

Exchangability

Y_1, \dots, Y_N ^{are} ~~is~~ exchangeable random variables if $\pi(y_1, \dots, y_N) = \pi(y_{\sigma_1}, \dots, y_{\sigma_N})$ for any permutation σ of $\{1, \dots, N\}$. (Def 1.8).

Thm 1.1 (de Finetti) If a sequence of random variables (Y_1, \dots, Y_N) is exchangeable, then its joint distribution can be written as.

$$\pi(y_1, \dots, y_N) = \int \left\{ \prod_{i=1}^N \pi(y_i | \theta) \right\} \pi(\theta) d\theta.$$

for some parameter θ and, some distribution on θ , and some sampling model $\pi(y_i | \theta)$.

Questions.

1. What does exchangeability mean?

Labels of random variables do not tell us anything about the outcomes.

2. What does de Finetti's theorem tell us about Bayesian inference?

It is a kind of existence theorem for Bayesian inference.

It says if we have exchangeable random variables, then a parameter θ must exist and ~~we can put~~ a subjective distribution must also exist for θ .

3. Let $X_1, \dots, X_N \sim \text{Ber}(p)$ be independent RVs. Show they are exchangeable.

We have ~~that~~ $\pi(x_i | p) = p^{x_i} (1-p)^{1-x_i}$. Let σ be a permutation of $\{1, \dots, N\}$. The joint distribution of x_1, \dots, x_N

$$\text{is } \pi(x_1, \dots, x_N) = \prod_{i=1}^N p^{x_i} (1-p)^{1-x_i}$$

$$= p^{\sum_{i=1}^N x_i} (1-p)^{\sum_{i=1}^N (1-x_i)}$$

$$= p^{\sum_{i=1}^N x_i} (1-p)^{N - \sum_{i=1}^N x_i}$$

$$= p^{\sum_{i \in \sigma} x_i} (1-p)^{N - \sum_{i \in \sigma} x_i}$$

(as addition is invariant to permutations).

$$= p^{\sum_{i \in \sigma} x_i} (1-p)^{\sum_{i \in \sigma} (1-x_i)}$$

$$= \prod_{i \in \sigma} p^{x_i} (1-p)^{1-x_i}$$

$$= \pi(x_{\sigma_1}, \dots, x_{\sigma_N})$$

□

Priors Let $x_1, \dots, x_n \sim \text{Bin}(M, p)$ How could we perform inference on p ? MLE. Conjugate prior. Non-informative prior. Prior elicitation.

~~Conjugate~~ Invariant prior distribution.

1. Show $p \sim \text{Beta}(\alpha, \beta)$ is conjugate.

By Bayes' theorem $\pi(p|\underline{x}) \propto \pi(\underline{x}|p)\pi(p)$

Likelihood $\pi(\underline{x}|p) \propto \prod_{i=1}^n \binom{M}{x_i} p^{x_i} (1-p)^{M-x_i}$

$$\propto p^{\sum x_i} (1-p)^{NM - \sum x_i}$$

Prior: $\pi(p) \propto p^{\alpha-1} (1-p)^{\beta-1}$

$$\Rightarrow \text{Posterior: } \pi(p|\underline{x}) \propto p^{\sum x_i} (1-p)^{NM - \sum x_i} p^{\alpha-1} (1-p)^{\beta-1}$$

$$\propto p^{\sum x_i + \alpha - 1} (1-p)^{NM - \sum x_i + \beta - 1}$$

$$\Rightarrow p|\underline{x} \sim \text{Beta}(\sum x_i + \alpha, NM - \sum x_i + \beta)$$

As both prior and posterior distributions have same functional form.

2. Derive the posterior distribution for p using a non-informative prior distribution.

A non-informative prior distribution is $p \sim \text{U}(0,1)$. $\pi(p) = \begin{cases} 1 & p \in [0,1] \\ 0 & \text{otherwise} \end{cases}$

$$\pi(p|\underline{x}) \propto p^{\sum x_i} (1-p)^{NM - \sum x_i}$$

$$\sim \text{Beta}(\sum x_i + 1, NM - \sum x_i + 1)$$

3. Derive an invariant prior distribution.

By Jeffrey's theorem $\pi(\theta) \propto \sqrt{I_\theta(y)}$ is an invariant prior distribution.

We have $I_\theta(y) = -E\left[\frac{d^2}{d\theta^2} \log \pi(y|\theta)\right]$. For the binomial distribution, we have.

$$\log \pi(\underline{x}|p) \propto \sum x_i \log p + (NM - \sum x_i) \log(1-p)$$

$$\Rightarrow \frac{d}{dp} \log \pi(\underline{x}|p) \propto \frac{\sum x_i}{p} - \frac{NM - \sum x_i}{1-p}$$

$$\Rightarrow \frac{d^2}{dp^2} \log \pi(\underline{x}|p) \propto -\frac{\sum x_i}{p^2} + \frac{NM - \sum x_i}{(1-p)^2}$$

$$\Rightarrow E\left(\frac{d^2 \log \pi(\underline{x}|p)}{dp^2}\right) = -\frac{E(\sum x_i)}{p^2} + \frac{NM - E(\sum x_i)}{(1-p)^2}$$

$$= -\frac{\sum E(x_i)}{p^2} + \frac{NM - \sum E(x_i)}{(1-p)^2}$$

$$= -\frac{NMp}{p^2} + \frac{NM - NMp}{(1-p)^2}$$

$$= NM \left(-\frac{1}{p} + \frac{1}{1-p} \right)$$

$$= \frac{NM}{p(1-p)}$$

$$\Rightarrow \pi(\theta) \propto p^{-1/2} (1-p)^{-1/2}$$