

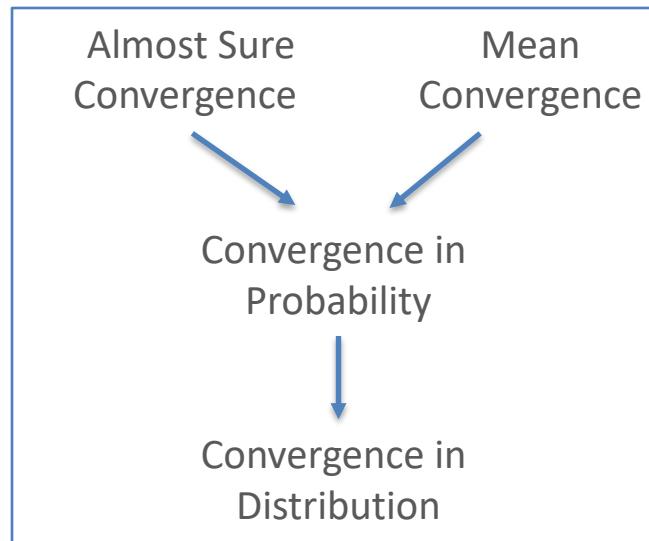
Economics 403A

Sampling Distributions and Limits

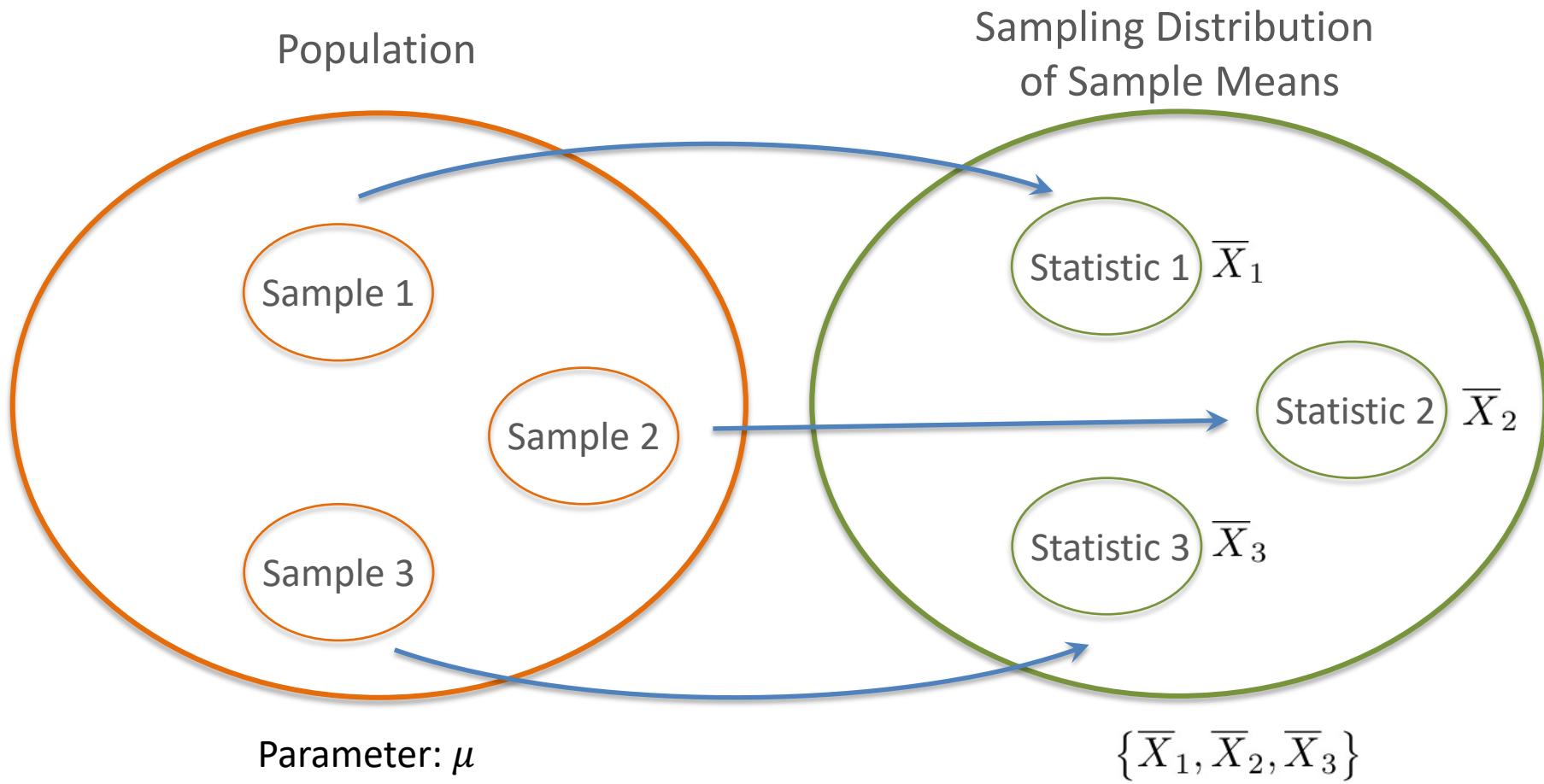
Dr. Randall R. Rojas

Today's Class

- Sampling Distributions
- Limits
 - Convergence in Probability
 - Weak Law of Large Numbers
 - Convergence in Probability 1
 - Strong Law of Large Numbers
 - Convergence in Distribution
 - Central Limit Theorem
- Monte Carlo Approximations
- Normal Distribution Theory
 - Chi-Squared Distribution
 - t-Distribution
 - F-Distribution



Sampling Distributions



Sampling Distributions

Anatomy of a Basic Sampling Algorithm

for i= 1:k

1. Choose a sample size n, and a statistic
2. Randomly draw a sample from the population with the same size
3. Estimate the statistic from the sample

end

Sampling Distributions

Def: Sampling Variability

= The variability among random samples from the same population.

Def: Sampling Distribution

= A probability distribution that characterizes some aspect of sampling variability.
= The distribution of a statistic given repeated sampling.

A sampling distribution tells us how close the resemblance between the sample and population is likely to be

Goal: We use probability theory to derive a distribution for a statistic, which allows us to make statistical inferences about population parameters.

Sampling Distributions

Suppose X_1, X_2, \dots, X_n is a sample from some distribution, f is a function of this sequence, and we want to estimate the distribution of the r.v., $Y = f(X_1, X_2, \dots, X_n)$, i.e., Y 's ‘sampling distribution’.

Q: How can we obtain the distribution of Y ?

A: Analytically (works well when n is small) or Approximately (through sampling)

Sampling Distributions

- Example: Given a sample X_1, X_2 of size $n = 2$, with p.m.f. p_X , and $Y_2 = (X_1 X_2)^{1/2}$, find the distribution of Y_2 .

$$p_X(x) = \begin{cases} 1/2 & x = 1 \\ 1/4 & x = 2 \\ 1/4 & x = 3 \\ 0 & \text{otherwise.} \end{cases}$$

Note: Y_2 is known as the **Geometric Average**, which is widely used in finance.

- Solution: We can compute it analytically

y	Sample	$p_{Y_2}(y)$
1	$\{(1, 1)\}$	$(1/2)(1/2) = 1/4$
$\sqrt{2}$	$\{(1, 2), (2, 1)\}$	$(1/2)(1/4) + (1/4)(1/2) = 1/4$
$\sqrt{3}$	$\{(1, 3), (3, 1)\}$	$(1/2)(1/4) + (1/4)(1/2) = 1/4$
2	$\{(2, 2)\}$	$(1/4)(1/4) = 1/16$
$\sqrt{6}$	$\{(2, 3), (3, 2)\}$	$(1/4)(1/4) + (1/4)(1/4) = 1/8$
3	$\{(3, 3)\}$	$(1/4)(1/4) = 1/16$

Sampling Distributions

- **Application:** Multiple-Period Realized Return

Suppose you invest \$1000 at the beginning of year 1, and get a 12% return at the end this year, and then get 8% at the end of year 2. What was your '**average**' return over the two years?

Option 1: Arithmetic Average: $(12+8)/2 = 10\%$

Q: Is this correct? **No**

Q: Why is this the **wrong** answer? **A:** Does not take into account interest over interest.

- If you start with \$1000, then at the end of 2 years you have
 $\$1000 (1.12)(1.08) = \textcolor{brown}{\$1209.6}$.
- The amount of money you would get if you invested \$1000 at 10%/year for two years is $1000*(1.1)^2 = \$1210$. They are not the same!

Option 2: Geometric Average: $[(1+0.12)(1+0.08)]^{1/2} - 1 = 0.09982 = 9.982\%$

The geometric average gives us the correct per annum return:
 $1000(1.09982)^2 = \textcolor{brown}{\$1209.6}$.

Multiple-Period Realized Return Arithmetic vs. Geometric Average

- Example: Time Series of HPR for the S&P 500

Period	HPR (decimal)	Gross HPR = 1+HPR	
2001	-0.1189	0.8811	\times
2002	-0.2210	0.7790	\cdot
2003	0.2869	1.2869	\cdot
2004	+ 0.1088	1.1088	\times
2005	<u>0.0491</u>	<u>1.0491</u>	

Arithmetic Average:

$$\begin{aligned} &= 1/5(-0.1189 + \dots + 0.0491) \\ &= 0.0210 \\ &= 2.10\% \end{aligned}$$

Geometric Average:

$$\begin{aligned} &= (0.8811 \times \dots \times 1.0491)^{1/5} - 1 \\ &= 0.0054 \\ &= 0.54\% \end{aligned}$$

The larger the swings in rates of return, the greater the discrepancy between the arithmetic and geometric averages.

Multiple-Period Realized Return Arithmetic vs. Geometric Average

- Example: S&P 500 annual returns from 2000-2010
- Q: Given the two averages, if you had invested \$10,000 in the S&P500 from 2000-2010, how much would you have in your account?
 - $R = \{-9.2, -11.9, -22.1, 28.7, 10.9, 4.9, 15.8, 5.5, -37.0, 26.5, 15.1\}$
- Arithmetic Avg = 2.47%, Geometric Avg = 0.41%

\$10,247.00	\$10,041.00	}	correct amount
- Reality?

After adjusting for inflation
→ Arithmetic (\$9,994), Geometric (**\$9,798**)

Multiple-Period Realized Return

Arithmetic vs. Geometric Average 3 of 3

Concluding Remarks

- Volatility lowers investment returns.
- Most returns are reported as an arithmetic average because this is the highest average that can be reported.
 - Arithmetic averages are accurate only if there is no volatility.
- Careful with e.g., mutual fund and hedge fund managers! –They usually quote arithmetic averages.

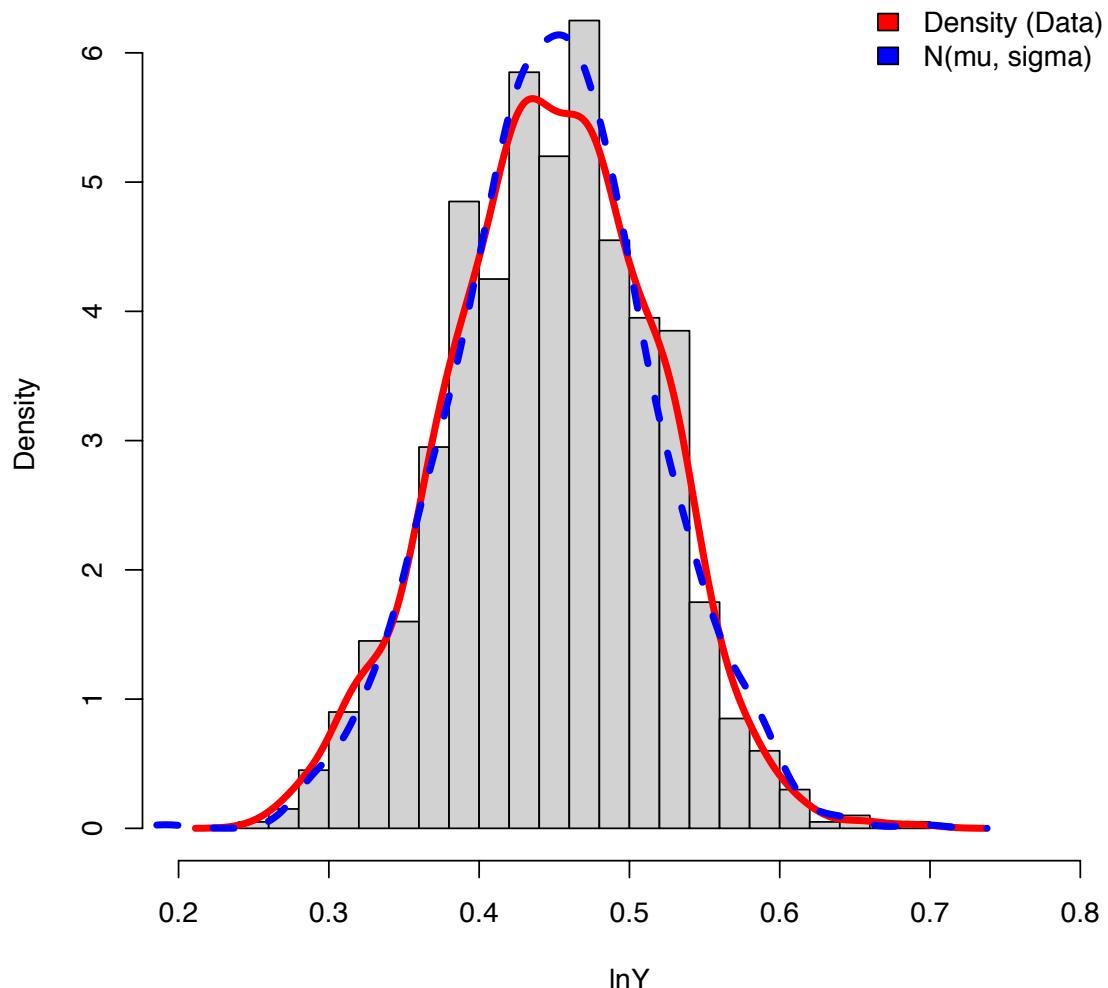
Sampling Distributions

- **Example:** Given a sample X_1, X_2, \dots, X_{20} of size $n = 20$, with p.m.f. p_X , and $Y_2 = (X_1 \cdots X_{20})^{1/2}$, find the distribution of Y_2 .
- **Note:** We *cannot* compute it analytically! Even computationally it would be difficult because of the large number of operations.

Q: How do we find Y_2 ?

A: Approximation: $\ln Y = \frac{1}{n} \sum_{i=1}^n \ln X_i$

Sampling Distributions



Limits

Convergence in Probability

A sequence of random variables $X_1, X_2, X_3, \dots, X_n$ converges in probability to a random variable X , i.e., $X_n \xrightarrow{p} X$ if:

$$\lim_{n \rightarrow \infty} P(|X_n - X| \geq \varepsilon) = 0, \quad \forall \varepsilon > 0$$

Example: Let $X_n \sim \text{Exp}(n)$, show that $X_n \xrightarrow{p} 0$. That is, the sequence $X_1, X_2, X_3, \dots, X_n$ converges in probability to the zero random variable X .

→ Proof:

$$\begin{aligned}\lim_{n \rightarrow \infty} P(|X_n - 0| \geq \varepsilon) &= \lim_{n \rightarrow \infty} P(X_n > \varepsilon) \\ &= \lim_{n \rightarrow \infty} e^{-n\varepsilon} \\ &= 0, \quad \forall \varepsilon > 0\end{aligned}$$

Limits

Convergence in Probability (Textbook Example)

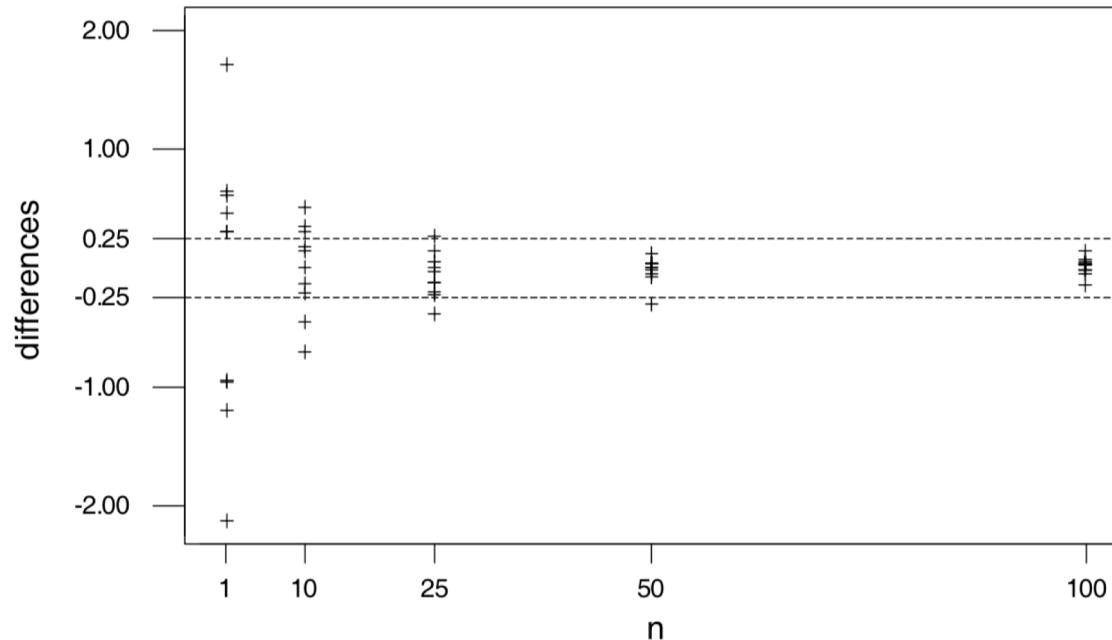


Figure 4.2.1: Plot of 10 replications of $\{X_n - Y\}$ illustrating the convergence in probability of X_n to Y .

Limits

Weak Law of Large Numbers (WLLN)

Let $X_1, X_2, X_3, \dots, X_n$ be i.i.d random variables with a finite expected value $E[X_i] = \mu < \infty$. Then for any $\varepsilon > 0$,

$$\lim_{n \rightarrow \infty} P(|\bar{X} - \mu| \geq \varepsilon) = 0.$$

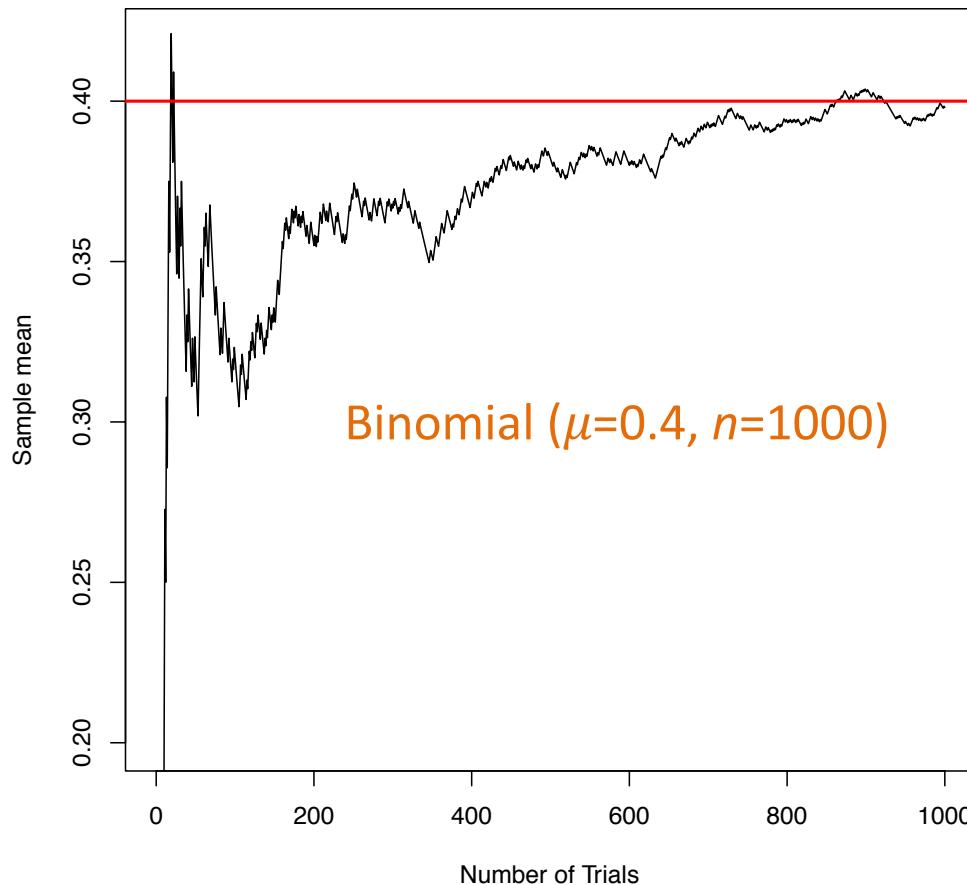
Note: A common notation is $M_n = \bar{X}$.

The WLLN is also known as **Bernoulli's Theorem**. In simpler terms, it states that: "*the mean of the results obtained from a large number of trials is close to the population mean*".

Limits

Weak Law of Large Numbers (WLLN)

Simulation of the Weak (and Strong) Law of Large Numbers

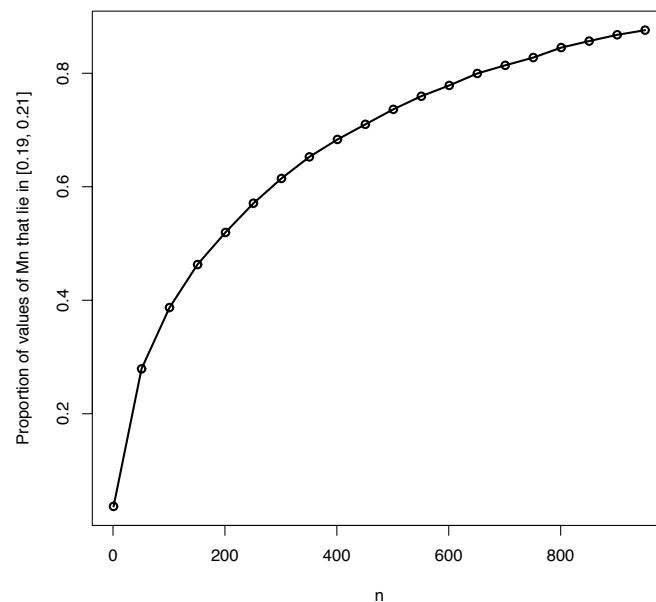


Limits

Weak Law of Large Numbers (WLLN)

Example (4.2.12): Generate i.i.d. X_1, \dots, X_n distributed Exponential(5) and compute M_n when $n = 20$. Repeat this N times, where N is large (if possible, take $N = 10^5$, otherwise as large as is feasible), and compute the proportion of values of M_n that lie between 0.19 and 0.21. Repeat this with $n = 50$.

→ Since $X \sim Exp(5)$, $\mu = 1/5 = 0.2$. Therefore, as n increases, the proportion of values that lie between 0.19 and 0.21 should increase.



Limits

Convergence in Probability 1 (“almost surely”)

A sequence of random variables X_1, X_2, \dots, X_n converges *almost surely* to a random variable X , i.e., $X_n \xrightarrow{a.s.} X$, if

$$P\left(\lim_{n \rightarrow \infty} X_n = X\right) = 1$$

Example: A fair coin is tossed once. The sample space is $S=\{H,T\}$. We define a sequence of random variables X_1, X_2, \dots, X_n on this sample space according to:

$$X_n(s) = \begin{cases} \frac{n}{n+1} & \text{if } s=H \\ (-1)^n & \text{if } s=T \end{cases} \rightarrow \begin{array}{l} s = H \rightarrow X_n = \{1/2, 2/3, \dots\} \rightarrow 1 \text{ as } n \rightarrow \infty \\ s = T \rightarrow X_n = \{-1, 1, \dots\} \text{ (does } \mathbf{not} \text{ converge)} \end{array}$$

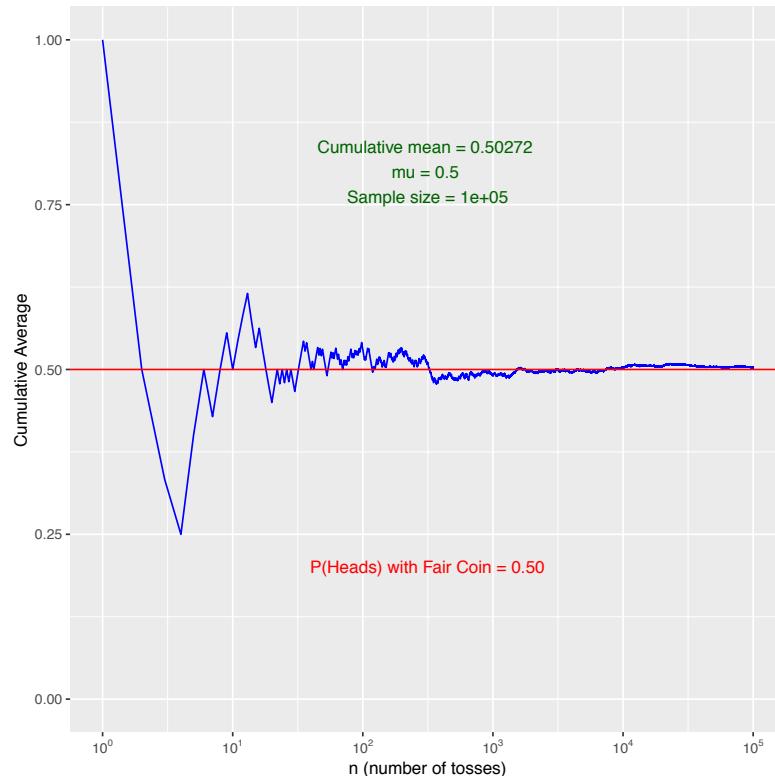
$$P\left(\lim_{n \rightarrow \infty} X_n = 1\right) = P(H) = 1/2$$

X_n converges if $s=H$ but not if $s=T$, but since the probability that X_n converges to X is equal to 1 as $n \rightarrow \infty$, then $X_n \xrightarrow{a.s.} X$

Limits

Strong Law of Large Numbers

Let X_1, X_2, \dots, X_n be i.i.d., random variables each having a finite mean $E[X_i] = \mu < \infty$, and $M_n = \frac{X_1 + \dots + X_n}{n}$, then $M_n \xrightarrow{a.s.} \mu$.



Limits

Strong Law of Large Numbers

Application: *Business Growth*

According to the LLN as additional units are added to a sample, the average of the sample converges to the average of the population.

→ The LLN implies that the more a company grows, the harder it is for the company to sustain that percentage of growth.

Company A: \$100M (Market Cap),
year 1 growth rate = 50%.

- A 50% growth rate for the next 6 years would grow the company from \$100M to \$1.2B.

Company B: \$20M (Market Cap)
year 1 growth rate = 50%.

- A 50% growth rate for the next 6 years would grow the company from \$20M to \$228M

Which company has more room for growth (stock appreciation)?

Limits

Convergence in Distribution

A sequence of random variables X_1, X_2, \dots, X_n converges in *distribution* to a random variable X , i.e., $X_n \xrightarrow{d} X$, if $\lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x)$ for all x at which $F_X(x)$ is continuous.

Example: Let X_1, X_2, \dots, X_n be a sequence of r.v's such that

$$F_{X_n}(x) = \begin{cases} 1 - \left(1 - \frac{1}{n}\right)^{nx} & x > 0 \\ 0 & \text{otherwise} \end{cases}$$

Show that X_n converges in distribution to $\text{Exp}(1)$.

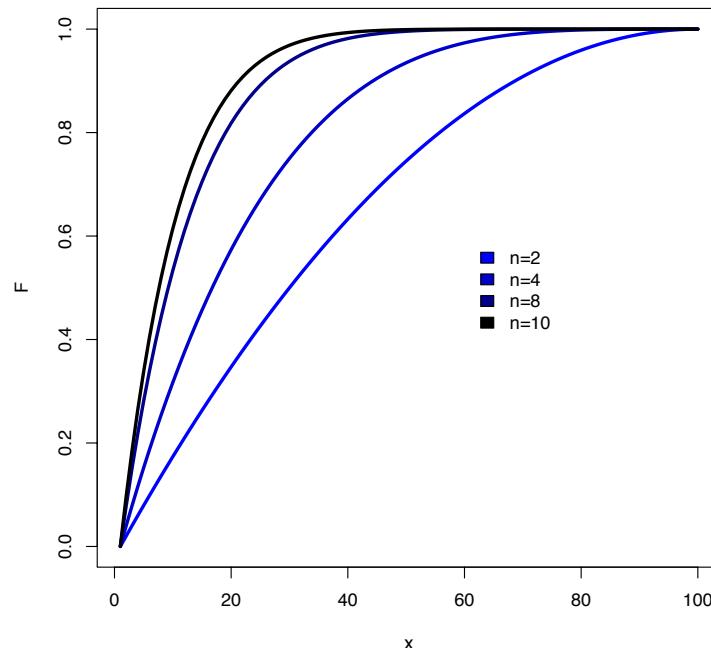
Limits

Convergence in Distribution

Solution: For $x \leq 0$, $F_{X_n}(x) = F_X(x) = 0$, for $n = 2, 3, \dots$

For $x \geq 0$, $\lim_{n \rightarrow \infty} F_{X_n}(x) = \lim_{n \rightarrow \infty} \left(1 - \left(1 - \frac{1}{n}\right)^{nx}\right) = \lim_{n \rightarrow \infty} 1 - e^{-x} = F_X(x), \quad \forall x$

$$\therefore X_n \xrightarrow{d} X$$



Limits

Central Limit Theorem

Let X_1, X_2, \dots, X_n be i.i.d r.v's with $E[X_i] = \mu < \infty$ and $0 < Var[X_i] = \sigma^2 < \infty$. Then, the r.v.

$$Z_n = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + X_2 + \dots + X_n - n\mu}{\sqrt{n}\sigma}$$

converges in distribution to the Standard Normal ($N(0,1)$) random variable, $\Phi(x)$ as $n \rightarrow \infty$, i.e.,

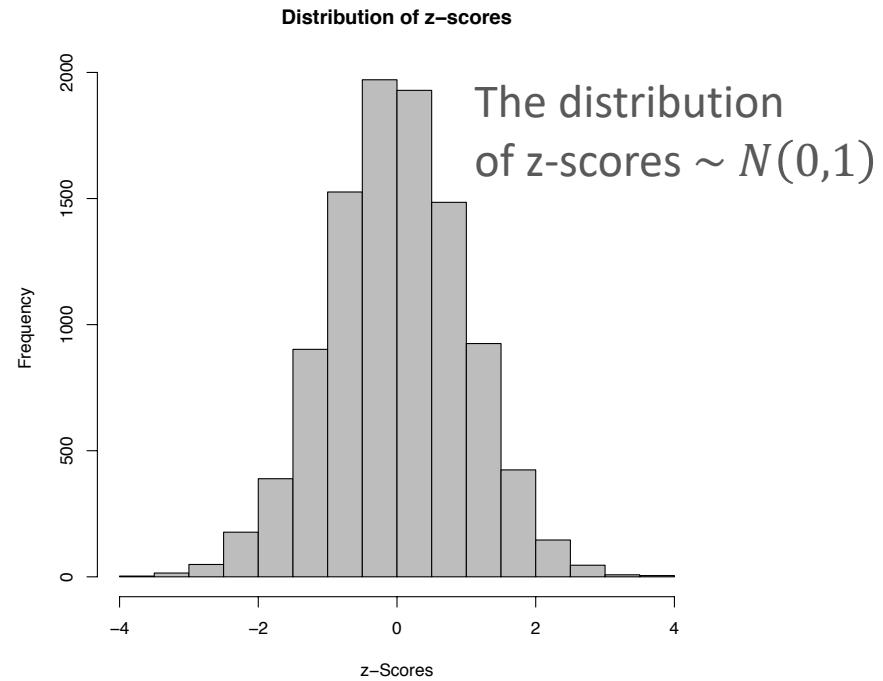
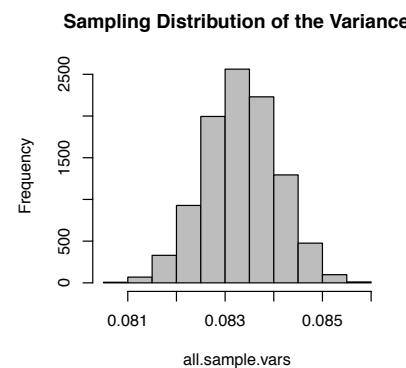
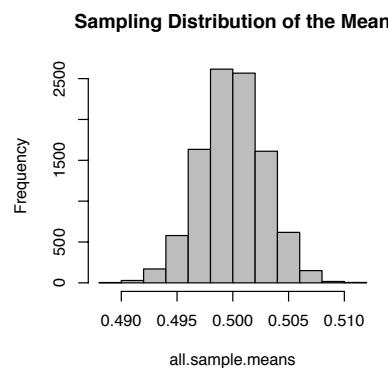
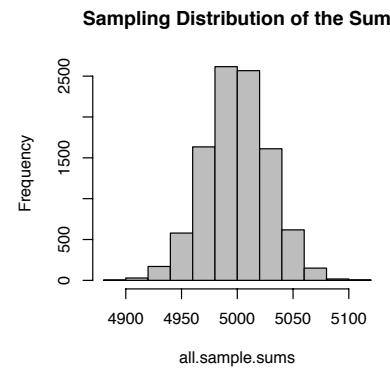
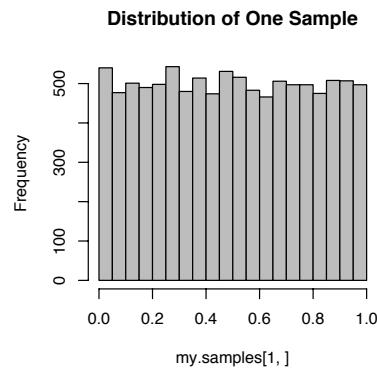
$$\lim_{n \rightarrow \infty} P(Z_n \leq x) = \Phi(x), \quad \forall x \in \mathbb{R}$$

Note: $\Phi(x)$ is the standard normal cdf.

Limits

Central Limit Theorem

Example: $n= 1000$ samples from a $\text{Unif}[0, 1]$



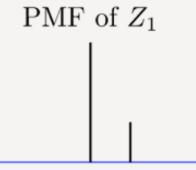
Limits

Assumptions:

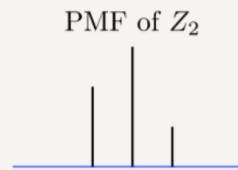
- $X_1, X_2 \dots$ are iid Bernoulli(p).
- $Z_n = \frac{X_1 + X_2 + \dots + X_n - np}{\sqrt{np(1-p)}}$.

We choose $p = \frac{1}{3}$.

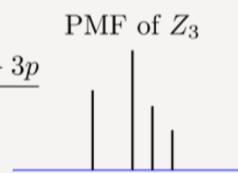
$$Z_1 = \frac{X_1 - p}{\sqrt{p(1-p)}}$$



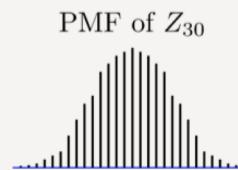
$$Z_2 = \frac{X_1 + X_2 - 2p}{\sqrt{2p(1-p)}}$$



$$Z_3 = \frac{X_1 + X_2 + X_3 - 3p}{\sqrt{3p(1-p)}}$$



$$Z_{30} = \frac{\sum_{i=1}^{30} X_i - 30p}{\sqrt{30p(1-p)}}$$



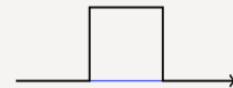
Limits

Assumptions:

- $X_1, X_2 \dots$ are iid Uniform(0,1).
- $Z_n = \frac{X_1 + X_2 + \dots + X_n - \frac{n}{2}}{\sqrt{\frac{n}{12}}}.$

$$Z_1 = \frac{X_1 - \frac{1}{2}}{\sqrt{\frac{1}{12}}}$$

PDF of Z_1



$$Z_2 = \frac{X_1 + X_2 - 1}{\sqrt{\frac{2}{12}}}$$

PDF of Z_2



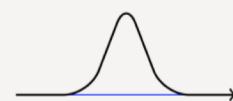
$$Z_3 = \frac{X_1 + X_2 + X_3 - \frac{3}{2}}{\sqrt{\frac{3}{12}}}$$

PDF of Z_3



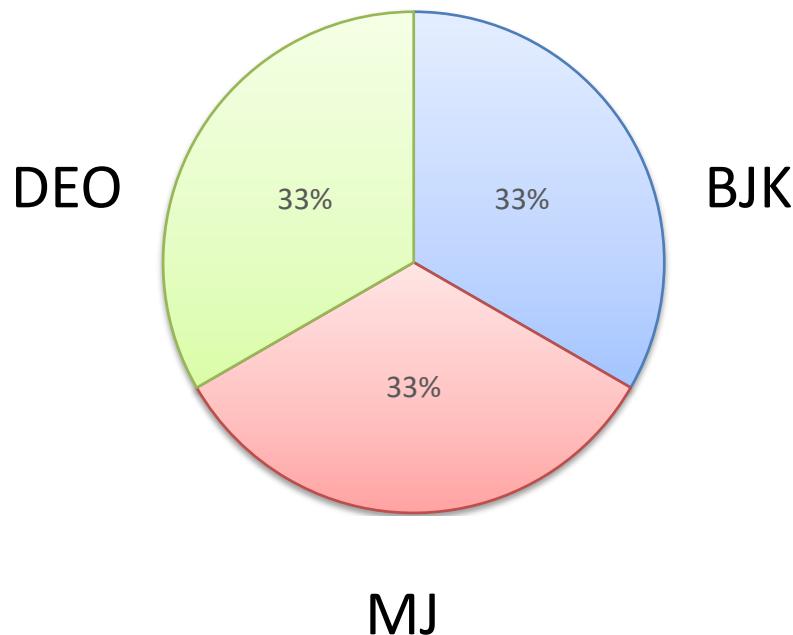
$$Z_{30} = \frac{\sum_{i=1}^{30} X_i - \frac{30}{2}}{\sqrt{\frac{30}{12}}}$$

PDF of Z_{30}



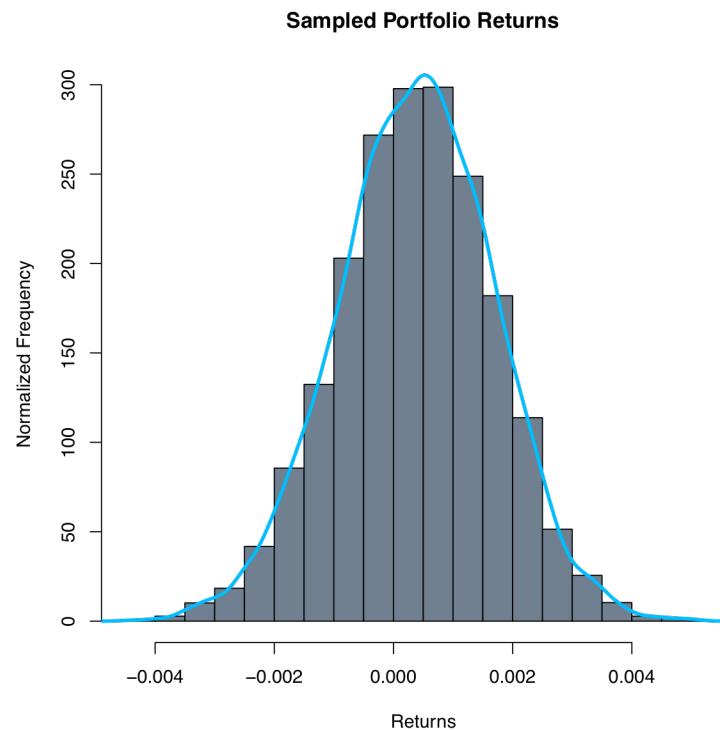
Limits

- **Application:** Investors of all types rely on the CLT to analyze stock returns, construct portfolios and manage risk.
- **Example:** Portfolio with 3 ETFs all equally weighted

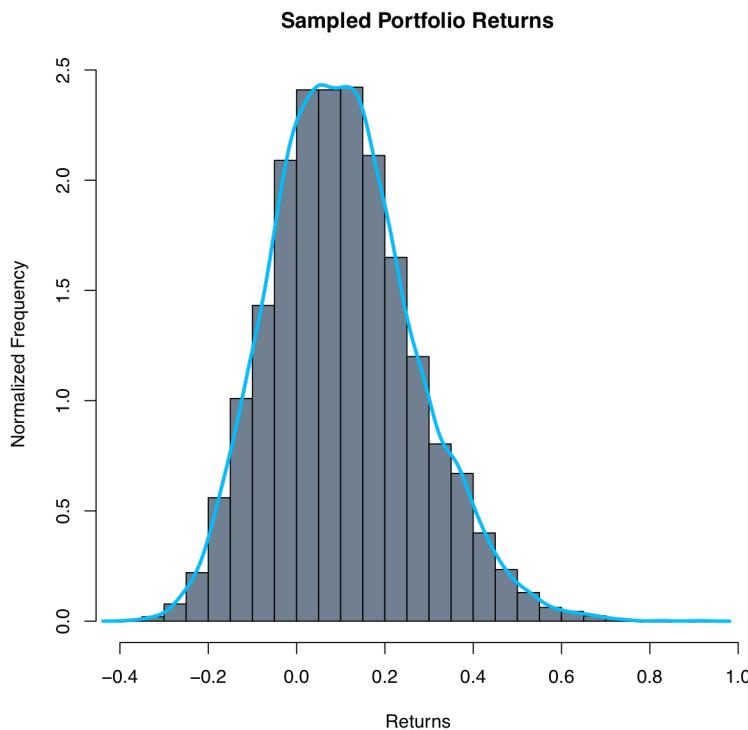


Limits

From Jan 1st, 2016 to September 1st, 2018



Daily



Annual

Note: See *Portfolio.R*

Monte Carlo Approximations

- Q: If X is a r.v where $X \sim f(X)$ then $g(X)$ is also a r.v., but what is the probability distribution of $g(X)$?

Example: Let $X \sim f(X) = \text{Unif}[1,4]$ and $Y = g(X) = (X - 3)^2$, find $E[Y]$.

Step 1: $p(X) = 1/4 \forall X$ values in $\{1,2,3,4\}$

Step 2: $g(X)$

$$X = 1 \rightarrow g(1) = 4$$

$$X = 2 \rightarrow g(2) = 1$$

$$X = 3 \rightarrow g(3) = 0$$

$$X = 4 \rightarrow g(4) = 1$$

Step 3: $P(g(X)) = P(Y)$

$$Y = 0, 1, 4$$

$$P(Y = 0) = 1/4, P(Y = 1) = 1/2, P(Y = 4) = 1/4$$

Step 4: $E[g(X)] = \sum Y P(Y)$
 $= 0(1/4) + 1(1/2) + 4(1/4)$
 $= 3/2$

Monte Carlo Approximations

- **Q:** What if we don't' know $P(Y)$? Do we need it?
- **Example:** Let $X \sim f(X) = \text{Unif}[1,4]$ and $Y = g(X) = (X - 3)^2$, find $E[Y]$.

$$\begin{aligned} E[Y] &= E[g(X)] = \sum Y P(Y) = \sum g(X) P_X(X) \\ &= (1-3)^2 P(X=1) + (2-3)^2 P(X=2) + (3-3)^2 P(X=3) + (4-3)^2 P(X=4) \\ &= 4(1/4) + 1(1/4) + 0(1/4) + 1(1/4) \\ &= 3/2 \rightarrow \text{same as before without knowing } P(Y) \smile \end{aligned}$$

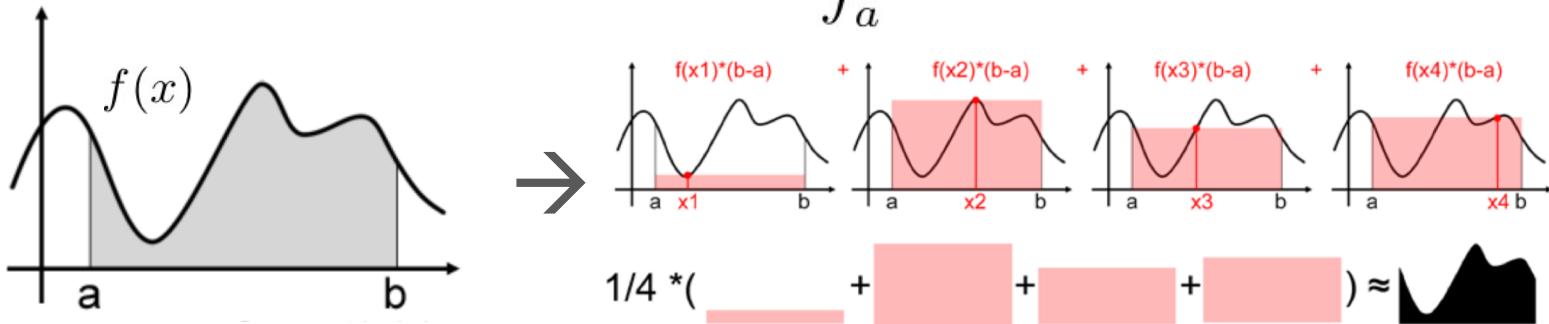
Conclusion: We do not need to know $P(Y)$ in order to compute $E[Y]$ provided we know $P(X)$.

= ***Foundation of MC Approximations!***

Monte Carlo Approximations

Application: Approximating Integrals

- How can we estimate: $F = \int_a^b f(x)dx$?



$$\langle F^N \rangle = (b - a) \frac{1}{N} \sum_{i=0}^{N-1} f(X_i)$$

Choose the points x at random between a and b ,
evaluate $f(x) \rightarrow$ compute the area $A = l \times w$

$$\rightarrow P(\lim_{x \rightarrow \infty} \langle F^N \rangle = F) = 1 \rightarrow E [\langle F^N \rangle] = F$$

(Converges in probability) (by the LLN)

Algorithmic Implementation of Sequential Bayesian Learning

Application: Approximating Integrals Generalization (arbitrary PDF)

Since $F = \int_D f(\vec{x}) d\vec{x} = \int_D \frac{f(\vec{x})}{p(\vec{x})} p(\vec{x}) d\vec{x}$ for any PDF p

but $\int_D \frac{f(\vec{x})}{p(\vec{x})} p(\vec{x}) d\vec{x} = E\left[\frac{f(\vec{x})}{p(\vec{x})}\right], x \sim p$

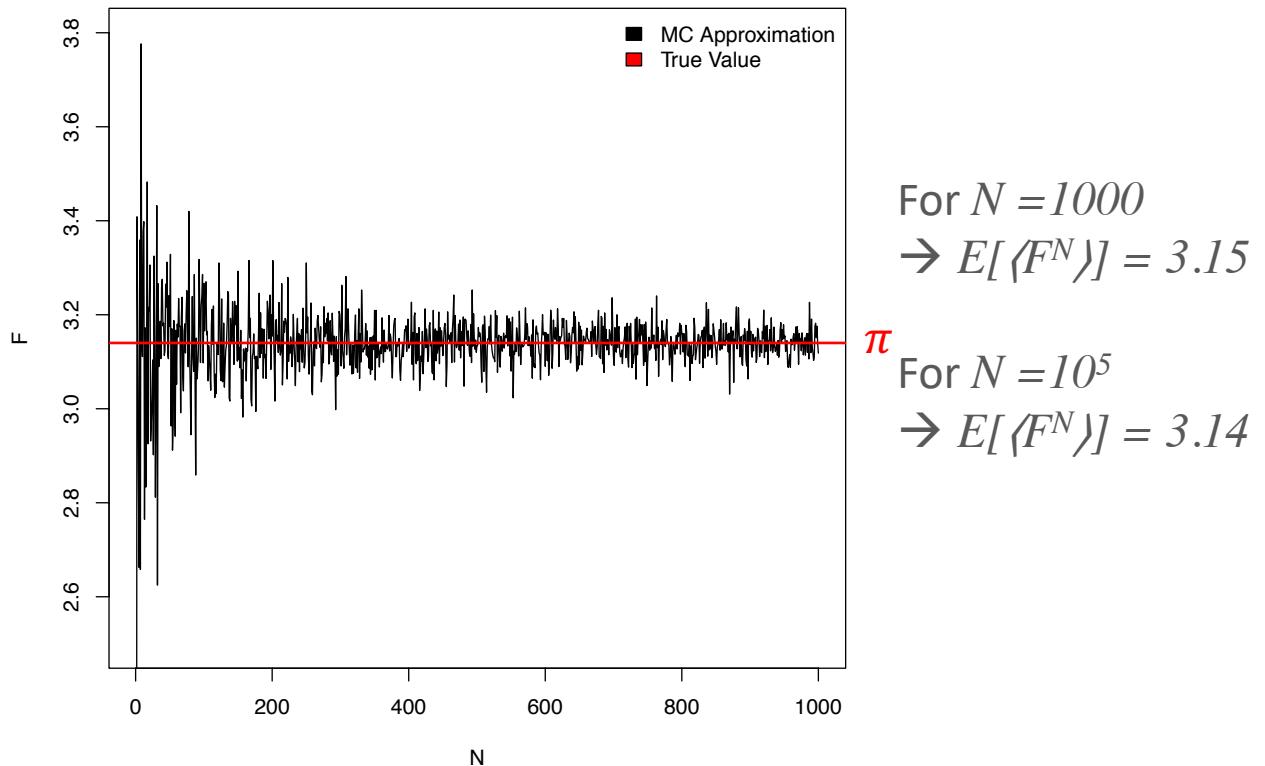
then $F_N = \frac{1}{N} \sum_{i=1}^N \frac{f(\vec{x}_i)}{p(\vec{x}_i)}$.

→ (by the LLN) as $N \rightarrow \infty, F_N \xrightarrow{a.s} E\left[\frac{f(\vec{x})}{p(\vec{x})}\right] = F$

Monte Carlo Approximations

- Example: Approximate the integral: $F = \int_0^1 4\sqrt{1 - x^2}dx = \pi$

Let X_1, X_2, \dots, X_n = i.i.d r.v.s. $\sim \text{unif}[0,1]$. $\rightarrow F^1 = f(X_1), \dots, F^N = f(X_N)$ are i.i.d. r.v.'s with mean $\langle F^N \rangle$ and $E[\langle F^N \rangle] = F = \pi$



Monte Carlo Approximations

- **Example:** Value at Risk
Given a current stock price (of your choice), what is the lowest it can go (worst-case scenario) 20 days from now with 99% probability of occurrence?
- **Model:** Geometric Brownian Motion $\rightarrow S_{t+1} = S_t \times (1 + \mu\Delta t + \sigma\varepsilon\sqrt{t})$
 - Standard model for stock price movements

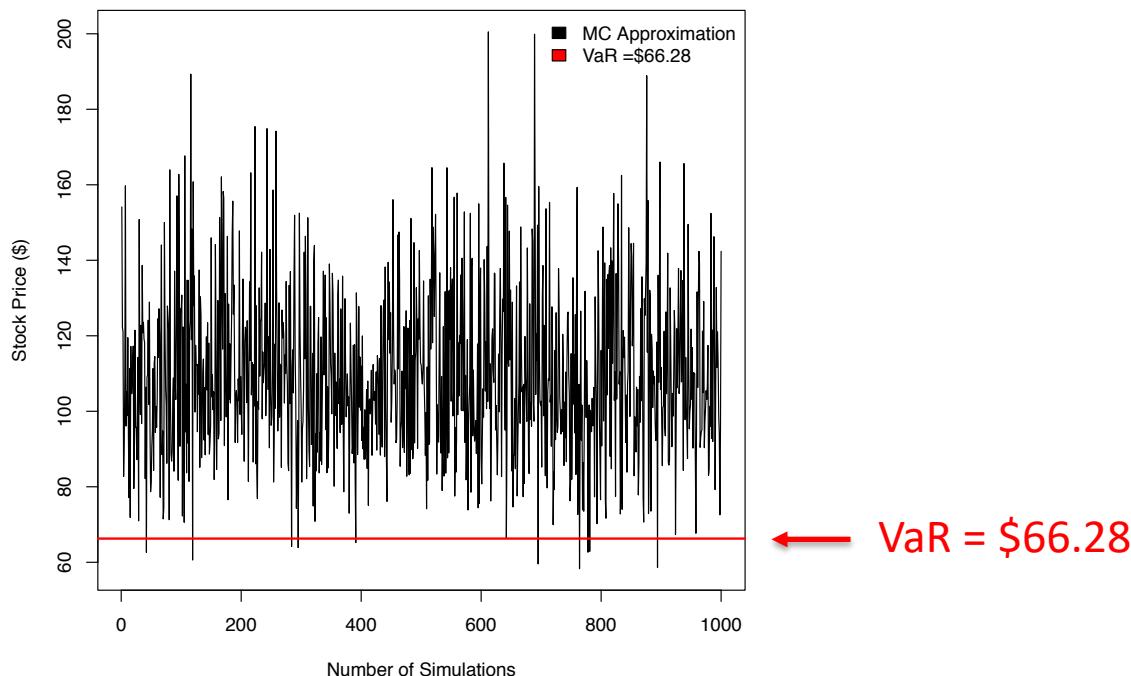
Assumptions:

$$S_0 = \$100$$

$$\sigma = 0.1 \text{ (10\% drift)}$$

$$\mu = 0.2 \text{ (20\% volatility)}$$

$$\varepsilon \sim \text{unif}[0,1]$$



Normal Distribution Theory

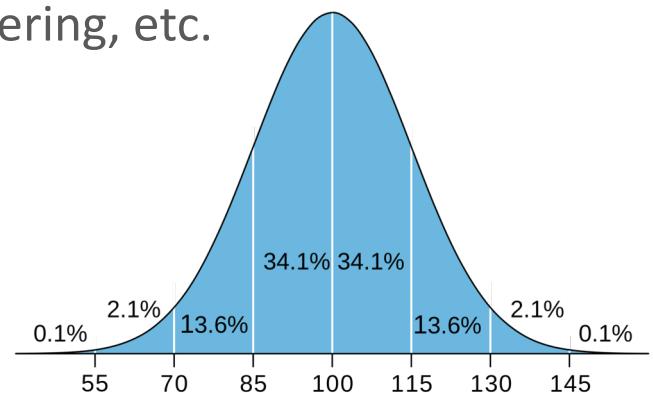
Normal Distribution $N(\mu, \sigma)$

- Suppose $X_i \sim N(\mu, \sigma_i^2)$, for $i = 1, 2, \dots, n$ and that they are independent r.v's. Let $Y = (\sum_i a_i X_i) + b$ for some constants $\{a_i\}$ and b . Then

$$Y \sim N\left(\sum_i a_i \mu_i + b, \sum_i a_i^2 \sigma_i^2\right).$$

Applications: finance models (e.g., mean variance analysis), business, economics, science, engineering, etc.

PDF: $f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}(\frac{x-\mu}{\sigma})^2}, -\infty < x < +\infty$

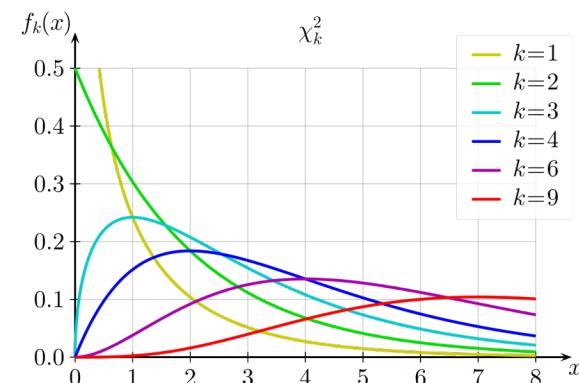


Normal Distribution Theory

χ^2 -Distribution

- The *chi-squared* distribution with ν degrees of freedom (or $\chi^2(\nu)$) is the distribution of the random variable $Z = X_1^2 + X_2^2 + \dots + X_\nu^2$, where X_1, \dots, X_n are i.i.d., each with the standard normal distribution $N(0,1)$.
- Applications:** population variance estimation, goodness-of-fit tests, contingency tables, etc.

- PDF: $f(x) = \frac{1}{2^{\nu/2}\Gamma(\nu/2)}x^{(\nu/2)-1}e^{-x/2}, x > 0$,
where, $\mu = \nu, \sigma^2 = 2\nu, M(t) = (1 - 2t)^{-\nu/2}$



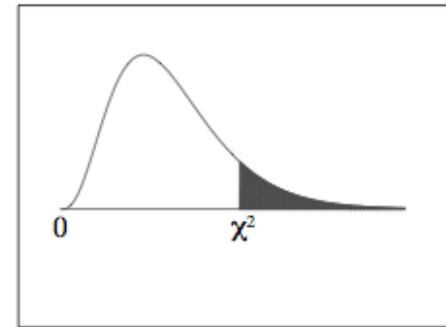
χ^2 -Distribution

Examples

- Find the value of χ^2_5 such that the area on the right tail is equal to 0.05.

Solution: $df = 5$, and $\alpha = 0.05$

$$\rightarrow \chi^2_5 = 11.070$$



The shaded area is equal to α for $\chi^2 = \chi^2_\alpha$.

df	$\chi^2_{.995}$	$\chi^2_{.990}$	$\chi^2_{.975}$	$\chi^2_{.950}$	$\chi^2_{.900}$	$\chi^2_{.100}$	$\chi^2_{.050}$	$\chi^2_{.025}$	$\chi^2_{.010}$	$\chi^2_{.005}$
1	0.000	0.000	0.001	0.004	0.016	2.706	3.841	5.024	6.635	7.879
2	0.010	0.020	0.051	0.103	0.211	4.605	5.991	7.378	9.210	10.597
3	0.072	0.115	0.216	0.352	0.584	6.251	7.815	9.348	11.345	12.838
4	0.207	0.297	0.484	0.711	1.064	7.779	9.488	11.143	13.277	14.860
5	0.412	0.554	0.831	1.145	1.610	9.236	11.070	12.833	15.086	16.750

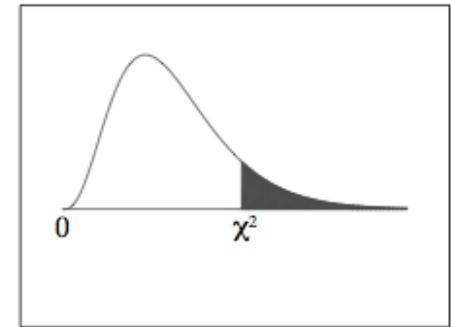
χ^2 -Distribution

Examples

2. Find the value of χ^2_3 that represents the 97.5th percentile.

Solution: $df = 3$, and $\alpha = 0.025$

$$\rightarrow \chi^2_3 = 9.348$$



The shaded area is equal to α for $\chi^2 = \chi^2_\alpha$.

df	$\chi^2_{.995}$	$\chi^2_{.990}$	$\chi^2_{.975}$	$\chi^2_{.950}$	$\chi^2_{.900}$	$\chi^2_{.100}$	$\chi^2_{.050}$	$\chi^2_{.025}$	$\chi^2_{.010}$	$\chi^2_{.005}$
1	0.000	0.000	0.001	0.004	0.016	2.706	3.841	5.024	6.635	7.879
2	0.010	0.020	0.051	0.103	0.211	4.605	5.991	7.378	9.210	10.597
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5	0.412	0.554	0.831	1.145	1.610	9.236	11.070	12.833	15.086	16.750

χ^2 -Distribution

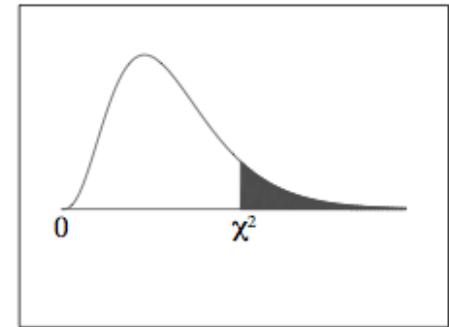
Examples

3. Given $X \sim \chi^2_4$, find $P(0.484 < X < 7.779)$.

Solution: $df = 4$.

$$\rightarrow P(0.484 < X < 7.779) = 0.975 - 0.100$$

$$= 0.874$$



The shaded area is equal to α for $\chi^2 = \chi^2_\alpha$.

df	$\chi^2_{.995}$	$\chi^2_{.990}$	$\chi^2_{.975}$	$\chi^2_{.950}$	$\chi^2_{.900}$	$\chi^2_{.100}$	$\chi^2_{.050}$	$\chi^2_{.025}$	$\chi^2_{.010}$	$\chi^2_{.005}$
1	0.000	0.000	0.001	0.004	0.016	2.706	3.841	5.024	6.635	7.879
2	0.010	0.020	0.051	0.103	0.211	4.605	5.991	7.378	9.210	10.597
3	0.072	0.115	0.216	0.352	0.584	6.251	7.815	9.348	11.345	12.838
4	0.207	0.297	0.484	0.711	1.064	7.779	9.488	11.143	13.277	14.860
5	0.412	0.554	0.831	1.145	1.610	9.236	11.070	12.833	15.086	16.750

Normal Distribution Theory

t-Distribution

- The *t-distribution* with n degrees of freedom (or Student-t), is the distribution of the random variable

$$Z = \frac{X}{\sqrt{(X_1^2 + \cdots + X_n^2)/n}}$$

where X_1, X_2, \dots, X_n are i.i.d., each with the standard normal distribution $N(0,1)$. Equivalently $Z = \frac{X}{\sqrt{Y/n}}$, where $Y \sim \chi^2(n)$.

Normal Distribution Theory

t-Distribution

- The *t-distribution* with n degrees of freedom (or Student-t), is the distribution of the random variable: $Z = \frac{X}{\sqrt{(X_1^2 + \dots + X_n^2)/n}}$

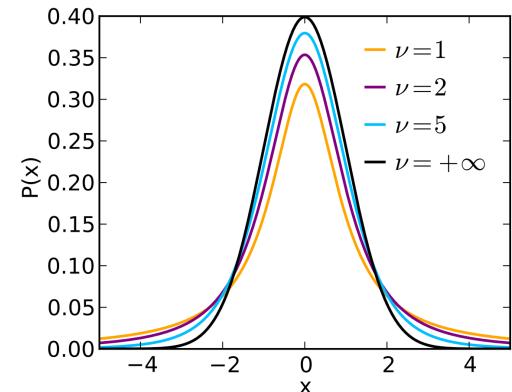
where X_1, X_2, \dots, X_n are i.i.d., each with the standard normal distribution $N(0,1)$. Equivalently $Z = \frac{X}{\sqrt{Y/n}}$, where $Y \sim \chi^2(\nu)$.

Applications: small sample sizes ($n \lesssim 30$), and/or limited information

PDF: $f(t) = \frac{\Gamma(\frac{\nu+1}{2})}{\sqrt{2\pi}\Gamma(\frac{\nu}{2})} \left(1 + \frac{t^2}{\nu}\right)^{-(\nu+1)/2}, -\infty < t < +\infty,$
where, $\mu = 0, \sigma^2 = \frac{\nu}{\nu-2}, \nu > 2$

t-Distribution

Examples



- If T follows a t_4 distribution, find t_0 such that:
 $P(|T| < t_0) = 0.90.$

Solution: df = 4, and $|T| < t_0 \rightarrow -t_0 < T < t_0$

(Area = 90%) $\rightarrow t_0 = 2.132$

t Table

cum. prob	$t_{.50}$	$t_{.75}$	$t_{.80}$	$t_{.85}$	$t_{.90}$	$t_{.95}$	$t_{.975}$	$t_{.99}$	$t_{.995}$	$t_{.999}$	$t_{.9995}$
one-tail	0.50	0.25	0.20	0.15	0.10	0.05	0.025	0.01	0.005	0.001	0.0005
two-tails	1.00	0.50	0.40	0.30	0.20	0.10	0.05	0.02	0.01	0.002	0.001
df											
1	0.000	1.000	1.376	1.963	3.078	6.314	12.71	31.82	63.66	318.31	636.62
2	0.000	0.816	1.061	1.386	1.886	2.920	4.303	6.965	9.925	22.327	31.599
3	0.000	0.765	0.978	1.250	1.638	2.353	3.182	4.541	5.841	10.215	12.924
4	0.000	0.741	0.941	1.190	1.533	2.132	2.776	3.747	4.604	7.173	8.610
5	0.000	0.727	0.920	1.156	1.476	2.015	2.571	3.365	4.032	5.893	6.869

t-Distribution

Examples

2. If T follows a t_2 distribution, find t_0 such that:
 $P(T < t_0) = 0.80.$

Solution: df = 2, and $T < t_0 \rightarrow$ Area = 80%

$$\rightarrow t_0 = 1.061$$

t Table

cum. prob one-tail	$t_{.50}$	$t_{.75}$	$t_{.80}$	$t_{.85}$	$t_{.90}$	$t_{.95}$	$t_{.975}$	$t_{.99}$	$t_{.995}$	$t_{.999}$	$t_{.9995}$
two-tails	0.50	0.25	0.20	0.15	0.10	0.05	0.025	0.01	0.005	0.001	0.0005
df											
1	0.000	1.000	1.376	1.963	3.078	6.314	12.71	31.82	63.66	318.31	636.62
2	0.000	0.816	1.061	1.386	1.886	2.920	4.303	6.965	9.925	22.327	31.599
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5	0.000	0.727	0.920	1.156	1.476	2.015	2.571	3.365	4.032	5.893	6.869

t-Distribution

Examples

3. Determine the t-critical value such that the lower-tail area =0.025 for a t_3 distribution.

Solution: df = 3, and $T < t_0 \rightarrow \text{Area} = 2.5\%$

$$\rightarrow t_0 = -3.182$$

***t* Table**

cum. prob one-tail	$t_{.50}$	$t_{.75}$	$t_{.80}$	$t_{.85}$	$t_{.90}$	$t_{.95}$	$t_{.975}$	$t_{.99}$	$t_{.995}$	$t_{.999}$	$t_{.9995}$
two-tails	0.50	0.25	0.20	0.15	0.10	0.05	0.025	0.01	0.005	0.001	0.0005
df											
1	0.000	1.000	1.376	1.963	3.078	6.314	12.71	31.82	63.66	318.31	636.62
2	0.000	0.816	1.061	1.386	1.886	2.920	4.303	6.965	9.925	22.327	31.599
3	0.000	0.765	0.978	1.250	1.638	2.353	3.182	4.541	5.841	10.215	12.924
4	0.000	0.741	0.941	1.190	1.533	2.132	2.776	3.747	4.604	7.173	8.610
5	0.000	0.727	0.920	1.156	1.476	2.015	2.571	3.365	4.032	5.893	6.869

Normal Distribution Theory

F-Distribution

- The *F* distribution with ν_1 and ν_2 degrees of freedom (or F_{ν_1, ν_2}) is the distribution of the random variable

$$Z = \frac{(X_1^2 + \cdots + X_{\nu_1}^2) / \nu_1}{(Y_1^2 + \cdots + Y_{\nu_2}^2) / \nu_2}$$

where $X_1, \dots, X_{\nu_1}, Y_1, \dots, Y_{\nu_2}$ are i.i.d., each with the standard normal distribution (i.e., $Z = \frac{X/\nu_1}{Y/\nu_2}$, where $X \sim \chi^2(\nu_1)$, $Y \sim \chi^2(\nu_2)$.)

- Applications:** population variance comparison (ANOVA), likelihood ratio tests, comparing statistical models, etc.

- PDF: $f(x) = \frac{\Gamma(\frac{\nu_1+\nu_2}{2})}{\Gamma(\frac{\nu_1}{2})\Gamma(\frac{\nu_2}{2})} \nu_1^{\nu_1/2} \nu_2^{\nu_2/2} x^{(\nu_1/2)-1} (\nu_2 + \nu_1 x)^{-(\nu_1+\nu_2)/2}$, $x \geq 0$
where, $\mu = \frac{\nu_2}{\nu_2 - 2}$

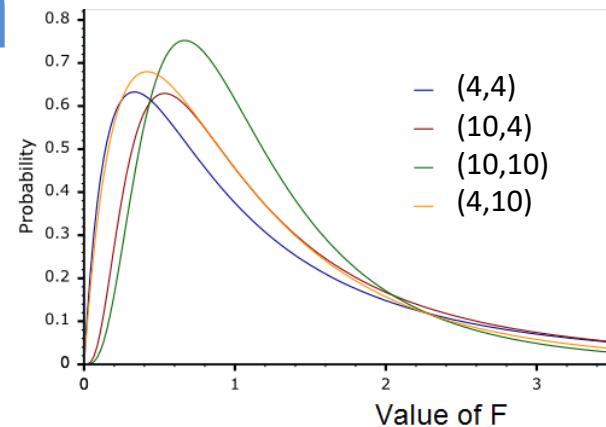
F-Distribution

v_2

v_1

$1-\alpha$

$$F_{\alpha, v_1, v_2}$$



F Distribution Table (Percentiles)

		1	2	3	4	5	6	7	8	9	10
0.95	1	161.5	199.5	215.7	224.6	230.2	234.0	236.8	238.9	240.5	241.9
0.975		647.8	799.5	864.2	899.6	921.8	937.1	948.2	956.7	963.3	968.7
0.99		4052.2	4999.5	5403.4	5624.6	5763.7	5859.0	5928.4	5981.07	6022.4	6055.9
0.95	2	18.51	19.00	19.16	19.25	19.30	19.33	19.35	19.37	19.38	19.40
0.975		38.51	39.00	39.17	39.25	39.30	39.33	39.36	39.37	39.39	39.40
0.99		98.50	99.00	99.17	99.25	99.30	99.33	99.36	99.37	99.39	99.40
0.95	3	10.13	9.55	9.28	9.12	9.01	8.94	8.89	8.85	8.81	8.79
0.975		17.44	16.04	15.44	15.10	14.88	14.73	14.62	14.54	14.47	14.42
0.99		34.12	30.82	29.46	28.71	28.24	27.91	27.67	27.49	27.35	27.23

F-Distribution

Examples

- Determine the value of $F_{0.05, 3, 2}$

Solution: $\alpha = 0.05$, $1 - \alpha = 0.95$, $v_1 = 3$, $v_2 = 2$

$$\rightarrow F_{0.05, 3, 2} = 19.16$$

		1	2	3	4	5	6	7	8	9	10
0.95	1	161.5	199.5	215.7	224.6	230.2	234.0	236.8	238.9	240.5	241.9
0.975		647.8	799.5	864.2	899.6	921.8	937.1	948.2	956.7	963.3	968.7
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0.975		38.51	39.00	39.17	39.25	39.30	39.33	39.36	39.37	39.39	39.40
0.99		98.50	99.00	99.17	99.25	99.30	99.33	99.36	99.37	99.39	99.40
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0.975		17.44	16.04	15.44	15.10	14.88	14.73	14.62	14.54	14.47	14.42
0.99		34.12	30.82	29.46	28.71	28.24	27.91	27.67	27.49	27.35	27.23

F-Distribution

Examples

2. Determine the value of $F_{0.95, 3, 2}$

$$F_{1-\alpha, \nu_1, \nu_2} = \frac{1}{F_{\alpha, \nu_2, \nu_1}}$$

Solution: $\alpha = 0.95, \nu_1 = 3, \nu_2 = 2$

$$\rightarrow F_{0.95, 3, 2} = 1/F_{0.05, 2, 3} = 1/9.55 = 0.1047$$

		1	2	3	4	5	6	7	8	9	10
0.95	1	161.5	199.5	215.7	224.6	230.2	234.0	236.8	238.9	240.5	241.9
0.975		647.8	799.5	864.2	899.6	921.8	937.1	948.2	956.7	963.3	968.7
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0.99		98.50	99.00	99.17	99.25	99.30	99.33	99.36	99.37	99.39	99.40
0.95	3	10.13	9.55	9.28	9.12	9.01	8.94	8.89	8.85	8.81	8.79
0.975		17.44	16.04	15.44	15.10	14.88	14.73	14.62	14.54	14.47	14.42
0.99		34.12	30.82	29.46	28.71	28.24	27.91	27.67	27.49	27.35	27.23

F-Distribution

Examples

$$F_{1-\alpha, \nu_1, \nu_2} = \frac{1}{F_{\alpha, \nu_2, \nu_1}}$$

3. Given that $X \sim F_{3,2}$, find $P(X < 0.0623)$

Solution: $\alpha = ?$, $\nu_1 = 3$, $\nu_2 = 2$, $1/0.0623 = 16.04$

$\rightarrow P(X < 0.0623) = P(1/X > 1/0.0623) = P(1/X > 16.04) = 0.025$

		1	2	3	4	5	6	7	8	9	10
0.95	1	161.5	199.5	215.7	224.6	230.2	234.0	236.8	238.9	240.5	241.9
0.975		647.8	799.5	864.2	899.6	921.8	937.1	948.2	956.7	963.3	968.7
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0.99		98.50	99.00	99.17	99.25	99.30	99.33	99.36	99.37	99.39	99.40
0.95	3	10.13	9.55	9.28	9.12	9.01	8.94	8.89	8.85	8.81	8.79
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