

# Week 2: Multiple Regression, Training Exercises

Coursera/Erasmus U., Econometric Methods and Applications

*Anthony Nguyen*

## Training Exercise 2.4.1

### Questions

By solving the questions of this exercise, you provide a proof of the Gauss-Markov theorem. We use the following notation:

- The OLS estimator is  $b = A_0 y$ , where  $A_0 = (X'X)^{-1}X'$ .
- Let  $\hat{\beta} = Ay$  be linear, unbiased, with  $A(k \times n)$  matrix.
- Define the difference matrix  $D = A - A_0$ .

**(a) Prove the following three results:**

- (i)  $\text{var}(\hat{\beta}) = \sigma^2 AA'$ .
- (ii)  $\hat{\beta}$  unbiased implies  $AX = I$  and  $DX = 0$ .
- (iii) Part (ii) implies  $AA' = DD' + (X'X)^{-1}$ .

**(b) Prove that part (a-iii) implies  $\text{var}(\hat{\beta}) = \sigma^2 DD'$ .**

**(c) Prove that part (b) implies  $\text{var}(\hat{\beta}) - \text{var}(b)$  is positive semidefinite (Gauss-Markov).**

**(d) Prove that  $\text{var}(\hat{\beta}) \geq \text{var}(b_j)$  for every  $j = 1, \dots, k$ .**

---

## Answers

(a) Prove the following three results:

(i)  $\text{var}(\hat{\beta}) = \sigma^2 AA'$ .

Given that  $y = X\beta + \varepsilon$ , then we find that:

$$\begin{aligned}\hat{\beta} &= Ay \\ &= A(X\beta + \varepsilon) \\ &= AX\beta + A\varepsilon\end{aligned}$$

Now given that  $A$ ,  $X$ , and  $\beta$  are fixed and  $E(\varepsilon) = 0$ , we find that:

$$\begin{aligned}E(\hat{\beta}) &= AX\beta + AE(\varepsilon) \\ &= AX\beta \\ \text{var}(\hat{\beta}) &= E((\hat{\beta} - E(\hat{\beta}))(\hat{\beta} - E(\hat{\beta}))') \\ &= E(A\varepsilon(A\varepsilon)') \\ &= E(A\varepsilon\varepsilon'A') \\ &= AE(\varepsilon\varepsilon')A' \\ &= A\sigma^2 IA' \\ &= \sigma^2 AA'\end{aligned}$$

(ii)  $\hat{\beta}$  unbiased implies  $AX = I$  and  $DX = 0$ .

$E(\hat{\beta})$  is unbiased if  $AX\beta = \beta$  for all  $\beta$ , or:

$$\begin{aligned}DX &= (A - A_0)X \\ &= AX - A_0X \\ &= I - (x'x)^{-1}X'X \\ &= I - I \\ &= 0\end{aligned}$$

(iii) Part (ii) implies  $AA' = DD' + (X'X)^{-1}$ .

$$\begin{aligned} AA' &= (D + A_0)(D + A_0)' \\ &= DD' + A_0D' + DA'_0 + A_0A'_0 \end{aligned}$$

$$DA'_0 = DX(X'X)^{-1} = 0 \cdot (X'X)^{-1} = 0$$

$$A_0D' = (DA'_0)^{-1} = 0$$

$$\begin{aligned} AA' &= DD' + A_0A'_0 \\ &= DD' + (X'X)^{-1}X'X(X'X)^{-1} \\ &= DD' + (X'X)^{-1} \end{aligned}$$

**(b) Prove that part (a-iii) implies  $\text{var}(\hat{\beta}) = \sigma^2 DD'$ .**

Given that  $\text{var}(b) = \sigma^2(X'X)^{-1}$ , we find that:

$$\begin{aligned} \text{var}(\hat{\beta}) &= \sigma^2 AA' \\ &= \sigma^2(DD' + (X'X)^{-1}) \\ &= \sigma^2 DD' + \sigma^2(X'X)^{-1} \\ &= \sigma^2 DD' + \text{var}(b) \end{aligned}$$

**(c) Prove that part (b) implies  $\text{var}(\hat{\beta}) - \text{var}(b)$  is positive semidefinite (Gauss-Markov).**

From (b) we've seen that  $\text{var}(\hat{\beta}) - \text{var}(b) = \sigma^2 DD'$ . Given that  $\sigma^2 > 0$ , then it remains to be shown that  $DD'$  is positive semidefinite (p.s.d.):

$D$  is a  $(k \times n)$  matrix

$D'$  is a  $(n \times k)$  matrix

$c$  is a  $(n \times 1)$  vector

$d = D'c$  is a  $(n \times 1)$  vector with components  $d_1, \dots, d_n$

$$\begin{aligned} \therefore c' DD' c &= (D'c)' D'c \\ &= d' d \\ &= \sum_{i=1}^n d_i^2 \geq 0 \quad \text{proving that } DD' \text{ is p.s.d.} \end{aligned}$$

**(d) Prove that  $\text{var}(\hat{\beta}) \geq \text{var}(b_j)$  for every  $j = 1, \dots, k$ .**

Let  $c_j$  be the  $(k \times 1)$   $j^{\text{th}}$  unit vector, which means  $c_j$  can be re-written as:

$$c_j = \begin{bmatrix} 0 \\ \vdots \\ 0 \\ 1 \\ 0 \\ \vdots \\ 0 \end{bmatrix} \quad \text{with the single element 1 at the } j^{th} \text{ entry}$$

from  $c$  we obtain that:

$$c_j'(var(\hat{\beta}) - var(b))c_j = \sigma^2 c_j' D D' c_j \geq 0$$

hence:

$$c_j' var(\hat{\beta}) c_j \geq c_j' var(b) c_j$$

or equivalently:

$$var(c_j' \hat{\beta}) \geq var(c_j' b)$$

As  $c_j' \hat{\beta} = \hat{\beta}_j$  and  $c_j' b = b_j$ , it follow that:

$$var(\hat{\beta}_j) \geq var(b_j) \quad \text{for every } j = 1, \dots, n$$