

Lecture 20

Wednesday, April 6, 2022 10:03 AM

Informative Datasets and Persistence of Excitation - Part I

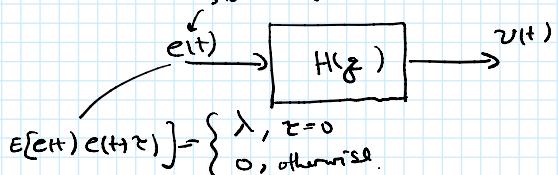
- Reading : Wood et al Asada 2007

Last time:

- Considerations of how to tell if a dataset is good enough.
- How accurate is the model?
- Power spectral analysis

Theorem

Let $H(\varphi)$ be a transfer function of BIBO process with a white noise input $e(t)$ w/ variance λ . stationary.



$$E[e(t)e(t+\tau)] = \begin{cases} \lambda, & \tau=0 \\ 0, & \text{otherwise.} \end{cases}$$

bounded input
bounded output

$$v(t) = H(\varphi) e(t).$$

The power spectral density of $v(t)$ is given by:

$$\overline{\omega}_v(\omega) = \lambda |H(e^{i\omega})|^2$$

variance



- For a stationary signal, $\{w(t)\}$ with spectrum $\overline{\omega}_w(\omega)$

$$\overline{w(t)} \xrightarrow{G(\varphi)} s(t) \Rightarrow s(t) = G(\varphi) w(t)$$

\Rightarrow The power spectra are given by:

$$\overline{\omega}_s(\omega) = \underbrace{|G(e^{i\omega})|}_{ss}^2 \overline{\omega}_w(\omega) \xleftarrow{\text{Fourier}} R_{ss}(\tau) = E[s(t)s(t+\tau)]$$

$$\overline{\omega}_{sw}(\omega) = G(e^{i\omega}) \overline{\omega}_w(\omega) \xleftarrow{\text{Fourier}} R_{sw}(\tau) = E[s(t)w(t+\tau)]$$

$$\text{Can find } G(e^{i\omega}): \quad G(e^{i\omega}) = \frac{\overline{\omega}_{sw}(\omega)}{\overline{\omega}_{sw}(\omega)}$$

$$|G(e^{i\omega})|$$



- Applying spectral analysis to system identification / parameter estimation!

$$y(t) = G(\varphi) w(t) + H(\varphi) e(t)$$

$y(t)$
noise

$$y(t) = G(\omega)u(t) + \underbrace{H(\omega)e(t)}_{v(t)}$$

noise.

Two issues to address

1) strictly speaking, the process is not stationary: the input $u(t)$ drives the system.

$$\begin{aligned} R_u(\tau) &= \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{t=1}^N u(t)u(t+\tau) \\ &\stackrel{?}{=} E[u(t)u(t+\tau)] \end{aligned}$$

properties
single changing
w.r.t. time.

In real system, $R_u(\tau)$ cannot be defined in general.

→ But we need covariance to exist.

=) Extend the definition for existence of $R_u(\tau)$

Extension: quasi-stationary (wide-sense stationary)

$$\Rightarrow R_u(\tau) = E[u(t)u(t+\tau)] \approx \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{t=1}^{\infty} u(t)u(t+\tau)$$

enables use of ergodicity.

- under this assumption, signal generated through stable filters $G(\omega)$

$$s(t) = G(\omega) e(t)$$

are also considered to be ergodic.

2) Informative datasets.

True system: $y(t) = G(\omega)u(t) + H(\omega)e(t)$

$$\begin{aligned} \text{predictor: } \hat{y}(t|t-1) &= \underbrace{H(\omega)G(\omega)u(t)}_{W_1(\omega)} + \underbrace{[I - H(\omega)G(\omega)]y(t)}_{W_2(\omega)} \\ &= \underbrace{[W_1(\omega) \quad W_2(\omega)]}_{W(\omega)} \begin{bmatrix} u(t) \\ y(t) \end{bmatrix} = W(\omega)z(t) \end{aligned}$$

Definition 1: Two models $W_1(\omega)$ and $W_2(\omega)$ are equal if frequency func.

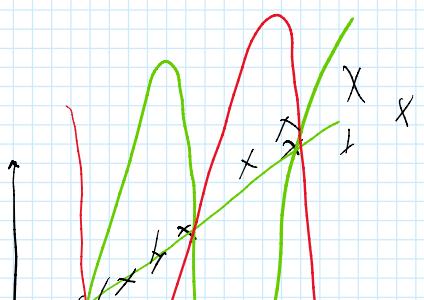
$$W_1(e^{i\omega}) = W_2(e^{i\omega}) \text{ for almost all } \omega$$

Definition 2: A quasi-stationary dataset z^∞ is informative with respect to the model structure, M , if for any two models in M :

$$\hat{y}_1(t, \theta_1) = W_1(\omega)z(t) \quad \hat{y}_2(t, \theta_2) = W_2(\omega)z(t)$$

then the condition

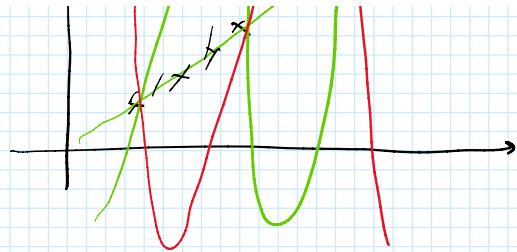
$$E[\hat{y}_1(t, \theta_1) - \hat{y}_2(t, \theta_2)]^2 = 0$$



$$E[(\hat{y}_1(t|\theta_1) - \hat{y}_2(t, \theta_2))^2] = 0$$

implies

$$w_1(e^{i\omega}) = w_2(e^{i\omega}) \text{ for almost all } \omega.$$



→ let's characterize the data:

$$\Phi_{zz}(\omega) = \begin{bmatrix} Euu(\omega) & Euy(\omega) \\ Eyu(\omega) & Eyy(\omega) \end{bmatrix} \in \mathbb{R}^{2 \times 2}$$

Theorem: A quasi-stationary dataset z^0 is informative if the spectrum matrix $\Phi_{zz}(\omega)$ is strictly positive definite for almost all ω .

proof → Insert

positive.

Proof:

$$\hat{y}_1(t|\theta_1) - \hat{y}_2(t, \theta_2) = [w_1(\theta) - w_2(\theta)] \cdot z(t) \sim [u(t) \quad y(t)]^\top$$

$$\begin{aligned} E[(\hat{y}_1(t|\theta_1) - \hat{y}_2(t, \theta_2))^2] &= E\left\{\underbrace{[(w_1 - w_2) z(t)]^\top}_{\sim [u(t) \quad y(t)]^\top}\right\}^2 \\ &= \frac{1}{2\pi} \int_{-\pi}^{\pi} [w_1(e^{i\omega}) - w_2(e^{i\omega})]^\top \underbrace{\Phi_{zz}(\omega)}_{\neq 0 \text{ for any } \omega} [w_1(e^{i\omega}) - w_2(e^{i\omega})] d\omega \end{aligned}$$

Since $\Phi_{zz}(\omega)$ is strictly positive definite for almost all ω , the above integral becomes zero only when $w_1 - w_2$ is zero for almost all ω .

$$\Rightarrow w_1(e^{i\omega}) = w_2(e^{i\omega}).$$

- consistency of prediction error estimate:

$$\hat{\theta} \rightarrow \theta_0 \text{ as } N \rightarrow \infty$$

\uparrow
true parameters

- For linear the invariant model structures based on minimizing squared error
 \Rightarrow estimates are consistent.

→ Liung Chap 8.

Next time: persistence of excitation.