

Lecture 6

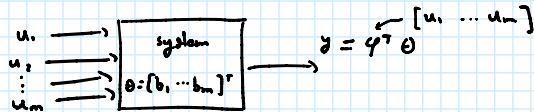
Monday, January 31, 2022 9:26 AM

Admin: HW1 due Feb 9th.

Question: about co-labs.

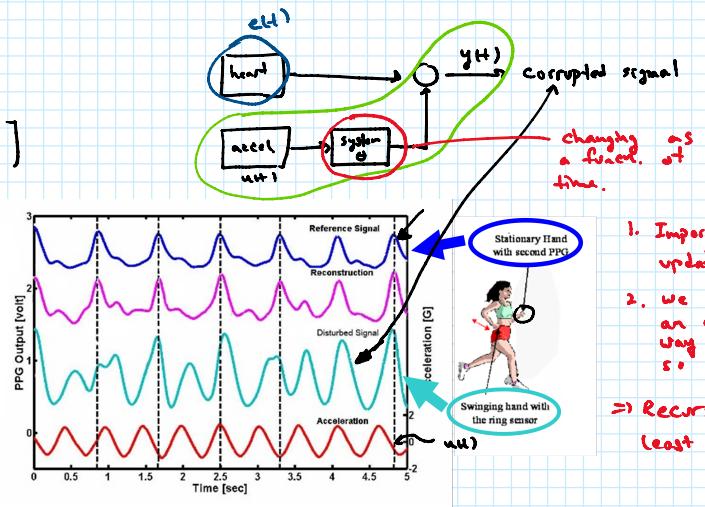
Recursive Least Squares Parameter Estimation

Last time: Least squares estimation



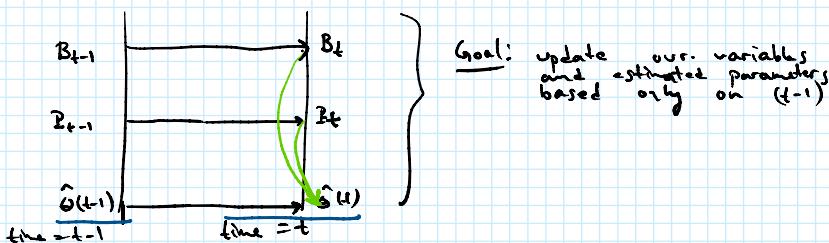
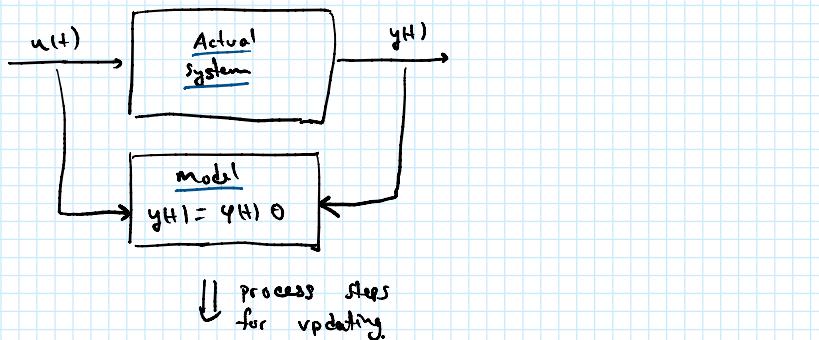
$$\begin{aligned} \text{- Find } \hat{\theta} \text{ by minimizing } V_N(\hat{\theta}) &= \frac{1}{N} \sum_{n=1}^N [\hat{y}(t_n|\hat{\theta}) - y(t_n)]^2 \\ \Rightarrow \hat{\theta} &= P B \text{ where } P = \left[\sum_{n=1}^N (p_n) p_n^T \right]^{-1} \text{ and } B = \left[\sum_{n=1}^N y_n p_n \right] \end{aligned}$$

auto-covariance



Recursive Least Squares Algorithm

- Least squares (above) is useful for batch processing where all data are already collected
=> no more to consider for estimating $\hat{\theta}$
- When we are estimating parameters in real-time (i.e., online updating $\hat{\theta}$ is required), a recursive computing method is preferred!
 - Reduce computational overhead => faster computation
 - No need to store all of the data. => higher frequency.



- Three steps for obtaining a recursive least squares (RLS)

algorithm.

- Step 1: Write B_t and I_t in terms of $B_{t-1} \notin I_{t-1}$

$$B_t = \sum_{i=1}^t y(i) \varphi(i) = \underbrace{\sum_{i=1}^{t-1} y(i) \varphi(i)}_{B_{t-1}} + y(t) \varphi(t)$$

$$\Rightarrow B_t = B_{t-1} + y(t) \varphi(t)$$

For P_t

$$I_t^{-1} = \sum_{i=1}^t [\varphi(i) \varphi^T(i)] = \underbrace{\sum_{i=1}^{t-1} [\varphi(i) \varphi^T(i)]}_{I_{t-1}^{-1}} + \varphi(t) \varphi^T(t)$$

$$\Rightarrow I_t^{-1} = I_{t-1}^{-1} + \varphi(t) \varphi^T(t)$$

→ Need to compute inverse of I_t^{-1}

- Step 2: Use Matrix Inversion Lemma to find $\underline{I_t}$

First, premultiply by I_t and postmultiply by I_{t-1}

$$\Rightarrow I_t \cancel{I_t^{-1}} I_{t-1} = I_t I_{t-1}^{-1} + I_t \varphi(t) \varphi^T(t) I_{t-1}$$

$$\Rightarrow I_{t-1} = I_t + \boxed{I_t \varphi(t) \varphi^T(t) I_{t-1}}$$

Next, postmultiply by $\varphi(t)$:

$$I_{t-1} \varphi(t) = I_t \cancel{\varphi(t)} + I_t \varphi(t) \varphi^T(t) I_{t-1} \varphi(t)$$
$$= I_t \varphi(t) [1 + \underbrace{\varphi^T(t) I_{t-1} \varphi(t)}_{(1 \times m)(m \times m)(m \times 1)}]$$

Solve for $I_t \varphi(t)$:

$$I_t \varphi(t) = \frac{I_{t-1} \varphi(t)}{1 + \varphi^T(t) I_{t-1} \varphi(t)}$$
$$\Rightarrow m \times 1$$
$$\Rightarrow (1 \times 1)$$

Finally, postmultiply by $\varphi^T(t) I_{t-1}$:

$$(I_t \varphi(t) \varphi^T(t) I_{t-1}) = \frac{I_{t-1} \varphi(t) \varphi^T(t) I_{t-1}}{1 + \varphi^T(t) I_{t-1} \varphi(t)}$$

$$I_{t-1} - I_t$$

update formula
for inverted
covariance matrix

$$\Rightarrow \text{so, } \boxed{I_t = I_{t-1} - \frac{I_{t-1} \varphi(t) \varphi^T(t) I_{t-1}}{1 + \varphi^T(t) I_{t-1} \varphi(t)}}$$

A matrix inversion is not required

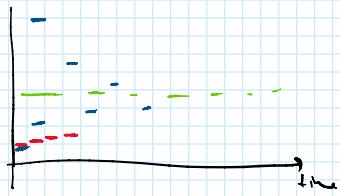
- Step 3 Reduce $\hat{\theta} = I_t B_t$ to the following recursive form:

$$\hat{\theta}(t) = \hat{\theta}(t-1) + K_t [y(t) - \underbrace{\varphi^T(t) \hat{\theta}(t-1)}_{\text{Error Correction Gain}}]$$

Error Correction
Gain Prediction using prior estimates
 Error

$$\hat{\theta}(t) = \mathbf{B}_t \mathbf{B}_{t-1}^{-1}; \quad \hat{\theta}(t-1) = \mathbf{B}_{t-1} \mathbf{B}_{t-1}^{-1} \quad \text{so,}$$

$$\hat{\theta}(t) - \hat{\theta}(t-1) = \underbrace{\mathbf{B}_t \mathbf{B}_{t-1}^{-1} - \mathbf{B}_{t-1} \mathbf{B}_{t-1}^{-1}}_{\text{we have update formulas for these}}$$



Algebra

$$\Rightarrow \hat{\theta}(t) - \hat{\theta}(t-1) = \underbrace{\frac{\mathbf{B}_{t-1} \mathbf{y}(t)}{1 + \mathbf{y}^T(t) \mathbf{B}_{t-1}^{-1} \mathbf{y}(t)}}_{K_t} \left[\mathbf{y}(t) - \mathbf{y}^T(t) \hat{\theta}(t-1) \right]$$

scalar

Now, we have all eqns needed for RLS:

$$\hat{\theta}(t) = \hat{\theta}(t-1) + \frac{\mathbf{B}_{t-1} \mathbf{y}(t)}{1 + \mathbf{y}^T(t) \mathbf{B}_{t-1}^{-1} \mathbf{y}(t)} \left[\mathbf{y}(t) - \mathbf{y}^T(t) \hat{\theta}(t-1) \right]$$

$$\mathbf{P}_t = \mathbf{P}_{t-1} - \frac{\mathbf{B}_{t-1} (\mathbf{y}(t) \mathbf{y}^T(t) \mathbf{B}_{t-1})}{1 + \mathbf{y}^T(t) \mathbf{B}_{t-1}^{-1} \mathbf{y}(t)}, \quad \text{for } t=1, 2, \dots$$

Assuming initial conditions:

$\hat{\theta}(0)$: arbitrary.

\mathbf{P}_0 : usually identity matrix [can be any positive definite matrix]

Initial Conditions for RLS

* How do ICS $\hat{\theta}(0)$ and \mathbf{P}_0 affect the quality of the parameter estimation?

- Even if ICS are sufficient, they may be suboptimal
- Comparing the "optimality" of different ICS is best done w/ a specific dataset, ^{it can theoretically show} that optimal $\hat{\theta}(0)$ can be derived from prelim. data.

Theorem: RLS algorithm minimizes the following cost funcn:

$$J_c(\theta) = \underbrace{\frac{1}{2} \sum_{i=1}^t [\mathbf{y}(i) - \mathbf{y}^T(i) \theta]^2}_{\text{Squared Estimation Error (A)}} + \underbrace{\frac{1}{2} (\theta - \hat{\theta}(0))^T \mathbf{P}_0^{-1} (\theta - \hat{\theta}(0))}_{\text{Weighted squared distance of parameters from } \hat{\theta}(0) \text{ (B)}}$$

Proof: See addendum

- 1) As time t gets larger more data obtained and term (A) gets much larger than (B) \Rightarrow Influence of ICS fades.
- 2) Early in time t , θ is pulled toward $\hat{\theta}(0)$ especially when the eigenvalues of \mathbf{P}_0^{-1} are large.
- 3) In contrast, if the eigenvalues of \mathbf{P}_0^{-1} are small, θ tends to respond more quickly to prediction error: $[\mathbf{y}(t) - \mathbf{y}^T(t) \theta]$
- 4) The initial matrix \mathbf{P}_0 represents the level of confidence in initial parameter value, $\hat{\theta}(0)$.

Proof: Differentiate $J_t(\Theta)$ and set to zero:

$$\frac{\partial J_t(\Theta)}{\partial \Theta} = -\sum_{i=1}^t (y(i) - \phi^T(i)\Theta) \phi(i) + P_0^{-1}(\Theta - \hat{\Theta}(0)) = 0 \quad (19)$$

- Collect terms: $\left[\sum_{i=1}^t \phi(i) \phi^T(i) + P_0^{-1} \right] \Theta = \sum_{i=1}^t y(i) \phi(i) + P_0^{-1} \hat{\Theta}(0)$

$$= P_t^{-1}$$

- Solving for Θ : $\hat{\Theta}(t) = P_t \left[\sum_{i=1}^t y(i) \phi(i) + P_0^{-1} \hat{\Theta}(0) \right]$

$$= P_t \left[y(t) \phi(t) + \sum_{i=1}^{t-1} y(i) \phi(i) + P_0^{-1} \hat{\Theta}(0) \right]$$

Current data Previous data

\Rightarrow The term for the previous data can be written as:

$$\sum_{i=1}^{t-1} y(i) \phi(i) + P_0^{-1} \hat{\Theta}(0) = \sum_{i=1}^{t-1} y(i) \phi(i) + B_0 \quad (20)$$

$B_0 \xrightarrow{\text{from}} \hat{\Theta}(0) = P_0 B_0$

Since we're using premim data, B_0 is the same as B_{t-1} , and $B_{t-1} = P_{t-1}^{-1} \hat{\Theta}(t-1)$

- Now recall that $P_t^{-1} = \phi(t) \phi^T(t) + P_{t-1}^{-1}$. We can write:

$$\begin{aligned} \hat{\Theta} &= P_t \left[P_t^{-1} \hat{\Theta}(t-1) + y(t) \phi(t) - \phi(t) \phi^T(t) \hat{\Theta}(t-1) \right] \\ &= \hat{\Theta}(t-1) + P_t \phi(t) \left[y(t) - \phi^T(t) \hat{\Theta}(t-1) \right] \end{aligned} \quad (21)$$

- If we postmultiply $\phi(t)$ to both sides of (19) in Lecture 4 we get the expression for $P_t \phi(t)$ from prior data

$$\begin{aligned} P_t \phi(t) &= P_{t-1} \phi(t) - \frac{P_{t-1} \phi(t) \phi^T(t) P_{t-1} \phi(t)}{(1 + \phi^T(t) P_{t-1} \phi(t))} \quad (22) \\ &= \frac{P_{t-1} \phi(t)}{(1 + \phi^T(t) P_{t-1} \phi(t))} \end{aligned}$$

- Plugging this into (21) we get:

$$\hat{\Theta}(t) = \hat{\Theta}(t-1) + \frac{P_{t-1} \phi(t)}{(1 + \phi^T(t) P_{t-1} \phi(t))} \left[y(t) - \phi^T(t) \hat{\Theta}(t-1) \right] \quad (23)$$