization

- We can express the total (top level,  $y_t$ ) via a summing matrix  $S$  of the low level elements  $y_t^L$

$$y_t = S y_t^L$$

- Then, we can express the whole system via a grouping matrix summing all the individual elements (also called base forecast  $y_t^b$ ), so have a consistent approach irrespective of the level structure

$$\tilde{y}_t = S * G * y_h^b$$

$$\begin{bmatrix} y_t \\ y_{A,t} \\ y_{B,t} \\ y_{AA,t} \\ y_{AB,t} \\ y_{AC,t} \\ y_{BA,t} \\ y_{BB,t} \end{bmatrix} = \begin{bmatrix} 1 & 1 & 1 & 1 & 1 \\ 1 & 1 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 1 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} y_{AA,t} \\ y_{AB,t} \\ y_{AC,t} \\ y_{BA,t} \\ y_{BB,t} \end{bmatrix} \quad G = \begin{bmatrix} 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 \end{bmatrix}$$

Credit: [otexts.com/fpp3/reconciliation.html](http://otexts.com/fpp3/reconciliation.html)

# The Min Trace Approach

## Forecasts Errors in a Hierarchical System

$$V_h = \mathbb{V}[y_{T+h} - \tilde{y}_h] = SGW_hG'S'$$

where  $W_h = \mathbb{V}[y_{T+h} - \tilde{y}_h^b]$  the variance-covariance matrix of the base vector

- The objective is to find a matrix  $G$  to minimize the variance-covariance matrix of the corresponding base forecast errors
- Because the error variances are on the diagonal of  $V_h$ , minimizing the forecasting errors is equivalent to minimizing the trace of  $V_h$
- Wickramasuriya et al. (2019) show that the minimized trace matrix is given by:

$$G = (S'W_h^{-1}S)^{-1}S'W_h^{-1}$$

- This approach is called the **MinT (or Minimum Trace) optimal reconciliation approach**