# Coordinate Descent and Coordinate Gradient Descent

It is generally believed that coordinate descent should perform better than first-order methods. We have our minimization problem  $\min_{\mathbf{x}} f(\mathbf{x})$  where f is convex.

We can use coordinate descent. This algorithm is also called 'coordinatewise minimization'.

let the initial value be  $\mathbf{x}^{(0)} \in \mathbb{R}^n$ , and repeat:  $x_i^{(k)} = \operatorname{argmin}_{x_i} f(x_1^{(k)}, \dots, x_{i-1}^{(k)}, x_i, x_{i+1}^{(k-1)}, \dots, x_n^{(k)})$ , for  $i = 1, \dots, n$  and  $k = 1, 2, 3, \dots$  Essentially, we minimize f with respect to one element  $x_i$ , plug it back in f, and move to the next index.

Some guidelines for coordinate descent:

- The order of cycle through coordinates is arbitrary, we can use any permutation of  $\{1, \ldots, n\}$ ;
- we can replace everywhere individual coordinates with blocks of coordinates (Just remember that this is possible if you see some such step in the future.)
- The "one-at-a-time" update scheme is critical, and "all-at-once" scheme does not necessarily converge;

Why is it used?

- Very simple and easy to implement.
- Careful implementations can achieve state-of-the-art solution.
- Scalable, e.g., don't need to keep full data in memory

## 1 Example: linear regression

Given  $\mathbf{y} \in \mathbb{R}^n$  and  $\mathbf{X}$  is  $n \times p$  with columns  $X_1, \dots, X_p$ , consider the linear regression problem:  $\min_{\boldsymbol{\beta}} \frac{1}{2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2$ .

Minimizing over  $\beta_i$  with all  $\beta_j, j \neq i$  fixed:  $0 = \frac{\partial f(\boldsymbol{\beta})}{\partial \beta_i} = X_i^T(\mathbf{X}\boldsymbol{\beta} - \mathbf{y}) = X_i^T(X_i\beta_i + \mathbf{X}_{-i}\boldsymbol{\beta}_{-i} - \mathbf{y})$  where  $\mathbf{X}_{-i}$  and  $\boldsymbol{\beta}_{-i}$  are original matrix or vector with i-th column or element removed, respectively.

Thus

$$\beta_i = \frac{X_i^T (\mathbf{y} - \mathbf{X}_{-i} \boldsymbol{\beta}_{-i})}{X_i^T X_i}.$$
 (1)

Like gradient descent, we will update  $\boldsymbol{\beta}$  sequentially. At k+1-th iteration, the update for  $\boldsymbol{\beta}$  be  $\boldsymbol{\beta}^{(k+1)}$ . Then according to the gradient descent algorithm  $\boldsymbol{\beta}^{(k+1)} = \boldsymbol{\beta}^{(k)} - L_n^2 \mathbf{X}^T (\mathbf{y} - \mathbf{X} \boldsymbol{\beta}^{(k)}) = \boldsymbol{\beta}^{(k)} - L_n^2 \mathbf{X}^T (\mathbf{r}^k)$ , where  $\mathbf{r}_k = \mathbf{y} - \mathbf{X} \boldsymbol{\beta}^{(k)}$  is the residual error vector.

In case of coordinate descent,  $\beta_i^{k+1} = \frac{X_i^T(\mathbf{y} - \mathbf{X}_{-i}\boldsymbol{\beta}_{-i} - \mathbf{X}_i\boldsymbol{\beta}_i^{(k)} + \mathbf{X}_i\boldsymbol{\beta}_i^{(k)})}{X_i^TX_i} = \beta_i^{(k)} + \frac{X_i^T\mathbf{r}^{(k)}}{X_i^TX_i}$ . Thus, the two updates look very similar. The only difference is in the step size. For gradient-descent, it is fixed at L/n. However, here, it is coordinate specific,  $1/X_i^TX_i$ .

Coordinate descent repeats this update for i = 1, 2, ..., p, Note that the computational cost for one cycle of coordinate descent is O(np) where O(n) to compute  $X_i^T(\mathbf{y} - \mathbf{X}_{-i}\boldsymbol{\beta}_{-i})$  for each update in a cycle (it is O(n) because we can precompute  $X_i^T X_i \beta_i$ ), which is the same as gradient descent.

```
x <- rnorm(1000)
z <- rnorm(1000)
w <- rnorm(1000)

y <- rnorm(1000, 2*x+3*z)

#We fit the model y=beta1*x+beta2*z+beta3*w + error</pre>
```

Apply coordinate descent in the above problem to estimate beta1, beta2 and beta3. Use (1) to get the updates of components of beta.

```
#Initialize
beta <- rep(0, 3)

#Design matrix
X = cbind(x, z, w)

for(itr in 1:10000){
   beta[1] <- t(X[,1])%*%(y-X[,-1]%*%beta[-1])/(t(X[,1])%*%X[,1])
   beta[2] <- t(X[,2])%*%(y-X[,-2]%*%beta[-2])/(t(X[,2])%*%X[,2])
   beta[3] <- t(X[,3])%*%(y-X[,-3]%*%beta[-3])/(t(X[,3])%*%X[,3])
}</pre>
```

#Least square estimate

betals <- solve(crossprod(X))%\*%crossprod(X, y)</pre>

#Compare with least square estimate
sum((beta-betals)^2)

### 2 Pathwise Coordinate Descent for Lasso

The objective function is given by,

$$\hat{\boldsymbol{\beta}} = \arg\min_{\boldsymbol{\beta}} (\frac{1}{n} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_{2}^{2} + \lambda \|\boldsymbol{\beta}\|_{1}),$$

where the part  $\lambda \|\boldsymbol{\beta}\|_1$  is called the LASSO penalty [2]. Here  $\|\boldsymbol{\beta}\|_1 = \sum_{j=1}^p |\beta_j|$  where  $\boldsymbol{\beta} = (\beta_1, \dots, \beta_p)$ .

If we minimize above expression with respect to  $\beta_i$ , keeping everything else fixed, the minimizer would be  $\hat{\beta}_i$ ?

First define,  $\beta_i^* = \frac{X_i^T(\mathbf{y} - \mathbf{X}_{-i}\boldsymbol{\beta}_{-i})}{X_i^T X_i}$ , the coordinate descent update in the least square case.

$$\hat{\beta}_{i} = \begin{cases} \beta_{i}^{*} + \frac{n}{2}\lambda & \text{if } \beta_{i}^{*} > 0, \frac{n}{2}\lambda \leq |\beta_{i}^{*}| \\ \beta_{i}^{*} - \frac{n}{2}\lambda & \text{if } \beta_{i}^{*} < 0, \frac{n}{2}\lambda \leq |\beta_{i}^{*}| \\ 0 & \text{if } \frac{n}{2}\lambda > |\beta_{i}^{*}|, \end{cases}$$

This above estimate help us to get the following way to get 'solution-path' of LASSO. Specifically, [1] proposed this the pathwise-coordinate descent method for Lasso problem. The algorithm runs over two loops. Outer Loop (pathwise strategy): • Compute the solution over a sequence  $\lambda_1 > \lambda_2 > \cdots > \lambda_r$  of tuning parameter values • For tuning parameter value  $\lambda_k$ , initialize coordinate descent algorithm at the computed solution for  $\lambda_{k+1}$  for warm start.

#### 3 Coordinate Gradient Descent

For a smooth function f, the iterations  $x_i^{(k)} = x_i^{(k-1)} - t_{ki} \nabla f(x_1^{(k)}, \dots, x_{i-1}^{(k)}, x_i, x_{i+1}^{(k-1)}, \dots, x_n^{(k-1)})$ ,  $i = 1, \dots, n$  for  $k = 1, 2, 3 \dots$  are called coordinate gradient descent.

## References

- [1] Jerome Friedman, Trevor Hastie, and Rob Tibshirani. Regularization paths for generalized linear models via coordinate descent. *Journal of statistical software*, 33(1):1, 2010.
- [2] R. Tibshirani. Regression shrinkage and selection via the lasso. *Journal of the Royal Statistical Society B*, 58:267–288, 1996.