CS258: Information Theory

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Outline

- ☐ Differential Entropy
- AEP for Continuous Random Variable
- Relative Entropy and Mutual Information
- Property of Differential Information Measures
- Information inequalities and applications

I(X;Y): Correlated Gaussian

(Mutual information between correlated Gaussian random variables with correlation ρ) Let $(X,Y) \sim \mathcal{N}(0,K)$, where

$$K = \begin{bmatrix} \sigma^2 & \rho \sigma^2 \\ \rho \sigma^2 & \sigma^2 \end{bmatrix}$$

I(X;Y)?

$$h(X) = h(Y) = \frac{1}{2} \log 2\pi e \sigma^{2}$$

$$h(X,Y) = \frac{1}{2} \log(2\pi e)^{2} |K| = \frac{1}{2} \log(2\pi e)^{2} \sigma^{4} (1 - \rho^{2})$$

$$I(X;Y) = h(X) + h(Y) - h(X,Y) = -\frac{1}{2} \log(1 - \rho^{2})$$

- ho = 0, X and Y are independent and I is 0
- lacksquare $ho=\pm1$, X and Y are perfectly correlated and I is ∞

Maximum Entropy with Constraints

■ Let the random variable $X \in R$ have mean μ and variance σ^2 . Then

$$h(X) \leq \frac{1}{2} \log 2\pi e \sigma^2$$

with equality iff $X \sim \mathcal{N}(\mu, \sigma^2)$

■ Let the random variable $X \in R$ satisfy $EX^2 \le \sigma^2$. Then

$$h(X) \leq \frac{1}{2} \log 2\pi e \sigma^2$$

with equality iff $X \sim \mathcal{N}(0, \sigma^2)$

1. Let $X_G \sim \mathcal{N}(\mu, \sigma^2)$. Consider

$$D(X||X_G) \geq 0$$

Then

$$\int f \log \frac{f}{g} \ge 0$$

$$h(X) = h(f) \le -\int f \log g = -\int f \log \frac{1}{\sqrt{2\pi\sigma^2}} + f\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$
$$h(X) \le \frac{1}{2}\log 2\pi\sigma^2 + \frac{1}{2} = \frac{1}{2}\log 2\pi e\sigma^2$$

2. $Var(X) = E(X^2) - E(X)^2 \le \sigma^2$. \Rightarrow Case 1.

 $E(X^2)$, Var(X) 给定的情况下, 高斯分布最大化微分熵

Maximum Entropy

Consider the following problem: Maximize the entropy h(f) over all probability densities f satisfying

- 1. $f(x) \ge 0$, with equality outside the support
- $2. \int_{S} f(x)dx = 1 \tag{++}$
- 3. $\int_{S} f(x)r_i(x)dx = \alpha_i$ for $1 \le i \le m$. $(r_i(x))$ is a function of x)

Thus, f is a density on support set S meeting certain moment constraints $\alpha_1, \alpha_2, \ldots, \alpha_m$.

Theorem 12.1.1 (Maximum entropy distribution) Let

$$f^*(x) = f_{\lambda}(x) = e^{\lambda_0 + \sum_{i=1}^m \lambda_i r_i(x)}$$

 $x \in S$, where $\lambda_0, \ldots, \lambda_m$ are chosen so that f^* satisfies (++). Then f^* uniquely maximizes h(f) over all probability densities f satisfying constraints (++).

- Let S = [a, b], with no other constraints. Then the maximum entropy distribution is the uniform distribution over this range.
- $S = [0, \infty)$ and $EX = \mu$. Then the entropy-maximizing distribution is

$$f(x) = \frac{1}{\mu} e^{-\frac{x}{\mu}}, \qquad x \ge 0$$

■ $S = (-\infty, \infty)$, $EX = \alpha_1$, and $EX^2 = \alpha_2$. The maximum entropy distribution is $\mathcal{N}(\alpha_1, \alpha_2 - \alpha_1^2)$

Hadamard's Inequality

K is a nonnegative definite symmetric $n \times n$ matrix. Let |K| denote the determinant of K.

Theorem (Hadamard) $|K| \leq \prod K_{ii}$, with equality iff $K_{ij} = 0$, $i \neq j$

Let
$$X \sim \mathcal{N}(0,K)$$
. Then
$$\frac{1}{2}\log(2\pi e)^n|K| = \mathbf{h}(X_1,X_2,...,X_n) \leq \sum \mathbf{h}(X_i) = \sum_{i=1}^n \frac{1}{2}\log 2\pi e|K_{ii}|$$
 with equality iff $X_1,X_2,...,X_n$ are independent (i.e., $K_{ij}=0, i\neq j$)

- \blacksquare $\log |K|$ is concave
- $log(|K_n|/|K_{n-p}|)$ is concave in K_n
- $\blacksquare |K_n|/|K_{n-1}|$ is concave in K_n

- A general technique to deal with nonnegative definite symmetric matrix K
- Ref. Ch. 17.9, 17.10, Cover

Balanced Information Inequality

Differences between inequalities of the discrete entropy and differential entropy

- Both $H(X,Y) \le H(X) + H(Y)$ and $h(X,Y) \le h(X) + h(Y)$ are valid
- $H(X,Y) \ge H(X)$ but neither $h(X,Y) \ge h(X)$ nor $h(X,Y) \le h(X)$ is valid

Take
$$H(X, Y, Z) \le \frac{1}{4}H(X) + \frac{1}{2}H(Y, Z) + \frac{3}{4}H(Z, X)$$
 for example.

Count the weights of random variables X, Y, Z in both sides

$$X: 1, 1; Y: 1, \frac{1}{2}; Z: 1, \frac{5}{4}$$

The net weights of X, Y, Z are $0, \frac{1}{2}, -\frac{1}{4}$

Balanced: If the net weights of X, Y, Z are all zero.

$$h(X,Y) \le h(X) + h(Y)$$
 and $h(X,Y,Z) \le \frac{1}{2}h(X,Y) + \frac{1}{2}h(Y,Z) + \frac{1}{2}h(Z,X)$

Let $[n]:=\{1,2,\ldots,n\}$. For any $\alpha\subseteq [n]$, denote $(X_i:i\in\alpha)$ by X_α . For example, $\alpha=\{1,3,4\}$, we denote X_1,X_3,X_4 by $X_{\{1,3,4\}}$ for simplicity.

- We could write any information inequality in the form $\sum_{\alpha} w_{\alpha} H(X_{\alpha}) \geq 0$ or $\sum_{\alpha} w_{\alpha} h(X_{\alpha}) \geq 0$.
- \blacksquare An information inequality is called balanced if for any $i \in [n]$, the net weight of X_i is zero.
- The linear continuous inequality $\sum_{\alpha} w_{\alpha} h(X_{\alpha}) \geq 0$ is valid if and only if its corresponding discrete counterpart $\sum_{\alpha} w_{\alpha} H(X_{\alpha}) \geq 0$ is valid and balanced.

Ref: Balanced Information Inequalities, T. H. Chan, IEEE Transactions on Information Theory, Vol. 49, No. 12, December 2003

Han's Inequality

■ Let $(X_1, X_2, ..., X_n)$ have a density, and for **every S** \subseteq {1, 2, ..., n}, denoted by X(S) the subset $\{X_i : i \in S\}$. Let

$$h_k^{(n)} = \frac{1}{\binom{n}{k}} \sum_{S:|S|=k} \frac{h(X(S))}{k}$$
$$g_k^{(n)} = \frac{1}{\binom{n}{k}} \sum_{S:|S|=k} \frac{h(X(S)|X(S^c))}{k}$$

When n = 3, $h_1^{(3)} = \frac{H(X_1) + H(X_2) + H(X_3)}{3} \ge h_2^{(3)} = \frac{H(X_1, X_2) + H(X_2, X_3) + H(X_3, X_1)}{6}$ $\ge h_3^{(3)} = H(X_1, X_2, X_3)/3$ $g_1^{(3)} = \frac{H(X_1|X_2, X_3) + H(X_2|X_1, X_3) + H(X_3|X_1, X_2)}{3}$ $\le g_2^{(3)} = \frac{H(X_1, X_2|X_3) + H(X_2, X_3|X_1) + H(X_3, X_1|X_2)}{6}$ $\le g_3^{(3)} = H(X_1, X_2, X_3)/3$

Han's inequality:

$$h_1^{(n)} \ge h_2^{(n)} \dots \ge h_n^{(n)} = H(X_1, X_2, \dots, X_n) / n = g_n^{(n)} \ge \dots \ge g_2^{(n)} \ge g_1^{(n)}$$

Information of Heat







 \blacksquare Heat equation (Fourier): Let x be the position and t be the time,

$$\frac{\partial}{\partial t}f(x,t) = \frac{1}{2}\frac{\partial^2}{\partial x^2}f(x,t)$$

Let X be any random variable with a density f(x). Let Z be an independent normally distributed random variable with zero mean and unit variance, $Z \sim \mathcal{N}(0,1)$. Let

$$Y_t = X + \sqrt{t}Z$$

The probability density function f(y;t) (f(y;t) is a function in y, not t) of Y_t satisfies heat equation

$$f(y;t) = \int f(x) \frac{1}{\sqrt{2\pi t}} e^{-\frac{(y-x)^2}{2t}} dx$$

Gaussian channel ↔ Heat equation

Entropy and Fisher Information

Fisher information: Let X be any random variable with density f(x). Its Fisher information is given by

$$I(X) = \int_{-\infty}^{+\infty} f(x) \left[\frac{\partial}{\partial x} f(x) \right]^{2} dx$$

Let X be any random variable with a density f(x). Let Z be an independent normally distributed random variable with zero mean and unit variance. Let $Y_t = X + \sqrt{t}Z$

$$\frac{\partial}{\partial t}h(Y_t) = \frac{1}{2}I(Y_t)$$

■ Let f(y,t) (or f) be the p.d.f of Y_t

$$\frac{\partial}{\partial t}h(Y_t) = \frac{1}{2}I(Y_t) = \frac{1}{2}\int \frac{f_y^2}{f}dy \ge 0$$

$$\frac{\partial^2}{\partial t^2}h(Y_t) = -\frac{1}{2} \int f\left(\frac{f_{yy}}{f} - \frac{f_y^2}{f^2}\right)^2 dy \le 0$$

■ When X is Gaussian $\mathcal{N}(0,1)$,

$$h(Y_t) = \frac{1}{2}\log 2\pi e(1+t)$$

All the derivatives alternate in signs: +, -, +, -, ...

Higher Order Derivatives of $h(Y_t)$

(Cheng 2015) Let X be any random variable with a density f(x). Let Z be an independent normally distributed random variable with zero mean and unit variance. Let $Y_t = X + \sqrt{t}Z$ and f(y,t) (or f) be the p.d.f of Y_t . Then

$$\frac{\partial^3}{\partial t^3}h(Y_t) \ge 0$$
 and $\frac{\partial^4}{\partial t^4}h(Y_t) \le 0$

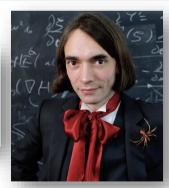
Conjecture: When n is even, $\frac{\partial^n}{\partial t^n}h(Y_t) \leq 0$, otherwise $\frac{\partial^n}{\partial t^n}h(Y_t) \geq 0$



A review of mathematical topics in collisional kinetic theory

Cédric Villani

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C. Villani 2010 Fields medalist $\frac{\partial^{3}}{\partial t^{3}}h(Y_{t}) = \frac{1}{2} \int f\left(\frac{f_{3}}{f} - \frac{f_{1}f_{2}}{f^{2}} + \frac{1}{3}\frac{f_{1}^{3}}{f^{3}}\right)^{2} + \frac{f_{1}^{6}}{45f^{5}}dy$ $\frac{\partial^{4}}{\partial t^{4}}h(Y_{t})$ $= -\frac{1}{2} \int f\left(\frac{f_{4}}{f} - \frac{6}{5}\frac{f_{1}f_{3}}{f^{2}} - \frac{7}{10}\frac{f_{2}^{2}}{f^{2}} + \frac{8}{5}\frac{f_{1}^{2}f_{2}}{f^{3}} - \frac{1}{2}\frac{f_{1}^{4}}{f^{4}}\right)^{2}$ $+ f\left(\frac{2}{5}\frac{f_{1}f_{3}}{f^{2}} - \frac{1}{3}\frac{f_{1}^{2}f_{2}}{f^{3}} + \frac{9}{100}\frac{f_{1}^{4}}{f^{4}}\right)^{2}$ $+ f\left(-\frac{4}{100}\frac{f_{1}^{2}f_{2}}{f^{3}} + \frac{4}{100}\frac{f_{1}^{4}}{f^{4}}\right)^{2}$ $+ \frac{1}{300}\frac{f_{2}^{4}}{f^{3}} + \frac{56}{90000}\frac{f_{1}^{4}f_{2}^{2}}{f^{5}} + \frac{13}{70000}\frac{f_{1}^{8}}{f^{7}}dy$

"This suggests that....., etc., but I could not prove it" (1966)

H. P. McKean

Ref: F. Cheng and Y. Geng, "Higher Order Derivatives in Costa's Entropy Power Inequality"

EPI and **FII**

(Shannon 1948, Entropy power inequality (EPI)) If X and Y are independent random n-vectors with densities, then

$$e^{\frac{2}{n}h(X+Y)} \ge e^{\frac{2}{n}h(X)} + e^{\frac{2}{n}h(Y)}$$

$$e^{2h(X+Y)} \ge e^{2h(X)} + e^{2h(Y)} (n = 1)$$

Fisher information inequality (FII)

$$\frac{1}{I(X+Y)} \ge \frac{1}{I(X)} + \frac{1}{I(Y)}$$

- Most profound result in Shannon's 1948 paper
- EPI can imply some very fundamental results
 - Uncertainty principle
 - Young's inequality
 - Nash's inequality
 - Cramer-Rao bound

Reference

- T. Cover, "Information theoretic inequalities," 1990
- O. Rioul, "Information Theoretic Proofs of Entropy Power Inequalities," 2011

Summary

Cover: 8.9, 12.1, 17.6, 17.7, 17.9, 17.10