

# Economics 403A

## Bayesian Inference

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# Today's Class

- Introduction to Bayesian Inference
  - Bayes Theorem
  - Historical Background
  - Bayesians vs Frequentists
  - Priors, Likelihood and Posterior
  - Credible Intervals
  - Predictive Inference
  - Hypothesis Testing

# Introduction to Bayesian Inference

## Bayes' Theorem

$$P(A|X) = \frac{P(X|A)P(A)}{P(X)} = \frac{P(X|A)P(A)}{P(X|A)P(\sim A) + P(X|\sim A)P(\sim A)}$$

## Posterior = Prior + Data

Given some phenomenon A (e.g., cancer) that we want to investigate, and an observation X (e.g., mammography test outcome) that is evidence about A, we can update the original probability of A, given the new evidence X.

The **posterior** captures all the information that  $X$  can provide about  $A$

# Introduction to Bayesian Inference

## Goals of Inference

The two main goals of inference:

1. What is a good guess of the population model (the true parameters)?
2. How do I quantify my uncertainty in the guess?

Bayesian inference answers both questions directly through the posterior.

# Introduction to Bayesian Inference

## Motivating Example

- Assume the following information:
  - 1% of women aged 40 have breast cancer
  - A mammography test has an 80% success rate
  - A mammography test has a 10% false alarm rate.
- Given the above, a women aged 40 receives a positive mammography test. What is the probability that she actually has cancer?
- **Q:** Do you think this probability is high or low?

# Introduction to Bayesian Inference

## Motivating Example

- **Solution:**
  1. Identify A and X:  
 $X = +\text{mamm. Test}$ ,  $\sim X = - \text{mamm. Test}$   
 $A = \text{Cancer present}$ ,  $\sim A = \text{Cancer not present}$
  2. Identify the relevant probabilities  
 $P(A) = \text{Prob. of having breast cancer} = 1\%$   
 $P(\sim A) = \text{Prob. of not having breast cancer} = 99\%$   
 $P(X|A) = \text{Prob. that she receives a +mamm. test, given that she has cancer} = 80\%$   
 $P(X|\sim A) = \text{Prob. that she receives a +mamm. test but does not have cancer} = 10\%$   
 $P(A|X) = \text{Prob. That she has cancer given that she received a positive mamm. test.}$
  3. Apply Bayes' Theorem:

$$P(A|X) = \frac{P(X|A)P(A)}{P(X|A)P(\sim A) + P(X|\sim A)P(\sim A)} = \frac{80 \times 1}{80 \times 1 + 10 \times 99} = 7.5\%$$

# Introduction to Bayesian Inference

## Bayes' Theorem (in General)

**Prior:** contains all the available information about the parameter values “prior” to observing the data.

**Likelihood:** assuming you know the parameters exactly, how “likely” are the data. Represents the data that are now available.

$$P(A|B_k) = \frac{P(A_k)P(B|A_k)}{\sum_{i=1}^m P(A_i)P(B|A_i)}$$

**Posterior:** expresses our knowledge about the parameters after seeing the data.

**Normalizing Constant**

# Introduction to Bayesian Inference

## Background



**James Bernoulli** (1713) asked whether the rules of probability (then being worked out to calculate odds in gambling) could also be used in the problem of inference or inductive logic.



**Rev. Thomas Bayes** answered this in the affirmative in an essay published posthumously in 1763.



**Pierre-Simon Laplace** reformulated Bayes' theorem in 1812. He expressed it with far greater clarity than Bayes and used it to solve problems in celestial mechanics and medical statistics.

- Laplace combined the available astronomical data to provide an estimate (and uncertainty) on the mass of Saturn. He stated "...it is a bet of 11,000 to 1 that the error in this result is not 1/100th of its value". Note: The modern estimate differs from Laplace's by 0.63%

# Introduction to Bayesian Inference

## Background

- Bayes and Laplace thought of probability as measuring a degree of belief in something. They used the language of odds and betting in their arguments.
  - For about 100 years, the Bayesian interpretation reigned supreme.
- However, with the push in the 19th and early 20th centuries towards more rigorous arguments in mathematics, this view was considered too subjective and a new definition of probability was developed.
  - In the beginning of the 20th century, the so-called **frequentist** interpretation emerged (Ronald Fisher, Jerzy Neyman, Egon Pearson)
  - Probability was defined solely in terms of the long-term frequency of events.
  - Note: This definition works fine for rolling dice and tossing coins but where does it leave Laplace's measurement of the mass of Saturn ?

# Introduction to Bayesian Inference

## Background

- During the second half of the 20th century, Bayesian statistics made a strong comeback.
- The frequentist approach is plagued by inconsistencies and limitations.
- Bayesian models are often analytically intractable and thus require methods based on simulation. Cheap and fast computers, and general-purpose software such as BUGS, resolved this issue.
- Hierarchical graphical models, such as hidden Markov models and Bayesian networks, provided a unified framework for Bayesian computation.
  - Bayesian approaches are very common in machine learning.
  - Recent work in cognitive science indicates that our minds may well operate by Bayesian inference
  - Bonus Fact: Alan Turing cracked the U-boats encrypted communications using Bayesian methods



# Introduction to Bayesian Inference

## Background

- **Frequentists** treat the parameters as **fixed** (deterministic).
  - Considers the training data to be a random draw from the population model.
  - Uncertainty in estimates is quantified through the sampling distribution: what is seen if the estimation procedure is repeated over and over again, over many sets of training data.
- **Bayesians** treat the parameters as **random**.
  - Key element is a prior distribution on the parameters.
  - Using Bayes' theorem, combine prior with data to obtain a posterior distribution on the parameters.
  - Uncertainty in estimates is quantified through the posterior distribution.

# Introduction to Bayesian Inference

## Key differences between Bayesian and Classical (Frequentist) Approaches

- Unknown parameter  $\theta$ :
  - **Frequentist**: a fixed number
  - **Bayesian**: a random variable
- Inference about  $\theta$ :
  - **Frequentist**: ad hoc (different types of estimators/tests are “best” for different problems, no unique algorithm)
  - **Bayesian**: given 3 choices (likelihood, prior, loss), there is a unique inferential procedure

# Introduction to Bayesian Inference

## Key differences between Bayesian and Classical (Frequentist) Approaches

- Interpretation of the conclusions: e.g., interval estimation of  $\theta$ :
  - **Frequentist**:  $(1 - \alpha)100\%$  confidence interval of  $\theta$ : among all such data sets  $y$ , in  $(1 - \alpha)100\%$  of them,  $\theta$  belongs to this interval
  - **Bayesian**:  $(1 - \alpha)100\%$  credible interval of  $\theta$ : for given data  $y$ ,  $\theta$  belongs to this interval with probability  $1 - \alpha$
- Subjectivity/objectivity of conclusions about  $\theta$ :
  - **Frequentist**: objective (some argue that subjectivity is often hidden in this approach)
  - **Bayesian**: subjective: use a priori information about  $\theta$

# Introduction to Bayesian Inference

## Posterior Predictive Distribution

If we wish to predict a new observation  $\tilde{x}$  on the basis of the sample  $x = (x_1, \dots, x_n)$ , we may use its posterior predictive distribution. This is defined to be the conditional distribution of  $\tilde{x}$  given  $x$ :

$$p(\tilde{x}|x) = \int p(\tilde{x}, \theta|x)d\theta = \int p(\tilde{x}|x, \theta)p(\theta|x)d\theta$$

where  $p(\tilde{x}|x, \theta)$  is the density of the predictive distribution. It is easy to simulate the posterior predictive distribution. First, draw simulations  $\theta_1, \dots, \theta_L$  from the posterior  $p(\theta|x)$ , then, for each  $i$ , draw  $\tilde{y}_i$  from the predictive distribution  $p(\tilde{x}|x, \theta_i)$ .

# Introduction to Bayesian Inference

## Posterior Predictive Distribution

$p(x|\theta)$

likelihood

$\pi(\theta)$

prior

$$p(x) = \int p(x|\theta)\pi(\theta) d\theta$$

marginal likelihood

$$p(\theta|x) = \frac{p(x|\theta)\pi(\theta)}{p(x)}$$

posterior probability

$$p(x_{new}|x) = \int p(x_{new}|\theta)\pi(\theta|x) d\theta$$

predictive probability

# Introduction to Bayesian Inference

## Posterior Predictive Distribution

- **Example:** Assume that we have a coin with unknown probability  $\theta$  of a head. If there occurs  $x$  heads among the first  $n$  tosses what is the probability of a head on the next throw?
- Let  $\tilde{x} = 1$  ( $\tilde{x} = 0$ ) indicate the event that the next throw is a head (tail). If the prior of  $\theta$  is a Beta( $\alpha, \beta$ ), then

$$\begin{aligned} p(\tilde{x}|x) &= \int_0^1 p(\tilde{x}|x, \theta)p(\theta|x)d\theta = \int_0^1 \theta^{\tilde{x}}(1-\theta)^{1-\tilde{x}} \frac{\theta^{\alpha+x-1}(1-\theta)^{\beta+n-x-1}}{B(\alpha+x, \beta+n-x)} d\theta \\ &= \frac{(\alpha+x)^{\tilde{x}}(\beta+n-x)^{1-\tilde{x}}}{\alpha+\beta+n} \end{aligned}$$

Therefore, for example  $p(\tilde{x}=1|x)=(\alpha+x)/(\alpha+\beta+n)$ . This tends to the sample proportion  $x/n$  as  $n \rightarrow \infty$ , so that the role of the prior information vanishes.

# Introduction to Bayesian Inference

## Prior

Where did the prior come from?

→ There are two schools of thought on the prior:

- Subjective Bayesian
  - The prior is a summary of our subjective beliefs about the data.
  - E.g., in the coin flipping example: the prior for  $p$  should be strongly peaked around 1/2.
- Objective Bayesian
  - The prior should be chosen in a way that is “uninformed”.
  - E.g., in the coin flipping example: the prior should be uniform on [0, 1].

# Introduction to Bayesian Inference

## Non-informative Priors

- We often do not have any prior information, although true Bayesian's would argue we always have some prior information!
- We would hope to have good agreement between the frequentist approach and the Bayesian approach with a non-informative prior.
- Diffuse or flat priors are often better terms to use as no prior is strictly non-informative!
- For our example of an unknown mean, candidate priors are a Uniform distribution over a large range or a Normal distribution with a huge variance.

# Introduction to Bayesian Inference

## Improper Priors

- The limiting prior of both the Uniform and Normal is a Uniform prior on the whole real line.
- Such a prior is defined as **improper** as it is not strictly a probability distribution and does not integrate to 1.
- Some care has to be taken with improper priors however in many cases they are acceptable provided they result in a proper posterior distribution.
- Uniform priors are often used as non-informative priors however it is worth noting that a uniform prior on one scale can be very informative on another.
- **For example:** If we have an unknown variance we may put a uniform prior on the variance, standard deviation or  $\log(\text{variance})$  which will all have different effects.

# Introduction to Bayesian Inference

## Conjugate Priors

- Computations can often be facilitated using **conjugate prior** distributions. We say that a prior is conjugate for the likelihood if the prior and posterior distributions belong to the same family.

Note: See e.g., [https://en.wikipedia.org/wiki/Conjugate\\_prior](https://en.wikipedia.org/wiki/Conjugate_prior) for a list of common conjugate priors.

# Introduction to Bayesian Inference

## Posterior

The posterior can be used in many ways to estimate the parameters. For example, we can compute:

- The mean
- The median
- The mode

Depending on context, these are all potentially useful ways to estimate model parameters from a posterior.

# Introduction to Bayesian Inference

## Example: Normal Distribution

- Consider a single-parameter model and assume  $x = (x_1, \dots x_n)$  is a sample from a normal distribution with unknown mean  $\theta$  and variance  $\sigma^2$ .
- Likelihood:  $p(x|\theta) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(x_i - \theta)^2} \propto e^{-\frac{n}{2\sigma^2}(\theta - \bar{x})^2}$
- We can find a conjugate prior by replacing  $\sigma^2/n$  by  $\tau_0^2$  and with  $\mu_0$

$$\rightarrow p(\theta) \propto e^{-\frac{1}{2\tau_0^2}(\theta - \mu_0)^2} \sim N(\mu_0, \tau_0^2)$$

# Introduction to Bayesian Inference

## Example: Normal Distribution

- Now we can easily compute the posterior  $p(\theta|x)$ :

$$p(\theta|x) \propto p(\theta)p(x|\theta) \propto e^{-\frac{1}{2\tau_0^2}(\theta-\mu_0)^2} e^{-\frac{n}{2\sigma_o^2}(\theta-\bar{x})^2}$$

$$\propto e^{-\frac{1}{2\tau_n^2}(\theta-\mu_n)^2} \sim N(\mu_n, \tau_n^2)$$

$$\text{where } \mu_n = \frac{\frac{1}{\tau_0^2}\mu_0 + \frac{n}{\sigma^2}\bar{x}}{\frac{1}{\tau_0^2} + \frac{n}{\sigma^2}} \text{ and } \tau_n^2 = \left( \frac{1}{\tau_0^2} + \frac{n}{\sigma^2} \right)^{-1}$$

# Introduction to Bayesian Inference

## Example: Normal Distribution

- The inverse of variance is called the precision. Therefore, we see that:  
 $\rightarrow \text{posterior precision} = \text{prior precision} + \text{data precision}$
- For example, for our previous case, the prior precision is  $1/\tau_0^2$  and data precision  $n/\sigma^2$  (i.e., the inverse of the variance of the sample mean)
- The **posterior mean** is a weighted average of the prior mean  $\mu$  and sample mean  $\bar{x}$  where the weights are the corresponding precisions. When  $n \rightarrow \infty$  (or when  $\tau_0 \rightarrow \infty$ ), the *role of the prior information vanishes*. Thus, for large values of  $n$ , approximately  $\theta|y \sim N(\bar{x}, \sigma^2/n)$ .

# Introduction to Bayesian Inference

## Example: Normal Distribution

- Next, we determine the posterior predictive distribution of a new observation  $\tilde{x}$ . The joint posterior distribution of  $\theta$  and  $\tilde{x}$  is:

$$\rightarrow p(\theta, \tilde{x}|x) = p(\theta|y)p(\tilde{x}|x, \theta)$$
$$\propto e^{-\left(\frac{1}{2\tau_n^2}(\theta-\mu_n)^2 + \frac{1}{2\sigma^2}(\tilde{x}-\theta)^2\right)}$$

$$\rightarrow E[\tilde{x}|x] = E[\theta|x] = \mu_n$$

$$\text{Var}[\tilde{x}|x] = E[\sigma^2|x] + \text{Var}[\theta|x] = \sigma^2 + \tau_n^2$$

- Therefore, the posterior predictive distribution is:  
 $p(\tilde{x}|x) = N(\tilde{x}|\mu_n, \tau_n^2)$

# Introduction to Bayesian Inference

## Point and Interval Estimation

- In Bayesian inference the outcome of interest for a parameter is its full posterior distribution however we may be interested in summaries of this distribution.
- A simple **point estimate** would be the **mean of the posterior** (although the median and mode are alternatives)
- **Interval estimates** are also easy to obtain from the posterior distribution and are given several names, for example **credible intervals**, Bayesian confidence intervals and Highest density regions (HDR). All of these refer to the same quantity.

# Introduction to Bayesian Inference

## Credible Intervals

- **Credible intervals** can be interpreted in the more natural way that there is a probability of 0.95 that the interval contains  $\mu$  rather than the frequentist conclusion that 95% of such intervals contain  $\mu$ .
- **Def:** A credible interval, for a real-valued parameter  $\psi(\theta)$ , is an interval  $C(s) = [l(s), u(s)]$  that we believe will contain the true value of  $\psi$ . As with the sampling theory approach, we specify a probability  $\gamma$  and then find an interval  $C(s)$  satisfying

$$\Pi(\psi(\theta) \in C(s) | s) = \Pi (\{\theta : l(s) \leq \psi(\theta) \leq u(s)\} | s) \geq \gamma$$

We then refer to  $C(s)$  as a  $\gamma$ -credible interval for  $\psi$ .

# Introduction to Bayesian Inference

## Credible Intervals

- A Bayesian credible interval of size  $1 - \alpha$  is an interval  $(a, b)$  such that

$$P(a \leq \theta \leq b|x) = 1 - \alpha$$
$$\int_a^b p(\theta|x) d\theta = 1 - \alpha$$

- **Note:** When you are calculating credible intervals, you will find the values of  $a$  and  $b$  by several means. You could be asked do the following:
  - Find the  $a, b$  using means of calculus to determine the credible interval or set.
  - Use a Standard Normal (Z-scores) when appropriate.
  - Use R to approximate the values of  $a$  and  $b$ .

# Introduction to Bayesian Inference

## Credible Intervals

- Our definition for the credible interval could lead to many choices of  $(a, b)$  for particular problems. Suppose that we required our credible interval to have equal probability  $\alpha/2$  in each tail. That is, we will assume:

$$P(\theta < a|x) = \alpha/2 \text{ and } P(\theta > b|x) = \alpha/2$$

- **Note:** Credible intervals are not necessarily unique.

# Introduction to Bayesian Inference

## Confidence Intervals vs. Credible Intervals

- A **confidence interval** is constructed to contain  $\theta$  a percentage of the time, say 95%.
  - Suppose our confidence level is 95% and our interval is  $(L, U)$ . Then we are 95% confident that the true value of  $\theta$  is contained in  $(L, U)$  in the long run.
  - In the long run means that this would occur nearly 95% of the time if we repeated our study millions and millions of times.
- Conceptually, **probability comes into play** in a frequentist confidence interval **before collecting the data**.
  - For example, there is a 95% probability that **we will collect data** that produces an interval that contains the true parameter value.

# Introduction to Bayesian Inference

## Confidence Intervals vs. Credible Intervals

- We would like to make statements about the probability that the interval contains the true parameter value given the **data that we actually observed**.
- Probability comes into play in a Bayesian credible interval **after collecting the data**
  - For example, based on the data, we now think there is a 95% probability that the true parameter value is in the interval.
- This is more natural because we want to make a probability statement regarding that data after we have observed it.

# Introduction to Bayesian Inference

## Confidence Intervals vs. Credible Intervals

- **Example:** Assume we are interested in the proportion of the population of American college students that sleep at least eight hours each night ( $\theta$ ).
- The Gamecock, at the USC printed an internet article “College Students Don’t Get Enough Sleep” (2004).
  - Most students spend six hours sleeping each night.
- University of Notre Dame’s paper, “Fresh Writing” (2003).
  - The article reported took random sample of 100 students: “approximately 70% reported to receiving only five to six hours of sleep on the weekdays
  - 28% receiving seven to eight and only 2% receiving the healthy nine hours for teenagers.”

# Introduction to Bayesian Inference

## Confidence Intervals vs. Credible Intervals

- Assumptions:
  - Suppose we collect a random sample of 27 students from UCLA, where 11 students recorded they slept at least eight hours per night.
- Based on this information, we are interested in estimating  $\theta$ 
  - From USC's and UND's studies, we believe it's probably true that most college students get less than eight hours of sleep.
  - We want our prior to assign most of the probability to values of  $\theta < 0.5$ .
  - From the information given, we decide that our best guess for  $\theta$  is 0.3, although we think it is very possible that  $\theta$  could be any value in  $[0, 0.5]$

# Introduction to Bayesian Inference

## Confidence Intervals vs. Credible Intervals

- Calculations:
- Lets assume that a Beta-Binomial Distribution is appropriate, such that:  $X|\theta \sim \text{Binomial}(n, \theta)$   
 $\theta \sim \text{Beta}(a, b)$

→ Need to find:  $\pi(\theta|X)$  and  $\theta|x$ :

$$\begin{aligned}\pi(\theta|x) &\propto p(x|\theta)p(\theta) \\ &\propto \binom{n}{x} \theta^x (1-\theta)^{n-x} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{a-1} (1-\theta)^{b-1} \\ &\propto \theta^x (1-\theta)^{n-x} \theta^{a-1} (1-\theta)^{b-1} \\ &\propto \theta^{x+a-1} (1-\theta)^{n-x+b-1} \implies\end{aligned}$$

$$\theta|x \sim \text{Beta}(x+a, n-x+b).$$

# Introduction to Bayesian Inference

## Confidence Intervals vs. Credible Intervals

- Calculations:
- Given that  $\theta|x \sim \text{Beta}(x+a, n-x+b)$ , we need to find  $a$  and  $b$ .
  - If given the studies' information, we believe that the median of  $\theta$  is 0.3 and the 90th percentile is 0.5, we can estimate the unknown values of  $a$  and  $b$  by solving:

$$\int_0^{0.3} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{a-1} (1-\theta)^{b-1} d\theta = 0.5$$
$$\int_0^{0.5} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{a-1} (1-\theta)^{b-1} d\theta = 0.9$$

} Use **BBsolve** in R

# Introduction to Bayesian Inference

## Confidence Intervals vs. Credible Intervals

- R Code:

```
## install the BBsolve package
install.packages("BB", repos="http://cran.r-project.org")
library(BB)
## using percentiles
myfn <- function(shape){
  test <- pbeta(q = c(0.3, 0.5), shape1 = shape[1],
    shape2 = shape[2]) - c(0.5, 0.9)
  return(test)
}
BBlolve(c(1,1), myfn)
## using quantiles
fn = function(x){qbeta(c(0.5,0.9),x[1],x[2])-c(0.3,0.5)}
BBlolve(c(1,1),fn)
```

→  $a = 3.3$  and  $b = 7.2$

# Introduction to Bayesian Inference

## Confidence Intervals vs. Credible Intervals

- Model:

Assume that the prior on  $\theta$  is a Beta(3.3,7.2):

$$X \mid \theta \sim \text{Binomial}(27, \theta)$$

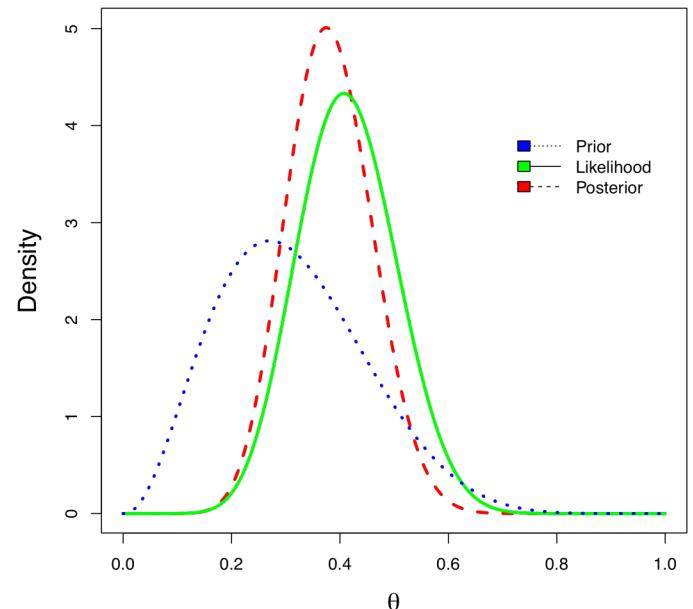
$$\theta \sim \text{Beta}(3.3, 7.2)$$

$$\theta \mid x \sim \text{Beta}(x + 3.3, 27 - x + 7.2)$$

$$\theta \mid 11 \sim \text{Beta}(14.3, 23.2)$$

- Thus, the posterior distribution is

$$\rightarrow \theta \mid 11 \sim \text{Beta}(11 + 3.3, 27 - 11 + 7.2) = \text{Beta}(14.3, 23.2).$$



# Introduction to Bayesian Inference

## Confidence Intervals vs. Credible Intervals

- R Code:

```
th = seq(0,1,length=500)
a = 3.3
b = 7.2
n = 27
x = 11
prior = dbeta(th,a,b)
like = dbeta(th,x+1,n-x+1)
post = dbeta(th,x+a,n-x+b)
quartz()
plot(th,post,type="l",ylab="Density",lty=2,lwd=3,xlab =
expression(theta),col="red",cex.lab=1.5)
lines(th,like,lty=1,lwd=3,col="green")
lines(th,prior,lty=3,lwd=3,col="blue")
legend(0.7,4,c("Prior","Likelihood","Posterior"),lty=c(3,1,2),lwd=c(1,1,1),fill =
c("blue","green","red"),cex=1,bty='n')
```

# Introduction to Bayesian Inference

## Confidence Intervals vs. Credible Intervals

- Based on the previous results, find a **90% credible interval**
- We need to solve for  $L$  (*lower*) and  $U$  (*upper*) such that:

$$P(\theta < L|x) = 0.05 \text{ and } P(\theta > U|x) = 0.05$$

→ Need to solve numerically:

$$\int_0^L \text{Beta}(14.3, 23.3)d\theta = 0.05 \text{ and } \int_U^1 \text{Beta}(14.3, 23.3)d\theta = 0.05$$

```
a = 3.3; b = 7.2
n = 27; x = 11
a.star = x+a
b.star = n-x+b
L = qbeta(0.05,a.star,b.star)
U = qbeta(1-0.05,a.star,b.star)
```

R Code gives us the values:  
 $L = 0.256$  and  $U = 0.514$

# Introduction to Bayesian Inference

## Hypothesis Testing

- The frequentist approach to hypothesis testing would compare a null hypothesis  $H_0$  with an alternative  $H_1$  through a test statistic  $T$  which typically obtains a larger value when  $H_1$  is true than when  $H_0$  is true. The null hypothesis is rejected with a level  $\alpha$  if the observed value of the test statistic,  $t_{obs}$ , is larger than the critical value  $t_c$  where  $Pr(T > t_c | H_0) = \alpha$ . The p-value,  $p = Pr(T \geq t_{obs} | H_0)$ .
- In frequentist statistics, we do not assign probabilities to hypotheses. In particular, the p-value cannot be interpreted as  $p(H_0)$ .

# Introduction to Bayesian Inference

## Hypothesis Testing

- In the Bayesian approach, we may assign the prior probabilities  $p(H_0)$  and  $p(H_1)$ , and, using Bayes' theorem, compute the posterior probabilities.

$$p(H_i|x) = \frac{p(H_i)p(x|H_i)}{p(H_0)p(x|H_0) + p(H_1)p(x|H_1)}, \quad i = 0, 1$$

- In the frequentist approach it is not absolutely necessary to specify an alternative hypothesis. Further, if an alternative is specified, the p-value is independent of it. In the Bayesian approach, the both hypotheses must be fully specified.

# Introduction to Bayesian Inference

## Hypothesis Testing

- In the frequentist approach it is not absolutely necessary to specify an alternative hypothesis. Further, if an alternative is specified, the p-value is independent of it. In the Bayesian approach, the both hypotheses must be fully specified.
- One usually computes the posterior odds

$$\frac{p(H_1|x)}{p(H_0|x)} = \frac{p(x|H_1)}{p(x|H_0)} \times \frac{p(H_1)}{p(H_0)}$$

which depends on the data  $y$  only through the **Bayes factor**

$$B = \frac{p(x|H_1)}{p(x|H_0)} \quad ]$$

B is the likelihood ratio of  $H_0$  against  $H_1$  i.e. the odds in favor of  $H_0$  against  $H_1$  that are given by the data.  
See: H. Jeffreys' scale interpretation

# Introduction to Bayesian Inference

## Hypothesis Testing (Example)

Visualization( <http://rpsychologist.com/d3/bayes/> )

# Introduction to Bayesian Inference

## Linear Regression Example (Example)

Visualization(<https://towardsdatascience.com/introduction-to-bayesian-linear-regression-e66e60791ea7>)