

Examen

2 Iunie 2018



Timp de lucru 2h30. Toate documentele, computerele personale, telefoanele mobile și/sau calculatoarele electronice de mână sunt autorizate. Orice modalitate de comunicare între voi este **strict interzisă**. Mult succes !

Exercițiul 1

10p

Fie X_1, X_2, \dots, X_n un eșantion de talie n dintr-o populație Poisson de parametru $\theta > 0$.

- Determinați estimatorul de verosimilitate maximă $\hat{\theta}$ și verificați dacă acesta este deplasat, consistent și eficient.
- Găsiți estimatorul de verosimilitate maximă pentru $\mathbb{P}_\theta(X_1 = 1 \mid X_1 > 0)$. Este acesta consistent ?
- Verificați dacă estimatorul aflat la punctul b) este sau nu nedepășat.

Exercițiul 2

10p

Fie X o variabilă aleatoare repartizată $\mathbb{P}_\theta(X = k) = A(k+1)\theta^k$, $k \in \mathbb{N}$ unde $\theta \in (0, 1)$ un parametru necunoscut și $A \in \mathbb{R}$ este o constantă.

- Determinați constanta A și calculați $\mathbb{E}[X]$ și $Var(X)$.

Dorim să estimăm pe θ plecând de la un eșantion X_1, X_2, \dots, X_n de talie n din populația dată de repartiția lui X .

- Determinați estimatorul $\tilde{\theta}$ a lui θ obținut prin metoda momentelor și calculați $\mathbb{P}_\theta(\tilde{\theta} = 0)$.
- Determinați estimatorul de verosimilitate maximă $\hat{\theta}$ a lui θ și verificați dacă acesta este bine definit.
- Studiați consistența estimatorului $\tilde{\theta}$ și determinați legea lui limită.

Exercițiul 3

10p

Calculați marginea Rao-Cramer pentru familia $\mathcal{N}(\mu, 1)$ unde μ este necunoscut. Determinați estimatorul obținut prin metoda momentelor și verificați dacă este eficient.

Exercițiul 4

10p

Considerăm următorul eșantion de talie 20 dintr-o populație Bernoulli de parametru $\theta \in (0, 1)$:

0 1 1 0 1 1 0 1 1 0 0 1 1 1 1 1 0 0 0 0

- Găsiți estimatorul de verosimilitate maximă $\hat{\theta}$ și determinați informația lui Fisher $I(\theta)$.
- Determinați estimatorul de verosimilitate maximă pentru $\mathbb{V}_\theta[X_1]$. Este acesta nedepășat? Dar consistent? Justificați răspunsul.
- Construiți un interval de încredere pentru $\hat{\theta}$ de nivel 95%.

Ex 1

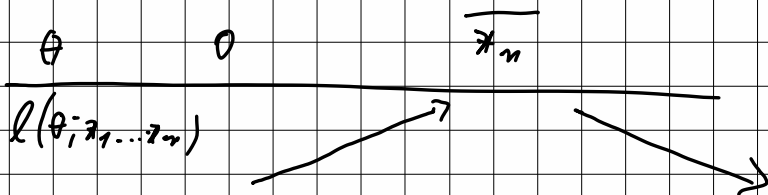
$$X_1, X_2, \dots, X_n \sim P(\theta)$$

$$f_\theta(h) = P(X=h) = e^{-\theta} \frac{\theta^h}{h!}, \quad h \in \mathbb{N}$$

$$L(\theta; x_1, \dots, x_n) = e^{-n\theta} \frac{\theta^{x_1 + \dots + x_n}}{(x_1 + \dots + x_n)!}$$

$$l(\theta; x_1, \dots, x_n) = (x_1 + \dots + x_n) \log(\theta) - n\theta - \ln(x_1 + \dots + x_n)!$$

$$\frac{dl}{d\theta}(\theta; x_1, \dots, x_n) = (x_1 + \dots + x_n) \cdot \frac{1}{\theta} - n$$



Derivates $l(\theta; x_1, \dots, x_n) = \overline{x}_n$

$$\hat{\theta} = \overline{x}_n$$

Déclarées:

$$\mathbb{E}[X_1] = \sum_{h \geq 0} h \cdot e^{-\theta} \cdot \frac{\theta^h}{h!}$$

$$\sum_{h \geq 0} e^{-\theta} \cdot \frac{\theta^h}{h!} = 1$$

$$\sum_{h \geq 0} \frac{\theta^h}{h!} = e^{\theta} \quad \Big| \quad \frac{d}{d\theta}$$

$$\sum_{h \geq 0} h \cdot \frac{\theta^{h-1}}{h!} = e^{\theta} \quad \Big| \cdot \theta$$

$$\sum_{h \geq 0} h \cdot e^{-\theta} \cdot \frac{\theta^h}{h!} = \theta$$

$$\text{Donc } \mathbb{E}[X_1] = \theta = \mathbb{E}(\bar{X}_n) = \mathbb{E}(\hat{\theta})$$

Donc $\hat{\theta}$ est non biaisé.

$$\hat{\theta} = \bar{X}_n \xrightarrow{IP} \mathbb{E}[X_1] = \theta \quad (L.N.M.) =,$$

$\Rightarrow \hat{\theta}$ est consistant

$$\text{Var}(\bar{X}_n) = \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i)$$

$$= \frac{1}{n} \text{Var}(X_1)$$

$$\text{Var}(X_1) = \sum_{h=0}^{\infty} h^2 \cdot e^{-\theta} \cdot \frac{\theta^h}{h!} - \theta^2$$

$$\sum_{h=0}^{\infty} h \cdot \frac{\theta^h}{h!} = e^{-\theta} \cdot \theta \left| \frac{d}{d\theta} \right|$$

$$\sum_{h=0}^{\infty} h^2 \cdot \frac{\theta^{h-1}}{h!} = e^{-\theta} \cdot \theta + e^{-\theta} \cdot \theta$$

$$\sum_{h=0}^{\infty} h^2 \cdot \frac{\theta^h}{h!} = e^{-\theta} (1 + \theta) \cdot \theta$$

$$\text{Var}(X_1) = (1 + \theta) \cdot \theta - \theta^2 = \theta$$

$$\text{Var}(\hat{\theta}) = \text{Var}(\bar{X}_n) = \frac{\theta}{n}$$

$$I_1 = E_{\theta} \left[\left(\frac{d}{d\theta} \log f_{\theta}(X_1) \right)^2 \right] = \sum_{h=0}^{\infty} \left(\frac{d}{d\theta} \log f_{\theta}(h) \right)^2 \cdot f_{\theta}(h)$$

$$= \sum_{h=0} \frac{(f'_\theta(h))^2}{(f_\theta(h))^2} \cdot f_\theta(h)$$

(derivative
in exp. w/ θ)

$$f_\theta(h) = e^{-\theta} \cdot \frac{\theta^h}{h!}$$

$$f'_\theta(h) = -e^{-\theta} \cdot \frac{\theta^h}{h!} + e^{-\theta} \cdot h \cdot \frac{\theta^{h-1}}{h!}$$

$$= -e^{-\theta} \cdot \frac{\theta^h}{h!} + e^{-\theta} \cdot \frac{\theta^h}{(h-1)!}$$

$$= e^{-\theta} \cdot \frac{\theta^h}{h!} \cdot \left(-1 + \frac{h}{\theta}\right)$$

$$I_1 = \sum_{h=0} \left(\frac{h}{\theta} - 1\right)^2 \cdot e^{-\theta} \cdot \frac{\theta^h}{h!}$$

$$= \frac{1}{\theta^2} \sum_{h=0} h^2 \cdot e^{-\theta} \cdot \frac{\theta^h}{h!} - \frac{2}{\theta} \cdot \sum_{h=0} h \cdot e^{-\theta} \cdot \frac{\theta^h}{h!} + \sum_{h=0} e^{-\theta} \cdot \frac{\theta^h}{h!}$$

$$= \frac{1}{\theta^2} \cdot \theta \cdot (\theta + 1) - \frac{2}{\theta} \cdot \theta + 1 = \frac{1}{\theta}$$

Ineq. Rao-Cramer :

$$\text{Var}(\hat{\theta}) \geq \frac{1}{I_n} = \frac{1}{n \cdot I_1}$$

$$\frac{\theta}{n} \geq \frac{1}{n \cdot \frac{1}{\theta}}$$

Avem egalitate \Rightarrow
 $\Rightarrow \hat{\theta}$ e eficient

b,

$$\begin{aligned} P_{\theta}(X_1=1 | X_1 \geq 0) &= \frac{P_{\theta}(X_1=1)}{P_{\theta}(X_1 \geq 0)} \\ &= \frac{e^{-\theta} \cdot \frac{\theta}{1!}}{1 - e^{-\theta} \cdot \frac{1}{0!}} \\ &= \frac{e^{-\theta} \cdot \theta}{1 - e^{-\theta}} \end{aligned}$$

$$g: (0, \infty) \rightarrow (0, \infty)$$

$$g(\theta) = \frac{e^{-\theta} \cdot \theta}{1 - e^{-\theta}}$$

\hat{X}_n e estimatorul de verosimilitate
maximă pt. $\theta \Rightarrow$

$\Rightarrow g(\bar{X}_n)$ e est. de ver. maximă
pt. $P_\theta(X_1=1 | X_1>0)$

Consistența: g continuă /
L.N.M: $\bar{X}_n \xrightarrow{IP} \theta$ / $T_{\text{apl. cont.}}$

$$g(\bar{X}_n) \xrightarrow{IP} g(\theta) = P_\theta(X_1=1 | X_1>0)$$

Deci g e consistent

$$\begin{aligned} c) \quad g(\bar{X}_n) &= \frac{e^{-\bar{X}_n} \cdot \bar{X}_n}{1 - e^{-\bar{X}_n}} = \frac{\bar{X}_n}{e^{\bar{X}_n} - 1} \\ &= \frac{X_1 + \dots + X_n}{n \cdot (e^{\frac{X_1 + \dots + X_n}{n}} - 1)} \end{aligned}$$

$$P(X_1 + \dots + X_n = l) =$$

$$= \sum_{i=0}^l P(X_1 + \dots + X_{n-1} = i) \cdot P(X_n = l-i)$$

$$\begin{aligned}
 P(X_1 + X_2 = l) &= \sum_{i=0}^l e^{-2\theta} \frac{\theta^i}{i!} \cdot \frac{\theta^{l-i}}{(l-i)!} \\
 &= \sum_{i=0}^l e^{-2\theta} \frac{\theta^l}{i! \cdot (l-i)!} \\
 &= \frac{e^{-2\theta}}{l!} \cdot \theta^l \cdot \underbrace{\sum_{i=0}^l \binom{l}{i}}_{2^l} \\
 &= e^{-2\theta} \cdot \frac{2^l \theta^l}{l!}
 \end{aligned}$$

$$\begin{aligned}
 P(X_1 + X_2 + X_3 = l) &= \\
 &= \sum_{i=0}^l e^{-3\theta} \cdot \frac{(2\theta)^i}{i!} \cdot \frac{\theta^{l-i}}{(l-i)!} \\
 &= e^{-3\theta} \cdot \frac{\theta^l}{l!} \cdot \underbrace{\sum_{i=0}^l 2^i \cdot \binom{l}{i}}_{\substack{\text{Binom Newton} \\ = (1+2)^l}} \\
 &= e^{-3\theta} \cdot \frac{(3\theta)^l}{l!}
 \end{aligned}$$

Prin inductie:

$$P(X_1 + \dots + X_n = h) = e^{-n\theta} \frac{(n\theta)^h}{h!}$$

$$E[g(X_n)] = \sum_{h=0}^{\infty} \frac{h}{n \cdot (e^{\frac{h}{n}} - 1)}$$

$$\cdot P(X_1 + \dots + X_n = h)$$

$$= \sum_{h=0}^{\infty} \frac{h}{n \cdot (e^{\frac{h}{n}} - 1)} \cdot e^{-n\theta} \cdot \frac{(n\theta)^h}{h!}$$

(c) 2

$$P_{\theta}(X=h) = A(h+1)\theta^h$$

o)

$$\sum_{h=0}^{\infty} A \cdot (h+1) \cdot \theta^h = A \cdot \sum_{h=0}^{\infty} (h+1) \theta^h$$

$$\sum_{h=0}^{\infty} \theta^{h+1} = \theta \cdot \frac{1}{1-\theta} \quad \Bigg| \quad \frac{d}{d\theta}$$

$$\sum_{h=0}^{\infty} (h+1) \theta^h = \frac{1}{(1-\theta)^2} \Rightarrow$$

$$\Rightarrow A = (1-\theta)^2$$

$$E[X] = (1-\theta)^2 \cdot \sum_{h=0}^{\infty} h \cdot (h+1) \cdot \theta^h$$

$$\sum_{h=0}^{\infty} \theta^h = \frac{1}{1-\theta} \quad \Bigg| \quad \frac{d}{d\theta}$$

$$\sum_{h=0}^{\infty} h \cdot \theta^{h-1} = \frac{1}{(1-\theta)^2} \Rightarrow \sum_{h=0}^{\infty} h \cdot \theta^h = \frac{\theta}{(1-\theta)^2}$$

Kaj derivām o datā:

$$\sum_{h=0}^{\infty} h^2 \theta^{h-1} = \frac{(1-\theta)^2 + 2\theta(1-\theta)}{(1-\theta)^4}$$

$$= \frac{1 - 2\theta + \theta^2 + 2\theta - 2\theta^2}{(1-\theta)^4}$$

$$= \frac{1 - \theta^2}{(1-\theta)^4} = \frac{1 + \theta}{(1-\theta)^3}$$

$$\sum_{h=0}^{\infty} h^2 \cdot \theta^h = \frac{\theta + \theta^2}{(1-\theta)^3} \quad \left| \frac{d}{d\theta} \right.$$

$$\sum_{h=0}^{\infty} h^3 \cdot \theta^{h-1} = \frac{(1+2\theta) \cdot (1-\theta)^3 + 3(\theta + \theta^2) \cdot (1-\theta)^2}{(1-\theta)^6}$$

$$= \frac{1 + 2\theta - \theta - 2\theta^2 + 3\theta + 3\theta^2}{(1-\theta)^4}$$

$$= \frac{1 + \theta + \theta^2}{(1-\theta)^4}$$

$$\sum_{h=0}^{\infty} h^3 \cdot \theta^h = \frac{\theta + \theta^2 + \theta^3}{(1-\theta)^3}$$

$$E[X] = (1-\theta)^2 \cdot \left(\frac{\theta}{(1-\theta)^2} + \frac{\theta + \theta^2}{(1-\theta)^3} \right)$$

$$= \theta + \frac{\theta + \theta^2}{1-\theta}$$

$$= \frac{\theta - \theta^2 + \theta + \theta^2}{1-\theta} = \frac{2\theta}{1-\theta}$$

Analog variante

$$2. \quad E[X] = \frac{2\theta}{1-\theta} =$$

$$\Rightarrow 2\theta = (1-\theta) E[X]$$

$$\theta(2 + E[X]) = E[X]$$

$$\theta = \frac{E[X]}{2 + E[X]} \Rightarrow$$

$$\Rightarrow \hat{\theta} = \frac{\overline{X_n}}{2 + \overline{X_n}} \quad (\text{metoda momentelor})$$

$$P_{\theta}(\tilde{\theta} = 0) =$$

$$= P(\overline{X}_n = 0)$$

X_1, \dots, X_n can do only values positive \Rightarrow

$$\Rightarrow \overline{X}_n = 0 \Leftrightarrow X_1 = X_2 = \dots = X_n = 0$$

$$\begin{aligned} P_{\theta}(\tilde{\theta} = 0) &= (P(X_1 = 0))^n \\ &= ((1-\theta)^2 \cdot (0+1))^n \\ &= (1-\theta)^{2n} \end{aligned}$$

$$3. L(\theta; x_1, \dots, x_n) = (1-\theta)^{2n} \cdot \prod_{i=1}^n (x_i + 1) \cdot \theta^{\sum_{i=1}^n x_i}$$

$$\begin{aligned} \ell(\theta; x_1, \dots, x_n) &= 2n \ln(1-\theta) + \sum_{i=1}^n \ln(x_i + 1) + \\ &\quad + \left(\sum_{i=1}^n x_i \right) \ln \theta \end{aligned}$$

$$\frac{dl}{d\theta} = -\frac{2n}{1-\theta} + \left(\sum_{i=1}^n x_i\right) \cdot \frac{1}{\theta}$$

$$\frac{dl}{d\theta} = 0 \Leftrightarrow \frac{2n}{1-\theta} = \frac{1}{\theta} \cdot \sum_{i=1}^n x_i \Leftrightarrow$$

$$\Leftrightarrow 2n\theta = \sum x_i - \theta \cdot \sum x_i \Leftrightarrow \theta(2n + \sum x_i) = \sum x_i$$

$$\theta = \frac{\sum x_i}{2n + \sum x_i}$$

$$l \nearrow \text{pt. } \theta < \frac{\sum x_i}{2n + \sum x_i}$$

$$l \searrow \text{pt. } \theta > \frac{\sum x_i}{2n + \sum x_i}$$

$$\text{Deci } \hat{\theta} = \frac{\sum x_i}{2n + \sum x_i} = \frac{\overline{X_n}}{2 + \overline{X_n}}$$

$$\overline{X_n} \geq 0 \Rightarrow 2 + \overline{X_n} \geq 2 \Rightarrow \hat{\theta} \leq 1$$

bine definit $g: (0, \infty) \rightarrow (0, \infty)$

$$4) \quad \hat{\theta} = \frac{\overline{X_n}}{2 + \overline{X_n}} \quad g(x) = \frac{x}{2+x} \quad \text{continuă}$$

Dim T. apl. continuă:

$$\overline{X_n} \xrightarrow{1P} E[X_1] = \frac{2\theta}{1-\theta} \Rightarrow$$

$$\Rightarrow \tilde{\theta}_n = g(\bar{X}_n) \xrightarrow{IP} g(E[X]) = \\ = g\left(\frac{2\theta}{1-\theta}\right) = \frac{\frac{2\theta}{1-\theta}}{2 + \frac{2\theta}{1-\theta}} = \theta \Rightarrow$$

$\Rightarrow \tilde{\theta}_n$ e consistent

Dim T.L.C $\sqrt{n}(\bar{X}_n - \frac{2\theta}{1-\theta}) \xrightarrow{d} N(0, \text{Var}(X_1))$

\uparrow
 calculate
 anterior

Dim T. privind metoda Delta:

$$\sqrt{n}(g(\bar{X}_n) - g(\frac{2\theta}{1-\theta})) \xrightarrow{d} g'(\frac{2\theta}{1-\theta}) \cdot N(0, \text{Var}(X_1))$$

$$\sqrt{n}(\tilde{\theta}_n - \theta) \xrightarrow{d} N(0, (g'(\frac{2\theta}{1-\theta}))^2 \cdot \text{Var}(X_1))$$

$$g'(x) = \frac{2+x-x}{(2+x)^2} = \frac{2}{(2+x)^2}$$

$$g'(\frac{2\theta}{1-\theta}) = \dots$$

Ex 3 $N(\mu, 1)$

$$X_1, X_2, \dots, X_n \sim N(\mu, 1)$$

$$f_\mu(x) = \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{(x-\mu)^2}{2}}$$

$$\begin{aligned} \frac{d}{d\mu} \log f_\mu(x) &= \frac{d}{d\mu} \left(-\frac{(x-\mu)^2}{2} \right) \\ &= x - \mu \end{aligned}$$

$$\begin{aligned} I_1 &= E \left[\left(\frac{d}{d\mu} \log f_\mu(x) \right)^2 \right] = E \left[(x - \mu)^2 \right] \\ &= \text{Var}(X) = 1 \end{aligned}$$

$$I_n = n \Rightarrow \text{MIRC} = \frac{1}{n}$$

$$E[X_1] = \mu \Rightarrow \tilde{\mu} = \bar{X}_n \rightarrow \text{metoda momentelor}$$

$$\begin{aligned} \text{Var}(\tilde{\mu}) &= \text{Var} \bar{X}_n = \frac{1}{n^2} \cdot n \cdot \text{Var}(X_1) \\ &= \frac{1}{n} \end{aligned}$$

$$\text{Var}(\tilde{\mu}) = \text{MIRC}, \text{ deci } \tilde{\mu} \text{ e eficient}$$

$$\text{Ex 1} \quad B(\theta) \quad \theta \in (0, 1)$$

$$f_{\theta}(x) = P(X=x) = \theta^x \cdot (1-\theta)^{1-x}$$

$$L(\theta; x_1, \dots, x_n) = \theta^{\sum x_i} \cdot (1-\theta)^{n - \sum x_i}$$

$$l(\theta; x_1, \dots, x_n) = \left(\sum x_i\right) \ln \theta + (n - \sum x_i) \ln(1-\theta)$$

$$\frac{dl}{d\theta} = \left(\sum x_i\right) \cdot \frac{1}{\theta} - (n - \sum x_i) \cdot \frac{1}{1-\theta}$$

$$\frac{dl}{d\theta} = 0 \Leftrightarrow \frac{1-\theta}{\theta} = \frac{n - \sum x_i}{\sum x_i} \quad (=)$$

$$\Leftrightarrow \sum x_i - \theta \cdot \sum x_i = n\theta - \theta \sum x_i \quad \Leftrightarrow \theta = \frac{\sum x_i}{n}$$

$$\begin{array}{l} \text{pt. } \theta < \bar{x}_n \Rightarrow l \nearrow \\ \theta > \bar{x}_n \Rightarrow l \searrow \end{array} \quad \left| \Rightarrow \hat{\theta} = \bar{x}_n \right. \\ \text{(verosim. maxim.)}$$

$$\log f_{\theta} = x \cdot \ln \theta + (1-x) \ln(1-\theta)$$

$$\frac{d}{d\theta} \log f_{\theta} = \frac{x}{\theta} - \frac{1-x}{1-\theta}$$

$$\frac{d^2}{d\theta^2} \log f_{\theta} = -\frac{x}{\theta^2} - \frac{1-x}{(1-\theta)^2}$$

$$I_1 = - E_{\theta} \left[\frac{d^2}{d\theta^2} \log f_{\theta}(x) \right]$$

$$= - E_{\theta} \left[-\frac{X}{\theta^2} - \frac{1-X}{1-\theta^2} \right]$$

$$= \frac{1}{\theta^2} E_{\theta}[X] + \frac{1}{(1-\theta)^2} \cdot E[1-X]$$

$$= \frac{1}{\theta^2} \cdot \theta + \frac{1}{(1-\theta)^2} \cdot (1-\theta)$$

$$= \frac{1}{\theta} + \frac{1}{1-\theta} = \frac{1}{\theta(1-\theta)} =,$$

$$\Rightarrow I_n = \frac{n}{\theta(1-\theta)}$$

$$b) V_{\theta}[X_1] = \theta(1-\theta)$$

$$\hat{\theta} = \bar{X}_n \Rightarrow T_n \text{ est. de vrais. max. pt. } V_{\theta}[X_1]$$

$$T_n = \bar{X}_n (1 - \bar{X}_n)$$

$$\begin{aligned} \text{Notation: } E[T_n] &= E[\bar{X}_n] - E[\bar{X}_n^2] \\ &= \theta - \left(\text{Var}(\bar{X}_n) + (E[\bar{X}_n])^2 \right) \end{aligned}$$

$$\begin{aligned}\text{Var}(\bar{X}_n) &= \frac{1}{n^2} \text{Var}(X_1 + \dots + X_n) \\ &= \frac{n}{n^2} \text{Var}(X_1) = \frac{1}{n} \theta(1-\theta)\end{aligned}$$

$$\begin{aligned}\mathbb{E}[T_n] &= \theta - \left(\frac{1}{n} \theta(1-\theta) + \theta^2 \right) \\ &= \theta(1-\theta) - \frac{1}{n} \theta(1-\theta) = \frac{n-1}{n} \theta(1-\theta)\end{aligned}$$

Deci $\mathbb{E}[T_n] \neq \theta(1-\theta) \Rightarrow T_n$ nu e
nedegrasat.

Consistent:

$$\bar{X}_n \xrightarrow{\mathbb{P}} \theta$$

Fie $g: [0,1] \rightarrow \mathbb{R}$
 $g(\theta) = \theta(1-\theta)$ continuă

Din T. opl. continue $\Rightarrow g(\bar{X}_n) \xrightarrow{\mathbb{P}} g(\theta)$
 $T_n \xrightarrow{\mathbb{P}} \theta(1-\theta) =,$
 $\Rightarrow T_n$ e consistent