Introduction to Machine Learning

Bayesian Regression

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1 Linear Regression

1.1 Problem Formulation

- There is one scalar **target** variable y (instead of hidden)
- There is one vector **input** variable x
- Inductive bias:

$$y = \mathbf{w}^{\top} \mathbf{x}$$

Linear Regression Learning Task

Learn w given training examples, $\langle \mathbf{X}, \mathbf{y} \rangle$.

The training data is denoted as $\langle \mathbf{X}, \mathbf{y} \rangle$, where \mathbf{X} is a $N \times D$ data matrix consisting of N data examples such that each data example is a D dimensional vector. \mathbf{y} is a $N \times 1$ vector consisting of corresponding target values for the examples in \mathbf{X} .

 \bullet y is assumed to be normally distributed

$$y \sim \mathcal{N}(\mathbf{w}^{\top}\mathbf{x}, \sigma^2)$$

• or, equivalently:

$$y = \mathbf{w}^{\top} \mathbf{x} + \epsilon$$

where $\epsilon \sim \mathcal{N}(0, \sigma^2)$

- \bullet y is a linear combination of the input variables
- Given ${\bf w}$ and σ^2 , one can find the probability distribution of y for a given ${\bf x}$

1.2 Learning Parameters

• Find w and σ^2 that maximize the likelihood of training data

$$\begin{aligned} \widehat{\mathbf{w}}_{MLE} &= & (\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{y} \\ \widehat{\sigma}_{MLE}^2 &= & \frac{1}{N}(\mathbf{y} - \mathbf{X}\mathbf{w})^{\top}(\mathbf{y} - \mathbf{X}\mathbf{w}) \end{aligned}$$

The derivation of the MLE estimates can be done by maximizing the loglikelihood of the data set. The likelihood of the training data set is given by:

$$L(\mathbf{w}) = \prod_{i=1}^{N} \frac{1}{\sqrt{2\pi}\sigma} exp(-\frac{(y_i - \mathbf{w}^{\top} \mathbf{x_i})^2}{2\sigma^2})$$

The log-likelihood is given by:

$$LL(\mathbf{w}) = -\frac{1}{2}\log 2\pi - \log \sigma - \frac{1}{2\sigma^2}\sum_{i=1}^{N}(y_i - \mathbf{w}^{\top}\mathbf{x}_i)^2$$

This can be rewritten in matrix notation as:

$$LL(\mathbf{w}) = -\frac{1}{2}\log 2\pi - \log \sigma - \frac{1}{2\sigma^2}(\mathbf{y} - \mathbf{X}\mathbf{w})^{\top}(\mathbf{y} - \mathbf{X}\mathbf{w})$$

To maximize the log-likelihood, we first compute its derivative with respect to ${\bf w}$ and σ .

$$\frac{\partial LL(\mathbf{w})}{\partial \mathbf{w}} = -\frac{1}{2\sigma^2} \frac{\partial}{\partial \mathbf{w}} (\mathbf{y} - \mathbf{X}\mathbf{w})^{\top} (\mathbf{y} - \mathbf{X}\mathbf{w})$$

$$= -\frac{1}{2\sigma^2} \frac{\partial}{\partial \mathbf{w}} (\mathbf{y}^{\top} \mathbf{y} + \mathbf{w}^{\top} \mathbf{X}^{\top} \mathbf{X} \mathbf{w} - 2\mathbf{y}^{\top} \mathbf{X} \mathbf{w})$$

Note that, we use the fact that $(\mathbf{X}\mathbf{w})^{\mathsf{T}}\mathbf{y} = \mathbf{y}^{\mathsf{T}}\mathbf{X}\mathbf{w}$, since both quantities are scalars and the transpose of a scalar is equal to itself. Continuing with the derivative:

$$\frac{\partial LL(\mathbf{w})}{\partial \mathbf{w}} = -\frac{1}{2\sigma^2} (2\mathbf{w}^{\top} \mathbf{X}^{\top} \mathbf{X} - 2\mathbf{y}^{\top} \mathbf{X})$$

Setting the derivative to 0, we get:

$$\begin{aligned} 2\mathbf{w}^{\top}\mathbf{X}^{\top}\mathbf{X} - 2\mathbf{y}^{\top}\mathbf{X} &= 0 \\ \mathbf{w}^{\top}\mathbf{X}^{\top}\mathbf{X} &= \mathbf{y}^{\top}\mathbf{X} \\ (\mathbf{X}^{\top}\mathbf{X})^{\top}\mathbf{w} &= \mathbf{X}^{\top}\mathbf{y} \text{ (Taking transpose both sides)} \\ (\mathbf{X}^{\top}\mathbf{X})\mathbf{w} &= \mathbf{X}^{\top}\mathbf{y} \\ \mathbf{w} &= (\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{y} \end{aligned}$$

In a similar fashion, one can set the derivative to 0 with respect to σ and plug in the the optimal value of \mathbf{w}

1.3 Issues with Linear Regression

- 1. Not truly Bayesian
- 2. Susceptible to outliers
- 3. Too simplistic Underfitting
- 4. No way to control overfitting
- 5. Unstable in presence of correlated input attributes
- 6. Gets "confused" by unnecessary attributes

2 Bayesian Linear Regression

3 Bayesian Regression

- "Penalize" large values of w
- A zero-mean Gaussian prior

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w}|0, \tau^2 I)$$

• What is posterior of w

$$p(\mathbf{w}|\mathcal{D}) \propto \prod_{i} \mathcal{N}(y_i|\mathbf{w}^{\top}\mathbf{x}_i, \sigma^2)p(\mathbf{w})$$

- Posterior is also Gaussian
- Regularized least squares estimate of w

$$\underset{\mathbf{w}}{\arg\max} \sum_{i=1}^{N} log \mathcal{N}(y_i | \mathbf{w}^{\top} \mathbf{x}_i, \sigma^2) + \log \mathcal{N}(\mathbf{w} | 0, \tau^2 I)$$

3.1 Estimating Bayesian Regression Parameters

• Prior for w

$$\mathbf{w} \sim \mathcal{N}(\mathbf{w}|\mathbf{0}, \tau^2 \mathbf{I}_D)$$

• Posterior for w

$$p(\mathbf{w}|\mathbf{y}, \mathbf{X}) = \frac{p(\mathbf{y}|\mathbf{X}, \mathbf{w})p(\mathbf{w})}{p(\mathbf{y}|\mathbf{X})}$$
$$= \mathcal{N}(\bar{\mathbf{w}} = (\mathbf{X}^{\top}\mathbf{X} + \frac{\sigma^2}{\tau^2}\mathbf{I}_N)^{-1}\mathbf{X}^{\top}\mathbf{y}, \sigma^2(\mathbf{X}^{\top}\mathbf{X} + \frac{\sigma^2}{\tau^2}\mathbf{I}_N)^{-1})$$

- \bullet Posterior distribution for \mathbf{w} is also Gaussian
- What will be MAP estimate for **w**?

The denominator term in the posterior above can be computed as the marginal likelihood of data by marginalizing \mathbf{w} :

$$p(\mathbf{y}|\mathbf{X}) = \int p(\mathbf{y}|\mathbf{X}, \mathbf{w}) p(\mathbf{w}) d\mathbf{w}$$

One can compute the posterior for \mathbf{w} as follows. We first show that the likelihood of \mathbf{y} , i.e., all target values in the training data, can be jointly modeled as a Gaussian as follows:

$$p(\mathbf{y}|\mathbf{X}, \mathbf{w}) = \prod_{i=1}^{N} \frac{1}{\sqrt{2\pi\sigma^2}} exp\left(-\frac{1}{2\sigma^2} (y_i - \mathbf{w}^{\top} \mathbf{x}_i)^2\right)$$
$$= \frac{1}{(2\pi\sigma^2)^{N/2}} exp\left(-\frac{1}{2\sigma^2} |\mathbf{y} - \mathbf{X} \mathbf{w}|^2\right)$$
$$= \mathcal{N}(\mathbf{X} \mathbf{w}, \sigma^2 \mathbf{I}_N)$$

Ignoring the denominator which does not depend on w:

$$p(\mathbf{w}|\mathbf{y}, \mathbf{X}) \propto exp(-\frac{1}{2\sigma^2}(\mathbf{y} - \mathbf{X}\mathbf{w})^{\top}(\mathbf{y} - \mathbf{X}\mathbf{w}))exp(-\frac{1}{2\tau^2}\mathbf{w}^{\top}\mathbf{w})$$
$$\propto exp(-\frac{1}{2}(\mathbf{w} - \bar{\mathbf{w}})^{\top}(\frac{1}{\sigma^2}\mathbf{X}^{\top}\mathbf{X} + \frac{1}{\tau^2}\mathbf{I}_N)(\mathbf{w} - \bar{\mathbf{w}}))$$

where
$$\bar{\mathbf{w}} = (\mathbf{X}^{\top}\mathbf{X} + \frac{\sigma^2}{\tau^2}\mathbf{I}_N)^{-1}\mathbf{X}\mathbf{y}$$
.

3.2 Prediction with Bayesian Regression

- For a new \mathbf{x}^* , predict y^*
- Point estimate of y*

$$y^* = \widehat{\mathbf{w}}_{MLE}^{\top} \mathbf{x}^*$$

 \bullet Treating y as a Gaussian random variable

$$p(y^*|\mathbf{x}^*) = \mathcal{N}(\widehat{\mathbf{w}}_{MLE}^{\top} \mathbf{x}^*, \widehat{\sigma}_{MLE}^2)$$

$$p(y^*|\mathbf{x}^*) = \mathcal{N}(\widehat{\mathbf{w}}_{MAP}^{\top} \mathbf{x}^*, \widehat{\sigma}_{MAP}^2)$$

• Treating y and \mathbf{w} as random variables

$$p(y^*|\mathbf{x}^*) = \int p(y^*|\mathbf{x}^*, \mathbf{w}) p(\mathbf{w}|\mathbf{X}, \mathbf{y}) d\mathbf{w}$$

• This is also Gaussian!

4 Handling Outliers in Regression

- Linear regression training gets impacted by the presence of outliers
- The square term in the exponent of the Gaussian pdf is the culprit
 - Equivalent to the square term in the loss
- How to handle this (Robust Regression)?
- Probabilistic:

- Use a different distribution instead of Gaussian for $p(y|\mathbf{x})$
- Robust regression uses Laplace distribution

$$p(y|\mathbf{x}) \sim Laplace(\mathbf{w}^{\top}\mathbf{x}, b)$$

- Geometric:
 - Least absolute deviations instead of least squares

$$J(\mathbf{w}) = \sum_{i=1}^{N} |y_i - \mathbf{w}^{\top} \mathbf{x}|$$

5 Generative vs. Discriminative Classifiers

• Probabilistic classification task:

$$p(Y = benign | \mathbf{X} = \mathbf{x}), p(Y = malicious | \mathbf{X} = \mathbf{x})$$

• How do you estimate $p(y|\mathbf{x})$?

$$p(y|\mathbf{x}) = \frac{p(y,\mathbf{x})}{p(\mathbf{x})} = \frac{p(\mathbf{x}|y)p(y)}{p(\mathbf{x})}$$

- $\bullet\,$ Two step approach Estimate generative model and then posterior for y (Naïve Bayes)
- Solving a more general problem [2, 1]
- Why not directly model $p(y|\mathbf{x})$? Discriminative approach
- Number of training examples needed to learn a PAC-learnable classifier \(\times VC-dimension \) of the hypothesis space
- Number of parameters for $p(y, \mathbf{x}) > \text{Number of parameters for } p(y|\mathbf{x})$

Discriminative classifiers need lesser training examples to for PAC learning than generative classifiers

- $y|\mathbf{x}$ is a Bernoulli distribution with parameter $\theta = sigmoid(\mathbf{w}^{\top}\mathbf{x})$
- When a new input \mathbf{x}^* arrives, we toss a coin which has $sigmoid(\mathbf{w}^{\top}\mathbf{x}^*)$ as the probability of heads
- If outcome is heads, the predicted class is 1 else 0
- Learns a linear boundary

Learning Task for Logistic Regression

Given training examples $\langle \mathbf{x}_i, y_i \rangle_{i=1}^D$, learn **w**

7 Logistic Regression - Training

- MLE Approach
- Assume that $y \in \{0, 1\}$
- What is the likelihood for a bernoulli sample?

- If
$$y_i = 1$$
, $p(y_i) = \theta_i = \frac{1}{1 + exp(-\mathbf{w}^{\top}\mathbf{x}_i)}$
- If $y_i = 0$, $p(y_i) = 1 - \theta_i = \frac{1}{1 + exp(\mathbf{w}^{\top}\mathbf{x}_i)}$

- In general,
$$p(y_i) = \theta_i^{y_i} (1 - \theta_i)^{1-y_i}$$

Log-likelihood

$$LL(\mathbf{w}) = \sum_{i=1}^{N} y_i \log \theta_i + (1 - y_i) \log (1 - \theta_i)$$

• No closed form solution for maximizing log-likelihood

To understand why there is no closed form solution for maximizing the log-likelihood, we first differentiate $LL(\mathbf{w})$ with respect to \mathbf{w} . We make use of the useful result for sigmoid:

$$\frac{d\theta_i}{d\mathbf{w}} = \theta_i (1 - \theta_i) \mathbf{x}_i$$

Using this result we obtain:

$$\frac{d}{d\mathbf{w}}LL(\mathbf{w}) = \sum_{i=1}^{N} \frac{y_i}{\theta_i} \theta_i (1 - \theta_i) \mathbf{x}_i - \frac{(1 - y_i)}{1 - \theta_i} \theta_i (1 - \theta_i) \mathbf{x}_i$$

$$= \sum_{i=1}^{N} (y_i (1 - \theta_i) - (1 - y_i) \theta_i) \mathbf{x}_i$$

$$= \sum_{i=1}^{N} (y_i - \theta_i) \mathbf{x}_i$$

Obviously, given that θ_i is a non-linear function of \mathbf{w} , a closed form solution is not possible.

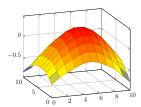
7.1 Using Gradient Descent for Learning Weights

- Compute gradient of LL with respect to w
- ullet A convex function of ${f w}$ with a unique global maximum

$$\frac{d}{d\mathbf{w}}LL(\mathbf{w}) = \sum_{i=1}^{N} (y_i - \theta_i)\mathbf{x}_i$$

• Update rule:

$$\mathbf{w}_{k+1} = \mathbf{w}_k + \eta \frac{d}{d\mathbf{w}_k} LL(\mathbf{w}_k)$$



7.2 Using Newton's Method

• Setting η is sometimes tricky

- Too large incorrect results
- Too small slow convergence
- Another way to speed up convergence:

Newton's Method

$$\mathbf{w}_{k+1} = \mathbf{w}_k + \eta \mathbf{H}_k^{-1} \frac{d}{d\mathbf{w}_k} LL(\mathbf{w}_k)$$

- Hessian or **H** is the second order derivative of the objective function
- Newton's method belong to the family of second order optimization algorithms
- For logistic regression, the Hessian is:

$$H = -\sum_{i} \theta_{i} (1 - \theta_{i}) \mathbf{x}_{i} \mathbf{x}_{i}^{\top}$$

7.3 Regularization with Logistic Regression

- Overfitting is an issue, especially with large number of features
- Add a Gaussian prior $\sim \mathcal{N}(\mathbf{0}, \tau^2)$
- Easy to incorporate in the gradient descent based approach

$$LL'(\mathbf{w}) = LL(\mathbf{w}) - \frac{1}{2}\lambda \mathbf{w}^{\mathsf{T}} \mathbf{w}$$
$$\frac{d}{d\mathbf{w}} LL'(\mathbf{w}) = \frac{d}{d\mathbf{w}} LL(\mathbf{w}) - \lambda \mathbf{w}$$
$$H' = H - \lambda I$$

where I is the identity matrix.

7.4 Handling Multiple Classes

- $p(y|\mathbf{x}) \sim Multinoulli(\boldsymbol{\theta})$
- Multinoulli parameter vector $\boldsymbol{\theta}$ is defined as:

$$\theta_j = \frac{exp(\mathbf{w}_j^{\top} \mathbf{x})}{\sum_{k=1}^{C} exp(\mathbf{w}_k^{\top} \mathbf{x})}$$

• Multiclass logistic regression has C weight vectors to learn

7.5 Bayesian Logistic Regression

- How to get the posterior for **w**?
- Not easy Why?

Laplace Approximation

- ullet We do not know what the true posterior distribution for ${f w}$ is.
- Is there a close-enough (approximate) Gaussian distribution?

One should note that we used a Gaussian prior for ${\bf w}$ which is not a conjugate prior for the Bernoulli distribution used in the logistic regression. In fact there is no convenient prior that may be used for logistic regression.

References

References

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