



Model Trees for Identifying Exceptional Players in the NHL Draft

Paper Track: Other Sports
Paper ID: 5724

Abstract

Drafting strong players is crucial for a team's success. We describe a new data-driven interpretable approach for assessing draft prospects in the National Hockey League. Successful previous approaches have 1) built a *predictive model* based on player features (e.g. Schuckers 2017 [1]), or 2) derived performance predictions from the observed performance of *comparable players* in a cohort (Weissbock 2015 [2]). This paper develops model tree learning, which incorporates strengths of both model-based and cohort-based approaches. A model tree partitions the feature space according to the values of discrete features, or learned thresholds for continuous features. Each leaf node in the tree defines a group of players, easily described to hockey experts, with its own group regression model. Compared to a single model, the model tree forms an ensemble that increases predictive power. Compared to cohort-based approaches, the groups of comparables are discovered from the data, without requiring a similarity metric. The performance predictions of the model tree are competitive with the state-of-the-art methods, which validates our model empirically. We show in case studies that the model tree player ranking can be used to highlight strong and weak points of players.

1. Introduction

Player ranking is one of the most studied subjects in sports analytics [17]. In this paper we consider predicting success in the National Hockey League (NHL) from junior league data, with the goal of supporting draft decisions. The publicly available junior league data aggregate a season's performance into a single set of numbers for each player. Our method can be applied to any data of this type, for example also to basketball NBA draft data [24]. Since our goal is to support draft decisions by teams, we ensure that the results of our data analysis method can be easily explained to and interpreted by sports experts.

Previous approaches for analyzing hockey draft data take a regression approach or a similarity-based approach. Regression approaches build a predictive model that takes as input a set of player features, such as demographics (age, height, weight) and junior league performance metrics (goals scored, plus-minus), and output a predicted success metric (e.g. number of games played in the professional league). The current state-of-the-art is a generalized additive model [1]. Cohort-based approaches divide players into groups of comparables and predict future success based on a player's cohort. For example, the PCS model [2] clusters players according to age, height, and scoring rates. One advantage of the cohort model is that predictions can be explained by reference to similar known players, which many domain experts find intuitive. For this reason, several commercial sports analytics systems, such as Sony's Hawk-Eye system, identify groups of comparables for each player. Our aim in this paper is to describe a new model for draft data that achieves the best of both approaches, regression-based and similarity-based.



Our method uses a *model tree* [8, 23]. Each node in the tree defines a new yes/no question, until a leaf is reached. Depending on the answers to the questions, each player is assigned a group corresponding to a leaf. The tree builds a different regression model for each leaf node. Figure 1 shows an example model tree. A model tree offers several advantages.

1. Compared to a single regression model, the tree defines an ensemble of regression models, based on non-linear thresholds. This increases the expressive power and predictive accuracy of the model. The tree can represent complex interactions between player features and player groups. For example, if the data indicate that players from different junior leagues are sufficiently different to warrant building distinct models, the tree can introduce a split to distinguish different leagues.
2. Compare to a similarity-based model, tree construction learns groups of players from the data, without requiring the analyst to specify a similarity metric. Because tree learning selects splits that increase predictive accuracy, the learned distinctions between the groups are guaranteed to be predictively relevant to future NHL success. Also, the tree creates a model, not a single prediction, for each group, which allows it to differentiate players from the same group.

A natural approach would be to build a linear regression tree to predict NHL success, which could be measured by the number of games a draft pick plays in the NHL. However, only about half the draft picks ever play a game in the NHL [19]. As observed by [1], this creates a zero-inflation problem that limits the predictive power of linear regression. We propose a novel solution to the zero-inflation problem, which applies logistic regression to predict whether a player will play at least one game in the NHL. We learn a logistic regression model tree, and rank players by the probability that the logistic regression model tree assigns to them playing at least one game. Intuitively, if we can be confident that a player will play at least one NHL game, we can also expect the player to play many NHL games. Empirically, we found that on the NHL draft data, the logistic regression tree produces a much more accurate player ranking than the linear regression tree.

Following [1], we evaluate the logistic regression ranking by comparing it to ranking players by their future success, measured as the number of NHL games they play after 7 years. The correlation of the logistic regression ranking with future success is competitive with that achieved by the generalized additive model of [1]. We show in case studies that the logistic model tree adds information to the NHL's Central Scouting Service Rank (CSS). For example, Stanley Cup winner Kyle Cumiskey was not ranked by the CSS in his draft year, but was ranked as the third draft prospect in his group by the model tree, just behind Brad Marchand and Mathieu Carle. Our case studies also show that the feature weights learned from the data can be used to explain the ranking in terms of which player features contribute the most to an above-average ranking. In this way the model tree can be used to highlight exceptional features of a player for scouts and teams to take into account in their evaluation.

Paper Outline. After we review related work, we show and discuss the model tree learned from the 2004-2006 draft data. The rank correlations are reported to evaluate predictive accuracy. We discuss in detail how the ensemble of group models represents a rich set of interactions between player features, player categories, and NHL success. Case studies give examples of strong players in different groups and show how the model can be used to highlight exceptional player features.



2. Related Work

Different approaches to player ranking are appropriate for different data types. For example, with dynamic play-by-play data, Markov models have been used to rank players [12, 13, 14, 18]. For data that record the presence of players when a goal is scored, regression models have also been applied to extend the classic plus-minus metric [15, 16]. In this paper, we utilize player statistics that aggregate a season's performance into a single set of numbers. While this data is much less informative than play-by-play data, it is easier to obtain, interpret, and process.

Regression Approaches. To our knowledge, this is the first application of model trees to hockey draft prediction, and the first model for predicting whether a draftee plays any games at all. The closest predecessor to our work is due to Schuckers [1], who uses a single generalized additive model to predict future NHL game counts from junior league data.

Similarity-Based Approaches assume a similarity metric and group similar players to predict performance. A sophisticated example from baseball is the nearest neighbour analysis in the PECOTA system [25]. For ice hockey, the Prospect Cohort Success (PCS) model [2], cohorts of draftees are defined based on age, height, and scoring rates. Model tree learning provides an automatic method for identifying cohorts with predictive validity. We refer to cohorts as groups to avoid confusion with the PCS concept. Because tree learning is computationally efficient, our model tree is able to take into account a larger set of features than age, height, and scoring rates. Also, it provides a separate predictive model for each group that assigns group-specific weights to different features. In contrast, PCS makes the same prediction for all players in the same cohort. So far, PCS has been applied to predict whether a player will score more than 200 games career total. Tree learning can easily be modified to make predictions for any game count threshold.

3. Dataset

Our data was obtained from public-domain on-line sources, including nhl.com, eliteprospects.com, and draftanalyst.com. We are also indebted to David Wilson for sharing his NHL performance dataset [3]. The full dataset is posted on the worldwide web [4]. We consider players drafted into the NHL between 1998 to 2008 (excluding goalies). Following [1], we took as our dependent variable **the total number of games g_i** played by a player i after 7 years under an NHL contract. The first seven seasons are chosen because NHL teams have at least seven-year rights to players after they are drafted [9]. Our dataset includes also the total time on ice after 7 years. The results for time on ice were very similar to number of games, so we discuss only the results for number of games. The independent variables include demographic factors (e.g. age), performance metrics for the year in which a player was drafted (e.g., goals scored), and the rank assigned to a player by the NHL Central Scouting Service (CSS). If a player was not ranked by the CSS, we assigned (1+ the maximum rank for his draft year) to his CSS rank value. Another preprocessing step was to pool all European countries into a single category. If a player played for more than one team in his draft year (e.g., a league team and a national team), we added up this counts from different teams. Table 1 lists all data columns and their meaning. Table 2 shows an excerpt from the dataset.

Variable Name	Description
id	nhl.com id for NHL players, otherwise Eliteprospects.com id
DraftAge	Age in Draft Year
Country	Nationality. Canada -> 'CAN', USA -> 'USA', countries in Europe -> 'EURO'
Position	Position in Draft Year. Left Wing -> 'L', Right Wing -> 'R', Center -> 'C', Defencemen -> 'D'
Overall	Overall pick in NHL Entry Draft
CSS_rank	Central scouting service ranking in Draft Year
rs_GP	Games played in regular seasons in Draft Year
rs_G	Goals in regular seasons in Draft Year
rs_A	Assists in regular seasons in Draft Year
rs_P	Points in regular seasons in Draft Year
rs_PIM	Penalty Minutes in regular seasons in Draft Year
rs_PlusMinus	Goal Differential in regular seasons in Draft Year
po_GP	Games played in playoffs in Draft Year
po_G	Goals in playoffs in Draft Year
po_A	Assists in playoffs in Draft Year
po_P	Points in playoffs in Draft Year
po_PIM	Penalty Minutes in playoffs in Draft Year
po_PlusMinus	Goal differential in playoffs in Draft Year
sum_7yr_GP	Total NHL games played in player's first 7 years of NHL career
sum_7yr_TOI	Total NHL Time on Ice in player's first 7 years of NHL career
GP_7yr_greater_than_0	Played a game or not in player's first 7 years of NHL career

Table 1 Player Attributes listed in dataset (excluding weight and height)

id	Player Name	Draft Age	Country	Height (in)	Weight (lbs)	Position	Overall	CSS_rank	rs_GP	rs_G	rs_A	rs_P	rs_PIM	rs_PlusMinus	sum_7yr_GP	sum_7yr_TOI	GP_7yr > 0
847-4141	Patrick Kane	19	USA	71	177	R	1	2	65	67	87	154	94	44	515	9927	yes
847-3419	Brad Marchand	18	CAN	69	181	L	71	80	68	29	37	66	83	40	218	3418	yes
27	Yared Hagos	18	EURO	73	218	C	70	24	43	11	26	37	24	1	0	0	no

Table 2 Sample Player Data for their draft year. rs = regular season. We use the same statistics for the playoffs (not shown).

4. Model Tree Construction

Model trees are a flexible formalism that can be built for any regression model. An obvious candidate for a regression model would be linear regression; alternatives include a generalized additive model [1], and a Poisson regression model specially built for predicting counts [22]. We introduce a different approach: a *logistic regression model to predict whether a player will play any games at all in the NHL ($g_i > 0$)*. The motivation is that many players in the draft never play any NHL



games at all (up to 50% depending on the draft year) [19]. This poses an extreme zero-inflation problem for any regression model that aims to predict directly the number of games played. In contrast, for the *classification problem* of predicting whether a player will play any NHL games, zero-inflation means that the data set is balanced between the classes. This classification problem is interesting in itself; for instance, a player agent would be keen to know what chances their client has to participate in the NHL. The logistic regression probabilities $p_i = P(g_i > 0)$ can be used not only to predict whether a player will play any NHL games, but also to *rank players* such that the ranking correlates well with the actual number of games played. Our method is therefore summarized as follows.

1. Build a tree whose leaves contain a logistic regression model.
2. The tree assigns each player i to a unique leaf node l_i , with a logistic regression model $m(l_i)$.
3. Use $m(l_i)$ to compute a probability $p_i = P(g_i > 0)$.

Figure 1 shows the logistic regression model tree learned for our second cohort by the LogiBoost algorithm (described in the appendix). It places CSS rank at the root as the most important attribute. Players ranked better than 12 form an elite group, of whom almost 82% play at least one NHL games. For players at rank 12 or below, the tree considers next their regular season points total. Players with rank and total points below 12 form an unpromising group: only 16% of them play an NHL game. Players with rank below 12 but whose points total is 12 or higher, are divided by the tree into three groups according to whether their regular season plus-minus score is positive, negative, or 0. (A three-way split is represented by two binary splits). If the plus-minus score is negative, the prospects of playing an NHL game are fairly low at about 37%. For a neutral plus-minus score, this increases to 61%. For players with a positive plus-minus score, the tree uses the number of playoff assists as the next most important attribute. Players with a positive plus-minus score and more than 10 playoff assists form a small but strong group that is 92% likely to play at least one NHL game.

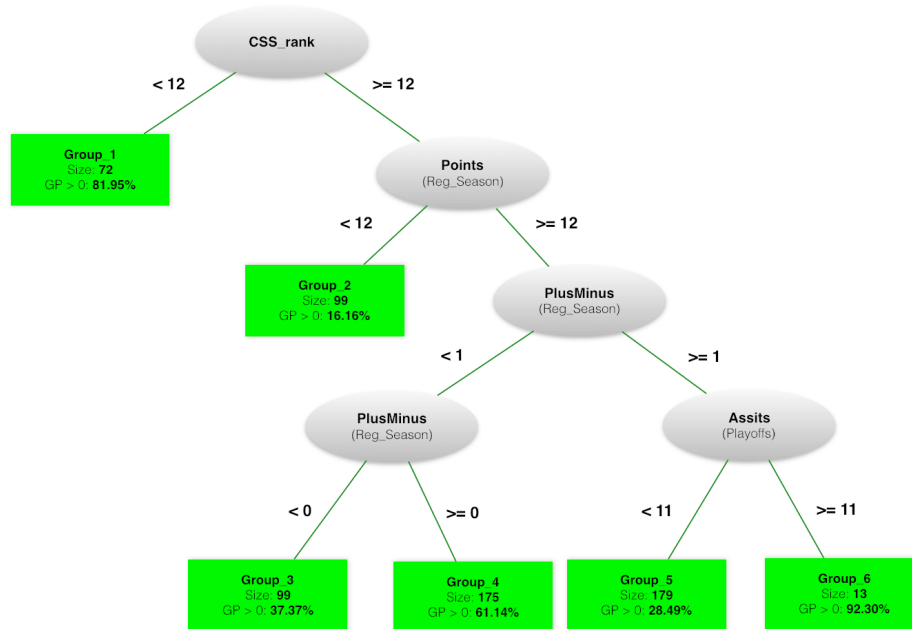


Figure 1 Logistic Regression Model Tree for the 2004,2005,2006 cohort. The Tree was built using the LogiBoost algorithm implemented in the LMT package of the Weka program [5, 6].



5. Results: Predictive Modelling

Following [1], we evaluated the predictive accuracy of the LMT model using the Spearman rank correlation (SRC) between two player rankings: i) the performance ranking based on the actual number of NHL games that a player played, and ii) the ranking of players based on the probability p_i of playing at least one game (Tree Model SRC). We also compared it with iii) the ranking of players based on the order in which they were drafted (Draft Order SRC). The draft order can be viewed as the ranking that reflects the judgment of NHL teams. We provide the formula for the Spearman correlation in the appendix. Table 3 shows the Spearman correlation for different rankings.

Training Data NHL Draft Years	Out of Sample Draft Years	Draft Order Spearman Rank Correlation	Tree Model classification accuracy	Tree Model Spearman Rank Correlation
1998, 1999, 2000	2001	0.43	82.27%	0.83
1998, 1999, 2000	2002	0.30	85.79%	0.85
2004, 2005, 2006	2007	0.46	81.23 %	0.84
2004, 2005, 2006	2008	0.51	63.56 %	0.71

Table 3 Predictive Performance (our model, over all draft ranking) using Spearman Rank Correlation. Bold indicates the best values.

Other Approaches. We also tried designs based on a linear regression model tree, using the M5P algorithm implemented in the Weka program. The result is a decision stump that splits on CSS rank only, which had substantially worse predictive performance (e.g. Spearman correlation of only 0.4 for the 2004-2006 cohort). For the generalized additive model (*gam*), the reported correlations were 2001: 0.53, 2002: 0.54, 2007: 0.69, 2008: 0.71 [1]. Our correlation is not directly comparable to the *gam* model because of differences in data preparation: the *gam* model was applied only to drafted players who played at least one NHL game, and the CSS rank was replaced by the Cescin conversion factors: for North American players, multiply CSS rank by 1.35, and for European players, by 6.27 [7]. The Cescin conversion factors represent an interaction between the player's country and the player's CSS rank. A model tree offers another approach to representing such interactions: by splitting on the player location node, the tree can build a different model for each location. Whether the data warrant building different models for different locations is a data-driven decision made by the tree building algorithm. The same point applies to other sources of variability, for example the draft year or the junior league. Including the junior league as a feature has the potential to lead to insights about the differences between leagues, but would make the tree more difficult to interpret; we leave this topic for future work. In the next section we examine the interaction effects captured by the model tree in the different models learned in each leaf.

6. Results: Learned Groups and Logistic Regression Models



We examine the learned group regression models, first in terms of the dependent success variable, then in terms of the player features.

6.1. Groups and the Dependent Variable

Figure 2 shows boxplots for the distribution of our dependent variable g_i . The strongest groups are, in order, 1, 6, and 4. The other groups show weaker performance on the whole, although in each group some players reach high numbers of games. Most players in Group 2&3&4&5 have GP equals to zero while Group 1&6 represent the strongest cohort in our prediction, where over 80% players played at least 1 game in NHL. The tree identifies that among the players who do not have a very high CSS rank (worse than 12), the combination of regular season Points ≥ 12 , PlusMinus > 0 , and play-off Assists > 10 is a strong indicator of playing a substantive number of NHL games (median $g_i = 128$).

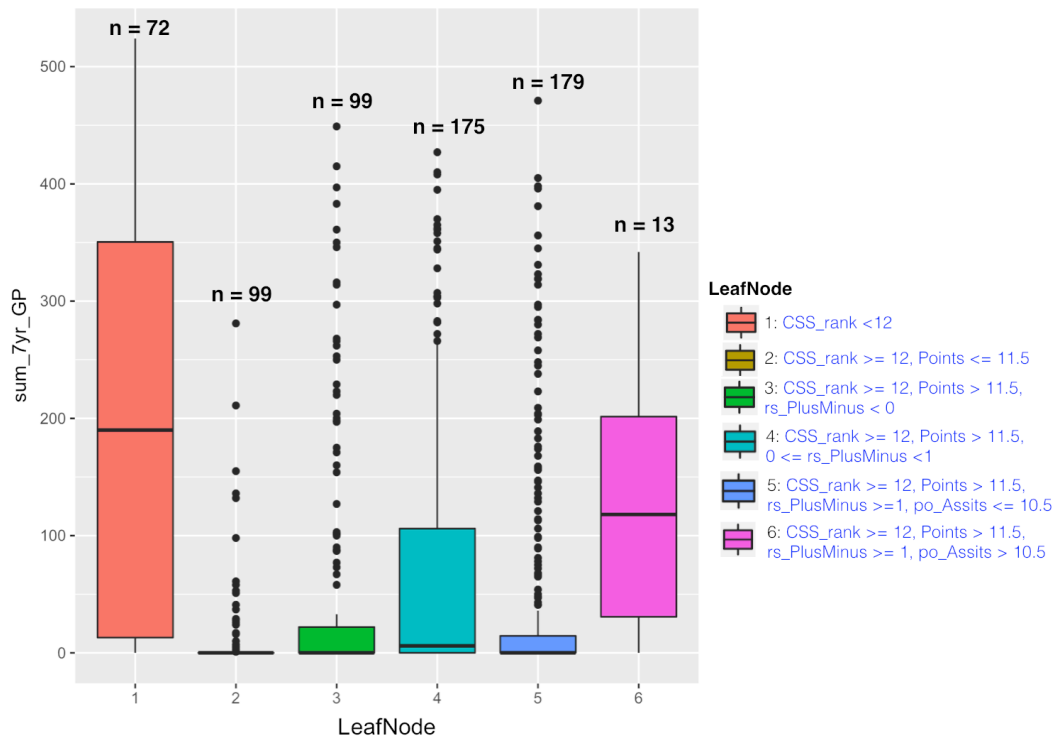


Figure 2 Boxplots for the dependent variable g_i , the total number of NHL games played after 7 years under an NHL contract. Each boxplot shows the distribution for one of the groups learned by the logistic regression model tree. The group size is denoted n .

6.2. Groups and the Independent Variables

Table 4 shows the average statistics by group and for all players. The CSS rank for Group 1 is by far the highest. The data validate the high ranking in that 82% players in this group went on to play an NHL game. Group 6 in fact attains an even higher proportion of 92%. The average statistics of this group are even more impressive than those of group 1 (e.g. 67 regular season points in group 6 vs.



47 for group 1). But the average CSS rank is the lowest of all groups. So this group may represent a small group of players ($n = 13$) overlooked by the scouts but identified by the tree. Other than Group 6, the group with the lowest CSS rank on average is Group 2. The data validate the low ranking in that only 16% of players in this group went on to play an NHL game. The group averages are also low (e.g. 6 regular season points is much lower than other groups).

Group	Mean Points											
	rs_P	rs_A	CSS_rank	po_A	Weight	Height	rs_Plus Minus	rs_GP	rs_PIM	Draft Age	po_GP	po_P
1	47	27	7	4	204	74	6	55	63	18	8	7
2	6	4	94	1	206	74	-2	38	56	18	8	1
3	40	23	76	2	201	73	-9	63	78	18	7	4
4	44	25	86	4	198	73	0	47	59	19	10	8
5	43	26	71	3	199	73	12	62	73	19	9	5
6	67	44	107	14	193	71	23	65	83	19	19	19
<i>all</i>	39	23	101	2	201	73	3	55	67	18	6	4

Table 4 Statistics for the average players in each group and all players.

7. Group Models and Variable Interactions

Table 5 illustrates logistic regression weights by group. A positive weight implies that an increase in the covariate value predicts a large increase in the probability of playing more than one game, compared to the probability of playing zero games. Conversely, a negative weight implies that an increase in the covariate value decreases the predicted probability of playing more than one game. Bold numbers show the groups for which an attribute is most relevant. The table exhibits many interesting interactions among the independent variables; we discuss only a few. Notice that *if the tree splits on an attribute, the attribute is assigned a high-magnitude regression weight by the logistic regression model for the relevant group*. Therefore our discussion focuses on the tree attributes.

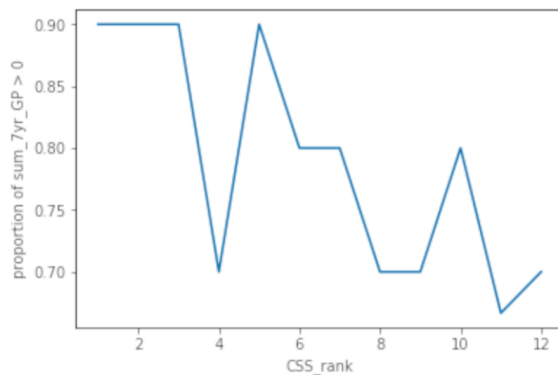
At the tree root, *CSS rank* receives a large negative weight of -17.9 for identifying the most successful players in Group 1, where all CSS ranks are better than 12. Figure 3a shows that the proportion of above-zero to zero-game players decreases quickly in Group 1 with worse CSS rank. However, the decrease is not monotonic. Figure 3b is a scatterplot of the original data for Group 1. We see a strong linear correlation ($p = -0.39$), and also a large variance within each rank. The proportion aggregates the individual data points at a given rank, thereby eliminating the variance. This makes the proportion a smoother dependent variable than the individual counts for a regression model.



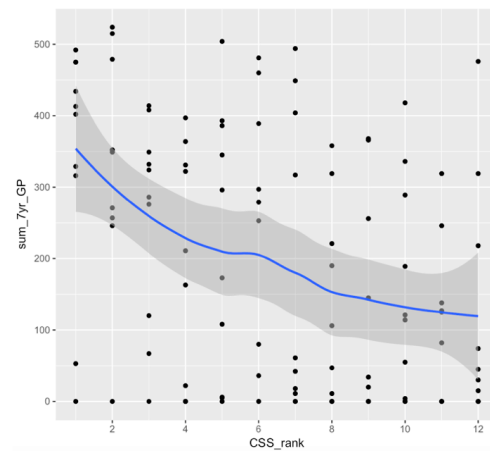
Met- rics Group	CSS_rank	Draft Age	Height	Weight	Country _group	Position	Games _played	Goals	Assists	Points	Penalty_ in_ Minutes	Plus Minus
1	-17.9	-3.91	-2.69	2.35	E: -0.77 C: 1.23 U: -0.49	D: -0.54 L: 2.09 R: -0.68	rs: -2.43 po: 4.15	rs: -0.03 po: -9.8	rs: 1.97 po: 8.89	rs: 1.73 po: 0.3	rs: 7.98 po: -7.6	rs: -0.45
2	-1.12	-1.1	-4.8	6.7	E: -0.13 C: 0.28 U: -0.22	D: -1.1 L: -0.45 R: 1.89	rs: 5.9 po: -14.1	rs: -2.17 po: -2.9	rs: 11.8 po: 21.6	rs: 14.2 po: 11.1	rs: -2.72 po: 5.2	rs: 1.57
3	-2.6	6.95	-7.4	6.7	E: -2.4 C: 1.04 U: 2.34	D: 0.39 L: 0.68 R: -0.4	rs: 3.21 po: -1.05	rs: -0.52 po: -0.6	rs: 1.36 po: -0.39	rs: 0.54 po: -2.6	rs: -1.88 po: 2.7	rs: 13.16
4	-2.4	5.2	-4.2	-0.52	E: 1.08 C: -0.40 U: -0.6	D: -0.03 L: -0.24 R: 0.14	rs: 3.58 po: -4.5	rs: -2.16 po: 1.58	rs: -0.12 po: 1.71	rs: -1.4 po: 1.6	rs: -2.72 po: 3.45	rs: 0
5	<u>-0.65</u>	-3.89	0.01	4.68	E: -1.26 C: 0.74 U: 0.47	D: 0.91 L: -0.64 R: 0.05	rs: 2.24 po: -0.25	rs: <u>3.59</u> po: -1.7	rs: -0.23 po: 0.33	rs: 2.19 po: -0.8	rs: -4.05 po: 6.86	rs: -0.73
6	-8.89	6.64	-14.91	0.34	E: -28.1 C: 5.9 U: 7.2	D: 3.32 L: 0.74 R: -28.12	rs: 16.7 po: 2.74	rs: 21.6 po: -9.7	rs: -0.34 po: -0.43	rs: -0.5 po: -0.4	rs: 1.3 po: -1.6	rs: 21.9

Table 5 Group 200(4 + 5 + 6 + 7 + 8) Weights Illustration. E = Europe, C = Canada, U = USA, rs = Regular Season, po = Playoff. Largest-magnitude weights are in bold. Underlined weights are discussed in the text.

Group 5 has the smallest logistic regression coefficient of -0.65. Group 5 consists of players whose CSS ranks are worse than 12, regular season points above 12, and plus-minus above 1. Figure 4a plots CSS rank vs. above-zero proportion for Group 5. As the proportion plot shows, the low weight is due to the fact that the proportion trends downward only at ranks worse than 200. The scatterplot in Figure 4b shows a similarly weak linear correlation of -0.12.



a)



b)

Figure 3 Proportion and scatter plots for CSS_rank vs. sum_7yr_GP in Group 1.

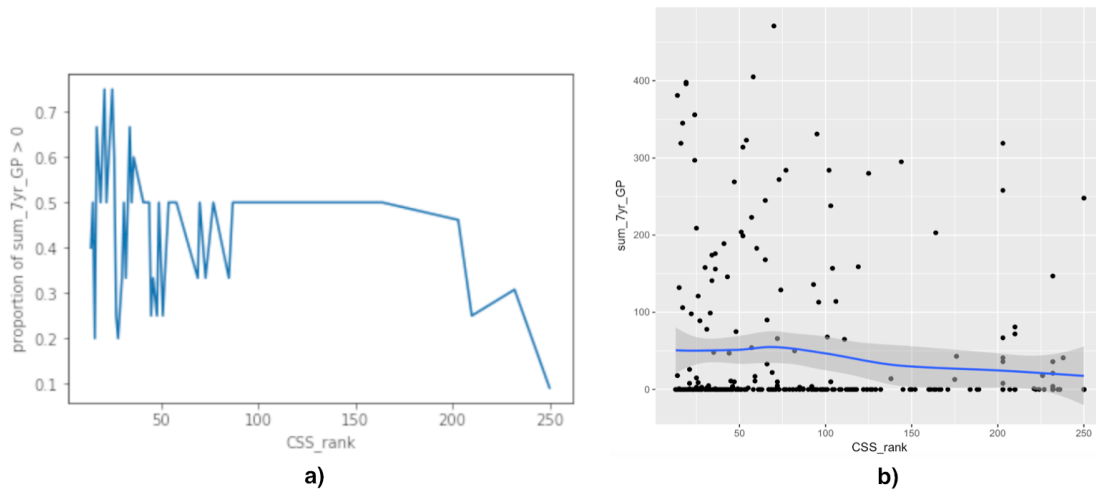


Figure 4 Proportion and scatter plots for *CSS_rank* vs. *sum_7yr_GP* in Group 5.

Regular season points are the most important predictor for Group 2, which comprises players with CSS rank worse than 12, and regular season points below 12. In the proportion plot Figure 5, we see a strong relationship between points and the chance of playing more than 0 games (logistic regression weight 14.2). In contrast in Group 4 (overall weight -1.4), there is essentially no relationship up to 65 points; for players with points between 65 and 85 in fact the chance of playing more than zero games slightly decreases with increasing points.

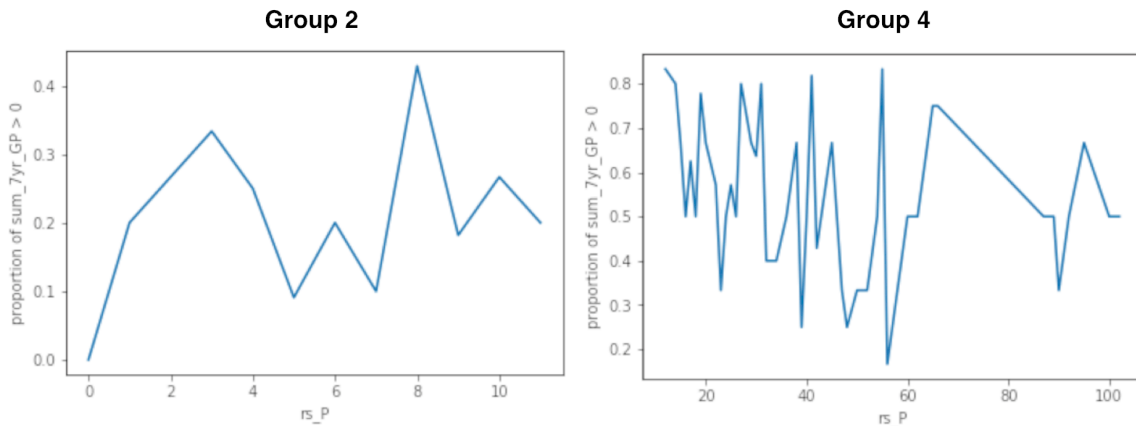


Figure 5 *Proportion_of_Sum_7yr_GP_greater_than_0* vs. *rs_P* in Group 2&4.

In Group 3, players are ranked at level 12 or worse, have collected at least 12 regular season points, and show a negative plus-minus score. The most important feature for Group 3 is the *regular season plus-minus score* (logistic regression weight 13.16), which is negative for all players in this group. In this group, the chances of playing an NHL game increase with plus-minus, but not monotonically, as Figure 6 shows.

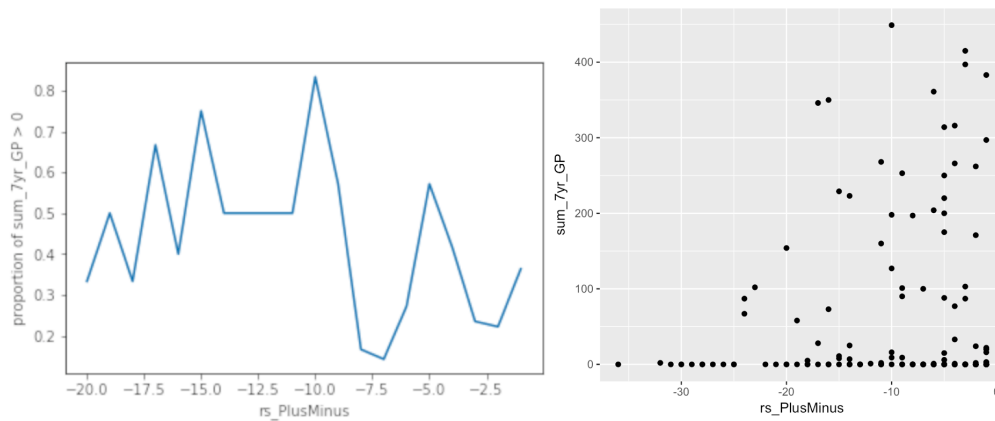


Figure 6 Proportion and scatter plots for rs_PlusMinus vs. sum_7yr_GP in group 3.

For *regular season goals*, Group 5 assigns a high logistic regression weight of 3.59. However, Group 2 assigns a surprisingly negative weight of -2.17. Group 5 comprises players at CSS rank worse than 12, regular season points 12 or higher, and positive plus-minus greater than 1. About 64.8% in this group are offensive players (see Figure 7). The positive weight therefore indicates that successful forwards score many goals, as we would expect.

Group 2 contains mainly defensemen (61.6%; see Figure 7). The typical strong defenseman scores 0 or 1 goals in this group. Players with more goals tend to be forwards, who are weaker in this group. In sum, the tree assigns weights to goals that are appropriate for different positions, using statistics that correlate with position (e.g. plus-minus), rather than the position directly.

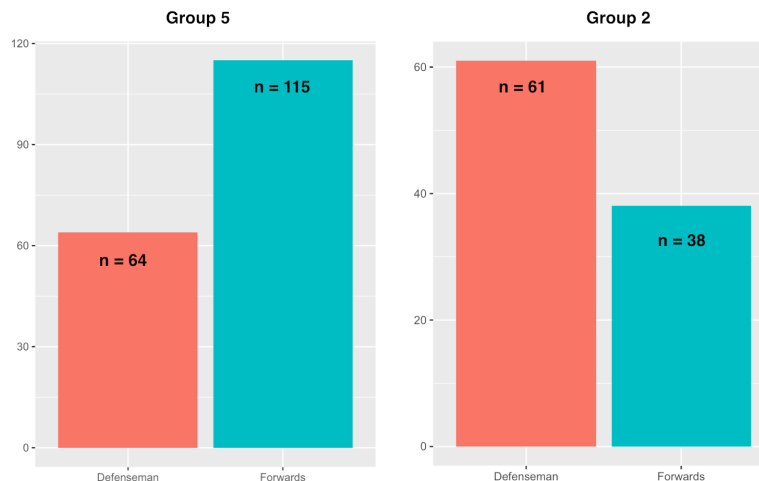


Figure 7 Distribution of Defenseman vs. Forwards in Group 5&2. The size is denoted as n .

8. Identifying Exceptional Players

Teams make drafting decisions not based on player statistics alone, but drawing on all relevant source of information, and with extensive input from scouts and other experts. As Cameron



Lawrence from the Florida Panthers put it, “the numbers are often just the start of the discussion” [11]. In this section we discuss how the model tree can be applied to support the discussion of individual players by highlighting their special strengths. The idea is that the learned weights can be used to identify which features of a highly-ranked player differentiate him the most from others in his group.

Explaining the Rankings: identify weak points and strong points

Our method is as follows. For each group, we find the average feature vector of the players in the group, which we denote by $\underline{x}_{g1}, \underline{x}_{g2}, \dots, \underline{x}_{gm}$ (see Table 4). We denote the features of player i as $\underline{x}_{i1}, \underline{x}_{i2}, \dots, \underline{x}_{im}$. Then given a weight vector (w_1, w_m) for the logistic regression model of group g , the log-odds difference between player i and a random player in the group is given by

$$\sum_{j=1}^m w_j (x_{ij} - \bar{x}_{gj}).$$

We can interpret this sum as a measure of how high the model ranks player i compared to other players in his group. This suggests defining as the player’s *strongest features* the x_{ij} that maximize $w_j(x_{ij} - \bar{x}_{gj})$, and as his *weakest features* those that minimize $w_j(x_{ij} - \bar{x}_{gj})$. This approach highlights features that are i) relevant to predicting future success, as measured by the magnitude of w_j , and ii) different from the average value in the player’s group of comparables, as measured by the magnitude of $x_{ij} - \bar{x}_{gj}$.

Case Studies. Table 6 shows, for each group, the three strongest points for the most highly ranked players in the group. We see that the ranking for individual players is based on different features, even within the same group. The table also illustrates how the model allows us to identify a group of comparables for a given player. We discuss a few selected players and their strong points. The most interesting cases are often those where the ranking differs from the scouts’ CSS rank. We therefore discuss the groups with lower rank first.

Among the players who were not ranked by CSS at all, our model ranks *Kyle Cumiskey* at the top. Cumiskey was drafted in place 222, played 132 NHL games in his first 7 years, represented Canada in the World Championship, and won a Stanley Cup in 2015 with the Blackhawks. His strongest points were being Canadian, and the number of games played (e.g., 27 playoff games vs. 19 group average).

In the lowest CSS-rank group 6 (average 107), our top-ranked player *Brad Marchand* received CSS rank 80, even below his Boston Bruin teammate’s *Lucic*’s. Given his Stanley Cup win and success representing Canada, arguably our model was correct to identify him as a strong NHL prospect. The model highlights his superior play-off performance, both in terms of games played and points scored. Group 2 (CSS average 94) is a much weaker group. *Matt Pelech* is ranked at the top by our model because of his unusual weight, which in this group is unusually predictive of NHL participation. In group 4 (CSS average 86), *Sami Lepisto* was top-ranked, in part because he did not suffer many penalties although he played a high number of games. In group 3 (CSS average 76), *Brandon McMillan* is ranked relatively high by our model compared to the CSS. This is because in this group, left-wingers and shorter players are more likely to play in the NHL. In our ranking, *Milan Lucic* tops Group 5 (CSS average 71). At 58, his CSS rank is above average in this group, but much



below the highest CSS rank player (*Legein* at 13). The main factors for the tree model are his high weight and number of play-off games played. Given his future success (Stanley Cup, NHL Young Stars Game), arguably our model correctly identified him as a star in an otherwise weaker group. The top players in Group 1 like *Sidney Crosby* and *Patrick Kane* are obvious stars, who have outstanding statistics even relative to other players in this strong group.

Group	Top Players	Strongest Points (\bar{x} = group mean)		
1	<u>Sidney Crosby</u>	rs_P	rs_A	CSS_rank
		188 (\bar{x} = 47)	110 (\bar{x} = 27)	1 (\bar{x} = 7)
	<u>Patrick Kane</u>	rs_P	rs_A	CSS_rank
		154 (\bar{x} = 47)	87 (\bar{x} = 27)	2 (\bar{x} = 7)
	Sam Gagner	rs_P	po_A	rs_A
		118 (\bar{x} = 47)	22 (\bar{x} = 4)	83 (\bar{x} = 27)
2	<u>Matt Pelech</u>	Weight	CSS_rank	rs_A
		230 (\bar{x} = 206)	41 (\bar{x} = 94)	4 (\bar{x} = 4)
	Adam Pineault	CSS_rank	rs_P	Height
		25 (\bar{x} = 94)	8 (\bar{x} = 6)	73 (\bar{x} = 74)
	Roman Wick	rs_P	CSS_rank	rs_PlusMinus
		10 (\bar{x} = 6)	36 (\bar{x} = 94)	0 (\bar{x} = -2)
3	A.J.Jenks	CSS_rank	Weight	Country
		20 (\bar{x} = 76)	205 (\bar{x} = 201)	USA
	Bill Sweatt	CSS_rank	Position	rs_PlusMinus
		27 (\bar{x} = 76)	L	-1 (\bar{x} = -2)
	<u>Brandon McMillan</u>	CSS_rank	Position	Height
		44 (\bar{x} = 76)	L	71 (\bar{x} = 73)
4	Sami Lepisto	CSS_rank	rs_GP	rs_PIM
		25 (\bar{x} = 86)	61 (\bar{x} = 47)	30 (\bar{x} = 59)
	Linus Omark	CSS_rank	Height	DraftAge
		55 (\bar{x} = 86)	70 (\bar{x} = 73)	20 (\bar{x} = 19)
	Oscar Moller	CSS_rank	Height	rs_GP
		20 (\bar{x} = 86)	70 (\bar{x} = 73)	68 (\bar{x} = 47)
5	<u>Milan Lucic</u>	Weight	po_GP	CSS_rank
		236 (\bar{x} = 199)	23 (\bar{x} = 9)	58 (\bar{x} = 71)
	Michael Del Zotto	Position	Country	po_GP
		D	CAN	15 (\bar{x} = 9)
	Steven Delisle	Weight	Country	po_GP
		234 (\bar{x} = 199)	CAN	19 (\bar{x} = 9)
6	<u>Brad Marchand</u>	Country	po_GP	po_P
		CAN	25 (\bar{x} = 19)	23 (\bar{x} = 19)
	Mathieu Carle	Country	CSS_rank	rs_GP
		CAN	53 (\bar{x} = 107)	67 (\bar{x} = 65)
	<u>Kyle Cumiskey</u>	Country	po_GP	rs_GP
		CAN	27 (\bar{x} = 19)	72 (\bar{x} = 65)

Table 6 Strongest Statistics for the top players in each group. Underlined players are discussed in the text.



Conclusion and Future Work

We have proposed building a regression model tree for ranking draftees in the NHL, or other sports, based on a list of player features and performance statistics. The model tree groups players according to the values of discrete features, or learned thresholds for continuous performance statistics. Each leaf node defines a group of players that is assigned its own regression model. Tree models combine the strength of both regression and cohort-based approaches, where player performance is predicted with reference to comparable players. An obvious approach is to use a linear regression tree for predicting our dependent variable, the number of NHL games played by a player within 7 NHL years. However, we found that a linear regression tree performs poorly due to the zero-inflation problem (many draft picks never play any NHL game). Instead, we introduced the idea of using a logistic regression tree to predict whether a player plays any NHL game within 7 years. Players are ranked according to the model tree probability that they play at least 1 game.

Key findings include the following. 1) The model tree ranking *correlates well with the actual success ranking* according to the actual number of games played: *better than draft order* and competitive with the state-of-the-art generalized additive model [1]. 2) The model predictions *complement the Central Scouting Service (CSS) rank*. For example, the tree identifies a group whose average CSS rank is only 107, but whose median number of games played after 7 years is 128, including several Stanley Cup winners. 3) The model tree can *highlight the exceptionally strong and weak points of draftees* that make them stand out compared to the other players in their group.

Tree models are flexible and can be applied to other prediction problems to discover groups of comparable players as well as predictive models. For example, we can predict future NHL success from past NHL success, similar to Wilson [3] who used machine learning models to predict whether a player will play more than 160 games in the NHL after 7 years. Another direction is to apply the model to other sports, for example drafting for the National Basketball Association.

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Appendix

Spearman Rank Correlation

Spearman's correlation measures the relevance and direction of monotonic association between two variables [10]. The standard formula for calculating is based on the squared rank differences:

(1) $p = 1 - \frac{6 \sum d_i^2}{n(n^2-1)}$, formula for no tied ranks. n = number of ranks, d_i = different in paired ranks.

This is the formula we applied in Table 3.

In the case of tied ranks, an alternative is to use the Pearson correlation between ranks, computed by the formula:

(2) $p = \frac{\sum_i (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum_i (x_i - \bar{x})^2 \sum_i (y_i - \bar{y})^2}}$, where x_i = rank of player i according to ranking x , ditto for y_i .

Players who have played zero NHL games are tied when ranked by the number of NHL games; this is the only case of ties. Table 7 repeats the calculation of Table 3 using the Pearson correlation among ranks (2) rather than the squared rank differences (1). With this measure also, the model ranking correlates more highly with actual number of games played than the team draft order.

Training Data NHL Draft Years	Out of Sample Draft Years	Draft Order Pearson Rank Correlation	Pearson Rank Correlation
1998, 1999, 2000	2001	0.43	0.69
1998, 1999, 2000	2002	0.45	0.72
2004, 2005, 2006	2007	0.48	0.60
2004, 2005, 2006	2008	0.51	0.58

Table 7 Spearman correlation as Pearson correlation of ranks



LogitBoost Algorithm

Ensemble methods provide a combination of classifiers to obtain a better predictive result than any of the standalone constituents [20, 21, 22]. While a model tree constructs a set of different classifiers, it also partitions the space of players, so the prediction for each player is based on exactly one model only, rather than a weighted majority vote. The LogitBoost algorithm [8] combines a model tree with weighted votes by building a separate model for each tree node (including non-leaf nodes). The prediction for a specific player is then the weighted vote of all models along the branch assigned to the player. This offers some of the advantages of a hierarchical shrinkage model in smoothing parameter values so that predictions from players from similar groups tend to use similar weights. In this paper, we used the tree structure learned from LogitBoost, with simple maximum likelihood estimates for the weights to make the weight for a leaf more interpretable by fitting the data in the leaf's group more closely. The GUIDE system [23] is a well-developed software package that supports building ensembles of model trees. While such an ensemble tends to have even higher predictive accuracy, we have in this paper built only a single model tree to maintain interpretability.