

BAYESUVIUS

A VISUAL DICTIONARY OF BAYESIAN
NETWORKS AND CAUSAL INFERENCE



ROBERT R. TUCCI

Bayesuvius,
a visual dictionary of Bayesian Networks and
Causal Inference

Robert R. Tucci
www.ar-tiste.xyz

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This book is constantly being expanded and improved. To download
the latest version, go to <https://github.com/rrtucci/Bayesuvius>

Bayesuvius

by Robert R. Tucci

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Figure 1: View of Mount Vesuvius from Pompeii



Figure 2: Mount Vesuvius and Bay of Naples

Contents

Foreword	16
Appendices	17
A Navigating the ocean of Judea Pearl’s Books	18
B CI-2-3 track	19
C Notational Conventions and Preliminaries	22
C.1 Some abbreviations frequently used throughout this book	22
C.2 $\mathcal{N}(!a)$	22
C.3 One hot vector	22
C.4 Special sets	23
C.5 Kronecker delta function	23
C.6 Dirac delta function	23
C.7 Indicator function (a.k.a. Truth function)	23
C.8 Majority function	23
C.9 Underlined letters indicate random variables	24
C.10 Probability distributions	24
C.11 Discretization of continuous probability distributions	24
C.12 Samples, i.i.d. variables	25
C.13 Expected Value and Variance	26
C.14 Conditional Expected Value	26
C.15 Notation for covariances	27
C.16 Conditional Covariance	27
C.17 Normal Distribution	28
C.18 Uniform Distribution	29
C.19 Softmax function (a.k.a. Boltzmann Distribution)	29
C.20 Sigmoid and log-odds functions	30
C.21 Estimand, Estimator (curve-fit, cfit), Estimate, Bias	31
C.22 Maximum Likelihood Estimate, Likelihood Ratio Test	32
C.23 Mean Square Error (MSE)	33
C.24 Cramer-Rao Bound	35
C.25 Bayes Rule, Bayesian Updating And Conjugate Priors	39

C.26 Linear regression, Ordinary Least Squares (OLS)	40
C.26.1 LR, assuming x_σ are non-random	41
Derivation of LR From Minimization of Error	42
Geometry of LR with non-random x_σ	43
LR Goodness of Fit, R^2	44
C.26.2 LR, assuming x_σ are random	46
Transforming expressions from non-random to random x_σ	46
Geometry of LR with random x_σ	48
Regression interpreted as differentiation of y	50
C.27 Logistic Regression (LoR)	51
C.28 Entropy, Kullback-Leibler divergence, Cross-Entropy	52
C.29 Definition of various entropies used in Shannon Information Theory .	52
C.30 Mean log likelihood asymptotic behavior	54
C.31 Arc Strength (Arc Force)	55
C.32 Pearson Chi-Squared Test	56
C.33 Demystifying Population and Sample Variances	57
C.34 Independence of $\hat{\mu}$ and $\hat{\sigma}^2$	58
C.35 Chi-square distribution	59
C.36 Student's t-distribution	60
C.37 Hypothesis testing and 3 classic test statistics (Likelihood, Score, Wald)	63
C.38 Error Bars	66
C.39 Confidence Interval	67
C.40 Score p-value	69
C.41 Convex/Concave functions, Jensen's Inequality	71
C.42 Chebyshev's inequality	72
C.43 Short Summary of Boolean Algebra	74
D Definition of a Bayesian Network	76
E Bayesian Networks, Causality and the Passage of Time	80
E.1 Unifying Principle of this book	80
E.2 You say tomato, I say tomato	81
E.3 A dataset is causal model free	81
E.4 What is causality?	82
E.5 Bayesian Networks and the passage of time	83
E.6 Advice for the DAG-phobic	84
1 AdaBoost	85
1.1 AdaBoost for general ensemble of w-classifiers	85
1.2 AdaBoost for ensemble of tree stumps	89

2 ANOVA	91
2.1 Law of Total Variance	91
2.2 Sum of Squares Estimates	92
2.3 F-statistic and hypothesis testing	94
3 ARACNE structure learning	96
4 Backdoor Adjustment Formula	98
4.1 Examples	99
5 Back Propagation (Automatic Differentiation)	103
5.1 General Theory	103
5.1.1 Jacobians	103
5.1.2 Bnets for function composition, forward propagation and back propagation	104
5.2 Application to Neural Networks	105
5.2.1 Absorbing b_i^λ into w_{ij}	105
5.2.2 Bnets for function composition, forward propagation and back propagation for NN	107
5.3 General bnets instead of Markov chains induced by layered structure of NNs	110
6 Basic Curve Fitting Using Gradient Descent	111
7 Bell and Clauser-Horne Inequalities in Quantum Mechanics	113
8 Berkson's Paradox	114
9 Binary Decision Diagrams	116
10 Chow-Liu Trees and Tree Augmented Naive Bayes (TAN)	120
10.1 Chow-Liu Trees	120
10.2 Tree Augmented Naive Bayes (TAN)	124
11 Counterfactual Reasoning	126
11.1 The 3 Rungs of Causal AI	126
11.2 Do operator	127
11.3 Imagine operator	127
12 Cross-Validation	131
13 Dataset Shift and Batch Normalization	134
13.1 Covariate Shift	135
13.2 Concept Shift	135

13.3 Batch Normalization	136
14 Decision Trees	137
14.1 Transforming a dtree into a bnet	139
14.2 Structure Learning for Dtrees	140
14.2.1 Information Gain, Gini	141
14.3 Information Gain Ratio	144
14.3.1 Pseudo-code	144
15 Decisions Based on Rungs 2 and 3: COMING SOON	147
16 Difference-in-Differences	148
16.1 John Snow, DID and a cholera transmission pathway	148
16.2 PO analysis	150
16.3 Linear Regression	152
17 Diffusion Models	155
17.1 Bnet for DM	155
17.2 Mean Values $M^{t-1}(x^t)$ and $M_\theta^{t-1}(x^t)$	158
17.3 Loss function \mathcal{L}	161
17.4 Algorithms for training and sampling DM	163
18 Digital Circuits	165
18.1 Mapping any dcircuit to a bnet	165
18.1.1 Option A of Fig.18.2	165
18.1.2 Option B of Fig.18.2	166
19 Do Calculus	167
19.1 3 Rules of Do Calculus	170
19.2 Parent Adjustment Formula	171
19.3 Backdoor Adjustment Formula	173
19.4 Frontdoor Adjustment Formula	174
19.5 Comparison of Backdoor and Frontdoor adjustment formulae	175
19.6 Do operator for DEN diagrams	176
20 Do Calculus proofs	179
21 D-Separation	197
22 D-Separation and Do Calculus Rules for LDEN: COMING SOON	200
23 D-Separation in Quantum Mechanics	201
24 Dynamical Bayesian Networks	202

25 Expectation Maximization	204
25.1 The EM algorithm:	205
25.1.1 Motivation	206
25.2 Minorize-Maximize (MM) algorithms	206
25.3 Examples	208
25.3.1 Gaussian mixture	208
25.3.2 Blood Genotypes and Phenotypes	209
25.3.3 Missing Data/Imputation	211
26 Factor Graphs	212
27 Frisch-Waugh-Lovell (FWL) theorem	215
27.1 FWL, assuming x^σ are non-random	215
27.2 FWL, assuming x^σ are random	216
28 Frontdoor Adjustment Formula	219
28.1 Examples	220
29 G-formula (Sequential Backdoor Adjustment Formula)	221
30 Gaussian Nodes with Linear Dependence on Parents	225
31 Generalized Linear Model (GLM)	228
31.1 Exponential Family of Distributions	228
31.2 GLM	230
32 Generative Adversarial Networks (GANs)	235
33 Goodness of Causal Fit	240
33.1 Abstract	240
33.2 Introduction	240
33.3 Goodness of Statistical Fit	241
33.4 GCF example 1	242
33.5 GCF example 2	244
33.6 GCF in general	245
34 Granger Causality	247
35 Hidden Markov Model	250
35.1 Calculating $P(x_t, v^n)$ and $P(x_t, x_{t+1}, v^n)$	252
35.2 Calculating \mathcal{F}_t and $\bar{\mathcal{F}}_t$	253
35.3 Calculating $P(x^n v^n)$	254
35.4 Calculating $P(v^n A, B, \pi)$	255
35.5 Calculating \hat{x}^n (Viterbi algorithm)	256

35.6 Calculating \widehat{A} , \widehat{B} , $\widehat{\pi}$ (Baum-Welch algorithm)	258
36 Influence Diagrams & Utility Nodes	260
37 Instrumental Inequality and beyond	262
37.1 I-inequality	262
37.1.1 I-inequality for binary z,d,y	264
37.2 Bounds on Effect of IV on treatment outcome y	265
38 Instrumental Variables	268
38.1 δ with unmeasured confounder	268
38.2 δ (with unmeasured confounder) can be inferred via IV	269
38.3 More general bnets with IVs	270
38.4 Instrumental Inequality	271
39 Jackknife Resampling	272
39.1 Case $A = A^n(\vec{x}) = \frac{1}{n} \sum_{\sigma} x^{\sigma}$	274
40 Junction Tree Algorithm	276
41 Kalman Filter	277
41.1 Prediction Problem	278
41.2 Solution	279
41.3 Simple Example	280
41.4 Invariants	281
41.5 Derivation of Solution	281
42 LATE (Local Average Treatment Effect)	283
43 Linear and Logistic Regression	288
43.1 Generalization to x with multiple components (features)	290
43.2 Alternative $V(b, m)$ for logistic regression	290
44 Linear Deterministic Bnets with External Noise	292
44.1 Example of LDEN diagram	292
44.2 Fully Connected LDEN diagrams	293
44.2.1 Fully connected LDEN diagram with $nx = 2$	294
44.2.2 Fully connected LDEN diagram with $nx = 3$	294
44.2.3 Fully connected LDEN diagram with arbitrary nx	295
44.3 Non-linear DEN diagrams	297
45 Markov Blankets	298

46 Markov Chain Monte Carlo (MCMC)	300
46.1 Inverse Cumulative Sampling	300
46.2 Rejection Sampling	302
46.3 Metropolis-Hastings Sampling	303
46.4 Gibbs Sampling	306
46.5 Importance Sampling	307
47 Markov Chains	309
48 Mediation Analysis	310
49 Message Passing and Bethe Free Energy	316
49.1 2MRFs	316
49.2 Message Passing Intuition	317
49.3 $-\ln Z_\theta$ = Free Energy (FE)	321
49.4 $-\ln Z_{\theta^*}$ = Minimum FE	322
49.5 $-\ln Z_\theta^{tree}$ =Tree FE (a.k.a. Bethe FE)	323
49.6 $-\ln Z_{\theta^*}^{tree}$ = Tree Minimum FE, and message passing	324
50 Message Passing, Pearl's theory	328
50.1 Distributed Soldier Counting	328
50.2 Spring Systems	330
50.3 BP for Markov Chains	330
50.4 BP Algorithm for Polytrees	337
50.4.1 How BP algo for polytrees reduces to the BP algo for Markov chains	341
50.5 Derivation of BP Algorithm for Polytrees	342
50.6 Example of BP algo for a Tree	345
50.7 Bipartite bnets	349
50.8 BP for bipartite bnets (BP-BB)	350
50.8.1 BP-BB and general BP agree on Markov chains	352
50.8.2 BP-BB and general BP agree on tree bnets.	354
50.9 BP-BB and sum-product decomposition	356
51 Message Passing (Belief Propagation) in Quantum Mechanics	357
52 Meta-learners for estimating ATE	358
53 Missing Data, Imputation	362
53.1 Imputation via EM	363
53.2 Imputation via MCMC	366
53.3 Multiple Imputations	367

54 Modified Treatment Policy	368
54.1 One time MTP	368
54.2 $\Delta_{ c}$ estimand	372
54.3 Estimates of $\Delta_{ c}$	375
54.3.1 Empirical estimate of $\Delta_{ c}$	375
54.3.2 OR estimate of $\Delta_{ c}$	375
54.4 Other Estimands besides $\Delta_{ c}$	378
54.5 Multi-time MTP	378
55 Monty Hall Problem	381
56 Multi-armed Bandits	383
56.1 Bnet for MAB	384
56.2 Reward functions	386
56.3 Regret functions	388
56.4 Strategies with random exploration	389
56.4.1 ϵ -greedy algorithm	389
56.4.2 ϵ_t -greedy algorithm	390
56.5 Strategies with nonrandom exploration	390
56.5.1 Upper Confidence Bounds (UCB) algorithms	390
Frequentist UCB (UCB1) algorithm	390
Bayesian UCB algorithm	391
56.5.2 Thompson Sampling MAB (TS-MAB) algorithm	392
Bnet for general TS-MAB algorithm	392
TS-MAB algorithm with Beta agent and Bernoulli environment	394
TS-MAB algorithm, skeletal reprise	395
56.5.3 Grad-MAB algorithm	396
57 Naive Bayes	399
58 Neural Networks	400
58.1 Activation Functions $\mathcal{A}_i^\lambda : \mathbb{R} \rightarrow \mathbb{R}$	401
58.2 Weight optimization via supervised training and gradient descent . .	402
58.3 Non-dense layers	404
58.4 Autoencoder NN	406
59 Noisy-OR gate	407
59.1 3 ways to interpret the parameters π_i	408
60 Non-negative Matrix Factorization	411
60.1 Bnet interpretation	411
60.2 Simplest recursive algorithm	412

61 Observationally Equivalent DAGs	413
61.1 Examples	413
62 Personalized Treatment Effects	416
62.1 Goal, Strategy and Rationale of PTE theory	417
62.2 Bnets for PTE theory	419
62.3 $ATE = PNS - AMM$	420
62.4 Probabilities Relevant to PTE theory	421
62.5 Symmetry	426
62.6 Linear Programming Problem	427
62.7 Special constraints	428
62.8 Matrix representation of probabilities	430
62.9 Bounds on Exp. Probs. imposed by Obs. Probs.	434
62.10 Bounds on PNS_3 for unspecified bnet	435
62.11 Bounds on PNS_3 for specific bnet families	441
62.12 Numerical Examples	446
63 Personalized Expected Utility	447
63.1 Goal of PEU Theory	448
63.2 Bnets for PEU Theory	449
63.3 Bounds on EU for unspecified bnet	449
63.4 Bounds on EU for specific bnet families	452
64 Potential Outcomes and Beyond	455
64.1 G and G_{den} bnets, the starting point bnets	456
64.2 G bnet with nodes $y^\sigma(0), y^\sigma(1)$ added to it.	458
64.3 Expected Values of treatment outcome y^σ	460
64.4 Translation Dictionary	460
64.5 $\mathcal{Y}_{ d,x} = \mathcal{Y}_{d x}$ (SUTVA)	461
64.6 Conditional Independence Assumption (CIA)	462
64.7 Treatment Effects	462
64.8 Insights into what makes treatment effects equal and $\mathcal{Y}_{1 0} = \mathcal{Y}_1$	465
64.9 G_{do+} bnet	466
64.10 $ACE = ATE$	467
64.11 Good, Bad Controls	468
64.12 PO Confounder Sensitivity Analysis	469
64.13 (SDO, ATE) space	470
64.14 Strata-Matching	473
64.14.1 Exact strata-matching	473
Estimates of Treatment Effects	473
Example, estimation of treatment effects	475
64.14.2 Approximate strata-matching	477
64.14.3 Unbiased strata-matching estimates	477

64.15	Propensities	479
64.16	Propensity based estimates of treatment effects	483
64.17	Positivity	484
64.18	Multi-time PO bnets (Panel Data)	485
65	Program evaluation and review technique (PERT)	488
65.1	Example	490
66	Random Forest and Bagging	494
66.1	Bagging (with fully-featured bags)	494
66.2	Bagging (with randomly-shortened bags)	496
67	Recurrent Neural Networks	497
67.1	Language Sequence Modeling	500
67.2	Other types of RNN	500
67.2.1	Long Short Term Memory (LSTM) unit (1997)	502
67.2.2	Gated Recurrence Unit (GRU) (2014)	504
68	Regression Discontinuity Design	506
68.1	PO analysis	506
68.2	Linear Regression	508
69	Reinforcement Learning (RL)	509
69.1	Exact RL bnet	512
69.2	Actor-Critic RL bnet	514
69.3	Q function learning RL bnet	516
70	Reliability Box Diagrams and Fault Tree Diagrams	518
70.1	Minimal Cut Sets	524
71	Restricted Boltzmann Machines	526
72	ROC curves	528
72.1	Terminology Table Adapted from Wikipedia Ref.[134]	531
73	Scoring the Nodes of a Learned Bnet	533
73.1	Probability Distributions and Special Functions	534
73.2	Single node with no parents	536
73.3	Multiple nodes with any number of parents	538
73.4	Bayesian Scores	540
73.5	Information Theoretic scores	540

74 Selection Bias Removal	542
74.1 Pre and Post Selection Nodes	543
74.2 Removing SB from passive query $P(y x)$	545
74.3 Removing SB from active query $P(y \mathcal{D}x)$	547
75 Shannon Information Theory	549
76 Shapley Explainability	550
76.0.1 Numerical examples of SHAP	553
77 Simpson's Paradox	556
77.1 Pearl Causality	558
77.2 Numerical Example	560
78 Structure and Parameter Learning for Bnets	561
78.1 Overview	561
78.2 Score based SL algorithms	563
78.3 Constraint based SL algorithms	564
78.4 Pseudo-code for some bnet learning algorithms	565
79 Support Vector Machines And Kernel Method	567
79.1 Learning Algorithm for SVM Classifier	568
79.2 Linear (dot-product) Kernel	569
79.3 Alternatives to Linear Kernel	573
79.4 Random Forest and Kernel Method	573
80 Survival Analysis	574
80.1 $S(t)$ estimates	576
80.1.1 No-censoring estimate of $S(t)$	576
80.1.2 Kaplan-Meier estimate of $S(t)$	576
80.2 $\lambda(t)$ models	581
80.2.1 $\lambda(t)$ independent of covariates Z	581
80.2.2 $\lambda(t)$ dependent on covariates Z	582
80.3 $S_0(t)$ estimates	585
81 Synthetic Controls	587
81.1 PO analysis	589
82 Targeted Estimator	591
82.1 Goal, Strategy, and Rationale of TE theory	591
82.2 Functional Calculus	593
82.3 Linear Approximation of $\Psi[P_N]$	595
82.4 ATE estimand	596
82.5 ATE estimates	597

82.5.1	Ψ^E	597
82.5.2	Ψ^G	598
82.5.3	Ψ^{IPW}	598
82.5.4	Ψ^{LIPW}	598
82.5.5	Ψ^{LIPW++} (a.k.a. Ψ^{TMLE})	602
82.6	Ψ^{TMLE} in practice	605
83	Time Series Analysis: ARMA and VAR	607
83.1	White noise	607
83.2	Backshift operator	608
83.3	Metrics	608
83.4	Definition of $ARMA(p, q)$, $AR(p)$ and $MA(q)$.	610
83.5	Solving $AR(p)$	612
83.6	Solving $MA(q)$	613
83.7	Solving $ARMA(p, q)$	614
83.8	Auto-correlation and partial auto-correlation	614
83.9	Generating function of auto-correlation	618
83.10	Impulse Response	619
83.11	$AR(p)$ and Yule-Walker equations	621
83.12	Forecasting	622
83.13	Model Learning	628
83.14	Differencing and $ARIMA(p, d, q)$	628
83.15	Parameter Learning	631
83.15.1	PL of $AR(p)$	632
83.15.2	PL of $MA(q)$	634
83.15.3	PL of $ARMA(p, q)$	636
83.16	$VAR(p)$	637
84	Transfer Learning	639
85	Transformer Networks	641
86	Transportability of Causal Knowledge	644
87	Turbo Codes	648
87.1	Decoding Algorithm	651
87.2	Message Passing Interpretation of Decoding Algorithm	653
88	Uplift Modelling	654
88.1	UP types	654
88.2	Some Relevant Technical Formulas from Chapter 64	655
88.3	UP Analysis	656
88.4	UP Decision Trees	658

88.4.1 Appendix, connection between Δ_c and $\Delta_{c j}$	663
89 Variational Bayesian Approximation for Medical Diagnosis	664
89.1 Convex/Concave Dual Functions	667
90 Variational Bayesian Approximation via D_{KL}	672
90.1 Free Energy $\mathcal{F}(\vec{x})$	674
91 XGBoost	677
91.1 Divergences	677
91.2 Minimizing Cost function for single tree	679
91.3 Leaf Splitting	682
91.4 Pruning	682
91.5 Feature Binning	684
91.6 Final estimate of target attribute	684
91.7 Bnet for XGBoost	685
92 Zero Information Transmission (Graphoid Axioms)	687
92.1 Consequences of Eq.(92.5)	688
Bibliography	690

Foreword

Welcome to Bayesuvius! a proto-book uploaded to github.

A different Bayesian network is discussed in each chapter. Each chapter title is the name of a Bnet. Chapter titles are in alphabetical order.

This is a volcano in its early stages. First version uploaded to a github repo called Bayesuvius on June 24, 2020. First version only covers 2 Bnets (Linear Regression and GAN). I will add more chapters periodically. Remember, this is a moon-lighting effort so I can't do it all at once.

For any questions about notation, please go to Notational Conventions section.
Requests and advice are welcomed.

Thanks for reading this

Robert R. Tucci

www.ar-tiste.xyz

ADDENDA

- **August 15, 2021:** At this point in time, the book has grown to 67 Chapters and 433 pages. Today, I am self-publishing it as an ebook at Amazon and similar outlets. It will still be free.

Appendices

Appendix A

Navigating the ocean of Judea Pearl's Books

Many of the greatest ideas in the bnet field were invented by Judea Pearl and his collaborators. Thus, this book is heavily indebted to those great scientists. Those ideas have had no clearer and more generous expositor than Judea Pearl himself.

Pearl has written 4 books that I have used in writing Bayesuvius. His 1988 book Ref.[44] dates back to his pre-causal period. That book I used to learn about topics such as d-separation, belief propagation, Markov-blankets, and noisy-ORs. 3 other books that he wrote later, in his causal period, are:

1. In 2000 (1st ed.), and 2013 (2nd ed.), Pearl published what is so far his most technical and exhaustive book on the subject of causality, Ref.[46].
2. In 2016, he released together with Glymour and Jewell, a less advanced “primer” on causality, Ref.[49].
3. In 2018, he released together with Mackenzie his lovely “The Book of Why”, Ref.[50].

Those 3 books I used to learn about causality topics such as Do Calculus, backdoor and frontdoor adjustment formulae, linear deterministic bnets with exogenous noise, and counterfactuals.

A micro poem written by me to celebrate Judea Pearl and his work:

I, Robot

Let other robots `talk()`,
while I,
`talk()`, `do()` and `imagine()`.

Appendix B

CI-2-3 track

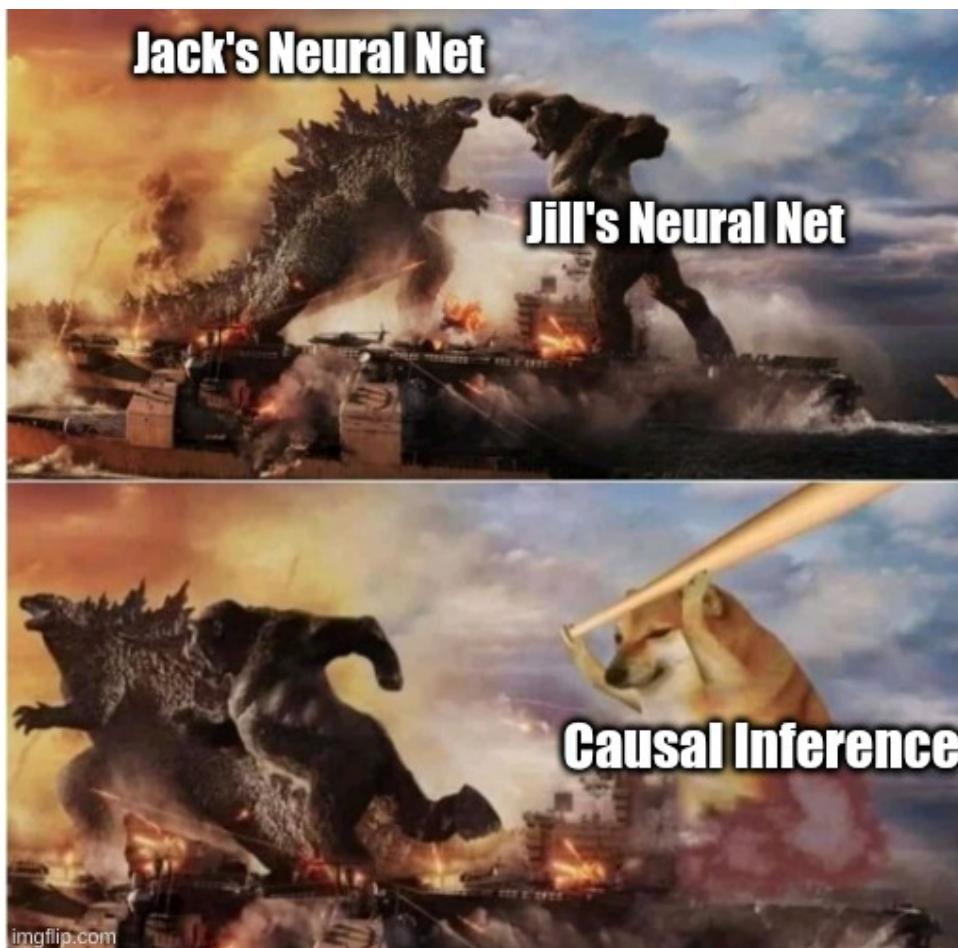


Figure B.1: CI meme

As discussed in Chapter 11, Judea Pearl has proposed 3 rungs of Causal Inference (CI). This book covers all 3 rungs.

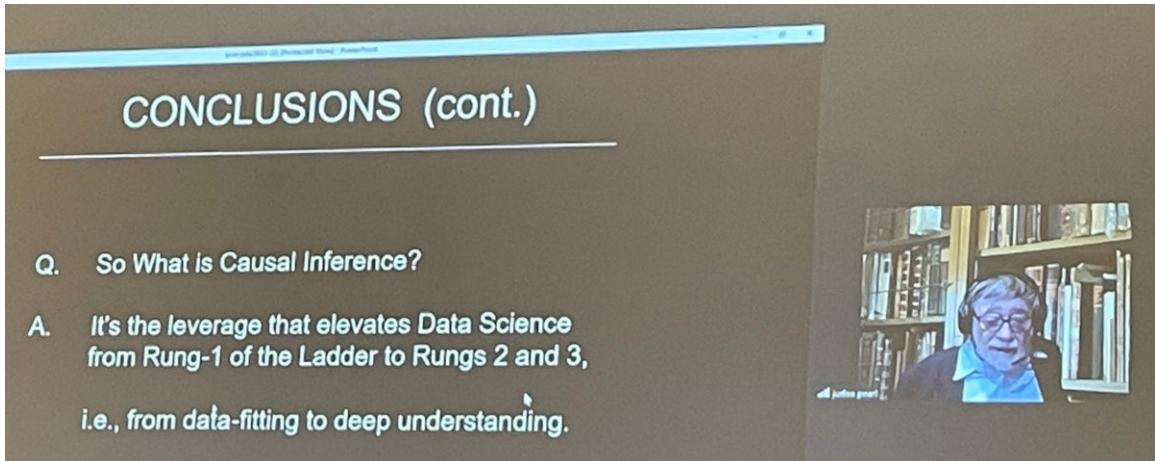


Figure B.2: Slide from Pearl talk at IJCAI-2022. Putting the joking of Fig.B.1 aside, let me emphasize that CI advocates are not trying to vanish NNs from AI. To us, NNs and bnets are different tools, like hammer and saw. We believe AI should use both tools. For those who are trying to do CI using a NN instead of a bnet, it looks to me like you are trying to use a hammer to cut wood. Why don't you cut with a saw instead? As Pearl says in this slide, CI elevates Data Science from Deep Learning (data-fitting) to Deep Understanding.

Confusingly, it has become common to use the term CI to refer to only the highest 2 rungs of the CI hierarchy; i.e, rung 2 (do operations) and rung 3 (imaging/counterfactual thinking). Also confusingly, rung 1 uses causal diagrams and is often referred to as “inference”, so it could reasonably have been defined as the whole of CI, but Pearl has defined the CI hierarchy to include two more rungs. To patch over this linguistic confusion, I sometimes refer to rung 1 as “prediction”, or as “predictive inference” instead of calling it merely “inference”. Also, when I want to be precise, I use the term “CI-2-3” to refer to CI restricted to only rungs 2 and 3.

Here is a subset of chapters that I call the CI-2-3 track, that are devoted mostly to rungs 2 and 3.

1. Backdoor Adjustment Formula
2. Berkson’s Paradox
3. Counterfactual Reasoning
4. Decisions Based on Rungs 2 and 3: COMING SOON
5. Difference-in-Differences
6. Do Calculus
7. Do Calculus proofs

8. D-Separation
9. D-Separation and Do Calculus Rules for LDEN: COMING SOON
10. Frisch-Waugh-Lovell (FWL) theorem
11. Frontdoor Adjustment Formula
12. G-formula (Sequential Backdoor Adjustment Formula)
13. Goodness of Causal Fit
14. Granger Causality
15. Instrumental Inequality and beyond
16. Instrumental Variables
17. LATE (Local Average Treatment Effect)
18. Mediation Analysis
19. Meta-learners for estimating ATE
20. Modified Treatment Policy
21. Personalized Treatment Effects
22. Personalized Expected Utility
23. Potential Outcomes and Beyond
24. Regression Discontinuity Design
25. Selection Bias Removal
26. Simpson's Paradox
27. Survival Analysis
28. Synthetic Controls
29. Targeted Estimator
30. Transportability of Causal Knowledge
31. Uplift Modelling

Appendix C

Notational Conventions and Preliminaries

C.1 Some abbreviations frequently used throughout this book

- bnet= Bnet= Bayesian Network
- CPT = Conditional Probabilities Table, same as TPM
- DAG = Directed Acyclic Graph
- i.i.d.= independent identically distributed.
- RCT= Randomized Controlled Trial, a.k.a. A/B testing.
- TPM= Transition Probability Matrix, same as CPT

C.2 $\mathcal{N}(!a)$

$\mathcal{N}(!a)$ will denote a normalization constant that does not depend on a . For example, $P(x) = \mathcal{N}(!x)e^{-x}$ where $\int_0^\infty dx P(x) = 1$.

C.3 One hot vector

A **one hot vector** is a vector with all entries equal to zero with the exception of a single entry which is one. A **one cold vector** is a vector with all entries equal to one with the exception of a single entry which is zero. For example, if $x^n = (x_0, x_1, \dots, x_{n-1})$ and $x_i = \delta(i, 0)$ then x^n is one hot.

Two types of sets that one frequently encounters are **categorical sets** (a.k.a. “nominal sets”, i.e., sets with “named” elements, with elements given a “name”)

and **numerical sets** (a.k.a. “ordinal sets”, i.e., sets with “ordered” elements). For example, $\{1, 2, 5\}$ is a numerical set because its elements have a natural order, and $\{\text{cat, dog, bird}\}$ is a categorical set because its elements don’t have a natural order.

In Machine Learning (ML), one often encodes categorical sets as one-hot vectors. For example, suppose we have 4 binary registers (i.e., nodes) x_3, x_2, x_1, x_0 and the categorical set $\{\text{cat, dog, canary}\}$. Then a possible **one-hot encoding** of the set is cat=0001, dog=0010 and canary=0100. This differs from a **binary encoding** of the set such as cat=0000, dog=0001, canary=0011. Clearly, a binary encoding requires fewer registers than a one-hot encoding to encode the same set, and the one-hot encoding of a set with n elements requires n or more registers.

C.4 Special sets

Define $\mathbb{Z}, \mathbb{R}, \mathbb{C}$ to be the integers, real numbers and complex numbers, respectively.

For $a < b$, define \mathbb{Z}_I to be the integers in the interval I , where $I = [a, b], [a, b), (a, b], (a, b)$ (i.e., I can be closed or open on either side).

$$A_{>0} = \{k \in A : k > 0\} \text{ for } A = \mathbb{Z}, \mathbb{R}.$$

C.5 Kronecker delta function

For x, y in discrete set S ,

$$\delta(x, y) = \begin{cases} 1 & \text{if } x = y \\ 0 & \text{if } x \neq y \end{cases} \quad (\text{C.1})$$

C.6 Dirac delta function

For $x, y \in \mathbb{R}$,

$$\int_{-\infty}^{+\infty} dx \delta(x - y) f(x) = f(y) \quad (\text{C.2})$$

C.7 Indicator function (a.k.a. Truth function)

$$\mathbb{1}(\mathcal{S}) = \begin{cases} 1 & \text{if } \mathcal{S} \text{ is true} \\ 0 & \text{if } \mathcal{S} \text{ is false} \end{cases} \quad (\text{C.3})$$

For example, $\delta(x, y) = \mathbb{1}(x = y)$.

C.8 Majority function

The **majority function** is defined as follows.

$$\text{majority}(L) = \begin{array}{l} \text{most common element of list } L \\ (\text{ties resolved by chance}) \end{array} \quad (\text{C.4})$$

Note that the majority function acts on lists, not sets. By definition, all elements of a set appear only once in the set. $\text{majority}(L)$ is usually used when the elements of L are categorical (i.e., not real numbers). When they are real numbers, it makes more sense to use, instead of $\text{majority}(L)$, a simple average of the elements of L .

C.9 Underlined letters indicate random variables

Random variables will be indicated by underlined letters and their values by non-underlined letters. Each node of a bnet will be labelled by a random variable. Thus, $\underline{x} = x$ means that node \underline{x} is in state x .

It is more conventional to use an upper case letter to indicate a random variable and a lower case letter for its state. Thus, $X = x$ means that random variable X is in state x . However, we have opted in this book to avoid that notation, because we often want to define certain lower case letters to be random variables or, conversely, define certain upper case letters to be non-random variables.

C.10 Probability distributions

$P_{\underline{x}}(x) = P(\underline{x} = x) = P(x)$ is the probability that random variable \underline{x} equals $x \in S_{\underline{x}}$. $S_{\underline{x}}$ is the set of states (i.e., values) that \underline{x} can assume and $n_{\underline{x}} = |S_{\underline{x}}|$ is the size (a.k.a. cardinality) of that set. Hence,

$$\sum_{x \in S_{\underline{x}}} P_{\underline{x}}(x) = 1 \quad (\text{C.5})$$

$$P_{\underline{x}, \underline{y}}(x, y) = P(\underline{x} = x, \underline{y} = y) = P(x, y) \quad (\text{C.6})$$

$$P_{\underline{x}|y}(x|y) = P(\underline{x} = x | \underline{y} = y) = P(x|y) = \frac{P(x, y)}{P(y)} \quad (\text{C.7})$$

C.11 Discretization of continuous probability distributions

The TPM of a node of a bnet can be either a discrete or a continuous probability distribution. To go from continuous to discrete, one replaces integrals over states of a node by sums over new states, and Dirac delta functions by Kronecker delta functions. More precisely, consider a function $f : [a, b] \rightarrow \mathbb{R}$. Express $[a, b]$ as a union of small, disjoint (except for one point) closed sub-intervals (bins) of length Δx . Name one

point in each bin to be the representative of that bin, and let $S_{\underline{x}}$ be the set of all the bin representatives. This is called discretization or binning. Then

$$\frac{1}{(b-a)} \int_{[a,b]} dx f(x) \rightarrow \frac{\Delta x}{(b-a)} \sum_{x \in S_{\underline{x}}} f(x) = \frac{1}{n_{\underline{x}}} \sum_{x \in S_{\underline{x}}} f(x). \quad (\text{C.8})$$

Both sides of last equation are 1 when $f(x) = 1$. Furthermore, if $y \in S_{\underline{x}}$, then

$$\int_{[a,b]} dx \delta(x-y)f(x) = f(y) \rightarrow \sum_{x \in S_{\underline{x}}} \delta(x,y)f(x) = f(y). \quad (\text{C.9})$$

As usual in this book, let $S_{\underline{x}}$ denote the set of values that the random variable \underline{x} can take. When $S_{\underline{x}} \subset \mathbb{R}$, we will assume that $S_{\underline{x}}$ for a probability distribution $P(x)$ can be either a discrete or a continuous subset of \mathbb{R} .¹ When $S_{\underline{x}}$ is a discrete subset of \mathbb{R} , $P(x)$ will denote a probability distribution for which $\sum_{x \in S_{\underline{x}}} P(x) = 1$, whereas when $S_{\underline{x}}$ is continuous, $P(x)$ will denote a probability density for which $\int_{x \in S_{\underline{x}}} dx P(x) = 1$.

C.12 Samples, i.i.d. variables

$$\vec{x} = (x[0], x[1], x[2] \dots, x[n_{\text{sam}}(\vec{x}) - 1]) = x[:] \quad (\text{C.10})$$

$n_{\text{sam}}(\vec{x})$ is the number of samples of \vec{x} . $x[\sigma] \in S_{\underline{x}}$ are i.i.d. (independent identically distributed) samples with

$$x[\sigma] \sim P_{\underline{x}} \quad (\text{i.e. } P_{x[\sigma]} = P_{\underline{x}}) \quad (\text{C.11})$$

$$P(\underline{x} = x) = \frac{1}{n_{\text{sam}}(\vec{x})} \sum_{\sigma} \mathbb{1}(x[\sigma] = x) \quad (\text{C.12})$$

Hence, for any $f : S_{\underline{x}} \rightarrow \mathbb{R}$,

$$\sum_x P(\underline{x} = x) f(x) = \frac{1}{n_{\text{sam}}(\vec{x})} \sum_{\sigma} f(x[\sigma]) \quad (\text{C.13})$$

If we use two sampled variables, say \vec{x} and \vec{y} , in a given bnet, their number of samples $n_{\text{sam}}(\vec{x})$ and $n_{\text{sam}}(\vec{y})$ need not be equal.

$$P(\vec{x}) = \prod_{\sigma} P(x[\sigma]) \quad (\text{C.14})$$

$$\sum_{\vec{x}} = \prod_{\sigma} \sum_{x[\sigma]} \quad (\text{C.15})$$

¹By a “continuous set” we mean a finite set of intervals each of which has non-zero length.

$$\partial_{\vec{x}} = [\partial_{x[0]}, \partial_{x[1]}, \partial_{x[2]}, \dots, \partial_{x[n sam(\vec{x})-1]}] \quad (\text{C.16})$$

$$P(\vec{x}) \approx [\prod_x P(x)^{P(x)}]^{n sam(\vec{x})} \quad (\text{C.17})$$

$$= e^{n sam(\vec{x}) \sum_x P(x) \ln P(x)} \quad (\text{C.18})$$

$$= e^{-n sam(\vec{x}) H(P_{\underline{x}})} \quad (\text{C.19})$$

C.13 Expected Value and Variance

Given a random variable \underline{x} with states $S_{\underline{x}}$ and a function $f : S_{\underline{x}} \rightarrow \mathbb{R}$, define

$$E_{\underline{x}}[f(\underline{x})] = E_{x \sim P(x)}[f(x)] = \sum_x P(x)f(x) \quad (\text{C.20})$$

$$Var_{\underline{x}}[f(\underline{x})] = E_{\underline{x}}[(f(\underline{x}) - E_{\underline{x}}[f(\underline{x})])^2] \quad (\text{C.21})$$

$$= E_{\underline{x}}[f(\underline{x})^2] - (E_{\underline{x}}[f(\underline{x})])^2 \quad (\text{C.22})$$

$$E[\underline{x}] = E_{\underline{x}}[\underline{x}] \quad (\text{C.23})$$

$$Var[\underline{x}] = Var_{\underline{x}}[\underline{x}] \quad (\text{C.24})$$

C.14 Conditional Expected Value

Given a random variable \underline{x} with states $S_{\underline{x}}$, a random variable \underline{y} with states $S_{\underline{y}}$, and a function $f : S_{\underline{x}} \times S_{\underline{y}} \rightarrow \mathbb{R}$, define

$$E_{\underline{x}|\underline{y}}[f(\underline{x}, \underline{y})] = \sum_x P(x|\underline{y})f(x, \underline{y}), \quad (\text{C.25})$$

$$E_{\underline{x}|\underline{y}=y}[f(\underline{x}, y)] = E_{\underline{x}|y}[f(\underline{x}, y)] = \sum_x P(x|y)f(x, y). \quad (\text{C.26})$$

Note that

$$E_{\underline{y}}[E_{\underline{x}|\underline{y}}[f(\underline{x}, \underline{y})]] = \sum_{x,y} P(x|y)P(y)f(x, y) \quad (\text{C.27})$$

$$= \sum_{x,y} P(x, y)f(x, y) \quad (\text{C.28})$$

$$= E_{\underline{x}, \underline{y}}[f(\underline{x}, \underline{y})]. \quad (\text{C.29})$$

C.15 Notation for covariances

Consider two random variables $\underline{x}, \underline{y}$.

- Mean value of \underline{x}

$$\langle \underline{x} \rangle = E_{\underline{x}}[\underline{x}] \quad (\text{C.30})$$

- Signed distance of \underline{x} to its mean value

$$\Delta \underline{x} = \underline{x} - \langle \underline{x} \rangle \quad (\text{C.31})$$

- Covariance of $(\underline{x}, \underline{y})$

$$Cov(\underline{x}, \underline{y}) = \langle \underline{x}, \underline{y} \rangle = \langle \Delta \underline{x} \Delta \underline{y} \rangle = \langle \underline{x} \underline{y} \rangle - \langle \underline{x} \rangle \langle \underline{y} \rangle \quad (\text{C.32})$$

$\langle \underline{x}, \underline{y} \rangle$ is symmetric (i.e., $\langle \underline{x}, \underline{y} \rangle = \langle \underline{y}, \underline{x} \rangle$) and bilinear (i.e., $\langle \sum_i \alpha_i \underline{x}_i, \underline{y} \rangle = \sum_i \alpha_i \langle \underline{x}_i, \underline{y} \rangle$, where $\alpha_i \in \mathbb{R}$ are non-random scalars and $\underline{x}_i, \underline{y} \in \mathbb{R}$ are real-valued random variables.)

- Variance of \underline{x}

$$Var(\underline{x}) = \langle \underline{x}, \underline{x} \rangle \quad (\text{C.33})$$

- Standard deviation or \underline{x}

$$\sigma_{\underline{x}} = \sqrt{\langle \underline{x}, \underline{x} \rangle} \quad (\text{C.34})$$

- Correlation Coefficient of $(\underline{x}, \underline{y})$

$$\rho_{\underline{x}, \underline{y}} = \frac{\langle \underline{x}, \underline{y} \rangle}{\sqrt{\langle \underline{x}, \underline{x} \rangle \langle \underline{y}, \underline{y} \rangle}} \quad (\text{C.35})$$

C.16 Conditional Covariance

Let $\underline{x}, \underline{y}, \underline{a}$ be random variables. The covariance $Cov(\underline{x}, \underline{y}|\underline{a})$ of \underline{x} and \underline{y} given \underline{a} , is defined the same way as $Cov(\underline{x}, \underline{y})$, except that all expected values are conditioned on \underline{a} .

$$Cov(\underline{x}, \underline{y}|\underline{a}) = \langle \underline{x}, \underline{y} \rangle_{|\underline{a}} = \left\langle (\underline{x} - \langle \underline{x} \rangle_{|\underline{a}})(\underline{y} - \langle \underline{y} \rangle_{|\underline{a}}) \right\rangle_{|\underline{a}} \quad (\text{C.36})$$

where

$$\langle \underline{x} \rangle_{|\underline{a}} = E_{\underline{x}|\underline{a}}[\underline{x}] . \quad (\text{C.37})$$

C.17 Normal Distribution

For $x, \mu, \sigma \in \mathbb{R}$, $\sigma > 0$, we define the Normal Distribution (see Fig.C.1) by

$$\mathcal{N}(x; \mu, \sigma^2) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}. \quad (\text{C.38})$$

For a **standard deviation** σ , the **precision** τ is defined as $\tau = \frac{1}{\sigma^2}$.

Claim 1 *If*

$$\underline{x}_1 \sim \mathcal{N}(\mu_1, \sigma_1^2) \quad (\text{C.39})$$

and

$$\underline{x}_2 \sim \mathcal{N}(\mu_2, \sigma_2^2) \quad (\text{C.40})$$

then

$$\underline{x} = \underline{x}_1 + \underline{x}_2 \sim \mathcal{N}(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2). \quad (\text{C.41})$$

proof:

$$P(\underline{x} = x) = \mathcal{N}(!x) \int_{-\infty}^{+\infty} dx_2 P(\underline{x}_1 + \underline{x}_2 = x | \underline{x}_2 = x_2) P(x_2) \quad (\text{C.42})$$

$$= \mathcal{N}(!x) \int_{-\infty}^{+\infty} dx_2 \mathcal{N}(x - x_2; \mu_1, \sigma_1^2) \mathcal{N}(x_2; \mu_2, \sigma_2^2) \quad (\text{C.43})$$

$$= \mathcal{N}(x; \mu_1 + \mu_2; \sigma_1^2 + \sigma_2^2) \quad (\text{C.44})$$

QED

The **Standard Normal Distribution** $P_{SND}(x)$ and its cumulative distribution $\Phi(x)$ are defined by

$$P_{SND}(x) = \mathcal{N}(x; \mu = 0, \sigma = 1) \quad (\text{C.45})$$

$$\Phi(x) = \int_{-\infty}^x dx' P_{SND}(x') \quad (\text{C.46})$$

The **error function** $\text{erf} : \mathbb{R} \rightarrow [-1, 1]$ is defined by

$$\text{erf}(x) = \frac{2}{\sqrt{\pi}} \int_0^x du e^{-\frac{u^2}{2}} \quad (\text{C.47})$$

Note that

$$\Phi(x) = \frac{1}{2} + \frac{1}{2}\text{erf}\left(\frac{x}{\sqrt{2}}\right) \quad (\text{C.48})$$

Eq.(C.48) is interpreted geometrically in Fig.C.2.

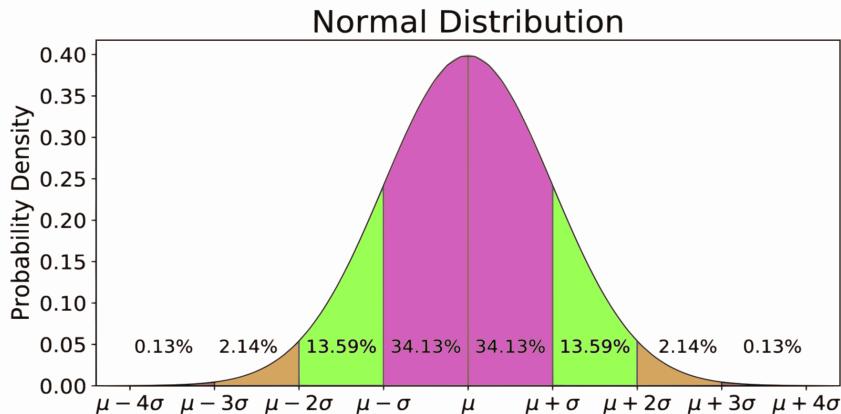


Figure C.1: Normal Distribution $\mathcal{N}(x; \mu, \sigma^2)$.

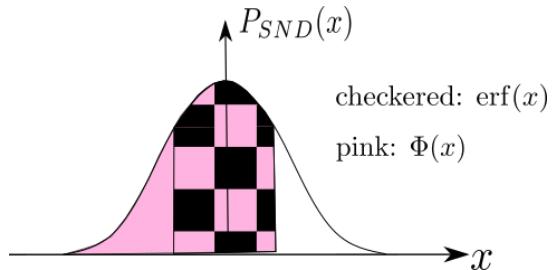


Figure C.2: Plot of Standard Normal Distribution $P_{SND}(x)$. Values of $\text{erf}(x)$ and $\Phi(x)$ equal indicated areas.

C.18 Uniform Distribution

For $a < b$, $x \in [a, b]$

$$\mathcal{U}(x; a, b) = \frac{1}{b - a} \quad (\text{C.49})$$

C.19 Softmax function (a.k.a. Boltzmann Distribution)

The Softmax function is defined by

$$P(x_i|x.) = \frac{e^{x_i}}{\sum_i e^{x_i}} = \text{softmax}(x.)_i \quad (\text{C.50})$$

The Boltzmann distribution is defined as

$$P(\underline{E}_a = E_a) = \frac{\exp(-\frac{E_a}{kT})}{\sum_a \exp(-\frac{E_a}{kT})} = P(\frac{-E_a}{kT} | E.) \quad (\text{C.51})$$

for a system with energies E_a and temperature T , where k is Boltzmann's constant.

The function softmax() is called softmax because if we approximate the exponentials, both in the numerator and denominator of Eq.(C.50), by the largest one of them or zero, we get

$$\text{softmax}(x.)_i \approx \mathbb{1}(i = \underset{k}{\text{argmax}} x_k) . \quad (\text{C.52})$$

Thus, softmax($x.$) returns a continuous function that approximates a one-hot vector that is 1 at the i th component, where $i = \underset{k}{\text{argmax}}(x_k)$, and zero at the other components.

Note that

$$\frac{\partial \ln P(x_i|x)}{\partial x_a} = \frac{\partial}{\partial x_a} \ln \left[\frac{e^{x_i}}{\sum_i e^{x_i}} \right] = \delta(a, i) - P(x_a|x) \quad (\text{C.53})$$

For 2 variables x_0, x_1 ,

$$P(x_0|x.) = \frac{e^{x_0}}{e^{x_0} + e^{x_1}} \quad (\text{C.54})$$

$$= \text{smoid}(x_0 - x_1) , \quad (\text{C.55})$$

$$P(x_1|x.) = \text{smoid}(x_1 - x_0) . \quad (\text{C.56})$$

C.20 Sigmoid and log-odds functions

The **sigmoid (a.k.a. exp-it, logistic) function** smoid: $\mathbb{R} \rightarrow [0, 1]$ is defined by

$$\text{smoid}(x) = \frac{1}{1 + e^{-x}} \quad (\text{C.57})$$

smoid() is monotonically increasing with $\text{smoid}(-\infty) = 0$, $\text{smoid}(0) = 1/2$ and $\text{smoid}(+\infty) = 1$. Note that for $x \ll 0$, $\text{smoid}(x) \approx e^x$, which is why "smoid" is also called "expit".

$$\text{smoid}(x) + \text{smoid}(-x) = \frac{1}{1 + e^{-x}} + \frac{1}{1 + e^x} \quad (\text{C.58})$$

$$= \frac{2 + e^x + e^{-x}}{2 + e^x + e^{-x}} \quad (\text{C.59})$$

$$= 1 \quad (\text{C.60})$$

The **log-odds (a.k.a. log-it) function** $\text{lodds}:[0, 1] \rightarrow \mathbb{R}$ is defined by

$$\text{lodds}(p) = \ln \frac{p}{1-p} \quad (\text{C.61})$$

Note that for $0 < p \ll 1$, $\text{lodds}(x) \approx \ln p$, which is why “lodds” is also called “logit”.

Note that for $x \ll 1$, $\text{smoid}(x) \approx e^x \ll 1$, so $\text{lodds}(e^x) \approx \ln(e^x) = x$. More generally, it is easy to check that for any $p \in [0, 1]$ and $x \in \mathbb{R}$,

$$\text{lodds}[\text{smoid}(x)] = x \quad (\text{C.62})$$

$$\text{smoid}[\text{lodds}(p)] = p \quad (\text{C.63})$$

Hence, $\text{lodds}()$ is the inverse of $\text{smoid}()$ and vice-versa.

Claim 2

$$\text{smoid}'(x) = \text{smoid}(x)[1 - \text{smoid}(x)] \quad (\text{C.64})$$

$$\text{smoid}''(x) = \text{smoid}'(x)[1 - 2\text{smoid}(x)] \quad (\text{C.65})$$

proof:

In this proof, we will abbreviate $\text{smoid}(x)$ by $s(x)$.

$$1 - s(x) = 1 - \frac{1}{1 + e^{-x}} = \frac{e^{-x}}{1 + e^{-x}} \quad (\text{C.66})$$

$$s'(x) = \frac{e^{-x}}{(1 + e^{-x})^2} = s(x)[1 - s(x)] \quad (\text{C.67})$$

$$s''(x) = s'(x)[1 - s(x)] + s(x)(-1)s'(x) \quad (\text{C.68})$$

$$= s'(x)[1 - 2s(x)] \quad (\text{C.69})$$

$$= s(x)[1 - s(x)][1 - 2s(x)] \quad (\text{C.70})$$

QED

C.21 Estimand, Estimator (curve-fit, cfit), Estimate, Bias

For an **estimand** θ , an **estimator (a.k.a. curve fit, cfit)** $\hat{\theta}$ gives **estimate** $E[\hat{\theta}(\theta)] = \theta + b$ with **bias** b . We say this estimate is an **unbiased estimate** if $b = 0$.

Note that, strictly speaking, an estimator is a function waiting to be averaged over and denoted by a letter with a hat, whereas an estimate is a real number denoted by a letter without a hat. Unfortunately, the words “estimator” and “estimate” are often used interchangeably, as if they were synonyms. And often the estimate $\theta + b$ is denoted by a letter with a hat too. In some sense, an estimator is an estimate of a curve, so it’s understandable that the terms “estimator” and “estimate” are used synonymously. In this book, we will bow to traditional practice and also use the terms “estimator” and “estimate” synonymously, and use a letter with a hat to denote either of them. This is not ambiguous as long as we don’t use the same letter with a hat to denote two different quantities, of course. When we need to distinguish semantically between the real value and the function, we will call the function a cfit, and the real value the estimate.

C.22 Maximum Likelihood Estimate, Likelihood Ratio Test

Given a bnet, let $P(x|\theta)$ be its full joint probability distribution, where x denotes the joint state of all the nodes and θ denotes all the parameters. $P(x|\theta)$ is often called the **likelihood function of θ** and is denoted by

$$L(\theta) = P(x|\theta) \quad (\text{C.71})$$

It’s called a likelihood of θ because, even though it’s a probability, it isn’t the probability of θ , but rather of x .

The value of θ that we obtain by maximizing $L(\theta)$ over θ is called the **maximum likelihood estimate (MLE) of θ** . Let us denote it by $\hat{\theta}$. Note that²

$$\sup_{\theta \in S} L(\theta) = L(\hat{\theta}) \quad (\text{C.72})$$

Let S_0, S_1 be disjoint sets such that $S = S_0 \cup S_1$. We’ll say the **null hypothesis** H_0 holds if $\theta \in S_0$, and the **alternative hypothesis** H_1 holds if $\theta \in S_1$. The **likelihood ratio (LR) test statistic** is defined by

$$R = -2 \ln \left(\frac{\sup_{\theta \in S_0} L(\theta)}{\sup_{\theta \in S} L(\theta)} \right) \quad (\text{C.73})$$

$R \geq 0$ and $R = 0$ if $S_0 = S$. For some small $c > 0$, if $R < c$, then we reject the alternative hypothesis, and if $R > c$, we accept it.

If $S_0 = \{\theta_0\}$, then

²“sup” stands for supremum. It’s a generalization of the function `max()` to arbitrary sets that might not be discrete or finite. If S is a finite set, then $\sup_{\theta \in S} f(\theta) = \max_{\theta \in S} f(\theta)$ for any function $f : S \rightarrow \mathbb{R}$. Likewise, “inf” stands for infimum, and it generalizes the `min()` function.

$$R = -2[\ln L(\theta_0) - \ln L(\hat{\theta})] \quad (\text{C.74})$$

C.23 Mean Square Error (MSE)

Suppose we are given $nsam$ samples $y^\sigma \in \mathbb{R}$ labeled by an index σ , and a cfit $\hat{y}^\sigma(a) \in \mathbb{R}$ that depends on a parameter $a \in \mathbb{R}$. Define the **Mean Square Error (MSE)** by

$$MSE(a) = \frac{1}{nsam} \sum_{\sigma} (y^\sigma - \hat{y}^\sigma(a))^2. \quad (\text{C.75})$$

For example, in Linear Regression (LR), we have $\hat{y}^\sigma = a_0 + a_1 x^\sigma$ where $a = (a_0, a_1)$ is a deterministic parameter. If the samples y^σ are i.i.d, then we can also write

$$MSE(a) = E_{|a}[(\underline{y} - \hat{y}(a))^2]. \quad (\text{C.76})$$

and for LR, $\hat{y}(a) = a_0 + a_1 \underline{x}$.

Define the **residual** $\Delta\underline{y}$ by:

$$\Delta\underline{y}(a) = \underline{y} - \hat{y}(a) \quad (\text{Hence } \underline{y} = \hat{y} + \Delta\underline{y}) \quad (\text{C.77})$$

In the rest of this section, we will discuss the case that $\hat{y}^\sigma(a)$ is independent of x^σ . I call this the **deterministic MSE (D-MSE)** model. Note that this is different from the LR case where $\hat{y}^\sigma(a)$ does depend on x^σ . In LR, we are trying to fit a line to a cigar-shaped 2-D scatter plot. Here, we are just trying to estimate the mean value (center of mass) of a scatter plot.

Claim 3 *MSE is minimized over all functions \hat{y} if*

$$\hat{y} = E_{|a}[\underline{y}] \quad (\text{C.78})$$

proof:

$$MSE = E_{|a}[\underline{y}^2] - 2\hat{y}E_{|a}[\underline{y}] + \hat{y}^2 \quad (\text{C.79})$$

$$0 = \frac{d}{d\hat{y}} MSE = 2(-E_{|a}[\underline{y}] + \hat{y}) \quad (\text{C.80})$$

Hence,

$$\hat{y} = E_{|a}[\underline{y}] \quad (\text{C.81})$$

QED

Sometimes, we will use the notation

$$\hat{y}_{MSE} = E_{|a=a_{MSE}}[\underline{y}]. \quad (\text{C.82})$$

Claim 4 Suppose $f(a)$ is a function of a . If $\hat{y} = E_{|a}[\underline{y}]$, then

$$E_{|a}[\Delta\underline{y}] = E[\Delta\underline{y}] = 0 \quad (\text{C.83})$$

$$E_{|a} [\Delta\underline{y}f(\underline{a})] = E [\Delta\underline{y}f(\underline{a})] = 0 \quad (\text{C.84})$$

proof:

$$E_{|a}[\Delta\underline{y}] = E_{|a} [y - E_{|a}[\underline{y}]] = E_{|a}[y] - E_{|a}[\underline{y}] = 0 \quad (\text{C.85})$$

$$E[\Delta\underline{y}] = E_{\underline{a}}[E_{|\underline{a}}[\Delta\underline{y}]] = 0 \quad (\text{C.86})$$

$$E_{|\underline{a}} [\Delta\underline{y}f(\underline{a})] = f(\underline{a}) \underbrace{E_{|\underline{a}}[\Delta\underline{y}]}_{=0} \quad (\text{C.87})$$

$$E[\Delta\underline{y}f(\underline{a})] = E_{\underline{a}}[E_{|\underline{a}}[\Delta\underline{y}f(\underline{a})]] = 0 \quad (\text{C.88})$$

QED

Claim 5 If $\hat{y} = E_{|a}[\underline{y}]$, then

$$\langle \Delta\underline{y}, \hat{y} \rangle_{|a} = 0 \quad (\text{C.89})$$

$$Var_{|a}[\underline{y}] = Var_{|a}[\hat{y}] + Var_{|a}[\Delta\underline{y}] \quad (\text{C.90})$$

The same results hold without the conditioning on a .

proof:

$$\langle \Delta\underline{y}, \hat{y} \rangle_{|a} = \underbrace{E_{|a}[\Delta\underline{y} \underbrace{\hat{y}}_{f(\underline{a})}]}_{=0} - \underbrace{E_{|a}[\Delta\underline{y}] E_{|a}[\hat{y}]}_{=0} \quad (\text{C.91})$$

$$Var_{|a}[\underline{y}] = \langle \hat{y} + \Delta\underline{y}, \hat{y} + \Delta\underline{y} \rangle_{|a} \quad (\text{C.92})$$

$$= \langle \hat{y}, \hat{y} \rangle_{|a} + \langle \Delta\underline{y}, \Delta\underline{y} \rangle_{|a} \quad (\text{by Eq.(C.89)}) \quad (\text{C.93})$$

$$= Var_{|a}[\hat{y}] + Var_{|a}[\Delta\underline{y}] \quad (\text{C.94})$$

The same proof holds if we remove all the $|a$ subscripts.

QED

Fig.C.3 illustrates how $\underline{y} = \hat{y} + \Delta\underline{y}$ and the variances of these quantities add.

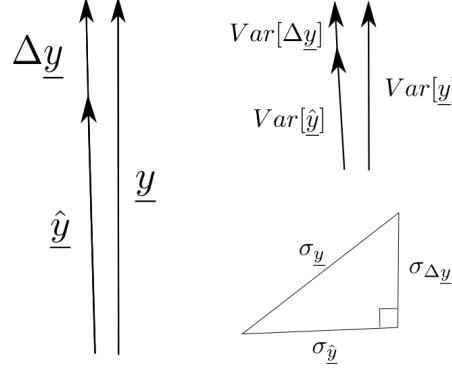


Figure C.3: $\underline{y} = \hat{y} + \Delta\underline{y}$ and the variances (not standard deviations) of these quantities add.

C.24 Cramer-Rao Bound

This discussion of the Cramer-Rao (CR) bound is based on Ref.[92].

Suppose \underline{x} is a random variable with values $x \in S_{\underline{x}}$ and $\theta \in \mathbb{R}$ is a parameter. For any function $f_{\underline{x},\theta} : S_{\underline{x}} \times \mathbb{R} \rightarrow \mathbb{R}$, define

$$\langle f_{\underline{x},\theta} \rangle = \sum_x P(x|\theta) f_{x,\theta} \quad (\text{C.95})$$

$$\Delta f_{\underline{x},\theta} = f_{\underline{x},\theta} - \langle f_{\underline{x},\theta} \rangle \quad (\text{C.96})$$

$$\langle f_{\underline{x},\theta}, f_{\underline{x},\theta} \rangle = \langle \Delta f_{\underline{x},\theta} \Delta f_{\underline{x},\theta} \rangle \quad (\text{C.97})$$

Define the **log likelihood function** by

$$LL_\theta = \ln P(x|\theta) \quad (\text{C.98})$$

Define the **Fisher information** by

$$I_\theta = \langle \partial_\theta LL_\theta, \partial_\theta LL_\theta \rangle \quad (\text{C.99})$$

Note that $LL_\theta \leq 0$. Let θ^* be the value of θ that maximizes LL_θ .

Note that $I_\theta \geq 0$ and $I_\theta = 0$ when $\theta = \theta^*$ because $\partial_\theta LL_\theta|_{\theta=\theta^*} = 0$. This suggests that I_θ measures the distance between θ and θ^* .

Note that

$$\langle \partial_\theta LL_\theta \rangle = \sum_x P(x|\theta) \frac{1}{P(x|\theta)} \partial_\theta P(x|\theta) \quad (\text{C.100})$$

$$= \partial_\theta \sum_x P(x|\theta) \quad (\text{C.101})$$

$$= 0 \quad (\text{C.102})$$

Therefore

$$I_\theta = \langle [\partial_\theta LL_\theta]^2 \rangle - \langle \partial_\theta LL_\theta \rangle^2 \quad (\text{C.103})$$

$$= \langle [\partial_\theta LL_\theta]^2 \rangle \quad (\text{C.104})$$

Claim 6

$$I_\theta = -\langle \partial_\theta^2 LL_\theta \rangle \quad (\text{C.105})$$

proof:

$$I_\theta = \langle [\partial_\theta LL_\theta]^2 \rangle \quad (\text{C.106})$$

$$= \sum_x P(x|\theta) \frac{1}{P(x|\theta)} \partial_\theta P(x|\theta) \partial_\theta \ln P(x|\theta) \quad (\text{C.107})$$

$$= -\sum_x P(x|\theta) \partial_\theta^2 \ln P(x|\theta) + \partial_\theta \sum_x P(x|\theta) \partial_\theta \ln P(x|\theta) \quad (\text{C.108})$$

$$= -\langle \partial_\theta^2 LL_\theta \rangle + \partial_\theta^2 \sum_x P(x|\theta) \quad (\text{C.109})$$

$$= -\langle \partial_\theta^2 LL_\theta \rangle \quad (\text{C.110})$$

QED

Claim 7 If $x = [x_i]_{i=1,2,\dots,\nu} \in \mathbb{R}^\nu$ are i.i.d., then

$$I_\theta = \nu \langle [\partial_\theta LL_{\theta,i}]^2 \rangle \quad (\text{C.111})$$

where

$$LL_{\theta,i} = \ln P(x_i|\theta) \quad (\text{C.112})$$

proof:

$$LL_\theta = \ln \prod_i P(x_i|\theta) \quad (\text{C.113})$$

$$= \sum_i LL_{\theta,i} \quad (\text{C.114})$$

$$I_\theta = \sum_i \sum_j \langle \partial_\theta LL_{\theta,i} \partial_\theta LL_{\theta,j} \rangle \quad (\text{C.115})$$

$$= \sum_i \langle [\partial_\theta LL_{\theta,i}]^2 \rangle \quad (\text{C.116})$$

$$= \nu \langle [\partial_\theta LL_{\theta,i}]^2 \rangle \quad (\text{C.117})$$

QED

A function $t_{\underline{x}} : S_{\underline{x}} \rightarrow \mathbb{R}$ is called a **test statistic** of random variable \underline{x} .

Claim 8 (*Cramer-Rao bound for single parameter $\theta \in \mathbb{R}$*)

$$\langle t_{\underline{x}}, t_{\underline{x}} \rangle I_{\theta} \geq [\partial_{\theta} \langle t_{\underline{x}} \rangle]^2 \quad (\text{C.118})$$

proof:

Cauchy-Schwartz inequality

For two vectors $\vec{a}, \vec{b} \in \mathbb{R}^n$:

$$\vec{a} \cdot \vec{b} = |\vec{a}| |\vec{b}| \cos \phi \leq |\vec{a}| |\vec{b}| \quad (\text{C.119})$$

For two real valued random variables $\underline{a}, \underline{b}$:

$$\langle \underline{a}, \underline{a} \rangle \langle \underline{b}, \underline{b} \rangle \geq |\langle \underline{a}, \underline{b} \rangle|^2 \quad (\text{C.120})$$

Replace

$$\underline{a} \rightarrow t_{\underline{x}}, \quad \underline{b} \rightarrow \partial_{\theta} LL_{\theta} \quad (\text{C.121})$$

Then

$$\langle t_{\underline{x}}, \partial_{\theta} LL_{\theta} \rangle = \langle t_{\underline{x}} \partial_{\theta} LL_{\theta} \rangle - \underbrace{\langle t_{\underline{x}} \rangle}_{=0} \langle \partial_{\theta} LL_{\theta} \rangle \quad (\text{C.122})$$

$$= \sum_x P(x|\theta) t_{\underline{x}} \frac{1}{P(x|\theta)} \partial_{\theta} P(x|\theta) \quad (\text{C.123})$$

$$= \partial_{\theta} \sum_x t_{\underline{x}} P(x|\theta) \quad (\text{C.124})$$

$$= \partial_{\theta} \langle t_{\underline{x}} \rangle \quad (\text{C.125})$$

QED

See Fig.C.4 for a pictorial representation of Eq.(C.118).

Now suppose the test statistic $t_{\underline{x}}$ equals an **estimator** $\hat{\theta}$ of θ with bias $b_{\underline{x}}$: $S_{\underline{x}} \rightarrow \mathbb{R}$.

$$t_{\underline{x}} = \hat{\theta}(\underline{x}) = \theta + b_{\underline{x}} \quad (\text{C.126})$$

$\hat{\theta}$ is said to be a **biased estimator** if $b_{\underline{x}} \neq 0$ and an **unbiased estimator** if $b_{\underline{x}} = 0$.

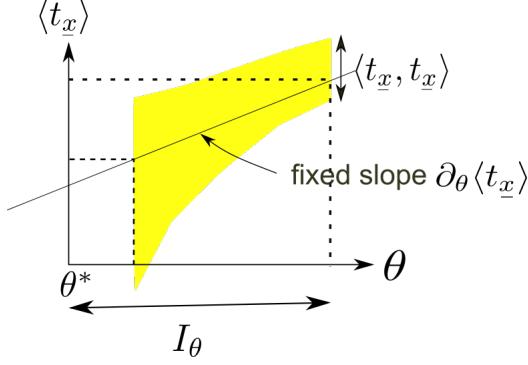


Figure C.4: In this drawing, θ^* is the value of θ that maximizes LL_θ . According to the CR bound, the product of the variance $\langle t_{\underline{x}}, t_{\underline{x}} \rangle$ and the distance I_θ must be greater or equal to $[\partial_\theta \langle t_{\underline{x}} \rangle]^2$. At fixed $[\partial_\theta \langle t_{\underline{x}} \rangle]^2$, if the variance increases, the distance decreases, and vice versa.

Claim 9

$$\langle \widehat{\theta} \rangle = \theta + \langle b_{\underline{x}} \rangle \quad (\text{C.127})$$

$$\langle \widehat{\theta}, \widehat{\theta} \rangle \geq \frac{[1 + \partial_\theta \langle b_{\underline{x}} \rangle]^2}{I_\theta} \quad (\text{C.128})$$

$$\langle [\widehat{\theta} - \theta]^2 \rangle \geq \frac{[1 + \partial_\theta \langle b_{\underline{x}} \rangle]^2}{I_\theta} + \langle b_{\underline{x}} \rangle^2 \quad (\text{C.129})$$

proof:

$$\partial_\theta \langle t_{\underline{x}} \rangle = \partial_\theta \langle \theta + b_{\underline{x}} \rangle \quad (\text{C.130})$$

$$= \partial_\theta [\theta + \langle b_{\underline{x}} \rangle] \quad (\text{C.131})$$

$$= 1 + \partial_\theta \langle b_{\underline{x}} \rangle \quad (\text{C.132})$$

Eq.(C.128) follows from Eq.(C.118) once we replace $t_{\underline{x}}$ by $\widehat{\theta}$.

Let

$$\Delta\widehat{\theta} = \widehat{\theta} - \langle \widehat{\theta} \rangle = \underbrace{(\widehat{\theta} - \theta)}_{\xi} - \langle b_{\underline{x}} \rangle \quad (\text{C.133})$$

Then

$$0 = \langle \Delta\widehat{\theta} \rangle = \langle \xi \rangle - \langle b_{\underline{x}} \rangle \quad (\text{C.134})$$

$$\frac{[1 + \partial_\theta \langle b_{\underline{x}} \rangle]^2}{I_\theta} \leq \langle [\Delta\widehat{\theta}]^2 \rangle \quad (\text{C.135})$$

$$= \langle \xi^2 - 2\xi \langle b_{\underline{x}} \rangle + \langle b_{\underline{x}} \rangle^2 \rangle \quad (\text{C.136})$$

$$= \langle \xi^2 \rangle - \langle b_{\underline{x}} \rangle^2 \quad (\text{C.137})$$

QED

Multi-dimensional case: parameter $\theta = [\theta_1, \theta_2, \dots, \theta_n]^T \in \mathbb{R}^n$ and test statistic $t_{\underline{x}} = [t_{\underline{x},1}, t_{\underline{x},2}, \dots, t_{\underline{x},n}]^T \in \mathbb{R}^n$ are column vectors.

Define **Fisher information matrix** by

$$[I_\theta]_{i,j} = \langle \partial_{\theta_i} LL_\theta, \partial_{\theta_j} LL_\theta \rangle = \langle \partial_{\theta_i} LL_\theta \partial_{\theta_j} LL_\theta \rangle \quad (\text{C.138})$$

CR bound for multi-dimensional parameter $\theta \in \mathbb{R}^n$:

$$\text{matrix} [\langle t_{\underline{x},i}, t_{\underline{x},j} \rangle] \geq \text{matrix} [\partial_{\theta_i} \langle t_{\underline{x},a} \rangle [I_\theta]_{a,b}^{-1} \partial_{\theta_j} \langle t_{\underline{x},b} \rangle] \quad (\text{C.139})$$

where we are using the Einstein summation convention (repeated indices are summed over). For two matrices $A, B \in \mathbb{R}^n$, $A \geq B$ means $A - B$ has non-negative eigenvalues.

C.25 Bayes Rule, Bayesian Updating And Conjugate Priors

Bayes Rule says:

$$P(\theta|x)P(x) = P(x|\theta)P(\theta) \quad (\text{C.140})$$

Expressed diagrammatically³, we have for $\underline{x} \in \mathbb{R}$:

$$\underline{\theta} \xleftarrow{} \underline{x} = \underline{\theta} \xrightarrow{} \underline{x} \quad (\text{C.141})$$

and for $\underline{x} = (x_1, \underline{x}_2) \in \mathbb{R}^2$:

$$\begin{array}{ccc} \begin{array}{c} \underline{\theta} \\ \swarrow \quad \searrow \\ x_1 \\ \downarrow \\ \underline{x}_2 \end{array} & = & \begin{array}{c} \underline{\theta} \\ \swarrow \quad \searrow \\ x_1 \\ \downarrow \\ \underline{x}_2 \end{array} \end{array} \quad (\text{C.142})$$

Note how Bayes rule allows us to reverse the direction of the arrows impinging on θ . We see from Bayes Rule that even though the directions of the arrows in a bnet can have causal motivation, a bnet with arrows reversed from their causally motivated directions can still be very useful as a calculation tool.

Another way of stating Bayes Rule is

³Two bnets are equated if their full probability distributions (i.e., their full instantiations) are equal numerically. For example,

$$\underline{a} \rightarrow \underline{b} \rightarrow \underline{c} = P(c|b)P(b|a)P(a) = \underline{a} \leftarrow \underline{b} \leftarrow \underline{c}$$

$$\underbrace{P(\theta|x)}_{\text{posterior}} = \mathcal{N}(!\theta) \underbrace{P(x|\theta)}_{\text{likelihood}} \underbrace{P(\theta)}_{\text{prior}} . \quad (\text{C.143})$$

If, for a given likelihood, the prior and posterior distributions belong to the same family (for instance, they are both Beta distributions), then we say that the prior is the **conjugate prior** of that likelihood.

For example, $\text{Beta} \sim \text{Bernoulli}^*\text{Beta}$. Hence, the Beta distribution⁴ is the conjugate prior of the Bernoulli distribution⁵. More explicitly, if

$$p_1 \sim \text{Beta}(p_1; \alpha, \beta) \quad (\text{C.144})$$

and

$$x|p_1 \sim \text{Bernoulli}(x; p_1) , \quad (\text{C.145})$$

where $p_1 = P(x = 1)$, then

$$p_1|x \sim \text{Beta}(p_1; \alpha', \beta') \quad (\text{C.146})$$

where

$$\alpha' = \alpha + x \quad (\text{C.147})$$

$$\beta' = \beta + (1 - x) \quad (\text{C.148})$$

Ref.[91] has a table of conjugate priors.

Conjugate priors facilitate Bayesian updating of the prior to posterior in a feedback loop(see Fig.C.5).

C.26 Linear regression, Ordinary Least Squares (OLS)

Wikipedia articles

1. Linear Regression (LR)

- linear regression, Ref.[119]
- simple linear regression, Ref.[137]
- errors in variable, Ref.[96]

2. Least squares (LS)

⁴See Ref.[82] for a discussion of the Beta distribution.

⁵See Ref.[81] for a discussion of the Bernoulli distribution

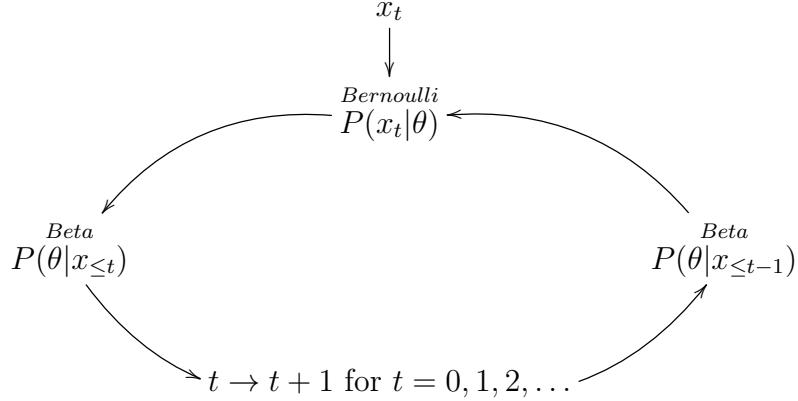


Figure C.5: Bayesian updating facilitated by conjugate prior. In this figure, $x_{\leq t} = (x_0, x_1, \dots, x_{t-1}, x_t)$.

- least squares, Ref.[117]
- ordinary least squares (OLS), Ref.[131]

Some nomenclature: In LR, the data consists of **independent x-variables** $x_1^\sigma, x_2^\sigma, \dots, x_n^\sigma$ and a **dependent y-variable** y^σ . We find a linear fit $\hat{y}^\sigma = \beta_0 + \sum_{i=1}^n \beta_i x_i^\sigma$ to the data. \hat{y}^σ is called the **estimate** of y^σ . The coefficients β_0, β_i are called **regression coefficients**. $y^\sigma - \hat{y}^\sigma = \epsilon^\sigma$ are called the **residuals**. $\mathcal{E} = \sum_\sigma (\epsilon^\sigma)^2$ is called the **error or cost**. We choose the regression coefficients so as to minimize the error.

Below, we consider two types of LR:

1. LR in which the independent x-variables are non-random.
2. LR in which the independent x-variables are random and i.i.d.

The term OLS is often used to refer to LR of type 1.

For LR of type 2, there is randomness in y coming from the randomness in x and in the residuals. For LR of type 1, there is randomness in y too, but it comes from the residuals only.

Once one assumes that certain variables are random, a “model” (i.e., a bnet, with probabilities expressed as TPMs) must be specified.

C.26.1 LR, assuming x_σ are non-random

Let

- $\sigma \in \{0, 1, 2, \dots, nsam - 1\}$: sample index
- $i_0 \in \{0, 1, 2, \dots, n\}$: index that can assume values 0 to n
- $i \in \{1, 2, \dots, n\}$: index that can assume values 1 to n . i is never equal to 0.

- $y_\sigma \in \mathbb{R}$: dependent y-variables
- $x_{\sigma i} \in \mathbb{R}$: independent x-variables
- $\epsilon_\sigma \in \mathbb{R}$: residuals
- $\beta_0, \beta_i \in \mathbb{R}$: regression coefficients

$$y_\sigma = \beta_0 + \sum_{i=1}^n x_{\sigma i} \beta_i + \epsilon_\sigma \quad (\text{C.149})$$

If we define

$$x_{\sigma 0} = 1 \quad (\text{C.150})$$

for all σ , then

$$y_\sigma = \sum_{i_0=0}^n x_{\sigma i_0} \beta_{i_0} + \epsilon_\sigma . \quad (\text{C.151})$$

If y and ϵ are $nsam$ dimensional column vectors and β is an $n+1$ dimensional column vector, and X is an $nsam \times (n+1)$ matrix, then we can write the previous equation in matrix form as:

$$y = X\beta + \epsilon . \quad (\text{C.152})$$

Derivation of LR From Minimization of Error

Let $W = [W_{\sigma,\sigma'}]$ be a symmetric matrix with non-negative diagonal elements $W_{\sigma,\sigma} \geq 0$ for all σ . W is called the **weight matrix**. The following claim describes the method of **Weighted LR** when $W \neq 1$ and of simple LR when $W = 1$.

Claim 10 *Assume the Einstein summation convention; i.e., implicit sum over repeated indices. The error function \mathcal{E} given by*

$$\mathcal{E} = \underbrace{(y_\sigma - X_{\sigma,j_0} \beta_{j_0})}_{\text{residual } \epsilon_\sigma} W_{\sigma,\sigma'} \underbrace{(y_{\sigma'} - X_{\sigma',k_0} \beta_{k_0})}_{\epsilon_{\sigma'}} , \quad (\text{C.153})$$

is minimized over β_{k_0} for all $k_0 \in \{0, 1, \dots, n\}$, if β_{k_0} is given by:

$$\hat{\beta} = (X^T W X)^{-1} X^T W y . \quad (\text{C.154})$$

When $W = 1$,

$$\hat{\beta} = (X^T X)^{-1} X^T y . \quad (\text{C.155})$$

proof:

At the minimum of \mathcal{E} , the variation $\delta\mathcal{E}$ must vanish:

$$0 = \delta\mathcal{E} = -2X_{\sigma j_0}(\delta\beta_{j_0})W_{\sigma,\sigma'}(y_{\sigma'} - X_{\sigma',k_0}\beta_{k_0}) . \quad (\text{C.156})$$

Thus,

$$X^T W y - X^T W X \beta = 0 \quad (\text{C.157})$$

which implies Eq.(C.154).

QED

Geometry of LR with non-random x_σ .

Recall that

$$y = X\beta + \epsilon . \quad (\text{C.158})$$

Define the **projection matrices**

$$\wedge = X(X^T X)^{-1} X^T , \quad \vee = 1 - \wedge \quad (\text{C.159})$$

A square matrix M is symmetric if $M^T = M$ and is idempotent if $M^2 = M$. \wedge is symmetric and idempotent and so is \vee . Note that \wedge and \vee also satisfy:

$$\vee \wedge = \wedge \vee = 0 \quad (\text{C.160})$$

and

$$\wedge X = X , \quad \vee X = 0 . \quad (\text{C.161})$$

One has

$$\beta = (X^T X)^{-1} X^T (y - \epsilon) . \quad (\text{C.162})$$

Define

$$\boxed{\widehat{\beta} = \underbrace{(X^T X)^{-1} X^T}_B y ,} \quad (\text{C.163a})$$

$$\widehat{y} = X\widehat{\beta} = \wedge y , \quad (\text{C.163b})$$

and

$$\widehat{\epsilon} = y - X\widehat{\beta} = y - \widehat{y} = (1 - \wedge)y = \vee y . \quad (\text{C.163c})$$

\wedge is sometimes called the **hat matrix**, because it gives y a hat.

Given any function $f = f(y, X, \epsilon)$ and a scalar factor $\xi \in \mathbb{R}$, suppose $f(\xi y, \xi X, \xi \epsilon) = \xi^{\mathcal{O}} f(y, X, \epsilon)$. Then we will say that $f(\cdot)$ is of **order \mathcal{O} under scaling**. Note that $\{\widehat{y}, \widehat{\epsilon}\}$ are all of order 1 under scaling, $\{\beta, \widehat{\beta}, \wedge, \vee\}$ are all of order 0 under scaling, and B is of order -1 under scaling. Thus, each cfit (i.e., symbol with a hat) scales

the same way as its estimand (i.e., same symbol without a hat). Furthermore, β , its cfit $\hat{\beta}$, and the projection matrices \wedge, \vee are invariant ($\mathcal{O} = 0$) under scaling.

Note that y can be expressed as a sum of 2 orthogonal estimates:

$$y = \underbrace{\hat{y}}_{\wedge y} + \underbrace{\hat{\epsilon}}_{\vee y}. \quad (\text{C.164})$$

Fig.C.6 shows triangles representing $y = X\beta + \epsilon$ and $y = \hat{y} + \hat{\epsilon}$.

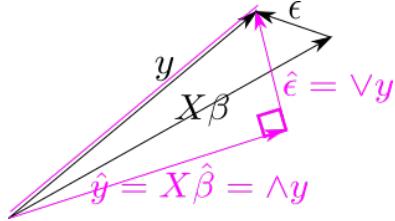


Figure C.6: Triangles representing $y = X\beta + \epsilon$ and $y = \hat{y} + \hat{\epsilon}$.

LR Goodness of Fit, R^2

Assume the components of ϵ are random with zero mean:

$$E[\epsilon] = \langle \epsilon \rangle = 0 \quad (\text{C.165})$$

Assume X and β are not random. This makes $\underline{y} = X\beta + \underline{\epsilon}$ and $\hat{\beta} = (X^T X)^{-1} X^T \underline{y}$ random. One finds that

$$\langle \underline{y} \rangle = X\beta \quad (\text{C.166a})$$

$$\langle \hat{y} \rangle = \wedge \underbrace{\langle \underline{y} \rangle}_{X\beta} = \langle \underline{y} \rangle \quad (\text{C.166b})$$

$$\langle \hat{\epsilon} \rangle = \vee \underbrace{\langle \underline{y} \rangle}_{X\beta} = 0 \quad (\text{C.166c})$$

$$\langle \hat{\beta} \rangle = \beta \quad (\text{C.166d})$$

So far, we have assumed a zero mean value for ϵ . Next, assume “**homoscedasticity**” (**homo-spread**)⁶, which means that

⁶I find the word “homoscedasticity” unnecessarily long, cryptic and easy to misspell so I like to replace it by “homo-spread”. The opposite of “homoscedasticity” is “heteroscedasticity”, which I like to replace with “hetero-spread”.

$$\langle \underline{\epsilon}, \underline{\epsilon}^T \rangle = \xi^2 I_{nsam} \quad (\text{C.166e})$$

where $\xi \geq 0$, and I_{nsam} is the $nsam \times nsam$ identity matrix. It follows that

$$\langle \underline{y}, \underline{y}^T \rangle = \langle \underline{\epsilon}, \underline{\epsilon}^T \rangle = \xi^2 I_{nsam}, \quad (\text{C.167a})$$

$$\langle \widehat{\underline{\epsilon}}, \widehat{\underline{\epsilon}}^T \rangle = \vee \langle \underline{y}, \underline{y}^T \rangle \vee^T = \xi^2 \vee, \quad (\text{C.167b})$$

$$\langle \widehat{\underline{y}}, \widehat{\underline{y}}^T \rangle = \wedge \langle \underline{y}, \underline{y}^T \rangle \wedge^T = \xi^2 \wedge \quad (\text{C.167c})$$

and

$$\langle \widehat{\underline{\beta}}, \widehat{\underline{\beta}}^T \rangle = B \langle \underline{y}, \underline{y}^T \rangle B^T = \xi^2 (X^T X)^{-1}. \quad (\text{C.167d})$$

For any random column vector \underline{a} , let

$$\| \underline{a} \|^2 = \underline{a}^T \underline{a} = \text{tr}(\underline{a} \underline{a}^T) \quad (\text{C.168})$$

and

$$\langle \| \underline{a} - \langle \underline{a} \rangle \|^2 \rangle = \langle \underline{a}^T \underline{a} \rangle - \langle \underline{a}^T \rangle \langle \underline{a} \rangle = \text{tr} \langle \underline{a}, \underline{a}^T \rangle. \quad (\text{C.169})$$

Define the following sums of squares (SS):

$$SS_{\underline{y}} = \langle \| \underline{y} - \langle \underline{y} \rangle \|^2 \rangle = \langle \underline{y}^T \underline{y} \rangle - \langle \underline{y}^T \rangle \langle \underline{y} \rangle = \text{tr} \langle \underline{y}, \underline{y}^T \rangle \quad (\text{C.170a})$$

$$SS_{\widehat{\underline{y}}} = \langle \| \widehat{\underline{y}} - \langle \widehat{\underline{y}} \rangle \|^2 \rangle = \langle \widehat{\underline{y}}^T \widehat{\underline{y}} \rangle - \langle \widehat{\underline{y}}^T \rangle \langle \widehat{\underline{y}} \rangle = \text{tr} \langle \widehat{\underline{y}}, \widehat{\underline{y}}^T \rangle \quad (\text{C.170b})$$

$$SS_{res} = \langle \| \underline{y} - \widehat{\underline{y}} \|^2 \rangle = \langle \| \widehat{\underline{\epsilon}} \|^2 \rangle = \text{tr} \langle \widehat{\underline{\epsilon}}, \widehat{\underline{\epsilon}}^T \rangle \quad (\text{C.170c})$$

Claim 11 *The following is true without homo-spread:*

$$\underbrace{\text{tr} \langle \underline{y}, \underline{y}^T \rangle}_{SS_{\underline{y}}} = \underbrace{\text{tr} \langle \widehat{\underline{y}}, \widehat{\underline{y}}^T \rangle}_{SS_{\widehat{\underline{y}}}} + \underbrace{\text{tr} \langle \widehat{\underline{\epsilon}}, \widehat{\underline{\epsilon}}^T \rangle}_{SS_{res}} \quad (\text{C.171})$$

This is like the Pythagorean Theorem for the magenta right triangle in Fig.C.6.

proof:

From Eqs.C.167 and C.170, we see that

$$SS_{\underline{y}} = \text{tr} \langle \underline{y}, \underline{y}^T \rangle \quad (\text{C.172})$$

$$SS_{\hat{y}} = \text{tr} \langle \hat{y}, \hat{y}^T \rangle = \text{tr} \langle \wedge \underline{y}, \underline{y}^T \rangle \quad (\text{C.173})$$

$$SS_{res} = \text{tr} \langle \hat{\epsilon}, \hat{\epsilon}^T \rangle = \text{tr} \langle \vee \underline{y}, \underline{y}^T \rangle \quad (\text{C.174})$$

Now use $\wedge + \vee = 1$.

QED

The goodness of fit for this model is often measured using the **coefficient of determination** R^2 . R^2 is defined by

$$R^2 = 1 - \frac{SS_{res}}{SS_{\underline{y}}} = \frac{SS_{\hat{y}}}{SS_{\underline{y}}} = \frac{\text{tr} \langle \hat{y}, \hat{y}^T \rangle}{\text{tr} \langle \underline{y}, \underline{y}^T \rangle} \quad (\text{C.175})$$

If homo-spread holds, then R^2 reduces to

$$R^2 = \frac{\text{tr} \wedge}{nsam} . \quad (\text{C.176})$$

C.26.2 LR, assuming x_σ are random

Let

- $i_0 \in \{0, 1, 2, \dots, n\}$: index that can assume values 0 to n
- $i \in \{1, 2, \dots, n\}$: index that can assume values 1 to n . i is never equal to 0.
- $y \in \mathbb{R}$: true value of dependent y-variable
- $\hat{y} \in \mathbb{R}$: cfit of dependent y-variable
- $\epsilon \in \mathbb{R}$: residual
- $\underline{x}_i \in \mathbb{R}$: independent x-variables for $i \in \{1, \dots, n\}$
- $\beta_0, \beta_i \in \mathbb{R}$: regression coefficients

$$\hat{y} = \beta_0 + \sum_{j=1}^n \beta_j \underline{x}_j \quad (\text{C.177})$$

$$= \sum_{j_0=0}^n \beta_{j_0} \underline{x}_{j_0} \quad (\text{Assume } \underline{x}_0 = 1.) \quad (\text{C.178})$$

$$\underline{y} = \hat{y} + \epsilon \quad (\text{C.179})$$

Transforming expressions from non-random to random x_σ

Define the following population averages:

$$E_\sigma[x^\sigma] = \frac{1}{nsam} \sum_\sigma x^\sigma , \quad (\text{C.180})$$

$$E_\sigma[x^\sigma y^\sigma] = \frac{1}{nsam} \sum_\sigma x^\sigma y^\sigma , \quad (\text{C.181})$$

$$\langle x^\sigma, y^\sigma \rangle_\sigma = E_\sigma[x^\sigma y^\sigma] - E_\sigma[x^\sigma] E_\sigma[y^\sigma] . \quad (\text{C.182})$$

Claim 12 If the x_σ are i.i.d. random variables,

$$E_\sigma[x^\sigma] = \langle \underline{x} \rangle \quad (\text{C.183})$$

$$E_\sigma[x^\sigma y^\sigma] = \langle \underline{xy} \rangle \quad (\text{C.184})$$

$$\langle x^\sigma, y^\sigma \rangle_\sigma = \langle \underline{x}, \underline{y} \rangle \quad (\text{C.185})$$

proof:

$$\frac{1}{nsam} \sum_\sigma x^\sigma = \frac{1}{nsam} \sum_{x \in S_{\underline{x}}} x \underbrace{\sum_\sigma \mathbb{1}(x^\sigma = x)}_{N(x^\sigma=x)} \quad (\text{C.186})$$

$$= \sum_x x P(x) \quad (\text{C.187})$$

$$= \langle \underline{x} \rangle \quad (\text{C.188})$$

$$\frac{1}{nsam} \sum_\sigma x^\sigma y^\sigma = \frac{1}{nsam} \sum_{x \in S_{\underline{x}}} \sum_{y \in S_{\underline{y}}} xy \underbrace{\sum_\sigma \mathbb{1}(x^\sigma = x, y^\sigma = y)}_{N(x^\sigma=x, y^\sigma=y)} \quad (\text{C.189})$$

$$= \sum_{x,y} xy P(x, y) \quad (\text{C.190})$$

$$= \langle \underline{xy} \rangle \quad (\text{C.191})$$

Eq.(C.185) follows from Eq.(C.183) and Eq.(C.184).

QED

Recall that

$$Y_\sigma = \beta_0 + \sum_{j=1}^n X_{\sigma,j} \beta_j + \epsilon_\sigma . \quad (\text{C.192})$$

Assume

$$E_\sigma[X_{\sigma,k} \epsilon_\sigma] = E_\sigma[X_{\sigma,k}] \underbrace{E_\sigma[\epsilon_\sigma]}_{=0} = 0 . \quad (\text{C.193})$$

Then we have

$$E_\sigma[X_{\sigma,k}Y_\sigma] = E_\sigma[X_{\sigma,k}]\beta_0 + \sum_{j=1}^n E_\sigma[X_{\sigma,k}X_{\sigma,j}]\beta_j + \underbrace{E_\sigma[X_{\sigma,k}\epsilon_\sigma]}_{=0} \quad (\text{C.194})$$

and

$$E_{\sigma'}[X_{\sigma',k}]E_\sigma[Y_\sigma] = E_{\sigma'}[X_{\sigma',k}]\beta_0 + \sum_{j=1}^n E_{\sigma'}[X_{\sigma',k}]E_\sigma[X_{\sigma,j}]\beta_j + \underbrace{E_{\sigma'}[X_{\sigma',k}]E_\sigma[\epsilon_\sigma]}_{=0} . \quad (\text{C.195})$$

Subtracting Eq.(C.195) from Eq.(C.194), we get

$$\langle X_{\sigma,k}, Y_\sigma \rangle_\sigma = \sum_{j=1}^n \langle X_{\sigma,k}, X_{\sigma,j} \rangle_\sigma \beta_j . \quad (\text{C.196})$$

Define the n dimensional covariance matrix C by

$$C_{k,j} = \langle X_{\sigma,k}, X_{\sigma,j} \rangle_\sigma . \quad (\text{C.197})$$

Then Eq.(C.196) implies

$$\beta_j = \sum_{k=1}^n C_{j,k}^{-1} \langle X_{\sigma,k}, Y_\sigma \rangle_\sigma \quad (\text{C.198})$$

for all $j = 1, 2, \dots, n$.

If we assume that the x_σ are i.i.d., then, by virtue of Claim 12, the matrix C tends to

$$C_{k,j} \rightarrow \langle \underline{x}_k, \underline{x}_j \rangle \quad (\text{C.199})$$

and Eq.(C.198) implies

$$\beta_j = \sum_{k=1}^n C_{j,k}^{-1} \langle \underline{x}_k, \underline{y} \rangle . \quad (\text{C.200})$$

Geometry of LR with random x_σ

Recall that

$$\underline{y} = \beta_0 + \underbrace{\sum_{j=1}^n \beta_j \underline{x}_j}_{\hat{\underline{y}}} + \underline{\epsilon} . \quad (\text{C.201})$$

Assume

$$\langle \underline{\epsilon} \rangle = 0 \quad (\text{C.202})$$

and

$$\langle \underline{x}_j, \underline{\epsilon} \rangle = 0 \quad (\text{C.203})$$

for all j .

For $k = 1, \dots, n$,

$$\langle \underline{x}_k, \underline{y} \rangle = \sum_{j=1}^n \beta_j \langle \underline{x}_k, \underline{x}_j \rangle . \quad (\text{C.204})$$

Let \underline{x}^n and β^n be n -dimensional column vectors. Then Eq.(C.204) can be represented in matrix notation by

$$\langle \underline{x}^n, \underline{y} \rangle = \langle \underline{x}^n, (\underline{x}^n)^T \rangle \beta^n . \quad (\text{C.205})$$

Hence,

$$\boxed{\beta^n = \langle \underline{x}^n, (\underline{x}^n)^T \rangle^{-1} \langle \underline{x}^n, \underline{y} \rangle .} \quad (\text{C.206})$$

For β_0 , use

$$\beta_0 = \langle \underline{y} \rangle - \langle \underline{x}^n \rangle^T \beta^n \quad (\text{C.207})$$

Notice that Eq.(C.206) for the regression coefficients is the same as Eq.(C.200). So we have rederived the same formula via a different method.

Next, we will write Eq.(C.206) for the special cases $n = 1$ and $n = 2$, where n is the number of independent x-variables \underline{x}_j

1. $n = 1$ (\underline{y} fitted by a line)

$$\underline{y} = \beta_0 + \beta_1 \underline{x} + \underline{\epsilon} \quad (\text{C.208})$$

Eq.(C.206) becomes

$$\beta_1 = \frac{\langle \underline{y}, \underline{x} \rangle}{\langle \underline{x}, \underline{x} \rangle} \quad (\text{C.209})$$

2. $n = 2$ (\underline{y} fitted by a plane)

$$\underline{y} = \beta_0 + \beta_1 \underline{x}_1 + \beta_2 \underline{x}_2 + \underline{\epsilon} \quad (\text{C.210})$$

Define

$$C_{i,j} = \langle \underline{x}_i, \underline{x}_j \rangle \quad (\text{C.211})$$

for all i, j . Then Eq.(C.206) becomes⁷

$$\begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix} = C^{-1} \begin{bmatrix} \langle \underline{y}, \underline{x}_1 \rangle \\ \langle \underline{y}, \underline{x}_2 \rangle \end{bmatrix} \quad (\text{C.212})$$

$$= \frac{1}{\det C} \begin{bmatrix} C_{22} & -C_{12} \\ -C_{21} & C_{11} \end{bmatrix} \begin{bmatrix} \langle \underline{y}, \underline{x}_1 \rangle \\ \langle \underline{y}, \underline{x}_2 \rangle \end{bmatrix} \quad (\text{C.213})$$

Hence,

$$\beta_1 = \frac{\partial \underline{y}}{\partial \underline{x}_1} = \frac{C_{22} \langle \underline{y}, \underline{x}_1 \rangle - C_{12} \langle \underline{y}, \underline{x}_2 \rangle}{C_{11} C_{22} - C_{12}^2} \quad (\text{C.214})$$

Eq.(C.214) agrees with the value of $\beta_{YX,Z}$ in Ref.[40] by Pearl, if we replace in Pearl's formulae $X \rightarrow \underline{x}_1$, $Y \rightarrow \underline{y}$, $Z \rightarrow \underline{x}_2$.

Regression interpreted as differentiation of y

Finding the derivative of y with respect to (wrt) X (a.k.a. “**Regressing y on X** ”) is finding $\widehat{\beta} = \frac{d}{dX} \underbrace{y}_{X\beta}$.

Recall that

$$\underline{y} = \beta_0 + \sum_{k=1}^n \underline{x}_k \beta_k + \epsilon. \quad (\text{C.215})$$

Therefore,

$$\langle \underline{x}_i, \underline{y} \rangle = \sum_{k=1}^n \langle \underline{x}_i, \underline{x}_k \rangle \beta_k \quad (\text{C.216})$$

$$= \langle \underline{x}_i, \underline{x}_i \rangle \beta_i + \sum_{k=1}^n \mathbb{1}(k \neq i) \langle \underline{x}_i, \underline{x}_k \rangle \beta_k. \quad (\text{C.217})$$

Hence,

$$\beta_i = \frac{\langle \underline{x}_i, \underline{y} \rangle}{\langle \underline{x}_i, \underline{x}_i \rangle} - \sum_{k=1}^n \mathbb{1}(k \neq i) \frac{\langle \underline{x}_i, \underline{x}_k \rangle}{\langle \underline{x}_i, \underline{x}_i \rangle} \beta_k. \quad (\text{C.218})$$

Let's represent the linear operator $\langle \underline{x}_i, \underline{x}_i \rangle^{-1} \langle \underline{x}_i, \cdot \rangle$ as a derivative:

$$\frac{d \cdot}{d \underline{x}_i} = \frac{\langle \underline{x}_i, \cdot \rangle}{\langle \underline{x}_i, \underline{x}_i \rangle}. \quad (\text{C.219})$$

⁷Recall that if $M = \begin{bmatrix} a & b \\ c & d \end{bmatrix}$ then $M^{-1} = \frac{1}{\det M} \begin{bmatrix} d & -b \\ -c & a \end{bmatrix}$

Eq.(C.218) can be expressed in derivative notation as:

$$\beta_i = \frac{d\underline{y}}{d\underline{x}_i} - \sum_{k=1}^n \mathbb{1}(k \neq i) \frac{d\underline{x}_k}{d\underline{x}_i} \beta_k \quad (\text{C.220})$$

Note that Eq.(C.220) evokes the formula for differentials:

$$d\underline{y} = \sum_{k=1}^n \beta_k d\underline{x}_k . \quad (\text{C.221})$$

Note also that, because of the linearity of the derivative operator, Eq.(C.220) implies:

$$\beta_i = \frac{d}{d\underline{x}_i} \left(\underbrace{\underline{y} - \sum_{k=1}^n \mathbb{1}(k \neq i) \underline{x}_k \beta_k}_{\underline{y} - \underline{x}_i \beta_i} \right) . \quad (\text{C.222})$$

Eq.(C.222) can be used to find $\hat{\beta}_i$ in two steps:

STEP 1: Regress $\underline{y} - \underline{x}_i \beta_i$ on $(\underline{x}_k)_{k \in \{1, 2, \dots, n\} - \{i\}}$. Get estimates $(\hat{\beta}_k)_{k \in \{1, 2, \dots, n\} - \{i\}}$.

STEP 2: Regress $\underline{y} - \sum_{k \neq i} \underline{x}_k \hat{\beta}_k$ on \underline{x}_i . Get estimate $\hat{\beta}_i$.

Of course, one can also find $\hat{\beta}_i$ by regressing \underline{y} on $(\underline{x}_k)_{k \in \{1, 2, \dots, n\}}$, to get estimates $(\hat{\beta}_k)_{k \in \{1, 2, \dots, n\}}$.

C.27 Logistic Regression (LoR)

Suppose $x_\sigma \in \mathbb{R}^n$, $y_\sigma \in \mathbb{R}$, and Σ is a population of individuals σ . In general, a **regression** is when we curve-fit a dataset $\{(x_\sigma, y_\sigma) : \sigma \in \Sigma\}$ with a function $\hat{y} = f(x)$. In **Linear Regression (LR)**, which we discussed earlier, $f(x)$ is a hyperplane in x . On the other hand, in **Logistic Regression (LoR)**, $y_\sigma \in [0, 1]$ and $f(x)$ is the sigmoid of a hyperplane in x .

More specifically, for LR we have Eq.(C.149) which reads as follows:

$$y_\sigma = \underbrace{\beta_0 + \sum_{i=1}^n x_{\sigma i} \beta_i}_{\hat{y}_\sigma} + \epsilon_\sigma \quad (\text{LR}) . \quad (\text{C.223})$$

For LoR, we have instead

$$p_\sigma = \text{smoid} \left(\beta_0 + \sum_{i=1}^n x_{\sigma i} \beta_i + \epsilon_\sigma \right) \quad (\text{LoR}) , \quad (\text{C.224})$$

or, equivalently,

$$\underbrace{\text{lodds}(p_\sigma)}_{\ln \frac{p_\sigma}{1-p_\sigma}} = \beta_0 + \underbrace{\sum_{i=1}^n x_{\sigma i} \beta_i}_{\hat{y}_\sigma} + \epsilon_\sigma \quad (\text{LoR}) , \quad (\text{C.225})$$

where we have used the fact that $\text{lodds}()$ is the inverse function of $\text{smoid}()$. Hence, an LoR fit can be calculated by collecting a dataset $\{(x_\sigma, p_\sigma) : \sigma \in \Sigma\}$, transforming that dataset to the dataset $\{(x_\sigma, \text{lodds}(p_\sigma)) : \sigma \in \Sigma\}$, and fitting the latter dataset with a hyperplane. Let $P(\underline{Y}_\sigma = 1) = p_\sigma \in [0, 1]$. LoR can be used for binary classification if we define the binary class variable $c_\sigma \in \{0, 1\}$ by

$$c_\sigma = \mathbb{1}(P(\underline{Y}_\sigma = 1) > \alpha) \quad (\text{C.226})$$

for some $0 < \alpha < 1$.

C.28 Entropy, Kullback-Leibler divergence, Cross-Entropy

For probability distributions $p(x), q(x)$ of $x \in S_x$

- Entropy:

$$H(p) = - \sum_x p(x) \ln p(x) \geq 0 \quad (\text{C.227})$$

- Kullback-Leibler divergence:

$$D_{KL}(p \parallel q) = \sum_x p(x) \ln \frac{p(x)}{q(x)} \geq 0 \quad (\text{C.228})$$

- Cross entropy:

$$CE(p \parallel q) = - \sum_x p(x) \ln q(x) \quad (\text{C.229})$$

$$= H(p) + D_{KL}(p \parallel q) \quad (\text{C.230})$$

C.29 Definition of various entropies used in Shannon Information Theory

- (plain) Entropy of \underline{x}

$$H(\underline{x}) = - \sum_x P(x) \ln P(x) \quad (\text{C.231})$$

This quantity measures the spread of $P_{\underline{x}}$. $H(\underline{x}) \geq 0$ and it vanishes iff $P(x) = \delta(x, x_0)$ (deterministic case)

- **Conditional Entropy of \underline{y} given \underline{x}**

$$H(\underline{y}|\underline{x}) = - \sum_{x,y} P(x,y) \ln P(y|x) \quad (\text{C.232})$$

$$= H(\underline{y}, \underline{x}) - H(\underline{x}) \quad (\text{C.233})$$

This quantity measures the conditional spread of \underline{y} given \underline{x} . $H(\underline{y}|\underline{x}) \geq 0$.

- **Mutual Information (MI) of \underline{x} and \underline{y} .**

$$H(\underline{y} : \underline{x}) = \sum_{x,y} P(x,y) \ln \frac{P(x,y)}{P(x)P(y)} \quad (\text{C.234})$$

$$= H(\underline{x}) + H(\underline{y}) - H(\underline{y}, \underline{x}) \quad (\text{C.235})$$

This quantity measures the correlation between \underline{x} and \underline{y} . $H(\underline{y} : \underline{x}) \geq 0$ and it vanishes iff $P(x,y) = P(x)P(y)$.

- **Conditional Mutual Information (CMI)⁸ of \underline{x} and \underline{y} given $\underline{\lambda}$**

$$H(\underline{y} : \underline{x} | \underline{\lambda}) = \sum_{x,y,\lambda} P(x,y,\lambda) \ln \frac{P(x,y|\lambda)}{P(x|\lambda)P(y|\lambda)} \quad (\text{C.236})$$

$$= H(\underline{x} | \underline{\lambda}) + H(\underline{y} | \underline{\lambda}) - H(\underline{y}, \underline{x} | \underline{\lambda}) \quad (\text{C.237})$$

This quantity measures the conditional correlation of \underline{x} and \underline{y} given $\underline{\lambda}$. $H(\underline{y} : \underline{x} | \underline{\lambda}) \geq 0$ and it vanishes iff $P(x,y|\lambda) = P(x|\lambda)P(y|\lambda)$.

An interesting special case occurs when $P(\lambda) = \delta(\lambda, \lambda_0)$ (the frequentist case of no λ prior.) In that case CMI reduces to

$$H(\underline{y} : \underline{x} | \lambda_0) = \sum_{x,y} P(x,y|\lambda_0) \ln \frac{P(x,y|\lambda_0)}{P(x|\lambda_0)P(y|\lambda_0)} \geq 0 \quad (\text{C.238})$$

- **Kullback-Leibler Divergence from $P_{\underline{x}}$ to $P_{\underline{y}}$.**

Assume random variables \underline{x} and \underline{y} have the same set of states $S_{\underline{x}} = S_{\underline{y}}$. Then

⁸CMI can be read as “see me”.

$$D_{KL}(P_{\underline{x}} \parallel P_{\underline{y}}) = \sum_x P_{\underline{x}}(x) \ln \frac{P_{\underline{x}}(x)}{P_{\underline{y}}(x)} \quad (\text{C.239})$$

This quantity measures a non-symmetric distance between the probability distributions $P_{\underline{x}}$ and $P_{\underline{y}}$. $D_{KL}(P_{\underline{x}} \parallel P_{\underline{y}}) \geq 0$ and it equals zero iff $P_{\underline{x}} = P_{\underline{y}}$.

C.30 Mean log likelihood asymptotic behavior

Define the log likelihood by

$$LL_{y|\theta} = \ln P(y|\theta). \quad (\text{C.240})$$

In this section, we will represent averages over $\underline{y}|\theta$ by angular brackets:

$$\langle f(y) \rangle = \sum_y P(y|\theta) f(y) = E_{\underline{y}|\theta}[f(y)]. \quad (\text{C.241})$$

Note that the mean log likelihood equals minus the entropy:

$$H(\underline{y}|\theta) = -\langle LL_{y|\theta} \rangle \quad (\text{C.242})$$

Claim 13

$$\langle \partial_\theta LL_{y|\theta} \rangle = 0 \quad (\text{C.243})$$

$$\langle \partial_\theta^2 LL_{y|\theta} \rangle = -\langle (\partial_\theta LL_{y|\theta})^2 \rangle \quad (\text{C.244})$$

proof:

$$\langle \partial_\theta LL_{y|\theta} \rangle = \sum_y P(y|\theta) \partial_\theta \ln P(y|\theta) \quad (\text{C.245})$$

$$= \sum_y \partial_\theta P(y|\theta) \quad (\text{C.246})$$

$$= 0 \quad (\text{C.247})$$

$$\langle \partial_\theta^2 LL_{y|\theta} \rangle = \sum_y P(y|\theta) \partial_\theta \left[\frac{1}{P(y|\theta)} \partial_\theta P(y|\theta) \right] \quad (\text{C.248})$$

$$= -\sum_y P(y|\theta) \frac{1}{P(y|\theta)^2} [\partial_\theta P(y|\theta)]^2 + \underbrace{\sum_y \partial_\theta^2 P(y|\theta)}_{=0} \quad (\text{C.249})$$

$$= -\sum_y P(y|\theta) [\partial_\theta \ln P(y|\theta)]^2 \quad (\text{C.250})$$

$$= -\langle (\partial_\theta LL_{y|\theta})^2 \rangle \quad (\text{C.251})$$

QED

Define

$$\Delta\theta = \theta' - \theta \quad (\text{C.252})$$

and

$$-\Delta H(\underline{y}|\theta) = \Delta \langle LL_{y|\theta} \rangle = \langle LL_{y|\theta'} \rangle - \langle LL_{y|\theta} \rangle \quad (\text{C.253})$$

If we expand $\langle LL_{y|\theta'} \rangle$ as a Taylor series to second order about the point $\theta' = \theta$, we get

$$\langle LL_{y|\theta'} \rangle = \langle LL_{y|\theta} \rangle + \Delta\theta \underbrace{\langle \partial_\theta LL_{y|\theta} \rangle}_{=0} + \frac{(\Delta\theta)^2}{2} \underbrace{\langle \partial_\theta^2 LL_{y|\theta} \rangle}_{-\langle (\partial_\theta LL_{y|\theta})^2 \rangle} + \mathcal{O}((\Delta\theta)^3) \quad (\text{C.254})$$

$$-\Delta H(\underline{y}|\theta) = \Delta \langle LL_{y|\theta} \rangle = -\frac{(\Delta\theta)^2}{2} \langle (\partial_\theta LL_{y|\theta})^2 \rangle + \mathcal{O}((\Delta\theta)^3) \quad (\text{C.255})$$

Thus, $\theta' = \theta$ maximizes the mean log likelihood $\langle LL_{y|\theta} \rangle$ (and minimizes the entropy $H(\underline{y}|\theta)$).

Note that

$$\Delta \langle LL_{y|\theta} \rangle = \left\langle \ln \frac{P(y|\theta')}{P(y|\theta)} \right\rangle. \quad (\text{C.256})$$

If we approximate the ratio of these 2 probabilities by a Gaussian,

$$\frac{P(y|\theta')}{P(y|\theta)} \approx \exp \left(-\frac{(\Delta\theta)^2}{2\sigma_\theta^2} \right), \quad (\text{C.257})$$

then

$$\sigma_\theta^2 = \langle (\partial_\theta LL_{y|\theta})^2 \rangle^{-1}. \quad (\text{C.258})$$

C.31 Arc Strength (Arc Force)

Given a bnet with an arc (i.e., arrow) $\underline{x} \rightarrow \underline{y}$, we define the **arc strength or arc force** of arc $\underline{x} \rightarrow \underline{y}$ to be $H(\underline{x} : \underline{y})$ (i.e., the mutual information between \underline{x} and \underline{y}). Evaluation of $H(\underline{x} : \underline{y})$ requires knowing $P(y|x)$, $P(x)$ and $P(y)$. $P(y|x)$ is the TPM of node y , so it is immediately available from the specification of the bnet. Calculating $P(x)$ and $P(y)$ is more involved, and requires marginalizing the full probability distribution of the bnet. Such marginalizations can be done using the junction tree algorithm described in Chapter 40.

C.32 Pearson Chi-Squared Test

The **Pearson divergence** (a.k.a. **Pearson Chi-squared test statistic**) for two probability distributions $PO(x)$ and $PE(x)$, where $x \in S_{\underline{x}}$, is defined as follows:

$$D_{\chi^2} = \sum_x \frac{[PO(x) - PE(x)]^2}{PE(x)} = \sum_x \frac{PO^2(x)}{PE(x)} - 1. \quad (\text{C.259})$$

Usually PO is the observed probability distribution and PE is the expected, theoretical one.

As the following claim shows, the Pearson divergence is closely related to the Kullback-Leibler divergence.

Claim 14 If $\left| \frac{PO(x)}{PE(x)} - 1 \right| << 1$ for all $x \in S_{\underline{x}}$, then

$$D_{KL}(PO \parallel PE) \approx D_{\chi^2}. \quad (\text{C.260})$$

proof:

$$D_{KL}(PO \parallel PE) = \sum_x PO(x) \ln \frac{PO(x)}{PE(x)} \quad (\text{C.261})$$

$$= \sum_x PO(x) \ln \left(1 + \frac{PO(x)}{PE(x)} - 1 \right) \quad (\text{C.262})$$

$$\approx \sum_x PO(x) \left(\frac{PO(x)}{PE(x)} - 1 \right) \quad (\text{C.263})$$

$$= \sum_x \frac{PO^2(x)}{PE(x)} - 1 \quad (\text{C.264})$$

$$= D_{\chi^2} \quad (\text{C.265})$$

QED

Let $nx = |S_{\underline{x}}|$. Let $P_{\chi^2}(y)$ be the χ^2 (with $nx - 1$ degrees of freedom) probability distribution, and let $F_{\chi^2}(\alpha)$ be its cumulative distribution. Find α such that

$$95\% = \int_0^\alpha dy P_{\chi^2}(y) = F_{\chi^2}(\alpha) \quad (\text{C.266})$$

If $D_{\chi^2} < \alpha$, then we say that $PO = PE$ to 95% significance level (SL), whereas if $D_{\chi^2} > \alpha$, we say that $PO \neq PE$ to 95% SL (i.e., SL = 95%). The higher SL becomes, the higher α becomes, and the bigger the divergence D_{χ^2} has to be, before we are willing to declare that $PO \neq PE$.

C.33 Demystifying Population and Sample Variances

Let $x[\sigma] = x^\sigma$. Given i.i.d. real variables $(x^\sigma)_{\sigma=0,1,\dots,n-1}$, let⁹

$$\hat{\mu} = \bar{x} = \frac{1}{n} \sum_{\sigma} x^\sigma \quad (\text{C.267})$$

$$(\hat{\sigma}^2)_\infty = \frac{1}{n} \sum_{\sigma} (x^\sigma - \mu)^2 \quad (\text{C.268})$$

$$\hat{\sigma}^2 = \frac{1}{n-1} \sum_{\sigma} (x^\sigma - \hat{\mu})^2 \quad (\text{C.269})$$

Statisticians¹⁰ call $(\hat{\sigma}^2)_\infty$ the “population variance”. I will call it the **population variance for fixed μ** . Note that it depends on the fixed parameter μ . Statisticians call $\hat{\sigma}^2$ the “sample variance”. Instead, I will call $\hat{\sigma}^2$ the **population variance for random μ** .

If one treats x^σ as a random variable, then one must treat $\hat{\mu}$ as a random variable too. Let

$$E[\underline{x}^\sigma] = \mu \quad (\text{C.270})$$

and

$$\left\langle \underline{x}^\sigma, \underline{x}^{\sigma'} \right\rangle = \delta(\sigma, \sigma') \sigma^2. \quad (\text{C.271})$$

Then one can show that

$$E[(\hat{\sigma}^2)_\infty] = \frac{1}{n} E \left[\sum_{\sigma} (\underline{x}^\sigma - \mu)^2 \right] \quad (\text{C.272})$$

$$= \sigma^2 \quad (\text{C.273})$$

and

$$E[\hat{\sigma}^2] = \frac{1}{n-1} E \left[\sum_{\sigma} (\underline{x}^\sigma - \hat{\mu})^2 \right] \quad (\text{C.274})$$

$$= \sigma^2 \quad (\text{C.275})$$

⁹Do not confuse the sample index σ and the standard deviation σ .

¹⁰In the language of Statisticians, a “population” is supposed to be so large that its μ does not fluctuate, and a “sample” is supposed to be a small subset of that population for which the μ is assumed to fluctuate. In this book, I use the word “population” to mean a set of any size containing individuals, I use the word “sub-population” to refer to a subset of the population, and I use the word “sample” (a.k.a. individual, observation, unit, record) to mean a single individual of the population.

This is the reason why we use an $n - 1$ instead of an n in $\widehat{\sigma^2}$. Because it makes $E[\widehat{\sigma^2}] = \sigma^2$ so $\widehat{\sigma^2}$ is an unbiased estimator of the single individual variance σ^2 .

The intuitive reason for why $\widehat{\sigma^2}$ is divided by $n - 1$ instead of n is that whereas μ in $(\widehat{\sigma^2})_\infty$ is kept fixed and is “quiet”, the $\widehat{\mu}$ in $\widehat{\sigma^2}$ is a random variable, noisy instead of quiet. The fluctuations in $\widehat{\mu}$ are strongly correlated with the fluctuations of the \underline{x}^σ , so they decrease the fluctuations in $\widehat{\sigma^2}$ compared to those in $(\widehat{\sigma^2})_\infty$. By dividing by $n - 1$ instead of n , we compensate for this decrease in fluctuations so that the ratio of the numerator and denominator of $\widehat{\sigma^2}$ equals σ^2 , instead of something *smaller* than σ^2 , as would happen if were to divide by n instead of $n - 1$. In terms of “degrees of freedom”(DOFs), $(\widehat{\sigma^2})_\infty$ has n DOFs (namely one for each \underline{x}^σ), whereas $\widehat{\sigma^2}$ has $n - 1$ DOFs. (the presence of $\widehat{\mu}$ subtracts one DOF). In both $(\widehat{\sigma^2})_\infty$ and $\widehat{\sigma^2}$, one divides by the number of DOFs.

C.34 Independence of $\widehat{\mu}$ and $\widehat{\sigma^2}$

Let $x[\sigma] = x^\sigma$. Consider i.i.d.real variables $(x^\sigma)_{\sigma=0,1,\dots,n-1}$ such that¹¹

$$E[\underline{x}^\sigma] = \mu \quad (\text{C.276})$$

$$\langle \underline{x}^\sigma, \underline{x}^{\sigma'} \rangle = \delta(\sigma, \sigma')\sigma^2 . \quad (\text{C.277})$$

$$\widehat{\mu} = \bar{x} = \frac{1}{n} \sum_{\sigma} x^\sigma \quad (\text{C.278})$$

$$(\widehat{\sigma^2})_\infty = \frac{1}{n} \sum_{\sigma} (x^\sigma - \mu)^2 \quad (\text{C.279})$$

$$\widehat{\sigma^2} = \frac{1}{n-1} \sum_{\sigma} (x^\sigma - \widehat{\mu})^2 \quad (\text{C.280})$$

Claim 15 Let

$$\underline{\Delta}^\sigma = \underline{x}^\sigma - \mu . \quad (\text{C.281})$$

For any $\sigma_1, \sigma_2, \sigma_3$,

$$\langle \underline{\Delta}^{\sigma_1} \underline{\Delta}^{\sigma_2}, \underline{\Delta}^{\sigma_3} \rangle = 0 . \quad (\text{C.282})$$

proof:

Suppose $\sigma_2 \neq \sigma_3$. Then

$$\langle \underline{\Delta}^{\sigma_1} \underline{\Delta}^{\sigma_2}, \underline{\Delta}^{\sigma_3} \rangle = \underbrace{\langle \underline{\Delta}^{\sigma_2} \rangle}_0 \langle \underline{\Delta}^{\sigma_1}, \underline{\Delta}^{\sigma_3} \rangle = 0 . \quad (\text{C.283})$$

¹¹Do not confuse the sample index σ and the standard deviation σ .

So assume $\sigma_2 = \sigma_3 = \sigma$ and evaluate $\langle \underline{\Delta}^{\sigma_1} \underline{\Delta}^\sigma, \underline{\Delta}^\sigma \rangle$.

Suppose $\sigma_1 \neq \sigma$. Then

$$\langle \underline{\Delta}^{\sigma_1} \underline{\Delta}^\sigma, \underline{\Delta}^\sigma \rangle = \underbrace{\langle \underline{\Delta}^{\sigma_1} \rangle}_0 \langle \underline{\Delta}^\sigma, \underline{\Delta}^\sigma \rangle = 0 . \quad (\text{C.284})$$

So suppose $\sigma_1 = \sigma$ and evaluate $\langle (\underline{\Delta}^\sigma)^2, \underline{\Delta}^\sigma \rangle$.

$$\langle (\underline{\Delta}^\sigma)^2, \underline{\Delta}^\sigma \rangle = \underbrace{\langle (\underline{\Delta}^\sigma)^3 \rangle}_0 - \langle (\underline{\Delta}^\sigma)^2 \rangle \underbrace{\langle \underline{\Delta}^\sigma \rangle}_0 = 0 . \quad (\text{C.285})$$

QED

Claim 16

$$\langle \widehat{\underline{\sigma}^2}, \widehat{\underline{\mu}} \rangle = 0 . \quad (\text{C.286})$$

proof:

$$\langle \widehat{\underline{\sigma}^2}, \widehat{\underline{\mu}} \rangle = \frac{1}{n(n-1)} \sum_{\sigma, \sigma'} \left\langle \left(\underline{x}^\sigma - \frac{1}{n} \sum_{\sigma''} \underline{x}^{\sigma''} \right)^2, \underline{x}^{\sigma'} \right\rangle \quad (\text{C.287})$$

$$= \frac{1}{n(n-1)} \sum_{\sigma, \sigma'} \left\langle \left(\underline{x}^\sigma - \frac{1}{n} \sum_{\sigma''} \underline{x}^{\sigma''} \right)^2, \underline{\Delta}^{\sigma'} \right\rangle \quad (\text{C.288})$$

$$= \frac{1}{n(n-1)} \sum_{\sigma, \sigma'} \left\langle \left(\underline{\Delta}^\sigma - \frac{1}{n} \sum_{\sigma''} \underline{\Delta}^{\sigma''} \right)^2, \underline{\Delta}^{\sigma'} \right\rangle \quad (\text{C.289})$$

$$= 0 \quad \text{by Claim 15 .} \quad (\text{C.290})$$

QED

C.35 Chi-square distribution

This section is based on Ref.[88].

$$\underline{q} \longleftarrow \underline{z}.$$

Figure C.7: Bnet used to define the Chi-square distribution.

Let $q \in \mathbb{R}$ and $\underline{z} = \{z_i\}_{i=0,1,\dots,\nu-1}$ where $z_i \in \mathbb{R}$. Consider the bnet of Fig.C.7. The TPMs, printed in blue, for that bnet, are as follows:¹²

¹²Don't confuse the q independent constant $\mathcal{N}(!q)$ with the normal probability distribution $\mathcal{N}(x; \mu, \sigma^2)$.

$$P(z_i) = \mathcal{N}(z_i; \mu = 0, \sigma^2 = 1) . \quad (\text{C.291})$$

We want

$$\underline{q} = \sum_{i=0}^{\nu-1} (\underline{z}_i)^2 \quad (\text{C.292})$$

so $P(q|z.)$ is a Dirac delta function:

$$P(q|z.) = \delta(q - \sum_{i=0}^{\nu-1} (z_i)^2) . \quad (\text{C.293})$$

Therefore

$$P(q) = \prod_{i=0}^{\nu-1} \left\{ \int dz_i P(z_i) \right\} P(q|z.) \quad (\text{C.294})$$

$$= \mathcal{N}(!q) q^{\frac{\nu}{2}-1} e^{-q/2} = \chi^2(q; \nu) , \quad (\text{C.295})$$

where $\mathcal{N}(!q)$ is a constant that does not depend on q and is adjusted so that $\int_0^\infty dq P(q) = 1$.

C.36 Student's t-distribution

This section is based on Ref.[140].

Let $x[\sigma] = x^\sigma$. Consider i.i.d.real variables $(x^\sigma)_{\sigma=0,1,\dots,n-1}$ such that¹³

$$E[\underline{x}^\sigma] = \mu \quad (\text{C.296})$$

$$\left\langle \underline{x}^\sigma, \underline{x}^{\sigma'} \right\rangle = \delta(\sigma, \sigma') \sigma^2 . \quad (\text{C.297})$$

$$\hat{\mu} = \bar{x} = \frac{1}{n} \sum_{\sigma} x^\sigma \quad (\text{C.298})$$

$$(\hat{\sigma}^2)_\infty = \frac{1}{n} \sum_{\sigma} (x^\sigma - \mu)^2 \quad (\text{C.299})$$

$$\hat{\sigma}^2 = \frac{1}{n-1} \sum_{\sigma} (x^\sigma - \hat{\mu})^2 . \quad (\text{C.300})$$

If we define

¹³Do not confuse the sample index σ and the standard deviation σ .

$$z = \frac{\hat{\mu} - \mu}{\sqrt{\frac{\sigma^2}{n}}} , \quad (\text{C.301})$$

then \underline{z} has a Standard Normal Distribution (SND):

$$P(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} = \mathcal{N}(z; \mu = 0, \sigma^2 = 1) \quad (\text{C.302})$$

But what if we allow the standard deviation σ to fluctuate in the expression Eq.(C.301) for \underline{z} ? Define

$$\underline{t} = \frac{\hat{\mu} - \mu}{\sqrt{\frac{\hat{\sigma}^2}{n}}} . \quad (\text{C.303})$$

Then one can show that \underline{t} has the **Student's t-distribution** $\text{Stud}(t; \nu = n - 1)$ given by:

$$P(t) = \mathcal{N}(!t)(1 + \frac{t^2}{\nu})^{-\frac{\nu+1}{2}} = \text{Stud}(t; \nu = n - 1) \quad (\text{C.304})$$

Note that if we use the approximation $e^x \approx 1 + x + \mathcal{O}(x^2)$, we can show that $\text{Stud}(t)$ tends to the SND when $n \gg 1$:

$$P(t) = \mathcal{N}(!t)(1 + \frac{t^2}{\nu})^{-\frac{\nu+1}{2}} \quad (\text{C.305})$$

$$\approx \mathcal{N}(!t)e^{-\frac{t^2}{2}\frac{\nu+1}{\nu}} \quad (\text{C.306})$$

$$\approx \mathcal{N}(t; \mu = 0, \sigma^2 = 1) . \quad (\text{C.307})$$

Partial derivation of the explicit form of $\text{Stud}(t)$.

Note that the \underline{z} definition Eq.(C.301) and the \underline{t} definition Eq.(C.303), imply that

$$\underline{t} = \underline{z} \underbrace{\sqrt{\frac{\sigma^2}{\hat{\sigma}^2}}}_{\varrho} , \quad (\text{C.308})$$

In the expression $\underline{t} = \underline{z}\varrho$, the random variables \underline{z} and ϱ are independent because, as shown in Section C.34, $\hat{\mu}$ and $\hat{\sigma}^2$ are independent. Therefore, the random variable \underline{t} can be defined using the bnet of Fig.C.8.

The TPMs, printed in blue, for the bnet Fig.C.8, are as follows:

$$P(t|z, \varrho) = \delta(t - z\varrho) \quad (\text{Dirac delta function}) \quad (\text{C.309})$$

$$P(z) = \mathcal{N}(z; \mu = 0, \sigma^2 = 1) \quad (\text{C.310})$$

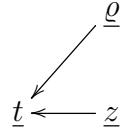


Figure C.8: Bnet used to define the Student's t-distribution.

$$P(\underline{\rho}) = \text{given by Eq.(C.323) below.} \quad (\text{C.311})$$

Note that

$$P(\underline{t} = t) = P(\underline{z}\underline{\rho} = t) \quad (\text{C.312})$$

$$= \int d\underline{\rho} P(\underline{z} = \frac{t}{\underline{\rho}} | \underline{\rho}) P(\underline{\rho}) \quad (\text{C.313})$$

$$= \int d\underline{\rho} \mathcal{N}(\frac{t}{\underline{\rho}}; 0, 1) P(\underline{\rho}) \quad (\text{C.314})$$

$$= \int d\underline{\rho} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(\frac{t}{\underline{\rho}})^2} P(\underline{\rho}) . \quad (\text{C.315})$$

If we define q by

$$q = \frac{n-1}{\underline{\rho}^2}, \quad (\text{C.316})$$

then

$$q = \frac{(n-1)\widehat{\sigma}^2}{\sigma^2} = \frac{1}{\sigma^2} \sum_{\sigma=0}^{n-1} (x^\sigma - \widehat{\mu})^2. \quad (\text{C.317})$$

As a consequence of “Cochran’s Theorem” (see Ref.[90]), q given by Eq.(C.317) must have a Chi-square probability distribution with $\nu = n - 1$ degrees of freedom.¹⁴

$$P(q) = \chi^2(q; \nu = n - 1) \quad (\text{C.318})$$

Henceforth, let $\nu = n - 1$. From the definition Eq.(C.316) of q , we get

$$dq = \frac{-2\nu}{\underline{\rho}^3} d\underline{\rho}. \quad (\text{C.319})$$

¹⁴Note that this q is a quadratic form $q = \vec{x}^T M \vec{x}$, where \vec{x} is an n dimensional column vector with components x^σ , and M is an $n \times n$ matrix. Cochran’s Theorem diagonalizes M and replaces the vectors \vec{x} by equivalent ones in a new basis. Then the number of *DOFs* (degrees of freedom) of the chi-square distribution is the number of non-zero diagonal elements in the diagonalized M (this number is called the rank of M). In the particular case of Eq.(C.317), *DOF* = $n - 1$.

Therefore,

$$P(\varrho)d\varrho = P(q)dq \quad (\text{C.320})$$

$$= \chi^2\left(\frac{\nu}{\varrho^2}; \nu\right) \frac{(-2\nu)}{\varrho^3} d\varrho \quad (\text{C.321})$$

$$= \mathcal{N}(!\varrho) \left(\frac{\nu}{\varrho^2}\right)^{\frac{\nu}{2}-1} e^{-\frac{\nu}{2\varrho^2}} \frac{d\varrho}{\varrho^3} \quad (\text{C.322})$$

$$= \mathcal{N}(!\varrho) \frac{d\varrho}{\varrho^{\nu+1}} e^{-\frac{\nu}{2\varrho^2}}. \quad (\text{C.323})$$

Hence,

$$P(t) = \mathcal{N}(!t) \int_0^\infty \frac{d\varrho}{\varrho^{\nu+1}} e^{-\frac{\nu}{2\varrho^2}} e^{-\frac{1}{2}(\frac{t}{\varrho})^2} \quad (\text{C.324})$$

$$= \mathcal{N}(!t) \int_0^\infty \frac{d\varrho}{\varrho^{\nu+1}} e^{-\frac{1}{2}\frac{t^2+\nu}{\varrho^2}}. \quad (\text{C.325})$$

C.37 Hypothesis testing and 3 classic test statistics (Likelihood, Score, Wald)

Suppose we have data $\vec{x} = [x^\sigma]_{\sigma=0,1,\dots,n_{\text{sam}}-1}$ which is distributed according to a probability distribution $P(\vec{x}; \theta)$ which depends on some parameters $\theta = (\theta_i)_{i=0,1,\dots,n-1} \in \mathbb{R}^n$. We define the

- **Likelihood of θ** by

$$L(\theta) = P(\vec{x}|\theta), \quad (\text{C.326})$$

- the **Maximum Likelihood estimate of θ** by

$$\hat{\theta} = \underset{\theta}{\operatorname{argmax}} L(\theta), \quad (\text{C.327})$$

- the **log likelihood of θ** by

$$LL(\theta) = \ln L(\theta), \quad (\text{C.328})$$

- the **Score or Lagrange Multiplier vector¹⁵** of θ by

¹⁵ $sc_i(\theta) = \partial_{\theta_i} LL(\theta)$ is called a Lagrange multiplier because to maximize $LL(\theta)$ over θ subject to the constraint $\theta = \theta^*$, we maximize the Lagrangian $\mathcal{L} = LL(\theta) - \sum_i \alpha_i [\theta_i - \theta_i^*]$ over (θ_i, α_i) for all i . The latter gives $\alpha_i = \partial_{\theta_i} LL(\theta^*)$ and $\theta = \theta^*$. For a general constrained optimization problem, $\mathcal{L} = LL(\theta) - \sum_i \alpha_i c_i(\theta)$ for some constraint functions $c_i(\theta)$. Hence, $\alpha_i = \partial_{\theta_i} LL / \partial_{\theta_i} c_i$ and $c_i(\theta) = 0$ for all i . So, in general, a Lagrange multiplier is a ratio of two partial derivatives.

$$sc(\theta) = [sc_i(\theta)] \in \mathbb{R}^n , \quad (C.329)$$

where

$$sc_i(\theta) = \partial_{\theta_i} LL(\theta) , \quad (C.330)$$

- and the **Fisher Information Matrix** of θ by

$$FI(\theta) = [FI_{i,j}(\theta)] \in \mathbb{R}^{n \times n} \quad (C.331)$$

where

$$FI_{i,j}(\theta) = -E_{\vec{x}|\theta}[\partial_{\theta_i} \partial_{\theta_j} \overbrace{\ln P(\vec{x}|\theta)}^{LL(\theta)}] . \quad (C.332)$$

Note that if

$$LL(\theta) = \ln P(x|\theta) = \frac{-(\theta - \hat{\theta})^2}{2\sigma^2} + \mathcal{N}(!\theta) , \quad (C.333)$$

then

$$sc(\theta) = \frac{-(\theta - \hat{\theta})}{\sigma^2} \quad (C.334)$$

and

$$FI(\theta) = \frac{1}{\sigma^2} \quad (C.335)$$

In hypothesis testing, one has a **null hypothesis** $H_0: \theta = \theta^*$, and an **alternative hypothesis** $H_1: \theta \neq \theta^*$. We use a **test statistic** $\lambda(\theta^*, \hat{\theta}) \geq 0$ which measures a kind of distance or separation between the estimate $\hat{\theta}$ and the value θ^* for the null Hypothesis H_0 . For some **confidence level** $C > 0$, if $\lambda(\theta^*, \hat{\theta}) > C$, H_0 is **rejected**, whereas if $\lambda(\theta^*, \hat{\theta}) \leq C$, H_0 is **accepted**.

$\alpha = 1 - C$ is called the **significance level**. Usually $\alpha \ll 1$ represents the area under both tails of a Normal distribution, and $C \approx 1$ represents all the area except the tails of a Normal distribution.

Henceforth in this section, we will occasionally use the Einstein summation convention; i.e., implicit sum over repeated indices.

Three classic test statistics are (See Fig.C.9):

1. **Likelihood Ratio test statistic** (Ref.[118].)

$$\lambda_{Li} = 2 \ln \left[\frac{L(\hat{\theta})}{L(\theta^*)} \right] = 2[LL(\hat{\theta}) - LL(\theta^*)] \quad (C.336)$$

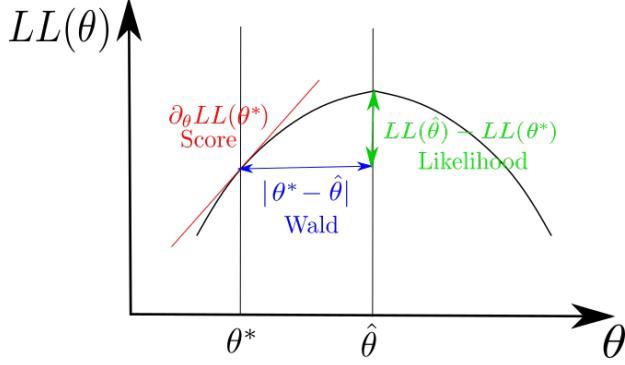


Figure C.9: For $n = 1$ and $\theta \in \mathbb{R}$, this figure shows the geometrical significance of certain quantities that characterize the 3 classic test statistics (Likelihood, Score, Wald) for hypothesis testing.

2. Score (a.k.a. Lagrange multiplier) test statistic (Ref.[136].)

$$\lambda_{Sc} = \partial_{\theta_i} LL(\theta^*) [FI(\theta^*)^{-1}]_{i,j} \partial_{\theta_j} LL(\theta^*) \quad (\text{C.337})$$

$$= \frac{[\partial_{\theta} LL(\theta^*)]^2}{FI(\theta^*)} \quad \text{if } n = 1 \quad (\text{C.338})$$

Doesn't depend on $\hat{\theta}$.

3. Wald test statistic (Ref.[150].)

$$\lambda_{Wa} = (\hat{\theta} - \theta^*)_i \left[\langle \hat{\theta}, \hat{\theta}^T \rangle^{-1} \right]_{i,j} (\hat{\theta} - \theta^*)_j \quad (\text{C.339})$$

$$= \frac{(\theta^* - \hat{\theta})^2}{\langle \hat{\theta}, \hat{\theta} \rangle} \quad \text{if } n = 1 \quad (\text{C.340})$$

More generally, one can replace $\theta^* \rightarrow R\theta^*$ and $\hat{\theta} \rightarrow R\hat{\theta}$ in Eq.(C.339), where θ^* and $\hat{\theta}$ are n dimensional column vectors, and $R \in \mathbb{R}^{n \times n}$. The null and alternative hypotheses become: $H_0 : R\theta = R\theta^*$ and $H_1 : R\theta \neq R\theta^*$. Note that ν is the number of constraints imposed by the null hypothesis. R is called a reparametrization of θ . The Wald test is not reparametrization invariant (i.e., R invariant), but the Likelihood Ratio test is.

Note that if $LL(\theta)$ is given by Eq.(C.333), then $\langle \hat{\theta}, \hat{\theta} \rangle = \sigma^2 = \frac{1}{FI(\theta)}$. Hence,

$$\lambda_{Li} = \lambda_{Sc} = \lambda_{Wa} = \frac{(\hat{\theta} - \theta^*)^2}{\sigma^2} \quad (\text{C.341})$$

Many other commonly used test statistics (or their squares) are special cases of one of the 3 classic test statistics. For example, the z-statistic used with normal distributions, the t-statistic used with the Student t-distribution, the F-statistic used in linear regression, the chi-squared statistic used to do Pearson's chi-squared test.

Asymptotic Behavior

If the data \vec{x} is i.i.d.,

$$P(\vec{x}|\theta) = \prod_{\sigma=0}^{nsam-1} P(x^\sigma|\theta) \quad (\text{C.342})$$

Hence, as $nsam \rightarrow \infty$,

$$LL(\theta) = \ln P(\vec{x}|\theta) \quad (\text{C.343})$$

$$= \sum_{\sigma} \ln P(x^\sigma|\theta) \quad (\text{C.344})$$

$$\rightarrow nsam \sum_x P(x|\theta) \ln P(x|\theta) \quad (\text{C.345})$$

$$= -nsam H(\underline{x}|\theta) \quad (\text{C.346})$$

Thus, *maximizing* the log likelihood $LL(\theta)$ and *minimizing* the entropy $H(\underline{x}|\theta)$ give the same estimate $\hat{\theta}$.

When the data is i.i.d. and $nsam \rightarrow \infty$, it is also possible to prove that the 3 test statistics defined above all tend to the same probability distribution, namely $\mathcal{X}^2(\theta^*; \nu)$, the chi-square distribution with ν degrees of freedom, where $\theta \in \mathbb{R}^n$, $R \in \mathbb{R}^{\nu \times n}$, and $\nu = n$ if $R = 1$.

C.38 Error Bars

Never report measurements without error bars!!

Assume a distribution with mean μ and standard deviation σ for a subpopulation with n samples.

$SE = \frac{\sigma}{\sqrt{n}}$ is called the **standard error**.

Some popular types of error bars:

- **Box and Whiskers plot (a.k.a. Boxplot)**

See Fig.C.10. IQR stands for **Intermediate Quantile Range**. Sometimes, the endpoints of the error bars are taken to be the minimum and maximum samples instead of $Q_1 - 1.5 * IQR$ and $Q_3 + 1.5 * IQR$. The points that fall in the intervals $[\min, Q_1 - 1.5 * IQR]$ and $[Q_3 + 1.5 * IQR, \max]$ are called **outliers**.

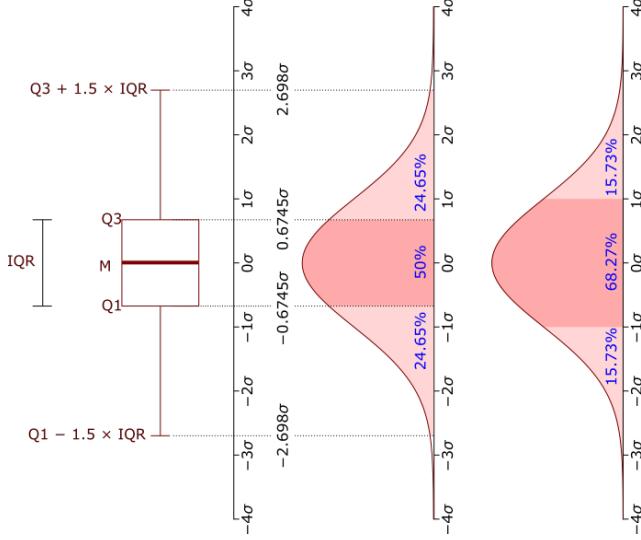


Figure C.10: Boxplot plot for Normal distribution $\mathcal{N}(\mu = 0, \sigma)$. Q_1 and Q_3 are the first and third quantiles, and M is the median (i.e., half-way point). For a non-normal skewed distribution, Q_1 and Q_3 are not equidistant from the median, and the median is not exactly equal to the mean.

- **Standard Deviation**

Error bar endpoints are located one standard deviation away from the mean.

$$\mu - \sigma < \mu < \mu + \sigma \quad (\text{C.347})$$

- **Confidence Interval**

$$\mu - |z^*|SE < \mu < \mu + |z^*|SE \quad (\text{C.348})$$

$|z^*| = 1.96$ for a confidence level of 95%.

The origin of Eq.(C.348) is explained in the next section entitled “Confidence Intervals”. Confidence intervals are derived from the Gaussian in Fig.C.11, which should not be confused with the Gaussian of Fig.C.10. They are different!

C.39 Confidence Interval

Normal distribution with mean μ and standard deviation σ :

$$\mathcal{N}(x; \mu, \sigma^2) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}. \quad (\text{C.349})$$

Standard Normal Distribution (SND):

$$P(z) = \mathcal{N}(z; 0, 1) \quad (\text{C.350})$$

Cumulative distribution for $P(z)$:

$$\Phi(z) = \int_{-\infty}^z dz' P(z') . \quad (\text{C.351})$$

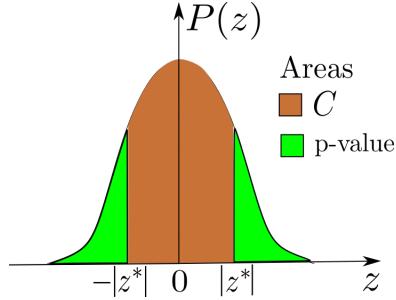


Figure C.11: Interpretation of confidence level C and p-value as areas under curve of the Standard Normal Distribution (SND).

Confidence Level C and corresponding $|z^*|$ value (see Fig.C.11):

$$C = \int_{-|z^*|}^{|z^*|} dz P(z) = \Phi(|z^*|) - \Phi(-|z^*|) = 2 \left(\Phi(|z^*|) - \frac{1}{2} \right) \quad (\text{C.352})$$

Equivalent definition:

$$C = P \left(\underbrace{\frac{|\underline{x} - \mu|}{\frac{\sigma}{\sqrt{n}}}}_{|z|} < |z^*| \right) \quad (\text{C.353})$$

For $C = 95\%$, $|z^*| = 1.960 \approx 2$. For $C = 99\%$, $|z^*| = 2.576$.

Area of each tail in Fig.C.11 is usually called α , and the area of both tails is called the **p-value**:

$$C + \underbrace{2\alpha}_{p\text{-value}} = 1 . \quad (\text{C.354})$$

Estimators¹⁶ of mean μ and standard deviation σ from measurements x^σ of a sub-population Σ_1 of size $n = |\Sigma_1|$:

$$\hat{\mu} = \bar{x} = \frac{1}{n} \sum_{\sigma \in \Sigma_1} x^\sigma \quad (\text{C.355})$$

$$\hat{\sigma}^2 = \frac{1}{n-1} \sum_{\sigma \in \Sigma_1} (x^\sigma - \bar{x})^2 \quad (\text{C.356})$$

¹⁶Don't confuse the sample index σ with the standard deviation σ .

We get from Eq.(C.353), the **Error bars (a.k.a. confidence intervals)** and **Error E (a.k.a. margin of error)**:

$$\text{estimate of } x \text{ with error bars} = \bar{x} \pm |z^*| \underbrace{\frac{\hat{\sigma}}{\sqrt{n}}}_{E} \quad (\text{C.357})$$

$$n = \left(\frac{|z^*| \hat{\sigma}}{E} \right)^2 \quad (\text{C.358})$$

So far, we have assumed that the sub-population (a.k.a. sample population) is normally distributed. This might be false for several reasons. Some red flags: (1) n is too small (according to a rule of thumb derived from Central Limit Theorem, n should be larger than 30 to insure a Normal Distribution). (2) Sub-population not truly random (i.i.d.) because was taken without replacement. In many cases, especially when $n < 30$, the Student's t-distribution models the sub-population statistics much better than the Normal distribution.

The Student's t-distribution $\text{Stud}(t; \nu = n - 1)$, depends on a parameter ν called the number of degrees of freedom. In the case being considered here, ν equals the sub-population size n minus one. When fitting the data with $\text{Stud}()$, variable t replaces variable z , and $\text{Stud}(t; \nu = n - 1)$ replaces the Standard Normal distribution (SND) $\mathcal{N}(z; \mu = 0, \sigma = 1)$. $\text{Stud}()$ is symmetric about the origin like SND, but its tails are fatter. When fitting the data with $\text{Stud}()$, the $|z^*|$ value is replaced by a $|t^*|$ value. Eq.(C.352) is replaced by

$$C = \int_{-|t^*|}^{|t^*|} dt \text{Stud}(t) = \Phi_S(|t^*|) - \Phi_S(-|t^*|) = 2 \left(\Phi_S(|t^*|) - \frac{1}{2} \right), \quad (\text{C.359})$$

where $\Phi_S()$ is the cumulative distribution for $\text{Stud}()$. Also, Eq.(C.357) is replaced by

$$\text{estimate of } x \text{ with error bars} = \bar{x} \pm |t^*| \underbrace{\frac{\hat{\sigma}}{\sqrt{n}}}_{E}. \quad (\text{C.360})$$

Tables of $|t^*|(C, \nu = n - 1)$ are available. Note that $|t^*|$ depends on both C and ν , whereas $|z^*|(C)$ depends only on C .

C.40 Score p-value

When defining error bars and confidence intervals in Fig.C.11, we defined a triplet of values (z, C, p) where $C + p = 1$. In this section, we will consider a different triplet of those values. We will refer to the triplet (z_{th}, C_{th}, p_{th}) used in error bars as the **threshold triplet**, and to the triplet (z_{sc}, C_{sc}, p_{sc}) introduced in this section as the

score triplet. When we do hypothesis testing, if $|z_{sc}| > |z_{th}|$ (i.e., if z_{sc} falls outside the error bars), we say that the null hypothesis is violated. Equivalently, we say the null hypothesis is violated if $p_{sc} < p_{th} = 1 - C_{th} = 0.05$ typically. Most statistics books do a poor job at distinguishing between the threshold and score triplets, and seldom use distinguishing subscripts like th and sc . In this book, we will often drop the sc subscripts, but we will try not to drop the th subscripts. Often, instead of a th subscript, we will use an asterisk superscript. For instance, instead of z_{th} , we might use z^* . When statisticians use the term “p-value”, they are usually referring to the score p-value, although not always.

Given a parameter θ , call $\theta = \theta_0$ (or $\theta < \theta_0$ or $\theta > \theta_0$) the **null hypothesis** h_0 , and call the negation of h_0 (i.e., $\theta \neq \theta_0$ (or $\theta \geq \theta_0$ or $\theta \leq \theta_0$)) the **alternative or opposite hypothesis** h_1 . Assume we are given data $\vec{x} = \{x^\sigma | \sigma \in \Sigma\}$. Assume also that we are given distributions $P(\underline{x} = x|h)$ for $h \in \{h_0, h_1\}$, and $P(\underline{x} = x)$. Now let

$$P(\vec{x}|h) = \prod_{\sigma} P(\underline{x} = x^\sigma|h) \quad (\text{C.361})$$

$$P(\vec{x}) = \prod_{\sigma} P(\underline{x} = x^\sigma) \quad (\text{C.362})$$

(so the x^σ are i.i.d.).

A Bayesian would assume that there is a prior $P(h)$, and use it to calculate $P(h|\vec{x}) = \frac{P(\vec{x}|h)P(h)}{P(\vec{x})}$. $P(\underline{h} = h_0|\vec{x})$ is the probability that the null hypothesis is true. A p-value is a monotonically increasing function of $P(\underline{h} = h_0|\vec{x})$, so Bayesians have no trouble saying that **a p-value is a measure of $P(\underline{h} = h_0|\vec{x})$** , i.e., **a measure of the probability that the null-hypothesis is true**.

Frequentists, on the other hand, believe that h is a “parameter”, not a random variable, so $P(\underline{h} = h_0|\vec{x})$ is undefined. Next, we explain the correct way of thinking about p-values, according to Frequentists. p-values were invented by Frequentists, so it’s worth hearing what they have to say about them. The Frequentist definition is not against Bayesianism, and Bayesians, unlike Frequentists, don’t accuse Frequentists of having a sinfully incorrect definition of p-values. A Bayesian would just say: our definition of p-values (shown in red above) is not incorrect, but the Frequentist definition is more precise than ours, and doesn’t assume a particular form for a prior. We welcome it.

Call the random variable \underline{t} the **test or score statistic** and let t^* be a user defined parameter. \underline{t} and t^* are defined so that when $\underline{t} = t^*$, the h_0 hypothesis is on the threshold between being and not being satisfied. Frequentists define the **p-value** p as

$$p = \begin{cases} P(\underline{t} \geq t^*|h_0) & \text{right-sided-tail, if } h_0 \text{ is } \theta < \theta_0 \\ P(\underline{t} \leq t^*|h_0) & \text{left-sided-tail, if } h_0 \text{ is } \theta > \theta_0 \\ P(|\underline{t}| > |t^*| |h_0) & \text{double-sided-tail, if } h_0 \text{ is } \theta = \theta_0 \end{cases} \quad (\text{C.363})$$

Thus, for a Frequentist, a p-value is a probabilistic weight of the region where the h_0 hypothesis is defined (by the user) to be violated. If that weight is large, then the region where the h_0 hypothesis is defined to be satisfied is small, which means the h_0 hypothesis is expected to be close to the truth. The larger the p-value, the closer h_0 is expected to be near the truth, just like the Bayesian definition says. Note that the p-value is a probability so it ranges in value from 0 to 1.

Suppose we are given a subpopulation with n samples, mean \bar{x} and variance $\hat{\sigma}^2$. Let $\theta_0 = \mu_0$. Define

$$\underline{t} = \underline{z} = \frac{\bar{x} - \mu_0}{\frac{\hat{\sigma}}{\sqrt{n}}} . \quad (\text{C.364})$$

For $n > 30$,

$$P(\underline{z} \geq z^* | h_0) = 1 - \Phi(z^*) = \Phi(-z^*) \quad \text{if } h_0 \text{ is } \mu < \mu_0 \quad (\text{C.365a})$$

$$P(\underline{z} \leq z^* | h_0) = \Phi(z^*) \quad \text{if } h_0 \text{ is } \mu > \mu_0 \quad (\text{C.365b})$$

$$P(|\underline{z}| \geq |z^*| | h_0) = 2\Phi(-|z^*|) \quad \text{if } h_0 \text{ is } \mu = \mu_0 \quad (\text{C.365c})$$

where $\Phi(x)$ is the cumulative distribution for the Standard Normal Distribution $\mathcal{N}(x; \mu = 0, \sigma = 1)$. For $n < 30$, $\Phi()$ is replaced by $\Phi_S()$, where $\Phi_S()$ is the cumulative distribution for the Student t-distribution $\text{Stud}(x; \nu = n - 1)$. Note that Eq.(C.365c) agrees with Eq.(C.353).

The quantity

$$\hat{z}_{sc} = \frac{\bar{x} - \mu_0}{\frac{\hat{\sigma}}{\sqrt{n}}} \quad (\text{C.366})$$

is called the **z score estimator**. If $|\hat{z}_{sc}| > |z^*|$, then the h_0 hypothesis is defined to be violated (for the double sided case).

C.41 Convex/Concave functions, Jensen's Inequality

Suppose $f : \mathbb{R} \rightarrow \mathbb{R}$. $f(x)$ is a **concave function** if looks like a cave (\cap) (i.e., $f''(x) > 0$ if differentiable) and it's a **convex function** if it looks like a valley (\cup) (i.e., $f''(x) < 0$ if differentiable). More generally, if $f : \mathbb{R}^a \rightarrow \mathbb{R}$, $f(x)$ is said to be concave if $f(\alpha x + \beta y) \geq \alpha f(x) + \beta f(y)$ and convex if $f(\alpha x + \beta y) \leq \alpha f(x) + \beta f(y)$ for $\alpha, \beta \in \mathbb{R}$ and $x, y \in \mathbb{R}^a$.

Suppose $f : \mathbb{R}^a \rightarrow \mathbb{R}$ is a convex function. Let α, β be non-negative numbers that sum to 1, and $x, y \in \mathbb{R}^a$. From the definition of convexity, it follows that (see Fig.C.12 for a geometrical representation of this when $a = 1$)

$$f(\alpha x + \beta y) \leq \alpha f(x) + \beta f(y) \quad (\text{C.367})$$

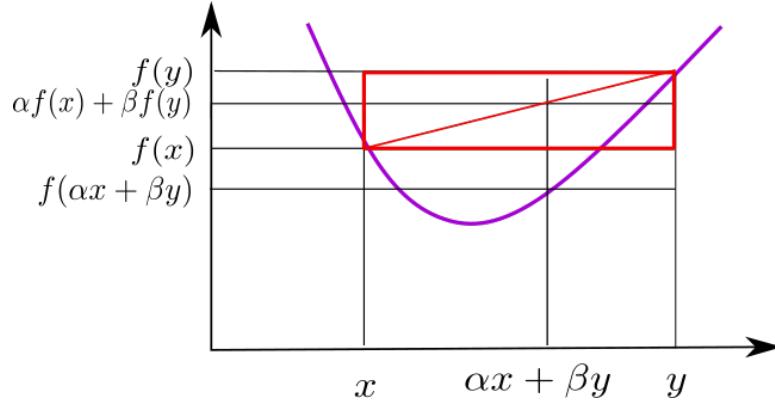


Figure C.12: Jensen's inequality for sum of 2 terms, when $f : \mathbb{R} \rightarrow \mathbb{R}$ is convex.

Jensen's inequality is a simple generalization of this inequality to sums of more than 2 terms. Let $\{p_i\}_{i=0}^n$ be non-negative numbers that sum to one. Also assume that $\{x_i\}_{i=0}^n$ are elements of \mathbb{R}^a . Then

$$f\left(\sum_{i=1}^n p_i x_i\right) \leq \sum_{i=1}^n p_i f(x_i) \quad (\text{C.368})$$

or, written in terms of expected values,

$$f(E[\underline{x}]) \leq E[f(\underline{x}]) \quad (\text{C.369})$$

The same result is true if f is concave instead of convex and we reverse the inequality signs.

C.42 Chebyshev's inequality

Chebyshev's inequality (CI) gives an upper bound for the area under the 2 tails (left and right ones) of a probability distribution.

Below, we follow the common practice of proving CI as a corollary of Markov's Inequality (MI).

Claim 17 (Markov's Inequality) *Let \underline{x} be a non-negative random variable and let $a > 0$. Then*

$$P(\underline{x} \geq a) \leq \frac{E[\underline{x}]}{a} \quad (\text{C.370})$$

proof:

$$\frac{E[\underline{x}]}{a} = \frac{\int_0^\infty dx \ x P(x)}{a} \quad (\text{C.371})$$

$$\geq \frac{\int_a^\infty dx \ x P(x)}{a} \quad (\text{C.372})$$

$$\geq \frac{\int_a^\infty dx \ a P(x)}{a} \quad (\text{C.373})$$

$$= P(\underline{x} \geq a) \quad (\text{C.374})$$

QED

Claim 18 (Chebyshev's inequality) Let \underline{x} be a random variable with mean μ and variance σ^2 . Then for any real number $k > 0$,

$$P(|\underline{x} - \mu| \geq k\sigma) \leq \frac{1}{k^2} \quad (\text{C.375})$$

proof:

$$P(|\underline{x} - \mu| \geq k\sigma) = P(|\underline{x} - \mu|^2 \geq k^2\sigma^2) \quad (\text{C.376})$$

$$\leq \frac{\sigma^2}{k^2\sigma^2} \quad (\text{by MI}) \quad (\text{C.377})$$

$$= \frac{1}{k^2} \quad (\text{C.378})$$

QED

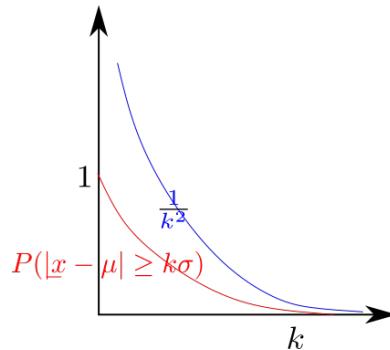


Figure C.13: Pictorial representation of Chebyshev's Inequality (CI). Markov's Inequality (MI) has a similar representation, but k is proportional to \sqrt{a} , so the upper bound for MI goes as $1/a$ instead of $1/k^2$.

See Fig.C.13 for a pictorial representation of CI. Note that the CI approximation is always a bad approximation for small k , and might not become good

even for large k . For example, if the \underline{x} distribution is a box centered at μ , then $P(|\underline{x} - \mu| \geq k\sigma) = 0$ for $k\sigma$ larger than the width of the box, so $1/k^2$ is a terrible approximation for large k too.

A (1 dimensional) Center of Mass (CM) for a unit mass object is an expectation value where the probability distribution is a mass distribution. Hence, it's not too surprising that MI can be interpreted in terms of CMs.

Physics Intuition for MI in terms of CMs:

$$\underbrace{\int_0^\infty dx xP(x)}_{E[\underline{x}] = CM} = \underbrace{\frac{\int_0^a dx xP(x)}{P(\underline{x} < a)}}_{E_{\underline{x} < a}[\underline{x}] = CM^-} \underbrace{P(\underline{x} < a)}_{f^-} + \underbrace{\frac{\int_a^\infty dx xP(x)}{P(\underline{x} > a)}}_{E_{\underline{x} > a}[\underline{x}] = CM^+} \underbrace{P(\underline{x} > a)}_{f^+} \quad (\text{C.379})$$

$$CM \underbrace{(f^+ + f^-)}_{=1} = CM^- f^- + CM^+ f^+ \quad (\text{C.380})$$

$$\geq CM^+ f^+ \quad (\text{C.381})$$

$$\geq af^+ \quad (\text{because } CM^+ \geq a) \quad (\text{C.382})$$

You can think of $CM^+ f^+$ as a torque with moment arm= CM^+ and force= f^+ . In this picture, MI is the approximation of a torque with moment arm= CM and force= 1, by a smaller torque with moment arm= any real $a > 0$ and force= $f^+ \leq 1$.

C.43 Short Summary of Boolean Algebra

See Ref.[85] for more info about this topic.

Suppose $x, y, z \in \{0, 1\}$. Define

$$x \text{ or } y = x \vee y = x + y - xy , \quad (\text{C.383})$$

$$x \text{ and } y = x \wedge y = xy , \quad (\text{C.384})$$

and

$$\text{not } x = \bar{x} = 1 - x , \quad (\text{C.385})$$

where we are using normal addition and multiplication on the right hand sides.¹⁷

Actually, since $x \wedge y = xy$, we can omit writing the symbol \wedge . The symbol \wedge is useful to exhibit the symmetry of the identities, and to remark about the analogous identities for sets, where \wedge becomes intersection \cap and \vee becomes union \cup . However, for practical calculations, \wedge is an unnecessary nuisance.

¹⁷Note the difference between \vee and modulus 2 addition \oplus . For \oplus (a.k.a. XOR): $x \oplus y = x + y - 2xy$.

Associativity	$x \vee (y \vee z) = (x \vee y) \vee z$ $x \wedge (y \wedge z) = (x \wedge y) \wedge z$
Commutativity	$x \vee y = y \vee x$ $x \wedge y = y \wedge x$
Distributivity	$x \wedge (y \vee z) = (x \wedge y) \vee (x \wedge z)$ $x \vee (y \wedge z) = (x \vee y) \wedge (x \vee z)$
Identity	$x \vee 0 = x$ $x \wedge 1 = x$
Annihilator	$x \wedge 0 = 0$ $x \vee 1 = 1$
Idempotence	$x \vee x = x$ $x \wedge x = x$
Absorption	$x \wedge (x \vee y) = x$ $x \vee (x \wedge y) = x$
Complementation	$x \wedge \bar{x} = 0$ $x \vee \bar{x} = 1$
Double negation	$\overline{(\bar{x})} = x$
De Morgan Laws	$\bar{x} \wedge \bar{y} = \overline{(x \vee y)}$ $\bar{x} \vee \bar{y} = \overline{(x \wedge y)}$

Table C.1: Boolean Algebra Identities

Since $x \in \{0, 1\}$,

$$P(\bar{x}) = 1 - P(x) . \quad (\text{C.386})$$

Clearly, from analyzing the simple event space $(x, y) \in \{0, 1\}^2$,

$$P(x \vee y) = P(x) + P(y) - P(x \wedge y) . \quad (\text{C.387})$$

Appendix D

Definition of a Bayesian Network

A **directed graph** $G = (V, E)$ consists of two sets, V and E . V contains the **vertices (nodes)** and E contains the **edges (arrows)**. An arrow $\underline{a} \rightarrow \underline{b}$ is an ordered pair $(\underline{a}, \underline{b})$ where $\underline{a}, \underline{b} \in V$.

The **parents** of a node \underline{x} are those nodes \underline{a} such that there are arrows $\underline{a} \rightarrow \underline{x}$. The **children** of a node \underline{x} are those nodes \underline{b} such that there are arrows $\underline{x} \rightarrow \underline{b}$. A **root node** is a node with no parents. A **leaf node** is a node with no children. The **neighbors** of a node \underline{x} is the set of parents and children of \underline{x} .

A **path** is a set of nodes that are connected by arrows, so that all nodes have 1 or 2 neighbors, but only two nodes (**open path**) or zero nodes (**closed path**) have only one neighbor. A **directed path** is a path in which all the arrows point in the same direction. A **loop** is a closed path; i.e., a path in which all nodes have exactly 2 neighbors. A **cycle** is a directed loop. A **Directed Acyclic Graph (DAG)** is a directed graph that has no cycles.

A **fully connected directed graph** is a directed graph in which every node has all other nodes as neighbors. Figs.D.1 and D.2 show 2 different ways of drawing the same directed graph, a fully connected graph with 4 nodes. Note that a convenient way to label the nodes of a fully connected directed graph with N nodes is to point arrows from \underline{x}_k to \underline{x}_j where $j = 0, 1, 2, \dots, N - 1$ and $k = j - 1, j - 2, \dots, 0$.



Figure D.1: Fully connected directed graph with 4 nodes, drawn as a line.

A **connected graph** is a graph for which there is no way of separating the nodes into two sets so that there is no arrow from one set to the other. A **tree** is a directed graph in which all nodes have a single parent except for a single node called the “root” node which has no parents. A **polytree** is a DAG with no loops.

A **Bayesian network (bnet)** consists of a DAG and a **Transition Probability Matrix (TPM)** associated with each node of the graph. A TPM is often called a **Conditional Probability Table (CPT)**. The **structure** of a bnet is its

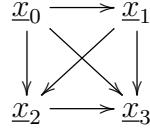


Figure D.2: Fully connected directed graph with 4 nodes, drawn as a square.

DAG alone, sans the TPMs. The **skeleton** of a bnet is the undirected graph beneath the bnet's DAG.

In this book, random variables are indicated by underlined letters and their values by non-underlined letters. We use $S_{\underline{x}}$ to denote the set of states (i.e., values) that a random variable \underline{x} can assume. Each node of a bnet is labelled by a random variable. Thus, $\underline{x} = x$ means that node \underline{x} is in state x .

Some sets of nodes associated with each node \underline{a} of a bnet

- $ch(\underline{a})$ = children of \underline{a} .
- $pa(\underline{a})$ = parents of \underline{a} .
- $nb(\underline{a}) = pa(\underline{a}) \cup ch(\underline{a})$ = neighbors of \underline{a} .
- $de(\underline{a}) = \cup_{n=1}^{\infty} ch^n(\underline{a}) = ch(\underline{a}) \cup ch \circ ch(\underline{a}) \cup \dots$, descendants of \underline{a} .
- $an(\underline{a}) = \cup_{n=1}^{\infty} pa^n(\underline{a}) = pa(\underline{a}) \cup pa \circ pa(\underline{a}) \cup \dots$, ancestors of \underline{a} .

In this book, we will use $\underline{a}.$ to indicate a **multi-node (node set, node array)** $\underline{a}. = (\underline{a}_j)_{j=0,1,\dots,na-1}$. We will often treat multinodes as if they were sets, and combine them with the usual set operators. For instance, for two multinodes $\underline{a}.$ and $\underline{b}.$, we define $\underline{a}.\cup\underline{b}.$, $\underline{a}.\cap\underline{b}.$, $\underline{a}.\dashv\underline{b}.$ and $\underline{a}.\subset\underline{b}.$ in the obvious way. We will indicate a singleton set (single node multi-node) $\underline{a}. = \{\underline{a}\}$ simply by $\underline{a}.. = \underline{a}.$ For instance, $\underline{a}..\dashv\underline{b} = \underline{a}.. - \{\underline{b}\}.$

The TPM of a node \underline{x} of a bnet is a matrix of probabilities $P(\underline{x} = x | pa(\underline{x}) = a.)$, where $x \in S_{\underline{x}}$ and $a. \in S_{pa(\underline{x})}$.

A **deterministic node** is a node such that its TPM is of the form

$$P(\underline{x} = x | pa(\underline{x}) = a.) = \delta(x, f(a.)) \quad (\text{D.1})$$

for some function $f : S_{pa(\underline{x})} \rightarrow S_{\underline{x}}$, where $\delta(x, y)$ is the Kronecker delta function.

A bnet with nodes $\underline{x}.$ represents a probability distribution

$$P(x.) = \prod_j P(\underline{x}_j = x_j | (\underline{x}_k = x_k)_{k: \underline{x}_k \in pa(\underline{x}_j)}). \quad (\text{D.2})$$

Note that for a fully connected bnet with N nodes, Eq.(D.2) becomes

$$P(x.) = \prod_{j=0}^{N-1} P(x_j | (x_k)_{k=j-1,j-2,\dots,0}). \quad (\text{D.3})$$

For example, if $N = 4$, Eq.(D.3) becomes

$$P(x_0, x_1, x_2, x_3) = P(x_3|x_2, x_1, x_0)P(x_2|x_1, x_0)P(x_1|x_0)P(x_0). \quad (\text{D.4})$$

We see that Eq.(D.3) is just the chain rule for conditional probabilities.

Of course, not all bnets are fully connected. So what determines whether in a bnet with nodes \underline{x} and \underline{y} , we draw or not draw, an arrow connecting the two nodes? Henceforth, we will say that there is a **causal correlation** between nodes \underline{x} and \underline{y} , if a correlation between \underline{x} and \underline{y} persists after we amputate the links of \underline{x} to all its neighbors, except \underline{y} . A bnet has a link between \underline{x} and \underline{y} (i.e., an arrow $\underline{x} \rightarrow \underline{y}$ or an arrow $\underline{y} \rightarrow \underline{x}$) iff there is a causal correlation between \underline{x} and \underline{y} . \underline{x} and \underline{y} can be correlated but not causally correlated. For example, in the bnet $\underline{x} \rightarrow \underline{m} \rightarrow \underline{y}$, there is no arrow between \underline{x} and \underline{y} , but \underline{x} and \underline{y} are correlated.

In this book, we use the following conventions for bnet diagrams:

Random variables are underlined and their values are not. For example, $\underline{a} = a$ means the random variable \underline{a} takes the value a . A diagram with all its nodes underlined represents a Bayesian Network (bnet), whereas the same diagram with the letters not underlined represents a specific **instantiation** of that bnet. For example $\underline{a} \rightarrow \underline{b} \rightarrow \underline{c}$ represents the bnet with full probability distribution $P(c|b)P(b|a)P(a)$, whereas $a \rightarrow b \rightarrow c$ represents $P(c|b)P(b|a)$. Note that, for convenience, we define $a \rightarrow b \rightarrow c$ to exclude the priors of root nodes such as $P(a)$.

If \underline{a} is a root node, then $\sum a$ signifies a weighted sum $\sum_a P(a)$. For example,

$$\sum a \rightarrow b \rightarrow c = \sum_a P(c|b)P(b|a)P(a) \quad (\text{D.5})$$

If \underline{a} is not a root node, then $\sum a$ signifies a simple unweighted sum \sum_a . For example,

$$x \rightarrow \sum a \rightarrow y = \sum_a P(y|a)P(a|x) \quad (\text{D.6})$$

Two bnets are equated if their full probability distributions (i.e., their full instantiations) are equal numerically. For example,

$$\underline{a} \rightarrow \underline{b} \rightarrow \underline{c} = P(c|b)P(b|a)P(a) = \underline{a} \leftarrow \underline{b} \leftarrow \underline{c} \quad (\text{D.7})$$

In this book, Unobserved nodes are indicated by enclosing them in a dashed circle. For example, (\underline{u}) .

In Chapter 33, we define a measure of goodness (G) of causal fit (CF) for each bnet of a finite set of bnets \mathcal{G} .

Given a dataset of samples for the random variables $(\underline{x}_i)_{i=0,1,\dots,N-1}$, and a finite set of possible bnets \mathcal{G} , set \mathcal{G} may contain several bnets (differing in the direction of some arrows) that fit the data well causally. However, the one with the highest GCF is most likely to be used by Nature. In this book, we will refer to that single one as

the **best CF bnet**. We will also refer to a bnet in \mathcal{G} that has a high (resp., low) value of GCF, though not necessarily the highest (resp., lowest), as a **good (resp., bad) CF bnet**.¹

It's important to realize that bnets that are not a good CF are far from useless; they are frequently used as intermediate calculational tools. They are incorrect causally, but they aren't incorrect numerically. For instance, in Bayes rule, we switch from a good CF bnet $P(x|\theta)P(\theta) = \underline{x} \leftarrow \underline{\theta}$ to a bad CF bnet $P(\theta|x)P(x) = \underline{x} \rightarrow \underline{\theta}$, where x is the data and θ are the parameters.

¹In this book, the term “causal bnet” will mean a good CF bnet. Pearl is fond of using the term “causal bnet”, but he uses it in a different sense that does not allude to a measure of goodness of causal fit.

Appendix E

Bayesian Networks, Causality and the Passage of Time

This chapter is based on a blog post (see Ref.[64]) from my blog “Quantum Bayesian Networks”.

E.1 Unifying Principle of this book

The unifying principle of this book is Bayesian Networks (bnets). The main goal of this book is to explain as much of Artificial Intelligence (AI) and Machine Learning (ML) as possible using bnets.

Bayesian Networks are a graphical representation of the chain rule for conditional probabilities. They are not a “heuristic algorithm” like XGBoost or Neural Nets. They are a very simple, intuitive, basic and general definition. I would say that the definition of a Bayesian Network is as important to Probability Theory as the definition of a Group is to Abstract Algebra. Algebraic groups are never going to go out of fashion and neither are B nets.

An Artificial Neural Net can be defined as a Bayesian Network with a layered structure, and such that all its nodes are deterministic¹. A decision tree is not exactly a Bayesian Network, but it can be trivially replaced by an equivalent B net that has the same tree structure (for more details about this equivalence, see the Chapter 14 on decision trees). The SCM diagrams favored by Pearl are just bnets whose internal nodes are deterministic and external ones are probabilistic. In fact, as I show in this book, most methods in AI can be understood in terms of B nets—just like many theorems in Abstract Algebra can be understood in terms of groups.

¹Neural Nets (NNs) are DAGs, but they contain a lot of spurious arrows whose direction or very existence has no causal motivation, and they could be missing other arrows which would have a causal motivation. So I like to say that NNs are acausal DAGs.

E.2 You say tomato, I say tomato

In this chapter, I will use the terms Bayesian Network (bnet), causal model and DAG as if they were synonymous. My justification for doing this is as follows.

A Bayesian Network is a DAG + probability tables. One can easily compute the probability tables from DAG + Dataset. Therefore,

You say DAG+Dataset, I say Bayesian Network.

The use of the terms “causal model” and “DAG”, as an alternative to the term “Bayesian Network”, has become popular in the last decade among economists, epidemiologists, AI researchers and even Judea Pearl himself. It seems some people think “causal models” and “DAGs” are revolutionary, whereas Bayesian Networks are a concept that was tried 25 years ago, and has been replaced since then by stuff that works better. But any time you have a Dataset, which is almost always true in practice in Economics, Epidemiology and AI, a DAG implies a Bayesian Network and vice versa.

E.3 A dataset is causal model free

Time and time again, Judea Pearl makes the point on Twitter to neural net advocates that they are trying to do a provably impossible task, to derive a causal model from data. I could be wrong, but this is what I think he means.

When Pearl says “data”, he is referring to what is commonly called a dataset. A dataset is a table of data, where all the entries of each column have the same units, and measure a single feature, and each row refers to one particular sample or individual. Datasets are particularly useful for estimating probability distributions and for training neural nets. When Pearl says a “causal model”, he is referring to a DAG (directed acyclic graph) or a bnet (Bayesian Network= DAG + probability table for each node of DAG).

Sure, you can try to derive a causal model from a dataset, but you’ll soon find out that you can only go so far.

The process of finding a partial causal model from a dataset is called structure learning (SL). SL can be done quite nicely with Marco Scutari’s open source program bnlearn. There are 2 main types of SL algorithms: score-based and constraint based. The first and still very competitive constraint-based SL algorithm was the Inductive Causation (IC) algorithm proposed by Pearl and Verma in 1991. So Pearl is quite aware of SL. The problem is that SL often cannot narrow down the causal model to a single one. It finds an undirected graph (UG), and it can determine the direction of some of the arrows in the UG, but it is often incapable, for well understood fundamental—not just technical—reasons, of finding the direction of ALL the arrows of the UG. So it often fails to fully specify a DAG model.

Let’s call the ordered pair (dataset, causal model) a dataset++ . Then what I believe Pearl is saying is that a dataset is causal model-free or causal model-less

(although sometimes one can find a partial causal model hidden in there). A dataset is not a dataset++.

E.4 What is causality?

What is Causality, really, and how do Bayesian Networks (a.k.a. Causal Models, DAGs) encode it? For me, Causality is a time-induced ordering between two events, the transmission of information (and its accompanying energy) from the earlier of the two events to the later one, and the physical response of the later event to the reception of that information.

Note that this definition of causality does not mention correlation. It is often said that even though correlation does not imply causation, causation does imply correlation. But the latter statement is false; there are scenarios, albeit unusual, “fine tuned” ones, in which there is causation without correlation. For example, consider a bnet with arrows $\underline{x} \rightarrow \underline{y}$ and $\underline{x} \leftarrow \underline{c} \rightarrow \underline{y}$. When we amputate the arrows entering \underline{x} , a correlation between \underline{x} and \underline{y} persists, so we say \underline{x} causes \underline{y} . Even though \underline{x} causes \underline{y} , it’s possible to tune the probabilities of the bnet so that the effect of the path $\underline{x} - \underline{c} - \underline{y}$ and the effect of the direct path $\underline{x} - \underline{y}$ cancel each other out and produce zero correlation between \underline{x} and \underline{y} .

Henceforth, we will say that there is a **causal correlation** between nodes \underline{x} and \underline{y} , if a correlation between \underline{x} and \underline{y} persists after we amputate the links of \underline{x} to all its neighbors, except \underline{y} . A bnet has a link between \underline{x} and \underline{y} (i.e., an arrow $\underline{x} \rightarrow \underline{y}$ or an arrow $\underline{y} \rightarrow \underline{x}$) iff there is a causal correlation between \underline{x} and \underline{y} . \underline{x} and \underline{y} can be correlated but not causally correlated. For example, in the bnet $\underline{x} \rightarrow \underline{m} \rightarrow \underline{y}$, there is no arrow between \underline{x} and \underline{y} , but \underline{x} and \underline{y} are correlated.

The nodes of a bnet represent random variables. Some of those random variables are clearly events (i.e., they occur at a definite time). For example, let $D=0$ if a patient is not given a drug, $D=1$ if he/she is given it. D occurs at a definite time. But other random variables represent qualities which do not occur at a definite time. For example, $G=\text{gender}=\text{male}, \text{female}$. G does not occur at a definite time. But even in the case of a quality like G , its value is first decided at birth, so one can ascribe to G a particular, albeit fuzzy time interval during which it is decided. If $M=0(\text{single}), 1(\text{married})$, then we can assign to M the day of the marriage. Both the time interval assigned to G and to M are somewhat ambiguous, but still, most people would say that G occurs before M (if a marriage occurs at all). Saying the opposite, that M occurs before G , seems pretty hard to understand. If two nodes A and B of a bnet have time intervals ascribed to them such that the time interval of A does not clearly occur before or after the time interval of B , and if also there is a large causal correlation between A and B , then it probably does not matter much whether one draws an arrow from A to B , or the opposite.

Now that we understand that the arrows in a bnet really do encode the direction of time, it becomes clear why a dataset does not fully specify a bnet. By a

dataset (think of a dataframe in Pandas or R), I mean an array of numbers where the columns refer to features and the rows refer to individuals in a population. The column labels of the dataset become the node names of the bnet. Nowhere in a dataset is there any indication of the time ordering of the features. Hence, it's impossible to create, from a dataset alone, a bnet, because bnets do carry such time-ordering information.

Chapter 34 discusses Granger Causality (GC). The critics of GC point out that it assumes, somewhat erroneously, that if event A precedes B and the two events are correlated, A must cause B. I agree. Most roosters crow in response to the stimulus of the sunrise light. A rooster could crow before sunrise if, for example, he had an alarm clock that woke him up 30 minutes before sunrise, but such cases are uncommon, and seem to involve other intermediate events. The moral is that time ordering and correlation are not sufficient conditions for causality. To establish causality with more certainty, one also needs a pinch of prior expert knowledge, or one must gain that expert knowledge through “do” operator experimentation.

E.5 Bayesian Networks and the passage of time

Now that we understand that a bnet's arrows are encoding roughly the passage of time, it becomes possible to glean from this insight a simple method, which, although not very rigorous, is really helpful to me. I will illustrate said method with the famous “Asia” bnet in Fig.E.1. In this bnet, all nodes have two possible values, 0 and 1.

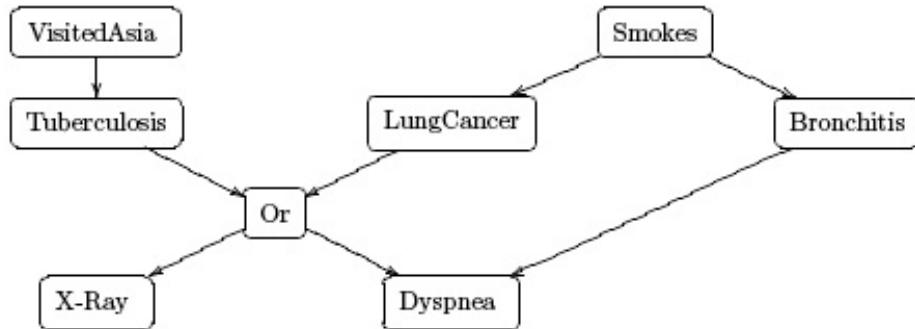


Figure E.1: Asia bnet. Dyspnea=trouble breathing

Given a dataset for this bnet, one can calculate the correlation between every 2 features of the dataset. The feature names become the node names, and links are drawn between any 2 nodes whose correlation is causal and greater in absolute value to some threshold value. This gives an undirected graph that can be obtained from bnet Fig.E.1 by erasing the directions of the arrows. So how can we guess the

directions of the arrows? Well, one uses a little bit of “expert knowledge” to conclude that

time(Visited Asia) < time(Tuberculosis) < time(Or) < time(X-Ray, Dyspnea)

Also

time(smokes) < time(LungCancer, Bronchitis) < time(Or) < time(Dyspnea)

If $\text{time}(A) < \text{time}(B)$, then $A \rightarrow B$. Like I said before, the times we ascribe to these events are somewhat fuzzy and open to debate, so this algorithm is far from being rigorous. But often, saying that $\text{time}(A) < \text{time}(B)$ makes much more sense than saying that $\text{time}(B) < \text{time}(A)$. When in doubt about the best direction to give to an arrow of an undirected graph, I recommend calculating a Goodness of Causal Fit metric (see Chapter 33) which makes use of “do” operator experimentation.

A dataset cannot fully specify a bnet because it lacks time ordering info. A dataset also cannot do the harder task of specifying a bnet that is a good causal fit to the problem, because it lacks time ordering info AND prior expert knowledge AND expert knowledge gained from posterior “do” operator experimentation.

E.6 Advice for the DAG-phobic

DAGs are your friends. DAGs should be easy and fun to dream up. After all, I am convinced that DAGs are an integral part of how humans think, so they should come naturally to us. Nevertheless, many people are scared of, or detest, DAGs. I think it’s because they fail to grasp the following 3 things:

1. DAGs are not unique. Stop thinking that you have to find the unique DAG for the situation being considered. You just have to find a DAG that is a good causal fit for the situation. If a DAG is too complicated, you can always simplify it by merging several nodes into a single more abstract one, or by summing over unwanted nodes.
2. The nodes of a DAG are roughly ordered from past to present. The arrows of a DAG roughly reflect the passage of time.
3. DAGs represent scientific hypotheses that can and should be tested with do experiments. Causal Inference is an application of the scientific method, which consists of the following steps: formulate hypothesis (DAG), devise experiment to test it, test it.

Chapter 1

AdaBoost

This chapter is based on Ref.[75].

Adaptive Boosting (AdaBoost) is a method of constructing a strong classifier function as a linear combination of an ensemble of weak classifier functions.

Below, we will abbreviate “ensemble classifiers” by “e-classifiers” and “weak classifiers” by “w-classifiers”.

Chapter 14 defines decision trees (dtrees) and explains how to construct them. A **tree stump** is a dtree with only one parent and 2 children nodes. Usually the w-classifier functions for AdaBoost come from tree stumps (because tree stumps are w-classifiers and simple to compute), but the core AdaBoost algorithm is oblivious to where the w-classifier functions came from.

Boosting (see this chapter on AdaBoost and Chapter 91 on XGBoost) and bagging (see Chapter 66 on Random Forest) are two methods of building a classifier function from an ensemble of classifier functions. These two methods are most commonly applied to dtrees: Boosting for an ensemble of small dtrees, and Bagging for a random forest (which is an ensemble of dtrees that are usually much more complicated than small dtrees).

1.1 AdaBoost for general ensemble of w-classifiers

In this section we discuss the core AdaBoost algorithm, valid for a generic ensemble of w-classifiers.

Let $L = [0, 1, 2, \dots, nsam - 1]$ be a list of individuals (samples) in a population. In this chapter, we will use the notation $A^\sigma = A[\sigma]$ and $\vec{A} = [A^\sigma : \sigma \in L]$ for a list (vector, 1-D array) indexed by L . We will refer to $DS = (\vec{x}, \vec{y})$ where $x^\sigma \in S_{\underline{x}}$, $y^\sigma \in S_y$, as a dataset. Let $T = \{0, 1, \dots, nt - 1\}$. Let $x^\sigma = (x_0^\sigma, x_1^\sigma, \dots, x_{nt-1}^\sigma) \in S_{\underline{x}_0} \times S_{\underline{x}_1} \times \dots \times S_{\underline{x}_{nt-1}} = S_{\underline{x}}$.

AdaBoost assumes that the classifier class set S_y and all the feature sets $S_{\underline{x}_t}$ are binary: $S_y = S_{\underline{x}_t} = \{-1, 1\}$ for all $t \in T$.

Suppose that we are given an ensemble of nt **w-classifiers** $Y_t : S_{\underline{x}} \rightarrow \{-1, 1\}$, where $t \in T$. Suppose we want to find **intermediate e-classifiers** $\mathcal{Y}_t : S_{\underline{x}} \rightarrow \mathbb{R}$ given

by

$$\mathcal{Y}_t(x^\sigma) = \sum_{t'=0}^t \alpha_{t'} Y_{t'}(x^\sigma) \in \mathbb{R} \quad (1.1)$$

for $t \in T$ and a **final e-classifier** given by

$$\mathcal{Y}_{fin}(x^\sigma) = \text{sign}(\mathcal{Y}_{nt-1}(x^\sigma)) \in \{-1, 1\}. \quad (1.2)$$

The AdaBoost algorithm yields a set of coefficients α_t for which the final e-classifier is much stronger (i.e., less error prone) than any of the w-classifiers of the ensemble.

Note that

$$\mathcal{Y}_t(x^\sigma) = \mathcal{Y}_{t-1}(x^\sigma) + \alpha_t Y_t(x^\sigma) \quad (1.3)$$

for $t \in T$ if we define $\mathcal{Y}_{-1} = 0$. Hence

$$\underbrace{e^{-y^\sigma \mathcal{Y}_t(x^\sigma)}}_{Z_t w_{t+1}^\sigma} = \underbrace{e^{-y^\sigma \mathcal{Y}_{t-1}(x^\sigma)}}_{Z_{t-1} w_t^\sigma} e^{-\alpha_t y^\sigma Y_t(x^\sigma)} \quad (1.4)$$

where the **weights** w_t^σ and the **partition function** Z_t are defined by

$$w_{t+1}^\sigma = \begin{cases} 1/nsam & \text{for } t = -1 \\ \frac{\exp(-y^\sigma \mathcal{Y}_t(x^\sigma))}{Z_t} & \text{for } t \geq 0 \end{cases} \quad (1.5)$$

and

$$Z_t = \sum_\sigma e^{-y^\sigma \mathcal{Y}_t(x^\sigma)}. \quad (1.6)$$

Note that the probability distribution $P(\sigma|t+1) = w_{t+1}^\sigma$ of weights at time $t+1$ emphasizes (i.e., gives higher probability to) the errors (i.e., occurrences of $y^\sigma \mathcal{Y}_t(x^\sigma) = -1$ for some population individual σ) of the previous (i.e., at time t) intermediate e-classifier \mathcal{Y}_t . In other words, every new intermediate e-classifier \mathcal{Y}_{t+1} concentrates on those individuals σ on which the previous e-classifier \mathcal{Y}_t performed poorly.

Note also that the partition function Z_t is a good measure of the **classification error** (i.e., occurrences of $y^\sigma \mathcal{Y}_t(x^\sigma) = -1$) of \mathcal{Y}_t . We will therefore use Z_t for that purpose, as an error measure.

For $t > 1$, we have

$$Z_t = \sum_\sigma e^{-y^\sigma \mathcal{Y}_t(x^\sigma)} \quad (1.7)$$

$$= \sum_\sigma \underbrace{e^{-y^\sigma \mathcal{Y}_{t-1}(x^\sigma)}}_{Z_{t-1} w_t^\sigma} e^{-\alpha_t y^\sigma Y_t(x^\sigma)} \quad (1.8)$$

$$= Z_{t-1} E_\sigma [e^{-\alpha_t y^\sigma Y_t(x^\sigma)}] \quad (1.9)$$

Define the **success rate** by

$$S_t = \sum_{\sigma} w_t^{\sigma} \mathbb{1}(y^{\sigma} Y_t(x^{\sigma}) = 1) \quad (1.10)$$

$$= E_{\sigma}[\underbrace{\mathbb{1}(y^{\sigma} Y_t(x^{\sigma}) = 1)}_{\text{iff } y^{\sigma} = Y_t(x^{\sigma})}] \quad (1.11)$$

and the **failure rate** by

$$F_t = \sum_{\sigma} w_t^{\sigma} \mathbb{1}(y^{\sigma} Y_t(x^{\sigma}) = -1) \quad (1.12)$$

$$= E_{\sigma}[\underbrace{\mathbb{1}(y^{\sigma} Y_t(x^{\sigma}) = -1)}_{\text{iff } y^{\sigma} \neq Y_t(x^{\sigma})}] . \quad (1.13)$$

Note that

$$S_t + F_t = 1 , \quad (1.14)$$

and

$$Z_t = Z_{t-1}(e^{-\alpha_t} S_t + e^{+\alpha_t} F_t) . \quad (1.15)$$

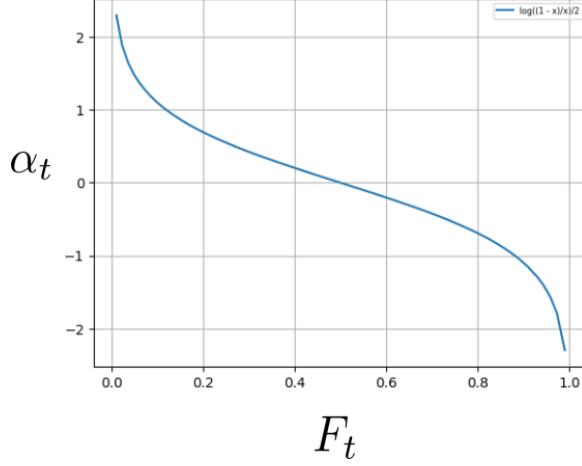


Figure 1.1: Plot of function $\alpha_t = \frac{1}{2} \ln \frac{1-F_t}{F_t}$.

We can find the α_t values that minimize the classification error Z_t , and then evaluate Z_t at those optimum α_t values:

$$\frac{dZ_t}{d\alpha_t} = Z_{t-1}(-e^{-\alpha_t} S_t + e^{\alpha_t} F_t) = 0 \quad (1.16)$$

$$e^{2\alpha_t} = \frac{S_t}{F_t} \quad (1.17)$$

$$\alpha_t = \frac{1}{2} \ln \frac{S_t}{F_t} = \frac{1}{2} \ln \frac{1 - F_t}{F_t} \quad (1.18)$$

$$\frac{Z_t}{Z_{t-1}} = 2\sqrt{S_t F_t} = 2\sqrt{(1 - F_t) F_t} \leq 1 \quad (1.19)$$

$f(x) = 2\sqrt{(1 - x)x}$ for $x \in [0, 1]$ is dome shaped and its maximum is 1, which is achieved iff $x = \frac{1}{2}$

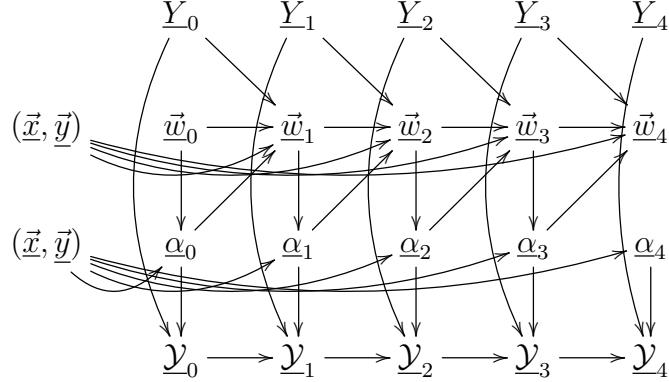


Figure 1.2: Bnet for AdaBoost with 5 w-classifiers, $nt = 5$. All nodes labelled (\vec{x}, \vec{y}) are the same node.

The AdaBoost algo described in this chapter is summarized by the bnet of Fig.1.2. The TPMs, printed in blue, for that bnet, are as follows:

$$P(w_0^\sigma) = \frac{1}{nsam} \quad (1.20)$$

for all σ .

$$P(\vec{w}_{t+1} | \vec{w}_t, \alpha_t, Y_t, \vec{x}, \vec{y}) = \prod_\sigma \mathbb{1}(\quad w_{t+1}^\sigma = \frac{w_t^\sigma e^{-\alpha_t y^\sigma Y_t(x^\sigma)}}{\sum_\sigma \text{numerator}} \quad) \quad (1.21)$$

for $t \geq 0$.

$$P(\alpha_t | \vec{w}_t, \vec{x}, \vec{y}) = \mathbb{1}(\quad \alpha_t = \frac{1}{2} \ln \frac{1 - F_t}{F_t} \text{ where } F_t = F_t(\vec{w}_t, \vec{x}, \vec{y}) \quad) \quad (1.22)$$

for $t \in T$.

$$P(\mathcal{Y}_t | \mathcal{Y}_{t-1}, \alpha_t, Y_t) = \mathbb{1}(\quad \mathcal{Y}_t = \mathcal{Y}_{t-1} + \alpha_t Y_t \quad) \quad (1.23)$$

for $t \in T$, where $\mathcal{Y}_{-1} = 0$.

1.2 AdaBoost for ensemble of tree stumps

Keep in mind that AdaBoost assumes $S_y = S_{x_t} = \{-1, 1\}$ for all $t \in T = \{0, 1, \dots, nt - 1\}$. In order to implement AdaBoost, we need to specify nt w-classifiers $Y_t : \{-1, 1\}^{nt} \rightarrow \{-1, 1\}$ for $t \in T$. One can either specify the nt w-classifiers a priori or build them on-the-fly.

- **w-classifiers specified a priori**

Define a classifier for each feature x_t where $t \in T$ by:

$$Y_t(x^\sigma) = x_t^\sigma \in \{-1, 1\} \quad (1.24)$$

Hence, for this classifier, $y^\sigma Y_t(x^\sigma) = y^\sigma x_t^\sigma$.

- **w-classifiers built on-the-fly**

Recall dataset $(\vec{x}, \vec{y}) = [(x^\sigma, y^\sigma) : \sigma \in L]$ is indexed by the list $L = [0, 1, \dots, nsam - 1]$. If L_j is a list (possibly with duplicate items) such that $set(L_j) \subset set(L)$, then define $DS_j = (\vec{x}, \vec{y})_{L_j} = ((x^\sigma)_{\sigma \in L_j}, (y^\sigma)_{\sigma \in L_j})$. We will refer to DS_j as the **restriction of (\vec{x}, \vec{y}) to L_j** .

The idea behind on-the-fly w-classifiers is to choose $Y_t(x^\sigma) = x_t^\sigma$, where x_t is the feature with the lowest Gini in the current dataset $(\vec{x}, \vec{y})_{L_t}$. To build $(\vec{x}, \vec{y})_{L_t}$, we select at random from $L = [0, 1, \dots, nsam - 1]$, a list L_t of the same length as L , using the probability distribution \vec{w}_{t-1} . By choosing L_t with probabilities \vec{w}_{t-1} , we emphasize individuals σ that are failing. Then we define $(\vec{x}, \vec{y})_{L_t}$ as the restriction of (\vec{x}, \vec{y}) to L_t .

Perhaps a causal diagram will make all these new steps clearer to the reader. The bnet Fig.1.3 is a modification of the bnet Fig.1.2 to include these new steps. The TPMs, printed in blue, for new or changed nodes, are as follows:

$$P(Y_t | (\vec{x}, \vec{y})_{L_t}) = \mathbb{1}(\text{ } Y_t(x^\sigma) = x_t^\sigma \text{ where } x_t \text{ is feature in } (\vec{x}, \vec{y})_{L_t} \text{ the with lowest Gini. }) \quad (1.25)$$

$$P(L_{t+1}^\sigma = \sigma' | \vec{w}_t) = w_t^{\sigma'} \quad (1.26)$$

$$P((\vec{x}, \vec{y})_{L_t} | L_t, (\vec{x}, \vec{y})) = \mathbb{1}(\text{ } (\vec{x}, \vec{y})_{L_t} = \text{ restriction of } (\vec{x}, \vec{y}) \text{ to } L_t \text{ }) \quad (1.27)$$

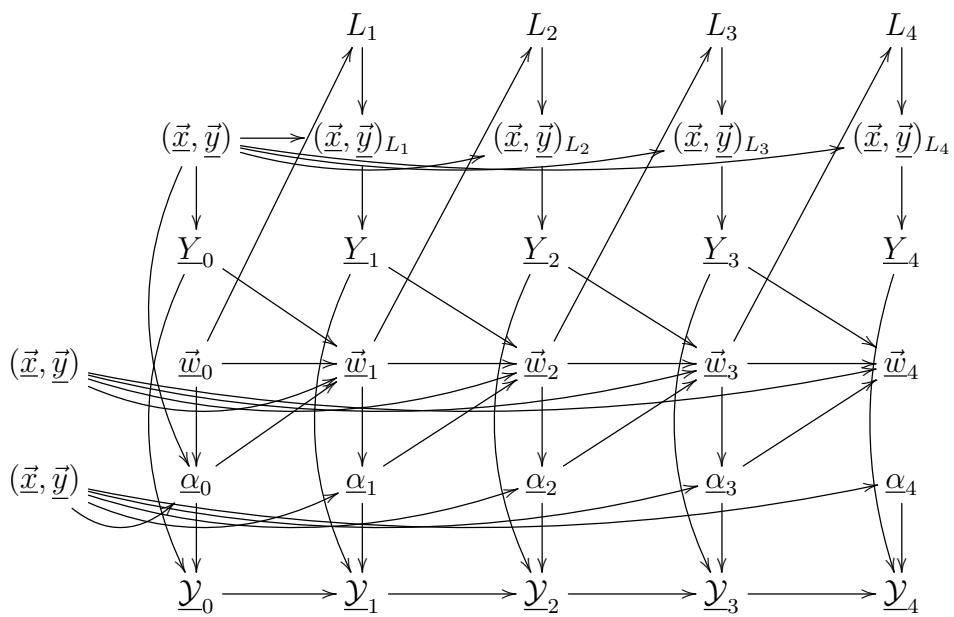


Figure 1.3: Modification of the bnet of Fig.1.2 to include on-the-fly generation of the w-classifiers.

Chapter 2

ANOVA

ANOVA stands for “Analysis of Variance”.

2.1 Law of Total Variance

Claim 19 (Law of Total Variance) Suppose $P : S_{\underline{x}} \times S_{\underline{y}} \rightarrow [0, 1]$ is a probability distribution. Suppose $f : S_{\underline{x}} \times S_{\underline{y}} \rightarrow \mathbb{R}$ and $f = f(x, y)$. Then

$$Var_{\underline{x}, \underline{y}}(f) = E_{\underline{y}}[Var_{\underline{x}|\underline{y}}(f)] + Var_{\underline{y}}(E_{\underline{x}|\underline{y}}[f]). \quad (2.1)$$

In particular,

$$Var_{\underline{x}}(x) = E_{\underline{y}}[Var_{\underline{x}|\underline{y}}(x)] + Var_{\underline{y}}(E_{\underline{x}|\underline{y}}[x]). \quad (2.2)$$

proof:

Let

$$A = \sum_y P(y) \left(\sum_x P(x|y) f \right)^2. \quad (2.3)$$

Then

$$Var_{\underline{x}, \underline{y}}(f) = \sum_{x,y} P(x, y) f^2 - \left(\sum_{x,y} P(x, y) f \right)^2 \quad (2.4)$$

$$= \begin{cases} \sum_{x,y} P(x, y) f^2 - A \\ + \left(A - \left(\sum_{x,y} P(x, y) f \right)^2 \right) \end{cases} \quad (2.5)$$

$$E_{\underline{y}}[Var_{\underline{x}|\underline{y}}(f)] = \sum_y P(y) \left(\sum_x P(x|y) f^2 - \left(\sum_x P(x|y) f \right)^2 \right) \quad (2.6)$$

$$= \sum_{x,y} P(x, y) f^2 - A \quad (2.7)$$

$$Var_{\underline{y}}(E_{\underline{x}|\underline{y}}[f]) = \sum_y P(y) \left(\sum_x P(x|y)f \right)^2 - \left(\sum_y P(y) \sum_x P(x|y)f \right)^2 \quad (2.8)$$

$$= A - \left(\sum_{x,y} P(x,y)f \right)^2 \quad (2.9)$$

QED

2.2 Sum of Squares Estimates

Consider a population Σ partitioned into groups Σ_g such that $\Sigma = \cup_{g=1}^{ng} \Sigma_g$, where the Σ_g are mutually disjoint.

Let

dof stand for “degrees of freedom”

SS stand for “Sum of Squares”

MS stand for “Mean Square”

$x_{\Sigma_g} = \{x_{\sigma|g}\}_{\sigma \in \Sigma_g}$

Define the total and group mean values by

$$\bar{x} = \frac{1}{|\Sigma|} \sum_{g=1}^{ng} \sum_{\sigma \in \Sigma_g} x_{\sigma|g} \quad (2.10)$$

$$\bar{x}_g = \frac{1}{|\Sigma_g|} \sum_{\sigma \in \Sigma_g} x_{\sigma|g} \quad (2.11)$$

Define the **SS Total** by

$$SS_T = \sum_{g=1}^{ng} \sum_{\sigma \in \Sigma_g} (x_{\sigma|g} - \bar{x})^2 \quad (2.12)$$

$$= |\Sigma| Var_{\underline{x}}(x) \quad (2.13)$$

with *dof* and *MS* given by

$$dof_T = |\Sigma| - 1, \quad MS_T = \frac{SS_T}{dof_T} \quad (2.14)$$

Define the **SS Within** by

$$SS_W = \sum_{g=1}^{ng} \sum_{\sigma \in \Sigma_g} (x_{\sigma|g} - \bar{x}_g)^2 \quad (2.15)$$

$$= |\Sigma| \sum_{g=1}^{ng} \frac{|\Sigma_g|}{|\Sigma|} \left\{ \frac{1}{|\Sigma_g|} \sum_{\sigma \in \Sigma_g} (x_{\sigma|g} - \bar{x}_g)^2 \right\} \quad (2.16)$$

$$= |\Sigma| E_{\underline{g}}[Var_{\underline{x}|g}(x)] \quad (\text{note this is a mean value}) \quad (2.17)$$

with dof and MS given by

$$dof_W = |\Sigma| - ng, \quad MS_W = \frac{SS_W}{dof_W} \quad (2.18)$$

Define the **SS Between** by

$$SS_B = \sum_{g=1}^{ng} |\Sigma_g| (\bar{x}_g - \bar{x})^2 \quad (2.19)$$

$$= |\Sigma| \sum_{g=1}^{ng} \frac{|\Sigma_g|}{|\Sigma|} (\underbrace{\bar{x}_g}_{E_{\underline{x}|g}[x]} - \bar{x})^2 \quad (2.20)$$

$$= |\Sigma| Var_{\underline{g}}(E_{\underline{x}|g}[x]) \quad (\text{note this is a variance}) \quad (2.21)$$

with dof and MS given by

$$dof_B = ng - 1, \quad MS_B = \frac{SS_B}{dof_B} \quad (2.22)$$

Claim 20

$$SS_T = SS_W + SS_B \quad (2.23)$$

$$dof_T = dof_W + dof_B \quad (2.24)$$

proof: By the just proven Law of Total Variance,

$$\underbrace{Var_{\underline{x}}(x)}_{SS_T/|\Sigma|} = \underbrace{E_{\underline{g}}[Var_{\underline{x}|g}(x)]}_{SS_W/|\Sigma|} + \underbrace{Var_{\underline{g}}(E_{\underline{x}|g}[x])}_{SS_B/|\Sigma|}. \quad (2.25)$$

QED

Fig.2.1 shows a bnet for calculating SS_T , SS_W and SS_B , The TPMs, printed in blue, for the bnet Fig.2.1, are as follows.

$$P(\bar{x} | \{x_{\Sigma_g}\}_{g=1}^{ng}) = \mathbb{1} \left(\bar{x} = \frac{1}{|\Sigma|} \sum_{g=1}^{ng} \sum_{\sigma \in \Sigma_g} x_{\sigma|g} \right) \quad (2.26)$$

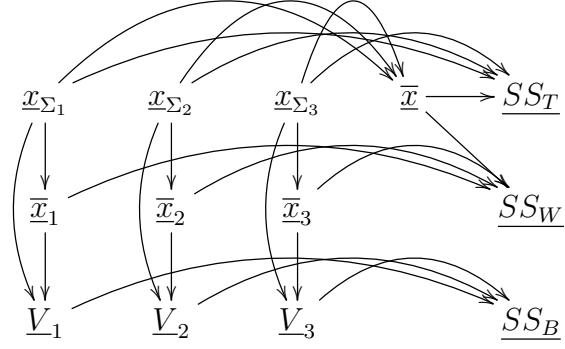


Figure 2.1: Bnet for calculating SS_T , SS_W and SS_B .

$$P(\bar{x}_g | x_{\Sigma_g}) = \mathbb{1} \left(\bar{x}_g = \frac{1}{|\Sigma_g|} \sum_{\sigma \in \Sigma_g} x_{\sigma|g} \right) \quad (2.27)$$

$$P(V_g | x_{\Sigma_g}, \bar{x}_g) = \mathbb{1} \left(V_g = \frac{1}{|\Sigma_g|} \sum_{\sigma \in \Sigma_g} (x_{\sigma|g} - \bar{x}_g)^2 \right) \quad (2.28)$$

$$P(SS_T | \{x_{\Sigma_g}\}_{g=1}^{ng}, \bar{x}) = \mathbb{1}(SS_T = SS_T(\{x_{\Sigma_g}\}_{g=1}^{ng}, \bar{x})) \quad (2.29)$$

$$P(SS_B | \{\bar{x}_g\}_{g=1}^{ng}, \bar{x}) = \mathbb{1}(SS_B = SS_B(\{\bar{x}_g\}_{g=1}^{ng}, \bar{x})) \quad (2.30)$$

$$P(SS_W | \{V_g\}_{g=1}^{ng}) = \mathbb{1}(SS_W = SS_W(\{V_g\}_{g=1}^{ng})) \quad (2.31)$$

2.3 F-statistic and hypothesis testing

The **F-statistic** for ANOVA is defined by

$$F = \frac{MS_B}{MS_W} \quad \left(= \frac{\text{variance}}{\text{mean value}} \right) \quad (2.32)$$

(Note that F is the ratio of two chi-square distributions)

Consider $ng = 3$ for definiteness. Let

h_0 = hypothesis that $\mu_1 = \mu_2 = \mu_3$ (null hypothesis)

h_1 = opposite of h_0 (alternative, opposite hypothesis)

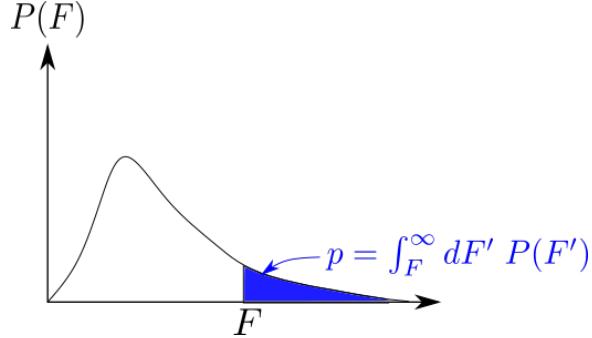


Figure 2.2: Probability distribution $P(F)$ for the F-statistic, at fixed dof_B and dof_W . See Wikipedia article Ref.[98] for more info about $P(F)$.

It's not hard to find on the Internet, tables and software calculators that give the score p-value¹

$$p_{sc} = \int_F^\infty dF' P(F'; dof_B, dof_W) \quad (2.33)$$

of the F statistic, for a given dof_B and dof_W . See Fig.2.2 for a portrait of $P(F)$ and p_{sc} . In Bayesian language, p_{sc} measures the probability that h_0 is true. Assume $p_{th} = 0.05$.² To determine if the difference between group means is “statistically significant”, we compare p_{sc} with p_{th} . If $p_{sc} < p_{th} = 0.05$, we reject the null hypothesis. Otherwise, we fail to reject (accept) the null hypothesis.

If h_0 is rejected, we can perform **post-hoc** tests to determine how much the groups differ from each other. There are several popular post-hoc tests, such as the Tukey test, Bonferroni test, and Scheffe test.

¹Some tables give $F = F(p_{sc}, dof_B, dof_W)$ instead of $p_{sc} = p_{sc}(F, dof_B, dof_W)$, but note that there is a 1-1/onto invertible map between F and p_{sc} . Some tables refer to p_{sc} as α .

²The difference between threshold (th) and score (sc) p-values is discussed in Section C.40.

Chapter 3

ARACNE structure learning

This chapter is based on Ref.[31].

The ARACNE algo is an algo for learning the structure of a bnet from data. The algo considers data samples for n random variables $(x_i)_{i=0,1,\dots,n-1}$, and estimates the mutual information $MI_{i,j} = H(\underline{x}_i : \underline{x}_j)$ between every pair of nodes. The set UG is initialized to contain all the edges of a fully connected undirected graph. Next the algo removes from UG every edge with $MI_{i,j} < \epsilon$ for some threshold $0 < \epsilon \ll 1$. Then the algo examines every triplet of edges in UG , and marks for removal the edge of the triplet with the smallest MI. Finally, the algo removes from UG all edges marked for removal. Each triplet is analyzed irrespective of whether its edges have been marked for removal when considering a prior triplet. Thus the network constructed by the algorithm is independent of the order in which the triplets are examined. Some of the unoriented edges in UG can be given an orientation using the same techniques used to orient edges in constraint based structure learning (see Chapter 78).

Ref.[31] incorrectly claims that removing the smaller of 3 MI's is "an application" of the Data Processing Inequality (DPI) of Shannon Information Theory. See Chapter 47 for more info about DPI. Note that DPI is only valid for a Markov chain, and not all triplets of random variables are in a Markov chain. Removing the smaller of 3 numbers does not require DPI.

Fig.3.1 gives an example of the application of the ARACNE algo.

See Chapter 10 on Chow-Liu trees (CLT). A CLT is just a maximum spanning tree where the weights are mutual informations $MI_{i,j}$ estimated from data.

Sometimes, the outcome of the ARACNE algo is a CLT. For example, Fig.3.1 (a) was considered in Chapter 10 on CLTs, and the CLT algo also gave Fig.3.1 (c) as the final structure.

According to Ref.[31], the ARACNE algo sometimes yields a polytree (i.e., a connected graph with no loops). It may even yield a structure with loops. Hence, it does not always yield a CLT.

By breaking all cliques (i.e., fully connected subgraphs) with 3 edges and 3 nodes, ARACNE breaks all cliques with 3 or more nodes. However, cliques are not uncommon in Nature, especially 3 node cliques. Cliques become less likely in Nature

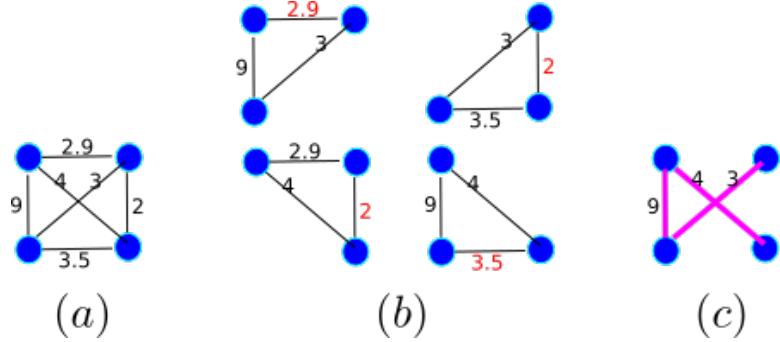


Figure 3.1: An example where the ARACNE algo gives a Chow-Liu tree. (a) Fully connected undirected graph with weights $MI_{i,j}$ along the edges. (b) All 4 possible triplets of edges with nonzero weights. Edges marked for removal have their weights printed in red.(c) Final structure.

the bigger the number of nodes they have *after* 3. Therefore, a nice generalization of ARACNE would be to list all 4 node cliques, and break each of them by eliminating their edge with the smallest MI. This will have the effect of breaking all cliques with 4 or more nodes but keeping 3 node cliques. One could also break some, not all, of the 3 node cliques, by consecutively removing the clique-breaking-edge with the smallest MI of all edges of all 3 node cliques. Let β stand for banned clique number of nodes. Then the current ARACNE has $\beta = 3$. We are suggesting that a β of 4 might be more likely to occur in Nature.

Chapter 4

Backdoor Adjustment Formula

The backdoor (BD) adjustment formula is proven in Chapter 19 from the rules of Do Calculus. The goal of this chapter is to give examples of the use of that theorem. We will restate the theorem in this chapter, sans proof. There is no need to understand the theorem's proof in order to use it. However, you will need to skim Chapter 19 in order to familiarize yourself with the notation used to state the theorem. This chapter also assumes that you are comfortable with the rules for checking for d-separation. Those rules are covered in Chapter 21.

Suppose that we have access to data that allows us to estimate a probability distribution $P(x., y., z.)$. Hence, the variables $\underline{x}.$, $\underline{y}.$, $\underline{z}.$ are ALL observed (i.e, not hidden). Then we say that the backdoor $\underline{z}.$ satisfies the **backdoor adjustment criterion** relative to $(\underline{x}., \underline{y}.)$ if

1. All backdoor paths from $\underline{x}.$ to $\underline{y}.$ are blocked by conditioning on $\underline{z}..$
2. $\underline{z}.. \cap de(\underline{x}.) = \emptyset$.

Claim 21 Backdoor Adjustment Formula

If $\underline{z}.$ satisfies the backdoor criterion relative to $(\underline{x}., \underline{y}.)$, then

$$P(y.|D\underline{x}. = x.) = \sum_{z.} P(y.|x., z.)P(z.) \quad (4.1)$$

$$= \sum_{z.} \begin{array}{c} z. \\ \searrow \\ x. \longrightarrow y. \end{array} \quad (4.2)$$

where $\sum z.$ means node $\underline{z}.$ is summed over.

proof: See Chapter 19.

QED

4.1 Examples

1.



BD criterion satisfied if $\underline{x}_. = \underline{x}$, $\underline{y}_. = \underline{y}$, $\underline{z}_. = \emptyset$. No adjustment necessary.

$$P(y|\mathcal{D}\underline{x} = x) = P(y|x) \quad (4.4)$$

2.



BD criterion satisfied if $\underline{x}_. = \underline{x}$, $\underline{y}_. = \underline{y}$, $\underline{z}_. = \emptyset$. No adjustment necessary.

$$P(y|\mathcal{D}\underline{x} = x) = P(y|x) \quad (4.6)$$

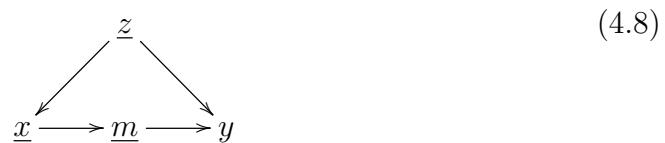
3.



BD criterion satisfied if $\underline{x}_. = \underline{x}$, $\underline{y}_. = \underline{y}$, $\underline{z}_. = \underline{z}$.

Note that here the backdoor formula adjusts the parents of $x..$

4.



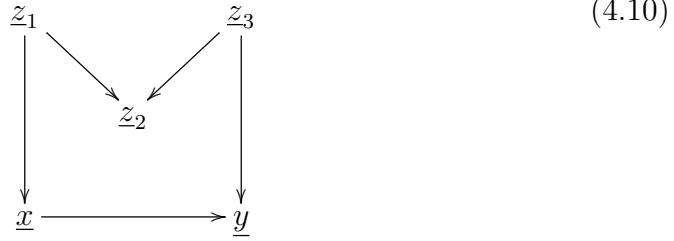
BD criterion satisfied if $\underline{x}_. = \underline{x}$, $\underline{y}_. = \underline{y}$, $\underline{z}_. = \underline{z}$.

5.



BD criterion satisfied if $\underline{x}_. = \underline{x}, \underline{y}_. = \underline{y}, \underline{z}_. = (\underline{z}_1, \underline{z}_2)$.

6.



BD criterion satisfied if $\underline{x}_. = \underline{x}, \underline{y}_. = \underline{y}$ and $\underline{z}_.$ = one of the following.

- \emptyset
 - \underline{z}_3
 - $\underline{z}_2, \underline{z}_3$
 - $\underline{z}_1, \underline{z}_2, \underline{z}_3$
 - \underline{z}_1
 - $\underline{z}_1, \underline{z}_2$
 - $\underline{z}_1, \underline{z}_3$
-

7.



BD criterion is impossible to satisfy if $\underline{x}_. = \underline{x}, \underline{y}_. = \underline{y}$. However, the frontdoor criterion can be satisfied. See Chapter 28.

8.



BD criterion satisfied if $\underline{x}_. = \underline{x}, \underline{y}_. = \underline{y}, \underline{z}_. = \underline{z}$. We are able to block the backdoor path by conditioning on \underline{z} .

9.



Conditioning on \underline{z} blocks backdoor path $\underline{x}-\underline{z}-\underline{y}$, but opens path $\underline{x}-\underline{e}-\underline{z}-\underline{a}-\underline{y}$ because \underline{z} is a collider for that path. That path is blocked if we also condition on \underline{a} , which is possible because \underline{a} is observed. In conclusion, the BD criterion is satisfied if $\underline{x}_. = \underline{x}, \underline{y}_. = \underline{y}$ and $\underline{z}_. = (\underline{z}, \underline{a})$.

Conditioning on the parents of $\underline{x}.$ is often enough to block all backdoor paths. However, sometimes some of the parents are unobserved and one must condition on other nodes that are not parents of $\underline{x}.$ in order to satisfy the BD criterion.

10.



No need to control anything because only possible backdoor path is blocked by not conditioning on collider $\underline{w}.$ Hence,

$$P(y|\mathcal{D}\underline{x} = x) = P(y|x) . \quad (4.15)$$

However, if for some reason we want to control $\underline{t},$ we can do so. We can't control \underline{w} though, because $\underline{w} \in de(\underline{x}).$ Thus, the BD criterion is satisfied if $\underline{x} = \underline{x}, \underline{y} = \underline{y}$ and $\underline{z} = \underline{t}.$ Therefore,

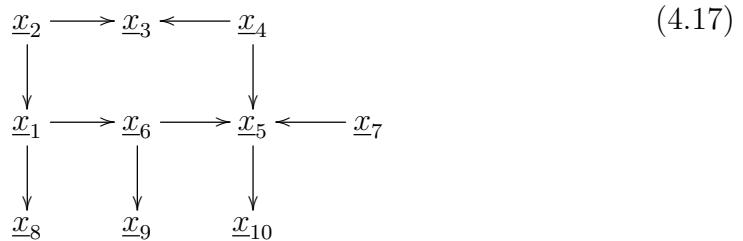
$$P(y|\mathcal{D}\underline{x} = x) = \sum_t P(y|x, t)P(t) . \quad (4.16)$$

11. Discuss what to do if several sets $\underline{z}.$ satisfy the BD criterion.

- Can evaluate $P(y|\mathcal{D}\underline{x} = x.)$ multiple ways and compare the results. This is a test that the causal bnet is correct.
 - It might be easier or less expensive to get data for some $\underline{z}.$ more than for others.
-

12. (Taken from online course notes Ref.[18])

Consider the bnet



If $\underline{x} = \underline{x}_1$ and $\underline{y} = \underline{x}_5,$ find all possible adjustment multinode $\underline{z}.$ that satisfy the BD criterion. Ans:

- \emptyset
- \underline{x}_4
- $\underline{x}_2, \underline{x}_3$
- $\underline{x}_2, \underline{x}_3, \underline{x}_4$
- \underline{x}_2
- $\underline{x}_2, \underline{x}_4$
- $\underline{x}_3, \underline{x}_4$

Add \underline{x}_7 to each of the previous 7 possible $\underline{z}..$. This gives a total of 14 possible adjustment multinodes $\underline{z}..$.

Chapter 5

Back Propagation (Automatic Differentiation)

5.1 General Theory

5.1.1 Jacobians

Suppose $f : \mathbb{R}^{nx} \rightarrow \mathbb{R}^{nf}$ and

$$y = f(x) . \quad (5.1)$$

Then the Jacobian $\frac{\partial y}{\partial x}$ is defined as the matrix with entries¹

$$\left[\frac{\partial y}{\partial x} \right]_{i,j} = \frac{\partial y_i}{\partial x_j} . \quad (5.2)$$

Jacobian of function composition. Suppose $f : \mathbb{R}^{nx} \rightarrow \mathbb{R}^{nf}$, $g : \mathbb{R}^{nf} \rightarrow \mathbb{R}^{ng}$. If

$$y = g \circ f(x) , \quad (5.3)$$

then

$$\frac{\partial y}{\partial x} = \frac{\partial g}{\partial f} \frac{\partial f}{\partial x} . \quad (5.4)$$

Right hand side of last equation is a product of two matrices so order of matrices is important.

$$\underline{f}^4 \longleftarrow \underline{f}^3 \longleftarrow \underline{f}^2 \longleftarrow \underline{f}^1 \longleftarrow \underline{f}^0$$

(a) Composition

$$\frac{\partial f^4}{\partial x} \longleftarrow \frac{\partial f^3}{\partial x} \longleftarrow \frac{\partial f^2}{\partial x} \longleftarrow \frac{\partial f^1}{\partial x} \longleftarrow 1$$

(b) Forward-p

$$1 \longrightarrow \frac{\partial y}{\partial f^3} \longrightarrow \frac{\partial y}{\partial f^2} \longrightarrow \frac{\partial y}{\partial f^1} \longrightarrow \frac{\partial y}{\partial f^0}$$

(c) Back-p

Figure 5.1: bnets for function composition, forward propagation and back propagation for $nf = 5$ nodes.

5.1.2 Bnets for function composition, forward propagation and back propagation

Let

$$y = f^4 \circ f^3 \circ f^2 \circ f^1(x) . \quad (5.5)$$

This function composition chain can be represented by the bnet Fig.5.1(a) with TPMs

$$P(f^\mu | f^{\mu-1}) = \mathbb{1}(f^\mu = f^\mu(f^{\mu-1})) \quad (5.6)$$

for $\mu = 1, 2, 3, 4$.

¹Mnemonic for remembering order of indices: i in numerator/ j in denominator becomes index i/j of Jacobian matrix.

Note that

$$\frac{\partial y}{\partial x} = \frac{\partial y}{\partial f^3} \frac{\partial f^3}{\partial f^2} \left[\frac{\partial f^2}{\partial f^1} \frac{\partial f^1}{\partial x} \right] \quad (5.7)$$

$$= \frac{\partial y}{\partial f^3} \left[\frac{\partial f^3}{\partial f^2} \frac{\partial f^2}{\partial x} \right] \quad (5.8)$$

$$= \left[\frac{\partial y}{\partial f^3} \frac{\partial f^3}{\partial x} \right] \quad (5.9)$$

$$= \frac{\partial y}{\partial x}. \quad (5.10)$$

This forward propagation can be represented by the bnet Fig.5.1(b) with node TPMs

$$P\left(\frac{\partial f^{\mu+1}}{\partial x} \mid \frac{\partial f^\mu}{\partial x}\right) = \mathbb{1}\left(\frac{\partial f^{\mu+1}}{\partial x} = \frac{\partial f^{\mu+1}}{\partial f^\mu} \frac{\partial f^\mu}{\partial x}\right) \quad (5.11)$$

for $\mu = 1, 2, 3$.

Note that

$$\frac{\partial y}{\partial x} = \left[\frac{\partial y}{\partial f^3} \frac{\partial f^3}{\partial f^2} \right] \frac{\partial f^2}{\partial f^1} \frac{\partial f^1}{\partial x} \quad (5.12)$$

$$= \left[\frac{\partial y}{\partial f^2} \frac{\partial f^2}{\partial f^1} \right] \frac{\partial f^1}{\partial x} \quad (5.13)$$

$$= \left[\frac{\partial y}{\partial f^1} \frac{\partial f^1}{\partial x} \right] \quad (5.14)$$

$$= \frac{\partial y}{\partial x}. \quad (5.15)$$

This back propagation can be represented by the bnet Fig.5.1(c) with node TPMs

$$P\left(\frac{\partial y}{\partial f^\mu} \mid \frac{\partial y}{\partial f^{\mu+1}}\right) = \mathbb{1}\left(\frac{\partial y}{\partial f^\mu} = \frac{\partial y}{\partial f^{\mu+1}} \frac{\partial f^{\mu+1}}{\partial f^\mu}\right) \quad (5.16)$$

for $\mu = 2, 1, 0$.

$\frac{\partial f^{\mu+1}}{\partial f^\mu}$ is a Jacobian matrix so the order of multiplication matters. In forward prop, it pre-multiplies, and in back prop it post-multiplies.

5.2 Application to Neural Networks

5.2.1 Absorbing b_i^λ into $w_{i|j}$.

The TPMs, printed in blue, for a NN bnet, as given in Chapter 58, are as follows.

For all hidden layers $\lambda = 0, 1, \dots, \Lambda - 2$,

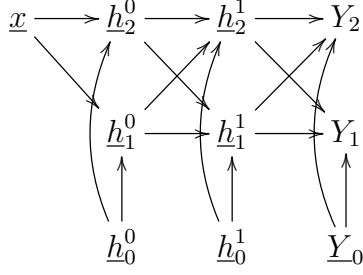


Figure 5.2: Nodes $\underline{h}_0^0, \underline{h}_0^1, \underline{Y}_0$ are all set to 1. They allow us to absorb b_i^λ into the first column of $w_{i|j}^\lambda$.

$$P(h_i^\lambda | h_{\cdot}^{\lambda-1}) = \delta \left(h_i^\lambda, \mathcal{A}_i^\lambda \left(\sum_j w_{i|j}^\lambda h_j^{\lambda-1} + b_i^\lambda \right) \right) \quad (5.17)$$

for $i = 0, 1, \dots, nh(\lambda) - 1$. For the output visible layer $\lambda = \Lambda - 1$:

$$P(Y_i | h_{\cdot}^{\Lambda-2}) = \delta \left(Y_i, \mathcal{A}_i^{\Lambda-1} \left(\sum_j w_{i|j}^{\Lambda-1} h_j^{\Lambda-2} + b_i^{\Lambda-1} \right) \right) \quad (5.18)$$

for $i = 0, 1, \dots, ny - 1$.

For each λ , replace the matrix $w_{\cdot|}^\lambda$ by the augmented matrix $[b_{\cdot|}^\lambda, w_{\cdot|}^\lambda]$ so that the new $w_{\cdot|}^\lambda$ satisfies

$$w_{i|0}^\lambda = b_i^\lambda \quad (5.19)$$

Let the nodes \underline{h}_0^λ for all λ and \underline{Y}_0 be root nodes (so no arrows pointing into them). For each λ , draw arrows from \underline{h}_0^λ to all other nodes in that same layer. Draw arrows from \underline{Y}_0 to all other nodes in that same layer.

After performing the above steps, the TPMs, printed in blue, for the NN bnet, are as follows:

For all hidden layers $\lambda = 0, 1, \dots, \Lambda - 2$,

$$P(h_0^\lambda) = \delta(h_0^\lambda, 1) , \quad (5.20)$$

and

$$P(h_i^\lambda | h_{\cdot}^{\lambda-1}, h_0^\lambda = 1) = \delta \left(h_i^\lambda, \mathcal{A}_i^\lambda \left(\sum_j w_{i|j}^\lambda h_j^{\lambda-1} \right) \right) \quad (5.21)$$

for $i = 1, \dots, nh(\lambda) - 1$. For the output visible layer $\lambda = \Lambda - 1$:

$$P(Y_0) = \delta(Y_0, 1) , \quad (5.22)$$

and

$$P(Y_i \mid h_{\cdot}^{\Lambda-2}, Y_0 = 1) = \delta \left(Y_i, \mathcal{A}_i^{\Lambda-1} \left(\sum_j w_{i|j}^{\Lambda-1} h_j^{\Lambda-2} \right) \right) \quad (5.23)$$

for $i = 1, 2, \dots, ny - 1$.

5.2.2 Bnets for function composition, forward propagation and back propagation for NN

$$\underline{\mathcal{A}}^3 \longleftarrow \underline{\mathcal{B}}^3 \longleftarrow \underline{\mathcal{A}}^2 \longleftarrow \underline{\mathcal{B}}^2 \longleftarrow \underline{\mathcal{A}}^1 \longleftarrow \underline{\mathcal{B}}^1 \longleftarrow \underline{\mathcal{A}}^0 \longleftarrow \underline{\mathcal{B}}^0 \longleftarrow \underline{x}$$

(a)

$$\underline{\frac{\partial \mathcal{A}^3}{\partial x}} \longleftarrow \underline{\frac{\partial \mathcal{B}^3}{\partial x}} \longleftarrow \underline{\frac{\partial \mathcal{A}^2}{\partial x}} \longleftarrow \underline{\frac{\partial \mathcal{B}^2}{\partial x}} \longleftarrow \underline{\frac{\partial \mathcal{A}^1}{\partial x}} \longleftarrow \underline{\frac{\partial \mathcal{B}^1}{\partial x}} \longleftarrow \underline{\frac{\partial \mathcal{A}^0}{\partial x}} \longleftarrow \underline{\frac{\partial \mathcal{B}^0}{\partial x}} \longleftarrow \underline{1}$$

(b)

$$\underline{1} \longrightarrow \underline{\frac{\partial Y}{\partial \mathcal{B}^3}} \longrightarrow \underline{\frac{\partial Y}{\partial \mathcal{A}^2}} \longrightarrow \underline{\frac{\partial Y}{\partial \mathcal{B}^2}} \longrightarrow \underline{\frac{\partial Y}{\partial \mathcal{A}^1}} \longrightarrow \underline{\frac{\partial Y}{\partial \mathcal{B}^1}} \longrightarrow \underline{\frac{\partial Y}{\partial \mathcal{A}^0}} \longrightarrow \underline{\frac{\partial Y}{\partial \mathcal{B}^0}} \longrightarrow \underline{\frac{\partial Y}{\partial x}}$$

(c)

Figure 5.3: bnets for (a) function composition, (b) forward propagation and (c) back propagation for a neural net with 4 layers (3 hidden and output visible).

From here on, we will rename y above by $Y = \hat{y}$ and consider samples $y[i]$ for $i = 0, 1, \dots, nsam - 1$. The Error (a.k.a. loss or cost function) is

$$\mathcal{E} = \frac{1}{nsam} \sum_{\sigma=0}^{nsam-1} \sum_{i=0}^{ny-1} |Y_i - y_i[\sigma]|^2 \quad (5.24)$$

To perform simple gradient descent, one uses:

$$(w_{i|j}^{\lambda})' = w_{i|j}^{\lambda} - \eta \frac{\partial \mathcal{E}}{\partial w_{i|j}^{\lambda}} . \quad (5.25)$$

One has

$$\frac{\partial \mathcal{E}}{\partial w_{i|j}^\lambda} = \frac{1}{nsam} \sum_{\sigma=0}^{nsam-1} \sum_{i=0}^{ny-1} 2(Y_i - y_i[\sigma]) \frac{\partial Y}{\partial w_{i|j}^\lambda}. \quad (5.26)$$

Define \mathcal{B}_i^λ thus

$$\mathcal{B}_i^\lambda(h^{\lambda-1}) = \sum_j w_{i|j}^\lambda h_j^{\lambda-1}. \quad (5.27)$$

Then

$$\frac{\partial Y}{\partial w_{i|j}^\lambda} = \frac{\partial Y}{\partial \mathcal{B}_i^\lambda} \frac{\partial \mathcal{B}_i^\lambda}{\partial w_{i|j}^\lambda} \quad (5.28)$$

$$= \frac{\partial Y}{\partial \mathcal{B}_i^\lambda} h_j^{\lambda-1} \quad (5.29)$$

$$\frac{\partial \mathcal{E}}{\partial w_{i|j}^\lambda} = \frac{\partial \mathcal{E}}{\partial \mathcal{B}_j^\lambda} \frac{\partial \mathcal{B}_j^\lambda}{\partial w_{i|j}^\lambda} \quad (5.30)$$

$$= \frac{\partial \mathcal{E}}{\partial \mathcal{B}_j^\lambda} h_j^{\lambda-1}. \quad (5.31)$$

This suggest that we can calculate the derivatives of the error \mathcal{E} with respect to the weights $w_{i|j}^\lambda$ in two stages, using an intermediate quantity δ_j^λ :

$$\begin{cases} \delta_j^\lambda = \frac{\partial \mathcal{E}}{\partial \mathcal{B}_j^\lambda} \\ \frac{\partial \mathcal{E}}{\partial w_{i|j}^\lambda} = \delta_j^\lambda h_j^{\lambda-1} \end{cases} \quad (5.32)$$

To apply what we learned in the earlier General Theory section of this chapter, consider a NN with 4 layers (3 hidden, and the output visible one). Define the functions f_i as follows:

$$f_i^0 = x_i \quad (5.33)$$

$$\text{Layer 0: } f_i^1 = \mathcal{B}_i^0(x_i), \quad f_i^2 = \mathcal{A}_i^0(\mathcal{B}_i^0) \quad (5.34)$$

$$\text{Layer 1: } f_i^3 = \mathcal{B}_i^1(\mathcal{A}_i^0), \quad f_i^4 = \mathcal{A}_i^1(\mathcal{B}_i^1) \quad (5.35)$$

$$\text{Layer 2: } f_i^5 = \mathcal{B}_i^2(\mathcal{A}_i^1), \quad f_i^6 = \mathcal{A}_i^2(\mathcal{B}_i^2) \quad (5.36)$$

$$\text{Layer 3: } f_i^7 = \mathcal{B}_i^3(\mathcal{A}_i^2), \quad f_i^8 = \mathcal{A}_i^3(\mathcal{B}_i^3) \quad (5.37)$$

See Fig.5.3. The TPMs, printed in blue, for the bnet (c) for back propagation, are as follows:

$$P\left(\frac{\partial Y}{\partial \mathcal{B}^\lambda} \mid \frac{\partial Y}{\partial \mathcal{B}^{\lambda+1}}\right) = \mathbb{1}\left(\frac{\partial Y}{\partial \mathcal{B}^\lambda} = \frac{\partial Y}{\partial \mathcal{B}^{\lambda+1}} \frac{\partial \mathcal{B}^{\lambda+1}}{\partial \mathcal{A}^\lambda} \frac{\partial \mathcal{A}^\lambda}{\partial \mathcal{B}^\lambda}\right). \quad (5.38)$$

One has

$$\frac{\partial \mathcal{A}_i^\lambda}{\partial \mathcal{B}_j^\lambda} = D\mathcal{A}_i^\lambda(\mathcal{B}_i^\lambda)\delta(i,j) \quad (5.39)$$

where $D\mathcal{A}_i^\lambda(z)$ is the derivative of $\mathcal{A}_i^\lambda(z)$.

From Eq.(5.27)

$$\mathcal{B}_i^{\lambda+1}(\mathcal{A}^\lambda) = \sum_j w_{i|j}^{\lambda+1} \mathcal{A}_j^\lambda \quad (5.40)$$

so

$$\frac{\partial \mathcal{B}_i^{\lambda+1}}{\partial \mathcal{A}_j^\lambda} = w_{i|j}^{\lambda+1}. \quad (5.41)$$

Therefore, Eq.(5.38) implies

$$P\left(\frac{\partial Y}{\partial \mathcal{B}_j^\lambda} \mid \frac{\partial Y}{\partial \mathcal{B}_j^{\lambda+1}}\right) = \mathbb{1}\left(\frac{\partial Y}{\partial \mathcal{B}_j^\lambda} = \sum_i \frac{\partial Y}{\partial \mathcal{B}_i^{\lambda+1}} D\mathcal{A}_j^\lambda(\mathcal{B}_j^\lambda) w_{i|j}^{\lambda+1}\right), \quad (5.42)$$

$$P\left(\frac{\partial \mathcal{E}}{\partial \mathcal{B}_j^\lambda} \mid \frac{\partial \mathcal{E}}{\partial \mathcal{B}_j^{\lambda+1}}\right) = \mathbb{1}\left(\frac{\partial \mathcal{E}}{\partial \mathcal{B}_j^\lambda} = \sum_i \frac{\partial \mathcal{E}}{\partial \mathcal{B}_i^{\lambda+1}} D\mathcal{A}_j^\lambda(\mathcal{B}_j^\lambda) w_{i|j}^{\lambda+1}\right), \quad (5.43)$$

$$P(\delta_j^\lambda \mid \delta_j^{\lambda+1}) = \mathbb{1}\left(\delta_j^\lambda = \sum_i \delta_i^{\lambda+1} D\mathcal{A}_j^\lambda(\mathcal{B}_j^\lambda) w_{i|j}^{\lambda+1}\right). \quad (5.44)$$

First delta of iteration, belonging to output layer $\lambda = \Lambda - 1$:

$$\delta_j^{\Lambda-1} = \frac{\partial \mathcal{E}}{\partial \mathcal{B}_j^{\Lambda-1}} \quad (5.45)$$

$$= \frac{1}{nsam} \sum_{\sigma=0}^{nsam-1} \sum_{i=0}^{ny-1} 2(Y_i - y_i[\sigma]) D\mathcal{A}_i^{\Lambda-1}(\mathcal{B}_i^{\Lambda-1}) \delta(i,j) \quad (5.46)$$

$$= \frac{1}{nsam} \sum_{\sigma=0}^{nsam-1} 2(Y_j - y_j[\sigma]) D\mathcal{A}_j^{\Lambda-1}(\mathcal{B}_j^{\Lambda-1}) \quad (5.47)$$

Cute expression for derivative of sigmoid function:

$$Dsmoid(x) = smoid(x)(1 - smoid(x)) \quad (5.48)$$

5.3 General bnets instead of Markov chains induced by layered structure of NNs

$$P(\delta_{\underline{x}} \mid (\delta_{\underline{a}})_{\underline{a} \in ch(\underline{x})}) = \mathbb{1}(\delta_{\underline{x}} = \sum_{\underline{a} \in ch(\underline{x})} \delta_{\underline{a}} D \mathcal{A}_{\underline{x}}(\mathcal{B}_{\underline{x}})) w_{\underline{a}|\underline{x}}) \quad (5.49)$$

Reverse arrows of original bnet and define the TPM of nodes of “time reversed” bnet by

$$P(\delta_{\underline{x}} \mid (\delta_{\underline{a}})_{\underline{a} \in pa(\underline{x})}) = \mathbb{1}(\delta_{\underline{x}} = \sum_{\underline{a} \in pa(\underline{x})} \delta_{\underline{a}} D \mathcal{A}_{\underline{x}}(\mathcal{B}_{\underline{x}})) w_{\underline{x}|\underline{a}}^T) \quad (5.50)$$

Chapter 6

Basic Curve Fitting Using Gradient Descent

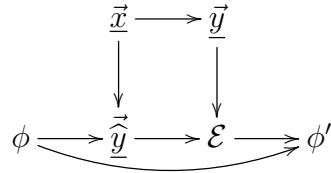


Figure 6.1: Basic curve fitting bnet.

Samples $(x^\sigma, y^\sigma) \in S_{\underline{x}} \times S_{\underline{y}}$ are given. $nsam(\vec{x}) = nsam(\vec{y})$. Estimator function $\hat{y}(x; \phi)$ for $x \in S_{\underline{x}}$ and $\phi \in \mathbb{R}$ is given.

Let

$$P_{\underline{x}, \underline{y}}(x, y) = \frac{1}{nsam(\vec{x})} \sum_{\sigma} \mathbb{1}(x = x^\sigma, y = y^\sigma). \quad (6.1)$$

Let

$$\mathcal{E}(\vec{x}, \vec{y}, \phi) = \frac{1}{nsam(\vec{y})} \sum_{\sigma} |y^\sigma - \hat{y}(x^\sigma; \phi)|^2 \quad (6.2)$$

\mathcal{E} is called the mean square error.

Best fit is parameters ϕ^* such that

$$\phi^* = \operatorname{argmin}_{\phi} \mathcal{E}(\vec{x}, \vec{y}, \phi). \quad (6.3)$$

The TPMs, printed in blue, for the basic curve fitting bnet Fig.6.1, are as follows.

$$P(\phi) = \text{given}. \quad (6.4)$$

The first time it is used, ϕ is arbitrary. After the first time, it is determined by previous stage.

$$P(\vec{x}) = \prod_{\sigma} P_{\underline{x}}(x^{\sigma}) \quad (6.5)$$

$$P(\vec{y}|\vec{x}) = \prod_{\sigma} P_{\underline{y}|\underline{x}}(y^{\sigma} | x^{\sigma}) \quad (6.6)$$

$$P(\hat{y}^{\sigma}|\phi, \vec{x}) = \delta(\hat{y}^{\sigma}, \hat{y}(x^{\sigma}; \phi)) \quad (6.7)$$

$$P(\mathcal{E}|\vec{y}, \vec{y}) = \delta(\mathcal{E}, \frac{1}{nsam(\vec{x})} \sum_{\sigma} |y^{\sigma} - \hat{y}^{\sigma}|^2) . \quad (6.8)$$

$$P(\phi'|\phi, \mathcal{E}) = \delta(\phi', \phi - \eta \partial_{\phi} \mathcal{E}) \quad (6.9)$$

$\eta > 0$ is the descent rate. If $\Delta\phi = \phi' - \phi = -\eta \frac{\partial \mathcal{E}}{\partial \phi}$, then $\Delta\mathcal{E} = \frac{-1}{\eta} (\Delta\phi)^2 < 0$ so this will minimize the error \mathcal{E} . This is called “gradient descent”.

Chapter 7

Bell and Clauser-Horne Inequalities in Quantum Mechanics

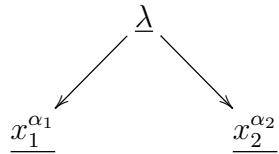


Figure 7.1: bnet used to discuss Bell and Clauser-Horne inequalities in Quantum Mechanics.

I wrote an article about this in 2008 for my blog “Quantum Bayesian Networks”. See Ref.[65].

Chapter 8

Berkson's Paradox

For more information about Berkson's Paradox (BP), see Ref.[80]

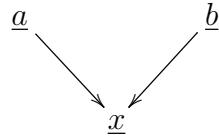


Figure 8.1: Bnet used to discuss Berkson's Paradox (BP). \underline{a} and \underline{b} are both causes of collider \underline{x} .

Consider the bnet of Fig.8.1. For that bnet, we have

$$P(a, b, x) = P(a)P(b)P(x|a, b) . \quad (8.1)$$

Summing Eq.(8.1) over x , we get

$$P(a, b) = P(a)P(b) \quad (8.2)$$

so \underline{a} and \underline{b} are independent. It follows that a can be ignored in calculating the probability of b ; i.e.,

$$\boxed{P(b|a) = P(b)} . \quad (8.3)$$

However, a cannot be ignored in calculating the probability of b , if x is being held fixed; i.e.,

$$\boxed{P(b|a, x) \neq P(b|x)} . \quad (8.4)$$

Indeed,

$$P(b|a, x) = \frac{P(b)P(x|a, b)}{\sum_b P(b)P(x|a, b)} \quad (8.5)$$

whereas

$$P(b|x) = \frac{\sum_a P(a)P(b)P(x|a,b)}{\sum_{a,b} P(a)P(b)P(x|a,b)} . \quad (8.6)$$

The two boxed equations are what is referred to as BP.

BP is also called **collider bias** because \underline{x} is a collider.

BP is also called **explaining away** in the special case that $\underline{a}, \underline{b}, \underline{x} \in \{\text{false} = 0, \text{true} = 1\}$. In that case, if \underline{x} is fixed to true, and the cause \underline{a} is known to be true, then the cause \underline{b} is less likely to be true. For example, suppose a car engine fails ($x = 1$) and the two most likely causes of the failure are alternator (\underline{a}) and battery (\underline{b}). Once we know that the alternator has failed ($\underline{a} = 1$), it is less likely that the battery is failing ($\underline{b} = 1$) than when the status of \underline{a} was not known; i.e., $P(\underline{b} = 1|x = 1, \underline{a} = 1) < P(\underline{b} = 1|x = 1)$.

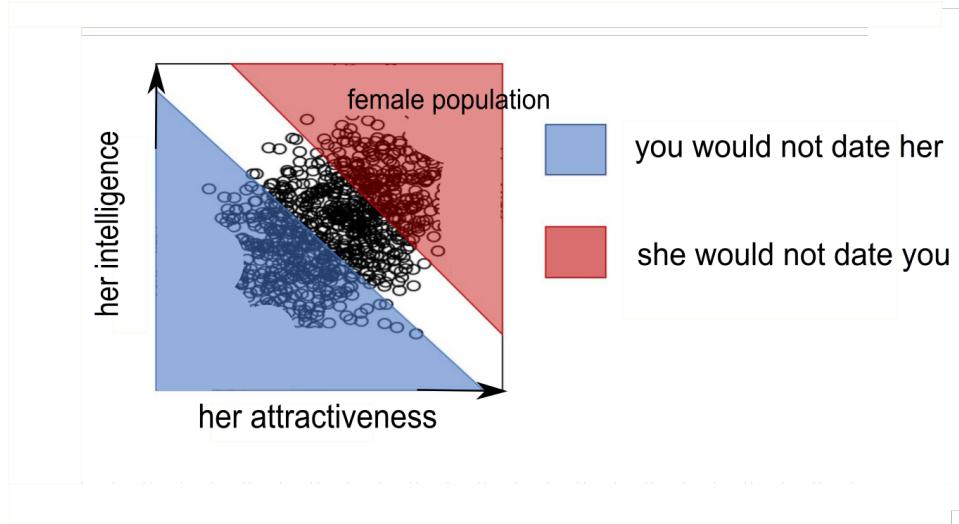


Figure 8.2: Example of Berkson's paradox (BP).

Fig.8.2 presents an example of BP. The figure consists of a scatter plot with axes a = her attractiveness, b = her intelligence, for a female population of possible dates for you, assuming you are a male person. Let $x \in \{\text{false} = 0, \text{true} = 1\} =$ she goes out on a date with you. For the full population,

$$(a, b) \sim P(a, b) = P(a)P(b) \quad (8.7)$$

whereas for the population in the white swath,

$$(a, b) \sim P(a, b|x) = P(b|a, x)P(a|x) \neq P(b|x)P(a|x) . \quad (8.8)$$

As shown by Fig.8.2, BP is an example of **selection bias**. Selection bias happens when a non-representative subset of the total population is considered (i.e., selected).

Chapter 9

Binary Decision Diagrams

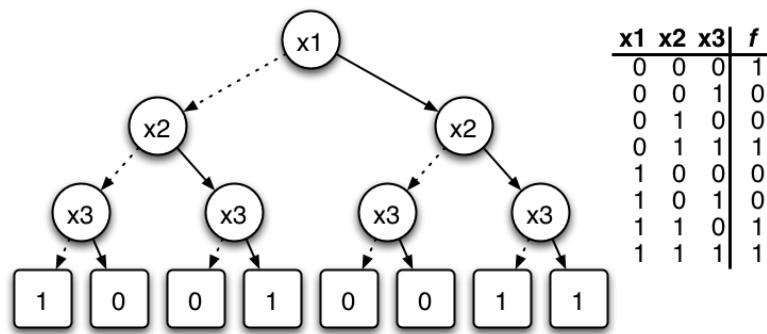


Figure 9.1: Binary decision tree and truth table for the function $f(x_1, x_2, x_3) = \bar{x}_1(x_2 + \bar{x}_3) + x_1x_2$

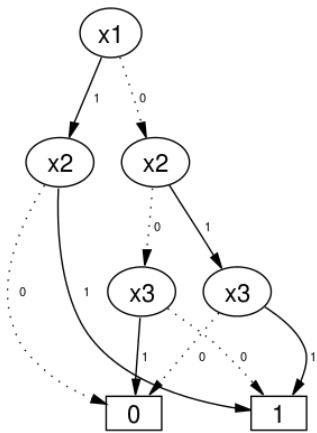


Figure 9.2: BDD for the function f of Fig.9.1.

This chapter is based on Wikipedia article Ref.[84].

Binary Decision Diagrams (BDDs) can be understood as a special case of Decision Trees (dtrees). We will assume that the reader has read Chapter 14 on dtrees before reading this chapter.

Both Figs.9.1 and 9.2 were taken from the aforementioned Wikipedia article. They give a simple example of a function $f : \{0, 1\}^3 \rightarrow \{0, 1\}$ represented in Fig.9.1 as a **binary decision tree** and in Fig.9.2 as a **binary decision diagram (BDD)**. It is possible to find, for each of those two figures, a bnet with the same graph structure. We show how to do this next.

We begin by noting that the function $f : \{0, 1\}^3 \rightarrow \{0, 1\}$ is a special case of a probability distribution $P : \{0, 1\}^3 \rightarrow [0, 1]$. In fact, if we restrict P to be deterministic, then $P_{det} : \{0, 1\}^3 \rightarrow \{0, 1\}$ has the same domain and range as f . Henceforth, we will refer to $f(x_1, x_2, x_3)$ as $P(x_1, x_2, x_3)$, keeping in mind that we are restricting our attention to deterministic probability distributions.

If we apply the chain rule for conditional probabilities to $P(x_1, x_2, x_3)$, we get

$$P(x_1, x_2, x_3) = P(x_3|x_1, x_2)P(x_2|x_1)P(x_1), \quad (9.1)$$

which can be represented by the bnet:

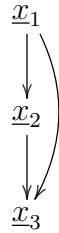


Figure 9.3: Most general 3 node bnet.

But in Chapter 14, we learned how to represent the dtree of Fig.9.1 as the image bnet Fig.9.4. The TPMs, printed in blue, for the image bnet Fig.9.4, are as follows. Note that the TPMs for Fig.9.4 can be constructed from the TPMs for the bnet Fig.9.3. If $x_1, x_2, x_3, x_4 \in \{0, 1, null\}$ and $a, b, c \in \{0, 1\}$, then

$$P(\underline{x}_1 = x_1) = \begin{cases} P_{x_1}(x_1) & \text{if } x_1 \in \{0, 1\} \\ 0 & \text{if } x_1 = \text{null} \end{cases} \quad (9.2)$$

$$P(\underline{x}_2|a = x_2 | \underline{x}_1 = x_1) = \begin{cases} P_{x_2|x_1}(x_2|a) & \text{if } x_1 = a \\ \mathbb{1}(x_2 = \text{null}) & \text{otherwise} \end{cases} \quad (9.3)$$

$$P(\underline{x}_3|a, b = x_3 | \underline{x}_2|a = x_2) = \begin{cases} P_{x_3|x_1, x_2}(x_3|a, b) & \text{if } x_2 = b \\ \mathbb{1}(x_3 = \text{null}) & \text{otherwise} \end{cases} \quad (9.4)$$

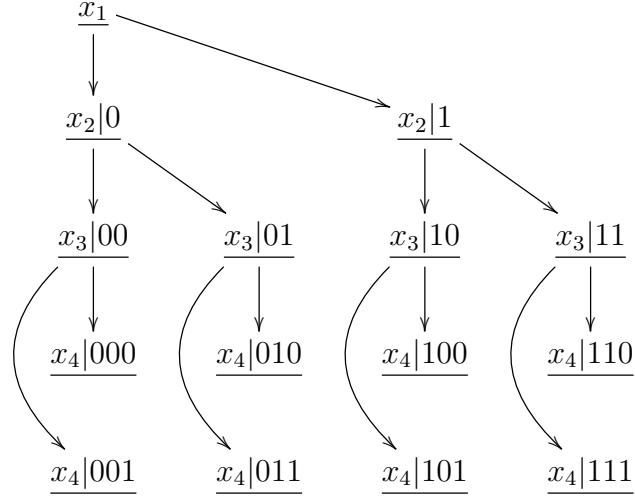


Figure 9.4: Image bnet for binary dtree of Fig.9.1.

$$P(\underline{x}_4|a, b, c = x_4 | \underline{x}_3|a, b = x_3) = \begin{cases} \delta(x_4, c) & \text{if } x_3 = c \\ \mathbb{1}(x_4 = \text{null}) & \text{otherwise} \end{cases} \quad (9.5)$$

Note that if $P_{\underline{x}_3|x_1, x_2} = P_{\underline{x}_3|x_2}$ in Eq.(9.4), then the bnet Fig.9.3 reduces to a Markov chain $\underline{x}_1 \rightarrow \underline{x}_2 \rightarrow \underline{x}_3$.

The BDD shown in Fig.9.2 emphasizes the fact that

$$P(x_1, x_2, x_3|x_1 = 1) = P(x_2|x_1 = 1) = x_2 . \quad (9.6)$$

The BDD of Fig.9.2 has as image bnet Fig.9.5. Define

$$pa(\underline{0}) = pa(\underline{1}) = (x_2|1, x_3|00, x_3|01) . \quad (9.7)$$

Let $pa(\underline{0}) = abc$ mean the same as $pa(\underline{0}) = (a, b, c)$. The TPMs of the image bnet Fig.9.5 are the same as those for the image bnet Fig.9.4 except for the TPMs of the nodes $\underline{0}$ and $\underline{1}$. For those two nodes, the TPMs, printed in blue, are as follows.

$$P(\underline{0} = x|pa(\underline{0})) = \begin{cases} \delta(x, 0) & \text{if } pa(\underline{0}) = 011 \\ \delta(x, \text{null}) & \text{otherwise} \end{cases} \quad (9.8)$$

$$P(\underline{1} = x|pa(\underline{1})) = \begin{cases} \delta(x, 1) & \text{if } pa(\underline{1}) = 101 \\ \delta(x, \text{null}) & \text{otherwise} \end{cases} \quad (9.9)$$

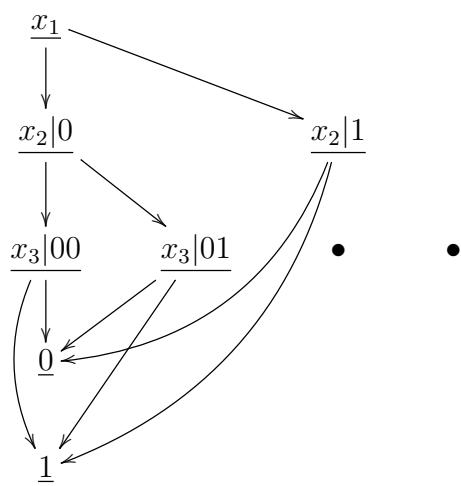


Figure 9.5: Image bnet for BDD of Fig.9.2.

Chapter 10

Chow-Liu Trees and Tree Augmented Naive Bayes (TAN)

This chapter is mostly based on chapter 8 of Pearl's 1988 book Ref.[44]. See also Ref.[89] and references therein.

This chapter uses various Shannon Information Theory entropies. Our notation for these entropies is described in Chapter C.

10.1 Chow-Liu Trees

Chow-Liu trees refers to an algorithm for finding a bnet tree that fits an a priori given probability distribution as closely as possible.

Consider a bnet with n nodes $\underline{x}^n = (\underline{x}_0, \underline{x}_1, \dots, \underline{x}_{n-1})$ such that $\underline{x}_i \in S_{\underline{x}_i}$ for all i . Let its total probability distribution be $P_{\underline{x}^n}$. For simplicity, we will abbreviate $P_{\underline{x}^n}$ by P . Hence

$$P(x^n) = P_{\underline{x}^n}(x^n). \quad (10.1)$$

Suppose we want to fit $P_{\underline{x}^n}$ by a tree bnet with nodes $\underline{t}^n = (\underline{t}_0, \underline{t}_1, \dots, \underline{t}_{n-1})$ such that $\underline{t}_i \in S_{\underline{t}_i} = S_{\underline{x}_i}$ for all i . For simplicity, we will abbreviate $P_{\underline{t}^n}$ by P_T . Hence

$$P_T(x^n) = P_{\underline{t}^n}(x^n). \quad (10.2)$$

Throughout this chapter, let $V = \{0, 1, \dots, n-1\}$, the set of vertices. Suppose μ is a function $\mu : V \rightarrow V$ such that $\mu(i) < i$. Let $T_\mu = \{\underline{t}_{\mu(i)} \rightarrow \underline{t}_i : i \in V - \{0\}\}$. Then T_μ is a tree that spans (i.e., it includes all nodes) \underline{t}^n . Its root node is \underline{t}_0 , because \underline{t}_0 has no parents. All other nodes \underline{t}_i have exactly one parent, namely $\underline{t}_{\mu(i)}$. Let P_T , the total probability distribution for the tree, be parameterized by the function μ as follows:

$$P_T(x^n) = \prod_{i=0}^{n-1} P_T(x_i | x_{\mu(i)}), \quad (10.3)$$

where, for the root node 0, $P_T(x_0|x_{\mu(0)}) = P_T(x_0)$.

Claim 22 $D_{KL}(P \parallel P_T)$ is minimized over all probability distributions P_T that are expressible as Eq.(10.3) iff

$$P_T(x_i|x_{\mu(i)}) = P(x_i|x_{\mu(i)}) \quad (10.4)$$

for all i , and

$$\sum_i H(\underline{x}_i : \underline{x}_{\mu(i)}) \quad (10.5)$$

is maximized over all μ .

proof:

$$D_{KL}(P \parallel P_T) = \sum_{x^n} P(x^n) \ln \frac{P(x^n)}{P_T(x^n)} \quad (10.6)$$

$$= - \sum_{x^n} \sum_i P(x^n) \ln P_T(x_i|x_{\mu(i)}) + \sum_i P(x^n) \ln P(x^n) \quad (10.7)$$

$$= - \sum_i \sum_{x_i, x_{\mu(i)}} P(x_i, x_{\mu(i)}) \ln P_T(x_i|x_{\mu(i)}) - H(\underline{x}^n) \quad (10.8)$$

$$= - \sum_i \sum_{x_{\mu(i)}} P(x_{\mu(i)}) \left[\sum_{x_i} P(x_i|x_{\mu(i)}) \ln P_T(x_i|x_{\mu(i)}) \right] - H(\underline{x}^n) . \quad (10.9)$$

Now note that

$$\sum_{x_i} P(x_i|x_{\mu(i)}) \ln \frac{P(x_i|x_{\mu(i)})}{P_T(x_i|x_{\mu(i)})} \geq 0 \quad (10.10)$$

and this inequality becomes an equality iff

$$P(x_i|x_{\mu(i)}) = P_T(x_i|x_{\mu(i)}) . \quad (10.11)$$

Therefore

$$D_{KL}(P \parallel P_T) \geq - \underbrace{\sum_i \sum_{x_{\mu(i)}} P(x_{\mu(i)}) \left[\sum_{x_i} P(x_i|x_{\mu(i)}) \ln P(x_i|x_{\mu(i)}) \right]}_{=H(\underline{x}_i|x_{\mu(i)})=H(\underline{x}_i:\underline{x}_{\mu(i)})-H(\underline{x}_i)} - H(\underline{x}^n) , \quad (10.12)$$

and this inequality becomes an equality iff Eq.(10.11) is satisfied.

Note from the last equation that

$$\operatorname{argmin}_{\mu} D_{KL}(P \parallel P_T) = \operatorname{argmax}_{\mu} \sum_i H(\underline{x}_i : \underline{x}_{\mu(i)}) . \quad (10.13)$$

QED

Claim 23

$$\operatorname{argmin}_{\mu} H(\underline{x}^n) = \operatorname{argmax}_{\mu} \sum_i H(\underline{x}_i : \underline{x}_{\mu(i)}) \quad (10.14)$$

proof:

$$H(\underline{x}^n) = - \sum_{x^n} P(x^n) \sum_i \ln P(x_i | x_{\mu(i)}) \quad (10.15)$$

$$= - \sum_i \sum_{x_i, x_{\mu(i)}} P(x_i, x_{\mu(i)}) \ln P(x_i | x_{\mu(i)}) \quad (10.16)$$

$$= - \sum_i \sum_{x_i, x_{\mu(i)}} P(x_i, x_{\mu(i)}) \left[\ln \frac{P(x_i | x_{\mu(i)})}{P(x_i)} + \ln P(x_i) \right] \quad (10.17)$$

$$= - \sum_i [H(\underline{x}_i : \underline{x}_{\mu(i)}) - H(\underline{x}_i)] \quad (10.18)$$

$$= \sum_i H(\underline{x}_i) - \sum_i H(\underline{x}_i : \underline{x}_{\mu(i)}) \quad (10.19)$$

QED

The meaning of Claims 22 and 23 is as follows. If $D_{KL}(P \parallel P_T)$ is minimized over all P_T , then

1. P_T inherits its TPMs from P , and
2. P_T gets its structure, which is being parameterized by the function μ , by maximizing the score given by

$$\text{score} = \sum_i H(\underline{x}_i : \underline{x}_{\mu(i)}) . \quad (10.20)$$

(mutual information $H(\underline{a} : \underline{b})$ measures correlation between \underline{a} and \underline{b}). Maximizing the score is the same as minimizing the entropy $H(\underline{x}^n)$ over all the structures μ . (i.e., finding least complex structure).

So far, we have studied the properties of those probability distributions P_T for a tree bnet that best approximates an a priori given probability distribution P , but we haven't yet described how to build a Chow-Liu tree based on empirical data. Next we give Chow-Liu's algorithm for doing so.

1. Find MST using Kruskal's algorithm¹. (see Fig.10.1)

Calculate weights $w_{i,j} = H(\underline{x}_i : \underline{x}_j)$ for all $i, j \in V$ and store them in a dictionary D that maps edges to weights.

Order D by weight size.

Let T be a list of the edges in the tree. Initialize T to empty.

Repeat this until T has $n - 1$ elements:

 Remove largest weight w from D and corresponding edge e .

 Add e to T if $\{e\} \cup T$ has no loops. Otherwise discard e and w .

2. Give directions to edges in T . (see Fig.10.2)

Let DT be a list of directed edges. Initialize DT to empty.

Choose any node as root node.

Point arrows along edges in T , away from root node.

Add new arrows to DT .

Repeat this until DT has $n - 1$ elements:

 Point arrows along edges in T , away from leaf nodes of current DT .

 Add new arrows to DT .

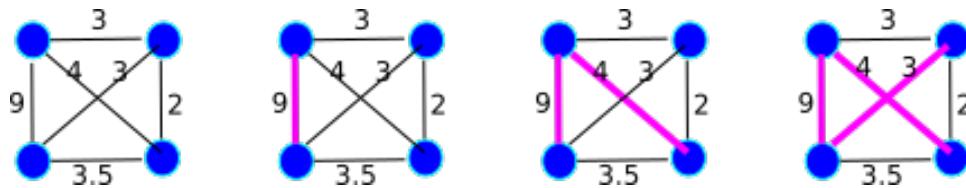


Figure 10.1: Example of finding MST (maximum spanning tree)

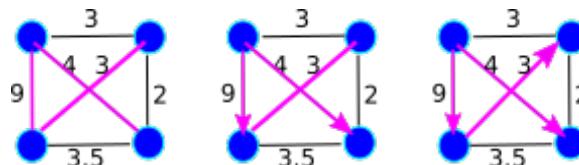


Figure 10.2: Example of giving directions to edges of spanning tree.

Nodes in a Chow-Liu tree can be rated in terms of their relative importance. Here are 2 possible metrics for measuring the importance of a node \underline{a} :

$$N_{nb}(\underline{a}) = \text{ number of neighbors of } \underline{a} \quad (10.21)$$

¹Kruskal's algorithm is one several famous algorithms (Prim's algo is another one) for finding an MST (maximum or minimum spanning tree). An MST algorithm takes an undirected graph with weights along its edges as input. It then finds a tree subgraph (i.e., subset of the edges of the graph with no loops) that (1) spans the graph (i.e., includes every vertex of the graph) and (2) maximizes (or minimizes) the sum of weights among all possible tree subgraphs. For more information, see Ref[123] and references therein, or any other of numerous explanations of MST in the Internet.

$$\text{traffic}(\underline{a}) = \sum_{n \in nb(\underline{a})} H(\underline{a} : \underline{n}) \quad (10.22)$$

For example, to get a tree with low depth, one can choose as the root node the node which has largest N_{nb} , and if there are several with the same largest N_{nb} , choose out of those the one with the largest traffic.

10.2 Tree Augmented Naive Bayes (TAN)

Recall from Chapter 57 that a Naive Bayes bnet consists of a class node \underline{c} with n children nodes \underline{x}^n , called the feature nodes. A Tree Augmented Naive Bayes (TAN) bnet is a Naive Bayes bnet with a tree grafted onto it like a chimera. More precisely, one starts with a Naive Bayes bnet and adds arrows between the feature nodes. The arrows are added in such a way that the TAN bnet sans node \underline{c} constitutes a tree. It's not the most well motivated bnet in human history, but at least it adds a bit of correlation between the feature nodes of the Naive Bayes bnet. Those nodes are independent at fixed \underline{c} in the Naive Bayes bnet, but are no longer so in the TAN bnet. See Figs.10.3 and 10.4 for an example of a TAN bnet.

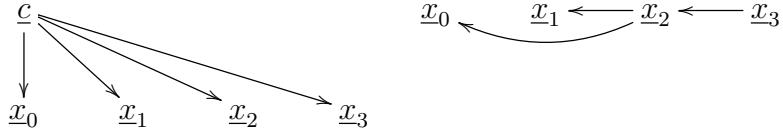


Figure 10.3: bnet for Naive Bayes with 4 feature nodes and another bnet for a tree made of the same feature nodes.

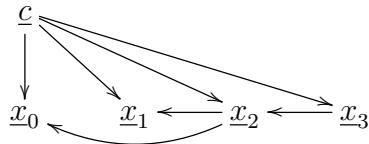


Figure 10.4: TAN bnet constructed by merging Naive Bayes bnet and tree bnet of Fig.10.3.

The total probability distribution P_{TAN} for a TAN bnet can be parameterized as follows.

$$P_{TAN}(x^n, c) = P_{TAN}(c) \prod_{i=0}^{n-1} P_{TAN}(x_i | x_{\mu(i)}, c). \quad (10.23)$$

As with Chow Liu trees, we can attempt to find a TAN bnet whose total probability $P_{TAN} = P_{\underline{x}^n, c}$ best approximates an a priori given probability distribution $P = P_{\underline{x}^n, c}$.

Note that

Claim 24

$$\operatorname{argmin}_{\mu} H(\underline{x}^n, c) = \operatorname{argmax}_{\mu} \sum_i H(\underline{x}_i : \underline{x}_{\mu(i)} | c) \quad (10.24)$$

proof:

$$H(\underline{x}^n, c) = - \sum_{x^n, c} P(x^n, c) \left[\ln P(c) + \sum_i \ln P(x_i | x_{\mu(i)}, c) \right] \quad (10.25)$$

$$= - \sum_{x^n, c} P(x^n, c) \left[\ln P(c) + \sum_i \ln \left(\frac{P(x_i, x_{\mu(i)} | c)}{P(x_i | c) P(x_{\mu(i)} | c)} P(x_i | c) \right) \right] \quad (10.26)$$

$$= \sum_i H(\underline{x}_i, c) - \sum_i H(\underline{x}_i : \underline{x}_{\mu(i)} | c) \quad (10.27)$$

QED

Following the same line of reasoning that we followed for Chow-Liu trees, we conclude that:

If $D_{KL}(P \| P_{TAN})$ is minimized over all P_{TAN} , then

1. P_{TAN} inherits its TPMs from P , and
2. P_{TAN} gets its structure, which is being parameterized by the function μ , by maximizing the score defined by

$$\text{score} = \sum_i H(\underline{x}_i : \underline{x}_{\mu(i)} | c) \quad (10.28)$$

One can build a TAN bnet from empirical data as follows:

Calculate a Chow-Liu Tree for each $c \in S_c$. For each of those trees, create a TAN bnet, and calculate its score given by Eq.(10.28). Keep the TAN bnet with the largest score.

Chapter 11

Counterfactual Reasoning

11.1 The 3 Rungs of Causal AI

According to Judea Pearl, there are 3 rungs in the ladder of causal AI. These are (as I see them):

1. **Observing Passively. Answering “What next?”:** Collecting data and fitting curves to it, without any plan designed to investigate Nature’s causal connections. Predicting the future.
2. **Doing causal experiments. Answering “Why?”:** Doing experiments consciously designed to elucidate Nature’s causal connections. Even cats do this!, but current AI doesn’t.
3. **Imagining counterfactual situations, Analogizing. Answering “What if?”:** Imagining gedanken experiments to further understand Nature’s causal connections, and to decide what future courses of action are more likely to succeed, even if those courses of action are unprecedented, and have never been taken before. Making predictions about events that have never happened (“counterfactuals”) is a very Bayesian concern, well out of the purview of frequentists. Nevertheless, humans do such “analogizing” all the time to great advantage. It becomes possible if there is some foreign but similar data that can be transported (transplanted, applied) to the situation of interest.

We will use the term **intervention operator (or simply “intervention”)** to refer to an operator that maps a bnet to another bnet. In Chapter 19, we introduced an intervention operator called the **do operator** $\mathcal{D}_{\underline{x}=x}$ (this is our notation for what Pearl symbolizes by $do(\underline{x}) = x$). The study of counterfactuals requires that we introduce a new kind of intervention operator that we will call an **imagine operator**, and denote by $\mathcal{I}_{\underline{x} \rightarrow \bar{y}}$. These 2 types of intervention operators will be defined in subsequent sections of this chapter. Usage of the do operator characterizes rung 2, and usage of the imagine operator characterizes rung 3.

Chapter 50 on message passing is about rung 1. Chapter 19 on Do Calculus is about rung 2. This chapter is dedicated to rung 3.

Judea Pearl is fond of discussing rung 3 solely in terms of SCM.¹ In this chapter, we define rung 3 without using SCM, using solely bnets. This gives a more general version of rung 3, because SCM are a subset of bnets.

11.2 Do operator

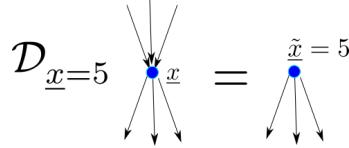


Figure 11.1: Action of “do” operator $\mathcal{D}_{\underline{x}=5}$ on node \underline{x} .

The do operator $\mathcal{D}_{\underline{x}=5}$ is defined graphically in Fig.11.1. The TPM, printed in blue, for node $\tilde{\underline{x}}$ of Fig.11.1, is as follows.²

$$P(\tilde{\underline{x}}; 5) = \delta(5, \tilde{\underline{x}}) \quad (11.1)$$

The do operator $\mathcal{D}_{\underline{x}=5}$ amputates the incoming arrows of node \underline{x} and sets the TPM of the new root node $\tilde{\underline{x}}$ to a delta function $\delta(\tilde{\underline{x}}, 5)$ (or some state of \underline{x} other than 5). Sometimes we call the new node $\mathcal{D}\underline{x}$ instead of $\tilde{\underline{x}}$.

The uses of the do operator are discussed in detail in Chapter 19.

11.3 Imagine operator

The imagine operator $\mathcal{I}_{\underline{x} \rightarrow y}(5)$ is defined graphically in Fig.11.2. Note that Fig.11.2 actually defines two types of imagine operators, the one with an argument: $\mathcal{I}_{\underline{x} \rightarrow y}(5)$, and the one without an argument: $\mathcal{I}_{\underline{x} \rightarrow y}$. The TPMs, printed in blue, for various nodes in Fig.11.2, are as follows.

- For $\mathcal{I}_{\underline{x} \rightarrow y}(\tilde{\underline{x}})G$

$$P(y|\tilde{\underline{x}}, a.) = P(y|\underline{x} = \tilde{\underline{x}}, a.) \quad (11.2)$$

¹SCM are what we call DEN. DEN (deterministic systems with external noise) are discussed in Chapter 44.

²The “;5” in the distribution indicates that 5 is a frequentist parameter of the distribution.

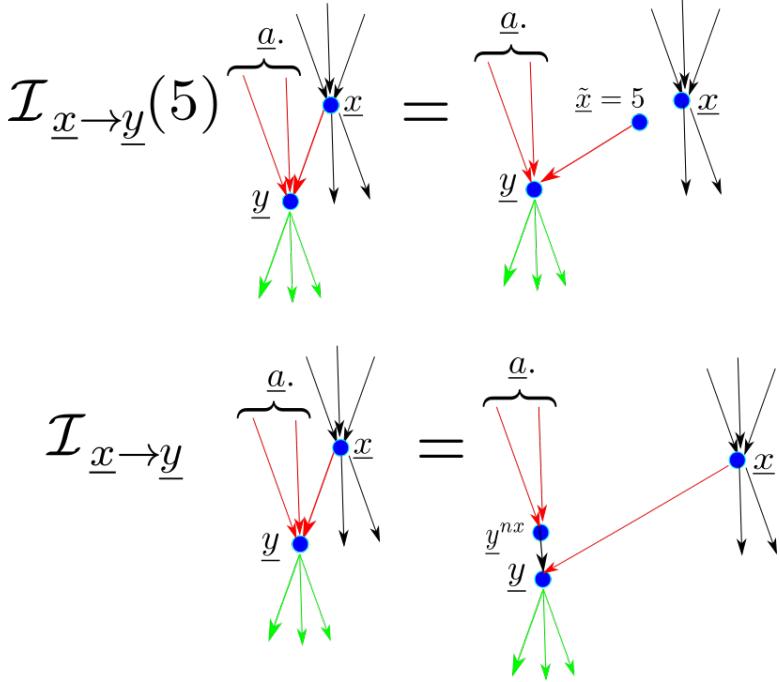


Figure 11.2: Action of “imagine” operators $\mathcal{I}_{\underline{x} \rightarrow \underline{y}}(5)$ and $\mathcal{I}_{\underline{x} \rightarrow \underline{y}}$ on arrow $\underline{x} \rightarrow \underline{y}$. In this figure, $y^{nx} = [y(x)]_{\forall x \in S_{\underline{x}}}$, where $nx = |S_{\underline{x}}|$ and $S_{\underline{x}}$ is the set of states of node \underline{x} .

$$P(\tilde{x}; 5) = \delta(\tilde{x}, 5) \quad (11.3)$$

- For $\mathcal{I}_{\underline{x} \rightarrow \underline{y}} G$

$$P(y^{nx}|a.) = \prod_{\tilde{x}} P(y(\tilde{x}) = y(\tilde{x})|a.) \quad (11.4)$$

$$P(y|y^{nx}, x) = \delta(y, y(x)) \quad (11.5)$$

The imagine operators $\mathcal{I}_{\underline{x} \rightarrow \underline{y}}(5)$ and $\mathcal{I}_{\underline{x} \rightarrow \underline{y}}$ operate on an arrow whereas the \mathcal{D} operator operates on a node. $\mathcal{I}_{\underline{x} \rightarrow \underline{y}}(5)$ deletes arrow $\underline{x} \rightarrow \underline{y}$ and creates a new root node \tilde{x} and a new arrow $\tilde{x} \rightarrow \underline{y}$. Sometimes we call the new node $\mathcal{I}_{\underline{y} \rightarrow \underline{x}}$ instead of \tilde{x} . $\mathcal{I}_{\underline{x} \rightarrow \underline{y}}$ creates a new node y^{nx} and an arrow $y^{nx} \rightarrow \underline{y}$.

Fig.11.3 shows how the imagine operator arises in Potential Outcomes (PO) theory. PO theory is discussed extensively in Chapter 64. As you can see, PO theory only uses a limited version of the 3 rungs of causal inference, because it doesn’t use the do-operator, and it only uses one of 2 possible types of imagine operators. Furthermore, it assumes a very limited triangular DAG.

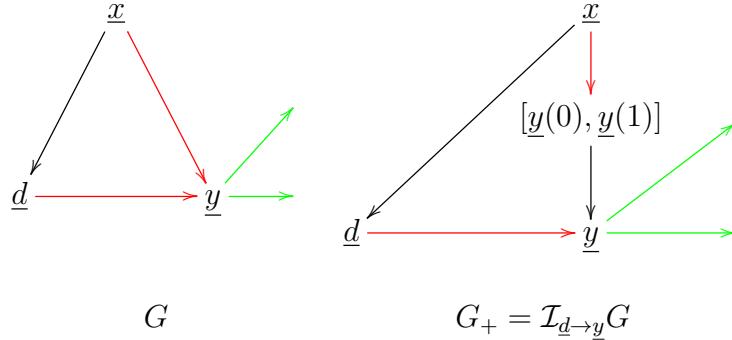


Figure 11.3: How imagine operator arises in Potential Outcomes (PO) theory.

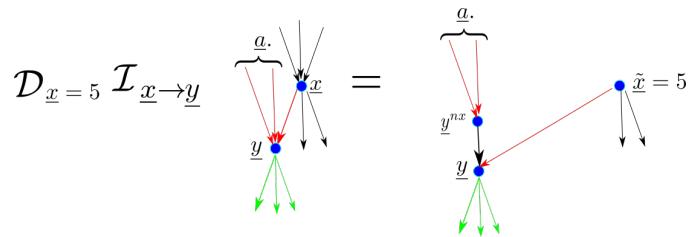


Figure 11.4: $\mathcal{D}_{\underline{x}=5} \mathcal{I}_{\underline{x} \rightarrow \underline{y}} G$ gives a connection between do and imagine operators.

Fig.11.4 gives a connection between do and imagine operators. We see from that figure that for $\mathcal{D}_{\underline{x}=\tilde{x}} \mathcal{I}_{\underline{x} \rightarrow \underline{y}} G$, we have³

$$P(y|\mathcal{D}\underline{x}=\tilde{x}, a.) = P(\underline{y}(\tilde{x})=y|a.) \quad (11.6)$$

One can define a **do-imagine-calculus** whose objective is to express probabilities such as $P(\underline{y}|\mathcal{D}\underline{r}=r, \mathcal{I}\underline{b}\underline{s}=s, t)$ in terms of observable probabilities that do not contain any do or imagine operators in them. As with Do Calculus, this reduction is not always possible, and we say a probability is **\mathcal{D} -identifiable**, **\mathcal{I} -identifiable** or **\mathcal{DI} -identifiable** if it can be expressed without do, imagine or both operators.

In causal inference, we often consider “counterfactual” random variables $\underline{y}(0) \in \{0, 1\}$ and $\underline{y}(1) \in \{0, 1\}$. They are called counterfactual variables because one of the 2 variables refers to an event that has occurred, whereas the other variable refers to a “counterfactual event”, i.e., an event that has never occurred. For some patients, $\underline{y}(0)$ has a value but $\underline{y}(1)$ doesn’t. For other patients, the opposite is the case. There is some disagreement in the community as to which algorithms perform rung 3 operations, and which don’t. This is the convention used in this book. We will say

³In the notation favored by Pearl, Eq.(11.6) would be

$$P(y|do(X)=\tilde{x}, a.) = P(Y_{\tilde{x}}=y|a.)$$

rung 3 operations are being performed if the counterfactual variables $\underline{y}(0)$ and $\underline{y}(1)$ are being used, or, equivalently, if a bnet is being used that includes nodes $\underline{y}(0)$, and $\underline{y}(1)$ that were produced by an imagine operator. Potential Outcomes (PO) theory (see Chapter 64) qualifies as rung 3 according to this convention. However, note that PO theory only does the bare minimum to reach rung 3. In PO theory, one usually evaluates $ATE = E[\underline{y}(1)] - E[\underline{y}(0)]$, which entails calculating $P(\underline{y}(0))$ and $P(\underline{y}(1))$. Pearl has extended the reach of rung 3 much further by calculating expected values that require knowledge of the joint distribution $P(\underline{y}(0), \underline{y}(1))$.

Chapter 12

Cross-Validation

This chapter is based on Ref.[93].

Cross-Validation (CV) is a method of calculating the “**out-of-training-set**” (**OOTS**) error for a classifier. What this means is that the classifier is trained on a training set, and its propensity to err is evaluated on a set different from the training set.

In **k -fold CV**, the most common CV method, and the only one we will discuss in this chapter, one partitions a dataset into k disjoint datasets of equal length. One uses $k - 1$ of those sub-datasets to train a model, and saves the last sub-dataset to validate the model just trained. One actually rotates which of the k sub-datasets is used for validation purposes, and calculates k validation errors \mathcal{E}_j for $j = 0, 1, \dots, k - 1$. Then one averages over the \mathcal{E}_j to obtain a final OOTS error \mathcal{E} .

CV strongly resembles Jackknife Resampling (JR) (see Chapter 39), but in JR the validation sub-dataset is never used for anything, whereas in CV, it is used for validation purposes, to calculate an OOTS error.

Next, we will explain k -fold CV more explicitly, using equations and a bnet.

Let $L = [0, 1, 2, \dots, nsam - 1]$ be a list of individuals (samples) in a population. In this chapter, we will use the notation $A^\sigma = A[\sigma]$ and $\vec{A} = [A^\sigma : \sigma \in L]$ for a list (vector, 1-D array) indexed by L . We will refer to $DS = (\vec{x}, \vec{y})$ where $x^\sigma \in S_x$, $y^\sigma \in S_y$, as a dataset. If L_j is a list (possibly with duplicate items) such that $set(L_j) \subset set(L)$, then define $DS_j = (\vec{x}, \vec{y})_{L_j} = ((x^\sigma)_{\sigma \in L_j}, (y^\sigma)_{\sigma \in L_j})$. We will refer to DS_j as the **restriction of (\vec{x}, \vec{y}) to L_j** .

Let $J = \{0, 1, 2, \dots, nj - 1\}$.

Define a **training list(TL), validation list(VL) pair** (TL, VL) to be a pair of lists such that $set(TL)$ and $set(VL)$ are disjoint subsets of $set(L)$. Let (TL_j, VL_j) for $j \in J$ be nj such TL-VL pairs.

Fig.12.1 shows the TL-VL pairs that are used when doing k -fold CV. In that figure, $k = nj = 4$. As you can see, in k -fold CV, one chooses $nj = k$ list pairs (TL_j, VL_j) such that all individuals $\sigma \in L$ appear exactly once, in either TL_j or VL_j , but not in both.

We will refer to a function $Y : S_x \rightarrow S_c$ as a classifier. It maps a vector of

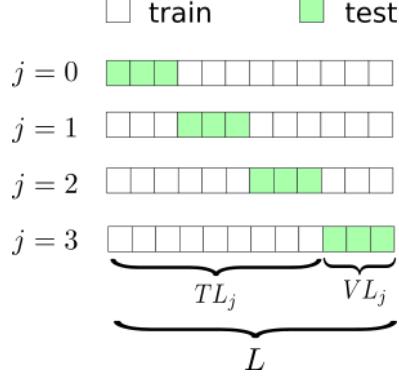


Figure 12.1: 4-fold CV with $|L| = 12$. For all j , $|VL_j| = 3$ and $|TL_j| = 9$. All individuals $\sigma \in L$ appear exactly once, in either TL_j or VL_j , but not in both.

features x to a class c . Let Y_j for $j \in J$ denote n_j classifiers.

If $Y_j : S_{\underline{x}} \rightarrow S_c$ and S_c is a discrete set (“categorical”), then define the **OOTs error for the j th classifier** as:

$$\mathcal{E}_j = \frac{1}{|VL_j|} \sum_{\sigma \in VL_j} \mathbb{1}(y^\sigma \neq Y_j(x^\sigma)) . \quad (12.1a)$$

On the other hand, if $S_c = \mathbb{R}$, it makes more sense to define \mathcal{E}_j as a mean square error:

$$\mathcal{E}_j = \frac{1}{|VL_j|} \sum_{\sigma \in VL_j} (y^\sigma - Y_j(x^\sigma))^2 . \quad (12.1b)$$

Finally, define the **final OOTs error** as

$$\mathcal{E} = \frac{1}{|J|} \sum_{j \in J} \mathcal{E}_j . \quad (12.2)$$

Fig.12.2 gives a bnet that represents the CV algorithm. The TPMs, printed in blue, for the bnet Fig.12.2, are as follows:

$$P((\vec{x}, \vec{y})_{TL_j}|(\vec{x}, \vec{y})) = \mathbb{1}(\text{ } (\vec{x}, \vec{y})_{TL_j} \text{ = restriction of } (\vec{x}, \vec{y}) \text{ to } TL_j.) \quad (12.3)$$

$$P((\vec{x}, \vec{y})_{VL_j}|(\vec{x}, \vec{y})) = \mathbb{1}(\text{ } (\vec{x}, \vec{y})_{VL_j} \text{ = restriction of } (\vec{x}, \vec{y}) \text{ to } VL_j.) \quad (12.4)$$

$$P(Y_j|(\vec{x}, \vec{y})_{TL_j}) = \mathbb{1}(\text{ } Y_j \text{ = classifier trained with } (\vec{x}, \vec{y})_{TL_j}) \quad (12.5)$$

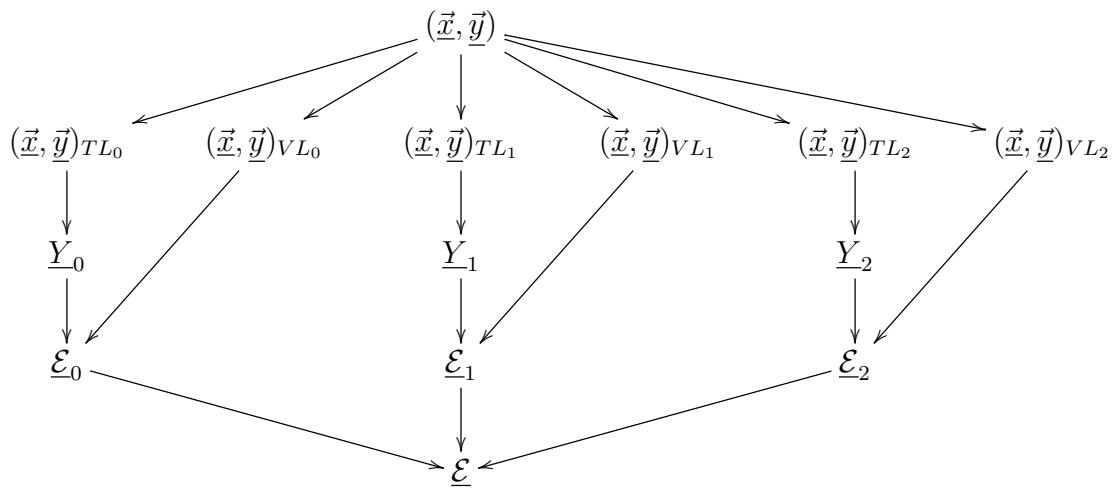


Figure 12.2: Bnet for 3-fold CV.

$$P(\mathcal{E}_j | Y_j, (\underline{\vec{x}}, \underline{\vec{y}})_{VL_j}) = \mathbb{1}(\mathcal{E}_j = \text{defined by Eqs.(12.1).}) \quad (12.6)$$

$$P(\mathcal{E} | (\mathcal{E}_j)_{j \in J}) = \mathbb{1}(\mathcal{E} = \text{defined by Eq.(12.2).}) \quad (12.7)$$

Chapter 13

Dataset Shift and Batch Normalization

In this chapter, we will represent Linear Regression (LR) as follows. We list a dataset; i.e., a set of tuples indexed by the individuals σ of a population Σ such that $|\Sigma| = nsam$. The independent variables of the LR (i.e., x^σ) are unboxed and the dependent variable (a.k.a. target feature) (i.e., y^σ) is shown inside a box. Then we show an arrow with the superscript “LR-fit”, followed by the fit function obtained by performing the LR.

$$\{(\sigma, x^\sigma = [x_i^\sigma], \boxed{y^\sigma}) : \sigma \in \Sigma\} \xrightarrow{\text{LR-fit}} \hat{y}(x) = \alpha + \sum_i x_i \beta_i \quad (13.1)$$

Analogously, we represent Supervised Machine Learning (ML) as follows.

$$\{(\sigma, x^\sigma, \boxed{y^\sigma}) : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \hat{y}(x) \quad (13.2)$$

When doing ML, we partition the full population Σ_{full} into two disjoint sets, the **training set** $\Sigma_{train} = \Sigma(\underline{s} = 0) = \Sigma$ and the **testing set** $\Sigma_{test} = \Sigma(\underline{s} = 1) = \Sigma^*$. Then we do two ML fits:

$$\begin{aligned} \text{training: } & \{(\sigma, x^\sigma, \boxed{y^\sigma}) : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \hat{y}(x) \\ \text{testing: } & \{(\sigma, x^\sigma, \boxed{y^\sigma}) : \sigma \in \Sigma^*\} \xrightarrow{\text{ML-fit}} \hat{y}^*(x) \end{aligned} \quad (13.3)$$

Ideally, $\hat{y}(x)$ and $\hat{y}^*(x)$, will be almost equal for all x . Dataset shift occurs when this is not the case. Equivalently, let

$$\begin{aligned} P_{train}(x, y) &= P(x, y | \underline{s} = 0) = P(x, y) \\ P_{test}(x, y) &= P(x, y | \underline{s} = 1) = P^*(x, y) \end{aligned} \quad (13.4)$$

We say there is a **dataset shift** if

$$P(x, y) \neq P^*(x, y) \quad (13.5)$$

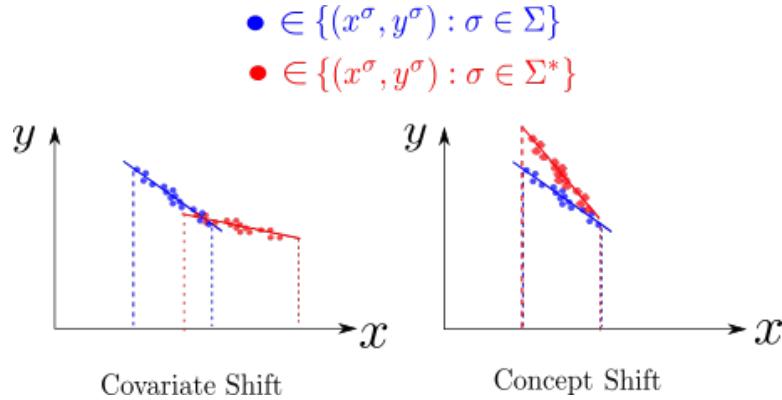


Figure 13.1: For Linear Regression, 2 types of dataset shift: covariate shift and concept shift.

13.1 Covariate Shift

We say there is a **covariate shift** if (see Fig.13.1)

$$P(y|x) = P^*(y|x) \text{ but } P(x) \neq P^*(x) \quad (13.6)$$

This can be represented in terms of bnets as follows¹

$$\begin{array}{ccc}
 \underline{s} = 0 & & \underline{s} = 1 \\
 \downarrow & & \downarrow \\
 \underline{x} \longrightarrow \underline{y} & \neq & \underline{x} \longrightarrow \underline{y} \\
 \underbrace{\hspace{2cm}}_{\underline{s} = 0} & & \\
 = & & \\
 \underline{x} \longrightarrow \underline{y} & &
 \end{array} \quad (13.7)$$

13.2 Concept Shift

We say there is a **concept shift** if (see Fig.13.1)

$$P(y|x) \neq P^*(y|x) \text{ but } P(x) = P^*(x) \quad (13.8)$$

This can be represented in terms of bnets as follows²

¹See See Chapter 74.

²See See Chapter 74.

$$\begin{array}{ccc}
\underline{s} = 0 & & \underline{s} = 1 \\
& \searrow & \searrow \\
& \underline{x} \longrightarrow \underline{y} & \neq \quad \underline{x} \longrightarrow \underline{y} \\
& \underbrace{\hspace{2cm}}_{\underline{s} = 0} & \\
& \uparrow & \\
& \underline{x} \longrightarrow \underline{y} &
\end{array} \tag{13.9}$$

13.3 Batch Normalization

Batch Normalization (BN) is a technique that is used to diminish dataset shift in Neural Nets.

Let $h_i^{\lambda, \sigma}$ be the output, for individual $\sigma \in \Sigma$, of the i th node of layer λ of a Neural Net (NN). Using the notation of Chapter 58,

$$h_i^{\lambda, \sigma} = \mathcal{A}_i^\lambda(z_i^{\lambda, \sigma}) \tag{13.10}$$

where

$$z_i^{\lambda, \sigma} = \sum_j w_{i|j}^\lambda h_j^{\lambda-1, \sigma} + b_i^\lambda \tag{13.11}$$

Activation functions $\mathcal{A}_i^\lambda : \mathbb{R} \rightarrow \mathbb{R}$ for NNs are discussed in Section 58.1. Suppose the population Σ is partitioned into disjoint batches $\Sigma^{(b)}$ for $b = 1, 2, \dots, B$. Let the set of points $\{z_i^{\lambda, \sigma} : \sigma \in \Sigma^{(b)}\}$ have mean $\mu_i^{\lambda(b)}$ and standard deviation $\sigma_i^{\lambda(b)}$. For any $z_i^{\lambda, \sigma}$ with $\sigma \in \Sigma$, define the BN activation function $\mathcal{A}_{i, BN}^\lambda(\cdot)$ by

$$\mathcal{A}_{i, BN}^\lambda(z_i^{\lambda, \sigma}) = \gamma_i^\lambda \left[\frac{z_i^{\lambda, \sigma} - \mu_i^{\lambda(b)}}{\sigma_i^{\lambda(b)}} \right] + \beta_i^\lambda \quad \text{if } \sigma \in \Sigma^{(b)}, \tag{13.12}$$

where the real valued parameters γ_i^λ and β_i^λ are learned during the optimization process. If node \underline{h}_i^λ of the NN has activation function $\mathcal{A}_i^\lambda : \mathbb{R} \rightarrow \mathbb{R}$, defined a new activation function $\mathcal{A}_{i, new}^\lambda : \mathbb{R} \rightarrow \mathbb{R}$ by the composition of functions

$$\mathcal{A}_{i, new}^\lambda = \mathcal{A}_i^\lambda \circ \mathcal{A}_{i, BN}^\lambda \tag{13.13}$$

Hence, the BN activation function is applied after the linear transformation Eq.(13.11), but before the nonlinear transformation \mathcal{A}_i^λ .

Intuition on why BN diminishes dataset shift: We discussed in Section 58.1 how nonlinear activation functions have a range that is smaller than their domain. Presumably, BN helps to make the range of the activation functions even more concentrated.

Chapter 14

Decision Trees

This chapter is based mainly on Ref.[60].

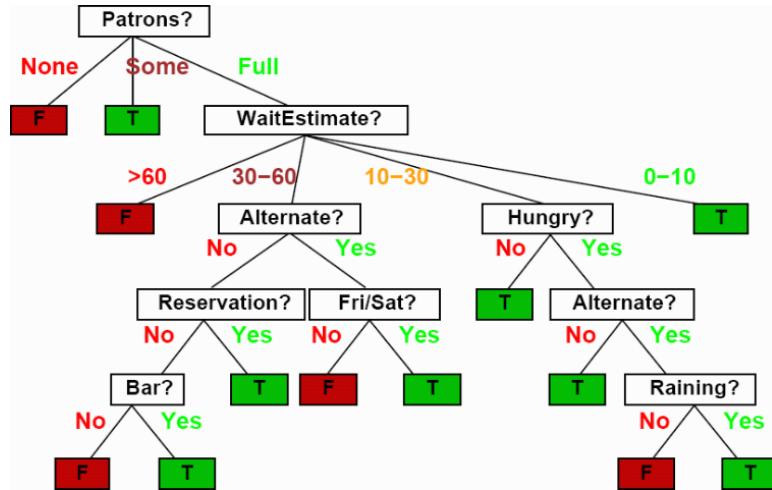


Figure 14.1: Example of dtree taken from Ref.[60]

Fig.14.1 shows a typical decision tree (dtree). This example was taken from Ref.[60], where it is analyzed in detail. As you can see, a dtree contains two main types of nodes: the non-leaf nodes, and the leaf nodes. The **non-leaf nodes** pose **questions**. In general, the **answers**¹ to those questions can be multiple choices with two or more choices. For each of those choices, a tree branch labeled by the choice comes down from the question node. The **leaf nodes** represent endpoints, goals, final conclusions, etc. Dtress can be viewed as classifiers. They take in a large amount of information about a population and compress that information to just a few classes. If S_c is the set of distinct leaf node labels, then we call each $c \in S_c$ a **class of the classifier**. In the case of Fig.14.1, $S_c = \{False, True\}$.

¹The **question-answer pairs** in dtrees are also called **attribute-value pairs**. Attributes are also called **features**.

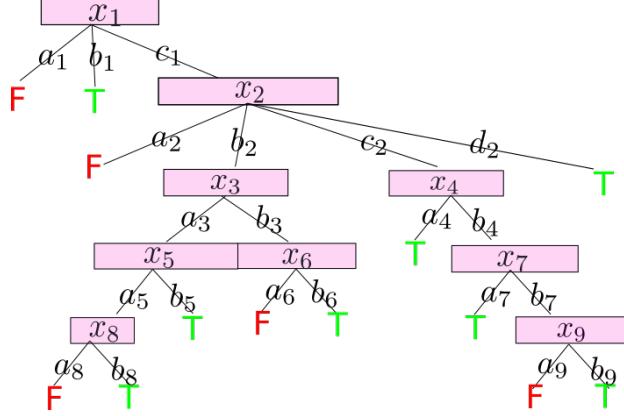


Figure 14.2: Fig.14.1 with abridged labels.

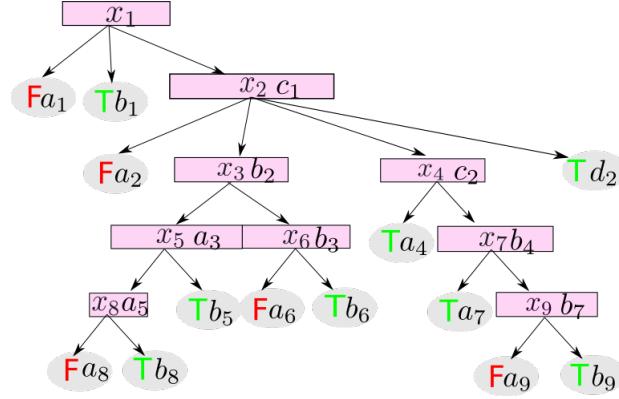


Figure 14.3: Fig.14.2 converted to a bnet.

Dtrees can be used with probabilities attached to each node, or without probabilities (as a plain undirected graph(UG)). This is analogous to bnets, which can be used with probabilities attached to each node (as DAGs with TPMs specified for each node) or without probabilities (as plain DAGs). Dtrees differ from bnets in that their tree branches are labelled, whereas bnet arrows aren't labelled. Also, whereas the nodes of a bnet carry a matrix of probabilities (the TPM), the nodes of a dtree carry just a column vector of probabilities which represents a single probability distribution. Henceforth, we will refer to the column vector of probabilities carried by each node of a dtree as its **Transition Probability Vector (TPV)**. Without the TPVs, a dtree can be used as a deterministic classifier, to classify inputs. With the TPVs, it can be used as a probabilistic sampler (to generate random samples.)

$P(x a)$	$a = a_0$	$a \in S_{\underline{a}}^- - \{a_0\}$	$null$
0	p_0	0	0
1	p_1	0	0
\vdots	\vdots	0	0
$N_{\underline{x}}^- - 1$	$p_{N_{\underline{x}}^- - 1}$	0	0
$null$	0	1	1

Table 14.1: TPM of a node of a dtree image bnet.

14.1 Transforming a dtree into a bnet

A trivial observation that is seldom made in the dtree pedagogical literature is that every dtree maps into a special bnet, let's call it its "image" bnet, in a very natural way. We use the dtree of Fig.14.1 as an example to show how to do this. As a first step, we go from Fig.14.1 to Fig.14.2 by replacing all the labels of the nodes and of the branches of the dtree by abridged symbols. Next, we go from Fig.14.2 to Fig.14.3, by replacing all tree branches by arrows pointing down, and by moving the tree branch labels down so that they become a suffix to the question that the tree branch leads to. At this point, we have created Fig.14.3, which constitutes the DAG of the image bnet. It remains for us to define a TPM for each node of this DAG.

Table 14.1 gives the TPM $P(x|a)$ for a node \underline{x} with single parent \underline{a} of a dtree image bnet. Say node \underline{x} has a set $S_{\underline{x}}^-$ of possible tree branches coming out of it. Let $N_{\underline{x}}^- = |S_{\underline{x}}^-|$. Let $S_{\underline{x}} = S_{\underline{x}}^- \cup \{null\}$ and $N_{\underline{x}} = |S_{\underline{x}}| = N_{\underline{x}}^- + 1$. Define $S_{\underline{a}}^-, N_{\underline{a}}^-, S_{\underline{a}}$ and $N_{\underline{a}}$ analogously for node \underline{a} . In Table 14.1, $S_{\underline{x}}^- = \{0, 1, \dots, N_{\underline{x}}^- - 1\}$ and a_0 is the value of node \underline{a} which labels the tree branch connecting nodes \underline{a} and \underline{x} . $\vec{p} = (p_0, p_1, \dots, p_{N_{\underline{x}}^- - 1})$ is a probability distribution associated with node \underline{x} , its TPV. TPVs can be learned from a dataset following the dtree Structure Learning (SL) algorithm discussed in Section 14.2.

Table 14.1 also applies when node \underline{x} is a leaf node, except that for leaf nodes, \vec{p} is one hot (i.e., all components are zero except for one component which is 1). Also, all leaf nodes \underline{x} have the same $S_{\underline{x}}^-$, namely S_c .

Adding a null state to the set of states (SOS) of each node of the image bnet is necessary because, once $null$ is added to the SOS of any node, it must be added to the SOS of all descendant nodes. $null$ must be added to the SOS of the children of the root node to take care of the situations when those first children don't receive the state they were expecting from their parent, i.e., the root node.

When drawing dtrees, some people put info like explanations and probabilities on the branches of the dtree. That info can all be preserved in the TPM and the node names and node state names of the image bnet nodes. One can also place info inside tool tips attached to the node name and node state names. Often, the pedagogical literature states that dtrees are more explicit and carry more info than their image bnets, but if one follows the above prescriptions, both can carry the same info.

A naive Bayes (NB) bnet (see Chapter 57) consists of a single “class node” with states S_c that fans out with arrows pointing to the “feature nodes”. If each leaf node of a NB bnet fans out into a set of new leaf nodes, and those new leaf nodes also fan out and so on, we get a generalized NB bnet. Let’s call this type of tree bnet an NB^* bnet. An NB^* bnet has the same graph structure as the image bnet of a dtree, but it’s more general, because its TPMs are more general. Each TPM of a NB^* bnet can have several non-trivial columns instead of just one $\text{TPV} = \vec{p}$.

14.2 Structure Learning for Dtrees

Let

$J_0 = \{0, 1, \dots, nj - 1\}$
 $\Sigma = \{0, 1, 2, \dots, nsam - 1\}$
 $DS = \{(\sigma, x^\sigma, c^\sigma) : \sigma \in \Sigma\}$ be a dataset
 $\sigma \in \Sigma$ be an individual (a sample) from a population,
 $x^\sigma \in S_{\underline{x}}$ be the **feature (attributes, questions) vector**. $S_{\underline{x}} = S_{x_0} \times S_{x_1} \times \dots \times S_{x_{nj-1}}$, $x = (x_0, x_1, \dots, x_{nj-1}) \in S_{\underline{x}}$, $x_j \in S_{x_j}$
 $c^\sigma \in S_c$ be a **classification class**
 We will assume $S_{\underline{x}}$ and S_c are finite sets.

Building a **classifier** Y (curve fit) for a dtree means finding a deterministic function $Y : S_{\underline{x}} \rightarrow S_c$ such that $c^\sigma \approx Y(x^\sigma)$ for all $\sigma \in \Sigma$. If we divide the population Σ into two large disjoint sets, a **training set** Σ_{train} and a **validation set** Σ_{vali} , and if $c^\sigma \approx Y(x^\sigma)$ very closely for $\sigma \in \Sigma_{train}$ but fits poorly for $\sigma \in \Sigma_{vali}$, then we say the classifier Y suffers from **overfitting**. We can learn the structure and TPVs of a dtree from a dataset DS , by using the dtree **Structure Learning (SL)** algorithm that we will discuss in detail later. However, that algorithm is prone to produce a classifier Y that overfits. Two techniques commonly used to reduce the effects of overfitting are **pruning** and **Random Forest (RF)** (see Chapter 66). Pruning just means somehow removing nodes that are too specific. An RF is an ensemble of dtrees that one averages over. In this chapter, we will only deal with a single dtree, not an ensemble of them.

Dtree SL was invented in 1984-1986 so it is fairly old. Many in the AI community consider dtrees old fashioned compared to neural nets. But dtrees are **interpretable** whereas neural nets aren’t.² Bnets are interpretable too.

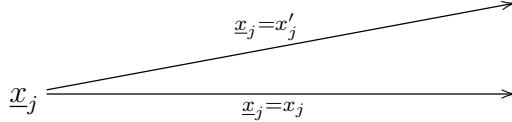
Below, we give the standard algorithm for SL of a dtree, in the form of pseudo-code. But first, we define two quantities, Information Gain and Gini, that are used in that pseudo-code.

²To be precise, only plain dtrees without boosting or bagging are interpretable. Dtrees used within boosting (see Chapter 1 on AdaBoost and Chapter 91 on XGBoost) or bagging (see Chapter 66 on Random Forest) gain much accuracy but lose **interpretability**.

14.2.1 Information Gain, Gini

This section uses various Shannon Information Theory entropies. Our notation for those entropies is described in Chapter C.

Call a **separation ability measure** (SAM) a measure used to decide, when constructing a dtree from a dataset, in what order to ask the questions about the feature vector x . The question order is decided by searching over all so far unused questions for the question with the largest SAM.³



$$\{N_j(c, x_j)\}_{c \in S_c, x_j \in S_{\underline{x}_j}}$$

$$\sum_{c \in S_c} N_j(c, x_j) = N_j(x_j)$$

$$\sum_{x_j \in S_{\underline{x}_j}} N_j(c, x_j) = N_j(c) \quad \sum_{c \in S_c} N_j(c) = \sum_{x_j \in S_{\underline{x}_j}} N_j(x_j) = N_j$$

Figure 14.4: Some population numbers associated with the nodes of a dtree. $N_j(c, x_j)$ is the number of individuals σ in the population that reaches node j , belonging to class c and having $\underline{x}_j = x_j$.

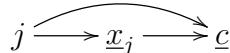


Figure 14.5: Bnet derived from population numbers in Fig.14.4

Fig.14.4 defines some population numbers associated with the nodes of a dtree. From these population numbers, we can define the bnet in Fig.14.5. The TPMs, printed in blue, for the (non-root) nodes of this bnet, are as follows

$$P(c|x_j, j) = \frac{N_j(c, x_j)}{N_j(x_j)} \quad (14.1)$$

$$P(x_j|j) = \frac{N_j(x_j)}{N_j} \quad (14.2)$$

³SAM is also called, somewhat confusingly, the splitting criterion and Gain.

where $j \in J_0$, $x_j \in S_{\underline{x}_j}$, and $c \in S_{\underline{c}}$ is a class node.

One can define the following information theory quantities⁴ associated with the bnet Fig.14.5.

$$INFO_fin_j = - \sum_c P(c|j) \ln P(c|j) \quad (14.3)$$

$$= H(\underline{c}|j) \quad (14.4)$$

$$(14.5)$$

$$INFO_init_j = \sum_{x_j \in S_{\underline{x}_j}} P(x_j|j) H(\underline{c}|x_j, j) \quad (14.6)$$

$$= H(\underline{c}|x_j, j) \quad (14.7)$$

$$IG_j = INFO_fin_j - INFO_init_j \quad (14.8)$$

$$= H(\underline{c}|j) - H(\underline{c}|x_j, j) \quad (14.9)$$

$$= H(\underline{c} : \underline{x}_j | j) \quad (14.10)$$

IG_j is called the **information gain for node \underline{x}_j** . Maximizing this mutual information produces a node \underline{x}_j that has a large correlation to a class c . If the goal is to reach a point where each leaf node is closely correlated to a different class, then maximizing the Information Gain of each new node is a greedy move towards that goal. Thus, Information Gain is a good SAM for dtree SL.

Note that if we approximate

$$\ln P(c|x_j) \approx \ln[1 + P(c|x_j) - 1] \quad (14.11)$$

$$\approx P(c|x_j) - 1 \quad (14.12)$$

in $H(\underline{c}|x_j)$, we get what is called the **Gini (or Gini Index) for node \underline{x}_k** :

$$H(\underline{c}|x_j) = - \sum_c P(c|x_j) \ln P(c|x_j) \quad (14.13)$$

$$\approx 1 - \sum_{c \in S_{\underline{c}}} P(c|x_j)^2 \stackrel{\text{def}}{=} Gini_{x_j} \quad (14.14)$$

⁴The average of $H(\underline{c} : b)$ over b is $H(\underline{c} : \underline{b}) = \sum_b P(b)H(\underline{c} : b)$. Likewise, the average of $H(\underline{c}|b)$ over b is $H(\underline{c}|b) = \sum_b P(b)H(\underline{c}|b)$. $H(\underline{c} : \underline{b})$ becomes $H(\underline{c} : b)$ and $H(\underline{c}|b)$ becomes $H(\underline{c}|b)$ when there is no b -prior (i.e., $P(b)$ is a delta function).

$Gini_{x_j}$ is a fairly good polynomial approximation to $H(\underline{c}|x_j)$. It is computationally much less expensive than $H(\underline{c}|x_j)$, because it does not require computing a log.

We say a probability distribution $P_{\underline{x}}$, is **pure (i.e., deterministic)** if $P_{\underline{x}}(x) = \delta(x, x_0)$. $Gini_{x_j}$ and $H(\underline{c}|x_j)$ are both always non-negative. They both vanish iff $P(c|x_j)$ is pure. Thus, $Gini_{x_j}$ and $H(\underline{c}|x_j)$ are both good measures of **class impurity**.

The **average Gini of node \underline{x}_j** is defined as

$$AGini_j = \sum_{x_j \in S_{\underline{x}_j}} P(x_j|j) Gini_{x_j} . \quad (14.15)$$

It measure the average impurity of the children of node \underline{x}_j .

In practice, the SL algorithm is done recursively. Each recursion step decides which feature x_j will be the root node of the current tree. For all “candidate” features (i.e., all \underline{x}_j that haven’t been used yet as tree nodes), one calculates IG_j , either exactly or approximately via Gini’s, using the following formula:

$$IG_j = \underbrace{H(\underline{c}|j)}_{\approx Gini_j} - \sum_{x_j \in S_{\underline{x}_j}} P(x_j|j) \underbrace{H(\underline{c}|x_j)}_{\approx Gini_{x_j}} \quad (14.16)$$

One then chooses $j = \underset{j}{\operatorname{argmax}} IG_j$. This maximizes the $\underline{c} - \underline{x}_j$ correlation.

Alternatively, some software programs use the average Gini $AGini_j$ as their SAM. They choose as root node $j = \underset{j}{\operatorname{argmin}} AGini_j$. This minimizes the average impurity of the children of node j . Since IG_j differs from $AGini_j$ by $H(\underline{c}|j)$, maximizing IG_j and minimizing $AGini_j$ might lead to different results.

Example of calculation of $AGini_j$ and IG_j

Suppose we deduce from a dataset the numbers in the yellow cells in Table 14.2. These numbers are repeated in Fig.14.6. Then we can calculate the white cells as follows:

$j = temp?$	$x_j = hot$	$x_j = med$	$x_j = cold$
$c = play$	3	3	3
$c = stay$	1	2	2
$Gini$	0.375	0.48	0.48
$AGini = 0.45$			
$H(\underline{c} x_j)$	≈ 0.375	≈ 0.48	≈ 0.48
$IG \approx 0$			

Table 14.2: Evaluating AGini and IG for node *temp*.

- $\text{Gini}(\text{hot}) = 1 - (3/4) - (1/4) = 0.375$

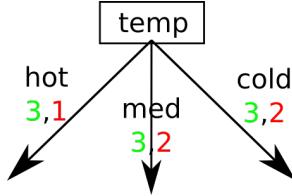


Figure 14.6: Tree stump corresponding to Table 14.2.

- $\text{Gini}(\text{med}) = 1 - (3/5)^2 - (2/5)^2 = 0.48$
- $\text{Gini}(\text{cold}) = 1 - (3/5)^2 - (2/5)^2 = 0.48$
- $\text{AGini} = (4/14)(0.375) + (5/14)(0.48) + (5/14)(0.48) = 0.45$
- $\text{IG} \approx [1 - (6/9)^2 - (3/9)^2] - \text{AGini} = 0.44 - 0.45 \approx 0$

This gives the AGini and IG for just one candidate root node. We would have to calculate AGini (or IG) for all possible candidates and choose the candidate with the lowest AGini (or highest IG).

14.3 Information Gain Ratio

The Information Gain Ratio (IGR) is an alternative SAM.

$$IGR_j = \frac{IG_j}{H(\underline{x}_j|j)} \quad (14.17)$$

$$= \frac{H(c : \underline{x}_j|j)}{H(\underline{x}_j|j)} \quad (14.18)$$

$$= \frac{H(\underline{x}_j|j) - H(\underline{x}_j|c, j)}{H(\underline{x}_j|j)} \quad (14.19)$$

$$= 1 - \frac{H(\underline{x}_j|c, j)}{H(\underline{x}_j|j)} \quad (14.20)$$

$$0 \leq IGR_j \leq 1$$

$$IGR_j = 0 \text{ iff } H(\underline{x}_j|c, j) = H(\underline{x}_j|j).$$

14.3.1 Pseudo-code

Below, we give the standard algorithm for SL of a dtree, in the form of pseudo-code. The strategy employed by the algo is to assume an incoming population into the current root node, then determine the feature x_j that best separates that incoming population. The feature x_j is chosen so as to maximize IG_j (or minimize $AGini_j$).

This process is repeated by nominating the end of each new branch to be the current root node. Features can appear as a node more than once, so the order in which nodes are split does not matter. In essence, what we are doing is performing a top-down, greedy search through the space of possible dtrees.

The pseudo-code below describes the following historically important software programs:

- CART (Classification and Regression Trees), invented by Breiman et al in 1984. Uses $AGini_j$ as SAM.
- ID3 (Iterative Dichotomiser 3) invented by Quinlan in 1986. Uses IG_j as SAM. C4.5/C5.0 are successors to ID3.

Thus, CART and ID were invented independently around the same time. The main difference between them is the SAM being used.

The pseudo-code below uses the majority function defined in Chapter C.

Algorithm 1: Pseudo-code for learning a dtree from a dataset

Input :

- dataset $DS = \{(\sigma, x^\sigma, c^\sigma) : \sigma \in \Sigma\}$
- set of currently available node indices J , where $J \subset J_0$

Output:

- tree T ,
- population numbers $\{(r, c, x_r, N_r(c, x_r)) : r \in J_0, c \in S_c, x_r \in S_{x_r}\}$ stored globally

From DS , calculate J_0, S_c, S_{x_i} for each i
 c^σ is called the target feature/attribute.

$$J \leftarrow J_0$$

Function `learn_dtreet(DS, J):`

```

 $\Sigma \leftarrow$  set of all  $\sigma$  in  $DS$ 
if  $\{c^\sigma : \sigma \in \Sigma\} = \{c\}$  then
     $T \leftarrow$  one node tree with leaf node label=  $c$ 
else if  $J = \emptyset$  then
     $T \leftarrow$  one node tree with leaf node label= majority( $[c^\sigma : \sigma \in \Sigma]$ )
else
     $r \leftarrow \underset{j \in J}{\operatorname{argmax}} IG_j(DS)$  // or replace  $\operatorname{argmax}_{j \in J} IG_j$  by  $\operatorname{argmin}_{j \in J} AGini_j$ 
    from  $DS$ , calculate  $\{(r, c, x_r, N_r(c, x_r)) : c \in S_c, x_r \in S_{x_r}\}$  and store it
        globally
    for  $v \in S_{x_r}$  do
        /* Notice that  $J$  is the same every time repeat this loop,
           so order in which  $v \in S_{x_r}$  are called does not matter.
           Furthermore, this means that multiple tree nodes may be
           labeled by same feature. */
        On current tree  $T$ , add a branch below  $x_r$  with label " $x_r = v$ "
         $DS|_{x_r=v} \leftarrow$  subset of  $DS$  with  $x_r = v$ 
        if  $DS|_{x_r=v} = \emptyset$  then
            below the new branch add a
            leaf node labeled = majority( $[c^\sigma : \sigma \in \Sigma]$ )
        else
            below the new branch add
            subtree = learn_dtreet( $DS|_{x_r=v}, J - \{r\}$ )
    return  $T$ 

```

Chapter 15

**Decisions Based on Rungs 2 and 3:
COMING SOON**

Chapter 16

Difference-in-Differences

This chapter is based on Ref.[8].

The Difference-in-Differences (DID) method was first used by John Snow in an 1854 report that argued that cholera in London was being transmitted by sewage polluted water rather than, as others at the time believed, by air (in fetid vapors called miasmas). In general, one can apply DID to discover causal effects in historical data. By **historical data** (a.k.a. a **natural experiment**. See Ref.[129]) we mean data that is collected long after the treatment (rather than during it) and is thus not subject to active intervention by the experimenter.

This chapter assumes that the reader has read Chapter 64 on Potential Outcomes (PO). The DID method applies the basic single-time PO theory described in Chapter 64, to 2 well separated times in which different conditions prevail.

16.1 John Snow, DID and a cholera transmission pathway

Let

$$\begin{aligned} d &\in \{0, 1\} \\ t &\in \{t_0, t_1\}, t_0 < t_1 \\ y &= f(d, t) \in \mathbb{R}. \end{aligned}$$

Define

$$\Delta_t f(d, t) = f(d, t_1) - f(d, t_0), \quad (16.1)$$

$$\Delta_d f(d, t) = f(1, t) - f(0, t), \quad (16.2)$$

$$DID = \delta = \Delta_d \Delta_t f(d, t). \quad (16.3)$$

DID is illustrated in Fig.16.1.

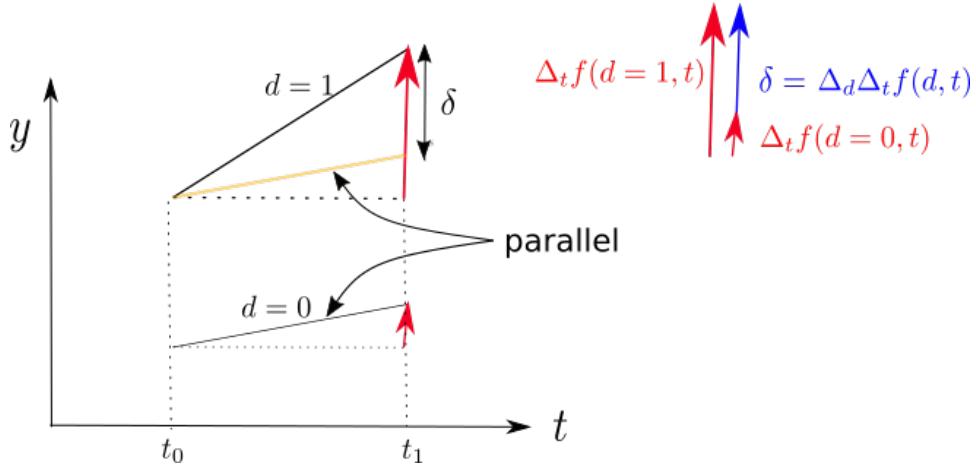


Figure 16.1: Pictorial representation of Difference-in-differences (DID) as a difference of two differences (i.e., a difference of two slopes).

A **time series** is any function of time for which the domain is a discrete set of times.

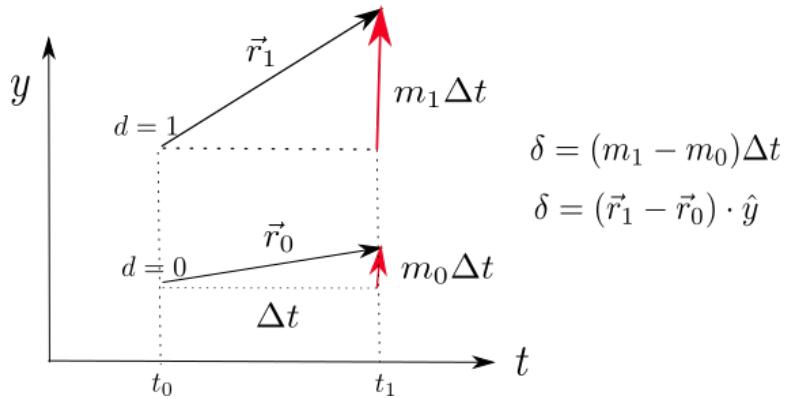


Figure 16.2: $DID = \delta$ expressed as difference of slopes or difference of vectors.

Note that, as shown Fig.16.2, $DID = \delta$ can also be expressed as a difference of 2 slopes times $\Delta t = t_1 - t_0$. Let \hat{y} be a unit in the y direction. δ can also be expressed as the dot product of a difference of 2 vectors dotted with \hat{y} .

A condensation of the data collected by John Snow in 1854 is given in Table 16.1. From that data, we find that

$$\delta = \Delta_d \Delta_t f(d, t) = (19 - 85) - (147 - 135) = -66 - 12 = -78 \quad (16.4)$$

	$t = t_0$ (1849)	$t = t_1$ (1854)
$d = 1$ (town 1)	85 deaths, polluted DW	19 deaths, unpolluted DW
$d = 0$ (town 0)	135 deaths, polluted DW	147 deaths, polluted DW

Table 16.1: A condensation of the data collected by John Snow in 1854, to test the hypothesis that cholera in London was being spread by polluted drinking water (DW).

16.2 PO analysis

In this section, we show how to analyze the DID method using the formalism of PO theory.

We will speak of a treatment outcome $\underline{y}_{t,g^\sigma}^\sigma(c^\sigma, x^\sigma)$ for individual σ that depends, not just on the treatment dose $c^\sigma \in \{0, 1\}$ and the confounder state x^σ , but also on a group parameter (i.e., which population or town) $g^\sigma \in \{0, 1\}$ and on a time parameter $t \in \{t_0, t_1\}$ (note t is independent of σ). Actually, we will assume $g^\sigma = c^\sigma$, so we will just speak of $\underline{y}_t^\sigma(c^\sigma, x^\sigma)$ with no explicit g^σ dependence. As usual for PO theory, we will consider expected values of y_t^σ :

$$E_{\sigma|d,x}[\underline{y}_t^\sigma(c)] = E_{\underline{y}_t(c)|d,x}[\underline{y}_t(c)] = \mathcal{Y}_{c|d,x}(t) \quad (16.5)$$

To calculate these expected values, we need a “model” with probability distributions. In this case, the needed model and probability distributions are provided by the bnets depicted in Fig.16.3. The TPMs, printed in blue, for the bnet $G_{t,+}$ in Fig.16.3, are as follows. Note that the TPMs for the bnet $G_{t,+}$ are defined in terms of the TPMs for the bnet G_t .

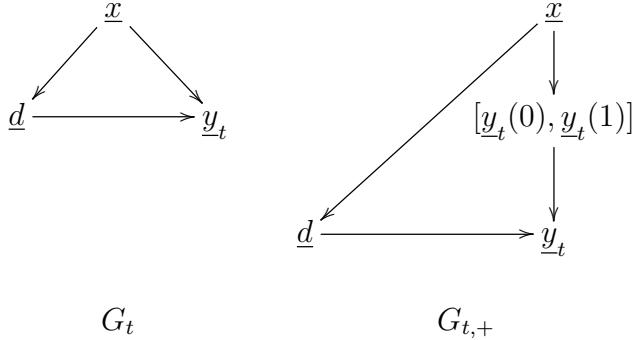


Figure 16.3: $t \in \{t_0, t_1\}$. Bnet $G_{t,+}$ is obtained by adding two new nodes $\underline{y}_t(0)$ and $\underline{y}_t(1)$ to bnet G_t .

$$P(x) = P_{\underline{x}}(x) \quad (16.6)$$

$$P(d|x) = P_{d|\underline{x}}(d|x) \quad (16.7)$$

$$P(y_t|y_t(0), y_t(1), d) = \mathbb{1}(y_t = y_t(d)) \quad (16.8)$$

$$P(y_t(c)|x) = P(y_t(c)|d, x) = \text{given} \quad (16.9)$$

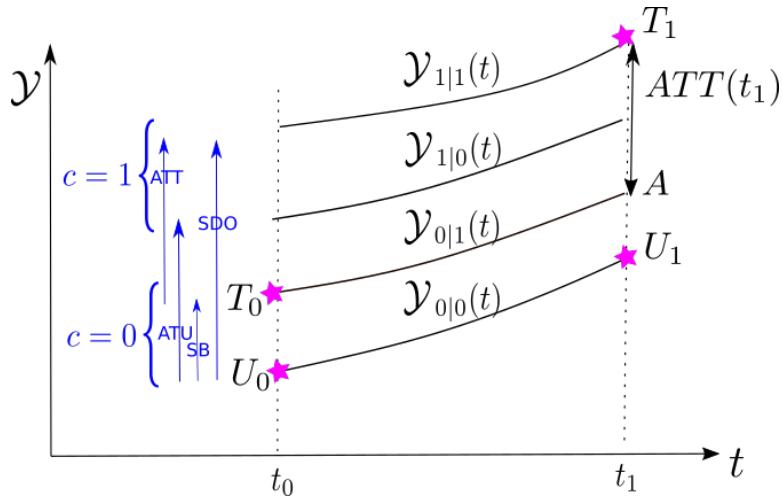


Figure 16.4: Four different time-dependent expected values $\mathcal{Y}_{c|d}(t)$ of y_t^σ for bnet $G_{t,+}$. The 4 magenta stars represent the 4 DID measurements.

We define the function $c(t)$ for $t = t_0, t_1$ by

$$c(t) = \begin{cases} 0 & \text{if } t = t_0 \\ d & \text{if } t = t_1 \end{cases} \quad (16.10)$$

Now we claim that the DID δ calculated in the previous section for John Snow's data, can be expressed in PO formalism as follows:

$$\delta = \Delta_d \Delta_t \mathcal{Y}_{c(t)|d}(t) . \quad (16.11)$$

Fig.16.4 depicts the four functions $\mathcal{Y}_{c|d}(t)$ for t in the interval $[t_0, t_1]$ and for $c, d \in \{0, 1\}$. The \mathcal{Y} coordinates of the four magenta stars in Fig.16.4 can be calculated using bnet G_t .

Define the **parallel trends** (PT) by

$$PT = \Delta_d \Delta_t \mathcal{Y}_{0|d}(t) . \quad (16.12)$$

We will say the **parallel trends assumption (PTA)** holds if $PT = 0$.

Next we prove that the DID δ equals the sum of an ATT¹ and PT.

$$\delta = \Delta_d \Delta_t \mathcal{Y}_{c(t)|d}(t) \quad (16.13)$$

$$= [\Delta_t \mathcal{Y}_{c(t)|1}(t) - \Delta_t \mathcal{Y}_{c(t)|0}(t)] \quad (16.14)$$

$$= \mathcal{Y}_{1|1}(t_1) - \mathcal{Y}_{0|1}(t_0) - \{\mathcal{Y}_{0|0}(t_1) - \mathcal{Y}_{0|0}(t_0)\} \quad (16.15)$$

$$= \mathcal{Y}_{1|1}(t_1) - \mathcal{Y}_{0|1}(t_0) - \{\mathcal{Y}_{0|0}(t_1) - \mathcal{Y}_{0|0}(t_0)\} + \underbrace{\{\mathcal{Y}_{0|1}(t_1) - \mathcal{Y}_{0|1}(t_1)\}}_{\text{zero}} \quad (16.16)$$

$$= \underbrace{\mathcal{Y}_{1|1}(t_1) - \mathcal{Y}_{0|1}(t_1)}_{ATT(t_1)} - \mathcal{Y}_{0|1}(t_0) - \{\mathcal{Y}_{0|0}(t_1) - \mathcal{Y}_{0|0}(t_0)\} + \mathcal{Y}_{0|1}(t_1) \quad (16.17)$$

$$= ATT(t_1) - \Delta_t \mathcal{Y}_{0|0}(t) + \Delta_t \mathcal{Y}_{0|1}(t) \quad (16.18)$$

$$= ATT(t_1) + \underbrace{\Delta_d \Delta_t \mathcal{Y}_{0|d}(t)}_{\text{zero if PTA holds}} \quad (16.19)$$

16.3 Linear Regression

In this section, we show how to apply linear regression (LR) to the PO analysis of DID.

As before, let $y_t^\sigma(c^\sigma)$ be the treatment outcome for individual σ , who receives a treatment dose c^σ at times $t \in \{t_0, t_1\}$. $y_t^\sigma(c^\sigma)$ can be fitted as follows. Here ϵ^σ is the residual for individual σ , and $b_0, m_0, b_1, m_1 \in \mathbb{R}$ are the fit parameters.

$$y_t^\sigma = [b_0 + m_0(t - t_0)](1 - c^\sigma) + [b_1 + m_1(t - t_0)]c^\sigma + \epsilon^\sigma. \quad (16.20)$$

Note that Eq.(16.20) yields a straight line in the $y_t^\sigma - t$ plane for $c^\sigma = 0$, and another straight line for $c^\sigma = 1$. We are using the standard symbols b to denote the y-intercept, and m to denote the slope of a straight line.

Taking the expected value of Eq.(16.20), we get

$$\mathcal{Y}_{c|d}(t) = [b_0 + m_0(t - t_0)](1 - d) + [b_1 + m_1(t - t_0)]d. \quad (16.21)$$

If $\Delta t = t_1 - t_0$, then

$$\mathcal{Y}_{d|d}(t_1) = [b_0 + m_0 \Delta t](1 - d) + [b_1 + m_1 \Delta t]d, \quad (16.22)$$

and

$$\mathcal{Y}_{0|d}(t_0) = b_0. \quad (16.23)$$

¹ATT stands for the average treatment effect of the treated. ATT is defined in Chapter 64

Thus,

$$\delta = \Delta_d \Delta_t \mathcal{Y}_{c(t)|d}(t) \quad (16.24)$$

$$= \Delta_d [\mathcal{Y}_{d|d}(t_1) - \mathcal{Y}_{0|d}(t_0)] \quad (16.25)$$

$$= (m_1 - m_0) \Delta t . \quad (16.26)$$

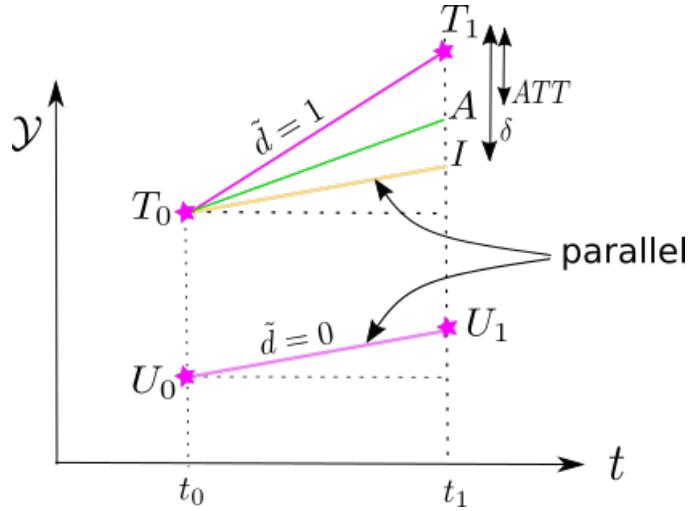


Figure 16.5: We use Linear Regression to fit a straight line between points U_0 and U_1 , and between points T_0 and T_1 . (U =untreated, T =treated, subscript refers to times t_0, t_1). U_0, T_0, U_1, T_1 are the measurement points. Point I is an image of point U_1 .

	$t = t_0$	$t = t_1$
$d = 1$	$\mathcal{Y}(T_0) = \mathcal{Y}_{0 1}(t_0)$	$\mathcal{Y}(T_1) = \mathcal{Y}_{1 1}(t_1)$
$d = 0$	$\mathcal{Y}(U_0) = \mathcal{Y}_{0 0}(t_0)$	$\mathcal{Y}(U_1) = \mathcal{Y}_{0 0}(t_1)$

Table 16.2: \mathcal{Y} coordinates of points U_0, T_0, U_1, T_1 in Figs.16.4 and 16.5.

Figs.16.4 and 16.5 define points U_0, T_0, U_1, T_1, I, A . The \mathcal{Y} coordinates of points U_0, T_0, U_1, T_1 are given by Table 16.2. The \mathcal{Y} coordinates of points A, I are given by Eqs.16.27.

$$\mathcal{Y}(A) = \mathcal{Y}_{0|1}(t_1) \quad (16.27a)$$

$$\mathcal{Y}(I) = \mathcal{Y}(U_1) + [\mathcal{Y}(T_0) - \mathcal{Y}(U_0)] \quad (16.27b)$$

We can express ATT and the δ for DID in terms of the \mathcal{Y} of the points U_0, T_0, U_1, T_1, I, A . Indeed,

$$\delta = \mathcal{Y}(T_1) - \mathcal{Y}(I) \quad (16.28)$$

$$= \mathcal{Y}(T_1) - \mathcal{Y}(U_1) - [\mathcal{Y}(T_0) - \mathcal{Y}(U_0)] \quad (16.29)$$

$$ATT = \mathcal{Y}(T_1) - \mathcal{Y}(A) \quad (16.30)$$

Hence,

$$\delta = ATT \iff \mathcal{Y}(I) = \mathcal{Y}(A) \iff \text{PTA holds} \quad (16.31)$$

Chapter 17

Diffusion Models

This paper is based on Ref.[74]

Diffusion Models (DM) are a way of generating fake images from an original image. They are a competitor to GANs (see Chapter 32), and are used in DALL-E (OpenAI’s computer program that generates images from text).

DM works by subjecting each degree of freedom of the image to a forward Markov chain of transformations labeled $t = 0, 1, 2, \dots, T$ and a reverse Markov chain of transformations labeled $t = T - 1, T - 2, \dots, 0$. Each step of the forward chain multiplies each degree of freedom of the image by a constant and adds white noise to it. The last image ($t = T$) of the forward chain is changed to the point that it looks like a normal distribution. The reverse chain tries to undo, as well as possible, the alterations to the original image done by the forward chain. Full, faithful restoration is impossible because the forward chain is an irreversible process.

Of course, a DM is only a small part of the magic of DALL-E. I haven’t studied DALL-E’s algorithm, but my guess is that it works roughly as follows. Given a text description of an image, such as “A hedgehog using a calculator painted in the style of Vincent van Gogh”, it uses a neural net trained on a vast corpus of words and images, to match separate words in the description with an image for each word. Then it uses a second neural net to create a pastiche from the set of images created in the first stage. Then, it probably modifies that pastiche by passing it through a stylistic filter that can be specified in the initial description. Finally, it uses a DM to smooth the transitions of the pastiche.

17.1 Bnet for DM

For $t \in 1, 2, \dots, T$, let

$$0 < \alpha^t < 1, \beta^t = 1 - \alpha^t \quad (17.1)$$

$$\pi_1^t \alpha = \prod_{\tau=1}^t \alpha^\tau \quad (17.2)$$

Let

$\underline{x}^t, \tilde{\underline{x}}^t \in \mathbb{R}^{df}$ for $t = 0, 1, 2, \dots, T$ be column vectors describing an image. Here df is the number of degrees of freedom of the image. \underline{x}^0 will denote the initial Original image, $\underline{x}^T = \tilde{\underline{x}}^T$ will denote an image which is very close to a normal distribution, and $\tilde{\underline{x}}^0$ will denote the final Fake image.

$I \in \mathbb{R}^{df \times df}$ is the identity matrix

$\underline{w}^t, \underline{a}^t, \underline{n}_\theta^t \in \mathbb{R}^{df}$

Fig.17.1 shows a bnet for DM, and 17.2 shows the same bnet in more detail.¹

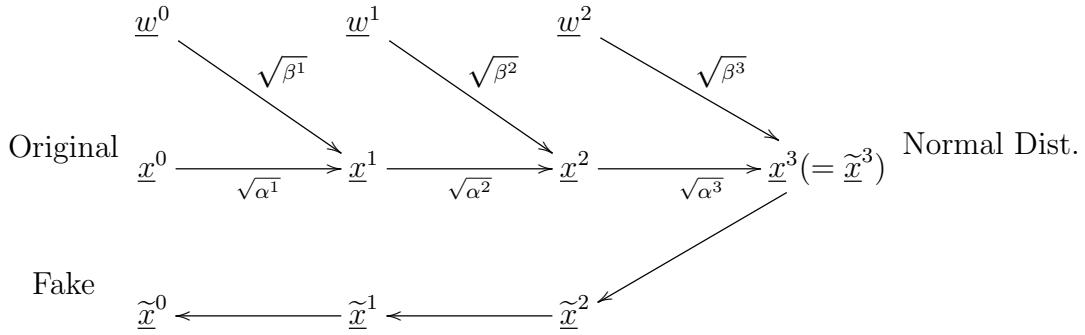


Figure 17.1: Bnet for DM with $T = 3$. See Chapter 44 for an explanation of LDEN notation.

The TPMs, printed in blue, for the bnet of Fig.17.1 are as follows:

$$P(x^0) = \delta(x^0, X^0) \quad (\text{original image}) \quad (17.3)$$

$$P(x^t | x^{t-1}, w^t) = \mathbb{1}(x^t = \sqrt{\alpha^t} x^{t-1} + \sqrt{\beta^t} w^{t-1}) \quad (17.4)$$

$$P(w^t) = \mathcal{N}(w^t; \mu = 0, \sigma^2 = I) \quad (\text{white noise}) \quad (17.5)$$

$$P(\tilde{x}^{t-1} = x^{t-1} | \tilde{x}^t = x^t) = \tilde{P}_\theta(x^{t-1} | x^t) \quad (17.6)$$

$$= \mathcal{N}(x^{t-1}; \mu = M_\theta^{t-1}(x^t), \sigma^2 = \Sigma_\theta^{t-1}(x^t)) \quad (17.7)$$

¹Ref.[74] uses \underline{z}^t instead of \underline{w}^t for white noise. Note that the time index of \underline{w}^t often does not matter because $\underline{w}^t \sim \mathcal{N}(0, I)$ for all t . This does not mean that we can drop the time index of \underline{w}^t , because \underline{w}^{t_1} and \underline{w}^{t_2} are uncorrelated for $t_1 \neq t_2$, so we can get into trouble if we assume $\underline{w}^{t_1} = \underline{w}^{t_2} = \underline{w}$.

We will assume that

$$\Sigma_{\theta}^{t-1}(x^t) = (\sigma^{t-1})^2 I \quad (17.8)$$

where $\sigma^{t-1} \in \mathbb{R}$.

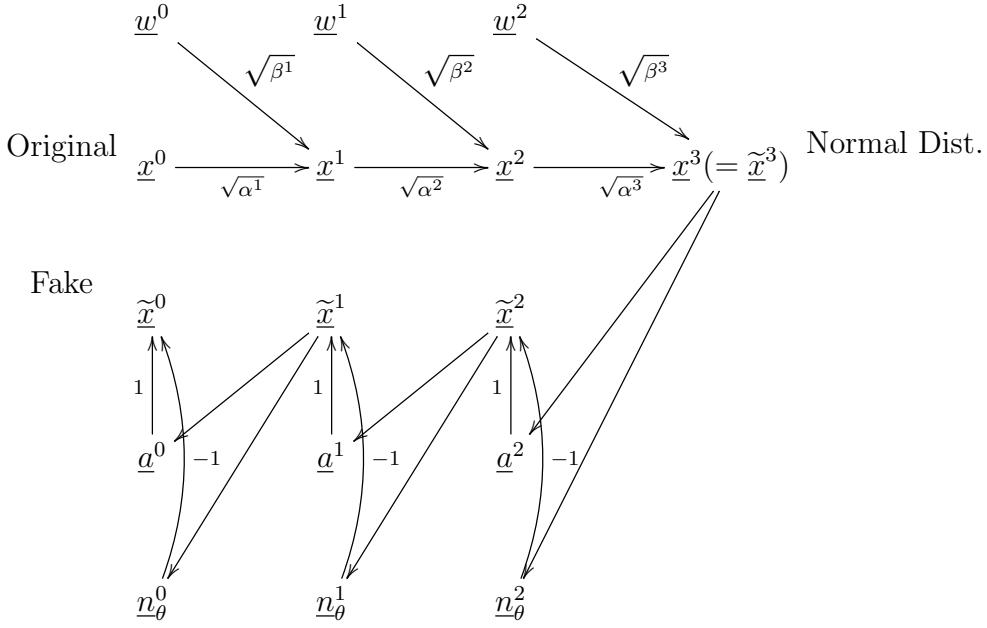


Figure 17.2: The bnet of Fig.17.1 shown in more detail.

The TPMs, printed in blue, for the bnet of Fig.17.2 are as follows:

$$P(n_{\theta}^t) = \mathcal{N}(n_{\theta}^t; \mu = M_{\theta}^t(x^{t+1}), \sigma = \sigma^t) \quad (17.9)$$

$$P(a^t | \tilde{x}^{t+1} = x^{t+1}) = \mathbb{1}(a^t = M^t(x^{t+1})) \quad (17.10)$$

$$P(\tilde{x}^t = x^t | a_t, n_{\theta}^t) = \mathbb{1}(x^t = a^t - n_{\theta}^t) \quad (17.11)$$

Note that these TPMs for the bnet Fig.17.2 imply that

$$\underbrace{P(\tilde{x}^t = x^t | \tilde{x}^{t+1} = x^{t+1})}_{\text{call this } \tilde{P}_{\theta}(x^t | x^{t+1})} = \mathcal{N}(x^t; \mu = M^t(x^{t+1}) - M_{\theta}^t(x^{t+1}), \sigma = \sigma^t) \quad (17.12)$$

17.2 Mean Values $M^{t-1}(x^t)$ and $M_\theta^{t-1}(x^t)$

Claim 25

$$\boxed{\underline{x}^t = \sqrt{\pi_1^t \alpha} \underline{x}^0 + \sqrt{1 - \pi_1^t \alpha} \underline{w}^t} \quad (17.13)$$

$$\begin{array}{ccc} & \underline{w}^t & \\ & \downarrow & \\ \underline{x}^0 & \xrightarrow[\sqrt{\pi_1^t \alpha}]{} & \underline{x}^t \end{array} \quad (17.14)$$

proof:

Suppose \underline{x}_1 and \underline{x}_2 are independent random variables with variances V_1 and V_2 , respectively. Then the variance V of $\underline{x} = \underline{x}_1 + \underline{x}_2$ is

$$V = \langle \underline{x}, \underline{x} \rangle \quad (17.15)$$

$$= \langle \underline{x}_1 + \underline{x}_2, \underline{x}_1 + \underline{x}_2 \rangle \quad (17.16)$$

$$= \langle \underline{x}_1, \underline{x}_1 \rangle + \langle \underline{x}_2, \underline{x}_2 \rangle \quad (17.17)$$

$$= V_1 + V_2 \quad (17.18)$$

By similar reasoning, the mean of \underline{x} equals the sum of the means of \underline{x}_1 and \underline{x}_2 . It's also true that if both \underline{x}_1 and \underline{x}_2 are normally distributed, then \underline{x} is too. We will refer to a sum of independent normals as a SIN.

Let $\underline{w} \sim \mathcal{N}(0, I)$.

$$\underline{x}^t = \sqrt{\alpha^t} \underline{x}^{t-1} + \sqrt{1 - \alpha^t} \underline{w}^{t-1} \quad (17.19)$$

$$= \sqrt{\alpha^t} [\sqrt{\alpha^{t-1}} \underline{x}^{t-2} + \sqrt{1 - \alpha^{t-1}} \underline{w}^{t-2}] + \sqrt{1 - \alpha^t} \underline{w}^{t-1} \quad (17.20)$$

$$= \sqrt{\alpha^t \alpha^{t-1}} \underline{x}^{t-2} + [\sqrt{\alpha^t (1 - \alpha^{t-1})} \underline{w}^{t-2} + \sqrt{1 - \alpha^t} \underline{w}^{t-1}] \quad (17.21)$$

$$= \sqrt{\alpha^t \alpha^{t-1}} \underline{x}^{t-2} + \sqrt{1 - \alpha^t \alpha^{t-1}} \underline{w} \quad (\text{because it's a SIN}) \quad (17.22)$$

$$= \dots \quad (17.23)$$

$$= \sqrt{\pi_1^t \alpha} \underline{x}^0 + \sqrt{1 - \pi_1^t \alpha} \underline{w} \quad (17.24)$$

Now replace \underline{w} by \underline{w}^t . This is justified because they are both $\mathcal{N}(0, I)$, and, when $t_1 \neq t_2$, we want the \underline{w} for \underline{x}^{t_1} to be independent from the \underline{w} for \underline{x}^{t_2} .

QED

Solving Eq.(17.13) for \underline{x}^0 , we get

$$\boxed{\underline{x}^0 = \frac{1}{\sqrt{\pi_1^t \alpha}} \underline{x}^t - \frac{\sqrt{1 - \pi_1^t \alpha}}{\sqrt{\pi_1^t \alpha}} \underline{w}^t} \quad (17.25)$$

$$\begin{array}{ccc}
& \frac{\sqrt{1-\pi_1^t \alpha}}{\sqrt{\pi_1^t \alpha}} & \underline{w}^t \\
& \swarrow & \nearrow \\
\underline{x}_0 & \xleftarrow[\frac{1}{\sqrt{\pi_1^t \alpha}}]{} & \underline{x}^t
\end{array} \tag{17.26}$$

Claim 26

$$P(x^{t-1}|x^t, x^0) = P(x^t|x^{t-1}) \frac{P(x^{t-1}|x^0)}{P(x^t|x^0)} \tag{17.27}$$

proof:

This is a simple consequence of Bayes rule. We will prove this in two ways: algebraically and graphically.

The algebraic proof goes as follows.

$$P(x^{t-1}|x^t, x^0)P(x^t|x^0) = \overbrace{P(x^t|x^{t-1}, x^0)}^{P(x^t|x^{t-1})} P(x^{t-1}|x^0) \tag{17.28}$$

The graphical proof, although longer, is more intuitive. Start by noticing that

$$\begin{aligned}
& \sum w^0 \quad \dots \quad \sum w^{t-2} \quad \sum w^{t-1} \\
& \downarrow \qquad \qquad \downarrow \qquad \qquad \downarrow \\
x^0 & \longrightarrow \sum x^1 \longrightarrow \dots \longrightarrow x^{t-1} \longrightarrow x^t \\
= & \quad x^0 \longrightarrow x^{t-1} \longrightarrow x^t \text{ (diagram A)} \tag{17.29} \\
= & \quad x^0 \xrightarrow[0]{\quad} x^{t-1} \xrightarrow{\quad} x^t \quad (\text{make fully connected}) \\
= & \quad x^0 \xrightarrow[0]{\quad} x^{t-1} \xleftarrow{\quad} x^t \quad (\text{reverse arrow}) \text{ (diagram B)}
\end{aligned}$$

Therefore,

$$\underbrace{P(x^{t-1}|x^t, x^0)P(x^t|x^0)}_{\text{diagram B of Eq.17.29}} = \underbrace{P(x^t|x^{t-1})P(x^{t-1}|x^0)}_{\text{diagram A of Eq.17.29}} \tag{17.30}$$

QED

Define the following mean value:

$$M^{t-1}(x^t) = \sum_{x^{t-1}} x^{t-1} P(x^{t-1}|x^t, x^0) \tag{17.31}$$

Claim 27

$$M^{t-1}(x^t) = \frac{\sqrt{\alpha^t}(1 - \pi_1^{t-1}\alpha)}{1 - \pi_1^t\alpha} x^t + \frac{\sqrt{\pi_1^{t-1}\alpha} \beta^t}{1 - \pi_1^t\alpha} x^0 \quad (17.32)$$

proof:

By Claim 26, we know that $\mathcal{Q} = P(x^{t-1}|x^t, x^0)$ can be expressed as a product of two Gaussians $P(x^t|x^{t-1})$ and $P(x^{t-1}|x^0)$, divided by a third Gaussian $P(x^t|x^0)$, and all 3 of these Gaussians have been calculated in closed form previously in this chapter. So it's just a matter of algebra to express \mathcal{Q} as a Gaussian, complete the square inside the exponent of \mathcal{Q} , and thus obtain its mean value $M^{t-1}(x^t)$. We leave the algebra to the reader. If in doubt, the algebra can be found in Ref.[74].

QED

Claim 28

$$M^{t-1}(x^t) = \frac{1}{\sqrt{\alpha^t}} \left(x^t - \frac{\beta^t}{\sqrt{1 - \pi_1^t\alpha}} w^t \right) \quad (17.33)$$

proof:

Use Eq.(17.25) to replace x^0 in Eq.(17.32).

QED

Let us parameterize $M_\theta^{t-1}(x^t)$ by

$$M_\theta^{t-1}(x^t) = \frac{1}{\sqrt{\alpha^t}} \left(x^t - \frac{\beta^t}{\sqrt{1 - \pi_1^t\alpha}} n_\theta^{t-1}(x^t) \right) \quad (17.34)$$

If we define

$$C^t = \frac{-\beta^t}{\sqrt{\alpha^t(1 - \pi_1^t\alpha)}} , \quad (17.35)$$

then

$$M^{t-1}(x^t) - M_\theta^{t-1}(x^t) = C^t [w^t - n_\theta^{t-1}(x^t)] \quad (17.36)$$

$$= C^t [w^t - n_\theta^{t-1}(\sqrt{\pi_1^t\alpha} x^0 + \sqrt{1 - \pi_1^t\alpha} w^t)] \quad (17.37)$$

17.3 Loss function \mathcal{L}

Note that²

$$0 \leq D_{KL}(P(x^{1:T}|x^0) \parallel \tilde{P}_\theta(x^{1:T}|x^0)) \quad (17.38)$$

$$= \sum_{x^{1:T}} P(x^{1:T}|x^0) \ln \frac{P(x^{1:T}|x^0)}{\tilde{P}_\theta(x^{1:T}|x^0)} \quad (17.39)$$

$$= \ln \tilde{P}_\theta(x^0) + \sum_{x^{1:T}} P(x^{1:T}|x^0) \ln \frac{P(x^{1:T}|x^0)}{\tilde{P}_\theta(x^{0:T})} \quad (17.40)$$

$$\underbrace{- \sum_{x^0} P(x^0) \ln \tilde{P}_\theta(x^0)}_{-E_{x^0}[\ln \tilde{P}_\theta(x^0)]} \leq \underbrace{\sum_{x^{0:T}} P(x^{0:T}) \ln \frac{P(x^{1:T}|x^0)}{\tilde{P}_\theta(x^{0:T})}}_{E_{x^{0:T}}\left[\ln \frac{P(x^{1:T}|x^0)}{\tilde{P}_\theta(x^{0:T})}\right]} = \mathcal{L}(\theta) = \text{loss function} \quad (17.41)$$

Henceforth, expected values $E_{x^{0:T}}$ are to be understood as being with respect to $P(x^{0:T})$.

The left hand side of the inequality Eq.(17.41) is expected to be a small positive constant, because we expect $\tilde{P}_\theta(x^0) \approx P(x^0) \approx \delta(x^0, X^0)$. Thus, we are justified in defining the right hand side of Eq.(17.41) as loss function $\mathcal{L} = \mathcal{L}(\theta)$ to be minimized.

Claim 29

$$\mathcal{L} = \sum_{t=0}^T \mathcal{L}^t \quad (17.42)$$

$$\mathcal{L}^0 = -\ln \tilde{P}_\theta(x^0|x^1) \quad (17.43)$$

$$\mathcal{L}^{t-1} = E_{x^{0:T}} \left[\ln \frac{P(x^{t-1}|x^t, x^0)}{\tilde{P}_\theta(x^{t-1}|x^t)} \right] \quad \text{for } t = 2, 3, \dots, T \quad (17.44)$$

$$\mathcal{L}^T = E_{x^{0:T}} \left[\ln \frac{P(x^T|x^0)}{\tilde{P}_\theta(x^T)} \right] \quad (17.45)$$

² D_{KL} is the Kullback-Leibler divergence. It's defined in Chapter C.

proof:

$$\mathcal{L} = E_{x^{0:T}} \left[\ln \frac{P(x^{1:T}|x^0)}{\tilde{P}_\theta(x^{0:T})} \right] \quad (17.46)$$

$$= E_{x^{0:T}} \left[\ln \frac{\prod_{t=1}^T P(x^t|x^{t-1})}{\tilde{P}_\theta(x^T) \prod_{t=1}^T \tilde{P}_\theta(x^{t-1}|x^t)} \right] \quad (17.47)$$

$$= E_{x^{0:T}} \left[\ln \frac{P(x^1|x^0) \prod_{t=2}^T \overbrace{P(x^t|x^{t-1})}^{P(x^t|x^{t-1}, x^0)}}{\tilde{P}_\theta(x^T) \tilde{P}_\theta(x^0|x^1) \prod_{t=2}^T \tilde{P}_\theta(x^{t-1}|x^t)} \right] \quad (17.48)$$

$$= E_{x^{0:T}} \left[\ln \frac{P(x^1|x^0) \prod_{t=2}^T P(x^{t-1}|x^t, x^0) \frac{P(x^t|x^0)}{P(x^{t-1}|x^0)}}{\tilde{P}_\theta(x^T) \tilde{P}_\theta(x^0|x^1) \prod_{t=2}^T \tilde{P}_\theta(x^{t-1}|x^t)} \right] \quad (17.49)$$

$$= E_{x^{0:T}} \left[\ln \frac{P(x^1|x^0) \frac{P(x^T|x^0)}{P(x^1|x^0)} \prod_{t=2}^T P(x^{t-1}|x^t, x^0)}{\tilde{P}_\theta(x^T) \tilde{P}_\theta(x^0|x^1) \prod_{t=2}^T \tilde{P}_\theta(x^{t-1}|x^t)} \right] \quad (17.50)$$

$$= E_{x^{0:T}} \left[\ln \frac{1}{\tilde{P}_\theta(x^0|x^1)} \cdot \frac{P(x^T|x^0)}{\tilde{P}_\theta(x^T)} \cdot \frac{\prod_{t=2}^T P(x^{t-1}|x^t, x^0)}{\prod_{t=2}^T \tilde{P}_\theta(x^{t-1}|x^t)} \right] \quad (17.51)$$

QED

We expect $\tilde{P}_\theta(x^0|x^1)$ in \mathcal{L}^0 to depend only very weakly on θ because it comes at the end of the reverse Markov chain. Likewise, we expect $\tilde{P}_\theta(x^T)$ in \mathcal{L}^T to be approximately a normal distribution independent of θ , because it comes at the beginning of the reverse Markov chain. Hence, we are justified in using $\mathcal{L}' = \sum_{t=2}^T \mathcal{L}^{t-1}$ as the new loss function to me minimized with respect to θ .

Claim 30

$$\mathcal{L}^{t-1} = E_{x^t} \left[\frac{1}{2(\sigma^{t-1})^2} [M^{t-1}(x^t) - M_\theta^{t-1}(x^t)]^2 \right] \quad (17.52)$$

$$= E_{x^0, w^t} \left[\frac{(C^t)^2}{2(\sigma^{t-1})^2} \left[w^t - n_\theta^{t-1} (\sqrt{\pi_1^\alpha} x^0 + \sqrt{1 - \pi_1^\alpha} w^t) \right]^2 \right] \quad (17.53)$$

proof:

Eq.(17.53) follows trivially from Eq.(17.37) and Eq.(17.52).

To show Eq.(17.52), recall that

$$\mathcal{L}^{t-1} = E_{x^{0:T}} \left[\ln \frac{P(x^{t-1}|x^t, x^0)}{\tilde{P}_\theta(x^{t-1}|x^t)} \right] \quad (17.54)$$

The denominator $\tilde{P}_\theta(x^{t-1}|x^t)$ is one of the TPMs for the bnet Fig.17.2 , and it's a normal distribution with mean $M_\theta^{t-1}(x^t)$. The numerator $P(x^{t-1}|x^t, x^0)$ is a normal distribution too; it was calculated in the proof of Claim 27, where its mean was called $M^{t-1}(x^t)$. Hence, after some algebra which is left to the reader, one can show Eq.(17.52).

QED

Eq.(17.53) for \mathcal{L}^{t-1} is usually simplified to

$$\mathcal{L}_{simple} = E_{x^0, w^t} \left[\left[w^t - n_\theta^{t-1}(\sqrt{\pi_1^t \alpha} x^0 + \sqrt{1 - \pi_1^t \alpha} w^t) \right]^2 \right] \quad (17.55)$$

17.4 Algorithms for training and sampling DM

Algorithm 2: Algorithm for training DM (i.e., finding optimum θ)

Input : $\{\beta^t\}_{t=1}^T$, $0 < \epsilon < 1$, $n_\theta^t(x)$ function, θ_{in} , $\Delta\theta_{in}$, $P(x^0)$, T , $\eta > 0$

Output: θ_{next} = optimal θ

$\Delta\theta = \Delta\theta_{in}$

$\theta_{next} = \theta_{in}$

while $|\Delta\theta| > \epsilon$ **do**

Choose $x^0 \sim P(x^0)$

Choose $t \sim Uniform(\{1, 2, \dots, T\})$

Choose $w \sim \mathcal{N}(0, I)$

$\theta = \theta_{next}$

// Gradient descent for simple loss function given by Eq.(17.55).

$\theta_{next} = \theta + \eta \partial_\theta \left[w - n_\theta^{t-1}(\sqrt{\pi_1^t \alpha} x^0 + \sqrt{1 - \pi_1^t \alpha} w) \right]^2$

$\Delta\theta = \theta_{next} - \theta$

return θ_{next}

Algorithm 3: Algorithm for sampling DM (i.e., finding fake image x^0)

Input : $n_\theta^t(x)$, where θ is optimal, T , $\{\alpha^t\}_{t=1}^T$, $\{\sigma^t\}_{t=1}^T$

Output: x^0 = fake image

Choose $x^T \sim \mathcal{N}(0, I)$

for $t = T, T - 1, \dots, 2, 1$ **do**

Choose $w \sim \mathcal{N}(0, I)$ if $t > 1$, else $w = 0$

// See Eq. 17.33

$$x^{t-1} = \frac{1}{\sqrt{\alpha^t}} \left(x^t - \frac{\beta^t}{\sqrt{1-\pi_1^t \alpha}} n_\theta^{t-1}(x^t) \right) + \sigma^t w$$

return x^0

Chapter 18

Digital Circuits

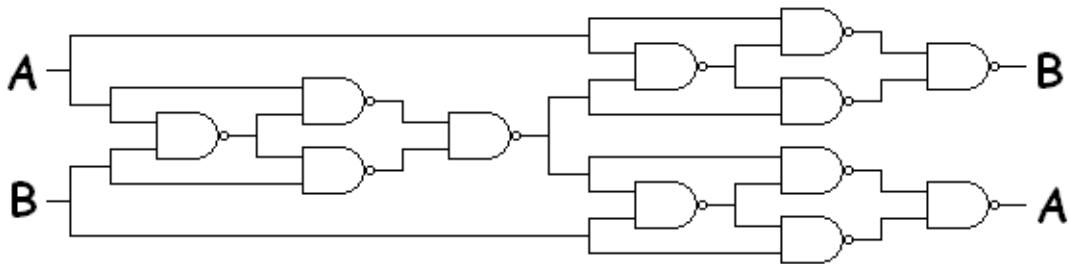


Figure 18.1: Typical digital circuit of NAND gates.

Digital (logic) gate: node with na input ports and nx output ports which represents a function

$$f : \{0, 1\}^{na} \rightarrow \{0, 1\}^{nx}. \quad (18.1)$$

Suppose

$$a^{na} = (a_i)_{i=0,1,\dots,na-1} \text{ where } a_i \in \{0, 1\},$$

$$x^{nx} = (x_i)_{i=0,1,\dots,nx-1} \text{ where } x_i \in \{0, 1\}.$$

f maps a^{na} into x^{nx} .

Digital circuit (dcircuit) = circuit of digital gates.

18.1 Mapping any dcircuit to a bnet

18.1.1 Option A of Fig.18.2

1. Replace every dcircuit gate described by Eq.(18.1) by nx bnet nodes \underline{x}_i for $i = 0, 1, \dots, nx - 1$ such that

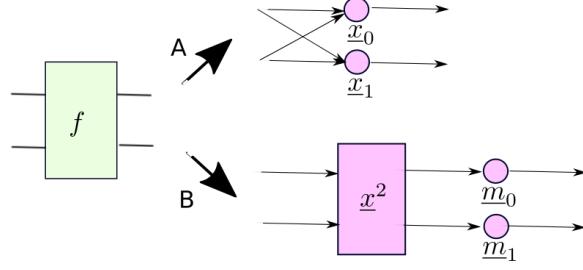


Figure 18.2: 2 options for mapping dcircuit node with multiple output ports into bnet.

$$P(x_i|a^{na}) = \delta(x_i, f_i(a^{na})) \quad (18.2)$$

2. Replace all connectors of the dcircuit by arrows pointing in the direction of the bit flow.

18.1.2 Option B of Fig.18.2

1. Replace every dcircuit gate described by Eq.(18.1) with one bnet node called x^{nx} and, if $nx > 0$, nx “marginalizer nodes” \underline{m}_i for $i = 0, 1, \dots, nx - 1$, such that

$$P(x^{nx}|a^{na}) = \delta(x^{nx}, f(a^{na})) , \quad (18.3)$$

and

$$P(m_i|x^{nx}) = \delta(m_i, x_i) . \quad (18.4)$$

2. Replace all connectors of the dcircuit by arrows pointing in the direction of the bit flow.

Options A and B don't work for digital circuits with feedback loops such as flip-flops. Those could probably be modeled with dynamical bnets.

Chapter 19

Do Calculus

The Do Calculus and associated ideas were invented by Judea Pearl and collaborators. This chapter is based on Judea Pearl's books (see Chapter A).

When doing Do Calculus, it is convenient to separate the nodes of a bnet into 2 types: **observed**, and **hidden** (i.e., **unobserved**, **latent**, **unmeasured**, **non-visible**), depending on whether data describing the state of that node is available (i.e., measured) or not. In this chapter, every hidden node will be indicated in a bnet diagram by either: (1) enclosing its random variable in a dashed circle or (2) making the arrows coming out of it dashed. Accordingly, the 3 diagrams in Fig.19.1 all mean the same thing.

A **confounder node** \underline{c} for nodes \underline{x} and \underline{y} is a root node with arrows pointing from it to both \underline{x} and \underline{y} . Thus, \underline{c} acts as a **common cause** of \underline{x} and \underline{y} . In general, confounders can be either observed or hidden nodes. The word “confounder” itself just means that it confuses the analysis. It says nothing about whether it is hidden or not. The node \underline{c} in Fig.19.1) is a **hidden confounder**.

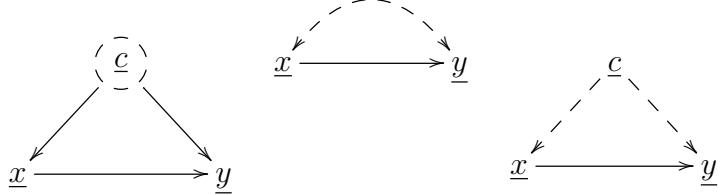


Figure 19.1: These 3 diagrams are equivalent. They mean that node \underline{c} is hidden. Node \underline{c} is implicit in the middle diagram.

Let $\mathcal{D}_{\underline{x}}$ be an operator that acts on a graph G with a node \underline{x} by deleting all the arrows entering \underline{x} , thus converting \underline{x} into a new node $\mathcal{D}\underline{x}$ that is a root node. Let $\mathcal{L}_{\underline{x}}$ be an operator that acts on a graph G with a node \underline{x} by deleting all the arrows leaving \underline{x} , thus converting \underline{x} into a new node $\mathcal{L}\underline{x}$ that is a leaf node. $\mathcal{D}_{\underline{x}}$ and $\mathcal{L}_{\underline{x}}$ are depicted in Fig.19.2. ¹

¹Pearl uses $\mathcal{D}_X G = G_{\overline{X}}$ and $\mathcal{L}_X G = G_{\underline{X}}$ for a random variable X in a graph G . The way I remember Pearl's notation is top-in (as in topping), and bottom-out (as in butt-out).

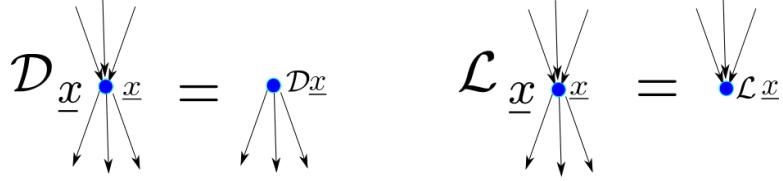


Figure 19.2: The do operator $\mathcal{D}_{\underline{x}}$ converts node \underline{x} into a root node $\mathcal{D}\underline{x}$. The leaf operator $\mathcal{L}_{\underline{x}}$ converts node \underline{x} into a leaf node $\mathcal{L}\underline{x}$.

If you don't know yet what we mean by a multi-node $\underline{a}.$, see Chapter D.

Given a bnet G , we define as follows the operators $\mathcal{D}_{\underline{a}.$ and $\mathcal{L}_{\underline{a}.$ for a multi-node $\underline{a}..$

$$\mathcal{D}_{\underline{a}.}G = \left[\prod_j \mathcal{D}_{\underline{a}_j} \right] G, \quad \mathcal{L}_{\underline{a}.}G = \left[\prod_j \mathcal{L}_{\underline{a}_j} \right] G. \quad (19.1)$$

Consider a bnet whose totality of nodes is labeled $\underline{X}..$. Recall that

$$P(X.) = \prod_j P(X_j | (X_k)_{k: \underline{X}_k \in pa(\underline{X}_j)}). \quad (19.2)$$

Define an operator \mathcal{D} that acts as follows²: Let $X. - a. = (X_k)_{k: \underline{X}_k \notin a..}$

$$P(X. - a. | \mathcal{D}\underline{a}. = a.) = \mathcal{N}(!(\underline{X} - a.)) \frac{P(X.)}{\prod_{j: \underline{X}_j \in a.} P(X_j | (X_k)_{k: \underline{X}_k \in pa(\underline{X}_j)})} \quad (19.3)$$

$$= \mathcal{N}(!(\underline{X} - a.)) \prod_{j: \underline{X}_j \notin a.} P(X_j | (X_k)_{k: \underline{X}_k \in pa(\underline{X}_j)}) \quad (19.4)$$

$$\neq P(X. - a. | \underline{a}. = a.). \quad (19.5)$$

Also,

$$P(\mathcal{D}\underline{a}. = a.) = \delta(a'., a.). \quad (19.6)$$

In words, we replace the TPM for multinode $\underline{a}.$ by a delta function.

For instance, for the bnet

$$\underline{r} \longrightarrow \underline{x} \longrightarrow \underline{y} \quad (19.7)$$

with

$$P(r, x, y) = P(y|x)P(x|r)P(r), \quad (19.8)$$

one has

²As usual, $\mathcal{N}(!x)$ denotes a constant that is independent of x .

$$P(r, y | \mathcal{D}\underline{x} = x) = P(y|x)P(r) \quad (19.9)$$

Hence,

$$P(y | \mathcal{D}\underline{x} = x) = P(y|x) \quad (19.10)$$

For the bnet



with

$$P(x, y, c) = P(y|x, c)P(x|c)P(c), \quad (19.12)$$

one has

$$P(y, c | \mathcal{D}\underline{x} = x) = P(y|x, c)P(c). \quad (19.13)$$

Hence,

$$P(y | \mathcal{D}\underline{x} = x) = \sum_c P(y|x, c)P(c). \quad (19.14)$$

This is called **adjusting the parents of \underline{x}** .

For $b. \subset \underline{X}. - \underline{a}.$, define

$$P(b. | \mathcal{D}\underline{a}. = a.) = \sum_{X. - a. - b.} P(X. - a. | \mathcal{D}\underline{a}. = a.), \quad (19.15)$$

and for $s. \subset \underline{X}. - \underline{a}. - \underline{b}.$, define

$$P(b. | \mathcal{D}\underline{a}. = a., s.) = \frac{P(b., s. | \mathcal{D}\underline{a}. = a.)}{P(s. | \mathcal{D}\underline{a}. = a.)}. \quad (19.16)$$

$P(b. | \mathcal{D}\underline{a}. = a., s.)$ is denoted by Pearl by $P(b. | do(\underline{a}. = a.), s.)$. I prefer to use \mathcal{D} instead of $do()$. I will still call \mathcal{D} a **do operator**.

In $P(y | \mathcal{D}\underline{x} = x)$, node \underline{x} is turned into a root node. This guarantees that there is no confounding node connecting \underline{x} and \underline{y} . Such confounding nodes are unwelcomed when calculating causal effects between the 2 variables \underline{x} and \underline{y} because they introduce non-causal correlations between the two. This is also what happens in a **Randomized Controlled Trial (RCT)**. In an RCT with treatment \underline{x} , the value of \underline{x} for each patient is determined by a coin toss, effectively turning \underline{x} into a root node. Hence, the do operator mimics an RCT.

$P(b | \mathcal{D}\underline{a} = a, s)$ is said to be **do-identifiable** (i.e., expressible without `do()`) if it can be expressed in terms of probability distributions that only depend on observed variables, and that have no do operators in them.³

For $\underline{x}, \underline{y} \in \{0, 1\}$, the **average causal effect (ACE)** is defined as

$$ACE = P(y = 1 | \mathcal{D}\underline{x} = 1) - P(y = 1 | \mathcal{D}\underline{x} = 0) \quad (19.17)$$

and the **Risk Difference (RD)** is defined as

$$RD = P(y = 1 | \underline{x} = 1) - P(y = 1 | \underline{x} = 0). \quad (19.18)$$

19.1 3 Rules of Do Calculus

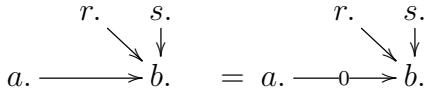
Throughout this section, suppose $\underline{a}, \underline{b}, \underline{r}, \underline{s}$ are disjoint multinodes in a bnet G .

Recall from Chapter 21 on d-separation, that $(\underline{b} \perp_G \underline{a} | \underline{r}, \underline{s})$ means that we have established from the d-separation rules that all paths in G from \underline{a} to \underline{b} are blocked if we condition on $\underline{r} \cup \underline{s}$. Recall also that:

Rule 0: Insertion or deletion of observations, without do operators.

If $(\underline{b} \perp \underline{a} | \underline{r}, \underline{s})$ in G , then

✓ $P(b | \underline{a}, r, s) = P(b | r, s)$ (i.e., $\underline{a} = a \leftrightarrow 1$)



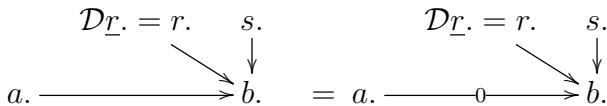
✓ $H(\underline{b} : \underline{a} | \underline{r}, \underline{s}) = 0$

Zeroing an arrow is the same as deleting it. The 3 rules of Do Calculus can be presented in the same format.

- **Rule 1:** Insertion or deletion of observations

If $(\underline{b} \perp \underline{a} | \underline{r}, \underline{s})$ in $\mathcal{D}_{\underline{r}}G$, then

✓ $P(b | \underline{a}, \mathcal{D}\underline{r} = r, s) = P(b | \mathcal{D}\underline{r} = r, s)$ (i.e., $\underline{a} = a \leftrightarrow 1$)



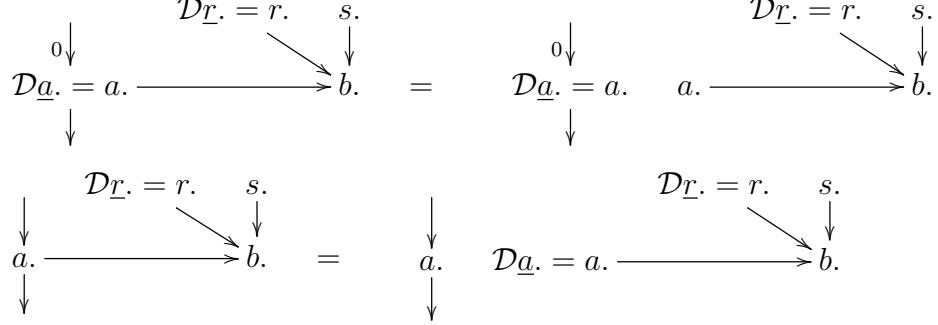
✓ $H(\underline{b} : \underline{a} | \mathcal{D}\underline{r}, \underline{s}) = 0$

- **Rule 2:** Action or observation exchange

If $(\underline{b} \perp \underline{a} | \underline{r}, \underline{s})$ in $\mathcal{L}_{\underline{a}}\mathcal{D}_{\underline{r}}G$, then

³In Statistics, one says a probability distribution $P(x; \theta)$ of x that depends on a parameter θ is **identifiable** if $P(x; \theta_1) = P(x; \theta_2)$ implies $\theta_1 = \theta_2$.

$$\checkmark P(b | \mathcal{D}\underline{a}_. = a., \mathcal{D}\underline{r}_. = r., s.) = P(b | a., \mathcal{D}\underline{r}_. = r., s.) \quad (\text{i.e., } \mathcal{D}\underline{a}_. = a. \leftrightarrow \underline{a}_. = a.)$$



In this rule, the node is split into a $\mathcal{D}\underline{a}_. = a.$ node and a $a.$ node. The original node keeps the arrows, and the new node is a root node.

$$\checkmark H(\underline{b}_. : \mathcal{D}\underline{a}_. | \mathcal{D}\underline{r}_., \underline{s}_.) = H(\underline{b}_. : \underline{a}_. | \mathcal{D}\underline{r}_., \underline{s}_.)$$

- **Rule 3:** Insertion and deletion of actions

If $(\underline{b}_. \perp \underline{a}_. | \underline{r}_., \underline{s}_.)$ in $\mathcal{D}_{\underline{a}_.-an(\underline{s}_.)} \mathcal{D}\underline{r}_. G$, then

$$\checkmark P(b | \mathcal{D}\underline{a}_. = a., \mathcal{D}\underline{r}_. = r., s.) = P(b | \mathcal{D}\underline{r}_. = r., s.) \quad (\text{i.e., } \mathcal{D}\underline{a}_. = a. \leftrightarrow 1)$$



$$\checkmark H(\underline{b}_. : \mathcal{D}\underline{a}_. | \mathcal{D}\underline{r}_., \underline{s}_.) = 0$$

See Fig.19.3 for a pictorial representation of these rules.

These rules have been proven to be sufficient for removing all do operators from an expression for which it is possible to do so.

Next we discuss two formulae that can be proven using Do Calculus: the backdoor and the frontdoor adjustment formulae.

The backdoor formula adjusts one multinode and the frontdoor formula adjusts two.

19.2 Parent Adjustment Formula

Suppose that $\underline{x}_., \underline{y}_., \underline{z}_.$ are disjoint multinodes and their union equals the totality of all nodes of a bnet. Suppose we have data available that allows us to estimate $P(\underline{x}_., \underline{y}_., \underline{z}_.).$ Hence, all nodes of the bnet are observable. Furthermore, suppose $\underline{z}_. = pa(\underline{x}_.).$ In other words, we are considering the bnet



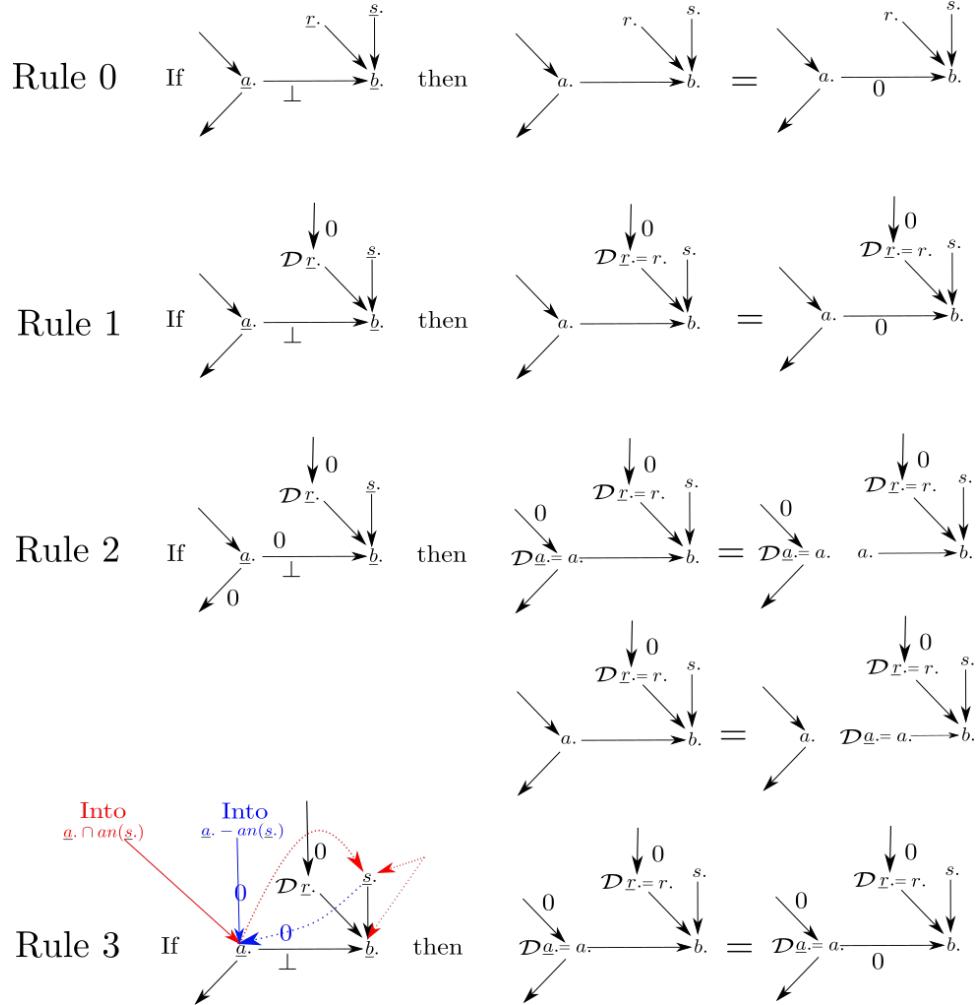


Figure 19.3: Pictorial representation of the rules of Do Calculus. $an(\underline{s})$ stands for the ancestors of \underline{s} . In Rule 2, we zero the arrows leaving \underline{a} . and ascertain that there is no flow of info between \underline{a} . and \underline{b} . under those circumstances. In Rule 3, as shown by the dotted arrows, conditioning on \underline{s} . is enough to block paths going from \underline{b} . into $\underline{a} - an(\underline{s})$, but it is not sufficient to block paths going from \underline{b} . into $\underline{a} \cap an(\underline{s})$. In the example with red dotted arrows, conditioning on \underline{s} . opens a path between \underline{a} . and \underline{b} . in which \underline{s} . is a collider. That is the reason for zeroing arrows going into $\underline{a} \cap an(\underline{s})$, but not zeroing arrows going into $\underline{a} \cup an(\underline{s})$.

Then

$$P(y., z.| \mathcal{D}\underline{x}. = x.) = P(y.| x., z.)P(z.) \quad (19.20)$$

so

$$P(y.|D\underline{x}. = x.) = \sum_{z.} P(y.|x., z.)P(z.). \quad (19.21)$$

This is called **adjusting the parents** of $\underline{x}.$.

We say that we are **adjusting or controlling a node \underline{a}** if we condition a probability on \underline{a} and then we average that probability over \underline{a} . More generally, we can adjust a whole multinode \underline{a} together.

Next, we will introduce a generalization of this parent adjustment formula called the backdoor adjustment formula. In a backdoor adjustment formula, the adjusted multinode is not necessarily the parents of $\underline{x}.$.

19.3 Backdoor Adjustment Formula

See Chapter 4 for examples of the use of the backdoor adjustment formula. In this section, we shall mainly be concerned with proving this theorem using Do Calculus.

For any two disjoint multinodes $\underline{x}.$ and $\underline{y}.$, we define a **backdoor path** from $\underline{x}.$ to $\underline{y}.$ as a path from $\underline{x}.$ and $\underline{y}.$ that starts with an arrow pointing into $\underline{x}.$,

Suppose that we have access to data that allows us to estimate a probability distribution $P(x., y., z.).$ Hence, the variables $\underline{x}., \underline{y}., \underline{z}.$ are ALL observed (i.e, not hidden). Then we say that the backdoor $\underline{z}.$ satisfies the **backdoor adjustment criterion** relative to $(\underline{x}., \underline{y}.)$ if

1. All backdoor paths from $\underline{x}.$ to $\underline{y}.$ are blocked by conditioning on $\underline{z}..$
2. $\underline{z}_. \cap de(\underline{x}.) = \emptyset.$

Motivation for BD criterion: Part 1 rules out paths from \underline{x} to \underline{y} containing a fork node (confounder) which, if not blocked by conditioning on $\underline{z}.$, would introduce a non-causal correlation (confounder bias). Part 2 rules out a directed path from \underline{x} to \underline{y} that has a mediator node blocked by conditioning on $\underline{z}.$ or a collider node unblocked by conditioning on $\underline{z}..$

Claim 31 (*Backdoor Adjustment Formula*)

If $\underline{z}.$ satisfies the backdoor criterion relative to $(\underline{x}., \underline{y}.)$, then

$$P(y.|D\underline{x}. = x.) = \sum_{z.} P(y.|x., z.)P(z.) \quad (19.22)$$

$$= \sum_{z.} \begin{array}{c} \nearrow \\ z. \end{array} \quad (19.23)$$

$x. \longrightarrow y.$

where $\sum z.$ means node $\underline{z}.$ is summed over.

proof:

For simplicity, let us omit the dots from the multinodes. If z satisfies the backdoor criterion relative to $(\underline{x}, \underline{y})$, then $\underline{x}, \underline{y}, \underline{z}$ might have the following structure.



See Claim 35 for a proof of this claim for the special case Eq.(19.24).

QED

Note that the backdoor adjustment formula can be written as

$$P(y | \mathcal{D}\underline{x} = x.) = \sum_{z.} P(y | x., z.) P(z.) \quad (19.25)$$

$$= \sum_{z.} \frac{P(y., x., z.)}{P(x.|z.)} \quad (19.26)$$

This assumes $P(x.|z.) \neq 0$ for all $x., z..$ This assumption is referred to as **positivity**, and is violated if $P(x.|z.) = \delta(x., x.(z.))$. $P(x.|z.)$ is called the **propensity score** of $x.$ given $z..$ This equation does **inverse probability weighting**. One can approximate $P(x.|z.)$ in this equation to get an approximation to $P(y | \mathcal{D}\underline{x} = x)$.

19.4 Frontdoor Adjustment Formula

See Chapter 28 for examples of the use of the frontdoor adjustment formula. In this section, we shall mainly be concerned with proving this theorem using Do Calculus.

Suppose that we have access to data that allows us to estimate a probability distribution $P(x., m., y.)$. Hence, the variables $\underline{x}, \underline{m}, \underline{y}$ are ALL observed (i.e, not hidden). Then we say that the frontdoor \underline{m} . satisfies the **frontdoor adjustment criterion** relative to $(\underline{x}, \underline{y})$ if

1. All directed paths from \underline{x} . to \underline{y} . are intercepted by (i.e., have a node in) $\underline{m}..$
2. All backdoor paths from \underline{x} . to \underline{m} . are blocked.
3. All backdoor paths from on \underline{m} . to \underline{y} . are blocked by conditioning on $\underline{x}..$

Claim 32 (Frontdoor Adjustment Formula)

If \underline{m} . satisfies the frontdoor criterion relative to $(\underline{x}, \underline{y})$, and $P(x., m.) > 0$, then

$$P(y_{\cdot} | \mathcal{D}\underline{x}_{\cdot} = x_{\cdot}) = \sum_{m_{\cdot}} \underbrace{\left[\sum_{x'_{\cdot}} P(y_{\cdot} | x'_{\cdot}, m_{\cdot}) P(x'_{\cdot}) \right]}_{P(y_{\cdot} | \mathcal{D}\underline{m}_{\cdot} = m_{\cdot})} \underbrace{P(m_{\cdot} | x_{\cdot})}_{P(m_{\cdot} | \mathcal{D}\underline{x}_{\cdot} = x_{\cdot})} \quad (19.27)$$

$$= \sum_{x'_{\cdot}} \begin{array}{c} \text{---} \\ x_{\cdot} \longrightarrow \sum m_{\cdot} \longrightarrow y_{\cdot} \end{array} \quad (19.28)$$

where $\sum x'_{\cdot}$. and $\sum m_{\cdot}$ means nodes \underline{x}'_{\cdot} . and \underline{m}_{\cdot} are summed over.

proof:

For simplicity, let us omit the dots from the multinodes. If \underline{m} satisfies the frontdoor criterion relative to $(\underline{x}, \underline{y})$, then $\underline{x}, \underline{m}, \underline{y}$ might have the following structure, where node \underline{c} is unobserved.



See Claim 36 for a proof of this claim for the special case Eq.(19.29).

See also Ref.[41] for a proof by Pearl of the Frontdoor Adjustment Formula without using Do Calculus.

QED

19.5 Comparison of Backdoor and Frontdoor adjustment formulae

Define a **direct effect path** for a query $P(y | \mathcal{D}\underline{x} = x, z_{\cdot})$ as a directed path that starts at \underline{x} and ends at \underline{y} . A backdoor path (i.e., one that connects \underline{x} and \underline{y} starting with an arrow pointing into \underline{x}), is not a direct effect path; it's an **indirect effect path**.

Note that in the backdoor AF (adjustment formula), we can find a possibly empty observed multinode \underline{z}_{\cdot} such that if we condition on \underline{z}_{\cdot} , all indirect effect paths are blocked. In the frontdoor AF, we can't find a multinode \underline{z}_{\cdot} that blocks all indirect effect paths. Despite this, in the frontdoor scenario, the do-query is identifiable and an adjustment formula exists. How is that possible? The frontdoor AF uses the backdoor AF once and then it uses the backdoor AF again, a second time, on the result of the first use. The frontdoor AF replaces a sum over an unobserved node by a sum over an observed one.

19.6 Do operator for DEN diagrams

Recall that the structural equations for a linear DEN, as given by Eq.(44.21) of Chapter 44, are:

$$\underline{x} = A\underline{x} + \underline{u}. \quad (19.30)$$

Therefore,

$$\underline{x} = (1 - A)^{-1}\underline{u} \quad (19.31)$$

which can be represented for both linear and non-linear DEN diagrams by:

$$\underline{x}_i = x_i(\underline{u}) \quad (19.32)$$

If now we apply the operator $\mathcal{D}_{\underline{a}=\underline{a}}$ to the diagram described by the structural equations Eqs.19.30, we get the following new structural equations:

$$\underline{x}_i^* = \begin{cases} \sum_{j < i} A_{i|j} \underline{x}_j^* + \underline{u}_i & \text{if } \underline{x}_i \neq \underline{a} \\ \underline{a} & \text{if } \underline{x}_i = \underline{a} \end{cases}, \quad (19.33)$$

where we are calling \underline{x}_i^* the nodes of the DEN diagram post intervention.

Eqs.(19.33) can be expressed in matrix notation as follows. Define $\pi_{\underline{a}}$ to be the $nx \times nx$ matrix with all entries equal to zero except for the (i_0, i_0) entry, which is 1. And define $e_{\underline{a}}$ to be the column vector with all entries zero except for the i_0 'th one, which is 1. Here i_0 is defined so that $\underline{x}_{i_0} = \underline{a}$. In other words, $\pi_{\underline{a}}$ and $e_{\underline{a}}$ are defined by

$$(\pi_{\underline{a}})_{i,j} = \mathbb{1}(i = j, \underline{a} = \underline{x}_i) \quad (19.34)$$

and

$$(e_{\underline{a}})_i = \mathbb{1}(\underline{a} = \underline{x}_i), \quad (19.35)$$

for $i, j \in \{0, 1, \dots, nx - 1\}$. Next define

$$\pi_{!\underline{a}} = 1 - \pi_{\underline{a}}, \quad (19.36)$$

$$A^* = \pi_{!\underline{a}} A, \quad (19.37)$$

and

$$\underline{u}_{!\underline{a}} = \pi_{!\underline{a}} \underline{u}. \quad (19.38)$$

The effect of pre-multiplying the matrix A and the column vector \underline{u} by $\pi_{!\underline{a}}$ is to leave all rows intact except for the i_0 row, which is set to zero. Here i_0 is defined by $\underline{a} = \underline{x}_{i_0}$.

Finally, using all of the variables just defined, we can express the structural equations of the linear DEN diagram, post intervention, as

$$\underline{x}^* = A^* \underline{x}^* + \underline{u}_{!a} + ae_a . \quad (19.39)$$

Thus,

$$\underline{x}^* = (1 - A^*)^{-1}(\underline{u}_{!a} + ae_a) . \quad (19.40)$$

which can be represented for both linear and non-linear DEN diagrams by:

$$\underline{x}_i^* = x_i^*(\underline{u}_{!a}, a) . \quad (19.41)$$

For any bnet,

$$P(\underline{y} = \underline{y} | \underline{x} = \underline{x}) = P_G(\underline{y} = \underline{y} | \underline{x} = \underline{x}) \quad (19.42)$$

$$P(\underline{y} = \underline{y} | \mathcal{D}\underline{x} = \underline{x}) = P_{\mathcal{D}_{\underline{x}=\underline{x}}G}(\underline{y} = \underline{y}) \quad (19.43)$$

Claim 33 *For a non-linear DEN diagram,*

$$P(\underline{y} | \mathcal{D}\underline{x} = \underline{x}) = E [\delta[\underline{y}, \underline{y}(\underline{u}_{!x}, \underline{x})]] . \quad (19.44)$$

proof:

$$P(\underline{y} = \underline{y} | \mathcal{D}\underline{x} = \underline{x}) = P_{\mathcal{D}_{\underline{x}=\underline{x}}G}(\underline{y} = \underline{y}) \quad (19.45)$$

$$= \sum_{u_{!x}} P(u_{!x}) P_{\mathcal{D}_{\underline{x}=\underline{x}}G}(\underline{y} = \underline{y} | u_{!x}) \quad (19.46)$$

$$= \sum_{u_{!x}} P(u_{!x}) \delta[\underline{y}, \underline{y}(u_{!x}, \underline{x})] \quad (19.47)$$

$$= E_{u_{!x}} [\delta[\underline{y}, \underline{y}(u_{!x}, \underline{x})]] \quad (19.48)$$

$$= E[\delta[\underline{y}, \underline{y}(u_{!x}, \underline{x})]] \quad (19.49)$$

QED

Claim 34 *For a nonlinear DEN diagram,*

$$E[\underline{y} | \mathcal{D}\underline{x} = \underline{x}] = E[\underline{y}(\underline{u}_{!x}, \underline{x})] . \quad (19.50)$$

proof:

$$E[\underline{y}|\mathcal{D}\underline{x} = x] = \sum_y y P(\underline{y} = y|\mathcal{D}\underline{x} = x) \quad (19.51)$$

$$= \sum_y y E[\delta[y, y(u_{!\underline{x}}, x)]] \quad (19.52)$$

$$= E[y(u_{!\underline{x}}, x)] \quad (19.53)$$

QED

For any bnet

$$P(y|\mathcal{D}\underline{x} = x, z) = \frac{P(y, z|\mathcal{D}\underline{x} = x)}{P(z|\mathcal{D}\underline{x} = x)} = P_{\mathcal{D}_{\underline{x}=x}G}(y|x, z) \quad (19.54)$$

For a nonlinear DEN diagram,

$$P(y, z|\mathcal{D}\underline{x} = x) = \sum_{u_{!\underline{x}}} P(u_{!\underline{x}}) \delta[y, y(u_{!\underline{x}}, x)] \delta[z, z(u_{!\underline{x}}, x)] \quad (19.55)$$

$$P(z|\mathcal{D}\underline{x} = x) = \sum_{u_{!\underline{x}}} P(u_{!\underline{x}}) \delta[z, z(u_{!\underline{x}}, x)]. \quad (19.56)$$

Chapter 20

Do Calculus proofs

In Chapter 19, we explained Do Calculus, but referred to this chapter for proofs of claims that use Do Calculus. In this chapter, we've aggregated all proofs, from throughout the book, of claims that use Do Calculus.

Note that even though the 3 rules of Do Calculus are great for proving adjustment formulae for general classes of DAGs, they are sometimes overkill for proving adjustment formulae for a single specific DAG. Indeed, since the 3 rules of Do Calculus are a consequence of the d-separation theorem, it follows that all adjustment formulae should be provable from first principles, assuming only the d-separation theorem and the standard rules of probability theory.

In this chapter, we use the following conventions for bnet diagrams.

Random variables are underlined and their values are not. For example, $\underline{a} = a$ means the random variable \underline{a} takes the value a . A diagram with all its nodes underlined represents a Bayesian Network (bnet), whereas the same diagram with the letters not underlined represents a specific **instantiation** of that bnet. For example $\underline{a} \rightarrow \underline{b} \rightarrow \underline{c}$ represents the bnet with full probability distribution $P(c|b)P(b|a)P(a)$, whereas $a \rightarrow b \rightarrow c$ represents $P(c|b)P(b|a)$. Note that, for convenience, we define $a \rightarrow b \rightarrow c$ to exclude the priors of root nodes such as $P(a)$.

If \underline{a} is a root node, then $\sum a$ signifies a weighted sum $\sum_a P(a)$. For example,

$$\sum a \rightarrow b \rightarrow c = \sum_a P(c|b)P(b|a)P(a) \quad (20.1)$$

If \underline{a} is not a root node, then $\sum a$ signifies a simple unweighted sum \sum_a . For example,

$$x \rightarrow \sum a \rightarrow y = \sum_a P(y|a)P(a|x) \quad (20.2)$$

Two bnets are equated if their full probability distributions (i.e., their full instantiations) are equal numerically. For example,

$$\underline{a} \rightarrow \underline{b} \rightarrow \underline{c} = P(c|b)P(b|a)P(a) = \underline{a} \leftarrow \underline{b} \leftarrow \underline{c} \quad (20.3)$$

Unobserved nodes are indicated by enclosing them in a dashed circle. For example, $\langle u \rangle$.

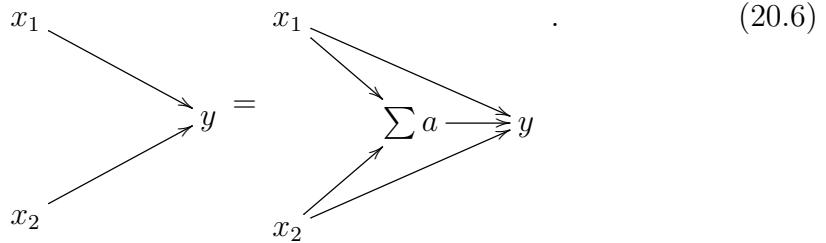
Selection diagrams with selection nodes are discussed in Chapter 86. In a selection diagram with a selection node $s \in \{0, 1\}$, if a node \underline{x} has parents $pa(\underline{x})$ where $s \notin pa(\underline{x})$, then the TPM of \underline{x} is $P(x|pa(\underline{x}))$. If, on the other hand, \underline{x} has parents $pa(\underline{x}) = pa'(\underline{x}) \cup s$, where $pa'(\underline{x}) = pa(\underline{x}) - s$, then the TPM of \underline{x} is

$$P(x|pa'(\underline{x}), s) = \begin{cases} P(x|pa'(\underline{x})) & \text{if } s = 0 \\ P^*(x|pa'(\underline{x})) & \text{if } s = 1 \end{cases} \quad (20.4)$$

Some identities that are used in this chapter:

1.

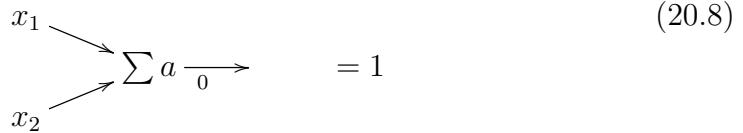
$$P(y|x_1, x_2) = \sum_a P(y|a, x_1, x_2)P(a|x_1, x_2). \quad (20.5)$$



One can describe this identity as “giving y a universal backdoor”, because $\sum a$ is a backdoor (i.e., input) to y , and $\sum a$ is universal in the sense that it is entered by every arrow that enters y except $\sum a$ itself.

2.

$$\sum_a P(a|x_1, x_2) = 1 \quad (20.7)$$



One can describe this identity as “summing over the values of a collider node which has no emerging arrows”¹. Eq.(20.8) can be understood as an edge case (when $y = \emptyset$) of Eq.(20.6).

3.

$$\sum_a P(x_2|a)P(a|x_1) = P(x_2|x_1) \quad (20.9)$$

¹A zeroed arrow means the same as no arrow.

$$x_1 \longrightarrow \sum a \longrightarrow x_2 = x_1 \longrightarrow x_2 \quad (20.10)$$

One can describe this identity as “summing over the values of a mediator node”.

4.

$$P(x) = \sum_a P(x|a)P(a) = \sum_b P(x|b)P(b) \quad (20.11)$$

$$P(x) = \longrightarrow_0 \sum a \longrightarrow x = \longrightarrow_0 \sum b \longrightarrow x \quad (20.12)$$

One can describe this identity as “averaging over different priors”. Eq.(20.12) can be understood as an edge case of Eq.(20.10).

A **do-adjustment formula** expresses a **do-query** (i.e., a conditional probability with do operators in its condition) by an equivalent expression without do operators. The equivalent expression must satisfy 2 constraints to be discussed below. If a do-adjustment formula exists for a particular do-query, then we say the do-query is **do-identifiable (DI)**. A **do-transport formula** is a relationship between 2 do-queries. This chapter deals with both do-adjustment and do-transport formulae.

See Fig.20.1 for some simple examples of bnets for which the do-query $P(y|\mathcal{D}\underline{x} = x)$ is DI and non-DI.

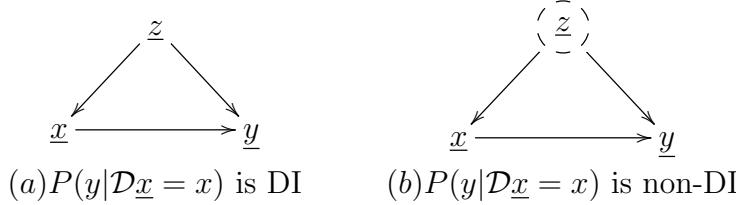


Figure 20.1: Examples of bnets for which the do-query $P(y|\mathcal{D}\underline{x} = x)$ is DI and non-DI.

In general, the following is always true, whether it applies to a bnet with or without hidden nodes:²

$$P(y|\mathcal{D}\underline{x} = x) = P(y|x) \quad (20.13)$$

However, the right hand side of this equation is seldom a valid adjustment formula for this query. We define a valid adjustment formula for query $P(y|\mathcal{D}\underline{x} = x)$ to be a bnet instantiation diagram that satisfies the following 2 constraints: Call the original bnet G_0 .

²Note that $P(y|\mathcal{D}\underline{x} = x) = P(y|x) \neq P_0(y|x)$, where we are using $P_0()$ (respectively, $P()$) to denote the full probability distribution for the bnet before (respectively, after) amputation of the arrows entering node \underline{x} . In fact, $P(y|\mathcal{D}\underline{x} = x) = P_0(y|x)$ iff there is no confounding, so $P(y|\mathcal{D}\underline{x} = x) \neq P_0(y|x)$ can be used to detect confounding.

1. (structural constraint)

The adjustment formula must be representable by a bnet instantiation that has a DAG structure identical to the DAG structure of G_0 , except that arrows entering node \underline{x} have been amputated. All nodes of that instantiation, except nodes x and y , must be summed over.

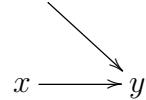
2. (probabilistic constraint)

- If G_0 has hidden nodes, these must be renamed and assigned a TPM that can be constructed from the *observable* TPMs of G_0 .
- The observable nodes of G_0 with hidden parents, must also be assigned a TPM that can be constructed from the *observable* TPMs of G_0 .
- The observable nodes of G_0 with no hidden parents, must be assigned the same TPM as they have in G_0 .

The reason for these 2 constraints is that we want an adjustment formula to utilize as much information as possible from the original bnet G_0 . Adjustment formulae are always numerically equal to $P(y|x)$, but when you effect that reduction, you throw away a lot of valuable information carried by G_0 .

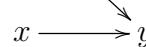
Based on these 2 constraints, we can easily see why the query $P(y|\mathcal{D}\underline{x} = x)$ is DI (resp., non-DI) for bnet (a) (resp., bnet (b)) of Fig.20.1. For bnet (a), after amputating arrow $z \rightarrow x$ and summing over node z , we get

$$P(y|\mathcal{D}\underline{x} = x) = \sum z \quad (20.14)$$



The right hand side of Eq.(20.14) is a valid adjustment formula because it satisfies both constraints. For bnet (b), if we amputate arrow $z \rightarrow x$ and sum over node z , we get

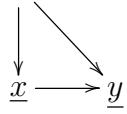
$$P(y|\mathcal{D}\underline{x} = x) = \left(\sum z \right) \quad (20.15)$$



The right hand side of Eq.(20.15) is not a valid adjustment formula because it violates the second constraint. Furthermore, try as we may, there is no way to replace the sum over hidden node z by a sum over an observed node, such that constraint 2 is satisfied.

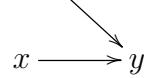
Claim 35 (*Backdoor Adjustment Formula*)

If \underline{z} then



$$P(y|\mathcal{D}\underline{x} = \underline{x}) = \sum_z P(y|x, z)P(z) \quad (20.16)$$

$$= \sum_z \begin{array}{c} z \\ \searrow \\ x \longrightarrow y \end{array} \quad (20.17)$$



proof:

* **proof 1:**

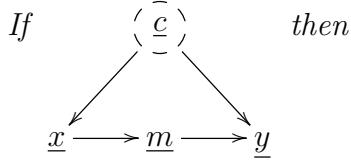
$$P(y|\mathcal{D}\underline{x} = \underline{x}) = \sum_z \begin{array}{c} z \\ \searrow \\ x \longrightarrow y \end{array} \quad (20.18)$$

* **proof 2:**

$$\begin{aligned} P(y|\mathcal{D}\underline{x} = \underline{x}) &= \sum_z P(y|\mathcal{D}\underline{x} = \underline{x}, z)P(z|\mathcal{D}\underline{x} = \underline{x}) \\ &\quad \text{by Probability Axioms} \\ &= \sum_z P(y|x, z)P(z|\mathcal{D}\underline{x} = \underline{x}) \\ &\quad P(y|\mathcal{D}\underline{x} = \underline{x}, z) \rightarrow P(y|x, z) \\ &\quad \text{by Rule 2: If } (\underline{b} \perp \underline{a} | \underline{r}, \underline{s}) \text{ in } \mathcal{L}_{\underline{a}} \mathcal{D}_{\underline{r}} G, \text{ then} \\ &\quad \mathcal{D}_{\underline{a}} = a \leftrightarrow \underline{a} = a. \\ &\quad \underline{y} \perp \underline{x} | \underline{z} \text{ in } \mathcal{L}_{\underline{x}} \mathcal{D}_{\emptyset} G : \begin{array}{c} \underline{z} \\ \downarrow \\ \underline{x} \end{array} \longrightarrow \underline{y} \\ &= \sum_z P(y|x, z)P(z) \\ &\quad P(z|\mathcal{D}\underline{x} = \underline{x}) \rightarrow P(z) \\ &\quad \text{by Rule 3: If } (\underline{b} \perp \underline{a} | \underline{r}, \underline{s}) \text{ in } \mathcal{D}_{\underline{a} - an(\underline{s})} \mathcal{D}_{\underline{r}} G, \text{ then} \\ &\quad \mathcal{D}_{\underline{a}} = a \leftrightarrow 1 \\ &\quad \underline{z} \perp \underline{x} \text{ in } \mathcal{D}_{\underline{x}} \mathcal{D}_{\emptyset} G : \begin{array}{c} \underline{z} \\ \searrow \\ \underline{x} \longrightarrow \underline{y} \end{array} \end{aligned}$$

QED

Claim 36 (*Frontdoor Adjustment Formula*)



$$P(y|\mathcal{D}\underline{x} = x) = \sum_m \left[\sum_{x'} P(y|x', m) P(x') \right] P(m|x) \quad (20.19)$$

$$= \sum_{x'} \text{---} \quad (20.20)$$

$$x \longrightarrow \sum m \longrightarrow y$$

proof:

* **proof 1:**

$$P(y|\mathcal{D}\underline{x} = x) = \sum_{\substack{\sum c \\ -}} \text{---} \quad (20.21)$$

$$= \sum_{x'} \sum_{\substack{\sum c \\ -}} \text{---} \quad (20.22)$$

$$= \sum_{x'} \text{---} \quad (20.23)$$

$$x \longrightarrow \sum m \longrightarrow y$$

* **proof 2:**

$$\begin{aligned} P(y|\mathcal{D}\underline{x} = x) &= \sum_m P(y|\mathcal{D}\underline{x} = x, m) P(m|\mathcal{D}\underline{x} = x) \\ &\quad \text{by Probability Axioms} \\ &= \sum_m P(y|\mathcal{D}\underline{x} = x, \mathcal{D}\underline{m} = m) P(m|\mathcal{D}\underline{x} = x) \\ &\quad P(y|\mathcal{D}\underline{x} = x, m) \rightarrow P(y|\mathcal{D}\underline{x} = x, \mathcal{D}m = m) \\ &\quad \text{by Rule 2: If } (\underline{b.} \perp \underline{a.} | \underline{r.}, \underline{s.}) \text{ in } \mathcal{L}_{\underline{a.}} \mathcal{D}_{\underline{r.}} G, \text{ then} \\ &\quad \mathcal{D}\underline{a.} = a. \leftrightarrow \underline{a.} = a. \end{aligned}$$

$$\begin{aligned}
& \underline{y} \perp \underline{m} | \underline{x} \text{ in } \mathcal{L}_{\underline{m}} \mathcal{D}_{\underline{x}} G : \quad \left(\frac{\underline{c}}{\underline{a}} \right) \\
& \qquad \qquad \qquad \swarrow \qquad \searrow \\
& \qquad \qquad \qquad \underline{x} \longrightarrow \underline{m} \qquad \underline{y} \\
& = \sum_m P(y | \mathcal{D}\underline{x} = \underline{x}, \mathcal{D}\underline{m} = m) P(m | \underline{x}) \\
& \qquad \qquad \qquad P(m | \mathcal{D}\underline{x} = \underline{x}) \rightarrow P(m | \underline{x}) \\
& \text{by Rule 2: If } (\underline{b} \perp \underline{a} | \underline{r}, \underline{s}) \text{ in } \mathcal{L}_{\underline{a}} \mathcal{D}_{\underline{r}} G, \text{ then} \\
& \qquad \qquad \qquad \mathcal{D}\underline{a} = a \leftrightarrow \underline{a} = a.
\end{aligned}$$

$$\begin{aligned}
& \underline{m} \perp \underline{x} \text{ in } \mathcal{L}_{\underline{x}} \mathcal{D}_{\emptyset} G : \quad \left(\frac{\underline{c}}{\underline{a}} \right) \\
& \qquad \qquad \qquad \swarrow \qquad \searrow \\
& \qquad \qquad \qquad \underline{x} \qquad \underline{m} \longrightarrow \underline{y} \\
& = \sum_m P(y | \mathcal{D}\underline{m} = m) P(m | \underline{x}) \\
& \qquad \qquad \qquad P(y | \mathcal{D}\underline{x} = \underline{x}, \mathcal{D}\underline{m} = m) \rightarrow P(y | \mathcal{D}\underline{m} = m) \\
& \text{by Rule 3: If } (\underline{b} \perp \underline{a} | \underline{r}, \underline{s}) \text{ in } \mathcal{D}_{\underline{a}-an(\underline{s})} \mathcal{D}_{\underline{r}} G, \text{ then} \\
& \qquad \qquad \qquad \mathcal{D}\underline{a} = a \leftrightarrow 1
\end{aligned}$$

$$\begin{aligned}
& \underline{y} \perp \underline{x} | \underline{m} \text{ in } \mathcal{D}_{\underline{x}} \mathcal{D}_{\underline{m}} G : \quad \left(\frac{\underline{c}}{\underline{a}} \right) \\
& \qquad \qquad \qquad \swarrow \qquad \searrow \\
& \qquad \qquad \qquad \underline{x} \qquad \underline{m} \longrightarrow \underline{y} \\
& = \sum_{x'} \sum_m P(y | \mathcal{D}\underline{m} = m, x') P(x' | \mathcal{D}\underline{m} = m) P(m | \underline{x}) \\
& \qquad \qquad \qquad \text{by Probability Axioms} \\
& = \sum_{x'} \sum_m P(y | m, x') P(x' | \mathcal{D}\underline{m} = m) P(m | \underline{x}) \\
& \qquad \qquad \qquad P(y | \mathcal{D}\underline{m} = m, x') \rightarrow P(y | m, x') \\
& \text{by Rule 2: If } (\underline{b} \perp \underline{a} | \underline{r}, \underline{s}) \text{ in } \mathcal{L}_{\underline{a}} \mathcal{D}_{\underline{r}} G, \text{ then} \\
& \qquad \qquad \qquad \mathcal{D}\underline{a} = a \leftrightarrow \underline{a} = a.
\end{aligned}$$

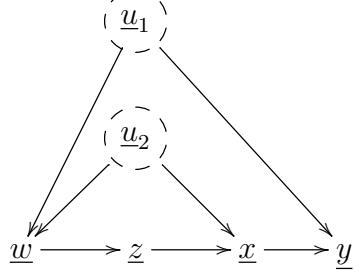
$$\begin{aligned}
& \underline{y} \perp \underline{m} | \underline{x} \text{ in } \mathcal{L}_{\underline{m}} \mathcal{D}_{\emptyset} G : \quad \left(\frac{\underline{c}}{\underline{a}} \right) \\
& \qquad \qquad \qquad \swarrow \qquad \searrow \\
& \qquad \qquad \qquad \underline{x} \longrightarrow \underline{m} \qquad \underline{y} \\
& = \sum_{x'} \sum_m P(y | m, x') P(x') P(m | \underline{x}) \\
& \qquad \qquad \qquad P(x' | \mathcal{D}\underline{m} = m) \rightarrow P(x') \\
& \text{by Rule 3: If } (\underline{b} \perp \underline{a} | \underline{r}, \underline{s}) \text{ in } \mathcal{D}_{\underline{a}-an(\underline{s})} \mathcal{D}_{\underline{r}} G, \text{ then} \\
& \qquad \qquad \qquad \mathcal{D}\underline{a} = a \leftrightarrow 1
\end{aligned}$$

$$\begin{aligned}
& \underline{x} \perp \underline{m} \text{ in } \mathcal{D}_{\underline{m}} \mathcal{D}_{\emptyset} G : \quad \left(\frac{\underline{c}}{\underline{a}} \right) \\
& \qquad \qquad \qquad \swarrow \qquad \searrow \\
& \qquad \qquad \qquad \underline{x} \qquad \underline{m} \longrightarrow \underline{y}
\end{aligned}$$

QED

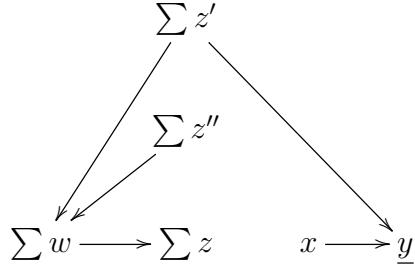
Claim 37 (*Napkin problem from Ref.[50]*)

If $\begin{array}{c} \text{---} \\ | \\ \text{---} \\ | \\ \text{---} \end{array}$ then



$$P(y|\mathcal{D}\underline{x} = x) = \sum_{w,z,z',z''} P(y|x,z')P(w|z')P(z')P(w|z'')P(z'')P(z|w) \quad (20.24)$$

$$= \sum_{\sum z'} \sum_{\sum z''} \quad (20.25)$$



proof:

$$P(y|\mathcal{D}\underline{x} = \underline{x}) = \begin{array}{c} \sum u_1 \\ \sum u_2 \\ \sum w \end{array} \xrightarrow{\quad} \underline{y} \quad (20.26)$$

$$= \begin{array}{c} \sum u_1 \\ \sum u_2 \\ \sum z' \end{array} \xleftarrow{\quad} \begin{array}{c} \sum w \xrightarrow{\quad} \sum z \\ \sum z' \end{array} \quad (20.27)$$

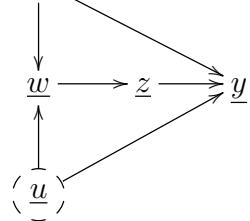
$$= \begin{array}{c} \sum u_1 \\ \sum u_2 \\ \sum z'' \end{array} \xleftarrow{\quad} \begin{array}{c} \sum w \xrightarrow{\quad} \sum z \\ \sum z'' \end{array} \quad (20.28)$$

$$\begin{array}{c} \sum z' \\ \sum z'' \\ \sum z \end{array} \xrightarrow{\quad} \underline{y}$$

QED

Claim 38 (from Ref.[50])

If $\begin{array}{c} x \\ \downarrow \\ w \xrightarrow{\quad} z \xrightarrow{\quad} y \end{array}$ then



$$P(y|\mathcal{D}\underline{z} = z, x) = \sum_w P(y|z, x, w)P(w) \quad (20.29)$$

$$= \begin{array}{c} x \\ \searrow \\ \sum w \xrightarrow{\quad} z \xrightarrow{\quad} y \end{array} \quad (20.30)$$

$$\sum w \xrightarrow{\quad} z \xrightarrow{\quad} y$$

proof:

$$P(y|\mathcal{D}\underline{z} = z, x) = \begin{array}{c} x \\ \downarrow \\ \sum w \\ \uparrow \\ (\sum u) \\ \bar{x} \\ \downarrow \\ \sum w \\ \uparrow \\ (\sum u) \end{array} \quad (20.31)$$

$$= \begin{array}{c} x \\ \downarrow \\ \sum w \\ \uparrow \\ (\sum u) \\ \bar{x} \\ \downarrow \\ \sum w \\ \uparrow \\ (\sum u) \end{array} \quad (20.32)$$

$$= \begin{array}{c} x \\ \downarrow \\ \sum w \\ \uparrow \\ (\sum u) \\ \bar{x} \\ \downarrow \\ \sum w' \\ \leftarrow \sum w' \\ \uparrow \\ (\sum u) \end{array} \quad (20.33)$$

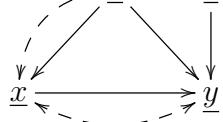
$$= \begin{array}{c} x \\ \downarrow \\ \sum w \\ \uparrow \\ (\sum u) \\ \bar{x} \\ \downarrow \\ \sum w \\ \uparrow \\ (\sum u) \end{array} \quad (20.34)$$

$$\sum w \xrightarrow{\quad z \quad} y$$

QED

Claim 39 (*Trivial Memoryless Transportability, from Ref.[48]*)

If $\begin{array}{ccc} & z & \\ \swarrow & & \searrow \\ x & \xrightarrow{\quad \dots \quad} & y \end{array}$ where $\underline{s} \in \{0, 1\}$ is a selection node, then

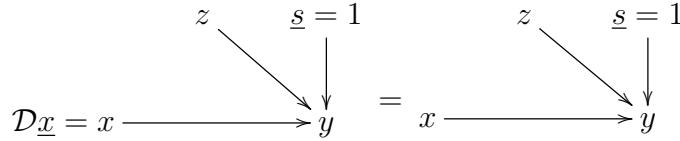


$$P^*(y|\mathcal{D}\underline{x} = x, z) = P^*(y|x, z) \quad (\text{replace } \mathcal{D} \text{ by } 1, \text{ keep } P^*) \quad (20.35)$$

$$\mathcal{D}\underline{x} = x \xrightarrow{\quad z \quad} y \quad \stackrel{\underline{s} = 1}{=} \quad x \xrightarrow{\quad z \quad} y \quad (20.36)$$

proof:

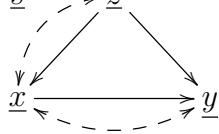
$$P(y|\mathcal{D}\underline{x} = x, z, \underline{s} = 1) = P(y|x, z, \underline{s} = 1)$$



QED

Claim 40 (*Direct Transportability, a.k.a. External Validity, from Ref.[48]*)

If $\underline{s} \xrightarrow{\quad} \underline{z}$ where $\underline{s} \in \{0, 1\}$ is a selection node, then



$$P^*(y|\mathcal{D}\underline{x} = x, z) = P(y|\mathcal{D}\underline{x} = x, z) \quad (\text{replace } P^* \text{ by } P, \text{ keep } \mathcal{D}) \quad (20.37)$$

$$\begin{array}{ccc} \underline{s} = 1 \longrightarrow z & & z \\ \swarrow & \searrow & \downarrow \\ \mathcal{D}\underline{x} = x \longrightarrow y & = & \mathcal{D}\underline{x} = x \longrightarrow y \end{array} \quad (20.38)$$

Furthermore,

$$P^*(y|\mathcal{D}\underline{x} = x) = \sum_z P(y|\mathcal{D}\underline{x} = x, z)P^*(z) \quad (20.39)$$

$$\begin{array}{ccc} \underline{s} = 1 & & \underline{s} = 1 \longrightarrow \sum z \\ \searrow & & \swarrow \\ \mathcal{D}\underline{x} = x \longrightarrow y & = & \mathcal{D}\underline{x} = x \longrightarrow y \end{array} \quad (20.40)$$

proof:

$$P(y|\mathcal{D}\underline{x} = x, z, \underline{s} = 1) = P(y|\mathcal{D}\underline{x} = x, z)$$

$$\begin{array}{ccc} \underline{s} = 1 \longrightarrow z & & z \\ \swarrow & \searrow & \downarrow \\ \mathcal{D}\underline{x} = x \longrightarrow y & = & \mathcal{D}\underline{x} = x \longrightarrow y \end{array} \quad \text{Because } \underline{s} \perp \underline{y} | \underline{z}$$

Furthermore,

$$P(y|\mathcal{D}\underline{x} = x, \underline{s} = 1) = \sum_z P(y|\mathcal{D}\underline{x} = x, z)P(z|\underline{s} = 1)$$

$\underline{s} = 1 \longrightarrow \sum z$

 $\mathcal{D}\underline{x} = x \longrightarrow y$

QED

Claim 41 (*S-Admissible Transportability, from Ref.[48]*)

If $\underline{s} \xrightarrow{\text{---}} (\underline{z}) \xrightarrow{\text{---}} \underline{a}$ where $\underline{s} \in \{0, 1\}$ is a selection node, then

$\underline{x} \xrightarrow{\text{---}} \underline{y}$

$$P^*(y|\mathcal{D}\underline{x} = x) = \sum_a P(y|\mathcal{D}\underline{x} = x, a)P^*(a) \quad (20.41)$$

$$\begin{array}{ccc} \underline{s} = 1 & & \underline{s} = 1 \longrightarrow \sum a \\ \searrow & & \downarrow \\ \mathcal{D}\underline{x} = x \longrightarrow y & = & \mathcal{D}\underline{x} = x \longrightarrow y \end{array} \quad (20.42)$$

proof:

$$P(y|\mathcal{D}\underline{x} = x, \underline{s} = 1) = \sum_a P(y|\mathcal{D}\underline{x} = x, a)P(a|\underline{s} = 1)$$

$\underline{s} = 1 \longrightarrow (\sum z) \longrightarrow \sum a$

 $\mathcal{D}\underline{x} = x \longrightarrow y$

QED

Claim 42 (*Non-transportability, from Ref.[48]*)

If $\underline{s} \xrightarrow{\text{---}} (\underline{h}) \xrightarrow{\text{---}} \underline{s}$ where $\underline{s} \in \{0, 1\}$ is a selection node, then

$\underline{x} \longrightarrow \underline{y}$

$$P^*(y|\mathcal{D}\underline{x} = x) = P^*(y|\mathcal{D}\underline{x} = x) \quad (20.43)$$

$$\begin{array}{ccc}
 & \underline{s} = 1 & \\
 & \downarrow & \\
 \mathcal{D}\underline{x} = x & \xrightarrow{\hspace{2cm}} & y = \text{same}
 \end{array} \tag{20.44}$$

proof:

$$P^*(y|\mathcal{D}\underline{x} = x) = P^*(y|\mathcal{D}\underline{x} = x)$$

$$\begin{array}{ccc}
 & \left(\sum h \right) & \underline{s} = 1 \\
 & \swarrow \searrow & \downarrow \\
 \mathcal{D}\underline{x} = x & \xrightarrow{\hspace{2cm}} & y = \mathcal{D}\underline{x} = x \xrightarrow{\hspace{2cm}} y \quad \underline{s} = 1 \downarrow
 \end{array}$$

Can't replace $\mathcal{D}\underline{x} = x$ by x because $y \not\perp \underline{x}$ in $\mathcal{L}_{\underline{x}}G$. Hence, Rule 2 not satisfied.

QED

Claim 43 (from Ref.[48])

$$\begin{array}{l}
 \text{If } \underline{s} \xrightarrow{\hspace{1cm}} \left(\sum h \right) \xrightarrow{\hspace{1cm}} z \text{ where } \underline{s} \in \{0, 1\} \text{ is a selection node, then} \\
 \begin{array}{c}
 \xrightarrow{\hspace{1cm}} \left(\sum h \right) \xrightarrow{\hspace{1cm}} z \\
 \swarrow \searrow \\
 x \xrightarrow{\hspace{1cm}} y
 \end{array}
 \end{array}$$

$$P^*(y|\mathcal{D}\underline{x} = x) = P(y|\mathcal{D}\underline{x} = x) \tag{20.45}$$

$$\underline{s} = 1 \quad \mathcal{D}\underline{x} = x \xrightarrow{\hspace{2cm}} y = \mathcal{D}\underline{x} = x \xrightarrow{\hspace{2cm}} y \tag{20.46}$$

$$\mathcal{D}\underline{x} = x \xrightarrow{\hspace{2cm}} y = \mathcal{D}\underline{x} = x \xrightarrow{\hspace{2cm}} y$$

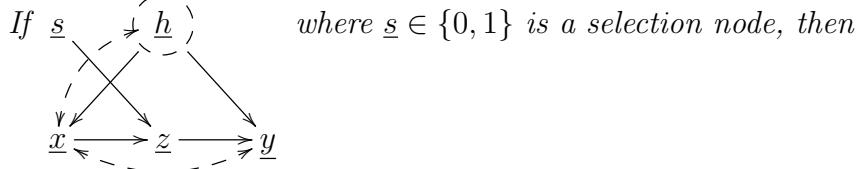
proof:

$$P(y|\mathcal{D}\underline{x} = x, \underline{s} = 1) = \sum_h P(y|\mathcal{D}\underline{x} = x, h)P(h)$$

$$\begin{array}{ccc}
 & \underline{s} = 1 \xrightarrow{\hspace{1cm}} \left(\sum h \right) \xrightarrow{\hspace{1cm}} \sum z & \\
 & & \left(\sum h \right) \xrightarrow{\hspace{1cm}} y \\
 & \mathcal{D}\underline{x} = x \xrightarrow{\hspace{2cm}} y & = \mathcal{D}\underline{x} = x \xrightarrow{\hspace{2cm}} y \\
 & = P(y|\mathcal{D}\underline{x} = x) & \\
 & = \mathcal{D}\underline{x} = x \xrightarrow{\hspace{2cm}} y &
 \end{array}$$

QED

Claim 44 (from Ref.[48])

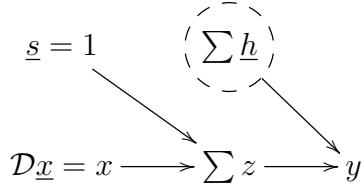


$$P^*(y|\mathcal{D}\underline{x} = x) = \sum_z P(y|\mathcal{D}\underline{x} = x, z)P^*(z|x) \quad (20.47)$$

$$\begin{array}{ccc} \underline{s} = 1 & \mathcal{D}\underline{x} = x & \underline{s} = 1 \\ \searrow & \downarrow & \searrow \\ & y & = \\ & x \longrightarrow \sum z \longrightarrow y & \end{array} \quad (20.48)$$

proof:

$$P(y|\mathcal{D}\underline{x} = x, \underline{s} = 1) = \sum_h \sum_z P(y|h, z)P(h)P(z|\mathcal{D}\underline{x} = x, \underline{s} = 1)$$

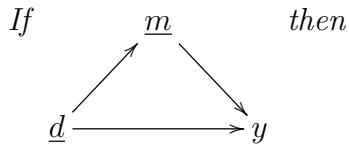


$$= \sum_h \sum_z P(y|h, z)P(h|\mathcal{D}\underline{x} = x)P(z|x, \underline{s} = 1)$$

$$\begin{aligned} & \underline{s} = 1 \quad (\sum h) \quad \mathcal{D}\underline{x} = x \\ & \searrow \qquad \qquad \swarrow \qquad \qquad \downarrow \\ & \mathcal{D}\underline{x} = x \longrightarrow \sum z \longrightarrow y \\ & = x \longrightarrow \sum z \longrightarrow y \\ & = \sum_z P(y|\mathcal{D}\underline{x} = x, z)P(z|x, \underline{s} = 1) \\ & = \underline{s} = 1 \quad \mathcal{D}\underline{x} = x \\ & = x \longrightarrow \sum z \longrightarrow y \end{aligned}$$

QED

Claim 45 (Unconfounded Mediation, from Ref.[47])



$$P(y|\mathcal{D}\underline{d} = d, \mathcal{I}\underline{m}\underline{d} = d') = \sum_m P(y|d, m)P(m|d') \quad (20.49)$$

$$\begin{array}{ccc} \mathcal{I}\underline{d} = d' & & \mathcal{I}\underline{d} = d' \longrightarrow \sum m \\ \searrow & & \swarrow \\ \mathcal{D}\underline{d} = d \longrightarrow y & = & \mathcal{D}\underline{d} = d \longrightarrow y \end{array} \quad (20.50)$$

proof:

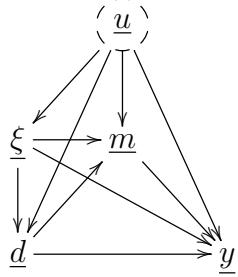
$$P(y|\mathcal{D}\underline{d} = d, \mathcal{I}\underline{d} = d') = \sum_m P(y|d, m)P(m|d')$$

$$\begin{array}{ccc} \mathcal{I}\underline{d} = d' \longrightarrow \sum m & & \\ \searrow & & \\ \mathcal{D}\underline{d} = d \longrightarrow y & & \end{array}$$

QED

Claim 46 (*Mediation with universal prior ξ and universal confounder \underline{u} , from Ref.[47]*)

If $\xi \rightarrow (\underline{u})$ then

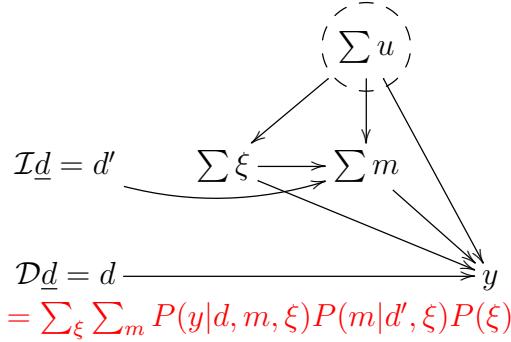


$$P(y|\mathcal{D}\underline{d} = d, \mathcal{I}\underline{m}\underline{d} = d') = \sum_{\xi} \sum_m P(y|d, m, \xi)P(m|d', \xi)P(\xi) \quad (20.51)$$

$$\begin{array}{ccc} \mathcal{I}\underline{d} = d' & & \mathcal{I}\underline{d} = d' \xrightarrow{\sum \xi} \sum m \\ \searrow & & \swarrow \\ \mathcal{D}\underline{d} = d \longrightarrow y & = & \mathcal{D}\underline{d} = d \longrightarrow y \end{array} \quad (20.52)$$

proof:

$$P(y|\mathcal{D}\underline{d} = d, \mathcal{I}\underline{d} = d') = \sum_{\xi,u} \sum_m P(y|d, m, \xi, u)P(m|d', \xi, u) \underbrace{P(\xi|u)P(u)}_{P(\xi,u)}$$



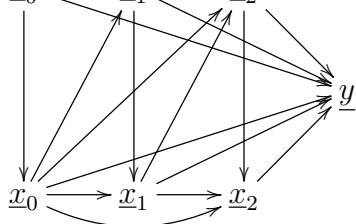
$$\begin{aligned} & \mathcal{I}\underline{d} = d' \\ & \quad \sum \xi \longrightarrow \sum m \\ & = \mathcal{D}\underline{d} = d - \sum_{\xi} \sum_m P(y|d, m, \xi)P(m|d', \xi)P(\xi) \end{aligned}$$

We switch from averaging over the prior of ξ, u to averaging over the prior of ξ .

QED

Claim 47 (Sequential backdoor (SBD) adjustment formula, from Ref.[51])

If $\underline{z}_0 \rightarrow \underline{z}_1 \rightarrow \underline{z}_2$ then



$$P(y|\mathcal{D}\underline{x}^3 = x^3) = \mathcal{Q}(y|x^3) \quad (20.53)$$

$$\sum z_0 \rightarrow \sum z_1 \rightarrow \sum z_2 \quad (20.54)$$

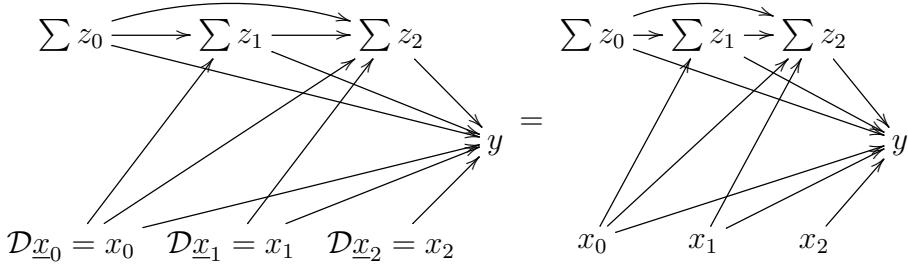
$$\mathcal{D}\underline{x}^3 = x^3 \rightarrow y =$$



The result shown here for $n = 3$ is true for any integer $n \geq 1$.

proof:

$$P(y|\mathcal{D}\underline{x}^3 = x^3) = \mathcal{Q}(y|x^3)$$



We can replace $\mathcal{D}\underline{x}_i = x_i$ by x_i once all nodes in bnet are observed nodes.

QED

Claim 48 (*Selection Bias (SB) Backdoor Adjustment Formula, from Ref.[2]*)

If $\underline{s} \xrightarrow{\quad} \underline{z}^{<\underline{x}} \xrightarrow{\quad} \underline{z}^{>\underline{x}}$ where $\underline{s} \in \{0, 1\}$ is a selection node and $\underline{z} = (\underline{z}^{<\underline{x}}, \underline{z}^{>\underline{x}})$,

then

$$P(y|\mathcal{D}\underline{x} = x, \underline{s} = 1) = \sum_z P(y|x, z)P(z) = P(y|x) \quad (20.55)$$

$$\begin{array}{ccc} \underline{s} = 1 & & \sum z \\ \searrow & & \downarrow \\ \mathcal{D}\underline{x} = x \longrightarrow y & = & x \longrightarrow y \end{array} = x \longrightarrow y \quad (20.56)$$

proof:

$$P(y|\mathcal{D}\underline{x} = x, \underline{s} = 1) = \sum_z P(y|\mathcal{D}\underline{x} = x, z)P(z^{<\underline{x}}|\underline{s} = 1)P(z^{>\underline{x}}|x, z^{<\underline{x}}, \underline{s} = 1)$$

$$\begin{aligned} & \underline{s} = 1 \xrightarrow{\quad} \sum z^{<\underline{x}} \xrightarrow{\quad} \sum z^{>\underline{x}} \\ & \mathcal{D}\underline{x} = x \longrightarrow y \\ &= \sum_{z^{<\underline{x}}} P(y|\mathcal{D}\underline{x} = x, z^{<\underline{x}})P(z^{<\underline{x}}|\underline{s} = 1) \\ & \underline{s} = 1 \longrightarrow \sum z^{<\underline{x}} \\ &= \mathcal{D}\underline{x} = x \longrightarrow y \\ &= \sum_z P(y|x, z)P(z|\underline{s} = 1) \end{aligned}$$

$$\begin{aligned}
 & \underline{s} = 1 \longrightarrow \sum z \\
 &= \quad \quad \quad \downarrow \\
 &= \quad \quad \quad x \longrightarrow y \\
 &= \sum_z P(y|x, z)P(z)
 \end{aligned}
 \qquad \begin{aligned}
 & \mathcal{D} \text{ can be removed because there are} \\
 & \text{no sums over unobserved nodes.} \\
 \\
 & \begin{aligned}
 & \underline{s} = 1 \text{ node can be removed} \\
 & \text{because this expression must} \\
 & \text{equal } P(y|x, \underline{s} = 1). \text{ Furthermore,} \\
 & \underline{y} \perp \underline{s} | (\underline{x}, \underline{z}) \text{ in the hypothesis bnet.} \\
 & \text{Hence, this expression must also} \\
 & \text{equal } P(y|x).
 \end{aligned}
 \end{aligned}$$

QED

Chapter 21

D-Separation

Before reading this chapter, I recommend that you read Chapter D.

A path γ that isn't a loop can have 3 types of intermediate nodes \underline{x} (an intermediate node of γ is a node in γ that isn't one of the two end nodes). Suppose $\underline{a}, \underline{b} \in \gamma$ are the two neighbors of \underline{x} . Then the 3 possible cases are:

1. **\underline{x} is a mediator node:** $(\underline{a} \leftarrow \underline{x} \leftarrow \underline{b})$ or $(\underline{a} \rightarrow \underline{x} \rightarrow \underline{b})$
2. **\underline{x} is a fork node:** $(\underline{a} \leftarrow \underline{x} \rightarrow \underline{b})$
3. **\underline{x} is a collider node:** $(\underline{a} \rightarrow \underline{x} \leftarrow \underline{b})$

We say that a non-loop path γ from \underline{a} to \underline{b} (i.e., with end nodes $\underline{a}, \underline{b}$) is **blocked** by conditioning on a multinode \underline{Z} . if one or more of the following statements is true:

1. There is a node $\underline{x} \in \underline{Z}$. which is a mediator or a fork of γ .
2. γ contains a collider node \underline{c} and $(\underline{c} \cup de(\underline{c})) \cap \underline{Z} = \emptyset$ (i.e., neither \underline{c} nor any of the descendants of \underline{c} is contained in \underline{Z} .)

This definition of a blocked path¹ is easy to remember if one thinks of the following analogy with pipes carrying a fluid. Think of path γ as if it were a pipe carrying a fluid. Think of the nodes of γ as junctions in the pipe. If \underline{Z} intersects γ at either a mediator or a fork junction, that blocks the pipe flow. A collider junction \underline{c} is like a blackhole or a huge leak. Its presence blocks passage of the fluid as long as neither \underline{c} nor any of the descendants of \underline{c} are in \underline{Z} . If, on the other hand, $\underline{c} \in \underline{Z}$, or $\underline{c}' \in \underline{Z}$ where $\underline{c}' \in de(\underline{c})$, then that acts as a complete (in the case of $\underline{c} \in \underline{Z}$) or a partial (in the case of $\underline{c}' \in \underline{Z}$) bridge across the blackhole.

See Fig.21.1 for some examples of paths that are blocked or not blocked by conditioning on a multinode \underline{Z} .

¹Note that we speak of blocked paths or info, not of blocked nodes. Nodes are not blocked; rather they are either conditioned upon or not.

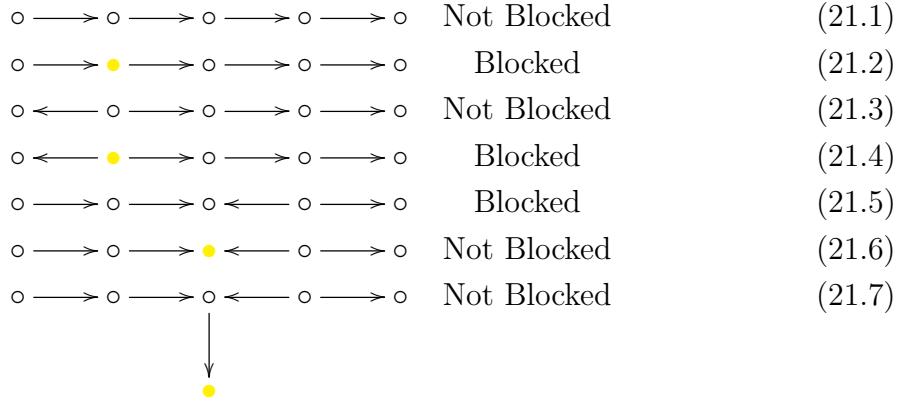


Figure 21.1: Examples of paths that are blocked or not blocked by conditioning on a multinode $\underline{Z}.$ Nodes belonging to $\underline{Z}.$ are colored yellow.

Given 3 disjoint multinode $\underline{A}., \underline{B}., \underline{Z}.$ of a graph G , we write “ $\underline{A}. \perp_G \underline{B}.|\underline{Z}.$ ” or say “ $\underline{A}.$ and $\underline{B}.$ are **d-separated** by $\underline{Z}.$ in G ” iff there exists no path γ from $a \in \underline{A}.$ to $b \in \underline{B}.$ which is not blocked by conditioning on $\underline{Z}..$ ²

The minimal Markov blanket (see Chapter 45) of a node a is the smallest multinode $\underline{Z}.$ such that $a \perp_G b|\underline{Z}.$ for all $b \notin a \cup \underline{Z}..$

We are finally ready to state the d-separation theorem, without proof.

A probability distribution P is **compatible with a DAG** G if P and G have the same random variables, and they can be combined to form a bnet without contradictions; i.e., one can calculate all the TPMs from P and multiply them together to obtain P again.

Claim 49 (d-separation Theorem)

Suppose $\underline{A}., \underline{B}., \underline{Z}.$ are disjoint multinode of a DAG G .

If $\underline{A}. \perp_G \underline{B}.|\underline{Z}.$ then $P(B.|A., Z.) = P(B.|Z.)$ for all $B., A., Z.,$ for all P compatible with G .

The full converse of the theorem can also be proven, but we won't be using it in this book.

Often, the right hand side of this theorem is stated as “ $\underline{A}. \perp_P \underline{B}.|\underline{Z}.$ for all P ”. Then the theorem is stated: “If $\underline{A}. \perp_G \underline{B}.|\underline{Z}.$ then $\underline{A}. \perp_P \underline{B}.|\underline{Z}.$ for all P .”

Note that the following are equivalent:

- $P(B.|A., Z.) = P(B.|Z.)$ for all $B., A., Z..$
- $\underline{A}. \perp_P \underline{B}.|\underline{Z}.$

² $\underline{Z}.$ are the nodes we are “conditioning on”. Unmeasured (i.e., hidden, unobserved) nodes cannot be conditioned on, because that would entail measuring them.

- $H(\underline{A}_. : \underline{B}_. | \underline{Z}_.) = 0$ (see Chapter C for definition of conditional mutual information (CMI))

Extra stuff: mostly only for pure mathematicians

Below, we will use the notation $nde(\underline{a})$ to denote all non-descendants, including \underline{a} itself, of a node \underline{a} in a DAG G ; i.e., all nodes of G that are not in $de(\underline{a}) \cup \underline{a}$, where $de(\underline{a})$ is defined in Chapter D.

Given a DAG G , define the following sets of d-separations:³

$$DS(G) = \{(\underline{A}_. \perp_G \underline{B}_. | \underline{Z}_.) : \underline{A}_., \underline{B}_., \underline{Z}_. \text{ are multinodes of } G\} . \quad (21.8)$$

$$DS_{min}(G) = \{(\underline{A}_. \perp_G nde(\underline{A}_.) | pa(\underline{A}_.)) : \underline{A}_. \text{ is a multinode of } G\} . \quad (21.9)$$

See Chapter 61 for an example where set $DS_{min}(G)$ is calculated for a particular DAG G .

Claim 50 *For all DAGs G , $DS(G) = DS_{min}(G)$.*

Given a probability distribution P , define the following set of conditional independencies:

$$CI(P) = \{(\underline{A}_. \perp_P \underline{B}_. | \underline{Z}_.) : \underline{A}_., \underline{B}_., \underline{Z}_. \text{ are multinodes of } P\} , \quad (21.10)$$

For a DAG G and a probability distribution P compatible with G , define a map ϕ by

$$\phi : DS_{min}(G) \rightarrow CI(P) \quad (21.11)$$

$$\phi : \underline{A}_. \perp_G nde(\underline{A}_.) | pa(\underline{A}_.) \mapsto \underline{A}_. \perp_P nde(\underline{A}_.) | pa(\underline{A}_.) \quad (21.12)$$

In general, this map is 1-1 but not onto.

Claim 51 *For a bnet with a DAG G and a total probability distribution P , the map ϕ is a bijection.*

$DS(G)$ does not fully specify a DAG. DAGs with the same $DS(G)$ are said to be **d-separation equivalent**. See Chapter 61 for more info about d-separation equivalence.

³Note that $(\underline{A}_. \perp_G nde(\underline{A}_.) | pa(\underline{A}_.))$ and $(\underline{A}_. \perp_G nde(\underline{A}_.) - pa(\underline{A}_.) | pa(\underline{A}_.))$ are equivalent because $H(\underline{a} : \underline{b}, \underline{c} | \underline{c}) = H(\underline{a} : \underline{b} | \underline{c})$.

Chapter 22

D-Separation and Do Calculus Rules for LDEN: COMING SOON

In this chapter, we will prove the D-Separation Theorem (see Chapter 21) and the Do Calculus Rules (see Chapter 19) for the special bnets that we call LDEN(see Chapter 44). Proving these theorems for LDEN rather than for general bnets is much simpler and quite instructive.

Chapter 23

D-Separation in Quantum Mechanics

See Ref.[68].

Chapter 24

Dynamical Bayesian Networks

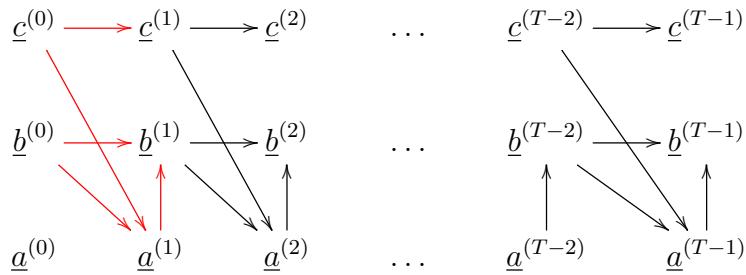


Figure 24.1: Example of a dynamical bnet. The pattern of red arrows is repeated $T - 1$ times.

A dynamical bnet is simply a time homogeneous Markov chain (see Chapter 47) for which each node is called a **time slice**, and each time slice represents at finer resolution a sub-DAG which is the same between any 2 successive time slices. Fig.24.1 gives an example of a dynamical bnet. In Fig.24.1, we've drawn the 3 nodes of each time slice vertically, and labeled them with a superscript $^{(t)}$, where $t \in \{0, 1, \dots, T-1\}$ is the time of the slice. To fully specify the dynamical bnet of Fig.24.1, we would also have to specify the TPMs

$$\begin{aligned} &P(c^{(0)}), \\ &P(b^{(0)}) \\ &P(c^{(1)}|c^{(0)}), \\ &P(b^{(1)}|b^{(0)}, a^{(1)}) \\ &P(a^{(1)}|b^{(0)}). \end{aligned}$$

Dynamical bnets are very common in AI and Data Science. Kalman filters (Chapter 41), Hidden Markov Models (Chapter 35) and Recurrent Neural Networks (Chapter 67) are famous examples of dynamical bnets.

Bnets are acyclic; they can't have cycles (i.e., closed directed paths). Yet feedback loops are an important concept in Science. So what is the equivalent of feedback loops in the bnet world? Dynamical bnets are. Fig.24.2 represents Fig.24.1

more compactly using feedback loops. Any bnet with feedback loops can be “unrolled” into a dynamical bnet.

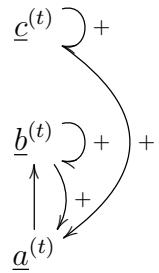


Figure 24.2: Dynamical bnet Fig.24.1 represented more compactly using feedback loops. Arrows labelled + point from nodes of the t time slice to nodes of the $t + 1$ time slice.

Chapter 25

Expectation Maximization

This chapter is based on Refs.[97] and [151].

The Expectation Maximization (EM) algorithm is commonly used in Data Science to find the maximum over an **unknown parameter** θ of a likelihood function

$$P(\vec{x}|\theta) = \sum_{\vec{h}} P(\vec{x}, \vec{h}|\theta), \quad (25.1)$$

where \vec{x} denotes the **observed variables**, and \vec{h} denotes the **hidden variables**. Both θ and \vec{h} are hidden (i.e., unobserved).¹

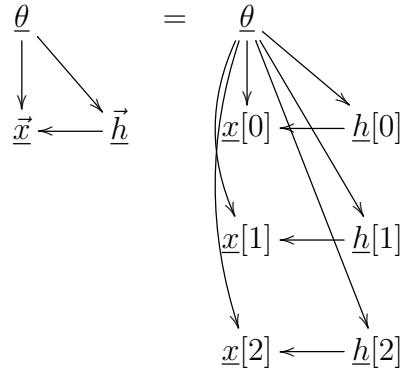


Figure 25.1: bnet for EM with $nsam = 3$. $x[\sigma] = x^\sigma$ and $h[\sigma] = h^\sigma$.

The bnet for the EM algorithm is given by Fig.25.1 for $nsam = 3$. Later on in this chapter, we will give the node TPMs for this bnet for the special case in which $P(x^\sigma | \theta)$ is a mixture (i.e., weighted sum) of Gaussians.

¹The term “unknown parameter” is mainly of frequentist origin. For Bayesians, θ is a random variable with a delta function prior, whereas for frequentists, it is not a random variable at all, just an unknown parameter with no randomness.

Note that if we erase the \underline{h}^σ nodes from Fig.25.1, we get the bnet for naive Bayes, which is used for classification into the states of $\underline{\theta}$. However, there is one big difference. With naive Bayes, the leaf nodes have different TPMs. Here, we will assume they are i.i.d. Naive Bayes is used for classification: i.e., given the states of the leaf nodes, we infer the state of the root node. EM is used for clustering; i.e., given many i.i.d. samples, we fit their distribution by a weighted sum of prob distributions, usually Gaussians.

Let

L = likelihood function.

$nsam$ = number of samples.

$\vec{x} = (x[0], x[1], \dots, x[nsam - 1])$, $x^\sigma = x[\sigma] \in S_{\underline{x}}$ for all σ .

$\vec{h} = (h[0], h[1], \dots, h[nsam - 1])$, $h^\sigma = h[\sigma] \in S_{\underline{h}}$ for all σ .

We assume that the samples (x^σ, h^σ) are i.i.d. for different σ at fixed θ . What this means is that there are probability distributions $P_{\underline{x}|\underline{h},\underline{\theta}}$ and $P_{\underline{h}|\underline{\theta}}$ such that

$$P(\vec{x}, \vec{h} | \theta) = \prod_{\sigma} [P_{\underline{x}|\underline{h},\underline{\theta}}(x^\sigma | h^\sigma, \theta) P_{\underline{h}|\underline{\theta}}(h^\sigma | \theta)] . \quad (25.2)$$

Definition of likelihood functions:

$$\underbrace{P(\vec{x} | \theta)}_{L(\theta; \vec{x})} = \sum_{\vec{h}} \underbrace{P(\vec{x}, \vec{h} | \theta)}_{L(\theta; \vec{x}, \vec{h})} \quad (25.3)$$

θ^* = maximum likelihood estimate of θ (no prior $P(\theta)$ assumed):

$$\theta^* = \operatorname{argmax}_{\theta} L(\theta; \vec{x}) \quad (25.4)$$

25.1 The EM algorithm:

1. **Expectation step:**²

$$Q(\theta | \theta^{(t)}) = E_{\vec{h} | \vec{x}, \theta^{(t)}} \ln P(\vec{x}, \vec{h} | \theta) \quad (25.5)$$

2. **Maximization step:**

$$\theta^{(t+1)} = \operatorname{argmax}_{\theta} Q(\theta | \theta^{(t)}) . \quad (25.6)$$

Claim: $\lim_{t \rightarrow \infty} \theta^{(t)} = \theta^*$.

Fig.25.2 portrays the recursive nature of the EM algo as a dynamical, recurrent bnet. For Fig.25.2, the TPMs, printed in blue, for the $\underline{\theta}^{(t)}$ nodes for $t = 1, 2, \dots$, are as follows:

²Note that the right hand side of Eq.(25.5) is expressible in the form $\sum_{\sigma} \sum_{h^\sigma} f(x^\sigma, h^\sigma)$.

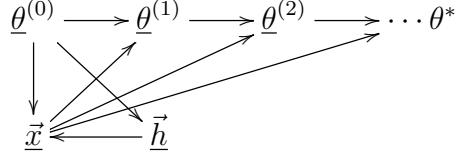


Figure 25.2: The EM algo generates a sequence of parameter estimates $(\theta^{(t)})_{t=0,1,2,\dots}$ that converges to the optimum (i.e., best-fit) parameter θ^* .

$$P(\theta^{(t+1)}|\vec{x}, \theta^{(t)}) = \delta(\theta^{(t+1)}, \operatorname{argmax}_{\theta} Q(\theta|\theta^{(t)})) . \quad (25.7)$$

25.1.1 Motivation

$$Q(\theta|\theta^{(t)}) = E_{\vec{h}|\vec{x}, \theta^{(t)}} \ln P(\vec{x}, \vec{h}|\theta) \quad (25.8)$$

$$= E_{\vec{h}|\vec{x}, \theta^{(t)}} [\ln P(\vec{h}|\vec{x}, \theta) + \ln P(\vec{x}|\theta)] \quad (25.9)$$

$$= -D_{KL} \left(P(\vec{h}|\vec{x}, \theta^{(t)}) \parallel P(\vec{h}|\vec{x}, \theta) \right) - H[P(\vec{h}|\vec{x}, \theta^{(t)})] + \ln P(\vec{x}|\theta) \quad (25.10)$$

When $\theta^{(t)} = \theta$, this becomes

$$Q(\theta|\theta) = -H[P(\vec{h}|\vec{x}, \theta)] + \ln P(\vec{x}|\theta) . \quad (25.11)$$

Hence,

$$\partial_{\theta} Q(\theta|\theta) = - \sum_{\vec{h}} \partial_{\theta} P(\vec{h}|\vec{x}, \theta) + \partial_{\theta} \ln P(\vec{x}|\theta) \quad (25.12)$$

$$= \partial_{\theta} \ln P(\vec{x}|\theta) \quad (25.13)$$

So if $\theta^{(t)} \rightarrow \theta$ and $Q(\theta|\theta)$ is max at $\theta = \theta^*$, then $\ln P(\vec{x}|\theta)$ is max at $\theta = \theta^*$ too.

For a more rigorous proof that $\lim_{t \rightarrow \infty} \theta^{(t)} = \theta^*$, see Wikipedia article Ref.[97] and references therein.

25.2 Minorize-Maximize (MM) algorithms

A function $\mu(\theta|\theta^{(t)})$ is said to **minorize a target function** $LL(\theta)$ iff for all θ at fixed $\theta^{(t)}$, it satisfies the “ $\mu \leq LL$ property”

$$\mu(\theta|\theta^{(t)}) \leq LL(\theta) , \quad (25.14)$$

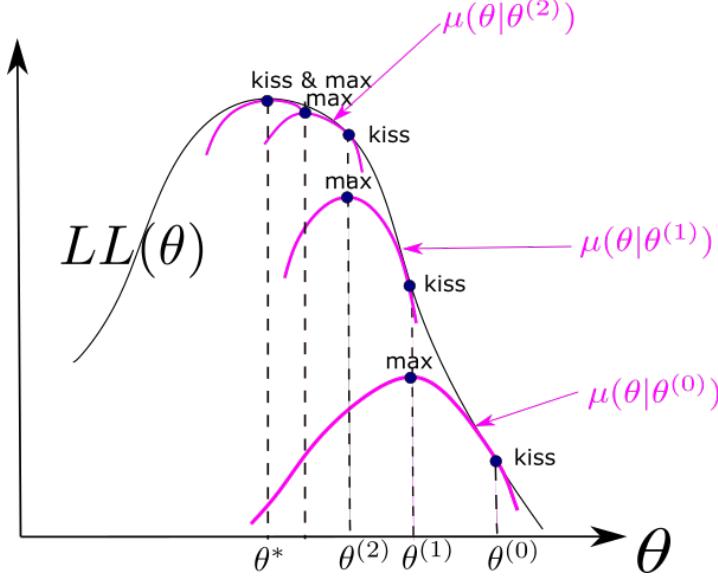


Figure 25.3: Function $\mu(\theta|\theta^{(t)})$ minorizes the function $LL(\theta)$. Note that $\mu(\theta|\theta^{(t)})$ is always below $LL(\theta)$. “max” indicates $\theta^{(t+1)} = \underset{\theta}{\operatorname{argmax}} \mu(\theta|\theta^{(t)})$. “kiss” indicates $\mu(\theta^{(t)}|\theta^{(t)}) = LL(\theta^{(t)})$.

and the “ $\mu = LL$ property”

$$\mu(\theta^{(t)}|\theta^{(t)}) = LL(\theta^{(t)}) . \quad (25.15)$$

We **recursively maximize a minorizing function** $\mu(\theta|\theta^{(t)})$ if we define a sequence $(\theta^{(t)})_{t=0,1,\dots}$ as follows:

$$\theta^{(t+1)} = \underset{\theta}{\operatorname{argmax}} \mu(\theta|\theta^{(t)}) . \quad (25.16)$$

The sequence $(LL(\theta^{(t)}))_{t=0,1,2,\dots}$ generated by recursively maximizing a minorizing function must be nondecreasing:

$$LL(\theta^{(t+1)}) \geq \mu(\theta^{(t+1)}|\theta^{(t)}) \geq \mu(\theta^{(t)}|\theta^{(t)}) = LL(\theta^{(t)}) . \quad (25.17)$$

A **minorize-maximize (MM) algorithm** is any algo that specifies a minorizing function $\mu(\theta|\theta^{(t)})$ for a particular target function $LL(\theta)$. One can also define a **majorize-minimize algo (also called MM)** by inverting the inequalities throughout.

The EM algo is an MM algo. Indeed, if we define

$$LL(\theta) = \ln P(\vec{x}|\theta) \quad (25.18)$$

and

$$\mu(\theta|\theta^{(t)}) = Q(\theta|\theta^{(t)}) + H(P(\vec{h}|\vec{x}, \theta^{(t)}), \quad (25.19)$$

then Eq.(25.10) establishes the $\mu \leq LL$ and $\mu = LL$ properties required of a minorizing function.

How an MM algo works is portrayed in Fig.25.3.

25.3 Examples

25.3.1 Gaussian mixture

$x^\sigma \in \mathbb{R}^d = S_{\underline{x}}$. $S_{\underline{h}}$ discrete and not too large. $n_{\underline{h}} = |S_{\underline{h}}|$ is number of Gaussians that we are going to fit the samples with.

Let

$$\theta = [w_h, \mu_h, \Sigma_h]_{h \in S_{\underline{h}}}, \quad (25.20)$$

where $[w_h]_{h \in S_{\underline{h}}}$ is a probability distribution of weights, and where $\mu_h \in \mathbb{R}^d$ and $\Sigma_h \in \mathbb{R}^{d \times d}$ are the mean value vector and covariance matrix of a d -dimensional Gaussian distribution.

The TPMs, printed in blue, for the bnet Fig.25.1, for the special case of a mixture of Gaussians, are as follows:

$$P(x^\sigma | h^\sigma | \theta) = \mathcal{N}_d(x^\sigma; \mu_{h^\sigma}, \Sigma_{h^\sigma}) \quad (25.21)$$

$$P(h^\sigma | \theta) = w_{h^\sigma} \quad (25.22)$$

Note that

$$P(x^\sigma | \theta) = \sum_h P(x^\sigma | h^\sigma = h, \theta) P(h^\sigma = h | \theta) \quad (25.23)$$

$$= \sum_h w_h \mathcal{N}_d(x^\sigma; \mu_h, \Sigma_h) \quad (25.24)$$

$$P(\vec{x}, \vec{h} | \theta) = \prod_\sigma [w_{h^\sigma} \mathcal{N}_d(x^\sigma; \mu_{h^\sigma}, \Sigma_{h^\sigma})] \quad (25.25)$$

$$= \prod_\sigma \prod_h [w_h \mathcal{N}_d(x^\sigma; \mu_h, \Sigma_h)]^{\mathbb{1}(h=h^\sigma)} \quad (25.26)$$

Old Faithful: See Wikipedia Ref.[97] for an animated gif of a classic example of using EM to fit samples with a Gaussian mixture. Unfortunately, could not include

it here because pdflatex does not support animated gifs. The gif shows samples in a 2 dimensional space (eruption time, delay time) from the Old Faithful geyser. In that example, $d = 2$ and $n_h = 2$. Two clusters of points in a plane are fitted by a mixture of 2 Gaussians.

K-means clustering is often presented as the main competitor to EM for doing **clustering (non-supervised learning)**. In K-means clustering, the sample points are split into K mutually disjoint sets S_0, S_1, \dots, S_{K-1} . The algorithm is easy to describe:

1. Initialize by choosing at random K data points $(\mu_k)_{k=0}^{K-1}$ called means or centroids and placing μ_k in S_k for all k .
2. **STEP 1:** For each data point, add it to the S_k whose centroid μ_k is closest to it.
3. **STEP 2:** Recalculate the centroids. Set μ_k equal to the mean value of set S_k .
4. Repeat steps 1 and 2 until the centroids stop changing by much.

Step 1 is analogous to the expectation step in EM, and Step 2 to the maximization step in EM (θ estimation versus μ_k estimation). We won't say anything further about K-means clustering because it isn't related to bnets in any way, and this is a book about bnets. For more info about K-means clustering, see Ref.[113].

25.3.2 Blood Genotypes and Phenotypes

Notation: $\vec{a} = (\underline{a}^\sigma)_{\sigma=0,1,\dots,n_{sam}-1}$, where n_{sam} is the number of samples.

Suppose $\vec{x} = (\vec{x}_0)$ (i.e., just one component)

$\vec{h} = (\vec{h}_0)$ (i.e., just one component)

$\underline{h}^\sigma \in S_h = \{AA, AO, BB, BO, OO, AB\}$ (the 6 blood genotypes)

$\underline{x}^\sigma \in S_x = \{A, B, O, AB\}$ (the 4 blood phenotypes)

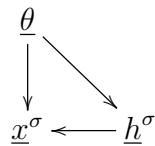


Figure 25.4: bnet for blood phenotypes x^σ and genotypes h^σ .

For the bnet Fig.25.4, the TPMs, printed in blue, are as follows:

$$P(h^\sigma | \theta) = \begin{array}{c|c} & \\ \hline AA & p_A^2 \\ AO & 2p_A p_O \\ BB & p_B^2 \\ BO & 2p_B p_O \\ OO & p_O^2 \\ AB & 2p_A p_B \end{array}, \quad (25.27)$$

where $p_A + p_B + p_O = 1$.

$$P(x^\sigma | h^\sigma, \theta) = \begin{array}{c|cccccc} & AA & AO & BB & BO & OO & AB \\ \hline A & 1 & 1 & 0 & 0 & 0 & 0 \\ B & 0 & 0 & 1 & 1 & 0 & 0 \\ O & 0 & 0 & 0 & 0 & 1 & 0 \\ AB & 0 & 0 & 0 & 0 & 0 & 1 \end{array} \quad (25.28)$$

$$\theta = (p_A, p_B) \quad (25.29)$$

Multiplying the TPMs in Eqs.(25.27 and (25.28), we get

$$P(x^\sigma | \theta) = \begin{array}{c|c} & \\ \hline A & p_A^2 + 2p_A p_O (= \pi_A) \\ B & p_B^2 + 2p_B p_O (= \pi_B) \\ O & p_O^2 (= \pi_O) \\ AB & 2p_A p_B (= \pi_{AB}) \end{array} \quad (25.30)$$

Note that

$$P(\vec{x} | \theta) = \prod_{\sigma} P(x^\sigma | \theta) \quad (25.31)$$

$$= (\pi_A)^{N_A} (\pi_B)^{N_B} (\pi_O)^{N_O} (\pi_{AB})^{N_{AB}}, \quad (25.32)$$

where N_x for $x \in S_x = \{A, B, O, AB\}$ are the counts from the data. We can get estimates for the parameters p_A and p_B right here without doing EM. Just note that

$$\hat{\pi}_x = \frac{N_x}{N_+} \quad (25.33)$$

for $x \in S_x$, where $N_+ = \sum_x N_x$. Eqs.(25.33) give 4 quadratic equations that can be solved for the parameters p_A, p_B in terms of the observed counts N_x for $x \in S_x$.

If, instead, you want to find the optimum parameters p_A, p_B using EM, note that

$$Q(\theta|\theta^{(t)}) = \sum_{\vec{h}} P(\vec{h}|\theta^{(t)}) \ln P(\vec{x}, \vec{h}|\theta) \quad (25.34)$$

$$= \sum_{\vec{h}} \left[\prod_{\sigma} P(h^{\sigma}|\theta^{(t)}) \right] \ln \left[\prod_{\sigma} P(x^{\sigma}, h^{\sigma}|\theta) \right] \quad (25.35)$$

$$= \sum_{\sigma} \sum_{h^{\sigma}} P(h^{\sigma}|\theta^{(t)}) \ln P(x^{\sigma}, h^{\sigma}|\theta) \quad (25.36)$$

$$= \sum_{\sigma} \sum_{h^{\sigma}} P(h^{\sigma}|\theta^{(t)}) [\ln P(x^{\sigma}|h^{\sigma}, \theta) + \ln P(h^{\sigma}|\theta)] \quad (25.37)$$

$$= nsam \sum_{h^{\sigma}} P(h^{\sigma}|\theta^{(t)}) \ln P(h^{\sigma}|\theta). \quad (25.38)$$

25.3.3 Missing Data/Imputation

The previous example on blood genotypes and phenotypes assumed no missing data in compiling the counts N_x . But what if there is missing data? Can one still apply the EM algo in that case? Yes! See Chapter 53.

Chapter 26

Factor Graphs

Suppose $x^{nx} = (x_0, x_1, \dots, x_{nx-1})$. Consider a product

$$g(x^{nx}) = \prod_{\alpha=0}^{nf-1} f_\alpha(x_{A_\alpha}) \quad (26.1)$$

of scalar functions $f_\alpha : \underline{x}_{A_\alpha} \rightarrow \mathbb{R}$, where $A_\alpha \subset \{0, 1, \dots, nx - 1\}$. For instance, consider $g : S_{\underline{x}_0} \times S_{\underline{x}_1} \times S_{\underline{x}_2} \rightarrow \mathbb{R}$ defined by:

$$g(x_0, x_1, x_2) = f_0(x_0)f_1(x_0, x_1)f_2(x_0, x_1, x_2)f_3(x_1, x_2). \quad (26.2)$$

The **factor graph** for this function g is given by Fig.26.1.

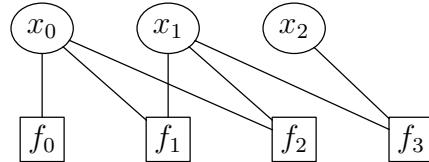


Figure 26.1: Factor graph for function g defined by Eq.(26.2).

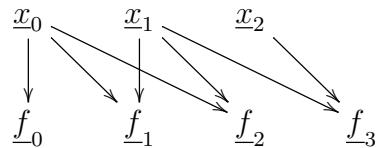


Figure 26.2: Bipartite bnet corresponding to factor graph Fig.26.1.

A **Markov Random Field (MRF)** is a statistical model whose probability distribution is of the form

$$P(x^{nx}) = \mathcal{N}(!x^{nx}) \prod_{\alpha} f_{\alpha}(x_{A_{\alpha}}) , \quad (26.3)$$

so it can be represented graphically by a factor graph. The factor functions f_{α} of a factor graph are called **potentials**.

One can map any factor graph (the “source”) to a special bipartite bnet (the “image”), as follows. Replace each x_i by $\underline{x}_i \in S_{x_i}$ for $i = 0, 1, \dots, nx - 1$ and each f_{α} by \underline{f}_{α} for $\alpha = 0, 1, \dots, nf - 1$. Then replace the connections (edges) of the factor graph by arrows from \underline{x}_i to \underline{f}_{α} .¹ For example, Fig.26.2 is the image bipartite bnet of the source factor graph Fig.26.1.

Let $\underline{x}^{nx} = (\underline{x}_0, \underline{x}_1, \dots, \underline{x}_{nx-1})$ and $\underline{f}^{nf} = (\underline{f}_0, \underline{f}_1, \dots, \underline{f}_{nf-1})$. Let² $f_{\alpha} \in \{0, 1\}$ for all α , and $y_{\alpha} = f_{\alpha}(x_{nb(\underline{f}_{\alpha})})$. Here we are using $nb(\underline{f}_{\alpha})$ to denote the neighborhood of node \underline{f}_{α} in the image bipartite bnet, and we are using x_S to denote $(x_i)_{i \in S}$. Without loss of generality, we will assume that $y_{\alpha} \in [0, 1]$ for all α . Then we define the TPMs, printed in blue, for the image bipartite bnet, as follows.

$$P(\underline{f}_{\alpha} | x_{nb(\underline{f}_{\alpha})}) = y_{\alpha} \delta(f_{\alpha}, 1) + [1 - y_{\alpha}] \delta(f_{\alpha}, 0) \quad (26.4)$$

for $\alpha = 0, 1, \dots, nf - 1$ and

$$P_{\underline{x}_i}(x_i) = \text{arbitrary prior} \quad (26.5)$$

for $i = 0, 1, \dots, nx - 1$. Note that

$$P(\underline{f}^{nf} = 1^{nf} | \underline{x}^{nx}) = \prod_{\alpha} y_{\alpha} . \quad (26.6)$$

Fig.26.3 gives another bipartite bnet, alternative to Fig.26.2, corresponding to factor graph Fig.26.1. In this new bnet, we replaced the $\underline{f}_{\alpha} \in \{0, 1\}$ nodes by $y_{\alpha} \in [0, 1]$ nodes. We also defined a new leaf node $\underline{y} \in \{0, 1\}$ with incoming arrows from all nodes \underline{y}_{α} . The TPMs, printed in blue, for the \underline{y}_{α} and \underline{y} nodes, are as follows.

$$P(\underline{y}_{\alpha} = y_{\alpha} | x_{nb(\underline{y}_{\alpha})}) = \delta(y_{\alpha}, f_{\alpha}(x_{nb(\underline{y}_{\alpha})})) \quad (26.7)$$

$$P(\underline{y} | \{y_{\alpha}\}_{\alpha=0}^{nf-1}) = y \prod_{\alpha} y_{\alpha} + (1 - y) \left(1 - \prod_{\alpha} y_{\alpha} \right) \quad (26.8)$$

¹Pointing arrows from the \underline{x}_i to the \underline{f}_{α} is more causal than the opposite because the \underline{x}_i preceed the \underline{f}_{α} in time.

²Note that we are using f_{α} to denote both a function $f_{\alpha}(\cdot)$ and a Boolean value. Which one we mean will be clear from context. f_{α} could also be used to denote, besides a function and a Boolean value, the real number $y_{\alpha} = f_{\alpha}(x_{nb(\underline{f}_{\alpha})})$. However, we won't be using it that third way in this chapter.

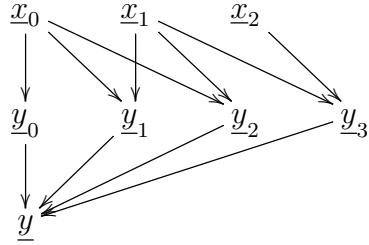


Figure 26.3: Another bipartite bnet, alternative to Fig.26.2, corresponding to factor graph Fig.26.1.

Note that

$$P(\underline{y} = 1 | \{y_\alpha\}_{\alpha=0}^{nf-1}) = \prod_{\alpha} y_\alpha \quad (26.9)$$

We've shown how to go from a factor graph to a bnet. Going from a bnet to a factor graph is also possible. Fig.26.4 gives a simple example. Bnet to factor graph conversion is used in the Junction Tree Algorithm (See Chapter 40).

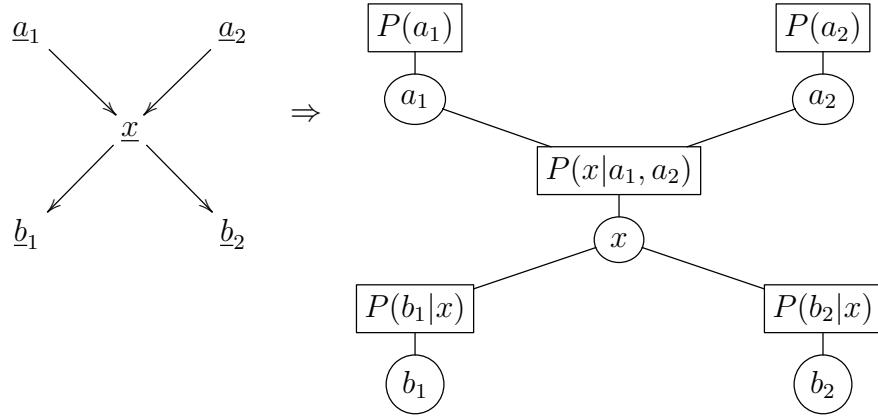


Figure 26.4: Example of converting a bnet to a factor graph.

Chapter 27

Frisch-Waugh-Lovell (FWL) theorem

The Frisch-Waugh-Lovell (FWL) theorem (see Ref.[99]) (mnemonic: FoWL Theorem) is a method used in Linear Regression (LR). It allows us to calculate a regression coefficient by doing LR with two residuals. Let's call them the **independent residual** and the **dependent or target residual**.¹ These two residuals are calculated with the help of two previously performed LR steps.

As in Section C.26 on LR, we will consider two cases: x^σ non-random, and x^σ random i.i.d..

27.1 FWL, assuming x^σ are non-random

Suppose

$$y = X_1\beta_1 + X_2\beta_2 + \epsilon \quad (27.1)$$

where

$$\begin{aligned} y, \epsilon &\in \mathbb{R}^{nsam} \\ X_a &\in \mathbb{R}^{nsam \times k_a}, \beta_a \in \mathbb{R}^{k_a} \text{ for } a = 1, 2 \\ \text{Define the matrices } U_1 \text{ and } A_1 \text{ by} \end{aligned}$$

$$U_1 = X_1(X_1^T X_1)^{-1} X_1^T \quad (27.2)$$

and

$$A_1 = I - U_1. \quad (27.3)$$

Note that

$$U_1 X_1 = X_1, \quad A_1 X_1 = 0 \quad (27.4)$$

¹Standard LR is performed with two features, the independent feature and the dependent or target feature.

(mnemonic: A_1 Annihilates X_1 , and U_1 acts like Unity on X_1).

Applying A_1 to Eq.(27.1) gives

$$A_1y = A_1X_2\beta_2 + A_1\epsilon \quad (27.5)$$

so we can estimate β_2 by

$$\boxed{\widehat{\beta}_2 = (A_1X_2)^{-1}A_1y} \quad (27.6)$$

27.2 FWL, assuming x^σ are random

Assume for simplicity that $k_1 = k_2 = 1$ in Eq.(27.1). Let $\beta_1 = \alpha \in \mathbb{R}$, $\beta_2 = \beta \in \mathbb{R}$. When the x^σ are random and i.i.d., and X_1, X_2 are replaced by the random variables $\underline{x}, \underline{d} \in \mathbb{R}$, Eq.(27.1) is equivalent to

$$\underline{y} = \alpha\underline{x} + \beta\underline{d} + \underline{u}_y. \quad (27.7)$$

As usual, we will denote the linear operator $\langle \underline{x}, \cdot \rangle^{-1} \langle \underline{x}, \cdot \rangle$ by a derivative operator:

$$\frac{\langle \underline{x}, \cdot \rangle}{\langle \underline{x}, \underline{x} \rangle} = \frac{d \cdot}{d\underline{x}}. \quad (27.8)$$

Define operators $U_{\underline{x}}$ and $A_{\underline{x}}$ by

$$U_{\underline{x}}(\cdot) = \underline{x} \frac{d \cdot}{d\underline{x}} \quad (27.9)$$

and

$$A_{\underline{x}} = 1 - U_{\underline{x}}. \quad (27.10)$$

Note that

$$U_{\underline{x}}\underline{x} = \underline{x}, \quad A_{\underline{x}}\underline{x} = 0. \quad (27.11)$$

If we apply $A_{\underline{x}}$ to Eq.(27.7), we get

$$A_{\underline{x}}\underline{y} = \beta A_{\underline{x}}\underline{d} + A_{\underline{x}}\underline{u}_y. \quad (27.12)$$

Hence,

$$\langle A_{\underline{x}}\underline{d}, A_{\underline{x}}\underline{y} \rangle = \beta \langle A_{\underline{x}}\underline{d}, A_{\underline{x}}\underline{d} \rangle. \quad (27.13)$$

Thus,

$$\boxed{\beta = \frac{d(A_{\underline{x}}y)}{d(A_{\underline{x}}d)}}. \quad (27.14)$$

Eqs.(27.6) and (27.14) constitute our statement of the FWL theorem. They are the main result of this chapter. But before ending this chapter, let us show how the FWL theorem for random x^σ can be interpreted graphical using LDEN diagrams (LDEN diagrams are discussed in Chapter 44).

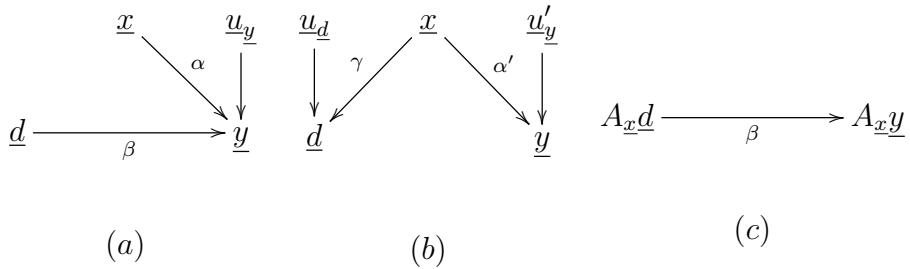


Figure 27.1: LDEN diagrams for discussing the FWL theorem.

Fig.27.1(a) is a graphical representation of the following equation:

$$y = \alpha \underline{x} + \beta \underline{d} + \underline{u}_y. \quad (27.15a)$$

Likewise, 27.1(b) represents

$$\begin{cases} \underline{d} = \gamma \underline{x} + \underline{u}_d \\ \underline{y} = \alpha' \underline{x} + \underline{u}'_y \end{cases} \quad (27.15b)$$

and 27.1(c) represents

$$A_{\underline{x}}y = \beta A_{\underline{x}}d. \quad (27.15c)$$

Applying $d/d\underline{x}$ to Eq.(27.15a) yields

$$\frac{dy}{d\underline{x}} = \alpha + \beta \frac{dd}{d\underline{x}}. \quad (27.16)$$

Applying $d/d\underline{x}$ to Eq.(27.15b) yields

$$\gamma = \frac{dd}{d\underline{x}}, \quad \alpha' = \frac{dy}{d\underline{x}}. \quad (27.17)$$

Therefore,

$$\alpha' = \alpha + \beta \gamma. \quad (27.18)$$

Applying $A_{\underline{x}}$ to Eq.(27.15b) yields

$$A_{\underline{x}}\underline{y} = \left(1 - \underline{x}\frac{d}{dx}\right)\underline{y} = \underline{y} - \alpha'\underline{x} \quad (27.19)$$

and

$$A_{\underline{x}}\underline{d} = \left(1 - \underline{x}\frac{d}{dx}\right)\underline{d} = \underline{d} - \gamma\underline{x}. \quad (27.20)$$

Therefore, Eq.(27.15c) is equivalent to

$$\underline{y} - \alpha'\underline{x} = \beta(\underline{d} - \gamma\underline{x}). \quad (27.21)$$

Chapter 28

Frontdoor Adjustment Formula

The frontdoor (FD) adjustment formula is proven in Chapter 19 from the rules of Do Calculus. The goal of this chapter is to give examples of the use of that theorem. We will restate the theorem in this chapter, sans proof. There is no need to understand the theorem's proof in order to use it. However, you will need to skim Chapter 19 in order to familiarize yourself with the notation used to state the theorem. This chapter also assumes that you are comfortable with the rules for checking for d-separation. Those rules are covered in Chapter 21.

Suppose that we have access to data that allows us to estimate a probability distribution $P(x., m., y.)$. Hence, the variables $\underline{x}., \underline{m}., \underline{y}.$ are ALL observed (i.e., not hidden). Then we say that the frontdoor $\underline{m}.$ satisfies the **frontdoor adjustment criterion** relative to $(\underline{x}., \underline{y}.)$ if

1. All directed paths from $\underline{x}.$ to $\underline{y}.$ are intercepted by (i.e., have a node in) $\underline{m}..$
2. All backdoor paths from $\underline{x}.$ to $\underline{m}.$ are blocked.
3. All backdoor paths from $\underline{m}.$ to $\underline{y}.$ are blocked by conditioning on $\underline{x}..$

Claim 52 *Frontdoor Adjustment Formula*

If $\underline{m}.$ satisfies the frontdoor criterion relative to $(\underline{x}., \underline{y}.)$, and $P(x., m.) > 0$, then

$$P(y.|D\underline{x}. = x.) = \sum_{m.} \underbrace{\left[\sum_{x'.} P(y.|x'., m.) P(x'.) \right]}_{P(y.|D\underline{m}. = m.)} \underbrace{P(m.|x.)}_{P(m.|D\underline{x}. = x.)} \quad (28.1)$$

$$= \sum_{x'.} \sum_{m.} \begin{array}{c} \diagdown \\ x. \longrightarrow \sum m. \longrightarrow y. \end{array} \quad (28.2)$$

where $\sum x'.$ and $\sum m.$ means nodes $\underline{x}'.$ and $\underline{m}.$ are summed over.

proof: See Chapter 19.

QED

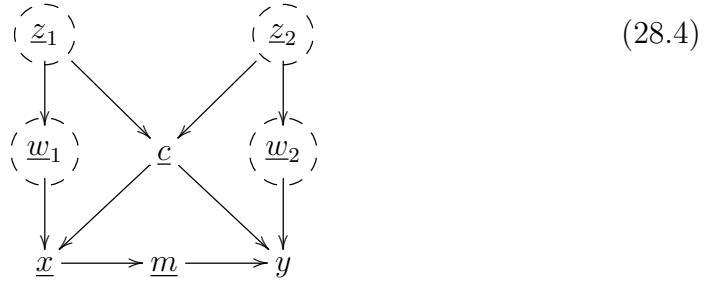
28.1 Examples

1.



If $x_+ = \underline{x}$, $m_+ = \underline{m}$ and $y_+ = \underline{y}$, then the FD criterion is satisfied. Can't satisfy backdoor criterion because \underline{c} unobserved so can't condition on it to block backdoor path $\underline{x} - \underline{c} - \underline{y}$.

2.



If $x_+ = \underline{x}$, $m_+ = \underline{m}$ and $y_+ = \underline{y}$, then the FD criterion is satisfied. Can't satisfy backdoor criterion because to block backdoor path $\underline{x} - \underline{c} - \underline{y}$, need to condition on \underline{c} but if this is true, then long path $\underline{x} - \underline{w}_1 - z_1 - \underline{c} - z_2 - \underline{w}_2 - \underline{y}$ becomes unblocked.

Chapter 29

G-formula (Sequential Backdoor Adjustment Formula)



Figure 29.1: Piscina Mirabilis, ancient Roman cistern in Naples

This chapter is based on Ref.[51] by Pearl and Robins.

A **g-formula**¹ is any formula that defines recursively the full probability distribution of a bnet. In other words, it's a recursive definition of a Dynamical Bayesian Network.²

The goal of this chapter is to generalize the backdoor adjustment formula (see Chapter 4) from a query $P(y|\mathcal{D}x = x)$ with a single do node to a query $P(y|\mathcal{D}\underline{x}^n = x^n)$

¹It's not clear from the literature what the "g" stands for. I assume it stands for "generating".

²Dynamical Bayesian Networks are discussed in Chapter 24.

with multiple do nodes. The resulting generalized adjustment formula is called a **sequential backdoor (SBD) adjustment formula** and it is associated with an **SBD g-formula**.

For $n = 1, 2, 3 \dots$, define

$$\mathcal{Q}(y|x^n) = \sum_{z^n} P(y|x^n, z^n) \prod_{t=0}^{n-1} P(z_t|x_{<t}, z_{<t}) \quad (29.1)$$

$$= \begin{array}{c} \sum_{z^n} z^n \\ \downarrow \\ y \\ \uparrow \\ x^n \end{array} \quad \begin{array}{c} z_{<t} \longrightarrow z_t \\ \nearrow \\ x_{<t} \end{array} \quad (29.2)$$

For $n = 1$,

$$\mathcal{Q}(y|x_0) = \begin{array}{c} \sum_{z_0} z_0 \\ \downarrow \\ y \\ \uparrow \\ x_0 \end{array} = \begin{array}{c} \sum z_0 \\ \searrow \\ y \\ \nearrow \\ x_0 \end{array} . \quad (29.3)$$

For $n = 2$,

$$\mathcal{Q}(y|x^2) = \begin{array}{c} \sum_{z^2} (z_0, z_1) P(z_0) z_0 \longrightarrow z_1 \\ \downarrow \\ y \\ \uparrow \\ (x_0, x_1) \end{array} \quad (29.4)$$

$$= \begin{array}{c} \sum z_0 \longrightarrow \sum z_1 \\ \nearrow \searrow \\ x_0 \quad x_1 \\ \nearrow \searrow \\ y \end{array} . \quad (29.5)$$

For $n = 3$,

$$\begin{aligned} & \sum_{z^3} (z_0, z_1, z_2) P(z_0) z_0 \rightarrow z_1 (z_0, z_1) \rightarrow z_2 \\ \mathcal{Q}(y|x^3) = & \downarrow y \\ & (x_0, x_1, x_2) \end{aligned} \quad (29.6)$$

$$\begin{aligned} & \sum_{z^3} (z_0, z_1, z_2) P(z_0) z_0 \xrightarrow{\text{double}} z_1 \xrightarrow{\text{double}} z_2 \\ = & \downarrow y \\ & (x_0, x_1, x_2) \end{aligned} \quad (29.7)$$

$$\begin{aligned} & \sum z_0 \xrightarrow{\text{double}} \sum z_1 \xrightarrow{\text{double}} \sum z_2 \\ = & \downarrow y \\ & x_0 \quad x_1 \quad x_2 \end{aligned} \quad (29.8)$$

Suppose that we have access to data that allows us to estimate a probability distribution $P(x^n, y, z^n)$. Hence, the variables $\underline{x}^n, \underline{y}, \underline{z}^n$ are ALL observed (i.e, not hidden). Then we say that the the multinode of “covariates” \underline{z}^n satisfies the **sequential backdoor (SBD) adjustment criterion** relative to $(\underline{x}^n, \underline{y})$ if for all $t \in \{0, 1, \dots, n-1\}$,

1. $\underline{y} \perp \underline{x}_t \mid \underbrace{(\underline{x}_0, \underline{x}_1, \dots, \underline{x}_{t-1}, \underline{z}_0, \underline{z}_1, \dots, \underline{z}_t)}_{\text{Past of } \underline{x}_t}$ in $\mathcal{L}_{\underline{x}_t} \mathcal{D}_{\underline{x}_{t+1}, \underline{x}_{t+2}, \dots, \underline{x}_{n-1}} G$.
2. $\underline{z}_t \cap de(\underline{x}_t) = \emptyset$.

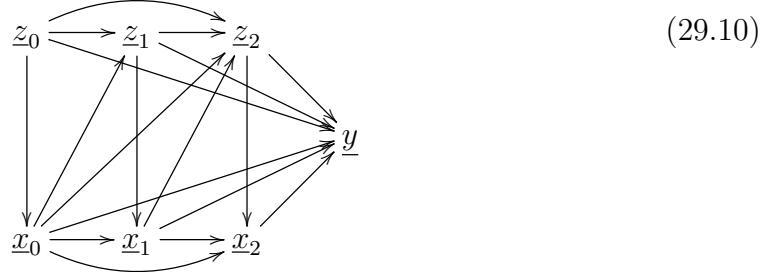
Claim 53 (SBD Adjustment Formula)

If \underline{z}^n satisfies the sequential backdoor criterion relative to $(\underline{x}^n, \underline{y})$, then

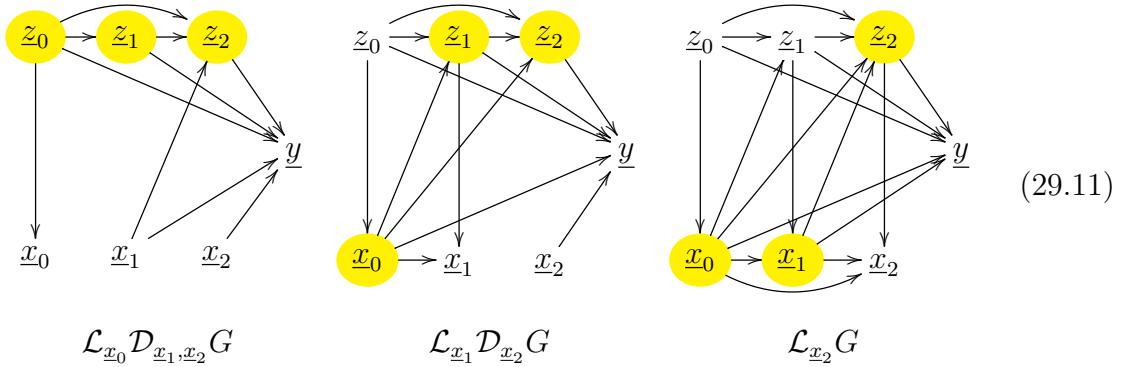
$$P(y | \mathcal{D} \underline{x}^n = x^n) = \mathcal{Q}(y | x^n), \quad (29.9)$$

where $\mathcal{Q}(y | x^n)$ is defined by Eq.(29.2).

proof: If z^n satisfies the SBD criterion relative to $(\underline{x}^n, \underline{y})$, then $\underline{x}^n, \underline{y}, \underline{z}^n$ might have the following structure for $n = 3$.



One can check using the following 3 auxiliary bnets that bnet Eq.(29.10) satisfies the SBD criterion. Note that conditioned nodes are shaded yellow.



See Claim 47 for a proof of this claim for the special case Eq.(29.10).

QED

Chapter 30

Gaussian Nodes with Linear Dependence on Parents

Bnet nodes that have a Gaussian TPM with a linear dependence on their parent nodes (GLP) are a very popular way of modeling continuous nodes of bnets. A convenient aspect of them is that their parents can be discrete or continuous nodes, and their children can be discrete or continuous nodes too. Also, they can be learned easily from the data because their parameters can be expressed in terms of two node covariances. For these reasons, they are commonly used when doing structure learning of bnets with continuous nodes (see Chapter 78).

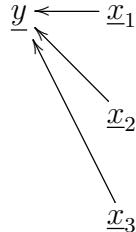


Figure 30.1: GLP node \underline{y} with 3 parent nodes $\underline{x}^3 = (\underline{x}_1, \underline{x}_2, \underline{x}_3)$.

Recall our notation for a Gaussian distribution:

$$\mathcal{N}(x; \mu, \sigma^2) = \frac{1}{\sigma\sqrt{2\pi}} e^{\frac{-(x-\mu)^2}{2\sigma^2}}, \quad (30.1)$$

where $x, \mu \in \mathbb{R}$ and $\sigma > 0$.

A GLP node \underline{y} with n parents $\underline{x}^n = (\underline{x}_1, \underline{x}_2, \dots, \underline{x}_n)$ has the following TPM:

$$P(\underline{y}|\underline{x}^n) = \mathcal{N}(\underline{y}; \beta_0 + \beta^{nT} \underline{x}^n, \sigma^2) \quad (30.2)$$

where $\underline{y}, \beta_0 \in \mathbb{R}$ and $\sigma^2 > 0$, and where $\underline{x}^n, \beta^n \in \mathbb{R}^n$ are **column vectors**. The T in β^{nT} stands for transpose. Any \underline{x}_i can have a discrete set of states as long as

they are real valued and ordinal (ordered by size). Fig.30.1 shows a diagrammatic representation of a GLP node with 3 parents.

Note that as $\sigma \rightarrow 0$, a GLP node becomes deterministic. In fact, it becomes a neural net node with a linear activation function.

An equivalent way of defining a GLP node \underline{y} is in terms of a random variable equation expressing \underline{y} as a hyperplane function of the parents \underline{x}^n plus a Gaussian noise variable. Define a cfit $\widehat{\underline{y}}$ of a “true value” \underline{y} by

$$\widehat{\underline{y}} = \beta_0 + \beta^{nT} \underline{x}^n \quad (30.3a)$$

and

$$\underline{y} = \widehat{\underline{y}} + \underline{\epsilon} \quad (30.3b)$$

where the residual $\underline{\epsilon}$ satisfies

$$P(\epsilon) = \mathcal{N}(\epsilon; 0, \sigma^2) \quad (30.3c)$$

and

$$\langle \underline{x}^n, \underline{\epsilon} \rangle = 0 . \quad (30.3d)$$

The notation $\langle \underline{x}, \underline{y} \rangle$ for the covariance of random variables \underline{x} and \underline{y} is explained in Chapter C.

Claim 54 *The parameters of a GLP node can be expressed in terms of 2-node covariances. Specifically,*

$$\beta^n = \langle \underline{x}^n, \underline{x}^{nT} \rangle^{-1} \langle \underline{y}, \underline{x}^n \rangle \quad (30.4)$$

$$\beta_0 = \langle \underline{y} \rangle - \beta^{nT} \langle \underline{x}^n \rangle \quad (30.5)$$

$$\sigma^2 = \langle \underline{y}, \underline{y} \rangle - \beta^{nT} \langle \underline{x}^n, \underline{y} \rangle \quad (30.6)$$

proof:

Note that $\langle \underline{x}^n, \underline{x}^{nT} \rangle^T = \langle \underline{x}^n, \underline{x}^{nT} \rangle$ and $\langle \underline{y}, \underline{x}^{nT} \rangle^T = \langle \underline{y}, \underline{x}^n \rangle$.

$$\langle \underline{y}, \underline{x}^{nT} \rangle = \beta^{nT} \langle \underline{x}^n, \underline{x}^{nT} \rangle \quad (30.7)$$

$$\langle \underline{y}, \underline{x}^n \rangle = \langle \underline{x}^n, \underline{x}^{nT} \rangle \beta^n \quad (30.8)$$

$$\beta^n = \langle \underline{x}^n, \underline{x}^{nT} \rangle^{-1} \langle \underline{y}, \underline{x}^n \rangle \quad (30.9)$$

$$\langle \underline{y} \rangle = \beta_0 + \beta^{nT} \langle \underline{x}^n \rangle \quad (30.10)$$

$$\langle \underline{y}, \underline{y} \rangle = \langle \beta_0 + \beta^{nT} \underline{x}^n + \underline{\epsilon}, \underline{y} \rangle \quad (30.11)$$

$$= \beta^{nT} \langle \underline{x}^n, \underline{y} \rangle + \sigma^2 \quad (30.12)$$

QED

Let D=Discrete, GLP=Gaussian with Linear dependence in Parents

The following arrows are possible in a bnet.

- $GLP \leftarrow GLP$

- $GLP \leftarrow D$

Pass to GLP a separate set of regression coefficients β_0, β^n and variance σ^2 for each state of D. If D is called \underline{d} , let

$$P(y|(x^n)_d, d) = \mathcal{N}(y; (\beta_0)_d + (\beta^{nT})_d(x^n)_d, \sigma_d^2) \quad (30.13)$$

for each $d \in S_{\underline{d}}$.

- $D \leftarrow GLP$

If D expects a continuous parent, no need to preprocess GLP output. If D expects a discrete parent, break the interval $[a, b]$ that contains most of the range of the GLP node into sub-intervals and assign a discrete label to each subinterval.

- $D \leftarrow D$

Chapter 31

Generalized Linear Model (GLM)

This chapter is based on chapter 4 of Ref.[1].

31.1 Exponential Family of Distributions

The **Exponential Family (EF)** of probability distributions is defined by

$$P(y|\theta, \phi) = \exp\left(\frac{\theta y - b(\theta)}{a(\phi)} + c(y, \phi)\right) \quad (31.1)$$

In this chapter, we will denote averages over $y|\theta, \phi$ by angular brackets

$$E_{y|\theta, \phi}[f(y)] = \sum_y P(y|\theta, \phi)f(y) = \langle f(y) \rangle \quad (31.2)$$

As usual in this book, let S_x denote the set of values that the random variable x can take. We will assume that S_y for EF can be either a discrete or a continuous subset of \mathbb{R} ,¹ but S_θ must be continuous. When S_y is a discrete subset of \mathbb{R} , $P(y|\theta, \phi)$ will denote a probability distribution, whereas when S_y is continuous, it will denote a probability density.

Claim 55

$$\mu = \langle y \rangle = b'(\theta) \quad (31.3)$$

$$\sigma^2 = \langle y, y \rangle = a(\phi)b''(\theta) \quad (31.4)$$

proof:

¹By a “continuous set” we mean a finite set of intervals each of which has non-zero length.

$$0 = \partial_\theta \int_{-\infty}^{\infty} dy P(y|\theta, \phi) \quad (31.5)$$

$$= \int_{-\infty}^{\infty} dy \frac{1}{a} [y - b'(\theta)] P(y|\theta, \phi) \quad (31.6)$$

$$= \frac{1}{a} [\langle \underline{y} \rangle - b'(\theta)] \quad (31.7)$$

$$0 = \partial_\theta^2 \int_{-\infty}^{\infty} dy P(y|\theta, \phi) \quad (31.8)$$

$$= \int_{-\infty}^{\infty} dy \left\{ \frac{-b''(\theta)}{a} + \frac{1}{a^2} [y - b'(\theta)]^2 \right\} P(y|\theta, \phi) \quad (31.9)$$

Hence,

$$\langle [y - b'(\theta)]^2 \rangle = ab''(\theta) \quad (31.10)$$

QED

Note that the Normal Distribution belongs to the EF. Indeed,

$$\mathcal{N}(y; \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y - \mu)^2}{2\sigma^2}\right) \quad (31.11)$$

$$= \exp\left(\frac{-y^2 + 2\mu y - \mu^2}{2\sigma^2} - \ln \sqrt{2\pi\sigma^2}\right) \quad (31.12)$$

$$= \exp\left(\frac{-\frac{1}{2}y^2 + \mu y - \frac{1}{2}\mu^2}{\sigma^2} - \ln \sqrt{2\pi\sigma^2}\right) \quad (31.13)$$

So, for the Normal Distribution, $\theta = \mu$, $a = \sigma^2$, $b = \mu^2/2$, $b' = \mu$, $b'' = 1$.

EF can be defined for an ensemble $\{y_\sigma : \sigma \in \Sigma\}$ of independent (but not necessarily identically distributed) random variables y_σ describing individuals σ of a population Σ .

$$P(\vec{y}|\vec{\theta}, \phi) = \prod_{\sigma} \exp\left(\frac{\theta_\sigma y_\sigma - b(\theta_\sigma)}{a(\phi)} + c(y_\sigma, \phi)\right) \quad (31.14)$$

$$\mu_\sigma = \langle \underline{y}_\sigma \rangle = \partial_{\theta_\sigma} b(\theta_\sigma) \quad (31.15)$$

$$\langle \underline{y}_\sigma, \underline{y}_{\sigma'} \rangle = \delta(\sigma, \sigma') a(\phi) \partial_{\theta_\sigma}^2 b(\theta_\sigma) \quad (31.16)$$

31.2 GLM

The Generalized Linear Model (GLM) is a statistical model for an ensemble of independent (but not necessarily identically distributed) random variables \underline{y}_σ . GLM consists of 3 parts:

1. Exponential Family

Model \underline{y}_σ by probability distribution of exponential family (EF).

2. Linear predictor

In EF, replace θ_σ by the **linear predictor** $X_\sigma^T \beta = \sum_i X_{\sigma,i} \beta_i$. $X_\sigma^T \beta$ is commonly denoted by η_σ , but I will avoid that notation because I think the results are much clearer when expressed in the more explicit notation $X_\sigma^T \beta$ instead of η_σ .

3. Link Function

$$\mu_\sigma = \langle \underline{y}_\sigma \rangle = \partial_{X_\sigma^T \beta} b \quad (31.17)$$

$$X_\sigma^T \beta = g(\mu_\sigma) = \hat{\theta}(\mu_\sigma) \quad (31.18)$$

$$\mu_\sigma = g^{-1}(X_\sigma^T \beta) = \hat{\mu}(X_\sigma^T \beta) \quad (31.19)$$

$\hat{\theta}() = g()$ is called the **link function**.

Note that in Linear Regression (LR), we consider independent (but not necessarily identically distributed) random variables $\underline{y}_\sigma \in \mathbb{R}$ that satisfy

$$\underline{y}_\sigma = X_\sigma^T \beta + \epsilon_\sigma \quad (31.20)$$

where

$$\langle \epsilon_\sigma \rangle = 0, \quad \langle \epsilon_\sigma, \epsilon_\sigma \rangle = \delta(\sigma, \sigma') \sigma_\sigma^2 \quad (31.21)$$

Equivalently, one states that

$$\underline{y}_\sigma \sim \mathcal{N}(\mu_\sigma = X_\sigma^T \beta, \sigma_\sigma^2) \quad (31.22)$$

Hence, for LR, the link function and its inverse are the identity map.

For the Bernoulli distribution with $y_\sigma \in \{0, 1\}$,

$$Ber(y_\sigma; p_\sigma) = p_\sigma^{y_\sigma} (1 - p_\sigma)^{1-y_\sigma} \quad (31.23)$$

$$= \exp(y_\sigma \ln p_\sigma + (1 - y_\sigma) \ln(1 - p_\sigma)) \quad (31.24)$$

$$= \exp \left(y_\sigma \underbrace{\ln \left(\frac{p_\sigma}{1 - p_\sigma} \right)}_{X_\sigma^T \beta} + \underbrace{\ln(1 - p_\sigma)}_{-b} \right) \quad (31.25)$$

$\mu_\sigma = p_\sigma$, and $X_\sigma^T \beta = \text{logit}(\mu_\sigma)$ so $\mu_\sigma = \text{smoid}(X_\sigma^T \beta)$.
 $g() = \text{logit}()$ and $g^{-1}() = \text{smoid}()$

$$a = 1 \quad (31.26)$$

$$b = -\ln(1 - \mu_\sigma) \quad (31.27)$$

$$= -\ln \left(1 - \frac{1}{1 + e^{-X_\sigma^T \beta}} \right) \quad (31.28)$$

$$= -\ln \left(\frac{e^{-X_\sigma^T \beta}}{1 + e^{-X_\sigma^T \beta}} \right) \quad (31.29)$$

$$= \ln(1 + e^{X_\sigma^T \beta}) \quad (31.30)$$

$$\partial_{X_\sigma^T \beta} b = \frac{e^{X_\sigma^T \beta}}{1 + e^{X_\sigma^T \beta}} \quad (31.31)$$

$$= \text{smoid}(X_\sigma^T \beta) \quad (31.32)$$

$$= \mu_\sigma \quad (31.33)$$

Table 31.1 gives various probability distributions and their natural link functions, for cases where the link function is simple and easy to invert.

GLM is a generalization of LR. Recall some of the main results of LR:

$$y = X\beta + \epsilon \quad (31.34)$$

$$\langle \epsilon \rangle = 0, \quad \langle \epsilon, \epsilon^T \rangle = \sigma^2 I_N \quad (31.35)$$

where I_N is the $N = |\Sigma|$ dimensional unit matrix.

The Maximum Likelihood Estimate (MLE) for β is

$$\hat{\beta} = (X^T X)^{-1} X^T y. \quad (31.36)$$

prob. distribution	link function $X_\sigma^T \beta = g(\mu_\sigma)$	$\mu_\sigma = g^{-1}(X_\sigma^T \beta)$
Normal $y_\sigma \in (-\infty, +\infty)$	$X_\sigma^T \beta = \mu_\sigma$ (g = identity map)	$\mu_\sigma = X_\sigma^T \beta$
Exponential $y_\sigma \in (0, +\infty)$	$X_\sigma^T \beta = -\frac{1}{\mu_\sigma}$	$\mu_\sigma = -\frac{1}{X_\sigma^T \beta}$
Poisson $y_\sigma \in \{0, 1, 2, \dots\}$	$X_\sigma^T \beta = \ln \mu_\sigma$	$\mu_\sigma = \exp(X_\sigma^T \beta)$
Bernoulli $y_\sigma \in \{0, 1\}$	$X_\sigma^T \beta = \text{lodds}(\mu_\sigma)$	$\mu_\sigma = \text{smoid}(X_\sigma^T \beta)$

Table 31.1: Various probability distributions and their natural link functions.

The covariance of $\hat{\beta}$ is

$$\langle \hat{\beta}, \hat{\beta}^T \rangle = \langle X^T X)^{-1} X^T y, y^T X (X^T X)^{-1} \rangle \quad (31.37)$$

$$= \sigma^2 (X^T X)^{-1} \quad (31.38)$$

Next we will try to find analogous results for GLM. We will give (1) a numerical method for calculating an estimate $\hat{\beta}$ and (2) an asymptotic expression for $\langle \hat{\beta}, \hat{\beta}^T \rangle$.

Let

$$LL = \sum_{\sigma} LL_{\sigma} \quad (31.39)$$

where

$$LL_{\sigma} = LL_{y_{\sigma} | \theta_{\sigma}} \quad (31.40)$$

$$= \frac{y_{\sigma} \theta_{\sigma} - b(\theta_{\sigma})}{a(\phi)} + c(y_{\sigma}, \phi) \quad (31.41)$$

The Newton-Raphson method for calculating an estimate $\hat{\beta}$ is as follows. Let $u^T = [\frac{\partial \langle LL \rangle}{\partial \beta_j}]_{j=0,1,2,\dots}$

$H = [H_{j,j'}]$, $H_{j,j'} = \frac{\partial^2 \langle LL \rangle}{\partial \beta_j \partial \beta_{j'}}$. H is called the **Hessian matrix**

For $t = 0, 1, 2, \dots$, consider the Taylor expansion to second order of $\langle LL \rangle(\beta)$ about the point $\beta = \beta^{(t)}$

$$\langle LL \rangle(\beta) \approx \langle LL \rangle(\beta^{(t)}) + u^{(t)T}(\beta - \beta^{(t)}) + \frac{1}{2}(\beta - \beta^{(t)})^T H^{(t)}(\beta - \beta^{(t)}) \quad (31.42)$$

By Section C.30, we have,

$$0 = \frac{\partial \langle LL \rangle(\beta)}{\partial \beta^T} = u^{(t)} + H^{(t)}(\beta - \beta^{(t)}) \quad (31.43)$$

This last equation suggests the recursion

$$\beta^{(t+1)} = \beta^{(t)} - (H^{(t)})^{-1}u^{(t)}. \quad (31.44)$$

Fig.31.1 gives a graphical representation of one cycle of this recursion.

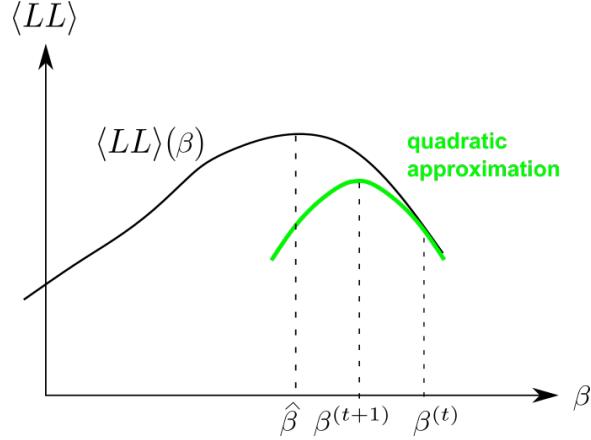


Figure 31.1: One cycle of Newton-Raphson method for calculating an estimate $\hat{\beta}$ for the GLM.

Claim 56

$$\frac{\partial LL_\sigma}{\partial \beta_j} = \frac{[y_\sigma - b'(X_\sigma^T \beta)]}{\langle y_\sigma, y_\sigma \rangle} \frac{\partial \hat{\mu}_\sigma}{\partial X_\sigma^T \beta} X_{\sigma,j} \quad (31.45)$$

proof:

$$\frac{\partial LL_\sigma}{\partial \beta_j} = \frac{\partial LL_\sigma}{\partial \theta_\sigma} \frac{\partial \theta_\sigma}{\partial \mu_\sigma} \frac{\partial \hat{\mu}_\sigma}{\partial X_\sigma^T \beta} \frac{\partial X_\sigma^T \beta}{\partial \beta_j} \quad (31.46)$$

$$\frac{\partial LL_\sigma}{\partial \theta_\sigma} = \frac{y_\sigma - b'(\theta_\sigma)}{a(\phi)} \quad (31.47)$$

$$\frac{\partial \mu_\sigma}{\partial \theta_\sigma} = b''(\theta_\sigma) = \frac{\langle y_\sigma, y_\sigma \rangle}{a(\phi)} \quad (31.48)$$

$$\frac{\partial \hat{\mu}_\sigma}{\partial X_\sigma^T \beta} = (g^{-1})'(X_\sigma^T \beta) \quad (31.49)$$

$$\frac{\partial X_\sigma^T \beta}{\partial \beta_j} = X_{\sigma,j} \quad (31.50)$$

$$\frac{\partial LL_\sigma}{\partial \beta_j} = \frac{[y_\sigma - b'(\theta_\sigma)]}{\langle y_\sigma, y_\sigma \rangle} \frac{\partial \hat{\mu}_\sigma}{\partial X_\sigma^T \beta} X_{\sigma,j} \quad (31.51)$$

QED

Claim 57 (*Asymptotic expression for $\langle \hat{\beta}, \hat{\beta}^T \rangle$ for GLM*)

$$\langle \hat{\beta}, \hat{\beta}^T \rangle \rightarrow \mathcal{I}^{-1} \quad (\text{asymptotic covariance}) \quad (31.52)$$

(Hence, more information \mathcal{I} yields a smaller variance.) where

$$\mathcal{I} = X^T W X \quad (31.53)$$

and

$$W_{\sigma,\sigma'} = \left[\left(\frac{\partial \hat{\mu}}{\partial X_\sigma^T \beta} \right)^2 \frac{\delta(\sigma, \sigma')}{\langle y_\sigma, y_\sigma \rangle} \right]_{\beta=\hat{\beta}} \quad (31.54)$$

\mathcal{I} is called the **information matrix**.

proof:

$$\left\langle \frac{\partial LL_\sigma}{\partial \beta_j} \frac{\partial LL_\sigma}{\partial \beta_{j'}} \right\rangle = \left\langle \frac{[y_\sigma - b'(X_\sigma^T \beta)]}{\langle y_\sigma, y_\sigma \rangle} \frac{\partial \hat{\mu}_\sigma}{\partial X_\sigma^T \beta} X_{\sigma,j} \frac{[y_\sigma - b'(X_\sigma^T \beta)]}{\langle y_\sigma, y_\sigma \rangle} \frac{\partial \hat{\mu}_\sigma}{\partial X_\sigma^T \beta} X_{\sigma,j'} \right\rangle \quad (31.55)$$

$$= \left[\frac{\partial \hat{\mu}_\sigma}{\partial X_\sigma^T \beta} \right]^2 \frac{X_{\sigma,j} X_{\sigma,j'}}{\langle y_\sigma, y_\sigma \rangle} \quad (31.56)$$

$$\sum_\sigma \left\langle \frac{\partial LL_\sigma}{\partial \beta_j} \frac{\partial LL_\sigma}{\partial \beta_{j'}} \right\rangle = (X^T W X)_{j,j'} \quad (31.57)$$

By Section C.30, we have

$$\left\langle \frac{\partial^2 LL_\sigma}{\partial \beta_j \partial \beta_{j'}} \right\rangle = - \left\langle \frac{\partial LL_\sigma}{\partial \beta_j} \frac{\partial LL_\sigma}{\partial \beta_{j'}} \right\rangle \quad (31.58)$$

Summing both sides of the last equation over σ , we find

$$\left\langle \frac{\partial^2 LL}{\partial \beta_j \partial \beta_{j'}} \right\rangle = -(X^T W X)_{j,j'} \quad (31.59)$$

According to Section C.30, we have

$$\langle \hat{\beta}, \hat{\beta}^T \rangle \rightarrow (X^T W X)^{-1} \quad (31.60)$$

QED

Chapter 32

Generative Adversarial Networks (GANs)

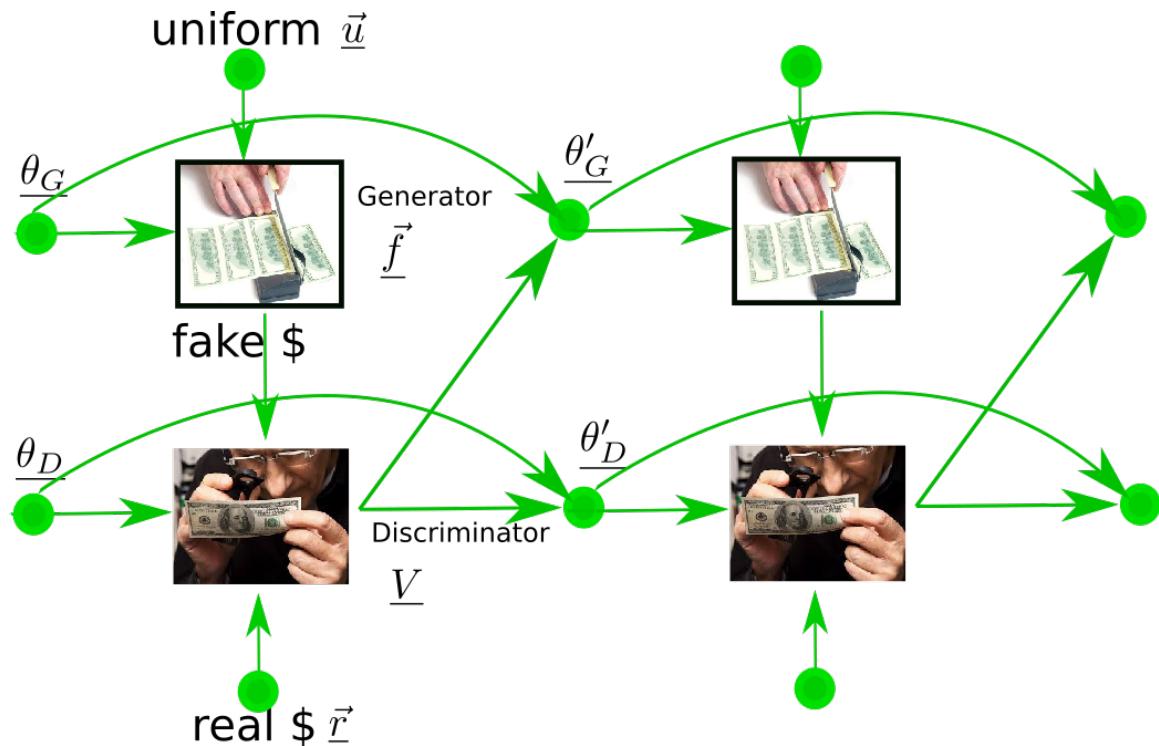


Figure 32.1: Generative Adversarial Network (GAN)

Original GAN, Ref.[14](2014).

Generator G (counterfeiter) generates samples \vec{f} of fake money and submits them to Discriminator D (Treasury agent). D also gets samples \vec{r} of real money. D submits verdict $V \in [0, 1]$. G depends on parameter θ_G and D on parameter θ_D . Verdict V and initial θ_G, θ_D are used to get new parameters θ'_G, θ'_D . Process is repeated

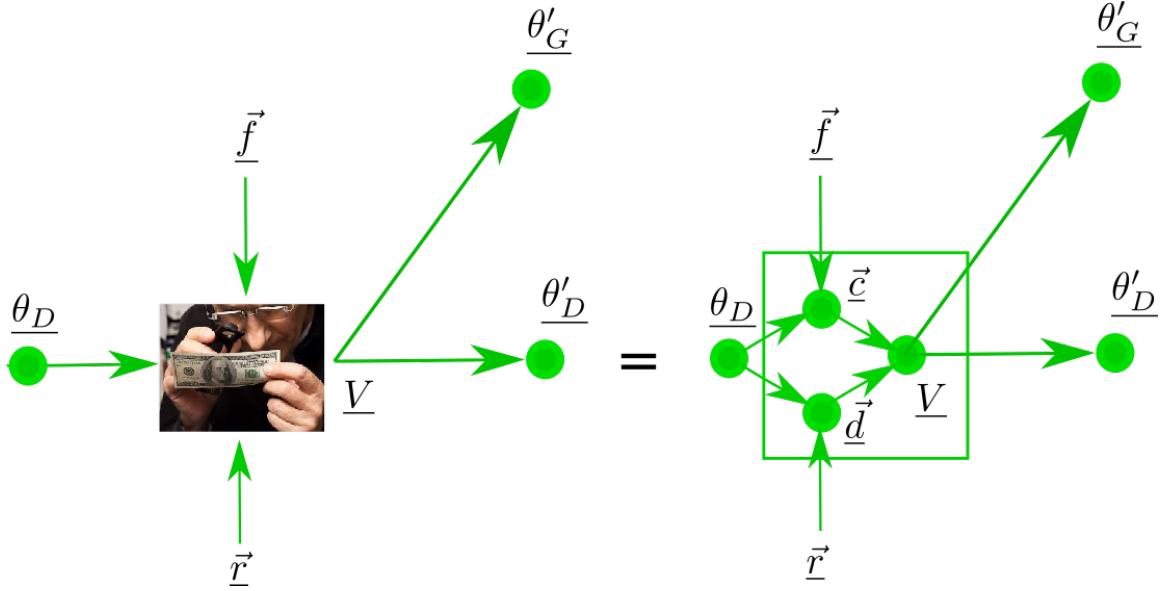


Figure 32.2: Discriminator node \underline{V} in Fig.32.1 can be split into 3 nodes \vec{c} , \vec{d} and \underline{V} .

(Dynamical Bayesian Network) until saddle point in $V(\theta_G, \theta_D)$ is reached. D makes G better and vice versa. Zero-sum game between D and G .

Let \mathcal{D} be the domain of $D(\cdot, \theta_D)$. Assume that for any $x \in \mathcal{D}$,

$$0 \leq D(x, \theta_D) \leq 1 . \quad (32.1)$$

For any $S \subset \mathcal{D}$, define

$$\sum_{x \in S} D(x, \theta_D) = \lambda(S, \theta_D) . \quad (32.2)$$

In general, $G(\cdot, \theta_G)$ need not be real valued.

Assume that for every $u \in S_u$, $G(u, \theta_G) = f \in S_f \subset \mathcal{D}$. Define

$$\overline{D}(f, \theta_D) = 1 - D(f, \theta_D) . \quad (32.3)$$

Note that

$$0 \leq \overline{D}(f, \theta_D) \leq 1 . \quad (32.4)$$

Define:

$$V(\theta_G, \theta_D) = \sum_r P(r) \ln D(r, \theta_D) + \sum_u P(u) \ln \overline{D}(G(u, \theta_G), \theta_D) . \quad (32.5)$$

We want the first variation of $V(\theta_G, \theta_D)$ to vanish.

$$\delta V(\theta_G, \theta_D) = 0 . \quad (32.6)$$

This implies

$$\partial_{\theta_G} V(\theta_G, \theta_D) = \partial_{\theta_D} V(\theta_G, \theta_D) = 0 \quad (32.7)$$

and

$$V_{opt} = \min_{\theta_G} \max_{\theta_D} V(\theta_G, \theta_D) . \quad (32.8)$$

The TPMs, printed in blue, for bnets Figs.32.1 and 32.2, are as follows:

$$P(\theta_G) = \text{ given} \quad (32.9)$$

$$P(\theta_D) = \text{ given} \quad (32.10)$$

$$P(\vec{u}) = \prod_i P(u[i]) \quad (\text{usually uniform distribution}) \quad (32.11)$$

$$P(\vec{r}) = \prod_i P(r[i]) \quad (32.12)$$

$$P(f[i] | \vec{u}, \theta_G) = \delta[f[i], G(u[i], \theta_G)] \quad (32.13)$$

$$P(c[i] | \vec{f}, \theta_D) = \delta(c[i], \overline{D}(f[i], \theta_D)) \quad (32.14)$$

$$P(d[j] | \vec{r}, \theta_D) = \delta(d[j], D(r[j], \theta_D)) \quad (32.15)$$

$$P(V|\vec{d}, \vec{c}) = \delta(V, \frac{1}{N} \ln \prod_{i,j} (c[i]d[j])) \quad (32.16)$$

where $N = nsam(\vec{r})nsam(\vec{u})$.

Let $\eta_G, \eta_D > 0$. Maximize V wrt θ_D , and minimize it wrt θ_G .

$$P(\theta'_G|V, \theta_G) = \delta(\theta'_G, \theta_G - \eta_G \partial_{\theta_G} V) \quad (32.17)$$

$$P(\theta'_D | V, \theta_D) = \delta(\theta'_D, \theta_D + \eta_D \partial_{\theta_D} V) \quad (32.18)$$

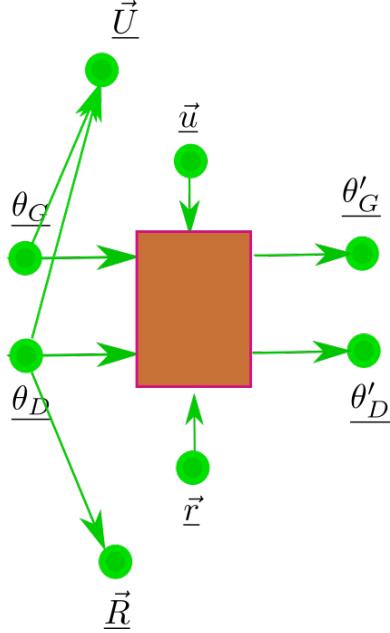


Figure 32.3: GAN, Constraining Bayesian Network

Constraining Bnet given in Fig.32.3. It adds 2 new nodes, namely \vec{U} and \vec{R} , to the bnet of Fig.32.1. The purpose of these 2 barren (childrenless) nodes is to constrain certain functions to be probability distributions.

The TPMs, printed in blue, for the 2 new nodes, are as follows.

$$P(U[i] | \theta_G) = \frac{\overline{D}(G(U[i], \theta_G), \theta_D)}{\bar{\lambda}(\theta_G, \theta_D)} \quad (32.19)$$

where $S_{\underline{U}[i]} = S_u$ and $\bar{\lambda}(\theta_G, \theta_D) = \sum_u \overline{D}(G(u, \theta_G), \theta_D)$.

$$P(R[i] | \theta_G, \theta_D) = \frac{D(R[i], \theta_D)}{\lambda(\theta_D)} \quad (32.20)$$

where $S_{\underline{R}[i]} = S_r$ and $\lambda(\theta_D) = \sum_r D(r, \theta_D)$.

$$P(V | \vec{u}, \vec{r}) = \delta(V, \frac{1}{N} \ln \prod_{i,j} (P(R[i] = r[i] | \theta_G, \theta_D) P(\underline{U}[i] = u[j] | \theta_G))) \quad (32.21)$$

where $N = nsam(\vec{r})nsam(\vec{u})$.

L = likelihood

$$L = P(\vec{r}, \vec{u} | \theta_G, \theta_D) \quad (32.22)$$

$$= \prod_{i,j} \left[\frac{D(r[i], \theta_D)}{\lambda(\theta_D)} \frac{\overline{D}(G(u[j], \theta_G), \theta_D))}{\overline{\lambda}(\theta_G, \theta_D)} \right] \quad (32.23)$$

$$\ln L = N[V(\theta_G, \theta_D) - \ln \lambda(\theta_D) - \ln \overline{\lambda}(\theta_G, \theta_D)] \quad (32.24)$$

Chapter 33

Goodness of Causal Fit

This chapter is an almost verbatim copy of a paper of mine, Ref.[66].

33.1 Abstract

We propose a Goodness (G) of Causal Fit (CF) measure which depends on Pearl “do” interventions. This is different from a measure of Goodness of Statistical Fit (GF), which does not use interventions. Given a DAG set \mathcal{G} , to find a good $G \in \mathcal{G}$, we propose plotting $GCF(G)$ versus $GF(G)$ for all $G \in \mathcal{G}$, and finding a graph $G \in \mathcal{G}$ with a large amount of both types of goodness.

33.2 Introduction

Frequently, when students first encounter Bayesian Networks (bnets) and Causal Inference (CI), they experience serious doubts about the usefulness of this theory, because they believe finding the underlying causal model (i.e., DAG) for most realistic physical situations is too difficult or impossible. I think that part of the problem is that these students are assuming, perhaps unconsciously, that there exists a unique DAG that fits Nature perfectly, and a mind-boggling number of possibilities to sift through to find that DAG. Rather than looking for a unique DAG, I think a better strategy is to write down a set \mathcal{G} of likely DAGs, and to calculate for each DAG in \mathcal{G} , a measure called Goodness of Causal Fit (GCF). Then use the DAGs with the highest GCF scores.

The goal of this paper is to propose a GCF measure. Such a measure is of course not unique, and someone may propose in the future a measure that is better than ours.

When designing a GCF measure, it is important to keep in mind the First Dictum¹ of CI: The data is causal model-less. In the First Dictum, when we say “data”, we are referring to what is commonly called a dataset. A dataset is a table

¹This is just my whimsical name for it.

of data, where all the entries of each column have the same units, and measure a single feature, and each row refers to one particular sample or individual. Datasets are particularly useful for estimating probability distributions and for training neural nets. In the First Dictum, when we say “causal model”, we are referring to a DAG (directed acyclic graph) or a bnet (Bayesian Network= DAG + probability table for each node of DAG).

You can try to derive a causal model from a dataset, but you’ll soon find out that you can only go so far. The process of finding a *partial* causal model from a dataset is called structure learning (SL). SL can be done quite nicely with Marco Scutari’s open source program `bnlearn` (Ref[56]). The problem is that SL often cannot narrow down the causal model to a single one. It finds an undirected graph (UG), and it can determine the direction of some of the arrows in the UG, but it is often incapable, for well understood fundamental—not just technical—reasons, of finding the direction of *all* the arrows of the UG. So it often fails to fully specify a DAG model.

Let’s call the ordered pair (dataset, causal model) a **dataset++**. Then what the First Dictum is saying is that a dataset is causal model-free or causal model-less (although sometimes one can find a partial causal model hidden in there). A dataset is not a dataset++.

Graphs which contain both directed and undirected edges are called **partially directed (PD) graphs**. `bnlearn` takes a dataset as input and returns a PD graph G_{pd} . Given a PD graph G_{pd} , let $\mathcal{G}(G_{pd})$ be the DAG set \mathcal{G} which is generated by giving directions to all undirected edges of G_{pd} in all possible ways. We will refer to $\mathcal{G}(G_{pd})$ as the **DAG set generated by G_{pd}** . and to any $\mathcal{G}' \subset \mathcal{G}(G_{pd})$ as a **DAG set partially generated by G_{pd}** . Once we define below our GCF measure, we will apply it to sets of the type $\mathcal{G}(G_{pd})$ as an example.

It’s clear that any measure of GCF will have to involve interventions such as the “do” intervention invented by Pearl et al for CI. Without interventions like “do”, it is impossible to distinguish causally the DAGs in a set $\mathcal{G}(G_{pd})$.

Henceforth, random variables will be indicated by underlining.

33.3 Goodness of Statistical Fit

Before trying to define a GCF measure, it is instructive to review the closely related, well established, measures of Goodness of Statistical Fit (GF).

Consider two probability distributions $PO(x)$ and $PE(x)$, where $x \in S_{\underline{x}}$. By a GF measure, we mean a measure of the difference between PO and PE . Usually PO is the observed probability distribution and PE is the expected, theoretical one.

Three popular measures of the difference between PO and PE are:

1. The **Kullback-Leibler divergence**:

$$D_{KL}(PO \parallel PE) = \sum_{x \in S_{\underline{x}}} PO(x) \ln \frac{PO(x)}{PE(x)}. \quad (33.1)$$

2. The **Pearson divergence** (a.k.a. **Pearson Chi-squared test statistic**):

$$D_{\chi^2}(PO \parallel PE) = \sum_{x \in S_{\underline{x}}} \frac{[PO(x) - PE(x)]^2}{PE(x)} = \sum_{x \in S_{\underline{x}}} \frac{PO^2(x)}{PE(x)} - 1. \quad (33.2)$$

It's easy to show using $\ln(1 + \delta) = \delta + \mathcal{O}(\delta^2)$ that if $\left| \frac{PO(x)}{PE(x)} - 1 \right| \ll 1$ for all $x \in S_{\underline{x}}$, then

$$D_{KL}(PO \parallel PE) \approx D_{\chi^2}(PO \parallel PE) \quad (33.3)$$

3. The **Euclidean distance squared**:

$$D_E(PO, PE) = \sum_{x \in S_{\underline{x}}} [PO(x) - PE(x)]^2 \quad (33.4)$$

Note that of these 3 measures, only $D_E(PO, PE)$ is symmetric in PO and PE .

Given any bnet G with full probability distribution² $P_G(x.)$ and a probability distribution³ $\tilde{P}(x.)$ derived empirically from a dataset, let

$$D(G) = \sum_{x.} \tilde{P}(x.) \ln \frac{\tilde{P}(x.)}{P_G(x.)} \quad (33.5)$$

$$= D_{KL}(\tilde{P}(\underline{x.}) \parallel P_G(\underline{x.})) \quad (33.6)$$

We define **Goodness of Statistical Fit (GF)** of the bnet G by

$$GF(G) = \ln \frac{1}{D(G)} \quad (33.7)$$

33.4 GCF example 1

For the first example of our measure of GCF, we consider $\mathcal{G}(G_{pd}) = \{G_1, G_2\}$ given by Fig.33.1. We will assume the following:

²We define $x.$ to be a vector with components x_i

³Empirical distributions will be denoted by P with a tilde over it.

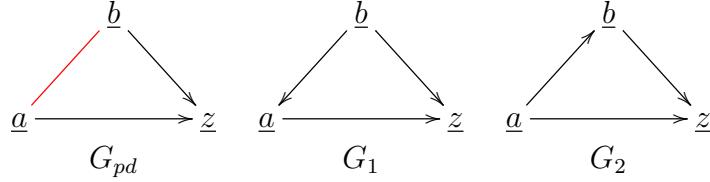


Figure 33.1: $\mathcal{G}(G_{pd}) = \{G_1, G_2\}$. The partially directed graph G_{pd} generates the DAGs G_1 and G_2 by giving directions to all undirected edges of G_{pd} in all possible ways. (In this case, there is only one undirected edge in G_{pd} .)

- First, we assume that we have collected a dataset from which we have extracted a full empirical distribution $\tilde{P}(z, a, b)$. From $\tilde{P}(z, a, b)$, we assume that the following have been calculated. $\tilde{P}(z, b|a)$ $\tilde{P}(z, a|b)$ $\tilde{P}(a)$, $\tilde{P}(b)$.
- Second, we assume that a dataset has been collected for which \underline{a} was held fixed to each of the possible values $a \in S_{\underline{a}}$ of \underline{a} . Furthermore, we assume that the distribution $\tilde{P}(z, b|do(\underline{a}) = a)$ has been extracted from that dataset.
- Third, we assume that a dataset has been collected for which \underline{b} was held fixed to each of the possible values $b \in S_{\underline{b}}$ of \underline{b} . Furthermore, we assume that the distribution $\tilde{P}(z, a|do(\underline{b}) = b)$ has been extracted from that dataset.

We will refer to $\tilde{P}(z, b|do(\underline{a}) = a)$ and $\tilde{P}(z, a|do(\underline{b}) = b)$ as **empirical do-probability distributions**.

Now define

$$D_a = \sum_{z,b} \tilde{P}(z, b|a) \ln \frac{\tilde{P}(z, b|a)}{\tilde{P}(z, b|do(\underline{a}) = a)} \quad (33.8)$$

$$= D_{KL}(\tilde{P}(z, \underline{b}|a) \parallel \tilde{P}(z, \underline{b}|do(\underline{a}) = a)) \quad (33.9)$$

$$D_{\underline{a}} = \sum_a \tilde{P}(a) D_a = E_a[D_a] \quad (33.10)$$

and

$$D_b = D_{KL}(\tilde{P}(z, \underline{a}|b) \parallel \tilde{P}(z, \underline{a}|do(b) = b)) \quad (33.11)$$

$$D_{\underline{b}} = \sum_a \tilde{P}(b) D_b = E_b[D_b]. \quad (33.12)$$

We will refer to $D_{\underline{a}}$ and $D_{\underline{b}}$ as **do-divergences**.

Note that

$$D_a(G_2) = 0 \text{ for all } a \text{ so } \underbrace{D_{\underline{a}}(G_2)}_0 \leq D_{\underline{b}}(G_2) \quad (33.13)$$

and

$$D_b(G_1) = 0 \text{ for all } b \text{ so } D_{\underline{a}}(G_1) \geq \underbrace{D_b(G_1)}_0 . \quad (33.14)$$

If $D_{\underline{a}} \leq D_b$, then $\underline{a} \rightarrow b$, and if $D_{\underline{a}} \geq D_b$ then $\underline{a} \leftarrow b$. Thus, the arrow and the inequality sign point in opposite directions. Alternatively, just remember that the arrow points to the larger of the two D 's.

If $D_{\underline{a}} \leq D_b$, then define $GCF(G_1) = -1$ and $GCF(G_2) = 1$.

If $D_b \leq D_{\underline{a}}$, then define $GCF(G_1) = 1$, $GCF(G_2) = -1$.

33.5 GCF example 2

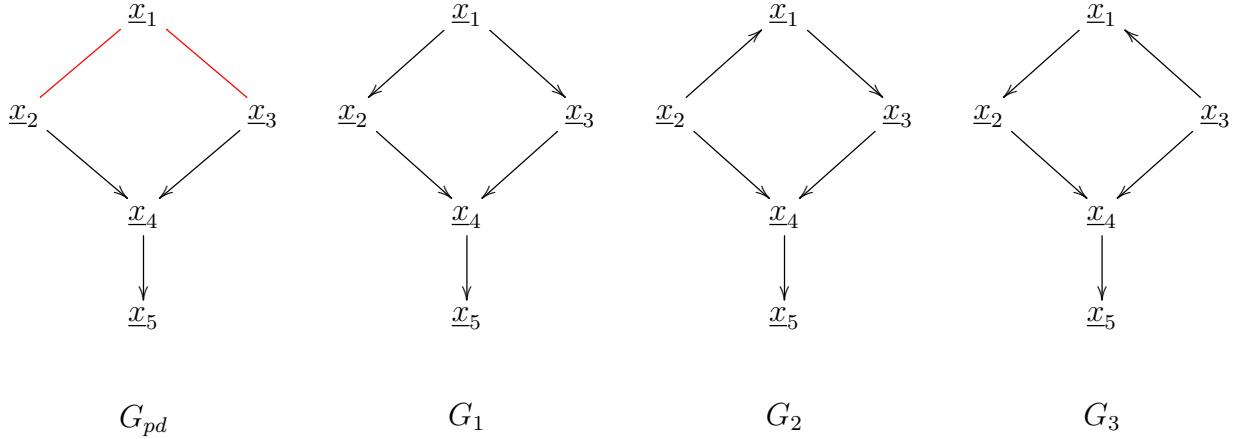


Figure 33.2: $\mathcal{G} = \{G_1, G_2, G_3\}$. \mathcal{G} is a set of observationally equivalent (OE) graphs. These are graphs that have the same value for GF, and are therefore indistinguishable from GF alone. For more info about OE graphs, see Chapter 61. Note that $\mathcal{G}(G_{pd})$ includes one more DAG, the one in which node \underline{x}_1 is a collider.

For the second example of our measure of GCF, consider $\mathcal{G} = \{G_1, G_2, G_3\}$ given by Fig.33.2.

Which of the do-divergences D_{x_2} , D_{x_1} and $D_{\underline{x}_3}$, is smallest, middle and highest, depends on the empirical do-probability distributions. For definiteness, suppose the sizes of these do-divergences are related as follows:

$$D_{\underline{x}_2} \leq D_{x_1} \leq D_{\underline{x}_3} . \quad (33.15)$$

For any two do-divergences $D_{\underline{a}}$ and D_b , define the distance

$$d_{b,a} = |D_b - D_{\underline{a}}| \quad (33.16)$$

If we abbreviate $D_{\underline{x}_j}$ by D_j , we can define the GCF for each of the graphs in \mathcal{G} by:

$$GCF(G_1) = \frac{-d_{2,1} + d_{1,3}}{d_{2,1} + d_{1,3}} \quad (33.17a)$$

$$GCF(G_2) = \frac{d_{2,1} + d_{1,3}}{d_{2,1} + d_{1,3}} = 1 \quad (33.17b)$$

$$GCF(G_3) = \frac{-d_{2,1} - d_{1,3}}{d_{2,1} + d_{1,3}} = -1 \quad (33.17c)$$

33.6 GCF in general

Eqs.(33.17) are a special case of the following formulae.

For any two do-divergences $D_{\underline{a}}$ and $D_{\underline{b}}$, define the **do-divergence distance** by

$$d_{\underline{b}, \underline{a}} = |D_{\underline{b}} - D_{\underline{a}}| . \quad (33.18)$$

For any $G_i \in \mathcal{G}$, define the **edge-sign function** by

$$\sigma_{G_i}(\underline{a} - \underline{b}) = \begin{cases} +1 & \text{if edge } \underline{a} - \underline{b} \text{ in } G_i \text{ points towards larger of } D_{\underline{a}} \text{ and } D_{\underline{b}}. \\ -1 & \text{otherwise} \end{cases} \quad (33.19)$$

Finally, suppose that \mathcal{G} is either partially or fully generated by a PD graph G_{pd} with undirected edges $\{\underline{a}_k - \underline{b}_k\}_{k=0,1,\dots,n_k-1}$. Then define the GCF of graph $G_i \in \mathcal{G}$ by

$$GCF(G_i) = \frac{\sum_{k=0}^{n_k-1} \sigma_{G_i}(\underline{a}_k - \underline{b}_k) d_{\underline{a}_k, \underline{b}_k}}{\sum_{k=0}^{n_k-1} d_{\underline{a}_k, \underline{b}_k}} . \quad (33.20)$$

Note that $-1 \leq GCF(G_i) \leq 1$.

If the DAG set \mathcal{G} contains only one DAG G , define $GCF(G) = 1$, because all arrows in G are in the correct direction, as far as we know.

Call the **skeleton** of a DAG, the undirected graph that one obtains by turning all its edges from directed to undirected ones.

So far, we have applied our measure of GCF to a DAG set \mathcal{G} which is either fully or partially generated by a PD graph G_{pd} , or is a singleton set. But what if we want a GCF that can score every DAG in a DAG set \mathcal{G} that contains DAGs with different skeletons? In that case, let $\{\underline{a}_k - \underline{b}_k\}_{k=0,1,\dots,n_k-1}$ be *all* the edges of graph $G_i \in \mathcal{G}$, and use the relative GCF given by Eq.(33.20), or use an absolute GCF defined by

$$GCF_a(G_i) = \sum_{k=0}^{n_k-1} \sigma_{G_i}(\underline{a}_k - \underline{b}_k) d_{\underline{a}_k, \underline{b}_k} . \quad (33.21)$$

So let \mathcal{G} be an arbitrary DAG set. Our GCF (or GCF_a) measure is not enough to decide the best possible G in \mathcal{G} , because there might be several graphs with $GCF \approx 1$. For this reason, we recommend plotting $GCF(G)$ (or $GCF_a(G)$) versus $GF(G)$ for all $G \in \mathcal{G}$. Then choose a G with a large amount of both types of goodness. It might even be advantageous to average over a small subset of DAGs in \mathcal{G} that have large amounts of both types of goodness. This would be similar to averaging over an ensemble of decision trees to get a random forest.

A plot of $GCF(G)$ versus $GF(G)$ agrees with the spirit of the First Dictum and datasets++, because also in those, we acknowledge a separation between the dataset (GF) and causal model (GCF) degrees of freedom.

Chapter 34

Granger Causality

This chapter is based on the Wikipedia article Ref.[105] and Scholarpedia article Ref.[55] on Granger causality and on the book Ref.[16] on time series analysis by Hamilton.

This chapter assumes the reader has read Chapter 83 on the stationary time series $ARMA(p, q)$ and $VAR(p)$.

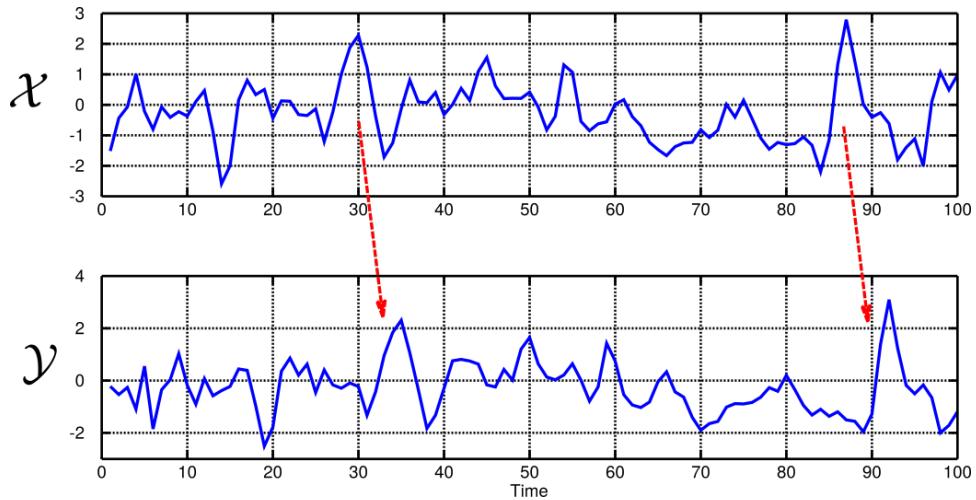


Figure 34.1: When t-series \mathcal{X} Granger-causes t-series \mathcal{Y} , the patterns in \mathcal{X} are approximately repeated in \mathcal{Y} after some time lag (two examples are indicated with arrows). Thus, past values of \mathcal{X} can be used for the prediction of future values of \mathcal{Y} . (image and caption from Ref.[105])

Let $\vec{x}_t = [x_t^A]_{\forall A} \in \mathbb{R}^{nr_1 \times 1}$ and $\vec{y}_t = [y_t^B]_{\forall B} \in \mathbb{R}^{nr_2 \times 1}$. Thus, \vec{x}_t for each t is a column vector with nr_1 rows, and \vec{y}_t for each t is a column vector with nr_2 rows. Let $nr_1 + nr_2 = nr$, the total number of rows. Consider a vector t-series $\{\vec{x}_t, \vec{y}_t\}_{\forall t}$ of type $VAR(p)$, as defined by Eq.(83.183). To simplify the notation of Eq.(83.183), we are replacing x^1 by x , x^2 by y , n^1 by n , and n^2 by w .

$$\begin{bmatrix} \underline{x}_t^A \\ \underline{y}_t^B \end{bmatrix} = \sum_{j=1}^p \begin{bmatrix} \alpha_j^{x|x;A,A'} & \alpha_j^{x|y;A,B'} \\ \alpha_j^{y|x;B,A'} & \alpha_j^{y|y;B,B'} \end{bmatrix} \begin{bmatrix} \underline{x}_{t-j}^{A'} \\ \underline{y}_{t-j}^{B'} \end{bmatrix} + \begin{bmatrix} \underline{n}_t^A \\ \underline{w}_t^B \end{bmatrix} \quad (34.1)$$

Hence,

$$E_{|\vec{x}_{<t}, \vec{y}_{<t}} \begin{bmatrix} \underline{y}_t^B \end{bmatrix} = \sum_{j=1}^p \alpha_j^{y|x;B,A'} \underline{x}_{t-j}^{A'} + \sum_{j=1}^p \alpha_j^{y|y;B,B'} \underline{y}_{t-j}^{B'} \quad (34.2)$$

Let $\mathcal{X} = \{\vec{x}_t\}_{\forall t}$ and $\mathcal{Y} = \{\vec{y}_t\}_{\forall t}$.

We say \mathcal{X} **does not G-cause** (or **does not G-forecast**) \mathcal{Y} , and symbolize this by $\mathcal{X} \dashv_G \mathcal{Y}$, if

$$\alpha_j^{y|x;B,A'} = 0 \quad \forall (j, B, A') \quad (34.3)$$

or, equivalently,

$$E_{|\vec{x}_{<t}, \vec{y}_{<t}} \begin{bmatrix} \underline{y}_t^B \end{bmatrix} = E_{|\vec{y}_{<t}} \begin{bmatrix} \underline{y}_t^B \end{bmatrix} \quad \forall (B, t) \quad (34.4)$$

We say \mathcal{X} **G-causes** (or **G-forecasts**) \mathcal{Y} and symbolize this by $\mathcal{X} \xrightarrow[G]{} \mathcal{Y}$, if $\mathcal{X} \dashv_G \mathcal{Y}$ is false.

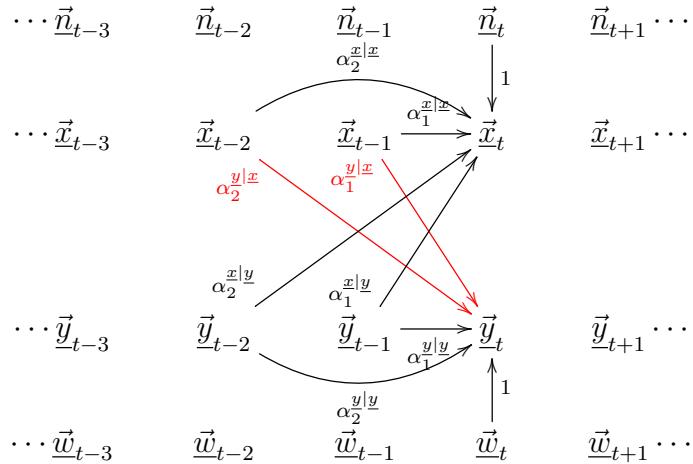


Figure 34.2: Bnet for $VAR(2)$. For clarity, we only show the arrows entering nodes \vec{x}_t and \vec{y}_t . Remove red arrows if \mathcal{X} does not G-cause \mathcal{Y} , and keep them if it does.

Eq.(34) describing a bipartite $VAR(p)$ t-series can be represented by the bnet Fig.34.2. The TPMs, printed in blue, for the two nodes \vec{x}_t and \vec{y}_t , are as follows:

$$P(\vec{x}_t | \vec{x}_{[t-p, t-1]}, \vec{y}_{[t-p, t-1]}, \vec{n}_t) = \prod_A \mathbb{1} \left(\underline{x}_t^A = \sum_{j=1}^p \alpha_j^{x|x;A,A'} \underline{x}_{t-j}^{A'} + \sum_{j=1}^p \alpha_j^{x|y;A,B'} \underline{y}_{t-j}^{B'} + n_t^A \right) \quad (34.5)$$

$$P(\vec{y}_t | \vec{x}_{[t-p,t-1]}, \vec{y}_{[t-p,t-1]}, \vec{w}_t) = \prod_B \mathbb{1} \left(\underline{y}_t^B = \sum_{j=1}^p \underbrace{\alpha_j^{y|x;B,A'}}_{0?} \vec{x}_{t-j}^{A'} + \sum_{j=1}^p \alpha_j^{y|y;B,B'} \vec{y}_{t-j}^{B'} + w_t^B \right) \quad (34.6)$$

Testing for GC

- Consider the datasets

$$\mathcal{D}_{\underline{x}} = \{(t, \vec{x}_{[t-p,t-1]}, \vec{y}_{[t-p,t-1]}, \boxed{\vec{x}_t}) : t\} \quad (34.7)$$

$$\mathcal{D}_{\underline{y}} = \{(t, \vec{x}_{[t-p,t-1]}, \vec{y}_{[t-p,t-1]}, \boxed{\vec{y}_t}) : t\} \quad (34.8)$$

One can do Linear Regression on $\mathcal{D}_{\underline{x}}$ with x-variables $(\vec{x}_{[t-p,t-1]}, \vec{y}_{[t-p,t-1]})$, y-variable \vec{x}_t , and regression coefficients $\alpha_{[1,p]}^{x|x}$, $\alpha_{[1,p]}^{x|y}$. One can also do Linear Regression on $\mathcal{D}_{\underline{y}}$ with x-variables $(\vec{x}_{[t-p,t-1]}, \vec{y}_{[t-p,t-1]})$, y-variable \vec{y}_t , and regression coefficients $\alpha_{[1,p]}^{y|x}$, $\alpha_{[1,p]}^{y|y}$. The LR software yields confidence intervals for the regression coefficients.

- One can do hypothesis testing using the Likelihood Ratio Test¹

Null hypothesis $H_0 : \mathcal{X} \xrightarrow[G]{\parallel} \mathcal{Y}$, Alternative hypothesis $H_1 : \mathcal{X} \xrightarrow[G]{} \mathcal{Y}$

- Test for both $\mathcal{X} \xrightarrow[G]{} \mathcal{Y}$ and $\mathcal{Y} \xrightarrow[G]{} \mathcal{X}$.

Limitations

It has been remarked that G-causality is not true causality because, even though an event A must precede an event B in order to cause it, that does not necessarily mean that A causes B. I think the problem with G-causality is that it uses a bnet that is a good fit for the dataset, but not necessarily also a good *causal* fit for the experiment. One can measure the goodness of causal fit of a bnet by doing do-intervention experiments (See Chapter 33).

¹The Likelihood Ratio Test is discussed in Section C.22.

Chapter 35

Hidden Markov Model

A dynamical Bayesian network (DBN) (see Chapter 24) is a generalization of a Hidden Markov Model (HMM), which in turn is a generalization of a Kalman Filter (KF) (see Chapter 41).

See Wikipedia article Ref.[106] to learn about the history and many uses of HMMs. This chapter is based on Refs.[39], [106], [149], [78].

In this chapter, we use the following conventions.

Random variables are underlined and their values are not. For example, $\underline{a} = a$ means the random variable \underline{a} takes the value a . A diagram with all its nodes underlined represents a Bayesian Network (bnet), whereas the same diagram with the letters not underlined represents a specific **instantiation** of that bnet. For example $\underline{a} \rightarrow \underline{b} \rightarrow \underline{c}$ represents the bnet with full probability distribution $P(c|b)P(b|a)P(a)$, whereas $a \rightarrow b \rightarrow c$ represents $P(c|b)P(b|a)$. Note that, for convenience, we define $a \rightarrow b \rightarrow c$ to exclude the priors of root nodes such as $P(a)$.

If \underline{a} is a root node, then $\sum a$ signifies a weighted sum $\sum_a P(a)$. For example,

$$\sum a \rightarrow b \rightarrow c = \sum_a P(c|b)P(b|a)P(a) \quad (35.1)$$

If \underline{a} is not a root node, then $\sum a$ signifies a simple unweighted sum \sum_a . For example,

$$x \rightarrow \sum a \rightarrow y = \sum_a P(y|a)P(a|x) \quad (35.2)$$

Two bnets are equated if their full probability distributions (i.e., their full instantiations) are equal numerically. For example,

$$\underline{a} \rightarrow \underline{b} \rightarrow \underline{c} = P(c|b)P(b|a)P(a) = \underline{a} \leftarrow \underline{b} \leftarrow \underline{c} \quad (35.3)$$

Unobserved nodes are indicated by enclosing them in a dashed circle. For example, $\langle \underline{u} \rangle$.

Suppose

$\underline{v}^n = (\underline{v}_0, \underline{v}_1, \dots, \underline{v}_{n-1})$ are n visible nodes that are measured, and

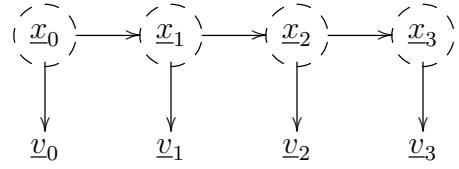


Figure 35.1: HMM bnet with $n = 4$.

$\underline{x}^n = (\underline{x}_0, \underline{x}_1, \dots, \underline{x}_{n-1})$ are the n hidden, unmeasurable state nodes of a system that is being monitored.

For the bnet of Fig.35.1, one has

$$P(\underline{x}^n, v^n) = \prod_{t=0}^{n-1} P(x_t | x_{t-1}) P(v_t | x_t), \quad (35.4)$$

where $x_{-1} = \emptyset$.

We assume that this bnet is stationary. The following notation emphasizes that fact:

$\pi(x)$ = prior probability

$$\pi(x) = P(\underline{x}_0 = x) \quad (35.5)$$

$A(x|x')$ = transition matrix

$$A(x|x') = P(\underline{x}_t = x | \underline{x}_{t-1} = x') \quad (35.6)$$

$B(v|x)$ = emission probability

$$B(v|x) = P(\underline{v}_t = v | \underline{x}_t = x) \quad (35.7)$$

Let $x_{<t} = (x_0, x_1, \dots, x_{t-1})$.

For $t = 0, 1, \dots, n-1$, define

$\mathcal{F}_t(x_t)$ =future measurements probability

$$\mathcal{F}_t(x_t) = P(v_{>t} | x_t) \quad (35.8)$$

$$= x_t \longrightarrow \sum x_{>t} \quad (35.9)$$

$$\downarrow \\ v_{>t}$$

$\bar{\mathcal{F}}_t(x_t)$ = past and present measurements probability

$$\bar{\mathcal{F}}_t(x_t) = P(v_{<t}, v_t, x_t) \quad (35.10)$$

$$= \sum x_{<t} \longrightarrow x_t \quad (35.11)$$



$\lambda_t(x_t)$ = present measurement probability, (a.k.a. emission probability $B(v_t|x_t)$). $\lambda_t(x_t)$ is the likelihood of x_t .

$$\lambda_t(x_t) = P(v_t|x_t) \quad (35.12)$$

$$= x_t \quad (35.13)$$

$$\begin{array}{c} \downarrow \\ v_t \end{array}$$

35.1 Calculating $P(x_t, v^n)$ and $P(x_t, x_{t+1}, v^n)$

Claim 58 For $t \geq 0$,

$$P(x_t, v^n) = \bar{\mathcal{F}}_t(x_t) \mathcal{F}_t(x_t) \quad (35.14)$$

$$= \sum x_{<t} \longrightarrow x_t \quad x_t \longrightarrow \sum x_{>t} \quad (35.15)$$

$$\begin{array}{ccc} \downarrow & & \downarrow \\ v_{<t} & & v_t & & v_{>t} \end{array}$$

For $t > 0$,

$$P(x_{t-1}, x_t, v^n) = \bar{\mathcal{F}}_{t-1}(x_{t-1}) P(x_t|x_{t-1}) \lambda_t(x_t) \mathcal{F}_t(x_t) \quad (35.16)$$

$$= \sum x_{<t-1} \longrightarrow x_{t-1} \quad x_{t-1} \longrightarrow x_t \quad x_t \longrightarrow \sum x_{>t} \quad (35.17)$$

$$\begin{array}{cccc} \downarrow & & \downarrow & \downarrow \\ v_{<t-1} & & v_{t-1} & v_t & v_{>t} \end{array}$$

proof:

$$P(x_t, v^n) = \sum_{x_{<i}} \sum_{x_{>i}} P(x^n, v^n) \quad (35.18)$$

$$= \sum_{x_{<i}} \sum_{x_{>i}} P(x^n, v^n | x_t) P(x_t) \quad (35.19)$$

$$= \sum_{x_{<i}} \sum_{x_{>i}} P(x_{<i}, v_{<i}, v_t | x_t) P(x_{>t}, v_{>i} | x_t) P(x_t) \quad (35.20)$$

$$= P(v_{<i}, v_t | x_t) P(v_{>i} | x_t) P(x_t) \quad (35.21)$$

$$= \overline{\mathcal{F}}_t(x_t) \mathcal{F}_t(x_t) \quad (35.22)$$

$$P(x_{t-1}, x_t, v^n) = \sum_{x_{<t-1}} \sum_{x_{>t}} P(x^n, v^n) \quad (35.23)$$

$$= \sum_{x_{<t-1}} \sum_{x_{>t}} P(x^n, v^n | x_{t-1}, x_t) P(x_{t-1}, x_t) \quad (35.24)$$

$$= \sum_{x_{<t-1}} \sum_{x_{>t}} P(x_{<t-1}, v_{<t-1}, v_{t-1} | x_{t-1}) P(v_t | x_t) P(x_{t-1}, x_t) P(x_{>i}, v_{>i} | x_t) \quad (35.25)$$

$$= P(v_{<t-1}, v_{t-1} | x_{t-1}) P(v_t | x_t) P(x_{t-1}, x_t) P(v_{>i} | x_t) \quad (35.26)$$

$$= \overline{\mathcal{F}}_{t-1}(x_{t-1}) \lambda_t(x_t) P(x_t | x_{t-1}) \mathcal{F}_t(x_t) \quad (35.27)$$

QED

35.2 Calculating \mathcal{F}_t and $\overline{\mathcal{F}}_t$

Claim 59 For $t > 0$, \mathcal{F}_t and $\overline{\mathcal{F}}_t$ can be calculated recursively as follows:

$$\overline{\mathcal{F}}_t(x_t) = \sum_{x_{t-1}} \overline{\mathcal{F}}_{t-1}(x_{t-1}) P(x_t | x_{t-1}) \lambda_t(x_t) \quad (35.28)$$

$$= \sum_{x_{t-1}} \sum_{x_{<t-1}} x_{<t-1} \longrightarrow x_{t-1} \quad x_{t-1} \longrightarrow x_t \quad x_t \\ \downarrow \qquad \qquad \qquad \downarrow \qquad \qquad \qquad \downarrow \\ v_{<t-1} \qquad \qquad \qquad v_{t-1} \qquad \qquad \qquad v_t \quad (35.29)$$

and

$$\mathcal{F}_{t-1}(x_{t-1}) = \sum_{x_t} P(x_t|x_{t-1})\lambda_t(x_t)\mathcal{F}_t(x_t) \quad (35.30)$$

$$= \sum_{x_t} x_{t-1} \longrightarrow x_t \quad x_t \quad x_t \longrightarrow \sum x_{>t} \quad (35.31)$$

\downarrow \downarrow
 v_t $v_{>t}$

proof:

$$\overline{\mathcal{F}}_t(x_t)\mathcal{F}_t(x_t) = P(x_t, v^n) \quad (35.32)$$

$$= \sum_{x_{t-1}} P(x_{t-1}, x_t, v^n) \quad (35.33)$$

$$= \sum_{x_{t-1}} \overline{\mathcal{F}}_{t-1}(x_{t-1})\lambda_t(x_t)P(x_t|x_{t-1})\mathcal{F}_t(x_t) \quad (35.34)$$

$$\overline{\mathcal{F}}_{t-1}(x_{t-1})\mathcal{F}_{t-1}(x_{t-1}) = P(x_{t-1}, v^n) \quad (35.35)$$

$$= \sum_{x_t} P(x_{t-1}, x_t, v^n) \quad (35.36)$$

$$= \sum_{x_t} \overline{\mathcal{F}}_{t-1}(x_{t-1})\lambda_t(x_t)P(x_t|x_{t-1})\mathcal{F}_t(x_t) \quad (35.37)$$

QED

35.3 Calculating $P(x^n|v^n)$

Claim 60

$$P(x_t|x_{t-1}, v^n) = \frac{P(x_t|x_{t-1})\lambda_t(x_t)\mathcal{F}_t(x_t)}{\mathcal{F}_{t-1}(x_{t-1})} \quad (35.38)$$

$$= \frac{x_{t-1} \longrightarrow x_t \quad x_t \quad x_t \longrightarrow \sum x_{>t}}{x_{t-1} \longrightarrow \sum x_t \longrightarrow \sum x_{>t}} \quad (35.39)$$

\downarrow \downarrow
 v_t $v_{>t}$
 \downarrow \downarrow
 v_t $v_{>t}$

Note that actually, $P(x_t|x_{t-1}, v^n) = P(x_t|x_{t-1}, v_{\geq t})$ by d-separation, but we won't use this fact.

$$P(x_{t-1}|x_t, v^n) = \frac{\bar{\mathcal{F}}_{t-1}(x_{t-1})P(x_t|x_{t-1})\lambda_t(x_t)}{\bar{\mathcal{F}}_t(x_t)} \quad (35.40)$$

$$= \frac{\sum x_{<t-1} \longrightarrow x_{t-1} \quad x_{t-1} \longrightarrow x_t \quad x_t}{\sum x_{<t-1} \longrightarrow \sum x_{t-1} \longrightarrow x_t} \quad (35.41)$$

Note that actually $P(x_{t-1}|x_t, v^n) = P(x_{t-1}|x_t, v_{\leq t-1})$ by d-separation, but we won't use this fact.

proof:

$$P(x_t|x_{t-1}, v^n) = \frac{P(x_{t-1}, x_t, v^n)}{P(x_{t-1}, v^n)} \quad (35.42)$$

$$= \frac{\bar{\mathcal{F}}_{t-1}(x_{t-1})\lambda_t(x_t)P(x_t|x_{t-1})\mathcal{F}_t(x_t)}{\bar{\mathcal{F}}_{t-1}(x_{t-1})\mathcal{F}_{t-1}(x_{t-1})} \quad (35.43)$$

Analogous proof for Eq.(35.41).

QED

$$P(x^n|v^n) = \prod_{t=1,\dots,n-1,n} P(x_t|x_{t-1}, v^n) \quad (\text{forward propagation}) \quad (35.44)$$

$$= \prod_{t=n+1,n,\dots,3,2} P(x_{t-1}|x_t, v^n) \quad (\text{backward propagation}) \quad (35.45)$$

35.4 Calculating $P(v^n|A, B, \pi)$

$P(v^n|A, B, \pi)$ can be calculated

- from $\bar{\mathcal{F}}_{n-1}(x_{n-1})$ (past of x_{n-1})

$$P(v^n|A, B, \pi) = \sum_{x_{n-1}} \underbrace{P(v_{<n-1}, v_{n-1}, x_{n-1})}_{\bar{\mathcal{F}}_{n-1}(x_{n-1})} \quad (35.46)$$

$$= \sum x_{<n-1} \longrightarrow \sum x_{n-1} \quad (35.47)$$

- from $\mathcal{F}_0(x_0)$ (future of \underline{x}_0)

$$P(v^n|A, B, \pi) = \sum_{x_0} P(x_0) P(v_0|x_0) \underbrace{P(v_{>0}|x_0)}_{\mathcal{F}_0(x_0)} \quad (35.48)$$

$$= \sum x_0 \longrightarrow \sum x_{>0} \quad (35.49)$$

\downarrow \downarrow
 v_0 $v_{>0}$

- from $\bar{\mathcal{F}}_t(x_t)$ and $\mathcal{F}_t(x_t)$ for some t (past and future of \underline{x}_t)

$$P(v^n|A, B, \pi) = \sum_{x_t} \underbrace{P(v_{<t}, v_t, x_t)}_{\bar{\mathcal{F}}_t(x_t)} \underbrace{P(v_{>t}|x_t)}_{\mathcal{F}_t(x_t)} \quad (35.50)$$

$$= \sum x_{<t} \longrightarrow \sum x_t \longrightarrow \sum x_{>t} . \quad (35.51)$$

\downarrow \downarrow \downarrow
 $v_{<t}$ v_t $v_{>t}$

35.5 Calculating \hat{x}^n (Viterbi algorithm)

x^n is not visible, only v^n is. Here is how to find an estimate \hat{x}^n of x^n .

Define

$$\bar{\mathcal{F}}_t^{max}(x_t) = \max_{x_{<t}} P(v_{<t}, v_t, x_{<t}, x_t) \quad (35.52)$$

$$= \max x_{<t} \longrightarrow x_t \quad (35.53)$$

\downarrow \downarrow
 $v_{<t}$ v_t

$$\hat{x}_{t-1}(x_t) = \operatorname{argmax}_{x_{t-1}} \bar{\mathcal{F}}_{t-1}^{max}(x_{t-1}) P(x_t|x_{t-1}) \lambda_t(x_t) \quad (35.54)$$

$$= \max x_{<t-1} \longrightarrow \operatorname{argmax} x_{t-1} \longrightarrow x_t \quad (35.55)$$

\downarrow \downarrow \downarrow
 $v_{<t-1}$ v_{t-1} v_t

For $t > 0$, $\bar{\mathcal{F}}_t^{max}$ and $\hat{x}_{t-1}(x_t)$ can be calculated recursively as follows:

$$\overline{\mathcal{F}}_t^{max}(x_t) = \max_{x_{t-1}} \overline{\mathcal{F}}_{t-1}^{max}(x_{t-1}) P(x_t|x_{t-1}) \lambda_t(x_t) \quad (35.56)$$

$$= \max_{x_{t-1}} \max x_{<t-1} \longrightarrow x_{t-1} \quad x_{t-1} \longrightarrow x_t \quad x_t \downarrow \\ v_{<t-1} \qquad \qquad \qquad v_{t-1} \qquad \qquad \qquad v_t$$

$$\hat{x}_{t-1}(x_t) = \operatorname{argmax}_{x_{t-1}} \overline{\mathcal{F}}_{t-1}^{max}(x_{t-1}) P(x_t|x_{t-1}) \lambda_t(x_t) \quad (35.58)$$

$$= \operatorname{argmax}_{x_{t-1}} \max x_{<t-1} \longrightarrow x_{t-1} \quad x_{t-1} \longrightarrow x_t \quad x_t \downarrow \\ v_{<t-1} \qquad \qquad \qquad v_{t-1} \qquad \qquad \qquad v_t$$

Claim 61 (*Viterbi Algorithm*)

If

$$\hat{x}^n = \operatorname{argmax}_{x^n} P(x^n|v^n), \quad (35.60)$$

then the components of \hat{x}^n can be calculated recursively from the last one \hat{x}_{n-1} to the first one \hat{x}_0 , as follows. Let

$$\hat{x}_{n-1} = \operatorname{argmax}_{x_{n-1}} \overline{\mathcal{F}}_{n-1}^{max}(x_{n-1}) \quad (35.61)$$

$$= \max x_{<n-1} \longrightarrow \operatorname{argmax} x_{n-1} \quad (35.62)$$

$$\downarrow \qquad \qquad \qquad \downarrow \\ v_{<n-1} \qquad \qquad \qquad v_{n-1}$$

and for $t < n - 1$, use

$$\hat{x}_{t-1} = \hat{x}_{t-1}(\hat{x}_t) \quad (35.63)$$

$$= \max x_{<t-1} \longrightarrow \operatorname{argmax} x_{t-1} \longrightarrow \hat{x}_t \quad (35.64)$$

$$\downarrow \qquad \qquad \qquad \downarrow \qquad \qquad \qquad \downarrow \\ v_{<t-1} \qquad \qquad \qquad v_{t-1} \qquad \qquad \qquad v_t$$

proof:
QED

35.6 Calculating \hat{A} , \hat{B} , $\hat{\pi}$ (Baum-Welch algorithm)

Let $\theta = (A, B, \pi)$. θ is a set of hidden parameters. Here is how to find an estimate $\hat{\theta}$ of θ .

If x^n and v^n were visible, we could use

$$\hat{\pi}(x) = \mathbb{1}(\underline{x}_0 = x) \quad (35.65)$$

$$\hat{A}(x'|x) = \frac{\sum_{t=0}^{n-2} \mathbb{1}(\underline{x}_t = x, \underline{x}_{t+1} = x')}{\sum_x \text{numerator}} \quad (35.66)$$

$$\hat{B}(v|x) = \frac{\sum_{t=0}^{n-1} \mathbb{1}(\underline{v}_t = v, \underline{x}_t = x)}{\sum_v \text{numerator}} \quad (35.67)$$

But x^n is not visible. So how can we estimate θ under those circumstances?

Define

$$\gamma_t(x_t) = P(x_t|v^n) \quad (35.68)$$

$$= \frac{P(x_t, v^n)}{P(v^n)} \quad (35.69)$$

$$= \frac{\mathcal{F}_t(x_t)\mathcal{F}_t(x_t)}{\sum_{x_t} \text{numerator}} \quad (35.70)$$

$$= \frac{\sum x_{<t} \longrightarrow x_t \ x_t \longrightarrow \sum x_{>t}}{\sum_{x_t} \text{numerator}} \quad (35.71)$$

$$\xi_t(x_t, x_{t+1}) = \frac{P(x_t, x_{t+1}, v^n)}{\sum_{x_t} \sum_{x_{t+1}} \text{numerator}} \quad (35.72)$$

$$= \frac{\mathcal{F}_{t-1}(x_{t-1})P(x_t|x_{t-1})\lambda_t(x_t)\mathcal{F}_t(x_t)}{\sum_{x_t} \sum_{x_{t+1}} \text{numerator}} \quad (35.73)$$

$$= \frac{\sum x_{<t-1} \longrightarrow x_{t-1} \ x_{t-1} \longrightarrow x_t \ x_t \longrightarrow \sum x_{>t}}{\sum_{x_t} \sum_{x_{t+1}} \text{numerator}} \quad (35.74)$$

Claim 62 (*Baum-Welch algorithm*)

If

$$\widehat{\theta} = \operatorname{argmax}_{\theta} P(v^n | \theta) , \quad (35.75)$$

then we can find $\widehat{\theta}$ using the following formulae:

$$\widehat{\pi}(x) = \overbrace{\gamma_0(x)}^{\sim P(x)} \quad (35.76)$$

$$\widehat{A}(x'|x) = \frac{\sum_{t=0}^{n-2} \overbrace{\xi_t(x, x')}^{\sim P(x, x')}}{\sum_{t=0}^{n-2} \underbrace{\gamma_t(x)}_{P(x)}} \quad (35.77)$$

$$\widehat{B}(v|x) = \frac{\sum_{t=0}^{n-1} \overbrace{\mathbb{1}(\underline{v}_t = v)}^{\sim P(v|x)} \overbrace{\gamma_t(x)}^{\sim P(x)}}{\sum_{t=0}^{n-1} \underbrace{\gamma_t(x)}_{\sim P(x)}} \quad (35.78)$$

proof:

QED

Chapter 36

Influence Diagrams & Utility Nodes

Influence diagrams are just arbitrary bnets enhanced with a new kind of node called an utility node. The rest of this brief chapter will be devoted to discussing utility nodes.

Suppose $U(x)$ is a deterministic function $U : S_x \rightarrow \mathbb{R}$ called the **utility function**. Then the **expected utility** is defined as

$$E_U[U] = \sum_U P(U)U \quad (36.1)$$

$$= \sum_x \sum_U \underbrace{P(U|x)}_{\delta[U,U(x)]} P(x)U \quad (36.2)$$

$$= \sum_x P(x)U(x) . \quad (36.3)$$

An **utility node** can be understood as a node composed of 3 simpler bnet nodes. This is illustrated in Fig.36.1.

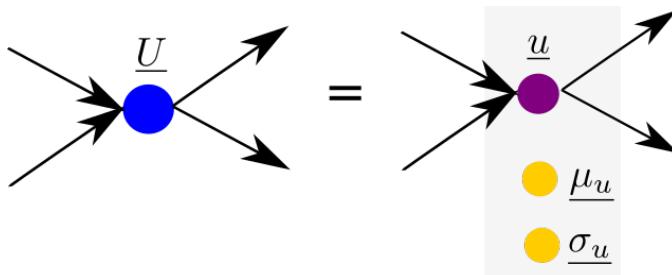


Figure 36.1: An utility node can be understood as a node composed of 3 simpler bnet nodes.

The TPMs, printed in blue, for the bnet Fig.36.1, are as follows:

$$P(U|pa(U)) = \delta[U, U(pa(U))], \quad (36.4)$$

where if $U : S_{\underline{x}} \rightarrow \mathbb{R}$, then $\underline{x} = pa(\underline{U})$.

$$P(u|pa(U)) = \delta[u, U(pa(U))] \quad (36.5)$$

Node $\underline{\mu}_u$ calculates the expected value (mean value) of \underline{u} :

$$P(\mu_u) = \delta(\mu_u, E_{\underline{u}}[\underline{u}]) \quad (36.6)$$

Node $\underline{\sigma}_u$ calculates the standard deviation of \underline{u} :

$$P(\sigma_u) = \delta(\sigma_u, \sqrt{E_{\underline{u}}[(\underline{u} - E_{\underline{u}}[\underline{u}])^2]}) \quad (36.7)$$

Note that in order to calculate expected values, it is necessary that $\underline{U}, \underline{u} \in \mathbb{R}$. Note that nodes $\underline{u}, \underline{\mu}_u, \underline{\sigma}_u$ must all 3 have access to the TPM $P(U|pa(U))$ of node \underline{U} . In fact, in order to calculate $E_{\underline{u}}[\cdot]$, it is necessary for nodes $\underline{\mu}_u$ and $\underline{\sigma}_u$ to have access not just to $P(U|pa(U))$ but also to $P(pa(U))$.

See Fig.36.2. An influence diagram may have multiple utility nodes (\underline{U}_1 and \underline{U}_2 in Fig.36.2). Then one can define a merging utility node \underline{U} that sums the values of all the other utility nodes.

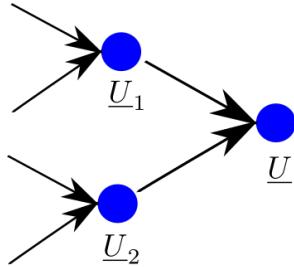


Figure 36.2: An influence diagram may have multiple utility nodes, say \underline{U}_1 and \underline{U}_2 . Then one can define an utility node $\underline{U} = \underline{U}_1 + \underline{U}_2$.

For the node \underline{U} of Fig.36.2,

$$P(U|U_1, U_2) = \delta(U, U_1 + U_2) \quad (36.8)$$

Chapter 37

Instrumental Inequality and beyond

This chapter is based on Refs. [9] and [42].

Instrumental Variables (IVs) are discussed in Chapter 38. This chapter will discuss the original Instrumental inequality (I-inequality) discovered by Pearl, and other related inequalities. The I-inequality arises in bnets that use an IV. The I-inequality bounds the effect that an IV \underline{z} can have on the outcome \underline{y} of a treatment $\underline{d} \rightarrow \underline{y}$. Since there is a path $\underline{z} \rightarrow \underline{d} \rightarrow \underline{y}$, the treatment dose \underline{d} acts as a mediator between the IV \underline{z} and the treatment outcome \underline{y} . The I-inequality is reminiscent of the data processing inequality $H(\underline{z} : \underline{y}) \leq H(\underline{d} : \underline{y})$ which is valid for a simple Markov chain bnet $\underline{z} \rightarrow \underline{d} \rightarrow \underline{y}$. The data processing inequality is saying that the endpoint \underline{y} receives more information from \underline{d} than from \underline{z} . This is reasonable, since \underline{y} is “closer” to \underline{d} than to \underline{z} .

37.1 I-inequality

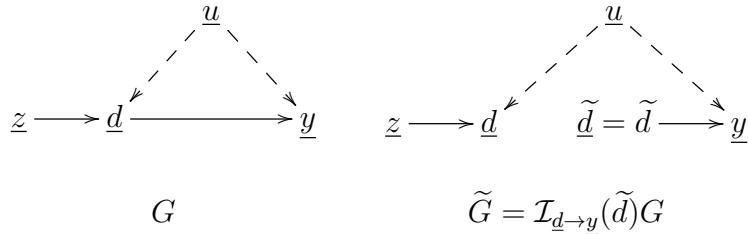


Figure 37.1: In bnet G , an IV \underline{z} acts on a treatment $\underline{d} \rightarrow \underline{y}$. Bnet \widetilde{G} is obtained by applying an imagine operator to arrow $\underline{d} \rightarrow \underline{y}$ of bnet G .

Claim 63 *The TPMs for the bnet G in Fig.37.1 satisfy*

$$\max_d \sum_y \max_z P(d, y|z) \leq 1 \quad (37.1)$$

proof:

Below, any probability that alludes to a value \tilde{d} refers to bnet \tilde{G} . Otherwise, if it doesn't allude to \tilde{d} , then it refers to G (or to \tilde{G} , since the TPMs of \tilde{G} are defined from those of G in a consistent manner.)

G satisfies

$$P(d, y|z) = \sum_u P(u)P(y|u, d)P(d|u, z), \quad (37.2)$$

and \tilde{G} satisfies

$$P(d, y|z, \tilde{d}) = \sum_u P(u)P(y|u, \tilde{d})P(d|u, z). \quad (37.3)$$

Note that Eqs.(37.2) and (37.3) imply that

$$P(d, y|z, d) = P(d, y|z) \quad (37.4)$$

and that

$$P(\tilde{d}, y|z, \tilde{d}) \leq \sum_d P(d, y|z, \tilde{d}) = P(y|\tilde{d}). \quad (37.5)$$

Thus,

$$\max_{\tilde{d}} \sum_y \max_z P(\tilde{d}, y|z, \tilde{d}) \leq \max_{\tilde{d}} \sum_y \max_z P(y|\tilde{d}) \quad (37.6)$$

$$\leq \max_{\tilde{d}} \sum_y P(y|\tilde{d}) \quad (37.7)$$

$$\leq \max_{\tilde{d}} 1 \quad (37.8)$$

$$\leq 1 \quad (37.9)$$

QED

As pointed out in Ref.[9] from which I learned the above proof, the above proof is highly generalizable.

Fig.37.2 gives a graphical representation of the boxed Eq.(37.5) which is crucial to the proof.

And here is a meta-description of the steps in the proof:

1. Use imagine operator to create a non-negative matrix $M_{d, \tilde{d}}$.
2. Use fact that row or column sum of $M_{d, \tilde{d}}$ is larger than diagonal element in sum:
 $\sum_d M_{d, \tilde{d}} \geq M_{\tilde{d}, \tilde{d}}$.

$$\sum_u \begin{array}{c} \xrightarrow{\quad \quad u \quad \quad} \\ \xleftarrow{\quad \quad \tilde{d} = \tilde{d} \quad \quad} \\ \underline{z} \longrightarrow \underline{d} = \underline{\tilde{d}} \end{array} \leq \sum_d \sum_u \begin{array}{c} \xrightarrow{\quad \quad u \quad \quad} \\ \xleftarrow{\quad \quad \tilde{d} = \tilde{d} \quad \quad} \\ \underline{z} \longrightarrow \underline{d} \end{array} = \begin{array}{c} \xrightarrow{\quad \quad u \quad \quad} \\ \xleftarrow{\quad \quad \tilde{d} = \tilde{d} \quad \quad} \\ \underline{\tilde{d}} = \underline{\tilde{d}} \longrightarrow \underline{y} \end{array} .$$

Figure 37.2: Graphical representation of the boxed equation Eq.(37.5).

37.1.1 I-inequality for binary $\underline{z}, \underline{d}, \underline{y}$

It is enlightening to write down the I-inequality for the special case that $\underline{z}, \underline{d}, \underline{y}$ are binary.

$$P(d=1, y|z) = \begin{matrix} & z=0 & z=1 \\ & \begin{matrix} y=0 & \begin{array}{|c|c|} \hline A & B \\ \hline C & D \\ \hline \end{array} \end{matrix} \\ & y=1 & \end{matrix}$$

$$A + D \leq 1$$

$$B + C \leq 1$$

Figure 37.3: I-inequality for binary $\underline{z}, \underline{d}, \underline{y}$. The same picture except with $d = 0$ is also true.

In the binary case, the I-inequality implies 4 different inequalities. These are as follows. One gets two inequalities by setting $d = 1$ in the next 2 equations.

$$\sum_{y=0}^1 \sum_{z=0}^1 \mathbb{1}(y=z) P(d, y|z) , \quad (37.10a)$$

$$\sum_{y=0}^1 \sum_{z=0}^1 \mathbb{1}(y \neq z) P(d, y|z) . \quad (37.10b)$$

One gets an additional 2 inequalities by setting $d = 0$ in Eqs.(37.10). These 4 inequalities are illustrated in Fig.37.3.

What do they mean? That at fixed \underline{d} , the correlation between \underline{z} and \underline{y} is limited.

37.2 Bounds on Effect of IV on treatment outcome \mathbf{y}

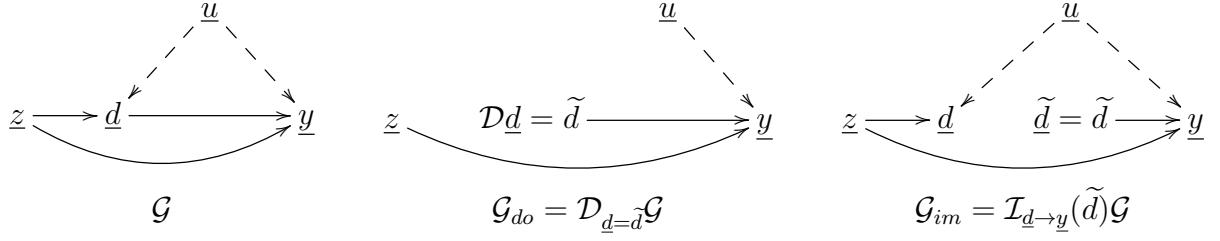


Figure 37.4: Bnet \mathcal{G} is obtained from the bnet G in Fig.37.1 by adding to G an arrow from the IV z to the treatment outcome y . Bnet \mathcal{G}_{do} is obtained by applying a do operator to node d of \mathcal{G} . Bnet \mathcal{G}_{im} is obtained by applying an imagine operator to arrow $d \rightarrow y$ of \mathcal{G} .

In this section, we will assume that random variables z, d, y are binary. Just as with the binary case of the I-inequality, we will find an inequality for each value of $d \in \{0, 1\}$.

Below, we will use the following 3 shorthand notations:

$$P_{y|z}(d) = P(d, y|z) , \quad (37.11)$$

$$P_{|z}(d) = \sum_y P(d, y|z) , \quad (37.12)$$

and

$$\pi_{|z}(d) = 1 - P_{|z}(d) . \quad (37.13)$$

For the bnet \mathcal{G}_{do} in Fig.37.4, define the IV effect at fixed $\mathcal{D}d = \tilde{d}$ by

$$IVE(\tilde{d}) = P(y = 1|z = 1, \mathcal{D}d = \tilde{d}) - P(y = 1|z = 0, \mathcal{D}d = \tilde{d}) . \quad (37.14)$$

Claim 64 *The TPMs for the bnet \mathcal{G}_{do} in Fig.37.4 satisfy*

$$\pi_{|0}(d) \leq [IVE(d) - \{P_{1|1}(d) - P_{1|0}(d)\}] \leq \pi_{|1}(d) \quad (37.15)$$

proof:

$$P(y|z, \mathcal{D}\underline{d} = \tilde{d}) = \sum_u P(u)P(y|u, z, \tilde{d}) \quad (37.16)$$

$$= \sum_u P(u) \sum_d P(d, y|u, z, \tilde{d}) \quad (37.17)$$

$$\geq \sum_u P(u)P(\tilde{d}, y|u, z, \tilde{d}) \quad (37.18)$$

$$= \sum_u P(u)P(\tilde{d}, y|u, z) \quad (37.19)$$

$$= P_{y|z}(\tilde{d}) \quad (37.20)$$

Next note that $P(d, y|z, \tilde{d}) \geq 0$, and $\sum_{d,y} P(d, y|z, \tilde{d}) = 1$. If we write a table for $P(d, y|z, \tilde{d})$ at fixed z, \tilde{d} with row and column indices (d, y) , then a partial sum of the entries of that table must be ≤ 1 :

$$\sum_{d \neq \tilde{d}} P(d, y|z, \tilde{d}) + \underbrace{\sum_{y'} P(\tilde{d}, y'|z, \tilde{d})}_{P_{|z}(\tilde{d})} \leq 1 . \quad (37.21)$$

Using the definitions of $P_{|z}$ and $\pi_{|z}$, we can rewrite the last equation as

$$\sum_{d \neq \tilde{d}} P(d, y|z, \tilde{d}) \leq \pi_{|z}(\tilde{d}) . \quad (37.22)$$

Next note that

$$P(y|z, \mathcal{D}\underline{d} = \tilde{d}) = \sum_u P(u)P(y|u, z, \tilde{d}) \quad (37.23)$$

$$= \sum_u P(u) \sum_d P(d, y|u, z, \tilde{d}) \quad (37.24)$$

$$= P(\tilde{d}, y|z, \tilde{d}) + \sum_{d \neq \tilde{d}} P(d, y|z, \tilde{d}) \quad (37.25)$$

$$= P_{y|z}(\tilde{d}) + \sum_{d \neq \tilde{d}} P(d, y|z, \tilde{d}) \quad (37.26)$$

$$\leq P_{y|z}(\tilde{d}) + \pi_{|z}(\tilde{d}) . \quad (37.27)$$

Hence,

$$P_{y|z}(\tilde{d}) \leq P(y|z, \mathcal{D}\underline{d} = \tilde{d}) \leq P_{y|z}(\tilde{d}) + \pi_{|z} \quad (37.28)$$

$$P_{1|1}(\tilde{d}) \leq P(y=1|z=1, \mathcal{D}\underline{d} = \tilde{d}) \leq P_{1|1}(\tilde{d}) + \pi_{|1} \quad (37.29)$$

$$-P_{1|0}(\tilde{d}) - \pi_{|0} \leq -P(y=1|z=0, \mathcal{D}\underline{d} = \tilde{d}) \leq -P_{1|0}(\tilde{d}) \quad (37.30)$$

QED

Chapter 38

Instrumental Variables

This chapter is based on Refs.[8] and [109].

The theory of potential outcomes (PO) discussed in Chapter 64 assumes that confounders can be ignored by conditioning on them. However, there are cases when that is not possible, as when there are some unmeasured (i.e., unobserved, hidden) confounder nodes in the bnet, because one can only condition on observed random variables, by definition. So what if confounders can't be ignored? Are we then precluded from using PO theory? Not necessarily. It might still be possible to use PO theory if one can find a suitable instrumental variable (IV) for the problem.

IVs were actually invented by Sewall Wright and his father Philip Wright long before PO theory was invented by Rubin. The reason why IVs save PO theory is greatly clarified by using Pearl causal DAGs and his d-separation theorem (see Chapter 21).

Most of the discussion in this chapter is limited to LDEN (linear deterministic bnets with external noise). These are discussed in Chapter 44. However, as will become obvious to the reader, IVs are also applicable and useful in general bnet modeling.

In this chapter, for any random variable \underline{x} , we will represent the linear operator $\langle \underline{x}, \cdot \rangle^{-1} \langle \underline{x}, \cdot \rangle$ as a derivative:

$$\frac{d}{d\underline{x}} \cdot = \frac{\langle \underline{x}, \cdot \rangle}{\langle \underline{x}, \underline{x} \rangle} . \quad (38.1)$$

38.1 δ with unmeasured confounder

In this section, we explain using LDENs why unmeasured confounders prejudice PO calculations.

Consider the LDEN bnet of Fig.38.1, For some $\delta, \mu \in \mathbb{R}$, we have

$$\underline{y} = \delta \underline{d} + \underbrace{\mu \underline{u} + \underline{u}_y}_{n_y} . \quad (38.2)$$

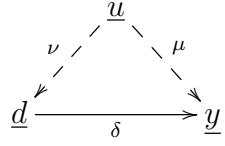


Figure 38.1: An LDEN bnet. The direct path $\underline{d} \rightarrow \underline{y}$ is confounded by a hidden variable \underline{u} .

Note that

$$\langle \underline{y}, \underline{d} \rangle = \delta \langle \underline{d}, \underline{d} \rangle + \langle \underline{n}_{\underline{y}}, \underline{d} \rangle . \quad (38.3)$$

Therefore,

$$\delta = \frac{d\underline{y}}{d\underline{d}} - \frac{d\underline{n}_{\underline{y}}}{d\underline{d}} . \quad (38.4)$$

If the confounder \underline{u} is measured, then we calculate the covariances at fixed $\underline{n}_{\underline{y}}$, and the conditional covariance $\langle \underline{n}_{\underline{y}}, \underline{d} \rangle_{|\underline{n}_{\underline{y}}} = 0$.

38.2 δ (with unmeasured confounder) can be inferred via IV

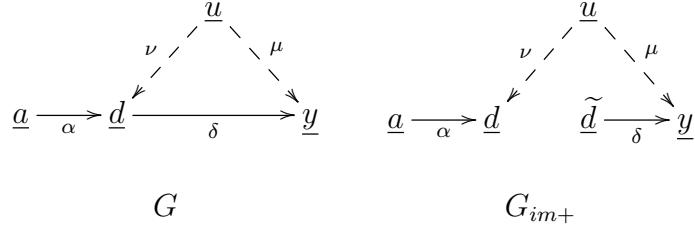


Figure 38.2: Two LDEN bnets. The direct path $\underline{d} \rightarrow \underline{y}$ is confounded by a hidden variable \underline{u} , but by using the IV \underline{a} , we are still able to identify (i.e., calculate) δ .

Now consider the two LDEN bnets shown in Fig.38.2. Note that there are no arrows $\underline{a} \rightarrow \underline{y}$ or $\underline{a} \rightarrow \underline{u}$. Note that node \underline{d} is a collider in the path $\underline{a} - \underline{d} - \underline{u} - \underline{y}$. Therefore, the only unblocked path from \underline{a} to \underline{y} in G is $\underline{a} \rightarrow \underline{d} \rightarrow \underline{y}$ and that path has been removed in G_{im+} . These observations are encapsulated in the following statements.

$$\underline{d} \perp_G \underline{y} = \text{false}, \quad \underline{a} \perp_G \underline{y} = \text{false} . \quad (38.5)$$

$$\underline{d} \perp_{G_{im+}} \underline{y} = \text{false}, \quad \underline{a} \perp_{G_{im+}} \underline{y} = \text{true}. \quad (38.6)$$

The following is true for G :

$$\underline{y} = \delta \underline{d} + \underbrace{\mu \underline{u} + \underline{u}_y}_{n_y} \quad (38.7)$$

$$\underline{d} = \alpha \underline{a} + \underbrace{\nu \underline{u} + \underline{u}_d}_{n_d}. \quad (38.8)$$

Since $\langle n_y, \underline{a} \rangle = \langle n_d, \underline{a} \rangle = 0$ is true in G , we have

$$\langle \underline{y}, \underline{a} \rangle = \delta \langle \underline{d}, \underline{a} \rangle \quad (38.9)$$

and

$$\langle \underline{d}, \underline{a} \rangle = \alpha \langle \underline{a}, \underline{a} \rangle. \quad (38.10)$$

Note that $\langle \underline{y}, \underline{a} \rangle = \delta = 0$ for G_{im+} but not for G , so we are speaking about G from here on. It follows that

$$\alpha = \frac{d\underline{d}}{d\underline{a}} \quad (38.11)$$

and

$$\delta = \frac{\langle \underline{y}, \underline{a} \rangle}{\langle \underline{d}, \underline{a} \rangle} \quad (38.12)$$

$$= \frac{\langle \underline{y}, \underline{a} \rangle}{\langle \underline{a}, \underline{a} \rangle} \frac{\langle \underline{a}, \underline{a} \rangle}{\langle \underline{d}, \underline{a} \rangle} \quad (38.13)$$

$$= \frac{\frac{dy}{da}}{\frac{dd}{da}} \quad (\text{equal to } \frac{dy}{dd}) \quad (38.14)$$

$$= \frac{dy}{d(\alpha a)}. \quad (38.15)$$

Eq.(38.14) is illustrated in Fig.38.3.

38.3 More general bnets with IVs

Figs.38.4 and 38.5 are examples of other bnets for which the effect δ is identifiable thanks to the IV \underline{a} .

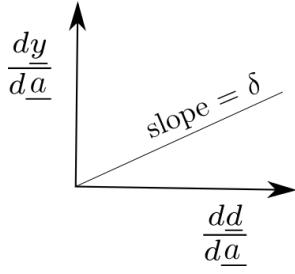


Figure 38.3: Effect $\delta = \frac{dy}{d\underline{d}}$ as slope of line.

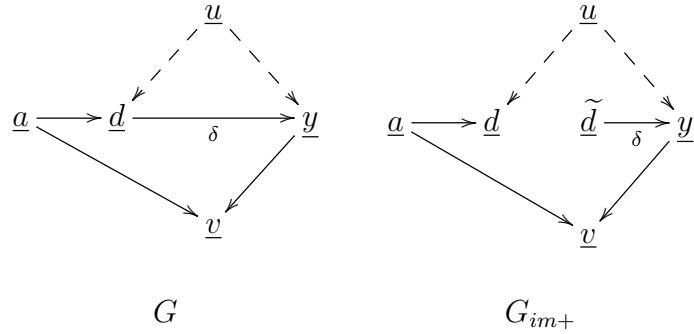


Figure 38.4: The 2 paths in G_{im+} from IV variable \underline{a} to \underline{y} are blocked by not conditioning on colliders \underline{v} and \underline{d} . Thus, $\underline{d} \perp_{G_{im+}} \underline{y} = \text{false}$, $\underline{a} \perp_{G_{im+}} \underline{y} = \text{true}$

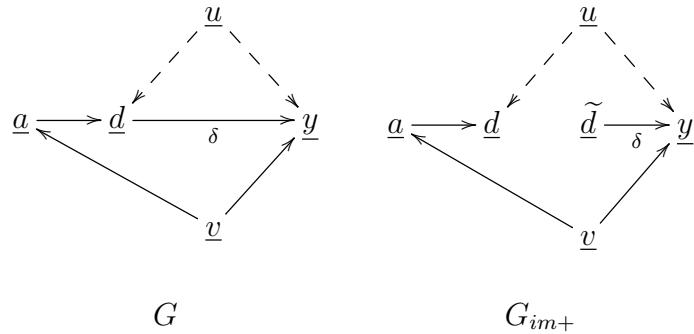


Figure 38.5: There are 2 paths in G_{im+} from IV variable \underline{a} to \underline{y} . One is blocked by not conditioning on the collider \underline{d} and the other can be blocked by conditioning on \underline{v} . Thus, $\underline{d} \perp_{G_{im+}} \underline{y}|v = \text{false}$, $\underline{a} \perp_{G_{im+}} \underline{y}|v = \text{true}$

38.4 Instrumental Inequality

Pearl's instrumental inequality and related inequalities are discussed in Chapter 37.

Chapter 39

Jackknife Resampling

This paper is based on Ref. [111].

Before reading this chapter, we recommend that you read the section entitled “Demystifying Population and Sample Variances”, in Chapter C.

Jackknife Resampling (JR) is a way of generating from an original list of n samples, a new list of n synthetic (i.e., man made, not occurring physically in Nature) samples obtained by deleting one of the samples from the original list and averaging over the rest.

Let $\Sigma = \{0, 1, 2, \dots, n - 1\}$. Let us consider the list of samples

$$\vec{x} = (x^\sigma)_{\sigma \in \Sigma}, \quad (39.1)$$

where the x^σ are assumed to be i.i.d. with

$$E[\underline{x}^\sigma] = \mu \quad (39.2)$$

and

$$\left\langle \underline{x}^\sigma, \underline{x}^{\sigma'} \right\rangle = V_1 \delta(\sigma, \sigma') . \quad (39.3)$$

If we define μ and V_1 estimators by

$$\hat{\mu} = \frac{1}{n} \sum_{\sigma} x^\sigma \quad (39.4a)$$

$$\hat{V}_1 = \frac{1}{n-1} \sum_{\sigma} (x^\sigma - \hat{\mu})^2 , \quad (39.4b)$$

then one can show that:

$$E[\hat{\mu}] = \mu , \quad E[\hat{V}_1] = V_1 \quad (39.5)$$

so both of these estimators are unbiased.

Now define lists of samples with one of the items in \vec{x} deleted:

$$\vec{x}_\xi = (x^\sigma)_{\sigma \in \Sigma - \{\xi\}} . \quad (39.6)$$

Suppose we are given functions of \vec{x} and \vec{x}_ξ

$$A = A^n(\vec{x}) , \quad (39.7)$$

$$A_\xi = A^{n-1}(\vec{x}_\xi) . \quad (39.8)$$

Then define a list \vec{A} by

$$\vec{A} = (A_\xi)_{\xi \in \Sigma} . \quad (39.9)$$

One can also define a list \vec{B} by:

$$\vec{B} = (B_\xi)_{\xi \in \Sigma} \quad (39.10)$$

where

$$B_\xi = nA - (n-1)A_\xi \quad (39.11)$$

$$= A_\xi - n[A_\xi - A] . \quad (39.12)$$

Later on, we will see why the list \vec{B} just defined is of interest.

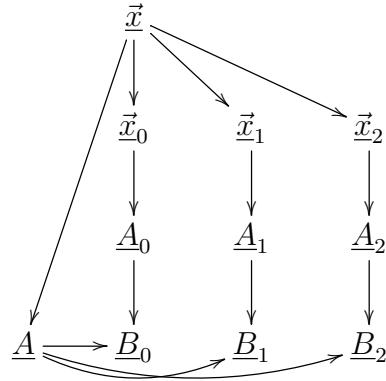


Figure 39.1: Bnet for jackknife resampling (JR).

Fig.39.1 is a bnet that encapsulates JR. The TPMs, printed in blue, for that bnet, are as follows:

$$P(\vec{x}_\xi | \vec{x}) = \mathbb{1}(\vec{x}_\xi = \text{defined by Eq.(39.6)}) \quad (39.13)$$

$$P(A_\xi | \vec{x}_\xi) = \mathbb{1}(A_\xi = \text{defined by Eq.(39.8)}) \quad (39.14)$$

$$P(B_\xi | \vec{A}_\xi, A) = \mathbb{1}(\quad B_\xi = \text{ defined by Eq.(39.12)} \quad) \quad (39.15)$$

39.1 Case $A = A^n(\vec{x}) = \frac{1}{n} \sum_{\sigma} x^{\sigma}$

Suppose

$$A = \frac{1}{n} \underbrace{\sum_{\sigma} x^{\sigma}}_{E_{\sigma}[x^{\sigma}]} \quad (39.16)$$

and

$$A_\xi = \underbrace{\frac{1}{n-1} \sum_{\sigma \in \Sigma - \{\xi\}} x^{\sigma}}_{E_{\sigma|\xi}[x^{\sigma}] \text{ where } P(\sigma|\xi) = \frac{1}{n-1}} . \quad (39.17)$$

Then

$$\frac{1}{n} \sum_{\xi} A_\xi = E_{\xi} E_{\sigma|\xi}[x^{\sigma}] = E_{\sigma}[x^{\sigma}] = A \quad (39.18)$$

Claim 65

$$E[\underline{A}_\xi] = \mu \quad (39.19)$$

$$\langle \underline{A}_\xi, \underline{A}_{\xi'} \rangle = V_1 \left[\frac{n-2-\delta(\xi, \xi')}{n-1} \right] \quad (39.20)$$

proof:

$$E[\underline{A}_\xi] = \frac{1}{n-1} \sum_{\sigma} \mathbb{1}(\xi \neq \sigma) E[x^{\sigma}] = \mu \quad (39.21)$$

$$\langle \underline{A}_\xi, \underline{A}_{\xi'} \rangle = \frac{1}{n-1} \sum_{\sigma} \sum_{\sigma'} [1 - \delta(\sigma, \xi)][1 - \delta(\sigma', \xi')] \langle x^{\sigma}, x^{\sigma'} \rangle \quad (39.22)$$

$$= \frac{V_1}{n-1} \sum_{\sigma} \sum_{\sigma'} [1 - \delta_{\xi}^{\sigma} - \delta_{\xi'}^{\sigma'} + \delta_{\xi}^{\sigma} \delta_{\xi'}^{\sigma'}] \delta_{\sigma'}^{\sigma} \quad (39.23)$$

$$= \frac{V_1}{n-1} [n-2 + \delta_{\xi'}^{\xi}] \quad (39.24)$$

QED

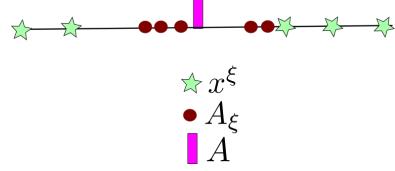


Figure 39.2: For each point x^ξ , there is a corresponding point A_ξ which is $n - 1$ times closer to the average A than x^ξ is.

Claim 66

$$A_\xi - A = \frac{A - x^\xi}{n - 1} \quad (39.25)$$

Hence, the distance of a point x^ξ to the mean value A is $n - 1$ times as large as the distance of A_ξ to A . (see Fig.39.2)

proof:

$$A_\xi - A = \frac{1}{n - 1} \left(\sum_{\sigma} x^\sigma - x^\xi \right) - \frac{1}{n} \sum_{\sigma} x^\sigma \quad (39.26)$$

$$= x^\xi \left(\frac{-1}{n - 1} \right) + \sum_{\sigma} x^\sigma \underbrace{\left(\frac{1}{n - 1} - \frac{1}{n} \right)}_{\frac{1}{n(n-1)}} \quad (39.27)$$

$$= \frac{A - x^\xi}{n - 1} \quad (39.28)$$

QED

Note that

$$(n - 1) \sum_{\xi} (A_\xi - E_{\xi}[A_\xi])^2 = \hat{V}_1 \quad (39.29)$$

by Eq.(39.25). Hence, we can estimate V_1 from \vec{A} instead of \vec{x} .

Note that

$$B_\xi = nA - (n - 1)A_\xi \quad (39.30)$$

$$= \sum_{\sigma} x^\sigma - \sum_{\sigma \neq \xi} x^\sigma = x^\xi \quad (39.31)$$

Since $B_\xi = x^\xi$, they have identical statistics. In particular, one can use for B_ξ the same μ and V_1 estimators that we defined in Eqs.(39.4) for x^σ .

Chapter 40

Junction Tree Algorithm

The Junction Tree (JT) algorithm is an algo for evaluating exact marginals of a bnet, including cases in which some nodes are fixed to a single state. (fixed nodes are called the a priori evidence.)

The JT algo starts by clustering the loops of a bnet into bigger nodes so as to transform the bnet into a polytree bnet. Then it applies Pearl Belief Propagation (see Chapter 50) to the ensuing polytree. The first breakthrough paper to achieve this agenda in full was Ref.[27] by Lauritzen, and Spiegelhalter in 1988. See the Wikipedia article Ref.[112] for more info and references on the JT algorithm.

I won't describe the JT algo any further here, because it would take too long for this brief book to give a complete treatment of it, including the mathematical proofs. If all you want to do is to code the JT algo, without delving into the mathematical theorems and proofs behind it, I strongly recommend Ref.[22]. Ref.[22] is an excellent cookbook for programmers of the JT algo. My open source program QuantumFog (see Ref.[69]) implements the JT algo in Python, following the recipe of Ref.[22].

Chapter 41

Kalman Filter

This chapter is based on Ref.[114], except we've replaced the variables F_t, w_t, v_t in that reference by A_t, ξ_t, ζ_t , respectively.

A Kalman Filter (KF) is a special case of a Hidden Markov Model. HMMs are discussed in Chapter 35.

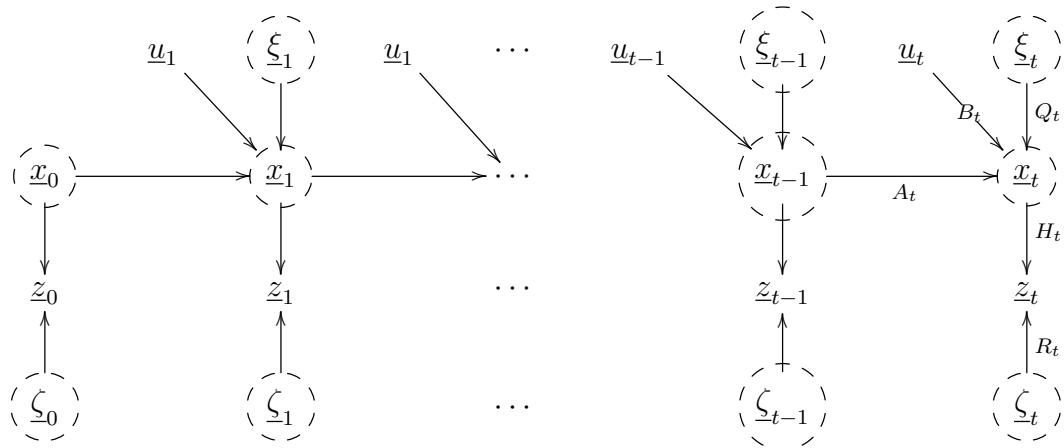


Figure 41.1: Kalman Filter (KF) bnet.

Let

$t = 0, 1, 2, \dots, T - 1$ be the time.

$\underline{\xi}_t \in \mathbb{R}^{nx}, \underline{\zeta}_t \in \mathbb{R}^{nz}$ be random variables that represent hidden (unobserved) external Gaussian white noise.

$\underline{x}_t \in \mathbb{R}^{nx}$ be random variables that represent the hidden (unobserved) true state of the system.

$\underline{u}_t \in \mathbb{R}^{nu}, \underline{z}_t \in \mathbb{R}^{nz}$ be random variables that represent the measured (observed) state of the system.

The TPMs, printed in blue, for the KF bnet Fig.41.1, are as follows:

$$P(\xi_t) = \mathcal{N}(\xi_t; 0, Q_t), \quad (41.1)$$

where Q_t is given.

$$P(x_t|x_{t-1}, u_t, \xi_t) = \mathbb{1}(x_t = A_t x_{t-1} + B_t u_t + \xi_t), \quad (41.2)$$

where $A_t, B_t u_t$ are given. $P(x_t|x_{t-1}, u_t, \xi_t)$ becomes $P(x_t)$ for $t = 0$.

$$P(\zeta_t) = \mathcal{N}(\zeta_t; 0, R_t), \quad (41.3)$$

where R_t is given.

$$P(z_t|x_t, \zeta_t) = \mathbb{1}(z_t = H_t x_t + \zeta_t), \quad (41.4)$$

where H_t is given.

41.1 Prediction Problem

Find \hat{x}_t (the best possible estimate of x_t) and P_t (the state of the filter at time t) in terms of

1. \hat{x}_{t-1} and P_{t-1} .
2. the 5 matrices $\mathcal{M}_t = (A_t, B_t, H_t, Q_t, R_t)$
3. the observed values of z_t and u_t .

See Fig.41.2. For that figure,

$$P(\hat{x}_t, P_t | \hat{x}_{t-1}, P_{t-1}, \mathcal{M}_t, z_t, u_t) = \delta(\hat{x}_t, ?) \delta(P_t, ?). \quad (41.5)$$

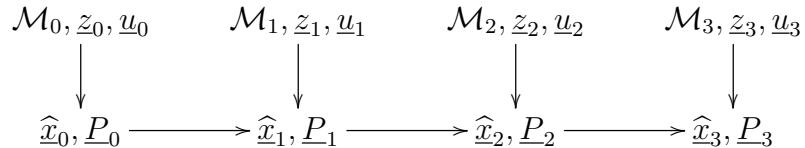


Figure 41.2: Evolution of \hat{x}_t, P_t for a KF.

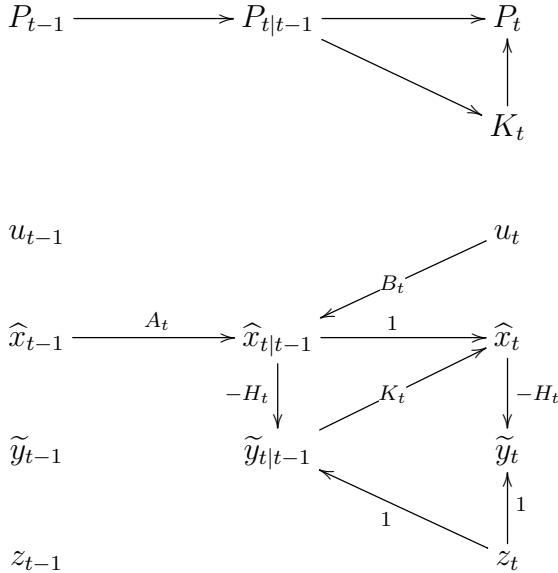


Figure 41.3: Bnet representation of the algebraic solution of the prediction problem for a KF.

41.2 Solution

The algebraic solution of the prediction problem for a KF is as follows. See Fig.41.3 for a bnet representation of this algebraic solution.

Define $\eta_{t|t} = \eta_t$ for $\eta = \hat{x}, P$.

- **Initial Conditions** \hat{x}_0, P_0

- **A priori estimates**

a priori state estimate

$$\hat{x}_{t|t-1} = \underbrace{A_t \hat{x}_{t-1} + B_t u_t}_{x_t - \xi_t} \quad (41.6)$$

a priori covariance estimate

$$P_{t|t-1} = A_t P_{t-1} A_t^T + Q_t \quad (41.7)$$

- **A posteriori estimates**

Optimal Kalman gain K_t

$$S_t = H_t P_{t|t-1} H_t^T + R_t \quad (41.8)$$

$$K_t = P_{t|t-1} H_t^T S_t^{-1} \quad (41.9)$$

$$= P_{t|t-1} H_t^T [H_t P_{t|t-1} H_t^T + R_t]^{-1} \quad (41.10)$$

a posteriori state estimate

$$\tilde{y}_{t|t-1} = z_t - H_t \hat{x}_{t|t-1} \quad (41.11)$$

$$\hat{x}_t = \hat{x}_{t|t-1} + K_t \tilde{y}_{t|t-1} \quad (41.12)$$

$$= (1 - K_t H_t) \hat{x}_{t|t-1} + K_t \underbrace{z_t}_{H_t x_t + \zeta_t} \quad (\text{interpolation formula}) \quad (41.13)$$

$$\tilde{y}_t = z_t - H_t \hat{x}_t \quad (41.14)$$

a posteriori covariance estimate

$$P_t = (I - K_t H_t) P_{t|t-1} \quad (41.15)$$

41.3 Simple Example

r_t position, v_t velocity, a_t acceleration of point particle.

$$x_t = Ax_{t-1} + Bu_t + \xi_t \quad (41.16)$$

$$x_t = \begin{bmatrix} r_t \\ v_t \end{bmatrix}, \quad u_t = a_t \quad (41.17)$$

$$A = \begin{bmatrix} 1 & \Delta t \\ 0 & 1 \end{bmatrix}, \quad B = \begin{bmatrix} \frac{1}{2}(\Delta t)^2 \\ \Delta t \end{bmatrix} \quad (41.18)$$

$$z_t = Hx_t + \zeta_t \quad (41.19)$$

$$H = \begin{bmatrix} 1 & 0 \end{bmatrix} \quad (41.20)$$

41.4 Invariants

Note that

$$x_t - \hat{x}_{t|t-1} = \xi_t \quad (41.21)$$

$$x_t - \hat{x}_t = (1 - K_t H_t)(x_t - \hat{x}_{t|t-1}) - K_t \zeta_t \quad (41.22)$$

$$= (1 - K_t H_t)\xi_t - K_t \zeta_t \quad (41.23)$$

$$\tilde{y}_{t|t-1} = \underbrace{z_t}_{H_t x_t + \zeta_t} - H_t \hat{x}_{t|t-1} \quad (41.24)$$

$$= H_t \xi_t + \zeta_t \quad (41.25)$$

$$\tilde{y}_t = \underbrace{z_t}_{H_t x_t + \zeta_t} - H_t \hat{x}_t \quad (41.26)$$

$$= H_t(x_t - \hat{x}_t) + \zeta_t \quad (41.27)$$

$$= H_t(1 - K_t H_t)\xi_t + (1 - H_t K_t)\zeta_t \quad (41.28)$$

$(x_t - \hat{x}_{t|t-1})$, $(x_t - \hat{x}_t)$, $\tilde{y}_{t|t-1}$ and \tilde{y}_t are called **residuals**. Since ξ_t and ζ_t have zero mean value,

$$E[x_t - \hat{x}_t] = E[x_t - \hat{x}_{t|t-1}] = 0 \quad (41.29)$$

$$E[\tilde{y}_t] = E[\tilde{y}_{t|t-1}] = 0 \quad (41.30)$$

These zero mean value identities are called **invariants**.

41.5 Derivation of Solution

First, some notational conventions. Let

$$Cov(\underline{a})_{i,j} = \langle \underline{a}_i, \underline{a}_j \rangle \quad (41.31)$$

$$Cov(\underline{a}) = \langle \underline{a}, \underline{a}^T \rangle = \langle \underline{a}, tp. \rangle \quad (41.32)$$

$$[A, B]_+ = AB + B^T A^T \quad (41.33)$$

$$A + tp. = A + A^T \quad (41.34)$$

$tp.$ stands for transpose.

Now define

$$P_t = Cov(\underline{x}_t - \hat{\underline{x}}_t) \quad (41.35)$$

$$P_{t|t-1} = Cov(\underline{x}_t - \hat{\underline{x}}_{t|t-1}) \quad (41.36)$$

$$S_t = Cov(\tilde{y}_{t|t-1}) \quad (41.37)$$

It follows that

$$P_t = \langle \underline{x}_t - \hat{\underline{x}}_t, tp. \rangle \quad (41.38)$$

$$= \left\langle (1 - K_t H_t)(\underline{x}_t - \hat{\underline{x}}_{t|t-1}) - K_t \underline{\zeta}_t, tp. \right\rangle \quad (\text{by Eq.(41.23)}) \quad (41.39)$$

$$= \left\langle (1 - K_t H_t)(\underline{x}_t - \hat{\underline{x}}_{t|t-1}), tp. \right\rangle + \left\langle K_t \underline{\zeta}_t, tp. \right\rangle \quad (\underline{\zeta}_t \text{ uncorrelated with } (\underline{x}_t - \hat{\underline{x}}_{t|t-1})) \quad (41.40)$$

$$= (1 - K_t H_t) \underbrace{\langle \underline{x}_t - \hat{\underline{x}}_{t|t-1}, tp. \rangle}_{P_{t|t-1}} (1 - K_t H_t)^T + K_t \underbrace{\langle \underline{\zeta}_t, tp. \rangle}_{R_t} K_t^T \quad (41.41)$$

$$= P_{t|t-1} - [K_t H_t, P_{t|t-1}]_+ + K_t \underbrace{(H_t P_{t|t-1} H_t^T + R_t)}_{S_t} K_t^T \quad (41.42)$$

Next we find the optimal Kalman gain K_t by minimizing with respect to K_t the following mean squared error.

$$\mathcal{E} = \sum_i E[(\underline{x}_t - \hat{\underline{x}}_t)_i^2] \quad (41.43)$$

$$= \text{tr} E[(\underline{x}_t - \hat{\underline{x}}_t)(\underline{x}_t - \hat{\underline{x}}_t)^T] \quad (41.44)$$

$$= \text{tr} P_t \quad (41.45)$$

$$= \text{tr}(P_{t|t-1} - [K_t H_t, P_{t|t-1}]_+ + K_t S_t K_t^T) \quad (41.46)$$

If we set to zero the variation of \mathcal{E} when K_t varies, we get

$$0 = \delta \mathcal{E} = \text{tr} [(-P_{t|t-1} H_t^T + K_t S_t) \delta K_t^T] + tp. \quad (41.47)$$

Hence

$$-P_{t|t-1} H_t^T + K_t S_t = 0 \quad (41.48)$$

$$K_t = P_{t|t-1} H_t^T S_t^{-1} \quad (41.49)$$

Chapter 42

LATE (Local Average Treatment Effect)

This chapter is based on Refs.[36, 10].

The **Local Average Treatment Effect (LATE)** is the same as the ATE estimand¹, but it only counts “compliers” (i.e., individuals that comply with the treatment they’ve been assigned). LATE assumes the same bnet that we considered when we discussed Instrumental Variables (IV)²

σ	a^σ	$d^\sigma(a^\sigma = 0)$	$d^\sigma(a^\sigma = 1)$	$y^\sigma(d^\sigma = 0)$	$y^\sigma(d^\sigma = 1)$
1	0	1			0
2	0	0		1	
3	0	1			1
4	0	0		0	
5	0	0		1	
6	1		1		0
7	1		0	0	
8	1		1		0
9	1		0	1	
10	1		1		1

Table 42.1: Hypothetical dataset for LATE problem. Pink cells indicate missing data.

Table 42.1 shows a hypothetical dataset for the LATE problem.

Consider the bnets G and G_+ of Fig.42.1.³

¹ATE is defined in the Potential Outcomes chapter, i.e., Chapter 64.

²Instrumental Variables are discussed in Chapters 38 and 37.

³If there were an arrow $\underline{a} \rightarrow \underline{y}$ in bnet G , then we could also apply the imagine operator $\mathcal{I}_{\underline{a} \rightarrow \underline{y}}$ to G . This would lead to nodes $[y(a', d')]_{a'=0}^1 |_{d'=0}^1$. Since there is no arrow $\underline{a} \rightarrow \underline{y}$, we would have $y(a, d) = y(d)$ for all a, d , i.e., y does not depend directly on the instrument a . This independence of $y(a, d)$ on a is called the **excludability assumption**. One could say the excludability assumption is built into our bnet G_+ .

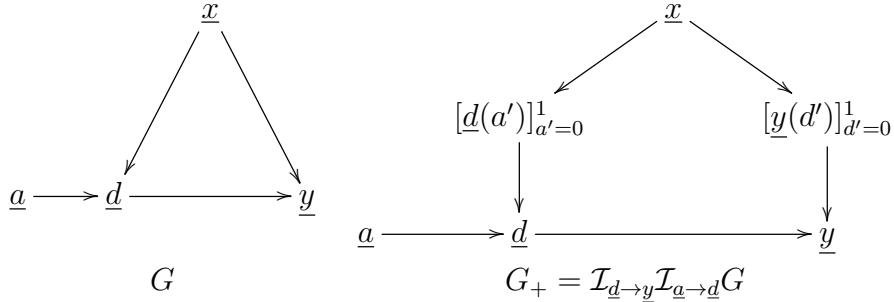


Figure 42.1: Bnet G_+ is bnet G after application of 2 imagine operators. Imagine operators are discussed in Chapter 11.

Let

$\underline{a} \in \{0, 1\}$, instrumental variable, initially assigned dose

$\underline{d} \in \{0, 1\}$, actual treatment dose

$\underline{y} \in \{0, 1\}$, treatment outcome

The definition of the imagine operators used in G_+ stipulates that nodes \underline{y} and \underline{d} in G_+ must be have the following deterministic TPMs (printed in blue below).

$$P(\underline{d}|\underline{a}, [d(a')]^1_{a'=0}) = \mathbb{1}(\underline{d} = \underbrace{\underline{ad}(1) + (1 - \underline{a})\underline{d}(0)}_{= \sum_{a'=0}^1 \delta(a, a')d(a')} = d(a)) \quad (42.1)$$

$$P(\underline{y}|\underline{d}, [y(d')]^1_{d'=0}) = \mathbb{1}(\underline{y} = y(d)) \quad (42.2)$$

	$d^\sigma(0)$	$d^\sigma(1)$
never-takers	0	0
compliers	0	1
defiers	1	0
always-takers	1	1

Table 42.2: Possible compliance behaviors for individual σ .

Table 42.2 gives a name to the 4 possible compliance behaviors that might be exhibited by an individual σ of a dataset. Below, we will use \mathcal{C} to denote the conditions that define a **complier**:

$$\mathcal{C} = \{\underline{d}(0) = 0, \underline{d}(1) = 1\} \quad (42.3)$$

Monotonicity is said to hold if

$$d^\sigma(1) \geq d^\sigma(0) \quad (42.4)$$

Note that monotonicity rules out defiers (i.e., $d^\sigma(1) = 0$, $d^\sigma(0) = 1$), but allows the other 3 compliance behaviors.

It is convenient to define the following expected values:

$$\mathcal{D}_{|a} = \sum_d d P(d|a) = E_{|a}[\underline{d}] \quad (42.5)$$

$$\mathcal{Y}_{|d,a} = \sum_y y P(y|d, a) = E_{|d,a}[\underline{y}] \quad (42.6)$$

$$\mathcal{Y}_{|\underline{d}=d} = \sum_y y P(y|d) = E_{|d}[\underline{y}] \quad (42.7)$$

$$\mathcal{Y}_{|\underline{a}=a} = \sum_y y P(y|a) = E_{|a}[\underline{y}] \quad (42.8)$$

Assume that $\mathcal{D}_{|1} \neq \mathcal{D}_{|0}$. Then LATE is defined by

$$LATE = \frac{\mathcal{Y}_{|\underline{a}=1} - \mathcal{Y}_{|\underline{a}=0}}{\mathcal{D}_{|1} - \mathcal{D}_{|0}} \quad (42.9)$$

Claim 67 *If monotonicity is satisfied, then*

$$P(\mathcal{C}) = \mathcal{D}_{|1} - \mathcal{D}_{|0} \quad (42.10)$$

proof:

$$\mathcal{D}_{|1} - \mathcal{D}_{|0} = \sum_{d(0), d(1)} P(d(0), d(1)) [d(1) - d(0)] \quad (42.11)$$

$$= \begin{cases} P(d(0) = 0, d(1) = 0) \underbrace{[d(1) - d(0)]}_0 \\ + P(d(0) = 0, d(1) = 1) \underbrace{[d(1) - d(0)]}_1 \\ + \underbrace{P(d(0) = 1, d(1) = 0)}_{=0 \text{ by monotonicity}} [d(1) - d(0)] \\ + P(d(0) = 1, d(1) = 1) \underbrace{[d(1) - d(0)]}_0 \end{cases} \quad (42.12)$$

$$= P(\mathcal{C}) \quad (42.13)$$

QED

Claim 68 Recall

$$ATE = E[\underline{y}(1) - \underline{y}(0)] \quad (42.14)$$

If monotonicity is satisfied, then

$$LATE = E_{|\mathcal{C}}[\underline{y}(1) - \underline{y}(0)] \quad (42.15)$$

proof:

$$\mathcal{Y}_{|\underline{a}=1} - \mathcal{Y}_{|\underline{a}=0} = E_{|\underline{a}=1}[\underline{y}] - E_{|\underline{a}=0}[\underline{y}] \quad (42.16)$$

$$= \sum_x \sum_{y(0), y(1)} \sum_{d(0), d(1)} \left\{ \begin{array}{l} [y(d(1)) - y(d(0))] \\ *P(y(0), y(1)|x) \\ *P(d(0), d(1)|x) \\ *P(x) \end{array} \right\} \quad (42.17)$$

$$= \sum_{y(0), y(1)} \sum_{d(0), d(1)} [y(d(1)) - y(d(0))] \left\{ \begin{array}{l} P(y(0), y(1)|d(0), d(1)) \\ *P(d(0), d(1)) \end{array} \right\} \quad (42.18)$$

$$= \sum_{y(0), y(1)} \left\{ \begin{array}{l} \underbrace{[y(0) - y(0)]}_0 \left\{ \begin{array}{l} P(y(0), y(1)|d(0) = 0, d(1) = 0) \\ *P(d(0) = 0, d(1) = 0) \end{array} \right\} \\ + [y(1) - y(0)] \left\{ \begin{array}{l} P(y(0), y(1)|d(0) = 0, d(1) = 1) \\ *P(d(0) = 0, d(1) = 1) \end{array} \right\} \\ + [y(0) - y(1)] \left\{ \begin{array}{l} P(y(0), y(1)|d(0) = 0, d(1) = 0) \\ *P(\underbrace{d(0) = 1, d(1) = 0}_{=0 \text{ monotonicity}}) \end{array} \right\} \\ + \underbrace{[y(1) - y(1)]}_0 \left\{ \begin{array}{l} P(y(0), y(1)|d(0) = 1, d(1) = 1) \\ *P(d(0) = 1, d(1) = 1) \end{array} \right\} \end{array} \right\} \quad (42.19)$$

$$= P(\mathcal{C}) \sum_{y(0), y(1)} [y(1) - y(0)] P(y(0), y(1)|\mathcal{C}) \quad (42.20)$$

$$= P(\mathcal{C}) E_{|\mathcal{C}}[\underline{y}(1) - \underline{y}(0)] \quad (42.21)$$

QED

It is instructive to evaluate LATE for the special case in which G of Fig.42.1 is an LDEN (Linear Deterministic with External Noise) bnet.⁴

Consider Fig.42.2. From that figure

⁴LDEN bnets are discussed in Chapter 44.

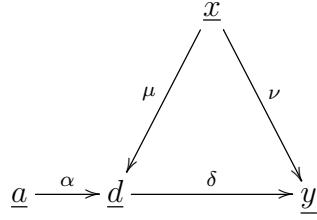


Figure 42.2: LDEN bnet for LATE

$$\underline{d} = \alpha \underline{a} + \mu \underline{x} \quad (42.22)$$

$$\underline{y} = \delta \underline{d} + \nu \underline{x} \quad (42.23)$$

Note that

$$\langle \underline{x}, \underline{a} \rangle = 0 \quad (42.24)$$

because both possible paths from \underline{a} to \underline{x} (i.e., $\underline{a} - \underline{d} - \underline{x}$ and $\underline{a} - \underline{d} - \underline{y} - \underline{x}$) are blocked by colliders.

Claim 69 *For the LDEN bnet of Fig.42.2,*

$$LATE = \delta = \frac{\langle \underline{y}, \underline{a} \rangle}{\langle \underline{d}, \underline{a} \rangle} \quad (42.25)$$

proof:

$$\langle \underline{d}, \underline{a} \rangle = \alpha \langle \underline{a}, \underline{a} \rangle \quad (42.26)$$

$$\langle \underline{y}, \underline{a} \rangle = \langle \delta(\alpha \underline{a} + \mu \underline{x}) + \nu \underline{x}, \underline{a} \rangle \quad (42.27)$$

$$= \alpha \delta \langle \underline{a}, \underline{a} \rangle \quad (42.28)$$

Hence,

$$\frac{\langle \underline{y}, \underline{a} \rangle}{\langle \underline{d}, \underline{a} \rangle} = \delta \quad (42.29)$$

$$E_{\underline{a}=1}[\underline{y}] - E_{\underline{a}=0}[\underline{y}] = E_{\underline{a}=1}[\delta \underline{d} + \nu \underline{x}] - E_{\underline{a}=0}[\delta \underline{d} + \nu \underline{x}] \quad (42.30)$$

$$= \delta(E_{\underline{a}=1}[\underline{d}] - E_{\underline{a}=0}[\underline{d}]) + \nu \underbrace{(E_{\underline{a}=1}[\underline{x}] - E_{\underline{a}=0}[\underline{x}])}_0 \quad (42.31)$$

QED

Chapter 43

Linear and Logistic Regression

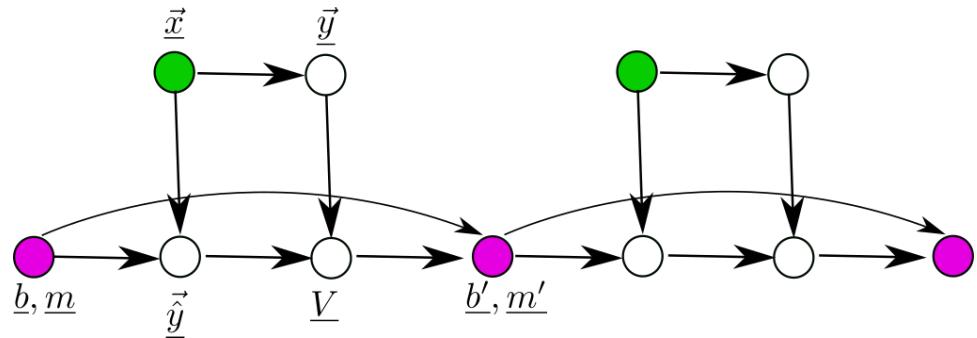


Figure 43.1: Linear Regression

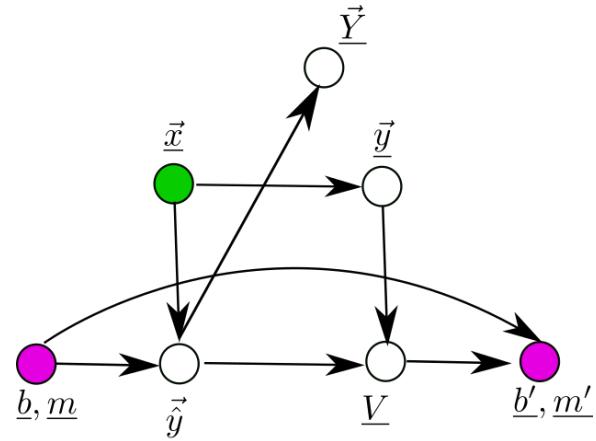


Figure 43.2: Bnet of Fig.43.1 with new \vec{Y} node.

Estimators \hat{y} for linear and logistic regression.

- **Linear Regression:** $y \in \mathbb{R}$. Note $\hat{y} \in \mathbb{R}$. $(x, \hat{y}(x))$ is the graph of a straight

line with y-intercept b and slope m .

$$\hat{y}(x; b, m) = b + mx \quad (43.1)$$

- **Logistic Regression:** $y \in \{0, 1\}$. Note $\hat{y} \in [0, 1]$. $(x, \hat{y}(x))$ is the graph of a sigmoid. Often in literature, b, m are replaced by β_0, β_1 .

$$\hat{y}(x; b, m) = \text{smoid}(b + mx) \quad (43.2)$$

Define

$$V(b, m) = \sum_{x,y} P(x, y) |y - \hat{y}(x; b, m)|^2. \quad (43.3)$$

We want to minimize $V(b, m)$ (called a cost or loss function) wrt b and m .

The TPMs, printed in blue, for the Bnet Fig.43.1, are as follows.

$$P(b, m) = \text{given} \quad (43.4)$$

The first time it is used, (b, m) is arbitrary. After the first time, it is determined by previous stage.

Let

$$P_{\underline{x}, \underline{y}}(x, y) = \frac{1}{nsam(\vec{x})} \sum_{\sigma} \mathbb{1}(x = x^{\sigma}, y = y^{\sigma}). \quad (43.5)$$

$$P(\vec{x}) = \prod_{\sigma} P(x^{\sigma}) \quad (43.6)$$

$$P(\vec{y}|\vec{x}) = \prod_{\sigma} P(y^{\sigma} | x^{\sigma}) \quad (43.7)$$

$$P(\vec{\hat{y}}|\vec{x}, b, m) = \prod_{\sigma} \delta(\hat{y}^{\sigma}, \hat{y}(x^{\sigma}, b, m)) \quad (43.8)$$

$$P(V|\vec{\hat{y}}, \vec{y}) = \delta(V, \frac{1}{nsam(\vec{x})} \sum_{\sigma} |y^{\sigma} - \hat{y}^{\sigma}|^2) \quad (43.9)$$

Let $\eta_b, \eta_m > 0$. For $x = b, m$, if $x' - x = \Delta x = -\eta \frac{\partial V}{\partial x}$, then $\Delta V \approx \frac{-1}{\eta} (\Delta x)^2 \leq 0$ for $\eta > 0$. This is called “gradient descent”.

$$P(b'|V, b) = \delta(b', b - \eta_b \partial_b V) \quad (43.10)$$

$$P(m'|V, m) = \delta(m', m - \eta_m \partial_m V) \quad (43.11)$$

43.1 Generalization to x with multiple components (features)

Suppose that for each sample σ , instead of x^σ being a scalar, it has n components called features:

$$x^\sigma = (x_0^\sigma, x_1^\sigma, x_2^\sigma, \dots, x_{n-1}^\sigma) . \quad (43.12)$$

Slope m is replaced by weights

$$w = (w_0, w_1, w_2, \dots, w_{n-1}) , \quad (43.13)$$

and the product of 2 scalars mx^σ is replaced by the inner vector product $w^T x^\sigma$.

43.2 Alternative $V(b, m)$ for logistic regression

For logistic regression, since $y^\sigma \in \{0, 1\}$ and $\hat{y}^\sigma \in [0, 1]$ are both in the interval $[0, 1]$, they can be interpreted as probabilities. Define probability distributions $p^\sigma(x)$ and $\hat{p}^\sigma(x)$ for $x \in \{0, 1\}$ by

$$p^\sigma(1) = y^\sigma, \quad p^\sigma(0) = 1 - y^\sigma \quad (43.14)$$

$$\hat{p}^\sigma(1) = \hat{y}^\sigma, \quad \hat{p}^\sigma(0) = 1 - \hat{y}^\sigma \quad (43.15)$$

Then for logistic regression, the following 2 cost functions $V(b, m)$ can be used as alternatives to the cost function Eq.(43.3) previously given.

$$V(b, m) = \frac{1}{nsam(\vec{x})} \sum_{\sigma} D_{KL}(p^\sigma \| \hat{p}^\sigma) \quad (43.16)$$

and

$$V(b, m) = \frac{1}{nsam(\vec{x})} \sum_{\sigma} CE(p^\sigma \| \hat{p}^\sigma) \quad (43.17)$$

$$= \frac{-1}{nsam(\vec{x})} \sum_{\sigma} \{y^\sigma \ln \hat{y}^\sigma + (1 - y^\sigma) \ln(1 - \hat{y}^\sigma)\} \quad (43.18)$$

$$= \frac{-1}{nsam(\vec{x})} \sum_{\sigma} \ln \{(\hat{y}^\sigma)^{y^\sigma} (1 - \hat{y}^\sigma)^{(1-y^\sigma)}\} \quad (43.19)$$

$$= \frac{-1}{nsam(\vec{x})} \sum_{\sigma} \ln P(\underline{Y} = y^\sigma | \hat{y} = \hat{y}^\sigma) \quad (43.20)$$

$$= - \sum_{x,y} P(x, y) \ln P(\underline{Y} = y | \hat{y} = \hat{y}(x, b, m)) \quad (43.21)$$

Above, we used

$$P(\underline{Y} = Y | \hat{y}) = \hat{y}^Y [1 - \hat{y}]^{1-Y} \quad (43.22)$$

for $Y \in S_{\underline{Y}} = \{0, 1\}$. (Bernoulli distribution).

There is no node corresponding to \underline{Y} in the Bnet of Fig.43.1. Fig.43.2 shows a new Bnet that has a new node called $\vec{\underline{Y}}$ compared to the Bnet of Fig.43.1. One defines the TPMs for all nodes of Fig.43.2 except $\vec{\underline{Y}}$ and \underline{V} the same as for Fig.43.1. For $\vec{\underline{Y}}$ and \underline{V} , one defines

$$P(Y^\sigma | \vec{\hat{y}}) = P(\underline{Y} = Y^\sigma | \hat{y}^\sigma) \quad (43.23)$$

$$P(V|\vec{\underline{Y}}, \vec{y}) = \delta(V, \frac{-1}{nsam(\vec{x})} \ln L), \quad (43.24)$$

where $L = \prod_\sigma P(\underline{Y} = y^\sigma | \hat{y}^\sigma)$ =likelihood.

Chapter 44

Linear Deterministic Bnets with External Noise

In this chapter, we will consider bnets which were referred to, prior to the invention of bnets, as: Sewall Wright's **Path Analysis (PA)** and **linear Structural Equations Models (SEM)**. Judea Pearl in his books calls them **linear Structural Causal Models (SCM)**, because they are very convenient for doing causal analysis. We will refer to them as linear Deterministic with External Noise (LDEN) diagrams. This chapter is devoted to LDEN diagrams, except that we will say a few words about non-linear DEN diagrams at the end.

A **DEN diagram** is a special kind of bnet. To build a DEN diagram, start with a deterministic bnet G . The deterministic nodes of G are called the **endogenous (internal) variables**. Now make a bigger bnet \bar{G} called a DEN diagram by adding to each node a of G a non-deterministic root node u_a pointing into a only. The nodes u_a are called the **exogenous (external) variables**. The exogenous variables make their children noisy. They are assumed to be unobserved and their TPMs are prior probability distributions. Since they are root nodes, they are mutually independent. When we draw a DEN diagram, we will never draw the exogenous nodes, leaving them implicit.

A **linear DEN diagram (LDEN)** is a DEN diagram whose deterministic nodes have a TPM that is a linear function of the states of the parent nodes.

44.1 Example of LDEN diagram

The TPMs, printed in blue, for the LDEN diagram Fig.44.1, are as follows.

$$P(y|w, z, u_y) = \mathbb{1}(y = \epsilon w + \delta z + u_y) \quad (44.1)$$

$$P(w|x, z, u_w) = \mathbb{1}(w = \beta x + \gamma z + u_w) \quad (44.2)$$

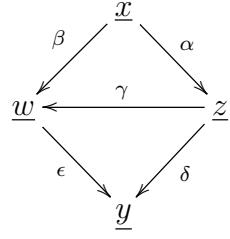


Figure 44.1: Example of a LDEN diagram wherein \underline{x} splits into two nodes \underline{z} and \underline{w} , then merges into node \underline{y} . There is also an arrow $\underline{z} \rightarrow \underline{w}$. Exogenous nodes are not shown. The Greek letters represent real numbers.

$$P(z|x, u_{\underline{z}}) = \mathbb{1}(z = \alpha x + u_{\underline{z}}) \quad (44.3)$$

$$P(x|u_{\underline{x}}) = \mathbb{1}(x = u_{\underline{x}}) \quad (44.4)$$

Hence,

$$y = \epsilon w + \delta z + u_{\underline{y}} \quad (44.5)$$

$$= \epsilon(\beta x + \gamma z + u_{\underline{w}}) + \delta z + u_{\underline{y}} \quad (44.6)$$

$$= (\epsilon\gamma + \delta)z + \epsilon\beta x + \epsilon u_{\underline{w}} + u_{\underline{y}} \quad (44.7)$$

$$= (\epsilon\gamma + \delta)z + \epsilon\beta u_{\underline{x}} + \epsilon u_{\underline{w}} + u_{\underline{y}} . \quad (44.8)$$

Therefore

$$\left(\frac{\partial y}{\partial z} \right)_{u_{\underline{w}}, u_{\underline{y}}} = \epsilon\gamma + \delta , \quad (44.9)$$

where the partial derivative holds fixed all exogenous variables except $u_{\underline{z}}$. Note that this partial derivative is a sum of terms, and that each of those terms represents a different directed path from \underline{z} to $\underline{y}(\underline{z})$. This is a general property of LDEN diagrams.

44.2 Fully Connected LDEN diagrams

The bnets that will be considered in this section will all be fully connected. Fully connected bnets are defined in Chapter D. This section uses the notation $\langle \underline{x}, \underline{y} \rangle$ for the covariance of any two random variables $\underline{x}, \underline{y}$. This $\langle \underline{x}, \underline{y} \rangle$ notation is defined in Chapter C.

Consider a LDEN diagram with deterministic nodes $\underline{x} = (\underline{x}_k)_{k=0,1,\dots,nx-1}$ and corresponding exogenous nodes $\underline{u} = (\underline{u}_k)_{k=0,1,\dots,nx-1}$. Assume $\langle \underline{u}_i, \underline{u}_j \rangle = 0$ if $i \neq j$. The strength of each connection $\underline{x}_i \rightarrow \underline{x}_j$ of the LDEN diagram is measured by a **structural coefficient** $\alpha_{j|i} \in \mathbb{R}$. Some of the $\alpha_{j|i}$ may be zero, in which case the corresponding arrow $i \rightarrow j$ would not be drawn.

44.2.1 Fully connected LDEN diagram with $nx = 2$

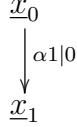


Figure 44.2: Fully connected LDEN diagram with two \underline{x}_j nodes (exogenous nodes \underline{u}_j not shown).

Consider the LDEN diagram of Fig.44.2. This diagram represents the following **structural equations**:

$$\underline{x}_0 = \underline{u}_0 \quad (44.10a)$$

$$\underline{x}_1 = \alpha_{1|0}\underline{x}_0 + \underline{u}_1. \quad (44.10b)$$

Eqs.44.10 constitute a system of 2 linear equations in 2 unknowns (the \underline{x} 's) so we can solve for the \underline{x} 's in terms of the α 's and \underline{u} 's.

Note also that

$$\langle \underline{x}_1, \underline{x}_0 \rangle = \alpha_{1|0} \langle \underline{x}_0, \underline{x}_0 \rangle. \quad (44.11)$$

Thus, $\alpha_{1|0}$ can be estimated from the covariances $\langle \underline{x}_i, \underline{x}_j \rangle$.

44.2.2 Fully connected LDEN diagram with $nx = 3$

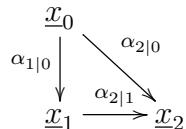


Figure 44.3: Fully connected LDEN diagram with three \underline{x}_j nodes (exogenous nodes \underline{u}_j not shown).

Consider the LDEN diagram of Fig.44.3. This diagram represents the following **structural equations**:

$$\underline{x}_0 = \underline{u}_0 \quad (44.12a)$$

$$\underline{x}_1 = \alpha_{1|0}\underline{x}_0 + \underline{u}_1 \quad (44.12b)$$

$$\underline{x}_2 = \alpha_{2|0}\underline{x}_0 + \alpha_{2|1}\underline{x}_1 + \underline{u}_2 . \quad (44.12c)$$

Eqs.44.12 constitute a system of 3 linear equations in 3 unknowns (the \underline{x} 's) so we can solve for the \underline{x} 's in terms of the α 's and \underline{u} 's.

Note also that

$$\langle \underline{x}_1, \underline{x}_0 \rangle = \alpha_{1|0} \langle \underline{x}_0, \underline{x}_0 \rangle \quad (44.13a)$$

$$\langle \underline{x}_2, \underline{x}_0 \rangle = \alpha_{2|0} \langle \underline{x}_0, \underline{x}_0 \rangle + \alpha_{2|1} \langle \underline{x}_1, \underline{x}_0 \rangle \quad (44.13b)$$

$$\langle \underline{x}_2, \underline{x}_1 \rangle = \alpha_{2|0} \langle \underline{x}_0, \underline{x}_1 \rangle + \alpha_{2|1} \langle \underline{x}_1, \underline{x}_1 \rangle \quad (44.13c)$$

Eqs.44.13 constitute a system of 3 linear equations in 3 unknowns (the α 's) so we can solve for the α 's in terms of covariances $\langle \underline{x}_i, \underline{x}_j \rangle$. This gives an estimate for the α 's.

44.2.3 Fully connected LDEN diagram with arbitrary nx

Let $\underline{x}_. = (x_i)_{i=0,1,\dots,nx-1}$ and $\underline{x}_{<i} = (x_k)_{k=0,1,\dots,i-1}$. Consider a fully connected LDEN diagram with deterministic nodes labeled \underline{x}_i . The \underline{x}_i labels are assumed to be in **topological order** (i.e., the parents of node \underline{x}_i are $\underline{x}_{<i}$). Let the TPMs, printed in blue, for the nodes \underline{x}_i of the LDEN diagram, be as follows.

$$P(x_i | \underline{x}_{<i}, \underline{u}_i) = \mathbb{1}(x_i = \sum_{k < i} \alpha_{i|k} x_k + u_i) , \quad (44.14)$$

for some parameters $\alpha_{i|k} \in \mathbb{R}$. The exogenous nodes $\underline{u}_.$ are assumed to be independent so

$$P(\underline{u}_.) = \prod_i P(u_i) \quad (44.15)$$

and

$$\langle \underline{u}_i, \underline{u}_j \rangle = 0 \text{ if } i \neq j . \quad (44.16)$$

Note that

$$P(x_.) = \sum_{\underline{u}_.} P(\underline{u}_.) \prod_i P(x_i | \underline{x}_{<i}, u_i) \quad (44.17)$$

$$= E_{\underline{u}_.} [\prod_i P(x_i | \underline{x}_{<i}, u_i)] . \quad (44.18)$$

In terms of random variables, this system is described by the following **structural equations**:

$$\underline{x}_i = \sum_{k < i} \alpha_{i|k} \underline{x}_k + \underline{u}_i . \quad (44.19)$$

The structural equations can be written in matrix form as follows. Define a strictly lower triangular matrix A with the connection strengths $\alpha_{i|k} \in \mathbb{R}$ as entries. For example, for $nx = 4$,

$$A = \begin{bmatrix} 0 & 0 & 0 & 0 \\ \alpha_{1|0} & 0 & 0 & 0 \\ \alpha_{2|0} & \alpha_{2|1} & 0 & 0 \\ \alpha_{3|0} & \alpha_{3|1} & \alpha_{3|2} & 0 \end{bmatrix} . \quad (44.20)$$

If we now represent the multinodes $\underline{x}_.$ and $\underline{u}_.$ as column vectors \underline{x} and \underline{u} , we get

$$\underline{x} = A\underline{x} + \underline{u} . \quad (44.21)$$

Note that

$$\underline{x} = (1 - A)^{-1} \underline{u} . \quad (44.22)$$

Therefore,

$$\underline{x}_i = f_i(\underline{u}_{\leq i}) . \quad (44.23)$$

Therefore, if $i > j$,

$$\langle \underline{u}_i, \underline{x}_j \rangle = \langle \underline{u}_i, f_j(\underline{u}_{\leq j}) \rangle = 0 . \quad (44.24)$$

Thus, if $i > j$,

$$\langle \underline{x}_i, \underline{x}_j \rangle = \sum_{k < i} \alpha_{i|k} \langle \underline{x}_k, \underline{x}_j \rangle + \langle \underline{u}_i, \underline{x}_j \rangle \quad (44.25)$$

$$= \sum_{k < i} \alpha_{i|k} \langle \underline{x}_k, \underline{x}_j \rangle . \quad (44.26)$$

In matrix notation, Eq.(44.26) becomes

$$\langle \underline{x}, \underline{x}^T \rangle_L = A[\langle \underline{x}, \underline{x}^T \rangle_L + \langle \underline{x}, \underline{x}^T \rangle_D] \quad (44.27)$$

where we are using $\langle \underline{x}, \underline{x}^T \rangle_{i,j} = \langle \underline{x}_i, \underline{x}_j \rangle$ and denoting the strictly lower triangular part and diagonal part of a matrix M by M_L and M_D . Thus,

$$A = \langle \underline{x}, \underline{x}^T \rangle_L [\langle \underline{x}, \underline{x}^T \rangle_L + \langle \underline{x}, \underline{x}^T \rangle_D]^{-1} . \quad (44.28)$$

This gives an estimate for the α 's in terms of the covariances $\langle \underline{x}_i, \underline{x}_j \rangle$.

44.3 Non-linear DEN diagrams

This chapter is dedicated to linear DEN diagrams. This implicitly assumes that the deterministic nodes \underline{x}_i of the DEN diagram have an interval of real values as their possible states. A trivial but very useful generalization of linear DEN diagrams is to replace Eq.(44.14) for the TPMs of the deterministic nodes of the diagram by

$$P(x_i|x_{<i}, u_i) = \mathbb{1}(x_i = f_i(x_{<i}, u_i)) , \quad (44.29)$$

with structural equations

$$\underline{x}_i = f_i(\underline{x}_{<i}, \underline{u}_i) , \quad (44.30)$$

for $i = 0, 1, \dots, nx - 1$. Here the f_i are possibly non-linear functions that depend the states $x_{<i}$ and u_i of nodes $\underline{x}_{<i}$ and \underline{u}_i . If a node \underline{x}_i has no arrows entering it (i.e., is a root node), then

$$P(x_i|x_{<i}, u_i) = P(x_i) = \delta(x_i, a) \quad (44.31)$$

and

$$\underline{x}_i = a \quad (44.32)$$

for some $a \in S_{\underline{x}_i}$.

With this generalization, we can make any $f_i()$ represent a continuous probability distribution such as a Gaussian, or a discrete-valued Boolean function such as an OR gate.

Eqs.(44.29) and (44.30) are the TPMs and structural equations for a fully connected, non-linear DEN diagram. For a non-fully connected diagram,

- replace the multinode $x_{<i}$ by a subset of itself, in Eqs.(44.29) and (44.30) , and
- delete the corresponding arrows from the graph.

Chapter 45

Markov Blankets

This chapter is based on the Wikipedia article, Ref.[121]. Markov blankets and Markov boundaries of bnets were apparently invented by Judea Pearl. His 1988 book Ref.[44], instead of a research paper, is usually given as the original reference.

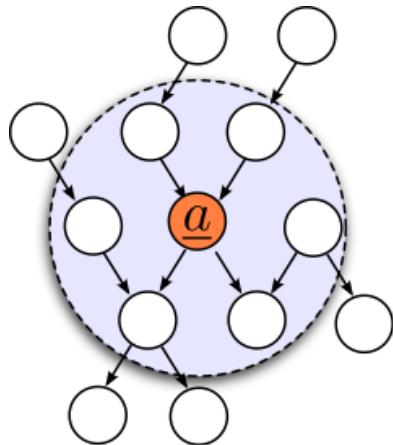


Figure 45.1: In a bnet, the minimal Markov blanket, a.k.a. Markov boundary, of node \underline{a} .

We will treat vectors of random variables as if they were sets when using the \in , \subset and $-$ operations. For example, if $\underline{x} = (\underline{x}_0, \underline{x}_1, \underline{x}_2, \underline{x}_3)$ and $\underline{b} = (\underline{x}_1, \underline{x}_2)$, then $\underline{x}_1 \in \underline{b} \subset \underline{x}$ and $\underline{x} - \underline{b} = (\underline{x}_0, \underline{x}_3)$.

Below, $H(\underline{a} : \underline{b} | \underline{c})$ denotes the conditional mutual information of random variables \underline{a} and \underline{b} conditioned on random variable \underline{c} . $H(\underline{a} : \underline{b} | \underline{c})$ is used in Shannon Information Theory, where it is defined by

$$H(\underline{a} : \underline{b} | \underline{c}) = \sum_{a,b,c} P(a, b, c) \ln \frac{P(a, b | c)}{P(a | c)P(b | c)}. \quad (45.1)$$

$H(\underline{a} : \underline{b} | \underline{c}) = 0$ iff \underline{a} and \underline{b} are independent (uncorrelated) when \underline{c} is held fixed.
Suppose $\underline{a} \in \underline{X}$, $\underline{B} \subset \underline{X}$, but $\underline{a} \notin \underline{B}$. Then \underline{B} is a **Markov blanket** of \underline{a} if

$$H(\underline{a} : \underline{X} - \underline{a} | \underline{B}) = 0 . \quad (45.2)$$

In other words, one may assume that \underline{a} depends on \underline{B} only, and is independent of all random variables in $\underline{X} - (\underline{a} \cup \underline{B})$.

The **minimal Markov blanket** is called the **Markov boundary**.

In a bnet, the Markov boundary of a node \underline{a} , contains:

1. the parents of \underline{a} ,
2. the children of \underline{a} ,
3. the parents, other than \underline{a} , of the children of \underline{a} .

This is illustrated in Fig.45.1.

From the d-separation theorem (see Chapter 21), we get an intuitive motivation for the definition of Markov boundary. By conditioning on the parents and children of node \underline{a} , we block almost all, but not all, information from reaching node \underline{a} . The reason not all info is blocked is that conditioning on a child \underline{c} of \underline{a} creates paths from a parent of \underline{c} to \underline{c} to \underline{a} , wherein node \underline{c} acts as a conditioned collider. To block such paths, we must condition on the parents of \underline{c} too.

Chapter 46

Markov Chain Monte Carlo (MCMC)

Monte Carlo methods are methods for using random number generation to sample probability distributions. The subject of Monte Carlo methods has many branches, as you can see from its Wikipedia category list, Ref.[124]. MCMC (Markov Chain Monte Carlo) is just one of those branches, albeit a major one. Metropolis-Hastings (MH) sampling is a very important MCMC method. Gibbs sampling is a special case of MH sampling. This chapter covers both, MH and Gibbs sampling. It also covers a few other types of sampling.

Throughout this chapter, we use $P_{\underline{x}} : S_{\underline{x}} \rightarrow [0, 1]$ to denote the target probability distribution that we wish to obtain samples from.

46.1 Inverse Cumulative Sampling

For more info about this topic and some original references, see Ref.[110].

This is one of the simplest methods for obtaining samples from a probability distribution $P_{\underline{x}}$, but it requires knowledge of the inverse cumulative distribution of $P_{\underline{x}}$, which is often not available.

The **cumulative distribution** function is defined by:

$$CUM_{\underline{x}}(x) = P(\underline{x} < x) = \int_{x' < x} dx' P_{\underline{x}}(x') . \quad (46.1)$$

Note that

$$P_{\underline{x}}(x) = \frac{d}{dx} CUM_{\underline{x}}(x) . \quad (46.2)$$

For $t = 0, 1, \dots, T - 1$, let

$\underline{u}^{(t)} \in [0, 1]$ = random variable, uniformly distributed over $[0, 1]$.

$\vec{\underline{x}}^{(t)} = (\underline{x}^{(t)}[\sigma])_{\sigma=0,1,\dots,nsam(t)-1}$ where $\underline{x}^{(t)}[\sigma] \in S_{\underline{x}}$ for all σ . Vector of samples collected up to time t .

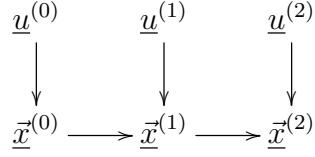


Figure 46.1: bnet for Inverse Cumulative Sampling

The TPMs, printed in blue, for bnet Fig.46.1, are as follows:

$$P(\underline{u}^{(t)}) = 1 \quad (46.3)$$

$$P(\underline{x}^{(t)} | \underline{x}^{(t-1)}, \underline{u}^{(t)}) = \delta(\underline{x}^{(t)}, [\underline{x}^{(t-1)}, CUM_{\underline{x}}^{-1}(\underline{u}^{(t)})]) \quad (46.4)$$

Motivation

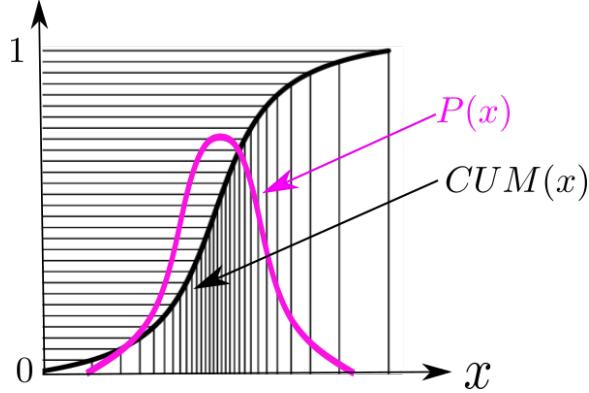


Figure 46.2: Motivation for Inverse Cumulative Sampling.

See Fig.46.2.

Note that if \underline{u} is uniformly distributed over the interval $[0, 1]$ and $a \in [0, 1]$, then

$$P(\underline{u} < a) = a . \quad (46.5)$$

Thus

$$P(CUM_{\underline{x}}^{-1}(\underline{u}) < x) = P(\underline{u} < CUM_{\underline{x}}(x)) \quad (46.6)$$

$$= CUM_{\underline{x}}(x) . \quad (46.7)$$

Therefore,

$$dP(CUM_{\underline{x}}^{-1}(\underline{u}) < x) = P_{\underline{x}}(x)dx . \quad (46.8)$$

46.2 Rejection Sampling

For more info about this topic and some original references, see Ref.[135].

This method samples from a “candidates” probability distribution $P_c : S_x \rightarrow [0, 1]$, in cases where sampling directly from the target probability distribution $P_{\underline{x}} : S_{\underline{x}} \rightarrow [0, 1]$ is not possible.

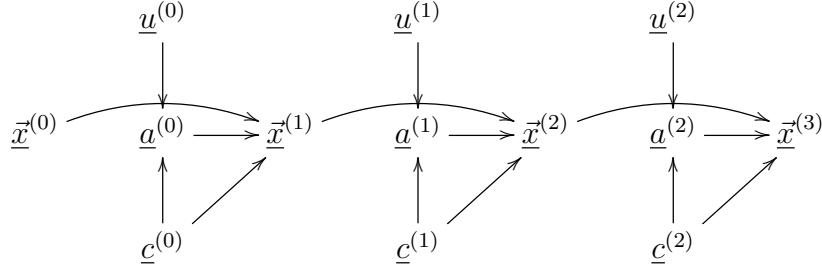


Figure 46.3: bnet for Rejection Sampling

For $t = 0, 1, \dots, T - 1$, let

$\underline{u}^{(t)} \in [0, 1]$ = random variable, uniformly distributed over $[0, 1]$.

$\underline{a}^{(t)} \in \{0, 1\}$ = accept candidate? (no=0, yes=1)

$\underline{c}^{(t)} \in S_x$ = sample that is a candidate for being accepted

$\vec{x}^{(t)} = (\underline{x}^{(t)}[\sigma])_{\sigma=0,1,\dots,nsam(t)-1}$ where $\underline{x}^{(t)}[\sigma] \in S_x$ for all σ . Vector of samples collected up to time t .

This algorithm requires a priori definition of a candidate probability distribution $P_c : S_x \rightarrow \mathbb{R}$ such that

$$P_{\underline{x}}(x) < \beta P_c(x) \quad (46.9)$$

for all $x \in S_x$, for some $\beta \in \mathbb{R}$.

The TPMs, printed in blue, for bnet Fig.46.3, are as follows:

$$P(u^{(t)} = u) = 1 \quad (46.10)$$

$$P(\underline{c}^{(t)} = c) = P_c(c) \quad (46.11)$$

$$P(\underline{a}^{(t)} = a | \underline{c}^{(t)} = c, \underline{u}^{(t)} = u) = \begin{cases} \delta(a, 0) & \text{if } u\beta P_c(c) \geq P_{\underline{x}}(c) \\ \delta(a, 1) & \text{if } u\beta P_c(c) < P_{\underline{x}}(c) \end{cases} \quad (46.12)$$

$$P(\vec{x}^{(t)} | \vec{x}^{(t-1)}, \underline{a}^{(t)} = a, \underline{c}^{(t)} = c) = \begin{cases} \delta(\vec{x}^{(t)}, \vec{x}^{(t-1)}) & \text{if } a = 0 \\ \delta(\vec{x}^{(t)}, [\vec{x}^{(t-1)}, c]) & \text{if } a = 1 \end{cases} \quad (46.13)$$

This last equation is only defined for $t > 0$. For $t = 0$, the left hand side reduces to $P(\vec{x}^{(0)})$ which must be specified a priori.

Motivation

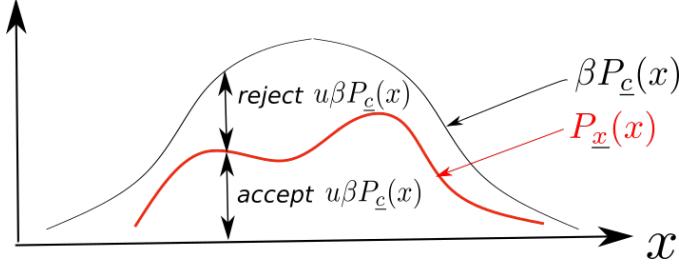


Figure 46.4: Motivation for Rejection Sampling.

See Fig.46.4.

46.3 Metropolis-Hastings Sampling

For more info about this topic and some original references, see Refs.[3] and [122].

An advantage of this method is that it can sample unnormalized probability distributions (*constant*) $P_{\underline{x}}$ because it only uses ratios of $P_{\underline{x}}$ at two different points. Another advantage of this method is that it scales much better than other sampling methods as the number of dimensions of the sampled variable \underline{x} increases.

This method produces samples that take a finite amount of time to reach steady state. The samples are also theoretically correlated instead of being i.i.d. as one desires. To mitigate for the steady state problem, one discards an initial set of samples (the “burn-in” period). To mitigate for the correlation problem, one calculates the autocorrelation between the samples and keeps only samples separated by a time interval after which the samples cease to be autocorrelated to a good approximation.

For $t = 0, 1, \dots, T - 1$, let

$\underline{u}^{(t)} \in [0, 1]$ = random variable, uniformly distributed over $[0, 1]$.

$\underline{a}^{(t)} \in \{0, 1\}$ = accept candidate? (no=0, yes=1)

$\underline{c}^{(t)} \in S_{\underline{x}}$ = sample that is a candidate for being accepted

$\underline{m}^{(t)} \in S_{\underline{x}}$ = memory of last accepted sample

$\vec{\underline{x}}^{(t)} = (\underline{x}^{(t)}[\sigma])_{\sigma=0,1,\dots,n_{sam}(t)-1}$ where $\underline{x}^{(t)}[\sigma] \in S_{\underline{x}}$ for all σ . Vector of samples collected up to time t .

A **proposal TPM** $P_{c|\underline{x}} : S_{\underline{x}}^2 \rightarrow [0, 1]$ must be specified a priori for this algorithm.

The TPMs, printed in blue, for bnet Fig.46.5, are as follows:

$$P(u^{(t)} = u) = 1 \quad (46.14)$$

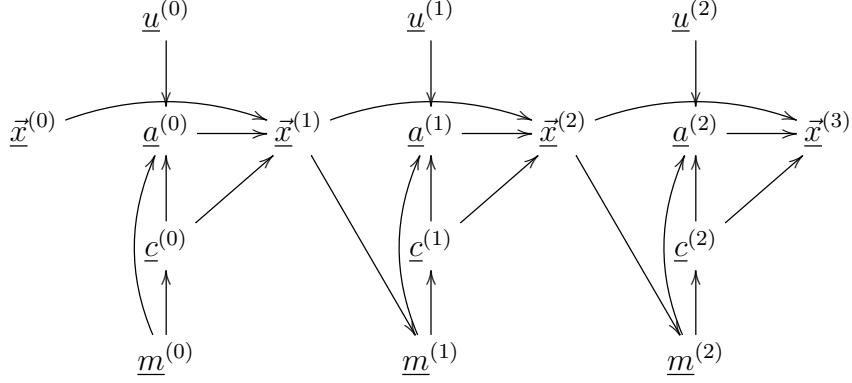


Figure 46.5: bnet for Metropolis-Hastings Sampling

$$P(\underline{c}^{(t)} = c | \underline{m}^{(t)} = m) = P_{\underline{c}|\underline{x}}(c|m) \quad (46.15)$$

$$P(\underline{a}^{(t)} = a | \underline{c}^{(t)} = c, \underline{u}^{(t)} = u, \underline{m}^{(t)} = m) = \begin{cases} \delta(a, 0) & \text{if } u \geq \alpha(c|m) \\ \delta(a, 1) & \text{if } u < \alpha(c|m) \end{cases} \quad (46.16)$$

where the **acceptance probability** α is defined as

$$\alpha(c|m) = \min \left(1, \frac{P_{\underline{c}|\underline{x}}(m|c) P_{\underline{x}}(c)}{P_{\underline{c}|\underline{x}}(c|m) P_{\underline{x}}(m)} \right). \quad (46.17)$$

Note that if the proposal distribution is symmetric, then

$$\alpha(c|m) = \min \left(1, \frac{P_{\underline{x}}(c)}{P_{\underline{x}}(m)} \right). \quad (46.18)$$

$$P(\vec{x}^{(t)} | \vec{x}^{(t-1)}, \underline{a}^{(t)} = a, \underline{c}^{(t)} = c) = \begin{cases} \delta(\vec{x}^{(t)}, \vec{x}^{(t-1)}) & \text{if } a = 0 \\ \delta(\vec{x}^{(t)}, [\vec{x}^{(t-1)}, c]) & \text{if } a = 1 \end{cases} \quad (46.19)$$

This last equation is only defined for $t > 0$. For $t = 0$, the left hand side reduces to $P(\vec{x}^{(0)})$ which must be specified a priori.

$$P(\underline{m}^{(t)} = m | \vec{x}^{(t)}) = \delta(m, \text{last component of } \vec{x}^{(t)}). \quad (46.20)$$

This last equation is only defined for $t > 0$. For $t = 0$, the left hand side reduces to $P(\underline{m}^{(0)} = m)$ which must be specified a priori.

Motivation

See Fig.46.6.

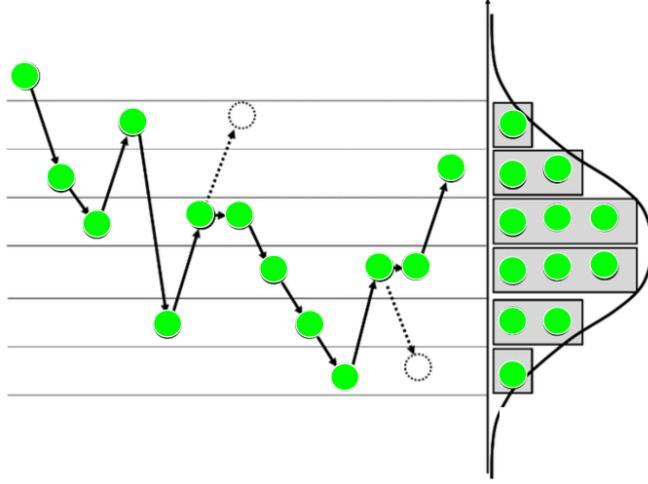


Figure 46.6: Motivation for Metropolis-Hastings Sampling.

Consider a time homogeneous (its TPM is the same for all times) Markov chain with TPM $P(x'|x) = [T]_{x',x}$. Its **stationary distribution**, if it exists, is defined as

$$\pi = \lim_{n \rightarrow \infty} T^n \pi_0 . \quad (46.21)$$

Suppose the prob distribution $P_{\underline{x}}(x)$ that we wish to sample from satisfies

$$P_{\underline{x}}(x) = \pi(x) . \quad (46.22)$$

Reversibility (detailed balance): For all $x, x' \in S_{\underline{x}}$,

$$P(x'|x)\pi(x) = P(x|x')\pi(x') . \quad (46.23)$$

Detailed balance is a sufficient (although not necessary) condition for a unique stationary prob distribution π to exist.¹

Let

$$P(x'|x) = P(\underline{a} = 1|x', x)P_{c|x}(x'|x) + \delta(x, x')P(\underline{a} = 0|x) , \quad (46.24)$$

where

$$P(\underline{a} = 0|x) = \sum_{x'} P(\underline{a} = 0|x', x)P_{c|x}(x'|x) . \quad (46.25)$$

Claim 70 *If*

¹As explained lucidly in Ref.[3], besides detailed balance, 2 other properties must also be satisfied by the Markov chain, irreducibility and aperiodicity. However, because of how it is constructed, the Metropolis-Hastings algorithm automatically produces a Markov chain that has those 2 properties.

$$P(\underline{a} = 1|x', x) = \alpha(x'|x), \quad (46.26)$$

then detailed balance is satisfied.

proof: Assume $x \neq x'$.

$$P(x'|x)P(x) = P(\underline{a} = 1|x', x)P_{\underline{c}|\underline{x}}(x'|x)P_{\underline{x}}(x) \quad (46.27)$$

$$= \min \left(1, \frac{P_{\underline{c}|\underline{x}}(x|x')P_{\underline{x}}(x')}{P_{\underline{c}|\underline{x}}(x'|x)P_{\underline{x}}(x)} \right) P_{\underline{c}|\underline{x}}(x'|x)P_{\underline{x}}(x) \quad (46.28)$$

$$= \min (P_{\underline{c}|\underline{x}}(x'|x)P_{\underline{x}}(x), P_{\underline{c}|\underline{x}}(x|x')P_{\underline{x}}(x')) \quad (46.29)$$

$$= P(x|x')P(x') \quad (46.30)$$

QED

46.4 Gibbs Sampling

For more info about this topic and some original references, see Ref.[104].

Gibbs sampling is a special case of Metropolis-Hastings sampling. Gibbs sampling is ideally suited for application to a bnet, because it is stated in terms of the conditional prob distributions of N random variables, and conditional prob distributions are part of the definition of a bnet.

Consider a bnet with nodes $\underline{x}_0, \underline{x}_1, \dots, \underline{x}_{N-1}$

Identify the random variable $\underline{x} = (\underline{x}_0, \underline{x}_1, \dots, \underline{x}_{N-1})$ with the random variable \underline{x} used in Metropolis-Hastings sampling. For Gibbs sampling, we use the following proposal distribution:

$$P_{\underline{c}|\underline{x}}(c|m) = \prod_{j=0}^{N-1} P(c_j | [m_i]_{i \neq j}). \quad (46.31)$$

Eq.(46.31) can be simplified using Markov Blankets (see Chapter 45) to the following:

$$P_{\underline{c}|\underline{x}}(c|m) = \prod_{j=0}^{N-1} P(c_j | [m_i : \forall i \ni \underline{x}_i \in MB(\underline{x}_j)]), \quad (46.32)$$

where, for any node \underline{a} , we denote its Markov blanket by $MB(\underline{a})$.

An alternative proposal distribution that leads to much faster convergence is as follows. The idea is to make the components $c_j^{(t)}$ of candidate sample $c^{(t)}$ depend on the previous components $(c_i^{(t)})_{i < j}$. See the bnet Fig.46.7. The TPM for the nodes of that bnet are

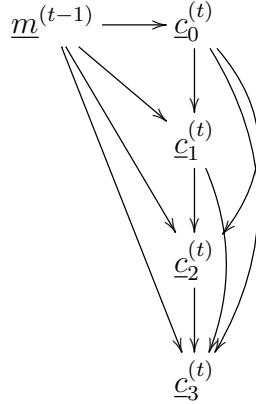


Figure 46.7: In Gibbs sampling, the proposal distribution $P_{\underline{c}|\underline{x}}$ can be defined by making the components $c_j^{(t)}$ of candidate sample $c^{(t)}$ depend on the previous components $(c_i^{(t)})_{i < j}$.

$$P(\underline{c}_j^{(t)} = c_j | (\underline{c}_i^{(t)})_{i < j} = (c_i)_{i < j}, \underline{m}^{(t-1)} = m) = P(c_j | (c_i)_{i < j}, (m_i)_{i > j}) \quad (46.33)$$

for $j = 0, 1, \dots, N - 1$. This implies

$$P_{\underline{c}|\underline{x}}(\underline{c}^{(t)} = c | \underline{m}^{(t-1)} = m) = \prod_{j=0}^{N-1} P(c_j | (c_i)_{i < j}, (m_i)_{i > j}). \quad (46.34)$$

As before, we can condition only on the Markov blanket of each node \underline{x}_j .

$$P_{\underline{c}|\underline{x}}(\underline{c}^{(t)} = c | \underline{m}^{(t-1)} = m) = \prod_{j=0}^{N-1} P(c_j | (c_i)_{i < j}, (m_i)_{i > j}, \text{ use only } c_i \text{ and } m_i \ni \underline{x}_i \in MB(\underline{x}_j)). \quad (46.35)$$

46.5 Importance Sampling

For more info about this topic and some original references, see Ref.[108].

Suppose random variables $\underline{x}[\sigma] \in S_{\underline{x}}$ for $\sigma = 0, 1, \dots, nsam - 1$ are i.i.d. with probability distribution $P_{\underline{x}}$. Then

$$E_{\underline{x}}[f(x)] \approx \frac{1}{nsam} \sum_{\sigma=0}^{nsam-1} f(x[\sigma]) \quad (46.36)$$

for any $f : S_{\underline{x}} \rightarrow \mathbb{R}$. Sometimes, instead of using i.i.d. samples $\underline{x}[\sigma] \in S_{\underline{x}}$ where $\underline{x}[\sigma] \sim P_{\underline{x}}$, we wish to use i.i.d. samples $\underline{y}[\sigma] \in S_{\underline{x}}$ where $\underline{y}[\sigma] \sim P_{\underline{y}}$.

$$E_{\underline{x}}[f(\underline{x})] = \sum_x P_{\underline{x}}(x) f(x) \quad (46.37)$$

$$= \sum_x P_{\underline{y}}(x) \frac{P_{\underline{x}}(x)}{P_{\underline{y}}(x)} f(x) \quad (46.38)$$

$$= E_{\underline{y}}\left[\frac{P_{\underline{x}}(y)}{P_{\underline{y}}(y)} f(y)\right] \quad (46.39)$$

Sampling from $P_{\underline{y}}(y)$ instead of $P_{\underline{x}}(x)$ might reduce (or increase) variance for a particular $f : S_{\underline{x}} \rightarrow \mathbb{R}$.

$$Var_{\underline{x}}[f(x)] = E_{\underline{x}}[(f(x))^2] - (E_{\underline{x}}[f(x)])^2 \quad (46.40)$$

$$Var_{\underline{y}}\left[\frac{P_{\underline{x}}(y)}{P_{\underline{y}}(y)} f(y)\right] = E_{\underline{y}}\left[\left(\frac{P_{\underline{x}}(y)}{P_{\underline{y}}(y)} f(y)\right)^2\right] - (E_{\underline{y}}\left[\frac{P_{\underline{x}}(y)}{P_{\underline{y}}(y)} f(y)\right])^2 \quad (46.41)$$

$$= E_{\underline{x}}\left[\frac{P_{\underline{x}}(x)}{P_{\underline{y}}(x)} (f(x))^2\right] - (E_{\underline{x}}[f(x)])^2 \quad (46.42)$$

Chapter 47

Markov Chains

A Markov Chain is simply a bnet with the graph structure of a chain. For example, Fig.47.1 shows a chain with $n = 4$ nodes.

$$\underline{x}_0 \longrightarrow \underline{x}_1 \longrightarrow \underline{x}_2 \longrightarrow \underline{x}_3$$

Figure 47.1: Markov chain with $n = 4$ nodes.

Because of its graph structure, the TPM of each node only depends on the state of the previous node:

$$P(x_t | (x_a)_{a \neq t}) = P(x_t | x_{t-1}), \quad (47.1)$$

where $(x_a)_{a \neq t}$ are all the nodes except x_t itself and $t = 1, 2, \dots, n - 1$.

If there exists a single TPM $P_{\underline{x}_1 | \underline{x}_0}$ such that

$$P(x_t | x_{t-1}) = P_{\underline{x}_1 | \underline{x}_0}(x_t | x_{t-1}) \quad (47.2)$$

for $t = 1, 2, \dots, n - 1$, then we say that the Markov chain is **time homogeneous**.

Claim 71 (*Data Processing Inequality (DPI)*)

Consider a Markov chain $\underline{x}_0 \rightarrow \underline{x}_1 \cdots \rightarrow \underline{x}_{n-1}$. Suppose $0 \leq a < m < b \leq n - 1$. Then

$$H(\underline{x}_a : \underline{x}_b) \leq \min[H(\underline{x}_a : \underline{x}_m), H(\underline{x}_m : \underline{x}_b)] \quad (47.3)$$

See Ref.[94] for references where the DPI is proven. This inequality confirms our intuitive expectations that the information transmitted (i.e., the mutual information(MI)) from a to b (or vice versa since MI is symmetric) is smaller or equal to the one transmitted from a to m or from m to b because a and b are “farther apart” and “some info can get lost during transmission through the mediator node m ”.

Chapter 48

Mediation Analysis

This chapter is mostly based on Refs.[47, 45] by Pearl.

To fully understand this chapter, the reader should first read Chapter 11 where do and imagine operators are defined.

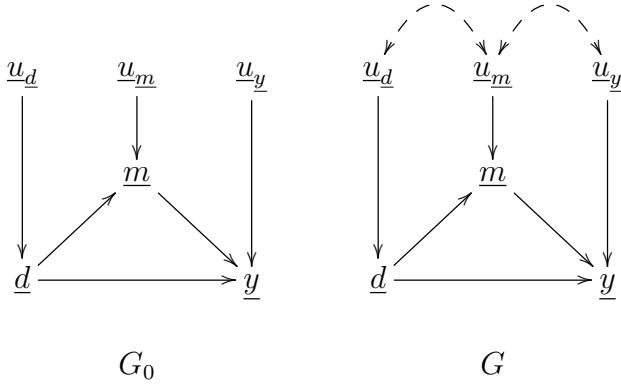


Figure 48.1: DEN bnets G_0 and G are used to discuss mediation analysis. In graph G_0 , the external variables are independent, whereas in graph G they are not.

The term “mediation analysis” (MA) refers to the analysis by Pearl of the DEN bnet G in Fig.48.1. We will discuss MA in terms of DEN bnets, just as Pearl does. However, note that much of what we will say applies also to the general (i.e., not necessarily a DEN) bnet G_{gen} show in Fig. 48.2.

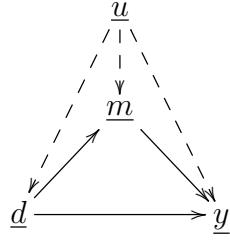
In the DEN bnet G , node \underline{d} influences node \underline{y} both directly and via the mediator node \underline{m} . The structural equations for G are of the form:

$$\underline{d} = \underline{u}_d \tag{48.1a}$$

$$\underline{m} = f_{\underline{m}}(\underline{d}, \underline{u}_m) \tag{48.1b}$$

$$\underline{y} = f_{\underline{y}}(\underline{d}, \underline{m}, \underline{u}_y). \tag{48.1c}$$

Thus,



$$G_{gen}$$

Figure 48.2: General bnet used to discuss mediation analysis. Node \underline{u} is a hidden confounder.

$$\underline{y} = f_{\underline{y}}(\underline{u}_d, f_{\underline{m}}(\underline{u}_d, \underline{u}_m), \underline{u}_y) . \quad (48.2)$$

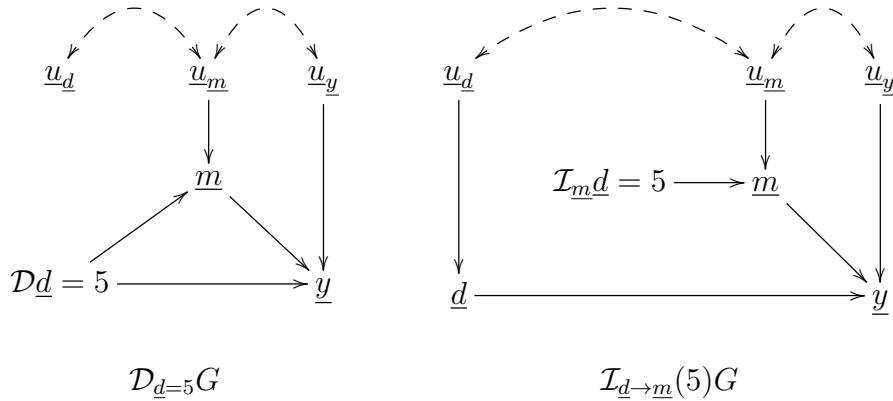


Figure 48.3: Graph G of Fig.48.1 after applying do operator $\mathcal{D}_{\underline{d}=5}$ and imagine operator $\mathcal{I}_{\underline{d} \rightarrow \underline{m}}(5)$.

If we apply $\mathcal{D}_{\underline{d}=5}G$ to Eqs.(48.1), we get

$$\underline{d} = 5 \quad (48.3a)$$

$$\underline{m} = f_{\underline{m}}(\underline{d}, \underline{u}_m) \quad (48.3b)$$

$$\underline{y} = f_{\underline{y}}(\underline{d}, \underline{m}, \underline{u}_y) . \quad (48.3c)$$

Eqs.48.3 are represented graphically in Fig.48.3. We will often denote the random variable \underline{y} in Eqs.(48.3) by the more explicit symbol $\underline{y}_{\mathcal{D}_{\underline{d}=5}G}$. Pearl often refers to this \underline{y} by the less explicit symbol Y_5 or $Y_5(u)$ where $u = (\underline{u}_m, \underline{u}_y) = u_{!d}$.

If we apply $\mathcal{I}_{\underline{d} \rightarrow \underline{m}}(5)G$ to Eqs.(48.1), we get

$$\underline{d} = \underline{u}_d \quad (48.4a)$$

$$\underline{m} = f_{\underline{m}}(5, \underline{u}_m) \quad (48.4b)$$

$$\underline{y} = f_{\underline{y}}(\underline{d}, \underline{m}, \underline{u}_y). \quad (48.4c)$$

Eqs.48.4 are represented graphically in Fig.48.3. We will often denote the random variable \underline{y} in Eqs.(48.4) by the more explicit symbol $\underline{y}_{\mathcal{I}_{\underline{d} \rightarrow \underline{m}}(5)G}$. Pearl often refers to this \underline{y} by the less explicit symbol Y_5 or $Y_5(u)$ where $u = (\underline{u}_d, \underline{u}_m, \underline{u}_y)$.

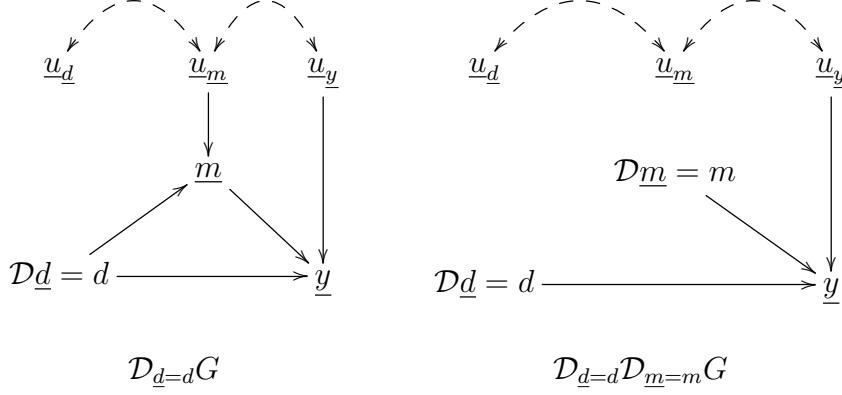


Figure 48.4: Graph G of Fig.48.1 after applying the do operators $\mathcal{D}_{\underline{d}=d}$ and $\mathcal{D}_{\underline{d}=d}\mathcal{D}_{\underline{m}=m}$.

Define

$$\mathcal{Y}_d = E[\underline{y}_{\mathcal{D}_{\underline{d}=d}G}] = \sum_y y P(y | \mathcal{D}\underline{d} = d) \quad (48.5)$$

and

$$\mathcal{Y}_d^m = E[\underline{y}_{\mathcal{D}_{\underline{d}=d}\mathcal{D}_{\underline{m}=m}G}] = \sum_y y P(y | \mathcal{D}\underline{d} = d, \mathcal{D}\underline{m} = m) \quad (48.6)$$

The two DEN diagrams $\mathcal{D}_{\underline{d}=d}G$ and $\mathcal{D}_{\underline{d}=d}\mathcal{D}_{\underline{m}=m}G$ used in the definitions of \mathcal{Y}_d and \mathcal{Y}_d^m are given in Fig.48.4.

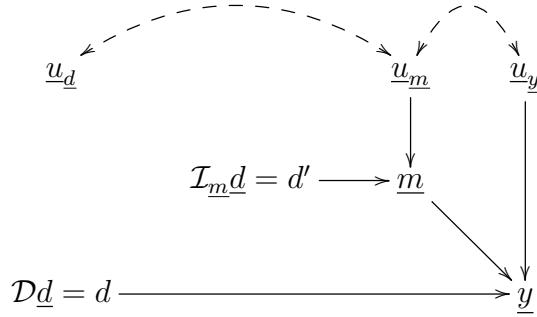
Now define the Total Effect (TE), and the Controlled Direct Effect (CDE) by

$$TE = \mathcal{Y}_1 - \mathcal{Y}_0 \quad (48.7)$$

$$CDE(m) = \mathcal{Y}_1^m - \mathcal{Y}_0^m \quad (48.8)$$

Define

$$\bar{\mathcal{Y}}_d^{d'} = E[\underline{y}_{\mathcal{D}_{\underline{d}=d}\mathcal{I}_{\underline{d} \rightarrow \underline{m}}(d')G}] = \sum_y y P(y | \mathcal{D}\underline{d} = d, \mathcal{I}_m \underline{d} = d') \quad (48.9)$$



$$\mathcal{D}_{\underline{d}=d} \mathcal{I}_{\underline{d} \rightarrow \underline{y}}(d') G$$

Figure 48.5: Graph G of Fig.48.1 after applying the imagine operator \mathcal{I} to arrow $\underline{d} \rightarrow \underline{m}$ and the do operator to node \underline{d} .

Fig.48.5 shows the diagram $\mathcal{D}_{\underline{d}=d} \mathcal{I}_{\underline{d} \rightarrow \underline{y}}(d') G$ used in the definition of $\bar{\mathcal{Y}}_d^{d'}$.

Note that

$$\bar{\mathcal{Y}}_d^{d'} = \sum_m \mathcal{Y}_d^m P(\underline{m} = m | d') \quad (48.10)$$

if there is no arrow $\underline{u}_m \rightarrow \underline{m}$. This expresses $P(y | \mathcal{D}\underline{d} = d, \mathcal{I}\underline{d} = d')$ in terms of $P(y | \mathcal{D}\underline{d} = d, \mathcal{D}\underline{m} = m)$.

Define

$$\bar{\mathcal{Y}}_d^{d'-} = \bar{\mathcal{Y}}_d^{d'} - \bar{\mathcal{Y}}_d^d \quad (48.11)$$

and

$$\bar{\mathcal{Y}}_{d-}^{d'} = \bar{\mathcal{Y}}_d^{d'} - \bar{\mathcal{Y}}_d^d \quad (48.12)$$

Now define the Natural Direct Effect (NDE), and the Natural Indirect Effect (NIE) by

$$NDE_d^{d'} = \bar{\mathcal{Y}}_d^{d'-} \quad (48.13)$$

$$NIE_d^{d'} = \bar{\mathcal{Y}}_{d-}^{d'} . \quad (48.14)$$

Note that

$$NDE_1^0 - NIE_1^0 = -\bar{\mathcal{Y}}_0^0 + \bar{\mathcal{Y}}_1^1 \quad (48.15)$$

$$= TE . \quad (48.16)$$

Linear Case

Consider the LDEN of Fig.48.6. One has

$$\underline{d} = \underline{u}_d \quad (48.17a)$$

$$\underline{m} = \alpha \underline{d} + \underline{u}_m \quad (48.17b)$$

$$\underline{y} = \gamma \underline{d} + \beta \underline{m} + \underline{u}_y . \quad (48.17c)$$

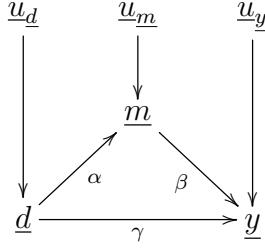


Figure 48.6: LDEN that is used to discuss mediation analysis.

$$\mathcal{Y}_d = (\gamma + \alpha\beta)d \quad (48.18)$$

$$\mathcal{Y}_d^m = \gamma d + \beta m \quad (48.19)$$

$$\bar{\mathcal{Y}}_d^{d'} = \gamma d + \alpha\beta d' \quad (48.20)$$

$$TE = \mathcal{Y}_1 - \mathcal{Y}_0 = \gamma + \alpha\beta \quad (48.21)$$

$$CDE(m) = \mathcal{Y}_1^m - \mathcal{Y}_0^m = \gamma \quad (48.22)$$

$$NDE_1^0 = \bar{\mathcal{Y}}_1^{0-} = \gamma \quad (48.23)$$

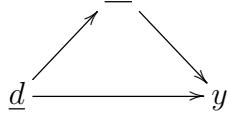
$$NIE_1^0 = \bar{\mathcal{Y}}_{1-}^0 = -\alpha\beta \quad (48.24)$$

As expected, $NDE_1^0 - NIE_1^0 = TE$.

TE and the “controlled effect” $CDE(m)$ contain do operators but no imagine operators so one can do do-intervention experiments to calculate them. On the other hand, the “natural effects” NDE_1^0 and NIE_1^0 contain imagine operators (and therefore counterfactual distributions) so it’s not obvious how to calculate them, or even whether it is possible to do so. The next claim shows how to calculate $\bar{\mathcal{Y}}_d^{d'}$ in certain cases. In technical jargon, the claim gives sufficient conditions for \mathcal{DI} -identifiability of $\bar{\mathcal{Y}}_d^{d'}$. NDE_1^0 and NIE_1^0 can be calculated once $\bar{\mathcal{Y}}_d^{d'}$ is known.

Claim 72 (*Unconfounded Mediation, from Ref.[47]*)

If $\begin{array}{c} m \\ \nearrow \quad \searrow \\ d & \xrightarrow{\hspace{1cm}} & y \end{array}$ then



$$P(y|\mathcal{D}\underline{d} = d, \mathcal{I}\underline{m}d = d') = \sum_m P(y|d, m)P(m|d') \quad (48.25)$$

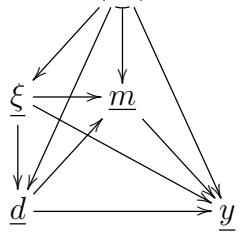
$$\begin{array}{ccc} \mathcal{I}\underline{d} = d' & & \mathcal{I}\underline{d} = d' \longrightarrow \sum m \\ \searrow & & \swarrow \\ \mathcal{D}\underline{d} = d \longrightarrow y & = & \mathcal{D}\underline{d} = d \longrightarrow y \end{array} \quad (48.26)$$

proof: See Claim 45.

QED

Claim 73 (*Mediation with universal prior ξ and universal confounder \underline{u} , from Ref.[47]*)

If $\begin{array}{c} (\underline{u}) \\ \downarrow \\ \xi \end{array}$ then



$$P(y|\mathcal{D}\underline{d} = d, \mathcal{I}_m\underline{d} = d') = \sum_{\xi} \sum_m P(y|d, m, \xi) P(m|d', \xi) P(\xi) \quad (48.27)$$

$$\begin{array}{ccc} \mathcal{I}\underline{d} = d' & & \mathcal{I}\underline{d} = d' \xrightarrow{\sum \xi} \sum m \\ \searrow & & \swarrow \\ \mathcal{D}\underline{d} = d \longrightarrow y & = & \mathcal{D}\underline{d} = d \longrightarrow y \end{array} \quad (48.28)$$

proof: See Claim 46.

QED

Actually,

$$\sum_m P(y|d, m) P(m|d') = \sum_{\xi} \sum_m P(y|d, m, \xi) P(m|d', \xi) P(\xi) \quad (48.29)$$

so the adjustment formulae in Claims 72 and 73 are equivalent, but if the available dataset contains info about ξ , then the adjustment formula that uses that ξ info should be used, as it will be more sensitive to deviations from the DAG model being hypothesized.

Chapter 49

Message Passing and Bethe Free Energy

This chapter is based on Refs. [72] and [153].

49.1 2MRFs

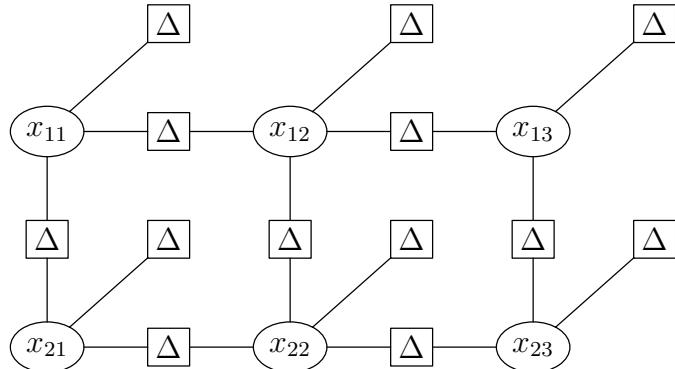


Figure 49.1: Example of factor graph for a 2MRF. In this figure, a boxed Δ between variables x_i and x_j , denotes the propagator function $\Delta(x_i, x_j)$. Also, a boxed Δ connected to a single variable x_i , denotes the function $\Delta(x_i, k)$, where k is a fixed number. If we view this factor graph as a representation of a 3 dimensional device with the leaf nodes Δ in the back part and everything else in the front part, then the back part can represent the light inputs from a scene, and the front part can represent the analysis being done with the light inputs.

Factor graphs are discussed in Chapter 26.

A pairwise Markov Random Field (2MRF) is a statistical model whose probability distribution is of the form

$$p(x) = \mathcal{N}(!x) \prod_{i-j} \Delta(x_i, x_j) \quad (49.1)$$

where $x = (x_1, x_2, \dots, x_n)$, where $i - j$ represents the edge that connects nodes i and j in an undirected graph G , and where the product is over all edges of G . $p(x)$ can be represented graphically by a factor graph.

Fig.49.1 shows an example of a 2MRF represented as a factor graph.

The **Ising model** is another prototypical example of a 2MRF. Let T be the temperature of a system,

$$\theta = \frac{1}{T}, \quad (49.2)$$

and $x = (x_1, x_2, \dots, x_n) \in \{-1, 1\}^{nx}$. The Ising model is defined for $J, h(x_i) \in \mathbb{R}$ by

$$\epsilon(x) = -J \sum_{i-j} x_i x_j - \sum_i h(x_i) \quad (49.3)$$

$$P(x) = \frac{e^{\frac{-\epsilon(x)}{T}}}{Z_\theta} \quad (49.4)$$

Note that a bnet can be easily converted to an equivalent 2MRF. You just set $\Delta(x_i, k) = p(x_i)$ if x_i is a root node of the bnet, and $\Delta(x_i, x_j) = p(x_i|x_j)$ if x_i isn't a root node. One must then choose the arrow directions of the bnet. A natural choice is to choose the bnet arrows so that they all point from an x_i to a Δ .

In what follows, we shall use **messages (a.k.a. cavity fields)** $m_{a \Rightarrow i}(x_i)$, Note that in our notation for messages, the letter i appears twice, and both occurrences are next to each other. This is always the case in our notation for messages.

Henceforth, we will refer to the functions $\Delta(x_i, x_j)$ as **propagators**, to make contact with Physics.

Henceforth, we shall use the notation ∂i to denote all nodes j that are neighbors of node i . ∂i is called the **neighborhood or boundary** of i .

49.2 Message Passing Intuition

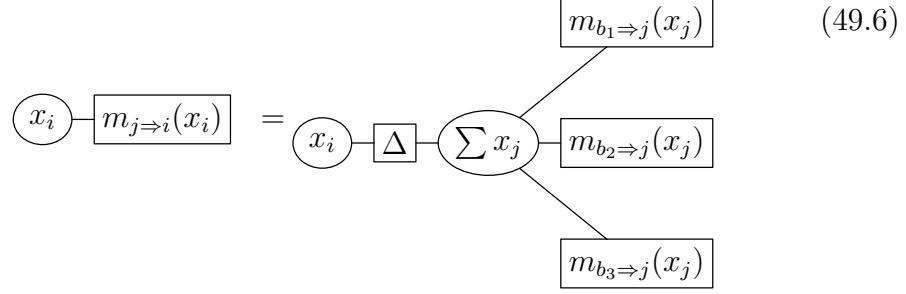
Next we present some results that will be derived later in the chapter, but which we will first present now without proof, and explain how they make sense intuitively. Assume that the functions $\Delta(x_i, x_j)$ are known for all edges $i - j$, and that the full probability distribution $p(x)$ of the model is given by Eq.(49.1).

1. Calculate message $m_{j \Rightarrow i}(x_i)$ in terms of messages leaving x_j . This is called the **message updating rule**.

One has that

$$m_{j \Rightarrow i}(x_i) = \sum_{x_j} \Delta(x_i, x_j) \prod_{b \in \partial j \setminus i} m_{b \Rightarrow j}(x_j) \quad (49.5)$$

This can be represented graphically by

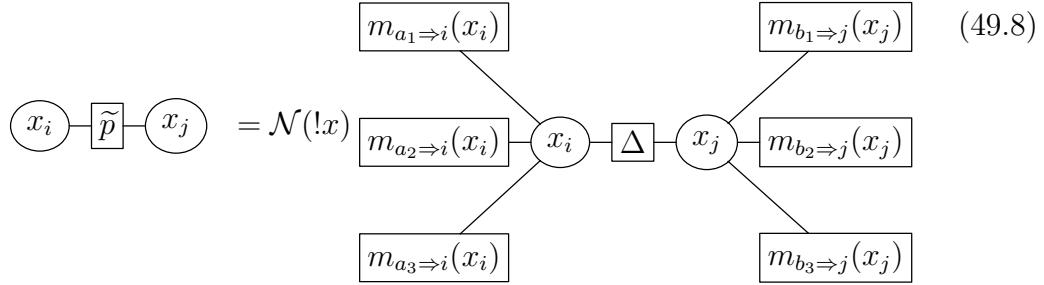


2. Calculate $\tilde{p}(x_i, x_j)$ for any edge $i - j$ in terms of messages entering x_i and leaving x_j

One has that

$$\tilde{p}(x_i, x_j) = \mathcal{N}(!x) \Delta(x_i, x_j) \prod_{a \in \partial i \setminus j} m_{a \Rightarrow i}(x_i) \prod_{b \in \partial j \setminus i} m_{b \Rightarrow j}(x_j) \quad (49.7)$$

This can be represented graphically by

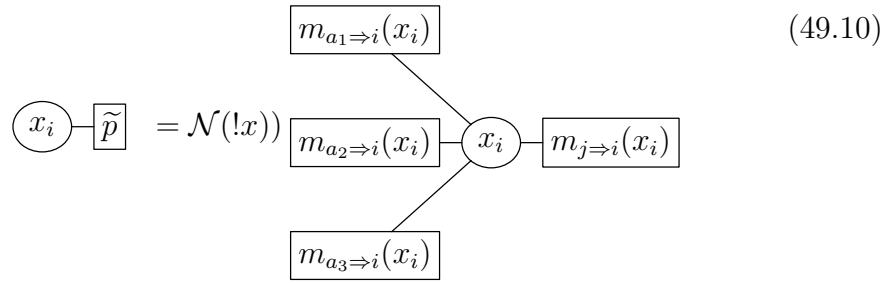


3. Calculate $\tilde{p}(x_i)$ in terms of messages entering x_i

One has that

$$\tilde{p}(x_i) = \sum_{x_j} \tilde{p}(x_i, x_j) \quad (49.9)$$

This can be represented graphically by



Why it works?

Suppose we approximate the propagator $\Delta(x_i, x_j)$ by

$$\Delta(x_i, x_j) \approx m_{j \Rightarrow i}(x_i)m_{i \Rightarrow j}(x_j) \quad (49.11)$$

This can be represented graphically by

$$(x_i) \rightarrow \Delta \rightarrow (x_j) \approx (x_i) \rightarrow [m_{j \Rightarrow i}(x_i)] \rightarrow [m_{i \Rightarrow j}(x_j)] \rightarrow (x_j) \quad (49.12)$$

Then the results 1,2,3 above can be justified as follows.

1.

$$m_{j \Rightarrow i}(x_i) \approx \sum_{x_j} \underbrace{m_{j \Rightarrow i}(x_i)m_{i \Rightarrow j}(x_j)}_{\Delta(x_i, x_j)} \prod_{b \in \partial j \setminus i} m_{b \Rightarrow j}(x_j) \quad (49.13)$$

$$= m_{j \Rightarrow i}(x_i) \sum_{x_j} \prod_{b \in \partial j} m_{b \Rightarrow j}(x_j) \quad (49.14)$$

$$= \mathcal{N}(!x)m_{j \Rightarrow i}(x_i) \quad (49.15)$$

2.

$$\tilde{p}(x_i, x_j) = \sum_{x \setminus x_i, x_j} \tilde{p}(x) \quad (49.16)$$

$$\approx \mathcal{N}(!x) \sum_{x \setminus x_i, x_j} \prod_{i=j} \underbrace{m_{j \Rightarrow i}(x_i)m_{i \Rightarrow j}(x_j)}_{\Delta(x_i, x_j)} \quad (49.17)$$

$$\approx \mathcal{N}(!x)\Delta(x_i, x_j) \prod_{a \in \partial i \setminus j} m_{a \Rightarrow i}(x_i) \prod_{b \in \partial j \setminus i} m_{b \Rightarrow j}(x_j) \quad (49.18)$$

3.

$$\tilde{p}(x_i) = \sum_{x_j} \tilde{p}(x_i, x_j) \quad (49.19)$$

$$\approx \mathcal{N}(!x) \sum_{x_j} \underbrace{m_{j \Rightarrow i}(x_i)m_{i \Rightarrow j}(x_j)}_{\Delta(x_i, x_j)} \prod_{a \in \partial i \setminus j} m_{a \Rightarrow i}(x_i) \prod_{b \in \partial j \setminus i} m_{b \Rightarrow j}(x_j) \quad (49.20)$$

$$\approx \mathcal{N}(!x) \prod_{a \in \partial i} m_{a \Rightarrow i}(x_i) \quad (49.21)$$

Henceforth, we will occasionally use the following more compressed notation for the factor graph of a 2MRF. Instead of using circles for variables and squares for functions, we will use arrows pointing from x_1 to x_2 and vice versa.

$$(x_1) \rightarrow [m_{2 \Rightarrow 1}(x_1)] \rightarrow [m_{1 \Rightarrow 2}(x_2)] \rightarrow (x_2) = x_1 \xrightarrow[m_{2 \Rightarrow 1}(x_1)]{m_{1 \Rightarrow 2}(x_2)} x_2 \quad (49.22)$$

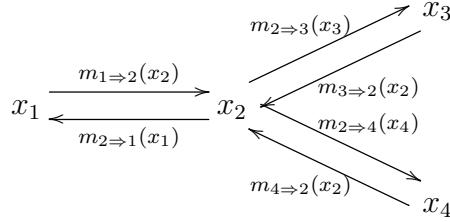


Figure 49.2: Example of compressed factor graph for a 2MRF.

Next, we give an example of how to use message passing to find the marginal $\tilde{p}(x_i)$ for any node x_i . Refer to Fig49.2 for this example.

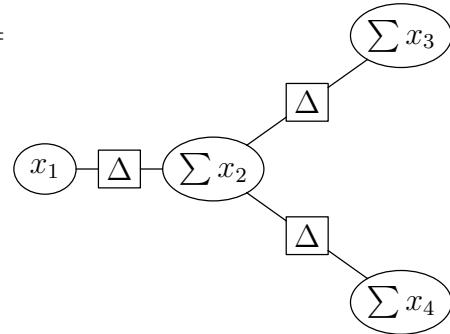
Using the message updating rule successively, we get

$$\tilde{p}(x_1) = \mathcal{N}(!x)m_{2\Rightarrow 1}(x_1) \quad (49.23)$$

$$= \mathcal{N}(!x) \sum_{x_2} \Delta(x_1, x_2) m_{3\Rightarrow 2}(x_2) m_{4\Rightarrow 2}(x_2) \quad (49.24)$$

$$= \mathcal{N}(!x) \sum_{x_2} \Delta(x_1, x_2) \sum_{x_3} \Delta(x_2, x_3) \sum_{x_4} \Delta(x_2, x_4) \quad (49.25)$$

$$= \quad (49.26)$$



This algorithm is guaranteed to work only for trees. In practice, one starts by calculating the messages pointing up from the leaf nodes of the tree. Then one calculates the messages pointing up from the parents of the leaf nodes of the tree. And so forth until one calculates all the messages pointing up from the leaf nodes to the root node of the tree. Then one goes in the opposite direction, first calculating messages pointing down from the root node to its children, from the children of the root node to their children. And so forth. By the end, all upward and downward pointing messages have been calculated. From this, $\tilde{p}(x_i)$ for all i can be calculated. $\tilde{p}(x_i)$ is the product of all messages pointing into x_i .

49.3 $-\ln Z_\theta$ = Free Energy (FE)

Suppose $\theta, \epsilon(x) \in \mathbb{R}^{ni}$ and $x = (x_1, x_2, \dots, x_{nx}) \in \mathbb{R}^{nx}$. Define the **partition function** Z_θ by

$$Z_\theta = \sum_x e^{-\theta^T \epsilon(x)} \quad (49.27)$$

and the probability distribution $P(x|h)$ by

$$\underbrace{P(x|\theta)}_{e^{LL_\theta(x)}} = \exp(-\theta^T \epsilon(x) - \ln Z_\theta) \quad (49.28)$$

$$= \frac{e^{-\theta^T \epsilon(x)}}{Z} \quad (49.29)$$

$\epsilon(x)$ is called a **sufficient statistic** for \underline{x} because $P(x|\theta)$ is a functional (i.e., a function of a function) of $\epsilon(x)$. Note that by taking first and higher order derivatives of $\ln Z_\theta$ with respect to θ_i , we can calculate the statistics of $\epsilon_i(x)$.

$$-\partial_{\theta_i} \ln Z_\theta = \frac{1}{Z_\theta} \sum_x \epsilon_i(x) e^{-\theta^T \epsilon(x)} \quad (49.30)$$

$$= E_{\underline{x}|\theta}[\epsilon_i(\underline{x})] = \langle \epsilon_i \rangle \quad (49.31)$$

$$\partial_{\theta_j} \partial_{\theta_i} \ln Z_\theta = \partial_{\theta_j} \frac{1}{Z_\theta} \sum_x -\epsilon_i(x) e^{-\theta^T \epsilon(x)} \quad (49.32)$$

$$= \left\{ \begin{array}{l} \frac{1}{Z_\theta} \sum_x \epsilon_j(x) \epsilon_i(x) e^{-\theta^T \epsilon(x)} \\ + \frac{-1}{Z_\theta^2} \left[\sum_x -\epsilon_j(x) e^{-\theta^T \epsilon(x)} \right] \left[\sum_x -\epsilon_i(x) e^{-\theta^T \epsilon(x)} \right] \end{array} \right\} \quad (49.33)$$

$$= \langle \epsilon_j \epsilon_i \rangle - \langle \epsilon_j \rangle \langle \epsilon_i \rangle \quad (49.34)$$

$$= \langle \epsilon_j, \epsilon_i \rangle \quad (49.35)$$

Define the log likelihood $LL_\theta(x)$ by

$$LL_\theta(x) = -\theta^T \epsilon(x) - \ln Z_\theta \quad (49.36)$$

and the **entropy** S by ¹

$$S = - \sum_x P(x|\theta) LL_\theta(x) \quad (49.37)$$

$$= - \sum_x P(x|\theta) \ln P(x|\theta) \quad (49.38)$$

¹In Thermodynamics, the entropy is denoted by the letter S . In Shannon Information Theory, and elsewhere in this book, it is denoted by the letter H .

Define the **internal energy** U by

$$U = \sum_x P(x|\theta) \epsilon_\theta(x) \quad (49.39)$$

$$-S = -\theta^T U - \ln Z_\theta \quad (49.40)$$

$$\partial_{U_i} S = \theta_i \quad (49.41)$$

S is concave. S and $-\ln Z_\theta = F/T$ are concave dual functions.²

Relationship to Thermodynamics

In Thermodynamics, U is the internal energy and S is the entropy of a system at **temperature** T . Define $\theta \in \mathbb{R}$ to be the inverse temperature

$$\theta = \frac{1}{T} \quad (49.42)$$

Define the **free energy** F by^a

$$F = -T \ln Z_\theta \quad (49.43)$$

$$= -T \ln \sum_x e^{-\frac{\epsilon(x)}{T}} \quad (49.44)$$

Then

$$U - TS = F \quad (49.45)$$

So the free energy equals the internal energy minus the energy held in disordered form.

^aIn this chapter, we call $-\ln Z$ the free energy too.

49.4 $-\ln Z_{\theta^*}$ = Minimum FE

Suppose we consider $P(x|\theta)$ for two different parameter θ and $\tilde{\theta}$. Define

$$p(x) = P(x|\theta), \quad \tilde{p}(x) = P(x|\tilde{\theta}) \quad (49.46)$$

The Kullback-Leibler divergence is always non-negative so:

²Concave dual functions are discussed in Chapter 89

$$0 \leq D_{KL}(\tilde{p}(x) \| p(x)) \quad (49.47)$$

$$= \sum_x \tilde{p}(x) \ln \frac{\tilde{p}(x)}{p(x)} \quad (49.48)$$

$$= -\tilde{S} - \sum_x \tilde{p}(x) [-(\theta)^T \epsilon(x) - \ln Z_\theta] \quad (49.49)$$

$$= -\tilde{S} + (\theta)^T \tilde{U} + \ln Z_\theta \quad (\tilde{S}, \tilde{U} \text{ correspond to parameter } \tilde{\theta}) \quad (49.50)$$

$$-\ln Z_\theta \leq -\tilde{S} + (\theta)^T \tilde{U} \quad (49.51)$$

Let

$$\theta^* = \operatorname{argmin}_\theta \left[-\tilde{S} + (\theta)^T \tilde{U} \right] \quad (49.52)$$

Henceforth, we will refer to $-\ln Z_\theta$ as the **FE (Free Energy)** and to $-\ln Z_{\theta^*}$ as the **minimum FE**.

Relationship to convex/concave dual functions³

$$\tilde{S} = \min_\theta \left[(\theta)^T \tilde{U} + \ln Z_\theta \right] \quad (49.53)$$

$$-\ln Z_\theta = \min_{\tilde{U}} \left[(\theta)^T \tilde{U} - \tilde{S} \right] \quad (49.54)$$

\tilde{S} and $-\ln Z_\theta$ are concave dual functions.

$$-\ln Z_{\theta^*} = (\theta^*)^T \tilde{U} - \tilde{S} \quad (49.55)$$

49.5 $-\ln Z_\theta^{tree}$ =Tree FE (a.k.a. Bethe FE)

$\tilde{p}(x)$ is said to satisfy **Mean Field Approximation (MFA)** or **independent variables approximation (IVA)** if its variables x_i are independently distributed:

$$\tilde{p}^{ind}(x) = \prod_k \tilde{p}(x_k) \quad (49.56)$$

In the MFA, the entropy is

³Convex/concave dual functions are discussed in Chapter 89

$$\tilde{S}^{ind} = - \sum_x \tilde{p}(x) \ln \prod_i \tilde{p}(x_i) \quad (49.57)$$

$$= - \sum_{x_i} \tilde{p}(x_i) \ln \tilde{p}(x_i) \quad (49.58)$$

$$= \sum_i \tilde{H}(\underline{x}_i) \quad (49.59)$$

A slightly more complicated case than the MFA is when $\tilde{p}(x)$ is defined on a tree graph G^{tree} with edges $i - j$. In such a case, it will take just a few examples of trees G^{tree} to convince the reader that in general, for any tree G^{tree} , $\tilde{p}(x)$ must have the following form:

$$\tilde{p}^{tree}(x) = \tilde{p}^{ind}(x) \prod_{i-j} \frac{\tilde{p}(x_i, x_j)}{\tilde{p}(x_i)\tilde{p}(x_j)} \quad (49.60)$$

Hence,

$$\tilde{S}^{tree} = - \sum_i \sum_{x_i} \tilde{p}^{tree}(x_i) \ln \tilde{p}^{tree}(x_i) \quad (49.61)$$

$$= \sum_i \tilde{H}(\underline{x}_i) - \sum_{i-j} \sum_{x_i, x_j} \tilde{p}(x_i, x_j) \ln \frac{\tilde{p}(x_i, x_j)}{\tilde{p}(x_i)\tilde{p}(x_j)} \quad (49.62)$$

$$= \sum_i \tilde{H}(\underline{x}_i) - \sum_{i-j} \tilde{H}(\underline{x}_i : \underline{x}_j) \quad (49.63)$$

Note that \tilde{S}^{tree} can be written in terms of the joint entropy $\tilde{H}(\underline{x}_i, \underline{x}_j)$ instead of the mutual entropy $\tilde{H}(\underline{x}_i : \underline{x}_j)$.

$$\tilde{S}^{tree} = - \sum_i (d_i - 1) \tilde{H}(\underline{x}_i) + \sum_{i-j} \tilde{H}(\underline{x}_i, \underline{x}_j) \quad (49.64)$$

where d_i is the number of neighbors of node i .

The following approximation is often called the **Bethe approximation**

$$-\ln Z_{\theta^*} \approx -\ln Z_{\theta^*}^{tree} \quad (49.65)$$

49.6 $-\ln Z_{\theta^*}^{tree}$ = Tree Minimum FE, and message passing

In this section, we will evaluate $-\ln Z_{\theta^*}^{tree}$ exactly using a message passing ansatz (an ansatz is an initial guess).

If we replace \tilde{S} by \tilde{S}^{tree} in Eq.(49.55), we get

$$-\ln Z_{\theta^*}^{tree} = \min_{\tilde{U}} \left[(\theta)^T \tilde{U} - \tilde{S}^{tree} \right] \quad (49.66)$$

But note that

$$(\theta)^T \tilde{U} = \sum_i \theta_i \sum_x \tilde{p}(x) \epsilon_i(x) \quad (49.67)$$

$$= \sum_x \tilde{p}(x) \underbrace{\sum_i \theta_i \epsilon_i(x)}_{\text{call } \Theta(x)} \quad (49.68)$$

so Eq.(49.66) becomes

$$-\ln Z_{\theta^*}^{tree} = \min_{\tilde{p}} \left[\sum_x \tilde{p}(x) \Theta(x) - \sum_i \tilde{H}(\underline{x}_i) + \sum_{i-j} \tilde{H}(\underline{x}_i : \underline{x}_j) \right] \quad (49.69)$$

subject to $\sum_x \tilde{p}(x) = 1$ and $\tilde{p}(x) \geq 0$ for all x .

Claim 74 $-\ln Z_{\theta^*}^{tree}$ is achieved if

$$\tilde{p}(x) = \mathcal{N}(!x) e^{-\Theta(x)} \quad (49.70)$$

$$\Theta(x) = \sum_i \Theta(x_i) + \sum_{i-j} \Theta(x_i, x_j) \quad (49.71)$$

(This form for $\tilde{p}(x)$ and $\Theta(x)$ agrees with Eq.(49.60))

$$m_{t \Rightarrow s}(x_s) = e^{\lambda_{t \Rightarrow s}(x_s)} \quad (49.72)$$

$$\tilde{p}(x_i) = \mathcal{N}(!x) e^{-\Theta(x_i)} \prod_{a \in \partial i} m_{a \Rightarrow i}(x_i) \quad (49.73)$$

$$\tilde{p}(x_i, x_j) = \mathcal{N}(!x) e^{-\Theta(x_i, x_j) - \Theta(x_i) - \Theta(x_j)} \left[\prod_{a \in \partial i \setminus j} m_{a \Rightarrow i}(x_i) \right] \left[\prod_{b \in \partial j \setminus i} m_{b \Rightarrow j}(x_j) \right] \quad (49.74)$$

proof:

We want to minimize the following Lagrangian with respect to variations $\delta \tilde{p}(x_i)$ of $\tilde{p}(x)$, for each i .

$$\mathcal{L} = \left\{ \begin{array}{l} \sum_x \tilde{p}(x) \Theta(x) \\ + \sum_i (1 - d_i) \sum_{x_i} \tilde{p}(x_i) \ln \tilde{p}(x_i) \\ + \sum_{i-j} \sum_{x_i, x_j} \tilde{p}(x_i, x_j) \ln \tilde{p}(x_i, x_j) \\ + \lambda [\sum_x \tilde{p}(x) - 1] \end{array} \right\} \quad (49.75)$$

The term proportional to the Lagrange multiplier λ enforces the constraint $\sum_x \tilde{p}(x) = 1$. Note that in general,

$$\delta \tilde{p}(x) = \sum_i \delta \tilde{p}(x_i) \quad (49.76)$$

so if we only vary $\tilde{p}(x_i)$,

$$\delta F(\tilde{p}(x)) = \frac{\partial F(\tilde{p}(x))}{\partial \tilde{p}(x)} \delta \tilde{p}(x) \quad (49.77)$$

$$= \frac{\partial F(\tilde{p}(x))}{\partial \tilde{p}(x)} \delta \tilde{p}(x_i) \quad (49.78)$$

for any well behaved function $F : \mathbb{R} \rightarrow \mathbb{R}$. Hence

$$\delta \mathcal{L} = \sum_x \delta \tilde{p}(x_i) \left\{ \begin{array}{l} \Theta(x) \\ + \sum_i (1 - d_i) [1 + \ln \tilde{p}(x_i)] \\ + \sum_{i-j} [1 + \ln \tilde{p}(x_i, x_j)] \\ + \lambda \end{array} \right\} \quad (49.79)$$

for any variation $\delta \tilde{p}(x_i)$. Setting the coefficient of $\delta \tilde{p}(x_i)$ to zero now yields

$$0 = \left\{ \begin{array}{l} \Theta(x) \\ + \sum_i [1 + \ln \tilde{p}(x_i)] \\ + \sum_{i-j} \left[1 + \ln \frac{\tilde{p}(x_i, x_j)}{\tilde{p}(x_i) \tilde{p}(x_j)} \right] \\ + \lambda \end{array} \right\} \quad (49.80)$$

for each i . If we now substitute the equations that are hypotheses to this claim, we get

$$0 = \left\{ \begin{array}{l} \Theta(x) \\ - \sum_i \Theta(x_i) \\ - \sum_{i-j} \Theta(x_i, x_j) \end{array} \right\} \quad (49.81)$$

$$0 = \left\{ \begin{array}{l} + \sum_i \ln \prod_{a \in \partial i} m_{a \Rightarrow i}(x_i) \\ + \sum_{i-j} \ln \frac{[\prod_{a \in \partial i \setminus j} m_{a \Rightarrow i}(x_i)] [\prod_{b \in \partial j \setminus i} m_{b \Rightarrow j}(x_j)]}{\prod_{a \in \partial i} m_{a \Rightarrow i}(x_i) \prod_{b \in \partial j} m_{b \Rightarrow j}(x_j)} \\ + \lambda \end{array} \right\} \quad (49.82)$$

$$= \left\{ \begin{array}{l} \sum_i \sum_{a \in \partial i} \lambda_{a \Rightarrow i}(x_i) \\ - \sum_{i-j} [\lambda_{j \Rightarrow i}(x_i) + \lambda_{i \Rightarrow j}(x_j)] \\ + \lambda \end{array} \right\} \quad (49.83)$$

$$= \lambda \quad (49.84)$$

So the hypotheses to this claim indeed do satisfy $\delta \mathcal{L} = 0$ for all variations $\delta \tilde{p}(x_i)$, for each i , with Lagrange multiplier $\lambda = 0$.

QED

Chapter 50

Message Passing, Pearl's theory

This chapter is mostly based on chapter 4 of Ref.[44] by Pearl. Refs.[79], and [37] were also helpful in writing this chapter.

In his book Ref.[44], Pearl explains two types of Message Passing (i.e., distributed computing in a bnet). In Chapter 4, he discusses one type of MP which he calls Belief Propagation (BP) or Belief Updating. In Chapter 5, he introduces a second type of MP which he calls Belief Revision, but which I prefer to call Explanation Optimization (EO). This chapter will be devoted to BP only.

BP was first proposed for bnets in 1982 Ref.[43] by Judea Pearl to simplify the exact evaluation of the probability of one node conditioned on other nodes of a bnet (exact inference). It gives exact results for trees and polytrees (i.e., bnets with a single connected component and no loops). For bnets with loops, it gives approximate results (loopy belief propagation), and it has been generalized to the junction tree algorithm (see Chapter 40) which can do exact inference for general bnets with loops. The basic idea behind the junction tree algorithm is to eliminate loops by clustering them into single nodes.

50.1 Distributed Soldier Counting

Consider a group of soldiers marching single file. Fig.50.1 shows several methods by which a member of the group can obtain a count of the soldiers without breaking the line to do global operations. This can be done in a distributed fashion, with every soldier doing only local operations (i.e., each soldier can only send messages to either the soldier in front or the one in back). Such distributed soldier counting is a rudimentary type of BP. In the next section, we will generalize this BP for soldiers to BP for a Markov chain.

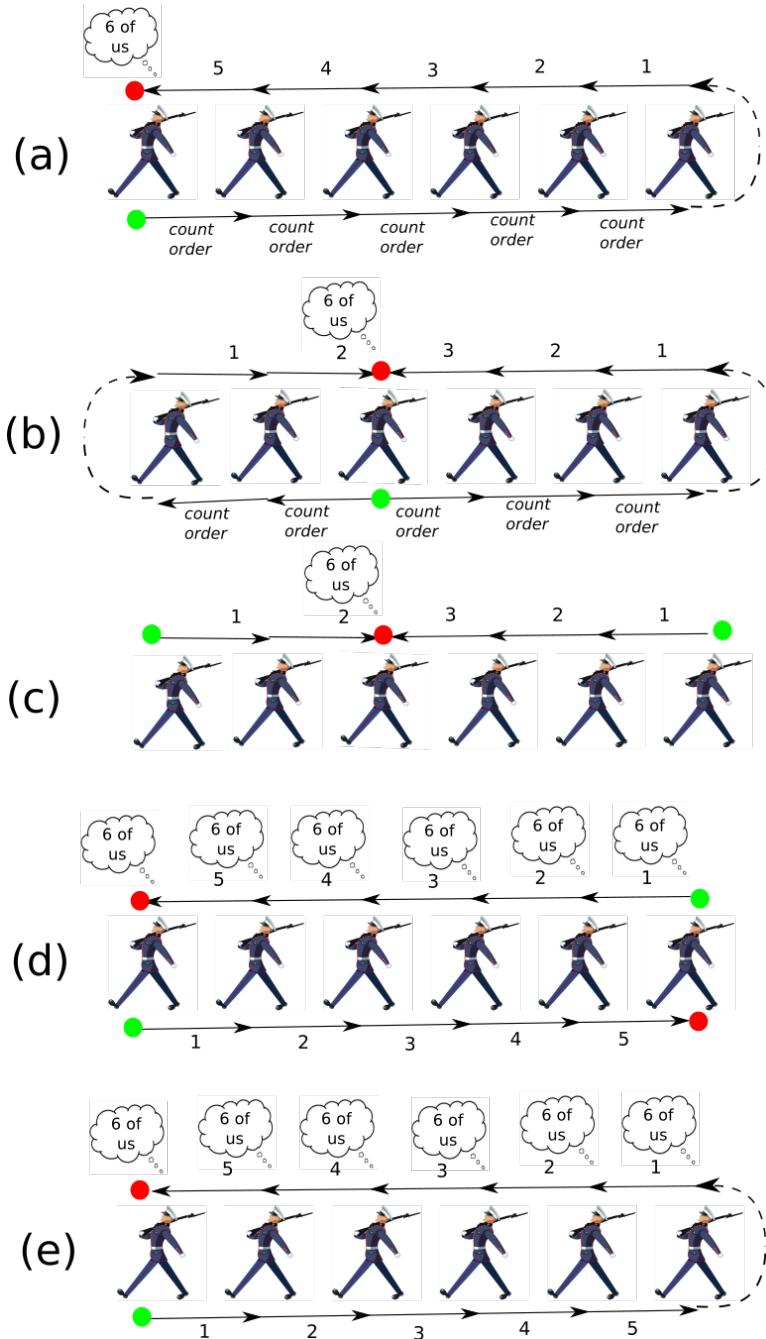


Figure 50.1: Distributed soldier counting (This example comes from Chapter 4 of Ref.[44]). Green dots indicate the beginning and red dots the end of a count. Only first soldier can calculate total count in (a). Only third soldier can calculate total count in (b,c). All soldiers can calculate the total count in (d,e). One starting point in (a,b,e). Two ends as starting points in (c,d).

50.2 Spring Systems

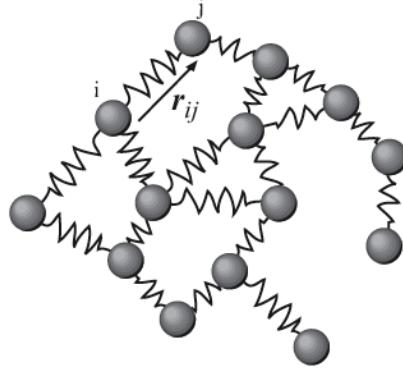


Figure 50.2: Spring system. Point masses connected by springs.

See Ref.[139] for an introduction to spring systems. Ideal springs between the point mass nodes would not be sufficient. One would have to add damping to the springs so as to reach an equilibrium. Time dependent forces (loads) pointing into or out of the page, applied to the point masses, would generate signals that would propagate like BP messages.

50.3 BP for Markov Chains

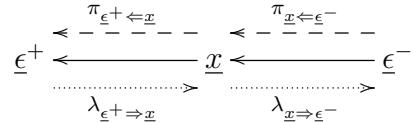


Figure 50.3: 3 node Markov chain $\underline{\epsilon}^+ \leftarrow \underline{x} \leftarrow \underline{\epsilon}^-$. The π messages (probability functions) travel downstream (i.e., they carry info in the direction of the graph arrows, towards the future) and are indicated by a dashed arrow or by a left double arrow \Leftarrow . The λ messages (likelihood functions) travel upstream (i.e., they carry info opposite to direction of the graph arrows, towards the past) and are indicated by a dotted arrow or by a right double arrow \Rightarrow . $\underline{\epsilon}^+$ stands for future evidence and $\underline{\epsilon}^-$ for past evidence.

Consider the 3 node Markov chain $\underline{\epsilon}^+ \leftarrow \underline{x} \leftarrow \underline{\epsilon}^-$ shown in Fig.50.3. Define¹

¹The pattern behind these definitions, in case it eludes you, is as follows: the π 's always carry information about the past and the λ 's about the future. But the past or future of what? Of the

$$\pi_{\underline{\epsilon}^+ \leftarrow \underline{x}}(x) = P(x|\epsilon^-) \text{ (past of } \underline{x}) \quad (50.1)$$

$$\lambda_{\underline{\epsilon}^+ \Rightarrow \underline{x}}(x) = P(\epsilon^+|x) \text{ (future of } \underline{x}) \quad (50.2)$$

$$\pi_{\underline{x} \leftarrow \underline{\epsilon}^-}(\epsilon^-) = P(\epsilon^-) = \delta(\epsilon^-, \epsilon_0^-) \text{ (past of } \underline{\epsilon}^-) \quad (50.3)$$

$$\lambda_{\underline{x} \Rightarrow \underline{\epsilon}^-}(\epsilon^-) = P(\epsilon^+|\epsilon^-) \text{ (future of } \underline{\epsilon}^-) . \quad (50.4)$$

Furthermore, define the Belief BEL in x to be

$$BEL_{\underline{x}}(x) = P(x|\epsilon) , \quad (50.5)$$

where

$$\epsilon = \underline{\epsilon}^+ \cup \underline{\epsilon}^- . \quad (50.6)$$

It follows that

$$BEL_{\underline{x}}(x) = P(x|\epsilon^+, \epsilon^-) = \quad (50.7)$$

$$= \mathcal{N}(!x)P(\epsilon^+, x, \epsilon^-) \quad (50.8)$$

$$= \mathcal{N}(!x)P(\epsilon^+|x)P(x|\epsilon^-) \quad (50.9)$$

$$= \mathcal{N}(!x)\lambda_{\underline{\epsilon}^+ \Rightarrow \underline{x}}(x)\pi_{\underline{\epsilon}^+ \leftarrow \underline{x}}(x) . \quad (50.10)$$

Note that Bayes rule would affirm that²

$$P(x|\epsilon^+) = \mathcal{N}(!x) \underbrace{P(\epsilon^+|x)}_{\lambda_{\underline{\epsilon}^+ \Rightarrow \underline{x}}(x)} P(x) . \quad (50.11)$$

Thus, Eq.(50.10) is like a 2-sided Janus Bayes rule.

Note that the π messages and λ messages propagate independently of each other, via the TPM $P(x|\epsilon^-)$:

$$\underbrace{\pi_{\underline{\epsilon}^+ \leftarrow \underline{x}}(x)}_{P(x|\epsilon_0^-)} = \sum_{\epsilon^-} P(x|\epsilon^-) \underbrace{\pi_{\underline{x} \leftarrow \underline{\epsilon}^-}(\epsilon^-)}_{\delta(\epsilon^-, \epsilon_0^-)} \quad (50.12a)$$

$$\underbrace{\lambda_{\underline{x} \Rightarrow \underline{\epsilon}^-}(\epsilon^-)}_{P(\epsilon^+|\epsilon^-)} = \sum_x P(x|\epsilon^-) \underbrace{\lambda_{\underline{\epsilon}^+ \Rightarrow \underline{x}}(x)}_{P(\epsilon^+|x)} \quad (50.12b)$$

argument of the function. Out of the two random variables in the subscript of the function, the one on the right hand side of the subscript, the one which is adjacent but beneath the argument, is always the argument.

²As usual in this book, $\mathcal{N}(!x)$ means a constant that is independent of x .

Eqs.(50.12) suggest that we define an edge bnet for the π and λ messages (these messages live in the edges between the nodes $\underline{\epsilon}^+, \underline{x}, \underline{\epsilon}^-$). Such an edge bnet, shown in Fig.50.4, is complementary to bnet for the nodes themselves. We will call it the **BP 2-track bnet** for the bnet Fig.50.3, because it has two “tracks”, one for π messages and another for λ ones. The TPMs, printed in blue, for bnet Fig.50.4, are as follows:

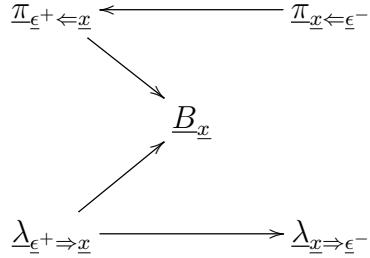


Figure 50.4: BP 2-track bnet for the bnet Fig.50.3.

$$P(\pi_{\underline{x} \leftarrow \underline{\epsilon}^-}) = \prod_{\epsilon^-} \mathbb{1}(\pi_{\underline{x} \leftarrow \underline{\epsilon}^-}(\epsilon^-) = P(\epsilon^-)) \quad (50.13)$$

$$P(\pi_{\underline{\epsilon}^+ \leftarrow \underline{x}} | \pi_{\underline{x} \leftarrow \underline{\epsilon}^-}) = \prod_x \mathbb{1} \left(\pi_{\underline{\epsilon}^+ \leftarrow \underline{x}}(x) = \sum_{\epsilon^-} P(x|\epsilon^-) \pi_{\underline{x} \leftarrow \underline{\epsilon}^-}(\epsilon^-) \right) \quad (50.14)$$

$$P(B_x | \pi_{\underline{\epsilon}^+ \leftarrow \underline{x}}, \lambda_{\underline{\epsilon}^+ \Rightarrow \underline{x}}) = \prod_x \mathbb{1}(B_x(x) = BEL_x(x)) \quad (50.15)$$

$$P(\lambda_{\underline{\epsilon}^+ \Rightarrow \underline{x}}) = \prod_x \mathbb{1}(\lambda_{\underline{\epsilon}^+ \Rightarrow \underline{x}}(x) = P(\epsilon^+|x)) \quad (50.16)$$

$$P(\lambda_{\underline{x} \Rightarrow \underline{\epsilon}^-} | \lambda_{\underline{\epsilon}^+ \Rightarrow \underline{x}}) = \prod_{\epsilon^-} \mathbb{1} \left(\lambda_{\underline{x} \Rightarrow \underline{\epsilon}^-}(\epsilon^-) = \sum_x P(x|\epsilon^-) \lambda_{\underline{\epsilon}^+ \Rightarrow \underline{x}}(x) \right) \quad (50.17)$$

So far in this section, we have considered Markov chains with 3 nodes. Before concluding our discussion of BP for Markov chains, let us consider BP for a slightly longer chain. Let us consider the 5 node Markov chain $\underline{\epsilon}^+ \leftarrow \underline{b} \leftarrow \underline{x} \leftarrow \underline{a} \leftarrow \underline{\epsilon}^-$ shown in Fig.50.5. We have already dealt with the end nodes of a Markov chain in the 3

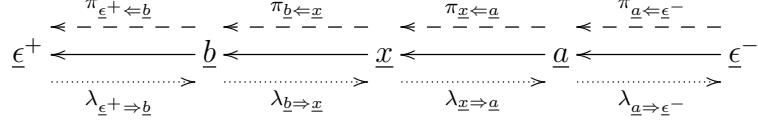


Figure 50.5: 5 node Markov chain

node Markov chain example above, so in the 5 node case, let us focus on the internal (i.e., not at an end) node \underline{x} and its neighbors \underline{a} and \underline{b} . Define

$$\pi_{\underline{b}\leftarrow\underline{x}}(x) = P(x|\epsilon^-) \text{ (past of } \underline{x}) , \quad (50.18)$$

$$\lambda_{\underline{b}\Rightarrow\underline{x}}(x) = P(\epsilon^+|x) \text{ (future of } \underline{x}) , \quad (50.19)$$

$$\pi_{\underline{x}\leftarrow\underline{a}}(a) = P(a|\epsilon^-) \text{ (past of } \underline{a}) \quad (50.20)$$

and

$$\lambda_{\underline{x}\Rightarrow\underline{a}}(a) = P(\epsilon^+|a) \text{ (future of } \underline{a}) . \quad (50.21)$$

Define the Belief BEL in x to be

$$BEL_{\underline{x}}(x) = P(x|\epsilon) , \quad (50.22)$$

where

$$\underline{\epsilon} = \underline{\epsilon}^+ \cup \underline{\epsilon}^- . \quad (50.23)$$

Then

$$BEL_{\underline{x}}(x) = \mathcal{N}(!x)P(\epsilon^+|x)P(x|\epsilon^-) \quad (50.24)$$

$$= \mathcal{N}(!x)\lambda_{\underline{b}\Rightarrow\underline{x}}(x)\pi_{\underline{b}\leftarrow\underline{x}}(x) . \quad (50.25)$$

In analogy with the case of BP for a 3 node Markov chain, we can define the bnet Fig.50.6, which we refer to as the **BP 2-track bnet** for Fig.50.5. The TPMs, printed in blue, for bnet Fig.50.6, are as follows:

$$P(\pi_{\underline{b}\leftarrow\underline{x}}|\pi_{\underline{x}\leftarrow\underline{a}}) = \prod_x \mathbb{1} \left(\pi_{\underline{b}\leftarrow\underline{x}}(x) = \sum_a P(x|a)\pi_{\underline{x}\leftarrow\underline{a}}(a) \right) \quad (50.26)$$

$$P(B_{\underline{x}}|\pi_{\underline{b}\leftarrow\underline{x}}, \lambda_{\underline{b}\Rightarrow\underline{x}}) = \prod_x \mathbb{1} (B_{\underline{x}}(x) = BEL_{\underline{x}}(x)) \quad (50.27)$$

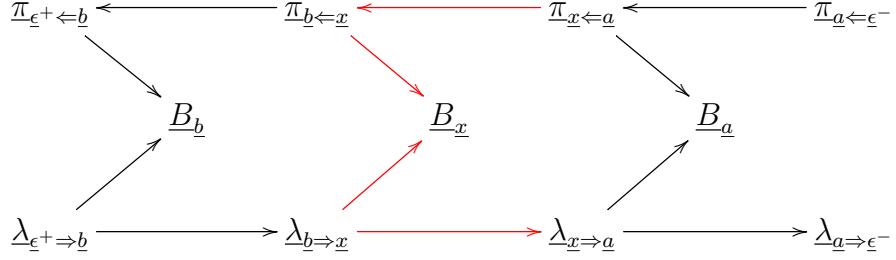


Figure 50.6: BP 2-track bnet for the bnet Fig.50.5.

$$P(\lambda_{\underline{x} \Rightarrow a} | \lambda_{b \Rightarrow \underline{x}}) = \prod_a \mathbb{1} \left(\lambda_{\underline{x} \Rightarrow a}(a) = \sum_x P(x|a) \lambda_{b \Rightarrow \underline{x}}(x) \right) \quad (50.28)$$

Let us represent the Markov chain of Fig.50.5 by $\underline{x}_{nx-1} \leftarrow \dots \leftarrow \underline{x}_2 \leftarrow \underline{x}_1 \leftarrow \underline{x}_0$ where $nx = 5$. For any node \underline{x}_i with parent $\underline{px}_i = \underline{x}_{i-1}$ and child $\underline{cx}_i = \underline{x}_{i+1}$, define the **memory matrix** $\mathcal{M}_{\underline{x}_i}$ for node \underline{x}_i as

$$\mathcal{M}_{\underline{x}_i} = [\mathcal{M}_{\underline{x}_i}^+, \mathcal{M}_{\underline{x}_i}^-], \quad (50.29)$$

where $+$ =future, $-$ =past, and

$$\mathcal{M}_{\underline{x}_i}^+ = \begin{bmatrix} \pi_{\underline{c}\underline{x}_i \leftarrow \underline{x}_i}(\cdot) \\ \lambda_{\underline{c}\underline{x}_i \rightarrow \underline{x}_i}(\cdot) \end{bmatrix}, \quad \mathcal{M}_{\underline{x}_i}^- = \begin{bmatrix} \pi_{\underline{x}_i \leftarrow \underline{p}\underline{x}_i}(\cdot) \\ \lambda_{\underline{x}_i \rightarrow \underline{p}\underline{x}_i}(\cdot) \end{bmatrix}. \quad (50.30)$$

Note that

$$\mathcal{M}_{\underline{x}_i}^- = \mathcal{M}_{\underline{p}\underline{x}_i}^+ \quad (50.31)$$

for all nodes \underline{x}_i . We will refer to Eqs.(50.31) as the **memory overlap conditions**.

We will also use a permuted version of the memory matrix

$$\mathcal{M}'_{\underline{x}_i} = [\mathcal{M}_{\underline{x}_i}^{OUT}, \mathcal{M}_{\underline{x}_i}^{IN}], \quad (50.32)$$

where

$$\mathcal{M}_{\underline{x}_i}^{OUT} = \begin{bmatrix} \pi_{\underline{c}\underline{x}_i \leftarrow \underline{x}_i}(\cdot) \\ \lambda_{\underline{x}_i \rightarrow \underline{p}\underline{x}_i}(\cdot) \end{bmatrix}, \quad \mathcal{M}_{\underline{x}_i}^{IN} = \begin{bmatrix} \pi_{\underline{x}_i \leftarrow \underline{p}\underline{x}_i}(\cdot) \\ \lambda_{\underline{c}\underline{x}_i \rightarrow \underline{x}_i}(\cdot) \end{bmatrix}. \quad (50.33)$$

$$\underline{\mathcal{M}}_{\underline{\epsilon}^+} \leftarrow \underline{\mathcal{M}}_b \leftarrow \underline{\mathcal{M}}_{\underline{x}} \leftarrow \underline{\mathcal{M}}_a \leftarrow \underline{\mathcal{M}}_{\underline{\epsilon}^-}$$

Figure 50.7: BP Memory Bnet for the bnet Fig.50.5.

Unfortunately, 2-track bnets cannot be generalized in any obvious way from Markov chains to more complicated DAGs. An alternative to 2-track bnets that still carries message info in its nodes, are memory bnets. An **BP memory bnet** is a bnet which takes each node of an original bnet and adds a local memory to it. More specifically, it keeps that DAG but replaces each node \underline{x}_i by a memory $\mathcal{M}_{\underline{x}_i}$. Fig.50.7 shows the memory bnet for the bnet Fig.50.5. The TPM, printed in blue, for the node $\mathcal{M}_{\underline{x}}$ of the memory bnet Fig.50.7, is as follows

$$P(\mathcal{M}_{\underline{x}_i} | \mathcal{M}_{n \in nb(\underline{x}_i)}) = AB , \quad (50.34)$$

where

$$A = \mathbb{1}(\mathcal{M}_{\underline{x}_i}^- = \mathcal{M}_{\underline{px}_i}^+) , \quad (50.35)$$

and

$$B = \mathbb{1}(\mathcal{M}_{\underline{x}_i}^{OUT} = \mathcal{C}(\mathcal{M}_{\underline{x}_i}^{IN})) . \quad (50.36)$$

The function \mathcal{C} , which we will call the **BP local computation**, maps $\mathcal{M}_{\underline{x}_i}^{IN}$ into $\mathcal{M}_{\underline{x}_i}^{OUT}$. More explicitly, \mathcal{C} is defined so that

$$B = \underbrace{P(\pi_{b \leftarrow \underline{x}} | \pi_{x \leftarrow a})}_{B_\pi} \underbrace{P(\lambda_{\underline{x} \Rightarrow a} | \lambda_{b \Rightarrow \underline{x}})}_{B_\lambda} , \quad (50.37)$$

where B_π and B_λ are given by Eqs.(50.26) and (50.28), respectively.

The BP memory bnet Fig. 50.7 is a deterministic bnet. A deterministic bnet is basically just a coupled system of equations (CSE) for some unknowns x_i . A CSE per se does not include with it a method for solving for the x_i . Such methods are not unique. For example, for the distributed soldier counting problem, the various methods that we described for counting soldiers are just different methods for solving the same CSE. One can describe a method for solving a CSE using a dynamical bnet.³ To solve the CSE represented by the memory bnet Fig.50.7, we will use the dynamical bnet Fig.50.8. Henceforth, we will refer to Fig.50.8 as an **BP dynamical bnet** for Fig.50.7.

Next, we will explain the meaning of Fig.50.8. Fig.50.8 is a step by step recipe (i.e., algorithm) for solving a CSE, where the unknowns are memory matrices. Each step encoded in Fig.50.8 corresponds to a specific message sending event, where the messages are sent along the edges of the Markov chain Fig.50.5. These message sending events are portrayed in chronological order in Fig.50.9. In that figure, π messages are indicated by dashed red arrows, and λ messages by dotted red arrows. These steps, or message sending events, lead to an updating of the memory matrices that we are solving for. Each step propagates information between the memory nodes.

³The term dynamical bnet was used in Chapter 24 to mean a time inhomogeneous Markov chain, but here we are stretching its meaning to include Markov chains that aren't time inhomogeneous.

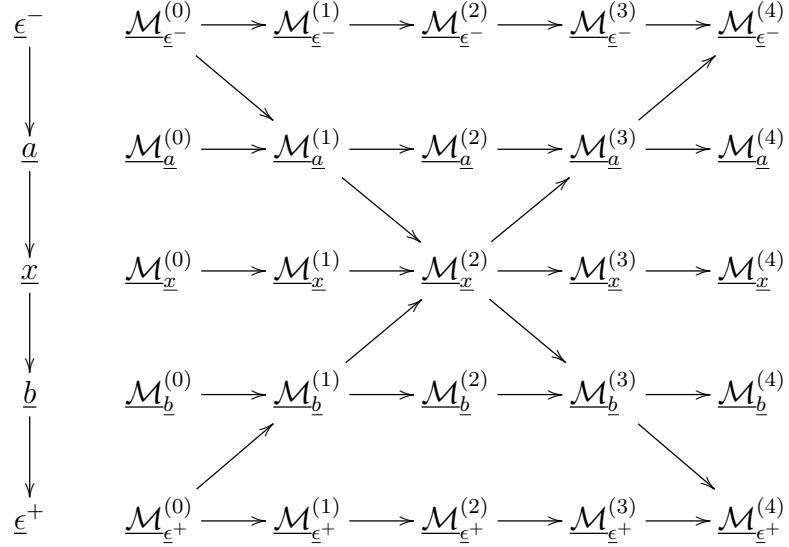


Figure 50.8: BP dynamical bnet for the bnet Fig.50.7.

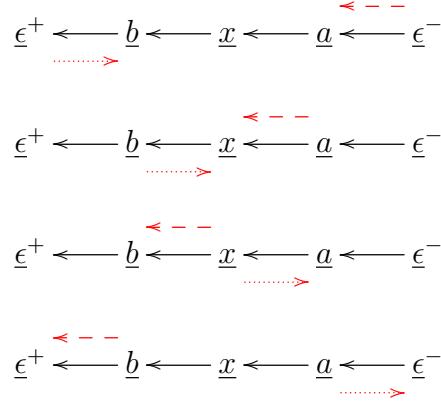


Figure 50.9: Steps encoded in the bnet Fig.50.8. Note the similarity of this figure to Fig.50.1 (d) for soldier counting.

In the usual Pearl BP algo, the evidence nodes initiate the BP chain of message passing events. These events continue until the memory matrices reach an equilibrium and the CSE is solved.

To use bnet Fig.50.8, we need to specify the **initial conditions** (i.e., the value of $\mathcal{M}_{\underline{x}_i}^{(0)}$ for all i). For that, one can use

$$\pi_{\underline{p}\underline{x}_0 \Leftarrow \underline{x}_0}^{(0)} = P(x_0) , \quad (50.38)$$

$$\lambda_{\underline{c}\underline{x}_i \Rightarrow \underline{x}_{nx-1}}^{(0)}(x_{nx-1}) = \delta(x_{nx-1}, x'_{nx-1}) . \quad (50.39)$$

All other $\mathcal{M}_{\underline{x}_i}^{(0)}$ entries for all i can be set to 1.

The TPMs, printed in blue, for bnet Fig.50.8, are as follows.

$$P(\mathcal{M}_{\underline{x}_i}^{(t)} | \mathcal{M}_{\underline{n} \in nb(\underline{x}_i)}^{(t-1)}, \mathcal{M}_{\underline{x}_i}^{(t-1)}) = AB, \quad (50.40)$$

where

$$A = \begin{cases} \mathbb{1}(\mathcal{M}_{\underline{x}_i}^{(t)-} = \mathcal{M}_{\underline{px}_i}^{(t-1)+}) & \text{if input from } \underline{px}_i \\ \mathbb{1}(\mathcal{M}_{\underline{x}_i}^{(t)+} = \mathcal{M}_{\underline{cx}_i}^{(t-1)-}) & \text{if input from } \underline{cx}_i \end{cases}, \quad (50.41)$$

and

$$B = \mathbb{1}(\mathcal{M}_{\underline{x}_i}^{(t)OUT} = \mathcal{C}(\mathcal{M}_{\underline{x}_i}^{(t)IN})). \quad (50.42)$$

The function \mathcal{C} , which we will call the **BP local computation**, maps $\mathcal{M}_{\underline{x}_i}^{(t)IN}$ into $\mathcal{M}_{\underline{x}_i}^{(t)OUT}$. More explicitly, \mathcal{C} is defined so that

$$B = B_\pi B_\lambda \quad (50.43)$$

where

$$B_\pi = \prod_x \mathbb{1} \left(\underbrace{\pi_{\underline{b} \Leftarrow \underline{x}}^{(t)}(x)}_{OUT} = \sum_a P(x|a) \underbrace{\pi_{\underline{x} \Leftarrow \underline{a}}^{(t)}(a)}_{IN} \right) \quad (50.44)$$

and

$$B_\lambda = \prod_a \mathbb{1} \left(\underbrace{\lambda_{\underline{x} \Rightarrow \underline{a}}^{(t)}(a)}_{OUT} = \sum_x P(x|a) \underbrace{\lambda_{\underline{b} \Rightarrow \underline{x}}^{(t)}(x)}_{IN} \right). \quad (50.45)$$

The basic idea behind Eq.(50.42), which we will call the **memory updating equation**, is simple: the memory overlap conditions translate the information from time $t - 1$ to t , and then the local computation translates IN to OUT at fixed time t .

50.4 BP Algorithm for Polytrees

Consider Fig.50.10, which illustrates a bnet node \underline{x} receiving and sending messages to its neighbors. The π messages (probability functions) travel downstream (i.e., they carry info in the direction of the graph arrows, towards the future) and are indicated by a dashed arrow or by a left double arrow \Leftarrow . The λ messages (likelihood functions) travel upstream (i.e., they carry info opposite to direction of the graph

arrows, towards the past) and are indicated by a dotted arrow or by a right double arrow \Rightarrow .

Note that argument arg of the $\pi(arg)$ and $\lambda(arg)$ functions is always the same as the letter in the subscript that is closest to the argument.

Note that in Fig.50.10, we indicate messages that travel “downstream” (resp., “upstream”), by arrows with dashed (resp., dotted) lines as shafts. Mnemonic: think of the shaft as a velocity vector field for the message. You travel faster when you swim downstream as opposed to upstream.

$$pa(\underline{x}) = \text{parents of node } \underline{x}$$

$$ch(\underline{x}) = \text{children of node } \underline{x}$$

$$nb(\underline{x}) = pa(\underline{x}) \cup ch(\underline{x}) = \text{neighbors of node } \underline{x}$$

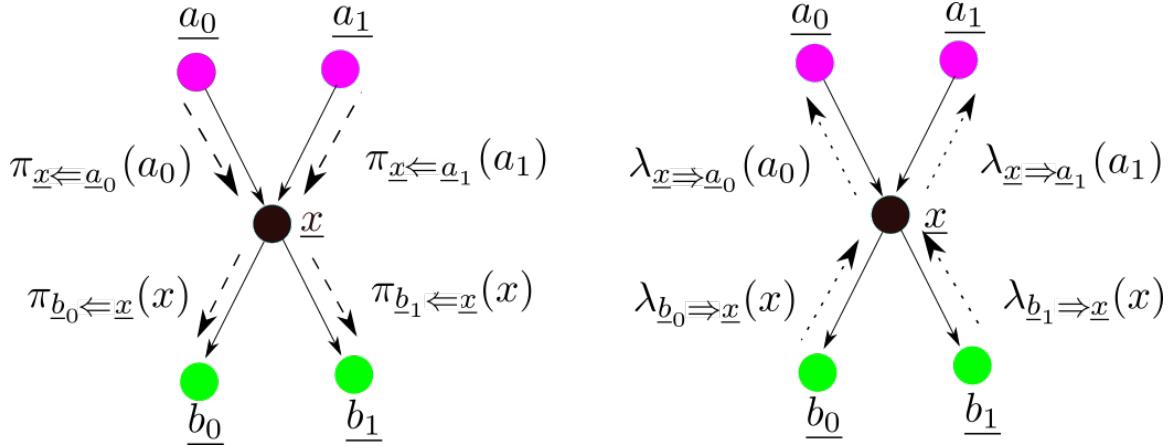


Figure 50.10: Node \underline{x} receiving and sending messages to its neighbors. (neighbors=parents and children).

We define a **memory matrix** $\mathcal{M}_{\underline{x}}$ for node \underline{x} as

$$\mathcal{M}_{\underline{x}} = [\mathcal{M}_{\underline{x}}^+, \mathcal{M}_{\underline{x}}^-] , \quad (50.46)$$

where $+$ =future, $-$ =past, and

$$\mathcal{M}_{\underline{x}}^+ = [\pi_{\underline{b} \leftarrow \underline{x}}(\cdot) \quad \lambda_{\underline{b} \Rightarrow \underline{x}}(\cdot)]_{\underline{b} \in ch(\underline{x})} = [\mathcal{M}_{\underline{b}, \underline{x}}^+]_{\underline{b} \in ch(\underline{x})} , \quad (50.47)$$

$$\mathcal{M}_{\underline{x}}^- = \left[\begin{array}{c} \pi_{\underline{x} \leftarrow \underline{a}}(\cdot) \\ \lambda_{\underline{x} \Rightarrow \underline{a}}(\cdot) \end{array} \right]_{\underline{a} \in pa(\underline{x})} = [\mathcal{M}_{\underline{x}, \underline{a}}^-]_{\underline{a} \in pa(\underline{x})} . \quad (50.48)$$

Note that

$$\mathcal{M}_{\underline{x}, \underline{a}}^- = \mathcal{M}_{\underline{a}, \underline{x}}^+ \quad (50.49)$$

for every arrow $\underline{x} \leftarrow \underline{a}$. We will refer to Eqs.(50.49) as the **memory overlap conditions**.

We will also use a permuted version of the memory matrix

$$\mathcal{M}'_{\underline{x}} = [\mathcal{M}_{\underline{x}}^{OUT}, \mathcal{M}_{\underline{x}}^{IN}] , \quad (50.50)$$

where

$$\mathcal{M}_{\underline{x}}^{OUT} = \left(\begin{array}{c} [\pi_{b \leftarrow \underline{x}}(\cdot)]_{b \in ch(\underline{x})} \\ [\lambda_{\underline{x} \Rightarrow a}(\cdot)]_{a \in pa(\underline{x})} \end{array} \right) = [\mathcal{M}_{\underline{x}, \underline{n}}^{OUT}]_{\underline{n} \in nb(\underline{x})} , \quad (50.51)$$

$$\mathcal{M}_{\underline{x}}^{IN} = \left(\begin{array}{c} [\pi_{x \leftarrow a}(\cdot)]_{a \in pa(\underline{x})} \\ [\lambda_{b \rightarrow \underline{x}}(\cdot)]_{b \in ch(\underline{x})} \end{array} \right) = [\mathcal{M}_{\underline{x}, \underline{n}}^{IN}]_{\underline{n} \in nb(\underline{x})} . \quad (50.52)$$

For times $t = 0, 1, \dots, T - 1$, we calculate $\mathcal{M}_{\underline{x}}^{(t)}$ in two steps: first we calculate $\mathcal{M}_{\underline{x}}^{(t)IN}$ from earlier memories at time $t - 1$, then we calculate $\mathcal{M}_{\underline{x}}^{(t)OUT}$:

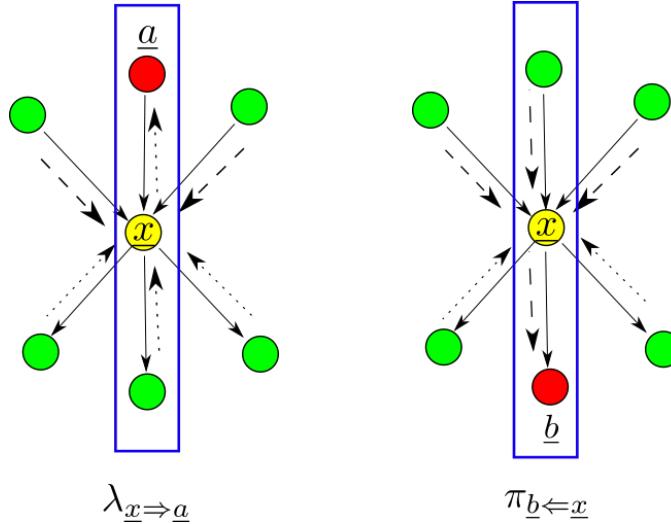


Figure 50.11: Subgraph of a bnet showing two cases (RULE 1 and RULE 2) of message info flow. The yellow node is a gossip monger. It receives messages from all the green nodes, and then it relays a joint message to the red node. Union of green nodes and the red node = full neighborhood of yellow node. There are two possible cases: the red node is either a parent or a child of the yellow one. As usual, we use arrows with dashed (resp., dotted) shafts for downstream (resp., upstream) messages. Blue boxes indicate Markov chain case.

An **evidence node** is a node whose TPM is a delta function set to a particular state of the node. We will assume, without loss of generality, that all evidence nodes are leaf nodes. If that is not the case, any evidence node \underline{e} that is not a leaf node, can be given a new companion leaf node \underline{l} connected to \underline{e} by an arrow $\underline{l} \leftarrow \underline{e}$, and such that \underline{l} has a delta function TPM.

1. Calculating $\mathcal{M}_{\underline{x}}^{(t)IN}$ from signals received from $\underline{n} \in nb(\underline{x})$, sent at earlier time $t - 1$:

Set

$$\mathcal{M}_{\underline{x}, \underline{a}}^{(t)-} |_{\pi} = \mathcal{M}_{\underline{a}, \underline{x}}^{(t-1)+} |_{\pi} , \quad (50.53)$$

for all $\underline{a} \in pa(\underline{x})$, and

$$\mathcal{M}_{\underline{b}, \underline{x}}^{(t)+} |_{\lambda} = \mathcal{M}_{\underline{x}, \underline{b}}^{(t-1)-} |_{\lambda} , \quad (50.54)$$

for all $\underline{b} \in ch(\underline{x})$. By $X|_{\lambda}$ (resp., $X|_{\pi}$) we mean the λ (resp., π) component of X .

2. Calculating $\mathcal{M}_{\underline{x}}^{(t)OUT}$ from already calculated $\mathcal{M}_{\underline{x}}^{(t)IN}$:

Let $\underline{a}^{na} = (\underline{a}_i)_{i=0,1,\dots,na-1}$ denote the parents of \underline{x} and $\underline{b}^{nb} = (\underline{b}_i)_{i=0,1,\dots,nb-1}$ its children.

Define

$$\pi_{\underline{x}}(x) = \sum_{a^{na}} P(x|a^{na}) \prod_i \pi_{\underline{x} \leftarrow \underline{a}_i}(a_i) \quad (50.55)$$

$$= E_{\underline{a}^{na}}[P(x|a^{na})] \quad (50.56)$$

(boundary case: if \underline{x} is a root node, use $\pi_{\underline{x}}(x) = P(x)$.) and

$$\lambda_{\underline{x}}(x) = \prod_i \lambda_{\underline{b}_i \Rightarrow \underline{x}}(x) . \quad (50.57)$$

(boundary case: if \underline{x} is a leaf node, use $\lambda_{\underline{x}}(x) = 1$.)

- RULE 1: (red parent)

From the $\lambda_{\underline{x} \Rightarrow \underline{a}}$ panel of Fig.50.11, we get

$$\underbrace{\lambda_{\underline{x} \Rightarrow \underline{a}_i}(a_i)}_{OUT} = \mathcal{N}(!a_i) \sum_x \left[\underbrace{\lambda_{\underline{x}}(x)}_{IN} \sum_{(a_k)_{k \neq i}} \left(P(x|a^{na}) \prod_{k \neq i} \underbrace{\pi_{\underline{x} \leftarrow \underline{a}_k}(a_k)}_{IN} \right) \right] \quad (50.58)$$

$$= \mathcal{N}(!a_i) \sum_x [\lambda_{\underline{x}}(x) E_{(\underline{a}_k)_{k \neq i}}[P(x|a^{na})]] \quad (50.59)$$

$$= \mathcal{N}(!a_i) E_{(\underline{a}_k)_{k \neq i}} E_{\underline{x}|a^{na}} \lambda_{\underline{x}}(x) \quad (50.60)$$

(boundary case: if \underline{x} is a root node, use $\lambda_{\underline{x} \Rightarrow \underline{a}_i}(a_i) = \mathcal{N}(!a_i)$.)

- RULE 2: (red child)

From the $\pi_{\underline{b} \leftarrow \underline{x}}$ panel of Fig.50.11, we get

$$\underbrace{\pi_{\underline{b}_i \leftarrow \underline{x}}(x)}_{OUT} = \mathcal{N}(!x) \underbrace{\pi_{\underline{x}}(x)}_{IN} \prod_{k \neq i} \underbrace{\lambda_{\underline{b}_k \Rightarrow \underline{x}}(x)}_{IN} \quad (50.61)$$

(boundary case: if \underline{x} is a leaf node, use $\pi_{\underline{b}_i \leftarrow \underline{x}}(x) = \mathcal{N}(!x) \pi_{\underline{x}}(x)$.)

In the above equations, if the range set of a product is empty, then define the product as 1; i.e., $\prod_{k \in \emptyset} F(k) = 1$.

Claim: Define

$$BEL^{(t)}(\underline{x}) = \mathcal{N}(!\underline{x}) \lambda_{\underline{x}}^{(t)}(\underline{x}) \pi_{\underline{x}}^{(t)}(\underline{x}) . \quad (50.62)$$

Then

$$\lim_{t \rightarrow \infty} BEL^{(t)}(\underline{x}) = P(\underline{x}|\epsilon) . \quad (50.63)$$

This says that the belief in $\underline{x} = \underline{x}$ converges to $P(\underline{x}|\epsilon)$ and it equals the product of messages received from all parents and children of $\underline{x} = \underline{x}$.

50.4.1 How BP algo for polytrees reduces to the BP algo for Markov chains

It is instructive to see how the BP algo for polytrees reduces to BP algo for Markov chains.

For a Markov chain, node \underline{x} has a single parent (i.e., ancestor) \underline{a} and a single child \underline{b} .

Therefore, Eqs.(50.55) and (50.57) reduce to

$$\pi_{\underline{x}}(\underline{x}) = \sum_a P(\underline{x}|a) \pi_{\underline{x} \leftarrow \underline{a}}(a) \quad (50.64)$$

and

$$\lambda_{\underline{x}}(\underline{x}) = \lambda_{\underline{b} \rightarrow \underline{x}}(\underline{x}) . \quad (50.65)$$

RULE 1 given by Eq.(50.58) reduces to

$$\lambda_{\underline{x} \rightarrow \underline{a}}(a) = \mathcal{N}(!a) \sum_x \lambda_{\underline{x}}(x) P(x|a) \quad (50.66)$$

$$= \mathcal{N}(!a) \sum_x \lambda_{\underline{b} \rightarrow \underline{x}}(x) P(x|a) \quad (50.67)$$

RULE 2 given by Eq.(50.61) reduces to

$$\pi_{\underline{b} \leftarrow \underline{x}}(\underline{x}) = \mathcal{N}(!\underline{x}) \pi_{\underline{x}}(\underline{x}) \quad (50.68)$$

$$= \sum_a P(\underline{x}|a) \pi_{\underline{x} \leftarrow \underline{a}}(a) . \quad (50.69)$$

50.5 Derivation of BP Algorithm for Polytrees

This derivation is taken from the 1988 book Ref.[44] by Judea Pearl, where it is presented very lucidly. We only made some minor changes in notation.

Notation

The BP algorithm yields an expansion for $P(x|\epsilon)$.

\underline{x} = the focus node, arbitrary node of bnet that we are focusing on to calculate its $P(x|\epsilon)$.

$(a_i)_{i=0,1,\dots,na-1}$. = parent nodes (mnemonic: a=ancestor) of \underline{x}

$(b_i)_{i=0,1,\dots,nb-1}$. = children nodes of \underline{x} .

$\underline{\epsilon}$ = set of nodes for which there is evidence; that is, $\underline{\epsilon} = \epsilon$, so the state of these nodes is fixed.

$$\epsilon_{\underline{x}}^- = \epsilon \cap an(\underline{x}) \text{ (evidence in past of } \underline{x})^4$$

$$\epsilon_{xa_i}^- = \epsilon_{\underline{x}}^- \cap an(a_i).$$

$$\text{Note that } \epsilon_{\underline{x}}^- = \cup_i \epsilon_{xa_i}^-$$

$$\epsilon_{\underline{x}}^+ = \epsilon \cap [de(\underline{x}) \cup \underline{x}] \text{ (evidence in future of } \underline{x})$$

$$\epsilon_{xb_i}^+ = \epsilon_{\underline{x}}^+ \cap [de(b_i) \cup b_i].$$

$$\text{Note that } \epsilon_{\underline{x}}^+ = \cup_i \epsilon_{xb_i}^+$$

$$\text{Note that } \underline{\epsilon} = \epsilon_{\underline{x}}^+ \cup \epsilon_{\underline{x}}^-$$

$$\pi_{\underline{x}}(x) = P(x|\epsilon_{\underline{x}}^-) \quad (50.70)$$

$$\pi_{\underline{x} \leftarrow a_i}(a_i) = P(a_i|\epsilon_{xa_i}^-) \quad (50.71)$$

$$\pi_{\underline{b}_i \leftarrow \underline{x}}(x) = P(x|\epsilon_{xb_i}^-) \quad (50.72)$$

$$\lambda_{\underline{x}}(x) = P(\epsilon_{\underline{x}}^+|x) \quad (50.73)$$

$$\lambda_{b_i \Rightarrow \underline{x}}(x) = P(\epsilon_{xb_i}^+|x) \quad (50.74)$$

$$\lambda_{\underline{x} \Rightarrow a_i}(a_i) = P(\epsilon_{xa_i}^+|a_i) \quad (50.75)$$

Expansions of $\lambda_{\underline{x}}(x)$ and $\pi_{\underline{x}}(x)$ into products of single node messages.

⁴Careful: Chapter 4 of Ref.[44] uses – indicate the future and + to indicate the past. This is the opposite of our notation.

$$\underbrace{P(x|\epsilon_{\underline{x}}^-)}_{\pi_{\underline{x}}(x)} = P(x| \cup_i \epsilon_{\underline{x}a_i}^-) \quad (50.76)$$

$$= \sum_{a^{na}} P(x|a^{na}) P(a^{na}| \cup_i \epsilon_{\underline{x}a_i}^-) \quad (50.77)$$

$$= \sum_{a^{na}} P(x|a^{na}) \prod_i \underbrace{P(a_i|\epsilon_{\underline{x}a_i}^-)}_{\pi_{\underline{x} \leftarrow a_i}(a_i)} \quad (50.78)$$

$$\underbrace{P(\epsilon_{\underline{x}}^+|x)}_{\lambda_{\underline{x}}(x)} = \prod_i \underbrace{P(\epsilon_{\underline{x}b_i}^+|x)}_{\lambda_{\underline{b}_i \Rightarrow \underline{x}}(x)} \quad (50.79)$$

Note that past and future evidences $\epsilon_{\underline{x}}^-$ and $\epsilon_{\underline{x}}^+$ that are causally connected to \underline{x} are conditionally independent at fixed \underline{x} :

$$P(\epsilon_{\underline{x}}^+, \epsilon_{\underline{x}}^-|x) = P(\epsilon_{\underline{x}}^+|x)P(\epsilon_{\underline{x}}^-|x). \quad (50.80)$$

This observation is key to the proof of the following claim:

Claim 75

$$P(x|\epsilon_{\underline{x}}^+, \epsilon_{\underline{x}}^-) = P(\epsilon_{\underline{x}}^+|x)P(x|\epsilon_{\underline{x}}^-) \frac{1}{P(\epsilon_{\underline{x}}^+|\epsilon_{\underline{x}}^-)} \quad (50.81)$$

$$= \mathcal{N}(!x)P(\epsilon_{\underline{x}}^+|x)P(x|\epsilon_{\underline{x}}^-) \quad (50.82)$$

$$= \mathcal{N}(!x) (\epsilon_{\underline{x}}^+ \leftarrow x \leftarrow \epsilon_{\underline{x}}^-) \quad (50.83)$$

$$= \mathcal{N}(!x)\lambda_{\underline{x}}(x)\pi_{\underline{x}}(x) \quad (50.84)$$

proof:

$$P(x|\epsilon_{\underline{x}}^+, \epsilon_{\underline{x}}^-) = P(\epsilon_{\underline{x}}^+, \epsilon_{\underline{x}}^-|x) \frac{P(x)}{P(\epsilon_{\underline{x}}^+, \epsilon_{\underline{x}}^-)} \quad (50.85)$$

$$= P(\epsilon_{\underline{x}}^+|x)P(\epsilon_{\underline{x}}^-|x) \frac{P(x)}{P(\epsilon_{\underline{x}}^+, \epsilon_{\underline{x}}^-)} \quad (50.86)$$

$$= P(\epsilon_{\underline{x}}^+|x)P(x|\epsilon_{\underline{x}}^-) \frac{P(\epsilon_{\underline{x}}^-)}{P(\epsilon_{\underline{x}}^+, \epsilon_{\underline{x}}^-)} \quad (50.87)$$

$$= P(\epsilon_{\underline{x}}^+|x)P(x|\epsilon_{\underline{x}}^-) \frac{1}{P(\epsilon_{\underline{x}}^+|\epsilon_{\underline{x}}^-)} \quad (50.88)$$

QED

Next we prove BP rules 1 and 2.

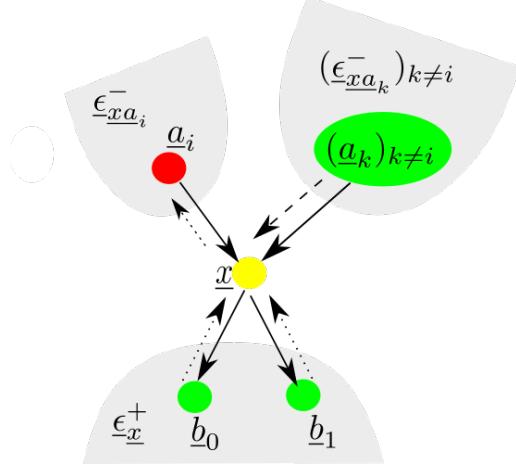


Figure 50.12: This figure is used in the derivation of the BP RULE 1.

- **RULE 1 (red parent)**

Note that

$$\epsilon_x^+ \cup \cup_{k \neq i} \epsilon_{xa_k}^- = (\epsilon_x^+ \cup \epsilon_x^-) - \epsilon_{xa_i}^- \quad (50.89)$$

$$= \epsilon_{xa_i}^+ \quad (50.90)$$

Let $y = (a_k)_{k \neq i}$ and $\epsilon_y^- = (\epsilon_{xa_k}^-)_{k \neq i}$.

$$\underbrace{P(\epsilon_{xa_i}^+ | a_i)}_{\lambda_{\underline{x} \Rightarrow \underline{a}_i}(a_i)} = P(\epsilon_x^+, \epsilon_y^- | a_i) \quad (50.91)$$

$$= \sum_x \sum_y P(\epsilon_x^+, \epsilon_y^- | x, y) P(x, y | a_i) \quad (50.92)$$

$$= \sum_x \sum_y P(\epsilon_x^+ | x) P(\epsilon_y^- | y) P(x | y, a_i) P(y | a_i) \quad (50.93)$$

$$= P(\epsilon_y^-) \sum_x \sum_y P(\epsilon_x^+ | x) \frac{P(y | \epsilon_y^-)}{P(y)} P(x | y, a_i) \underbrace{P(y | a_i)}_{=P(y)} \quad (50.94)$$

$$= \mathcal{N}(!a_i) \sum_x \sum_y P(\epsilon_x^+ | x) P(x | \underbrace{y, a_i}_{a^{na}}) P(y | \epsilon_y^-) \quad (50.95)$$

$$= \mathcal{N}(!a_i) \sum_x \underbrace{P(\epsilon_x^+ | x)}_{\lambda_{\underline{x}}(x)} \sum_{(a_k)_{k \neq i}} P(x | a^{na}) \prod_{k \neq i} \underbrace{P(a_k | \epsilon_{xa_k}^-)}_{\pi_{\underline{x} \leftarrow \underline{a}_k}(a_k)} \quad (50.96)$$

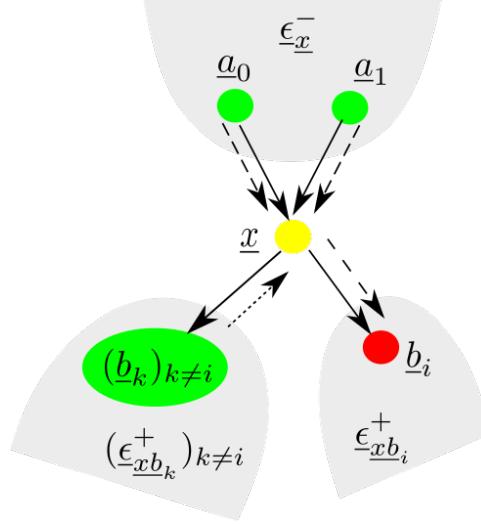


Figure 50.13: This figure is used in the derivation of the BP RULE 2.

- **RULE 2 (red child)**

Note that

$$(\cup_{k \neq i} \epsilon_{xb_k}^+) \cup \epsilon_x^- = (\epsilon_x^+ \cup \epsilon_x^-) - \epsilon_{xb_i}^+ \quad (50.97)$$

$$= \epsilon_{xb_i}^- \quad (50.98)$$

$$\underbrace{P(x|\epsilon_{xb_i}^-)}_{\pi_{b_i \leftarrow x}(x)} = P(x|(\epsilon_{xb_k}^+}_{k \neq i})_{k \neq i}, \epsilon_x^- \quad (50.99)$$

$$= \mathcal{N}(!x) P((\epsilon_{xb_k}^+}_{k \neq i}|x) P(x|\epsilon_x^-) \quad (50.100)$$

$$= \mathcal{N}(!x) \left(\prod_{k \neq i} \underbrace{P(\epsilon_{xb_k}^+|x)}_{\lambda_{b_k \Rightarrow x}(x)} \right) \underbrace{P(x|\epsilon_x^-)}_{\pi_x(x)} \quad (50.101)$$

50.6 Example of BP algo for a Tree

In this section, we describe how to apply the BP algo to the tree bnet Fig.50.14. In Fig.50.14, if we replace each integer i by the random variable \underline{A}_i , we get an **original bnet**, and if we replace each i by $\underline{M}_{\underline{A}_i}$, we get the **BP memory bnet** of the original bnet. In Fig.50.14, the magenta nodes are evidence nodes and the green ones aren't.

We want to solve for the memory matrices of the memory bnet. To do so, we use the **BP dynamical bnet** Fig.50.15. The steps encoded in the dynamical bnet are shown in Fig.50.16. Fig.50.16 has frames in chronological order, showing the direction of travel of the $\pi\&\lambda$ information. This sequence of frames also indicates the order in which we solve for the entries of the memory matrices. The information first emanates from the evidence nodes. It propagates generally upstream, although some nodes can generate downstream flow. Some of the info reaches the root node and is reflected there. The root node is the only one that is capable of reflection (i.e., instant output along an arrow, in response to input along that arrow). Eventually, all info reaches the leaf nodes via downstream propagation and is absorbed there.

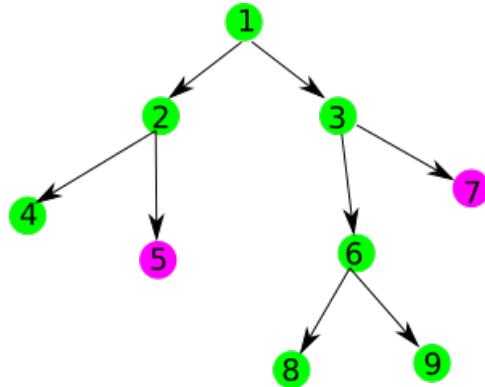


Figure 50.14: Example tree bnet used to illustrate BP.

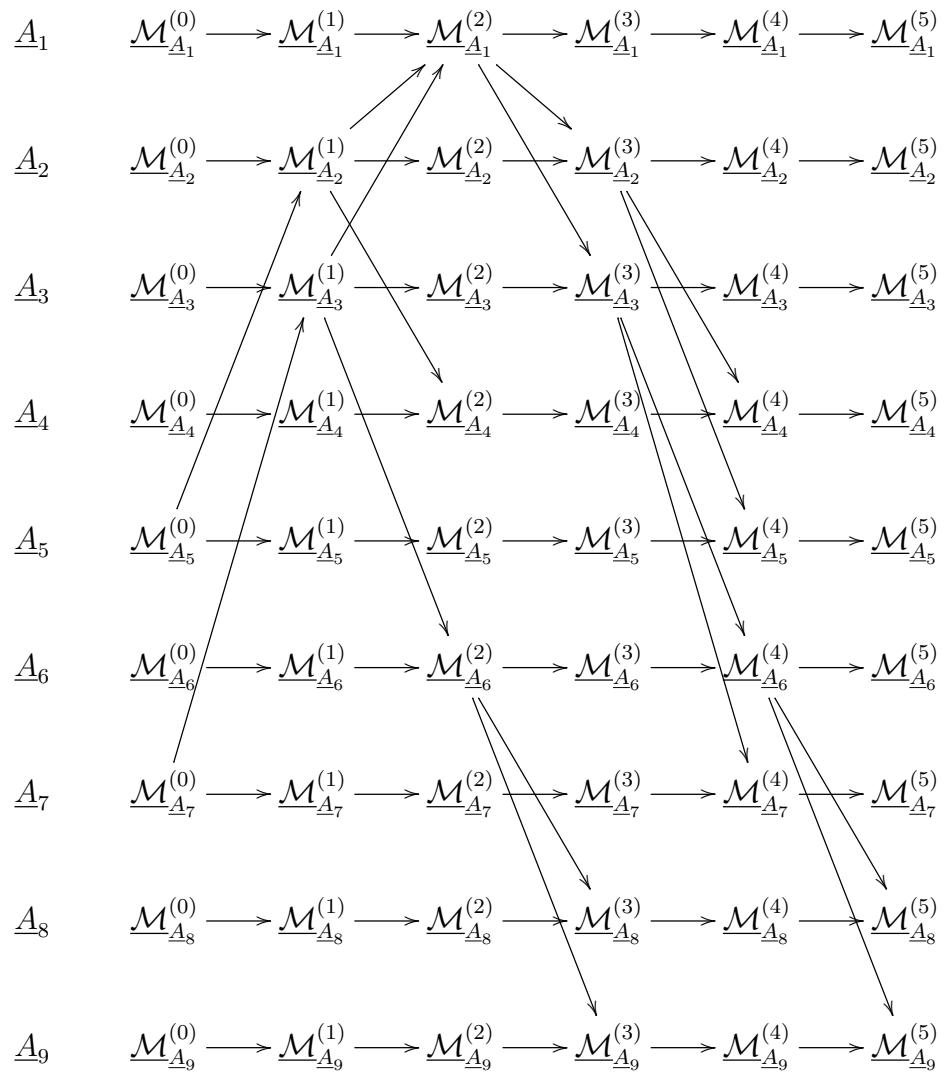


Figure 50.15: BP dynamical bnet for the bnet Fig.50.14.

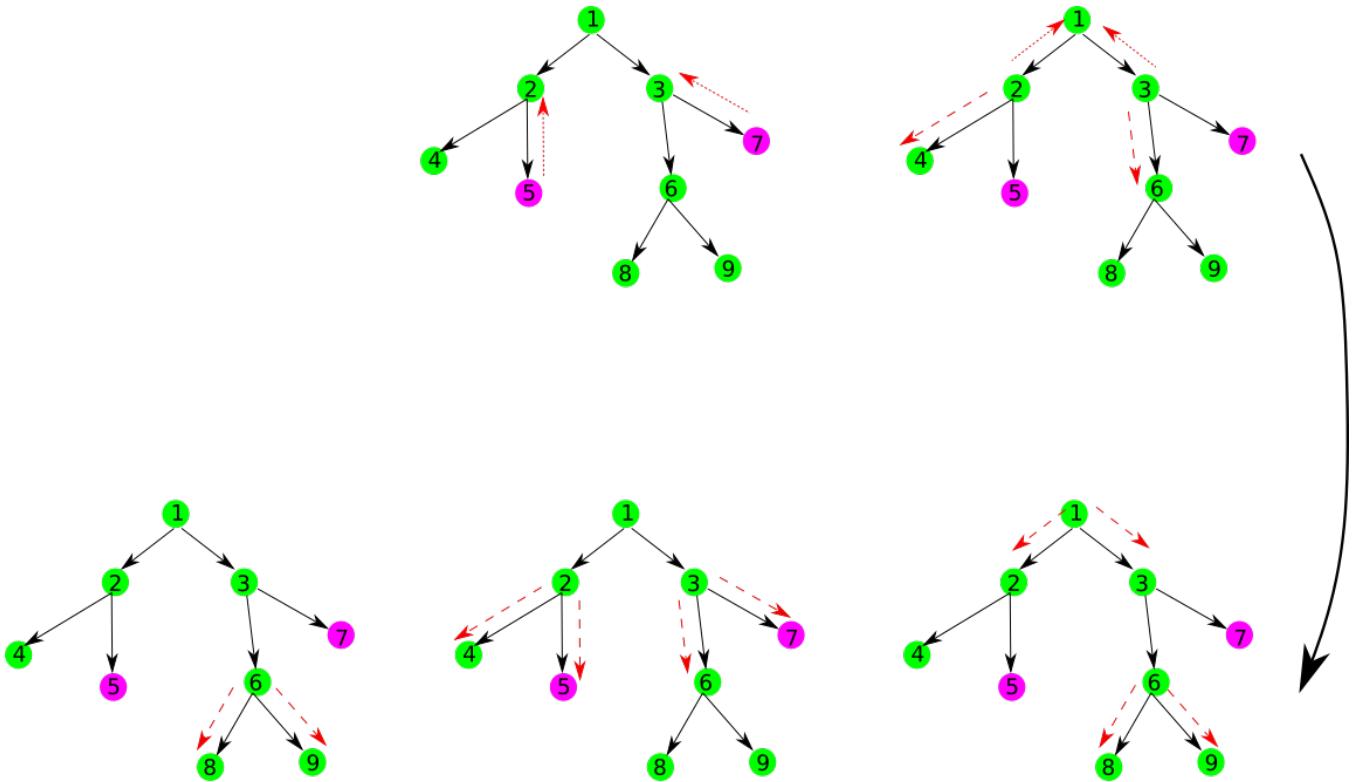


Figure 50.16: Steps encoded in the bnet Fig.50.15.

50.7 Bipartite bnets

By a **bipartite bnet** we will mean a bnet in which all nodes are either root nodes (parentless) or leaf nodes (childless). BP simplifies when dealing with bipartite bnets. Next, we will explain how it simplifies. But before doing so, let us define tree bnets and show how these can be replaced by equivalent bipartite bnets.

A **tree bnet** is a bnet for which all nodes have exactly one parent except for the apex root node which has none. A tree bnet is very much like the filing system in a computer.

One can map a tree bnet (the “source”) into an equivalent bipartite bnet (the “image”) as follows. Replace each arrow

$$\underline{x} \longrightarrow \underline{y} \quad (50.102)$$

of the tree bnet by

$$\underline{x} \longrightarrow P_{\underline{y}|\underline{x}} \longleftarrow \underline{y} . \quad (50.103)$$

For example, the tree bnet Fig.50.17 has the image bipartite bnet given by Fig.50.18. The bnet Fig.50.19 is just a different way of drawing the bnet Fig.50.18.

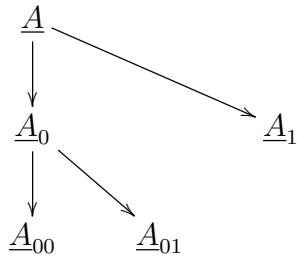


Figure 50.17: Example of a tree bnet.

The TPMs, printed in blue, for the image bipartite bnet Fig.50.18, are as follows. We express the TPMs of the image bnet in terms of the TPMs of the source bnet Fig.50.17. Let

$$P(P_{\underline{y}|\underline{x}}|x, y) = P_{\underline{y}|\underline{x}}(y|x)\delta(P_{\underline{y}|\underline{x}}, 1) + (1 - P_{\underline{y}|\underline{x}}(y|x))\delta(P_{\underline{y}|\underline{x}}, 0) \quad (50.104)$$

for all the leaf nodes $P_{\underline{y}|\underline{x}} \in \{0, 1\}$ of the image bipartite bnet. Also, let

$$P_y(y) = \text{arbitrary prior} \quad (50.105)$$

for all the root nodes \underline{y} of the image bipartite bnet except when \underline{y} corresponds to the root node \underline{A} of the source tree bnet. In that exceptional case,

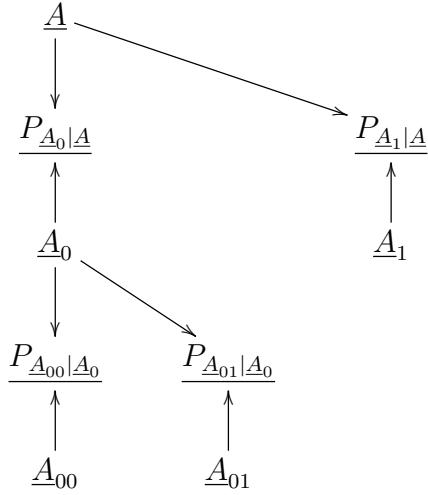


Figure 50.18: Bipartite bnet corresponding to tree bnet Fig.50.17.

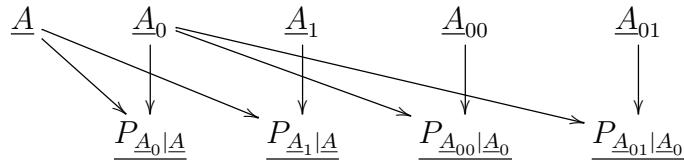


Figure 50.19: Different way of drawing the bnet Fig.50.18.

$$P_y(y) = P_{\underline{A}}(y) . \quad (50.106)$$

50.8 BP for bipartite bnets (BP-BB)

For a bipartite bnet as defined above, with root nodes \underline{x}_i and leaf nodes \underline{f}_α , let

$$nb(i) = \{\alpha : \underline{f}_\alpha \in nb(\underline{x}_i)\} , \quad (50.107)$$

$$nb(\alpha) = \{i : \underline{x}_i \in nb(\underline{f}_\alpha)\} , \quad (50.108)$$

$$m_{\alpha \Leftarrow i}(x_i) = \pi_{\underline{f}_\alpha \Leftarrow \underline{x}_i}(x_i) , \quad (50.109)$$

$$m_{\alpha \Rightarrow i}(x_i) = \lambda_{\underline{f}_\alpha \Rightarrow \underline{x}_i}(x_i) , \quad (50.110)$$

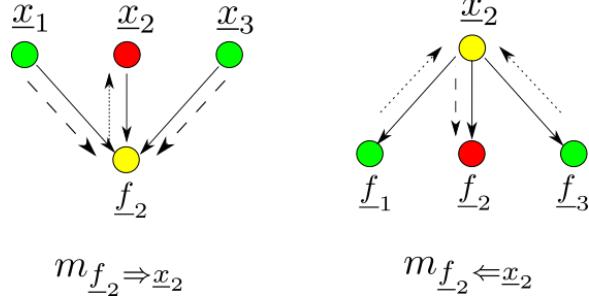


Figure 50.20: Fig.50.11 becomes this figure for the special case of a bipartite bnet. Union of green nodes and the red node = full neighborhood of yellow node. There are two possible cases: the red node is either a parent or a child of the yellow node.

Next we will show how to find $m_{\alpha \Leftarrow i}^{(t)}$ and $m_{\alpha \Rightarrow i}^{(t)}$ from $m_{\alpha \Leftarrow i}^{(t-1)}$ and $m_{\alpha \Rightarrow i}^{(t-1)}$.

1. Traversing an x (i.e., root) node.

See the $m_{f_2 \Leftarrow x_2}$ panel of Fig.50.20.

For $i = 0, 1, \dots, nx - 1$, if $\alpha \in nb(i)$, then,

$$m_{\alpha \Leftarrow i}^{(t)}(x_i) = \prod_{\beta \in nb(i) - \alpha} m_{\beta \Rightarrow i}^{(t-1)}(x_i), \quad (50.111)$$

whereas if $\alpha \notin nb(i)$

$$m_{\alpha \Leftarrow i}^{(t)}(x_i) = m_{\alpha \Leftarrow i}^{(t-1)}(x_i). \quad (50.112)$$

2. Traversing an f (i.e., leaf) node.

See the $m_{f_2 \Rightarrow x_2}$ panel of Fig.50.20.

For $\alpha = 0, 1, \dots, nf - 1$, if $i \in nb(\alpha)$, then

$$m_{\alpha \Rightarrow i}^{(t)}(x_i) = \sum_{(x_k)_{k \in nb(\alpha) - i}} f_\alpha(x_{nb(\alpha)}) \prod_{k \in nb(\alpha) - i} m_{\alpha \Leftarrow k}^{(t-1)}(x_k) \quad (50.113)$$

$$= E_{(x_k)_{k \in nb(\alpha) - i}}^{(t-1)} [f_\alpha(x_{nb(\alpha)})], \quad (50.114)$$

whereas if $i \notin nb(\alpha)$

$$m_{\alpha \Rightarrow i}^{(t)}(x_i) = m_{\alpha \Rightarrow i}^{(t-1)}(x_i). \quad (50.115)$$

In the above equations, if the range set of a product is empty, then define the product as 1; i.e., $\prod_{k \in \emptyset} F(k) = 1$.

Claim:

$$P(x_i|\epsilon) = \lim_{t \rightarrow \infty} \mathcal{N}(!x_i) \prod_{\alpha \in nb(i)} m_{\alpha \Rightarrow i}^{(t)}(x_i) \quad (50.116)$$

and

$$P(x_{nb(\alpha)}|\epsilon) = \lim_{t \rightarrow \infty} \mathcal{N}(!x_{nb(\alpha)}) f_\alpha(x_{nb(\alpha)}) \prod_{k \in nb(\alpha)} m_{\alpha \Leftarrow k}^{(t)}(x_k). \quad (50.117)$$

50.8.1 BP-BB and general BP agree on Markov chains

It is instructive to compare the belief values (i.e., $P(x_i|\epsilon)$) obtained by applying the general (i.e., polytree) BP and BP-BB algorithms to a Markov chain. Next we show that both algorithms yield the same belief values.

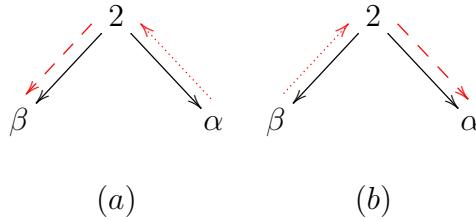


Figure 50.21: Traversing a root node of a Markov chain (a)Propagation towards left (i.e., towards future). (b)Propagation towards right (i.e., towards past).

Consider the BP-BB rule for traversing a root node. When traveling towards the left as in Fig.50.21 (a), it implies that

$$m_{\alpha \Rightarrow 2}(x_2) = m_{\beta \Leftarrow 2}(x_2), \quad (50.118)$$

and when traveling towards the right as in Fig.50.21 (b), it implies that

$$m_{\beta \Rightarrow 2}(x_2) = m_{\alpha \Leftarrow 2}(x_2). \quad (50.119)$$

Now consider the BP-BB rule for traversing a leaf node. When traveling to the left as in Fig.50.22 (a), it implies that

$$\underbrace{m_{\alpha \Rightarrow 2}(x_2)}_{\lambda} = \sum_{x_1} P(x_2|x_1) \underbrace{m_{\alpha \Leftarrow 1}(x_1)}_{\pi}. \quad (50.120)$$

One can rewrite the left and right hand sides (LHS, RHS) of Eq.(50.120) as follows

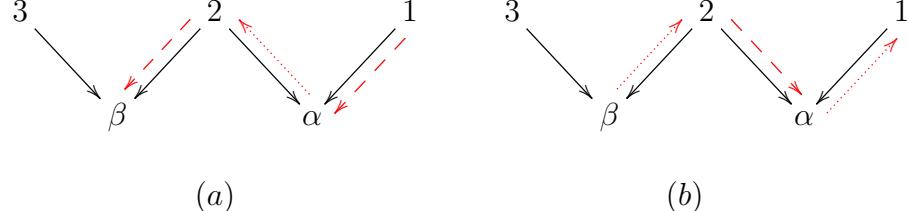


Figure 50.22: Traversing a leaf node of a Markov chain (a)Propagation towards left (i.e., towards future). (b)Propagation towards right (i.e., towards past).

$$RHS = \sum_{x_1} P(x_2|x_1) \pi_{\alpha \Leftarrow 1}(x_1) , \quad (50.121)$$

and

$$LHS = m_{\alpha \Rightarrow 2}(x_2) = m_{\beta \Leftarrow 2}(x_2) = \pi_{\beta \Leftarrow 2}(x_2) , \quad (50.122)$$

Therefore

$$\pi_{\beta \Leftarrow 2}(x_2) \sum_{x_1} P(x_2|x_1) \pi_{\alpha \Leftarrow 1}(x_1) . \quad (50.123)$$

Once again, consider the BP-BB rule for traversing a leaf node. When traveling to the right as in Fig.50.22 (b), it implies that

$$\underbrace{m_{\alpha \Rightarrow 1}(x_1)}_{\lambda} = \sum_{x_2} P(x_2|x_1) \underbrace{m_{\alpha \Leftarrow 2}(x_2)}_{\pi} . \quad (50.124)$$

One can rewrite the left and right hand sides (LHS, RHS) of Eq.(50.124) as follows

$$RHS = \sum_{x_2} P(x_2|x_1) \pi_{\alpha \Leftarrow 2}(x_2) \quad (50.125)$$

$$= \sum_{x_2} P(x_2|x_1) \lambda_{\beta \Rightarrow 2}(x_2) , \quad (50.126)$$

and

$$LHS = \lambda_{\alpha \Rightarrow 1}(x_1) . \quad (50.127)$$

Therefore,

$$\lambda_{\alpha \Rightarrow 1}(x_1) = \sum_{x_2} P(x_2|x_1) \lambda_{\beta \Rightarrow 2}(x_2) . \quad (50.128)$$

Finally, note that Eq.(50.116) becomes

$$P(x_2|\epsilon) = \mathcal{N}(!x_2)m_{\beta \Rightarrow 2}(x_2)m_{\alpha \Rightarrow 2}(x_2) \quad (50.129)$$

$$= \mathcal{N}(!x_2)m_{\alpha \Leftarrow 2}(x_2)m_{\alpha \Rightarrow 2}(x_2) \quad (50.130)$$

$$= \mathcal{N}(!x_2)\pi_{\alpha \Leftarrow 2}(x_2)\lambda_{\alpha \Rightarrow 2}(x_2) \quad (50.131)$$

$$= \mathcal{N}(!x_2)P(x_2|\epsilon^-)P(x_2|\epsilon^+) \quad (50.132)$$

and Eq.(50.117) becomes

$$P(x_2, x_1) = \mathcal{N}(!x_2, !x_1)P(x_2|x_1)m_{\alpha \Leftarrow 1}(x_1)m_{\alpha \Leftarrow 2}(x_2) \quad (50.133)$$

$$= \mathcal{N}(!x_2, !x_1)P(x_2|x_1)\pi_{\alpha \Leftarrow 1}(x_1)\pi_{\alpha \Leftarrow 2}(x_2). \quad (50.134)$$

50.8.2 BP-BB and general BP agree on tree bnets.

It is instructive to compare the belief values (i.e., $P(x_i|\epsilon)$) obtained by applying the general (i.e., polytree) BP and BP-BB algorithms to a tree bnet. Next we show that both algorithms yield the same belief values.

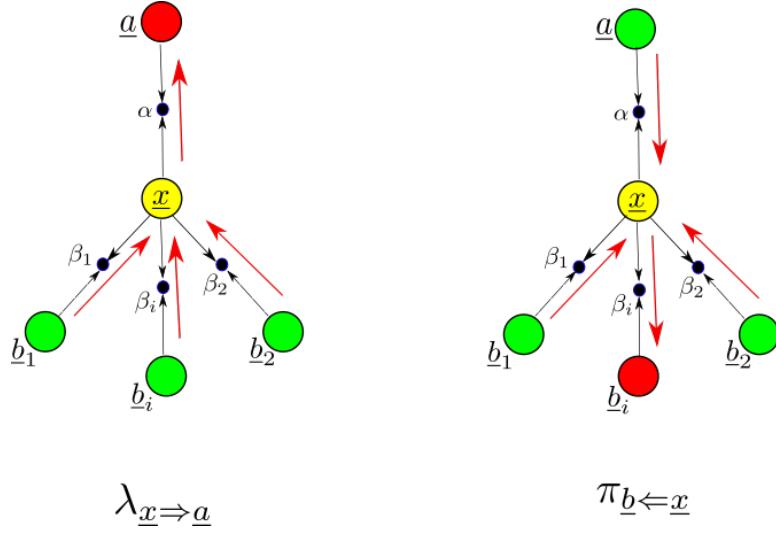


Figure 50.23: Subgraph of a tree bnet. This is the same as Fig.50.11, except that here the yellow node has a single parent because this is a subgraph of a tree bnet, not of an arbitrary bnet like Fig.50.11. The subgraph has been converted to a subgraph of a bipartite bnet by inserting a collider leaf node, labeled by a Greek letter, at the center of each edge of the tree bnet. Red arrows indicate the direction of message info flow.

Applying to the left panel of Fig.50.23 the BP-BB rule for traversing a root node, we get

$$m_{\alpha \Leftarrow \underline{x}}(x) = \prod_i m_{\beta_i \Rightarrow \underline{x}}(x) . \quad (50.135)$$

Applying to the left panel of Fig.50.23 the BP-BB rule for traversing a leaf node, we get

$$m_{\alpha \Rightarrow \underline{a}}(a) = \mathcal{N}(!a) \sum_x m_{\alpha \Leftarrow \underline{x}}(x) P(x|a) . \quad (50.136)$$

Combining Eqs.(50.135) and (50.136), we get

$$m_{\alpha \Rightarrow \underline{a}}(a) = \mathcal{N}(!a) \sum_x P(x|a) \prod_i m_{\beta_i \Rightarrow \underline{x}}(x) , \quad (50.137)$$

which can be rewritten as

$$\lambda_{\underline{x} \Rightarrow \underline{a}}(a) = \mathcal{N}(!a) \sum_x P(x|a) \underbrace{\prod_i \lambda_{\underline{b}_i \Rightarrow \underline{x}}(x)}_{\lambda_{\underline{x}}(x)} . \quad (50.138)$$

Eq.(50.138) is just RULE 1 for general BP.

Applying to the right panel of Fig.50.23 the BP-BB rule for traversing a root node, we get

$$m_{\beta_i \Leftarrow \underline{x}}(x) = \mathcal{N}(!x) m_{\alpha \Rightarrow \underline{x}}(x) \prod_{k \neq i} m_{\beta_k \Rightarrow \underline{x}}(x) \quad (50.139)$$

Applying to the right panel of Fig.50.23 the BP-BB rule for traversing a leaf node, we get

$$m_{\alpha \Rightarrow \underline{x}}(x) = \sum_a P(x|a) m_{\alpha \Leftarrow \underline{a}}(a) \quad (50.140)$$

$$= \sum_a P(x|a) \pi_{\underline{x} \Leftarrow \underline{a}}(a) \quad (50.141)$$

$$= \pi_{\underline{x}}(x) . \quad (50.142)$$

Combining Eqs.(50.139) and (50.142), we get

$$\pi_{\underline{b}_i \Leftarrow \underline{x}}(x) = \mathcal{N}(!x) \pi_{\underline{x}}(x) \prod_{k \neq i} \lambda_{\underline{b}_k \Rightarrow \underline{x}}(x) . \quad (50.143)$$

Eq.(50.143) is just RULE 2 of general BP.

50.9 BP-BB and sum-product decomposition

BP-BB yields what is often referred to as a **sum-product decomposition**. I don't like that name because it is unnecessarily confusing, and it fails to convey the recursive nature⁵ of the decomposition. I prefer to call it a **recursive sum of products (RSOP) decomposition**, and will call it so henceforth in this chapter.

Expressing the marginals of a bnet as RSOPs, which is what BP does, reduces the complexity of the calculation. (i.e., the total number of additions and multiplications that need to be performed) That makes using the BP algo very advantageous. For instance, consider a Markov chain $\underline{x}_{n-1} \leftarrow \dots \leftarrow \underline{x}_1 \leftarrow \underline{x}_0$, where $x_i \in \{0, 1, 2\}$ for all i . Note that if we calculate $P(x_{n-1})$ as follows

$$P(x_{n-1}) = \left[\sum_{x_{n-2}} P(x_{n-1}|x_{n-2}) \dots \left[\sum_{x_1} P(x_2|x_1) \left[\sum_{x_0} P(x_1|x_0)P(x_0) \right] \dots \right] \right], \quad (50.144)$$

we need to perform $2(n - 1)$ additions and $3(n - 1)$ multiplications. On the other hand, if we calculate $P(x_{n-1})$ as follows

$$P(x_{n-1}) = \sum_{x_{n-2}} \dots \sum_{x_1} \sum_{x_0} P(x_{n-1}|x_{n-2}) \dots P(x_2|x_1)P(x_1|x_0)P(x_0), \quad (50.145)$$

we need to perform $3^n - 1$ additions and $3^n(n - 1)$ multiplications.

⁵By "recursive nature", we mean bootstrapped definitions that lead to nested sums. The recursive nature of BP is evident from RULES 1 and 2 that define λ 's and π 's in terms of other λ 's and π 's.

Chapter 51

Message Passing (Belief Propagation) in Quantum Mechanics

See Ref.[68].

Chapter 52

Meta-learners for estimating ATE

This section is based on the final 2 chapters of Ref.[10].

The Average Treatment Effect (ATE) is defined in Chapter 64.

Economists are huge fans of Linear Regression (LR), and traditionally calculate ATE using LR. But in recent times, they have begun to calculate ATE using Machine Learning (ML) instead. This chapter describes various methods that economists and others have devised for calculating ATE with ML. These methods are called **meta-learners** because they involve multiple ML or LR steps.

Using ML to calculate ATE captures non-linear trends whose exclusion might sometimes lead to a poor result. On the other hand, ML is more expensive computationally than LR, and it introduces the danger of overfitting, a danger which is nonexistent with LR.

Below, we represent each Linear Regression (LR) step as follows. We list a dataset; i.e., a set of tuples indexed by the individuals σ of a population Σ such that $|\Sigma| = nsam$. The independent variables of the LR (i.e., x^σ) are unboxed and the dependent variable (a.k.a. target feature) (i.e., y^σ) is shown inside a box. Then we show an arrow with the superscript “LR-fit”, followed by the fit function obtained by performing the LR.¹

$$\{(\sigma, x^\sigma = [x_i^\sigma], \boxed{y^\sigma}) : \sigma \in \Sigma\} \xrightarrow{\text{LR-fit}} \hat{y}(x) = \alpha + x_i \beta_i \quad (52.1)$$

Analogously, below, we represent each Supervised Machine Learning (ML) step as follows.

$$\{(\sigma, x^\sigma, \boxed{y^\sigma}) : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \hat{y}(x) \quad (52.2)$$

Henceforth, we will use $\delta(x) (\approx \text{ATE})$ to denote the treatment effect.

- **S (Single)-learner**

¹In this chapter, we will occasionally use the Einstein summation convention; i.e., implicit sum over repeated indices.

$$\{(\sigma, x^\sigma, d^\sigma, \boxed{y^\sigma}) : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \hat{y}(d, x) \quad (52.3)$$

$$\delta(x) = \hat{y}(1, x) - \hat{y}(0, x) \quad (52.4)$$

- **T (Twin)-learner**

$$\{(\sigma, x^\sigma, d^\sigma = 0, \boxed{y^\sigma}) : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \hat{y}_0(x) \quad (52.5)$$

$$\{(\sigma, x^\sigma, d^\sigma = 1, \boxed{y^\sigma}) : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \hat{y}_1(x) \quad (52.6)$$

$$\delta(x) = \hat{y}_1(x) - \hat{y}_0(x) \quad (52.7)$$

- **X (Cross)-learner**

do T-learner, get $\hat{y}_0(x), \hat{y}_1(x)$

$$\{(\sigma, x^\sigma, d^\sigma = 0, \boxed{y^\sigma - \hat{y}_1(x^\sigma)}) : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \mathcal{Y}_0(x) \quad (52.8)$$

$$\{(\sigma, x^\sigma, d^\sigma = 1, \boxed{y^\sigma - \hat{y}_0(x^\sigma)}) : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \mathcal{Y}_1(x) \quad (52.9)$$

$$\delta(x) = \frac{1}{2}[\mathcal{Y}_1(x) - \mathcal{Y}_0(x)] \quad (52.10)$$

- **De-biased (a.k.a. Orthogonal) ML**

Standard supervised ML is performed with two features, the independent feature and the dependent or target feature. Supervised ML has a target feature. Unsupervised ML doesn't.

In **De-biased LR**, we do LR with two residuals. Let's call them the **independent residual** and the **dependent or target residual**. These two residuals are calculated with the help of two previously performed LR steps.

In **De-biased ML**, we do ML or LR (either one) with two residuals. These two residuals are calculated with the help of two previously performed ML steps (instead of two LR steps).

The FWL theorem discussed in Chapter 27 shows how to do De-biased LR. Next, we will describe how to do De-biased ML.

We start by doing two ML steps:

$$\{(\sigma, x^\sigma, \boxed{d^\sigma}) : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \hat{d}(x) \quad (52.11)$$

$$\{(\sigma, x^\sigma, \boxed{y^\sigma}) : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \hat{y}(x) \quad (52.12)$$

After these two ML steps, we do either LR or ML to get $\delta(x)$.

Let

$$\Delta d^\sigma = d^\sigma - \hat{d}(x^\sigma) \quad (52.13)$$

$$\Delta y^\sigma = y^\sigma - \hat{y}(x^\sigma) \quad (52.14)$$

Options for the final learning step that calculates $\delta(x)$:

1. Do a standard LR to get a $\delta(x)$ that is a constant (i.e., x independent):

$$\{(\sigma, \Delta d^\sigma, \boxed{\Delta y^\sigma}) : \sigma \in \Sigma\} \xrightarrow{\text{LR-fit}} \mathcal{Y}(\Delta d) = \alpha + \Delta d \delta \quad (52.15)$$

$$\delta = \text{coefficient of } \Delta d \text{ in } \mathcal{Y}(\Delta d). \quad (52.16)$$

2. Do a LR with a $x * d$ cross term to get a $\delta(x)$ that is linear in x :

$$\{(\sigma, x^\sigma, \Delta d^\sigma, \boxed{\Delta y^\sigma}) : \sigma \in \Sigma\} \xrightarrow{\text{LR-fit}} \mathcal{Y}(\Delta d, x) = \alpha_0 + x\beta_0 + \Delta d(\alpha_1 + x\beta_1) \quad (52.17)$$

$$\delta(x) = \mathcal{Y}(\Delta d|_{d=1}, x) - \mathcal{Y}(\Delta d|_{d=0}, x) = \alpha_1 + x\beta_1 \quad (52.18)$$

3. Do ML with a weighted cost function to get a general fit $\hat{\delta}(x)$.

The cost function used in standard ML (see Eq.(52.2)) is:

$$\mathcal{C} = \frac{1}{nsam} \sum_{\sigma} [y^\sigma - \hat{y}(x^\sigma)]^2 \quad (52.19)$$

Define the cost function in this case as

$$\mathcal{C} = \frac{1}{nsam} \sum_{\sigma} \left[\Delta y^\sigma - \hat{\delta}(x^\sigma) \Delta d^\sigma \right]^2 \quad (52.20)$$

$$= \frac{1}{nsam} \sum_{\sigma} [\Delta d^\sigma]^2 \left[\frac{\Delta y^\sigma}{\Delta d^\sigma} - \hat{\delta}(x^\sigma) \right]^2 \quad (52.21)$$

This is a weighted cost function with weights $[\Delta d^\sigma]^2$.

$$\{(\sigma, x^\sigma, \boxed{\frac{\Delta y^\sigma}{\Delta d^\sigma}}) : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \hat{\delta}(x) \quad (52.22)$$

4. Do ML with Δd^σ as an independent feature to get a general fit $\hat{\delta}(\Delta d, x)$:

$$\{(\sigma, x^\sigma, \Delta d^\sigma, \boxed{\frac{\Delta y^\sigma}{\Delta d^\sigma}}) : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \hat{\delta}(\Delta d, x) \quad (52.23)$$

$$\delta(x) = \hat{\delta}(\Delta d|_{d=1}, x) \quad (52.24)$$

Chapter 53

Missing Data, Imputation

This chapter assumes that the reader has read some parts of Chapter 25 on the Expectation Maximization (EM) algo and Chapter 46 on Markov Chain Monte Carlo (MCMC).

	h_0	x_0	x_1	x_2	
1	NA	0	1	1	
2	NA	0	0	0	
3	NA	1	1	0	
4	NA	NA	1	NA	
5	NA	0	NA	1	
6	NA	0	0	1	

	h_0	x_0	x_1	x_2	m
1	NA	0	1	1	(0,0,0)
2	NA	0	0	0	(0,0,0)
3	NA	1	1	0	(0,0,0)
4	NA	0		0	
		0		1	
		1		0	(1,0,1)
		1		1	
5	NA	0	0	1	(0,1,0)
6	NA	0	0	1	(0,0,0)

Table 53.1: **Left Table:** Dataset with $nsam = 6$ and some missing entries, for 4 binary variables h_0, x_0, x_1, x_2 . NA=not available. The h_0 column is completely missing because h_0 is an unobserved variable. **Right Table:** All possibilities for $x_i = NA$ cells of left table have been enumerated. A new column labeled m has been added. $m_i = \mathbb{1}(x_i \text{ is missing})$ for $i = 0, 1, 2$.

Suppose that you have compiled a **dataset** $\vec{x} = (x[\sigma])_{\sigma=0,1,\dots,nsam-1}$ where $x = (x_0, x_1, \dots, x_{nx-1})$ from a study or survey. It consists of $nsam$ number of samples (sample= row), and nx columns (each column is a different feature, or observation). Suppose that some of the cells in this matrix are empty. Throwing away all the incomplete rows is okay if the number of incomplete rows is much smaller than $nsam$. If not, throwing them away would throw away a substantial amount of information contained in all the filled cells in those incomplete rows, plus it might bias your dataset. This chapter deals with how to fill those empty cells with plausible fake data. A fancy name for this process is **imputation**. There is no unique way of

fabricating fake data, but some fakes are better than others by some metrics. This chapter will consider two popular ways (EM and MCMC) of filling those empty cells with their “most likely” values based on the cells of the dataset that aren’t missing, and also based on some bnet model that is expected to describe well the dataset.

Notation: $\vec{a} = (\underline{a}[\sigma])_{\sigma=0,1,\dots,nsam-1}$, where $nsam$ is the number of samples. Will sometimes denote $a[\sigma]$ by a^σ .

For concreteness, we will apply the concepts of this chapter to the dataset with missing data given by Table 53.1.

53.1 Imputation via EM

We begin by augmenting Fig.25.1 (the first figure in Chapter 25) by adding to it a new node \vec{m} called the **missingness variable**. Recall that node θ represents the **unknown parameters**, node \vec{x} represents the **observed variables**, and node \vec{h} represents the **hidden variables**. Both θ and \vec{h} are hidden (i.e., unobserved). Fig.53.1 shows 3 popular ways of connecting node \vec{m} to the other nodes in the graph Fig.25.1.

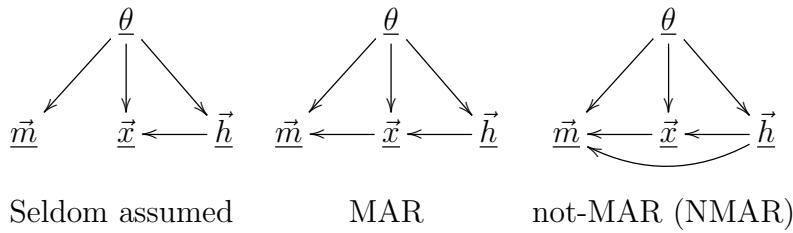


Figure 53.1: The left bnet is seldom assumed. The middle bnet is referred to as the MAR (missing at random) assumption. The right bnet is referred to as the not-MAR (NMAR) assumption.

From Fig.53.1, we have

$$P(\vec{m}|\vec{x}, \vec{h}, \theta) = \begin{cases} P(\vec{m}|\theta) & \text{Seldom assumed. Called missing-CAR (MCAR)} \\ P(\vec{m}|\vec{x}, \theta) & \text{MAR} \\ P(\vec{m}|\vec{x}, \vec{h}, \theta) & \text{not-MAR (NMAR)} \end{cases}. \quad (53.1)$$

For doing imputation via EM, we connect node \vec{m} as shown in the middle bnet (called MAR) of Fig.53.1.

For the example of Table 53.1, we have variables \vec{m} , \vec{x} and \vec{h} whose values range over the following sets:

$$\begin{aligned} \vec{x} &= (\vec{x}_0, \vec{x}_1, \vec{x}_2) \\ \vec{h} &= (\vec{h}_0) \\ \underline{h}_0[\sigma] &\in \{0, 1\}, \end{aligned}$$

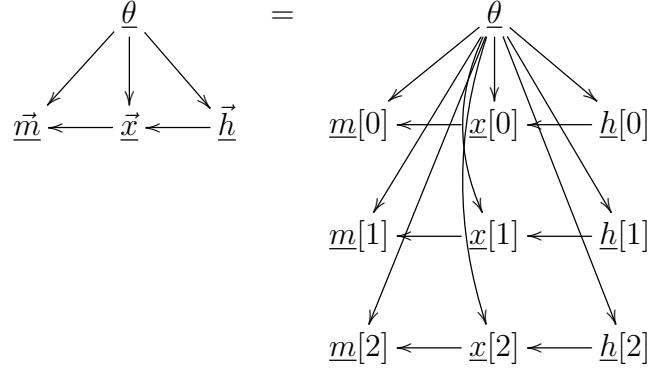


Figure 53.2: MAR bnet with $nsam = 3$.

$$\begin{aligned} \underline{x}_i[\sigma] &\in \{0, 1\} \text{ for } i = 0, 1, 2, \\ \underline{m}_i[\sigma] &\in \{0, 1\} \text{ for } i = 0, 1, 2. \end{aligned}$$

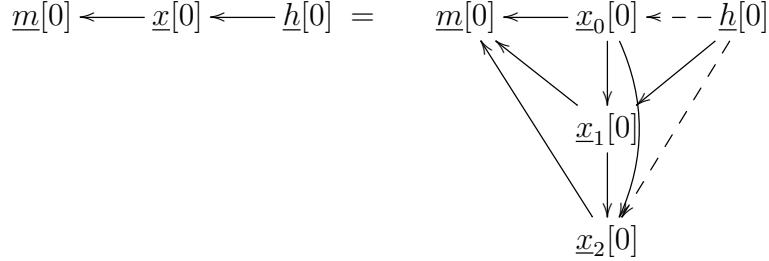


Figure 53.3: Our example for imputation via EM assumes this bnet between nodes $m[\sigma], \underline{x}[\sigma], \underline{h}[\sigma]$.

For concreteness, we will assume that the Markov chain $\underline{m}[\sigma] \leftarrow \underline{x}[\sigma] \leftarrow \underline{h}[\sigma]$ has a finer grained DAG structure given by Fig.53.3. where we will omit the dashed arrows. If one doesn't want to assume that the data can be fitted well by the bnet of Fig.53.3 without the dashed arrows, one can include those arrows too, at the expense of more unknown parameters (i.e., degrees of freedom) to be lumped into θ . We will parameterize the TPMs corresponding to Fig.53.3 using a Categorical Distribution for each column of the TPMs. We will thus assume that the TPMs, printed in blue, for bnet Fig.53.3, are as follows.

$$P(h_0^\sigma | \theta) = \frac{1 - \theta_0}{1 + \theta_0} \quad (53.2)$$

$$P(x_0^\sigma | \theta) = \begin{array}{c|cc} & & \\ & 0 & 1 - \theta_1 \\ \hline 0 & & \\ & 1 & \theta_1 \end{array} \quad (53.3)$$

$$P(x_1^\sigma | x_0^\sigma, h^\sigma, \theta) = \begin{array}{c|cccc} & 00 & 01 & 10 & 11 \\ \hline 0 & 1 - \theta_2 & 1 - \theta_3 & 1 - \theta_4 & 1 - \theta_5 \\ 1 & \theta_2 & \theta_3 & \theta_4 & \theta_5 \end{array} \quad (53.4)$$

$$P(x_2^\sigma | x_1^\sigma, x_0^\sigma, h^\sigma, \theta) = \begin{array}{c|cccc} & 00 & 01 & 10 & 11 \\ \hline 0 & 1 - \theta_6 & 1 - \theta_7 & 1 - \theta_8 & 1 - \theta_9 \\ 1 & \theta_6 & \theta_7 & \theta_8 & \theta_9 \end{array} \quad (53.5)$$

$$P(m^\sigma | x^\sigma, \theta) = \frac{1}{nsam} P((x_i)_{\forall i \ni m_i=1} | (x_i)_{\forall i \ni m_i=0}, \theta) \quad (53.6)$$

Eq.(53.6) can be illustrated as follows. In Table 53.2, we added a $P(m)$ column to Table 53.1.

	h_0	x_0	x_1	x_2	m	$P(m)$
1	NA	0	1	1	(0,0,0)	$\frac{1}{nsam}$
2	NA	0	0	0	(0,0,0)	$\frac{1}{nsam}$
3	NA	1	1	0	(0,0,0)	$\frac{1}{nsam}$
4	NA	0		0		$\frac{1}{nsam} P(x_0 = 0, x_2 = 0 x_1 = 1, \theta)$
		0		1		$\frac{1}{nsam} P(x_0 = 0, x_2 = 1 x_1 = 1, \theta)$
		1		0	(1,0,1)	$\frac{1}{nsam} P(x_0 = 1, x_2 = 0 x_1 = 1, \theta)$
		1		1		$\frac{1}{nsam} P(x_0 = 1, x_2 = 1 x_1 = 1, \theta)$
5	NA	0	0			$\frac{1}{nsam} P(x_1 = 0 x_0 = 0, x_2 = 1, \theta)$
		1		1	(0,1,0)	$\frac{1}{nsam} P(x_1 = 1 x_0 = 0, x_2 = 1, \theta)$
6	NA	0	0	1	(0,0,0)	$\frac{1}{nsam}$

Table 53.2: $P(m)$ column added to Table 53.1. Note that $\sum_m P(m) = 1$.

$$\theta = (\theta_i)_{i=0,1,\dots,9} \quad (53.7)$$

$$P(m^\sigma, x^\sigma, h^\sigma | \theta) = P(m^\sigma | x^\sigma, \theta) P(x^\sigma | h^\sigma, \theta) P(h^\sigma | \theta) \quad (53.8)$$

$$P(x^\sigma | h^\sigma, \theta) = P(x_2^\sigma | x_1^\sigma, x_0^\sigma, \theta) P(x_1^\sigma | x_0^\sigma, h^\sigma, \theta) P(x_0^\sigma | \theta) \quad (53.9)$$

$$P(x_1^\sigma | x_0^\sigma, \theta) = \sum_h P(x_1^\sigma | x_0^\sigma, h^\sigma, \theta) P(h^\sigma | \theta) \quad (53.10)$$

$$P(x^\sigma|\theta) = P(x_2^\sigma|x_1^\sigma, x_0^\sigma, \theta)P(x_1^\sigma|x_0^\sigma, \theta)P(x_0^\sigma|\theta) \quad (53.11)$$

$$Q(\theta|\theta^{(t)}) = \sum_{\vec{m}, \vec{h}} P(\vec{m}, \vec{h} | \vec{x}, \theta^{(t)}) \ln P(\vec{m}, \vec{x}, \vec{h}|\theta) \quad (53.12)$$

$$= \sum_{\vec{m}, \vec{h}} \left[\prod_{\sigma} P(m^\sigma, h^\sigma | x^\sigma, \theta^{(t)}) \right] \ln \left[\prod_{\sigma} P(m^\sigma, x^\sigma, h^\sigma|\theta) \right] \quad (53.13)$$

$$= \sum_{\sigma} \sum_{m^\sigma, h^\sigma} P(m^\sigma, h^\sigma | x^\sigma, \theta^{(t)}) \ln P(m^\sigma, x^\sigma, h^\sigma|\theta) \quad (53.14)$$

$$= \sum_{\sigma} \sum_{m^\sigma, h^\sigma} \frac{P(m^\sigma, h^\sigma, x^\sigma | \theta^{(t)})}{P(x^\sigma | \theta^{(t)})} \ln P(m^\sigma, x^\sigma, h^\sigma|\theta) \quad (53.15)$$

Once you find optimal parameters θ^* by recursing this $Q(\theta|\theta^{(t)})$, you can evaluate numerically the $P(m)$ column of Table 53.2. In Table 53.2, out of the 4 sub-rows for row 4, choose the one with the highest probability. Similarly, out of the 2 sub-rows for row 5, choose the one with the highest probability.

53.2 Imputation via MCMC

A simple and popular way to do imputation via MCMC is described in Ref.[61]. It goes as follows.

Let

$$\underline{H}[\sigma] = (\underline{h}[\sigma], \underline{m}[\sigma]) \quad (53.16)$$

for $\sigma = 0, 1, \dots, nsam - 1$. Initialize $\theta^{(0)}$ to a random value within the allowed ranges. Do the following 2 steps, for $t = 0, 1, \dots, T - 1$, where T is large enough that $\theta^{(t)}$ has reached a steady value that is independent of $\theta^{(0)}$. To do the sampling, use a standard sampling technique such as Gibbs sampling.

- **STEP 1:** For $\sigma = 0, 1, \dots, nsam - 1$, find a sample

$$(H^\sigma)^{(t+1)} \sim P(H^\sigma | x^\sigma, \theta^{(t)}) . \quad (53.17a)$$

- **STEP 2:** Find a sample

$$\theta^{(t+1)} \sim P^{(t+1)}(\theta) \quad (53.17b)$$

where

$$P^{(t+1)}(\theta) = \mathcal{N}(!\theta)P(\vec{x}, \vec{H}^{(t+1)}|\theta) \quad (53.17c)$$

$$= \mathcal{N}(!\theta) \prod_{\sigma} P(x^\sigma, (H^\sigma)^{(t+1)}|\theta) . \quad (53.17d)$$

Fig.53.4 illustrates this two step recursive process using a bnet.

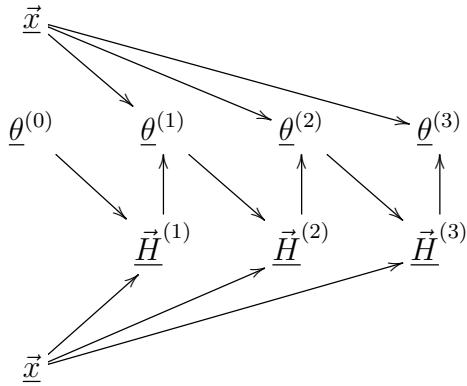


Figure 53.4: bnet illustrating Eqs.(53.17) for doing imputation via MCMC. The **same** node \underline{x} appears twice to make the graph clearer.

53.3 Multiple Imputations

Multiple imputations means calculating θ^* (i.e., the optimum θ) and the concomitant dataset \vec{x}^*, \vec{H}^* , via any method (such as EM or MCMC), a large number of times, starting from different, randomly chosen $\theta^{(0)}$ initial parameters. Then calculating the average and the variance of $\theta^*, \vec{x}^*, \vec{H}^*$ and functions thereof.

Chapter 54

Modified Treatment Policy

This chapter is based primarily on Ref. [17]. Refs. [19] and [20] were also very helpful.

A **Modified Treatment Policy (MTP)** is a generalization of the Potential Outcomes (PO)¹ model so as to modify the original treatment. This is accomplished by adding a modified treatment dose node \tilde{x} acting as a mediator between the treatment dose node x and the treatment effect node y (i.e., by considering $x \rightarrow \tilde{x} \rightarrow y$).

54.1 One time MTP

Consider a typical PO bnet G and its corresponding imagined bnet G_{im2} (see Fig.54.1)². A MTP adds a Bayesian prior to the node \tilde{x} in G_{im2} . The prior depends on the treatment dose variable (a.k.a. exposure variable) x and its parents $pa(x) = c$. This adds to G_{im2} new arrows $c \rightarrow \tilde{x}$ and $x \rightarrow \tilde{x}$ (see Fig.54.2).

The TPM, printed in blue, for node \tilde{x} of the imagined bnet G_{im2} in Fig.54.2, is as follows

$$P(\tilde{x} = \tilde{x}; x) = \delta(\tilde{x}, x) \quad (54.1)$$

The TPM, printed in blue, for node \tilde{x} of the modified imagined bnet $G_{im2,mod}$ in Fig.54.2, is as follows

$$P(\tilde{x} = \tilde{x}|x = x, c) = \text{prior} \quad (54.2)$$

Hence, node \tilde{x} has a totally informative, parametric (frequentist) prior in G_{im2} , and it has a more general (Bayesian) prior in $G_{im2,mod}$.

The following assumptions will be made about bnet $G_{im2,mod}$:

1. Consistency (a.k.a. SUTVA)

¹PO theory is discussed in Chapter 64.

²Ref.[17] uses the notation $L = c$, $A = x$, $Y = y$, $Q = \tilde{x}$, $Y_q = y(\tilde{x})$. Furthermore, it does not display a DAG.

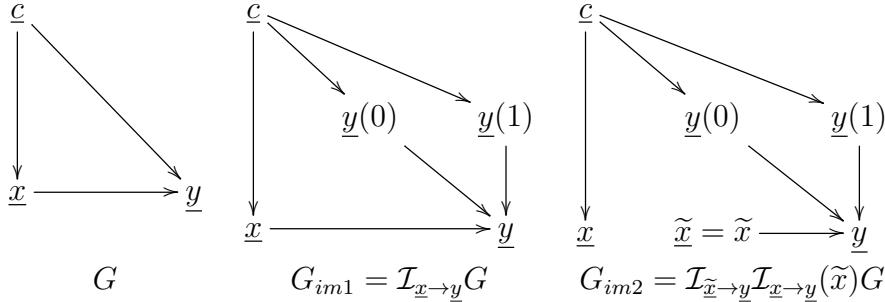


Figure 54.1: G is one of the simplest possible bnets considered in PO theory. In the usual PO theory, one only considers the bnet G_{im1} . In MTP theory, one considers the bnet G_{im2} . For G_{im2} , $\tilde{\underline{x}} \in S_{\tilde{\underline{x}}} = S_{\underline{x}} = \{0, 1\}$.
 (As usual in this book, $S_{\underline{a}}$ denotes the set of values that a random variable \underline{a} can assume. Imagine operators such as $\mathcal{I}_{\underline{x} \rightarrow \underline{y}}$ and $\mathcal{I}_{\underline{x} \rightarrow \underline{y}}(\tilde{\underline{x}})$ are discussed in Chapter 11.)

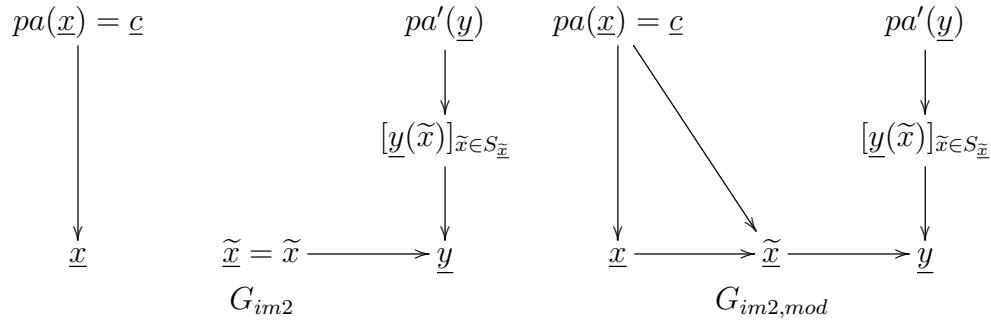


Figure 54.2: In this figure, G_{im2} is the generalization of the G_{im2} in Fig.54.1. Note that $pa'(\underline{y})$ are the parents in G of \underline{y} , excluding \underline{x} . Note that $S_{\tilde{\underline{x}}} \subset S_{\underline{x}}$.

The TPM, printed in blue, for node \underline{y} of the modified imagined bnet $G_{im2,mod}$ in Fig.54.2, is as follows

$$P(y | \tilde{x}, \{y(\tilde{x}')\}_{\tilde{x}' \in S_{\tilde{x}}}) = \mathbb{1}(y = \sum_{\tilde{x}' \in S_{\tilde{x}}} y(\tilde{x}') \mathbb{1}(\tilde{x}' = \tilde{x})) \quad (54.3)$$

When $S_{\tilde{\underline{x}}} = \{0, 1\}$, this becomes the more familiar statement from standard PO theory:

$$P(y | \tilde{x}, \{y(0), y(1)\}) = \mathbb{1}(y = y(0)(1 - \tilde{x}) + y(1)\tilde{x}) \quad (54.4)$$

An immediate consequence of this deterministic TPM for node \underline{y} is that

$$P(y = y | \tilde{x}, pa'(\underline{y}) = \xi) = P(y(\tilde{x}) = y | \tilde{x}, pa'(\underline{y}) = \xi) \quad (54.5)$$

2. Ignorability (a.k.a. conditional independence assumption (CIA))

$$P(\underline{y}(\tilde{x}') = y | \tilde{x}, pa'(\underline{y}) = \xi) = P(\underline{y}(\tilde{x}') = y | pa'(\underline{y}) = \xi) \quad (54.6)$$

CIA follows from the structure of the DAG $G_{im2,mod}$, because in that DAG, fixing the value of $pa'(\underline{y})$ blocks messages from \underline{x} to $\underline{y}(\tilde{x})$ (i.e., for all $\tilde{x} \in S_{\tilde{x}}$, we have $\tilde{x} \perp \underline{y}(\tilde{x}) | pa'(\underline{y})$.)

3. Identifiability

We must have

$$S_{\tilde{x}} \subset S_{\underline{x}} \quad (54.7)$$

in $G_{im2,mod}$ or else queries of the type $P(\underline{y}(\tilde{x}) = y | pa'(\underline{y}) = \xi)$ are not identifiable. This is clear because if $\tilde{x} \notin S_{\underline{x}}$, then we have no information of the type

$$P(\underline{y} = y | \tilde{x}, pa'(\underline{y}) = \xi) = P(\underline{y}(\tilde{x}) = y | pa'(\underline{y}) = \xi). \quad (54.8)$$

4. Positivity

Define the following 2 propensities for $x \in S_{\underline{x}}$ and $\tilde{x} \in S_{\tilde{x}}$:

$$g_{x|c} = P(\underline{x} = x | c) \quad (54.9)$$

$$\tilde{g}_{\tilde{x}|c} = P(\tilde{x} = \tilde{x} | c) \quad (54.10)$$

Positivity for $G_{im2,mod}$ is the requirement that

$$0 < g_{\tilde{x}|c} < 1 \text{ for all } \tilde{x} \in S_{\tilde{x}} \text{ and } c \in S_c \quad (54.11)$$

If $g_{\tilde{x}|c}$ is deterministic, then this requirement is not satisfied, and we cannot do IPW (i.e., inverse propensity weighing).

Note that using a MTP can allow IPW to be performed in cases when the propensity $g_{x|c}$ is anomalous (i.e., is either not defined or violates positivity) for some $x \in S_{\underline{x}}$, but is not anomalous for all $\tilde{x} \in S_{\tilde{x}}$, where $S_{\tilde{x}}$ is some proper subset of $S_{\underline{x}}$.

The probability distribution $P(\tilde{x}|x, c)$ (i.e., the \tilde{x} prior for $G_{im2,mod}$) is called the **MTP**. An MTP can be either deterministic or probabilistic. An MTP that depends (resp., does not depend) on x is said to be **dynamic** (resp., **static**), because \tilde{x} depends (resp., does not depend) on the previous value of x .

1. Deterministic MTP

Suppose $S_{\tilde{x}} \subset S_x$ and we are given a function $\tilde{x}_c(\cdot) : S_x \rightarrow S_{\tilde{x}}$. Then let the TPM, printed in blue, for node \tilde{x} , be as follows:

$$P(\tilde{x} = \tilde{x}|x = x, c) = \delta(\tilde{x}, \tilde{x}_c(x)) \quad (54.12)$$

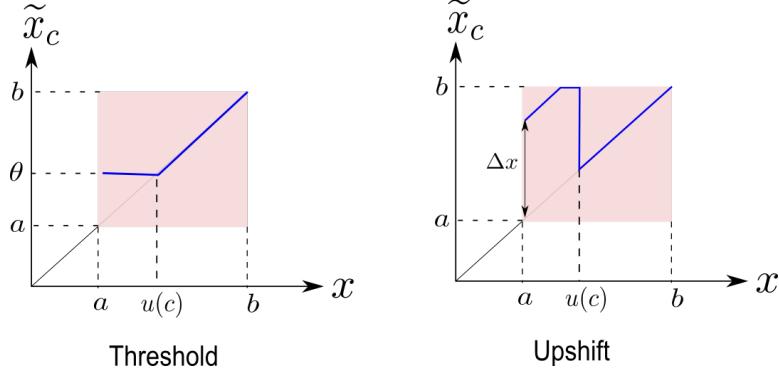


Figure 54.3: Two possible maps $\tilde{x}_c : [a, b] \rightarrow [a, b]$ for a deterministic MTP.

Examples of $\tilde{x}_c(x)$:

- threshold in the value of x for small values of x . (See Fig.54.3)

Let $S_x = [a, b]$ where $a < b$. For some $\theta \in [a, b]$ and $u(c) \in [a, b]$, set

$$\tilde{x}_c(x) = \begin{cases} \theta & \text{if } x \leq u(c) \\ x & \text{if } x > u(c) \end{cases} \quad (54.13)$$

- upshift in the value of x for small values of x . (See Fig.54.3)

Let $S_x = [a, b]$ where $a < b$. For some $\Delta x > 0$, and $u(c) \in [a, b]$, set

$$\tilde{x}_c(x) = \begin{cases} (x + \Delta x) & \text{if } x < u(c) \text{ and } (x + \Delta x) \in [a, b] \\ x & \text{if } x > u(c) \\ b & \text{if } (x + \Delta x) > b \end{cases} \quad (54.14)$$

2. Stochastic MTP

For some convenient, user specified, probability distribution $P(\tilde{x}|x', x, c)$, where $\tilde{x} \in S_{\tilde{x}}$, $x, x' \in S_x$, and $c \in S_c$, let the TPM, printed in blue, for node \tilde{x} , be as follows:

$$P(\tilde{x}|x, c) = \sum_{x' \in S_x} P(\tilde{x}|x', x, c) \underbrace{P(x = x'|c)}_{\text{unmodified propensity}} \quad (54.15)$$

Examples

- Suppose we are given a function $x_c : S_{\tilde{x}} \rightarrow S_x$. ($x_c(\tilde{x})$ could be the inverse, if it exists, of the function $\tilde{x}_c(x)$ defined in the deterministic TMP case.)

$$P(\tilde{x}|x, c) = \frac{P(\underline{x} = x_c(\tilde{x})|c)}{\sum_{\tilde{x}} \text{numerator}} \quad (54.16)$$

- Suppose $S_{\underline{x}} = S_{\tilde{x}} = \{0, 1\}$, $\theta \in \mathbb{R}$, and let

$$P(\tilde{x}|x, c) = [\pi(x|c)]^{\tilde{x}}[1 - \pi(x|c)]^{1-\tilde{x}} \quad (54.17)$$

where

$$\pi(x|c) = \frac{\theta^x P(\underline{x} = x|c)}{\sum_x \text{numerator}} \quad (54.18)$$

54.2 $\Delta_{|c}$ estimand

Let

$$\mathcal{Y}_{|x,c} = \sum_y y P(\underline{y}(x) = y | x, c) \quad (54.19)$$

$$= \sum_y y P(\underline{y} = y | x, c) \quad (54.20)$$

$$\tilde{\mathcal{Y}}_{|\tilde{x}=\tilde{x},c} = \sum_y y P(\underline{y}(\tilde{x}) = y | \tilde{x}, c) \quad (54.21)$$

$$= \sum_y y P(\underline{y} = y | \tilde{x}, c) \quad (54.22)$$

$$\tilde{\mathcal{Y}}_{|\underline{x}=x,c} = \sum_{\tilde{x}} \tilde{\mathcal{Y}}_{|\tilde{x}=\tilde{x},c} P(\tilde{x}|x, c) \quad (54.23)$$

$$\mathcal{Y}_{|c} = \sum_x \mathcal{Y}_{|x,c} P(x|c), \quad \mathcal{Y} = \sum_c P(c) \mathcal{Y}_{|c} \quad (54.24)$$

$$\tilde{\mathcal{Y}}_{|c} = \sum_x \tilde{\mathcal{Y}}_{|\underline{x}=x,c} P(x|c), \quad \tilde{\mathcal{Y}} = \sum_c P(c) \tilde{\mathcal{Y}}_{|c} \quad (54.25)$$

Note that there are two $\tilde{\mathcal{Y}}$, namely $\tilde{\mathcal{Y}}_{|\underline{x}=x,c}$ and $\tilde{\mathcal{Y}}_{|\tilde{x}=\tilde{x},c}$. Define the $\Delta_{|c}$ estimand by:

$$\Delta_{|c} = \mathcal{Y}_{|c} - \tilde{\mathcal{Y}}_{|c}, \quad \Delta = \sum_c P(c) \Delta_{|c} \quad (54.26)$$

$\Delta_{|c}$ measures the difference between the real world represented by $\mathcal{Y}_{|c}$ and a modified world represented by $\tilde{\mathcal{Y}}_{|c}$.

Claim 76 Consider the deterministic MTP case $P(\tilde{x}|x, c) = \delta(\tilde{x}, \tilde{x}_c(x))$, where the function $\tilde{x} = \tilde{x}_c(x)$ is invertible with inverse $x = x_c(\tilde{x})$ for some $x \in \mathbb{X} \subset \mathbb{R}$. Then,

$$\tilde{\mathcal{Y}}_{|\underline{x}=x,c} = \mathcal{Y}_{|x_c(\tilde{x}),c} \quad (54.27)$$

proof:

$$\tilde{\mathcal{Y}}_{|\underline{x}=x,c} = \sum_{\tilde{x}} \sum_y y P(\underline{y} = y | \underline{x} = \tilde{x}, c) \delta(\tilde{x}, \tilde{x}_c(x)) \quad (54.28)$$

$$= \sum_y y P(\underline{y} = y | \underline{x} = \tilde{x}_c(x), c) \quad (54.29)$$

$$= \sum_y y P(\underline{y} = y | \underline{x} = x_c(\tilde{x}), c) \quad (54.30)$$

$$= \mathcal{Y}_{|x_c(\tilde{x}),c} \quad (54.31)$$

QED

Henceforth in this chapter, we will restrict our attention to the deterministic MTP case in which $P(\tilde{x}|x, c) = \delta(\tilde{x}, \tilde{x}_c(x))$. We will also assume that the domain \mathbb{X}_c of $\tilde{x}_c(x)$ is a union of disjoint sets \mathbb{X}_c^j for $j = 1, 2, \dots, nj(c)$, and that on each set \mathbb{X}_c^j , $\tilde{x}_c(x)$ is invertible and differentiable.

$$\begin{cases} \mathbb{X} = S_{\underline{x}} \\ \mathbb{X}_c^j \cap \mathbb{X}_c^{j'} = \emptyset \text{ for } j \neq j' \\ \mathbb{X}_c = \cup_{j=1}^{nj(c)} \mathbb{X}_c^j \subset \mathbb{X} \end{cases} \quad (54.32)$$

$$\tilde{x}_c(x) = \sum_{j=1}^{nj(c)} \mathbb{1}(x \in \mathbb{X}_c^j) \tilde{x}_c^j(x) \quad (54.33)$$

$$P(\tilde{x}|x, c) = \delta(\tilde{x}, \tilde{x}_c(x)) = \sum_j \mathbb{1}(x \in \mathbb{X}_c^j) \delta(\tilde{x}, \tilde{x}_c^j(x)) \quad (54.34)$$

$$(\tilde{x}_c^j)^{-1}(\tilde{x}) = x^j(\tilde{x}) \quad (54.35)$$

$$\tilde{\mathbb{X}}_c^j = \tilde{x}_c^j(\mathbb{X}_c^j) = \{\tilde{x}_c^j(x) : x \in \mathbb{X}_c^j\} \quad (54.36)$$

$$P_{\mathbb{X}_c} = P(\underline{x} \in \mathbb{X}_c) \quad (54.37)$$

Claim 77

$$\tilde{\mathcal{Y}}_{|c} = \frac{1}{P_{\mathbb{X}_c}} E_{\tilde{x}|c} [\lambda_c(\tilde{x}) \mathcal{Y}_{|x_c(\tilde{x}),c}] , \quad (54.38)$$

$$\mathcal{Y}_{|c} = \frac{1}{P_{\mathbb{X}_c}} E_{\tilde{x}|c} [\lambda_c(\tilde{x}) \mathcal{Y}_{|\tilde{x},c}] \quad (54.39)$$

and

$$P_{\mathbb{X}_c} = E_{\tilde{x}|c} [\lambda_c(\tilde{x})] \quad (54.40)$$

where³

$$\lambda_c(\tilde{x}) = \sum_j \mathbb{1}(\tilde{x} \in \tilde{\mathbb{X}}_c^j) \frac{dx_c^j(\tilde{x})}{d\tilde{x}} \underbrace{\frac{P(\underline{x} = x_c^j(\tilde{x})|c)}{P(\underline{x} = \tilde{x}|c)}}_{=1} \quad (54.41)$$

proof:

$$\tilde{\mathcal{Y}}_{|c} = \frac{1}{P_{\mathbb{X}_c}} \int_{x \in \mathbb{X}_c} dx P(x|c) \sum_y y P(y|\underline{x} = \tilde{x}_c(x), c) \quad (54.42)$$

$$= \frac{1}{P_{\mathbb{X}_c}} \sum_j \int_{x \in \mathbb{X}_c^j} dx P(x|c) \sum_y y P(y|\underline{x} = \tilde{x}_c^j(x), c) \quad (54.43)$$

$$= \frac{1}{P_{\mathbb{X}_c}} \sum_j \int_{\tilde{x} \in \tilde{\mathbb{X}}_c^j} d\tilde{x} \frac{dx_c^j(\tilde{x})}{d\tilde{x}} P(\tilde{x}|c) \underbrace{\frac{P(\underline{x} = x_c^j(\tilde{x})|c)}{P(\tilde{x}|c)}}_{\mathcal{Y}_{|x_c^j(\tilde{x}),c}} \sum_y y P(y|\underline{x} = x_c^j(\tilde{x}), c) \quad (54.44)$$

$$= \frac{1}{P_{\mathbb{X}_c}} \int_{\tilde{x} \in \tilde{\mathbb{X}}_c} d\tilde{x} P(\tilde{x}|c) \lambda_c(\tilde{x}) \mathcal{Y}_{|x_c(\tilde{x}),c} \quad (\text{see Eq(54.41)}) \quad (54.45)$$

$$= \frac{1}{P_{\mathbb{X}_c}} E_{\tilde{x}|c} [\lambda_c(\tilde{x}) \mathcal{Y}_{|x_c(\tilde{x}),c}] \quad (54.46)$$

To get Eq.(54.39), replace $\mathcal{Y}_{|x_c(\tilde{x}),c}$ by $\mathcal{Y}_{|\tilde{x},c}$ in Eq.(54.38).

To get Eq.(54.40), replace $\mathcal{Y}_{|x_c(\tilde{x}),c}$ by 1 in Eq.(54.38).

QED

From the last claim, it follows that

$$\Delta_{|c} = \frac{E_{y,\tilde{x}|c} [\lambda_c(\tilde{x})(y - \mathcal{Y}_{|x_c(\tilde{x}),c})]}{E_{\tilde{x}|c} [\lambda_c(\tilde{x})]} \quad (54.47)$$

³ $\lambda_c(\tilde{x})$ is just a piecewise Jacobian. In Ref.[17], the part of Eq.(54.41) which is marked as being equal to 1 is incorrectly assumed to be different from 1.

54.3 Estimates of $\Delta_{|c}$

54.3.1 Empirical estimate of $\Delta_{|c}$

Consider a population Σ of individuals $\sigma \in \Sigma$, with $|\Sigma| = N$.

Let

$$N_c = \sum_{\sigma} \delta(c_{\sigma}, c) , \quad (54.48)$$

$$P(\sigma|c) = \frac{\delta(c_{\sigma}, c)}{N_c} , \quad (54.49)$$

and

$$E_{\sigma|c}[\xi_{\sigma}] = \sum_{\sigma} P(\sigma|c) \xi_{\sigma} . \quad (54.50)$$

Then set

$$\widehat{\Delta}_{|c} = \frac{E_{\sigma|c} [\lambda_c(\tilde{x}_{\sigma})(y_{\sigma} - \mathcal{Y}_{|x_c(\tilde{x}_{\sigma}), c})]}{E_{\sigma|c} [\lambda_c(\tilde{x}_{\sigma})]} \quad (54.51)$$

54.3.2 OR estimate of $\Delta_{|c}$

Outcome Regression (OR) estimate of $\Delta_{|c}$.

Calculate the following estimates in this order: $\widehat{\tau} \rightarrow \widehat{\Delta}_{|c}$.

Steps:

1. Calculate $\widehat{\tau}$

Let

$$X^T = [1, \tilde{x}, c], \quad X_{\sigma}^T = [1, \tilde{x}_{\sigma}, c_{\sigma}] . \quad (54.52)$$

Use Generalized Linear Modeling (GLM)⁴ to approximate y_{σ} :

$$y_{\sigma} \approx \widehat{y}(X_{\sigma}^T \widehat{\tau}) \quad (54.53)$$

$$\widehat{y}(X_{\sigma}^T \widehat{\tau}) = g^{-1}(X_{\sigma}^T \widehat{\tau}) \quad (54.54)$$

where $g()$ is the link function.

2. Calculate $\widehat{\Delta}_{|c}$

$$\widehat{\Delta}_{|c} = \frac{E_{\sigma|c} [\lambda_c(\tilde{x}_{\sigma})(y_{\sigma} - \widehat{y}(X_{\sigma}^T \widehat{\tau}))]}{E_{\sigma|c} [\lambda_c(\tilde{x}_{\sigma})]} \quad (54.55)$$

⁴GLM is discussed in Chapter 31.

Claim 78 (*Asymptotic behavior of OR estimate*)

As $N \rightarrow \infty$,

$$\sqrt{N_c}(\widehat{\Delta}_{|c} - \Delta_{|c}^*) \rightarrow \mathcal{N}(0, \mathcal{V}) \quad (54.56)$$

where

$$\mathcal{V} = E_{\sigma|c}[(IF_\sigma)^2] \quad (54.57)$$

$$IF_\sigma = \frac{R_\sigma(\widehat{\tau}, \widehat{\Delta}_{|c}) + B(\widehat{\tau})}{C} \quad (54.58)$$

$$R_\sigma(\tau, \Delta_{|c}) = \lambda_c(\tilde{x}_\sigma)(y_\sigma - \widehat{y}(X_\sigma^T \tau) - \Delta_{|c}) \quad (54.59)$$

$$C = -E_{\sigma|c}[\lambda_c(\tilde{x}_\sigma)] \quad (54.60)$$

$$B(\tau) = \sum_i E_{\sigma|c} \left[\lambda_c(\tilde{x}_\sigma) \frac{\partial \widehat{y}(X_\sigma^T \tau)}{\partial \tau_i} \right] \frac{E_{\sigma|c}[S_{i,\sigma}(\tau)]}{E_{\sigma|c} \left[\frac{\partial S_{i,\sigma}(\tau)}{\partial \tau_i} \right]} \quad (54.61)$$

(Note that index i labels the components of the column vector τ)

proof:

Eq.(54.55) that defines the OR estimate can be rewritten as

$$0 = E_{\sigma|c} \underbrace{[\lambda_c(\tilde{x}_\sigma)(y_\sigma - \widehat{y}(X_\sigma^T \widehat{\tau}) - \widehat{\Delta}_{|c})]}_{R_\sigma(\widehat{\tau}, \widehat{\Delta}_{|c})} \quad (54.62)$$

If we Taylor expand $R_\sigma(\widehat{\tau}, \widehat{\Delta}_{|c})$ to first order in each of its 2 estimator arguments, we get

$$E_{\sigma|c}[R_\sigma(\widehat{\tau}, \widehat{\Delta}_{|c})] = \begin{cases} \underbrace{E_{\sigma|c}[R_\sigma(\tau^*, \Delta_{|c}^*)]}_A \\ + \sum_i \underbrace{E_{\sigma|c} \left[\frac{\partial R_\sigma(\tau^*, \Delta_{|c}^*)}{\partial \tau_i^*} \right]}_{B_i} (\widehat{\tau}_i - \tau_i^*) \\ + \underbrace{E_{\sigma|c} \left[\frac{\partial R_\sigma(\tau^*, \Delta_{|c}^*)}{\partial \Delta_{|c}^*} \right]}_{-C} (\widehat{\Delta}_{|c} - \Delta_{|c}^*) \end{cases} \quad (54.63)$$

Solving the last equation for $\widehat{\Delta}_{|c} - \Delta_{|c}^*$ yields

$$\widehat{\Delta}_{|c} - \Delta_{|c}^* = \frac{A + \overbrace{\sum_i^B B_i(\widehat{\tau}_i - \tau_i^*)}^{C}}{C} \quad (54.64)$$

Assuming $y_\sigma \in \{0, 1\}$, define the Cross Entropy and its first derivative with respect to τ_i by

$$CE_\sigma = \sum_{y_\sigma=0,1} y_\sigma \ln \widehat{y}(X_\sigma^T \tau) \quad (54.65)$$

$$S_{i,\sigma}(\tau) = \frac{\partial CE_\sigma}{\partial \tau_i} \quad (54.66)$$

In this notation, according to the maximum likelihood principle, the best choice of parameters τ_i must satisfy

$$0 = E_{\sigma|c}[S_{i,\sigma}(\widehat{\tau})] \quad (54.67)$$

If we Taylor expand $S_{i,\sigma}(\widehat{\tau})$ to first order in its estimator argument, we get

$$E_{\sigma|c}[S_{i,\sigma}(\widehat{\tau})] = E_{\sigma|c}[S_{i,\sigma}(\tau^*)] + E_{\sigma|c}\left[\frac{\partial S_{i,\sigma}(\tau^*)}{\partial \tau_i}\right](\widehat{\tau}_i - \tau_i^*) \quad (54.68)$$

Hence,

$$\widehat{\tau}_i - \tau_i^* = -\frac{E_{\sigma|c}[S_{i,\sigma}(\tau^*)]}{E_{\sigma|c}\left[\frac{\partial S_{i,\sigma}(\tau^*)}{\partial \tau_i}\right]}. \quad (54.69)$$

Now putting our previous results together, we get

$$A = E_{\sigma|c}[R_\sigma(\tau^*, \Delta_{|c}^*)] \quad (54.70)$$

$$= E_{\sigma|c}[\lambda_c(\tilde{x}_\sigma)(y_\sigma - \widehat{y}(X_\sigma^T \tau^*) - \Delta_{|c}^*)] \quad (54.71)$$

$$B_i = -E_{\sigma|c}\left[\lambda_c(\tilde{x}_\sigma)\frac{\partial \widehat{y}(X_\sigma^T \tau^*)}{\partial \tau_i^*}\right] \quad (54.72)$$

$$B(\tau^*) = \sum_i B_i(\widehat{\tau}_i - \tau_i^*) \quad (54.73)$$

$$= \sum_i E_{\sigma|c}\left[\lambda_c(\tilde{x}_\sigma)\frac{\partial \widehat{y}(X_\sigma^T \tau^*)}{\partial \tau_i^*}\right] \frac{E_{\sigma|c}[S_{i,\sigma}(\tau^*)]}{E_{\sigma|c}\left[\frac{\partial S_{i,\sigma}(\tau^*)}{\partial \tau_i}\right]} \quad (54.74)$$

$$C = -E_{\sigma|c}[\lambda_c(\tilde{x}_\sigma)] \quad (54.75)$$

$$IF_\sigma = \frac{R_\sigma(\hat{\tau}, \hat{\Delta}_{|c}) + B(\hat{\tau})}{C} \quad (54.76)$$

$$\mathcal{V} = E_{\sigma|c}[(IF_\sigma)^2] \quad (54.77)$$

QED

54.4 Other Estimands besides $\Delta_{|c}$

Suppose $\tilde{x}_0, \tilde{x}_1 \in S_{\tilde{x}}$. Define the **Modified ATE (MATE)** by

$$MATE_{|c} = \tilde{\mathcal{Y}}_{|\tilde{x}=\tilde{x}_1, c} - \tilde{\mathcal{Y}}_{|\tilde{x}=\tilde{x}_0, c} \quad (54.78)$$

$$MATE = \sum_c P(c) MATE_{|c} \quad (54.79)$$

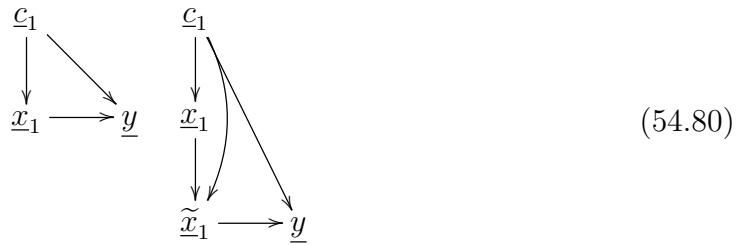
54.5 Multi-time MTP

A g-formula is any formula that defines recursively the full probability distribution of a bnet. In other words, it's a recursive definition of a Dynamical Bayesian Network.⁵

Let's define a g-formula⁶ for Multi-time (a.k.a. longitudinal) MTP (LMTP).

Consider $t = 1, 2, \dots, nt$,

For $nt = 1$,



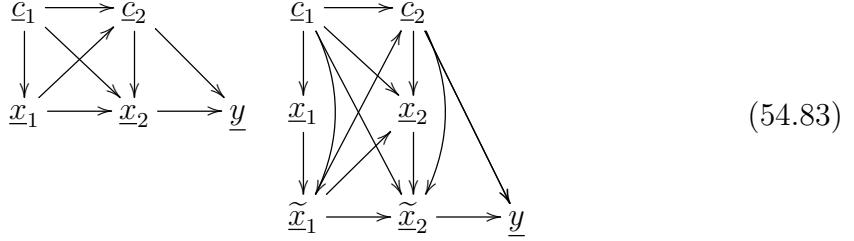
$$P(y, x_1, c_1) = P(y|x_1, c_1)P(x_1|c_1)P(c_1) \quad (54.81)$$

$$P(y, \tilde{x}_1, x_1, c_1) = P(y|\tilde{x}_1, c_1)P(x_1|c_1)P(\tilde{x}_1|x_1, c_1)P(c_1) \quad (54.82)$$

For $nt = 2$,

⁵Dynamical Bayesian Networks are discussed in Chapter 24.

⁶Chapter 29 defines a different g-formula that we call a sequential backdoor g-formula.

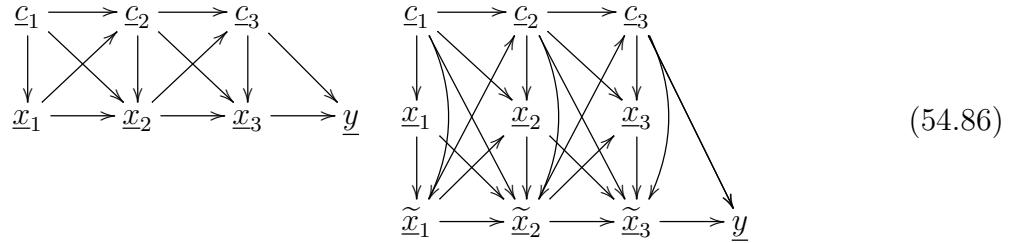


(54.83)

$$P(y, x_{\leq 2}, c_{\leq 2}) = \begin{cases} P(x_1|c_1)P(c_1) \\ *P(x_2|x_1, c_{2,1})P(c_2|x_1, c_1) \\ *P(y|x_2, c_2) \end{cases} \quad (54.84)$$

$$P(y, \tilde{x}_{\leq 2}, x_{\leq 2}, c_{\leq 2}) = \begin{cases} P(x_1|c_1)P(\tilde{x}_1|x_1, c_1)P(c_1) \\ *P(x_2|\tilde{x}_1, c_{2,1})P(\tilde{x}_2|x_2, c_{2,1})P(c_2|\tilde{x}_1, c_1) \\ *P(y|\tilde{x}_2, c_2) \end{cases} \quad (54.85)$$

For $nt = 3$,



(54.86)

$$P(y, x_{\leq 3}, c_{\leq 3}) = \begin{cases} P(x_1|c_1)P(c_1) \\ *P(x_2|x_1, c_{2,1})P(c_2|x_1, c_1) \\ *P(x_3|x_2, c_{3,2})P(c_3|x_2, c_2) \\ *P(y|x_3, c_3) \end{cases} \quad (54.87)$$

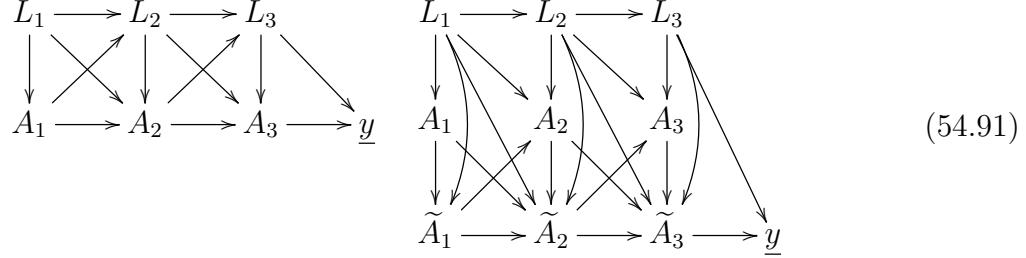
$$P(y, \tilde{x}_{\leq 3}, x_{\leq 3}, c_{\leq 3}) = \begin{cases} P(x_1|c_1)P(\tilde{x}_1|x_1, c_1)P(c_1) \\ *P(x_2|\tilde{x}_1, c_{2,1})P(\tilde{x}_2|\tilde{x}_1, x_{2,1}, c_{2,1})P(c_2|\tilde{x}_1, c_1) \\ *P(x_3|\tilde{x}_2, c_{3,2})P(\tilde{x}_3|\tilde{x}_2, x_{3,2}, c_{3,2})P(c_3|\tilde{x}_2, c_2) \\ *P(y|\tilde{x}_3, c_3) \end{cases} \quad (54.88)$$

In general,

$$P(y, x_{\leq nt}, c_{\leq nt}) = \begin{cases} \prod_{t=1}^{nt} \{P(x_t|x_{t-1}, c_{t,t-1})P(c_t|x_{t-1}, c_{t-1})\} \\ *P(y|x_{nt}, c_{nt}) \end{cases} \quad (54.89)$$

$$P(y, \tilde{x}_{\leq nt}, x_{\leq nt}, c_{\leq nt}) = \begin{cases} \prod_{t=1}^{nt} P(x_t|\tilde{x}_{t-1}, c_{t,t-1})P(\tilde{x}_t|\tilde{x}_{t-1}, x_{t,t-1}, c_{t,t-1})P(c_t|\tilde{x}_{t-1}, c_{t-1}) \\ *P(y|\tilde{x}_{nt}, x_{nt}, c_{nt}) \end{cases} \quad (54.90)$$

Note that Ref.[19] by Hernán and Robins (HR) uses the notation $L_t = \underline{c}_t$, $A_t = \underline{x}_t$ for $t = 1, 2, \dots, nt$. Hence, the $nt = 3$ bnet Eq.(54.86), written in the HR notation, is as follows (before and after adding the \tilde{A} nodes):



HR doesn't actually display the \tilde{A} nodes in its DAGs. Furthermore, in HR, nodes A_t and L_t have parent nodes that are time steps $0, 1, 2, 3, \dots$ in the past. I, on the other hand, do not draw or consider parents that are more than one time steps in the past, because I assume the effect of those parents is negligible to first approximation.

Chapter 55

Monty Hall Problem

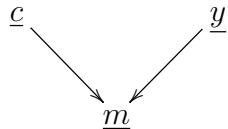


Figure 55.1: Monty Hall Problem.

Mr. Monty Hall, host of the game show “Let’s Make a Deal”, hides a car behind one of three doors and a goat behind each of the other two. The contestant picks Door No. 1, but before opening it, Mr. Hall opens Door No. 2 to reveal a goat. Should the contestant stick with No. 1 or switch to No. 3?

The Monty Hall problem can be modeled by the bnet Fig.55.1, where

- \underline{c} = the door behind which the car actually is.
- \underline{y} = the door opened by you (the contestant), on your first selection.
- \underline{m} = the door opened by Monty (game host)

We label the doors 1,2,3 so $S_c = S_y = S_m = \{1, 2, 3\}$.

The TPMs, printed in blue, for this bnet, are as follows:

$$P(c) = \frac{1}{3} \text{ for all } c \quad (55.1)$$

$$P(y) = \frac{1}{3} \text{ for all } y \quad (55.2)$$

$$P(m|c, y) = \mathbb{1}(m \neq c) \left[\frac{1}{2} \mathbb{1}(y = c) + \mathbb{1}(y \neq c) \mathbb{1}(m \neq y) \right] \quad (55.3)$$

It's easy to show that the above node probabilities imply that

$$P(c = 1|m = 2, y = 1) = \frac{1}{3} \quad (55.4)$$

$$P(c = 3|m = 2, y = 1) = \frac{2}{3} \quad (55.5)$$

So you are twice as likely to win if you switch your final selection to be the door which is neither your first choice nor Monty's choice.

The way I justify this to myself is: Monty gives you a piece of information. If you don't switch your choice, you are wasting that info, whereas if you switch, you are using the info.

Chapter 56

Multi-armed Bandits

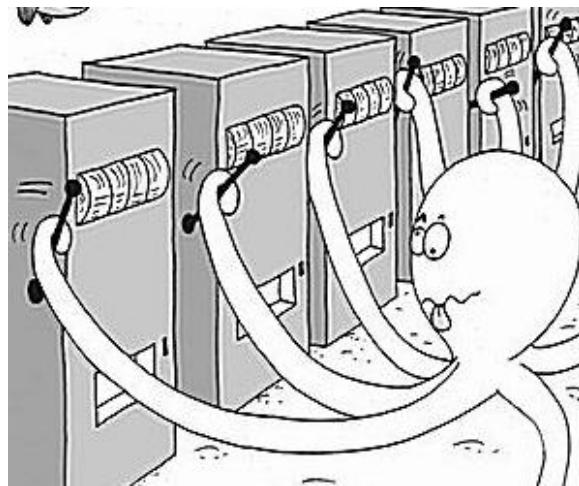


Figure 56.1: Multi-armed bandit (MAB).

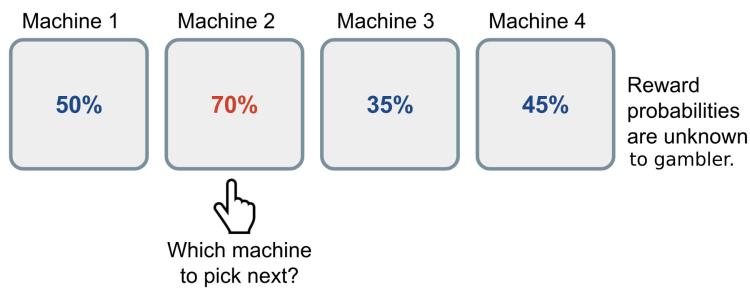


Figure 56.2: Bernoulli MAB (Bern-MAB) with 4 arms. For a Bern-MAB, the conditional probability distribution $P(r|a)$ for reward r , given arm a , is initially totally unknown to the gambler (agent), but it is known by the environment to be a Bernoulli distribution $P(r|a) = \mu_a^r(1 - \mu_a)^{1-r}$ for $r \in \{0, 1\}$, where $0 < \mu_a < 1$ for each a . The percentages shown are μ_a for each arm a .

This chapter is mostly based on Refs. [52] and [73].

Multi-armed Bandits (MABs) are a simple version of Reinforcement Learning (RL). RL is discussed in Chapter 69.

The term “one-armed-bandit” is a humorous term for what is also called a slot machine. A slot machine is a gambling device which has a slot into which you put coins or tokens for the privilege of being allowed to pull down a lever (arm) on one side of the device. This action generates a random combination of three shapes, which may or may not, depending on their combination, entitle the player to a money award.

Multi-armed bandit (MAB) is the name given to the optimization problem that considers an agent (gambler) that is playing multiple one-armed-bandits, each with a possibly different odds of winning. The optimization problem is to determine an efficient schedule whereby the gambler can converge on the device with the highest odds of winning.

MABs are often used in marketing as an alternative to A/B testing. These 2 methods yield different information but overlap in that they both can discover consumer preferences.

The MAB problem is an optimization problem (i.e., finding the maximum of a reward function or the minimum of a cost function). As with any minimization problem, an algorithm to solve it runs the danger of converging to a local minimum that isn’t the global (i.e., the overall) minimum. This danger can be diminished by doing both **exploration and exploitation**. Algorithms that do no exploration, only exploitation, are said to be **greedy**, and they are at the highest risk of converging to a non-global minimum.

56.1 Bnet for MAB

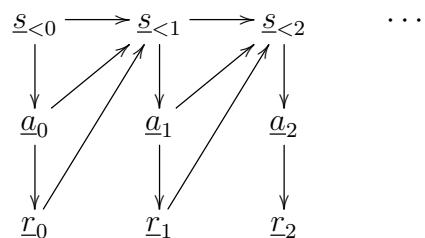


Figure 56.3: Bnet for a multi-armed bandit (MAB).

Let

$t \in \{0, 1, 2, \dots\}$ be the **time** slice (step),

$\underline{a}_t \in \{0, 1, \dots, N_a - 1\} = S_a$ be the **arm** of bandit that is pulled at time t , out of N_a arms. In RL language, it's also the **action** taken.

$\underline{r}_t \in \mathbb{R}$ be the **reward** at time t ,

$s_{<t} = \underbrace{[(a_\tau, r_\tau) : \tau < t]}_{s_\tau}$ be the **state** at time t .

Fig.56.3 shows a bnet for a MAB. The TPMs, printed in blue, for this bnet, are as follows.

$$P(s_{<t}|s_{<t-1}, a_{t-1}, r_{t-1}) = \mathbb{1}(s_{<t} = s_{<t-1} \cup (a_{t-1}, r_{t-1})) \quad (56.1)$$

For $t = 0$, $s_{<0} = \emptyset$, so choose a random a_0 .

$$P(a_t|s_{<t}) = \mathbb{1}(a_t = a_t^*) \text{ (agent's response)} , \quad (56.2)$$

where a_t^* depends on the **strategy (a.k.a. policy)** being used by the **gambler (agent)**. We consider various strategies below. a_t^* is defined solely in terms of empirical data (i.e., a_t, r_t values) collected from previous time slices. That is the only type of information that the gambler is privy to. a_0^* is chosen at random. The formulae for a_t^* that we give below for various strategies should only be used for $t > 0$.

$$P(r_t|a_t) = P_{r|a}(r_t|a_t) \text{ (environment's response)} . \quad (56.3)$$

$P_{r|a}$ is a probabilistic model that models the **environment** in which the agent lives. This assumes that the a_t (and the r_t) are i.i.d. $P_{r|a}$ depends on parameters whose values are known by the environment, but are not known a priori by the gambler. In fact, the goal of this exercise is for the gambler to find ever more accurate estimates of those parameters, using only the empirical data he/she can collect from the past, starting from total ignorance about those parameters.

For a **Bernoulli MAB (Bern-MAB)**, the conditional probability distribution $P_{r|a}$ is a Bernoulli distribution

$$P(r|a) = \mu_a^r (1 - \mu_a)^{1-r} \quad (56.4)$$

for $r \in \{0, 1\}$, where $0 < \mu_a < 1$ for each a . The parameters μ_a are initially unknown to the gambler. Note that $E[r|a] = \sum_r r P(r|a) = P(r=1|a) = \mu_a$.

For a **Gaussian MAB**, the conditional probability distribution $P_{r|a}$ is a Normal (Gaussian) distribution

$$P(r|a) = \mathcal{N}(r; \mu_a, \sigma_a^2) . \quad (56.5)$$

for $r \in \mathbb{R}$. The parameters μ_a, σ_a^2 are initially unknown to the gambler.

56.2 Reward functions

For $a \in S_{\underline{a}}$, define the **Long term Average Reward** for action a by

$$\mu_a = Q(a) = E_{|a}[\underline{r}] = \sum_r r P(r|a) . \quad (56.6)$$

Let

$$a^* = \operatorname{argmax}_a Q(a) \quad (56.7)$$

$$\mu^* = \max_a Q(a) = Q(a^*) \quad (56.8)$$

$$\Delta_a = \mu^* - Q(a) \quad (56.9)$$

μ_a, a^*, μ^* and Δ_a are not known to the gambler.

Define the **Instantaneous (at time t) Average Reward** for action a by

$$Q_t(a) = \frac{1}{N_t(a)} \sum_{\tau=0}^t r_\tau \mathbb{1}(a_\tau = a) \quad (56.10a)$$

where

$$N_t(a) = N_{in} + \sum_{\tau=0}^t \mathbb{1}(a_\tau = a) \quad (56.10b)$$

$N_{in} > 0$ insures that we never divide by zero. Note that Eqs.(56.10) can be stated recursively as

$$N_t(a)Q_t(a) = N_{t-1}(a)Q_{t-1}(a) + r_t \mathbb{1}(a_t = a) \quad (56.11a)$$

$$N_t(a) = N_{t-1}(a) + \mathbb{1}(a_t = a) \quad (56.11b)$$

with $Q_{-1}(a) = 0$ and $N_{-1}(a) = N_{in}$ for all a . We assume that at large t , $N_{in} \ll N_t(a)$ for all a . For instance, one can use $N_{in} = 1$.

$Q(a)$ is not known to the gambler but $N_t(a)$ and $Q_t(a)$ are because they are empirical.

We will write a hat over random variables that are defined by an empirical probability distribution.¹ Whereas we assume that \underline{a}_t and \underline{r}_t are i.i.d., we will not assume that \widehat{a}_t and \widehat{r}_t are i.i.d. for finite times t . What we will assume is that as $t \rightarrow \infty$,

¹An alternative convention is to not distinguish between \widehat{a}_t and \underline{a}_t , or between \widehat{r}_t and \underline{r}_t , but to distinguish between $P(r_t|a_t)$ and $\widehat{P}(r_t|a_t)$, where $\widehat{P}()$ is an empirical estimate of $P()$.

$$\widehat{a}_t \rightarrow \underline{a}, \quad \widehat{r}_t \rightarrow \underline{r} \quad (56.12)$$

Note that

$$\sum_a \frac{N_t(a)}{t+1} = 1 \quad (56.13)$$

so define

$$P(\widehat{a}_t = a) = \frac{N_t(a)}{t+1}. \quad (56.14)$$

Thus

$$E[Q(\widehat{a}_t)] = \sum_a P(\widehat{a}_t = a) Q(a). \quad (56.15)$$

Claim 79

$$Q_t(a) = E_{|\widehat{a}_t=a}[\widehat{r}_t] \quad (56.16)$$

proof:

Note that

$$\sum_r \sum_{\tau=0}^t \frac{\mathbb{1}(a_\tau = a, r_\tau = r)}{N_t(a)} = 1. \quad (56.17)$$

so define

$$P(\widehat{r}_t = r | \widehat{a}_t = a) = \sum_{\tau=0}^t \frac{\mathbb{1}(a_\tau = a, r_\tau = r)}{N_t(a)}. \quad (56.18)$$

Thus

$$Q_t(a) = \sum_r r \sum_{\tau=0}^t \frac{\mathbb{1}(a_\tau = a, r_\tau = r)}{N_t(a)} \quad (56.19)$$

$$= \sum_r r P(\widehat{r}_t = r | \widehat{a}_t = a) \quad (56.20)$$

$$= E_{|\widehat{a}_t=a}[\widehat{r}_t] \quad (56.21)$$

QED

Claim 80 As $t \rightarrow \infty$,

$$Q_t(a) \rightarrow E_{a=a}[\underline{r}] = Q(a) \quad (56.22)$$

$$E[Q(\widehat{a}_t)] \rightarrow E[Q(\underline{a})] \quad (56.23)$$

proof: This is clear from $\hat{a}_t \rightarrow a$ and $\hat{r}_t \rightarrow r$ and Claim 79.

QED

56.3 Regret functions

Define the **Instantaneous (at time t) Average Regret (I-Regret)** by

$$IReg_t = \mu^* - E[Q(\hat{a}_t)] = E[\Delta_{\hat{a}_t}] \quad (56.24)$$

and the **Cumulative (for times $\leq t$) Average Regret (C-Regret)** by

$$CReg_t = \sum_{\tau=0}^t IReg_\tau . \quad (56.25)$$

Note that

$$CReg_t = \sum_{\tau=0}^t E[\Delta_{\hat{a}_\tau}] \quad (56.26)$$

$$= \sum_a \Delta_a \sum_{\tau=0}^t \frac{N_\tau(a)}{\tau+1} \quad (56.27)$$

Let

$$\rho_t = \frac{1}{t+1} \sum_{t=0}^t r_t . \quad (56.28)$$

Define the **Cumulative Average Reward (C-Reward)** by

$$CRew_t = E_{|\hat{a}_t=a}[(t+1)\hat{\rho}_t] . \quad (56.29)$$

It can be shown that minimizing the C-Regret is equivalent to maximizing the C-Reward.

Claim 81 *As $t \rightarrow \infty$,*

$$IReg_t \rightarrow E[\Delta_a] \quad (56.30)$$

proof: This is clear from $\hat{a}_t \rightarrow a$ and $\hat{r}_t \rightarrow r$.

QED

56.4 Strategies with random exploration

In this section, we consider MAB algorithms that explore all values of the action a at random. In the section following this one, we consider MAB algorithm that do a more deliberate search of the action space.

56.4.1 ϵ -greedy algorithm

Recall that $\underline{a}_t \in \{0, 1, \dots, N_{\underline{a}} - 1\} = S_{\underline{a}}$. The user of the algorithm specifies an $\epsilon \in [0, 1]$ which measures the amount of exploration to be conducted.

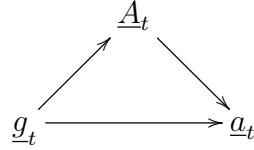


Figure 56.4: Extra structure added to MAB Bnet Fig.56.3 for ϵ -greedy algorithm.

For each t , add extra the structure shown in Fig.56.4 to the MAB bnet Fig.56.3. $g_t \in \{0, 1\}$ and $\underline{a}_t, \underline{A}_t \in S_{\underline{a}}$. (“g” stands for greedy). The TPMs, printed in blue, for the new nodes, are as follows

$$P(g_t) = \begin{cases} \epsilon \text{ (exploration)} & \text{if } g_t = 0 \\ 1 - \epsilon \text{ (exploitation)} & \text{if } g_t = 1 \end{cases}. \quad (56.31)$$

$$P(\underline{A}_t|g_t) = \begin{cases} \frac{1}{N_{\underline{a}}} \text{ (exploration)} & \text{if } g_t = 0 \\ \mathbb{1}(\underline{A}_t = 0) \text{ (exploitation)} & \text{if } g_t = 1 \end{cases} \quad (56.32)$$

$$P(\underline{a}_t|\underline{A}_t, g_t) = \begin{cases} \mathbb{1}(\underline{a}_t = \underline{A}_t) \text{ (exploration)} & \text{if } g_t = 0 \\ \mathbb{1}(\underline{a}_t = \underline{a}_t^*) \text{ (exploitation)} & \text{if } g_t = 1 \end{cases} \quad (56.33)$$

where \underline{a}_t^* is defined as follows:

$$\boxed{\underline{a}_t^* = \operatorname{argmax}_a Q_{t-1}(a)}. \quad (56.34)$$

As $t \rightarrow \infty$,

$$\frac{N_t(a)}{t+1} = P(\widehat{\underline{a}}_t = a) \quad (56.35)$$

$$\rightarrow P(a, g_t = 0)\epsilon + P(a, g_t = 1)(1 - \epsilon) \quad (56.36)$$

$$\geq P(a, g_t = 0)\epsilon \quad (56.37)$$

$$= \frac{\epsilon}{N_{\underline{a}}} \quad (56.38)$$

Hence

$$IReg_t = \sum_a \Delta_a \sum_{\tau=0}^t \frac{N_\tau(a)}{\tau + 1} \quad (56.39)$$

$$\geq \frac{\epsilon}{N_{\underline{a}}} \sum_a \Delta_a \quad (56.40)$$

and $CReg_t \geq \mathcal{N}(!t)(t + 1)$.

56.4.2 ϵ_t -greedy algorithm

Replace time-independent constant ϵ in the ϵ -greedy algorithm by a time dependent function ϵ_t .

56.5 Strategies with nonrandom exploration

56.5.1 Upper Confidence Bounds (UCB) algorithms

A MAB algorithm that maximizes merely $Q_t(a)$ to get a_t^* is totally greedy (i.e., does no exploration, only exploitation). This doesn't work too well because once the algo finds a particular a_t^* , it sticks with it. For times t after that, the $Q_t(a)$ for all a stay more or less the same. The $N_t(a)$ for $a \neq a_t^*$ also stay the same. Only $N_t(a_t^*)$ increases. To avoid this problem, **Upper Confidence Bounds (UCB) algorithms** maximize an effective $Q_t(a)_{\text{eff}} = Q_t(a) + U_t(a)$, where $U_t(a) > 0$, instead of maximizing merely $Q_t(a)$ to get a_t^* . $U_t(a)$ is proportional to $1/\sqrt{N_t(a)}$ (that's a property of UCBs). Adding $U_t(a)$ to $Q_t(a)$ gives those a with low $N_a(a)$ a $Q_t(a)_{\text{eff}} \gg Q_t(a)$. This encourages exploration of a different from $a_t^* = \operatorname{argmax}_a Q_t(a)$. In conclusion, for all UCB algorithms, we have

$$a_t^* = \operatorname{argmax}_a [Q_{t-1}(a) + U_{t-1}(a)] \quad (56.41)$$

Different UCB algos differ only in the definition of $U_t(a)$.

Next, we consider two UCB algorithms, frequentist UCB (UCB1), and Bayesian UCB.

Frequentist UCB (UCB1) algorithm

Claim 82 (Hoeffding's Inequality- HI) Let $\underline{x}_0, \underline{x}_1, \dots, \underline{x}_{T-1}$ be i.i.d. random variables such that $0 \leq \underline{x}_t \leq 1$ for all t . Let $\underline{m}_{T-1} = \frac{1}{T} \sum_{t=0}^{T-1} \underline{x}_t$ denote the sample mean. Then for $u > 0$, we have:

$$P(E[\underline{x}] > \underline{m}_{T-1} + u) \leq e^{-2Tu^2}. \quad (56.42)$$

proof: See Ref.[107].

QED

Let²

$$\underline{x}^T = \{\underline{x}_t : t \in \mathbb{Z}_{[0, T-1]}\} \quad (56.43)$$

and

$$\underline{r}_{\leq t-1}(a) = \{\underline{r}_\tau : \underline{a}_\tau = a, \tau \in \mathbb{Z}_{[0, t-1]}\}. \quad (56.44)$$

Note that \underline{x}^T has T components and $\underline{r}_{\leq t-1}(a)$ has $N_{t-1}(a)$ components. If we apply the HI with \underline{x}^T replaced by $\underline{r}_{\leq t-1}(a)$, we get

$$P(Q(a) > Q_{t-1}(a) + U_{t-1}(a)) \leq e^{-2N_{t-1}(a)[U_{t-1}(a)]^2} \quad (56.45)$$

If we define a threshold probability p by

$$p = e^{-2N_{t-1}(a)[U_{t-1}(a)]^2}, \quad (56.46)$$

then

$$U_{t-1}(a) = \sqrt{\frac{-\ln p}{2N_{t-1}(a)}} \quad (56.47)$$

If we choose $p = (t-1)^{-\alpha}$,

$$a_t^* = \underset{a}{\operatorname{argmax}} \left[Q_{t-1}(a) + \sqrt{\frac{\alpha(t-1)}{2N_{t-1}(a)}} \right] \quad (56.48)$$

Bayesian UCB algorithm

Claim 83 ³ ⁴ (*Bayesian updating of mean and deviation. See Fig.56.5.*)
Suppose

$$x^n = [x_i], \quad x_i \text{ are i.i.d. with } \underline{x}_i | \mu, \tau \sim \mathcal{N}(\mu, \tau) \quad (56.49)$$

$$\underline{\mu} | \tau \sim \mathcal{N}(\mu_0, n_0 \tau) \quad (56.50)$$

$$\underline{\tau} \sim \text{Gamma}(\alpha, \beta). \quad (56.51)$$

Then the posterior is

²As usual, we define $\mathbb{Z}_{[a,b]} = \{a, a+1, a+2, \dots, b\}$ for $a < b$.

³For a **standard deviation** σ , the **precision** τ is defined as $\tau = \frac{1}{\sigma^2}$.

⁴ $\underline{x} | \lambda \sim \mathcal{D}(\lambda)$ means that $P(x | \lambda) = \mathcal{D}(x; \lambda)$. $\mathcal{N}()$ stands for the Normal distribution and $\text{Gamma}()$ for the Gamma distribution. See Ref.[101] for a discussion of the Gamma distribution.

$$\underline{\mu}|\tau, \underline{x}^n \sim \mathcal{N}\left(\frac{n\bar{x} + n_0\mu_0}{n + n_0}, (n + n_0)\tau\right) \quad (56.52)$$

$$\underline{\tau}|x^n \sim \text{Gamma}\left(\alpha + \frac{n}{2}, \beta + \frac{1}{2}\sum_i(x_i - \bar{x})^2 + \frac{nn_0}{2(n + n_0)}(\bar{x} - \mu_0)^2\right) \quad (56.53)$$

proof: See Ref.[24].

QED

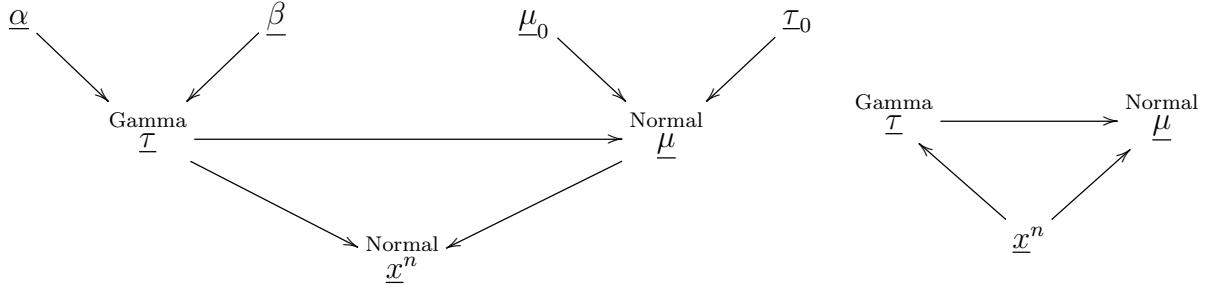


Figure 56.5: Prior and Posterior BNets in Claim 83.

In the Bayesian UCB algorithm, we use

$$a_t^* = \underset{a}{\operatorname{argmax}} E_{\mu_a, \sigma_a | s_{<t}} \left[\mu_a + c \frac{\sigma_a}{\sqrt{N_{t-1}(a)}} \right] \quad (56.54)$$

for some $c > 0$.

Eq.56.54 requires that we know $P(\mu_a, \sigma_a | s_{<t})$; i.e., the posterior distribution of μ_a, σ_a assuming the prior history $s_{<t}$. This follows from Bayes theorem if we assume a Normal distribution for the likelihood $P(s_{<t} | \mu_a, \sigma_a)$ and we assume conjugate priors for $P(\mu_a, \sigma_a)$. More precisely, if we replace \underline{x}^n by $\underline{s}_{<t}$ in Claim 83, then

$$P(\mu_a, \tau_a | s_{<t}) = P(\mu_a | \tau_a, s_{<t})P(\tau_a | s_{<t}) \quad (56.55)$$

where $P(\mu_a | \tau_a, s_{<t})$ and $P(\tau_a | s_{<t})$ are given by Claim 83.

56.5.2 Thompson Sampling MAB (TS-MAB) algorithm

Bnet for general TS-MAB algorithm

The Thompson Sampling MAB (TS-MAB) algorithm is described by the bnet Fig.56.6. This bnet differs from the bnet Fig.56.3 in that it includes new nodes λ_t . The TPMs, printed in blue, for bnet Fig.56.6, are as follows.

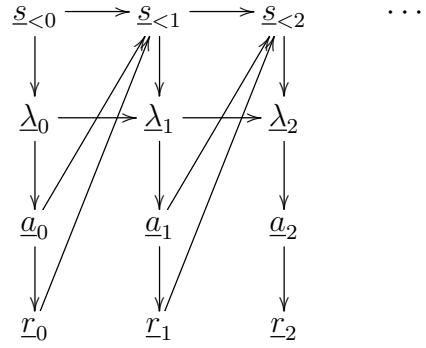


Figure 56.6: Bnet for TS-MAB algorithm.

$$P(s_{<t}|s_{<t-1}, a_{t-1}, r_{t-1}) = \text{ same as for Fig.56.3} \quad (56.56)$$

$$P(\lambda_t|s_{<t}, \lambda_{t-1}) = \mathbb{1}(\lambda_t = \lambda_t^*(s_{<t}, \lambda_{t-1})) \quad (56.57)$$

where λ_t^* is a function to be defined below.

$$P(a_t|\lambda_t) = \mathbb{1}[a_t = a_t(\lambda_t)] \quad (\text{Agent's response}) \quad (56.58)$$

where $a_t(\lambda_t)$ is a function to be described below.

$$P(r_t|a_t) = \text{ same as for Fig.56.3. (Environment's response)} \quad (56.59)$$

Let

$$Q_t(a, \lambda_t) = \sum_r r P(\hat{r}_t = r | \hat{a}_t = a, \lambda_t) \quad (56.60)$$

$$= E_{|\hat{a}_t=a, \lambda_t} [\hat{r}_t]. \quad (56.61)$$

Define

$a_t(\lambda_t) = \operatorname{argmax}_a Q_t(a, \lambda_t).$

(56.62)

Note that

$$P(a_t|s_{<t}) = \sum_{\lambda_t} P(\lambda_t|s_{<t}) \mathbb{1}[Q_t(a_t, \lambda_t) = \max_a Q_t(a, \lambda_t)] \quad (56.63a)$$

$$= \sum_{\lambda_t} P(\lambda_t|s_{<t}) \mathbb{1}[a_t = a_t(\lambda_t)] \quad (56.63b)$$

$$= E_{\lambda_t|s_{<t}} \{ \mathbb{1}[a_t = a_t(\lambda_t)] \}. \quad (56.63c)$$

If we further assume that $P(\lambda_t|s_{<t})$ is a delta function, then Eqs.(56.63) reduce to

$$P(a_t|s_{<t}) = \mathbb{1}[a_t = a_t(\lambda_t^*)] \quad (56.64)$$

TS-MAB algorithm with Beta agent and Bernoulli environment

The Beta distribution $\text{Beta}(x; \alpha, \beta)$ (see Ref.[82]) is defined for $\alpha > 0, \beta > 0$ and $x \in [0, 1]$. Since $x \in [0, 1]$, x can be interpreted as a probability. The mean and variance of the Beta distribution are

$$E[\underline{x}] = \frac{\alpha}{\alpha + \beta}, \quad (56.65)$$

$$\langle \underline{x}, \underline{x} \rangle = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}. \quad (56.66)$$

From this mean and variance, we see that if we increase α by one and leave β the same, the mean moves towards 1 and the variance decreases. Likewise, if we increase β by 1 and leave α the same, the mean moves towards 0 and the variance decreases.

Let

$$\alpha_t^a = \sum_{\tau=0}^t \mathbb{1}(r_\tau = 1, a_\tau = a) \quad (56.67)$$

$$\beta_t^a = \sum_{\tau=0}^t \mathbb{1}(r_\tau = 0, a_\tau = a) \quad (56.68)$$

$$\lambda_t^a = (\alpha_{t-1}^a, \beta_{t-1}^a) \quad (56.69)$$

$$\lambda_t = [(\alpha_{t-1}^a, \beta_{t-1}^a) : a \in S_a] \quad (56.70)$$

$$Q_t(a, \lambda_t) = \sum_r r \text{Beta}(r; \alpha_{t-1}^a, \beta_{t-1}^a) \quad (56.71a)$$

$$= \frac{\alpha_{t-1}^a}{\alpha_{t-1}^a + \beta_{t-1}^a} \quad (56.71b)$$

The TS-MAB algorithm for a Beta agent and Bernoulli environment can be described by the bnet Fig.56.6. For this special case of bnet Fig.56.6, the TPMs, printed in blue, are as follows.

$$P(s_{<t}|s_{<t-1}, a_{t-1}, r_{t-1}) = \text{ same as for Fig.56.6} \quad (56.72)$$

$$P(\lambda_t^a|s_{<t}, \lambda_{t-1}) = \mathbb{1}(\lambda_t^a = \lambda_{t-1}^a) = \begin{cases} \mathbb{1}(\lambda_t^a = \lambda_{t-1}^a) & \text{if } a \neq a_{t-1} \\ \mathbb{1}(\lambda_t^a = (\alpha_{t-2}^a + 1, \beta_{t-2}^a)) & \text{if } a_{t-1} = a \text{ and } r_{t-1} = 1 \\ \mathbb{1}(\lambda_t^a = (\alpha_{t-2}^a, \beta_{t-2}^a + 1)) & \text{if } a_{t-1} = a \text{ and } r_{t-1} = 0 \end{cases} \quad (56.73)$$

$$P(a_t|\lambda_t) = \mathbb{1}(a_t = a_t(\lambda_t)) \quad (56.74)$$

$$= \mathbb{1}(a_t = \operatorname{argmax}_a \underbrace{Q_t(a, \lambda_t)}_{\text{see Eq.(56.71)}}) \quad (\text{Beta agent response}) \quad (56.75)$$

$$P(r_t|a_t) = \mu_{a_t}^{r_t} (1 - \mu_{a_t})^{r_t} \quad (\text{Bernoulli environment response}) . \quad (56.76)$$

μ_a known to environment but unknown to agent.

TS-MAB algorithm, skeletal reprise

The TS-MAB algorithm is not very complicated but explaining it with precision requires nightmarishly many indices. Here is a pedagogical reprise of what we have said so far, where we have stripped out some of the inessential indices.

$$P(\lambda) = \mathbb{1}(\lambda, \lambda^*) \quad (56.77)$$

$$P(a|\lambda) = \mathbb{1}(a = a(\lambda)) \quad (56.78)$$

$$= \mathbb{1}(a = \operatorname{argmax}_a \sum_r r P(\hat{r} = r | \hat{a} = a, \lambda)) \quad (56.79)$$

$$= \mathbb{1}(a = \operatorname{argmax}_a \sum_r r \text{Beta}(r; \lambda^a)) \quad (\text{Beta agent response}) \quad (56.80)$$

Define $q()$ by

$$q(r; \lambda^a) = P(\hat{r} = r | \hat{a} = a, \lambda) \quad (56.81)$$

- PRIOR:

$$P(a) = \sum_{\lambda} P(\lambda)P(a|\lambda) \quad (56.82)$$

$$= P(a|\lambda^*) \quad (56.83)$$

$$= \mathbb{1}(a = \operatorname{argmax}_a \sum_r r q(r; \lambda^{*a})) \quad (56.84)$$

1. Use fact that $q = \text{Beta}$

$$\sum_r r \text{Beta}(r; \lambda^{*a}) = \frac{\alpha^{*a}}{\alpha^{*a} + \beta^{*a}} \quad (56.85)$$

2. Don't use fact that $q = \text{Beta}$.

For each a , get samples $r^\sigma \sim q(r; \lambda^{*a})$ for $\sigma = 0, 1, \dots, nsam - 1$ and estimate

$$\sum_r r q(r; \lambda^{*a}) \approx \frac{1}{nsam} \sum_{\sigma} r^\sigma q(r^\sigma; \lambda^{*a}). \quad (56.86)$$

This sampling is why TS is called a sampling.

- LIKELIHOOD:

$$P(r|a) = \mu_a^r (1 - \mu_a)^{r'} \quad (\text{Bernoulli environment response}) \quad (56.87)$$

56.5.3 Grad-MAB algorithm

Let

$$\lambda_{t+1}(a) = \lambda_t(a) + \eta r_t [\mathbb{1}(a_t = a) - \pi_t(a)] \quad (56.88)$$

for some $\eta > 0$, where $\pi_t(a)$ is defined by

$$\pi_t(a) = P(\underline{a}_t = a | \lambda_t) = \underbrace{\frac{e^{\lambda_t(a)}}{\sum_a e^{\lambda_t(a)}}}_{\text{softmax}(\lambda_t)_a}. \quad (56.89)$$

The $\lambda_t(a)$ are called **scores** at time t . Let

$a_t^* = \operatorname{argmax}_a \pi_t(a)$

(56.90)

The Gradient MAB (Grad-MAB) algorithm is described by the bnet Fig.56.7. This bnet differs from the bnet Fig.56.3 in that it includes new nodes $\underline{\lambda}_t$. The TPMs, printed in blue, for bnet Fig.56.7, are as follows.

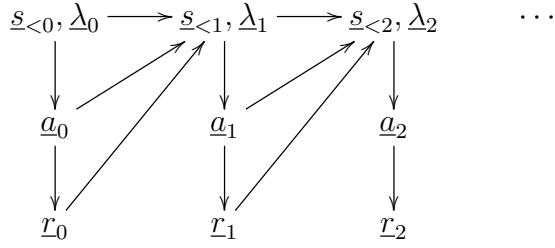


Figure 56.7: Bnet for Grad-MAB algorithm.

$$P(\underline{\lambda}_{t+1} = \lambda | a_t, r_t, \underline{\lambda}_t) = \mathbb{1}(\lambda = \lambda_{t+1} \text{ given by Eq.(56.88)}). \quad (56.91)$$

$$P(a_t | \underline{\lambda}_t) = \mathbb{1}(a_t = a_t^*) \quad \text{or, alternatively, } = P(\underline{a}_t = a_t | \underline{\lambda}_t) = \pi_t(a_t) \quad (\text{agent's response}) \quad (56.92)$$

$$P(r_t | a_t) = P_{\underline{r}|\underline{a}}(r_t | a_t) \quad (\text{environment's response}) . \quad (56.93)$$

Motivation:

Define the **Instantaneous Average Reward** by

$$\mathcal{R}_t(\underline{\lambda}_t) = \sum_a P(\underline{a}_t = a | \underline{\lambda}_t) E_{r_t | a_t = a}[\underline{r}_t] = E_{r_t, \underline{a}_t | \underline{\lambda}_t}[\underline{r}_t] . \quad (56.94)$$

We will assume that as $t \rightarrow \infty$, $\underline{a}_t \rightarrow \underline{a}$, $\underline{r}_t \rightarrow \underline{r}$ and $\underline{\lambda}_t \rightarrow \underline{\lambda}$. Therefore, $\mathcal{R}_t(\underline{\lambda}_t) \rightarrow E_{\underline{r}, \underline{a} | \underline{\lambda}}[\underline{r}] = E_{\underline{r} | \underline{\lambda}}[\underline{r}]$.

Note that if $E_{r_t | a_t = a}[\underline{r}_t] = B$, where B is independent of a , then $\mathcal{R}_t(\underline{\lambda}_t) = B$ and $\frac{\partial \mathcal{R}_t}{\partial \lambda_t(a)} = 0$.

Claim 84 *The gradient of $\mathcal{R}_t(\underline{\lambda}_t)$ is*

$$\frac{\partial \mathcal{R}_t}{\partial \lambda_t(a)} = E_{r_t, \underline{a}_t | \lambda_t}[g(r_t, \underline{a}_t | a, \lambda_t)] \quad (56.95)$$

where

$$g(r_t, \underline{a}_t | a, \lambda_t) = \underline{r}_t [\mathbb{1}(\underline{a}_t = a) - \pi_t(a)] \quad (56.96)$$

proof:

$$\frac{\partial \mathcal{R}_t}{\partial \lambda_t(a)} = \sum_{a'} \frac{\partial \pi_t(a')}{\partial \lambda_t(a)} E_{\underline{r}_t | \underline{a}_t = a'} [\underline{r}_t] \quad (56.97)$$

$$= \sum_{a'} \pi_t(a') \frac{\partial \ln \pi_t(a')}{\partial \lambda_t(a)} E_{\underline{r}_t | \underline{a}_t = a'} [\underline{r}_t] \quad (56.98)$$

$$= \sum_{a'} \pi_t(a') \frac{\partial}{\partial \lambda_t(a)} \left[\ln \frac{\exp[\lambda_t(a')]}{\sum_a \exp[\lambda_t(a)]} \right] E_{\underline{r}_t | \underline{a}_t = a'} [\underline{r}_t] \quad (56.99)$$

$$= \sum_{a'} \pi_t(a') [\mathbb{1}(a' = a) - \pi_t(a)] E_{\underline{r}_t | \underline{a}_t = a'} [\underline{r}_t] \quad (56.100)$$

$$= \sum_{a'} P(\underline{a}_t = a' | \lambda_t) E_{\underline{r}_t | \underline{a}_t = a'} [\underline{r}_t (\mathbb{1}(a' = a) - \pi_t(a))] \quad (56.101)$$

$$= E_{\underline{r}_t, \underline{a}_t | \lambda_t} [g(\underline{r}_t, \underline{a}_t | a, \lambda_t)] \quad (56.102)$$

QED

Eq.(56.88) can be written as

$$\lambda_{t+1}(a) = \lambda_t(a) + \eta g(r_t, a_t | a, \lambda_t) \quad (56.103)$$

Chapter 57

Naive Bayes

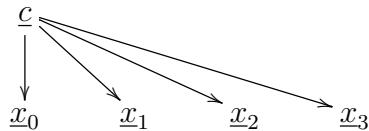


Figure 57.1: bnet for Naive Bayes with 4 features

Class node $\underline{c} \in S_{\underline{c}}$. $|S_{\underline{c}}| = n_{\underline{c}}$ = number of classes.

Feature nodes $\underline{x}_i \in S_{\underline{x}_i}$ for $i = 0, 1, 2, \dots, F - 1$. F =number of features.

Define

$$x. = [x_0, x_1, \dots, x_{F-1}] . \quad (57.1)$$

For the bnet of Fig.57.1,

$$P(c, x.) = P(c) \prod_{i=0}^{F-1} P(x_i|c) . \quad (57.2)$$

Given $x.$ values, find most likely class $c \in S_{\underline{c}}$.

Maximum a Posteriori (MAP) estimate:

$$c^* = \operatorname{argmax}_c P(c|x.) \quad (57.3)$$

$$= \operatorname{argmax}_c \frac{P(c, x.)}{P(x.)} \quad (57.4)$$

$$= \operatorname{argmax}_c P(c, x.) . \quad (57.5)$$

Chapter 58

Neural Networks

In this chapter, we discuss Neural Networks (NNs) of the feedforward kind, which is the most popular kind. In their plain, vanilla form, NNs only have deterministic nodes. But the nodes of a bnet can be deterministic too, because the TPM of a node can reduce to a delta function. Hence, NNs should be expressible as bnets. We will confirm this in this chapter.

Henceforth in this chapter, if we replace an index of an indexed quantity by a dot, it will mean the collection of the indexed quantity for all values of that index. For example, $\underline{x}.$ will mean the array of x_i for all i .

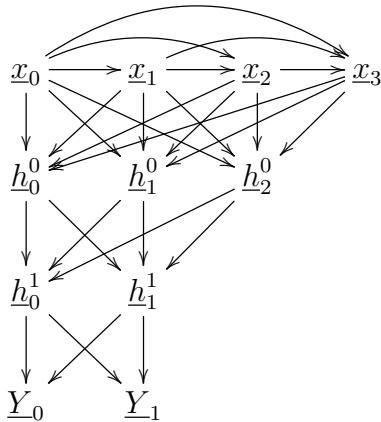


Figure 58.1: Neural Network (feed forward) with 4 layers: input layer $\underline{x}.$, 2 hidden layers $\underline{h}^0.$, $\underline{h}^1.$ and output layer $\underline{Y}.$

Consider Fig.58.1.

$\underline{x}_i \in \{0, 1\}$ for $i = 0, 1, 2, \dots, nx - 1$ is the **input layer**.

$\underline{h}_i^\lambda \in \mathbb{R}$ for $i = 0, 1, 2, \dots, nh(\lambda) - 1$ is the λ -th **hidden layer**. $\lambda = 0, 1, 2, \dots, \Lambda - 2$. A NN is said to be **deep** if $\Lambda > 2$; i.e., if it has more than one hidden layer.

$\underline{Y}_i \in \mathbb{R}$ for $i = 0, 1, 2, \dots, ny - 1$ is the **output layer**. We use a upper case y here because in the training phase, we will use pairs $(x.[\sigma], y.[\sigma])$ where $y_i[\sigma] \in \{0, 1\}$

for $i = 0, 1, \dots, ny - 1$. $Y = \hat{y}$ is an estimate of y . Note that lower case y is either 0 or 1, but upper case Y may be any real. Often, the activation functions are chosen so that $Y \in [0, 1]$.

The number of nodes in each layer and the number of layers are arbitrary. Fig.58.1 is fully connected (a.k.a. dense), meaning that every node of a layer is impinged arrow coming from every node of the preceding layer. Later on in this chapter, we will discuss non-dense layers.

Let $w_{i|j}^\lambda, b_i^\lambda \in \mathbb{R}$ be given, for $i \in \mathbb{Z}_{[0, nh(\lambda)]}$, $j \in \mathbb{Z}_{[0, nh(\lambda-1)]}$, and $\lambda \in \mathbb{Z}_{[0, \Lambda]}$.

The TPMs, printed in blue, for bnet Fig.58.1, are as follows.

$$P(x_i | x_{i-1}, x_{i-2}, \dots, x_0) = \text{given} \quad (58.1)$$

$$P(h_i^\lambda | h_{\cdot}^{\lambda-1}) = \delta \left(h_i^\lambda, \mathcal{A}_i^\lambda \left(\sum_j w_{i|j}^\lambda h_j^{\lambda-1} + b_i^\lambda \right) \right), \quad (58.2)$$

where $P(h_i^0 | h^{-1}) = P(h_i^0 | x)$.

$$P(Y_i | h_{\cdot}^{\Lambda-2}) = \delta \left(Y_i, \mathcal{A}_i^{\Lambda-1} \left(\sum_j w_{i|j}^{\Lambda-1} h_j^{\Lambda-2} + b_i^{\Lambda-1} \right) \right). \quad (58.3)$$

58.1 Activation Functions $\mathcal{A}_i^\lambda : \mathbb{R} \rightarrow \mathbb{R}$

Activation functions must be nonlinear. Why? Because if they were all linear, the NN mapping would be a bijection (1-1 onto map), and its domain and range would be the same. That is not what you want for a classifier. For a classifier, you want the range to be much smaller than the domain.

- Step function (Perceptron)

$$\mathcal{A}(x) = \mathbb{1}(x > 0) \quad (58.4)$$

Zero for $x \leq 0$, one for $x > 0$.

- Sigmoid function

$$\mathcal{A}(x) = \frac{1}{1 + e^{-x}} = \text{smoid}(x) \quad (58.5)$$

Smooth, monotonically increasing function. $\text{smoid}(-\infty) = 0, \text{smoid}(0) = 1/2, \text{smoid}(\infty) = 1$.

- **Hyperbolic tangent**

$$\mathcal{A}(x) = \tanh(x) = \frac{e^x - e^{-x}}{e^x + e^{-x}} \quad (58.6)$$

Smooth, monotonically increasing function. $\tanh(-\infty) = -1$, $\tanh(0) = 0$, $\tanh(\infty) = 1$.

Odd function:

$$\tanh(-x) = -\tanh(x) \quad (58.7)$$

Whereas smoid(x) $\in [0, 1]$, $\tanh(x) \in [-1, 1]$.

- **ReLU (Rectified Linear Unit)**

$$\mathcal{A}(x) = \underbrace{x \mathbb{1}(x > 0)}_{x_+} = \max(0, x) . \quad (58.8)$$

Compare this to the step function $\mathbb{1}(x > 0)$.

- **Swish**

$$\mathcal{A}(x) = x \text{ smoid}(x) \quad (58.9)$$

- **Softmax**

$$\mathcal{A}(x_i|x.) = \frac{e^{x_i}}{\sum_i e^{x_i}} = \text{softmax}(x.)_i \quad (58.10)$$

The softmax definition implies that the bnet nodes within a softmax layer are fully connected by arrows to form a “clique”.

58.2 Weight optimization via supervised training and gradient descent

The bnet of Fig.58.1 is used for classification of a single data point $x..$. It assumes that the weights $w_{ij}^\lambda, b_i^\lambda$ are given.

To find the optimum weights via supervised training and gradient descent, one uses the bnet Fig.58.2.

In Fig.58.2, the nodes in Fig.58.1 become sampling space vectors. For example, $\underline{x}..$ becomes $\vec{x}_..$, where the components of $\vec{x}_..$ in sampling space are $\underline{x}..[\sigma] \in \{0, 1\}^{nx}$ for $\sigma = 0, 1, \dots, nsam(\vec{x}) - 1$.

$nsam(\vec{x})$ is the number of samples used to calculate the gradient during each stage (a.k.a. iteration) of Fig.58.2. We will also refer to $nsam(\vec{x})$ as the **mini-batch size**. A **mini-batch** is a subset of the training dataset.

To train a bnet with a data set (d-set), the standard procedure is to split the d-set into 3 parts:

1. **training d-set**,
2. **testing1 d-set**, for tuning of hyperparameters like $nsam(\vec{x})$, Λ , and $nh(i)$ for each i .
3. **testing2 d-set**, for measuring how well the model tuned with the testing1 d-set performs.

The training d-set is itself split into mini-batches. An **epoch** is a pass through all the training d-set.

Define

$$W_{i|j}^\lambda = [w_{i|j}^\lambda, b_i^\lambda] . \quad (58.11)$$

The TPMs, printed in blue, for bnet Fig.58.2, are as follows.

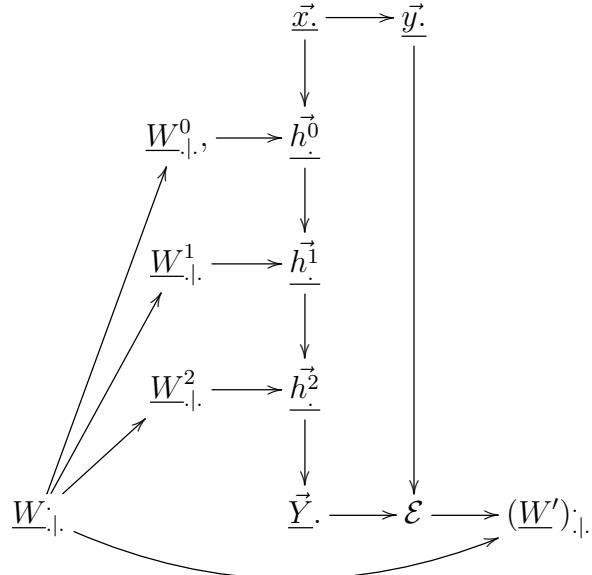


Figure 58.2: bnet for finding optimum weights of the bnet Fig.58.1 via supervised training and gradient descent.

$$P(x.[\sigma]) = \text{given} . \quad (58.12)$$

$$P(y.[\sigma] | x.[\sigma]) = \text{given} . \quad (58.13)$$

$$P(h_i^\lambda[\sigma] | h_j^{\lambda-1}[\sigma]) = \delta \left(h_i^\lambda[\sigma], \mathcal{A}_i^\lambda \left(\sum_j w_{i|j}^\lambda h_j^{\lambda-1}[\sigma] + b_i^\lambda \right) \right) \quad (58.14)$$

$$P(Y_i[\sigma] | h_{\cdot}^{\Lambda-2}[\sigma]) = \delta \left(Y_i[\sigma], \mathcal{A}_i^{\Lambda-1} \left(\sum_j w_{i|j}^{\Lambda-1} h_j^{\Lambda-2}[\sigma] + b_i^{\Lambda-1} \right) \right) \quad (58.15)$$

$$P(W_{\cdot|\cdot}) = \text{given} \quad (58.16)$$

The first time it is used, $W_{\cdot|\cdot}$ is arbitrary. After the first time, it is determined by previous stage.

$$P(W_{\cdot|\cdot}^{\lambda} | W_{\cdot|\cdot}) = \delta(W_{\cdot|\cdot}^{\lambda}, (W_{\cdot|\cdot})^{\lambda}) \quad (58.17)$$

$$P(\mathcal{E} | \vec{y}_{\cdot}, \vec{Y}_{\cdot}) = \frac{1}{nsam(\vec{x})} \sum_{\sigma} \sum_i d(y_i[\sigma], Y_i[\sigma]) , \quad (58.18)$$

where

$$d(y, Y) = |y - Y|^2 . \quad (58.19)$$

If $y, Y \in [0, 1]$, one can use this instead

$$d(y, Y) = XE(y \rightarrow Y) = -y \ln Y - (1 - y) \ln(1 - Y) . \quad (58.20)$$

$$P((W')_{i|j}^{\lambda} | \mathcal{E}, W_{\cdot|\cdot}) = \delta((W')_{i|j}^{\lambda}, W_{i|j}^{\lambda} - \eta \partial_{W_{i|j}^{\lambda}} \mathcal{E}) \quad (58.21)$$

$\eta > 0$ is called the learning rate. This method of minimizing the error \mathcal{E} is called gradient descent. $W' - W = \Delta W = -\eta \partial_W \mathcal{E}$ so $\Delta \mathcal{E} = \frac{-1}{\eta} (\Delta W)^2 < 0$.

58.3 Non-dense layers

The TPM for a non-dense layer is of the form:

$$P(h_i^{\lambda}[\sigma] | h_{\cdot}^{\lambda-1}[\sigma]) = \delta(h_i^{\lambda}[\sigma], H_i^{\lambda}[\sigma]) , \quad (58.22)$$

where $H_i^{\lambda}[\sigma]$ will be specified below for each type of non-dense layer.

- **Dropout Layer**

The dropout layer was invented in Ref.[58]. To dropout nodes from a fixed layer λ : For all i of layer λ , define a new node $\underline{r}_i^{\lambda}$ with an arrow $\underline{r}_i^{\lambda} \rightarrow \underline{h}_i^{\lambda}$. For $r \in \{0, 1\}$, and some $p \in (0, 1)$, define

$$P(r_i^\lambda = r) = [p]^r [1-p]^{1-r} \text{ (Bernoulli dist.) .} \quad (58.23)$$

Now one has

$$P(h_i^\lambda[\sigma] | h_j^{\lambda-1}[\sigma], r_i^\lambda) = \delta(h_i^\lambda[\sigma], H_i^\lambda[\sigma]) , \quad (58.24)$$

where

$$H_i^\lambda[\sigma] = \mathcal{A}_i^\lambda(r_i^\lambda \sum_j w_{i|j}^\lambda h_j^{\lambda-1}[\sigma] + b_i^\lambda) . \quad (58.25)$$

This reduces overfitting. Overfitting might occur if the weights follow too closely several similar minibatches. This dropout procedure adds a random component to each minibatch making groups of similar minibatches less likely.

The random r_i^λ nodes that induce dropout are only used in the training bnet Fig.58.2, not in the classification bnet Fig.58.1. We prefer to remove the r_i^λ stochasticity from classification and for Fig.58.1 to act as an average over sampling space of Fig.58.2. Therefore, if weights $w_{i|j}^\lambda$ are obtained for a dropout layer λ in Fig.58.2, then that layer is used in Fig.58.1 with no r_i^λ nodes but with weights $\langle r_i^\lambda \rangle w_{i|j}^\lambda = pw_{i|j}^\lambda$.

Note that dropout adds non-deterministic nodes to a NN, which in their vanilla form only have deterministic nodes.

- **Convolutional Layer**

- **1-dim**

Filter function $\mathcal{F} : \{0, 1, \dots, nf - 1\} \rightarrow \mathbb{R}$.

σ =stride length

For $i \in \{0, 1, \dots, nh(\lambda) - 1\}$, let

$$H_i^\lambda[\sigma] = \sum_{j=0}^{nf-1} h_{j+i\sigma}^{\lambda-1}[\sigma] \mathcal{F}(j) . \quad (58.26)$$

For the indices not to go out of bounds in Eq.(58.26), we must have

$$nh(\lambda - 1) - 1 = nf - 1 + (nh(\lambda) - 1)\sigma \quad (58.27)$$

so

$$nh(\lambda) = \frac{1}{\sigma}[nh(\lambda - 1) - nf] + 1 . \quad (58.28)$$

- **2-dim**

$h_i^\lambda[\sigma]$ becomes $h_{(i,j)}^\lambda[\sigma]$. Do 1-dim convolution along both i and j axes.

- **Pooling Layers (MaxPool, AvgPool)**

Here each node i of layer λ is impinged by arrows from a subset $Pool(i)$ of the set of all nodes of the previous layer $\lambda - 1$. Partition set $\{0, , 1, \dots, nh(\lambda - 1) - 1\}$ into $nh(\lambda)$ mutually disjoint, nonempty sets called $Pool(i)$, where $i \in \{0, 1, \dots, nh(\lambda) - 1\}$.

- AvgPool

$$H_i^\lambda[\sigma] = \frac{1}{|Pool(i)|} \sum_{j \in Pool(i)} h_j^{\lambda-1}[\sigma] \quad (58.29)$$

- MaxPool

$$H_i^\lambda[\sigma] = \max_{j \in Pool(i)} h_j^{\lambda-1}[\sigma] \quad (58.30)$$

58.4 Autoencoder NN

If the sequence

$$nx, nh(0), nh(1), \dots, nh(\Lambda - 2), ny \quad (58.31)$$

first decreases monotonically up to layer λ_{min} , then increases monotonically until $ny = nx$, then the NN is called an **autoencoder NN**. Autoencoders are useful for unsupervised learning and feature reduction. In this case, Y estimates x . The layers before layer λ_{min} are called the **encoder**, and those after λ_{min} are called the **decoder**. Layer λ_{min} is called the **code**.

Chapter 59

Noisy-OR gate

The Noisy-OR gate was first proposed by Judea Pearl in his 1988 book Ref.[44].

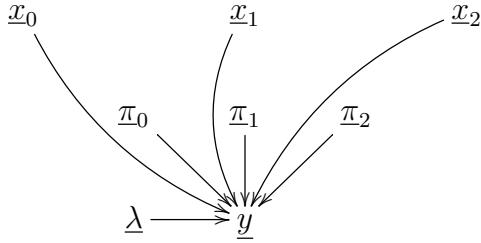


Figure 59.1: Noisy-OR gate $\underline{y} \in \{0, 1\}$ with $n = 3$, Boolean inputs $(\underline{x}_i)_{i=0,1,2}$ and parameters $\underline{\lambda}, (\underline{\pi})_{i=0,1,2}$.

Let

$\underline{\lambda} \in [0, 1] = \text{gate lea.k.a.ge.}$

$\underline{y} \in \{0, 1\} = \text{gate output}$

$\underline{x}^n = (\underline{x}_i)_{i=0,1,\dots,n-1}$, where $\underline{x}_i \in \{0, 1\}$ are gate inputs.

$\underline{\pi}^n = (\underline{\pi}_i)_{i=0,1,\dots,n-1}$, where $\underline{\pi}_i \in [0, 1]$ are gate parameters.

The TPM, printed in blue, for the Noisy-OR gate \underline{y} shown in Fig.59.1, is as follows.

$$P(\underline{y} = 1 | \underline{x}^n, \underline{\lambda}, \underline{\pi}^n) = 1 - (1 - \underline{\lambda}) \prod_i [1 - \underline{\pi}_i \underline{x}_i] \quad (59.1)$$

$$P(\underline{y} = 0 | \underline{x}^n, \underline{\lambda}, \underline{\pi}^n) = 1 - P(\underline{y} = 1 | \underline{x}^n, \underline{\lambda}, \underline{\pi}^n) \quad (59.2)$$

Note that if $\lambda = 0$ and $\pi_i = 1$ for all i , then this becomes a deterministic OR-gate. Indeed,

$$P(\underline{y} = 1 | \underline{x}^n, \lambda = 0, \underline{\pi}^n = 1^n) = 1 - \prod_i [1 - \underline{x}_i] = \vee_{i=0}^{n-1} \underline{x}_i , \quad (59.3)$$

so

$$P(y|x^n, \lambda = 0, \pi^n = 1^n) = \delta(y, \vee_{i=0}^{n-1} x_i) . \quad (59.4)$$

59.1 3 ways to interpret the parameters π_i

1. Note that if $\lambda = 0$ and x^n is one hot (i.e., $x^n = e_i^n$, where e_i^n is the vector with all components zero except for the i -th component which equals 1), then

$$P(y = 1|x^n = e_i^n, \lambda = 0, \pi^n) = 1 - [1 - \pi_i] = \pi_i . \quad (59.5)$$

This gives an interpretation to the parameters π_i .

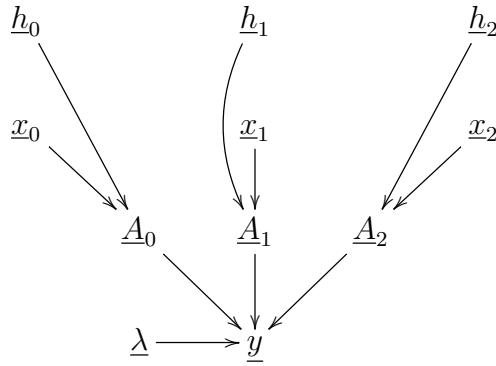


Figure 59.2: Fig.59.1 after replacing parameters $(\pi_i)_{i=0,1,2}$ by hidden nodes $(h_i)_{i=0,1,2}$.

2. Another way of interpreting the parameters π_i is to associate each of them with a hidden variable $h_i \in \{0, 1\}$ whose average equals π_i . More precisely, consider Fig.59.2.

Let $x_i, h_i, A_i, y \in \{0, 1\}$.

The TPMs, printed in blue, for the bnet Fig.59.2, are as follows:

$$P(h_i) = \pi_i \delta(h_i, 1) + (1 - \pi_i) \delta(h_i, 0) \quad (59.6)$$

$$P(A_i|h_i, x_i) = \delta(A_i, h_i \wedge x_i) = \delta(A_i, h_i x_i) \quad (59.7)$$

$$P(y = 1|A^n) = 1 - (1 - \lambda) \wedge_{i=0}^{n-1} \bar{A}_i \quad (59.8)$$

$$= 1 - (1 - \lambda) \prod_i (1 - A_i) \quad (59.9)$$

$$P(y = 0|A^n) = 1 - P(y = 1|A^n) \quad (59.10)$$

Note that

$$P(y = 1|x^n, \lambda) = \sum_{h^n} \sum_{A^n} \left[1 - (1 - \lambda) \prod_i (1 - A_i) \right] [\prod_i \delta(A_i, h_i x_i)] P(h^n) \quad (59.11)$$

$$= E_{\underline{h}^n} \left[[1 - (1 - \lambda) \prod_i (1 - h_i x_i)] \right]. \quad (59.12)$$

But

$$E_{\underline{h}_i}[h_i x_i] = \sum_{h_i=0,1} P(h_i) h_i x_i = \pi_i x_i \quad (59.13)$$

so

$$P(y = 1|x^n, \lambda) = 1 - (1 - \lambda) \prod_i (1 - \pi_i x_i). \quad (59.14)$$

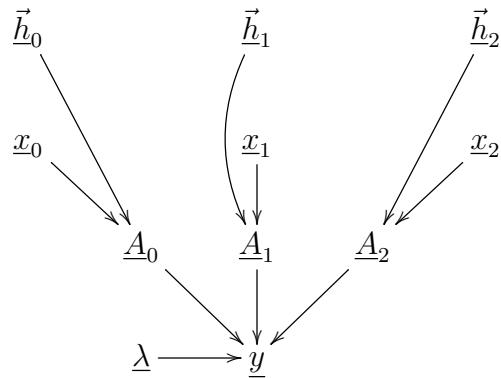


Figure 59.3: Fig.59.2 after replacing the hidden nodes $(\underline{h}_i)_{i=0,1,2}$ by vectors of samples $(\vec{h}_i)_{i=0,1,2}$.

3. Another way to interpret the parameters π_i is to associate each of them with a vector of samples \vec{h}_i whose average is π_i . More precisely, consider Fig.59.3.

Suppose $\underline{h}_i \in \{0, 1\}$ and define

$$P_{\underline{h}_i}(h_i) = \pi_i \delta(h_i, 1) + (1 - \pi_i) \delta(h_i, 0) . \quad (59.15)$$

Suppose $\vec{h}_i = (\underline{h}_i[\sigma])_{s=0,1,\dots,n_{sam}-1}$ and the Boolean samples $\underline{h}_i[\sigma] \in \{0, 1\}$ are i.i.d. with $\underline{h}_i[\sigma] \sim P_{\underline{h}_i}$ for all σ .

Note that for each i , an estimate $\widehat{P}_{\underline{h}_i}(h_i)$ of $P_{\underline{h}_i}(h_i)$ can be obtained from the vector of samples \vec{h}_i as follows:

$$\widehat{P}_{\underline{h}_i}(h_i) = \frac{1}{n_{sam}} \sum_{\sigma=0}^{n_{sam}-1} \mathbb{1}(\underline{h}_i[\sigma] = h_i) . \quad (59.16)$$

Let $x_i, \underline{h}_i[\sigma], A_i, y \in \{0, 1\}$.

The TPMs, printed in blue, for the bnet Fig.59.3, are as follows:

$$P(\vec{h}_i) = \prod_{\sigma=0}^{n_{sam}-1} P_{\underline{h}_i}(\underline{h}_i[\sigma]) \quad (59.17)$$

$$P(A_i | \vec{h}_i, x_i) = \delta(A_i, \frac{1}{n_{sam}} \sum_{\sigma} \underline{h}_i[\sigma] \wedge x_i) \quad (59.18)$$

$$= \delta(A_i, \pi_i x_i) \quad (59.19)$$

$$P(y = 1 | A^n) = 1 - (1 - \lambda) \wedge_{i=0}^{n-1} \overline{A}_i \quad (59.20)$$

$$= 1 - (1 - \lambda) \prod_i (1 - A_i) \quad (59.21)$$

$$P(y = 0 | A^n) = 1 - P(y = 1 | A^n) \quad (59.22)$$

Note that

$$P(y = 1 | x^n, \lambda, \vec{h}^n) = \sum_{A^n} \left[1 - (1 - \lambda) \prod_i (1 - A_i) \right] \prod_i \delta(A_i, \pi_i x_i) \quad (59.23)$$

$$= 1 - (1 - \lambda) \prod_i (1 - \pi_i x_i) . \quad (59.24)$$

Chapter 60

Non-negative Matrix Factorization

Based on Ref.[130].

Given matrix V , factor it into product of two matrices

$$V = WH , \quad (60.1)$$

where all 3 matrices have non-negative entries.

$V \in \mathbb{R}_{\geq 0}^{nv \times na}$: visible info matrix

$W \in \mathbb{R}_{\geq 0}^{nv \times nh}$: weight info matrix

$H \in \mathbb{R}_{\geq 0}^{nh \times na}$: hidden info matrix

Usually, $nv > nh < na$ so compression of information (a.k.a. dimensional reduction, clustering)

60.1 Bnet interpretation

Express node \underline{v} as a chain of two nodes.

$$\underline{v} \longleftarrow \underline{a} \quad = \quad \underline{w} \longleftarrow \underline{h} \longleftarrow \underline{a}$$

Figure 60.1: Bnet interpretation of non-negative matrix factorization.

The TPMs, printed in blue, for bnet Fig.60.1, are as follows.

$$P(\underline{v} = w | \underline{a}) = \frac{V_{w,a}}{\sum_w V_{w,a}} \quad (60.2)$$

$$P(w|h) = \frac{W_{w,h}}{\sum_w W_{w,h}} \quad (60.3)$$

$$P(h|a) = \frac{\sum_w W_{w,h}}{\sum_w V_{w,a}} H_{h,a} \quad (60.4)$$

60.2 Simplest recursive algorithm

Initialize: Choose nh . Choose $W^{(0)}$ and $H^{(0)}$ that have non-negative entries.

Update: For $n = 0, 1, \dots$, do

$$H_{i,j}^{(n+1)} \leftarrow H_{i,j}^{(n)} \frac{[(W^{(n)})^T V]_{i,j}}{[(W^{(n)})^T \underbrace{W^{(n)} H^{(n)}}_{\approx V}]_{i,j}} \quad (60.5)$$

and

$$W_{i,j}^{(n+1)} \leftarrow W_{i,j}^{(n)} \frac{[V(H^{(n+1)})^T]_{i,j}}{\underbrace{[W^{(n)} H^{(n+1)} (H^{(n+1)})^T]_{i,j}}_{\approx V}}. \quad (60.6)$$

After each step, record error defined by

$$\mathcal{E}^{(n)} = \| V - W^{(n)} H^{(n)} \|_2. \quad (60.7)$$

Using 2-norm, a.k.a. Frobenius matrix norm. Continue until reach acceptable error.

Can also use Kullback-Leibler divergence for error:

$$\mathcal{E} = \sum_a P(a) D_{KL}(P(\underline{v} = w|a) \| \sum_h P(w|h)P(h|a)), \quad (60.8)$$

for some arbitrary choice of prior $P(a)$. For example, can choose $P(a)$ uniform.

Chapter 61

Observationally Equivalent DAGs

This chapter is based on Chapter 1 of Ref.[46] and on a blog post by Bruno Gonçalves (Ref.[13]).

A probability distribution P is **compatible with a DAG** G if P and G have the same random variables, and they can be combined to form a bnet without contradictions; i.e., one can calculate all the TPMs from P and multiply them together to obtain P again. Let

$$\mathcal{P}(G) = \{P : P \text{ is compatible with } G\} . \quad (61.1)$$

Two DAGs G and G' are observationally equivalent (OE) if $\mathcal{P}(G) = \mathcal{P}(G')$. Hence, any total probability distribution that is compatible with one of them is compatible with the other. For example, $\underline{a} \rightarrow \underline{b}$ and $\underline{a} \leftarrow \underline{b}$ are OE because

$$P(a|b)P(b) = P(a, b) = P(b|a)P(a) . \quad (61.2)$$

We'll say two bnets are OE if their DAGs are OE.

Two DAGs G and G' are **d-separation equivalent** if $DS(G) = DS(G')$. See Chapter 21 for definition of $DS(G)$.

Claim 85 *Two DAGs are OE iff their DAGs are d-separation equivalent.*

The **skeleton** of a DAG is its underlying undirected graph.

A **v-structure** in a DAG consists of two arrows converging to a node and such that their tails are not connected by a third arrow. Fig.61.1 shows in red all the v-structures of a particular DAG.

Claim 86 *Observational Equivalence Theorem (by Verma and Pearl, 1990)*

Two DAGs are OE iff they have the same skeletons and the same v-structures.

61.1 Examples

The 3 DAGs in Fig.61.2 are OE. They form an equivalence class of OE DAGs that represent the same probability distribution. This equivalence class of DAGs can be

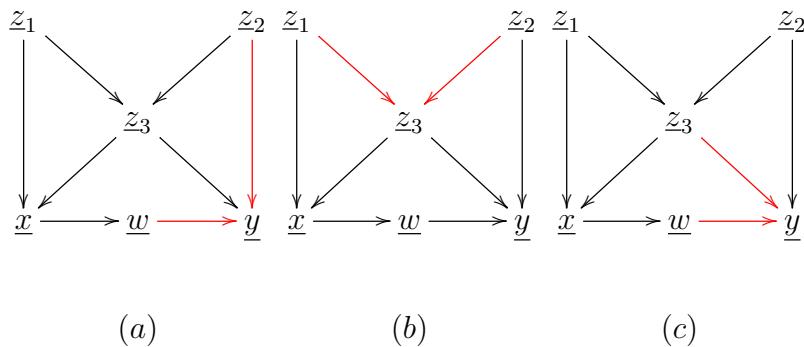


Figure 61.1: Example showing in red all v-structures of a particular DAG.

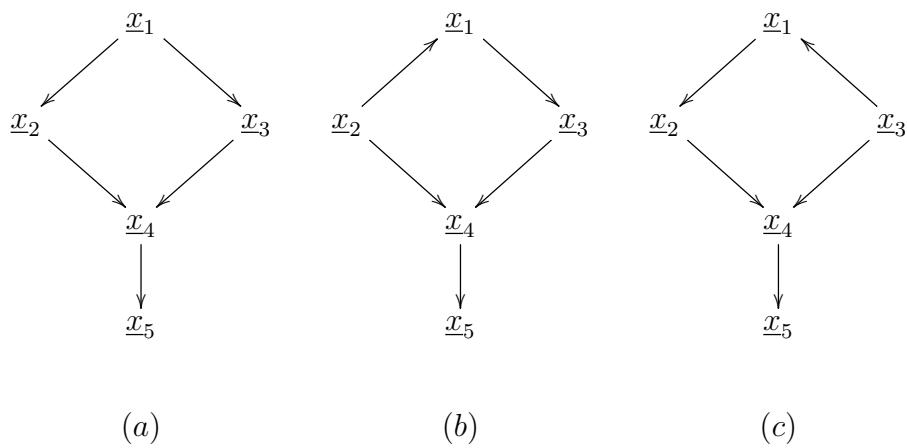


Figure 61.2: These 3 DAGs are observationally equivalent (OE).

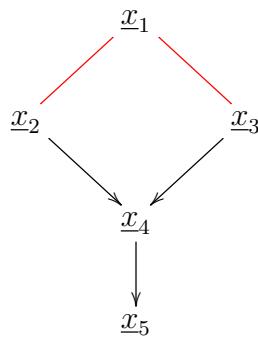


Figure 61.3: This partially directed graph represents the 3 DAGs in Fig.61.2.

represented by the partially directed graph Fig.61.3. These 3 DAGs can be proven to be OE in the following 3 ways:

1. Write the generic probability distributions represented by the 3 DAGs, and show that they are equal, as we did in Eq.(61.2). That is the low brow way of proving OE.
2. Use d-separation (see Chapter 21). Consider DAG (a) first. Rename the nodes as $\underline{\tau}_j$ with $j = 1, 2, \dots$ so that the names are in topological order (i.e., so that the parents of $\underline{\tau}_j$ have indices that are smaller than j). The node names x_j of DAG (a) are already in topological order, so we skip this step for DAG (a). Now write down its total probability distribution and notice which parents of a fully connected DAG were omitted.

$$P(x_1, x_2, x_3, x_4, x_5) = \underbrace{P(x_5|x_4)}_{x_3, x_2, x_1 \text{ omitted}} \underbrace{P(x_4|x_3, x_2)}_{x_1 \text{ omitted}} \underbrace{P(x_3|x_1)}_{x_2 \text{ omitted}} P(x_2|x_1)P(x_1) \quad (61.3)$$

The observations of which parents were omitted can be stated in d-separation lingo as the following 3 orthogonality relations:¹

$$\underline{x}_3 \perp_P \underline{x}_2 \mid \underline{x}_1 \quad (61.4a)$$

$$\underline{x}_4 \perp_P \underline{x}_1 \mid \underline{x}_2, \underline{x}_3 \quad (61.4b)$$

$$\underline{x}_5 \perp_P (\underline{x}_1, \underline{x}_2, \underline{x}_3) \& \mid \underline{x}_4. \quad (61.4c)$$

Going through the same procedure for the other 2 DAGs yields, for each of them, an equivalent set of 3 orthogonality equations.²

This is enough to conclude that the 3 DAGs of Fig.61.2 are OE.

Note that Eqs.(61.4) encompass all that there is to say about the observability of DAG (a). These 3 equations can be checked empirically to assess how well the DAG fits the data. For example, one can do OLS (ordinary least squares) regression $\underline{x}_5 \sim \underline{x}_1 + \underline{x}_2 + \underline{x}_3 + \underline{x}_4$ on the data, i.e., try to fit $x_5 = \beta_0 + \sum_{i=1}^4 \beta_i x_i$ to the data, and find that, to a good approximation, $\beta_1 = \beta_2 = \beta_3 = 0$.

3. Use the OE Theorem. All three DAGs have the same skeleton, and the same single v-structure $\underline{x}_2 \rightarrow \underline{x}_4 \leftarrow \underline{x}_3$.

¹Normally, if we had changed from the original node names to the $\underline{\tau}_j$ node names, these orthogonality relations would first be stated in terms of the $\underline{\tau}_j$ names, and we could translate them so that they were stated in terms of the original node names. But for DAG (a) there was no need to use the $\underline{\tau}_j$ names.

²The \underline{x}_j node names are no longer in topological order for DAGs (b) and (c) so for them you should go through the intermediate step of renaming the nodes $\underline{\tau}_j$, and then, after obtaining the orthogonality relations in terms of the $\underline{\tau}_j$ names, translating them back to the original \underline{x}_j names.

Chapter 62

Personalized Treatment Effects

This chapter is based on the work of Pearl et al, as reported in Refs. [63] and [35].

Recall from Chapter 64 on Potential Outcomes (PO) and Beyond, that the **Average Treatment Effect (ATE)** is defined as

$$ATE = E[\underline{y}_1 - \underline{y}_0] \quad (62.1)$$

$$= P(\underline{y}_1 = 1) - P(\underline{y}_0 = 1) \quad (62.2)$$

The **Conditional ATE (CATE)** is defined as the conditional expected value

$$ATE_z = E_{|z}[\underline{y}_1 - \underline{y}_0] \quad (62.3)$$

$$= P(\underline{y}_1 = 1|z) - P(\underline{y}_0 = 1|z) . \quad (62.4)$$

Note that

$$ATE = \sum_z P(z)ATE_z \quad (62.5)$$

Personalized Treatment Effect (PTE) theory as envisioned by Pearl is the study of bounds for “personalized” treatment effects such as

$$PNS = P(\underline{y}_1 - \underline{y}_0 = 1) \quad (62.6)$$

and

$$PNS_z = P(\underline{y}_1 - \underline{y}_0 = 1|z) . \quad (62.7)$$

They are said to be **personalized** because they are averages over a *single* ensemble (i.e., population) of individuals σ with probability $P(y_0^\sigma, y_1^\sigma)$. ATE , on the other hand, is not personalized, because it equals the difference of two of those averages.

If the conditioning z is fine grained enough to pick out a single individual σ (i.e., if $z = \sigma$), then we get

$$PNS_\sigma = \mathbb{1}(y_1^\sigma - y_0^\sigma = 1) \quad (62.8)$$

whereas

$$ATE_\sigma = \mathbb{1}(y_1^\sigma = 1) - \mathbb{1}(y_0^\sigma = 1) \quad (62.9)$$

ATE and *PNS* measure different things. *PNS* measures the probability that a *single person* will switch outcomes from 0 to 1 when he/she switches treatments from 0 to 1. *ATE* measures the difference in populations between those who survive taking the drug and those who survive without it.

One very promising field in which PTE theory can be applied is in **Personalized Causal Medicine**. For example, suppose we want to use PNS_z , where the conditioning is on the sex of the patient (i.e., $z = \text{male}, \text{female}$), to advice a female patient whether to take a cancer drug or not.

62.1 Goal, Strategy and Rationale of PTE theory

In this section, we described briefly the goal, strategy and rationale behind PTE theory.

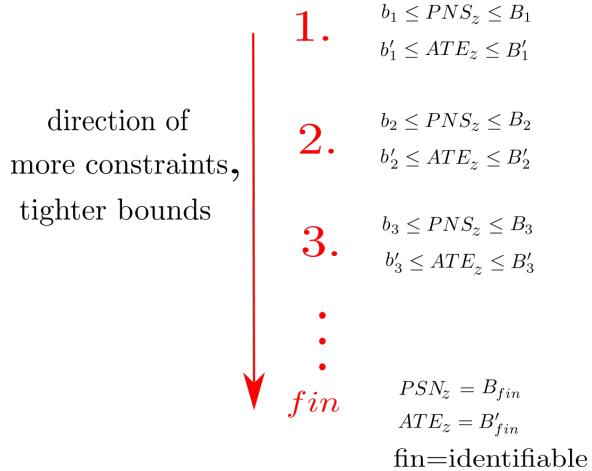


Figure 62.1: We use patient data to calculate at each step, increasingly tighter bounds b_j, b'_j, B_j, B'_j for PNS_z and ATE_z , ending in point bounds for both quantities.

The goal of PTE theory, as described by Fig.62.1, is to find increasingly tighter bounds for PTEs such as *PNS*, PNS_z , *ATE*, ATE_z , etc..

We say bounds, because it is not always possible to give a point estimate for *PNS* and other PTEs. If a point estimate for a PTE named Q is achievable, we say Q is **identifiable**. The same definition of identifiability has been used before in this book for *ATE*. It's possible that *ATE* is identifiable, but *PNS* isn't, or vice versa, for certain kinds of data; i.e., both quantities need not become identifiable simultaneous at the same step above.

The bounds given by PTE theory are as tight as possible, depending on the available data, and on what bnet model assumptions the user is willing to make. We

will consider two types of bounds: (1) Bounds for an unspecified bnet, (2) bounds for specific bnet families. Bounds for (2) will be tighter than bounds for (1).

The bounds are calculated from two types of data: **Observational Data (OD)** and **Experimental Data (ED)**.

For OD, one allows the patient to choose whether to take a drug or not ($x = 0, 1$), and then we conduct a **survey** to record his/her value for x and whether the treatment worked or not ($y = 0, 1$).

For ED, one conducts a **RCT** (Randomized Controlled Trial) instead of a survey. In the case of ED, we record, as in OD, the (x, y) for each patient, but the value of x for each patient is selected by the experimenter, at random, and, once selected, it is compulsory for the patient.

Unlike ED, OD is likely to be confounded, but it can still shed additional information that serves to tighten the bounds on PNS_z or other PTEs.

OD is usually collected first because it is easier and cheaper to collect than ED. Sometimes ED is too expensive or difficult or even impossible to collect, so only OD is available. Sometimes several ED (resp., several OD) need to be merged, before merging the merged ED with the merged OD. Pearl's PTE theory allows us to fuse together all the available data in all these types of situations.

Some purists such as the advocates of **EBM (Evidence Based Medicine)** advocate the use of only ED (i.e., only RCTs), no OD. This is reasonable in certain administrative professions to keep partisanship and chicanery out of the decision making process, but in other professions, throwing away OD would be foolish. A case in the history of medicine where OD was essential, was the case of John Snow (see Chapter 16). John Snow didn't have an RCT, but he was able to use OD to determine the driver of the cholera epidemic in London. He saved countless lives by doing so. An EBM purist would have thrown out his OD because it wasn't an RCT.

In the usual case, OD is collected first to aid in the design of an RCT, and then an RCT is conducted to collect ED. In this case, an EBM purist would throw away the OD, and only calculate an estimate of ATE. What PTE theory suggests is to calculate *both*, an estimate of ATE and bounds for PNS. Why? Because ATE and PNS measure different things. ATE utilizes only ED whereas PNS utilizes both OD and ED. Also, PNS is more personalized than ATE.

An RCT is called that because one chooses "at random", a subset Σ of a huge population Σ_∞ , and then one splits Σ , again at random, into 2 subsets, $\Sigma_{control}$ and $\Sigma_{treated}$. All patients in $\Sigma_{control}$ (resp., $\Sigma_{treated}$) are not treated (resp., treated). **Randomization** (i.e. choosing Σ , $\Sigma_{control}$, $\Sigma_{treated}$), if perfect, averages out (i.e., cancels out) the effect of all confounders. The problem is that in practice, ideal randomization is hard to achieve. Hence, ideal RCTs are hard to achieve (and costly), plus they don't shed light on the confounders and mechanism involved. By mechanism, I mean a DAG.¹ A DAG is a hypothesis, and RCTs don't test it, whereas observational

¹Note that DAGs are not unique. Some are a better causal fit than others for the physical situation being considered. See Chapter 33 for a discussion of Goodness of Causal Fit.

studies do. So it's a good idea to do a series of well designed observational studies (much cheaper than an RCT), with a hypothesis DAG, before doing the RCT. This elucidates the DAG, and helps design an RCT that achieves a good approximation to randomization.

Some people object to CI (Causal Inference) in general on the grounds that the causal DAGs are adhoc, arbitrary, a sort of unscientific voodoo. I think it's because they fail to grasp the true meaning and purpose of DAGs. I discuss this further in Section E.6.

62.2 Bnets for PTE theory

Let $\bar{0} = 1$ and $\bar{1} = 0$.

Whenever we write $P(\underline{a} = \underline{x}, b)$, we mean $P(\underline{a} = \underline{x}, \underline{b} = b)$.

In this chapter, we will not use the notation $P(y)$ and $P(y')$ used by Pearl to discuss PTE theory. Instead, we will use $P(\underline{y} = 1)$ and $P(\underline{y} = 0)$ to denote his $P(y)$ and $P(y')$, respectively.

On the other hand, in this chapter we will change the names of variables $\underline{d}, \underline{y}, \underline{z}, \underline{y}(d)$ used in Chapter 64 on Potential Outcomes (PO) to the names favored by Pearl. Hence, we will replace $\underline{d}, \underline{y}, \underline{x}, \underline{y}(d)$ by $\underline{x}, \underline{y}, \underline{z}, \underline{y}_x$, respectively.

PTE theory considers bnets of the form Fig.62.2. The bnet considered in Rubin's PO theory is a very simple special case of this where the box labeled "multiple nodes" is absent and there is an arrow pointing from \underline{z} to \underline{x} .

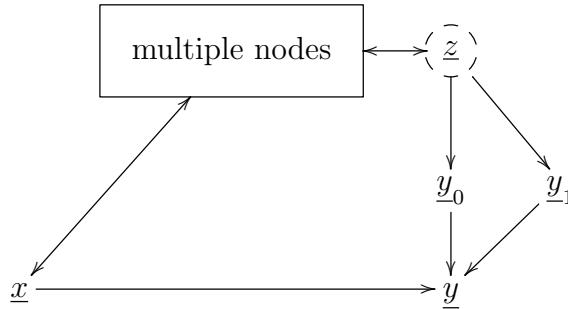


Figure 62.2: Type of Bnet considered in PTE theory. The box labeled "multiple nodes" contains various observed and hidden nodes with arrows to or from node \underline{x} and to or from node \underline{z} . \underline{z} can be a multinode. \underline{z} is shown as hidden but could be observed instead.

The TPM, printed in blue, for node \underline{y} of bnet Fig.62.2 is as follows:

$$P(\underline{y}|\underline{y}_0, \underline{y}_1, \underline{x}) = \delta(\underline{y}, \underline{y}_x) \quad (62.10)$$

This TPM is used frequently in PTE theory. If x, y_0, y_1 are arguments of $P()$, this TPM implies that one can swap $\underline{y}_x = y_x$ and $\underline{y} = y_x$ inside $P()$. For example,

$$P(y_0, y_1, x) = P(\underline{\textcolor{red}{y}}_{\underline{x}} = y_x, \underline{y}_{\bar{x}}, x) \quad (62.11)$$

$$= P(\underline{\textcolor{red}{y}} = y_x, \underline{y}_{\bar{x}}, x) \quad (62.12)$$

According to Pearl, the defining property of y_x is that

$$P(\underline{y}_x = y) = P(\underline{y} = y | \mathcal{D}\underline{x} = x) \quad (62.13)$$

Pearl likes to call Eq.(62.13) the 1st Law. The 1st Law is also a consequence of bnet Fig.62.2 and the TPM Eq.(62.10). Indeed, $\mathcal{D}\underline{x} = x$ means one should amputate all arrows entering node \underline{x} , and one should set the TPM of \underline{x} to a delta function centered at x . If that is done, then the values of \underline{y} and \underline{y}_x must be equal, because the TPM at node \underline{y} is a delta function that enforces this equality.

62.3 $ATE = PNS - AMM$

Define the **Probability of Necessity and Sufficiency**

$$PNS = P(\underline{y}_0 = 0, \underline{y}_1 = 1) \quad (62.14)$$

and the **Amonotonicity Measure** by

$$AMM = P(\underline{y}_0 = 1, \underline{y}_1 = 0) \quad (62.15)$$

Claim 87

$$PNS = P(\underline{y}_1 - \underline{y}_0 = 1) \quad (62.16)$$

$$AMM = P(\underline{y}_1 - \underline{y}_0 = -1) \quad (62.17)$$

proof:

$$\underline{y}_1 - \underline{y}_0 = 1 \text{ iff } (\underline{y}_1 = 1 \text{ and } \underline{y}_0 = 0).$$

$$\underline{y}_1 - \underline{y}_0 = -1 \text{ iff } (\underline{y}_1 = 0 \text{ and } \underline{y}_0 = 1).$$

QED

Claim 88

$$ATE = PNS - AMM \quad (62.18)$$

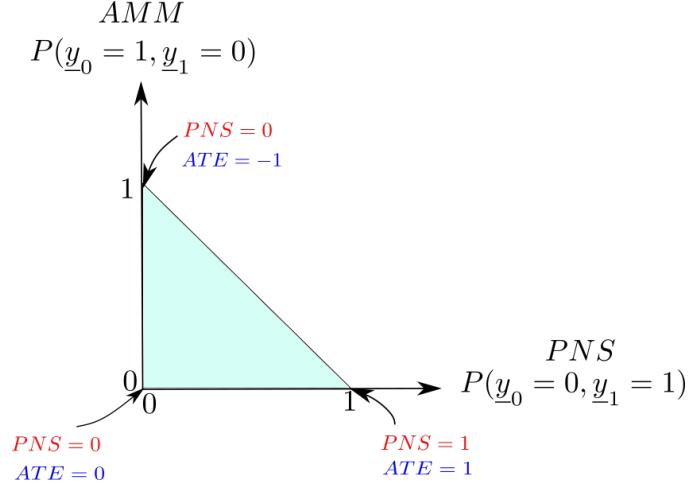


Figure 62.3: Probability simplex $\{(x, y) : x \geq 0, y \geq 0, x+y \leq 1\}$ with $x = PNS$ and $y = AMM$. Values of PNS and ATE at corners are given. Points on line $x+y=1$ can only be achieved if $P(1,1)=P(0,0)=0$.

proof:

$$ATE = E_\sigma[y_1^\sigma - y_0^\sigma] \quad (62.19)$$

$$= E[y_1 - y_0] \quad (62.20)$$

$$= \sum_y y [P(\underline{y}_1 = y) - P(\underline{y}_0 = y)] \quad (62.21)$$

$$= P(\underline{y}_1 = 1) - P(\underline{y}_0 = 1) \quad (62.22)$$

$$= \sum_{y_0} P(y_0, \underline{y}_1 = 1) - \sum_{y_1} P(\underline{y}_0 = 1, y_1) \quad (62.23)$$

$$= \underbrace{P(\underline{y}_0 = 0, \underline{y}_1 = 1)}_{PNS} - \underbrace{P(\underline{y}_0 = 1, \underline{y}_1 = 0)}_{AMM} \quad (62.24)$$

QED

See Fig.62.3 for a visualization of the extreme values of PNS and the corresponding values of ATE .

62.4 Probabilities Relevant to PTE theory

Note²

Let $x, y \in \{0, 1\}$ and $z \in S_z$ for some finite, not necessarily binary set S_z .

²To translate this section from our notation to the notation used by Tian and Pearl in Ref.[63], replace $P(\underline{y}_1 = i, \underline{y}_0 = j, \underline{x} = k) \rightarrow p_{i,j,k}$, $O_{1,0} \rightarrow P(x, y')$, $O_{0|1} \rightarrow P(y'|x)$, $E_{0|1} \rightarrow P(y'_x)$, etc.

Define

$$P_{x,y}^{x',y'} = P(\underline{y}_{x'} = y' | x, y) \quad (62.25)$$

observational (non-causal) probabilities

$$O_{x,y} = P(\underline{x} = x, \underline{y} = y) \quad (62.26)$$

$$O_{y|x} = P(\underline{y} = y | \underline{x} = x) \quad (62.27)$$

$$\pi_x = P(\underline{x} = x) \quad (62.28)$$

experimental (causal) probabilities

$$E_{y|x} = P(\underline{y}_x = y) \quad (62.29)$$

Average Treatment Effect (ATE)

$$ATE = P(\underline{y}_1 = 1) - P(\underline{y}_0 = 1) \quad (62.30)$$

$$= E_{1|1} - E_{1|0} \quad (62.31)$$

$$= E_{1|1} + E_{0|0} - 1 \quad (62.32)$$

Conditional ATE (CATE)

$$ATE_z = P(\underline{y}_1 = 1 | z) - P(\underline{y}_0 = 1 | z) \quad (62.33)$$

$$= E_{1|1,z} + E_{0|0,z} - 1 \quad (62.34)$$

Average Causal Effect

$$ACE = P(\underline{y} = 1 | \mathcal{D}\underline{x} = 1) - P(\underline{y} = 1 | \mathcal{D}\underline{x} = 0) \quad (62.35)$$

Note that

$$ACE = \underbrace{P(\underline{y} = 1 | \mathcal{D}\underline{x} = 1)}_{P(\underline{y}_1 = 1)} - \underbrace{P(\underline{y} = 1 | \mathcal{D}\underline{x} = 0)}_{P(\underline{y}_0 = 1)} = ATE \quad (62.36)$$

Conditional ACE (CACE)

This is what is called ACE_z in Chapter 64. Note that $ACE_z = ATE_z$.

Effect of Treatment on the Treated (ETT) (i.e., ATE for the treated)

$$ETT = \underbrace{P(\underline{y}_1 = 1 | \underline{x} = 1)}_{\mathcal{E}_1} - \underbrace{P(\underline{y}_0 = 1 | \underline{x} = 1)}_{\mathcal{E}_0} \quad (62.37)$$

Note that

$$\mathcal{E}_1 \pi_1 = P(\underline{y}_1 = 1, \underline{x} = 1) \quad (62.38)$$

$$= O_{1,1} \quad (62.39)$$

and

$$\mathcal{E}_0\pi_1 = P(\underline{y}_0 = 1, \underline{x} = 1) \quad (62.40)$$

$$= P(\underline{y}_0 = 1) - \underbrace{P(\underline{x} = 0, \underline{y}_0 = 1)}_{P(\underline{x}=0, \underline{y}=1)} \quad (62.41)$$

$$= E_{1|0} - O_{0,1} \quad (62.42)$$

so

$$ETT\pi_1 = \sum_x O_{x,1} - E_{1|0} \quad (62.43)$$

Probability of Necessity (PN)³

$$PN = P_{1,1}^{0,0} \quad (62.44)$$

Probability of Sufficiency (PS)⁴

$$PS = P_{0,0}^{1,1} \quad (62.45)$$

Probability of Necessity and Sufficiency (PNS)

$$PNS = P(\underline{y}_0 = 0, \underline{y}_1 = 1) \quad (62.46)$$

Henceforth, we will use $PNS3$ to denote the trio

$$PNS3 = (PNS, PN, PS) . \quad (62.47)$$

Amonotonicity Measure (AMM)

$$AMM = P(\underline{y}_0 = 1, \underline{y}_1 = 0) \quad (62.48)$$

Risk ratio or relative risk (RR)

$$RR = \frac{O_{1|1}}{O_{1|0}} \quad (62.49)$$

Excess Risk Ratio (ERR)

$$ERR = \frac{O_{1|1} - O_{1|0}}{O_{1|1}} = 1 - \frac{1}{RR} \quad (62.50)$$

³I like to call PN the Probability of Nullifying, because it goes from 11 to 00

⁴I like to call PS the Probability of Surging, because it goes from 00 to 11

Corrected ERR (CERR)

$$CERR = ERR + \frac{O_{1|0} - E_{1|0}}{O_{1,1}} \quad (62.51)$$

You might be wondering how $P(\underline{y}_x = y | \underline{x} = 0)$ and $P(\underline{y}_x = y | \underline{x} = 1)$ for $x, y \in \{0, 1\}^2$ are related to $O_{y|x}$ and $E_{y|x}$. The following claim shows how.

Claim 89

$$P(\underline{y}_x = y | \underline{x} = x) = O_{y|x} \quad (62.52a)$$

$$P(\underline{y}_x = y | \underline{x} = \bar{x}) = \frac{E_{y|x} - O_{y|x}\pi_x}{\pi_{\bar{x}}} \quad (62.52b)$$

proof: Before we begin the proof, note that summing both sides of Eqs.(62.52) gives $1 = 1$, so these 2 equations pass that test.

$$P(\underline{y}_x = y | \underline{x} = x) = P(\underline{y} = y | \underline{x} = x) = O_{y|x} \quad (62.53)$$

$$P(\underline{y}_x = y | \underline{x} = \bar{x}) = \frac{P(\underline{y}_x = y, \underline{x} = \bar{x})}{\pi_{\bar{x}}} \quad (62.54)$$

$$= \frac{P(\underline{y}_x = y) - P(\underline{y}_x = y, \underline{x} = x)}{\pi_{\bar{x}}} \quad (62.55)$$

$$= \frac{P(\underline{y}_x = y) - P(\underline{y}_x = y | \underline{x} = x)\pi_x}{\pi_{\bar{x}}} \quad (62.56)$$

$$= \frac{E_{y|x} - O_{y|x}\pi_x}{\pi_{\bar{x}}} \quad (62.57)$$

QED

Claim 90

$$P(y_0, y_1 | x) = P(y_{\bar{x}} | x, \underline{y} = y_x)P(\underline{y} = y_x | x) \quad (62.58)$$

$$= \begin{array}{c} y_{\bar{x}} \\ \nearrow \\ x \longrightarrow \underline{y} = y_x \end{array} \quad (62.59)$$

proof:

$$P(y_0, y_1|x) = \frac{P(y_0, y_1, x)}{P(x)} \quad (62.60)$$

$$= \frac{P(y_{\bar{x}}, x, \underline{y} = y_x)}{P(x)} \quad (62.61)$$

$$= \frac{P(y_{\bar{x}}|x, \underline{y} = y_x)P(x, \underline{y} = y_x)}{P(x)} \quad (62.62)$$

$$= P(y_{\bar{x}}|x, \underline{y} = y_x)P(\underline{y} = y_x|x) \quad (62.63)$$

QED

Claim 91

$$\underbrace{PNS}_{P(\underline{y}_0=0, \underline{y}_1=1)} = PN * O_{1,1} + PS * O_{0,0} \quad (62.64)$$

Hence, if we know any two of (PN, PS, PNS) , we can calculate the third.

proof:

$$P(y_0, y_1) = \sum_x P(y_0, y_1|x)P(x) \quad (62.65)$$

$$= \sum_x P(y_{\bar{x}}|x, \underline{y} = y_x)P(x, \underline{y} = y_x) \quad (\text{see Claim 90.}) \quad (62.66)$$

$$= \begin{cases} P(y_1|\underline{x} = 0, \underline{y} = y_0)P(\underline{x} = 0, \underline{y} = y_0) \\ + P(y_0|\underline{x} = 1, \underline{y} = y_1)P(\underline{x} = 1, \underline{y} = y_1) \end{cases} \quad (62.67)$$

Thus,

$$P(\underline{y}_0 = 0, \underline{y}_1 = 1) = \begin{cases} P(\underline{y}_1 = 1|\underline{x} = 0, \underline{y} = 0)P(\underline{x} = 0, \underline{y} = 0) \\ + P(\underline{y}_0 = 0|\underline{x} = 1, \underline{y} = 1)P(\underline{x} = 1, \underline{y} = 1) \end{cases} \quad (62.68)$$

$$= PS * O_{0,0} + PN * O_{1,1} \quad (62.69)$$

QED

Note that PNS refers to both \underline{y}_0 and \underline{y}_1 , PN refers only to \underline{y}_0 and PS refers only to \underline{y}_1 . Thus, PN and PS serve to separate the pair of variables $(\underline{y}_0, \underline{y}_1)$ and to isolate them individually. Claim 91 is a quantitative expression of that separation.

Pearl likes to say that PNS belongs to Rung 3 because it's a probability that involves both \underline{y}_0 and \underline{y}_1 , and one of those two must be a counterfactual (an event that never occurred). On the other hand, PN, PS, ATE , etc., are defined in terms of probabilities that involve either \underline{y}_0 or \underline{y}_1 but not both. Probabilities that involve

only one of them, can be expressed with the do operator, so they belong to Rung 2. Conditional probabilities like $P(y|x)$ that involve neither \underline{y}_0 nor \underline{y}_1 belong to Rung 1.⁵

62.5 Symmetry

Define \sim to be an operator that swaps zeros and ones in $P(\underline{y}_0 = y, \underline{y}_1 = y', x)$. Hence

$$[P(\underline{y}_0 = y, \underline{y}_1 = y', \underline{x} = x)]^\sim = P(\underline{y}_1 = \bar{y}, \underline{y}_0 = \bar{y}', \underline{x} = \bar{x}) \quad (62.70)$$

$$(P_{x,y}^{x',y'})^\sim = P_{\bar{x},\bar{y}}^{\bar{x}',\bar{y}'} \quad (62.71)$$

$$(O_{x,y})^\sim = O_{\bar{x},\bar{y}} \quad (62.72)$$

$$(O_{y|x})^\sim = O_{\bar{y}|\bar{x}} \quad (62.73)$$

$$(\pi_x)^\sim = \pi_{\bar{x}} \quad (62.74)$$

$$(E_{y|x})^\sim = E_{\bar{y}|\bar{x}} \quad (62.75)$$

$$(PN)^\sim = PS, \quad (PS)^\sim = PN \quad (62.76)$$

$$(PNS)^\sim = PNS \quad (62.77)$$

$$(RR)^\sim = \frac{O_{0|0}}{O_{0|1}} \quad (62.78)$$

Recall

$$ERR = \frac{O_{1|1} - O_{1|0}}{O_{1|1}} = 1 - \frac{1}{RR} \quad (62.79)$$

Therefore, define

$$(ERR)^\sim = \frac{O_{0|0} - O_{0|1}}{O_{0|0}} = 1 - \frac{1}{(RR)^\sim} \quad (62.80)$$

Note that

⁵Probabilities such as $P(\underline{y}_0 = 1, \underline{y} = 0, \underline{x} = 1) = P(\underline{y}_0 = 1, \underline{y}_1 = 0, \underline{x} = 1)$ are considered Rung 3.

$$O_{1|1} - O_{1|0} = O_{1|1} + O_{0|0} - 1 \quad (62.81)$$

$$= O_{0|0} - O_{0|1} \quad (62.82)$$

$$= (O_{1|1} - O_{1|0})^\sim \quad (62.83)$$

62.6 Linear Programming Problem

Probability Simplex

$$\mathcal{S} = \left\{ P(y_0, y_1, x) : \begin{array}{l} P(y_0, y_1, x) \geq 0 \\ y_0, y_1, x \in \{0, 1\} \\ \sum_{y_0=0}^1 \sum_{y_1=0}^1 \sum_{x=0}^1 P(y_0, y_1, x) = 1 \end{array} \right\} \quad (62.84)$$

Note that \mathcal{S} has 7 degrees of freedom (dofs).

1. observational (non-causal) constraints on \mathcal{S} (3 constraints)

$$O_{x,y} = P(x, y) = \sum_{y'} P(\underline{y}_x = y, \underline{y}_{\bar{x}} = y', x) \quad \text{for } (x, y) \in \{(0, 1), (1, 0), (1, 1)\} \quad (62.85)$$

2. experimental (causal) constraints on \mathcal{S} (2 constraints)⁶

$$E_{1|x} = P(\underline{y}_x = 1) = \sum_{x'} \sum_{y'} P(\underline{y}_x = 1, \underline{y}_{\bar{x}} = y', x = x') \quad \text{for } x \in \{0, 1\} \quad (62.86)$$

\mathcal{S} has 7 dofs but the 3 observational constraints reduce the number of dofs to 4. The 5 observational and experimental constraints reduce the number of dofs to 2. \mathcal{S} is embedded in \mathbb{R}^8 and then the 6 constraints (unit probability, 2 observational and 3 experimental) reduce it to the interior of a 6 or less sided figure in \mathbb{R}^2 .

Henceforth, we will refer to the observational and experimental constraints together as the **minimal constraints**.

This is half of a linear programming problem. Recall that a **linear programming problem** can be stated as finding the column vector ξ that minimizes a cost $\mathcal{C} = c^T \xi$ subject to $A\xi = b$ and $\xi \geq 0$. Here we have no cost function or minimization, but we have $A\xi = b$ and $\xi \geq 0$ where $\xi = [P(y_0, y_1, x)]_{\forall y_0, y_1, x}$. Also $b = [1, O_{0,1}, O_{1,0}, O_{1,1}, E_{1|0}, E_{1|1}]^T$, and A = a matrix of zeros and ones.

⁶ $E_{0|x}$ follows from $E_{0|x} = 1 - E_{1|x}$.

Note that

$$PNS = P(\underline{y}_0 = 0, \underline{y}_1 = 1) \quad (62.87)$$

$$= \sum_x P(\underline{y}_0 = 0, \underline{y}_1 = 1, x) \quad (62.88)$$

$$PN = P(\underline{y}_0 = 0 | \underline{x} = 1, \underline{y} = 1) \quad (62.89)$$

$$= \frac{P(\underline{y}_0 = 0, \underline{y}_1 = 1, \underline{x} = 1)}{O_{1,1}} \quad (62.90)$$

$$PS = P(\underline{y}_1 = 1 | \underline{x} = 0, \underline{y} = 0) \quad (62.91)$$

$$= \frac{P(\underline{y}_0 = 0, \underline{y}_1 = 1, \underline{x} = 0)}{O_{0,0}} \quad (62.92)$$

62.7 Special constraints

- **Exogeneity (a.k.a. no-confounding)** holds for simplex \mathcal{S} if $\underline{y}_x \perp \underline{x}$ for $x \in \{0, 1\}$. Hence

$$\underbrace{P(\underline{y}_x = 1)}_{E_{1|x}} = P(\underline{y}_x = 1 | x) = \underbrace{P(\underline{y} = 1 | x)}_{O_{1|x}} \quad \text{for } x \in \{0, 1\} \quad (62.93)$$

Exogeneity gives **2 constraints**.

Note that exogeneity is the same thing as identifiability of $P(\underline{y}_x = y) = P(\underline{y} = y | \mathcal{D}\underline{x} = x)$. But identifiability (i.e., do-identifiability) is a more general concept. One can speak of the identifiability of $P(y_x | z)$ or of $P(y_0, y_1, a)$, etc. In general, any expression with do operators is do-identifiable if it can be expressed as an expression without do-operators.

- **Strong Exogeneity** holds for simplex \mathcal{S} if $(\underline{y}_0, \underline{y}_1)_{joint} \perp \underline{x}$. Hence

$$P(y_0, y_1 | x) = P(y_0, y_1) \quad (62.94)$$

Strong exogeneity gives **3 constraints**. Strong exogeneity implies exogeneity but not the converse.⁷

⁷In this book, we use both $(\underline{a}, \underline{b})_\& \perp \underline{x}$ and $(\underline{a}, \underline{b})_{joint} \perp \underline{x}$. When we write $(\underline{a}, \underline{b})_\& \perp \underline{x}$, we mean that $\underline{a} \perp \underline{x}$ and $\underline{b} \perp \underline{x}$, or, equivalently, $P(a|x) = P(a)$ and $P(b|x) = P(b)$. When we write $(\underline{a}, \underline{b})_{joint} \perp \underline{x}$ or simply $(\underline{a}, \underline{b}) \perp \underline{x}$, we mean $P(a, b|x) = P(a, b)$. Summing $P(a, b|x) = P(a, b)$ over a gives $P(b|x) = P(b)$, and summing it over b gives $P(a|x) = P(a)$. Hence we see that $(\underline{a}, \underline{b})_{joint} \perp \underline{x}$ implies $(\underline{a}, \underline{b})_\& \perp \underline{x}$ but not the converse. If \underline{a} and \underline{b} are independent at fixed \underline{x} so that $P(a, b|x) = P(a|x)P(b|x)$ then $(\underline{a}, \underline{b})_{joint} \perp \underline{x}$ iff $(\underline{a}, \underline{b})_\& \perp \underline{x}$

- **Monotonicity**⁸ holds for simplex \mathcal{S} if

$$AMM = P(\underline{y}_0 = 1, \underline{y}_1 = 0) = 0 \quad (62.95)$$

. Equivalently,

$$\sum_x P(\underline{y}_0 = 1, \underline{y}_1 = 0, x) = 0 \quad (62.96)$$

which is true iff

$$P(\underline{y}_0 = 1, \underline{y}_1 = 0, x) = 0 \quad \text{for } x \in \{0, 1\} \quad (62.97)$$

Monotonicity gives **2 constraints**.

Note that when $AMM = 0$, $PNS = ATE$.

Claim 92 *Monotonicity and exogeneity together imply strong exogeneity.*

proof:

This proof is presented here for completeness. Later on, we will give a much simpler proof of this result. I advise the reader to skip this proof on first reading of this chapter.

Let $a \in \{0, 1\}$.

$$P(\underline{y}_a = \bar{a}, x) = \sum_{a'} P(\underline{y}_a = \bar{a}, \underline{y}_{\bar{a}} = a', x) \quad (62.98)$$

$$= P(\underline{y}_a = \bar{a}, \underline{y}_{\bar{a}} = \bar{a}, x) \quad (\text{by monotonicity}) \quad (62.99)$$

Thus,

$$P(\underline{y}_a = \bar{a} | \underline{x} = a) = P(\underline{y}_a = \bar{a}, \underline{y}_{\bar{a}} = \bar{a} | \underline{x} = a) \quad (62.100)$$

and

$$P(\underline{y}_a = \bar{a}) = P(\underline{y}_a = \bar{a}, \underline{y}_{\bar{a}} = \bar{a}) \quad (62.101)$$

By exogeneity, the left hand side of Eq.(62.100) and the left hand side of Eq.(62.101) are equal, so the right hand sides of those equations must be equal too.

$$P(\underline{y}_a = \bar{a}, \underline{y}_{\bar{a}} = \bar{a} | \underline{x} = a) = P(\underline{y}_a = \bar{a}, \underline{y}_{\bar{a}} = \bar{a}) \quad (62.102)$$

This immediately implies⁹

⁸This property is called monotonicity because it's equivalent to the statement that $y_0^\sigma \leq y_1^\sigma$ for all individuals σ in the population. Indeed, $y_0^\sigma \leq y_1^\sigma$ includes the 3 cases $(y_0^\sigma, y_1^\sigma) = (0, 0), (1, 1), (0, 1)$, but excludes the only other case, namely $(1, 0)$.

⁹If $x \in \{0, 1\}$ and $P(a|\underline{x} = 0) = P(a)$, then $P(a) - P(a, \underline{x} = 0) = P(a)[1 - P(\underline{x} = 0)]$, so $P(a, \underline{x} = 1) = P(a)P(\underline{x} = 1)$. Hence, $P(a|\underline{x} = 1) = P(a)$.

$$P(\underline{y}_a = \bar{a}, \underline{y}_{\bar{a}} = \bar{a} | x) = P(\underline{y}_a = \bar{a}, \underline{y}_{\bar{a}} = \bar{a}) \quad (62.103)$$

for $x \in \{0, 1\}$. This gives for $a = 0$

$$P(\underline{y}_0 = 1, \underline{y}_1 = 1 | x) = P(\underline{y}_0 = 1, \underline{y}_1 = 1) \quad (62.104)$$

and for $a = 1$

$$\underbrace{P(\underline{y}_1 = 0, \underline{y}_0 = 0 | x)}_{P(\underline{y}_0 = 0, \underline{y}_1 = 0 | x)} = \underbrace{P(\underline{y}_1 = 0, \underline{y}_0 = 0)}_{P(\underline{y}_0 = 0, \underline{y}_1 = 0)} \quad (62.105)$$

Monotonicity itself gives

$$P(\underline{y}_0 = 1, \underline{y}_1 = 0 | x) = 0 = P(\underline{y}_0 = 1, \underline{y}_1 = 0) \quad (62.106)$$

The remaining strong exogeneity constraint, given by

$$P(\underline{y}_0 = 0, \underline{y}_1 = 1 | x) = P(\underline{y}_0 = 0, \underline{y}_1 = 1), \quad (62.107)$$

follows because

$$\sum_{y_0, y_1} P(y_0, y_1 | x) = \sum_{y_0, y_1} P(y_0, y_1) = 1 \quad (62.108)$$

QED

62.8 Matrix representation of probabilities

Suppose $\epsilon_x(y_0, y_1)$ for $x, y_0, y_1 \in \{0, 1\}$ are eight vectors in a vector space V . $\epsilon_x(y_0, y_1)$ will only be used at position (y_0, y_1) of a 2×2 matrix, where the positions are defined as follows:

$$\begin{array}{ccc} & y_1 & \\ & \uparrow & \\ 1 & (0, 1) & (1, 1) \\ | & & \\ 0 & (0, 0) & (1, 0) \\ | & & \\ \hline & 0 & 1 \longrightarrow y_0 \end{array} \quad (62.109)$$

When $\epsilon_x(y_0, y_1)$ is used inside a 2×2 matrix, we will not write the argument (y_0, y_1) , leaving it implicit, unless confusion may arise. Thus, for example, we will represent

$$\mathcal{P} \begin{bmatrix} \epsilon_0(0, 1) & 0 \\ 0 & 0 \end{bmatrix} \text{ by } \mathcal{P} \begin{bmatrix} \epsilon_0 & 0 \\ 0 & 0 \end{bmatrix}.$$

For $x \in \{0, 1\}$, let

$$P_{\underline{y}_0, \underline{y}_1, x}(0, 0, x) = \mathcal{P} \begin{bmatrix} 0 & 0 \\ \epsilon_x & 0 \end{bmatrix} \quad (62.110a)$$

$$P_{\underline{y}_0, \underline{y}_1, \underline{x}}(0, 1, x) = \mathcal{P} \begin{bmatrix} \epsilon_x & 0 \\ 0 & 0 \end{bmatrix} \quad (62.110b)$$

$$P_{\underline{y}_0, \underline{y}_1, \underline{x}}(1, 0, x) = \mathcal{P} \begin{bmatrix} 0 & 0 \\ 0 & \epsilon_x \end{bmatrix} \quad (62.110c)$$

$$P_{\underline{y}_0, \underline{y}_1, \underline{x}}(1, 1, x) = \mathcal{P} \begin{bmatrix} 0 & \epsilon_x \\ 0 & 0 \end{bmatrix} \quad (62.110d)$$

Define a map $\mathcal{P}[\quad] : V \rightarrow \mathbb{R}$ that we will call the **matrix representation of probabilities (MRP)** (mnemonic, Mr. P). We will assume that the map $\mathcal{P}[\quad]$ is linear. Hence, for instance,

$$\mathcal{P} \begin{bmatrix} 4\epsilon_0 + 3\epsilon_1 & -5\epsilon_0 \\ 0 & 0 \end{bmatrix} = 4\mathcal{P} \begin{bmatrix} \epsilon_0 & 0 \\ 0 & 0 \end{bmatrix} + 3\mathcal{P} \begin{bmatrix} \epsilon_1 & 0 \\ 0 & 0 \end{bmatrix} - 5\mathcal{P} \begin{bmatrix} 0 & \epsilon_0 \\ 0 & 0 \end{bmatrix} \quad (62.111)$$

Henceforth, we will use the abbreviation

$$\epsilon_+ = \epsilon_0 + \epsilon_1 \quad (62.112)$$

Note that

$$\mathcal{P} \begin{bmatrix} \epsilon_+ & \epsilon_+ \\ \epsilon_+ & \epsilon_+ \end{bmatrix} = 1 \quad (62.113)$$

$$\pi_x = \mathcal{P} \begin{bmatrix} \epsilon_x & \epsilon_x \\ \epsilon_x & \epsilon_x \end{bmatrix} \quad (62.114)$$

$$E_{0|0} = \mathcal{P} \begin{bmatrix} \epsilon_+ & 0 \\ \epsilon_+ & 0 \end{bmatrix}, \quad O_{0,0} = \mathcal{P} \begin{bmatrix} \epsilon_0 & 0 \\ \epsilon_0 & 0 \end{bmatrix} \quad (62.115)$$

$$E_{0|1} = \mathcal{P} \begin{bmatrix} 0 & 0 \\ \epsilon_+ & \epsilon_+ \end{bmatrix}, \quad O_{1,0} = \mathcal{P} \begin{bmatrix} 0 & 0 \\ \epsilon_1 & \epsilon_1 \end{bmatrix} \quad (62.116)$$

$$E_{1|0} = \mathcal{P} \begin{bmatrix} 0 & \epsilon_+ \\ 0 & \epsilon_+ \end{bmatrix}, \quad O_{0,1} = \mathcal{P} \begin{bmatrix} 0 & \epsilon_0 \\ 0 & \epsilon_0 \end{bmatrix} \quad (62.117)$$

$$E_{1|1} = \mathcal{P} \begin{bmatrix} \epsilon_+ & \epsilon_+ \\ 0 & 0 \end{bmatrix}, \quad O_{1,1} = \mathcal{P} \begin{bmatrix} \epsilon_1 & \epsilon_1 \\ 0 & 0 \end{bmatrix} \quad (62.118)$$

$$P(\underline{y} = 0) = O_{0,0} + O_{1,0} = \mathcal{P} \begin{bmatrix} \epsilon_0 & 0 \\ \epsilon_+ & \epsilon_1 \end{bmatrix} \quad (62.119)$$

$$P(\underline{y} = 1) = O_{0,1} + O_{1,1} = \mathcal{P} \begin{bmatrix} \epsilon_1 & \epsilon_+ \\ 0 & \epsilon_0 \end{bmatrix} \quad (62.120)$$

$$PN * O_{1,1} = \mathcal{P} \begin{bmatrix} \epsilon_0 & 0 \\ 0 & 0 \end{bmatrix} \quad (62.121)$$

$$PS * O_{0,0} = \mathcal{P} \begin{bmatrix} \epsilon_1 & 0 \\ 0 & 0 \end{bmatrix} \quad (62.122)$$

$$PNS = \mathcal{P} \begin{bmatrix} \epsilon_+ & 0 \\ 0 & 0 \end{bmatrix} \quad (62.123)$$

$$AMM = \mathcal{P} \begin{bmatrix} 0 & 0 \\ 0 & \epsilon_+ \end{bmatrix} \quad (62.124)$$

$$ATE = PNS - AMM = \mathcal{P} \begin{bmatrix} \epsilon_+ & 0 \\ 0 & -\epsilon_+ \end{bmatrix} \quad (62.125)$$

When $AMM = 0$,

$$P(\underline{y} = 1)|_{AMM=0} = \mathcal{P} \begin{bmatrix} \epsilon_1 & \epsilon_+ \\ 0 & 0 \end{bmatrix} \quad (62.126)$$

The special constraints of exogeneity, strong exogeneity and monotonicity have a very simple in the MRP.

Suppose that $\pi_0, \pi_1 \geq 1$ and $\pi_0 + \pi_1 = 1$. Note that

$$\left. \begin{array}{l} \epsilon_+ = \frac{\epsilon_0}{\pi_0} \\ \text{or} \\ \epsilon_+ = \frac{\epsilon_1}{\pi_1} \end{array} \right\} \implies \epsilon_+ = \frac{\epsilon_0}{\pi_0} = \frac{\epsilon_1}{\pi_1} \quad (62.127)$$

Below, we will abbreviate

$$\hat{\epsilon}_x = \frac{\epsilon_x}{\pi_x} \quad (62.128)$$

for $x \in \{0, 1\}$.

Claim 93

(a) *Exogeneity holds iff (“4 sides can change color”)*

$$\mathcal{P} \begin{bmatrix} \hat{\epsilon}_0 & 0 \\ \hat{\epsilon}_0 & 0 \end{bmatrix} = \mathcal{P} \begin{bmatrix} \hat{\epsilon}_1 & 0 \\ \hat{\epsilon}_1 & 0 \end{bmatrix} = \mathcal{P} \begin{bmatrix} \epsilon_+ & 0 \\ \epsilon_+ & 0 \end{bmatrix} \quad (62.129a)$$

$$\mathcal{P} \begin{bmatrix} \hat{\epsilon}_0 & \hat{\epsilon}_0 \\ 0 & 0 \end{bmatrix} = \mathcal{P} \begin{bmatrix} \hat{\epsilon}_1 & \hat{\epsilon}_1 \\ 0 & 0 \end{bmatrix} = \mathcal{P} \begin{bmatrix} \epsilon_+ & \epsilon_+ \\ 0 & 0 \end{bmatrix} \quad (62.129b)$$

$$\mathcal{P} \begin{bmatrix} 0 & \hat{\epsilon}_0 \\ 0 & \hat{\epsilon}_0 \end{bmatrix} = \mathcal{P} \begin{bmatrix} 0 & \hat{\epsilon}_1 \\ 0 & \hat{\epsilon}_1 \end{bmatrix} = \mathcal{P} \begin{bmatrix} 0 & \epsilon_+ \\ 0 & \epsilon_+ \end{bmatrix} \quad (62.129c)$$

$$\mathcal{P} \begin{bmatrix} 0 & 0 \\ \hat{\epsilon}_0 & \hat{\epsilon}_0 \end{bmatrix} = \mathcal{P} \begin{bmatrix} 0 & 0 \\ \hat{\epsilon}_1 & \hat{\epsilon}_1 \end{bmatrix} \mathcal{P} \begin{bmatrix} 0 & 0 \\ \epsilon_+ & \epsilon_+ \end{bmatrix} \quad (62.129d)$$

(b) Monotonicity holds iff

$$\mathcal{P} \begin{bmatrix} 0 & 0 \\ 0 & \epsilon_x \end{bmatrix} = 0 \quad (62.130)$$

for $x \in \{0, 1\}$.

(c) Strong exogeneity holds iff (“4 corners can change color”)

$$\mathcal{P} \begin{bmatrix} \hat{\epsilon}_0 & 0 \\ 0 & 0 \end{bmatrix} = \mathcal{P} \begin{bmatrix} \hat{\epsilon}_1 & 0 \\ 0 & 0 \end{bmatrix} = \mathcal{P} \begin{bmatrix} \epsilon_+ & 0 \\ 0 & 0 \end{bmatrix} \quad (62.131a)$$

$$\mathcal{P} \begin{bmatrix} 0 & \hat{\epsilon}_0 \\ 0 & 0 \end{bmatrix} = \mathcal{P} \begin{bmatrix} 0 & \hat{\epsilon}_1 \\ 0 & 0 \end{bmatrix} = \mathcal{P} \begin{bmatrix} 0 & \epsilon_+ \\ 0 & 0 \end{bmatrix} \quad (62.131b)$$

$$\mathcal{P} \begin{bmatrix} 0 & 0 \\ 0 & \hat{\epsilon}_0 \end{bmatrix} = \mathcal{P} \begin{bmatrix} 0 & 0 \\ 0 & \hat{\epsilon}_1 \end{bmatrix} = \mathcal{P} \begin{bmatrix} 0 & 0 \\ 0 & \epsilon_+ \end{bmatrix} \quad (62.131c)$$

$$\mathcal{P} \begin{bmatrix} 0 & 0 \\ \hat{\epsilon}_0 & 0 \end{bmatrix} = \mathcal{P} \begin{bmatrix} 0 & 0 \\ \hat{\epsilon}_1 & 0 \end{bmatrix} = \mathcal{P} \begin{bmatrix} 0 & 0 \\ \epsilon_+ & 0 \end{bmatrix} \quad (62.131d)$$

proof:

(a) This follows from the definitions of $E_{y|x}$ and $O_{x,y}$ for $x, y \in \{0, 1\}$, as stated above in the MRP.

(b) This follows from the identity

$$P_{\underline{y}_0, \underline{y}_1, \underline{x}}(1, 0|x) = \mathcal{P} \begin{bmatrix} 0 & 0 \\ 0 & \frac{\epsilon_x}{\pi_x} \end{bmatrix} \quad (62.132)$$

for $x \in \{0, 1\}$.

(c) This follows from definitions of $P_{y_0, y_1, x}(y_0, y_1, x)$ for $y_0, y_1, x \in \{0, 1\}$, as stated above in the MRP.

QED

Note that it is obvious from Claim 93 that strong exogeneity implies exogeneity.

Claim 93 also makes it easy peasy to prove that exogeneity and monotonicity imply strong exogeneity. Indeed, here is a much simpler proof than the one we presented earlier.

Claim 94 *Exogeneity and monotonicity imply strong exogeneity.*

proof: Let's call $E(a), E(b), E(c), E(d)$ the four Eqs.(62.129) that define Exogeneity, and $SE(a), SE(b), SE(c), SE(d)$ the four Eqs.(62.131) that define Strong Exogeneity.

- Monotonicity implies the lower right entry is zero, so $SE(c)$ is true.
- Setting to zero the lower right entry in $E(c)$ and $E(d)$ implies $SE(b)$ and $SE(d)$.
- Subtracting $SE(d)$ from $E(a)$ gives $SE(a)$.

QED

62.9 Bounds on Exp. Probs. imposed by Obs. Probs.

Claim 95 *In general,*

$$O_{x,1} \leq E_{1|x} \leq 1 - O_{x,0} \quad \text{for } x \in \{0, 1\}. \quad (62.133)$$

In other words,

$$O_{1,1} \leq E_{1|1} \leq 1 - O_{1,0} \quad (62.134)$$

$$O_{0,1} \leq E_{1|0} \leq 1 - O_{0,0} \quad (62.135)$$

With monotonicity, we get the tighter bounds

$$\overbrace{O_{1,1} + \underbrace{O_{0,1}}_{new}}^{P(y=1)} \leq E_{1|1} \leq 1 - O_{1,0} \quad (62.136)$$

$$O_{0,1} \leq E_{1|0} \leq \overbrace{1 - O_{0,0} - \underbrace{O_{1,0}}_{new}}^{P(y=1)} \quad (62.137)$$

proof:

$$O_{1,1} \leq E_{1|1} \leq \underbrace{1 - O_{1,0}}_{O_{1,1} + O_{0,0} + O_{0,1}} \quad (62.138)$$

has the MRP

$$\mathcal{P} \begin{bmatrix} \epsilon_1 & \epsilon_1 \\ 0 & 0 \end{bmatrix} \leq \mathcal{P} \begin{bmatrix} \epsilon_+ & \epsilon_+ \\ 0 & 0 \end{bmatrix} \leq \mathcal{P} \begin{bmatrix} \epsilon_1 & \epsilon_1 \\ 0 & 0 \end{bmatrix} + \mathcal{P} \begin{bmatrix} \epsilon_0 & 0 \\ \epsilon_0 & 0 \end{bmatrix} + \mathcal{P} \begin{bmatrix} 0 & \epsilon_0 \\ 0 & \epsilon_0 \end{bmatrix} \quad (62.139)$$

which is obviously true.

$$O_{0,1} \leq E_{1|0} \leq \underbrace{1 - O_{0,0}}_{O_{1,1} + O_{0,1} + O_{1,0}} \quad (62.140)$$

has the MRP

$$\mathcal{P} \begin{bmatrix} 0 & \epsilon_0 \\ 0 & \epsilon_0 \end{bmatrix} \leq \mathcal{P} \begin{bmatrix} 0 & \epsilon_+ \\ 0 & \epsilon_+ \end{bmatrix} \leq \mathcal{P} \begin{bmatrix} \epsilon_1 & \epsilon_1 \\ 0 & 0 \end{bmatrix} + \mathcal{P} \begin{bmatrix} 0 & \epsilon_0 \\ 0 & \epsilon_0 \end{bmatrix} + \mathcal{P} \begin{bmatrix} 0 & 0 \\ \epsilon_1 & \epsilon_1 \end{bmatrix} \quad (62.141)$$

which is obviously true.

$$P(\underline{y} = 1)|_{AMM=0} \leq E_{1|1} \quad (62.142)$$

has the MRP

$$\mathcal{P} \begin{bmatrix} \epsilon_1 & \epsilon_+ \\ 0 & 0 \end{bmatrix} \leq \mathcal{P} \begin{bmatrix} \epsilon_+ & \epsilon_+ \\ 0 & 0 \end{bmatrix} \quad (62.143)$$

which is obviously true.

$$E_{1|0}|_{AMM=0} \leq P(\underline{y} = 1)|_{AMM=0} \quad (62.144)$$

has the MRP

$$\mathcal{P} \begin{bmatrix} 0 & \epsilon_+ \\ 0 & 0 \end{bmatrix} \leq \mathcal{P} \begin{bmatrix} \epsilon_1 & \epsilon_+ \\ 0 & 0 \end{bmatrix} \quad (62.145)$$

which is obviously true.

QED

62.10 Bounds on $PNS3$ for unspecified bnet

Claim 96 *If $P(a, b)$ is a probability distribution, then*

$$\max\{0, P(a) + P(b) - 1\} \leq P(a, b) \leq \min\{P(a), P(b)\} \quad (62.146)$$

proof: $P(a|b) \leq 1$ and $P(b|a) \leq 1$ implies $P(a, b) \leq P(b)$ and $P(a, b) \leq P(a)$. Hence, $P(a, b) \leq \min\{P(a), P(b)\}$.

$$P(a) + P(b) - 1 = \sum_{a', b'} P(a', b') \underbrace{[\delta(a, a') + \delta(b, b') - 1]}_T \quad (62.147)$$

$$\leq P(a, b) \quad (T \text{ biggest when } a = a' \text{ and } b = b') \quad (62.148)$$

QED

Claim 97 *Minimal bounds (see Fig.62.4)*

If the minimal constraints hold for simplex \mathcal{S} , then

$$\max \left\{ \begin{array}{l} 0 \\ E_{1|1} + E_{0|0} - 1 \\ E_{0|0} - \sum_x O_{x,0} \\ E_{1|1} - \sum_x O_{x,1} \end{array} \right\} \leq PNS \leq \min \left\{ \begin{array}{l} E_{1|1} \\ E_{0|0} \\ O_{1,1} + O_{0,0} \\ E_{1|1} + E_{0|0} - O_{1,1} - O_{0,0} \end{array} \right\} \quad (62.149)$$

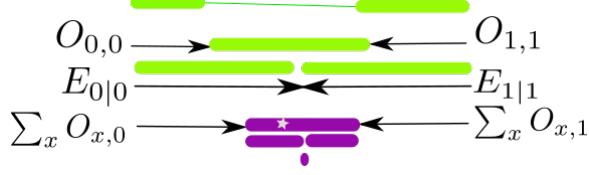


Figure 62.4: Minimal bounds for PNS . PNS is larger than maximum of the purple segments and smaller than the minimum of the green ones. The purple segment of almost zero length represents zero. The green segment interrupted by a thin green line represents $O_{0,0} + O_{1,1}$. Note that $\sum_x O_{x,0} + \sum_x O_{x,1} = 1$. When monotonicity holds, PNS equals $E_{1|1} + E_{0|0} - 1$ which is the purple segment marked with a star.

$$\max \left\{ \frac{E_{0|0} - \sum_x O_{x,0}}{O_{1,1}} \right\} \leq PN \leq \min \left\{ \frac{1}{E_{0|0} - O_{0,0}} \right\} \quad (62.150)$$

$$\max \left\{ \frac{E_{1|1} - \sum_x O_{x,1}}{O_{0,0}} \right\} \leq PS \leq \min \left\{ \frac{1}{E_{1|1} - O_{1,1}} \right\} \quad (62.151)$$

proof: These bounds were copied directly from Ref.[63], except the notation was changed. In certain cases, we have slightly rewritten the bounds from Ref.[63] to exhibit more explicitly their symmetry under swaps of zeros and ones. For example, instead of using $(E_{1|0}, E_{1|1})$ as parameters, we use $(E_{1|1}, E_{0|0})$.

These bounds are obviously true in the MRP. Indeed, in the MRP, the following holds. For PNS , we have

$$\begin{bmatrix} 0 \\ E_{1|1} + E_{0|0} - 1 \\ E_{0|0} - \sum_x O_{x,0} \\ E_{1|1} - \sum_x O_{x,1} \end{bmatrix} = \begin{bmatrix} \mathcal{P} \begin{bmatrix} \epsilon_+ & 0 \\ 0 & -\epsilon_+ \end{bmatrix} \\ \mathcal{P} \begin{bmatrix} \epsilon_1 & 0 \\ 0 & -\epsilon_1 \end{bmatrix} \\ \mathcal{P} \begin{bmatrix} \epsilon_0 & 0 \\ 0 & -\epsilon_0 \end{bmatrix} \end{bmatrix} \quad (62.152)$$

$$PNS = \mathcal{P} \begin{bmatrix} \epsilon_+ & 0 \\ 0 & 0 \end{bmatrix} \quad (62.153)$$

$$\begin{bmatrix} E_{1|1} \\ E_{0|0} \\ O_{1,1} + O_{0,0} \\ E_{1|1} + E_{0|0} - O_{1,1} - O_{0,0} \end{bmatrix} = \begin{bmatrix} \mathcal{P} \begin{bmatrix} \epsilon_+ & 0 \\ \epsilon_+ & 0 \end{bmatrix} \\ \mathcal{P} \begin{bmatrix} \epsilon_+ & \epsilon_+ \\ 0 & 0 \end{bmatrix} \\ \mathcal{P} \begin{bmatrix} \epsilon_+ & \epsilon_1 \\ \epsilon_0 & 0 \end{bmatrix} \\ \mathcal{P} \begin{bmatrix} \epsilon_+ & \epsilon_0 \\ \epsilon_1 & 0 \end{bmatrix} \end{bmatrix} \quad (62.154)$$

For PN , we have

$$\begin{bmatrix} 0 \\ E_{1|1} - \sum_x O_{x,1} \end{bmatrix} = \begin{bmatrix} 0 \\ \mathcal{P} \begin{bmatrix} \epsilon_0 & 0 \\ 0 & -\epsilon_0 \end{bmatrix} \end{bmatrix} \quad (62.155)$$

$$PN * O_{1,1} = \mathcal{P} \begin{bmatrix} \epsilon_0 & 0 \\ 0 & 0 \end{bmatrix} \quad (62.156)$$

$$\begin{bmatrix} 1 \\ E_{1|1} - O_{1,1} \end{bmatrix} = \begin{bmatrix} 1 \\ \mathcal{P} \begin{bmatrix} \epsilon_0 & \epsilon_0 \\ 0 & 0 \end{bmatrix} \end{bmatrix} \quad (62.157)$$

For PS , we have

$$\begin{bmatrix} 0 \\ E_{0|0} - \sum_x O_{x,0} \end{bmatrix} = \begin{bmatrix} 0 \\ \mathcal{P} \begin{bmatrix} \epsilon_1 & 0 \\ 0 & -\epsilon_1 \end{bmatrix} \end{bmatrix} \quad (62.158)$$

$$PS * O_{0,0} = \mathcal{P} \begin{bmatrix} \epsilon_1 & 0 \\ 0 & 0 \end{bmatrix} \quad (62.159)$$

$$\begin{bmatrix} 1 \\ E_{0|0} - O_{0,0} \end{bmatrix} = \begin{bmatrix} 1 \\ \mathcal{P} \begin{bmatrix} \epsilon_1 & \epsilon_1 \\ 0 & 0 \end{bmatrix} \end{bmatrix} \quad (62.160)$$

The first two lines of the bound for PNS are a simple consequence of Claim 96. Indeed, Claim 96 implies

$$\max \left\{ \frac{0}{P(y_0) + P(y_1) - 1} \right\} \leq P(y_0, y_1) \leq \min \left\{ \frac{P(y_0)}{P(y_1)} \right\} \quad (62.161)$$

When $y_0 = 0, y_1 = 1$, we get

$$\max \left\{ \frac{0}{E_{0|0} + E_{1|1} - 1} \right\} \leq PNS \leq \min \left\{ \frac{E_{0|0}}{E_{1|1}} \right\} \quad (62.162)$$

QED

In Claim 97, note that

- $\sum_x O_{x,y} = P(\underline{y} = y)$ for $y = 0, 1$.
- Let

$$E'_{0|0} = E_{0|0} - \sum_x O_{x,0} = \mathcal{P} \begin{bmatrix} \epsilon_1 & 0 \\ 0 & -\epsilon_1 \end{bmatrix} \quad (62.163)$$

$$E'_{1|1} = E_{1|1} - \sum_x O_{x,1} = \mathcal{P} \begin{bmatrix} \epsilon_0 & 0 \\ 0 & -\epsilon_0 \end{bmatrix} \quad (62.164)$$

$$ATE = E_{1|1} + E_{0|0} - 1 = E'_{1|1} + E'_{0|0} = \mathcal{P} \begin{bmatrix} \epsilon_+ & 0 \\ 0 & -\epsilon_+ \end{bmatrix} \quad (62.165)$$

- Claim 97 applies if we have both observational data (OD) and experimental data (ED). If we only have OD (resp., ED), ignore all bounds that involve an E (resp., involve an O) probability.

Hence, with only OD¹⁰

$$0 \leq PNS \leq O_{1,1} + O_{0,0} \quad (62.166)$$

and $PN, PS \in [0, 1]$, whereas with only ED,

$$\max \left\{ \frac{0}{E_{1|1} + E_{0|0} - 1} \right\} \leq PNS \leq \min \left\{ \frac{E_{1|1}}{E_{0|0}} \right\} \quad (62.167)$$

and $PN, PS \in [0, 1]$.

- At first blush, there seem to be 3 possible cases to consider: (1) Only OD. (2) Both OD and ED. (3) Only ED. Actually, Case (1) plus the assumption of exogeneity is the same as Case (3). Indeed, recall that exogeneity means no confounding, and ED (i.e., an RCT) also has no confounding. So we don't have to consider Case (3) if we consider Case (1) without and with exogeneity. Exogeneity is built into case (2), so case (2) without exogeneity is meaningless.

Claim 98 *If minimal and exogeneity constraints hold for simplex \mathcal{S} , then*

$$\max \left\{ \frac{0}{\frac{O_{1|1} + O_{0|0} - 1}{\frac{O_{0|0} - \sum_x O_{x,0}}{O_{1|1} - \sum_x O_{x,1}}}} \right\} \leq PNS \leq \min \left\{ \frac{\overbrace{O_{1|1}}^1}{\underbrace{O_{1|1} + O_{0|0} - O_{1|1} - O_{0,0}}^0} \right\} \quad (62.168)$$

$$\max \left\{ \frac{0}{ERR} \right\} \leq PN \leq \min \left\{ \frac{1}{\frac{O_{0|0}}{O_{1|1}}} \right\} \quad (62.169)$$

$$\max \left\{ \frac{0}{(ERR)^\sim} \right\} \leq PS \leq \min \left\{ \frac{1}{\frac{O_{1|1}}{O_{0|0}}} \right\} \quad (62.170)$$

¹⁰The bounds Eq.(62.166) assume exogeneity does not hold. For the case when exogeneity does hold, stronger bounds will be given later on in the chapter.

proof: Just replace $E_{y|x}$ by $O_{y|x}$ in the minimal bounds given in Claim 97.

The canceled terms do not improve the bounds and can be dropped. We show this next using MRP.

In the MRP, the minimal bounds for PNS are

$$\begin{bmatrix} 0 \\ E_{1|1} + E_{0|0} - 1 \\ E_{0|0} - \sum_x O_{x,0} \\ E_{1|1} - \sum_x O_{x,1} \end{bmatrix} = \begin{bmatrix} \mathcal{P} \begin{bmatrix} \epsilon_+ & 0 \\ 0 & -\epsilon_+ \end{bmatrix} \\ \mathcal{P} \begin{bmatrix} \epsilon_1 & 0 \\ 0 & -\epsilon_1 \end{bmatrix} \\ \mathcal{P} \begin{bmatrix} \epsilon_0 & 0 \\ 0 & -\epsilon_0 \end{bmatrix} \end{bmatrix} \quad (62.171)$$

$$PNS = \mathcal{P} \begin{bmatrix} \epsilon_+ & 0 \\ 0 & 0 \end{bmatrix} \quad (62.172)$$

$$\begin{bmatrix} E_{1|1} \\ E_{0|0} \\ O_{1,1} + O_{0,0} \\ E_{1|1} + E_{0|0} - O_{1,1} - O_{0,0} \end{bmatrix} = \begin{bmatrix} \mathcal{P} \begin{bmatrix} \epsilon_+ & 0 \\ \epsilon_+ & 0 \end{bmatrix} \\ \mathcal{P} \begin{bmatrix} \epsilon_+ & \epsilon_+ \\ 0 & 0 \end{bmatrix} \\ \mathcal{P} \begin{bmatrix} \epsilon_+ & \epsilon_1 \\ \epsilon_0 & 0 \end{bmatrix} \\ \mathcal{P} \begin{bmatrix} \epsilon_+ & \epsilon_0 \\ \epsilon_1 & 0 \end{bmatrix} \end{bmatrix} \quad (62.173)$$

If exogeneity holds, we can replace all E 's by O 's on the left hand sides. We can also use $\epsilon_+ = \hat{\epsilon}_0 = \hat{\epsilon}_1$ on the right hand sides to conclude that

$$O_{1|1} + O_{0|0} - 1 = \mathcal{P} \begin{bmatrix} \epsilon_+ & \epsilon_+ - \epsilon_+ \\ 0 & -\epsilon_+ \end{bmatrix} = \mathcal{P} \begin{bmatrix} \hat{\epsilon}_x & \hat{\epsilon}_x - \hat{\epsilon}_x \\ 0 & -\hat{\epsilon}_x \end{bmatrix} = \frac{1}{\pi_x} \mathcal{P} \begin{bmatrix} \epsilon_x & 0 \\ 0 & -\epsilon_x \end{bmatrix} \quad (62.174)$$

for both $x \in \{0, 1\}$. Note that $\frac{1}{\pi_x} \mathcal{P} \begin{bmatrix} \epsilon_x & 0 \\ 0 & -\epsilon_x \end{bmatrix}$ will always be greater or equal to $\mathcal{P} \begin{bmatrix} \epsilon_x & 0 \\ 0 & -\epsilon_x \end{bmatrix}$ when $\mathcal{P} \begin{bmatrix} \epsilon_x & 0 \\ 0 & -\epsilon_x \end{bmatrix} \geq 0$ (if it's negative, the 0 bound takes precedence). Hence, the two canceled lower bound terms can be dropped. A similar argument shows that the two canceled upper bound terms can be dropped too.

QED

Claim 99 *If the minimal and strong exogeneity constraints hold for simplex \mathcal{S} , then the inequalities for exogeneity Claim 98 hold. In addition,*

$$PN = \frac{PNS}{O_{1|1}} \quad (62.175)$$

and

$$PS = \frac{PNS}{O_{0|0}} \quad (62.176)$$

proof:

$$PN * O_{1|1} = P(\underline{y}_0 = 0, \underline{y} = 1 | \underline{x} = 1) \quad (62.177)$$

$$= P(\underline{y}_0 = 0, \underline{y}_1 = 1) \quad (62.178)$$

$$= PNS \quad (62.179)$$

$$PS * O_{0|0} = P(\underline{y}_1 = 1, \underline{y} = 0 | \underline{x} = 0) \quad (62.180)$$

$$= P(\underline{y}_0 = 0, \underline{y}_1 = 1) \quad (62.181)$$

$$= PNS \quad (62.182)$$

QED

Claim 100 *If the minimal and monotonicity constraints hold for simplex \mathcal{S} , then*

$$PNS = E_{1|1} + E_{0|0} - 1 \quad (62.183)$$

$$PN = \frac{E_{0|0} - \sum_x O_{x,0}}{O_{1,1}} = \frac{\sum_x O_{x,1} - E_{1|0}}{O_{1,1}} = CERR \quad (62.184)$$

$$PS = \frac{E_{1|1} - \sum_x O_{x,1}}{O_{0,0}} = \frac{\sum_x O_{x,0} - E_{0|1}}{O_{0,0}} = (CERR)^\sim \quad (62.185)$$

proof: These results can be easily proven using the MRP.

Note that

$$\frac{\sum_x O_{x,1} - E_{1|0}}{O_{1,1}} = \frac{O_{1|1}\pi_1 + O_{1|0}(1 - \pi_1) - E_{1|0}}{O_{1|1}\pi_1} \quad (62.186)$$

$$= 1 - \frac{O_{1|0}}{O_{1|1}} + \frac{O_{1|0} - E_{1|0}}{O_{1,1}} \quad (62.187)$$

$$= CERR \quad (62.188)$$

QED

Claim 101 *If the minimal, exogeneity and monotonicity constraints hold for simplex \mathcal{S} , then PNS, PN, PS are identifiable, and*

$$PNS = O_{1|1} + O_{0|0} - 1 \quad (62.189)$$

$$PN = \frac{O_{0|0} - \sum_x O_{x,0}}{O_{1,1}} = \frac{\sum_x O_{x,1} - O_{1|0}}{O_{1,1}} = ERR \quad (62.190)$$

$$PS = \frac{O_{1|1} - \sum_x O_{x,1}}{O_{0,0}} = \frac{\sum_x O_{x,0} - O_{0|1}}{O_{0,0}} = (ERR)^\sim \quad (62.191)$$

proof: Set $E_{y|x} = O_{y|x}$ in Claim 100.

QED

62.11 Bounds on PNS_3 for specific bnet families

Define

$$E_{*,*,z} = E_{0|0,z} + E_{1|1,z} = ATE_z + 1 \quad (62.192)$$

$$O_{*,*,z} = O_{0,0|z} + O_{1,1|z} \quad (62.193)$$

Claim 102 *If $\underline{z} \cap de(\underline{x}) = \emptyset$, then*

$$\max \left\{ \begin{array}{c} 0 \\ \frac{E_{*,*,z} - 1}{E_{0|0,z} - \sum_x O_{x,0|z}} \\ \frac{E_{1|1,z} - \sum_x O_{x,1|z}}{E_{*,*,z} - O_{*,*,z}} \end{array} \right\} \leq PNS_z \leq \min \left\{ \begin{array}{c} E_{1|1,z} \\ \frac{E_{0|0,z}}{O_{*,*,z}} \\ \frac{E_{0|0,z} - O_{0,0|z}}{E_{1|1,z} - O_{1,1|z}} \end{array} \right\} \quad (62.194)$$

$$\max \left\{ \frac{0}{\frac{E_{0|0,z} - \sum_x O_{x,0|z}}{O_{1,1|z}}} \right\} \leq PN_z \leq \min \left\{ \frac{1}{\frac{E_{0|0,z} - O_{0,0|z}}{O_{1,1|z}}} \right\} \quad (62.195)$$

$$\max \left\{ \frac{0}{\frac{E_{1|1,z} - \sum_x O_{x,1|z}}{O_{0,0|z}}} \right\} \leq PS_z \leq \min \left\{ \frac{1}{\frac{E_{1|1,z} - O_{1,1|z}}{O_{0,0|z}}} \right\} \quad (62.196)$$

proof: These are just the minimal bounds where everything is conditioned on z .

QED

In the minimal bounds, node \underline{z} is not mentioned because the minimal bounds don't assume that \underline{z} exists. Once we assume that node \underline{z} exists, then we can use it, as Claim 102 does.

Claim 103 If \underline{z} satisfies the backdoor adjustment criterion relative to $(\underline{x}, \underline{y})$, then

$$\max \left\{ \frac{0}{\frac{O_{0|0,z} - \sum_x O_{x,0|z}}{O_{1|1,z} - \sum_x O_{x,1|z}}} \right\} \leq PNS_z \leq \min \left\{ \frac{O_{1|1,z}}{\frac{O_{0|0,z}}{O_{*,*,z} - O_{*,*,|z}}}, \frac{O_{0|0,z}}{O_{*,*,z} - O_{*,*,|z}} \right\} \quad (62.197)$$

$$\max \left\{ \frac{0}{\frac{O_{0|0,z} - \sum_x O_{x,0|z}}{O_{1,1|z}}} \right\} \leq PN_z \leq \min \left\{ \frac{1}{\frac{O_{0|0,z} - O_{0,0|z}}{O_{1,1|z}}} \right\} \quad (62.198)$$

$$\max \left\{ \frac{0}{\frac{O_{1|1,z} - \sum_x O_{x,1|z}}{O_{0,0|z}}} \right\} \leq PS_z \leq \min \left\{ \frac{1}{\frac{O_{1|1,z} - O_{1,1|z}}{O_{0,0|z}}} \right\} \quad (62.199)$$

proof: Satisfaction of the backdoor criterion implies that

$$\underbrace{P(y|x, z)}_{O_{y|x,z}} = \underbrace{P(\underline{y}_{\underline{x}} = y|z)}_{E_{y|x,z}} \quad (62.200)$$

Replace $E_{y|x,z}$ by $O_{y|x,z}$ in Claim 102. These are just the minimal bounds, when exogeneity holds, that were considered previously. As before, the canceled terms do not affect the bounds so they can be dropped.

QED

So far we have considered 2 families of bnets: (1) $\underline{z} \cap de(\underline{x}) = \emptyset$ and (2) \underline{z} satisfies backdoor adjustment criterion. Actually, (2) is a subset of (1). Next, we will consider bnets in which \underline{z} is a mediator relative to $(\underline{x}, \underline{y})$. A mediator \underline{z} is impinged by an arrow $\underline{x} \rightarrow \underline{z}$, so it definitely does not satisfy (1). Bnets that do not satisfy (1) also fall outside the purview of Potential Outcomes (PO) theory.

We say \underline{z} is a **partial mediator** relative to $(\underline{x}, \underline{y})$ if $\underline{y}_{\underline{x}} \perp (\underline{x}, \underline{z}_{\underline{x}}) \& | \underline{z}_x$ for $x \in \{0, 1\}$. See Fig.62.5 for an example. In that figure, the TPMs, printed in blue, for nodes \underline{z} and \underline{y} , are as follows.

$$P(z|x, z_0, z_1) = \delta(z, z_x) \quad (62.201)$$

$$P(y|x, y_0, y_1) = \delta(y, y_x) \quad (62.202)$$

If \underline{z} is a partial mediator relative to $(\underline{x}, \underline{y})$ and there is no direct path from \underline{x} to \underline{y} , we say \underline{x} is a **pure mediator** relative to $(\underline{x}, \underline{y})$.

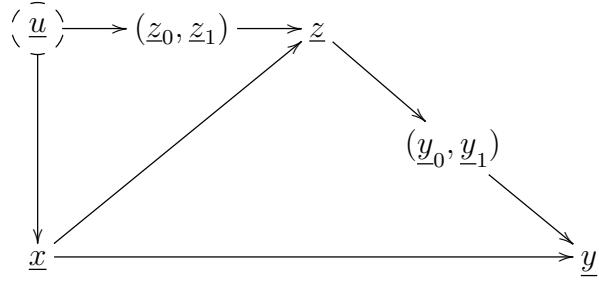


Figure 62.5: Example of a bnet in which \underline{z} is a partial mediator relative to $(\underline{x}, \underline{y})$. If $\underline{x} = 1$, then conditioning on \underline{z}_1 blocks the paths from \underline{y}_1 to \underline{x} and from \underline{y}_1 to \underline{z}_0 . If \underline{z} is a partial mediator relative to $(\underline{x}, \underline{y})$, and, in addition, there is no arrow $\underline{x} \rightarrow \underline{y}$, we say \underline{z} is a pure mediator relative to $(\underline{x}, \underline{y})$.

Claim 104 (*Bounds if \underline{z} is partial mediator*)

If \underline{z} is a partial mediator relative to $(\underline{x}, \underline{y})$, and $\underline{x}, \underline{y}, \underline{z} \in \{0, 1\}$, then

$$\max \left\{ \begin{array}{l} 0 \\ \frac{E_{*|*} - 1}{E_{0|0} - \sum_x O_{x,0}} \\ \frac{E_{1|1} - \sum_x O_{x,1}}{E_{0|0}} \end{array} \right\} \leq PNS \leq \min \left\{ \begin{array}{l} \frac{E_{1|1}}{E_{0|0}} \\ \frac{O_{*,*}}{E_{*|*} - O_{*,*}} \\ \frac{new}{new} \end{array} \right\} \quad (62.203)$$

where

$$new = \sum_{z_0, z_1} \min \left\{ \begin{array}{l} O_{1|1, z_1} \\ O_{0|0, z_0} \end{array} \right\} \min \left\{ \begin{array}{l} P(z_1) \\ P(z_0) \end{array} \right\} \quad (62.204)$$

proof:

$$PNS = P(\underline{y}_1 = 1, \underline{y}_0 = 0) \quad (62.205)$$

$$= \sum_{z_0, z_1} P(\underline{y}_1 = 1, \underline{y}_0 = 0 | z_1, z_0) P(z_1, z_0) \quad (62.206)$$

$$\leq \sum_{z_0, z_1} \min \left\{ \begin{array}{l} P(\underline{y}_1 = 1 | z_1, z_0) \\ P(\underline{y}_0 = 0 | z_1, z_0) \end{array} \right\} \min \left\{ \begin{array}{l} P(z_1) \\ P(z_0) \end{array} \right\} \quad (62.207)$$

$$= \sum_{z_0, z_1} \min \left\{ \begin{array}{l} P(\underline{y} = 1 | z_1) \\ P(\underline{y} = 0 | z_0) \end{array} \right\} \min \left\{ \begin{array}{l} P(z_1) \\ P(z_0) \end{array} \right\} (\text{because } \underline{y}_x \perp \underline{z}_x | \underline{z}_x) \quad (62.208)$$

$$= \sum_{z_0, z_1} \min \left\{ \begin{array}{l} P(\underline{y} = 1 | \underline{x} = 1, z_1) \\ P(\underline{y} = 0 | \underline{x} = 0, z_0) \end{array} \right\} \min \left\{ \begin{array}{l} P(z_1) \\ P(z_0) \end{array} \right\} (\text{because } \underline{y}_x \perp \underline{x} | \underline{z}_x) \quad (62.209)$$

$$= \sum_{z_0, z_1} \min \left\{ \begin{array}{l} P(\underline{y} = 1 | \underline{x} = 1, \underline{z} = z_1) \\ P(\underline{y} = 0 | \underline{x} = 0, \underline{z} = z_0) \end{array} \right\} \min \left\{ \begin{array}{l} P(z_1) \\ P(z_0) \end{array} \right\} \quad (62.210)$$

(because $\underline{z} = \underline{z}_x$ at fixed x)

$$= \sum_{z_0, z_1} \min \left\{ \begin{array}{l} O_{1|1, z_1} \\ O_{0|0, z_0} \end{array} \right\} \min \left\{ \begin{array}{l} P(z_1) \\ P(z_0) \end{array} \right\} \quad (62.211)$$

QED

Claim 105 (*Bounds if \underline{z} is a pure mediator*)

If \underline{z} is a pure mediator relative to $(\underline{x}, \underline{y})$, and $\underline{x}, \underline{y}, \underline{z} \in \{0, 1\}$, then Claim 104 for a partial mediator is still valid, except that now “new” equals

$$\text{new} = \sum_{z_0, z_1} \mathbb{1}(z_0 \neq z_1) \min \left\{ \begin{array}{l} O_{1|\underline{x}=z_1} \\ O_{0|\underline{x}=z_0} \end{array} \right\} \min \left\{ \begin{array}{l} P(\underline{z} = z_1 | \underline{x} = 1) \\ P(\underline{z} = z_0 | \underline{x} = 0) \end{array} \right\} \quad (62.212)$$

proof:

$$PNS = P(\underline{y}_1 = 1, \underline{y}_0 = 0) \quad (62.213)$$

$$= \sum_{z_0, z_1} \mathbb{1}(z_0 \neq z_1) P(\underline{y}_{z_1} = 1, \underline{y}_{z_0} = 0) P(z_1, z_0) \quad (62.214)$$

(because $(\underline{y}_0, \underline{y}_1) \perp (\underline{z}_0, \underline{z}_1)$ and $\underline{y}_0 \neq \underline{y}_1$)

$$\leq \sum_{z_0, z_1} \mathbb{1}(z_0 \neq z_1) \min \left\{ \begin{array}{l} P(\underline{y}_{z_1} = 1) \\ P(\underline{y}_{z_0} = 0) \end{array} \right\} \min \left\{ \begin{array}{l} P(z_1) \\ P(z_0) \end{array} \right\} \quad (62.215)$$

$$= \sum_{z_0, z_1} \mathbb{1}(z_0 \neq z_1) \min \left\{ \begin{array}{l} P(\underline{y} = 1 | \underline{x} = z_1) \\ P(\underline{y} = 0 | \underline{x} = z_0) \end{array} \right\} \min \left\{ \begin{array}{l} P(z_1 | \underline{x} = 1) \\ P(z_0 | \underline{x} = 0) \end{array} \right\} \quad (62.216)$$

(because $\underline{y} = \underline{y}_x$ at fixed \underline{x} and $(z_0, z_1) \perp \underline{x}$)

$$= \sum_{z_0, z_1} \mathbb{1}(z_0 \neq z_1) \min \left\{ \begin{array}{l} O_{1|\underline{x}=z_1} \\ O_{0|\underline{x}=z_0} \end{array} \right\} \min \left\{ \begin{array}{l} P(\underline{z} = z_1 | \underline{x} = 1) \\ P(\underline{z} = z_0 | \underline{x} = 0) \end{array} \right\} \quad (62.217)$$

QED

62.12 Numerical Examples

I've written an open source Python program (See Ref[67]) that calculates the bounds given in this chapter. The program is called "JudeasRx", in honor of Judea Pearl. Fig.62.6 is an example of its interface with data entered by Boris Sobolev from a real life case. JudeasRx considers $z = g = \text{gender} \in \{m, f\}$.

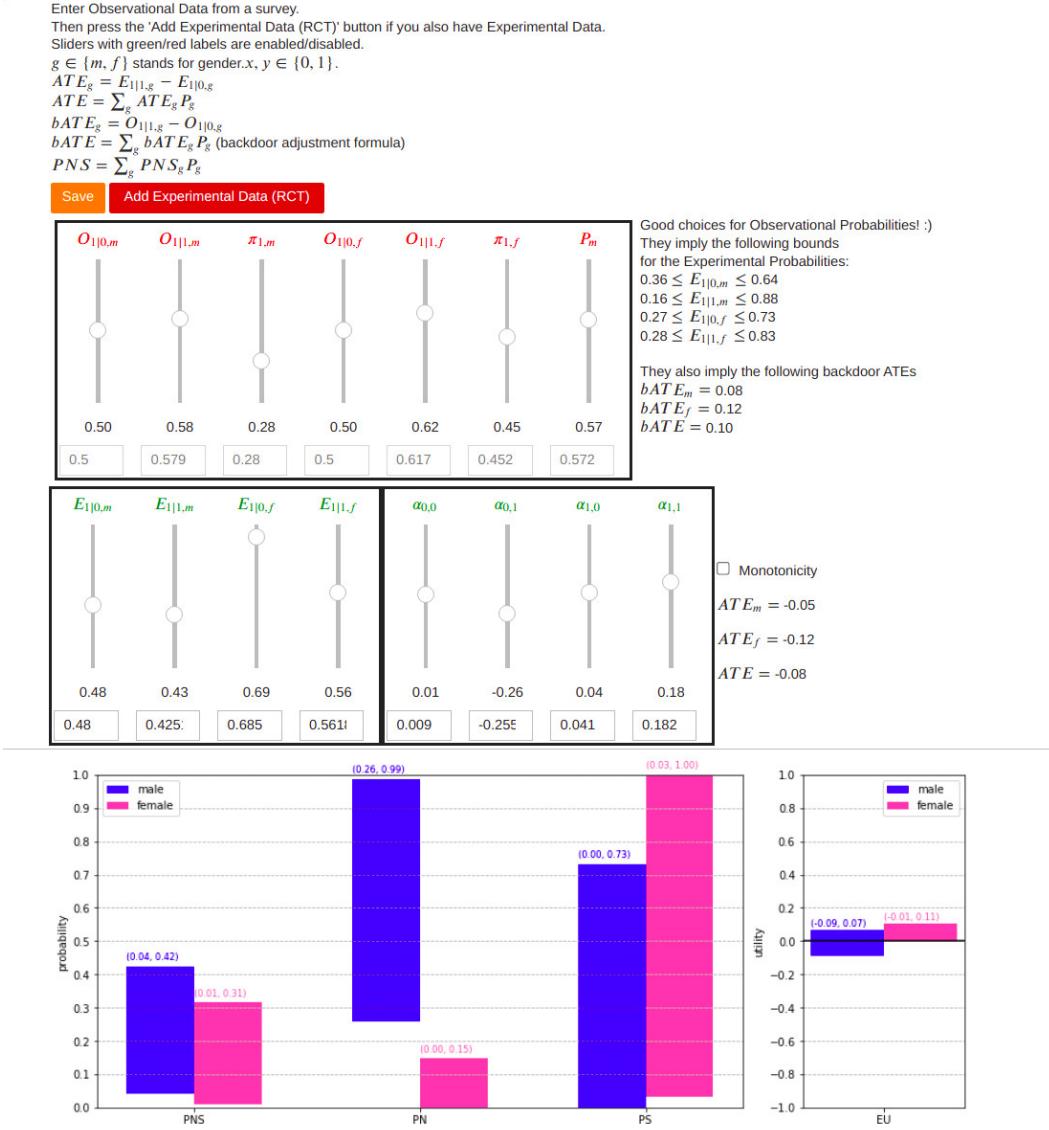


Figure 62.6: Interface for the Python program JudeasRx, with data entered by Boris Sobolev.

Chapter 63

Personalized Expected Utility

This chapter is based on Ref.[29].

This chapter assumes that the reader has already read Chapter 62 on Personalized Treatment Effects. Whereas Chapter 62 is concerned with finding bounds for PNS_z , this chapter will find bounds for EU_z , which is called the Personalized Expected Utility (PEU).

For $y_0, y_1 \in \{0, 1\}$, let us denote the **conditional joint experimental (causal, counterfactual) distribution** by:

$$P_{y_0, y_1 | z} = P(\underline{y}_0 = y_0, \underline{y}_1 = y_1 | z). \quad (63.1)$$

Suppose we are given a **Utility function**¹

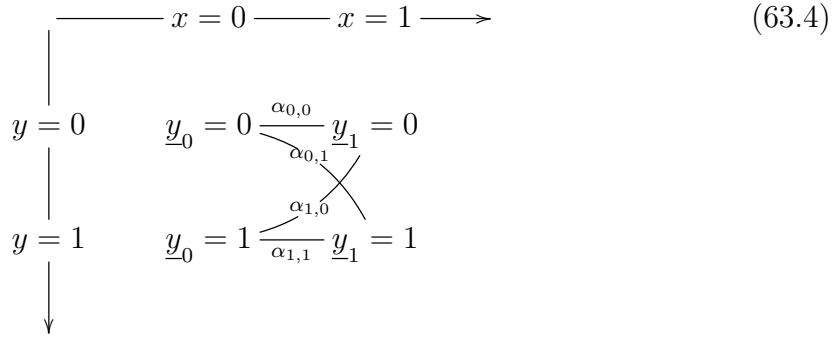
$$\alpha_{y_0, y_1} : \{0, 1\}^2 \rightarrow \mathbb{R} \quad (63.2)$$

Let²

$$\begin{aligned} \beta &= \alpha_{0,1} & (y_0 = 0, y_1 = 1) && \text{compliers} \\ \gamma &= \alpha_{1,1} & (y_0 = 1, y_1 = 1) && \text{always takers} \\ \theta &= \alpha_{0,0} & (y_0 = 0, y_1 = 0) && \text{never takers} \\ \delta &= \alpha_{1,0} & (y_0 = 1, y_1 = 0) && \text{defiers} \end{aligned} \quad (63.3)$$

¹Ref.[29] refers to the utility function as the benefit function.

²The notation $\beta, \gamma, \theta, \delta$ won't be used again in this chapter. We mention it here so the reader can translate to and from our equations and the equations in Ref.[29] where this $\beta, \gamma, \theta, \delta$ notation is used.



Define the **personalized expected utility (PEU)** by

$$EU = E[\alpha_{\underline{y}_0, \underline{y}_1}] = \alpha_{i,j} P_{i,j} \quad (63.5)$$

and the **conditional PEU** by

$$EU_z = E_{|z}[\alpha_{\underline{y}_0, \underline{y}_1}] = \alpha_{i,j} P_{i,j|z} \quad (63.6)$$

Note that

$$EU = \sum_z P(z) EU_z. \quad (63.7)$$

Above, we are using the Einstein summation convention (repeated indices are to be summed over) for the indices $i, j \in \{0, 1\}$.

Compare the definition of EU_z with that of two other causal effects used in his book:

$$PNS_z = P_{0,1||z} = P(\text{compliers}|z) \quad (63.8)$$

$$Uplift = ATE_z = E_{1|1,z} - E_{1|0,z} \quad (63.9)$$

As we shall see, EU_z contains the information in PNS_z and ATE_z , plus much more.

One can find the stratum z^* such that $z^* = \operatorname{argmax} ATE_z$ (as is done A/B=RCT testing) or $z^* = \operatorname{argmax} EU_z$ or $z^* = \operatorname{argmax} PNS_z$. This is called the **unit selection problem**. It's called "unit selection" rather than "stratum selection" because once the stratum z^* is found, one can find a unit (i.e., individual) within that stratum.

63.1 Goal of PEU Theory

Everything that we said in Chapter 62 in the section entitled "Goal of PTE Theory" applies to PEU theory too, if we just replace PNS_z by EU_z . As in Chapter 62, we

will consider two types of bounds (for EU_z instead of PNS_z): (1) Bounds for an unspecified bnet, (2) Bounds for specific bnet families.

Here is an explanation of why EU_z varies within a bounded region, something that might not be obvious to the beginner. In Fig.63.1, we represent the utility function α_{y_0,y_1} by a unit vector $\hat{\alpha}$, the probability distribution $P_{y_0,y_1|z}$ by a vector \vec{P} , and EU_z by the dot product $\vec{P} \cdot \hat{\alpha}$. When the probability vector \vec{P} varies within a bounded region (shown in green), this causes the dot product $EU_z = \vec{P} \cdot \hat{\alpha}$ to vary within a bounded interval (also shown in green).

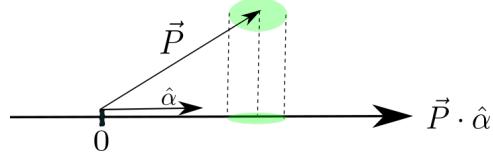


Figure 63.1: Bounds (shown in green) on the probability vector \vec{P} induce bounds (shown in green) on the dot product $EU_z = \vec{P} \cdot \hat{\alpha}$. Here $\hat{\alpha}$ is a unit vector that stands for a normalized utility function. $\hat{\alpha}$ does not vary but \vec{P} does.

63.2 Bnets for PEU Theory

Everything that we said in Chapter 62 in the section entitled “Bnets for PTE Theory” applies to PEU theory too, if we just replace PNS_z by EU_z .

63.3 Bounds on EU for unspecified bnet

Define the **balanced utility** by

$$\alpha_B = \alpha_{0,0} + \alpha_{1,1} , \quad (63.10)$$

the **unbalanced utility** by

$$\alpha_U = \alpha_{1,0} + \alpha_{0,1} , \quad (63.11)$$

and their difference by:

$$\sigma = \alpha_U - \alpha_B . \quad (63.12)$$

We will also use the abbreviations:

$$\alpha_{1-0,j} = \alpha_{1,j} - \alpha_{0,j} \quad (63.13)$$

and

$$\alpha_{j,1-0} = \alpha_{j,1} - \alpha_{j,0} . \quad (63.14)$$

Claim 106 *In general,*

$$\begin{cases} \max p_{[1\dots 4]} \leq EU_z \leq \min p_{[5\dots 8]} & \text{if } \sigma < 0 \\ \max p_{[5\dots 8]} \leq EU_z \leq \min p_{[1\dots 4]} & \text{if } \sigma > 0 \end{cases} \quad (63.15)$$

where³

$$\begin{cases} p_1 = \alpha_{0,1-0} E_{1|1,z} + \alpha_{j,0} E_{j|0,z} \\ p_2 = \alpha_{0-1,1} E_{0|0,z} + \alpha_{1,j} E_{j|1,z} \\ p_3 = p_5 + \sigma O_{*,*|z} \\ p_4 = p_1 - \sigma E_{1|0,z} + \sigma(1 - O_{*,*|z}) \end{cases} \quad (63.16)$$

$$\begin{cases} p_5 = \alpha_{1,1-0} E_{1|1,z} + \alpha_{j,0} E_{j|0,z} \\ p_6 = p_1 - \sigma E_{1|0,z} \\ p_7 = p_5 - \sigma E_{1|0,z} + \sigma P(\underline{y} = 1|z) \\ p_8 = p_1 - \sigma P(\underline{y} = 1|z) \end{cases} \quad (63.17)$$

proof: See Ref[29] or use the MRP (Matrix Representation of Probabilities) defined in Chapter 62.

QED

Later on, we will see that under monotonicity, these bounds collapse to a point estimate of EU_z . But what if monotonicity doesn't hold? In such cases, one can use the midpoint of the bounds as a non-rigorous point estimate of EU_z .

Claim 107 *In general,*

$$P_{0,0|z} = E_{0|1,z} - P_{1,0|z} \quad (63.18)$$

$$P_{1,1|z} = E_{1|0,z} - P_{1,0|z} \quad (63.19)$$

$$P_{0,1|z} = 1 - P_{0,0|z} - P_{1,1|z} - P_{1,0|z} \quad (63.20)$$

$$= 1 - E_{0|1,z} - E_{1|0,z} + P_{1,0|z} \quad (63.21)$$

proof:

$$P_{0,0|z} + P_{1,0|z} = P(\underline{y}_1 = 0|z) = E_{0|1,z} \quad (63.22)$$

³ $p_{[a\dots b]} = (p_a, p_{a+1}, \dots, p_b)$ for integers a, b such that $a < b$.

$$P_{1,1|z} + P_{1,0|z} = P(\underline{y}_0 = 1|z) = E_{1|0,z} \quad (63.23)$$

QED

Recall the definition of monotonicity. Monotonicity holds iff

$$P(\underline{y}_0 = 1, \underline{y}_1 = 0) = 0. \quad (63.24)$$

Claim 108 *In general,*

$$EU_z = \alpha_{0,0}P_{0,0|z} + \alpha_{1,1}P_{1,1|z} + \alpha_{0,1}P_{0,1|z} + \alpha_{1,0}P_{1,0|z} \quad (63.25)$$

Hence, if monotonicity holds, or $\alpha_{1,0} = 0$, then

$$EU_z = \alpha_{0,0}P_{0,0|z} + \alpha_{1,1}P_{1,1|z} + \alpha_{0,1}P_{0,1|z} \quad (63.26)$$

proof: Trivial

QED

Claim 108 is trivial, but it provides a good motivation and inspiration for the following less trivial claim.

Claim 109 *In general,*

$$EU_z = \alpha_{0,0} \underbrace{E_{0|1,z}}_{=1-E_{1|1,z}} + \alpha_{1,1}E_{1|0,z} + \alpha_{0,1} \underbrace{(1 - E_{0|1,z} - E_{1|0,z})}_{=E_{1|1,z}-E_{1|0,z}=ATE_z} + \sigma P_{1,0|z} \quad (63.27)$$

Hence, if monotonicity holds, or $\sigma = 0$, then

$$(EU_z)_{\sigma=0} = \alpha_{0,0}E_{0|1,z} + \alpha_{1,1}E_{1|0,z} + \alpha_{0,1}ATE_z \quad (63.28)$$

proof: In general,

$$EU_z = \begin{cases} \alpha_{0,0}P_{0,0|z} \\ + \alpha_{1,1}P_{1,1|z} \\ + \alpha_{0,1}P_{0,1|z} \\ + \alpha_{1,0}P_{1,0|z} \end{cases} = \begin{cases} \alpha_{0,0}(E_{0|1,z} - P_{1,0|z}) \\ + \alpha_{1,1}(E_{1|0,z} - P_{1,0|z}) \\ + \alpha_{0,1}(1 - E_{0|1,z} - E_{1|0,z} + P_{1,0|z}) \\ + \alpha_{1,0}P_{1,0|z} \end{cases} \quad (63.29)$$

Hence,

$$EU_z = (EU_z)_{\sigma=0} + \sigma P_{1,0|z} \quad (63.30)$$

QED

Perhaps this will make Claim 109 less mysterious to you. Note that Claims 108 and 109 imply the following:

$$\left[\frac{\partial EU_z}{\partial \alpha_{1,0}} \right]_{\alpha_{0,0}, \alpha_{1,1}, \alpha_{0,1}} = \left[\frac{\partial EU_z}{\partial \sigma} \right]_{\alpha_{0,0}, \alpha_{1,1}, \alpha_{0,1}} = P_{1,0|z} \quad (63.31)$$

63.4 Bounds on EU for specific bnet families

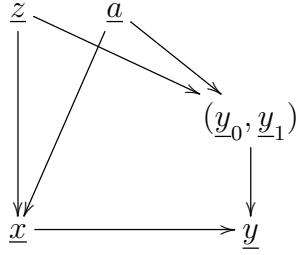


Figure 63.2: Bnet for $\underline{x} \rightarrow \underline{y}$ with 2 observed confounders $\underline{z}, \underline{a}$.

Fig.63.2 gives an example of a bnet that satisfies $(\underline{z} \cup \underline{a}) \cap de(\underline{x}) = \emptyset$.

Claim 110 *If $(\underline{z} \cup \underline{a}) \cap de(\underline{x}) = \emptyset$, then*

$$L \leq PNS_z \leq U \quad (\text{see Claim 97}) \quad (63.32)$$

$$\begin{cases} p_1 + \sigma U \leq EU_z \leq p_1 + \sigma L & \text{if } \sigma < 0 \\ p_1 + \sigma L \leq EU_z \leq p_1 + \sigma U & \text{if } \sigma > 0 \end{cases} \quad (63.33)$$

where

$$p_1 = \alpha_{0,1-0} E_{1|1,z} + \alpha_{j,0} E_{j|0,z} \quad (63.34)$$

$$L = \sum_a P(a|z) \max \left\{ \begin{array}{c} 0 \\ E_{1|1,a,z} - E_{1|0,a,z} \\ E_{0|0,a,z} - P(\underline{y} = 0|a, z) \\ E_{1|1,a,z} - P(\underline{y} = 1|a, z) \end{array} \right\} \quad (63.35)$$

$$U = \sum_a P(a|z) \min \left\{ \begin{array}{c} E_{1|1,a,z} \\ E_{0|0,a,z} \\ O_{*,*|a,z} \\ E_{*,*,a,z} - O_{*,*,a,z} \end{array} \right\} \quad (63.36)$$

proof:

QED

In Chapter 62, we defined what we mean when we say that \underline{z} is a partial or pure mediator relative to $(\underline{x}, \underline{y})$. Fig.63.3 is a copy of Fig.62.5.

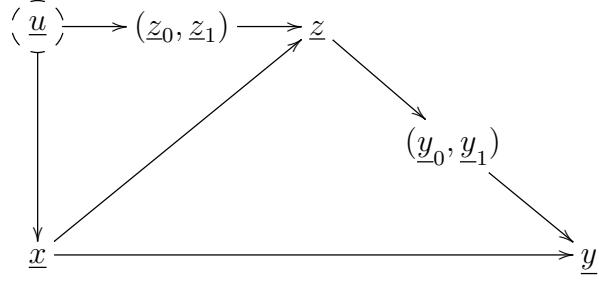


Figure 63.3: Example of a bnet in which \underline{z} is a partial mediator relative to $(\underline{x}, \underline{y})$. If \underline{z} is a partial mediator relative to $(\underline{x}, \underline{y})$, and, in addition, there is no arrow $\underline{x} \rightarrow \underline{y}$, we say \underline{z} is a pure mediator relative to $(\underline{x}, \underline{y})$.

Claim 111 (*Bounds if \underline{z} is partial mediator*)

If \underline{z} is a partial mediator relative to $(\underline{x}, \underline{y})$, and $\underline{x}, \underline{y}, \underline{z} \in \{0, 1\}$, then

$$L \leq PNS_z \leq U \quad (\text{see Claim 104}) \quad (63.37)$$

$$\begin{cases} p_1 + \sigma U \leq EU_z \leq p_1 + \sigma L & \text{if } \sigma < 0 \\ p_1 + \sigma L \leq EU_z \leq p_1 + \sigma U & \text{if } \sigma > 0 \end{cases} \quad (63.38)$$

where

$$p_1 = \alpha_{0,1-0} E_{1|1,z} + \alpha_{j,0} E_{j|0,z} \quad (63.39)$$

$$L = \max \left\{ \begin{array}{c} 0 \\ E_{1|1,z} - E_{1|0,z} \\ E_{0|0,z} - P(\underline{y} = 0|z) \\ E_{1|1,z} - P(\underline{y} = 1|z) \end{array} \right\} \quad (63.40)$$

$$U = \min \left\{ \begin{array}{c} E_{1|1,z} \\ E_{0|0,z} \\ O_{*,*|z} \\ E_{*,*|z} - O_{*,*|z} \\ new \end{array} \right\} \quad (63.41)$$

$$new = \sum_{z_0, z_1} \min \left\{ \begin{array}{c} O_{1|1,z_1,z} \\ O_{0|0,z_0,z} \end{array} \right\} \min \left\{ \begin{array}{c} P(z_1|z) \\ P(z_0|z) \end{array} \right\} \quad (63.42)$$

Claim 112 (*Bounds if \underline{z} is pure mediator*)

If \underline{z} is a pure mediator relative to $(\underline{x}, \underline{y})$, and $\underline{x}, \underline{y}, \underline{z} \in \{0, 1\}$, then Claim 111 for a partial mediator is still valid, except that now “new” equals

$$new = \sum_{z_0, z_1} \mathbb{1}(z_0 \neq z_1) \min \left\{ \begin{array}{c} O_{1|\underline{x}=z_1, z} \\ O_{0|\underline{x}=z_0, z} \end{array} \right\} \min \left\{ \begin{array}{c} P(\underline{z} = z_1 | \underline{x} = 1, z) \\ P(\underline{z} = z_0 | \underline{x} = 0, z) \end{array} \right\} \quad (63.43)$$

Chapter 64

Potential Outcomes and Beyond

This chapter is based on Ref.[8], a book by Stephen Cunningham entitled “Causal inference: the mixtape”.

The theory of potential outcomes (PO) was for the most part invented in a seminal 1974 paper by Donald B. Rubin. Rubin has also made important extensions to PO theory since 1974. However, he does not use Pearl’s causal DAGs to discuss PO theory. Pearl has shown that PO theory can be substantially clarified and extended by using the language of causal DAGs. The d-separation theorem and do operator that we discuss in Chapters 21 and 64 are especially useful in this regard. In this chapter, we stress the connection of PO theory to Pearl’s causal DAGs and bnets.

σ	d^σ	y^σ	$y^\sigma(0)$	$y^\sigma(1)$
Edith	0	5	5	.
Frank	0	7	7	.
George	0	8	8	.
Hank	0	10	10	.
Andy	1	10	.	10
Ben	1	5	.	5
Chad	1	16	.	16
Daniel	1	3	.	3

Table 64.1: PO dataset describing whether individual σ took a treatment drug ($d^\sigma = 1$) or didn’t ($d^\sigma = 0$). The treatment outcome is measured by the real number y^σ .

Suppose a **population of individuals** $\sigma = 0, 1, 2, \dots, nsam - 1$ is given ($d^\sigma = 1$) or is not given ($d^\sigma = 0$) a **treatment decision** d^σ , and that the **treatment outcome (i.e., response)** is measured by a real number y^σ . Table 64.1 gives a possible **PO dataset** for this scenario. As you can see from that table, each individual either takes a drug or doesn’t, but not both. PO theory can be viewed as a **missing data (MD) problem**. MD problems are discussed in Chapter 53. However, the PO MD problem is much more specialized than the generic MD problems discussed in Chapter 53. In the PO MD problem, we can fill in the blank cells by matching each

individual that took the drug with another *similar* individual that didn't. We will have much more to say about this matching strategy later in this chapter.

One can define similar individuals as individuals that have the same value for nx features $x^\sigma = (x_i^\sigma)_{i=0,1,\dots,nx-1}$. One can add to Table 64.1 nx extra columns giving the value of the feature vector x^σ for each individual. Members of a population with the same x^σ are referred to as a **subpopulation or stratum (i.e., layer)**.

In a **randomized controlled trial (RCT)**¹, the effect of the variable x^σ on the value of d^σ is eliminated by randomizing the population and therefore making the effect of x^σ on d^σ average out to zero. However, there are many situations in which carrying out an RCT is not possible or desirable. PO theory is a way of predicting the result of an RCT in situations where doing a real RCT is not possible or desirable.

In this chapter, x^σ will be called the confounders. Implicit throughout this chapter is the assumption that there are **no unmeasured confounders**. Because if there are some unmeasured confounders, those can send secret messages that influence the value that d^σ takes. This would ruin the predictions of someone trying to predict the results of an RCT without being privy to those secret messages. When there are **some unmeasured confounders**, it might still be possible to predict the effect of an RCT. This might be possible using instrumental variables. See Chapter 38 for a discussion of **instrumental variables**.

64.1 G and G_{den} bnets, the starting point bnets

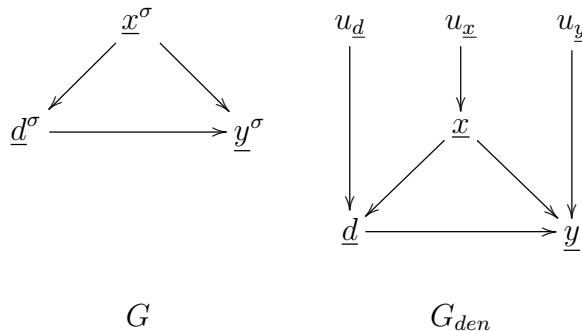


Figure 64.1: Bnets G and G_{den} are our starting point in discussing PO theory. G is for a single individual σ of the population. Bnet G_{den} is the DEN counterpart to G . DEN (Deterministic with External Noise) bnets are discussed in Chapter 44.

In this chapter, we will abbreviate $\underline{X}[\sigma] = \underline{X}^\sigma$ for $X \in \{d, x, y\}$, where $\sigma \in \{0, 1, 2, \dots, nsam - 1\}$.

¹The term **A/B test** is often used to mean an RCT where A and B are the treated and control groups. However, sometimes the term is used to refer to an experiment that conditions on confounders, which violates the definition of an RCT, and is the same as a PO test.

For each individual (a.k.a. unit, sample) $\sigma = 0, 1, 2, \dots, nsam - 1$, let:
 $\underline{d}^\sigma \in \{0, 1\}$ be the treatment decision or drug dose. It equals 1 if treated and 0 if untreated.

$y^\sigma \in \mathbb{R}$ be the treatment potential outcome

\underline{x}^σ be the column vector of treatment confounders (a.k.a. covariates because they are often used as covariates (i.e., independent variables) in linear regression.)

Consider bnets G and G_{den} in Fig.64.1. G reflects the language used in Ref.[8] to discuss PO theory. And G_{den} reflects the language that Judea Pearl prefers to use to discuss PO theory. Both languages are equivalent. To go from one language to the other, one need only perform the following swaps, where \underline{u} is the external noise of the DEN bnet.

$$\underline{X}^\sigma \leftrightarrow \underline{X}(\underline{u}) \text{ for } X \in \{d, x, y\}.$$

$$P(\sigma) = \frac{1}{nsam} \leftrightarrow P(u)$$

$$\sum_\sigma P(\sigma)(\cdot) \leftrightarrow \sum_u P(u)(\cdot)$$

The TPMs, printed in blue, for the bnet G in Fig.64.1, are as follows:

$$P(x^\sigma) = P_{\underline{x}}(x^\sigma) \quad (64.1)$$

$$P(d^\sigma|x^\sigma) = P_{\underline{d}|\underline{x}}(d^\sigma|x^\sigma) \quad (64.2)$$

$$P(y^\sigma|d^\sigma, x^\sigma) = P_{y|d,x}(y^\sigma|d^\sigma, x^\sigma) \quad (64.3)$$

Now let:

$\underline{d} \in \{0, 1\}$ be the treatment decision. It equals 1 if treated and 0 if untreated

$\underline{y} \in \mathbb{R}$ be the treatment potential outcome

\underline{x} be the column vector of treatment confounders (a.k.a. covariates)

$\underline{u} = (\underline{u}_d, \underline{u}_x, \underline{u}_y)$ be the external noise

The TPMs, printed in blue, for the bnet G_{den} in Fig.64.1, are as follows:

$$P(x|u_{\underline{x}}) = \mathbb{1}(x = u_{\underline{x}}) \quad (64.4)$$

$$P(d|x, u_d) = \mathbb{1}(d = f_d(x, u_d)) \quad (64.5)$$

$$P(y|d, x, u_y) = \mathbb{1}(y = f_y(d, x, u_y)) \quad (64.6)$$

If we linearize f_y in Eq.(64.6), we get

$$\underline{y} = \delta \underline{d} + \beta \underline{x} + \underline{u}_y , \quad (64.7)$$

where $\delta, \beta \in \mathbb{R}$. Assuming that $\underline{x}, \underline{y} \in \mathbb{R}$ and $\underline{d} \in \{0, 1\}$, Eq.(64.7) can be plotted. The resulting plot is given in Fig.64.2. This plot is a very special case of the PO problem, but it gives a crude idea of the “effects” $\delta = y(1) - y(0)$ that PO theory gives estimates for. Any individual participating in the experiment experiences either $y(1)$ or $y(0)$, but not both.

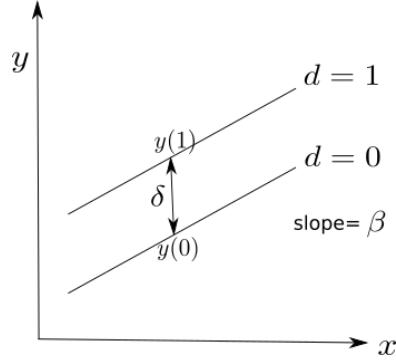


Figure 64.2: Plot of Eq.(64.7)

64.2 G bnet with nodes $y^\sigma(0), y^\sigma(1)$ added to it.

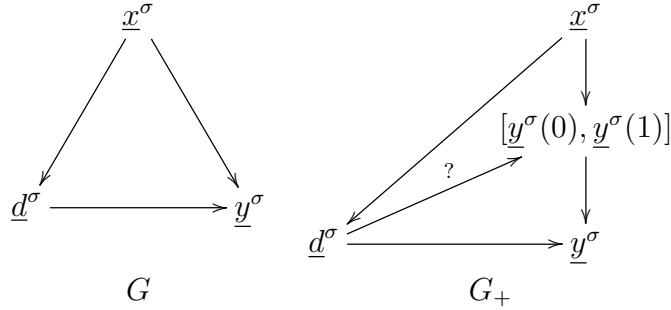


Figure 64.3: Bnet G_+ is bnet G with two new nodes $\underline{y}^\sigma(0)$ and $\underline{y}^\sigma(1)$ added to it. The tuple node $[\underline{y}^\sigma(0), \underline{y}^\sigma(1)]$ can also be represented by two nodes $\underline{c} \rightarrow \underline{y}(\underline{c})$, where $\underline{c} \in \{0, 1\}$.

Consider Fig.64.3. Bnet G_+ was obtained by adding two new nodes $\underline{y}^\sigma(0)$ and $\underline{y}^\sigma(1)$ to bnet G . The TPMs, printed in blue, for bnet G_+ , are as follows. Note that we define them in terms of the TPMs for bnet G .

$$P(x^\sigma) = P_{\underline{x}}(x^\sigma) \quad (64.8)$$

$$P(d^\sigma|x^\sigma) = P_{\underline{d}|\underline{x}}(d^\sigma|x^\sigma) \quad (64.9)$$

For $c \in \{0, 1\}$,

$$P(y^\sigma(c)|d^\sigma, x^\sigma) = \begin{cases} P_{\underline{y}(c)|\underline{d}, \underline{x}}(y^\sigma(c)|d^\sigma, x^\sigma) & \text{if INCLUDE arrow with question mark} \\ P_{\underline{y}(c)|\underline{x}}(y^\sigma(c)|x^\sigma) & \text{if EXCLUDE arrow with question mark} \end{cases} \quad (64.10)$$

$$P(y^\sigma|y^\sigma(0), y^\sigma(1), d^\sigma) = \mathbb{1}(y^\sigma = d^\sigma y^\sigma(1) + (1 - d^\sigma)y^\sigma(0)) \quad (64.11a)$$

$$= \mathbb{1}(y^\sigma = y^\sigma(d^\sigma)) \quad (64.11b)$$

Eq.(64.11) is often referred to as the **SUTVA assumption**.

If we sum over the nodes $\underline{y}(0)$ and $\underline{y}(1)$ of this bnet, we should get the bnet G . This is easy to check. Indeed,

$$P(y^\sigma|d^\sigma, x^\sigma) = \sum_{y^\sigma(0)} \sum_{y^\sigma(1)} \mathbb{1}(y^\sigma = y^\sigma(d^\sigma)) P(y^\sigma(0)|d^\sigma, x^\sigma) P(y^\sigma(1)|d^\sigma, x^\sigma) \quad (64.12)$$

$$= \begin{cases} P_{\underline{y}(0)|\underline{d}, \underline{x}}(y^\sigma|d^\sigma, x^\sigma) & \text{if } d^\sigma = 0 \\ P_{\underline{y}(1)|\underline{d}, \underline{x}}(y^\sigma|d^\sigma, x^\sigma) & \text{if } d^\sigma = 1 \end{cases}. \quad (64.13)$$

Henceforth, we will refer to the case where the question mark arrow is included as the general case, and to the case when it's excluded as the **weak-d limit**. Henceforth, we will first present results for the general case, and then describe how those results change for the weak-d limit. Rubinologists always assume the weak-d limit, but we find that with little effort, we can derive many results for general case, and then compare those results to their weak-d limit. I find such comparisons instructive.

Note that in the general case, $P(\underline{y}(c) = y|\underline{d} = d, x)$ for $c, d \in \{0, 1\}$ are four different probability distributions, and that $P(\underline{y} = y|d = d, x)$ is defined in terms of two of them, the so called **factual distributions** with $c = d$. By measuring y , we can't access the other 2 probability distributions, the so called **counter-factual distributions** with $c \neq d$.

In the weak-d limit, $P(\underline{y}(c) = y|\underline{d} = d, x) = P(\underline{y}(c) = y|x)$ are two probability distributions, and they both can be accessed by measuring y .

64.3 Expected Values of treatment outcome y^σ

It is convenient to define the following expected values of \underline{y}^σ in terms of the TPMs of bnet G_+ :

$$\mathcal{Y}_{c|d,x} = E_{\sigma|d,x}[\underline{y}^\sigma(c)] \rightarrow E_{\underline{y}(c)|d,x}[\underline{y}(c)] = \sum_y y P(\underline{y}(c) = y | \underline{d} = d, x) \quad (64.14)$$

$$\mathcal{Y}_{c|d} = E_{\sigma|d}[\underline{y}^\sigma(c)] \rightarrow E_{\underline{y}(c)|d}[\underline{y}(c)] = \sum_x \mathcal{Y}_{c|d,x} P(x|d) \quad (64.15)$$

$$\mathcal{Y}_{c|x} = E_{\sigma|x}[\underline{y}^\sigma(c)] \rightarrow E_{\underline{y}(c)|x}[\underline{y}(c)] = \sum_d \mathcal{Y}_{c|d,x} P(d|x) \quad (64.16)$$

$$\mathcal{Y}_c = E_\sigma[\underline{y}^\sigma(c)] \rightarrow E_{\underline{y}(c)}[\underline{y}(c)] = \sum_{x,d} \mathcal{Y}_{c|d,x} P(x, d) \quad (64.17)$$

Note that in the weak-d limit,

$$\mathcal{Y}_c = \sum_x \mathcal{Y}_{c|d,x} P(x) . \quad (64.18)$$

Note also that in the weak-d limit, $\mathcal{Y}_{c|d,x}$ is independent of d , but $\mathcal{Y}_{c|d}$ can depend on d if $P(x|d)$ depends on d .

$\mathcal{Y}_{0|0}, \mathcal{Y}_{1|1}$ are said to be **factual** (indicating compliant patients) whereas $\mathcal{Y}_{0|1}, \mathcal{Y}_{1|0}$ are said to be **counterfactual** (indicating non-compliant patients).

Also let

$$\mathcal{Y}_{|d,x} = E_{\sigma|d,x}[\underline{y}^\sigma] \rightarrow E_{\underline{y}|d,x}[\underline{y}] = \sum_y y P(\underline{y} = y | \underline{d} = d, x) \quad (64.19)$$

$$\mathcal{Y}_{|d} = E_{\sigma|d}[\underline{y}^\sigma] \rightarrow E_{\underline{y}|d}[\underline{y}] = \sum_x \mathcal{Y}_{|d,x} P(x|d) \quad (64.20)$$

$$\mathcal{Y}_{|x} = E_{\sigma|x}[\underline{y}^\sigma] \rightarrow E_{\underline{y}|x}[\underline{y}] = \sum_d \mathcal{Y}_{|d,x} P(d|x) \quad (64.21)$$

$$\mathcal{Y} = E_\sigma[\underline{y}^\sigma] \rightarrow E_{\underline{y}}[\underline{y}] = \sum_d \mathcal{Y}_{|d,x} P(d, x) \quad (64.22)$$

In the weak-d limit, $\mathcal{Y}_{|d,x}$ is independent of d , but $\mathcal{Y}_{|d}$ can still depend if $P(x|d)$ depends on d , then $\mathcal{Y}_{|d}$ depends on d too.

64.4 Translation Dictionary

Table 64.2 gives a dictionary for translating from the standard PO notation of Ref.[8] to our notation.

In standard PO notation	In our notation
i , individual (i.e., unit, sample) index	σ
$D_i = d_i$, treatment decision	$\underline{d}^\sigma = d^\sigma$
$Y_i = y_i$, treatment outcome	$\underline{y}^\sigma = y^\sigma$
$X_i = x_i$, treatment confounders	$\underline{x}^\sigma = x^\sigma$
$E[Y_i(c)]$	$E_\sigma[\underline{y}^\sigma(c)] = \mathcal{Y}_c$
$E[Y_i(c) D_i = d]$	$E_{\sigma d}[\underline{y}^\sigma(c)] = \mathcal{Y}_{c d}$
$E[Y_i(c) D_i = d, X_i = x]$	$E_{\sigma d,x}[\underline{y}^\sigma(c)] = \mathcal{Y}_{c d,x}$
$E[Y_i]$	$E_\sigma[\underline{y}^\sigma] = \mathcal{Y}$
$E[Y_i D_i = d]$	$E_{\sigma d}[\underline{y}^\sigma] = \mathcal{Y}_{ d}$
$E[Y_i D_i = d, X_i = x]$	$E_{\sigma d,x}[\underline{y}^\sigma] = \mathcal{Y}_{ d,x}$

Table 64.2: Dictionary for translating from standard PO notation of Ref.[8] to our notation. $c, d \in \{0, 1\}$.

64.5 $\mathcal{Y}_{|d,x} = \mathcal{Y}_{d|d,x}$ (SUTVA)

Claim 113 ²

$$\mathcal{Y}_{|d,x} = \mathcal{Y}_{d|d,x} \quad (64.23)$$

$$\mathcal{Y}_{|d} = \mathcal{Y}_{d|d} \quad (64.24)$$

proof:

$$\mathcal{Y}_{|d,x} = \sum_y y P(\underbrace{\underline{y}}_{=y(d) \text{ by Eq.(64.11)}} = y | d, x) \quad (64.25)$$

$$= \mathcal{Y}_{d|d,x}. \quad (64.26)$$

Applying $\sum_x P(x|d)$ to both sides of Eq.(64.23) gives Eq.(64.24).

QED

²In the standard PO notation, this is the frequently used identity

$$E[Y|D = d, x] = E[Y(d)|D = d, x]$$

.

64.6 Conditional Independence Assumption (CIA)

The **Conditional Independence Assumption** (CIA) is said to hold if

$$(\underline{y}^\sigma(0), \underline{y}^\sigma(1), \underline{y}^\sigma)_{\&} \perp_P \underline{d}^\sigma | \underline{x}^\sigma. \quad (64.27)$$

This is satisfied by G_+ in the weak-d limit. To prove this, check that

$$(\underline{y}^\sigma(0), \underline{y}^\sigma(1), \underline{y}^\sigma) \perp_{G_+} \underline{d}^\sigma | \underline{x}^\sigma \quad (64.28)$$

and then invoke the d-separation theorem (see Chapter 21).

I think CIA only makes sense if the individuals are treatment blind (i.e., have no knowledge of whether they are in the treated or control groups.) Otherwise, that extra knowledge becomes a confounder not being included in x .

A **Randomized Controlled Trial (RCT)** is defined to satisfy Eq.(64.27) without the \underline{x}^σ conditioning; i.e., it satisfies

$$(\underline{y}^\sigma(0), \underline{y}^\sigma(1), \underline{y}^\sigma)_{\&} \perp_P \underline{d}^\sigma. \quad (64.29)$$

This means that in an RCT, the arrow from \underline{x}^σ to \underline{d}^σ in G_+ is omitted.

Note that if we assume both CIA (i.e., weak-d limit) and SUTVA, we get

$$\mathcal{Y}_{|d,x} = E_{\sigma|d,x}[\underline{y}^\sigma] \quad (64.30a)$$

$$= E_\sigma[\underline{y}^\sigma | \underline{d}^\sigma = d, \underline{x}^\sigma = x] \quad (64.30b)$$

$$= E_\sigma[\underline{y}^\sigma(d) | \underline{d}^\sigma = d, \underline{x}^\sigma = x] \text{ (by SUTVA)} \quad (64.30c)$$

$$= E_\sigma[\underline{y}^\sigma(d) | \underline{x}^\sigma = x] \text{ (by CIA)} \quad (64.30d)$$

$$= E_{\sigma|x}[\underline{y}^\sigma(d)] \quad (64.30e)$$

$$= \mathcal{Y}_{d|x}. \quad (64.30f)$$

In an RCT, Eq.(64.30f) is valid without the x conditioning.

64.7 Treatment Effects

A **treatment effect** is a difference of two $\mathcal{Y}_{c|d}$. It is convenient to define the following treatment effects. See Figs.64.4 and 64.5.

- average treatment effect (ATE).

$$\textcolor{red}{ATE} = \mathcal{Y}_1 - \mathcal{Y}_0 = \delta \quad (64.31)$$

- average treatment effect of the treated (ATT)

$$\textcolor{red}{ATT} = \mathcal{Y}_{1|1} - \mathcal{Y}_{0|1} \quad (64.32)$$

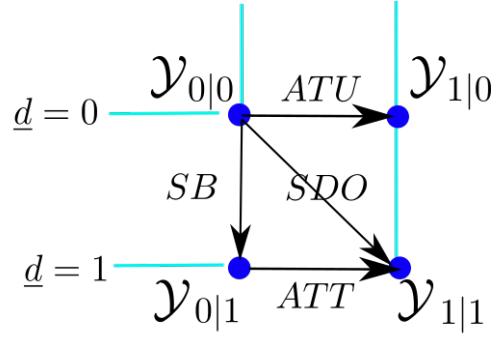


Figure 64.4: Different treatment effects. A treatment effect is a difference of two $\mathcal{Y}_{c|d}$.

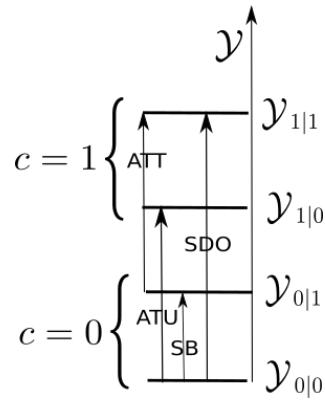


Figure 64.5: Alternative representation of the same information that is contained in Fig.64.4.

- average treatment effect of the untreated (ATU)

$$\textcolor{red}{ATU} = \mathcal{Y}_{1|0} - \mathcal{Y}_{0|0} \quad (64.33)$$

- selection bias (SB)

$$\textcolor{red}{SB} = \mathcal{Y}_{0|1} - \mathcal{Y}_{0|0} \quad (64.34)$$

- simple difference in outcomes (SDO)

$$\textcolor{red}{SDO} = \mathcal{Y}_{1|1} - \mathcal{Y}_{0|0} \quad (64.35)$$

Let

$$\pi_d = P(\underline{d} = d) \quad (64.36)$$

for $d \in \{0, 1\}$.

Note that there exist some linear constraints between these treatment effects.

$$\underbrace{\mathcal{Y}_1 - \mathcal{Y}_0}_{ATE} = \underbrace{\mathcal{Y}_{1|1}\pi_1 + \mathcal{Y}_{1|0}\pi_0}_{\mathcal{Y}_1} - \underbrace{(\mathcal{Y}_{0|1}\pi_1 + \mathcal{Y}_{0|0}\pi_0)}_{\mathcal{Y}_0} \quad (64.37)$$

$$= \underbrace{(\mathcal{Y}_{1|1} - \mathcal{Y}_{0|1})\pi_1}_{ATT} + \underbrace{(\mathcal{Y}_{1|0} - \mathcal{Y}_{0|0})\pi_0}_{ATU} \quad (64.38)$$

$$\underbrace{\mathcal{Y}_{1|1} - \mathcal{Y}_{0|0}}_{SDO} = \underbrace{(\mathcal{Y}_{1|1} - \mathcal{Y}_{0|1})}_{ATT} + \underbrace{\mathcal{Y}_{0|1} - \mathcal{Y}_{0|0}}_{SB} \quad (64.39)$$

$$\begin{aligned} \underbrace{\mathcal{Y}_{1|1} - \mathcal{Y}_{0|0}}_{SDO} &= \underbrace{(\mathcal{Y}_{1|1} - \mathcal{Y}_{0|1})\pi_1 + (\mathcal{Y}_{1|0} - \mathcal{Y}_{0|0})\pi_0}_{ATE} \\ &\quad + \underbrace{\mathcal{Y}_{0|1} - \mathcal{Y}_{0|0}}_{SB} \\ &\quad + \underbrace{(\mathcal{Y}_{1|1} - \mathcal{Y}_{0|1})}_{ATT}\pi_0 \\ &\quad - \underbrace{(\mathcal{Y}_{1|0} - \mathcal{Y}_{0|0})}_{ATU}\pi_0 \end{aligned} \quad (64.40)$$

By virtue of Eq.(64.38),

$$ATT = ATU \implies ATT = ATU = ATE \quad (64.41)$$

and

$$ATE = 0 \iff \frac{ATU}{ATT} = -\left(\frac{\pi_1}{\pi_0}\right). \quad (64.42)$$

Whenever $ATT = ATU$, we will say there is **T-U symmetry**.

In general, $SDO = ATT + SB$, but if there is T-U symmetry, then $SDO = ATE + SB$.

If there is T-U symmetry and zero bias $SB = 0$, then $SDO = ATE = ATT = ATU$.

If there is a null result for an RCT (i.e., $ATE = 0$), T-U symmetry and zero bias $SB = 0$, then $SDO = ATE = ATT = ATU = 0$.

Let

$$\mathcal{Y}_{c,d|x} = \mathcal{Y}_{c|d,x}P(d|x) \quad (64.43)$$

For each $\mathcal{E} \in \{ATE, ATT, ATU, SB, SDO\}$, we can define its restriction \mathcal{E}_x to a fixed stratum x by replacing each $\mathcal{Y}_{c|d}$ with $\mathcal{Y}_{c,d|x}$. For example,

$$ATT = \mathcal{Y}_{1|1} - \mathcal{Y}_{0|1} \text{ so } ATT_x = \mathcal{Y}_{1,1|x} - \mathcal{Y}_{0,1|x}. \quad (64.44)$$

We can calculate \mathcal{E} from \mathcal{E}_x using

$$\mathcal{E} = E_x[\mathcal{E}_x] = \sum_x P(x)\mathcal{E}_x . \quad (64.45)$$

64.8 Insights into what makes treatment effects equal and $\mathcal{Y}_{1|0} = \mathcal{Y}_1$

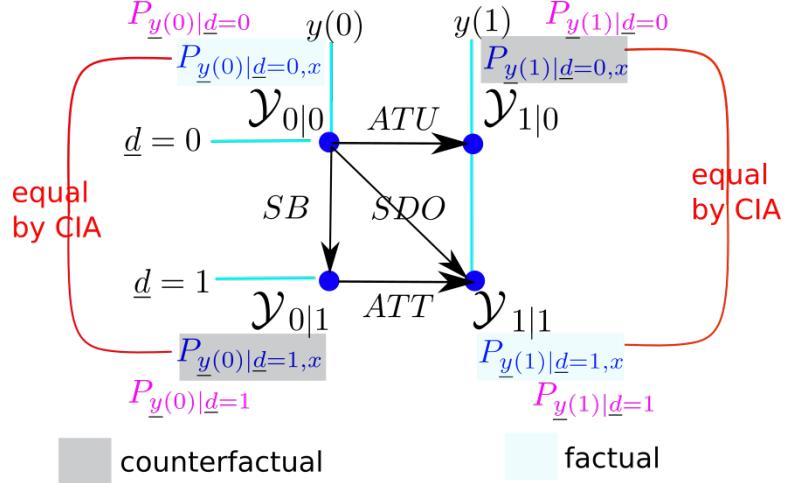


Figure 64.6: Figure 64.4 with added information about probability distributions used to obtain each expected value $\mathcal{Y}_{c|d}$.

1. Is it possible for $SDO = 0$ but $ATE \neq 0$ or vice versa, and what is going on when this is true?
2. What is going on when two treatment effects are equal; for instance, when $ATT = ATU$?
3. When is $\mathcal{Y}_{1|0} = \mathcal{Y}_1$, and what is going on when this is true?

Fig.64.6 gives some intuition about what is going on when any of these things happen.

Recall that each expected value $\mathcal{Y}_{c|d}$ is associated with a probability distribution $P_{\underline{y}(c)|\underline{d},x}$.

$$\mathcal{Y}_{c|d} = \sum_y y \underbrace{\sum_x P_{\underline{y}(c)|\underline{d},x}(y|d, x)P(x|d)}_{P_{\underline{y}(c)|\underline{d}}(y|d)} \quad (64.46)$$

for $c, d \in \{0, 1\}$. Fig.64.6 reminds us of which P is used to generate each \mathcal{Y} . From this figure, we see that

1. A sufficient condition for $SDO = 0$ is that $P_{\underline{y}(1)|\underline{d}=1} = P_{\underline{y}(0)|\underline{d}=0}$. A sufficient condition for $ATE = 0$ is that $P_{\underline{y}(1)|x} = P_{\underline{y}(0)|x}$.
2. A sufficient condition for $ATT = ATU$ is that $P_{\underline{y}(1)|\underline{d}=1,x} - P_{\underline{y}(0)|\underline{d}=1,x}$ equals $P_{\underline{y}(1)|\underline{d}=0,x} - P_{\underline{y}(0)|\underline{d}=0,x}$.
3. A sufficient condition for $\mathcal{Y}_{1|0} = \mathcal{Y}_1$ is that $P_{\underline{y}(1)|\underline{d}=0} = P_{\underline{y}(1)}$. Note that the CIA implies that $P_{\underline{y}(1)|\underline{d}=0,x} = P_{\underline{y}(1)|x}$ always, but this does not imply that $P_{\underline{y}(1)|\underline{d}=0} = P_{\underline{y}(1)}$.

64.9 G_{do+} bnet

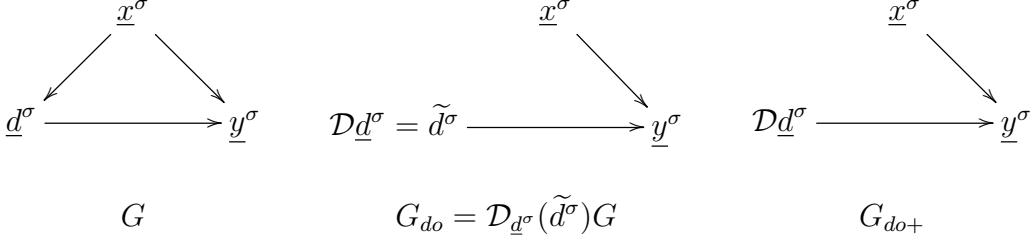


Figure 64.7: Bnet $G_{do} = \mathcal{D}_{\underline{d}^\sigma}(\tilde{d}^\sigma)G$ is obtained by applying the do operator to node \underline{d}^σ of bnet G . Bnet G_{do+} is obtained by adding a prior probability distribution $P(\tilde{d}^\sigma)$ to node $\mathcal{D}\underline{d}^\sigma$ of bnet G_{do} .

Fig.64.7 shows how bnet G_{do} is obtained by applying the do operator to bnet G , and how bnet G_{do+} is obtained by adding a prior probability distribution to one of the nodes of G_{do} . In bnet G_{do} , node \underline{d}^σ has been stripped of all outside influences and fixed to a specific state \tilde{d}^σ . This is what an RCT does.

The TPMs, printed in blue, for the bnets G_{do} and G_{do+} , are as follows. Note that the TPMs for bnets G_{do} and G_{do+} are defined in terms of the TPMs of bnet G .

$$P(x^\sigma) = P_{\underline{x}}(x^\sigma) \quad (64.47)$$

$$P_{\mathcal{D}\underline{d}}(\tilde{d}) = \sum_x P_{\underline{d}|\underline{x}}(\tilde{d}|x)P_{\underline{x}}(x) \quad (64.48)$$

$$P(\tilde{d}^\sigma) = \begin{cases} \delta(\tilde{d}^\sigma, \tilde{d}^\sigma) & \text{for } G_{do} \\ P_{\mathcal{D}\underline{d}}(\tilde{d}^\sigma) & \text{for } G_{do+} \end{cases} \quad (64.49)$$

$$P(y^\sigma | \tilde{d}^\sigma, x^\sigma) = P_{\underline{y}|\underline{d}, \underline{x}}(y^\sigma | \tilde{d}^\sigma, x^\sigma) \quad (64.50)$$

Note that in G_{do} ,

$$P(\underline{y} = y | \mathcal{D}\underline{d} = d, \underline{x} = x) = P(y | \underline{d} = d, x) \quad (64.51)$$

because, by the d-separation theorem, when we condition on the confounder \underline{x} , we block information from being transmitted from \underline{d} to \underline{y} through \underline{x} , and this is equivalent to amputating the arrow $\underline{x} \rightarrow \underline{d}$.

Using Eq.(64.51), we get

$$P(\underline{y} = y | \mathcal{D}\underline{d} = d) = \sum_x P(\underline{y} = y | \mathcal{D}\underline{d} = d, x) P(x | \mathcal{D}\underline{d} = d) \quad (64.52)$$

$$= \sum_x P(y | d, x) P(x) \quad (64.53)$$

Eq.(64.53) is called the **backdoor adjustment formula**. It allows us to express a probability with a do operator in its definition in terms of probabilities without do operators.

64.10 $ACE = ATE$

Define the Average Causal Effect (ACE) by

$$ACE = \sum_y y [P(y | \mathcal{D}\underline{d} = 1) - P(y | \mathcal{D}\underline{d} = 0)] \quad (64.54)$$

$$= \sum_x P(x) \sum_y y [P(y | \underline{d} = 1, x) - P(y | \underline{d} = 0, x)] . \text{ (by Eq.(64.53))} \quad (64.55)$$

Claim 114 *If we assume both SUTVA and CIA (i.e., weak-d limit), then*

$$ACE = ATE \quad (64.56)$$

proof:

$$ACE = \sum_x P(x) \sum_y y [P(\underline{y} = y | \underline{d} = 1, x) - P(\underline{y} = y | \underline{d} = 0, x)] \quad (64.57)$$

$$= \sum_x P(x) [\mathcal{Y}_{1|x} - \mathcal{Y}_{0|x}] \text{ (by SUTVA)} \quad (64.58)$$

$$= \sum_x P(x) [\mathcal{Y}_{1|x} - \mathcal{Y}_{0|x}] \text{ (by CIA)} \quad (64.59)$$

$$= \mathcal{Y}_1 - \mathcal{Y}_0 \quad (64.60)$$

$$= ATE \quad (64.61)$$

QED

We will say there is a **null result in an RCT** when $ACE = 0$. By the previous claim, this is true iff $ATE = 0$ (assuming weak-d limit).

64.11 Good, Bad Controls

The bnet G_+ of Fig.64.3, —the cornerstone of Rubin’s PO theory— is limited in scope and is easily misapplied, leading to incorrect results. The problem is some features that are available and could be conditioned on shouldn’t because they would introduce spurious contributions to ATE. Such features are called “bad controls”³

Examples:

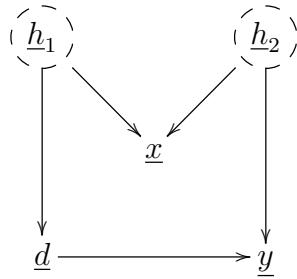


Figure 64.8: In this bnet, \underline{x} is a bad control (i.e., should not be conditioned on). Nodes \underline{h}_1 and \underline{h}_2 are hidden and therefore cannot be conditioned on.

1. Consider the bnet Fig.64.8, which Pearl calls M-bias, because it looks like an M. In that figure, \underline{x} is a “bad control” because calculating ATE by conditioning on it, and using the formula $ATE = \sum_x P(x)[\mathcal{Y}_{1|x} - \mathcal{Y}_{0|x}]$, yields a value of ATE

³In this section, the word “controls” refers to the covariates (i.e., independent variables), other than \underline{d} , in a regression with \underline{y} as target (i.e., dependent) variable. This should not be confused with the control (i.e., untreated) individuals of a RCT.

that is different from ACE . This value for ATE is unacceptable because it does not give the result of an RCT whereas ACE defined in terms of do operators always does. The reason⁴ $ATE \neq ACE$ for this figure is that in it, \underline{x} is a collider node, and conditioning on it allows rather than prevents information to flow from \underline{d} to \underline{y} via the path $\underline{d} - \underline{h}_1 - \underline{x} - \underline{h}_2 - \underline{y}$.

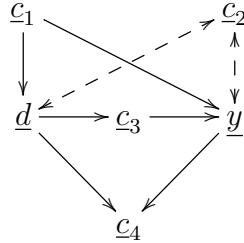


Figure 64.9: In this bnet, node c_1 is a good control and nodes c_2, c_3, c_4 are bad ones.

2. In Fig.64.9, node c_1 is a good control and nodes c_2, c_3, c_4 are bad ones.

Conditioning on c_1 blocks path $\underline{d} - c_1 - \underline{y}$, good

Conditioning on c_2 opens path $\underline{d} - c_2 - \underline{y}$, bad

Conditioning on c_3 blocks path $\underline{d} - c_3 - \underline{y}$, bad

Conditioning on c_4 opens path $\underline{d} - c_4 - \underline{y}$, bad

Pearl et al have a paper (Ref.[7]) that I highly recommend that gives dozens of examples of good, neutral and bad controls in an ACE calculation.

64.12 PO Confounder Sensitivity Analysis

Consider the LDEN bnet of Fig.64.10. whose structure equations, printed in blue, are as follows:

$$\underline{d} = \alpha \underline{x} + \alpha' \underline{u} \quad (64.62)$$

$$\underline{y} = \beta \underline{x} + \beta' \underline{u} + \delta \underline{d} \quad (64.63)$$

Therefore,

$$\underline{u} = \frac{1}{\alpha'} (\underline{d} - \alpha \underline{x}) \quad (64.64)$$

⁴We are using here arguments based on the d-separation theorem which is discussed in Chapter 21.

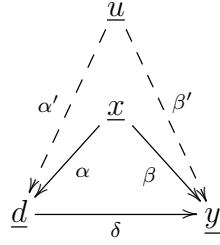


Figure 64.10: LDEN bnet used to do PO confounder sensitivity analysis. Node \underline{u} is an unobserved confounder.

$$\underline{y} = \left(\beta - \frac{\alpha\beta'}{\alpha'} \right) \underline{x} + \left(\delta + \frac{\beta'}{\alpha'} \right) \underline{d} \quad (64.65)$$

$$\mathcal{Y}_{d|x,u} = E[\underline{y}(d)|x, u] = E[\underline{y}|d, x, u] = \beta x + \beta' u + \delta d \quad (64.66)$$

$$\mathcal{Y}_{d|x} = E[\underline{y}(d)|x] = E[\underline{y}|d, x] = \left(\beta - \frac{\alpha\beta'}{\alpha'} \right) x + \left(\delta + \frac{\beta'}{\alpha'} \right) d \quad (64.67)$$

$$ATE = E_x[ATE_x] = E_x[\mathcal{Y}_{1|x} - \mathcal{Y}_{0|x}] = \delta + \frac{\beta'}{\alpha'} \quad (64.68)$$

$$ATE|_{\beta'=0} = \delta = \frac{dy}{d\underline{d}} \quad (64.69)$$

$$ATE - ATE|_{\beta'=0} = \frac{\beta'}{\alpha'} = -\frac{\frac{dy}{d\underline{u}}}{\frac{d\underline{d}}{d\underline{u}}} \quad (64.70)$$

64.13 (SDO, ATE) space

If we substitute $y^\sigma \rightarrow y^\sigma(d^\sigma)$ and $y^{m(\sigma)} \rightarrow y^\sigma(1 - d^\sigma)$ into the estimate Eq.(64.83) for ATE and the estimate Eq.(64.89) for SDO , we get

$$\widehat{ATE}_x = \frac{1}{N_x} \sum_{\sigma \in A_x} (2d^\sigma - 1)[y^\sigma(d^\sigma) - y^\sigma(1 - d^\sigma)] \quad (64.71)$$

$$= \frac{1}{N_x} \sum_{\sigma \in A_x} [y^\sigma(1) - y^\sigma(0)] \quad (64.72)$$

and

$$\widehat{SDO}_x = \frac{1}{N_{1,x}} \sum_{\sigma \in A_x} d^\sigma y^\sigma(d^\sigma) - \frac{1}{N_{0,x}} \sum_{\sigma \in A_x} (1 - d^\sigma) y^\sigma(d^\sigma) \quad (64.73)$$

$$= \frac{1}{N_{1,x}} \sum_{\sigma \in A_{1,x}} y^\sigma(1) - \frac{1}{N_{0,x}} \sum_{\sigma \in A_{0,x}} y^\sigma(0). \quad (64.74)$$

Recall that $\widehat{\mathcal{E}} = E_x[\widehat{\mathcal{E}}_x] = \sum_x \frac{N_x}{N} \widehat{\mathcal{E}}_x$ for $\mathcal{E} \in \{ATE, SDO\}$.

Recall also that $ACE = ATE = 0$ is the null result in an RCT.

Suppose that the treatment outcome y^σ has only two possible values, 0 and 1. Then, $-1 \leq ATE \leq 1$ and $-1 \leq SDO \leq 1$. But does $ATE = 0$ imply $SDO = 0$ or vice versa? Next, we answer that question and more by finding the region of accessibility in the (SDO, ATE) plane, assuming $y^\sigma \in \{0, 1\}$.

σ	d^σ	$y^\sigma(0)$	$y^\sigma(1)$
1	0	1	0
2	0	1	0
3	0	1	0
4	1	1	0
5	1	1	0
6	1	1	0

(a) $ATE = -1$ ($SDO = -1$)
point A

σ	d^σ	$y^\sigma(0)$	$y^\sigma(1)$
1	0	0	1
2	0	0	1
3	0	0	1
4	1	0	0
5	1	0	0
6	1	0	0

(b) $ATE = \frac{1}{2}$ ($SDO = 0$)
point B

σ	d^σ	$y^\sigma(0)$	$y^\sigma(1)$
1	0	0	1
2	0	0	1
3	0	0	1
4	1	0	1
5	1	0	1
6	1	0	1

(c) $ATE = 1$ ($SDO = 1$)
point C

Figure 64.11: Examples of PO datasets. Exploring ATE extremes.

σ	d^σ	$y^\sigma(0)$	$y^\sigma(1)$
1	0	1	1
2	0	1	1
3	0	1	1
4	1	0	0
5	1	0	0
6	1	0	0

(a) $SDO = -1$ ($ATE = 0$)
point D

σ	d^σ	$y^\sigma(0)$	$y^\sigma(1)$
1	0	1	0
2	0	1	0
3	0	1	0
4	1	1	1
5	1	1	1
6	1	1	1

(b) $SDO = 0$ ($ATE = -\frac{1}{2}$)
point E

σ	d^σ	$y^\sigma(0)$	$y^\sigma(1)$
1	0	0	0
2	0	0	0
3	0	0	0
4	1	1	1
5	1	1	1
6	1	1	1

(c) $SDO = 1$ ($ATE = 0$)
point F

Figure 64.12: Examples of PO datasets. Exploring SDO extremes.

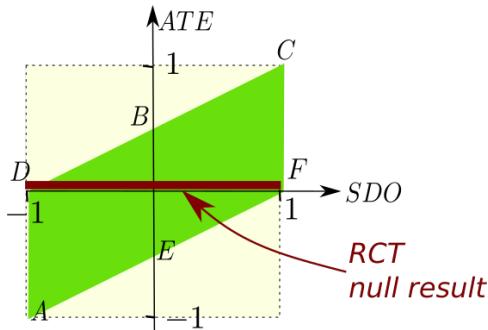


Figure 64.13: Green parallelogram is accessible region in (SDO, ATE) plane, assuming $y^\sigma \in \{0, 1\}$. Each of the six points A, B, ... F corresponds to one of the six tables in Figs. 64.11 and 64.12. Segment DF corresponds to the null result in an RCT.

64.14 Strata-Matching

For a situation described by the bnet G_+ in the weak-d limit, we can match *similar* individuals to fill the blank cells of Table 64.1. By “similar”, we mean that they have the same or almost the same value of x^σ .

The reason the weak-d limit is required is because it implies that $P(y(c)|d = 0, x) = P(y(c)|d = 1, x)$ for $c \in \{0, 1\}$. Hence, we can sample from a known factual ($c = d$) distribution to fill the missing data in the unknown counterfactual ($c \neq d$) distribution.

64.14.1 Exact strata-matching

Estimates of Treatment Effects

For $d \in \{0, 1\}$ and all strata x , define the sets of individuals $A_{d,x} = \{\sigma : d^\sigma = d, x^\sigma = x\}$, $A_x = A_{0,x} \cup A_{1,x}$ and $A = \cup_x A_x$. Let $N_{d,x} = |A_{d,x}|$, $N_x = |A_x|$ and $N = |A|$.

In an exact strata-matching, we match each individual with $d^\sigma = d, x^\sigma = x$ with exactly one individual with $d^\sigma = 1 - d, x^\sigma = x$. Define a map $m : A \rightarrow A$ such that, for each x , and for $d \in \{0, 1\}$, if $\sigma \in A_{d,x}$, then $m(\sigma) \in A_{1-d,x}$. This assumes $A_{0,x}$ and $A_{1,x}$ are non-empty for all x . The purpose of map $m()$ is to fill in the missing data in the PO dataset. See Fig.64.3 for a pictorial representation of this.

	$y^\sigma(0)$	$y^\sigma(1)$
$d^\sigma = 0$	y^σ	$y^{m(\sigma)}$
$d^\sigma = 1$	$y^{m(\sigma)}$	y^σ

Table 64.3: Illustration of the purpose of the map $m()$. Note that $y^\sigma = y^\sigma(d^\sigma)$ and $y^{m(\sigma)} = y^\sigma(1 - d^\sigma)$.

Note that

$$\sum_{\sigma \in A_x} \frac{d^\sigma}{N_{1,x}} = \sum_{\sigma \in A_{1,x}} \frac{1}{N_{1,x}} = 1. \quad (64.75)$$

Thus

$$\sum_{\sigma \in A_x} \frac{d^\sigma}{N_{1,x}} y^\sigma = E_{\sigma|d=1,x}[y^\sigma] = \mathcal{Y}_{1|x} \quad (64.76)$$

Table 64.4 gives estimates of $\mathcal{Y}_{c|d,x}$

Recall that

$$\mathcal{Y}_{c,d|x} = \mathcal{Y}_{c|d,x} P(d|x) \quad (64.77)$$

Hence,

	$y^\sigma(0)$	$y^\sigma(1)$
$d^\sigma = 0$	$\frac{1}{N_{0,x}} \sum_{\sigma \in A_x} (1 - d^\sigma) y^\sigma = \mathcal{Y}_{0 0,x}$	$\frac{1}{N_{0,x}} \sum_{\sigma \in A_x} (1 - d^\sigma) y^{m(\sigma)} = \mathcal{Y}_{1 0,x}$
$d^\sigma = 1$	$\frac{1}{N_{1,x}} \sum_{\sigma \in A_x} d^\sigma y^{m(\sigma)} = \mathcal{Y}_{0 1,x}$	$\frac{1}{N_{1,x}} \sum_{\sigma \in A_x} d^\sigma y^\sigma = \mathcal{Y}_{1 1,x}$

Table 64.4: Estimates of $\mathcal{Y}_{c|d,x}$.

	$y^\sigma(0)$	$y^\sigma(1)$
$d^\sigma = 0$	$\frac{1}{N_x} \sum_{\sigma \in A_x} (1 - d^\sigma) y^\sigma = \mathcal{Y}_{0,0 x}$	$\frac{1}{N_x} \sum_{\sigma \in A_x} (1 - d^\sigma) y^{m(\sigma)} = \mathcal{Y}_{1,0 x}$
$d^\sigma = 1$	$\frac{1}{N_x} \sum_{\sigma \in A_x} d^\sigma y^{m(\sigma)} = \mathcal{Y}_{0,1 x}$	$\frac{1}{N_x} \sum_{\sigma \in A_x} d^\sigma y^\sigma = \mathcal{Y}_{1,1 x}$

Table 64.5: Estimates of $\mathcal{Y}_{c,d|x}$.

$$\mathcal{Y}_{c,d|x} = (N_{d,x} \mathcal{Y}_{c|d,x}) \frac{P(d|x)}{N_{d,x}} \quad (64.78)$$

$$= (N_{d,x} \mathcal{Y}_{c|d,x}) \frac{1}{N_x} \quad (64.79)$$

Table 64.5 gives estimates of $\mathcal{Y}_{c,d|x}$

The treatment effects $\mathcal{E} \in \{ATE, ATT, ATU, SB, SDO\}$ can be estimated from the data via the following estimates.

$$\widehat{ATE}_x = \overbrace{\mathcal{Y}_{1|1,x} P(1|x) + \mathcal{Y}_{1|0,x} P(0|x)}^{\mathcal{Y}_{1|x}} - \overbrace{\mathcal{Y}_{0|1,x} P(1|x) + \mathcal{Y}_{0|0,x} P(0|x)}^{\mathcal{Y}_{0|x}} \quad (64.80)$$

$$= \frac{1}{N_x} [\widehat{ATT}_x N_{1,x} + \widehat{ATU}_x N_{0,x}] \quad (64.81)$$

$$= \frac{1}{N_x} \left[\sum_{\sigma \in A_x} d^\sigma [y^\sigma - y^{m(\sigma)}] + \sum_{\sigma \in A_x} (1 - d^\sigma) [y^{m(\sigma)} - y^\sigma] \right] \quad (64.82)$$

$$= \frac{1}{N_x} \sum_{\sigma \in A_x} (2d^\sigma - 1) [y^\sigma - y^{m(\sigma)}] \quad (64.83)$$

$$\widehat{ATT}_x = \overbrace{\frac{1}{N_x} \sum_{\sigma \in A_x} d^\sigma y^\sigma}^{\mathcal{Y}_{1,1|x}} - \overbrace{\frac{1}{N_x} \sum_{\sigma \in A_x} d^\sigma y^{m(\sigma)}}^{\mathcal{Y}_{0,1|x}} \quad (64.84)$$

$$= \frac{1}{N_x} \sum_{\sigma \in A_x} d^\sigma [y^\sigma - y^{m(\sigma)}] \quad (64.85)$$

$$\widehat{ATU}_x = \overbrace{\frac{1}{N_x} \sum_{\sigma \in A_x} (1 - d^\sigma) y^{m(\sigma)}}^{\mathcal{Y}_{1,0|x}} - \overbrace{\frac{1}{N_x} \sum_{\sigma \in A_x} (1 - d^\sigma) y^\sigma}^{\mathcal{Y}_{0,0|x}} \quad (64.86)$$

$$= \frac{1}{N_x} \sum_{\sigma \in A_x} (1 - d^\sigma) [y^{m(\sigma)} - y^\sigma] \quad (64.87)$$

$$\widehat{SB}_x = \overbrace{\frac{1}{N_x} \sum_{\sigma \in A_x} d^\sigma y^{m(\sigma)}}^{\mathcal{Y}_{0,1|x}} - \overbrace{\frac{1}{N_x} \sum_{\sigma \in A_x} (1 - d^\sigma) y^\sigma}^{\mathcal{Y}_{0,0|x}} \quad (64.88)$$

$$\widehat{SDO}_x = \overbrace{\frac{1}{N_x} \sum_{\sigma \in A_x} d^\sigma y^\sigma}^{\mathcal{Y}_{0,1|x}} - \overbrace{\frac{1}{N_x} \sum_{\sigma \in A_x} (1 - d^\sigma) y^\sigma}^{\mathcal{Y}_{0,0|x}} \quad (64.89)$$

Suppose we do linear regression to fit a hyperplane $y(x)$ to the dataset set $\{(x^\sigma, y^\sigma) : \sigma\}$, and then we calculate $\frac{dy}{dd} = \delta$. Out of all the treatment effects, this δ is probably (?) closest to $ACE = ATE$. Note also that the linear regression method of estimating δ does imputation (guesses missing values) by doing a linear fit. One can also use machine learning to do a non-linear fit. In contrast, the estimates of treatment effects presented in this section do imputation by non-linear strata-matching.

Example, estimation of treatment effects

For $\sigma \in \{1, 2, \dots, 10\}$, define

$$m(\sigma) = \begin{cases} \sigma + 5 & \text{if } \sigma \leq 5 \\ \sigma - 5 & \text{if } \sigma > 5 \end{cases} \quad (64.90)$$

Let $N(\mathcal{S})$ be the number of individuals σ that satisfy condition \mathcal{S} . For example, $N(\underline{d^\sigma} = d)$ is the number of individuals such that $\underline{d^\sigma} = d$.

$$N_1 = N(d^\sigma = 1) = 5 \quad (64.91)$$

$$N_0 = N(d^\sigma = 0) = 5 \quad (64.92)$$

$$N = N_0 + N_1 = 10 \quad (64.93)$$

$$\mathcal{Y}_{1|1} = \frac{1}{N_1} \sum_{\sigma} d^\sigma y^\sigma = \frac{4}{5} \quad (64.94)$$

σ	d^σ	y^σ	$d^\sigma y^\sigma$	$(1 - d^\sigma)y^\sigma$	$d^\sigma y^{m(\sigma)}$	$(1 - d^\sigma)y^{m(\sigma)}$
1	0	0	0	0	0	0
2	0	0	0	0	0	1
3	0	1	0	1	0	1
4	0	1	0	1	0	1
5	0	1	0	1	0	1
6	1	0	0	0	0	0
7	1	1	1	0	0	0
8	1	1	1	0	1	0
9	1	1	1	0	1	0
10	1	1	1	0	1	0

Table 64.6: Estimates of treatment effects are calculated for this example.

$N(d, y)$	$y = 0$	$y = 1$
$d = 0$	2	3
$d = 1$	1	4

Table 64.7: $N(\underline{d}^\sigma = d, \underline{y}^\sigma = y)$ for the data in Table 64.6.

$$\mathcal{Y}_{0|0} = \frac{1}{N_0} \sum_{\sigma} (1 - d^\sigma) y^\sigma = \frac{3}{5} \quad (64.95)$$

$$\mathcal{Y}_{0|1} = \frac{1}{N_1} \sum_{\sigma} d^\sigma y^{m(\sigma)} = \frac{3}{5} \quad (64.96)$$

$$\mathcal{Y}_{1|0} = \frac{1}{N_0} \sum_{\sigma} (1 - d^\sigma) y^{m(\sigma)} = \frac{4}{5} \quad (64.97)$$

$$ATT = \mathcal{Y}_{1|1} - \mathcal{Y}_{0|1} = \frac{1}{5} \quad (64.98)$$

$$ATE = ATT = ATU = SDO = \frac{1}{5}, \quad SB = 0 \quad (64.99)$$

This example is unusual in that it has a single stratum x , and for that stratum, the treated and untreated populations are **balanced** (of equal size). Also, the map $m()$ is 1-1 onto. If, for instance, $m(\sigma) = 6$ for all $\sigma \in A_0$ and $m(\sigma) = 5$ for $\sigma \in A_1$, then ATE, ATT, ATU, SDO would not all be same, and SB would not be zero.

In fact, whenever there is a single balanced stratum and the map $m()$ is 1-1 onto, Eq.(64.99) can be proven to be true using the methods of section 64.8.

64.14.2 Approximate strata-matching

It is very often the case that one can't find for a given individual σ another individual that has opposite d^σ but exactly the same value of x^σ . In such cases, one can discard all matchless individuals. But that would entail a loss of precious information. Instead of discarding orphans, a better way is to relax our demands and match individual σ with another individual $m(\sigma)$ such that x^σ and $x^{m(\sigma)}$ are very close in some metric. Alternatively, the matching individual might not be real; it might be a composite of individuals.

More precisely, for some arbitrary parameter $\epsilon > 0$, and an individual σ , define the **strata-matching set** $\mathcal{M}_\epsilon(\sigma)$ by⁵

$$\mathcal{M}_\epsilon(\sigma) = \{m : d^m = 1 - d^\sigma, \text{dist}(x^\sigma, x^m) \leq \epsilon\}, \quad (64.100)$$

where

$$\text{dist}(x^\sigma, x^m) = [x^\sigma]^T [\Sigma]^{-1} x^m, \quad (64.101)$$

where $\Sigma = \langle x^\sigma, [x^m]^T \rangle$. This metric $\text{dist}(x^\sigma, x^m)$ is called the **Mahalanobis distance**. We will call the case $\epsilon = 0$ an **exact strata-matching**, and the case $\epsilon \neq 0$ an **approximate strata-matching**. To do an approximate strata-matching, replace $y^{m(\sigma)}$ by $\langle y \rangle^{\mathcal{M}_\epsilon(\sigma)}$ in the estimates given above for an exact strata-matching. $\langle y \rangle^{\mathcal{M}_\epsilon(\sigma)}$ is defined by

$$\langle y \rangle^{\mathcal{M}_\epsilon(\sigma)} = \frac{1}{|\mathcal{M}_\epsilon(\sigma)|} \sum_{m \in \mathcal{M}_\epsilon(\sigma)} y^m. \quad (64.102)$$

64.14.3 Unbiased strata-matching estimates

The estimates we obtained via strata-matching are biased because strata-matching, due to its non-uniqueness, introduces noise into the estimate. However, one can define new bias-corrected estimates. Following Ref.[8], we will next find an unbiased estimate of ATT_x using the biased estimate of ATT_x that we obtained by strata-matching.

Let $\widehat{\mathcal{Y}}_{|d,x}$ be an estimate of $\mathcal{Y}_{|d,x} = E_{|d,x}[y]$ that is obtained, for example, via Linear Regression.

Claim 115 *The quantity*

$$\widehat{ATT}_x^{unbi} = \frac{1}{N_x} \sum_{\sigma \in A_x} d^\sigma \left[(y^\sigma - y^{m(\sigma)}) - (\widehat{\mathcal{Y}}_{|0,x^\sigma} - \widehat{\mathcal{Y}}_{|0,x^{m(\sigma)}}) \right] \quad (64.103)$$

⁵One can use an ϵ that depends on σ . For example, let $\epsilon(\sigma, 5)$ satisfy $|\mathcal{M}_{\epsilon(\sigma, 5)}(\sigma)| = 5$.

is an unbiased estimate of ATT_x .

proof:

We begin by assuming a special case of SUTVA. Let

$$\underline{y}^\sigma = \underline{y}(d^\sigma) = \widehat{\mathcal{Y}}_{|d^\sigma, x^\sigma} + \underline{\epsilon}^\sigma \quad (64.104)$$

where

$$\left\langle \widehat{\mathcal{Y}}_{|d^\sigma, x^\sigma}, \underline{\epsilon}^\sigma \right\rangle_\sigma = 0 \quad (64.105)$$

Recall that the biased estimate of ATT_x obtained by strata-matching is

$$\widehat{ATT}_x^{bi} = \frac{1}{N_x} \sum_{\sigma \in A_x} d^\sigma (y^\sigma - y^{m(\sigma)}) \quad (64.106)$$

where σ and $m(\sigma)$ are matched (i.e., $x^\sigma \approx x^{m(\sigma)}$ and $d^{m(\sigma)} = 1 - d^\sigma$).

$$\widehat{ATT}_x^{bi} = \frac{1}{N_x} \sum_{\sigma \in A_x} d^\sigma \left(\widehat{\mathcal{Y}}_{|1, x^\sigma} - \widehat{\mathcal{Y}}_{|0, x^{m(\sigma)}} \right) \quad (64.107)$$

$$+ \frac{1}{N_x} \sum_{\sigma \in A_x} d^\sigma (\epsilon^\sigma - \epsilon^{m(\sigma)}) \quad (64.108)$$

$$\widehat{ATT}_x^{bi} = \underbrace{\frac{1}{N_x} \sum_{\sigma \in A_x} d^\sigma \left(\widehat{\mathcal{Y}}_{|1, x^\sigma} - \widehat{\mathcal{Y}}_{|0, x^\sigma} \right)}_{ATT_x^{unbi}} \quad (64.109)$$

$$+ \underbrace{\frac{1}{N_x} \sum_{\sigma \in A_x} d^\sigma \left(\widehat{\mathcal{Y}}_{|0, x^\sigma} - \widehat{\mathcal{Y}}_{|0, x^{m(\sigma)}} \right)}_{\Delta ATT_x} \quad (64.110)$$

$$+ \underbrace{\frac{1}{N_x} \sum_{\sigma \in A_x} d^\sigma (\epsilon^\sigma - \epsilon^{m(\sigma)})}_{\mathcal{E}_x} \quad (64.111)$$

$$ATT_x^{unbi} = \underbrace{\widehat{ATT}_x^{bi} - \Delta ATT_x}_{\widehat{ATT}_x^{unbi}} - \mathcal{E}_x \quad (64.112)$$

By the Central Limit Theorem, for large N_x , this sum over $\sigma \in A_x$ of i.i.d. summands is normally distributed

$$\sqrt{N_x} ATT_x^{unbi} \sim \mathcal{N}(x; 0, var) \quad (64.113)$$

The reason for the $\sqrt{N_x}$ normalization is that we want the variance to be proportional to N_x .

$$var = N_x \langle ATT_x^{unbi}, ATT_x^{unbi} \rangle \quad (64.114)$$

$$= N_x \langle \widehat{ATT}_x^{unbi}, \widehat{ATT}_x^{unbi} \rangle + N_x \langle \mathcal{E}_x, \mathcal{E}_x \rangle \quad (64.115)$$

QED

64.15 Propensities

It is often the case that the discrete vector \underline{x}^σ has too many possible values to make matching possible. In such cases, it is convenient to map the space of vectors \underline{x}^σ to the real line. One very convenient choice for that map is the **propensity score**, which is defined as

$$g(\underline{x}^\sigma) = P(\underline{d}^\sigma = 1 | \underline{x}^\sigma). \quad (64.116)$$

$P(\underline{d}^\sigma = 1 | \underline{x}^\sigma)$ is easy to calculate from the dataset. $g(x) = N_{1,x}/N_x$.

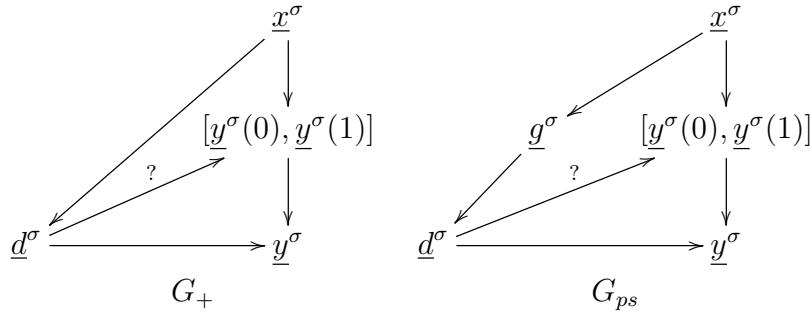


Figure 64.14: Bnet G_{ps} used when doing propensity scoring.

To use the propensity score, one replaces the bnet G_+ by the bnet G_{ps} as shown in Fig.64.14. The TPMs, printed in blue, for the 2 nodes of G_{ps} that differ from the nodes of G_+ , are as follows:

$$P(g^\sigma | x^\sigma) = \delta(g^\sigma, g(x^\sigma)) \quad (64.117)$$

$$P(d^\sigma | g^\sigma) = g^\sigma d^\sigma + (1 - g^\sigma)(1 - d^\sigma) \quad (64.118)$$

Note that these TPMs are self-consistent because

$$P(d|x) = \sum_g P(d|g)P(g|x) \quad (64.119)$$

$$= g(x)d + [1 - g(x)](1 - d) \quad (64.120)$$

$$= P(\underline{d} = 1|x)d + [1 - P(\underline{d} = 1|x)](1 - d) \quad (64.121)$$

$$= P(d|x) \quad (64.122)$$

We would like to do **propensity score strata-matching** by matching g-strata instead of x-strata. PO calculations for x- strata-matching use the TPMs for $P(d|x)$, $P(x)$ and $P(y|d, x)$. To do g- strata-matching using the same equations, but with x replaced by g , we would need to solve for $P(d|g)$, $P(g)$ and $P(y|d, g)$ in terms of $P(d|x)$, $P(x)$ and $P(y|d, x)$. We solve for those next.

From the TPMs for G_{ps} , one has

$$\boxed{P(d|g) = gd + (1 - g)(1 - d)} \quad (64.123)$$

and

$$\boxed{P(g) = \sum_x \overbrace{\delta(g, g(x))}^{P(g|x)} P(x)} . \quad (64.124)$$

Next, note that

$$P(y|d, g) = \sum_x P(y|d, x)P(x|g) \quad (64.125)$$

so we need to find $P(x|g)$. Since

$$P(x|g) = \frac{P(g|x)P(x)}{P(g)} \quad (64.126)$$

$$= \frac{\delta(g, g(x))P(x)}{P(g)} \quad (64.127)$$

we finally get

$$\boxed{P(y|d, g) = \sum_x P(y|d, x) \frac{\delta(g, g(x))P(x)}{P(g)}} . \quad (64.128)$$

Eq.(64.128) looks complicated, but all it is saying is that

$$P(y|d, g)P(g) = \sum_{x \in g^{-1}(g)} P(y|d, x)P(x) , \quad (64.129)$$

where $g^{-1}(g) = \{x : g(x) = g\}$. In general,

$$\sum_x \delta(g, g(x)) = \sum_{x \in g^{-1}(g)} \delta(g, g(x)) \quad (64.130)$$

Recall that for any treatment effect $\mathcal{E} \in \{ATE, ATT, ATU, SB, SDO\}$, we can estimate \mathcal{E}_x , and then calculate \mathcal{E} from it, using

$$\mathcal{E} = \sum_x P(x) \mathcal{E}_x. \quad (64.131)$$

Hence

$$\mathcal{E} = \sum_g \sum_x \delta(g, g(x)) P(x) \mathcal{E}_x \quad (64.132)$$

$$= \sum_g \sum_{x \in g^{-1}(g)} P(x) \mathcal{E}_x \quad (64.133)$$

Let

$$g_{d|x} = g_d(x) = P(d|x) \quad (64.134)$$

for $d \in \{0, 1\}$. Note that $g_{0|x} + g_{1|x} = 1$. We will refer to $g_d(x)$ for $d \in \{0, 1\}$ as a two **propensities** and to $g_1(x)$ as the **propensity score**.

Define

$$\delta_{y|x} = P(y|d = 1, x) - P(y|d = 0, x) \quad (64.135)$$

Note that

$$\delta_{y|x} = \frac{P(y, d = 1|x)}{g_{1|x}} - \frac{P(y, d = 0|x)}{g_{0|x}} \quad (64.136)$$

$$= \frac{P(y, d = 1|x)g_{0|x} - P(y, d = 0|x)g_{1|x}}{g_{0|x}g_{1|x}} \quad (64.137)$$

$$= \frac{P(y, d = 1|x) - P(y|x)g(x)}{g(x)(1 - g(x))} \quad (64.138)$$

Dividing each term in a sum over $d \in \{0, 1\}$ by $g_d(x)$ is often called **inverse probability (or propensity) weighting** (IPW). $g_d(x) = P(\underline{d} = d|x)$ is the likelihood of strata x , so dividing each term by this likelihood increases the contribution to the sum by less likely strata and decreases the contribution by more likely strata.

Note that the backdoor adjustment formula can be expressed as an IPW:

$$P(y|\mathcal{D}\underline{d} = d) = \sum_x P(y|d, x) P(x) \quad (64.139)$$

$$= \sum_x \frac{P(y, d, x)}{P(d|x)} \quad (64.140)$$

Note that $ATE = ACE$ can be expressed in terms of propensities:

$$ACE = \sum_y y [P(y|\mathcal{D}\underline{d} = 1) - P(y|\mathcal{D}\underline{d} = 0)] \quad (64.141)$$

$$= \sum_x P(x) \sum_y y [P(y|d = 1, x) - P(y|d = 0, x)] \quad (64.142)$$

$$= \sum_x P(x) \underbrace{\sum_y y \delta_{y|x}}_{ACE_x} \quad (64.143)$$

Note that

$$P(y(c)|d) = \sum_x P(y(c)|d, x) P(x|d) \quad (64.144)$$

$$= \sum_x P(y(c)|c, x) P(x|d) \text{ (by CIA)} \quad (64.145)$$

$$= \sum_x P(y|c, x) P(x|d) \text{ (by SUTVA)} \quad (64.146)$$

It's instructive to express all the other treatment effects besides ATE in terms of propensities:

$$ATT = \sum_y y [P(\underline{y}(1) = y|d = 1) - P(\underline{y}(0) = y|d = 1)] \quad (64.147)$$

$$= \sum_x P(x|d = 1) \sum_y y [P(y|\underline{d} = 1, x) - P(y|\underline{d} = 0, x)] \text{ (by Eq.(64.146))} \quad (64.148)$$

$$= \sum_x P(x|d = 1) \sum_y y \delta_{y|x} \quad (64.149)$$

$$= \sum_x P(x) \underbrace{\frac{1}{P(\underline{d} = 1)} \sum_y y g_{1|x} \delta_{y|x}}_{ATT_x} \quad (64.150)$$

$$ATU = \sum_x P(x) \underbrace{\frac{1}{P(\underline{d} = 0)} \sum_y y g_{0|x} \delta_{y|x}}_{ATU_x} \quad (64.151)$$

$$SB = \sum_y y [P(\underline{y}(0) = y | d = 1) - P(\underline{y}(0) = y | d = 0)] \quad (64.152)$$

$$= \sum_y y \sum_x P(y | \underline{d} = 0, x) [P(x | d = 1) - P(x | d = 0)] \text{ (by Eq.(64.146))} \quad (64.153)$$

$$= \sum_x P(x) \underbrace{\sum_y y \frac{P(y, \underline{d} = 0 | x)}{g_{0|x}} \left[\frac{g_{1|x}}{P(\underline{d} = 1)} - \frac{g_{0|x}}{P(\underline{d} = 0)} \right]}_{SB_x} \quad (64.154)$$

$$SDO = \sum_y y [P(\underline{y}(1) = y | d = 1) - P(\underline{y}(0) = y | d = 0)] \quad (64.155)$$

$$= \sum_x P(x) \underbrace{\sum_y y \left[\frac{P(y, d = 1 | x)}{P(\underline{d} = 1)} - \frac{P(y, d = 0 | x)}{P(\underline{d} = 0)} \right]}_{SDO_x} \text{ (by SUTVA)} \quad (64.156)$$

64.16 Propensity based estimates of treatment effects

In the strata-matching section 64.14, we gave an estimate of ACE_x . In strata-matching, one fills in the missing counterfactual values via the map $m()$. This is justified because, by CIA (weak-d limit), the counterfactual distributions are assumed to equal the factual distributions (see Fig. 64.6). In this section, we give estimates of treatment effects that are based on the formulae derived in Section 64.15. The estimates given in this section do no require the map $m()$, because the identification of counterfactual distributions with factual ones was used to derive the formulae of Section 64.15.

Eq.(64.143) suggests the estimate

$$\widehat{ACE}_x = \frac{1}{N_x} \sum_{\sigma \in A_x} y^\sigma \left[\frac{d^\sigma - g(x)}{g(x)(1 - g(x))} \right] \quad (64.157)$$

$$= \frac{1}{N_x} \sum_{\sigma \in A_x} \left[\frac{d^\sigma y^\sigma}{g_{1|x}} - \frac{(1 - d^\sigma) y^\sigma}{g_{0|x}} \right] \quad (64.158)$$

This estimate is unbiased (because it doesn't have a source of noise like the strata-matching estimates do), but it is still possible to improve it. We next define another ACE estimate called **Doubly Robust Estimate** (DRE) that is also unbiased and has smaller variance.

Let $\widehat{\mathcal{Y}}|_{d,x}$ be an estimate of $\mathcal{Y}|_{d,x} = E_{y|d,x}[\underline{y}]$ that is obtained, for example, via Linear Regression. Define the DRE

$$\widehat{ACE}_x^{DRE} = \widehat{\mathcal{Y}}_{1|x}^{DRE} - \widehat{\mathcal{Y}}_{0|x}^{DRE} \quad (64.159)$$

where

$$\widehat{\mathcal{Y}}_{1|x}^{DRE} = \frac{1}{N_x} \sum_{\sigma \in A_x} \left[\frac{d^\sigma(y^\sigma - \widehat{\mathcal{Y}}_{|1,x})}{g_{1|x}} + \widehat{\mathcal{Y}}_{|1,x} \right] \quad (64.160)$$

and

$$\widehat{\mathcal{Y}}_{0|x}^{DRE} = \frac{1}{N_x} \sum_{\sigma \in A_x} \left[\frac{(1-d^\sigma)(y^\sigma - \widehat{\mathcal{Y}}_{|0,x})}{g_{0|x}} + \widehat{\mathcal{Y}}_{|0,x} \right]. \quad (64.161)$$

The DRE $\widehat{\mathcal{Y}}_{1|x}^{DRE}$ requires first estimating 2 preparatory quantities, $\widehat{\mathcal{Y}}_{|1,x}$ and $g_{1|x}$. It's called doubly robust because it remains unbiased even if one of the estimates of those 2 preparatory quantities is wrong, but not if both are wrong. Let's check this.

- Suppose the propensity $g_{1|x}$ is slightly wrong. So what because

$$E_{\sigma|1,x} \left[\frac{d^\sigma(y^\sigma - \widehat{\mathcal{Y}}_{|1,x})}{g_{1|x}} \right] = 0 \quad (64.162)$$

- Suppose $\widehat{\mathcal{Y}}_{|1,x}$ is slightly wrong. So what because

$$E_{\sigma|1,x} \left[\frac{-d^\sigma \widehat{\mathcal{Y}}_{|1,x} + g_{1|x} \widehat{\mathcal{Y}}_{|1,x}}{g_{1|x}} \right] = 0 \quad (64.163)$$

A similar argument can be used to show that $\widehat{\mathcal{Y}}_{0|x}^{DRE}$ is doubly robust too.

64.17 Positivity

Positivity or non-zero overlap is defined as the requirement that for all layers x ,

$$0 < \underbrace{P(\underline{d}^\sigma = 1 | \underline{x}^\sigma = x)}_{g_{1|x}} < 1 \quad (64.164)$$

or, equivalently,

$$\underbrace{P(\underline{d}^\sigma = 1 | \underline{x}^\sigma = x)}_{g_{1|x}} > 0 \quad \text{and} \quad \underbrace{P(\underline{d}^\sigma = 0 | \underline{x}^\sigma = x)}_{g_{0|x}} > 0. \quad (64.165)$$

In other words, for each layer x , there is a non-zero probability of being both treated and untreated. See Fig.64.15 for a pictorial representation of positivity.

If positivity is violated for some layer x , then

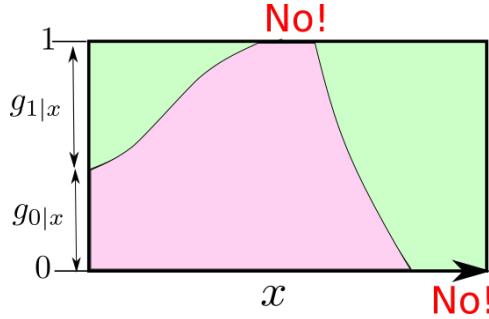


Figure 64.15: Pictorial representation of positivity. $g_{0|x} + g_{1|x} = 1$. $g_{0|x} = 0$ and $g_{1|x} = 0$ are forbidden.

- the propensity based estimate Eq.(64.158) for ACE_x (which equals ATE_x) is undefined.
- all strata-matching estimates of \mathcal{E}_x that use the matching function $m()$ are undefined because that function is undefined if $A_{0,x} = \emptyset$ or $A_{1,x} = \emptyset$.

If a quantity (estimand) can be estimated, it is said to be **do-identifiable** (i.e., expressible without `do()` operators). If positivity is violated, then $ACE = ATE$ is not identifiable.

When $P(d|x)$ becomes 0 or 1 for some x , the arrow $\underline{x} \rightarrow \underline{d}$ becomes deterministic for that x . This situation is the very antithesis of RCTs, wherein the influence exerted by \underline{x}^σ on \underline{d}^σ is uniformly random and therefore ignorable. Hence, it is perhaps not too surprising that a violation of positivity makes $ACE = ATE$ not identifiable.

64.18 Multi-time PO bnets (Panel Data)

In this section, we will discuss Multi-time PO bnets (MT-PO).

A **time-series** is a function $f : D \rightarrow \mathbb{R}$ whose domain D is a discrete set of times. A time-series usually describes a single unit σ (i.e., an individual) in a population.

An **observational study (or analysis or model)** can be cross-sectional or longitudinal. A **cross-sectional study** collects and analyzes a **cross-sectional dataset**; i.e., a dataset for a population at a single time. A **longitudinal study or panel study** collects and analyzes a **longitudinal dataset**; i.e., a dataset for a population at multiple times. Thus, a longitudinal study consists of one or more time-series.

Let $\mathcal{T} = \{t_0, t_1, \dots, t_{nt-1}\}$. For any time-series $a_t : \mathcal{T} \rightarrow \mathbb{R}$, define

$$E_t a_t = \frac{1}{nt} \sum_{t \in \mathcal{T}} a_t \quad (64.166)$$

$$\Delta_t a_t = a_t - E_t a_t \quad (64.167)$$

$$\langle a_t, b_t \rangle_t = E_t \Delta_t a_t \Delta_t b_t \quad (64.168)$$

Consider a quantity a_t^σ that is a function of the time t and of the particular unit σ in a population. a_t^σ is said to be a **t-constant effect** if it is t -independent. a_t^σ is said to be a **homogeneous effect** (antonym: **heterogeneous effect**) if it is σ -independent. Henceforth, we will avoid using the word “effect” for these because that word has already been used for “treatment effect” in PO theory. Instead, we will use the word “quantity”.

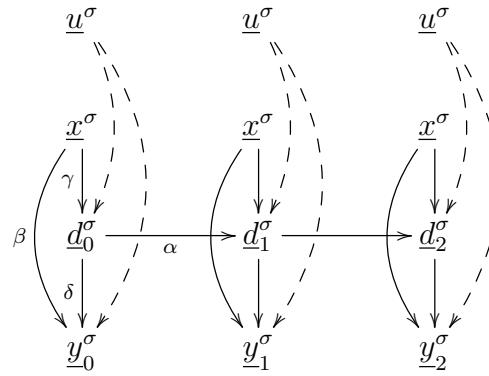


Figure 64.16: Example of multi-time PO bnet with t-constant quantities x^σ, u^σ . The 3 nodes x^σ should be identified as a single node. Likewise, the 3 nodes u^σ should be identified as a single node.

Fig.64.16 gives an example of a multi-time PO bnet (MT-PO). Note that in this example, x^σ and u^σ are t-constant quantities. u^σ is an unobserved confounder and x^σ is an observed confounder. For convenience and simplicity, we will assume linear deterministic TPMs. The TPMs, printed in blue, for the bnet Fig.64.16, are as follows:

$$P(x^\sigma) = P_{\underline{x}}(x^\sigma) \quad (64.169)$$

$$P(u^\sigma) = P_{\underline{u}}(u^\sigma) \quad (64.170)$$

$$P(y_t^\sigma | d_t^\sigma, x^\sigma, u^\sigma) = \mathbb{1}(y_t^\sigma = \delta d_t^\sigma + \beta x^\sigma + u^\sigma) \quad (64.171)$$

$$P(d_{t+1}^\sigma | d_t^\sigma, x^\sigma, u^\sigma) = \mathbb{1}(d_{t+1}^\sigma = \alpha d_t^\sigma + \gamma x^\sigma + u^\sigma) \quad (64.172)$$

Taking time averages of the treatment decision and treatment outcome, we get

$$E_t \underline{y}_t^\sigma = \delta E_t \underline{d}_t^\sigma + \beta \underline{x}^\sigma + \underline{u}^\sigma, \quad (64.173)$$

$$E_t \underline{d}_{t+1}^\sigma = \alpha E_t \underline{d}_t^\sigma + \gamma \underline{x}^\sigma + \underline{u}^\sigma. \quad (64.174)$$

Subtracting the time averages from the quantities being averaged, we get

$$\Delta_t \underline{y}_t^\sigma = \delta \Delta_t \underline{d}_t^\sigma, \quad (64.175)$$

$$\Delta_t \underline{d}_{t+1}^\sigma = \alpha \Delta_t \underline{d}_t^\sigma. \quad (64.176)$$

This allows us to find estimates for δ and α :

$$E_\sigma \langle \underline{y}_t^\sigma, \underline{y}_t^\sigma \rangle_t = \delta E_\sigma \langle \underline{y}_t^\sigma, \underline{d}_t^\sigma \rangle_t \quad (64.177)$$

$$\delta = \frac{E_\sigma \langle \underline{y}_t^\sigma, \underline{y}_t^\sigma \rangle_t}{E_\sigma \langle \underline{y}_t^\sigma, \underline{d}_t^\sigma \rangle_t} \quad (64.178)$$

$$E_\sigma \langle \underline{d}_{t+1}^\sigma, \underline{d}_{t+1}^\sigma \rangle_t = \alpha E_\sigma \langle \underline{d}_{t+1}^\sigma, \underline{d}_t^\sigma \rangle_t \quad (64.179)$$

$$\alpha = \frac{E_\sigma \langle \underline{d}_{t+1}^\sigma, \underline{d}_{t+1}^\sigma \rangle_t}{E_\sigma \langle \underline{d}_{t+1}^\sigma, \underline{d}_t^\sigma \rangle_t} \quad (64.180)$$

As shown in Fig.64.17, subtraction of time averages from each node removes the confounder nodes from the bnet of Fig.64.16 (However, this assumes that the confounders are t-constant and that the TPMs are linear deterministic, two very strong assumptions).

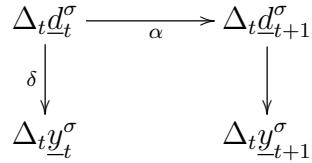


Figure 64.17: time-average-subtracted (TAS) bnet for the bnet of Fig.64.16.

Chapter 65

Program evaluation and review technique (PERT)

This chapter is based on Refs.[62] and [132].

PERT diagrams are used for scheduling a project consisting of a series of interdependent activities and estimating how long it will take to finish the project. PERT diagrams were invented by the NAVY in 1958 to manage a submarine project. Nowadays they are taught in many business and management courses.

A **PERT diagram** is a Directed Acyclic Graph (DAG) with the following properties. (See Fig.65.2 for an example of a PERT diagram). The nodes \underline{E}_i for $i = 1, 2, \dots, ne$ of a PERT diagram are called **events**. The edges $i \rightarrow j$ of a PERT diagram are called **activities**. An event represents the starting (kickoff) date of one or more activities. A PERT diagram has a single root node ($i = 1$, start event) and a single leaf node ($i = ne$, end event).

The PERT diagram user must initially provide a **Duration Times (DT) table** which gives $(DO_{i \rightarrow j}, DP_{i \rightarrow j}, DM_{i \rightarrow j})$ for each activity $i \rightarrow j$, where

$DO_{i \rightarrow j}$ = optimistic duration time of activity $i \rightarrow j$

$DP_{i \rightarrow j}$ = pessimistic duration time of activity $i \rightarrow j$

$DM_{i \rightarrow j}$ = median duration time of activity $i \rightarrow j$

From the DT table, one calculates:

Duration time of activity $i \rightarrow j$

$$D_{i \rightarrow j} = \frac{1}{6}(DO_{i \rightarrow j} + DP_{i \rightarrow j} + 4DM_{i \rightarrow j}) \quad (65.1)$$

Duration Variance of activity $i \rightarrow j$

$$V_{i \rightarrow j} = \left(\frac{DO_{i \rightarrow j} - DP_{i \rightarrow j}}{DM_{i \rightarrow j}} \right)^2 \quad (65.2)$$

Often, it is convenient to define “dummy” edges with $D_{i \rightarrow j} = 0$. That is perfectly fine.

Define:

TES_i = Earliest start time for event i
 TLS_i = Latest start time for event i
 $slack_i = TLS_i - TES_i$ = slack for event i
 $TEF_{i \rightarrow j} = TES_i + D_{i \rightarrow j}$ = Earliest finish time for activity $i \rightarrow j$.
 $TLF_{i \rightarrow j} = TLS_j - D_{i \rightarrow j}$ = Latest finish time for activity $i \rightarrow j$. See footnote below.¹

A **critical path** is a directed path (i.e., a chain of connected arrows, all pointing in the same direction) going from the start to the end node, such that slack equals zero at every node visited. In a DAG, the neighbors of a node is the union of its parent and children nodes. A critical path must also have all other nodes as neighbors; i.e, the union of the neighbors of every node in the path plus the nodes in the path itself, equals all nodes in the graph.

GOAL of PERT analysis: The main goal of PERT analysis is to find, based on the data of the DT table, the interval $[TES_i, TLS_i]$ giving a lower and an upper bound to the starting time of each node i . Another goal is to find a critical path for the PERT diagram (which represents an entire project). By adding the $D_{i \rightarrow j}$ of each edge of the critical path, one can get the mean value of the total duration of the entire project, and by adding the variances of each edge along the critical path, one can get an estimate of the total variance of the total duration. Knowing the mean and variance of the total duration and assuming a Normal distribution, one can predict the probability that the actual duration will deviate by a certain amount from its mean.

To calculate the interval $[TES_i, TLS_i]$, one follows the following two steps.

1. Assume $TES_1 = 0$ and solve

$$TES_i = \max_{a \in pa(i)} (\underbrace{TES_a + D_{a \rightarrow i}}_{TEF_{a \rightarrow i}}) \quad (65.3)$$

for $i \in [2, ne]$. This recursive equation is solved by what is called “forward propagation”, wherein one moves up the list of nodes i in order of increasing i starting at $i = 1$ with $TES_1 = 0$.

2. Assume $TLS_{ne} = TES_{ne}$ and solve

$$TLS_i = \min_{b \in ch(i)} (\underbrace{TLS_b - D_{i \rightarrow b}}_{TLF_{i \rightarrow b}}) \quad (65.4)$$

¹In the popular educational literature, the edge variables $TEF_{i \rightarrow j}$ and $TLF_{i \rightarrow j}$ are sometimes associated with the nodes, but they are clearly edge variables. This makes things confusing. The reason this is done is that some software draws PERT diagrams as trees whereas other software draws them as DAGs. For trees, storing $TEF_{i \rightarrow j}$ and $TLF_{i \rightarrow j}$ in a node makes some sense but not for DAGs. You will notice that giving specific names to the variables $TEF_{i \rightarrow j}$ and $TLF_{i \rightarrow j}$ is unnecessary. It is possible to delete all mention of their names from this chapter without losing any details. I only declare their names in this chapter so as tell the reader what they are in case he/she hears them mentioned and wonders what they are equal to in our notation.

for $i \in [1, ne - 1]$. This recursive equation is solved by what is called “backward propagation”, wherein one moves down the list of nodes i in order of decreasing i starting at $i = ne$ with $TLS_{ne} = TES_{ne}$. TES_{ne} is known from step 1.

Eqs.(65.3) and (65.4) are illustrated in Fig.65.1.

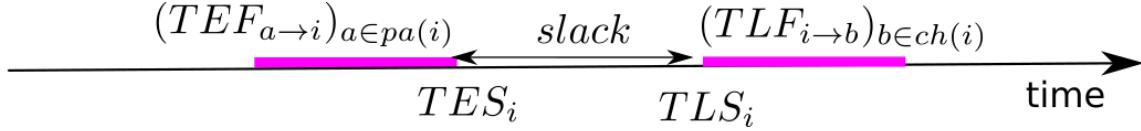


Figure 65.1: TES_i defined from info received from parents of i and TLS_i defined from info received from children of i .

65.1 Example

To illustrate PERT analysis, we end with an example. We present the example in the form of an exercise question and then provide the answer. This example comes from Ref.[62], except for part (e) about bnets, which is our own.

Question: For the PERT diagram of Fig.65.2, calculate the following:

- (a) Interval $[TES_i, TLS_i]$ for all i .
- (b) A critical path for this PERT diagram.
- (c) The mean and variance of the total duration of the critical path.
- (d) The probability that the total duration will be 225 days or less.
- (e) A bnet interpretation of this problem.

Answer to (a) $[TES_i, TLS_i]$ are given by Fig.65.3.

Answer to (b) The critical path is given in red in Fig.65.3. Note that this path does indeed have zero slack at each node it visits and the union of its neighborhood and the path itself encompasses all nodes.

Answer to (c) The mean and variance of the total duration are calculated in Table 65.1.

Answer to (d)

$$P(\underline{x} < 225) = P\left[\frac{\underline{x} - \mu}{\sigma} \leq \frac{225 - 220}{\sqrt{7.73}}\right] \quad (65.5)$$

$$= P[\underline{z} \leq 1.80] \quad (65.6)$$

$$= 0.9641 \quad (65.7)$$

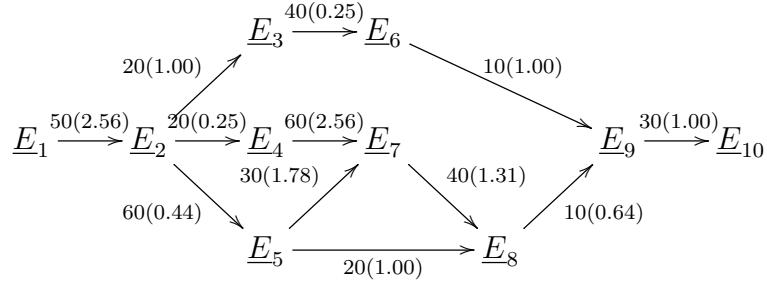


Figure 65.2: Example of a PERT diagram. The numbers attached to the arrows are the duration times $D_{i \rightarrow j}$ in days followed by, enclosed in parentheses, the variance $V_{i \rightarrow j}$ of that duration. The info given in this PERT diagram was derived from a DT table in Ref.[62]. The info in this PERT diagram is sufficient for calculating TES_i and TLS_i for each node i . The results of that calculation are given in Fig.65.3.

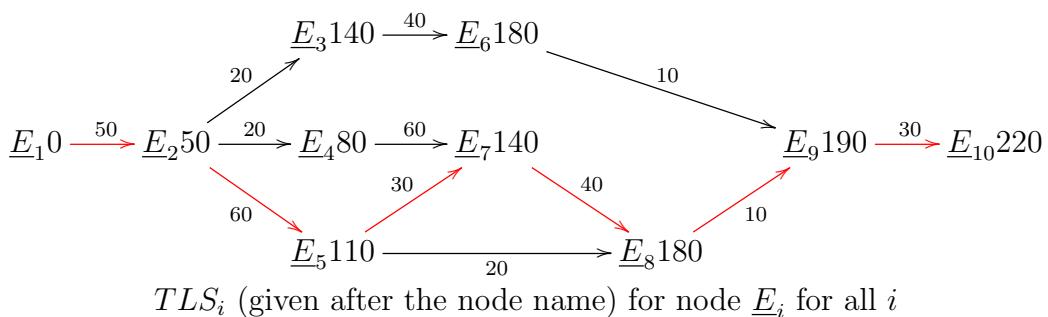
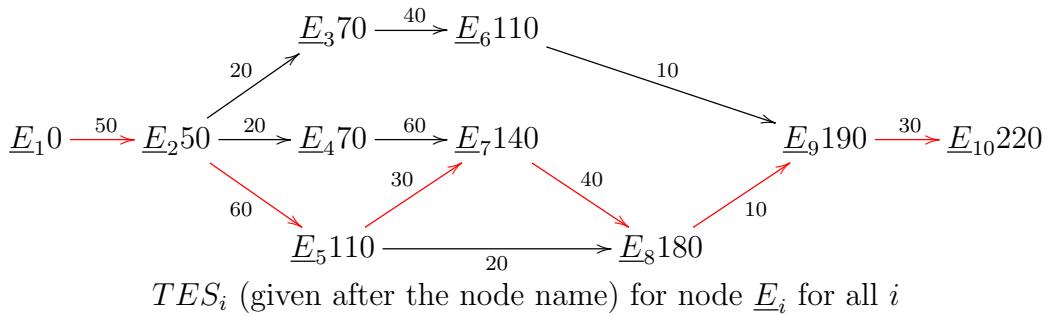


Figure 65.3: Results of calculating TES_i for all i via a forward pass, followed by calculating TLS_i for all i via a backward pass. Critical path indicated in red.

Answer to (e) Define 2 bnets.

1. The first PERT bnet is for calculating TES_i for all i and is given by Fig.65.4.

edge $i \rightarrow j$	duration $D_{i \rightarrow j}$	variance $V_{i \rightarrow j}$
A (1 → 2)	50	2.56
D (2 → 5)	60	0.44
G (5 → 7)	30	1.78
J (7 → 8)	40	1.31
K (8 → 9)	10	0.64
L (9 → 10)	30	1.00
Total	220	7.73

Table 65.1: Calculation of mean and variance of total duration along critical path.

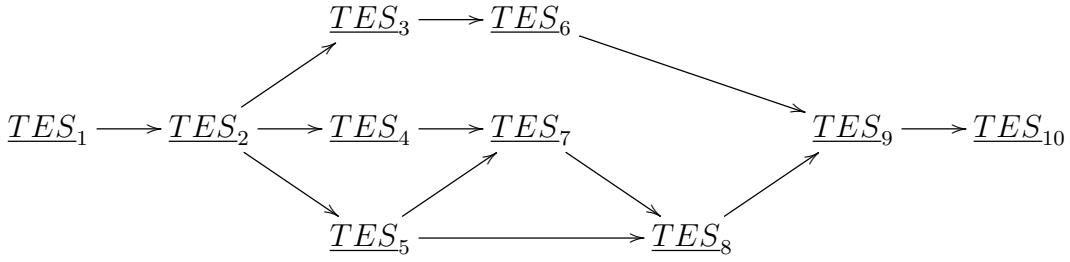


Figure 65.4: bnet for TES_i calculation.

The TPMs, printed in blue, for the bnet Fig.65.4, are as follows (this equation is to be evaluated recursively by a forward pass through the bnet):

$$P(TES_i | (TES_a)_{a \in pa(i)}) = \delta(TES_i, \max_{a \in pa(i)} (TES_a + D_{a \rightarrow i})) \quad (65.8)$$

2. The second PERT bnet is for calculating TLS_i for all i and is given by Fig.65.5. Note that the directions of all the arrows in the PERT diagram Fig.65.2 have been reversed so Fig.65.5 is a time reversed graph.

The TPMs, printed in blue, for the bnet Fig.65.5, are as follows (this equation is to be evaluate recursively by a backward pass through the bnet):

$$P(TLS_i | (TLS_b)_{b \in pa(i)}) = \delta(TLS_i, \min_{b \in pa(i)} (TLS_b - D_{b \rightarrow i}^T)) , \quad (65.9)$$

where $D_{i \rightarrow j}^T = D_{j \rightarrow i}$.

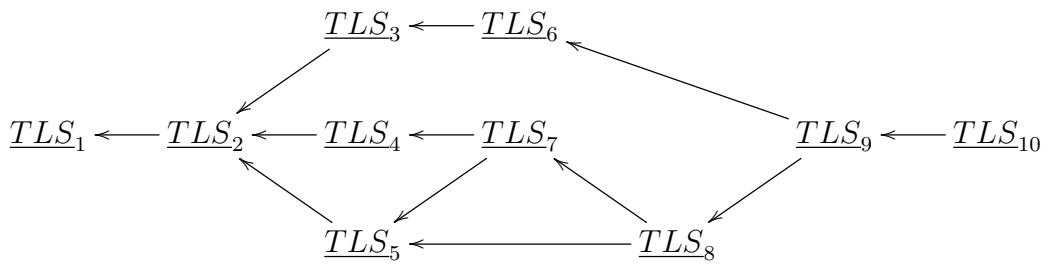


Figure 65.5: bnet for TLS_i calculation.

Chapter 66

Random Forest and Bagging

This chapter is based on Refs.[86] and [133].

Chapter 14 defines decision trees (dtrees) and explains how to construct them. A Random Forest (RF) is an ensemble of dtrees. The RF algorithm is a method of, given a dataset, constructing a RF and averaging over the classifier functions of the RF to produce an ensemble classifier.

The RF algo uses the method of Bootstrap Aggregating (a.k.a. Bagging), which is discussed in detail below. Bagging is a method of constructing an ensemble of datasets (called **bootstrap datasets or bags**) that are fairly uncorrelated. Each of these bags is used to train a **bag-classifier**. The bag-classifiers are averaged over to produce an **ensemble classifier, or e-classifier** for short. Bagging can be used to train any type of bag-classifier, but it was invented with dtrees in mind, and is still most commonly used to train dtrees.

Boosting (see Chapter 1 on AdaBoost and Chapter 91 on XGBoost) is another method, besides Bagging, of constructing a classifier function from an ensemble of classifier functions. These two methods are most commonly applied to dtrees: Boosting for an ensemble of small dtrees, and Bagging for a random forest (which is an ensemble of dtrees that are usually much more complicated than small dtrees).

66.1 Bagging (with fully-featured bags)

In this section we discuss the bagging algorithm. As already mentioned, bagging is usually used to train dtrees. In this section, we explain bagging in general, not just for dtrees.

Let $L = [0, 1, 2, \dots, nsam - 1]$ be a list of individuals (samples) in a population. In this chapter, we will use the notation $A^\sigma = A[\sigma]$ and $\vec{A} = [A^\sigma : \sigma \in L]$ for a list (vector, 1-D array) indexed by L . We will refer to $DS = (\vec{x}, \vec{y})$ where $x^\sigma \in S_x$, $y^\sigma \in S_y$, as a dataset. If L_j is a list (possibly with duplicate items) such that $set(L_j) \subset set(L)$, then define $DS_j = (\vec{x}, \vec{y})_{L_j} = ((x^\sigma)_{\sigma \in L_j}, (y^\sigma)_{\sigma \in L_j})$. We will refer to DS_j as the **restriction of (\vec{x}, \vec{y}) to L_j** .

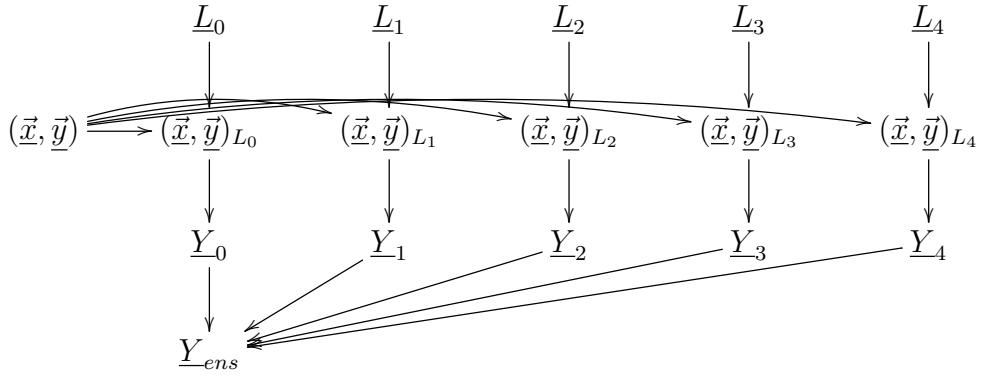


Figure 66.1: Bnet for Random Forest (RF) with 5 bags.

We will refer to a function $Y : S_x \rightarrow S_c$ as a classifier. It maps vector of vector of features $x \in S_x$ to a class $c \in S_c$. Below, Y_b for all b and Y_{ens} are classifiers.

Fig.66.1 is a bnet that encapsulates the RF algo. The TPMs, printed in blue, for the bnet Ref.66.1, are as follows.

Let $b \in \{0, 1, 2, \dots, nbags - 1\} = B$ and $\sigma \in L$. Let $L_b^\sigma \in L$ and

$$P(L_b^\sigma) = 1/nsam \quad (66.1)$$

In other words, each item in list L_b is chosen from the items of list L , uniformly at random with replacements. $|L_b| = |L|$ (same size as original). L_b can have duplicate items and be missing items from L .

$$P((\vec{x}, \vec{y})_{L_b} | (\vec{x}, \vec{y}), L_b) = \mathbb{1}(\text{ } (\vec{x}, \vec{y})_{L_b} = \text{restriction of } (\vec{x}, \vec{y}) \text{ to } L_b \text{ }) \quad (66.2)$$

We will refer to (\vec{x}, \vec{y}) as the **original dataset** and to the $(\vec{x}, \vec{y})_{L_b}$ for $b \in B$ as the **bootstrap datasets** or **bags**.

$$P(Y_b | (\vec{x}, \vec{y})_{L_b}) = \mathbb{1}(\text{ } Y_b(\cdot) = \text{classifier trained on dataset } (\vec{x}, \vec{y})_{L_b}. \text{ }) \quad (66.3)$$

$$P(Y_{ens} | (Y_b)_{b \in B}) = \prod_{\sigma} \mathbb{1}(\text{ } Y_{ens}(x^\sigma) = \text{majority}(\{Y_b(x^\sigma) : b \in B\}) \text{ }) \quad (66.4)$$

The `majority()` function can be replaced by an average $\frac{1}{nbags} \sum_b$ in case the set of classes S_c equals \mathbb{R} rather than a finite set. We will refer to Y_{ens} as the **ensemble classifier (e-classifier)** and to the Y_b as the **bag-classifiers**.

Define (these definitions are illustrated in Fig.66.2)

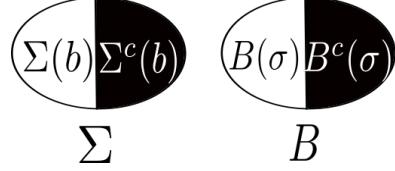


Figure 66.2: $\Sigma(b)$ and $\Sigma^c(b)$ are disjoint sets whose union is Σ . $B(\sigma)$ and $B^c(\sigma)$ are disjoint sets whose union is B .

$$\Sigma = \text{set}(L), \quad \Sigma(b) = \text{set}(L_b), \quad \Sigma^c(b) = \Sigma - \Sigma(b) \quad (66.5)$$

and

$$B(\sigma) = \{b \in B : \sigma \in \Sigma(b)\}, \quad B^c(\sigma) = B - B(\sigma) \quad (66.6)$$

$\Sigma(b)$ is the set of **in-the-b-bag individuals** and $\Sigma^c(b)$ is the set of **out-of-the-b-bag (OOB) individuals**. $B(\sigma)$ is the set of bags that contain individual σ , and $B^c(\sigma)$ is the set of bags that don't.

The **OOB error** is defined as

$$err = \sum_{\sigma \in L} \mathbb{1}(B^c(\sigma) \neq \emptyset) \mathbb{1}(y^\sigma \neq \text{majority}([Y_b(x^\sigma) : b \in B^c(\sigma)])). \quad (66.7)$$

Empirical results supposedly show that OOB error is comparable in accuracy to the error calculated by doing cross-validation (CV) (see Chapter 12), although CV error is considered more dependable.

66.2 Bagging (with randomly-shortened bags)

Suppose the feature vector x^σ in the dataset $DS = (\vec{x}, \vec{y})$ has nf components; i.e., $x^\sigma = (x_0^\sigma, x_1^\sigma, \dots, x_{nf-1}^\sigma) \in S_{\underline{x}_0} \times S_{\underline{x}_1} \times \dots \times S_{\underline{x}_{nf-1}} = S_{\underline{x}}$.

For each bag DS_b , one chooses at random $nf' = \sqrt{nf}$ out of the nf features, and discards the remaining features from DS_b , thus producing a new, **randomly-shortened-bag (rs-bag)** DS'_b . Each rs-bag is then used to train a bag-classifier, usually a dtree, using the methods for dtree SL described in Chapter 14. Using rs-bags is called the **random subspace method**. The reason for using rs-bags is that they further decorrelate the set of bags used to train bag-classifiers.

Chapter 67

Recurrent Neural Networks

This chapter is mostly based on Ref.[38].

This chapter assumes you are familiar with the material and notation of Chapter 58 on plain Neural Nets.

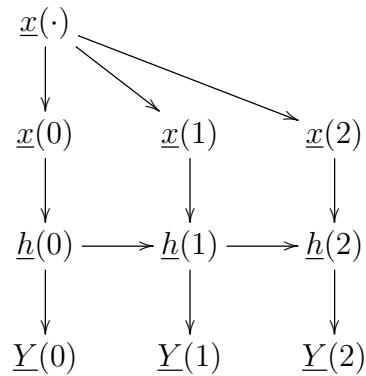


Figure 67.1: Simple example of RNN with $T = 3$

Suppose

T is a positive integer.

$t = 0, 1, \dots, T - 1$,

$\underline{x}_i(t) \in \mathbb{R}$ for $i = 0, 1, \dots, nx - 1$,

$\underline{h}_i(t) \in \mathbb{R}$ for $i = 0, 1, \dots, nh - 1$,

$\underline{Y}_i(t) \in \mathbb{R}$ for $i = 0, 1, \dots, ny - 1$,

$W^{h|x} \in \mathbb{R}^{nh \times nx}$,

$W^{h|h} \in \mathbb{R}^{nh \times nh}$,

$W^{y|h} \in \mathbb{R}^{ny \times nh}$,

$b^y \in \mathbb{R}^{ny}$,

$b^h \in \mathbb{R}^{nh}$.

Henceforth, $x(\cdot)$ will mean the array of $x(t)$ for all t .

The simplest kind of recurrent neural network (RNN) has the bnet Fig.67.1 with arbitrary T . The node TPMs, printed in blue, for this bnet, are as follows.

$$P(x(\cdot)) = \text{given} \quad (67.1)$$

$$P(x(t)) = \delta(x(t), [x(\cdot)]_t) \quad (67.2)$$

$$P(h(t) | h(t-1), x(t)) = \delta(h(t), \mathcal{A}(W^{h|x}x(t) + W^{h|h}h(t-1) + b^h)) , \quad (67.3)$$

where $h(-1) = 0$.

$$P(Y(t) | h(t)) = \delta(Y(t), \mathcal{A}(W^{y|h}h(t) + b^y)) \quad (67.4)$$

Define

$$W^h = [W^{h|x}, W^{h|h}, b^h] , \quad (67.5)$$

and

$$W^y = [W^{y|h}, b^y] . \quad (67.6)$$

The bnet of Fig.67.1 can be used for classification once its parameters W^h and W^y have been optimized. To optimize those parameters via gradient descent, one can use the bnet of Fig.67.2.

Let $\sigma = 0, 1, \dots, nsam(\vec{x}) - 1$ be the labels for a minibatch of samples. Below, we will write $A^\sigma = A[\sigma]$ for the σ component of any vector \vec{A} . The TPMs, printed in blue, for bnet Fig.67.2, are as follows.

$$P(x(\cdot)^\sigma) = \text{given} \quad (67.7)$$

$$P(x(t)^\sigma) = \delta(x(t)^\sigma, [x(\cdot)]_t^\sigma) \quad (67.8)$$

$$P(h(t)^\sigma | h(t-1)^\sigma, x(t)^\sigma) = \delta(h(t)^\sigma, \mathcal{A}(W^{h|x}x(t)^\sigma + W^{h|h}h(t-1)^\sigma + b^h)) \quad (67.9)$$

$$P(Y(t)^\sigma | h(t-1)^\sigma) = \delta(Y(t)^\sigma, \mathcal{A}(W^{y|h}h(t-1)^\sigma + b^y)) \quad (67.10)$$

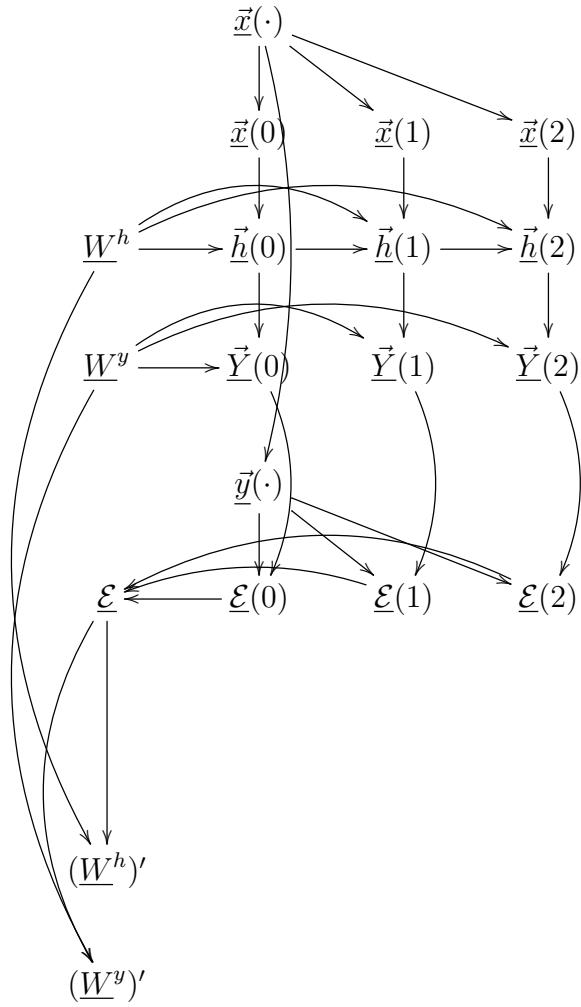


Figure 67.2: RNN bnet used to optimize parameters W^h and W^y of RNN bnet Fig.67.1.

$$P(y(\cdot)^\sigma | x(\cdot)^\sigma) = \text{given} \quad (67.11)$$

$$P(\mathcal{E}(t) | \vec{y}(t), \vec{Y}(t)) = \frac{1}{nsam(\vec{x})} \sum_{\sigma} d(y(t)^\sigma, Y(t)^\sigma) , \quad (67.12)$$

where

$$d(y, Y) = |y - Y|^2 . \quad (67.13)$$

If $y, Y \in [0, 1]$, one can use this instead

$$d(y, Y) = XE(y \rightarrow Y) = -y \ln Y - (1 - y) \ln(1 - Y). \quad (67.14)$$

$$P(\mathcal{E} | \mathcal{E}(\cdot)) = \delta(\mathcal{E}, \sum_t \mathcal{E}(t)) \quad (67.15)$$

For $a = h, y$,

$$P(W^a) = \text{given}. \quad (67.16)$$

The first time it is used, W^a is arbitrary. Afterwards, it is determined by previous horizontal stage.

$$P((W^a)' | \mathcal{E}, W^a) = \delta((W^a)', W^a - \eta^a \partial_{W^a} \mathcal{E}). \quad (67.17)$$

$\eta^a > 0$ is the learning rate for W^a .

67.1 Language Sequence Modeling

Estimate $P(x(\cdot))$ empirically. We can use this to:

- predict the probability of a sentence,
example: Get $P(x(0), x(1), x(2))$.
- predict the most likely next word in a sentence,
example: Get $P(x(2)|x(0), x(1))$.
- generate fake sentences.
example:
Get $x(0) \sim P(x(0))$.
Next get $x(1) \sim P(x(1)|x(0))$.
Next get $x(2) \sim P(x(2)|x(0), x(1))$.

67.2 Other types of RNN

Let $\mathcal{T} = \{0, 1, \dots, T - 1\}$, and $\mathcal{T}^x, \mathcal{T}^y \subset \mathcal{T}$. Above, we assumed that $\underline{x}(t)$ and $\underline{Y}(t)$ were both defined for all $t \in \mathcal{T}$. More generally, they might be defined only for subsets of \mathcal{T} : $\underline{x}(t)$ for $t \in \mathcal{T}^x$ and $\underline{Y}(t)$ for $t \in \mathcal{T}^y$. If $|\mathcal{T}^x| = 1$ and $|\mathcal{T}^y| > 1$, we say the RNN bnet is of the **1 to many** kind. In general, can have **1 to 1**, **1 to many**, **many to 1**, **many to many** RNN bnets.

Plain RNNs can suffer from the **vanishing or exploding gradients problem**. There are various ways to mitigate this (e.g., good choice of initial W^h and

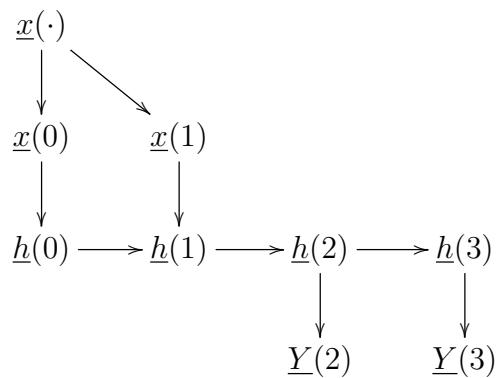


Figure 67.3: RNN bnet of the many to many kind. This one can be used for translation. $x(0)$ and $x(1)$ might denote two words of an English sentence, and $Y(2)$ and $Y(3)$ might be their Italian translation.

W^y , good choice of activation functions, regularization). Or by using GRU or LSTM (discussed below). **GRU and LSTM** were designed to mitigate the vanishing or exploding gradients problem. They are very popular in NLP (Natural Language Processing).

67.2.1 Long Short Term Memory (LSTM) unit (1997)

This section is based on Wikipedia article Ref.[120]. In this section, \odot will denote the Hadamard matrix product (elementwise product).

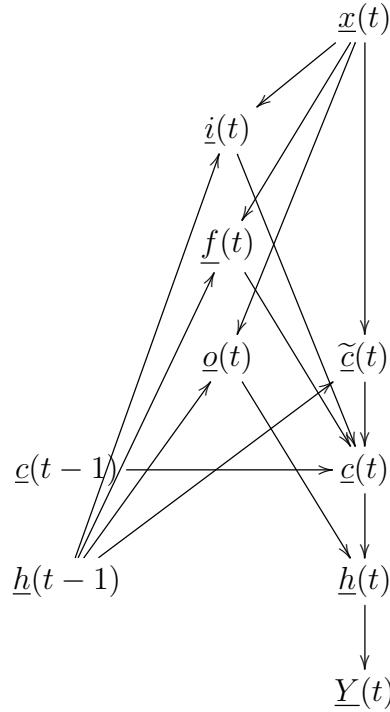


Figure 67.4: bnet for a Long Short Term Memory (LSTM) unit.

Let

$\underline{x}(t) \in \mathbb{R}^{nx}$: **input state vector** to the LSTM unit

$\underline{f}(t) \in \mathbb{R}^{nh}$: **forget activation** vector

$\underline{i}(t) \in \mathbb{R}^{nh}$: **input activation** vector

$\underline{o}(t) \in \mathbb{R}^{nh}$: **output activation** vector

$\underline{h}(t) \in \mathbb{R}^{nh}$: **hidden state vector**

$\underline{\tilde{c}}(t) \in \mathbb{R}^{nh}$: **cell activation** vector

$\underline{c}(t) \in \mathbb{R}^{nh}$: **cell state vector**

$\underline{Y}(t) \in \mathbb{R}^{ny}$: **classification** of $\underline{x}(t)$.

$W \in \mathbb{R}^{nh \times nx}$, $U \in \mathbb{R}^{nh \times nh}$ and $b \in \mathbb{R}^{nh}$: weight matrices and bias vectors, parameters learned by training.

$\mathcal{W}^{y|h} \in \mathbb{R}^{ny \times nh}$: weight matrix

Fig.67.4 is a bnet for a LSTM unit. The TPMs, printed in blue, for this bnet, are as follows.

$$P(f(t)|x(t), h(t-1)) = \mathbb{1}(\quad f(t) = \text{smoid}(W^{f|x}x(t) + U^{f|h}h(t-1) + b^f) \quad) , \quad (67.18)$$

where $h(-1) = 0$.

$$P(i(t)|x(t), h(t-1)) = \mathbb{1}(\quad i(t) = \text{smoid}(W^{i|x}x(t) + U^{i|h}h(t-1) + b^i) \quad) \quad (67.19)$$

$$P(o(t)|x(t), h(t-1)) = \mathbb{1}(\quad o(t) = \text{smoid}(W^{o|x}x(t) + U^{o|h}h(t-1) + b^o) \quad) \quad (67.20)$$

$$P(\tilde{c}(t)|x(t), h(t-1)) = \mathbb{1}(\quad \tilde{c}(t) = \tanh(W^{c|x}x(t) + U^{c|h}h(t-1) + b^c) \quad) \quad (67.21)$$

$$P(c(t)|f(t), c(t-1), i(t), \tilde{c}(t)) = \mathbb{1}(\quad c(t) = f(t) \odot c(t-1) + i(t) \odot \tilde{c}(t) \quad) \quad (67.22)$$

$$P(h(t)|o(t), c(t)) = \mathbb{1}(\quad h(t) = o(t) \odot \tanh(c(t)) \quad) \quad (67.23)$$

$$P(Y(t)|h(t)) = \mathbb{1}(\quad Y(t) = \mathcal{A}(\mathcal{W}^{y|h}h(t) + b^y) \quad) \quad (67.24)$$

67.2.2 Gated Recurrence Unit (GRU) (2014)

This section is based on Wikipedia article Ref.[103]. In this section, \odot will denote the Hadamard matrix product (elementwise product).

GRU is a more recent (17 years later) attempt at simplifying LSTM.

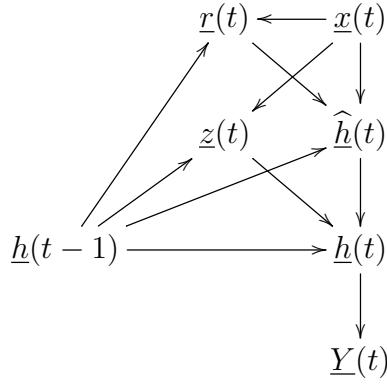


Figure 67.5: bnet for a Gated Recurrent Unit (GRU).

Let

$x(t) \in \mathbb{R}^{nx}$: **input state** vector

$h(t) \in \mathbb{R}^{nh}$: **hidden state** vector

$\hat{h}(t) \in \mathbb{R}^{nh}$: **hidden activation** vector

$z(t) \in \mathbb{R}^{nh}$: **update activation** vector

$r(t) \in \mathbb{R}^{nh}$: **reset activation** vector

$Y(t) \in \mathbb{R}^{ny}$: **classification** of $x(t)$.

$W \in \mathbb{R}^{nh \times nx}$, $U \in \mathbb{R}^{nh \times nh}$ and $b \in \mathbb{R}^{nh}$: weight matrices and bias vectors, parameters learned by training.

$\mathcal{W}^{y|h} \in \mathbb{R}^{ny \times nh}$: weight matrix

Fig.67.5 is a bnet for a GRU. The TPMs, printed in blue, for this bnet, are as follows.

$$P(z(t)|x(t), h(t-1)) = \mathbb{1}(\quad z(t) = \text{smoid}(W^{z|x}x(t) + U^{z|h}h(t-1) + b^z) \quad) , \quad (67.25)$$

where $h(-1) = 0$.

$$P(r(t)|x(t), h(t-1)) = \mathbb{1}(\quad r(t) = \text{smoid}(W^{r|x}x(t) + U^{r|h}h(t-1) + b^r) \quad) \quad (67.26)$$

$$P(\hat{h}(t)|x(t), r(t), h(t-1)) = \mathbb{1}(\quad \hat{h}(t) = \tanh(W^{h|x}x(t) + U^{h|h}(r(t) \odot h(t-1)) + b^h) \quad) \quad (67.27)$$

$$P(h(t)|z(t), h(t-1), \widehat{h}(t)) = \mathbb{1}(\quad h(t) = (1 - z(t)) \odot h(t-1) + z(t) \odot \widehat{h}(t) \quad) \quad (67.28)$$

$$P(Y(t)|h(t)) = \mathbb{1}(\quad Y(t) = \mathcal{A}(\mathcal{W}^{y|h}h(t) + b^y) \quad) \quad (67.29)$$

Chapter 68

Regression Discontinuity Design

This chapter is based on Ref.[8].

This chapter assumes that the reader has read Chapter 64 on Potential Outcomes (PO).

In Regression Discontinuity Design (RDD), one switches the treatment dose \underline{d} from 0 when $\underline{x} < \xi$ to 1 where $\underline{x} > \xi$, where \underline{x} is an observed confounder (call it the **switch confounder**) and ξ is a threshold value for \underline{x} . One measures the jump δ in the treatment outcome \underline{y} as \underline{x} passes through $\underline{x} = \xi$. Then one makes the very reasonable assumption that δ equals¹ $\mathcal{Y}_{1|x=\xi} - \mathcal{Y}_{0|x=\xi} = ATE|_{x=\xi}$ for an imaginary experiment in which the confounder \underline{x} acts as a normal confounder that doesn't switch the treatment dose \underline{d} .

For example, d^σ might be whether an individual is admitted to Harvard Univ., y^σ might be how much money the individual earns for the first 20 years after graduating from Harvard, and x^σ might be his SAT scores. We assume Harvard only admits students with an SAT score higher than ξ .

68.1 PO analysis

The TPMs, printed in blue, for the bnet G_{disc} shown in Fig.68.1, are as follows. Note that the TPMs for the bnet G_{disc} are defined in terms of the TPMs for the bnet G .

$$P(x^\sigma) = \delta(x^\sigma, x) \quad (68.1)$$

$$P(d^\sigma|x^\sigma = x) = \begin{cases} \delta(d^\sigma, 1) & \text{for } x > \xi \\ \delta(d^\sigma, 0) & \text{for } x < \xi \end{cases} \quad (68.2)$$

$$P(y^\sigma|y^\sigma(0), y^\sigma(1), d^\sigma) = \mathbb{1}(y^\sigma = y^\sigma(d^\sigma)) \quad (68.3)$$

¹ATE, which stands for “average treatment effect”, is defined in Chapter 64.

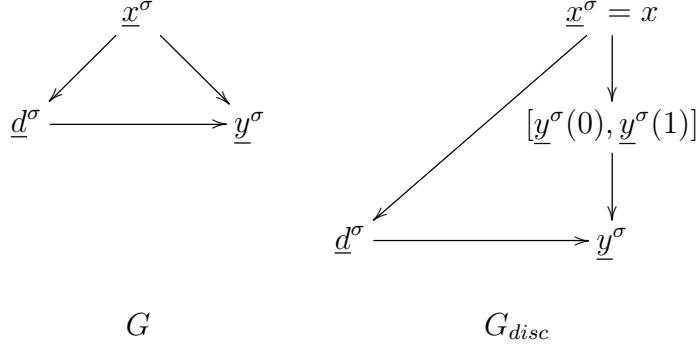


Figure 68.1: 2 bnets used in the PO analysis of RDD. The TPMs for G_{disc} are defined in terms of the TPMs for G . The TPM $P(d^{\sigma}|x^{\sigma})$ for G_{disc} is discontinuous in x^{σ} .

$$P(y^{\sigma}(c)|x) = P(y^{\sigma}(c)|d^{\sigma}, x) = \text{given} \quad (68.4)$$

Define

$$E_{\sigma|x}[y^{\sigma}(c)] = E_{y(c)|x}[\underline{y}(c)] = \mathcal{Y}_{c|x} \quad (68.5)$$

and

$$\xi^{\pm} = \xi \pm \epsilon \quad (68.6)$$

for some infinitesimal $\epsilon > 0$.

See Fig.68.2. In RDD, we assume that if we define the following 2 δ 's, one for bnet G and the other for bnet G_{disc} , then the two δ 's are equal, and they equal a conditional ATE.

$$\delta_{G_{disc}} = \mathcal{Y}_{1|x=\xi+} - \mathcal{Y}_{0|x=\xi-} \quad (68.7)$$

$$\delta_G = \mathcal{Y}_{1|x=\xi} - \mathcal{Y}_{0|x=\xi} \quad (68.8)$$

$$\delta_G = \delta_{G_{disc}} = \delta \quad (68.9)$$

$$\delta = ATE|_{x=\xi} \quad (68.10)$$

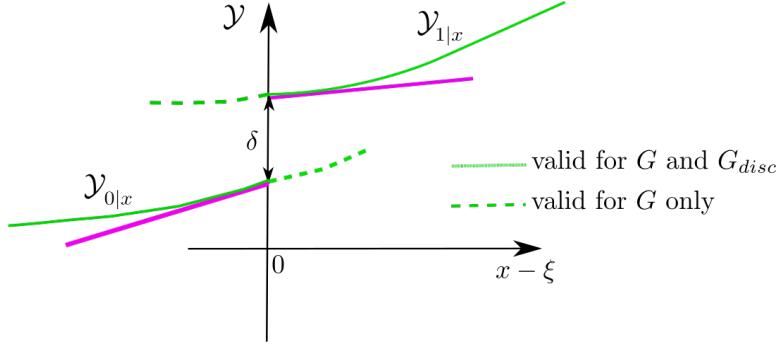


Figure 68.2: The jump δ between $\mathcal{Y}_{1|x}$ and $\mathcal{Y}_{0|x}$ is the same for G and G_{disc} .

68.2 Linear Regression

In this section, we show how to apply linear regression (LR) to the PO analysis of RDD.

y^σ can be fitted as a function of $x \in \mathbb{R}$, for $c^\sigma \in \{0, 1\}$, as follows. Here ϵ^σ is the residual for individual σ and $b_0, m_0, b_1, m_1 \in \mathbb{R}$ are the fit parameters.

$$y^\sigma = [b_0 + m_0(x - \xi)](1 - c^\sigma) + [b_1 + m_1(x - \xi)]c^\sigma + \epsilon^\sigma. \quad (68.11)$$

Note that Eq.(68.11) yields a straight line in the $y^\sigma - x$ plane for $c^\sigma = 0$, and another straight line for $c^\sigma = 1$. These 2 lines are colored magenta in Fig.68.2. We are using the standard symbols b to denote the y-intercept, and m to denote the slope of a straight line.

Taking the expected value of Eq.(68.11), we get

$$\mathcal{Y}_{c|x} = [b_0 + m_0(x - \xi)](1 - c) + [b_1 + m_1(x - \xi)]c. \quad (68.12)$$

Hence,

$$\mathcal{Y}_{1|x=\xi+} = b_1, \quad \mathcal{Y}_{0|x=\xi-} = b_0 \quad (68.13)$$

$$\delta = b_1 - b_0 \quad (68.14)$$

Chapter 69

Reinforcement Learning (RL)

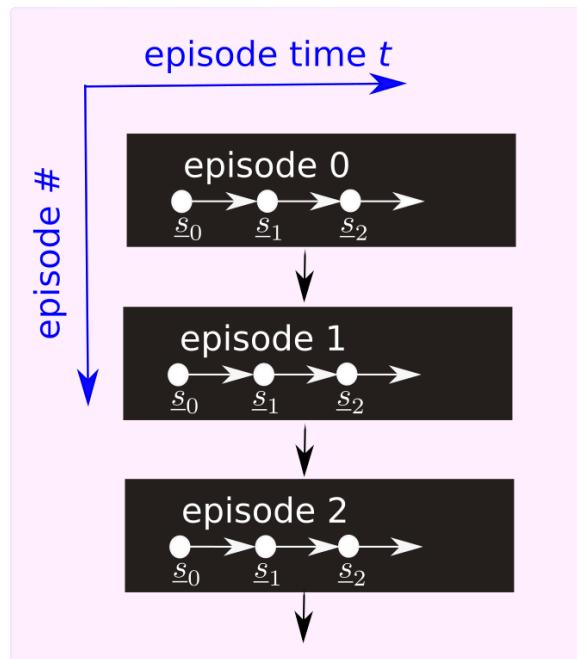


Figure 69.1: Axes for episode time and episode number.

I based this chapter on the following references. Refs.[12][28]

In RL, we consider an “agent” or robot that is learning.

Let $T \in \mathbb{Z}_{>0}$ be the duration time of an **episode** of learning. If $T = \infty$, we say that the episode has an infinite time horizon. A learning episode will evolve towards the right, for times $t = 0, 1, \dots, T - 1$. We will consider multiple learning episodes. The episode number will evolve from top to bottom. This is illustrated in Fig.69.1.

Let $\underline{s}_t \in S_s$ for $t \in \mathbb{Z}_{[0,T-1]}$ be random variables that record the **state** of the agent at various times t .

Let $\underline{a}_t \in S_a$ for $t \in \mathbb{Z}_{[0,T-1]}$ be random variables that record the **action** of the agent at various times t .

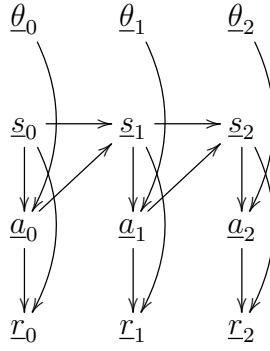


Figure 69.2: State-Action-Reward dynamical bnet

Let $\underline{\theta}_t \in S_{\underline{\theta}}$ for $t \in \mathbb{Z}_{[0,T-1]}$ be random variables that record the **policy parameters** at various times t .

For $\underline{X} \in \{\underline{s}, \underline{a}, \underline{\theta}\}$, define \underline{X} followed by a dot to be the vector

$$\underline{X}_{\cdot} = [\underline{X}_0, \underline{X}_1, \dots, \underline{X}_{T-1}] . \quad (69.1)$$

Also let

$$\underline{X}_{\geq t} = [\underline{X}_t, \underline{X}_{t+1}, \dots, \underline{X}_{T-1}] . \quad (69.2)$$

Fig.69.2 shows the basic State-Action-Reward bnet for an agent that is learning. The TPMs, printed in blue, for the bnet Fig.69.2, are as follows.

$$P(a_t|s_t, \theta_t) = \text{given.} \quad (69.3)$$

$P(a_t|s_t, \theta_t)$ is called a **policy with parameter** θ_t .

$$P(s_t|s_{t-1}, a_{t-1}) = \text{given.} \quad (69.4)$$

$P(s_t|s_{t-1}, a_{t-1})$ is called the **TPM of the model**. $P(s_t|s_{t-1}, a_{t-1})$ reduces to $P(s_0)$ when $t = 0$.

$$P(r_t|s_t, a_t) = \delta(r_t, r(s_t, a_t)) . \quad (69.5)$$

$r : S_{\underline{s}} \times S_{\underline{a}} \rightarrow \mathbb{R}$ is a given **one-time reward function**.

Note that

$$P(s_{\cdot}, a_{\cdot} | \theta_{\cdot}) = \prod_{t=0}^{T-1} \{P(s_t|s_{t-1}, a_{t-1})P(a_t|s_t, \theta_t)\} . \quad (69.6)$$

Define the **all times reward** Σ by

$$\Sigma(s., a.) = \sum_{t=0}^{T-1} \gamma^t r(s_t, a_t) . \quad (69.7)$$

Here $0 < \gamma < 1$. γ , called the **discount rate**, is included to assure convergence of Σ when $T \rightarrow \infty$. If $r(s_t, a_t) < K$ for all t , then $\Sigma < K \frac{1}{1-\gamma}$.

Define the **objective (i.e., goal) function** $E\Sigma(\theta.)$ by

$$E\Sigma(\theta.) = E_{\underline{s}.,\underline{a}.|\theta.} \Sigma(\underline{s}.,\underline{a}.) = \sum_{s.,a.} P(s.,a.|\theta.) \Sigma(s.,a.) \quad (69.8)$$

The goal of RL is to maximize the objective function over its parameters $\theta..$. The parameters θ^* . that maximize the objective function are the optimum strategy:

$$\theta.^* = \operatorname{argmax}_{\theta.} E\Sigma(\theta.) \quad (69.9)$$

Define a **future reward** for times $\geq t$ as:

$$\Sigma_{\geq t}((s_{t'}, a_{t'})_{t' \geq t}) = \sum_{t'=t}^{T-1} \gamma^{t'-t} r(s_{t'}, a_{t'}) \quad (69.10)$$

Define the following **expected conditional future rewards** (rewards for times $\geq t$, conditioned on certain quantities having given values):

$$v_t = v(s_t, a_t; \theta.) = E_{\underline{s}.,\underline{a}.|s_t,a_t,\theta.} [\Sigma_{\geq t}] \quad (69.11)$$

$$V_t = V(s_t; \theta.) = E_{\underline{s}.,\underline{a}.|s_t,\theta.} [\Sigma_{\geq t}] = E_{\underline{a}_t|s_t,\theta.} [v(s_t, \underline{a}_t; \theta.)] \quad (69.12)$$

v is usually called Q in the literature. We will refer to Q as v in order to follow a convention wherein an \underline{a}_t -average changes a lower case letter to an upper case one.

We will sometimes write $v(s_t, a_t)$ instead of $v(s_t, a_t; \theta.).$

Since $E\Sigma_{\geq t}$ only depends on $\theta_{\geq t}$, $v(s_t, a_t; \theta.) = v(s_t, a_t; \theta_{\geq t})$, and $V(s_t; \theta.) = V(s_t; \theta_{\geq t})$.

Note that the objective function $E\Sigma$ can be expressed in terms of v_0 by averaging over its unaveraged parameters:

$$E\Sigma(\theta.) = E_{\underline{s}_0,\underline{a}_0|\theta_0} v(\underline{s}_0, \underline{a}_0; \theta.) \quad (69.13)$$

Define a **one-time reward** and an **expected conditional one-time reward** as:

$$r_t = r(s_t, a_t) \quad (69.14)$$

$$R_t = R(s_t; \theta_t) = E_{\underline{a}_t|s_t,\theta_t} [r(s_t, \underline{a}_t)] . \quad (69.15)$$

Note that

$$\Sigma_{\geq t} = r_t + \gamma r_{t+1} + \gamma^2 r_{t+2} + \dots + \gamma^{T-1-t} r_{t+(T-1-t)} \quad (69.16)$$

$$= r_t + \gamma \Sigma_{\geq t+1}; \quad (69.17)$$

If we take $E_{s.,a.|s_t,a_t,\theta.}[\cdot]$ of both sides of Eq.(69.17), we get

$$v_t = r_t + \gamma E_{s_{t+1},a_{t+1}|\theta.}[v_{t+1}] . \quad (69.18)$$

If we take $E_{s.,a.|s_t,\theta.}[\cdot]$ of both sides of Eq.(69.17), we get

$$V_t = R_t + \gamma E_{s_{t+1}|\theta.}[V_{t+1}] . \quad (69.19)$$

Note that

$$\Delta r_t = r_t - R_t \quad (69.20)$$

$$= r_t - (V_t - \gamma E_{s_{t+1}|\theta.}[V_{t+1}]) \quad (69.21)$$

$$= r_t + \gamma E_{s_{t+1}|\theta.}[V_{t+1}] - V_t . \quad (69.22)$$

Define

$$\Delta v_t = v_t - V_t . \quad (69.23)$$

Note that

$$\Delta v_t = \Delta r_t . \quad (69.24)$$

Next, we will discuss 3 RL bnets

- exact RL bnet (exact, assumes policy is known)
- Actor-Critic RL bnet (approximate, assumes policy is known)
- Q function learning RL bnet (approximate, assumes policy is NOT known)

69.1 Exact RL bnet

An exact RL bnet is given by Fig.69.3.

Fig.69.3 is the same as Fig.69.2 but with more nodes added in order to optimize the policy parameters. The TPMs, printed in blue, for the nodes not already discussed in connection to bnet Fig.69.2, are as follows.

$$P(\theta_t|\theta.) = \delta(\theta_t, (\theta.)_t) \quad (69.25)$$

$$\forall(s_t, a_t) : P(v_t(s_t, a_t)|r_t, v_{t+1}(\cdot), \theta.) = \delta(v_t(s_t, a_t), r_t + \gamma E_{s_{t+1},a_{t+1}|\theta.}[v_{t+1}]) \quad (69.26)$$

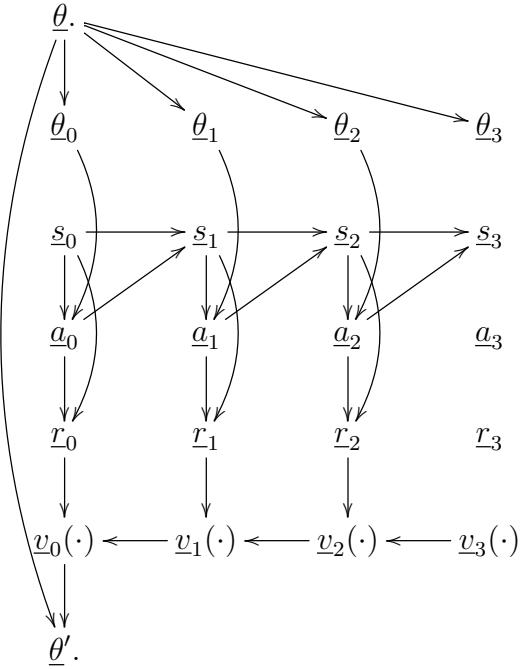


Figure 69.3: Exact RL bnet. $v_t(\cdot)$ means the array $[v_t(s_t, a_t)]_{\forall s_t, a_t}$ The following arrows are implicit: for all t , arrow from $\underline{\theta}.$ $\rightarrow \underline{v}_t(\cdot).$ We did not draw those arrows so as not to clutter the diagram.

$$P(\theta' | \theta, v_0(\cdot)) = \delta(\theta', \theta + \alpha \partial_{\theta} \underbrace{E_{\underline{s}_0, \underline{a}_0 | \theta_0} v(\underline{s}_0, \underline{a}_0; \theta)}_{E\Sigma(\theta)}) \quad (69.27)$$

$\alpha > 0$ is called the **learning rate**. This method of improving θ . is called gradient ascent.

Concerning the gradient of the objective function, note that

$$\partial_{\theta_t} E\Sigma(\theta) = \sum_{s., a.} \partial_{\theta_t} P(s., a. | \theta.) \Sigma(s., a.) \quad (69.28)$$

$$= \sum_{s., a.} P(s., a. | \theta.) \partial_{\theta_t} \ln P(s., a. | \theta.) \Sigma(s., a.) \quad (69.29)$$

$$= E_{\underline{s}, \underline{a} | \theta.} \{ \partial_{\theta_t} \ln P(a_t | s_t, \theta_t) \Sigma(s., a.) \} . \quad (69.30)$$

If we run the agent $nsam(\vec{s}_t)$ times and obtain samples $s_t[i], a_t[i]$ for all t and for $i = 0, 1, \dots, nsam(\vec{s}_t) - 1$, we can express this gradient as follows:

$$\partial_{\theta_t} E\Sigma(\theta) \approx \frac{1}{nsam(\vec{s}_t)} \sum_i \sum_{t=0}^{T-1} \partial_{\theta_t} \ln P(a_t[i] | s_t[i], \theta_t) r(s_t[i], a_t[i]) . \quad (69.31)$$

The exact RL bnet Fig.69.3 is difficult to use to calculate the optimum parameters θ^* . The problem is that \underline{s}_t propagates towards the future and the $\underline{v}_t(\cdot)$ propagates towards the past, so we don't have a Markov Chain with a chain link for each t (i.e., a dynamical bnet) in the episode time direction. Hence, people have come up with approximate RL bnets that are doubly dynamical (i.e., dynamical along the episode time and episode number axes.) We discuss some of those approximate RL bnets next.

69.2 Actor-Critic RL bnet

For the actor-critic RL bnet, we approximate Eq.(69.31) by

$$\partial_{\theta_t} E\Sigma(\theta) \approx \frac{1}{nsam(\vec{s})} \sum_i \sum_{t=0}^{T-1} \underbrace{\partial_{\theta_t} \ln P(a_t[i] | s_t[i], \theta_t)}_{Actor} \underbrace{\Delta r_t(s_t[i], a_t[i])}_{Critic} \quad (69.32)$$

The actor-critic RL bnet is given by Fig.69.4. This bnet is approximate and assumes that the policy is known. The TPMs, printed in blue, for this bnet, are as follows.

$$P(\theta_t) = \text{given} \quad (69.33)$$

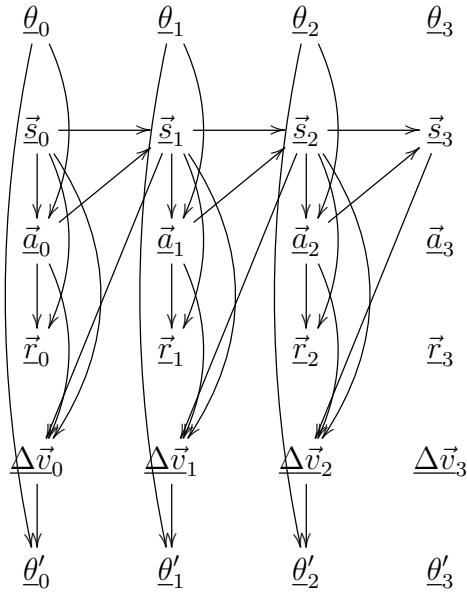


Figure 69.4: Actor-Critic RL bnet.

$$P(s_t[i] | s_{t-1}[i], a_{t-1}[i]) = \text{ given} \quad (69.34)$$

$$P(a_t[i] | s_t[i], \theta_t) = \text{ given} \quad (69.35)$$

$$P(r_t[i] | s_t[i], a_t[i]) = \delta(r_t[i], r(s_t[i], a_t[i])) \quad (69.36)$$

$r : S_{\underline{s}} \times S_{\underline{a}} \rightarrow \mathbb{R}$ is given.

$$P(\Delta v_t[i] | s_t[i], a_t[i], s_{t+1}[i]) = \delta(\Delta v_t[i], r(s_t[i], a_t[i]) + \gamma \widehat{V}(s_{t+1}[i]; \phi') - \widehat{V}(s_t[i]); \phi) . \quad (69.37)$$

$$P(\theta'.) = \delta(\theta'., \theta_t + \alpha \partial_{\theta_t} \sum_i \ln P(a_t[i] | s_t[i], \theta_t) \Delta v_t[i]) \quad (69.38)$$

$\widehat{V}(s_t[i]; \phi)$ is obtained by curve fitting (see Chapter 6) using samples $(s_t[i], a_t[i])$ $\forall t, i$ with

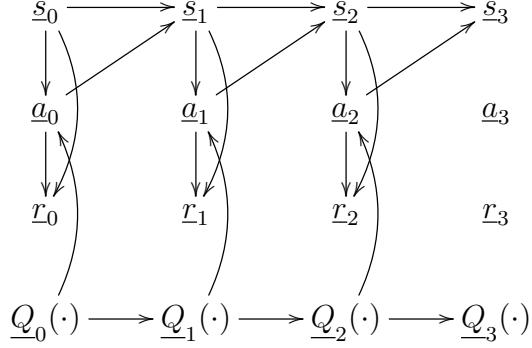


Figure 69.5: Q function learning RL bnet.

$$y[i] = \sum_{t'=t}^T r(s_{t'}[i], a_{t'}[i]) \quad (69.39)$$

and

$$\hat{y}[i] = \hat{V}(s_t[i]; \phi) . \quad (69.40)$$

Eq.(69.39) is an approximation because $(s_{t'}, a_{t'})_{t'>t}$ are averaged over in the exact expression for $V(s_t)$. $\hat{V}(s_{t+1}[i]); \phi'$ is obtained in the same way as $\hat{V}(s_t[i]); \phi$ but with t replaced by $t + 1$ and ϕ by ϕ' .

69.3 Q function learning RL bnet

The Q-function learning RL bnet is given by Fig.69.5. This bnet is approximate and assumes that the policy is NOT known. The TPMs, printed in blue, for this bnet, are as follows. (Remember that $Q = v$).

$$P(s_t|s_{t-1}, a_{t-1}) = \text{given} \quad (69.41)$$

$$P(a_t|s_t, v_t(\cdot)) = \delta(a_t, \operatorname{argmax}_a v_t(s_t, a)) \quad (69.42)$$

$$P(r_t|s_t, a_t) = \delta(r_t, r(s_t, a_t)) \quad (69.43)$$

$r : S_s \times S_a \rightarrow \mathbb{R}$ is given.

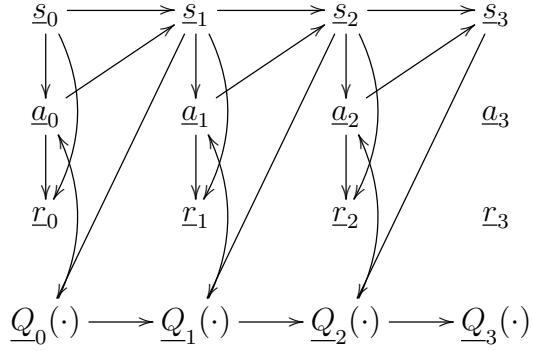


Figure 69.6: Q function learning RL bnet. Same as Fig.69.5 but with new arrow passing \$s_t\$ to \$\underline{Q}_{t-1}\$.

$$\begin{aligned} \forall(s_t, a_t) : P(v_t(s_t, a_t) | v_{t-1}(\cdot)) = \\ = \delta(v_t(s_t, a_t), r(s_t, a_t) + \gamma \max_a E_{s_{t+1}|s_t, a_t} v_{t-1}(s_{t+1}, a)) \end{aligned} \quad (69.44)$$

This value for \$v_t(s_t, a_t)\$ approximates \$v_t = r_t + \gamma E_{s_{t+1}|s_t, a_t} v_{t+1}\$.

Some people use the bnet of Fig.69.6 instead of Fig.69.5 and replace Eq.(69.44) by

$$\begin{aligned} \forall(s_t, a_t) : P(v_t(s_t, a_t) | s_{t+1}, v_{t-1}(\cdot)) = \\ = \delta(v_t(s_t, a_t), r(s_t, a_t) + \gamma \max_a v_{t-1}(s_{t+1}, a)) . \end{aligned} \quad (69.45)$$

Chapter 70

Reliability Box Diagrams and Fault Tree Diagrams

This chapter is based on Refs.[53] and [71].

In this chapter, we assume that reader is familiar with Boolean Algebra. See Chapter C for a quick review of what we recommend that you know about Boolean Algebra to fully appreciate this chapter.

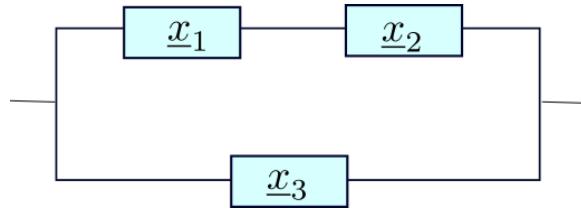


Figure 70.1: Example of rbox diagram.

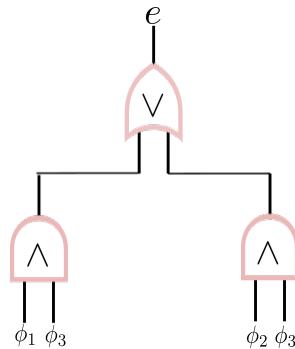


Figure 70.2: An ftree diagram equivalent to Fig.70.1. It represents $e = (\phi_1 \wedge \phi_3) \vee (\phi_2 \wedge \phi_3)$.

Complicated devices with a large number of components such as airplanes or rockets can fail in many ways. If their performance depends on some components

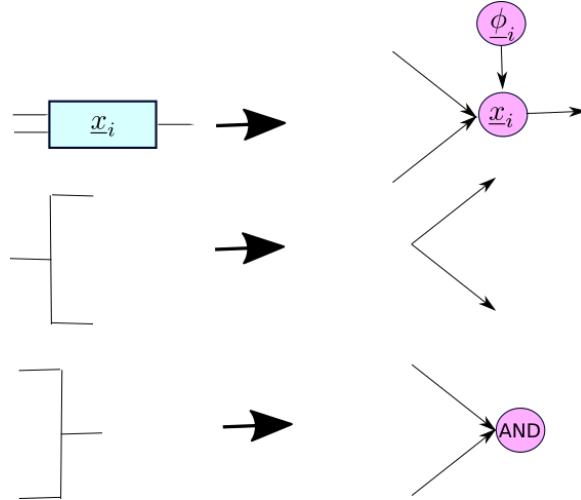


Figure 70.3: How to map an rbox diagram to a bnet.

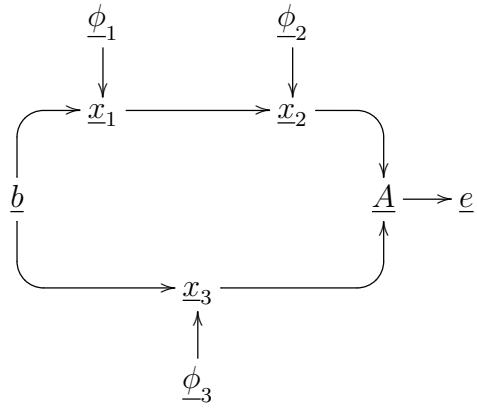


Figure 70.4: bnet corresponding the rbox diagram Fig.70.1.

working in series and one of the components in the series fails, this may lead to catastrophic failure. To avert such disasters, engineers use equivalent components connected in parallel instead of in series, thus providing multiple backup systems. They analyze the device to find its weak points and add backup capabilities there. They also estimate the average time to failure for the device.

The two most popular diagrams for finding the failure modes and their rates for large complicated devices are

- rbox diagrams = Reliability Box diagrams. See Fig.70.1 for an example.
- ftree diagrams = Fault Tree Diagrams. See Fig.70.2 for an example.

In an ftree diagram, several nodes might stand for the same component of a physical

device. In an rbox diagram, on the other hand, each node represents a distinct component in a device. Hence, rbox diagrams resemble the device they are addressing whereas ftree diagrams don't. Henceforth, we will refer to this desirable property as **physical resemblance**.

As we will show below with an example, it is pretty straightforward to translate an rbox to an ftree diagram. Going the other way, translating an ftree to an rbox diagram is much more difficult.

Next we will define a new kind of bnet that we will call a failure bnet that has physical resemblance. Then we will describe a simple method of translating (i.e., mapping) any rbox diagram to a failure bnet. Then we will show how a failure bnet can be used to do all the calculations that are normally done with an rbox or an ftree diagram. In that sense, failure bnets seem to afford all the benefits of both ftree and rbox diagrams.

A **failure bnet** contains nodes of 5 types, labeled \underline{b} , \underline{e} , \underline{x}_i , $\underline{\phi}_i$, and \underline{A}_i . All nodes have only two possible states $S = Success = 0$, $F = Failure = 1$.

1. The bnet has a beginning node labeled \underline{b} which is always set to success. The \underline{b} node and the $\underline{\phi}_i$ nodes are the only root nodes of the bnet.
2. The bnet has a single leaf node, the end node, labeled \underline{e} . \underline{e} is fixed. In rbox diagrams, $\underline{e} = S$ whereas in ftree diagrams, $\underline{e} = F$.
3. $\underline{x}^{nx} = (\underline{x}_0, \underline{x}_1, \dots, \underline{x}_{nx-1})$. $\underline{x}_i \in \{S, F\}$ for all i .

Suppose \underline{x}_i has parents $\underline{\phi}_i$ and $\underline{a}^{na} = (a_0, a_1, \dots, a_{na-1})$. Then the TPM of node \underline{x}_i is defined to be

$$P(x_i | \phi_i, a^{na}) = \delta(x_i, \phi_i \vee \vee_{i=0}^{na-1} a_i) \quad (70.1)$$

4. For each node \underline{x}_i , the bnet has a “performance” root node $\underline{\phi}_i \in \{0, 1\}$ with an arrow pointing from it to \underline{x}_i (i.e., $\underline{\phi}_i \rightarrow \underline{x}_i$). For all i ,

$$P(\phi_i) = \epsilon_i \delta(\phi_i, F) + \bar{\epsilon}_i \delta(\phi_i, S). \quad (70.2)$$

ϵ_i is the failure probability and $\bar{\epsilon}_i = 1 - \epsilon_i$ is the success probability. We name the failure probability ϵ_i because it is normally very small. It is usually set to $1 - e^{-\lambda_i t} \approx \lambda_i t$ when $\lambda_i t \ll 1$, where λ_i is the failure rate for node \underline{x}_i and t stands for time. The rblock literature usually calls $\bar{\epsilon}_i = R_i$ the **reliability** of node \underline{x}_i , and $\epsilon_i = (1 - R_i) = F_i$ its **unreliability**.

5. The nodes $\underline{A}_i \in \{0, 1\}$ are simply AND gates. If \underline{A}_i has inputs $\underline{y}^{ny} = (\underline{y}_0, \underline{y}_1, \dots, \underline{y}_{ny-1})$, then the TPM of \underline{A}_i is

$$P(A_i|y^{ny}) = \delta(A_i, \wedge_{i=0}^{ny-1} y_i) . \quad (70.3)$$

An instance (instantiation) of a bnet is the bnet with all nodes set to a specific state. A **realizable instance (r-instance)** of a bnet is one which has non-zero probability.

Fig.70.3 shows how to translate any rbox diagram to a failure bnet. To illustrate this procedure, we translated the rbox diagram Fig.70.1 into the failure bnet Fig.70.4.

For the failure bnet Fig.70.4, one has:

$$\begin{aligned} P(b) &= \mathbb{1}(b = 0) \\ P(x_1|\phi_1, b) &= \mathbb{1}(x_1 = \phi_1 \vee b) \\ P(x_2|\phi_2, x_1) &= \mathbb{1}(x_2 = \phi_2 \vee x_1) \\ P(x_3|\phi_3, b) &= \mathbb{1}(x_3 = \phi_3 \vee b) \\ P(A|x_2, x_3)e &= \mathbb{1}(x_2 \wedge x_3) \\ P(e|A) &= \mathbb{1}(e = A) \end{aligned} . \quad (70.4)$$

Therefore, all r-instances of this bnet must satisfy

$$e = (\phi_1 \vee \phi_2) \wedge \phi_3 \quad (70.5)$$

$$= (\phi_1 \wedge \phi_3) \vee (\phi_2 \wedge \phi_3) . \quad (70.6)$$

Eq.(70.6) proves that Fig.70.2 is indeed a representation of Fig.70.1.

Next, we consider r-instances of this bnet for two cases: $e = S$ and $e = F$.

- **rblock analysis:** $e = S = 0$.

Table 70.1 shows the probability of all possible r-instances that end in success for the failure bnet Fig.70.4. (These r-instances are the main focus of rblock analysis). The first 4 of those probabilities (those with $\phi_3 = 0$) sum to $\bar{\epsilon}_3$ so the sum $P(e = S)$ of all 5 is

$$P(e = S) = \bar{\epsilon}_3 + \bar{\epsilon}_1 \bar{\epsilon}_2 \epsilon_3 , \quad (70.7)$$

or, expressing it in reliability language in which $\bar{\epsilon} = R$,

$$P(e = S) = R_3 + R_1 R_2 \bar{R}_3 . \quad (70.8)$$

- **ftree analysis:** $e = F = 1$.

Table 70.2 shows the probability of all possible r-instances that end in failure for the failure bnet Fig.70.4. (These r-instances are the main focus of ftree analysis). If we set $\epsilon_i = \epsilon$ and $\bar{\epsilon}_i \approx 1$ for $i = 1, 2, 3$, then the first two of

instance	probability
	$\bar{\epsilon}_1 \epsilon_2 \bar{\epsilon}_3$
	$\epsilon_1 \bar{\epsilon}_2 \bar{\epsilon}_3$
	$\epsilon_1 \epsilon_2 \bar{\epsilon}_3$
	$\bar{\epsilon}_1 \bar{\epsilon}_2 \bar{\epsilon}_3$
	$\bar{\epsilon}_1 \bar{\epsilon}_2 \epsilon_3$

Table 70.1: Probabilities of all possible r-instances with $e = S = 0$ for failure bnet Fig.70.4.

instance	probability
	$\bar{\epsilon}_1 \epsilon_2 \epsilon_3$
	$\epsilon_1 \bar{\epsilon}_2 \epsilon_3$
	$\epsilon_1 \epsilon_2 \epsilon_3$

Table 70.2: Probabilities of all possible r-instances with $e = F = 1$ for the failure bnet Fig.70.4.

those r-instances have probabilities of $order(\epsilon^2)$ and the third has probability of $order(\epsilon^3)$. The two lowest order ($order(\epsilon^2)$) r-instances are called the “minimal cut sets” of the ftree. We will have more to say about minimal cut sets later on. For now, just note from Eq.(70.6) that the ftree Fig.70.2 is just the result of joining together with ORs two expressions, one for each of the two minimal cut sets.

More general \underline{x}_i .

Failure bnets can actually accommodate \underline{x}_i nodes of a more general kind than what we first stipulated. Here are some possibilities:

For any $a^n \in \{0, 1\}^n$, let

$$len(a^n) = \sum_i a_i \quad (70.9)$$

- **OR gate**

$$P(x_i|\phi_i, a^{na}) = \delta(x_i, \phi_i \vee \vee_j a_j) \quad (70.10)$$

$$= \delta(x_i, \phi_i \vee \mathbb{1}(\text{len}(a^{na}) > 0)) \quad (70.11)$$

- **AND gate**

$$P(x_i|\phi_i, a^{na}) = \delta(x_i, \phi_i \vee \wedge_j a_j) \quad (70.12)$$

$$= \delta(x_i, \phi_i \vee \mathbb{1}(\text{len}(a^{na}) = na)) \quad (70.13)$$

- **Fail if least K failures (less than K successes)**

$$P(x_i|\phi_i, a^{na}) = \delta(x_i, \phi_i \vee \mathbb{1}(\text{len}(a^{na}) \geq K)) \quad (70.14)$$

- **Fail if less than K failures (at least K successes)**

$$P(x_i|\phi_i, a^{na}) = \delta(x_i, \phi_i \vee \mathbb{1}(\text{len}(a^{na}) < K)) \quad (70.15)$$

- **Fail if exactly one failure**

$$P(x_i|\phi_i, a^{na}) = \delta(x_i, \phi_i \vee \mathbb{1}(\text{len}(a^{na}) = 1)) \quad (70.16)$$

This equals an XOR (exclusive OR) gate when $na = 2$.

- **General gate**

$$f : \{0, 1\}^{na} \rightarrow \{0, 1\}$$

$$P(x_i|\phi_i, a^{na}) = \delta(x_i, \phi_i \vee f(a^{na})) \quad (70.17)$$

70.1 Minimal Cut Sets

Suppose $x \in \{0, 1\}$ and $f : \{0, 1\} \rightarrow \{0, 1\}$. Then by direct evaluation, we see that

$$f(x) = [\bar{x}f(0)] \vee [xf(1)]. \quad (70.18)$$

Let

$$\begin{aligned} !x &= 1 - x, \\ !^0 x &= x, \\ !^1 x &= !x \end{aligned} \tag{70.19}$$

Then Eq.(70.18) can be rewritten as

$$f(x) = \vee_{a \in \{0,1\}} [(!^{\bar{a}} x) f(a)] . \tag{70.20}$$

Now suppose $x^n \in \{0,1\}^n$ and $f : \{0,1\}^n \rightarrow \{0,1\}$. Eq.(70.20) generalizes to

$$f(x^n) = \vee_{a^n \in \{0,1\}^n} \left[\prod_i (!^{\bar{a}_i} x_i) f(a^n) \right] . \tag{70.21}$$

Eq.(70.21) is called an **ors-of-ands** normal form expansion. There is also an **ands-of-orss** normal form expansion obtained by swapping multiplication and \vee in Eq.(70.21), but we won't need it here.

A **cut set** is a set of ϕ_i 's such that if they are all equal to F , then $e = F$ for all the r-instances. A **minimal cut set** is a cut set such that there are no larger cut sets that contain it. From the failure bnet, we can always find a function $f : \{0,1\}^{nx} \rightarrow \{0,1\}$ such that $e = f(\phi^{nx})$ for all the r-instances. We did that for our example failure bnet and obtained Eq.(70.6). We can then express $f(\phi^{nx})$ as an ors-of-ands expansion to find all the minimal cut sets. The ands terms in that ors-of-ands expansion each gives a different minimal cut set, after some simplification. The ors-of-ands expression is not unique and it may be necessary to simplify (using the Boolean Algebra identities given in Chapter C) to remove those redundancies.

Chapter 71

Restricted Boltzmann Machines

In what follows, we will abbreviate "restricted Boltzmann machine" by rebo.

Let

$$v \in \{0, 1\}^{nv}$$

$$h \in \{0, 1\}^{nh}$$

$b \in \mathbb{R}^{nv}$ (mnemonic, v and b sound the same)

$$a \in \mathbb{R}^{nh}$$

$$W^{v|h} \in \mathbb{R}^{nv \times nh}$$

Energy:

$$E(v, h) = -(b^T v + a^T h + v^T W^{v|h} h) \quad (71.1)$$

Boltzmann distribution:

$$P(v, h) = \frac{e^{-E(v, h)}}{Z} \quad (71.2)$$

Partition function:

$$Z = \sum_{v, h} e^{-E(v, h)} = Z(a, b, W^{v|h}) \quad (71.3)$$

$$P(v|h) = \frac{e^{b^T v + a^T h + v^T W^{v|h} h}}{\sum_v e^{b^T v + a^T h + v^T W^{v|h} h}} \quad (71.4)$$

$$= \frac{e^{b^T v + v^T W^{v|h} h}}{\sum_v e^{b^T v + v^T W^{v|h} h}} \quad (71.5)$$

$$= \prod_i \frac{e^{v_i(b_i + \sum_j W_{i,j}^{v|h} h_j)}}{\sum_{v_i=0,1} e^{v_i(b_i + \sum_j W_{i,j}^{v|h} h_j)}} \quad (71.6)$$

$$= \prod_i P(v_i|h) \quad (71.7)$$

$$P(v_i|h) = \frac{e^{v_i(b_i + \sum_j W_{i,j}^{v|h} h_j)}}{Z_i(h)} \quad (71.8)$$

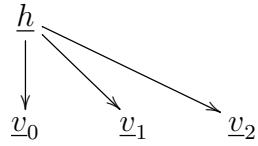


Figure 71.1: bnet for a Restricted Boltzmann Machine (rebo) with $nv = 3$

Eq.(71.8) implies that a rebo can be represented by the bnet Fig.71.1.

Let

$$x_i = b_i + \sum_j W_{ij}^{vh} h_j . \quad (71.9)$$

Then

$$P(v_i = 1|h) = \frac{e^{x_i}}{1 + e^{x_i}} \quad (71.10)$$

$$= \frac{1}{1 + e^{-x_i}} \quad (71.11)$$

$$= \text{smoid}(x_i) . \quad (71.12)$$

One could also expand the node h in Fig.71.1 into nh nodes. But note that $P(h) \neq \prod_j P(h_j)$ so there would be arrows among the h_j nodes.

Note that the rebo bnet is a special case of Naive Bayes (See Chapter 57) with $v_i, h_i \in \{0, 1\}$ and specific $P(h)$ and $P(v_i|h)$ node matrices.

Chapter 72

ROC curves

This chapter is based on Ref.[134].

ROC stands for **Receiver Operating Characteristic**. ROC curves are used in binary classification (BC).

To do BC, we are given the value $x \in \mathbb{R}$ for an individual. From this, we want to decide whether that individual has $a = 0$ or $a = 1$. The decision will depend on the value of a **threshold parameter** $\tau \in \mathbb{R}$.

$$\underline{x} \longleftarrow \underline{a}$$

Figure 72.1: bnet for BC.

Fig.72.1 shows the bnet used for BC.

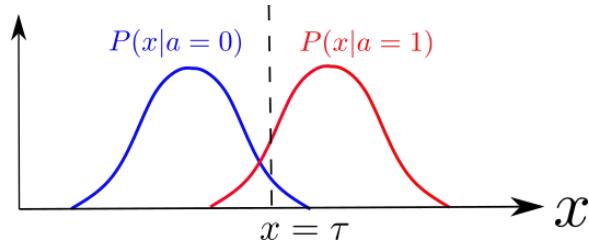


Figure 72.2: x -distribution for two hypotheses $a = 0, 1$.

Fig.72.2 is a plot of $P(x|a)$, i.e., the TPM for node \underline{x} of the bnet in Fig.72.1. Whereas a is binary, x is continuous. But we can replace x by a binary variable

$$b = \mathbb{1}(x > \tau). \quad (72.1)$$

$P(b|a)$ for $b, a \in \{0, 1\}$ is called the **confusion matrix** or **contingency table** for BC. The confusion matrix can be calculated from the TPM $P(x|a)$. Fig.72.3 illustrates

		Actual Value (a)	
		0	1
Predicted Value (x)	0	$TNR \triangleleft$	$FNR \curvearrowright$
	1	$FPR \curvearrowleft$	$TPR \triangleup$

Figure 72.3: The confusion matrix $P(b|a)$ for BC.

the confusion matrix $P(b|a)$ for BC. In that figure, the rates R are defined as follows.¹

- **True Negative Rate (TNR)**

$$R_{0|0}(\tau) = P(x < \tau | a = 0) = \int_{x < \tau} dx P(x | a = 0) \quad (72.2)$$

- **False Positive Rate (FPR)**

$$R_{1|0}(\tau) = 1 - R_{0|0}(\tau) \quad (72.3)$$

$$= P(x > \tau | a = 0) = \int_{x > \tau} dx P(x | a = 0) \quad (72.4)$$

In Hypothesis Testing, $R_{1|0}$ is called the **p-value that $x > \tau$ assuming curve 0 is the null hypothesis**.

- **False Negative Rate (FNR)**

$$R_{0|1}(\tau) = P(x < \tau | a = 1) = \int_{x < \tau} dx P(x | a = 1) \quad (72.5)$$

In Hypothesis Testing, $R_{0|1}$ is called the **p-value that $x < \tau$ assuming curve 1 is the null hypothesis**.

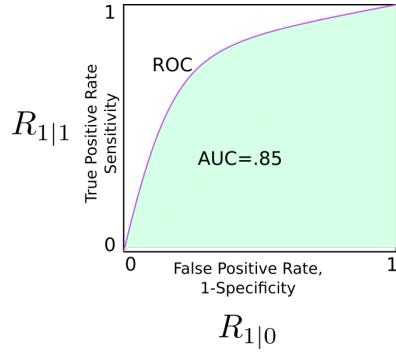
- **True Positive Rate (TPR)**

$$R_{1|1}(\tau) = 1 - R_{0|1}(\tau) \quad (72.6)$$

$$= P(x > \tau | a = 1) = \int_{x > \tau} dx P(x | a = 1) \quad (72.7)$$

The **Receiver Operating Characteristic (ROC)** is a parametric plot with $X = R_{1|0}(\tau)$ and $Y = R_{1|1}(\tau)$, where $\tau \in \mathbb{R}$. The **Area Under the Curve (AUC)** is the area under the ROC. Fig.72.4 shows an example of a ROC and its AUC.

¹I find the notation $x|a$ where $x, a \in \{0, 1\}$ much clearer than $\alpha\beta$ where $\alpha = T, F$ and $\beta = N, P$. Note that $\alpha = \mathbb{1}(x = a)$ and $\beta = x$, if we identify $0 = F = N$ and $1 = T = P$.



$$R_{1|0}$$

Figure 72.4: Example of ROC. Green shaded area is the AUC of the ROC.

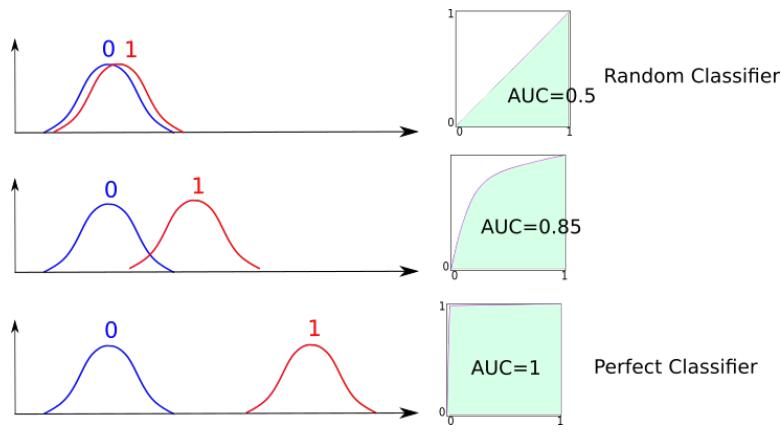


Figure 72.5: ROC curves for 3 different separations between the 0 and 1 x-distributions.

Fig.72.5 shows situations that give $AUC=.5$ (random classifier), $AUC=.85$, and $AUC=1$ (perfect classifier). It's also possible to get an $AUC \in [0, 0.5]$, but we will ignore those models because they are useless for BC.

Note that

$$AUC = \int_{x=0}^1 d\tau R_{1|1}(\tau) \frac{dR_{1|0}(\tau)}{d\tau} \quad (72.8)$$

$$= \int_{-\infty}^{-\infty} d\tau \left\{ \int_{-\infty}^{\infty} dx \mathbb{1}(x > \tau) P(x|a=1) \right\} (-1) P(x = \tau|a=0) \quad (72.9)$$

$$= \int_{-\infty}^{\infty} dx' \int_{-\infty}^{\infty} dx \mathbb{1}(x > x') P(x|a=1) P(x'|a=0) . \quad (72.10)$$

72.1 Terminology Table Adapted from Wikipedia Ref.[134]

Let $N_{x|a}$ be numbers (counts) so that

$$P(x|a) = \frac{N_{x|a}}{\sum_{x'} N_{x'|a}} \quad (72.11)$$

for all $x, a \in \{0, 1\}$.

condition positive (P): $N_{|1} = \sum_x N_{x|1}$, the number of real positive cases in the data

condition negative (N): $N_{|0} = \sum_x N_{x|0}$, the number of real negative cases in the data

true positive (TP): $N_{1|1}$, hit

true negative (TN): $N_{0|0}$, correct rejection

false positive (FP): $N_{1|0}$, false alarm, type I error or underestimation

false negative (FN): $N_{0|1}$, miss, type II error or overestimation

sensitivity, recall, hit rate, or true positive rate (TPR):

$$TPR = R_{1|1} = \frac{N_{1|1}}{N_{|1}} = \frac{N_{1|1}}{N_{1|1} + N_{0|1}} = 1 - R_{0|1} \quad (72.12)$$

specificity, selectivity or true negative rate (TNR):

$$TNR = R_{0|0} = \frac{N_{0|0}}{N_{|0}} = \frac{N_{0|0}}{N_{0|0} + N_{1|0}} = 1 - R_{1|0} \quad (72.13)$$

precision or positive predictive value (PPV):

$$PPV = \frac{N_{1|1}}{N_{1|1} + N_{1|0}} = 1 - FDR \quad (72.14)$$

negative predictive value (NPV):

$$NPV = \frac{N_{0|0}}{N_{0|0} + N_{0|1}} = 1 - FOR \quad (72.15)$$

miss rate or false negative rate (FNR):

$$FNR = R_{0|1} = \frac{N_{0|1}}{N_{|1}} = \frac{N_{0|1}}{N_{0|1} + N_{1|1}} = 1 - R_{1|1} \quad (72.16)$$

fall-out or false positive rate (FPR):

$$FPR = R_{1|0} = \frac{N_{1|0}}{N_{|0}} = \frac{N_{1|0}}{N_{1|0} + N_{0|0}} = 1 - R_{0|0} \quad (72.17)$$

false discovery rate (FDR):

$$FDR = \frac{N_{1|0}}{N_{1|0} + N_{1|1}} = 1 - PPV \quad (72.18)$$

false omission rate (FOR):

$$FOR = \frac{N_{0|1}}{N_{0|1} + N_{0|0}} = 1 - NPV \quad (72.19)$$

accuracy (ACC):

$$ACC = \frac{N_{1|1} + N_{0|0}}{N_{1|1} + N_{0|0}} = \frac{N_{1|1} + N_{0|0}}{N_{1|1} + N_{0|0} + N_{1|0} + N_{0|1}} \quad (72.20)$$

balanced accuracy (BA):

$$BA = \frac{R_{1|1} + R_{0|0}}{2} \quad (72.21)$$

F1 score is the harmonic mean of precision and sensitivity:

$$F_1 = 2 \times \frac{PPV \times R_{1|1}}{PPV + R_{1|1}} = \frac{2N_{1|1}}{2N_{1|1} + N_{1|0} + N_{0|1}} \quad (72.22)$$

Chapter 73

Scoring the Nodes of a Learned Bnet

Chapter 78 discusses how to learn a bnet from data. Many algorithms for doing this require scoring how well a particular bnet fits the data. This chapter is an introduction to such scoring.

Normally, each node of a bnet is scored separately, and then those node scores are summed to get the bnet score.

In this chapter, scores are defined so that a higher score means a better fit. By taking the negative of such a score, one can always get a score such that a lower score means a better fit.

There are 2 main types of bnet scores: Maximum Likelihood (ML) scores, and Shannon Information Theory (SIT) scores. ML scores consist of the log of a maximum likelihood function $P(\vec{x}|\theta)$ for i.i.d. samples $\vec{x} = (x^\sigma)_{\sigma=0,1,\dots,nsam-1}$, where $x^\sigma \sim P_{\underline{x}|\theta}(x|\theta)$:

$$\text{ML-score} = \ln(P(\vec{x}|\theta)) \quad (73.1)$$

$$= \ln \prod_{\sigma} P(x^\sigma | \theta) \quad (73.2)$$

$$= \sum_{\sigma} \ln P(x^\sigma | \theta) \quad (73.3)$$

$$\approx nsam \sum_x P(x|\theta) \ln P(x|\theta) \quad (73.4)$$

$$= -nsam H(P_{\underline{x}|\theta}) , \quad (73.5)$$

and SIT scores consist of a negative entropy:

$$\text{Info-score} = -H(P_{\underline{x}|\theta}) . \quad (73.6)$$

Thus, up to a factor of $nsam$, they are the same thing. Maximizing a log likelihood function for i.i.d. samples or minimizing the corresponding entropy, are the same

thing, and they both yield a good estimate of the hidden parameters θ .

73.1 Probability Distributions and Special Functions

While writing this chapter, I briefly consulted the following Wikipedia articles about the definitions and properties of certain probability distributions and special functions.

- Categorical Distribution, Ref.[87]
- Multinomial Distribution, Ref.[126]
- Dirichlet Distribution, Ref.[95]
- Multivariate Normal Distribution, Ref.[128]
- Beta function, Ref.[83]
- Multinomial Coefficients, Ref.[127]
- Gamma Function Ref.[102]

Here are a few results from those Wikipedia articles that we will use later on in this chapter.

Below, we will abbreviate $q_+ = \sum_i q_i$, and $q_\cdot = (q_0, q_1, \dots, q_{nq-1})$ for various quantities q

Gamma function. If $n > 0$ is an integer,

$$\Gamma(n+1) = n! \quad (73.7)$$

The **multivariate Beta function** is defined by

$$B(\alpha_\cdot) = \frac{\prod_k \Gamma(\alpha_k)}{\Gamma(\alpha_+)} \quad (73.8)$$

where $\alpha_k > 0$ for all k .

The **multinomial coefficient** is defined by

$$C(N_\cdot) = \frac{N_+!}{\prod_k N_k!} \quad (73.9)$$

where N_k are non-negative integers.

The **inverse of the multinomial coefficient** will be denoted by

$$CI(N_\cdot) = \frac{1}{C(N_\cdot)} = \frac{\prod_k N_k!}{N_+!} \quad (73.10)$$

The **Categorical Distribution** is defined by

$$Cat(x; \pi.) = \pi_x = \prod_k \pi_k^{\mathbb{1}(k=x)} \quad (73.11)$$

for $k, x \in S_x$, where $\pi.$ is a probability dist.(i.e., $\pi_k \geq 0$ for all k , and $\pi_+ = 1$).

The **Multinomial Distribution** is defined by

$$Mul(N.; \pi., N) = C(N.) , \quad (73.12)$$

where N_k is a non-negative integer for all k , $N_+ = N$, and $\pi.$ is a probability dist. $Mul()$ satisfies:

$$E[N_k] = N\pi_k . \quad (73.13)$$

The **Dirichlet Distribution** is defined by

$$Dir(\pi.; \alpha.) = \frac{1}{B(\alpha.)} \prod_k \pi_k^{\alpha_k - 1} \quad (73.14)$$

where $\alpha_k > 0$ for all k , and $\pi.$ is a probability dist. The $\alpha.$ are called **concentration parameters** or **hyperparameters**. $Dir()$ satisfies:

$$E[\pi_k] = \frac{N_k}{N_+} . \quad (73.15)$$

$Dir()$ is **conjugate prior** of $Mul()$

Note that

$$Mul(N.; \pi., N)Dir(\pi.; \alpha.) = \mathcal{K}(N., \alpha.)Dir(\pi.; N. + \alpha.) , \quad (73.16)$$

where

$$\mathcal{K}(N., \alpha.) = \frac{B(N. + \alpha.)}{CI(N.)B(\alpha.)} . \quad (73.17)$$

$Dir()$ is replaceable by a $Mul()$ for large concentration parameters

Note that if N_k is a positive integer and $\alpha_k = N_k + 1$ for all k , then

$$Dir(\pi.; \alpha_k = N_k + 1) = C(N.) \prod_k \pi_k^{N_k} \quad (73.18)$$

$$= Mul(N.; \pi., N_+) . \quad (73.19)$$

73.2 Single node with no parents

In this section, we consider a learned bnet consisting of a single node with no parents. We will consider arbitrary learned bnets in the next section. But we start with this simplified case so as to reduce the number of indices in most quantities from 3 to 1. All the results that we derive in this section will be used in the next section after adding the extra indices. This way, we will avoid carrying the extra indices throughout the intermediate steps of many derivations.

For state $k \in \{0, 1, \dots, nk - 1\}$ of a single node \underline{x} , let

\underline{N}_k = current count number (an integer, data)

$\underline{\pi}_.$ = a probability dist, the TPM for the node

$\underline{\alpha}_k$ = prior count number

$$\underline{N}_. \longleftrightarrow \underline{\pi}_. \longleftrightarrow \underline{\alpha}_.$$

Figure 73.1: For a bnet consisting of a single node with no parents, this is a Markov chain of current counts ($\underline{N}_.$), TPM ($\underline{\pi}_.$), and prior counts ($\underline{\alpha}_.$).

Consider the Markov chain bnet of Fig.73.1 with the following TPMs, printed in blue.

$$P(N_+ | \underline{\pi}_.) = \text{Mul}(N_+; \underline{\pi}_., N_+) \quad (73.20)$$

$$P(\underline{\pi}_+ | \underline{\alpha}_.) = \text{Dir}(\underline{\pi}_+; \underline{\alpha}_.) \quad (73.21)$$

It follows that

$$P(N_+, \underline{\pi}_+ | \underline{\alpha}_.) = P(N_+ | \underline{\pi}_.) P(\underline{\pi}_+ | \underline{\alpha}_.) \quad (73.22)$$

$$= \text{Mul}(N_+; \underline{\pi}_., N_+) \text{Dir}(\underline{\pi}_+; \underline{\alpha}_.) \quad (73.23)$$

$$= \mathcal{K}(N_+, \underline{\alpha}_.) \text{Dir}(\underline{\pi}_+; N_+ + \underline{\alpha}_.). \quad (73.24)$$

From Eq.(73.15) for the expected value of $\text{Dir}()$, we get

$$\hat{\pi}_+ = E[\underline{\pi}_+] = \frac{N_+ + \underline{\alpha}_+}{N_+ + \underline{\alpha}_+}. \quad (73.25)$$

Integrating both sides of Eq.(73.24) over $\underline{\pi}_.$, we find that

$$P(N_+ | \underline{\alpha}_.) = \mathcal{K}(N_+, \underline{\alpha}_.). \quad (73.26)$$

If $N_k \gg 1$ for all k , then the $\text{Dir}()$ in Eq.(73.24) can be replaced by a $\text{Mul}()$

$$P(N., \pi.| \alpha.) \approx \mathcal{K}(N., \alpha.) Mul(N. + \alpha.; \pi., N_+ + \alpha_+) . \quad (73.27)$$

Therefore,

$$P(N.| \pi., \alpha.) = \frac{P(N., \pi.| \alpha.)}{P(N.| \alpha.)} \quad (73.28)$$

$$= Mul(N. + \alpha.; \pi., N_+ + \alpha_+) . \quad (73.29)$$

Claim 116

$$\ln P(N.| \widehat{\pi}_., \alpha.) = -(N_+ + \alpha_+) H\left(\frac{N. + \alpha.}{N_+ + \alpha_+}\right) + \ln C(N. + \alpha.) \quad (73.30)$$

$$> -(N_+ + \alpha_+) H\left(\frac{N. + \alpha.}{N_+ + \alpha_+}\right) - \frac{1}{2}(nk - 1) \ln N_+ \quad (73.31)$$

proof:

$$\ln P(N.| \widehat{\pi}_., \alpha.) = \sum_k (N_k + \alpha_k) \ln \widehat{\pi}_k + \ln C(N. + \alpha.) \quad (73.32)$$

$$= \sum_k (N_k + \alpha_k) \ln \frac{N_k + \alpha_k}{N_+ + \alpha_+} + \ln C(N. + \alpha.) \quad (73.33)$$

$$= -(N_+ + \alpha_+) H\left(\frac{N. + \alpha.}{N_+ + \alpha_+}\right) + \ln C(N. + \alpha.) \quad (73.34)$$

Recall Stirling's approximation of a factorial, valid for large integers n :

$$\ln n! \approx (n + \frac{1}{2}) \ln n - n . \quad (73.35)$$

Assume $N_k \gg 1$ for all k . Applying Stirling's approximation to all factorials in $C(N)$, we get

$$\ln C(N.) \approx (N_+ + \frac{1}{2}) \ln N_+ - N_+ - \sum_k \left[(N_k + \frac{1}{2}) \ln N_k - N_k \right] \quad (73.36)$$

$$= (N_+ + \frac{1}{2}) \ln N_+ - \sum_k (N_k + \frac{1}{2}) \ln N_k . \quad (73.37)$$

Next assume that

$$N_k \approx \frac{N_+}{nk} . \quad (73.38)$$

Then

$$\ln C(N.) = (N_+ + \frac{1}{2}) \ln N_+ - nk(\frac{N_+}{nk} + \frac{1}{2})[\ln N_+ - \ln nk] \quad (73.39)$$

$$= -\frac{1}{2}(nk - 1) \ln N_+ + (N_+ + \frac{nk}{2}) \ln nk \quad (73.40)$$

$$> -\frac{1}{2}(nk - 1) \ln N_+ . \quad (73.41)$$

QED

73.3 Multiple nodes with any number of parents

In the previous section, we considered a bnet consisting of a single node with no parents, so we only needed a single index k for the states of the single node. In this section, we consider an arbitrary bnet with multiple nodes each of which may have multiple parents. Most of the results in the previous section are valid for the general case if we make the following replacements: $\pi.$ $\rightarrow \underline{\pi^i}_{|\mu}$ $N.$ $\rightarrow \underline{N^i}_{,\mu}$ $\alpha.$ $\rightarrow \underline{\alpha^i}_{,\mu}$. Upon this replacement, Fig.73.1 becomes Fig.73.2. The TPMs, printed in blue, of the new Markov chain, are as follows:

$$\underline{N^i}_{,\mu} \longleftarrow \underline{\pi^i}_{|\mu} \longleftarrow \underline{\alpha^i}_{,\mu}$$

Figure 73.2: Generalization of Fig.73.1. For a bnet with multiple nodes each of which may have multiple parents, this is a Markov chain of current counts ($\underline{N^i}_{,\mu}$), TPM ($\underline{\pi^i}_{|\mu}$), and prior counts ($\underline{\alpha^i}_{,\mu}$) .

$$P(\underline{N^i}_{,\mu} | \underline{\pi^i}_{|\mu}) = \text{Mul}(\underline{N^i}_{,\mu}; \underline{\pi^i}_{|\mu}, \underline{N^i}_{+, \mu}) \quad (73.42)$$

$$P(\underline{\pi^i}_{|\mu} | \underline{\alpha^i}_{,\mu}) = \text{Dir}(\underline{\pi^i}_{|\mu}; \underline{\alpha^i}_{,\mu}) \quad (73.43)$$

In these TPMs,

$i \in S_i = \{0, 1, \dots, ni - 1\}$ = node index

$\underline{x}.$ = $(\underline{x}_i)_{i \in S_i}$ = the nodes of the learned bnet.

$k \in S_k^i = \{0, 1, \dots, nk^i - 1\}$ = states of node x_i

$\mu \in S_\mu^i = \{0, 1, \dots, n\mu^i - 1\}$ = states of parents of node \underline{x}_i .

In the previous section, we assumed a single node ($ni = 1$) with no parents ($n\mu^0 = 1$) so that we could drop the i, μ indices. In this section, we eliminate that restriction.

It is convenient to define the magnitude of a bnet B to equal the sum over nodes of the number of free parameters in each TPM:

$$|B| = \sum_i (nk^i - 1)n\mu^i . \quad (73.44)$$

Suppose that we are given $nsam$ samples $\vec{x}_i = (\underline{x}_i^\sigma)_{\sigma=0,1,\dots,nsam-1}$ of our learned bnet. The count numbers $N_{k,\mu}^i$ are defined in terms of those samples as follows:

$$N_{k,\mu}^i = \sum_\sigma \mathbb{1}(\underline{x}_i^\sigma = k, pa(\underline{x}_i^\sigma) = \mu) . \quad (73.45)$$

It is also convenient to defined count number ratios

$$N_{k|\mu}^i = \frac{N_{k,\mu}^i}{N_{+, \mu}^i} . \quad (73.46)$$

Note that $N_{k,\mu}^i$ is a positive integer whereas $N_{k|\mu}^i \in [0, 1]$.

Let's denote the components of the TPMs by $\pi_{k|\mu}^i$:

$$\pi_{k|\mu}^i = P(\underline{x}_i = k \mid pa(\underline{x}_i) = \mu) \approx N_{k|\mu}^i . \quad (73.47)$$

The rest of this section lists equations that we obtained from the previous section, by adding the new indices i, μ :

$$\mathcal{K}(N_{\cdot, \mu}^i, \alpha_{\cdot, \mu}^i) = \frac{B(N_{\cdot, \mu}^i + \alpha_{\cdot, \mu}^i)}{CI(N_{\cdot, \mu}^i)B(\alpha_{\cdot, \mu}^i)} \quad (73.48)$$

$$\widehat{\pi}_{k|\mu}^i = \frac{N_{k,\mu}^i + \alpha_{k,\mu}^i}{N_{+, \mu}^i + \alpha_{+, \mu}^i} \quad (73.49)$$

$$P(N_{\cdot, \mu}^i | \alpha_{\cdot, \mu}^i) = \mathcal{K}(N_{\cdot, \mu}^i, \alpha_{\cdot, \mu}^i) \quad (73.50)$$

$$P(N_{\cdot, \mu}^i | \pi_{\cdot|\mu}^i, \alpha_{\cdot, \mu}^i) \approx Mul(N_{\cdot, \mu}^i + \alpha_{\cdot, \mu}^i; \pi_{\cdot|\mu}^i, N_{+, \mu}^i + \alpha_{+, \mu}^i) \quad (73.51)$$

Claim 117

$$\ln P(N_{\cdot, \mu}^i | \widehat{\pi}_{\cdot|\mu}^i, \alpha_{\cdot, \mu}^i) = \sum_k (N_{k,\mu}^i + \alpha_{k,\mu}^i) \ln \left(\frac{N_{k,\mu}^i + \alpha_{k,\mu}^i}{N_{+, \mu}^i + \alpha_{+, \mu}^i} \right) + \ln C(N_{\cdot, \mu}^i + \alpha_{\cdot, \mu}^i) \quad (73.52)$$

$$> \sum_k (N_{k,\mu}^i + \alpha_{k,\mu}^i) \ln \left(\frac{N_{k,\mu}^i + \alpha_{k,\mu}^i}{N_{+, \mu}^i + \alpha_{+, \mu}^i} \right) - \frac{1}{2}(nk^i - 1) \ln N_{+, \mu}^i \quad (73.53)$$

73.4 Bayesian Scores

- Bayesian Information Criterion (BIC)

$$\text{BIC-score} = - \sum_i \sum_{k,\mu} N_{k,\mu}^i \ln \left(\frac{N_{k,\mu}^i}{N_{+, \mu}^i} \right) + \underbrace{\left[-\frac{|B|}{2} \ln N_{+,+}^+ \right]}_{\sum_i \sum_\mu \ln C(N_{\cdot,\mu}^i) \text{ would be more accurate}} \quad (73.54)$$

$$\approx \sum_i \sum_\mu \ln P(N_{\cdot,\mu}^i | \widehat{\pi}_{\cdot|\mu}^i, \alpha_{\cdot,\mu}^i = 0) \quad (73.55)$$

- Bayesian Dirichlet (BD)

$$\text{BD-score} = \sum_i \sum_\mu \ln \frac{B(N_{\cdot,\mu}^i + \alpha_{\cdot,\mu}^i)}{B(N_{\cdot,\mu}^i)} \quad (73.56)$$

$$= \sum_i \sum_\mu \ln [CI(N_{\cdot,\mu}^i) P(N_{\cdot,\mu}^i | \alpha_{\cdot,\mu}^i)] \quad (73.57)$$

- BD equivalent (BDe)

$$\text{BDe-score} = \text{BD-score} (\alpha_{k,\mu}^i = \alpha' N_{k,\mu}^i) , \quad (73.58)$$

where α' is a free parameter.

- BD equivalent unified (BDeu)

$$\text{BDeu-score} = \text{BD-score} \left(\alpha_{k,\mu}^i = \frac{\alpha'}{n k^i n \mu^i} \right) , \quad (73.59)$$

where α' is a free parameter. The BDeu score satisfies **score equivalence**; i.e., it is the same for all DAGs in an equivalence class of observational equivalent DAGs. See Chapter 61 for more information about observational equivalence.

73.5 Information Theoretic scores

- Maximum likelihood

$$\text{ML-score} = \sum_i \sum_{k,\mu} N_{k,\mu}^i \ln N_{k|\mu}^i \quad (73.60)$$

$$= - \sum_i H(\underline{k}^i | \underline{\mu}^i) , \quad (73.61)$$

where $P_{\underline{k}^i | \underline{\mu}^i}(k | \mu) = N_{k|\mu}^i$ and $P_{\underline{k}^i, \underline{\mu}^i}(k, \mu) = N_{k,\mu}^i$.

- Bayesian Information Criterion (BIC), a.k.a. Minimum Description Length (MDL)

$$\text{BIC-score} = \text{ML-score} - \frac{|B|}{2} \ln N_{+,+}^+ \quad (73.62)$$

$$\approx \sum_i \sum_{k,\mu} N_{k,\mu}^i \ln \frac{N_{k|\mu}^i}{\sqrt{N_{+,+}^+}} \quad (73.63)$$

- Akaike Information Criterion (AIC)

$$\text{AIC-score} = \text{ML-score} - |B| \quad (73.64)$$

$$\approx \sum_i \sum_{k,\mu} N_{k,\mu}^i [\ln N_{k|\mu}^i - 1] \quad (73.65)$$

Chapter 74

Selection Bias Removal

This chapter is based on Ref.[2].

Selection bias (SB) occurs when one samples from an atypical subset of a population, producing a biased dataset. Are such biased datasets useless? Not necessarily. It is possible to add auxiliary features to the biased dataset, and to sample those new features in an unbiased way, from the whole population. Then one can merge the original biased dataset with the auxiliary, unbiased one, to obtain an enhanced dataset. It is sometimes possible to do this so that the enhanced dataset is provably unbiased. It's like making horizontal the surface of a table that was not initially horizontal. The theory of Bayesian Networks and Causal Inference tells us WHEN this is possible, and HOW to do it when it is possible.

Consider the bnet Fig.74.1.

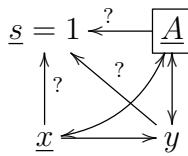


Figure 74.1: Bnet considered for selection bias (SB) removal. Arrows with question marks may or may not be present.

Let

$\underline{s} \in \{0, 1\}$ be the **selection node**. $\underline{s} = 1$ means there is SB in the parent nodes.

\underline{x} = **class features**.¹

\underline{y} = **target feature**.

\underline{A} = **auxiliary features**. This is a set of nodes that may contain arrows entering or exiting it, as indicated by the double arrows.

$\underline{E} = \{\underline{y}, \underline{x}\} \cup \underline{A}$ = **Enhanced feature set**.

¹A feature is the same as a node in a bnet.

Σ = unbiased population of individuals σ

Σ_o = biased sub-population of individuals, $\Sigma_o \subset \Sigma$.

$OD = \{(\sigma_o, \underline{x}^{\sigma_o}, \underline{y}^{\sigma_o}, \underline{s}^{\sigma_o} = 1) : \sigma_o \in \Sigma_o\}$ = **Original Dataset**, dataset for $(\underline{x}, \underline{y})$ features with $\underline{s} = 1$. Gives empirical distribution $P(\underline{y}|\underline{x}, \underline{s} = 1)$. (Ref.[2] calls this dataset the **biased study**.)

$AD = \{(\sigma, \underline{x}^\sigma, \underline{A}^\sigma) : \sigma \in \Sigma\}$ = **Auxiliary Dataset**, dataset for $(\underline{x}, \underline{A})$ features. Gives empirical distribution $P(\underline{A}|\underline{x})$. (Ref.[2] calls this dataset the **population level study**.)

$ED = \{(\sigma_o, \underline{x}^{\sigma_o}, \underline{A}^{\sigma_o}, \underline{y}^{\sigma_o}, \underline{s}^{\sigma_o} = 1) : \sigma_o \in \Sigma_o\}$ = **Enhanced Dataset**, dataset for $(\underline{x}, \underline{y}, \underline{A})$ features for $\underline{s} = 1$. Obtained by merging OD and AD . Gives empirical distribution $P(\underline{y}|\underline{x}, \underline{A}, \underline{s} = 1)$.

74.1 Pre and Post Selection Nodes

Selection nodes can be a source of confusion, so it's worth spending some time thinking about them.

This book uses 2 types of selection nodes: pre-selection nodes and post-selection nodes.

pre-selection nodes are used in Chapter 86 entitled “Transportability of Causal Knowledge”. Suppose that the nodes of a bnet can be either blue or colorless. Then we can define a pre-selection node \underline{s}_{blue} such that $\underline{s}_{blue} = 1$ points to the nodes that are blue, and nothing points to the nodes that are colorless. If we were to change the value of node \underline{s}_{blue} to zero, then all the nodes of the bnet would be colorless. Suppose instead that the nodes of a bnet could be either blue, red or colorless. Then one could define two pre-selection nodes, \underline{s}_{blue} and \underline{s}_{red} , and draw arrows from $\underline{s}_{blue} = 1$ to the blue nodes and from $\underline{s}_{red} = 1$ to the red nodes, and no arrow directed into the colorless nodes. So, basically, pre-selection nodes are binary **root** nodes that indicate whether the nodes pointed to belong to a set or not.

Most of the DAG literature, including Ref.[2], on which this chapter is based, define SB using a **post-selection node**. A post-selection node is a binary **leaf** node of the graph.

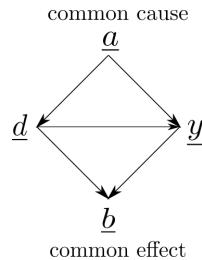


Figure 74.2: Common Cause and Effect for nodes $\underline{d}, \underline{y}$.

Note that in Potential Outcomes (PO) theory (see Chapter 64), pre-selection

nodes such as \underline{a} in Fig.74.2 are called **common cause (confounder, fork) nodes for nodes $\underline{d}, \underline{y}$** . Furthermore, post-selection nodes such as \underline{b} in Fig.74.2 are called **common effect (selection bias (SB), collider) nodes for nodes $\underline{d}, \underline{y}$** . Hence, in PO theory, SB is indicated by a leaf node, just as we do in this chapter.

Note that **Simpson's paradox** (see Chapter 77) arises from an indirect effect caused by not conditioning on a confounder, whereas **Berkson's paradox** (see Chapter 8) arises from an indirect effect caused by conditioning on a collider.

It's possible to replace a pre-selection node by a post selection node, or vice versa, as follows. Suppose that we start with a bnet G_0 that is fully connected, and we add to it a selection node \underline{s} that is a root node that points to all nodes of G_0 . Call the resulting bnet $\underline{s} \rightarrow G_0$. We can use Bayes rule to reverse the direction of the arrows emanating from \underline{s} so that they enter node \underline{s} rather than exit it. Call the resulting bnet $\underline{s} \leftarrow G_0$. In general, Bayes rule allows us to translate from $\underline{s} \rightarrow G_0$ to $\underline{s} \leftarrow G_0$, or in the opposite direction, without having to change the directions of any of the arrows in G_0 . If G_0 is not fully connected, then going from $\underline{s} \rightarrow G_0$ to $\underline{s} \leftarrow G'_0$ will often require that G'_0 have the same arrows in the same directions as G_0 plus some extra arrows new to G_0 . Likewise, going from $\underline{s} \leftarrow G_0$ to $\underline{s} \rightarrow G''_0$ may require that G''_0 have the same arrows as G_0 plus some new arrows.

So far, we have been intentionally vague in specifying the graphs G'_0 and G''_0 . In Fig.74.3 we give a trick for determining possible candidates for graphs G'_0 and G''_0 . In Fig.74.3, we consider 3 panels going from left to right, depicting the cases where \underline{s} has either 1,2 or 3 neighbors. The top graph $G_{post\text{-}sel}$, which has a post-selection node \underline{s} , is converted to a graph which is numerically equal to it, namely the bottom graph $G_{pre\text{-}sel}$, which has a pre-selection node \underline{s} . The magenta arrows represent any number of arrows exiting (but none entering) a node. If we start with a graph $\underline{s} \leftarrow G_0$, we find the biggest subset X of the nodes of G_0 such that $\underline{s} \leftarrow X$ only has nodes exiting it (i.e., only magenta nodes). Then we add enough arrows to $\underline{s} \leftarrow X$ to make it a fully connected graph $\underline{s} \leftarrow X'$. Now we can reverse the incoming arrows to \underline{s} and make them all outgoing and call the resulting graph $\underline{s} \rightarrow X'$.

Recall that in Chapter D, we made a distinction between a good CF bnet a bad CF bnet, and we pointed out that bad CF bnets are often useful as a numerical tool. Recall also from Chapter 61 that two bnets can be “observationally equivalent”. That is what is happening here. We are faced with the choice of making selection nodes either leaf nodes or root nodes. Both choices lead to observationally equivalent bnets. One of the two choices leads to a good CF bnet, and the other to a bad CF bnet. Both choices are numerically correct.

In this chapter, we will consider first a DAG $G_{post\text{-}sel}$ with a post-selection node, then we will transform that DAG to a numerically equal DAG $G_{pre\text{-}sel}$ with a pre-selection node. The latter DAG is more convenient for our needs. Why? Because in the SB theory that we present in this chapter, the $\underline{s} = 1$ appears as a condition in a conditional probability, as in $P(\dots | \dots, \underline{s} = 1)$. Such probabilities are represented in a more straightforward manner if arrows exit rather than enter node \underline{s} .

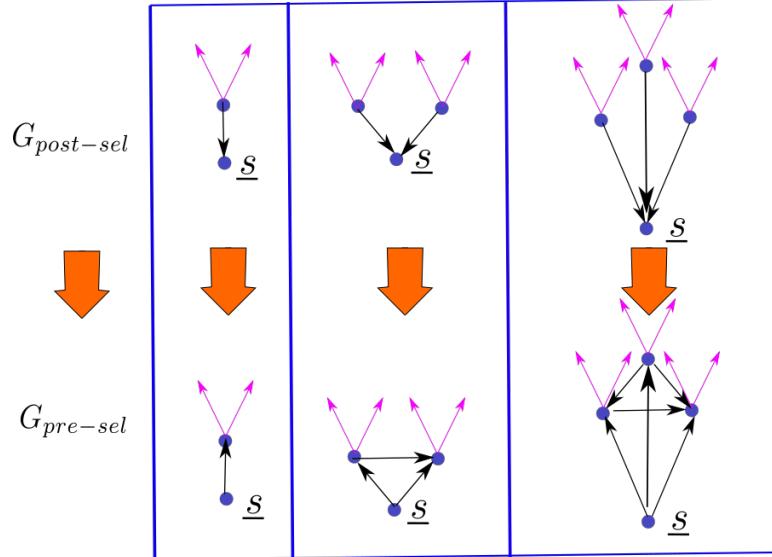


Figure 74.3: Switching from a post-selection node to a pre-selection node.

74.2 Removing SB from passive query $P(y|x)$

The **passive query** $Q = P(y|x, \underline{s} = 1)$ is **SB-recoverable** if it is independent of \underline{s} .

1. **Query $P(y|x)$ is SB-recoverable** with $\underline{A} = \emptyset$; SB can be removed by conditioning on \underline{x} .

If $\underline{y} \perp \underline{s}|\underline{x}$, then

$$P(y|x, \underline{s} = 1) = P(y|x). \quad (74.1)$$

For example,

$$\begin{array}{ccc} \underline{s} & & \underline{s} \\ \uparrow & & \downarrow \\ \underline{x} \longrightarrow \underline{y} & = & \underline{x} \longrightarrow \underline{y} \end{array} \quad (74.2a)$$

$$\begin{array}{ccc} \underline{s} \leftarrow \underline{a} & & \underline{s} \longrightarrow \underline{a} \\ \uparrow & & \downarrow \\ \underline{x} \longrightarrow \underline{y} & = & \underline{x} \longrightarrow \underline{y} \end{array} \quad (74.2b)$$

$$\begin{array}{ccc} \underline{s} \leftarrow \underline{a} & & \underline{s} \longrightarrow \underline{a} \\ \uparrow & & \downarrow \\ \underline{x} \longrightarrow \underline{y} & = & \underline{x} \longrightarrow \underline{y} \end{array} \quad (74.2c)$$

The bnets on the left hand sides of Eqs.(74.2) satisfy $y \perp \underline{s} | \underline{x}$.

2. **Query $P(y|x)$ is SB-recoverable** via \underline{a} ; SB can be removed by conditioning on \underline{x} and \underline{a} . Here \underline{a} is a nonempty subset of \underline{A}

Claim 118 *There exists $\underline{a} \subset \underline{A}$ such that $y \perp \underline{s} | (\underline{x}, \underline{a})$ and $\underline{a} \perp \underline{s} | \underline{x}$ iff*

$$P(y|x, \underline{s} = 1) = \sum_{\underline{a}} \underbrace{P(y|x, a, \underline{s} = 1)}_{P(y|x,a)} \underbrace{P(a|x, \underline{s} = 1)}_{P(a|x)} = P(y|x) \quad (74.3)$$

$$\begin{array}{c} \underline{s} = 1 \\ \searrow \\ x \longrightarrow y \end{array} = \begin{array}{c} \sum a \\ \nearrow \quad \downarrow \\ x \longrightarrow y \end{array} = \begin{array}{c} x \longrightarrow y \end{array} \quad (74.4)$$

proof:

The \Rightarrow part of this claim is obvious. For a proof of the \Leftarrow part, see Ref.[2].
QED

For example,

$$\begin{array}{c} s \\ \uparrow \\ x \longrightarrow y \end{array} = \begin{array}{c} s \\ \downarrow \\ x \longrightarrow y \end{array} \quad (74.5a)$$

The bnets on the left hand sides of Eqs.(74.5) satisfy $y \perp \underline{s} | (x, a)$ and $a \perp \underline{s} | x$.

3. **Query $P(y|x)$ is not SB-recoverable;** SB cannot be removed.

For example,

$$\begin{array}{c} s \\ \swarrow \quad \searrow \\ x \longrightarrow y \end{array} = \begin{array}{c} s \\ \downarrow \\ x \longrightarrow y \end{array} \quad (74.6)$$

$$\begin{array}{c} s \\ \uparrow \quad \searrow \\ x \longrightarrow y \end{array} = \begin{array}{c} s \\ \downarrow \\ x \longrightarrow y \end{array} \quad (74.7)$$

$$\begin{array}{ccc}
 \begin{array}{c} \underline{s} \leftarrow \underline{a} \\ \uparrow \\ \underline{x} \xrightarrow{\quad} \underline{y} \\ \uparrow \\ \underline{w} \end{array} & = & \begin{array}{c} \underline{s} \longrightarrow \underline{a} \\ \downarrow \\ \underline{x} \xrightarrow{\quad} \underline{y} \\ \uparrow \\ \underline{w} \end{array}
 \end{array} \tag{74.8}$$

74.3 Removing SB from active query $P(y|\mathcal{D}\underline{x})$

The **active query** (i.e., **do query**) $Q = P(y|\mathcal{D}\underline{x} = \underline{x}, \underline{s} = 1)$ is

- (a) **SB-recoverable** if it equals $P(y|\mathcal{D}\underline{x} = \underline{x})$,
- (b) **do-identifiable** if it equals $P(y|\underline{x}, \underline{s} = 1)$.
- (c) both **SB-recoverable and do-identifiable** if it equals $P(y|\underline{x})$.

Examples

- $$\begin{array}{ccc}
 \begin{array}{c} \underline{s} \leftarrow \underline{a} \\ \downarrow \\ \underline{x} \xrightarrow{\quad} \underline{y} \end{array} & = & \begin{array}{c} \underline{s} \longrightarrow \underline{a} \\ \downarrow \\ \underline{x} \xrightarrow{\quad} \underline{y} \end{array}
 \end{array} \tag{74.9}$$

SB-recoverable: NO
do-identifiable: YES

$Q = P(y|\mathcal{D}\underline{x} = \underline{x}, \underline{s} = 1)$ is do-identifiable because the bnet contains no hidden variables. It's s-recoverable because the bnet on the left hand side of Eq.(74.9) satisfies $\underline{y} \perp \underline{s} | (\underline{x}, \underline{a})$ but Not $\underline{a} \perp \underline{s} | \underline{x}$.

- $$\begin{array}{ccc}
 \begin{array}{c} \underline{s} \\ \uparrow \\ \underline{x} \xrightarrow{\quad} \underline{y} \end{array} & = & \begin{array}{c} \underline{s} \\ \downarrow \\ \underline{x} \xrightarrow{\quad} \underline{y} \end{array}
 \end{array} \tag{74.10}$$

SB-recoverable: YES
do-identifiable: NO

Let

V = set of nodes in graph

$V^{<\underline{x}}$ = non-descendants of \underline{x} (excluding \underline{x})

$V^{>\underline{x}}$ = descendants of \underline{x} (excluding \underline{x})

$\underline{z}^{<\underline{x}} = \underline{z} \cap V^{<\underline{x}}$

$\underline{z}^{>\underline{x}} = \underline{z} \cap V^{>\underline{x}}$

Suppose $\underline{z} \cup \{\underline{x}, \underline{y}\} \subset E$ and $\underline{z} \subset \underline{A}$. We say \underline{z} satisfies the **selection bias (SB) backdoor criterion** with respect to $(\underline{x}, \underline{y})$ if

1. all backdoor paths from \underline{x} to \underline{y} are blocked by conditioning on $\underline{z}^{<\underline{x}}$
2. $\underline{z}^{>\underline{x}} \perp \underline{y} | (\underline{x}, \underline{z}^{<\underline{x}})$
3. $\underline{y} \perp \underline{s} | (\underline{x}, \underline{z})$

Claim 119 (*SB Backdoor Adjustment Formula*)

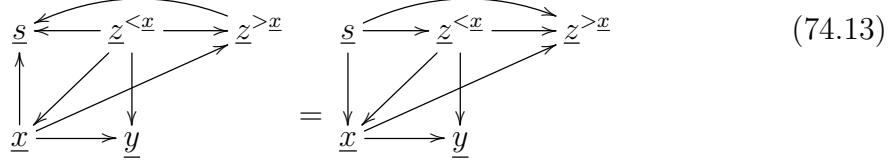
If \underline{z} satisfies the SB backdoor criterion relative to $(\underline{x}, \underline{y})$, then

$$P(y | \mathcal{D}\underline{x} = x, \underline{s} = 1) = \sum_z P(y|x, z)P(z) = P(y|x) \quad (74.11)$$

$$\begin{array}{c} \underline{s} = 1 \\ \searrow \\ \mathcal{D}\underline{x} = x \longrightarrow y \end{array} = \sum_z \begin{array}{c} z \\ \downarrow \\ x \longrightarrow y \end{array} = x \longrightarrow y \quad (74.12)$$

proof:

If z satisfies the SB backdoor criterion relative to $(\underline{x}, \underline{y})$, then $\underline{x}, \underline{y}, \underline{z}$ might have the following structure.



See Claim 48 for a proof of this claim for the special case Eq.(74.13).

QED

Chapter 75

Shannon Information Theory

Throughout this book, we often use the definitions of entropy, mutual information, conditional mutual information, cross entropy, etc. All these definitions originated from the seminal work of Claude Shannon. Note that the connection between bnets and Shannon Information Theory (SIT) goes much deeper than that. Most of SIT can be expressed graphically using bnets, as we explain in Ref.[70].

Chapter 76

Shapley Explainability

This chapter is based on Refs.[32] and [33], which I highly recommended.

“AI” is an ill-defined term. So is the term “explainability”. So the term “Explainable AI (XAI)” is doubly ill-defined. In 2018, the European Union codified the need for XAI — because of an individual’s “right to explanation” — into a law called the General Data Protection Right (GDPR). This EU law was a strong motivation for Neural Net and boosted decision tree practitioners to come up with a way to enhance their machine learning algorithms so that these comply with that law. Shapley explainability (SX) is one of the most popular methods for doing XAI.

So what does SX do? It ranks, for each individual of a population, the features (for example, race) of a dataset, in the order of how influential those features were in arriving at the decision the classifier made for that individual.

In my opinion, the goal of XAI is accomplished better with bnets than with SX enhanced NNs. SX “explains”, *a posteriori*, the outcome of a model, whereas bnets reveal the *a priori* process whereby that outcome was reached. Thus, SX can tell you that a model is racist, but it can’t suggest how to fix it. On the other hand, if a bnet is acting racist, you don’t have to throw it away. It can be fixed. Another weakness of SX is that it is quite expensive computationally. Bnets have explainability built into them. For bnets, explainability is not an additional, posterior and quite onerous, calculation. That is why I like to call bnets the gold standard of XAI.

Let

F be the feature set. For example, $F = \{age, gender, job\}$.

$\mathcal{P}(A) = \{S : S \subset A\}$ be the power set of the set A (*i.e.*, the set of all subsets of A , including the empty set \emptyset).

$$|\mathcal{P}(A)| = \sum_{k=0}^{|A|} \binom{|A|}{k} = \sum_{k=0}^{|A|} \binom{|A|}{k} 1^k 1^{|A|-k} = (1+1)^{|A|} = 2^{|A|}$$

$\mathcal{P}_f(F) = \{S \in \mathcal{P}(F) : f \in S\}$ for $f \in F$, be all sets in $\mathcal{P}(F)$ containing feature f . Note that $\mathcal{P}_f(F) = \mathcal{P}(F) - \mathcal{P}(F - \{f\})$

$\mathcal{P}_{!f}(F) = \{S \in \mathcal{P}(F) : f \notin S\}$ for $f \in F$, be all sets in $\mathcal{P}(F)$ not containing feature f . Note that $\mathcal{P}_{!f}(F) = \mathcal{P}(F - \{f\})$

Fig.76.1 shows a graph of $\mathcal{P}(F)$ for $F = \{age, gender, job\}$. Henceforth, we will refer to the generalization of Fig.76.1 to an arbitrary finite set F , as the **power**

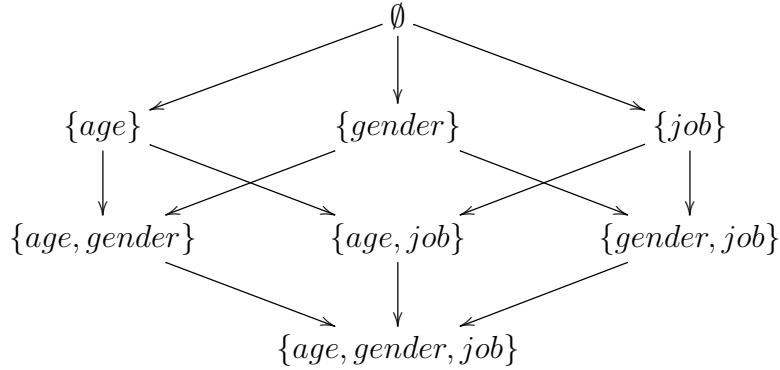


Figure 76.1: Graph of $\mathcal{P}(F)$ for $F = \{\text{age}, \text{gender}, \text{job}\}$. An arrow $H \leftarrow T$, where $H \in \mathcal{P}(F)$ is the head of the arrow and $T \in \mathcal{P}(F)$ is the tail of the arrow, means $H \supset T$ and $|H| = |T| + 1$.

set graph of F .¹

For any finite set F , consider its power set graph. Let $S \in \mathcal{P}(F)$ be a node of that graph. Let

$$N_{arr/nd}(S) = |S| = \text{number of arrows entering node } S.$$

$$N_{nds}(|S|) = \binom{|F|}{|S|} = \text{number of nodes in level } |S| \text{ (i.e., row } |S|).$$

$N_{arr}(S) = N_{arr/nd}(S)N_{nds}(|S|) = |S| \binom{|F|}{|S|} = \text{number of arrows going from row } |S| - 1 \text{ to row } |S|.$

$$P(S) = \frac{1}{N_{arr}(S)} = \frac{1}{|S| \binom{|F|}{|S|}} = \frac{1}{|F| \binom{|F|-1}{|S|-1}}. \quad (76.1)$$

Claim 120

$$\sum_{S \in \mathcal{P}_f(F)} P(S) = 1 \quad (76.2)$$

proof:

¹Note that a power set graph is a DAG. We won't define TPMs for its nodes in this chapter, so it's a DAG but not a bnet, in this chapter at least.

$$\sum_{S \in \mathcal{P}_f(F)} P(S) = \sum_{S \in \mathcal{P}_f(F)} \frac{1}{|F| \binom{|F|-1}{|S|-1}} \quad (76.3)$$

$$= \sum_{k=1}^{|F|} \sum_{S \in \mathcal{P}_f(F)} \mathbb{1}(|S|=k) \frac{1}{|F| \binom{|F|-1}{k-1}} \quad (76.4)$$

$$= \sum_{k=1}^{|F|} \frac{1}{|F| \binom{|F|-1}{k-1}} \underbrace{\sum_{S \in \mathcal{P}_f(F)} \mathbb{1}(|S|=k)}_{\binom{|F|-1}{k-1}} \quad (76.5)$$

$$= 1 \quad (76.6)$$

QED

Consider any $S \in \mathcal{P}(F)$. Henceforth, we will represent a Machine Learning (ML) model ML_S as follows. ML_S can be a Linear Regression (LR_S) model, a Neural Net model (NN_S), or any other type of ML model. We will list a dataset; i.e., a set of tuples indexed by the individuals σ of a population Σ such that $|\Sigma| = nsam$. The independent variables of ML_S (i.e., $x_S^\sigma = [x_f^\sigma : f \in S]$) will be shown unboxed and the dependent variable (a.k.a. target feature) (i.e., y^σ) will be shown inside a box. Then we will show an arrow with the superscript “ML-fit”, followed by the fit function obtained by performing ML_S .

$$ML_S : \{(\sigma, x_S^\sigma, \boxed{y^\sigma}) : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \hat{y}(x_S) \quad (76.7)$$

For any feature $f \in F$ and individual $\sigma \in \Sigma$, define the **Shapley Value** (SHAP) by

$$SHAP_f^\sigma = \sum_{S \in \mathcal{P}_f(F)} P(S) [\hat{y}(x_S^\sigma) - \hat{y}(x_{S-\{f\}}^\sigma)] \quad (76.8)$$

$$= E_S[\hat{y}(x_S^\sigma) - \hat{y}(x_{S-\{f\}}^\sigma)] \quad (76.9)$$

Hence,

- $SHAP_f^\sigma$ is an average over the ensemble $\{ML_S : S \in \mathcal{P}_f(F)\}$ for each individual $\sigma \in \Sigma$.
- $SHAP_f^\sigma$ averages the change in output \hat{y} when we change the model from one without feature f to one with feature f .
- $SHAP_f^\sigma$ can be negative or positive. Zero $SHAP_f^\sigma$ for individual σ means feature f does not influence how the decision \hat{y} was arrived at for individual σ .

- An exact calculation of $SHAP_f^\sigma$, for all f and for a single σ , requires training a different ML model for each node of the power set graph of F , so it requires training $2^{|F|}$ models. Yikes! As $|F|$ grows, this quickly becomes unfeasible. For large $|F|$, one must resort to using sampling and approximations to get an approximation of the SHAP.

76.0.1 Numerical examples of SHAP

Next we present 2 numerical examples of SHAP. The figures and numerical values in this section were taken directly from Refs. [32] and [33].

1. Predicting Income from $F = \{age, gender, job\}$.

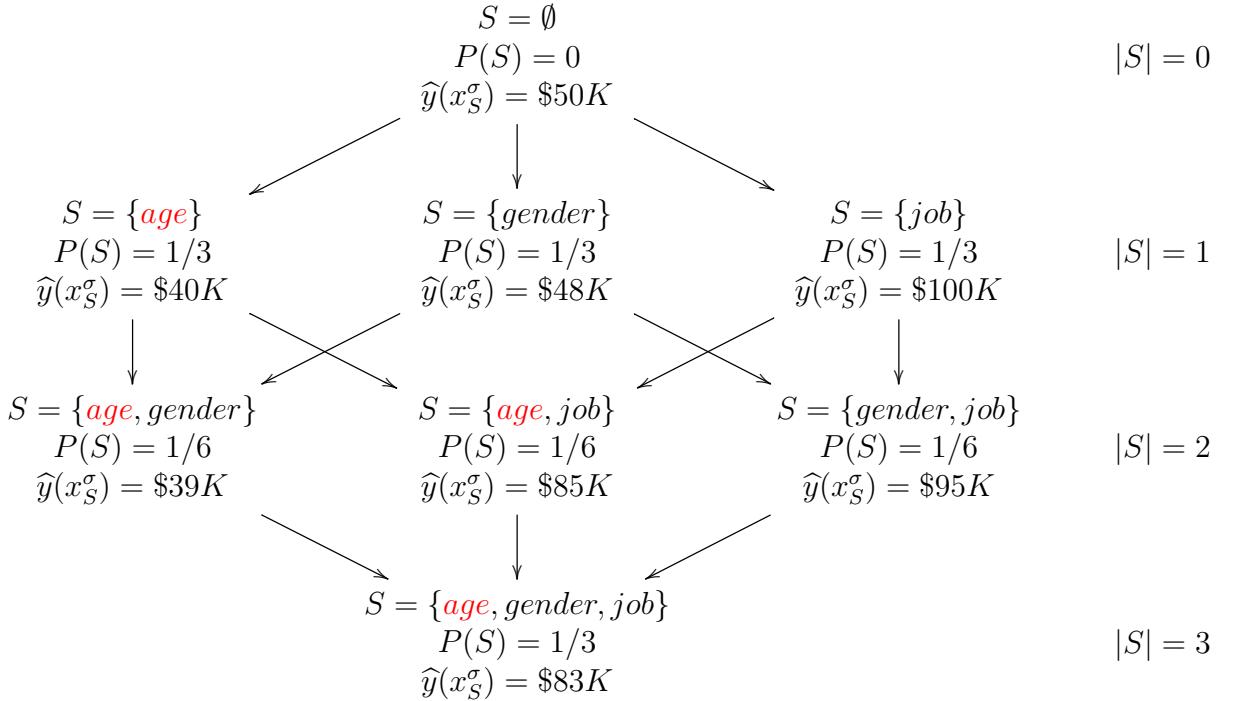


Figure 76.2: Same as Fig.76.1, but with added information for a specific individual σ . This figure contains enough information to evaluate $SHAP_{age}^\sigma$.

Consider the problem of predicting the income of a person based on the feature set $F = \{age, gender, job\}$. Suppose we are given a dataset for this problem. We can train models ML_S for each $S \in \mathcal{P}_{age}(F)$ where \hat{y} is the income. Then we can calculate the matrix $SHAP_{age}^\sigma$. Fig.76.2 gives all the information necessary to calculate $SHAP_{age}^\sigma$ for a single individual σ .

$$P(S) = \frac{1}{3^{\binom{|S|}{2}}} = \begin{cases} \frac{1}{3^{\binom{2}{0}}} = \frac{1}{3} & \text{if } |S| = 1 \\ \frac{1}{3^{\binom{2}{1}}} = \frac{1}{6} & \text{if } |S| = 2 \\ \frac{1}{3^{\binom{2}{2}}} = \frac{1}{3} & \text{if } |S| = 3 \end{cases} \quad (76.10)$$

$$SHAP_{age}^\sigma = \underbrace{P(age)}_{1/3} [\widehat{y}(x_{age}^\sigma) - \widehat{y}(x_\emptyset^\sigma)] \quad (76.11)$$

$$+ \underbrace{P(age, job)}_{1/6} [\widehat{y}(x_{age, job}^\sigma) - \widehat{y}(x_{job}^\sigma)] \quad (76.12)$$

$$+ \underbrace{P(age, gender)}_{1/6} [\widehat{y}(x_{age, gender}^\sigma) - \widehat{y}(x_{gender}^\sigma)] \quad (76.13)$$

$$+ \underbrace{P(age, gender, job)}_{1/3} [\widehat{y}(x_{age, gender, job}^\sigma) - \widehat{y}(x_{gender, job}^\sigma)] \quad (76.14)$$

$$= -\$11.33K \quad (76.15)$$

2. Predicting passenger survival in the Titanic disaster.

	Pclass	Sex	Age	SibSp	Parch	Fare	Embarked
PassengerId							
1	3	male	22.0	1	0	7.2500	S
2	1	female	38.0	1	0	71.2833	C
3	3	female	26.0	0	0	7.9250	S
4	1	female	35.0	1	0	53.1000	S
5	3	male	35.0	0	0	8.0500	S

Figure 76.3: First five rows (passengers) of an abridged version of the Titanic Dataset available at kaggle.com. This figure shows (σ, x_F^σ) for $\sigma = 1, 2, \dots, 5$. It doesn't show the column $y^\sigma \in \{died, survived\}$.

Consider the problem of predicting whether an individual will survive or not based on a Titanic Dataset. We can train models ML_S for each $S \in \mathcal{P}(F)$ where $\widehat{y} \in \{died, survived\}$. Then we can calculate the matrix $SHAP_f^\sigma$ for all individuals σ and features f . Fig.76.0.1 shows the first 5 rows of an abridged² version of the Titanic Dataset available at kaggle.com. Fig.76.4 displays $SHAP_f^\sigma$ in tabular form and Fig.76.5 displays $SHAP_f^\sigma$ in graphical form (in what is called a beeswarm plot).

²The Titanic Dataset available at kaggle.com has 891 rows and 15 columns, including columns for passenger ID and for $y^\sigma = survived? \in \{0, 1\}$. This abridged version has 8 columns.

	Pclass	Sex	Age	SibSp	Parch	Fare	Embarked
PassengerId							
1	-0.36	-0.76	0.11	0.05	-0.03	-0.3	-0.08
2	1.22	2.18	0.1	0.06	0.1	0.78	0.42
3	-0.75	1.62	0.07	0.1	-0.02	-0.01	-0.13
4	1.15	2.07	0.11	0.05	0.08	0.89	-0.13
5	-0.41	-0.73	-0.06	0.08	-0.03	-0.15	-0.08

Figure 76.4: For the Titanic dataset, this is a table of $SHAP_f^\sigma$, where $\sigma \in \{1, 2, \dots, 5\}$ and $f \in F$. Cells with positive SHAP are colored green, and those with negative SHAP are colored red. The colors are not an indication of whether the passenger died or survived. Note that the table of a dataset and the matrix $SHAP_f^\sigma$ have the same shape ($|\Sigma|, |F|$).

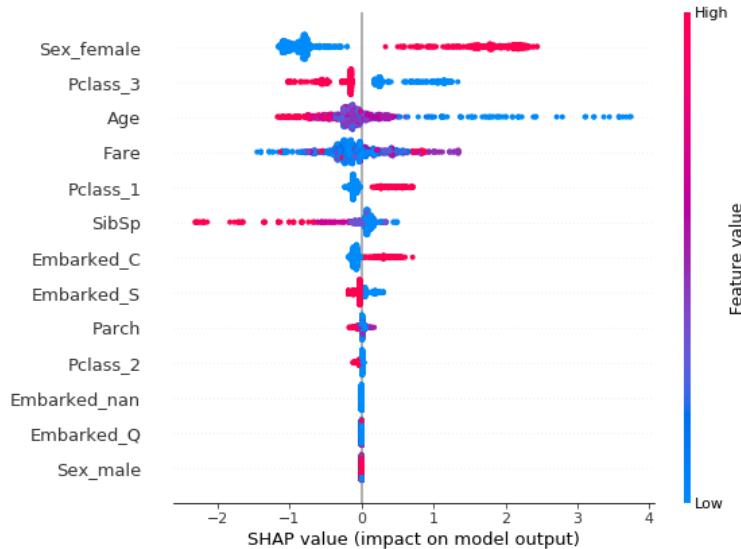


Figure 76.5: For the Titanic Dataset, this is a so called “beeswarm” plot of $SHAP_f^\sigma$, where $\sigma \in \{1, 2, \dots, 891\}$ and $f \in F$. In a beeswarm plot, the thickness of each row is proportional to how many individuals of the population have that value of the x coordinate. This plot comes from Ref.[32], where it was generated using the Titanic Dataset from kaggle.com and the wonderful Python library “SHAP”. The SHAP library can plot Shapley Values in many other styles besides this one.

Chapter 77

Simpson's Paradox

This chapter is based on Chapter 6 of “The Book of Why”, Ref.[50]. See also Ref.[138] and references therein.

Simpson’s paradox is a recurring nightmare for all statisticians overseeing a clinical trial for a medicine. It is possible that if they leave out a certain ”confounding” variable from a study, the study’s conclusion on whether a medicine is effective or not, might be, without measuring that confounding variable, the opposite of what it would have been had that variable been measured.

Simpson’s Paradox is greatly clarified by Judea Pearl’s theory of causality. At the end of this chapter, we explain how.

Here is a simple example of Simpson’s Paradox.

An equal number of patients of male and female genders are given a heart medicine or a placebo in a double blind study. Some subsequently have a heart attack. Let

\underline{a} = heart attack? No=0, Yes=1

\underline{t} = took medicine? No=0, Yes=1

\underline{g} = gender? Female=F, Male=M

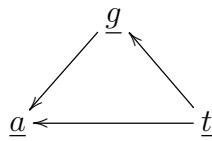


Figure 77.1: bnet for a simple example of Simpson’s paradox. Here node \underline{g} is a chain junction and a mediator.

This situation can be modeled by either bnet Fig.77.1. or bnet Fig.77.2. The two bnets are probabilistically equivalent (i.e., they both represent the same probability distribution $P(a, t, g)$) because

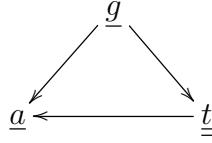


Figure 77.2: bnet that is probabilistically but not physically equivalent to bnet Fig.77.1. Here node \underline{g} is a fork junction and a confounder.

$$P(g|t)P(t) = P(g,t) = P(t|g)P(g) . \quad (77.1)$$

For the bnet Fig.77.1, one has

$$P(a,g,t) = P(a|g,t)P(g|t)P(t) . \quad (77.2)$$

Therefore,

$$P(a = 1|t) = \sum_g P(a = 1|t, g)P(g|t) = E_{g|t}P(a = 1|t, \underline{g}) , \quad (77.3)$$

where $E_{g|t}$ is a conditional expected value (a kind of weighted average).

Suppose q_0, q_1 are non-negative real numbers. For the vector $\vec{q} = (q_0, q_1)$:

Define a negative outcome (or failure or q_t increasing with t) if $q_0 \leq q_1$.

Define a positive outcome (or success or q_t decreasing with t) if $q_0 \geq q_1$.

Let

$$\vec{q}^g = [P(\underline{a} = 1|t, g)]_{t=0,1} \quad (77.4)$$

for $g = M, F$, and

$$\vec{q}^* = [P(\underline{a} = 1|t)]_{t=0,1} . \quad (77.5)$$

It is possible (see Fig.77.3 for a graphical explanation of how) to find perverse cases in which $P(a = 1|t, g = M)$ and $P(a = 1|t, g = F)$ increase with t but $P(a = 1|t)$ decreases with t . So it is possible to conclude that the medicine is a failure for each of the two g populations considered separately, yet the medicine is a success when both populations are “amalgamated”. The lesson is that a “trend reversal” is possible upon amalgamation. Trends are not necessarily preserved when we do a weighted average of type $E_{g|t}$. $E_{g|t}$ is an expected value on the random variable \underline{g} conditioned on the root random variable \underline{t} .

So far we have assumed $a \in \{0, 1\}$. Suppose that instead we assume a is a continuous variable taking values in the interval $[0, 1]$. This could reflect a continuum of possible attacks from none to a deadly one. Likewise, suppose the treatment variable t takes on values in the interval $[0, 1]$. This might reflect a continuum of

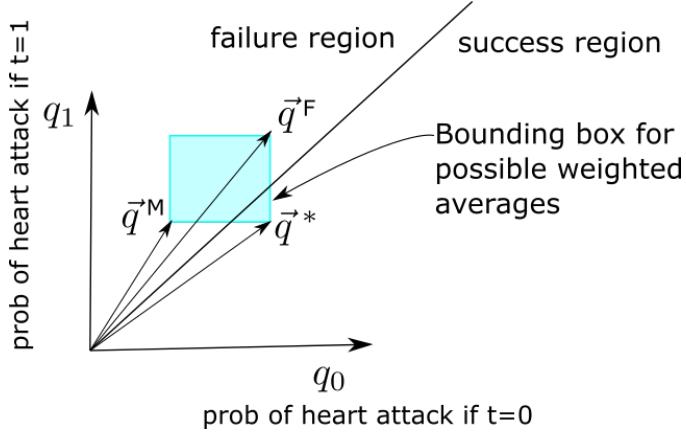


Figure 77.3: \vec{q}^M , \vec{q}^F vectors and bounding box for vector \vec{q}^* .

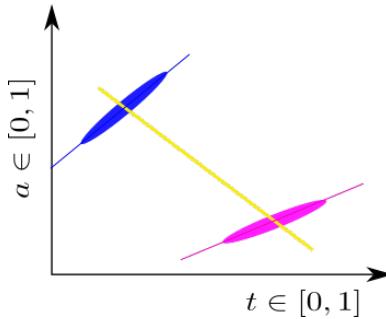


Figure 77.4: Illustrative example of Simpson's paradox, assuming t and a are continuous. The pink (for $g = F$) and blue (for $g = M$) elliptical regions are clouds of sample points. The lines are the result of doing linear regression. For $g = M, F$ separately, increasing treatment increases the probability of attack, but the opposite trend occurs if we amalgamate genders.

possible doses of a medicine. Fig. 77.4 gives an illustrative example of Simpson's paradox for this case of continuous a and t .

So far, we have proven that probabilistically, the drug can be a failure for the populations of both sexes considered separately, but a success for the aggregate population.

77.1 Pearl Causality

Pearl Causality would add the following two important insights to this problem:

1. bnets Fig.77.1 and Fig.77.2, although they are probabilistically equivalent, do not represent the same physical situation. In fact, only Fig.77.2 occurs in this case.

2. To decide whether the medicine is effective, we must apply a $do()$ operator to the \underline{t} variable in Fig.77.2. The effect of that $do()$ operator is to erase the arrow going from g to \underline{t} . This in turn means that the average $E_{g|\underline{t}}$ in our equation for $P(a = 1|\underline{t})$ becomes a simpler average E_g which is independent of \underline{t} . But for such an average, the bounding box in Fig.77.3 degenerates to its diagonal line that connects the tips of the two vectors \vec{q}^M and \vec{q}^F . The vector \vec{q}^* must now fall on that diagonal line and must therefore also fall in the success region.

In conclusion, as Judea Pearl would say, if we ask the right question to Nature, i.e., what is $P[a = 1|do(\underline{t} = t)]$ for $t = 0, 1$, we get as an answer that the aggregate population preserves rather than reverses the unanimous trend of the two gendered populations.

77.2 Numerical Example

(a, t, g)	number of patients segregated by gender	number of patients of either gender
0,0,M	19	47
0,0,F	28	
0,1,M	37	49
0,1,F	12	
1,0,M	1	13
1,0,F	12	
1,1,M	3	11
1,1,F	8	

Table 77.1: Data for numerical example of Simpson’s Paradox. This fictitious data was taken directly from Table 6.4, page 210 of “The Book of Why”, Ref.[50].

$$P(a|t, g) = \begin{array}{c|cccc} & 0,M & 0,F & 1,M & 1,F \\ \hline 0 & 19/20 & 28/40 & 37/40 & 12/20 \\ 1 & 1/20 & 12/40 & 3/40 & 8/20 \end{array} \quad (77.6)$$

$$P(a|t) = \begin{array}{c|cc} & 0 & 1 \\ \hline 0 & 47/60 & 49/60 \\ 1 & 13/60 & 11/60 \end{array} \quad (77.7)$$

$$\frac{P(a=1,t=1,g=M)}{\sum_a P(a,t=1,g=M)} = P(a = 1|t = 1, g = M) = \frac{3}{40} \quad (77.8)$$

$$\frac{P(a=1,t=0,g=M)}{\sum_a P(a,t=0,g=M)} = P(a = 1|t = 0, g = M) = \frac{1}{20} = \frac{2}{40}$$

$$\frac{P(a=1,t=1,g=F)}{\sum_a P(a,t=1,g=F)} = P(a = 1|t = 1, g = F) = \frac{8}{20} = \frac{16}{40} \quad (77.9)$$

$$\frac{P(a=1,t=0,g=F)}{\sum_a P(a,t=0,g=F)} = P(a = 1|t = 0, g = F) = \frac{12}{40}$$

$$\frac{\sum_g P(a=1,t=1,g)}{\sum_g \sum_a P(a,t=1,g)} = P(a = 1|t = 1) = \frac{11}{60} \quad (77.10)$$

$$\frac{\sum_g P(a=1,t=0,g)}{\sum_g \sum_a P(a,t=0,g)} = P(a = 1|t = 0) = \frac{13}{60}$$

Note that the right hand side of Eq.(77.8) is higher for $t = 1$ than for $t = 0$. Same trend occurs in Eqs.77.9 but is reversed in Eqs.77.10.

Chapter 78

Structure and Parameter Learning for Bnets

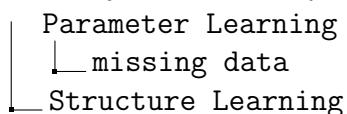
Learning a bnet from data is a computationally intensive NP-complete problem. Therefore, the best one can hope for is for heuristic algorithms that solve this problem approximately. A huge number of such algorithms have been tried and continue to be tried. Luckily, there exists a free open source software library called **bnlearn** that covers many of them. The goal of this chapter is to give a brief overview of the subject of bnet learning, after which we recommend to those readers who want to pursue this subject further, to learn **bnlearn**.

This chapter is based on the **bnlearn** website Ref.[56], and on a 2019 survey paper [57] by Scutari et al. I highly recommend looking at both. Refs. [5] and [30] were also helpful to me in understanding this subject.

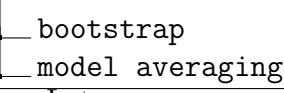
bnlearn (Ref.[56]) (free, open source) is very comprehensive and well maintained. It is written mostly in C with an R front-end. It was developed by Marco Scutari and collaborators over a time period of more than 10 years, and is still under active development. How things stand in the field of bnet learning software reminds me of how things stand in the field of linear algebra (LA) software. Perfecting and optimizing LA software takes many years so I would not advise you to write your own LA software library starting from scratch. There is no need to do so. Instead, you can use LAPACK (free, open source), which has been perfected and expanded for decades by world experts. I view **bnlearn** as the LAPACK of bnet learning.

78.1 Overview

To give the reader an overview of the subject and of **bnlearn** itself, here is a highly simplified tree, compiled from the **bnlearn** website and documentation, of some of the subjects covered by **bnlearn**.



- tree-like structures given a priori
 - Naive Bayes
 - Chow-Liu tree
 - Tree Augmented Naive Bayes (TAN)
 - ARACNE
- score based
 - algorithms
 - hill climbing (HC)
 - HC with random restarts
 - HC with Tabu list (Tabu)
 - simulated annealing
 - genetic algorithms
 - scoring functions
 - Information Theoretic scores
 - Bayesian Information Criterion (BIC)
 - Bayesian Dirichlet (BD) family
- constraint based
 - algorithms
 - PC family
 - Grow-Shrink (GS)
 - Incremental Association Markov Blanket (IAMB) family
 - conditional independence tests
 - mutual information (parametric, semiparametric and permutation tests)
 - shrinkage-estimate for the mutual information
- hybrid
 - Max-Min Hill Climbing (MMHC)
 - Hybrid HPC (H2PC)
 - General 2-Phase Restricted Maximization (RSMAX2)
- parallel mode structure learning
- node types
 - all-discrete
 - all-continuous
 - mixed
- utility functions
 - model comparison and manipulation
 - random data generation
 - arc orientation testing
 - simple and advanced plots
 - parameter estimation (maximum likelihood and Bayesian)
 - inference, conditional probability queries
 - cross-validation



-
- Let
- PL=parameters learning (i.e, learning the TPMs)
 - SL= structure learning (i.e., learning the DAG)
 - ML= model (or bnet) learning, SL+PL

PL is easy, once the structure is known. PL assuming no missing data goes as follows. Using the notation of Chapter 73, define

$$\pi_{k|\mu}^i = P(\underline{x}_i = k \mid pa(\underline{x}_i) = \mu). \quad (78.1)$$

Then $\pi_{k|\mu}^i$ can be estimated from the data $N_{k,\mu}^i$ using:

$$\pi_{k|\mu}^i \approx N_{k|\mu}^i = \frac{N_{k,\mu}^i}{N_{+, \mu}^i}. \quad (78.2)$$

PL described by Eq.(78.2) is only for discrete nodes with no missing data. **bnlearn** can also do PL with missing data and continuous (Gaussian linear only) nodes. See Chapter 53 on missing data and Chapter 30 on Gaussian linear nodes. SL actually does PL and SL at the same time.

There are 3 main types of SL: score based, constraint based, and hybrid. **bnlearn** can perform many algorithms of each of these 3 types of SL. It can perform most of them with either all-discrete, or all-continuous or mixed nodes. It can perform many of them in parallel mode. The 2019 survey paper Ref.[57] by Scutari et al compares the performance of many different bnet learning algorithms.

78.2 Score based SL algorithms

Score based SL algorithms require scoring bnets (with either all-discrete, all-continuous or mixed nodes). See Chapter 73 for an introduction to scoring bnets. The BIC score explained in that chapter is very popular and works for all-discrete, all-continuous or mixed nodes.

Score-based SL algorithms apply standard optimisation techniques. In the Hill Climbing algorithm, the current best bnet is changed slightly and then given a score that measures how well it fits the data. The bnet with the highest (=best) score so far, as well as that highest score, are stored. (Hence, this is called a greedy search). The process continues until the latest highest score stops changing. The problem with being greedy all the time is that the answer might converge to a local maximum. To mitigate this problem and allow some probability of visiting more than one local maximum, one uses a Tabu Table, random restarts, simulated annealing, genetic algorithms, etc.

78.3 Constraint based SL algorithms

To fully understand constraint based SL algorithms, the reader is advised to read Chapters 21 and 61 first.

Constraint based SL algorithms require estimating from the data the conditional independence $\underline{x} \perp_P \underline{y} | \underline{a}$. for any 3 disjoint multinodes $\underline{x}, \underline{y}, \underline{a}$. This can be done by estimating the conditional mutual information (CMI) $H(\underline{x} : \underline{y} | \underline{a})$. `bnlearn` can calculate CMI and other metrics of $\underline{x} \perp_P \underline{y} | \underline{a}$. All these metrics are very similar; they all measure how close $P(x|y, a)$ and $\bar{P}(x|a)$ are.

The first constraint-based SL algorithm was the Inductive Causation (IC) algorithm proposed by Pearl and Verma in 1991. Incremental improvements have been proposed since then, such as the PC family of algorithms, Grow-Shrink and the Incremental Association Markov Blanket (IAMB) family of algorithms.

78.4 Pseudo-code for some bnet learning algorithms

Algorithm 4: Pseudo-code for Hill Climbing algorithm

Input : Data D , Vertices V

Output: a bnet $B = (G, T)$, where $G = (V, E)$ is a DAG, where V are its vertices (nodes) and E are its edges (arrows). T are all its Transition Probability Matrices (TPMs) $T = TPMs(G, D)$.

```

 $E \leftarrow \emptyset$ 
 $T \leftarrow \emptyset$ 
 $B \leftarrow (V, E, T)$ 
 $maxscore \leftarrow -\infty$ 
// DE= all possible directed edges
 $DE = \{\underline{x} \rightarrow \underline{y} \in V \times V : \underline{x} \neq \underline{y}\}$ 
 $again \leftarrow True$ 
while  $again$  do
    for all  $\underline{x} \rightarrow \underline{y} \in DE$  do
        // add arrow
         $E_+ \leftarrow E \cup \{\underline{x} \rightarrow \underline{y}\}$ 
        // delete arrow
         $E_- \leftarrow E - \{\underline{x} \rightarrow \underline{y}\}$ 
        // reverse arrow
         $E_R \leftarrow E_- \cup \{\underline{y} \rightarrow \underline{x}\}$ 
        for  $E' = E_+, E_-, E_R$  do
            if  $E' \neq E$  and  $G' = (V, E')$  is a legal DAG then
                 $T' \leftarrow TPMs(G', D)$ 
                 $B' \leftarrow (G', T')$ 
                 $newscore = \text{BIC-score}(B')$ 
                if  $newscore > maxscore$  then
                     $B \leftarrow B'$ 
                     $maxscore \leftarrow newscore$ 
                else
                     $again \leftarrow False$ 
    return  $B$ 

```

Algorithm 5: Pseudo-code for PC-Stable algorithm

Input : Data D , Vertices (nodes) V , tolerance in CMI $\epsilon > 0$

Output: partially oriented acyclic graph $G = (V, E, UE)$, where V are the vertices (nodes), E are the oriented edges (arrows) and UE are the unoriented edges.

```

 $E \leftarrow \emptyset$ 
// initialize UE to fully-connected undirected graph
 $UE \leftarrow \{\underline{x} - \underline{y} \in V \times V : \underline{x} - \underline{y} = \underline{y} - \underline{x}, \underline{x} \neq \underline{y}\}$ 
// Shrink phase. Deletes edges from E.
 $\text{for } \lambda = 0, 1, 2, \dots, |V| - 2 \text{ do}$ 
     $\quad \text{for all } \underline{x} - \underline{y} \in UE \text{ do}$ 
         $\quad \quad \text{for all } S = \{a \in V : \underline{x} - \underline{a} \in UE, \underline{a} \neq \underline{x}, \underline{y}\} \ni |S| = \lambda \text{ do}$ 
             $\quad \quad \quad \text{if } H(\underline{x} : \underline{y}|S) < \epsilon \text{ then}$ 
                 $\quad \quad \quad \quad /* \text{ If there were an arrow between } \underline{x} \text{ and } \underline{y}, \text{ then}$ 
                 $\quad \quad \quad \quad \text{conditioning on } S \text{ would not be enough to interrupt}$ 
                 $\quad \quad \quad \quad \text{info transmission } H(\underline{x} : \underline{y}|S) \text{ between } \underline{x} \text{ and } \underline{y} */$ 
                 $\quad \quad \quad \quad UE \leftarrow UE - \{\underline{x} - \underline{y}\}$ 
                 $\quad \quad \quad \quad S(\underline{x} - \underline{y}) \leftarrow S$ 
     $\quad // Growth phase. Adds v structures to E.$ 
     $\quad \text{for all } \underline{x}, \underline{y}, \underline{a} \text{ such that } \underline{x} - \underline{a} \in UE, \underline{a} - \underline{y} \in UE, \underline{x} - \underline{y} \notin UE, \underline{a} \notin S(\underline{x} - \underline{y}) \text{ do}$ 
         $\quad \quad /* \text{ If there were no collider at } \underline{a}, \text{ then there would be info}$ 
         $\quad \quad \text{transmission between } \underline{x} \text{ and } \underline{y} */$ 
         $\quad \quad UE \leftarrow UE - \{\underline{x} - \underline{a}, \underline{a} - \underline{y}\}$ 
         $\quad \quad E \leftarrow E \cup \{\underline{x} \rightarrow \underline{a}, \underline{y} \rightarrow \underline{a}\}$ 
// Orienting edges.
 $again \leftarrow True$ 
 $size \leftarrow |UE|$ 
 $\text{while } again \text{ do}$ 
     $\quad \text{for all } \underline{x} - \underline{y} \in UE \text{ do}$ 
         $\quad \quad \text{if } \underline{x} \rightarrow \underline{y} \in E, \underline{y} - \underline{z} \in UE, \underline{x} - \underline{z} \notin UE, \nexists \underline{w} \ni \underline{w} \rightarrow \underline{y} \in E \text{ then}$ 
             $\quad \quad \quad // \text{ to avoid introducing new v structure}$ 
             $\quad \quad \quad UE \leftarrow UE - \{\underline{y} - \underline{z}\}$ 
             $\quad \quad \quad E \leftarrow E \cup \{\underline{y} \rightarrow \underline{z}\}$ 
         $\quad \quad \text{if } \underline{x} \rightarrow \underline{y} \in E \text{ and there is directed path from } \underline{x} \text{ to } \underline{y} \text{ in } E \text{ then}$ 
             $\quad \quad \quad // \text{ to avoid introducing cycles}$ 
             $\quad \quad \quad UE \leftarrow UE - \{\underline{x} - \underline{y}\}$ 
             $\quad \quad \quad E \leftarrow E \cup \{\underline{x} \rightarrow \underline{y}\}$ 
     $\quad newsize \leftarrow |UE|$ 
     $\quad \text{if } size == newsize \text{ then}$ 
         $\quad \quad again \leftarrow False$ 
     $\quad \text{else}$ 
         $\quad \quad size \leftarrow newsize$ 


---



566



return  $G = (V, E, UE)$


```

Chapter 79

Support Vector Machines And Kernel Method

This chapter is based on Refs.[115]. [141] and [116].

The Support Vector Machines (SVM) method was first invented with a linear kernel, but was later generalized to arbitrary kernels. We will use the terms SVM method and Kernel Method indistinguishably.

The SVM method is a fairly general method for calculating, via supervised learning, a *binary* classifier. The SVM method finds a continuous surface that separates a space into two disjoint parts.

Let $\Sigma = [0, 1, 2, \dots, nsam - 1]$ be a list of individuals (samples) in a population. In this chapter, we will use the notation $A^\sigma = A[\sigma]$ and $\vec{A} = [A^\sigma : \sigma \in \Sigma]$ for a list (vector, 1-D array) indexed by Σ . We will refer to $DS = (\vec{x}, \vec{y})$ where $x^\sigma \in S_{\underline{x}}$, $y^\sigma \in \{-1, 1\}$, as a dataset. Let $x^\sigma = (x_0^\sigma, x_1^\sigma, \dots, x_{nf-1}^\sigma) \in S_{\underline{x}_0} \times S_{\underline{x}_1} \times \dots \times S_{\underline{x}_{nf-1}} = S_{\underline{x}}$. When $x_j^\sigma \in \mathbb{R}$ for all j , we will take $x^\sigma \in \mathbb{R}^{nf}$ to be a column vector. x^σ is the feature vector for individual σ , and its components x_i^σ for $i = 0, 1, \dots, nf - 1$ are the features. $y^\sigma \in \{-1, 1\}$ is the binary class to which x^σ belongs.

Let $\hat{y}(x^{\sigma_0}) \in \{-1, 1\}$ be an estimate of $y^{\sigma_0} \in \{-1, 1\}$. The **SVM classifier** is defined as

$$\hat{y}(x^{\sigma_0}) = \text{sign}(Y(x^{\sigma_0})) \quad (79.1)$$

where¹

$$Y(x^{\sigma_0}) = \sum_{\sigma} \alpha^\sigma y^\sigma K(x^\sigma, x^{\sigma_0}) . \quad (79.2)$$

The **binary weight coefficients** $\alpha^\sigma \in \{0, 1\}$ for all $\sigma \in \Sigma$ are found by training, via an algorithm to be described below.

The function $K : S_{\underline{x}} \times S_{\underline{x}} \rightarrow \mathbb{R}$ is called the **Kernel or Similarity function**. We assume that $K(x^\sigma, x^{\sigma_0})$ grows bigger when its two arguments x^σ and x^{σ_0} become

¹Define $\text{sign}(0) = 1$.

more “similar”. We also assume that $K(x^\sigma, x^{\sigma_0})$ is symmetric in its two arguments.

79.1 Learning Algorithm for SVM Classifier

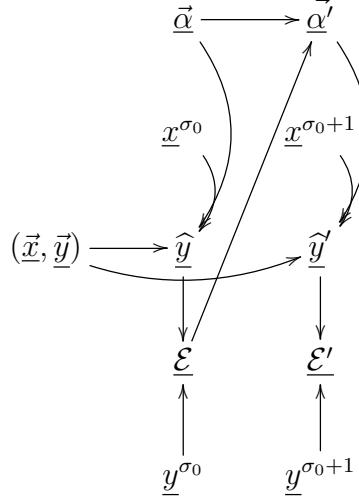


Figure 79.1: Time slice σ_0 of dynamical bnet for learning binary weights $\vec{\alpha}$ of SVM classifier.

Given a kernel function K and a dataset (\vec{x}, \vec{y}) , the SVM classifier is fully specified except for its binary weights $\vec{\alpha}$. Those weights can be learned via the algorithm represented as a causal diagram in Fig.79.1. That figure shows one time slice of a dynamical bnet. The TPMs, printed in blue, of bnet Fig.79.1, are as follows:

$$P(\hat{y}|\vec{\alpha}, (\vec{x}, \vec{y}), x^{\sigma_0}) = \mathbb{1}(\hat{y} = \text{ given by Eq.(79.1).}) \quad (79.3)$$

$$P(\mathcal{E}|\hat{y}, y^{\sigma_0}) = \mathbb{1}(\mathcal{E} = \mathbb{1}(\hat{y} \neq y^{\sigma_0})) \quad (79.4)$$

The first (but not the second, third , etc.) $\vec{\alpha}$ node of Fig.79.1 is a root node. The TPM for that root node should set all components of $\vec{\alpha}$ to zero:

$$P(\vec{\alpha}) = \prod_{\sigma} \mathbb{1}(\alpha^\sigma = 0) . \quad (79.5)$$

After that initialization,

$$P(\vec{\alpha}'|\vec{\alpha}, \mathcal{E}) = \mathbb{1}((\alpha')^{\sigma_0} = \alpha^{\sigma_0} + \mathcal{E}) \prod_{\sigma \neq \sigma_0} \mathbb{1}((\alpha')^\sigma = \alpha^\sigma) \quad (79.6)$$

Why this learning algorithm works.

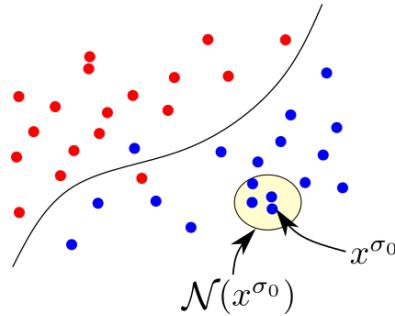


Figure 79.2: Define the neighborhood of x^{σ_0} by $\mathcal{N}(x^{\sigma_0}) = \{x^\sigma : |K(x^\sigma, x^{\sigma_0})| < \epsilon\}$ for some $\epsilon > 0$.

$K(x^\sigma, x^{\sigma_0})$ sets to zero any contribution to $Y(x^{\sigma_0})$ from points x^σ outside the neighborhood $\mathcal{N}(x^{\sigma_0})$ of x^{σ_0} . (See Fig. 79.2). If $\hat{y}(x^{\sigma_0}) = y^{\sigma_0}$, keep $\alpha^{\sigma_0} = 0$ because the neighbors of x^{σ_0} are giving the correct $\hat{y}(x^{\sigma_0})$ when they are polled and the majority wins. If, on the other hand, $\hat{y}(x^{\sigma_0}) \neq y^{\sigma_0}$, then switch α^{σ_0} from 0 to 1, which means x^{σ_0} gets to vote by adding y^{σ_0} to $Y(x^{\sigma_0})$. So we start off with all $\alpha^\sigma = 0$ and we end with most of them still zero except for a select few. If we were to set all α^σ equal to one, we would get overfitting and a very jagged separation between the two classes. The fact that we end with only a select few α^σ equal to 1, and the rest equal to 0, helps make the demarcation between the two classes less jagged.

79.2 Linear (dot-product) Kernel

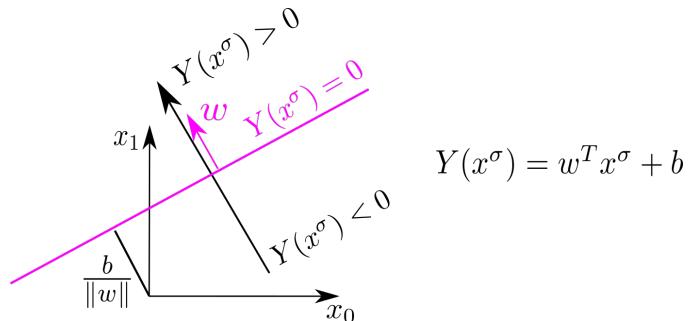


Figure 79.3: Graph of line $Y(x^\sigma) = 0$ splits plane into regions with $Y < 0$, $Y = 0$ and $Y > 0$.

So far, we have discussed the SVM method for an arbitrary kernel. This section is devoted to the **Linear (a.k.a. dot-product) Kernel**. Said kernel is defined as

$$K(x^\sigma, x^{\sigma_0}) = (x^\sigma)^T x^{\sigma_0} . \quad (79.7)$$

For this kernel, Eq.(79.2) specializes to

$$Y(x^{\sigma_0}) = \sum_{\sigma} \alpha^{\sigma} y^{\sigma} K(x^{\sigma}, x^{\sigma_0}) + b \quad (79.8)$$

$$= w^T x^{\sigma_0} + b \quad (79.9)$$

where

$$w = \sum_{\sigma} \alpha^{\sigma} y^{\sigma} x^{\sigma}. \quad (79.10)$$

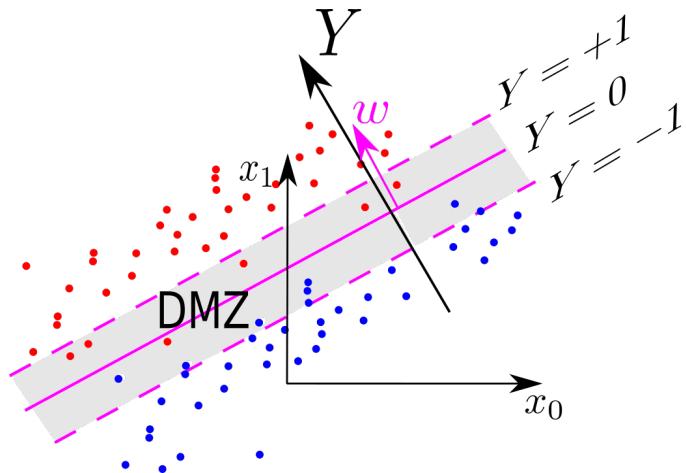


Figure 79.4: We refer to the gray shaded region with $-1 < Y < 1$, where $Y = w^T x + b$, as the DMZ.

We started this chapter by pulling the SVM classifier out of a hat. We did give reasons why it works, but we did not derive it from a more general minimization principle. Such a derivation is possible, at least in the linear kernel case, and we give it next.

Consider the following 3 straight lines:

$$w^T x^{\sigma} + b = +A \quad (79.11)$$

$$w^T x^{\sigma} + b = 0 \quad (79.12)$$

$$w^T x^{\sigma} + b = -A \quad (79.13)$$

where $w, x^{\sigma} \in \mathbb{R}^{nf}$, and $b, A \in \mathbb{R}$. We can re-scale the vector w and scalar b so as to get rid of the A . (i.e., replace $w \rightarrow wA$ and $b \rightarrow bA$ and divide common factor A out of equations). This rescaling does not affect the graphs (i.e., x loci) of these 3 lines. Now we have:

$$w^T x^\sigma + b = +1 \quad (79.14)$$

$$w^T x^\sigma + b = 0 \quad (79.15)$$

$$w^T x^\sigma + b = -1 \quad (79.16)$$

If Y stands for

$$Y = w^T x^\sigma + b, \quad (79.17)$$

then we define the **DMZ (demilitarized zone)** to be the region

$$DMZ = \{x^\sigma : |Y(x^\sigma)| < 1\}. \quad (79.18)$$

The lines $Y = \pm 1$ will be called the **borders (a.k.a. margins)** of the DMZ, and line $Y = 0$ will be called the **line of demarcation** of the DMZ. The DMZ is illustrated in Fig.79.4.

Let D_{DMZ} be the **DMZ width** (i.e., the distance from one border of the DMZ to the other.) Position vectors pointing from the origin to either of the two DMZ borders are called **support vectors**. Suppose $X^+, X^- \in \mathbb{R}^{nf}$ are two support vectors on opposite DMZ borders with $|X^+ - X^-| = D_{DMZ}$. Then

$$w^T X^+ + b = 1 \quad (79.19)$$

$$w^T X^- + b = -1 \quad (79.20)$$

so

$$D_{DMZ} = \frac{2}{|w|}. \quad (79.21)$$

For any $a \in \mathbb{R}$, let the **positive a_+ and negative a_- parts of a** be given by

$$a = \underbrace{a_+}_{a \mathbb{1}(a>0)} + \underbrace{a_-}_{a \mathbb{1}(a \leq 0)}. \quad (79.22)$$

When $Y = \pm 1$, an error in $Y(x^\sigma)$ occurs iff $y^\sigma Y(x^\sigma) = -1$. But how should we define errors when Y is a real number? Define the **Cost of erring for sample σ** to be

$$CE^\sigma(x^\sigma, y^\sigma) = (1 - y^\sigma Y(x^\sigma))_+. \quad (79.23)$$

CE^σ is shown in Fig.79.5. As you can see, there is a penalty for living on the incorrect side, and even a penalty for living on the correct side but too close to the DMZ.

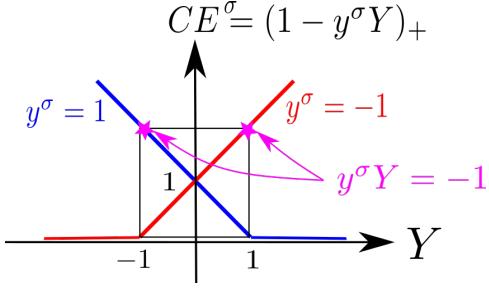


Figure 79.5: Plot of CE^σ versus Y .

Note that the line of demarcation should have the lowest CE^σ for all possible w, b . So to find that line, we want to minimize CE^σ with respect to w, b . But note that $CE^\sigma \geq 0$, and it can be zero for an appropriately chosen w . So we need to add another cost in order to get a non-zero total cost. Define the **DMZ cost** as

$$CZ = \frac{1}{2}|w|^2 = \frac{2}{D_{DMZ}^2} \quad (79.24)$$

Note that $CZ \rightarrow \infty$ as $D_{DMZ} \rightarrow 0$, so CZ penalizes DMZ's that are too narrow.

Now define a Lagrangian \mathcal{L} to be the sum of these 2 contributions.

$$\mathcal{L} = CZ + \sum_{\sigma} CE^\sigma \quad (79.25)$$

$$= \frac{1}{2}|w|^2 + \sum_{\sigma} (1 - y^{\sigma}Y(x^{\sigma}))_{+} \quad (79.26)$$

This particular choice of CZ is not unique, but it isn't totally arbitrary either. We want it to be independent of the sample σ , and to depend on a geometrical aspect of the DMZ, like its width $D_{DMZ} = 2/|w|$. Note that CE^σ behaves, when $|w| \rightarrow \infty$, linearly in $|w|$. We are going to differentiate \mathcal{L} with respect to $|w|$ to find an optimum. But straight lines have no optima, so we need CZ to behave, when $|w| \rightarrow \infty$, as $|w|^p$ for some integer $p > 1$.

Setting the variation $\delta\mathcal{L}$ to zero, we get

$$0 = \delta\mathcal{L} = \delta w_i \left\{ w_i - \sum_{\sigma} \mathbb{1}(y^{\sigma}Y(x^{\sigma}) < 1) y^{\sigma} x_i^{\sigma} \right\} \quad (79.27)$$

so

$$w = \sum_{\sigma} \underbrace{\mathbb{1}(y^{\sigma}Y(x^{\sigma}) < 1)}_{\alpha^{\sigma}} y^{\sigma} x^{\sigma}. \quad (79.28)$$

79.3 Alternatives to Linear Kernel

Sometimes it is convenient to replace the dot-product kernel given above by a different kernel. Other popular kernels are:

- **Radial Basis Function (RBF) Kernel.** In this case, K is a radial function; i.e., a function that depends only on the magnitude (radius, Euclidean distance, L^2 norm) of a vector. An example of an RBF kernel is the **Gaussian Kernel**

$$K(x^\sigma, x^{\sigma_0}) = e^{-\gamma|x^\sigma - x^{\sigma_0}|^2} \quad (79.29)$$

for some free parameter $\gamma > 0$.

- **Inhomogeneous Polynomial Kernel**

$$K(x^\sigma, x^{\sigma_0}) = [(x^\sigma)^T x^{\sigma_0} + 1]^d \quad (79.30)$$

for some positive integer d .

- **Homogeneous Polynomial Kernel**

$$K(x^\sigma, x^{\sigma_0}) = [(x^\sigma)^T x^{\sigma_0}]^d \quad (79.31)$$

for some positive integer d .

- **"Kernel trick" Kernel.** Consider a map $\Phi : \mathbb{R}^{nf} \rightarrow \mathbb{R}^N$. Usually $N > nf$. Φ can be a rectangular matrix $A \in \mathbb{R}^{N \times nf}$ so that $\Phi(x^\sigma) = Ax^\sigma \in \mathbb{R}^N$. Let

$$K(x^\sigma, x^{\sigma_0}) = [\Phi(x^\sigma)]^T \Phi(x^{\sigma_0}) \quad (79.32)$$

Although the constant surfaces of this kernel are hyperplanes in \mathbb{R}^N , its constant surfaces in \mathbb{R}^{nf} can be curved and even closed compact surfaces (e.g. spheres).

79.4 Random Forest and Kernel Method

Chapter 80

Survival Analysis

This chapter is based on Refs.[152] and [142].

Survival Analysis (SA) is used for curve-fitting, to fit a curve $S(t)$ to data indicating the number of patients surviving after receiving a treatment for time t . Alternatively, it can be used with data indicating the number of devices that haven't failed after running for time t . SA can also be used to compare two such $S(t)$ curves—for example, one for treated patients, and another for untreated patients. Hence, SA can be used to analyze the data of an RCT.

Let

$\sigma \in \Sigma$, individual (e.g., patient) in population Σ .

$N = |\Sigma|$, size of Σ , nsam, number of samples

Note: A subset of Σ , (a.k.a, sub-population or stratum) is often called a **cohort** in epidemiology.

$\tau^\sigma \geq 0$, time to **final event** such as death of an organism, or failure of a device, duration of stay, follow-up time period, time period, lifetime.

$\tau^\sigma = b^\sigma - a^\sigma$ for some absolute times a^σ, b^σ satisfying $a^\sigma < b^\sigma$

U^σ = **censoring upper bound**

L^σ = **censoring lower bound**

$B^\sigma = \min(b^\sigma, U^\sigma)$, **right censoring**

$A^\sigma = \max(a^\sigma, L^\sigma)$, **left censoring**

$T^\sigma = B^\sigma - A^\sigma$, where $A^\sigma < B^\sigma$

$d^\sigma = \mathbb{1}(b^\sigma < U^\sigma)$, equals 1 if death (i.e., no censoring), equals 0 if no death (i.e., censoring)

$e^\sigma = \mathbb{1}(a^\sigma > L^\sigma)$, equals 1 if death (i.e., no censoring), equals 0 if no death (i.e., censoring)

Will only use right censoring in this chapter.

For $t \geq 0$, define

- **Survival function** $S(t)$ and **Cumulative hazard function** $\Lambda(t)$

$$S(t) = P(\tau \geq t) = e^{-\Lambda(t)} \quad (80.1)$$

Note that we define in this chapter $S(t) = P(\tau \geq t)$ (those present among

survivors (PAS)) instead of $P(\underline{\tau} > t)$ (not PAS), like other authors do. For continuous functions, these 2 definitions of $S(t)$ are the same.

Note that, since $S(t)$ is a probability and $t, \tau \in [0, \infty]$, $\Lambda(t)$ must be a monotonically increasing function with $\Lambda(0) = 0$ and $\Lambda(\infty) = \infty$.

- **hazard function (a.k.a. instantaneous failure rate)** $\lambda(t)$

$$\Lambda(t) = \int_0^t du \lambda(u) \quad (80.2)$$

$$\lambda(t) = \Lambda'(t) \quad (80.3)$$

Note that $\lambda(t) \geq 0$ and its integral over $[0, \infty]$ must be ∞ . $\lambda(t)$ is in fact a conditional probability as we show next.

$$\lambda(t) = \frac{P(\underline{\tau} = t)}{S(t)} \quad (\text{shown below}) \quad (80.4)$$

$$= \frac{P(\underline{\tau} = t)}{P(\underline{\tau} \geq t)} \quad (80.5)$$

$$= \frac{P(\underline{\tau} = t, \underline{\tau} \geq t)}{P(\underline{\tau} \geq t)} \quad (\text{because } P(A \wedge B) = P(A) \text{ if } A \text{ implies } B) \quad (80.6)$$

$$= P(\underline{\tau} = t | \underline{\tau} \geq t) \quad (80.7)$$

- **$\underline{\tau}$ density function** $P_{\underline{\tau}}(t)$

$$P_{\underline{\tau}}(t) = -S'(t) = \lambda(t)e^{-\Lambda(t)} = \lambda(t)S(t) \quad (80.8)$$

- **$\underline{\tau}$ cumulative distribution function** $\Phi_{\underline{\tau}}(t)$

$$\Phi_{\underline{\tau}}(t) = P(\underline{\tau} < t) = 1 - S(t) \quad (80.9)$$

$$P_{\underline{\tau}}(t) = \Phi'_{\underline{\tau}}(t) \quad (80.10)$$

- **mean survival time** μ

$$\mu = \int_0^\infty dt tP_{\underline{\tau}}(t) \quad (80.11)$$

- **median survival time** τ_{med}

$$S(\tau_{med}) = 0.5 \quad (80.12)$$

80.1 $S(t)$ estimates

80.1.1 No-censoring estimate of $S(t)$

Let

$$r^\sigma(t) = \mathbb{1}(\tau^\sigma \geq t) \quad (80.13)$$

$r^\sigma(t)$ equals 1 iff individual σ at risk (i.e., still alive and not censored, not out) at time t

$$\widehat{S}(t) = \frac{1}{N} \sum_{\sigma} r^\sigma(t) \quad (80.14)$$

$\{\underline{r}^\sigma(t) : \sigma \in \Sigma\}$ are i.i.d.

$\underline{r}^\sigma(t) = \underline{x}$ is a Bernoulli random variable $Bern(p = S(t))$ (i.e., simple coin flip with $P(\underline{x} = 1) = p$, $P(\underline{x} = 0) = q$, mean p and variance pq , where $p = S(t)$)

$$\widehat{S}(t) \rightarrow \mathcal{N} \left(\mu = S(t), \sigma^2 = \frac{S(t)[1 - S(t)]}{N} \right) \text{ as } N \rightarrow \infty \quad (80.15)$$

(convergence in probability) A sanity check for Eq.(80.15) is to note that

$$\begin{aligned} \langle \widehat{S}(t) \rangle &= \frac{1}{N} \sum_{\sigma} \underbrace{\langle \underline{r}^\sigma(t) \rangle}_{S(t)} \end{aligned} \quad (80.16)$$

$$= S(t) \quad (80.17)$$

$$\langle \widehat{S}(t), \widehat{S}(t') \rangle = \frac{1}{N^2} \sum_{\sigma} \sum_{\sigma'} \langle \underline{r}^\sigma(t), \underline{r}^{\sigma'}(t') \rangle \quad (80.18)$$

$$= \frac{1}{N^2} \sum_{\sigma} \langle \underline{r}^\sigma(t), \underline{r}^\sigma(t') \rangle \quad (80.19)$$

$$= \frac{S(t)[1 - S(t)]}{N} \quad (80.20)$$

80.1.2 Kaplan-Meier estimate of $S(t)$

Let

$[\tau^j]_{j=1,2,\dots,J}$, times at which there is a final event, all measured from the same time origin, and ordered so that $\tau^j < \tau^{j+1}$

n_D^j = number of individuals that die at time τ^j

n_C^j = number of individuals that are censored at time τ^j

$n_O^j = n_D^j + n_C^j$ number of individuals that drop-Out at time τ^j , either because they die or are censored

$n_R^j = \sum_{k \geq j} n_O^k$, number of individuals that are at risk at or after time τ^j (i.e., still alive and not censored, not out, surviving before time τ^j)

$d^\sigma \in \{0, 1\}$, it equals 1 iff individual σ dies at any time.

$d^\sigma(t) \in \{0, 1\}$, it equals 1 iff individual σ dies at time t .

$d^\sigma(\tau^j) \in \{0, 1\}$, it equals 1 iff individual σ dies at time τ^j .

$d^\sigma(\tau^j) \rightarrow d(\tau^j)$ since i.i.d.

$c^\sigma = 1 - d^\sigma$, it equals 1 iff individual σ is censored at any time.

$c^\sigma(t) = 1 - d^\sigma(t)$, it equals 1 iff individual σ is censored at time t .

$c^\sigma(\tau^j) = 1 - d^\sigma(\tau^j)$, it equals 1 iff individual σ is censored at time τ^j .

$c^\sigma(\tau^j) \rightarrow c(\tau^j)$ since i.i.d.

$o^\sigma(t) = \mathbb{1}(\tau^\sigma = t)$, it equals 1 iff individual σ drops out at time t .

$o^\sigma(\tau^j) = \mathbb{1}(\tau^\sigma = \tau^j)$, it equals 1 iff individual σ is drops out at time τ^j .

$o^\sigma(\tau^j) \rightarrow o(\tau^j)$ since i.i.d.

Note that¹

$$\underbrace{\bigwedge_{k \leq j} \{o^\sigma(\tau^k) = 0\}}_{\sigma \text{ did not drop out in past or present } \tau^j} = \bigoplus_{k > j} \{o^\sigma(\tau^k) = 1\} \quad (80.21)$$

or, replacing o^σ by \underline{o} ,

$$\bigwedge_{k \leq j} \{\underline{o}(\tau^k) = 0\} = \bigoplus_{k > j} \{\underline{o}(\tau^k) = 1\} \quad (80.22)$$

¹notation: $\wedge_{i=1,2} a_i = (a_1 \wedge a_2) = (a_1 \text{ And } a_2)$,
 $\vee_{i=1,2} a_i = (a_1 \vee a_2) = (a_1 \text{ Or } a_2)$,
 $\oplus_{i=1,2} a_i = (a_1 \oplus a_2) = \text{exclusive or} = \text{modulus 2 addition} = \text{binary addition with } (1 \oplus 1) = 0$.

Kaplan-Meier (KM) estimate of $S(t)$ (see²)

$$\widehat{S}(\tau^j) = P \left(\bigoplus_{k>j} \{\underline{\varrho}(\tau^k) = 1\} \right) \quad (80.23)$$

$$= P \left(\bigwedge_{k \leq j} \{\underline{\varrho}(\tau^k) = 0\} \right) \quad (80.24)$$

$$= \prod_{k \leq j} P \left(\underline{o}(\tau^k) = 0 \mid \bigwedge_{k' < k} \{\underline{\varrho}(\tau^{k'}) = 0\} \right) \quad (\text{chain rule}) \quad (80.25)$$

$$= \begin{cases} \prod_{k \leq j} P \left(\underline{d}(\tau^k) = 0 \mid \bigwedge_{k' < k} \{\underline{\varrho}(\tau^{k'}) = 0\} \right) \\ + \underbrace{\prod_{k \leq j} P \left(\underline{c}(\tau^k) = 0 \mid \bigwedge_{k' < k} \{\underline{\varrho}(\tau^{k'}) = 0\} \right)}_{=0 \text{(see footnote)}} \end{cases} \quad (80.26)$$

$$= \prod_{k \leq j} \left[1 - P \left(\underline{d}(\tau^k) = 1 \mid \bigwedge_{k' < k} \{\underline{\varrho}(\tau^{k'}) = 0\} \right) \right] \quad (80.27)$$

$$= \prod_{k \leq j} \left[1 - \frac{n_D^k}{n_R^k} \right] \quad (80.28)$$

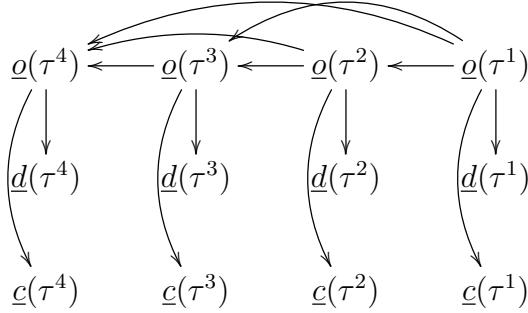


Figure 80.1: Bnet for KM estimate of $S(t)$ for $N = 4$.

Fig.80.1 gives a bnet for the KM estimate of $S(t)$. The TPMs, printed in blue, for that bnet, must be as follows:

$$P(\underline{\varrho}(\tau^k) = 1 \mid \bigwedge_{k' < k} \{\underline{\varrho}(\tau^{k'}) = 0\}) = \frac{n_O^k}{n_R^k} \quad (80.29)$$

²Even though $S(t)$ has been defined as $P(\underline{\tau} \geq t)$, a more precise definition in the presence of censoring is $P(\underline{\tau} \geq t \mid \text{patient dies rather than being censored})$

$$P(\underline{d}(\tau^k) = 1 \mid \underline{\varrho}(\tau^k) = 1) = \frac{n_D^k}{n_O^k} \quad (80.30)$$

$$P(\underline{c}(\tau^k) = 1 \mid \underline{\varrho}(\tau^k) = 1) = \frac{n_C^k}{n_O^k} \quad (80.31)$$

Intuition: Let

$$\widehat{\lambda}^k = \frac{n_D^k}{n_R^k} \quad (80.32)$$

If $\widehat{\lambda}^k \ll 1$, then, since $e^x \approx 1 + x$ for $|x| \ll 1$,

$$\widehat{S}(\tau^j) \approx \prod_{k \leq j} e^{-\widehat{\lambda}^k} \quad (80.33)$$

$$= \exp\left[-\sum_{k \leq j} \widehat{\lambda}^k\right] \quad (80.34)$$

$$\approx \exp\left[-\int_0^{\tau_j} dt \widehat{\lambda}(t)\right] \quad (80.35)$$

Note that:

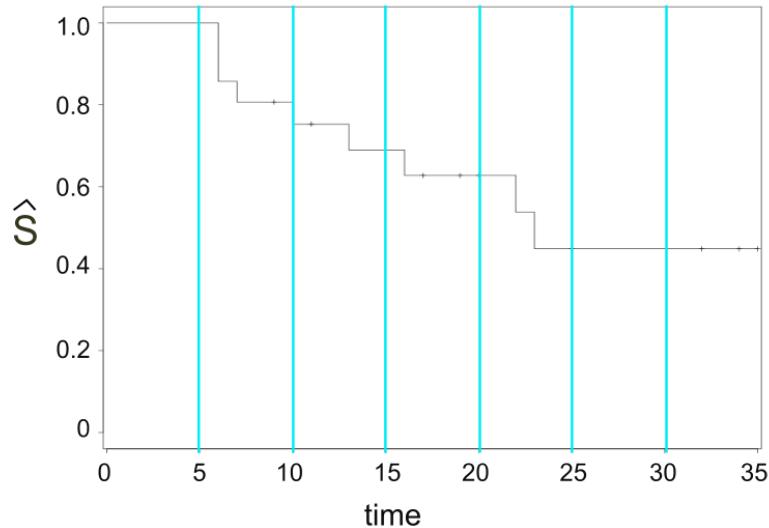
- $\widehat{S}(t)$ only changes when there is a death.
- $\widehat{S}(t) = 1$ before the first death time
- $\widehat{S}(t)$ only goes to 0 if the last observation is a death, so $n_D^j/n_R^j = 1$.
- When there is no censoring, the KM estimate equals the no censoring estimate given earlier.

Greenwood's formula for variance of KM estimate of $S(t)$

$$\left\langle \widehat{S}(t), \widehat{S}(t) \right\rangle = [S(t)]^2 \sum_{k: \tau^{k-1} \leq t} \frac{n_D^k}{n_C^k n_R^k} \quad (80.36)$$

Fig.80.2 and Table 80.1 give a numerical example of the KM estimate.

Fig.80.3 shows relevant parameters for steps τ^{j-1} and τ^j in a plot of a KM estimate.



6	6	6	6C	7	9C	10
10C	11C	13	16	17C	19C	20C
22	23	25C	32C	32C	34C	35C

Figure 80.2: Plot of KM estimate for $N = 21$ patients with τ^{σ} given in table below the plot. τ 's with C are censored.

j	τ^j	n_R^j	n_D^j	n_C^j
1	6	21	3	1
2	7	17	1	0
3	9	16	0	1
4	10	15	1	1
5	11	13	0	1

Table 80.1: Parameters $j, \tau^j, n_R^j, n_D^j, n_C^j$ for first five lifetimes for Fig.80.2.

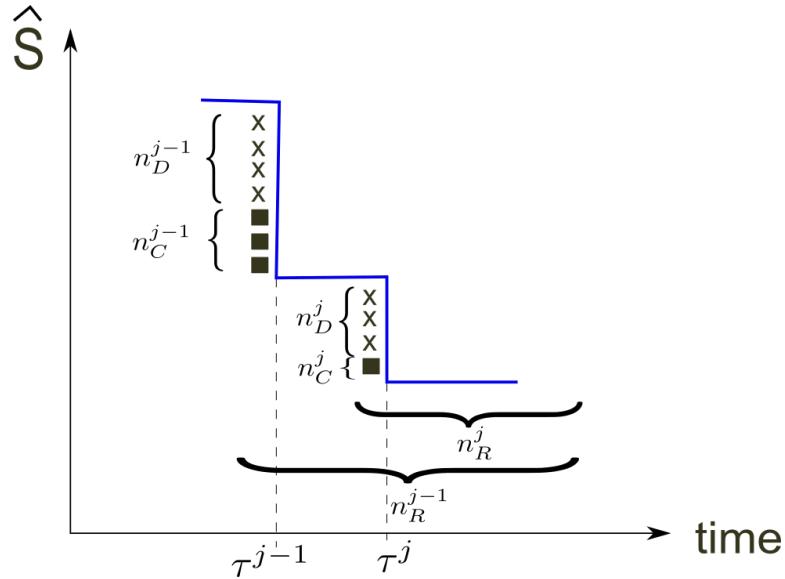


Figure 80.3: Steps τ^{j-1} and τ^j in a plot of a KM estimate. An **X** indicates a dead individual and **■** a censored one. Note that

$$n_O^j = n_D^j + n_C^j,$$

$$n_R^j = n_R^{j-1} - n_O^j \text{ and}$$

$$n_R^j = \sum_{k \geq j} n_O^k.$$

80.2 $\lambda(t)$ models

80.2.1 $\lambda(t)$ independent of covariates Z

Recall $t \geq 0$

- $\lambda(t)$ is independent of time

$$\lambda(t) = \lambda \quad (80.37)$$

where $\lambda \geq 0$.

$$\Lambda(t) = \lambda t \quad (80.38)$$

$$P_\tau(t) = \lambda e^{-\lambda t} \quad (\text{Exponential distribution}) \quad (80.39)$$

- $\lambda(t)$ is proportional to power of t

$$\lambda(t) = \kappa \lambda t^{\kappa-1} \quad (80.40)$$

where $\lambda, \kappa \geq 0$.

$$\Lambda(t) = \lambda t^\kappa \quad (80.41)$$

$$P_{\underline{\tau}}(t) = \kappa \lambda t^{\kappa-1} e^{-\lambda t^\kappa} \quad (\text{Weibull distribution}) \quad (80.42)$$

- $\lambda(t) = a + bt$ for $a, b \geq 0$
- $\lambda(t)$ is piecewise constant for $t \in [0, \infty]$
- etc.

Maximum Likelihood Estimate (MLE) of λ for $\lambda(t) = \lambda$, allowing censoring: Likelihood function is

$$L(\lambda) = \prod_j [P_{\underline{\tau}}(\tau^j)]^{d(\tau^j)} [S(\tau^j)]^{c(\tau^j)} \quad (80.43)$$

$$= \prod_j [\lambda e^{-\lambda \tau^j}]^{d(\tau^j)} [e^{-\lambda \tau^j}]^{1-d(\tau^j)} \quad (80.44)$$

$$= \prod_j \lambda^{d(\tau^j)} e^{-\lambda \tau^j} \quad (80.45)$$

Hence

$$\ln L(\lambda) = (\ln \lambda) \underbrace{\sum_j d(\tau^j)}_D - \lambda \underbrace{\sum_j \tau^j}_T \quad (80.46)$$

Setting $\partial_\lambda \ln L = 0$, we get

$$\hat{\lambda} = \frac{D}{T} \quad (80.47)$$

80.2.2 $\lambda(t)$ dependent on covariates Z

Suppose $\beta, Z \in \mathbb{R}^{nind}$ are column vectors, where $nind$ is number of independent variables (covariates) in a regression.

Cox Proportional Hazards (PH) model for $\lambda(t)$

$$\lambda(t) = \lambda(t|Z) = \lambda_0(t) e^{\beta^T Z} \quad (80.48)$$

where $\lambda_0(t) \geq 0$. $\lambda_0(t)$ is called the **baseline hazard function**. This $\lambda(t)$ model is called a PH model because $\lambda(t|Z_1)/\lambda(t|Z_2) = \exp[\beta^T(Z_1 - Z_2)]$ is independent of time. In a **Cox hazards model with time-dependent covariates (TC)**, one assumes that the covariates $Z(t)$ depend on time.

Note that since exponentials are always positive, the components of β and Z can range over all real values. If we had chosen $\lambda(t) = \lambda_0(t)\beta^T Z$ instead, then we would have to constrain $\beta^T Z \geq 0$.

If we define

$$\Lambda_0(t) = \int_0^t du \lambda_0(u) \quad (80.49)$$

then

$$P_{\underline{\tau}}(t) = \lambda_0(t)e^{\beta^T Z} \exp[-\Lambda_0(t)e^{\beta^T Z}] \quad (80.50)$$

$P_{\underline{\tau}}(t)$ is called the **Cox PH distribution**. It's a **semi-parametric distribution** because it depends on both, a prior parameter β and a prior function $\lambda_0(t)$ (a function is like an infinite dimensional parameter). A **parametric distribution** depends only on a prior parameter and a **non-parametric distribution** depends only on a prior function.

Recall that we defined earlier

n_D^j = number of individuals that die at time τ^j

n_C^j = number of individuals that are censored at time τ^j

To define the Cox partial likelihood function, we will assume that $n_D^j + n_C^j = 1$, i.e., each time τ^j has either a single death, or a single censorship, but not both. Since every individual eventually dies or is censored (but, we will assume, not both), there is a 1-1 onto map $j(\sigma)$ mapping Σ to the set of indices j of τ^j . So we can label the population individuals σ by their index j , or vice versa.

Let

$$L^j(\beta) = \frac{e^{\beta^T Z^j}}{\sum_{k \geq j} e^{\beta^T Z^k}} \quad (80.51)$$

Then define the **Cox partial likelihood function** by

$$L(\beta) = \prod_j [L^j(\beta)]^{d(\tau^j)} \quad (80.52)$$

Cox's approximate method for finding the best fit for β is to set $\frac{\partial \ln L(\beta)}{\partial \beta_a} = 0$ for all a . This does not determine the baseline hazard function, however.

Recall that

$$\lambda(\tau^j | Z^k) = \lambda_0(\tau^j) e^{\beta^T Z^k} = P(\underline{\tau}^k = \tau^j | \underline{\tau}^k \geq \tau^j) \quad (80.53)$$

Therefore

$$L^j(\beta) = \frac{\lambda_0(\tau^j)e^{\beta^T Z^j}}{\sum_{k \geq j} \lambda_0(\tau^j)e^{\beta^T Z^k}} \quad (80.54)$$

$$= \frac{\lambda(\tau^j | Z^j)}{\sum_{k \geq j} \lambda(\tau^j | Z^k)} \quad (80.55)$$

Next, we try to justify Cox's partial likelihood function. We will give two arguments.

1. Bayesian argument

Assume that Z^k is a random variable with a non-informative prior $P(Z^k) = \mathcal{N}(!k)$. Then

$$P(Z^k | \tau^j) = f(\tau^j) \underbrace{P(\tau^j | Z^k)}_{\lambda(\tau^j | Z^k)} \quad (80.56)$$

for some function $f : \mathbb{R} \rightarrow \mathbb{R}$. Hence

$$L^j(\beta) = \frac{P(Z^j | \tau^j)}{\sum_{k \geq j} P(Z^k | \tau^j)} \quad (80.57)$$

$$= \frac{P(Z^j | \tau^j)}{P(\bigvee_{k \geq j} \{\underline{Z}^k = Z^k\} | \tau^j)} \quad (\text{because the } Z^k \text{ are independent}) \quad (80.58)$$

$$= \frac{P(Z^j, \bigvee_{k \geq j} \{\underline{Z}^k = Z^k\}, \tau^j)}{P(\bigvee_{k \geq j} \{\underline{Z}^k = Z^k\}, \tau^j)} \quad (\text{because } P(A \wedge B) = P(A) \text{ if } A \implies B) \quad (80.59)$$

$$= P(Z^j | \bigvee_{k \geq j} \{\underline{Z}^k = Z^k\}, \tau^j) \quad (80.60)$$

Note that $L^j(\beta)$ depends on $\{Z^k : k \geq j\}$ because at time j , the past Z^k are fixed already, so the only ones we are allowed to optimize (remember, we are acting as Bayesians here, so we can optimize parameters) are the present and future ones.

2. Maximum Likelihood argument

$$L^j(\beta) = \begin{cases} P(\underline{\tau} \geq \tau^j) = S(\tau^j) & \text{if } d(\tau^j) = 0 \text{ (i.e., } c(\tau^j) = 1) \\ P(\underline{\tau} = \tau^j) = \lambda(\tau^j | Z^j)S(\tau^j) & \text{if } d(\tau^j) = 1 \end{cases} \quad (80.61)$$

$$= \lambda(\tau^j | Z^j)^{d(\tau^j)} S(\tau^j) \quad (80.62)$$

$$= \underbrace{\left[\frac{\lambda(\tau^j | Z^j)}{\sum_{k \geq j} \lambda(\tau^k | Z^j)} \right]^{d(\tau^j)}}_{L_1} \underbrace{\left[\sum_{k \geq j} \lambda(\tau^k | Z^j) \right]^{d(\tau^j)} S(\tau^j)}_{L_2} \quad (80.63)$$

Cox argued that L_2 varies very slowly with β .

80.3 $S_0(t)$ estimates

$$S_Z(t) = e^{-\Lambda(t)} \quad (80.64)$$

$$= e^{-\Lambda_0(t) \exp(\beta^T Z)} \quad (80.65)$$

$$= S_{Z=0}(t)^{\exp(\beta^T Z)} \quad (80.66)$$

Claim 121 (*Breslow $S_0(t)$ estimate*)

$$\widehat{S}_0(t) = e^{-\widehat{\Lambda}_0(t)} \quad (80.67)$$

where

$$\widehat{\Lambda}_0(t) = \sum_{j: \tau^j < t} \left[\frac{n_D^j}{\sum_{k \geq j} e^{\beta^T Z^k}} \right] \quad (80.68)$$

proof:

$$\frac{n_D^j}{\Delta t} \approx \sum_{k \geq j} P(\underline{\tau}^k = \tau^j | \underline{\tau}^k \geq \tau^j) \quad (80.69)$$

$$= \sum_{k \geq j} \lambda(\tau^j | Z^k) \quad (80.70)$$

$$= \lambda_0(\tau^j) \sum_{k \geq j} e^{\beta^T Z^k} \quad (80.71)$$

Hence

$$\lambda_0(\tau^j) \Delta t = \frac{n_D^j}{\sum_{k \geq j} e^{\beta^T Z^k}} \quad (80.72)$$

If we now apply $\sum_{j:\tau^j < t}$ to both sides of the last equation, we get Eq.(80.68).

QED

Note that

$$\widehat{S}_0(t) = \prod_{j:\tau^j < t} \exp \left[\frac{-n_D^j}{\sum_{k \geq j} e^{\beta^T Z^k}} \right] \quad (80.73)$$

$$\approx \prod_{j:\tau^j < t} \left[1 - \underbrace{\frac{n_D^j}{\sum_{k \geq j} e^{\beta^T Z^k}}}_{\widehat{\lambda}^j} \right] \text{ (because } e^x \approx 1 + x \text{ for } |x| \ll 1) \quad (80.74)$$

Chapter 81

Synthetic Controls

This chapter is based on Refs.[10] and [8].

This chapter assumes that the reader has read Chapter 16 on the Difference-in-Differences (DID) method.

The Synthetic Controls (SC) method is a simple enhancement of the DID method. SC enhances DID in two simple yet powerful ways:

1. **Better time resolution.** DID considers just 2 time-snapshots (i.e., a time-series with only 2 times) whereas SC considers arbitrarily many time-snapshots (i.e., a time-series with more than 2 times).
2. **Weighted average of controls.** DID divides the population of individuals into just 2 kinds: the treated and the untreated (a.k.a. controls). SC divides the total population into treated and controls just like DID does, but it goes further and divides the control population into multiple subpopulations, and calculates a weighted average, called a “synthetic control”, of those subpopulations. The weights of the synthetic control are chosen so that it mimics as closely as possible the behavior of the treated population for all times measured before the treatment was applied.

Let us describe these two enhancements more precisely.

- **timing:** Let t_k for $k = 0, 1, \dots, n_{pre} - 1$ be the pre-treatment times at which a measurement occurs. Let t_k for $k = n_{pre}, n_{pre} + 1, \dots, n_t - 1$ be the post-treatment times at which a measurement occurs. Note that $n_{pre} + n_{post} = n_t$. Note that $t_* = t_{n_{pre}+1}$ is the first measurement time after the treatment is applied, t_0 is the first measurement time, and $t_{fin} = t_{n_t-1}$ is the last one.
- **subpopulations:** Let $S_1 = \{\sigma_1\}$ be the set of treated units (just one). Let $S_0 = \{\sigma : \sigma \neq \sigma_1\}$ be the set of untreated units (i.e., controls). Let $nsam$ = number of all units σ , $n_1 = |S_1| = 1$, and $n_0 = |S_0| = nsam - 1$.
- **weights:**

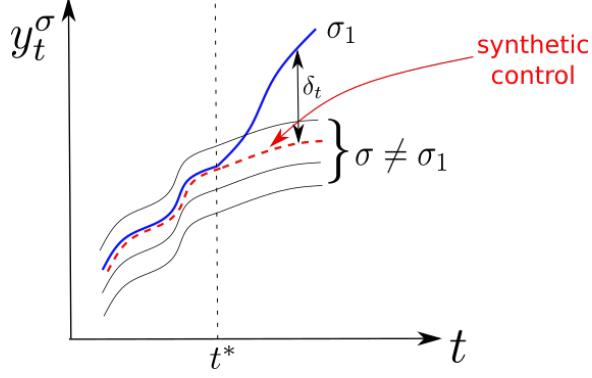


Figure 81.1: Pictorial representation of the Synthetic Controls (SC) method. The outcome y of the synthetic control unit is colored red and that of the treated unit is colored blue. They roughly agree for $t < t_*$.

We want to define a time-independent weight w^σ for each unit σ in such a way that the output y_t^σ for the synthetic control unit behaves like the output for the treated unit σ_1 for $t < t_*$.

Let

$$w^{\sigma_1} = 0 \quad (81.1)$$

and

$$w^{n_0} = \{w^\sigma\}_{\sigma \neq \sigma_1}. \quad (81.2)$$

Define a cost function \mathcal{C} :

$$\mathcal{C}(w^{n_0}) = \sum_{t < t_*} \left(y_t^{\sigma_1} - \sum_{\sigma \neq \sigma_1} w^\sigma y_t^\sigma \right)^2 \quad (81.3)$$

Then calculate w^{n_0} by minimizing the cost function, subject to the constraint that w^{n_0} be a probability distribution:

$$w^{n_0} = \underset{W^{n_0}}{\operatorname{argmin}} \left\{ \mathcal{C}(W^{n_0}) : W^\sigma \geq 0, \sum_{\sigma \neq \sigma_1} W^\sigma = 1 \right\}. \quad (81.4)$$

Now that we have defined a weight w^σ for every unit σ , we can define for $c \in \{0, 1\}$,

$$y_t^{\sigma_1}(c) = \begin{cases} y_t^{\sigma_1} & \text{if } c = 1 \\ \sum_{\sigma \neq \sigma_1} w^\sigma y_t^\sigma & \text{if } c = 0 \end{cases} \quad (81.5)$$

$$\mathcal{Y}_c(t) = E_\sigma[y_t^\sigma(c)] \quad (81.6)$$

and

$$\delta_t = \mathcal{Y}_1(t) - \mathcal{Y}_0(t) \quad (81.7)$$

δ_t is illustrated in Fig.81.1. It approximates $ATE(t)$.

81.1 PO analysis

In this section, we show how to analyze the SC method using the formalism of PO theory.

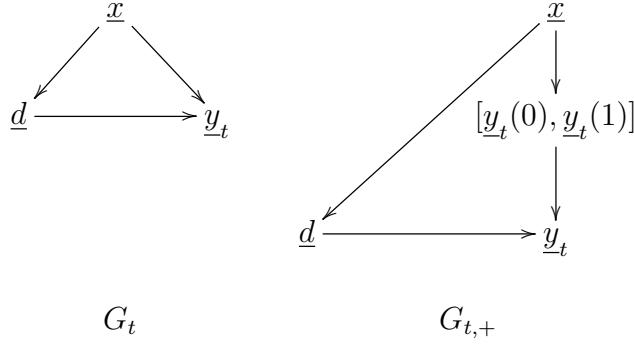


Figure 81.2: $t \in \{t_0, t_1, \dots, t_{fin}\}$. Bnet $G_{t,+}$ is obtained by adding two new nodes $y_t(0)$ and $y_t(1)$ to bnet G_t .

As usual for PO theory, we will consider expected values of y_t^σ :

$$E_{\sigma|d,x}[y_t^\sigma(c)] = E_{y_t(c)|d,x}[y_t(c)] = \mathcal{Y}_{c|d,x}(t) \quad (81.8)$$

To calculate these expected values, we need a “model” with probability distributions. In this case, the needed model and probability distributions are provided by the bnets depicted in Fig.81.2. The TPMs, printed in blue, for the bnet $G_{t,+}$ in Fig.81.2, are as follows. Note that the TPMs for the bnet $G_{t,+}$ are defined in terms of the TPMs for the bnet G_t .

$$P(x) = P_{\underline{x}}(x) \quad (81.9)$$

$$P(d|x) = P_{d|\underline{x}}(d|x) \quad (81.10)$$

$$P(y_t|y_t(0), y_t(1), d) = \mathbb{1}(y_t = y_t(d)) \quad (81.11)$$

$$P(y_t(c)|x) = P(y_t(c)|d, x) = \text{given} \quad (81.12)$$

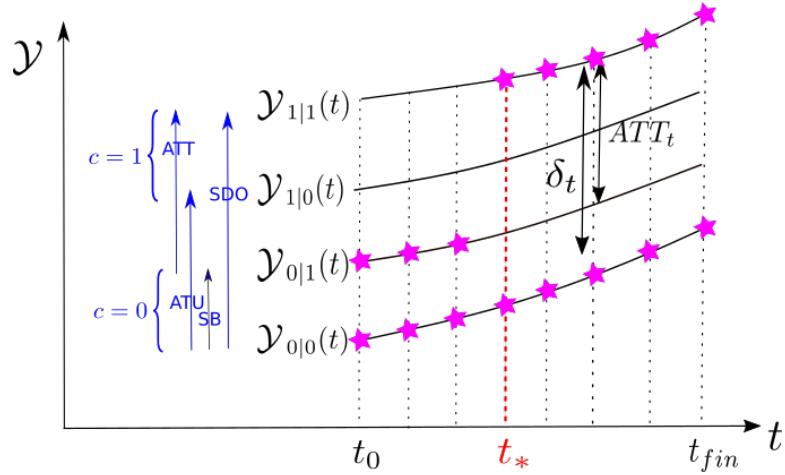


Figure 81.3: Four different time-dependent expected values $\mathcal{Y}_{c|d}(t)$ of y_t^σ for bnet $G_{t,+}$. The $2 * nt$ magenta stars represents the $2 * nt$ SC measurements.

Fig.81.3 depicts the four functions $\mathcal{Y}_{c|d}(t)$ for t in the interval $[t_0, t_{fin}]$ and for $c, d \in \{0, 1\}$. The \mathcal{Y} coordinates of the $2 * nt$ magenta stars in Fig.81.3 can be calculated using bnet G_t . Note that in Fig.81.3, we display a large gap between the curves $\mathcal{Y}_{0|d}(t)$ for $d \in \{0, 1\}$. In reality, $P(y_t(0)|d)$ has been constructed so as to make that gap as small as possible. Thus, to a good(?) approximation,

$$\delta_t \approx ATE_t \quad (81.13)$$

Unlike in the DID method, in the SC method, to a good(?) approximation, we don't have to worry about parallel trends.

Chapter 82

Targeted Estimator

This chapter is based on Refs.[4] and [21].

Targeted Estimator (TE) theory addresses the following concerns. Suppose $\Psi[P]$ is an estimator that depends on the full probability distribution P of a fixed bayesian network. $\Psi[P]$ is a functional (i.e., a function of a function) of P . If P is perturbed by a small amount δP , we get $\Psi[P + \delta P]$. $\Psi[P + \delta P]$ can be expanded in powers of δP . The term linear in δP defines the “influence function”; it also defines the “functional derivative” of Ψ with respect to P . *Why are influence functions useful?* The influence function measures, to first order in δP , how the estimator Ψ responds to a perturbation δP in P . In general, Ψ does not have to be a counterfactual estimator, but it might be one, like an estimator of ATE, or PNS or whatever. *So what is this good for?* It is a way of generating linear ”targeted” estimators that are less noisy and converge more quickly. It measures the sensitivity of an estimator to perturbations in P . It does not, however, measure sensitivity to changes in the DAG (the DAG is fixed throughout). And it does not generate new estimands.

The goal of TE and the strategy one uses to achieve it, is explained more precisely in the next section.

82.1 Goal, Strategy, and Rationale of TE theory

Let $\underline{b} = (\underline{b}_1, \underline{b}_2, \dots, \underline{b}_n)$ denote the n nodes of a Bayesian Network, and let $P_{\underline{b}}(\underline{b})$ for $\underline{b} \in S_{\underline{b}}$ denote the full probability distribution of the bnet.

Consider a population Σ of individuals $\sigma \in \Sigma$ with $N = |\Sigma|$. The **empirical probability distribution** $P_N : S_{\underline{b}} \rightarrow [0, 1]$ for this bnet is defined by

$$P_N(\underline{b}) = \frac{1}{N} \sum_{\sigma} \delta(\underline{b}, \underline{b}_{\sigma}) \quad (82.1)$$

$$\sum_{\underline{b}} P_N(\underline{b}) f(\underline{b}) = \frac{1}{N} \sum_{\sigma} f(\underline{b}_{\sigma}) \quad (82.2)$$

As $N \rightarrow \infty$, $P_N(\underline{b})$ tends to the probability distribution $P_{\underline{b}}(\underline{b})$ of the bnet.

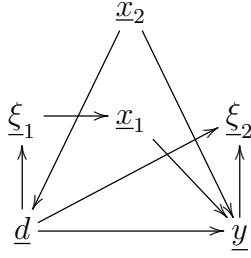


Figure 82.1: Example of bnet considered in TE theory.

Let $\underline{b} = (\underline{X}, \underline{\xi})$ and $\underline{X} = (\underline{d}, \underline{x}, \underline{y})$, where node $\underline{d} \in \{0, 1\}$ denotes a decision to treat a patient, node $\underline{y} \in \{0, 1\}$ denotes the treatment outcome, multi-node \underline{x} denotes the covariates that are good controls, and multi-node $\underline{\xi}$ denotes the covariates that we don't want to control. See Fig.82.1 for an example of a bnet that fits this description.

The cfit \hat{y} of y is a function $\hat{y} : S_d \times S_x \rightarrow \mathbb{R}$ that minimizes the **loss** (a.k.a. **loss functional**) \mathcal{L} given by

$$\mathcal{L}[P, \hat{y}] = \sum_X P(X) \hat{\mathcal{L}}[y, \hat{y}(d, x)]^2, \quad (82.3)$$

where $\hat{\mathcal{L}}(a, b)$, the **loss cfit**, is a non-negative function that vanishes when $a = b$. The function $\hat{\mathcal{L}}$ is designed to minimize a particular kind of error. In this chapter, the $\hat{\mathcal{L}}$ is designed to reduce ATE error.

The estimate $\Psi[P, \hat{y}]$ for the cfit \hat{y} is defined by

$$\Psi[P, \hat{y}] = \sum_X P(X) \hat{y}(d, x) \quad (82.4)$$

Unfortunately, the words “estimator” and “estimate” are often used interchangeably. See Section C.21. In this chapter, we use (\hat{y}, Ψ) for our (estimator, estimate) pair, and refer to estimators as cfits.

Let

$$\delta P(X) = P(X) - P_{in}(X) \quad (82.5)$$

where $P, P_{in} : S_X \rightarrow [0, 1]$ are probability distributions.

Define $\delta\hat{y}(X)$ by

$$\delta\hat{y}(X) = \frac{\hat{y}(d, x)\delta P(X)}{P(X)} \quad (82.6)$$

Hence

$$P(X)\delta\hat{y}(X) = \hat{y}(d, x)\delta P(X) \quad (82.7)$$

Since $\Psi[P, \hat{y}]$ is linear in P and \hat{y} , it follows that

$$\Psi[P, \hat{y} + \delta\hat{y}] = \Psi[P + \delta P, \hat{y}] \quad (82.8)$$

Suppose P_{in++} satisfies

$$P_{in++} = \operatorname{argmin}_{P_{in}} \mathcal{L}[P_{in}, \hat{y}] \quad (82.9)$$

and

$$\lim_{N \rightarrow \infty} \Psi[P_N, \hat{y} + \delta\hat{y}] = \lim_{N \rightarrow \infty} \underbrace{\Psi[P_N + \delta P, \hat{y}]}_{P_{in++}} = \lim_{N \rightarrow \infty} \Psi[P_N, \hat{y}] . \quad (82.10)$$

Eq.(82.10) is illustrated in Fig.82.2.

The goal of TE theory is to find, given a cfit \hat{y} , a new cfit $\hat{y} + \delta\hat{y}$ so that the estimate $\Psi[P_N, \hat{y} + \delta\hat{y}]$ has better behavior as $N \rightarrow \infty$ than $\Psi[P_N, \hat{y}]$ (i.e., converges faster, has smaller bias and variance).

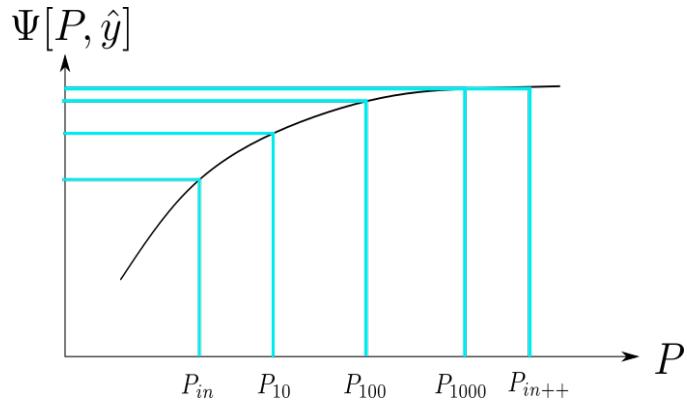


Figure 82.2: Plot of $\Psi[P, \hat{y}] \in \mathbb{R}$ versus P at fixed \hat{y} . In reality, the P are not real numbers but functions.

82.2 Functional Calculus

Define the Hilbert space of square integrable functions over $\underline{X} \in S_{\underline{X}}$ by

$$\mathcal{H}_{\underline{X}} = \{h : (h : S_{\underline{X}} \rightarrow \mathbb{R}) \text{ and } \sum_{X \in S_{\underline{X}}} [h(X)]^2 < \infty\} \quad (82.11)$$

For any $f, g \in \mathcal{H}_{\underline{X}}$, define the dot product (a.k.a. inner product) of f and g by

$$f \cdot g = \sum_X f(X)g(X) \quad (82.12)$$

and the norm

$$\| f \|_P = \sqrt{\sum_X [f(X)]^2}. \quad (82.13)$$

Suppose $P : S_{\underline{X}} \rightarrow [0, 1]$ is a probability distribution. Note that $P \in \mathcal{H}_{\underline{X}}$. For any $f, g \in \mathcal{H}_{\underline{X}}$, define the P expected value by

$$\langle f \rangle_P = P \cdot f \quad (82.14)$$

and the P covariance by

$$\langle f, g \rangle_P = \langle fg \rangle_P - \langle f \rangle_P \langle g \rangle_P \quad (82.15)$$

Suppose $\Psi[\eta] \in \mathbb{R}$ is a real valued function that depends on a function $\eta \in \mathcal{H}_{\underline{X}}$. $\Psi[\eta]$ is said to be a **functional** of η . Define the **functional derivative or gradient**¹ of $\Psi[\eta]$ with respect to η , as follows

$$\frac{\delta \Psi[\eta]}{\delta \eta(a)} = \lim_{\epsilon \rightarrow 0} \frac{\Psi[\eta(x) + \epsilon \frac{\delta(x,a)}{\Delta x}] - \Psi[\eta(x)]}{\epsilon} \quad (82.16)$$

where $\delta(x, a)$ is the Kronecker delta function. For example,

$$\frac{\delta}{\delta \eta(a)} \sum_x \Delta x \eta(x) h(x) = \sum_x \Delta x \frac{\delta(x, a)}{\Delta x} h(x) = h(a) \quad (82.17)$$

Let $\delta(x - a)$ denote the Dirac delta function. If we replace $\frac{\delta(x,a)}{\Delta x} \rightarrow \delta(x - a)$ and $\sum_x \Delta x \rightarrow \int dx$, we go from the discrete to the continuous version of the functional derivative.

$$\sum_x \Delta x \frac{\delta(x, a)}{\Delta x} \rightarrow \int dx \delta(x - a) \quad (82.18)$$

In this chapter, we will use only the discrete version. The Δx in the numerator and the one in the denominator, always cancel each other when we go from discrete to continuous, so we can set $\Delta x = 1$ with impunity.

We will also use the notation

$$\nabla \Psi[\eta](a) = \frac{\delta \Psi[\eta]}{\delta \eta(a)}. \quad (82.19)$$

Suppose $\eta, \eta_0 \in \mathcal{H}_{\underline{X}}$. Define the **functional Taylor expansion** of $\Psi[\eta]$ at η_0 , as follows:

¹Functional derivatives are commonly used in Physics especially in Quantum Field Theory. See Ref.[100] for more information about them.

$$\Psi[\eta] = \Psi[\eta_0] + \sum_x \left[\frac{\delta \Psi[\eta]}{\delta \eta(x)} \right]_{\eta=\eta_0} \delta \eta(x) + \frac{1}{2!} \sum_{x,x'} \left[\frac{\delta^2 \Psi[\eta]}{\delta \eta(x) \delta \eta(x')} \right]_{\eta=\eta_0} \delta \eta(x) \delta \eta(x') + \dots \quad (82.20)$$

where we abbreviate $\eta = \eta(x)$, $\eta_0 = \eta_0(x)$ and $\delta \eta(x) = \eta(x) - \eta_0(x)$.

Define the **functional directional derivative of $\Psi[\eta]$ in the $h \in \mathcal{H}_{\underline{X}}$ direction** by $h \cdot \frac{\delta \Psi[\eta]}{\delta \eta}$. If one compares functional calculus with vector calculus, we see that $\frac{\delta \Psi[\eta]}{\delta \eta}$ corresponds to a gradient $\nabla f(\vec{x})$ and $h \cdot \frac{\delta \Psi[\eta]}{\delta \eta}$ corresponds to a directional derivative $\vec{d} \cdot \nabla f(\vec{x})$. For $|\vec{d}| \ll 1$, $\vec{d} \cdot \nabla f(\vec{x})$ approximates the change in $f(\vec{x})$ when \vec{x} moves from \vec{x} to $\vec{x} + \vec{d}$

82.3 Linear Approximation of $\Psi[P_N]$

Consider probability distributions $P, P_{in} : S_{\underline{X}} \rightarrow [0, 1]$. The **linear approximation (a.k.a. one-step-approximation)** to the Taylor expansion of $\Psi[P]$ at P_{in} is given by

$$\underbrace{\Psi[P] - \Psi[P_{in}]}_{\delta \Psi[P, P_{in}]} = \sum_X \underbrace{\left[\frac{\delta \Psi[P]}{\delta P(X)} \right]_{P=P_{in}}}_{\nabla \Psi[P_{in}](X)} \underbrace{\delta P(X)}_{P(X) - P_{in}(X)} + \mathcal{R}[P, P_{in}] \quad (82.21)$$

If we set

$$\nabla \Psi_{in} = \langle \nabla \Psi[P_{in}] \rangle_{P_{in}}, \quad (82.22)$$

then Eq.(82.21) becomes

$$\delta \Psi[P, P_{in}] = \sum_X P(X) \underbrace{\{\nabla \Psi[P_{in}](X) - \nabla \Psi_{in}\}}_{\lambda(X)} + \mathcal{R}[P, P_{in}] \quad (82.23)$$

$\lambda(X)$ is called the **efficient influence curve (EIF)**.

Recall the Cauchy-Schwartz (CS) inequality for $\vec{a}, \vec{b} \in \mathbb{R}^n$:

$$\vec{a} \cdot \vec{b} = |\vec{a}| |\vec{b}| \cos \theta \leq |\vec{a}| |\vec{b}| \quad (82.24)$$

Note that

$$\langle \lambda \rangle_P = \sum_X \lambda(X) \sqrt{P(X)} \sqrt{P(X)} \quad (82.25)$$

$$= (\sqrt{P} \lambda) \cdot \sqrt{P} \quad (82.26)$$

$$\leq \sqrt{\sqrt{P} \lambda \cdot \sqrt{P} \lambda} \underbrace{\sqrt{\sqrt{P} \cdot \sqrt{P}}}_{=1} \quad (\text{by CS inequality}) \quad (82.27)$$

$$\leq \sqrt{\langle \lambda^2 \rangle_P} \quad (82.28)$$

When $P = P_N$,

$$\Psi[P_N] \approx \Psi[P_{in}] + \frac{1}{N} \sum_{\sigma} \lambda(X_{\sigma}) \quad (82.29)$$

If the random variables $\{X_{\sigma}\}_{\sigma \in \Sigma}$ are i.i.d. with probability distribution $P_{\underline{X}}(X)$, then, as $N \rightarrow \infty$, $\Psi[P_N]$ tends to a normally distributed random variable with mean

$$\langle \Psi[P_N] \rangle_{P_{\underline{X}}} = \Psi[P_{in}] + \frac{1}{N} \sum_{\sigma} \langle \lambda(X_{\sigma}) \rangle_{P_{\underline{X}}} \quad (82.30)$$

$$= \Psi[P_{in}] + \langle \lambda \rangle_{P_{\underline{X}}} \quad (82.31)$$

and variance

$$\langle \Psi[P_N], \Psi[P_N] \rangle_{P_{\underline{X}}} = \frac{1}{N^2} \sum_{\sigma} \sum_{\sigma'} \left\langle \lambda(X_{\sigma}), \lambda(X^{\sigma'}) \right\rangle_{P_{\underline{X}}} \quad (82.32)$$

$$= \frac{1}{N} \underbrace{\langle \lambda, \lambda \rangle_{P_{\underline{X}}}}_{\langle \lambda^2 \rangle_{P_{\underline{X}}} - \langle \lambda \rangle_{P_{\underline{X}}}^2} \quad (82.33)$$

Later on, we will discuss the so called TMLE estimate, for which P_{in} takes a special value P_{in++} that makes $\langle \lambda \rangle_{P_{\underline{X}}} = 0$, so $\Psi[P_N]$ tends to a normally distributed random variable with mean $\Psi[P_{in++}]$ and variance $\frac{1}{N} \langle \lambda^2 \rangle_{P_{\underline{X}}}$.

82.4 ATE estimand

If we set

$$\mathcal{Y}_{|d,x}[P] = \sum_y y P(y|d, x) = P(\underline{y} = 1|d, x) \quad (82.34)$$

and

$$\mathcal{Y}_{|d}[P] = \sum_x \mathcal{Y}_{|d,x}[P] P(x), \quad (82.35)$$

then the **Average Treatment Effect (ATE)** is defined as follows

$$ATE = \mathcal{Y}_{|1}[P_N] - \mathcal{Y}_{|0}[P_N] \quad (82.36)$$

The rest of this chapter is devoted to discussing ATE estimates.

In discussing ATE, it is convenient to define the **propensity** $g[P]$:

$$g[P](y) = P(\underline{d} = 1 | y) \quad (82.37)$$

and the Kronecker difference function $\Delta(d)$ for $d \in \{0, 1\}$:

$$\Delta(d) = \delta(d, 1) - \delta(d, 0) \quad (82.38)$$

$$= (2d - 1) \mathbb{1}(d \in \{0, 1\}). \quad (82.39)$$

Claim 122

$$\left\langle y \frac{\delta(d', d)}{P(d|x)} \right\rangle_P = \mathcal{Y}_{|d'}[P] \quad (82.40)$$

proof:

$$\left\langle y \frac{\delta(d', d)}{P(d|x)} \right\rangle_P = \sum_d \sum_y \sum_x P(y|d, x) \cancel{P(d|x)} P(x) y \frac{\delta(d', d)}{\cancel{P(d|x)}} \quad (82.41)$$

$$= \sum_y \sum_x P(y|\underline{d} = d', x) P(x) y \quad (82.42)$$

$$= \mathcal{Y}_{|\underline{d}=d'}[P] \quad (82.43)$$

QED

82.5 ATE estimates

82.5.1 Ψ^E

Empirical (E) estimate Ψ^E .

$$P_N(y, d, x) = \frac{1}{N} \sum_{\sigma} \delta(y, y_{\sigma}) \delta(d, d_{\sigma}) \delta(x, x_{\sigma}) \quad (82.44)$$

$$P_N(d, x) = \frac{1}{N} \sum_{\sigma} \delta(d, d_{\sigma}) \delta(x, x_{\sigma}) \quad (82.45)$$

$$P_N(x) = \frac{1}{N} \sum_{\sigma} \delta(x, x_{\sigma}) \quad (82.46)$$

$$P_N(y|d, x) = \frac{P_N(y, d, x)}{P_N(d, x)} \quad (82.47)$$

$$\Psi^E = \sum_x P_N(x) \sum_y y \sum_d P_N(y|d, x) \Delta(d) \quad (82.48)$$

82.5.2 Ψ^G

G estimate Ψ^G (a.k.a. g-computing or g-formula estimate).

Use Generalized Linear Modeling (GLM)² to approximate y_{σ} :

$$y_{\sigma} \approx E_{\underline{y}|d_{\sigma}, x_{\sigma}; \hat{\beta}}[\underline{y}] , \quad (82.49)$$

where $\hat{\beta}$ are the best curve fit parameters.

$$\Psi^G = \sum_x P_N(x) \{ E_{\underline{y}|d=1, x; \hat{\beta}}[\underline{y}] - E_{\underline{y}|d=0, x; \hat{\beta}}[\underline{y}] \} \quad (82.50)$$

82.5.3 Ψ^{IPW}

Inverse Propensity Weighted (IPW) estimate Ψ^{IPW} (a.k.a. Inverse Probability of Treatment Weighted (IPTW) estimate).

Assume propensity $P(\underline{d} = 1|x)$ is known.

Define

$$\Psi^{IPW}[P] = \left\langle y \frac{\Delta(d)}{P(d|x)} \right\rangle_P \quad (82.51)$$

$$\Psi^{IPW} = \Psi^{IPW}[P_N] = \left\langle y \frac{\Delta(d)}{P(d|x)} \right\rangle_{P_N} = \frac{1}{N} \sum_{\sigma} y_{\sigma} \frac{\Delta(d_{\sigma})}{P(d_{\sigma}|x_{\sigma})} \quad (82.52)$$

82.5.4 Ψ^{LIPW}

Linearized IPW (LIPW) estimate Ψ^{LIPW} .

Ψ^{LIPW} is the linear approximation of $\Psi^{IPW}[P_N]$ at point P_{in} :

$$\Psi^{LIPW} = \Psi^{IPW}[P_{in}] + \left\langle \nabla \Psi^{IPW}[P_{in}] \right\rangle_{P_N} - \nabla \Psi^{IPW}_{in} \quad (82.53)$$

²GLM is discussed in Chapter 31.

Claim 123

$$\nabla \Psi^{IPW}[P](X) = \mathcal{Y}_{|1,x}[P] - \mathcal{Y}_{|0,x}[P] + \frac{\Delta(d)}{P(d|x)}(y - \mathcal{Y}_{|d,x}[P]) \quad (82.54)$$

proof:

$$\frac{\delta \Psi^{IPW}[P]}{\delta P(X)} = \sum_{X'} y' \Delta(d') \frac{\delta}{\delta P(X)} \frac{P(X')}{P(d'|x')} \quad (82.55)$$

$$\frac{\delta}{\delta P(X)} \frac{P(X')}{P(d'|x')} = \frac{\delta(X, X')}{P(d'|x')} - \frac{P(X')}{[P(d'|x')]^2} \frac{\delta P(d'|x')}{\delta P(X)} \quad (82.56)$$

$$= \frac{\delta(X, X')}{P(d'|x')} - \frac{P(X')}{P(d'|x')} \frac{\delta \ln P(d'|x')}{\delta P(X)} \quad (82.57)$$

$$= \underbrace{\frac{\delta(X, X')}{P(d'|x')}}_{\delta^3/P(d'|x')} - P(y'|d', x') P(x') \frac{\delta \ln P(d'|x')}{\delta P(X)} \quad (82.58)$$

$$\frac{\delta \ln P(d', x')}{\delta P(X)} = \frac{1}{P(d', x')} \sum_{y'} \frac{\delta P(X')}{\delta P(X)} \quad (82.59)$$

$$= \underbrace{\frac{\delta(d, d') \delta(x, x')}{P(d', x')}}_{\delta^2/P(d', x')} \quad (82.60)$$

$$\frac{\delta \ln P(x')}{\delta P(X)} = \underbrace{\frac{\delta(x, x')}{P(x')}}_{\delta^1/P(x')} \quad (82.61)$$

$$\frac{\delta \ln P(d'|x')}{\delta P(X)} = \frac{\delta^2}{P(d', x')} - \frac{\delta^1}{P(x')} \quad (82.62)$$

$$\frac{\delta}{\delta P(X)} \frac{P(X')}{P(d'|x')} = \frac{\delta^3}{P(d'|x')} - P(y'|d', x') P(x') \left[\frac{\delta^2}{P(d', x')} - \frac{\delta^1}{P(x')} \right] \quad (82.63)$$

$$\sum_{X'} y' \Delta(d') \left[\frac{\delta^3}{P(d'|x')} \right] = \boxed{\frac{\Delta(d)}{P(d|x)} y} \quad (82.64)$$

$$\sum_{X'} y' \Delta(d') \left[\frac{-P(y'|d', x')\delta^2}{P(d'|x')} \right] = - \sum_{y'} y' \Delta(d) \frac{P(y'|d, x)}{P(d|x)} \quad (82.65)$$

$$= \boxed{\frac{\Delta(d)}{P(d|x)} (-\mathcal{Y}_{|d,x}[P])} \quad (82.66)$$

$$\sum_{X'} y' \Delta(d') [P(y'|d', x')\delta^1] = \sum_{y'} \sum_{d'} y' \Delta(d') P(y'|d', x) \quad (82.67)$$

$$= \boxed{\mathcal{Y}_{|1,x}[P] - \mathcal{Y}_{|0,x}[P]} \quad (82.68)$$

QED

Claim 124

$$\langle \nabla \Psi^{IPW}[P] \rangle_P = \Psi^{IPW}[P] \quad (82.69)$$

Hence,

$$\nabla \Psi_{in}^{IPW} = \langle \nabla \Psi^{IPW}[P_{in}] \rangle_{P_{in}} = \Psi^{IPW}[P_{in}] \quad (82.70)$$

proof:

$$\nabla \Psi^{IPW}[P](X) = \mathcal{Y}_{|1,x}[P] - \mathcal{Y}_{|0,x}[P] + \frac{\Delta(d)}{P(d|x)} (y - \mathcal{Y}_{|d,x}[P]) \quad (82.71)$$

$$\langle \mathcal{Y}_{|1,x}[P] - \mathcal{Y}_{|0,x}[P] \rangle_P = \sum_x P(x) (\mathcal{Y}_{|1,x}[P] - \mathcal{Y}_{|0,x}[P]) \quad (82.72)$$

$$= \Psi^{IPW}[P] \quad (82.73)$$

$$\left\langle \frac{\Delta(d)}{P(d|x)} y \right\rangle_P = \Psi^{IPW}[P] \quad (82.74)$$

$$\left\langle \frac{\Delta(d)}{P(d|x)} \mathcal{Y}_{|d,x}[P] \right\rangle_P = \sum_y \sum_x \sum_d P(y|d, x) P(x) \Delta(d) \mathcal{Y}_{|d,x}[P] \quad (82.75)$$

$$= \sum_x \sum_d P(x) \Delta(d) \mathcal{Y}_{|d,x}[P] \quad (82.76)$$

$$= \Psi^{IPW}[P] \quad (82.77)$$

QED

Note the following cancellation:

$$\Psi^{IPW}[P] = \Psi^{IPW}[P_{in}] + \langle \nabla \Psi^{IPW}[P_{in}] \rangle_P - \nabla \Psi_{in}^{IPW} + \mathcal{R}^{IPW}[P, P_{in}] \quad (82.78)$$

$$= \cancel{\Psi^{IPW}[P_{in}]} + \langle \nabla \Psi^{IPW}[P_{in}] \rangle_P - \cancel{\Psi^{IPW}[P_{in}]} + \mathcal{R}^{IPW}[P, P_{in}] \quad (82.79)$$

$$= \langle \nabla \Psi^{IPW}[P_{in}] \rangle_P + \mathcal{R}^{IPW}[P, P_{in}] \quad (82.80)$$

Claim 125

$$\mathcal{R}^{IPW}[P, P_{in}] = - \sum_x P(x) \sum_d \Delta(d) (P(d|x) - P_{in}(d|x)) \left(\frac{\mathcal{Y}_{|d,x}[P] - \mathcal{Y}_{|d,x}[P_{in}]}{P_{in}(d|x)} \right) \quad (82.81)$$

proof:

$$\mathcal{R}^{IPW}[P, P_{in}] = \Psi^{IPW}[P] - \langle \nabla \Psi^{IPW}[P_{in}] \rangle_P \quad (82.82)$$

$$= \left\langle y \frac{\Delta(d)}{P(d|x)} - \left(\mathcal{Y}_{|1,x}[P_{in}] - \mathcal{Y}_{|0,x}[P_{in}] + \frac{\Delta(d)}{P_{in}(d|x)} (y - \mathcal{Y}_{|d,x}[P_{in}]) \right) \right\rangle_P \quad (82.83)$$

$$= \begin{cases} \sum_x P(x) (-\mathcal{Y}_{|1,x}[P_{in}] + \mathcal{Y}_{|0,x}[P_{in}]) \\ + \sum_{d,x} P(d, x) \left(\frac{\Delta(d)}{P_{in}(d|x)} \mathcal{Y}_{|d,x}[P_{in}] \right) \\ + \sum_{y,d,x} P(y, d, x) \left(\frac{1}{P(d|x)} - \frac{1}{P_{in}(d|x)} \right) y \Delta(d) \end{cases} \quad (82.84)$$

$$= \sum_x P(x) \sum_d \Delta(d) \begin{cases} \left(\frac{P(d|x)}{P_{in}(d|x)} - 1 \right) \mathcal{Y}_{|d,x}[P_{in}] \\ + \sum_y P(y|d, x) \left(\frac{P_{in}(d|x) - P(d|x)}{P_{in}(d|x)} \right) y \end{cases} \quad (82.85)$$

$$= \sum_x P(x) \sum_d \Delta(d) \left(\frac{P(d|x) - P_{in}(d|x)}{P_{in}(d|x)} \right) (\mathcal{Y}_{|d,x}[P_{in}] - \mathcal{Y}_{|d,x}[P]) \quad (82.86)$$

QED

Claim 125 allows us to put a bound on the absolute value of the remainder \mathcal{R}^{IPW} :

$$|\mathcal{R}^{IPW}[P, P_{in}]| \leq \sum_d \sum_x \underbrace{\sqrt{P(x)} |P(d|x) - P_{in}(d|x)|}_{A(d,x)} \underbrace{\sqrt{P(x)} \left| \frac{\mathcal{Y}_{|d,x}[P] - \mathcal{Y}_{|d,x}[P_{in}]}{P_{in}(d|x)} \right|}_{B(d,x)}$$

(because $|\Delta(d)| = 1$, and $\left| \sum_i a_i \right| \leq \sum_i |a_i|$) (82.87)

$$\leq \sum_d \underbrace{\sqrt{\sum_x A^2(d,x)}}_{A_P(d)} \underbrace{\sqrt{\sum_x B^2(d,x)}}_{B_P(d)} \quad (\text{by CS inequality}) . \quad (82.88)$$

Define

$$A_P(d') = \sqrt{\sum_x P(x) (P(d'|x) - P_{in}(d'|x))^2} \quad (82.89)$$

$$= \sqrt{\langle (P(d'|x) - P_{in}(d'|x))^2 \rangle_P} \quad (82.90)$$

and

$$B_P(d') = \sqrt{\sum_x P(x) \left(\frac{\mathcal{Y}_{|d',x}[P] - \mathcal{Y}_{|d',x}[P_{in}]}{P_{in}(d'|x)} \right)^2} \quad (82.91)$$

$$= \sqrt{\left\langle \left(\frac{\mathcal{Y}_{|d',x}[P] - \mathcal{Y}_{|d',x}[P_{in}]}{P_{in}(d'|x)} \right)^2 \right\rangle_P} . \quad (82.92)$$

Then

$$|\mathcal{R}^{LIPW}[P_N, P_{in}]| \leq \sum_{d'=0}^1 A_{P_N}(d') B_{P_N}(d') \quad (82.93)$$

If either $A_{P_N}(1) = A_{P_N}(0) = 0$ (i.e., zero error in the propensities) or $B_{P_N}(0) = B_{P_N}(1) = 0$ (i.e., zero bias), then $\mathcal{R}^{LIPW}[P_N, P_{in}] = 0$. This property of Ψ^{LIPW} is referred to as **double robustness**.

82.5.5 Ψ^{LIPW++} (a.k.a. Ψ^{TMLE})

LIPW++ estimate Ψ^{LIPW++} (a.k.a, targeted minimum loss estimate (TMLE)) .

Ψ^{LIPW++} is the linear approximation of $\Psi^{IPW}[P_N]$ at the point P_{in++} , where the linear term of its Taylor expansion vanishes:

$$\Psi^{LIPW++} = \Psi^{TMLE} = \Psi^{IPW}[P_{in++}] + \underbrace{\langle \nabla \Psi^{IPW}[P_{in++}] \rangle_{P_N}}_{=0} - \nabla \Psi^{IPW}_{in++} \quad (82.94)$$

This property of Ψ^{TMLE} that the linear term in its Taylor expansion at P_{in++} vanishes is referred to as **substitution invariance**, and Ψ^{TMLE} is said to be a **substitution estimate**. A substitution estimate is very desirable because its absolute value is bounded, unlike the value of Ψ^{LIPW} .

Ψ^{TMLE} is both a doubly robust estimate and a substitution estimate.

Claim 123 allows us express more explicitly the constraint that defines P_{in++} :

$$0 = P_N \cdot \nabla \Psi^{IPW}[P_{in++}] - \nabla \Psi^{IPW}_{in++} \quad (82.95)$$

$$= -\nabla \Psi^{IPW}_{in++} + \left\{ \begin{array}{l} \overbrace{\frac{1}{N} \sum_{\sigma} (\mathcal{Y}_{|1,x_{\sigma}}[P_{in++}] - \mathcal{Y}_{|0,x_{\sigma}}[P_{in++}])}^{\approx \nabla \Psi^{IPW}[P_{in++}]} \\ + \frac{1}{N} \sum_{\sigma} \frac{\Delta(d_{\sigma})}{P_{in++}(d_{\sigma}|x_{\sigma})} (y_{\sigma} - \mathcal{Y}_{|d_{\sigma},x_{\sigma}}[P_{in++}]) \end{array} \right\} \quad (82.96)$$

$$= \frac{1}{N} \sum_{\sigma} \underbrace{\frac{\Delta(d_{\sigma})}{P_{in++}(d_{\sigma}|x_{\sigma})} (y_{\sigma} - \mathcal{Y}_{|d_{\sigma},x_{\sigma}}[P_{in++}])}_{\lambda(X_{\sigma})} \quad (82.97)$$

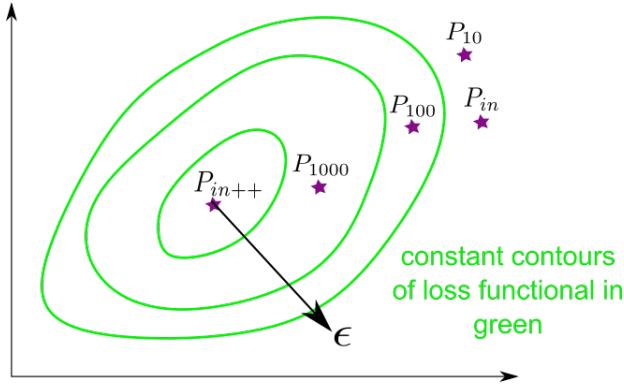


Figure 82.3: This figure portrays the space of functions \mathcal{H}_X as if it were the real plane \mathbb{R}^2 , and the functional $\mathcal{L}[P, \hat{y} = \text{fixed}] : \mathcal{H}_X \rightarrow \mathbb{R}$ as if it were a real valued function on \mathbb{R}^2 . It shows the constant contours of the loss functional $\mathcal{L}[P, \hat{y}]$ at fixed \hat{y} in green. P_N for $N = 10, 100, 1000$ represent empirical distributions. The loss is non-negative and it equals zero when $\epsilon = 0$ and $P = P_{in++}$.

Note that function $\lambda(X_{\sigma})$ defined in Eq.(82.97) can be positive or negative. Hence, it can't be defined as the loss cfit, because a loss cfit must be non-negative.

However, it can be defined as the derivative of a loss cfit. Suppose we define an $\epsilon \in \mathbb{R}$ parameterized family of probability distributions $P_\epsilon : S_{\underline{X}} \rightarrow [0, 1]$, and we expand the loss $\mathcal{L}[P_\epsilon, \hat{y}]$ in powers of ϵ :

$$\mathcal{L}[P_\epsilon, \hat{y}] = \mathcal{L}[P_0, \hat{y}] + \epsilon \{\partial_\epsilon \mathcal{L}[P_\epsilon, \hat{y}]\}_{\epsilon=0} + \mathcal{O}(\epsilon^2) \quad (82.98)$$

The parameter ϵ is called the **fluctuation parameter**. The function $\mathcal{L}[P_\epsilon, \hat{y}]$ is obviously not unique because all we know about it is the value of its ϵ derivative in the vicinity of $\epsilon = 0$. Next we will pick a convenient $\mathcal{L}[P_\epsilon, \hat{y}]$ that satisfies

$$\mathcal{L}[P_\epsilon, \hat{y}] \geq 0, \quad \mathcal{L}[P_0, \hat{y}] = 0, \quad \partial_\epsilon \{\mathcal{L}[P_\epsilon, \hat{y}]\}_{\epsilon=0} = 0 \quad (82.99)$$

Use as loss cfit $\hat{\mathcal{L}}$ the Cross Entropy $CE(p \parallel q)$ for $p, q \in [0, 1]$

$$\hat{\mathcal{L}} = CE(p \parallel q) = -[p \ln q + (1-p) \ln(1-q)] \quad (82.100)$$

$\hat{\mathcal{L}} \geq 0$ and attains its minimum when $p = q$. When $p = q$, it equals the entropy of p , i.e., when $p = q$, $\hat{\mathcal{L}} = -\sum_{x \in \{0,1\}} P(x) \ln P(x) = H(P)$, where $P(0) = p, P(1) = 1-p$.

For some $y \in \{0, 1\}$ and $\epsilon \in \mathbb{R}$, make the following substitutions in the loss cfit $\hat{\mathcal{L}}$ and call it $\hat{\mathcal{L}} = \hat{\mathcal{L}}(\beta, y, \hat{y}, \epsilon)$.³

$$p = y, \quad q = \text{expit}[\text{logit}(\hat{y}) + \epsilon\beta] \quad (82.101)$$

Note that since $p = y$ is binary, the minimum of this loss cfit is zero.

Recall that in Section C.20, we proved that the derivative of $\text{expit}(x)$ satisfies

$$\text{expit}'(x) = \text{expit}(x)[1 - \text{expit}(x)]. \quad (82.102)$$

Hence,

$$\lim_{\epsilon \rightarrow 0} \partial_\epsilon \hat{\mathcal{L}}(\beta, y, \hat{y}, \epsilon) = \lim_{\epsilon \rightarrow 0} \left[-\frac{p}{q} + \frac{1-p}{1-q} \right] \partial_\epsilon q \quad (82.103)$$

$$= \left[-\frac{y}{\hat{y}} + \frac{1-y}{1-\hat{y}} \right] \lim_{\epsilon \rightarrow 0} \left\{ \begin{array}{l} \text{expit}[\text{logit}(\hat{y}) + \epsilon\beta] \\ * \{1 - \text{expit}[\text{logit}(\hat{y}) + \epsilon\beta]\} \beta \end{array} \right\} \beta \quad (82.104)$$

$$= \left[-\frac{y}{\hat{y}} + \frac{1-y}{1-\hat{y}} \right] \hat{y}(1-\hat{y})\beta \quad (82.105)$$

$$= \beta[\hat{y} - y] \quad (82.106)$$

If we define $\hat{\mathcal{L}}(X)$ by

³To agree with the TE literature, we are using $\text{expit}(x)$ (resp., $\text{logit}(p)$) to denote the sigmoid function (resp., log-odds function) which we normally denote in this book by $\text{smoid}(x)$ (resp., $\text{lodds}(p)$).

$$\widehat{\mathcal{L}}(X) = \widehat{\mathcal{L}} \left(\begin{array}{l} \beta = \frac{\Delta(d)}{P_{in++}(d|x)}, \\ y = y, \\ \widehat{y} = \mathcal{Y}_{|d,x}[P_{in++}], \\ \epsilon = \epsilon \end{array} \right), \quad (82.107)$$

and \mathcal{L} by

$$\mathcal{L} = \frac{1}{N} \sum_{\sigma} \widehat{\mathcal{L}}(X_{\sigma}) = P_N \cdot \widehat{\mathcal{L}}, \quad (82.108)$$

then

$$\begin{cases} \mathcal{L} = P_N \cdot \widehat{\mathcal{L}} \geq 0 \\ \mathcal{L}_{\epsilon=0} = P_N \cdot \widehat{\mathcal{L}}_{\epsilon=0} = 0 \\ \{\partial_{\epsilon}\mathcal{L}\}_{\epsilon=0} = P_N \cdot \{\partial_{\epsilon}\widehat{\mathcal{L}}\}_{\epsilon=0} = 0 \end{cases} \quad (82.109)$$

82.6 Ψ^{TMLE} in practice

This section is based on Ref.[21].

In practice, one can calculate Ψ^{TMLE} by performing the following steps.

Below, “ML-fit” denotes a curve fitting obtained using any valid machine learning method, such as linear regression, a Neural Net, a decision tree, etc. The TE software often uses a “Super-Learner”, a program that merges the results of multiple fits obtained via various ML methods and also does cross validation.

Below, $\{(\sigma, d_{\sigma}, x_{\sigma}, \boxed{y_{\sigma}}) : \sigma \in \Sigma\}$ represents a dataset. The dependent variable y_{σ} is boxed, the independent ones d_{σ}, x_{σ} (a.k.a. covariates) aren’t.

1. Find cfit $\widehat{y}(d, x) = \mathcal{Y}_{|d,x}$ of outcome y

$$\{(\sigma, d_{\sigma}, x_{\sigma}, \boxed{y_{\sigma}}) : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \widehat{y}(d, x) \quad (82.110)$$

2. Estimate propensity $g(x)$

$$g(x) = P(\underline{d} = 1|x) \approx \frac{\sum_{\sigma} \delta(d_{\sigma}, 1)\delta(x_{\sigma}, x)}{\sum_{\sigma} \delta(x_{\sigma}, x)} \quad (82.111)$$

$$P(d|x) = dg(x) + (1 - d)[1 - g(x)] \quad (82.112)$$

$$\beta(d, x) = \frac{\Delta(d)}{P(d|x)} \quad (82.113)$$

3. Estimate fluctuation parameter ϵ

$$\eta(d, x) = \underbrace{\text{logit}[\hat{y}(d, x)]}_{\lambda(d, x)} + \epsilon \beta(d, x) \quad (82.114)$$

$$\{(\sigma, \lambda(d_\sigma, x_\sigma), \beta(d_\sigma, x_\sigma), \boxed{\eta(d_\sigma, x_\sigma)} : \sigma \in \Sigma\} \xrightarrow{\text{ML-fit}} \widehat{\eta}(d, x) \quad (82.115)$$

4. Estimate ATE

$$ATE = \frac{1}{N} \sum_{\sigma} \{\text{expit}[\widehat{\eta}(d = 1, x_\sigma)] - \text{expit}[\widehat{\eta}(d = 0, x_\sigma)]\} \quad (82.116)$$

Chapter 83

Time Series Analysis: ARMA and VAR

This chapter is based mostly on the book Ref.[16] on time series analysis by Hamilton, and on the lectures Ref. [26] by Chung-Ming Kuan. In writing this chapter, we also profited greatly from numerous Wikipedia entries on time series analysis, such as the entries on time series (Ref.[143]), ARMA time series (Ref.[77]), AR time series (Ref.[76]), MA time series (Ref.[125]), and VAR time series (Ref.[148]).

We cover only a small fraction of the treasures covered in those sources, and only cover stationary time-series. Non-stationary time series we don't even touch. But we hope to have covered enough to pique our readers's interest in time series analysis, and make him/her appreciate how bnets make time series much more intuitive and fun. The time-series considered in this chapter can be represented by one of the simplest types of bnets, namely, the LDEN bnets introduced in Chapter 44.

As usual, for $t, t_a, t_b \in \mathbb{Z}$, let $\mathbb{Z}_{<t} = \{t-1, t-2, t-3, \dots\}$, $\mathbb{Z}_{[t_a, t_b]} = \{t_a, t_a+1, \dots, t_b\}$, etc.

Let $\underline{x}_t \in \mathbb{R}$. A **time series (t-series)**, denoted variously by $\{\underline{x}_1, \underline{x}_2, \dots, \underline{x}_{nt}\} = \{\underline{x}_t\}_{t=1}^{nt} = \{\underline{x}_t\}_{t \in \mathbb{Z}}$, is a set of real numbers index by a discrete set of times $\mathbb{Z}_{[0, nt]}$.

For $t_a < t_b$, let $\underline{x}_{[t_a, t_b]} = (\underline{x}_{t_a}, \underline{x}_{t_a+1}, \dots, \underline{x}_{t_b})$. Let $x_{<t} = (\dots, x_{t-2}, \underline{x}_{t-1})$.

83.1 White noise

By **white noise** $\{\underline{n}_t\}_{\forall t} \sim WN(0, \sigma^2)$ we mean a t-series $\{\underline{n}_t\}_{\forall t}$ that satisfies

$$E[\underline{n}_t] = 0 \tag{83.1}$$

and

$$\langle \underline{n}_t, \underline{n}_{t'} \rangle = \sigma^2 \delta(t, t') . \tag{83.2}$$

Gaussian white noise $\{\underline{n}_t\}_{\forall t} \sim WN(0, \sigma^2)_{\mathcal{N}}$ is white noise such that also $\underline{n}_t \sim \mathcal{N}(0, \sigma^2)$.

83.2 Backshift operator

\mathcal{B} is the backshift (a.k.a. lag) operator. For any t-series $\{\underline{x}_t\}_{\forall t}$, $\{\underline{y}_t\}_{\forall t}$ and scalars $a, b \in \mathbb{R}$,

$$\mathcal{B}\underline{x}_t = \underline{x}_{t-1} \quad (83.3)$$

$$\mathcal{B}(ax_t + by_t) = a\mathcal{B}(x_t) + b\mathcal{B}(y_t) \quad (83.4)$$

A \mathcal{B} Polynomial with coefficients $\alpha_{[0,p]}$:

$$\alpha(\mathcal{B}) = \alpha_0 + \alpha_1\mathcal{B} + \alpha_2\mathcal{B}^2 + \dots + \alpha_p\mathcal{B}^p \quad (83.5)$$

\mathcal{B}^{-1} is the inverse of the backshift operator (a.k.a. frontshift operator)

$$\mathcal{B}^{-1}\underline{x}_t = \underline{x}_{t+1} \quad (83.6)$$

The following two Taylor expansions prove useful in finding the inverse of backshift operator polynomials:

- $$\frac{1}{1-z} = 1 + z + z^2 + \dots \quad (83.7)$$

converges for $z \in \mathbb{C}$ with $|z| < 1$. We will use this expansion with z replaced by $\alpha\mathcal{B}$, where $\alpha \in \mathbb{R}$.

- $$\frac{1}{1-z} = (-z^{-1}) \left[\frac{1}{1-z^{-1}} \right] = (-z^{-1}) [1 + z^{-1} + z^{-2} + \dots] \quad (83.8)$$

converges for $z \in \mathbb{C}$ with $|z| > 1$. We will use this expansion with z^{-1} replaced by $(\alpha\mathcal{B})^{-1}$, where $\alpha \in \mathbb{R}$.

83.3 Metrics

Consider a t-series $\{\underline{x}_t\}_{\forall t}$.

In general, if we have a metric like Auto-covariance (ACov) that is defined for $\tau = 1, 2, 3, \dots$, it is conventional in time series analysis to refer to the plot of that metric for all values of τ as the Auto-covariance *Function* (ACovF).

- Expected value and Variance

$$E[\underline{x}_t] \quad (83.9)$$

$$Var[\underline{x}_t] = \langle \underline{x}_t, \underline{x}_t \rangle \quad (83.10)$$

- **Auto-covariance (ACov)**

$$\gamma_{t,t+\tau} = \langle \underline{x}_t, \underline{x}_{t+\tau} \rangle \quad (83.11)$$

- **Auto-correlation (ACorr)** (assumes w-stationarity)

$$\rho(\tau) = \frac{\gamma(\tau)}{\gamma(0)} \quad (83.12)$$

- **Generating function of auto-covariance** (assumes w-stationarity)

$$\tilde{\gamma}(z) = \sum_{\tau=-\infty}^{\infty} \gamma(\tau) z^{\tau} \quad (83.13)$$

Note that this transform is double sided. Fourier Transform if $z = e^{-i\omega\tau}$.

- **Expected value and variance conditioned on all past information**

For $\tau = 1, 2, 3, \dots$,

$$E_{|x_{\leq t}}[\underline{x}_{t+\tau}] \quad (83.14)$$

$$Var_{|x_{\leq t}}[\underline{x}_{t+\tau}] = \langle \underline{x}_{t+\tau}, \underline{x}_{t+\tau} \rangle_{|x_{\leq t}} \quad (83.15)$$

- **Partial auto-covariance (PACov)**

Assume w-stationarity. For $\tau = 1, 2, 3, \dots$

$$\gamma^{part}(\tau) = \langle \underline{x}_t, \underline{x}_{t+\tau} \rangle_{|x_{\leq t}, \underline{x}_{t+\tau}} \quad (83.16)$$

The idea is that we set to zero the nodes $\underline{x}_{(t,t+\tau)} = \{\underline{x}_{t+1}, \underline{x}_{t+2}, \dots, \underline{x}_{t+\tau-1}\}$ that lie between (but not including) \underline{x}_t and $\underline{x}_{t+\tau}$.

- **Partial auto-correlation (PACorr)**

$$\rho^{part}(\tau) = \frac{\gamma^{part}(\tau)}{\gamma^{part}(0)} \quad (83.17)$$

weak stationarity (w-stationarity) means that $E[\underline{x}_t] = \mu$ and $\gamma_{t,t+\tau} = \gamma(\tau)$ are both independent of t . If we have w-stationarity, then

$$\gamma(-\tau) = \langle \underline{x}_t, \underline{x}_{t-\tau} \rangle \quad (83.18)$$

$$= \langle \underline{x}_{t-\tau}, \underline{x}_t \rangle \quad (83.19)$$

$$= \gamma(\tau) \quad (83.20)$$

We will often abbreviate $\gamma(\tau)$ by γ_τ .

Example of various metrics. If $\{\underline{n}_t\}_{\forall t} \sim WN(0, \sigma^2)$ then

$$E[\underline{n}_t] = 0 \quad (83.21a)$$

$$\gamma(\tau) = \sigma^2 \delta(\tau, 0) \quad (83.21b)$$

$$\gamma(0) = \sigma^2 \quad (83.21c)$$

$$\tilde{\gamma}(z) = \gamma(0) \quad (83.21d)$$

For $\tau > 0$,

$$E_{|\underline{n}_{\leq t}}[\underline{n}_{t+\tau}] = E[\underline{n}_{t+\tau}] = 0 \quad (83.21e)$$

$$\langle \underline{n}_{t+\tau}, \underline{n}_{t+\tau} \rangle_{|\underline{n}_{\leq t}} = E[\underline{n}_{t+\tau}^2] = \sigma^2 \quad (83.21f)$$

83.4 Definition of $ARMA(p, q)$, $AR(p)$ and $MA(q)$.

Suppose $\{\underline{y}_t\}_{\forall t}$ is a zero mean t-series. Hence $\underline{y}_t = \underline{Y}_t - \mu$, $E[\underline{Y}_t] = \mu$, $E[\underline{y}_t] = 0$. \underline{y}_t is said to be the **demeaned** version of \underline{Y}_t .

Suppose also that $\{\underline{n}_t\}_{\forall t} \sim WN(0, \sigma^2)$.

Then we define the **Auto-Regressive Moving-Average t-series** $ARMA(p, q)$ by

$$\underline{y}_t = \underbrace{\sum_{j=1}^p \alpha_j \underline{y}_{t-j}}_{\mathcal{Y}_t^{AR(p)}} + \underline{n}_t + \underbrace{\sum_{j=1}^q \nu_j \underline{n}_{t-j}}_{\mathcal{Y}_t^{MA(q)}} \quad (83.22)$$

(α stands for the first letter of “auto-regressive”. ν stands for first letter of “noise”.)

Special cases

1. Auto-Regressive t-series $AR(p)$

$$\underline{y}_t = \mathcal{Y}_t^{AR(p)} + \underline{n}_t \quad (83.23)$$

2. Moving-Average t-series $MA(q)$

$$\underline{y}_t = \underline{n}_t + \mathcal{Y}_t^{MA(q)} \quad (83.24)$$

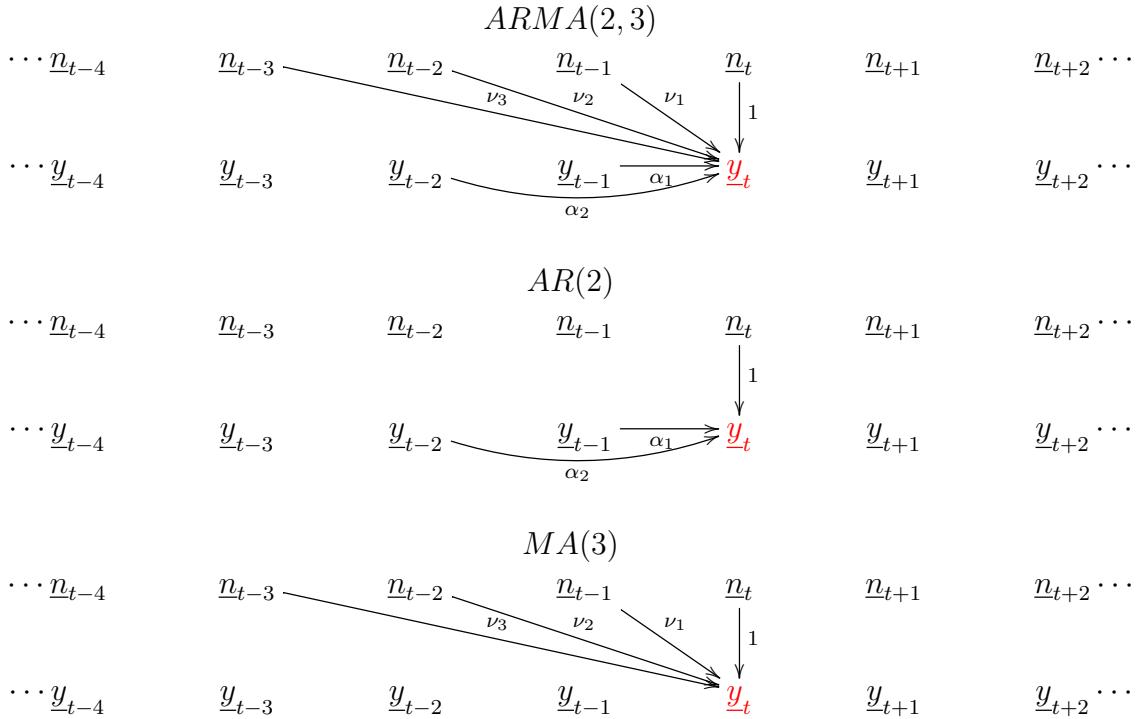


Figure 83.1: $ARMA(2,3)$, $AR(2)$ and $MA(3)$ bnets. For clarity, we show only the arrows entering node \underline{y}_t . The full bnet has the same structural pattern of incoming arrows (including the weights α_j, ν_j) for each node $\underline{y}_{t'}$ for all t' .

Fig.83.1 shows the bnets for $ARMA(2,3)$, $AR(2)$ and $MA(3)$. The TPM, printed in blue, for node \underline{y}_t in those bnets, is as follows:

For $AR(p)$,

$$P(y_t | y_{[t-p, t-1]}, n_{[t-q, t]}) = \mathbb{1}(y_t = \text{see Eq.83.22})) \quad (83.25)$$

For $AR(p)$,

$$P(y_t | y_{[t-p, t-1]}, n_t) = \mathbb{1}(y_t = \text{see Eq.83.23})) \quad (83.26)$$

For $MA(q)$,

$$P(y_t | n_{[t-q, t]}) = \mathbb{1}(y_t = \text{see Eq.83.24})) \quad (83.27)$$

The \underline{n}_t variable is variously referred to as the **external noise**, **impulse**, **shock**, **innovation** at time t .

83.5 Solving $AR(p)$

Suppose $\{\underline{y}_t\}_{\forall t}$ is an $AR(p)$ t-series. Hence

$$\underline{y}_t = \sum_{j=1}^p \alpha_j \underline{y}_{t-j} + \underline{n}_t \quad (83.28)$$

$AR(0)$ satisfies:

$$\underline{y}_t = \underline{n}_t \quad (83.29)$$

See Fig.83.2. This is just white noise.

$AR(1)$ satisfies:

$$\underline{y}_t = \alpha_1 \underline{y}_{t-1} + \underline{n}_t \quad (83.30)$$

See Fig.83.2. This is a Markov chain with external i.i.d. noise injected to each node. $AR(1)$ is the discrete form of the so called **Ornstein-Uhlenbeck t-series** (a.k.a. as the **Langevin Equation**). When $\alpha_1 = 1$, it is called a **random walk**.

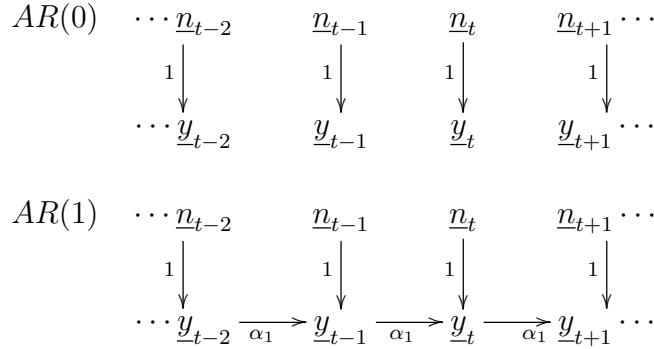


Figure 83.2: BNets for $AR(0)$ and $AR(1)$.

Note that

$$\alpha^-(\mathcal{B}) \underline{y}_t = \underline{n}_t \quad (83.31)$$

where

$$\alpha^-(\mathcal{B}) = 1 - \sum_{j=1}^p \alpha_j \mathcal{B}^j \quad (83.32)$$

Note that we can get $AR(p)$ from $AR(\infty)$ by setting $\alpha_{>p} = 0$.

If $\alpha^-(\beta)$ is invertible, then, using the Taylor expansion Eq.(83.7), we get

$$\underline{y}_t = \alpha'(\mathcal{B}) \underline{n}_t \quad (83.33)$$

where

$$\alpha'(\mathcal{B}) = \frac{1}{\alpha^-(\mathcal{B})} = 1 + \sum_{k=1}^{\infty} \left[\sum_{j=1}^p \alpha_j \mathcal{B}^j \right]^k = \sum_{j=0}^{\infty} \alpha'_j \mathcal{B}^j \quad (83.34)$$

where $\alpha'_0 = 1$.

83.6 Solving $MA(q)$

Suppose $\{\underline{y}_t\}_{\forall t}$ is an $MA(q)$ t-series. Hence,

$$\underline{y}_t = \underline{n}_t + \sum_{j=1}^q \nu_j \underline{n}_{t-j} = \sum_{j=0}^q \nu_j \underline{n}_{t-j} \quad (83.35)$$

where $\nu_0 = 1$. Thus,

$$\underline{y}_t = \nu(\mathcal{B}) \underline{n}_t \quad (83.36)$$

where

$$\nu(\mathcal{B}) = 1 + \sum_{j=1}^q \nu_j \mathcal{B}^j = \sum_{j=0}^q \nu_j \mathcal{B}^j . \quad (83.37)$$

Note that we can get $MA(q)$ from $MA(\infty)$ by setting $\nu_{>q} = 0$.

Claim 126 *If $\{\underline{y}_t\}_{\forall t}$ is an $MA(q)$ t-series, then*

$$E[\underline{y}_t] = 0 \quad (83.38a)$$

For $\tau \geq 0$,

$$\gamma(\tau) = \mathbb{1}(\tau \leq q) \sigma^2 \sum_{j=0}^{q-\tau} \nu_j \nu_{\tau+j} \quad (83.38b)$$

$$\gamma(0) = \sigma^2 \sum_{j=0}^q \nu_j^2 \quad (83.38c)$$

proof:

$$\gamma(\tau) = \left\langle \underline{y}_t, \underline{y}_{t+\tau} \right\rangle \quad (83.39)$$

$$= \left\langle \sum_{j=0}^q \nu_j \underline{n}_{t-j}, \sum_{k=0}^q \nu_k \underline{n}_{t+\tau-k} \right\rangle \quad (83.40)$$

$$= \sigma^2 \sum_{j=0}^q \sum_{k=0}^q \nu_j \nu_k \delta(t-j, t+\tau-k) \quad (83.41)$$

$$= \sigma^2 \sum_{j=0}^q \sum_{k=0}^q \nu_j \nu_k \delta(k, \tau+j) \quad (83.42)$$

$$= \mathbb{1}(\tau \leq q) \sigma^2 \sum_{j=0}^{q-\tau} \nu_j \nu_{\tau+j} \quad (83.43)$$

QED

83.7 Solving $ARMA(p, q)$

Suppose $\{\underline{y}_t\}_{\forall t}$ is an $ARMA(p, q)$ t-series. Hence, using Eqs.(83.31) and (83.36), \underline{y}_t satisfies

$$\alpha^-(\mathcal{B}) \underline{y}_t = \nu(\mathcal{B}) \underline{n}_t \quad (83.44)$$

If $\alpha^-(\beta)$ is invertible, then, using the Taylor expansion Eq.(83.7), we get

$$\underline{y}_t = \frac{\nu(\mathcal{B})}{\alpha^-(\mathcal{B})} \underline{n}_t = \nu(\mathcal{B}) \alpha'(\mathcal{B}) \underline{n}_t \quad (83.45)$$

We see that, if the $\alpha^-(\mathcal{B})$ operator is invertible, an $AR(p)$ or an $ARMA(p, q)$ t-series can be represented as an $MA(\infty)$ t-series. $MA(\infty)$ is often called **Wold's Decomposition**. Furthermore, if the $\nu(\mathcal{B})$ operator is invertible, an $MA(q)$ or an $ARMA(p, q)$ t-series can be represented as an $AR(\infty)$ t-series.

The polynomials $\alpha^-(z)$ and $\nu(z)$ can be expressed in factored form $\alpha^-(z) = \prod_{j=1}^p (z - z_j^\alpha)$ and $\nu(z) = \prod_{j=1}^q (z - z_j^\nu)$. If these two polynomials have a root z_0 in common, both polynomials should be divided by $(z - z_0)$. This reduces an $ARMA(p, q)$ t-series to an $ARMA(p-1, q-1)$ t-series. The bnet for $ARMA(p-1, q-1)$ has one less α_j arrow and one less ν_j arrow than the bnet for $AR(p, q)$.

83.8 Auto-correlation and partial auto-correlation

Note from Eq.(83.38b) that if $\{\underline{y}_t\}_{\forall t}$ is a $MA(q)$ t-series, then $\gamma(\tau) = 0$ for all $\tau > q$. Fig.83.3 gives a graphical proof, using bnets and the d-separation theorem, that for

a $MA(2)$ t-series, $\gamma(\tau) = 0$ for $\tau > 2$. As a consequence of this, a plot of $\gamma(\tau)$ versus τ for a typical $MA(2)$ t-series looks like Fig.83.4.

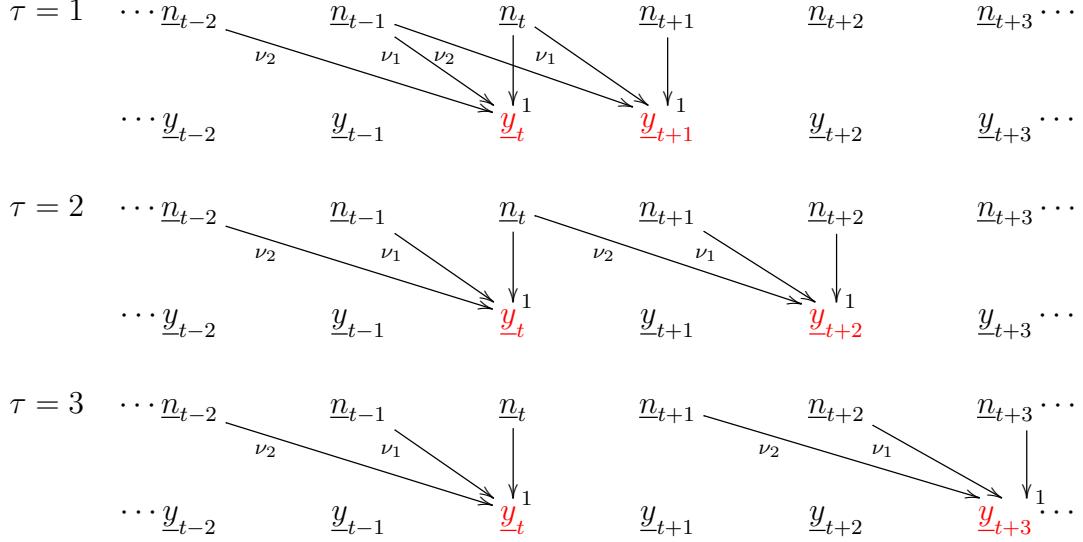


Figure 83.3: $MA(2)$ bnet. For clarity, we show only arrows entering nodes \underline{y}_t and $\underline{y}_{t+\tau}$ for $\tau = 1, 2, 3$. For $\tau = 1, 2$, there is a path through which information can flow from node \underline{y}_t to node $\underline{y}_{t+\tau}$. For $\tau = 2$, there is no such path.

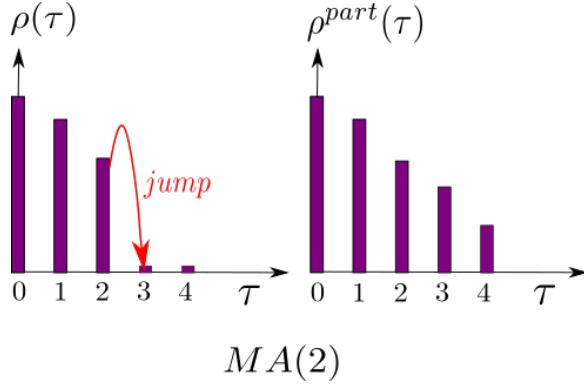


Figure 83.4: Plot of auto-correlation function (ACorrF) $\rho(\tau)$ and partial auto-correlation function (PACorrF) $\rho^{part}(\tau)$ for an instance of $MA(2)$. Note that $\rho(\tau)$ vanishes for $\tau > 2$.

Claim 127 If $\{\underline{y}_t\}_{\forall t}$ is an $AR(p)$ t-series, then $\gamma^{part}(\tau) = 0$ for all $\tau > p$.

proof:

Let

$$\underline{\xi} = \underline{y}_{\leq t}, \underline{y}_{t+\tau} \quad (83.46)$$

Recall that

$$\gamma^{part}(\tau) = \left\langle \underline{y}_t, \underline{y}_{t+\tau} \right\rangle_{|\xi} \quad (83.47)$$

$$= \left\langle \underline{y}_t \underline{y}_{t+\tau} \right\rangle_{|\xi} - \left\langle \underline{y}_t \right\rangle_{|\xi} \left\langle \underline{y}_{t+\tau} \right\rangle_{|\xi} \quad (83.48)$$

$$= \left\langle \underline{y}_t \underline{y}_{t+\tau} \right\rangle_{|\xi} \quad (83.49)$$

Define

$$Z_{(t,t+\tau)} = \mathbb{1}(\underline{y}_{(t,t+\tau)} = 0) \quad (83.50)$$

Hence, the operator $Z_{(t,t+\tau)}$ sets all $\underline{y}_{t'}$ with $t' < t + \tau$ equal to zero. Note that we can express the PACov as

$$\gamma^{part}(\tau) = \left\langle Z_{(t,t+\tau)} (\underline{y}_t \underline{y}_{t+\tau}) \right\rangle. \quad (83.51)$$

Since

$$\underline{y}_{t+\tau} = \sum_{j=1}^p \alpha_j \underline{y}_{t+\tau-j} + \underline{n}_{t+\tau} \quad (83.52)$$

$$= \alpha_p \underline{y}_{t+\tau-p} + \dots + \alpha_2 \underline{y}_{t+\tau-2} + \alpha_1 \underline{y}_{t+\tau-1} + \underline{n}_{t+\tau}, \quad (83.53)$$

we get

$$Z_{(t,t+\tau)} \underline{y}_{t+\tau} = \underline{n}_{t+\tau} \quad \text{if } \tau > p. \quad (83.54)$$

Hence, for $\tau > p$,

$$\gamma^{part}(\tau) = \left\langle \underline{y}_t \underline{n}_{t+\tau} \right\rangle = 0 \quad (83.55)$$

QED

Fig.83.5 gives a graphical proof, using bnets and the d-separation theorem, that for an AR(2) t-series, $\gamma^{part}(\tau) = 0$ for $\tau > 2$. As a consequence of this, a plot of $\gamma^{part}(\tau)$ versus τ for a typical AR(2) t-series looks like Fig.83.6.

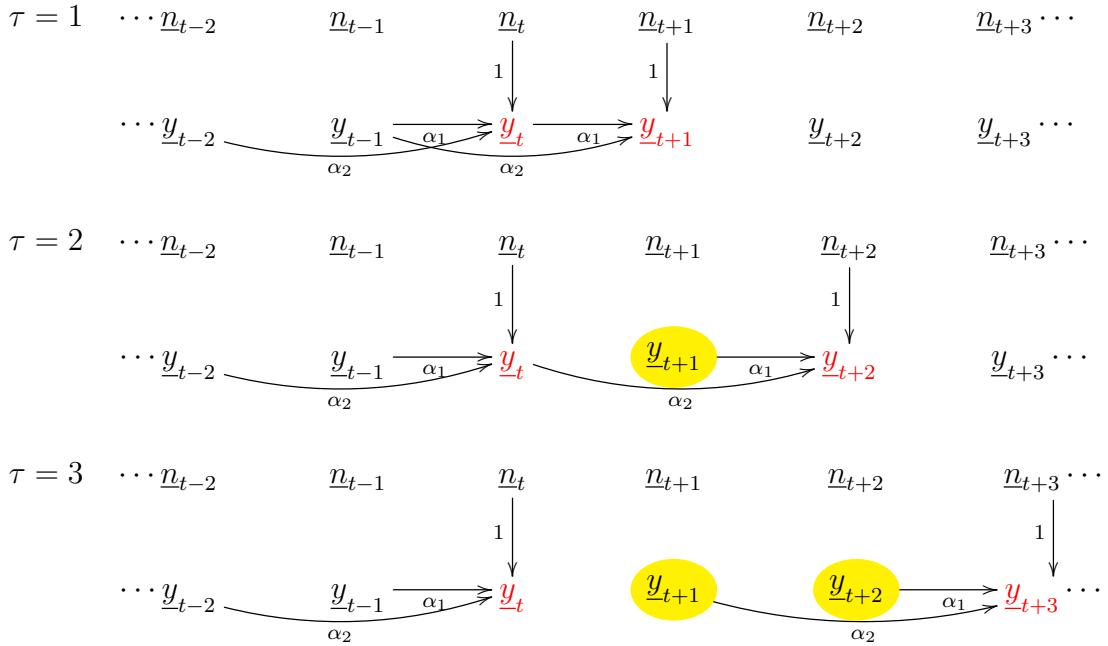


Figure 83.5: $AR(2)$ bnet. For clarity, we show only arrows entering nodes \underline{y}_t and $\underline{y}_{t+\tau}$ for $\tau = 1, 2, 3$. Yellow nodes are in set $\underline{y}_{(t,t+\tau)}$. They are conditioned on, so information can't flow through them to node $\underline{y}_{t+\tau}$ by the d-separation theorem.

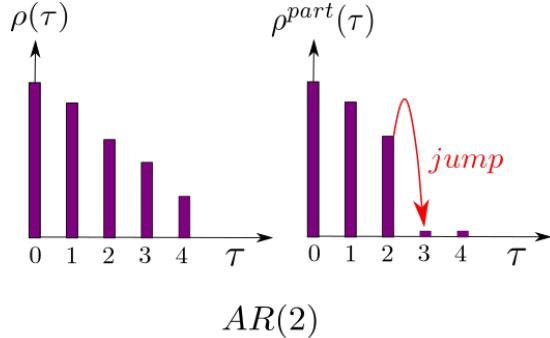


Figure 83.6: Plot of auto-correlation function (ACorrF) $\rho(\tau)$ and partial auto-correlation function (PACorrF) $\rho^{part}(\tau)$ for an instance of $AR(2)$. Note that $\rho^{part}(\tau)$ vanishes for $\tau > 2$.

	ACorr	PACorr
$AR(p)$	tapers off	jumps to zero for $\tau > p$
$MA(q)$	jumps to zero for $\tau > q$	tapers off

Table 83.1: Detecting $AR(p)$ and $MA(q)$ using auto-correlation and partial auto-correlation.

83.9 Generating function of auto-correlation

Claim 128 *If*

$$\alpha(z) = \sum_{\tau=-\infty}^{\infty} \alpha_{\tau} z^{\tau} \quad (83.56)$$

then

$$\alpha(z)\alpha(z^{-1}) = \sum_{\tau=-\infty}^{\infty} z^{\tau} \sum_{j=-\infty}^{\infty} \alpha_j \alpha_{j+\tau} \quad (83.57)$$

proof:

$$\alpha(z)\alpha(z^{-1}) = \sum_{j'=-\infty}^{\infty} z^{j'} \alpha_{j'} \sum_{j=-\infty}^{\infty} z^{-j} \alpha_j \quad (83.58)$$

$$= \sum_{\tau=-\infty}^{\infty} \sum_{j'=-\infty}^{\infty} \sum_{j=-\infty}^{\infty} z^{\tau} \alpha_{j'} \alpha_j \delta(j' - j, \tau) \quad (83.59)$$

$$= \sum_{\tau=-\infty}^{\infty} z^{\tau} \sum_{j=-\infty}^{\infty} \alpha_j \alpha_{j+\tau} \quad (83.60)$$

QED

As an example of this claim, note that

$$(\alpha_0 + \alpha_1 z)(\alpha_0 + \alpha_1 z^{-1}) = \alpha_0 \alpha_1 z^{-1} + (\alpha_0^2 + \alpha_1^2) + \alpha_0 \alpha_1 z \quad (83.61)$$

Claim 129 *If $\{\underline{y}_t\}_{\forall t}$ is an $AR(p)$ t-series, then*

$$\tilde{\gamma}(z) = \alpha'(z) \sigma^2 \alpha'(z^{-1}) \quad (83.62)$$

proof: Left to reader. See Claim 128.

QED

Claim 130 *If $\{\underline{y}_t\}_{\forall t}$ is an $MA(q)$ t-series, then*

$$\tilde{\gamma}(z) = \nu(z) \sigma^2 \nu(z^{-1}) \quad (83.63)$$

proof:

$$\underline{y}_t = \sum_{j=-\infty}^{\infty} \nu_j \underline{n}_{t-j} \quad (83.64)$$

$$\tilde{\gamma}(z) = \sum_{\tau=-\infty}^{\infty} \langle \underline{y}_0, \underline{y}_{\tau} \rangle z^{\tau} \quad (83.65)$$

$$= \sum_{\tau=-\infty}^{\infty} z^{\tau} \sum_{j=-\infty}^{\infty} \sum_{j'=-\infty}^{\infty} \nu_j \nu_{j'} \underbrace{\langle \underline{n}_{-j}, \underline{n}_{\tau-j'} \rangle}_{\sigma^2 \delta(j', \tau+j)} \quad (83.66)$$

$$= \sigma^2 \sum_{\tau=-\infty}^{\infty} z^{\tau} \sum_{j=-\infty}^{\infty} \nu_j \nu_{j+\tau} \quad (83.67)$$

Now use Claim 128.

QED

Claim 131 *If $\{\underline{y}_t\}_{\forall t}$ is an ARMA(p, q) t-series, then*

$$\tilde{\gamma}(z) = \alpha'(z) \nu(z) \sigma^2 \nu(z^{-1}) \alpha'(z^{-1}) \quad (83.68)$$

proof: Left to reader. See Claim 128.

QED

83.10 Impulse Response

The derivatives

$$IR_{\tau} = \frac{\partial y_{t+\tau}}{\partial n_t} \quad (83.69)$$

are called **impulse responses** or **dynamical multipliers**. Note that this derivative depends on τ but not t by w-stationarity. A plot of IR_{τ} versus τ is called the **Impulse Response Function (IRF)**. Examples:

- **Claim 132** *For MA(q),*

$$\frac{\partial y_{t+\tau}}{\partial n_t} = \nu_{\tau} \mathbb{1}(0 \leq \tau \leq q) \quad (83.70)$$

where $\nu_0 = 1$. (See Fig.83.7)

proof:

$$\underline{y}_{t+\tau} = \sum_{j=0}^q \nu_j \underline{n}_{t+\tau-j} \quad (83.71)$$

so Eq.(83.70) follows.

QED

- **Claim 133** *For AR(1),*

$$\frac{\partial y_{t+\tau}}{\partial n_t} = (\alpha_1)^\tau \mathbb{1}(\tau \geq 0). \quad (83.72)$$

(See Fig.83.8)

proof:

$$(1 - \alpha_1 \mathcal{B}) \underline{y}_{t+\tau} = \underline{n}_{t+\tau} \quad (83.73)$$

Therefore, using the Taylor expansion Eq.(83.7),

$$\underline{y}_{t+\tau} = 1 + \sum_{j=1}^{\infty} \alpha_1^j \mathcal{B}^j \underline{n}_{t+\tau} \quad (83.74)$$

$$= 1 + \sum_{j=1}^{\infty} \alpha_1^j \underline{n}_{t+\tau-j} \quad (83.75)$$

so Eq.(83.72) follows.

QED

In general, IR_τ equals a sum over paths from \underline{n}_t to $\underline{y}_{t+\tau}$. Each path contributes the product of the weights of all the arrows in the path.

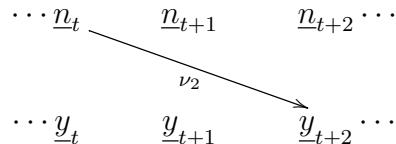


Figure 83.7: Pictorial representation of impulse response $IR_2 = \nu_2$ for a $MA(q)$ t-series with $q \geq 2$.

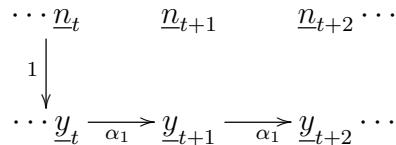


Figure 83.8: Pictorial representation of impulse response $IR_2 = \alpha_1^2$ for an $AR(1)$ t-series.

83.11 AR(p) and Yule-Walker equations

Claim 134 (*Yule-Walker equations*) If $\{\underline{y}_t\}_{\forall t}$ is an AR(p) t-series,

$$\gamma_\tau = \sum_{j=1}^p \gamma_{\tau-j} \alpha_j + \sigma^2 \delta(\tau, 0) \quad (83.76)$$

for all $\tau \in \mathbb{Z}$, where we are abbreviating $\gamma(j) = \gamma_j$ for all j .

proof:

Recall that

$$\underline{y}_\tau = \sum_{j=1}^p \underline{y}_{\tau-j} \alpha_j + \underline{n}_\tau \quad (83.77)$$

Therefore

$$\underbrace{\langle \underline{y}_0, \underline{y}_\tau \rangle}_{\gamma_\tau} = \sum_{j=1}^p \underbrace{\langle \underline{y}_0, \underline{y}_{\tau-j} \rangle}_{\gamma_{\tau-j}} \alpha_j + \underbrace{\langle \underline{y}_0, \underline{n}_\tau \rangle}_{\sigma^2 \delta(\tau, 0)} \quad (83.78)$$

QED

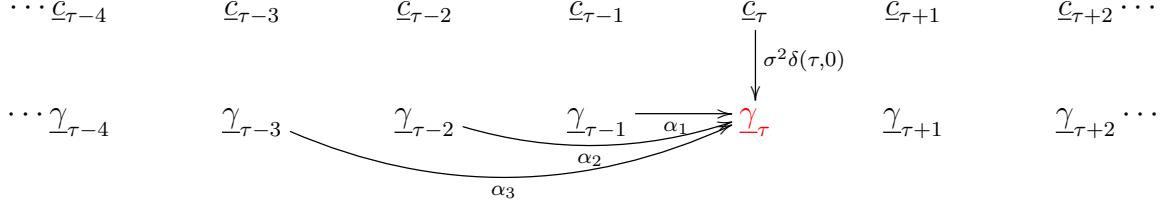


Figure 83.9: Bnet representing Yule-Walker (Christmas Walker) Eqs.(83.76) with $p = 3$. For clarity, we show only arrows entering node γ_τ . The full bnet has the same structural pattern of incoming arrows (including the α_j) for each node γ_τ' for all τ' .

Fig.83.9 shows a bnet for the Yule-Walker Eqs.(83.76). The TPMs, printed in blue, for the τ time slice of that bnet, are as follows:

$$P(\gamma_\tau) = \mathbb{1} (\gamma_\tau = \text{ See Eq.(83.76)}) \quad (83.79)$$

$$P(c_\tau) = \mathbb{1}(c_\tau = \sigma^2 \delta(\tau, 0)) \quad (83.80)$$

The Yule-Walker Eqs.(83.76) imply that

$$\underbrace{\begin{bmatrix} \gamma_1 \\ \gamma_2 \\ \vdots \\ \gamma_p \end{bmatrix}}_{\gamma^p} = \underbrace{\begin{bmatrix} \gamma_0 & \gamma_{-1} & \gamma_{-2} & \cdots & \gamma_{1-p} \\ \gamma_1 & \gamma_0 & \gamma_{-1} & \cdots & \gamma_{2-p} \\ \vdots & & & & \\ \gamma_{p-1} & \gamma_{p-2} & \gamma_{p-3} & \cdots & \gamma_0 \end{bmatrix}}_{\Gamma} \underbrace{\begin{bmatrix} \alpha_1 \\ \alpha_2 \\ \vdots \\ \alpha_p \end{bmatrix}}_{\alpha^p} \quad (83.81)$$

Hence

$$\gamma^p = \Gamma \alpha^p . \quad (83.82)$$

If Γ is invertible,

$$\alpha^p = \Gamma^{-1} \gamma^p . \quad (83.83)$$

Once we know α^p , we can solve for σ^2

$$\gamma_0 = \sum_{j=1}^p \gamma_{-j} \alpha_j + \sigma^2 , \quad (83.84)$$

i.e., Eq.(83.76) for $\tau = 0$.

83.12 Forecasting

Before submerging the reader in the messy details of t-series forecasting, and running the risk of quickly losing him/her, let me explain to the reader that what we are about to do in this section, is, in the final analysis, quite trivial, and boils down to solving simple systems of linear equations. That is all there is to it!! How hard can that be? In reading this section, don't lose sight of that fact and you'll be okay.

Suppose $\{\underline{y}_t\}_{\forall t}$ is an $AR(\infty)$ t-series. For $\tau > 0$, the “orthogonal projection” $\underline{y}_{t+\tau} | \underline{y}_{\leq t}$ is defined as the linear combination of $\underline{y}_{\leq t}$ obtained by doing Linear Regression with x-variables $\underline{y}_{\leq t}$ and y-variable $\underline{y}_{t+\tau}$. Thus

$$E \left[\underline{y}_{t+\tau} | \underline{y}_{\leq t} \right] = E_{|\underline{y}_{\leq t}} [\underline{y}_{t+\tau}] \quad (83.85)$$

and

$$Var \left[\underline{y}_{t+\tau} | \underline{y}_{\leq t} \right] = E \left[\left(\underline{y}_{t+\tau} | \underline{y}_{\leq t} - E[\underline{y}_{t+\tau} | \underline{y}_{\leq t}] \right)^2 \right] \quad (83.86)$$

$$= Var_{|\underline{y}_{\leq t}} [\underline{y}_{t+\tau}] . \quad (83.87)$$

For example, for $\tau = 1$,

$$\underline{y}_{t+1} = \sum_{j=1}^p \alpha_j \underline{y}_{t+1-j} + \underline{n}_{t+1} \quad (83.88)$$

$$\underline{y}_{t+1} | \underline{y}_{\leq t} = \sum_{j=1}^p \alpha_j \underline{y}_{t+1-j} \quad (83.89)$$

$$\underline{y}_{t+1} - \underline{y}_{t+1 | \underline{y}_{\leq t}} = \underline{n}_{t+1} \quad (83.90)$$

$$E \left[\underline{y}_{t+1} | \underline{y}_{\leq t} \right] = E_{|\underline{y}_{\leq t}} [\underline{y}_{t+1}] \quad (83.91)$$

$$= 0 \quad (83.92)$$

$$Var \left[\underline{y}_{t+1} | \underline{y}_{\leq t} \right] = Var_{|\underline{y}_{\leq t}} [\underline{y}_{t+1}] \quad (83.93)$$

$$= \sigma^2 \quad (83.94)$$

If $\{\underline{y}_t\}_{\forall t}$ is an $AR(\infty)$ t-series and $\tau > 0$, by **forecasting** $\underline{y}_{t+\tau} | \underline{y}_{\leq t}$, we will mean finding $\underline{y}_{t+\tau} | \underline{y}_{\leq t}$, given $\underline{y}_{\leq t}$ as input data or some other input data from which $\underline{y}_{\leq t}$ can be derived. Fig.83.10 illustrates 3 types of input data that we will consider next.

Note that if $\{\underline{y}_t\}_{\forall t}$ is an $AR(\infty)$ t-series, then $\alpha^-(\mathcal{B})\underline{y}_t = \underline{n}_t$. Thus, $\underline{n}_{\leq t}$ can be expressed in terms of $\underline{y}_{\leq t}$, yielding a function $\underline{n}_{\leq t}(\underline{y}_{\leq t})$. If the operator $\alpha^-(\mathcal{B})$ is invertible, the opposite is also true: $\underline{y}_{\leq t}$ can be expressed in terms of $\underline{n}_{\leq t}$, yielding a function $\underline{y}_{\leq t}(\underline{n}_{\leq t})$. Thus,

$$\underline{y}_{t+\tau} | \underline{y}_{\leq t} = \underline{y}_{t+\tau} | \underline{y}_{\leq t}(\underline{n}_{\leq t}) = \underline{y}_{t+\tau} | \underline{n}_{\leq t} \quad (83.95)$$

We've seen that if we assume invertibility of $\alpha^-(\mathcal{B})$, then, in order to forecast $\underline{y}_{t+\tau} | \underline{y}_{\leq t}$, we have the option of considering $\{\underline{y}_t\}_{\forall t}$ to be either an $AR(\infty)$ or an $MA(\infty)$ t-series. We will assume the latter, because this simplifies the algebra necessary to calculate both $\underline{y}_{t+\tau} | \underline{y}_{\leq t}$ and its variance.

(a) **find** $\underline{y}_{t+\tau} | \underline{y}_{\leq t}$ **given** $\underline{y}_{\leq t}$

Suppose $\{\underline{y}_t\}_{\forall t}$ is an $MA(\infty)$ t-series. Then

$$\underline{y}_{t+\tau} = \sum_{j=0}^{\infty} \nu_j \underline{n}_{t+\tau-j} \quad (83.96)$$

$$= \underbrace{\sum_{j=0}^{\tau-1} \nu_j \underline{n}_{t+\tau-j}}_{\underline{y}_{t+\tau} - \underline{y}_{t+\tau} | \underline{y}_{\leq t}} + \underbrace{\sum_{j=\tau}^{\infty} \nu_j \underline{n}_{t+\tau-j}}_{\underline{y}_{t+\tau} | \underline{y}_{\leq t}} \quad (83.97)$$

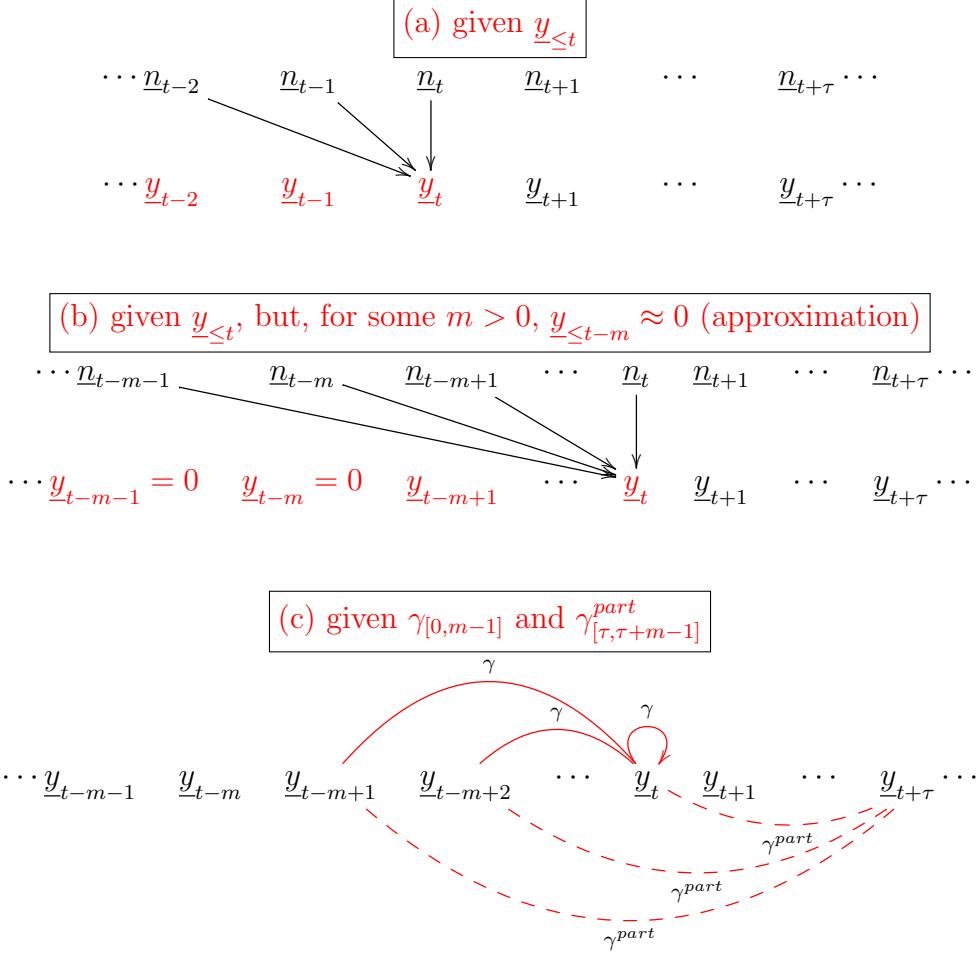


Figure 83.10: Bnet for $MA(\infty)$ for forecasting $\underline{y}_{t+\tau} | \underline{y}_{\leq t}$. Input data in red. For clarity, we show only the arrows entering node \underline{y}_t .

$$\underline{y}_{t+\tau} | \underline{y}_{\leq t} = \sum_{j=\tau}^{\infty} \nu_j \underline{n}_{t+\tau-j} \quad (83.98)$$

$$E_{|\underline{y}_{\leq t}}[\underline{y}_{t+\tau}] = E[\underline{y}_{t+\tau} | \underline{y}_{\leq t}] = 0^1 \quad (83.99)$$

$$\left\langle \underline{y}_{t+\tau} \underline{y}_{t+\tau} \right\rangle_{|\underline{y}_{\leq t}} = \left\langle \underline{y}_{t+\tau} - \underline{y}_{t+\tau} | \underline{y}_{\leq t}, \underline{y}_{t+\tau} - \underline{y}_{t+\tau} | \underline{y}_{\leq t} \right\rangle \quad (83.100)$$

$$= \sum_{j=0}^{\tau-1} \nu_j^2 \quad (83.101)$$

¹Remember that \underline{y}_t has zero mean by definition. $Y_t = y_t + \mu$ for some time independent constant μ .

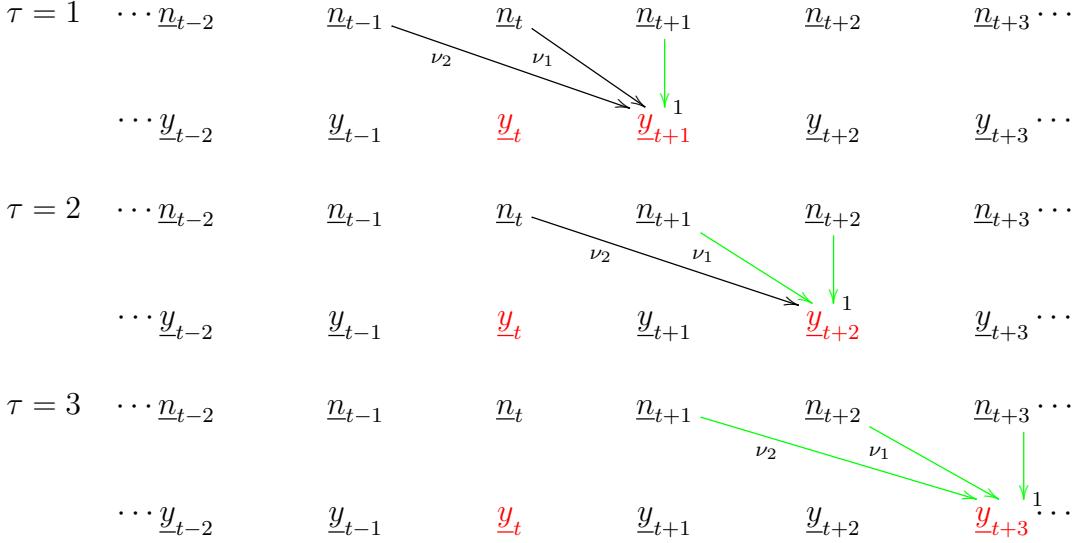


Figure 83.11: $MA(2)$ bnet. For clarity, we show only arrows entering node $\underline{y}_{t+\tau}$ for $\tau = 1, 2, 3$. Green arrows originate at a time after time t .

For $MA(q)$, because $\nu_j = 0$ for $j > q$, this reduces to

$$\underline{y}_{t+\tau} | y_{\leq t} = \mathbb{1}(\tau \leq q) \sum_{j=\tau}^q \nu_j \underline{n}_{t+\tau-j} \quad (83.102)$$

$$E_{|y_{\leq t}}[\underline{y}_{t+\tau}] = E[\underline{y}_{t+\tau} | y_{\leq t}] = 0 \quad (83.103)$$

$$\left\langle \underline{y}_{t+\tau}, \underline{y}_{t+\tau} \right\rangle_{|y_{\leq t}} = \sum_{j=0}^{\min(\tau-1, q)} \nu_j^2 \quad (83.104)$$

Eq.(83.104) is motivated by Fig.83.11). Only the green arrows entering $\underline{y}_{t+\tau}$ and originating in a node $\underline{n}_{t'}$ with $t' > t$ contribute to the right hand side of Eq.(83.104).

Thus, the mean of $\underline{y}_{t+\tau}$ can be predicted to remain zero for all $\tau > 0$ (forever). The error in that prediction increases with τ until τ reaches q . After that, the error remains constant.

Note that

$$\underline{y}_{t+\tau} | y_{\leq t} = \sum_{j=\tau}^{\infty} \nu_j \underline{n}_{t+\tau-j} \quad (83.105)$$

$$= \sum_{j=\tau}^{\infty} \nu_j \mathcal{B}^{j-\tau} \underline{n}_t \quad (83.106)$$

$$= \left[\sum_{j=0}^{\infty} \nu_j \mathcal{B}^{j-\tau} \right]_{\mathcal{B} \geq 0} \underline{n}_t \quad (83.107)$$

$$= [\mathcal{B}^{-\tau} \nu(\mathcal{B})]_{\mathcal{B} \geq 0} \underline{n}_t \quad (83.108)$$

$\mathcal{B} \geq 0$ means only the non-negative powers of \mathcal{B} are kept. Eq.(83.108) is just Eq.(83.98) in fancy notation.

In $MA(\infty)$, if $\nu(\mathcal{B})$ is invertible, then

$$\underline{n}_t = \nu(\mathcal{B})^{-1} \underline{y}_t \quad (83.109)$$

so

$$\boxed{\underline{y}_{t+\tau} | y_{\leq t} = \mathcal{L}_{WK} \underline{y}_t} \quad (83.110)$$

where

$$\mathcal{L}_{WK} = [\mathcal{B}^{-\tau} \nu(\mathcal{B})]_{\mathcal{B} \geq 0} \nu(\mathcal{B})^{-1} \quad (83.111)$$

Eq.(83.110) is called the **Wiener-Kolgomorov prediction formula (WKPF)**. Next, we apply WKPF to $AR(1)$ and $MA(1)$ as examples of its usage.

- $AR(1)$

$$\underbrace{(1 - \alpha_1 \mathcal{B})}_{\nu(\mathcal{B})^{-1}} \underline{y}_t = \underline{n}_t \quad (83.112)$$

$$\underline{y}_t = \underbrace{\left[\sum_{j=0}^{\infty} (\alpha_1 \mathcal{B})^j \right]}_{\nu(\mathcal{B})} \underline{n}_t \quad (83.113)$$

$$\nu_j = (\alpha_1)^j \quad (83.114)$$

$$[\mathcal{B}^{-\tau} \nu(\mathcal{B})]_{\mathcal{B} \geq 0} = (\alpha_1)^{\tau} \nu(\mathcal{B}) \quad (83.115)$$

$$\underline{y}_{t+\tau} | y_{\leq t} = (\alpha_1)^{\tau} \underline{y}_t \quad (83.116)$$

Hence, $\underline{y}_{t+\tau} | y_{\leq t}$ decreases geometrically as τ grows.

- $MA(1)$

$$\underline{y}_t = \underbrace{(1 + \nu_1 \mathcal{B})}_{\nu(\mathcal{B})} \underline{n}_t \quad (83.117)$$

$$[\mathcal{B}^{-\tau} \nu(\mathcal{B})]_{\mathcal{B} \geq 0} = \delta(\tau, 1) \nu_1 \quad (83.118)$$

$$\begin{aligned} \underline{y}_{t+\tau} | \underline{y}_{\leq t} &= \delta(\tau, 1) \nu_1 \underbrace{\left[\sum_{j=0}^{\infty} (-\nu_1 \mathcal{B})^j \right]}_{\nu(\mathcal{B})^{-1}} \underline{y}_t \\ &= \delta(\tau, 1) \nu_1 \underline{n}_t \end{aligned} \quad (83.119) \quad (83.120)$$

(b) find $\underline{y}_{t+\tau} | \underline{y}_{\leq t}$ given $\underline{y}_{\leq t}$, but, for some $m > 0$, $\underline{y}_{\leq t-m} \approx 0$ (approximation)

$$\underline{y}_{t+\tau} | \underline{y}_{\leq \tau} \approx [\mathcal{L}_{WK}]_{\mathcal{B}^{<m}} \underline{y}_t \quad (83.121)$$

(c) find $\underline{y}_{t+\tau} | \underline{y}_{\leq t}$ given $\gamma_{[0, m-1]}$ and $\gamma_{[\tau, \tau+m-1]}^{part}$

$$\underline{y}_{t+\tau} | \underline{y}_{\leq t} = \sum_{j=0}^{m-1} \underline{y}_{t-j} \beta_j \quad (83.122)$$

Suppose that $k \in \mathbb{Z}_{[0, m-1]}$.

Note that

$$\underline{y}_{t-k} | \underline{y}_{\leq t} = \underline{y}_{t-k} . \quad (83.123)$$

$$\underbrace{\langle \underline{y}_{t-k} | \underline{y}_{\leq t}, \underline{y}_{t+\tau} | \underline{y}_{\leq t} \rangle}_{\gamma_{\tau+k}^{part}} = \sum_{j=0}^{m-1} \underbrace{\langle \underline{y}_{t-k}, \underline{y}_{t-j} \rangle}_{\gamma_{k-j}} \beta_j \quad (83.124)$$

$$\begin{bmatrix} \gamma_{\tau}^{part} \\ \gamma_{\tau+1}^{part} \\ \vdots \\ \gamma_{\tau+m-1}^{part} \end{bmatrix} = \underbrace{\begin{bmatrix} \gamma_0 & \gamma_{-1} & \gamma_{-2} & \dots & \gamma_{1-m} \\ \gamma_1 & \gamma_0 & \gamma_{-1} & \dots & \gamma_{2-m} \\ \vdots & & & & \\ \gamma_{m-1} & \gamma_{m-2} & \gamma_{m-3} & \dots & \gamma_0 \end{bmatrix}}_{\Gamma} \underbrace{\begin{bmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_{m-1} \end{bmatrix}}_{\beta^m} \quad (83.125)$$

Solve Eq.(83.125) for β^m and plug β^m into Eq.(83.122).

83.13 Model Learning

Box-Jenkins Method

1. Filtering Data (FD)

Removing trends and periodicity (seasonality) via differencing, so as to achieve a w-stationary t-series. We deal with FD in Section 83.14.

2. Model Selection

Finding best possible p, q for $ARMA(p, q)$ using various statistical tests and metrics.

3. Parameter Learning (PL)

Finding best possible coefficients α_j and ν_j . We deal with PL in Section 83.15.

4. Testing

Measuring the goodness of fit.

83.14 Differencing and $ARIMA(p, d, q)$

Let $\alpha \in (0, 1)$. We say that $\{\underline{s}_t\}_{t \in \mathbb{N}}$ is a **Simple Exponential Smoothing (SES) t-series** if it satisfies

$$s_t = (1 - \alpha)s_{t-1} + \alpha x_t \quad (83.126)$$

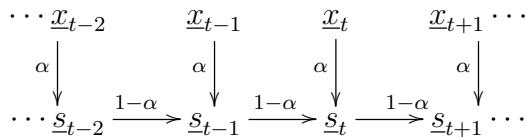


Figure 83.12: Bnet for Simple Exponential Smoothing t-series. $\alpha \in (0, 1)$

A SES t-series can be represented by the bnet of Fig.83.12. The TPM, printed in blue, for node s_t , is as follows

$$P(s_t) = \mathbb{1}(s_t = \text{ see Eq.(83.126)}) \quad (83.127)$$

One has

$$(1 - (1 - \alpha)\mathcal{B})\underline{s}_t = \alpha\underline{x}_t \quad (83.128)$$

so

$$\underline{s}_t = \frac{\alpha}{1 - (1 - \alpha)\mathcal{B}} \underline{x}_t \quad (83.129)$$

$$= \alpha \sum_{j=0}^{\infty} (1 - \alpha)^j \mathcal{B}^j \underline{x}_t \quad (\text{by Eq.(83.7)}) \quad (83.130)$$

Note from Fig.83.12 and Eq.(83.130) that each shock \underline{x}_{t-j} with $j > 0$ contributes to \underline{s}_t proportionally to the product $\alpha(1 - \alpha)^j$ of arrow weights in the path from \underline{x}_{t-j} to \underline{s}_t . For example, \underline{x}_{t-2} contributes to \underline{s}_t in the proportion $\alpha(1 - \alpha)^2$. The contributions of \underline{x}_{t-j} to \underline{s}_t decrease geometrically as the lag j increases.

Henceforth, we will abbreviate “differencing” by “diff”.

- First order diff operator

$$\Delta x_t = x_t - x_{t-1} = (1 - \mathcal{B})x_t \quad (83.131)$$

Seasonal first order diff:

$$\Delta_{sea} x_t = x_t - x_{t-ts} \quad (83.132)$$

where ts is length of season.

- Second order diff operator

$$\Delta^2 x_t = \Delta x_t - \Delta x_{t-1} \quad (83.133)$$

$$= x_t - 2x_{t-1} + x_{t-2} \quad (83.134)$$

$$= (1 - 2\mathcal{B} + \mathcal{B}^2)x_t \quad (83.135)$$

$$= (1 - \mathcal{B})^2 x_t \quad (83.136)$$

- d -th order diff operator

$$\Delta^d \underline{x}_t = (1 - \mathcal{B})^d \underline{x}_t \quad (83.137)$$

Recall that if $\{\underline{y}_t\}_{\forall t}$ is an $ARMA(p + d, q)$ t-series,

$$\left(1 - \sum_{j=1}^{p+d} \alpha_j \mathcal{B}^j\right) \underline{y}_t = \left(1 + \sum_{k=1}^q \nu_k \mathcal{B}^k\right) \underline{n}_t \quad (83.138)$$

We will say that $\{\underline{y}_t\}_{\forall t}$ is an **Autoregressive Integrated Moving Average $ARIMA(p, d, q)$ t-series** if

$$\left(1 - \sum_{j=1}^p \alpha_j \mathcal{B}^j\right) \underbrace{\Delta^d \underline{y}_t}_{(1-\mathcal{B})^d \underline{y}_t} = \left(1 + \sum_{k=1}^q \nu_k \mathcal{B}^k\right) \underline{n}_t \quad (83.139)$$

So an $ARIMA(p, d, q)$ t-series is a special case of an $ARMA(p + d, q)$ t-series.

Define \mathcal{O}_{BEF} and \mathcal{O}_{AFT} by

$$\mathcal{O}_{BEF} = 1 - \alpha_1 \mathcal{B} - \alpha_2 \mathcal{B}^2 - \alpha_3 \mathcal{B}^3 \quad (83.140)$$

$$\mathcal{O}_{AFT} = \mathcal{O}_{BEF}(1 - \mathcal{B}) \quad (83.141)$$

$$= \left\{ \begin{array}{cccc} 1 & -\alpha_1 \mathcal{B} & -\alpha_2 \mathcal{B}^2 & -\alpha_3 \mathcal{B}^3 \\ & -\mathcal{B} & +\alpha_1 \mathcal{B}^2 & +\alpha_2 \mathcal{B}^3 \\ \end{array} \right. + \alpha_3 \mathcal{B}^4 \quad (83.142)$$

$$= 1 - \underbrace{(\alpha_1 + 1)}_{\alpha'_1} \mathcal{B} - \underbrace{(\alpha_2 - \alpha_1)}_{\alpha'_2} \mathcal{B}^2 - \underbrace{(\alpha_3 - \alpha_2)}_{\alpha'_3} \mathcal{B}^3 - \underbrace{(0 - \alpha_3)}_{\alpha'_4} \mathcal{B}^4 \quad (83.143)$$

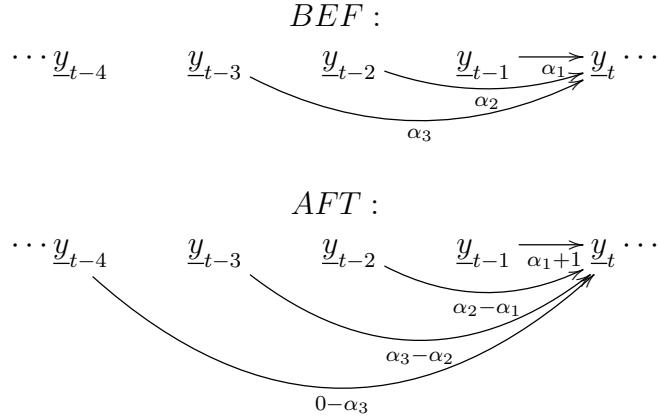


Figure 83.13: Effect on \underline{y}_t of going from before (BEF) with a t-series $ARIMA(3, 0, 0)$ to after (AFT) with a t-series $ARIMA(3, 1, 0)$.

$\mathcal{O}_{BEF}\underline{y}_t$ and $\mathcal{O}_{AFT}\underline{y}_t$ are illustrated in Fig.83.13. Note that in going from BEF to AFT, a new arrow is introduced with weight $\alpha'_4 = -\alpha_3$. The intermediate length arrows $\underline{y}_{t-3} \rightarrow \underline{y}_t$ and $\underline{y}_{t-2} \rightarrow \underline{y}_t$ have weights $\alpha'_3 = \alpha_3 - \alpha_2$ and $\alpha'_2 = \alpha_2 - \alpha_1$ so if $\alpha_j \approx \alpha_{j-1}$, then those 2 arrows have negligible weights. Only the the longest arrow $\underline{y}_{t-4} \rightarrow \underline{y}_t$ and the shortest arrow $\underline{y}_{t-1} \rightarrow \underline{y}_t$ have non-negligible weights.

Suppose

$$\underline{y}_t \approx a + bt + ct^2 + dt^3 \quad (83.144)$$

This smooth, low-order polynomial-fit to the t-series is called its **trend**.

$$\Delta \underline{y}_t \approx \Delta t \frac{d}{dt} \underline{y}_t = b + 2ct + 3dt^2 \quad (83.145)$$

$$\Delta^2 \underline{y}_t \approx (\Delta t)^2 \frac{d^2}{dt^2} \underline{y}_t = 2c + 6dt \quad (83.146)$$

$$\Delta^3 \underline{y}_t \approx (\Delta t)^3 \frac{d^2}{dt^3} \underline{y}_t = 6d \quad (83.147)$$

Note that the mean $6d$ in $\Delta^3 \underline{y}_t$ can be adjusted to zero (demeaning). From this example, we see that every time we apply a diff operator to a time series, we reduce the order of its polynomial trend by one.

Note that if $\{\underline{y}_t\}_{vt}$ is an $ARIMA(0, 1, 1)$ t-series,

$$\Delta \underline{y}_t = (1 + \nu_1 \mathcal{B}) \underline{n}_t \quad (83.148)$$

$$\underline{y}_t = \underline{y}_{t-1} + \underline{n}_t + \nu_1 \underline{n}_{t-1} \quad (83.149)$$

Claim 135 $ARIMA(0, 1, 1)$ is equivalent to a simple exponential smoothing t-series.

proof: Setting

$$\underline{n}_t = \underline{y}_t - \widehat{\underline{y}}_t \quad (83.150)$$

in Eq.(83.148), we get

$$(1 - \mathcal{B}) \underline{y}_t = (1 - \nu_1 \mathcal{B}) (\underline{y}_t - \widehat{\underline{y}}_t) \quad (83.151)$$

$$(1 - \nu_1 \mathcal{B}) \widehat{\underline{y}}_t = (1 - \nu_1) \mathcal{B} \underline{y}_t \quad (83.152)$$

$$\widehat{\underline{y}}_t = \nu_1 \widehat{\underline{y}}_{t-1} + (1 - \nu_1) \underline{y}_{t-1} \quad (83.153)$$

Now replace $\widehat{\underline{y}}_t \rightarrow \underline{s}_t$, $\underline{y}_{t-1} \rightarrow \underline{n}_t$ and $\nu_1 \rightarrow 1 - \alpha$.

Note that this proof has established the equivalence of two different bnets. Those two bnets are pictured in Fig.83.14.

QED

83.15 Parameter Learning

In this section, we will show how to find an estimate $\widehat{\theta}$ for the parameters $\theta = (\{\alpha_j\}_{\forall j}, \{\nu_j\}_{\forall j}, \sigma^2)$ used in $ARMA(p, q)$. We do this **parameter learning (PL)** by postulating a “quasi” log likelihood function $LL(\theta)$, and maximizing that over θ . The estimate $\widehat{\theta}$ that we obtain is called the **Quasi-Maximum Likelihood Estimate (QMLE)**. The reason for using the prefix “quasi” is that the likelihood probability involved doesn’t really satisfy the i.i.d assumption.

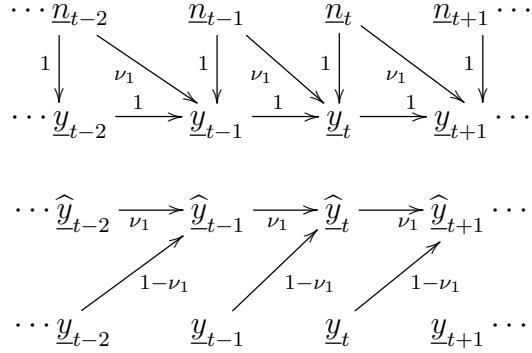


Figure 83.14: Bnets for two t-series that were proven equivalent in Claim 135. Note that the 2 bnets have NO arrows in common!

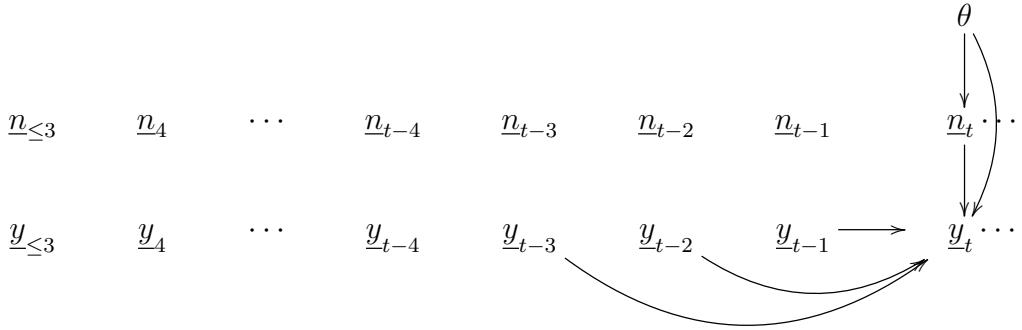


Figure 83.15: Bnet for PL of $AR(3)$. For clarity, we show only the arrows entering nodes y_t and n_t .

83.15.1 PL of $AR(p)$

Recall that if $\{y_t\}_{\forall t}$ is an $AR(p)$ t-series,

$$y_t = \sum_{j=1}^p \alpha_j y_{t-j} + n_t \quad (83.154)$$

Let $\theta = (\sigma^2, \alpha_{[1,p]})$. We want to learn the parameters θ of the $AR(p)$ bnet Fig.83.15, assuming the TPMs, printed in blue, are as follows.

For $t \geq p+1$

$$P(y_t|y_{\leq t}, n_t, \theta) = \mathbb{1}(y_t = \text{see Eq.(83.154)}) \quad (83.155)$$

$$P(n_t|\theta) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left[\frac{-(n_t)^2}{2\sigma^2}\right] \quad (83.156)$$

Note that

$$P(y_{\leq t}|\theta) = \sum_{n_{[p+1,t]}} \left\{ \prod_{\tau \in [p+1,t]} P(y_\tau|y_{[\tau-p,\tau-1]}, n_\tau) P(n_\tau|\theta) \right\} P(y_{\leq p}|\theta) \quad (83.157)$$

$$= \prod_{\tau \in [p+1,t]} \left\{ \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left[\frac{-\left(y_\tau - \sum_{j=1}^p \alpha_j y_{\tau-j}\right)^2}{2\sigma^2} \right] \right\} P(y_{\leq p}|\theta) \quad (83.158)$$

The log likelihood function $LL_t(\theta)$ for this bnet is defined by

$$LL_t(\theta) = \frac{1}{t} \ln P(y_{\leq t}|\theta). \quad (83.159)$$

Hence,

$$LL_t(\theta) = \underbrace{\frac{1}{t} \ln P(y_{\leq p}|\theta)}_{LL_t^{prior}(\theta)} + LL'_t(\theta) \quad (83.160a)$$

where

$$LL'_t(\theta) = \frac{-(t-p)}{t} \ln \sqrt{2\pi\sigma^2} - \frac{1}{2t\sigma^2} \sum_{\tau=p+1}^t \left(y_\tau - \sum_{j=1}^p \alpha_j y_{\tau-j} \right)^2. \quad (83.160b)$$

Henceforth, we will assume that $y_{[1,t]}$ is known.

We start by assuming that the prior log likelihood $LL_t^{prior}(\theta)$ is zero. This means that we have uniform (uninformative) priors. Later on, we will consider instead a Gaussian prior that uses the info that $y_{\leq p}$ is known a priori.

If we assume that $y_{[1,t]}$ is known, then we can plug this info into $LL'_t(\theta)$ and use a non-linear optimization method to maximize $LL'_t(\theta)$ and find the optimum θ .

Alternatively, one can use LR. Define the strange looking dataset

$$\mathcal{D} = \{(t, y_{[t-p,t-1]}, \boxed{y_t}) : t\} \quad (83.161)$$

This dataset is strange because it replaces sampling from a population with sampling in time. (i.e., a “stochastic” average by an “ergodic” average.) The datasets that we use to do LR usually satisfy the i.i.d. assumption, but that assumption does not necessarily hold for this one. Luckily, that assumption is not necessary for doing LR with x-variables $y_{[t-p,t-1]}$, y-variable y_t , and regression coefficients $\alpha_{[1,p]}$. The LR software gives estimates $\hat{\alpha}_{[1,p]}$ that we can use to calculate residuals

$$\epsilon_\tau = y_\tau - \sum_{j=1}^p \hat{\alpha}_j y_{\tau-j} \quad (83.162)$$

One can estimate σ^2 from those residuals using

$$\widehat{\sigma}^2 = \frac{1}{t-p} \sum_{\tau=p+1}^t \epsilon_\tau^2. \quad (83.163)$$

So far, we have assumed an uninformative prior. Alternatively, one can assume the Gaussian prior

$$P(y_{\leq p}|\theta) = \frac{1}{(2\pi\sigma^2)^{\frac{p}{2}}} \exp \left[-y_{\leq p}^T \frac{\Gamma^{-1}}{2} y_{\leq p} \right] \quad (83.164)$$

where

$$\Gamma_{i,j} = \left\langle \underline{y}_i, \underline{y}_j \right\rangle = \gamma_{j-i} \quad (83.165)$$

for $i, j \in \mathbb{Z}_{[1,p]}$. We are assuming that all \underline{y}_τ have been demeaned; i.e., the mean

$$\widehat{\mu} = \frac{1}{t} \sum_{\tau=1}^t y_\tau \quad (83.166)$$

has been subtracted from each y_τ for $\tau \in \mathbb{Z}_{[1,t]}$.

83.15.2 PL of $MA(q)$

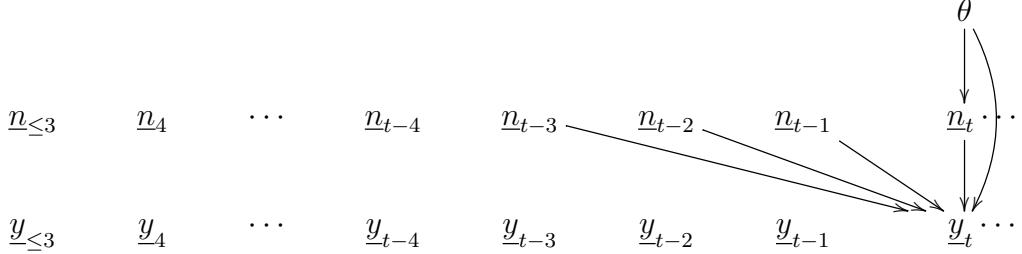


Figure 83.16: Bnet for PL of $MA(3)$. For clarity, we show only the arrows entering nodes \underline{y}_t and n_t .

Recall that if $\{\underline{y}_t\}_{\forall t}$ is an $MA(q)$ t-series,

$$y_t = \sum_{j=1}^q \nu_j n_{t-j} + n_t \quad (83.167)$$

Let $\theta = (\sigma^2, \nu_{[1,q]})$. We want to learn the parameters θ of the $MA(q)$ bnet Fig.83.16, assuming the TPMs, printed in blue, are as follows.

For $t \geq q + 1$, let

$$P(y_t | n_{\leq t}, \theta) = \mathbb{1}(y_t = \text{see Eq.(83.167)}) \quad (83.168)$$

$$P(n_t | \theta) = \frac{1}{\sigma \sqrt{2\pi}} \exp \left[\frac{-(n_t)^2}{2\sigma^2} \right] \quad (83.169)$$

Note that

$$P(y_{\leq t} | \theta) = \sum_{n_{[q+1,t]}} \left\{ \prod_{\tau \in [q+1,t]} P(y_\tau | n_{[\tau-q,\tau-1]}, n_\tau) P(n_\tau | \theta) \right\} P(n_{\leq q} | \theta) \quad (83.170)$$

$$= \prod_{\tau \in [q+1,t]} \left\{ \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left[\frac{-\left(y_\tau - \sum_{j=1}^q \nu_j n_{\tau-j} \right)^2}{2\sigma^2} \right] \right\} P(n_{\leq q} | \theta) \quad (83.171)$$

The log likelihood function $LL_t(\theta)$ for this bnet is defined by

$$LL_t(\theta) = \frac{1}{t} \ln P(y_{\leq t} | \theta) \quad (83.172)$$

$$LL_t(\theta) = \underbrace{\frac{1}{t} \ln P(y_{\leq q} | \theta)}_{LL_t^{prior}(\theta)} + LL'_t(\theta) \quad (83.173a)$$

where

$$LL'_t(\theta) = \frac{-(t-q)}{t} \ln \sqrt{2\pi\sigma^2} - \frac{1}{2t\sigma^2} \sum_{\tau=q+1}^t \left(y_\tau - \sum_{j=1}^q \nu_j n_{\tau-j} \right)^2 \quad (83.173b)$$

Henceforth, we will assume that $n_{\leq q} = 0$. By Eq.(83.167), this implies that $y_{\leq q} = 0$, so the prior log likelihood $LL_t^{prior}(\theta)$ is independent of θ .

Eq.(83.173) expresses $LL'_t(\theta)$ in terms of y_τ 's and n_τ 's, but the input data consist of only y_τ 's. Luckily, Eq.(83.167) can be used to express all the $n_{\leq t}$ in terms of the $y_{\leq t}$, assuming the input $n_{\leq q} = 0$. Next, we will explain how to do this. For definiteness and simplicity, we will assume $q = 2$ and $\nu_1 = \nu_2 = 1$, and we will use Fig.83.17 as a pedagogical aid.

For $q = 2$, Eq.(83.167) gives y_τ as a linear combination of 3 \underline{n} nodes, $\underline{n}_\tau, \underline{n}_{\tau-1}$ and $\underline{n}_{\tau-2}$. But since it's a linear equation, it can be used to express the latest \underline{n} node, i.e., \underline{n}_τ , in terms of y_τ and the 2 previous \underline{n} nodes.

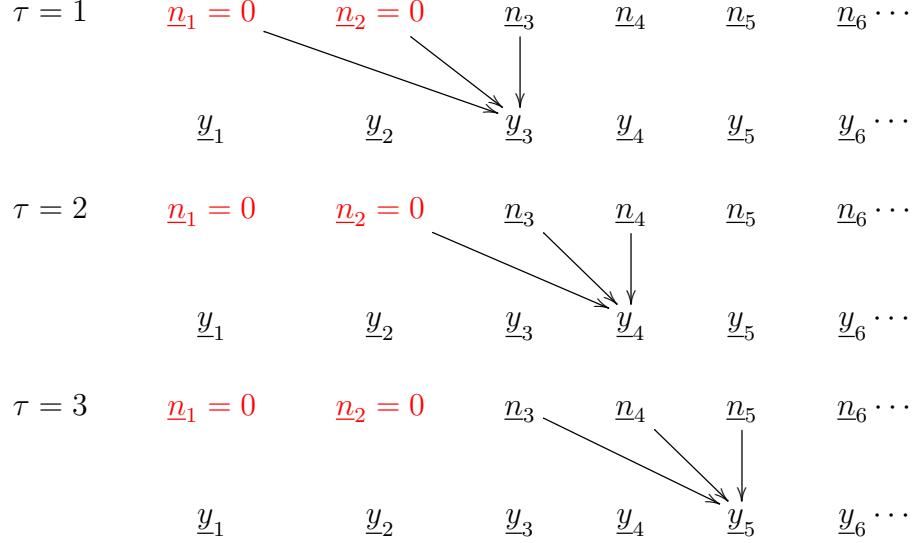


Figure 83.17: Bnet for PL of $MA(2)$. We assume $\underline{n}_1 = \underline{n}_2 = 0$. For clarity, we show only the arrows entering node $\underline{y}_{2+\tau}$ for $\tau = 1, 2, 3$.

Hence, we see from Fig.83.17 that:

From $\underline{n}_1 = \underline{n}_2 = 0$, we get $\underline{y}_1 = \underline{y}_2 = 0$.

From the $\tau = 1$ bnet, we get

$$\underline{n}_3 = \underline{y}_3 \quad (83.174)$$

From the $\tau = 2$ bnet, we get

$$\underline{n}_4 = \underline{y}_4 - \underline{n}_3 \quad (83.175)$$

From the $\tau = 3$ bnet, we get

$$\underline{n}_5 = \underline{y}_5 - \underline{n}_3 - \underline{n}_4 \quad (83.176)$$

83.15.3 PL of $ARMA(p, q)$

Recall that if $\{\underline{y}_t\}_{\forall t}$ is an $AR(p, q)$ t-series, then

$$y_t = \sum_{j=1}^p \alpha_j y_{t-j} + n_t + \sum_{j=1}^q \nu_j y_{t-j} \quad (83.177)$$

Let $\theta = (\sigma^2, \alpha_{[1,p]}, \nu_{[1,q]})$. We want to learn the parameters θ of the $ARMA(p, q)$ bnet Fig.83.18, in which the TPMs, printed in blue, are as follows.

For $t \geq \max(p+1, q+1)$, let

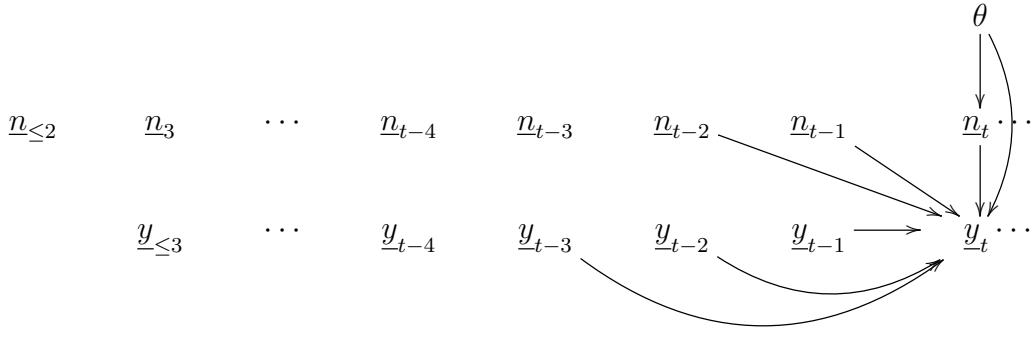


Figure 83.18: Bnet for PL of $AR(3, 2)$. For clarity, we show only the arrows entering nodes y_t and n_t .

$$P(y_t | y_{<t}, n_{\leq t}, \theta) = \mathbb{1} (y_t = \text{see Eq.(83.177)}) \quad (83.178)$$

$$P(n_t | \theta) = \frac{1}{\sigma \sqrt{2\pi}} \exp \left[\frac{-(n_t)^2}{2\sigma^2} \right] \quad (83.179)$$

As we did for $AR(p)$ and $MA(q)$, we can now calculate a log likelihood function and maximize it to obtain the parameters θ . But first we will have to assume that $n_{\leq q} = 0$ and $y_{[1,t]}$ is known, and we will have to express the unknown $n_{[q+1,t]}$ in terms of the known $y_{[1,t]}$ using the t-series definition Eq.(83.177).

83.16 VAR(p)

Let $\vec{x}_t = (\underline{x}_t^1, \underline{x}_t^2, \dots, \underline{x}_t^{nr})^T \in \mathbb{R}^{nr \times 1}$. Thus, \vec{x}_t for each t is a column vector with nr rows. A **vector time series**, denoted variously by $\{\vec{x}_1, \vec{x}_2, \dots, \vec{x}_{nt}\} = \{\vec{x}_t\}_{t=1}^{nt} = \{\vec{x}_t\}_{\forall t}$, is a set of $\mathbb{R}^{nr \times 1}$ vectors index by a discrete set of times $\mathbb{Z}_{[0,nt]}$.

Henceforth, we will use the Einstein summation convention for uppercase indices like $A, B, C \in \mathbb{Z}_{[1,nr]}$; i.e., they should be summed over from 1 to nr if they are repeated in a product. For example, $x^A y^A = \sum_{A=1}^{nr} x^A y^A$.

By **vector white noise** $\{\vec{n}_t\}_{\forall t} \sim WN(0, \Sigma)$, where $\Sigma \in \mathbb{R}^{nr \times nr}$, we mean a vector t-series $\{\vec{n}_t\}_{\forall t}$ that satisfies, for $A, B \in \mathbb{Z}_{[1,nr]}$ and $t, t' \in \mathbb{Z}_{[1,nt]}$,

$$E[n_t^A] = 0 \quad (83.180)$$

$$\langle \underline{n}_t^A, \underline{n}_{t'}^B \rangle = \Sigma^{A,B} \delta(t, t') \quad (83.181)$$

Suppose $\{\vec{y}_t\}_{\forall t}$ is a zero mean vector t-series. Hence $\underline{y}_t^A = \underline{Y}_t^A - \mu^A$, $E[\underline{Y}_t^A] = \mu^A$, $E[\underline{y}_t^A] = 0$. Suppose also that $\{\vec{n}_t\}_{\forall t} \sim WN(0, \Sigma)$. Then we define a **Vector Auto-Regressive t-series VAR(p)** by

$$\underline{y}_t^A = \sum_{j=1}^p \alpha_j^{A,B} \underline{y}_{t-j}^B + \underline{n}_t^A \quad (83.182)$$

The bnet for $VAR(p)$ is the same as the bnet for $AR(p)$ except that now, for all t , the nodes \underline{y}_t and \underline{n}_t are replaced by \vec{y}_t and \vec{n}_t and now the node weights α_t are $nr \times nr$ matrices with entries $\alpha_t^{A,B}$. For example, for $nr = 2$, we have

$$\begin{bmatrix} \underline{y}_t^1 \\ \underline{y}_t^2 \end{bmatrix} = \sum_{j=1}^p \begin{bmatrix} \alpha_j^{1,1} & \alpha_j^{1,2} \\ \alpha_j^{2,1} & \alpha_j^{2,2} \end{bmatrix} \begin{bmatrix} \underline{y}_{t-j}^1 \\ \underline{y}_{t-j}^2 \end{bmatrix} + \begin{bmatrix} \underline{n}_t^1 \\ \underline{n}_t^2 \end{bmatrix} \quad (83.183)$$

Chapter 84

Transfer Learning

Historically, the term Transfer Learning (TL) has been used in AI mostly by Artificial Neural Net (ANN) proponents (see the Wikipedia article on TL, Ref.[144]). Recently, however, a theory of causal TL has begun to emerge (see Chapter 86).

TL in AI is a fairly wide topic that is somewhat ambiguously defined. Some subjects that can be lumped under the heading of TL in AI are:

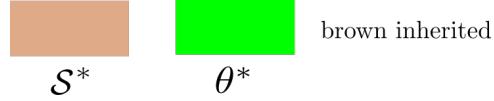
- data fusion/combining models
- model generalization
- transportability of causal knowledge, external validity (see Chapter 86)

Most AI researchers will agree that it is highly desirable to have TL in AI, because the human brain obviously does plenty of TL to great advantage. Although reams of papers have been written about the subject of TL in AI, a systematic theory of TL in AI that is universally accepted and popular remains elusive. The current theory of TL for ANN looks to me like a grab bag of heuristic approaches that are fragile, meaning they can be easily spoofed. The theory of TL for bnets (see Chapter 86) seems to me to be in much better shape: it's more elaborate, systematic, and it yields more robust results.

In this brief chapter, I will limit myself to describing a possible way of classifying, from a bnet perspective, the various approaches to TL in AI. Note that this method of classification even works for TL for ANNs, because ANNs can be viewed as bnets with deterministic nodes and a layered structure. One can describe TL in AI as a systematic way of defining a bnet \mathcal{B}^* using a bnet \mathcal{B} and other information. Bnet \mathcal{B} is associated with a dataset \mathcal{D} , and bnet \mathcal{B}^* is associated with a dataset \mathcal{D}^* . A bnet $\mathcal{B} = (\mathcal{S}, \theta)$ comprises a DAG structure \mathcal{S} and a TPM for each node of the DAG. We'll denote the TPMs (a.k.a. parameters) of \mathcal{B} by θ . So let's classify the various approaches to TL in AI by specifying what parts of the structure \mathcal{S} and parameters θ of \mathcal{B} are transferred to the structure \mathcal{S}^* and parameters θ^* of \mathcal{B}^* , and what parts of $\mathcal{B}^* = (\mathcal{S}^*, \theta^*)$ are new.

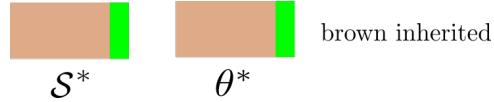
1. **Fine tune parameters.** $\mathcal{S}^* = \mathcal{S}$, $\theta^* \approx \theta$.

In this approach, we use the dataset \mathcal{D}^* associated with bnet \mathcal{B}^* to adjust slightly the parameters from θ to θ^* .

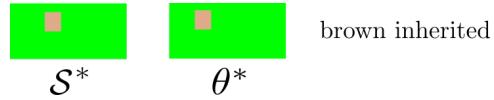


2. **Replace final layers of \mathcal{S} and of θ by new ones.** $\mathcal{S}^* = \mathcal{S}$ and $\theta^* = \theta$ except for final layers,

For example, in an ANN, replace the final layers by new ones, and use \mathcal{D}^* to find the parameters of those new final layers.



3. **Transfer only the TPM of a single node y of \mathcal{B} .** \mathcal{S}^* and θ^* are new except $P^*(y|pa(y)) = P(y|pa(y))$.



Chapter 85

Transformer Networks

This chapter is based on Refs.[25] and [145].

Transformer Networks (TN) have been taking the field of Natural Language Processing (NLP) by storm in recent years. They were introduced in 2017 and already are the basis of BERT (Bidirectional Encoder Representations from Transformers) and GPT (Generative Pre-trained Transformer), two TN libraries that have been trained with huge databases such as all of English Wikipedia (2,500M words).

Recurrent Neural Nets (RNNs) are discussed in Chapter 67. TNs are quickly displacing RNNs, an older method, in NLP. TNs are better than RNNs for doing NLP in several important ways. Whereas RNNs analyze the tokens (words) of a sentence sequentially (like a Kalman Filter), TNs analyze them in parallel, and thus are more amenable to parallel computing. Also, because RNNs analyze the words of a sentence sequentially, they tend to give more importance to the end of a sentence than to its beginning. That's because RNNs start forgetting the beginning of a sentence by the time they reach its end, like a patient with Alzheimer's. TNs do not suffer from this malady.

Dynamical bnets are discussed in Chapter 24. In Chapter 67, we showed that RNNs are dynamical bnets. The goal of this chapter is to define TNs, and to show that they too are dynamical bnets.

Let

\mathcal{S} be the set of words in a sentence,

$h_i^t \in \mathbb{R}^{nh}$ be an nh dimensional column vector for word $i \in \mathcal{S}$ at time t .

$Q^t, K^t, V^t \in \mathbb{R}^{nh \times nh}$ be the **prior-attention matrices for time slice t** .

These matrices are learned by training the net. The letters Q, K, V stand for Query, Key and Value, respectively.

Fig.85.1 represents a TN of a 3-word sentence as a dynamical bnet. The TPMs, printed in blue, for bnet Fig.85.1, are as follows:

$$P(V^t) = \delta(V^t, V_0^t) \quad (85.1a)$$

$$\begin{array}{ccc}
\underline{V}^t & \underline{Q}^t & \underline{K}^t \\
\downarrow & \downarrow & \downarrow \\
\{\underline{v}_i^t\}_{i=0,1,2} & \{\underline{q}_i^t\}_{i=0,1,2} & \{\underline{k}_i^t\}_{i=0,1,2}
\end{array}$$

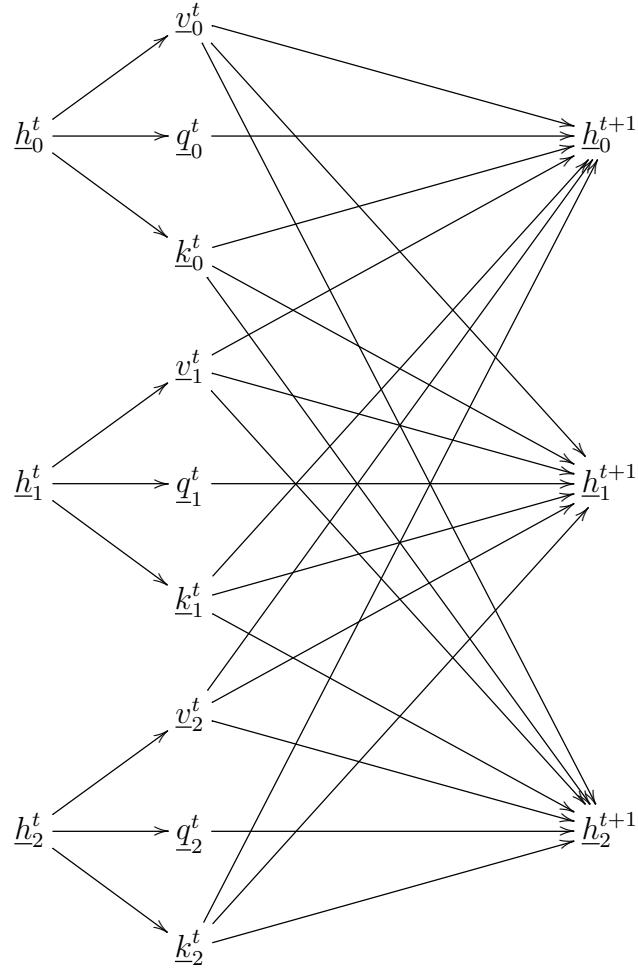


Figure 85.1: Time slice t of dynamical bnet for a transformer network (TN) of a 3-word sentence. The following arrows were not drawn explicitly for clarity: arrows pointing from node \underline{V}^t to nodes \underline{v}_i^t for $i = 0, 1, 2$, from node \underline{Q}^t to nodes \underline{q}_i^t for $i = 0, 1, 2$, and from node \underline{K}^t to nodes \underline{k}_i^t for $i = 0, 1, 2$. Note that \underline{k}_i^t for all i points to \underline{h}_j^{t+1} for all j . Likewise, \underline{v}_i^t for all i points to \underline{h}_j^{t+1} for all j .

$$P(Q^t) = \delta(Q^t, Q_0^t) \quad (85.1b)$$

$$P(K^t) = \delta(K^t, K_0^t) \quad (85.1c)$$

$$P(v_i^t | V^t, h_i^t) = \mathbb{1}(-v^t = V^t h_i^t) \quad (85.2)$$

$$P(q_i^t | Q^t, h_i^t) = \mathbb{1}(-q^t = Q^t h_i^t) \quad (85.3)$$

$$P(k_i^t | K^t, h_i^t) = \mathbb{1}(-k^t = K^t h_i^t) \quad (85.4)$$

$$P(h_i^{t+1} | v_i^t, q_i^t, k_i^t) = \mathbb{1}\left(-h_i^{t+1} = \sum_{j \in \mathcal{S}} v_j^t \underbrace{\frac{e^{(q_i^t)^T k_j^t}}{\sum_{j' \in \mathcal{S}} e^{(q_i^t)^T k_{j'}^t}}}_{\substack{w_{j|i} = [\text{softmax}((q_i^t)^T k_j^t)]_j \\ E_{j|i}[v_j^t] = \text{Attention}}} \right) \quad (85.5)$$

In Eqs.85.1, the 3 root nodes $\underline{V}^t, \underline{Q}^t, \underline{K}^t$ have delta function priors. To improve stability of the dynamical bnet, it is more common instead to use for these priors, multiple delta functions (a.k.a. **multi-heads**) labeled $c = 0, 1, \dots, nc - 1$, as follows.

$$P(V^t | c) = \delta(V^t, V_c^t) \quad (85.6a)$$

$$P(Q^t | c) = \delta(Q^t, Q_c^t) \quad (85.6b)$$

$$P(K^t | c) = \delta(K^t, K_c^t) \quad (85.6c)$$

for $c = 0, 1, \dots, nc - 1$. Even better, one can go fully Bayesian, and make these 3 prior distributions arbitrary categorical distributions to be determined by net training, instead of mere delta functions.

Chapter 86

Transportability of Causal Knowledge

This chapter is mostly based on Refs.[48] and [36].

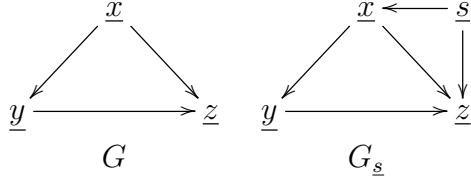


Figure 86.1: Example of selection bnet $G_{\underline{s}}$ created from bnet G .

Suppose one wants to transfer causal knowledge from a **source population** Σ to a **target population** Σ^* .

Given a bnet G , define a **selection diagram** $G_{\underline{s}}$ as a bnet formed by adding to G a new root node \underline{s} and new arrows pointing from **selection node** \underline{s} to one or more **target nodes** of G . We'll call the set of target nodes of \underline{s} the **target set** $T_{\underline{s}}$. $\underline{s} = 0$ corresponds to population Σ and $\underline{s} = 1$ to population Σ^* . For bnet G , the TPM for a node \underline{x} with parents $pa(\underline{x})$, is given by:

$$P_G(\underline{x}|pa(\underline{x})) = P(\underline{x}|pa(\underline{x})) \quad (86.1)$$

For bnet $G_{\underline{s}}$, nodes \underline{x} with parents $pa(\underline{x})$, where $\underline{s} \notin pa(\underline{x})$, have TPMs:

$$P_{G_{\underline{s}}}(\underline{x}|pa(\underline{x})) = P(\underline{x}|pa(\underline{x})). \quad (86.2)$$

Nodes \underline{x} with parents $pa(\underline{x}) \cup \underline{s}$, have TPMs:

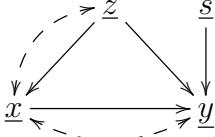
$$P_{G_{\underline{s}}}(\underline{x}|pa(\underline{x}), \underline{s}) = \begin{cases} P(\underline{x}|pa(\underline{x}), \underline{s} = 0) = P(\underline{x}|pa(\underline{x})) & \text{if } \underline{s} = 0 \\ P(\underline{x}|pa(\underline{x}), \underline{s} = 1) = P^*(\underline{x}|pa(\underline{x})) & \text{if } \underline{s} = 1 \end{cases} \quad (86.3)$$

Fig.86.1 shows an example of a selection diagram $G_{\underline{s}}$. In that figure, the target set of \underline{s} is $\{\underline{x}, \underline{z}\}$.

All this can be generalized so as to have more than one selection node, with the target sets of the selection nodes being disjoint, and such that a selection node can have more than 2 states.

Do-transport formulae

Claim 136 (*Trivial Memoryless Transportability, from Ref.[48]*)

If  where $\underline{s} \in \{0, 1\}$ is a selection node, then

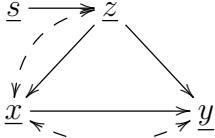
$$P^*(y|\mathcal{D}\underline{x} = x, z) = P^*(y|x, z) \quad (\text{replace } \mathcal{D} \text{ by } 1, \text{ keep } P^*) \quad (86.4)$$

$$\begin{array}{ccc} z & \xrightarrow{\underline{s}=1} & y \\ \swarrow & & \downarrow \\ \mathcal{D}\underline{x} = x & \xrightarrow{\quad} & y \end{array} = \begin{array}{ccc} z & \xrightarrow{\underline{s}=1} & y \\ \swarrow & & \downarrow \\ x & \xrightarrow{\quad} & y \end{array} \quad (86.5)$$

proof: See Claim 39.

QED

Claim 137 (*Direct Transportability, a.k.a. External Validity, from Ref.[48]*)

If  where $\underline{s} \in \{0, 1\}$ is a selection node, then

$$P^*(y|\mathcal{D}\underline{x} = x, z) = P(y|\mathcal{D}\underline{x} = x, z) \quad (\text{replace } P^* \text{ by } P, \text{ keep } \mathcal{D}) \quad (86.6)$$

$$\begin{array}{ccc} \underline{s}=1 & \longrightarrow & z \\ & \searrow & \\ \mathcal{D}\underline{x} = x & \xrightarrow{\quad} & y \end{array} = \begin{array}{ccc} z & \longrightarrow & y \\ & \searrow & \\ \mathcal{D}\underline{x} = x & \xrightarrow{\quad} & y \end{array} \quad (86.7)$$

Furthermore,

$$P^*(y|\mathcal{D}\underline{x} = x) = \sum_z P(y|\mathcal{D}\underline{x} = x, z)P^*(z) \quad (86.8)$$

$$\begin{array}{ccc} \underline{s}=1 & \longrightarrow & \sum z \\ & \searrow & \\ \mathcal{D}\underline{x} = x & \xrightarrow{\quad} & y \end{array} = \begin{array}{ccc} \underline{s}=1 & \longrightarrow & \sum z \\ & \searrow & \\ \mathcal{D}\underline{x} = x & \xrightarrow{\quad} & y \end{array} \quad (86.9)$$

proof: See Claim 40.

QED

Claim 138 (*S-Admissible Transportability, from Ref.[48]*)

If $\underline{s} \xrightarrow{\text{---}} \underline{z} \xrightarrow{\text{---}} \underline{a}$ where $\underline{s} \in \{0, 1\}$ is a selection node, then

$$P^*(y|\mathcal{D}\underline{x} = x) = \sum_a P(y|\mathcal{D}\underline{x} = x, a)P^*(a) \quad (86.10)$$

$$\begin{array}{ccc} \underline{s} = 1 & & \underline{s} = 1 \longrightarrow \sum a \\ \searrow & & \downarrow \\ \mathcal{D}\underline{x} = x \longrightarrow y & = & \mathcal{D}\underline{x} = x \longrightarrow y \end{array} \quad (86.11)$$

proof: See Claim 41.

QED

Claim 139 (*Non-transportability, from Ref.[48]*)

If $\underline{s} \xrightarrow{\text{---}} \underline{h} \xrightarrow{\text{---}} \underline{z}$ where $\underline{s} \in \{0, 1\}$ is a selection node, then

$$P^*(y|\mathcal{D}\underline{x} = x) = P^*(y|\mathcal{D}\underline{x} = x) \quad (86.12)$$

$$\begin{array}{ccc} \underline{s} = 1 & & \\ \downarrow & & \\ \mathcal{D}\underline{x} = x \longrightarrow y & = same & \end{array} \quad (86.13)$$

proof: See Claim 42.

QED

Claim 140 (*from Ref.[48]*)

If $\underline{s} \xrightarrow{\text{---}} \underline{h} \xrightarrow{\text{---}} \underline{z}$ where $\underline{s} \in \{0, 1\}$ is a selection node, then

$$P^*(y|\mathcal{D}\underline{x} = x) = P(y|\mathcal{D}\underline{x} = x) \quad (86.14)$$

$$\begin{array}{ccc}
 \underline{s} = 1 & & (86.15) \\
 \searrow & & \\
 \mathcal{D}\underline{x} = x \longrightarrow y & = & \mathcal{D}\underline{x} = x \longrightarrow y
 \end{array}$$

proof: See Claim 43.

QED

Claim 141 (*from Ref.[48]*)

If \underline{s} is a selection node, then

$$P^*(y|\mathcal{D}\underline{x} = x) = \sum_z P(y|\mathcal{D}\underline{x} = x, z)P^*(z|x) \quad (86.16)$$

$$\begin{array}{ccccc}
 \underline{s} = 1 & \mathcal{D}\underline{x} = x & \underline{s} = 1 & \mathcal{D}\underline{x} = x & (86.17) \\
 \searrow & \downarrow & \searrow & \downarrow & \\
 y & = & x \longrightarrow \sum z \longrightarrow y
 \end{array}$$

proof: See Claim 44.

QED

Chapter 87

Turbo Codes

This chapter is based on Ref.[34].

In this chapter, vectors with n components will be indicated by an n superscript. For example, $a^n = (a_0, a_1, \dots, a_{n-1})$.

Consider an n -letter message $u^n = (u_0, u_1, \dots, u_{n-1})$, where for all i , $u_i \in \mathcal{A}$ is an element of an alphabet \mathcal{A} , and where for all i , the u_i are i.i.d.. Suppose u^n is encoded deterministically in two different ways, $e_1(u^n)$ and $e_2(u^n)$. After passing through the same memoryless channel, the variables u^n, e_1, e_2 become $\tilde{u}^n, \tilde{e}_1, \tilde{e}_2$, respectively. The letter u stands for unencoded, and e for encoded. Quantities with a tilde $\tilde{u}^n, \tilde{e}_1, \tilde{e}_2$ occur after channel passage and are visible (measurable). Quantities without a tilde u^n, e_1, e_2 are hidden (unmeasurable).

The situation just described can be represented by the bnet Fig.87.1, or by its abridged version Fig.87.2. But note that the abridged version does not show explicitly that the u_i are i.i.d. or that the channel is memoryless (i.e., that the u_i for all i pass independently through the channel).

Define

$$x = (u^n, e_1, e_2) \quad (87.1)$$

and

$$\tilde{x} = (\tilde{u}^n, \tilde{e}_1, \tilde{e}_2) . \quad (87.2)$$

Fig.87.1 implies that

$$P(x, \tilde{x}) = P(\tilde{u}^n | u^n) \left[\prod_{r=1,2} P(\tilde{e}_r | e_r) P(e_r | u^n) \right] P(u^n) . \quad (87.3)$$

Because the u^n are i.i.d.,

$$P(u^n) = \prod_i P(u_i) . \quad (87.4)$$

Because the channel is memoryless,

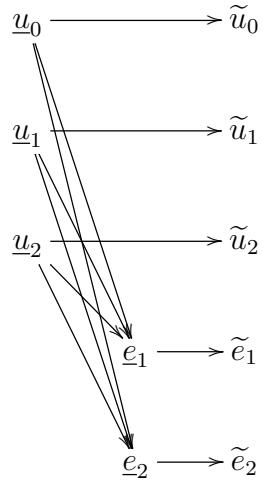


Figure 87.1: Turbo coding Bnet representing a message being encoded two different ways and then the original message and the 2 encodings pass through a memoryless channel.

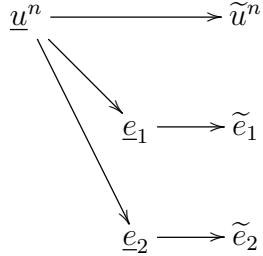


Figure 87.2: Abridged version of Fig.87.1.

$$P(\tilde{u}^n | u^n) = \prod_i P(\tilde{u}_i | u_i) . \quad (87.5)$$

Because the encoding is deterministic, we must have for $r = 1, 2$

$$P(e_r | u^n) = \delta(e_r, e_r(u^n)) . \quad (87.6)$$

Define the belief functions

$$BEL_i = BEL_i(\underline{u}_i = a) = P(\underline{u}_i = a | \tilde{x}) . \quad (87.7)$$

The best estimate of u_j given all visible evidence \tilde{x} is

$$\hat{u}_i = \operatorname{argmax}_{u_i} BEL_i(u_i) . \quad (87.8)$$

Define the probability functions

$$\pi_i = \pi_i(u_i) = P(u_i) , \quad (87.9)$$

and the likelihood functions

$$\lambda_i = \lambda_i(u_i) = P(\tilde{u}_i|u_i) . \quad (87.10)$$

For $r = 1, 2$, define the Kernel functions

$$K_r = K_r(u^n) = P(\tilde{e}_r|e_r = e_r(u^n)) . \quad (87.11)$$

In this book, $\mathcal{N}(!a)$ denotes a normalization constant that does not depend on a . Define

$$\mathcal{N}_i = \mathcal{N}(!u_i) . \quad (87.12)$$

Claim 142

$$BEL_i = \mathcal{N}_i \lambda_i \pi_i \mathcal{T}_i^{K_1 K_2} [\prod_{j \neq i} \lambda_j \pi_j] , \quad (87.13)$$

where $\mathcal{T}_i^K(\cdot)$ with $K = K_1 K_2$ is an operator (transform) that acts on functions of u^n :

$$\mathcal{T}_i^K(\cdot) = \sum_{u^n} \delta(u_i, a) K(u^n)(\cdot) . \quad (87.14)$$

proof:

$$\begin{aligned} P(\underline{u}_i = a | \tilde{x}) &= \\ &= \sum_x \delta(u_i, a) P(x | \tilde{x}) \end{aligned} \quad (87.15)$$

$$= \sum_x \delta(u_i, a) \frac{P(\tilde{x}|x)P(x)}{P(\tilde{x})} \quad (87.16)$$

$$= \mathcal{N}(!a) \sum_x \delta(u_i, a) P(\tilde{x}|x)P(x) \quad (87.17)$$

$$= \mathcal{N}(!a) \sum_x \delta(u_i, a) P(u^n) \left[\prod_{r=1,2} P(\tilde{e}_r|e_r) \delta(e_r, e_r(u^n)) \right] \prod_j P(\tilde{u}_j|u_j) \quad (87.18)$$

$$= \mathcal{N}(!a) \lambda_i(a) \pi_i(a) R , \quad (87.19)$$

where

$$R = \sum_{u^n} \delta(u_i, a) \left[\prod_{r=1,2} P(\tilde{e}_r | e_r(u^n)) \right] \prod_{j \neq i} P(\tilde{u}_j | u_j) P(u_j) \quad (87.20)$$

$$= \sum_{u^n} \delta(u_i, a) \left[\prod_{r=1,2} K_r(u^n) \right] \prod_{j \neq i} \lambda_j(u_j) \pi_j(u_j) \quad (87.21)$$

$$= \mathcal{T}_i^{K_1 K_2} \left[\prod_{j \neq i} \lambda_j(u_j) \pi_j(u_j) \right]. \quad (87.22)$$

Hence

$$BEL_i(a) = \mathcal{N}(!a) \lambda_i(a) \pi_i(a) \mathcal{T}_i^{K_1 K_2} \left[\prod_{j \neq i} \lambda_j(u_j) \pi_j(u_j) \right]. \quad (87.23)$$

QED

87.1 Decoding Algorithm

The Turbo algorithm for decoding the encode message is as follows. For $m = 0$, let

$$\pi_j^{(0)}(u_j) = \frac{1}{n_{\underline{u}_j}}. \quad (87.24)$$

Then for $m = 1, 2, \dots$, let

$$\pi_i^{(m)} = \mathcal{N}_i \mathcal{T}_i^{K_m \% 2} \left[\prod_{j \neq i} \lambda_j \pi_j^{(m-1)} \right], \quad (87.25)$$

where $m \% 2 = 1$ if m is odd and $m \% 2 = 2$ if m is even. Furthermore, for $m > 0$, let

$$BEL_i^{(m)} = \mathcal{N}_i \lambda_i \pi_i^{(m-1)} \pi_i^{(m)} \quad (87.26)$$

$$= \mathcal{N}_i \lambda_i \pi_i^{(m-1)} \mathcal{T}_i^{K_m \% 2} \left[\prod_{j \neq i} \lambda_j \pi_j^{(m-1)} \right]. \quad (87.27)$$

As $m \rightarrow \infty$, $BEL_i^{(m)}$ given by Eq.(87.27) is expected to converge to the the exact BEL_i given by Eq.(87.13).

Turbo decoding can be represented by the bnets Figs.87.3 and 87.4.

The TPMs, printed in blue, for bnet Fig.87.3, are as follows.

$$P(d_i^{(m)} = a | \tilde{u}^n, \tilde{e}_{m \% 2}) = BEL_i^{(m)}(a). \quad (87.28)$$

The TPMs, printed in blue, for bnet Fig.87.4, are as follows.

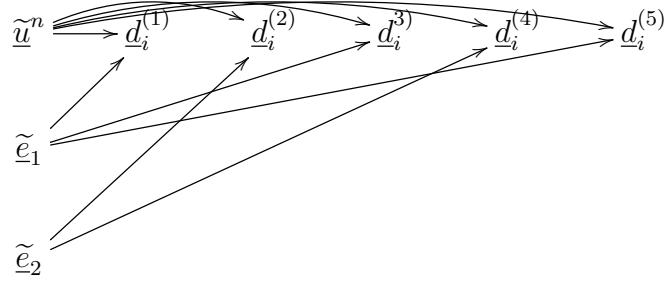


Figure 87.3: Bnet describing Turbo code generation of $BEL_i^{(m)}(a)$ for $m = 1, 2, \dots$

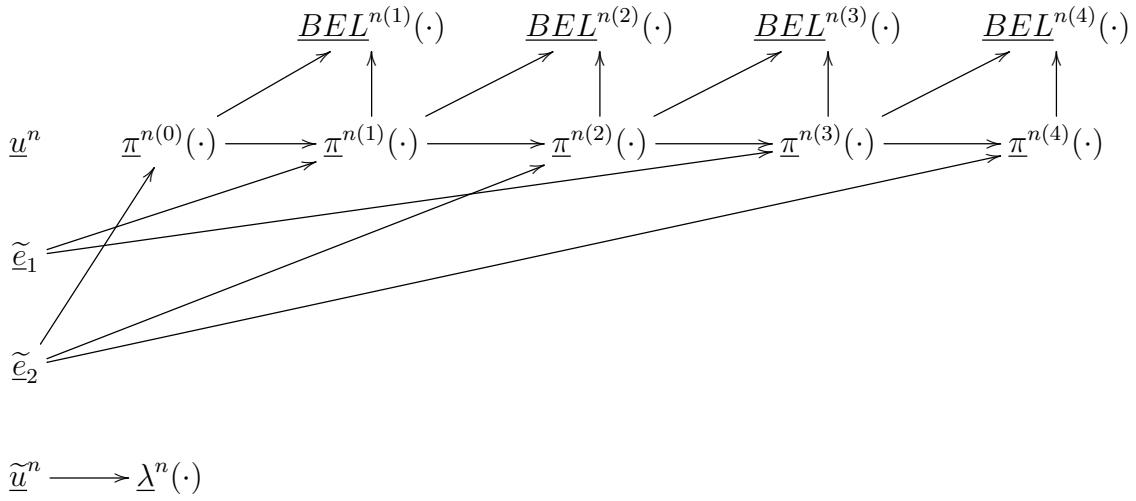


Figure 87.4: Bnet describing Turbo code generation of $BEL^{n(m)}(\cdot)$ and $\pi^{n(m)}(\cdot)$ for $m = 0, 1, 2, \dots$. The following arrows were not drawn for clarity: Arrows pointing from node $\lambda^n(\cdot)$ to nodes $\pi^{n(m)}(\cdot)$ and $BEL^{n(m)}(\cdot)$ for $m = 0, 1, 2, \dots$

$$P((\lambda^n)'(\cdot) | \tilde{u}^n) = \delta((\lambda^n)'(\cdot), \lambda^n(\cdot)) \quad (87.29)$$

$$P(\pi^{n(m)}(\cdot) | \lambda^n(\cdot), \pi^{n(m-1)}(\cdot), \tilde{e}_{m \% 2}) = \prod_i \prod_{u_i} \delta(\pi_i^{(m)}(u_i), \mathcal{N}_i \mathcal{T}_i^{K_m \% 2} [\prod_{j \neq i} \lambda_j \pi_j^{(m-1)}]) \quad (87.30)$$

$$P(BEL^{n(m)}(\cdot) | \lambda^n(\cdot), \pi^{n(m)}(\cdot), \pi^{n(m-1)}(\cdot)) = \prod_i \prod_{u_i} \delta(BEL_i(u_i), \mathcal{N}_i \lambda_i \pi_i^{(m-1)} \pi_i^{(m)}) \quad (87.31)$$

87.2 Message Passing Interpretation of Decoding Algorithm

Ref.[34] shows that the Turbo code decoding algo can be interpreted as an application of Message Passing. We leave all talk of Message Passing to a separate chapter, Chapter 50.

Chapter 88

Uplift Modelling

This chapter is based on many references, including Ref.[15, 11, 146, 54].

Uphill Modelling (UP) deals with the application of Rubin’s Theory of Potential Outcomes (PO) to advertisement and marketing.

PO, which is discussed in Chapter 64, is a subset of Pearl’s Causal Inference. Besides UP, other applications of PO theory that are discussed in this book are: Regression Discontinuity (Chapter 68), Difference-in-Differences (Chapter 16) and Synthetic Controls (Chapter 81).

In UP, each **participant person** is interrogated at two well anticipated, fairly closely spaced times t_0 and t_1 (as opposed to Difference-in-Differences (DID), where t_0 and t_1 might be years apart, and long before the DID analysis is attempted.). In between those two times, a treatment which we will refer to as the **UP diagnostic test** is applied. For example, at times t_0 and t_1 , every participant might be asked how important he/she rates climate change on a scale of 1 to 10. In between times t_0 and t_1 , every participant might be sent a brochure on climate change. In UP, as in all other PO applications, each **sample** σ is in the treated or control groups, but not both. But in UP, the same participant can be in both the treated and control groups. If so, that participant is considered two different samples σ ; for example, $\sigma = \text{treatedBob}, \text{controlBob}$. In UP, the samples are aware of which of those groups they are in, so they are not “treatment blind”.

88.1 UP types

Let $y_t^B \in \mathbb{R}$ for $t = t_0, t_1$ be the treatment response at time t for participant B . (We are using here the same notation as in Chapter 64). Call $\delta^B = y_{t_1}^B - y_{t_0}^B$ the **participant uplift** for participant B . As shown in Fig.88.1, UP classifies participants into 4 **UP-types**: Persuadables, SureThings, LostCauses, and SleepyDogs. The UP-type of a participant depends on the changes that are induced on that participant by an **UP-diagnostic-test**.

- For a **Persuadable** participant, $\delta^B > 0$.

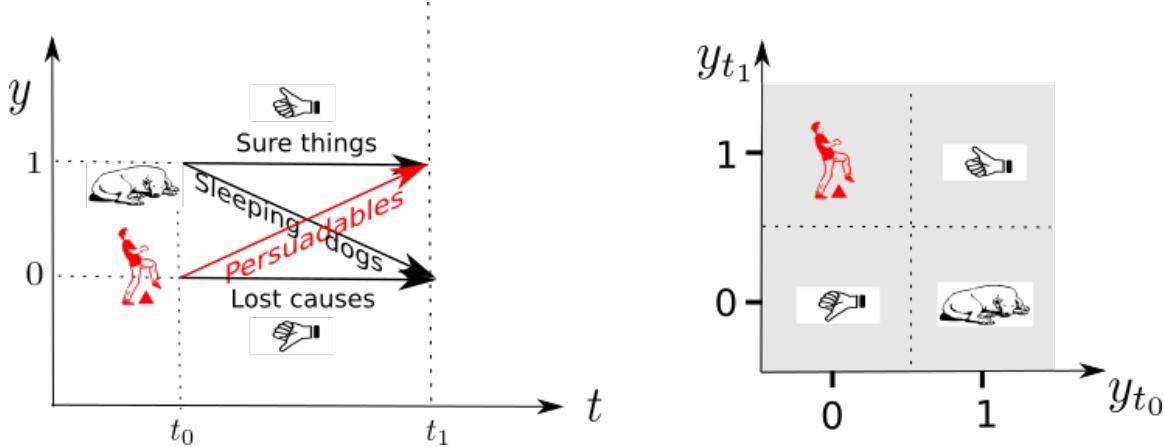


Figure 88.1: UP diagnostic test can be used to classify all participants of the population into 4 UP-types. This figure assumes $y \in \{0, 1\}$. More generally, $y \in \mathbb{R}$. t represents time. $t = t_0$ corresponds to $d = 0 = \text{untreated}$, and $t = t_1$ corresponds to $d = 1 = \text{treated}$.

- For a **SleepyDogs** participant, $\delta^B < 0$.
- For a **SureThings** participant, $\delta^B \approx 0$ and $y_{t_0}^B$ is high.
- For a **LostCauses** participant, $\delta^B \approx 0$ and $y_{t_0}^B$ is low.

Suppose B belongs to stratum A_x . What is commonly called the **uplift** is the **stratum-uplift** $\delta_x = ACE_x$. Strata can also be classified into the 4 UP-types, depending on the sign and size of their δ_x . A participant may not be typical for his stratum and may have different **participant and stratum UP-types**. For example, he may have positive participant uplift and therefore have a Persuadable participant UP-type, but his stratum-uplift might be negative, so he has the SleepyDogs stratum UP-type.

Advertisers are very interested in finding the Persuadable strata in a population so as to focus their resources on them. For example, UP was used very successfully during the Obama presidential campaigns. Team Obama conducted UP-diagnostic tests much like the climate change one described earlier. This allowed them to identify voters who might be sitting on the fence on whether to vote for Obama or not. Then Team Obama spent the lion share of resources on those fence-sitters.

88.2 Some Relevant Technical Formulas from Chapter 64

Recall the following technical formulae that were proven in Chapter 64:

- Recall Eq.(64.143):

$$ACE = \sum_x P(x) \underbrace{\sum_y y [P(y|d=1, x) - P(y|d=0, x)]}_{ACE_x} \quad (88.1)$$

If $y \in \{0, 1\}$, then

$$\underbrace{ACE_x}_{\delta_x} = \underbrace{\frac{P_{y|\underline{d}, \underline{x}}(1|1, x)}{Y_x^1}} - \underbrace{\frac{P_{y|\underline{d}, \underline{x}}(1|0, x)}{Y_x^0}} . \quad (88.2)$$

- Recall Eq.(64.158):

$$\widehat{ACE}_x = \underbrace{\frac{1}{N_x} \sum_{\sigma \in A_x} \frac{d^\sigma y^\sigma}{g_{1|x^\sigma}}}_{Y_x^1} - \underbrace{\frac{1}{N_x} \sum_{\sigma \in A_x} \frac{(1-d^\sigma)y^\sigma}{g_{0|x^\sigma}}}_{Y_x^0} \quad (88.3)$$

88.3 UP Analysis

The input to UP is a PO dataset $DS = \{(\sigma, d^\sigma, x^\sigma, y^\sigma) : \sigma = 0, 1, 2, \dots, nsam - 1\}$. where $d^\sigma \in \{0, 1\}$, $x^\sigma \in S_{\underline{x}}$, $y^\sigma \in \mathbb{R}$. A participant B is assigned two different σ if he/she belongs to both the treated and control groups. We will assume $S_{\underline{x}}$ is a finite set. In general, $x = (x_0, x_1, \dots, x_{n-1})$ is an n dimensional vector of features x_i . If any of the x_i is a priori continuous, we will assume it has been binned into a finite number of bins.

Starting with DS , UP performs the following steps. Fig.88.2 is a pictorial representation of the quantities that are calculated during these steps.

1. Find A_x for each observed $x \in S_{\underline{x}}$. Set $A_x = \emptyset$ for unobserved $x \in S_{\underline{x}}$.
2. Calculate δ_x for each $x \in S_{\underline{x}}$. Set $\delta_x = 0$ if $A_x = \emptyset$.
3. Calculate the set

$$\{\Delta_c\}_{c=0,1,\dots,nc-1} = \{\delta_x : x \in S_{\underline{x}}\} \quad (88.4)$$

of distinct uplifts δ_x . The class labels c should be assigned so that the sequence of Δ_c is monotonic and non-increasing; i.e.,

$$\Delta_0 \geq \Delta_1 \geq \dots \geq \Delta_{nc-1} . \quad (88.5)$$

Now calculate

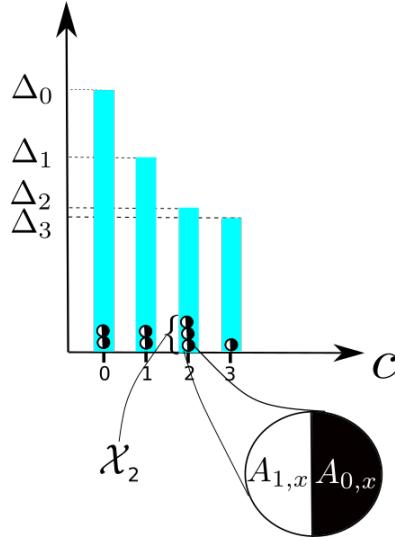


Figure 88.2: Pictorial representation of the sequence $\{(\mathcal{X}_c, \Delta_c)\}_{c=0,1,\dots,nc-1}$.

$$\mathcal{X}_c = \{x : \delta_x = \Delta_c\} \quad (88.6)$$

for each c . By the end of this step, we will have calculated $\{(\mathcal{X}_c, \Delta_c)\}_{c=0,1,\dots,nc-1}$. We will refer to the \mathcal{X}_c as **strata-bins**. Note that

$$\Delta_c = \frac{1}{|\mathcal{X}_c|} \sum_{x \in \mathcal{X}_c} \delta_x \quad (88.7)$$

$$= \underbrace{\frac{1}{|\mathcal{X}_c|} \sum_{x \in \mathcal{X}_c} Y_x^1}_{Y_c^1} - \underbrace{\frac{1}{|\mathcal{X}_c|} \sum_{x \in \mathcal{X}_c} Y_x^0}_{Y_c^0} . \quad (88.8)$$

4. For each c , calculate

$$\Sigma_{d,c} = \cup_{x \in \mathcal{X}_c} A_{d,x} \quad (88.9)$$

for $d \in \{0, 1\}$ and

$$\Sigma_c = \Sigma_{0,c} \cup \Sigma_{1,j} . \quad (88.10)$$

Fig.88.3 is a way of plotting the results of UP in an intuitive way that even a business type can understand. UP software often plots something called a Qini curve, but I find Qini curves opaque, confusingly defined in the literature, unnecessary and not very well motivated. So I don't use them.

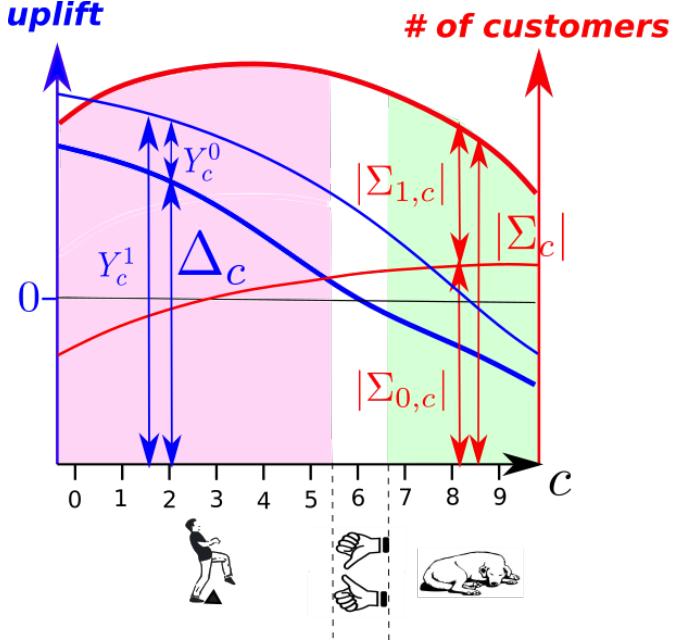


Figure 88.3: Plot of UP results. Alternative to Qini curves.

88.4 UP Decision Trees

In this section, we will describe how to build UP decision trees (UP dtrees), and explain why they are needed for UP.

Generic dtrees are described in Chapter 14. This section complements rather than replaces that chapter so the reader is advised to read that chapter first.

Ref.[54] is an excellent paper on the use of dtrees in UP.

The analysis described previously in Section 88.3, although theoretically correct, will work very poorly in practice. The strata-bins of Section 88.3 correspond to the classification classes of a dtree. But strata-bins are very specific so they severely overfit the data. Although dtrees can also suffer from overfitting, there are known methods of preventing or mitigating overfitting in dtrees.

There are also tasks that dtrees can do well and the methods explained so far cannot do well. For example, suppose we have a classless dataset $DS^- = \{(\sigma, x^\sigma) : \sigma \in \Sigma^-\}$ and we want to predict the class c^σ and uplift Δ_{c^σ} for each of these individuals $\sigma \in \Sigma^-$. A dtree can easily do that. The alternative is to use the classy dataset $DS = \{(\sigma, x^\sigma, c^\sigma) : \sigma \in \Sigma\}$ to prepare a dictionary that orders the elements of S_x and gives a class c and an uplift value Δ_c for each feature vector $x \in S_x$. But such a dictionary overfits and says nothing for feature vectors x that do not show up in the classy dataset DS ; i.e., the dictionary doesn't guess (interpolate). Dtrees, on the other hand, do guess.

So, without further ado, let us describe how to modify the results of Chapter 14 on generic dtrees to the case of UP dtrees. The main difference, as we will explain

in detail next, is that the Information Gain metric used for generic dtrees needs to be replaced by another metric.

$$\begin{aligned} & \text{Diagram showing population numbers and their summation.} \\ & \text{Top: Two arrows point from } \underline{x}_j \text{ to } x_j. \text{ The upper arrow is labeled } \underline{x}_j = x_j' \text{ and the lower arrow is labeled } \underline{x}_j = x_j. \\ & \text{Below: A bracket groups these two terms with the label } \{N_j^d(c, x_j)\}_{c \in S_c, x_j \in S_{\underline{x}_j}}. \\ & \text{Further down: The sum of these terms is shown as } \sum_{c \in S_c} N_j^d(c, x_j) = N_j^d(x_j). \\ & \text{Finally: The total population is given as } \sum_{x_j \in S_{\underline{x}_j}} N_j^d(c, x_j) = N_j^d(c) \text{ and } \sum_{c \in S_c} N_j^d(c) = \sum_{x_j \in S_{\underline{x}_j}} N_j^d(x_j) = N_j^d. \end{aligned}$$

Figure 88.4: Fig.14.4 with d dependence added. $d \in \{0, 1\}$ is the treatment dose.

Fig.88.4 was obtained from Fig.14.4 by adding d dependence. $d \in \{0, 1\}$ is the treatment dose. Note that in UP, we build a dtree in which every node carries a double ($d = 0, 1$) TPV. This is in contrast to the generic dtrees built in Chapter 14, in which each node carries a single TPV. $N_j^d(c, x_j)$ is the number of individuals σ in the population that reaches node x_j with $d \in \{0, 1\}$, belonging to class $c \in S_c$ and having $\underline{x}_j = x_j$. From these population numbers, we can define the bnet in Fig.88.5. The TPMs, printed in blue, for the (non-root) nodes of this bnet, are as follows

$$P(c|x_j, j, d) = \frac{N_j^d(c, x_j)}{N_j^d(x_j)} \quad (88.11)$$

$$P(x_j|j, d) = \frac{N^d(x_j)}{N_j^d} \quad (88.12)$$

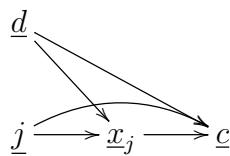


Figure 88.5: Bnet derived from population numbers in Fig.88.4

In Chapter 14, we used Information Gain (a mutual information) as the SAM (Separation Ability Measure) in SL (Structure Learning) of dtrees (Decision Trees). Information Gain is a bad SAM for SL of UP dtrees, because it knows nothing about $d = 0, 1$ and the double TPVs of nodes in UP dtrees. For UP dtrees, we need a SAM specifically designed to separate $d = 0, 1$, and generate classes that are uplift bins (i.e., uplift intervals).

Ref.[54] proposes and studies the following 3 SAMs for doing SL of UP dtrees.

1. SAM_DD (DD=Delta Delta)

For $d \in \{0, 1\}$ and $c, c' \in S_{\underline{c}}$, define the increments

$$\partial_d f(d) = f(1) - f(0) \quad (88.13)$$

and

$$\partial_{c',c} f(c) = f(c') - f(c) . \quad (88.14)$$

Let

$$\Delta_{c|j} = P(c|j, 1) - P(c|j, 0) \quad (88.15)$$

$$= \partial_d P(c|j, d) \quad (88.16)$$

$$SAM_DD_j = \max_{c, c'} |\partial_{c',c} \partial_d P(c|j, d)| \quad (88.17)$$

$$= \max_{c, c'} |\partial_{c',c} \Delta_{c|j}| \quad (88.18)$$

2. SAM_KL (KL=Kullback Leibler)

$$SAM_KL_j = \left[\sum_{x_j \in S_{\underline{x}_j}} P(x_j|j) D_{KL}(P_{\underline{c}|x_j,j,1} \| P_{\underline{c}|x_j,j,0}) \right] - D_{KL}(P_{\underline{c}|j,1} \| P_{\underline{c}|j,0}) \quad (88.19)$$

$$= \left[\sum_{x_j \in S_{\underline{x}_j}} P(x_j|j) \sum_{c \in S_{\underline{c}}} P(c|x_j, j, 1) \ln \frac{P(c|x_j, j, 1)}{P(c|x_j, j, 0)} \right] - \sum_{c \in S_{\underline{c}}} P(c|j, 1) \ln \frac{P(c|j, 1)}{P(c|j, 0)} \quad (88.20)$$

SAM_KL_j can be negative.

3. SAM_E (E=Euclidean)

SAM_E_j is defined the same way as SAM_KL_j except with the KL divergence $D_{KL}(P \parallel Q)$ in SAM_KL replaced by the Euclidean distance squared.

$$D(P, Q) = \sum_x (P(x) - Q(x))^2 \quad (88.21)$$

The intuitive reason for using these quantities as SAMs is that they maximize the change in uplift between successive tree levels, so that the uplift increases as quickly as possible as we descend down the UP tree. In the case of generic dtrees for which we use Information Gain as SAM, we are maximizing the correlation between classes and nodes as we descend down the tree. These two goals are related. In fact, in the limit where the number of control individuals becomes zero, SAM_KL_j and IG_j become the same, as will be shown later.

Next we show that SAM_KL_j satisfies the following 3 axioms¹

Claim 143 .

1. SAM_KL_j is minimum iff $P(c|x_j, j, 0) = P(c|x_j, j, 1)$ for all c and x_j .
2. If $P(c|j, d) = P(c|d)$ for all c, d , then $SAM_KL_j = 0$.
3. Suppose $N_r^0 = 0$ for all nodes $r \in J_0$ (i.e., no control population) and we use the Laplace Correction when warranted. Then

$$SAM_KL_j = H(\underline{c} : \underline{x}_j | j, 1) \quad (88.22)$$

$$= IG_j \quad \text{for treated population .} \quad (88.23)$$

proof:

The proof of items 1 and 2 follow by inspection of Eq.(88.20). Item 3 is proven in Claim 144 below.

QED

Let $N_{\underline{c}} = |S_{\underline{c}}|$. Define the uniform probability distribution

$$U_{\underline{c}}(c) = \frac{1}{N_{\underline{c}}} \quad (88.24)$$

for all $c \in S_{\underline{c}}$.

Eq.(88.11) for the TPM of node \underline{c} in the bnet Fig.88.5 can be "Laplace Corrected" as follows so that it is no longer undefined when its denominator vanishes:

¹We won't show it here, but according to Ref.[54], SAM_E_j also satisfies these 3 axioms, but SAM_DD_j satisfies only the first two.

$$P(c|j, d) = \begin{cases} \frac{N_j^d(c)}{N_j^d} & \text{if } N_j^d > 0 \\ U_{\underline{c}}(c) & \text{if } N_j^d = 0 \text{ (Laplace Correction)} \end{cases} \quad (88.25)$$

Claim 144 Suppose $N_r^0 = 0$ for all dtree nodes $r \in J_0$ and we use the Laplace Correction when warranted. Then

$$SAM_KL_j = H(\underline{c} : \underline{x}_j | j, 1) . \quad (88.26)$$

proof:

For all nodes $r \in J_0$, we must have

$$P_{\underline{c}|r,0} = U_{\underline{c}} \quad (88.27)$$

so

$$D_{KL}(P_{\underline{c}|r,1} \| P_{\underline{c}|r,0}) = D_{KL}(P_{\underline{c}|r,1} \| U_{\underline{c}}) \quad (88.28)$$

$$= \ln(N_{\underline{c}}) - H(\underline{c}|r, 1) . \quad (88.29)$$

For all $x_j \in S_{\underline{x}_j}$, we must also have

$$N_j = N_j^1, N(x_j) = N^1(x_j) \quad (88.30)$$

so

$$P(x_j | j) = P(x_j | j, 1) . \quad (88.31)$$

Now using Eqs.(88.29) and (88.31), we get

$$SAM_KL_j = - \left[\sum_{x_j \in S_{\underline{x}_j}} P(x_j | j) H(\underline{c} | x_j, j, 1) \right] + H(\underline{c} | j, 1) \quad (88.32)$$

$$= - \left[\sum_{x_j \in S_{\underline{x}_j}} P(x_j | j, 1) H(\underline{c} | x_j, j, 1) \right] + H(\underline{c} | j, 1) \quad (88.33)$$

$$= -H(\underline{c} | \underline{x}_j, j, 1) + H(\underline{c} | j, 1) \quad (88.34)$$

$$= H(\underline{c} : \underline{x}_j | j, 1) \quad (88.35)$$

QED

88.4.1 Appendix, connection between Δ_c and $\Delta_{c|j}$

Recall Eq.(88.8):

$$\Delta_c = \underbrace{\frac{1}{|\mathcal{X}_c|} \sum_{x \in \mathcal{X}_c} Y_x^1 - \frac{1}{|\mathcal{X}_c|} \sum_{x \in \mathcal{X}_c} Y_x^0}_{Y_c^1 - Y_c^0} \quad (88.36)$$

$$= \partial_d Y_c^d. \quad (88.37)$$

Compare that to Eq.(88.16):

$$\Delta_{c|j} = P(c|j, 1) - P(c|j, 0) \quad (88.38)$$

$$= \partial_d P(c|j, d) \quad (88.39)$$

What is the connection between these 2 deltas, Δ_c and $\Delta_{c|j}$? Are they equal?

First off, notice that $\Delta_{c|j}$ is defined for all nodes j of the dtree. Let $j(c)$ be the leaf node for which $\Delta_c \approx \Delta_{c|j(c)}$. Assume $y^\sigma \in \{0, 1\}$. Then

$$P(c|j = j(c), d) = \frac{N_{j(c)}^d(c)}{N_{j(c)}^d} \approx Y_c^d \quad (88.40)$$

So the two deltas are indeed approximately equal when $y^\sigma \in \{0, 1\}$ and $j = j(c)$.

Chapter 89

Variational Bayesian Approximation for Medical Diagnosis

This chapter is based on Ref.[23].

A Variational Bayesian Approximation (VBA) is when we approximate a probability distribution by another probability distribution that depends on a continuous “variational parameter”. This parameter is adjusted within its range of possible values, to make the approximation as good as possible. There are many VBA methods. VBA methods are inspired by ancient methods used in Calculus of Variations applied to Physics and Engineering problems.

In this chapter, we do VBA via Jensen’s inequality and convex/concave dual functions.

Ref.[23], on which this chapter is based, applies VBA methods to the problem of diagnostic inference using the Quick Medical Reference (QMR) bipartite Bayesian Network. According to Ref.[23] the maximal clique size of the QMR bnet is 150 nodes, which rules out exact methods of inference like the Junction Tree Algorithm (see Chapter 40). For such high complexity cases, one is forced to use either a VBA or a Monte Carlo method.

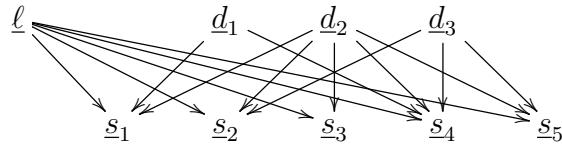


Figure 89.1: Typical bnet (bipartite, 2 level graph) for medical diagnosis to which we will apply VBA methods. In this case, $nd = 3$ and $ns = 5$. According to Ref.[23], for QMR, $nd \approx 600$ and $ns \approx 4000$.

Fig.89.1 gives a typical bnet for medical diagnosis to which we will apply VBA

methods. $\underline{d}_i \in \{0, 1\}$ for $i = 1, 2, \dots, nd$ are the possible diseases, $\underline{s}_\sigma \in \{0, 1\}$ for $\sigma = 1, 2, \dots, ns$ are the possible symptoms, and $\underline{\ell} \in \{0, 1\}$ is the leakage due to possible error in the parents of the symptoms. Note that the arrows point from diseases to symptoms because diseases precede in time the symptoms.

Let

$$\mathbb{Z}_{[1,n]} = \{1, 2, \dots, n\} \quad (89.1)$$

$$pa_\sigma = \{i \in \mathbb{Z}_{[1,nd]} : \underline{d}_i \in pa(\underline{s}_\sigma)\} = \text{parents of } \underline{s}_\sigma \quad (89.2)$$

Note that pa_σ does not include $\underline{\ell}$, which is also a parent of \underline{s}_σ .

$$ch_i = \{\sigma \in \mathbb{Z}_{[1,ns]} : \underline{s}_\sigma \in ch(\underline{d}_i)\} = \text{children of } \underline{d}_i \quad (89.3)$$

$$\underline{d}_A = \{\underline{d}_k : k \in A\} \quad (89.4)$$

$$\underline{d}^{md} = \{\underline{d}_k : k \in \mathbb{Z}_{[1,nd]}\} \quad (89.5)$$

$$\underline{d}_{!j} = \{\underline{d}_k : k \in \mathbb{Z}_{[1,nd]} - \{j\}\} \quad (89.6)$$

The TPMs, printed in blue, for the bnet Fig.89.1, are as follows:

$$P(\underline{d}_j) = \text{given} \quad (89.7)$$

$$P(\underline{\ell}) = \text{given} \quad (89.8)$$

$$P(\underline{s}_\sigma = 0 | d_{pa_\sigma}, \underline{\ell}) = \underbrace{P(\underline{s}_\sigma = 0 | \underline{\ell})}_{e^{-\theta_{\sigma|0}}} \prod_{j \in pa_\sigma} \underbrace{P(\underline{s}_\sigma = 0 | d_j)}_{e^{-\theta_{\sigma|j} d_j}} \quad (89.9)$$

$$= e^{-\theta_{\sigma|0} - \sum_{j \in pa_\sigma} \theta_{\sigma|j} d_j} \quad (89.10)$$

where $\theta_{\sigma|0}, \theta_{\sigma|j} > 0$. This $P(\underline{s}_\sigma = 0 | d_{pa_\sigma}, \underline{\ell})$ corresponds to the noisy-or model (See Chapter 59.).

$$P(\underline{s}_\sigma = 1 | d_{pa_\sigma}) = 1 - e^{-\theta_{\sigma|0} - \sum_{j \in pa_\sigma} \theta_{\sigma|j} d_j} \quad (89.11)$$

Define

$$x_\sigma = \theta_{\sigma|0} + \sum_{j \in pa_\sigma} \theta_{\sigma|j} d_j \quad (89.12)$$

Suppose $A \subset \mathbb{Z}_{[1,ns]}$, $A^c = \mathbb{Z}_{[1,ns]} - A$. Then

$$P(d_j | \underline{s}_A = 0, \underline{s}_{A^c} = 1) = \frac{P(\underline{s}_A = 0, \underline{s}_{A^c} = 1 | d_j) P(d_j)}{P(\underline{s}_A = 0, \underline{s}_{A^c} = 1)} \quad (89.13)$$

$$= \mathcal{N}(!d_j) \sum_{d_{!j}} P(\underline{s}_A = 0, \underline{s}_{A^c} = 1 | d^{nd}) P(d^{nd}) \quad (89.14)$$

$$= \mathcal{N}(!d_j) \sum_{d_{!j}} P(\underline{s}_{A^c} = 1 | d^{nd}) P(\underline{s}_A = 0 | d^{nd}) P(d^{nd}) \quad (89.15)$$

$$= \mathcal{N}(!d_j) \sum_{d_{!j}} \left\{ \begin{array}{l} \overbrace{\prod_{\sigma \in A^c} (1 - e^{-x_\sigma})}^{\text{call } \mathcal{P}_\sigma} \\ \prod_{\sigma \in A} e^{-\theta_{\sigma|0}} \prod_{j \in pa_\sigma} [e^{-\theta_{\sigma|j}}]^{d_j} \\ \prod_{j \in \mathbb{Z}_{[1,nd]}} P(d_j) \end{array} \right\} \quad (89.16)$$

Summing over $d_{!j}$ seems crazy, because $nd \gg 1$, but we will approximate the summand so that the sum can be done in closed form.

Define

$$f(x_\sigma) = \ln(1 - e^{-x_\sigma}) \quad (89.17)$$

and

$$\mathcal{P}_\sigma = 1 - e^{-x_\sigma} = e^{f(x_\sigma)} \quad (89.18)$$

$f(x_\sigma)$ is a concave function. See Fig.89.4 for a plot of it.

Next, we shall find a lower and upper bound for \mathcal{P}_σ .

To find the upper bound, we will use dual functions, which are discussed in Section 89.1.

Let $F(X_\sigma)$ be the dual function of $f(x_\sigma) = \ln(1 - e^{-x_\sigma})$. In Section 89.1, we show that

$$F(X_\sigma) = \min_{x_\sigma} (x_\sigma X_\sigma - f(x_\sigma)) \quad (89.19)$$

$$= -X_\sigma \ln X_\sigma + (1 + X_\sigma) \ln(1 + X_\sigma) \quad (89.20)$$

and

$$f(x_\sigma) \leq x_\sigma X_\sigma - F(X_\sigma) . \quad (89.21)$$

Therefore

$$\mathcal{P}_\sigma = e^{f(x_\sigma)} \quad (89.22)$$

$$\leq e^{x_\sigma X_\sigma - F(X_\sigma)} \quad (89.23)$$

$$= e^{-F(X_\sigma)} e^{-\theta_{\sigma|0} X_\sigma} \underbrace{\prod_{j \in pa_\sigma} [e^{-\theta_{\sigma|j} X_\sigma}]^{d_j}}_{\text{call } \mathcal{B}(X_\sigma)} \quad (89.24)$$

To find a lower bound for P_σ , we will use Jensen's inequality, which is discussed in Section C.41. Let $q_{j|\sigma} \in [0, 1]$ satisfy $\sum_j q_{j|\sigma} = 1$. Then

$$\mathcal{P}_\sigma = e^{f(x_\sigma)} \quad (89.25)$$

$$= e^{f(\theta_{\sigma|0} + \sum_j \theta_{\sigma|j} d_j)} \quad (89.26)$$

$$= e^{f\left(\theta_{\sigma|0} + \sum_j q_{j|\sigma} \frac{\theta_{\sigma|j} d_j}{q_{j|\sigma}}\right)} \quad (89.27)$$

$$\geq e^{\sum_j q_{j|\sigma} f\left(\theta_{\sigma|0} + \frac{\theta_{\sigma|j} d_j}{q_{j|\sigma}}\right)} \quad (89.28)$$

$$= e^{\sum_j q_{j|\sigma} \left[d_j f\left(\theta_{\sigma|0} + \frac{\theta_{\sigma|j} d_j}{q_{j|\sigma}}\right) + (1-d_j) f(\theta_{\sigma|0}) \right]} \quad (89.29)$$

$$= \underbrace{e^{f(\theta_{\sigma|0}) + \sum_j q_{j|\sigma} d_j \left[f\left(\theta_{\sigma|0} + \frac{\theta_{\sigma|j} d_j}{q_{j|\sigma}}\right) - f(\theta_{\sigma|0}) \right]}}_{\text{call } \mathcal{A}(q_{\cdot|\sigma})} \quad (89.30)$$

In conclusion,

$$\boxed{\mathcal{A}(q_{\cdot|\sigma}) \leq \mathcal{P}_\sigma \leq \mathcal{B}(X_\sigma)} \quad (89.31)$$

with variational parameters $q_{\cdot|\sigma}$ and X_σ .

89.1 Convex/Concave Dual Functions

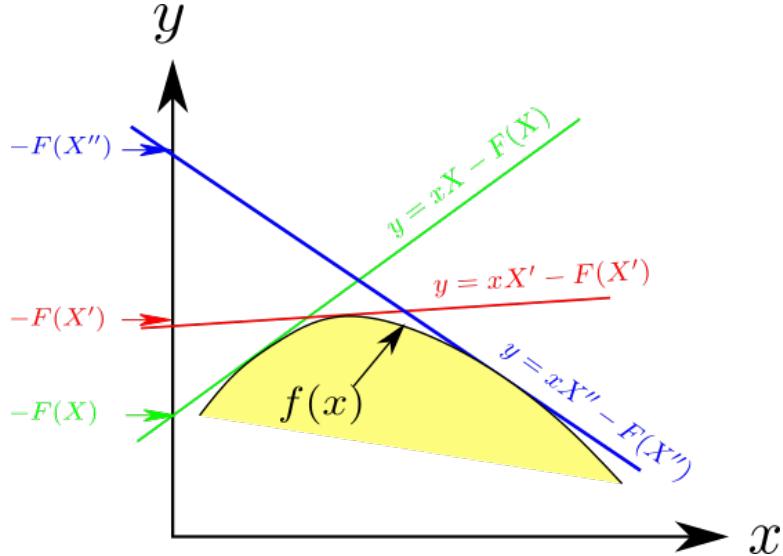


Figure 89.2: The dual function $f(x)$ of a concave function $F(X)$ is the osculating (kissing) locus of the family of lines $y = Xx - F(X)$ for all X .

If $f(x)$ is a concave function, we define its **concave dual (a.k.a. concave conjugate) function** by

$$F(X) = \min_x (X^T x - f(x)) . \quad (89.32)$$

Claim 145 *If $f(x)$ is a concave function with dual $F(X)$, then*

$$f(x) = \min_X (X^T x - F(X)) \quad (89.33)$$

$$\nabla_X F(X(x)) = (\nabla_x f)^{-1}(x) \quad (89.34)$$

Eqs.(89.32) and (89.33) are also true if we replace the words “concave” with “convex” and “min” with “max”.

proof:

Let x^* be the value of x which minimizes $F(X) = X^T x - f(x)$ with respect to x . Hence,

$$F(X) = X^T x^* - f(x^*) . \quad (89.35)$$

Rearranging terms in Eq.(89.35), we get

$$f(x^*) = X^T x^* - F(X) \quad (89.36)$$

Let X^* be the value of X which minimizes $f(x) = X^T x - F(X)$ with respect to X . Hence,

$$f(x) = (X^*)^T x - F(X^*) \quad (89.37)$$

Replacing X^* by X and x by x^* in Eq.(89.37) yield Eq.(89.36).

Note that minimization with respect to x is achieved if

$$X_i = \partial_{x_i} f(x), \quad X = \nabla_x f(x) \quad (89.38)$$

whereas minimization with respect to X is achieved if

$$x_i = \partial_{X_i} F(X), \quad x = \nabla_X F(X) \quad (89.39)$$

Hence,

$$x = \nabla_X F(\nabla_x f(x)) \quad (89.40)$$

QED

See Fig.89.2 for a geometrical representation of Eq.(89.33).

We end this section on dual functions with some examples.

1. Find the dual function of

$$\boxed{f(x) = e^x} . \quad (89.41)$$

$$X = \partial_{x^*} f \quad (89.42)$$

$$= e^{x^*} \quad (89.43)$$

$$x^* = \ln X \quad (89.44)$$

$$f(x^*) = X \quad (89.45)$$

$$F(X) = x^* X - f(x^*) \quad (89.46)$$

$$= X \ln X - X \quad (89.47)$$

2. Find the dual function of

$$\boxed{f(x) = f + f'x + \frac{1}{2}f''x^2} . \quad (89.48)$$

$$X = \partial_{x^*} f \quad (89.49)$$

$$= f' + f''x^* \quad (89.50)$$

$$x^* = \frac{X - f'}{f''} \quad (89.51)$$

$$f(x^*) = f + f' \left[\frac{X - f'}{f''} \right] + \frac{1}{2}f'' \left[\frac{X - f'}{f''} \right]^2 \quad (89.52)$$

$$= \left[f - \frac{(f')^2}{2f''} \right] + X^2 \left[\frac{1}{2f''} \right] \quad (89.53)$$

$$F(X) = Xx^* - f(x^*) \quad (89.54)$$

$$= X \left[\frac{X - f'}{f''} \right] - f(x^*) \quad (89.55)$$

$$= \left[-f + \frac{(f')^2}{2f''} \right] + X \left[\frac{-f'}{f''} \right] + X^2 \left[\frac{1}{2f''} \right] \quad (89.56)$$

Note that when $f = f' = 0$, we get

$$f(x) = \frac{f''x^2}{2}, \quad F(X) = \frac{X^2}{2f''} \quad (89.57)$$

Note that if f is convex (resp., concave), F is convex too (resp., concave too).

3. Find the dual function of

$$\boxed{f(x) = \ln(1 - e^{-x})} \quad (89.58)$$

$$X = \partial_{x^*} f(x^*) \quad (89.59)$$

$$= \frac{e^{-x^*}}{1 - e^{-x^*}} \quad (89.60)$$

$$X = (1 + X)e^{-x^*} \quad (89.61)$$

$$x^* = \ln \frac{1 + X}{X} \quad (89.62)$$

$$f(x^*) = \ln \left(1 - \frac{X}{1 + X} \right) = -\ln(1 + X) \quad (89.63)$$

$$F(X) = Xx^* - f(x^*) \quad (89.64)$$

$$= X \ln \frac{1 + X}{X} + \ln(1 + X) \quad (89.65)$$

$$= -X \ln X + (1 + X) \ln(1 + X) \quad (89.66)$$

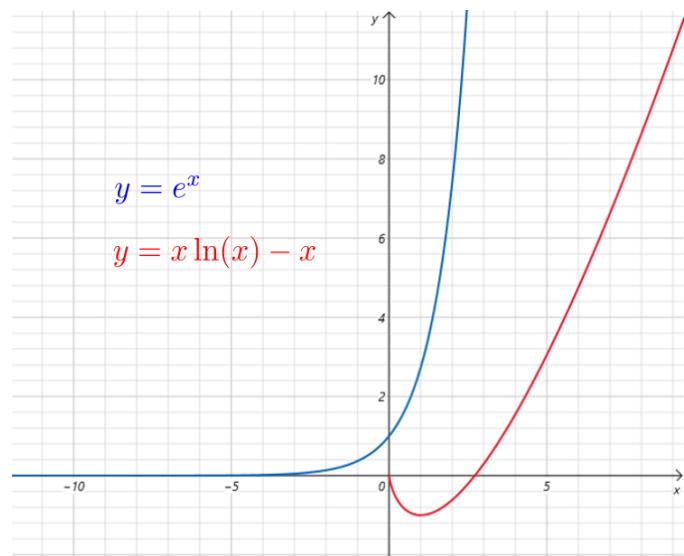


Figure 89.3: $f(x) = e^x$ and its dual.

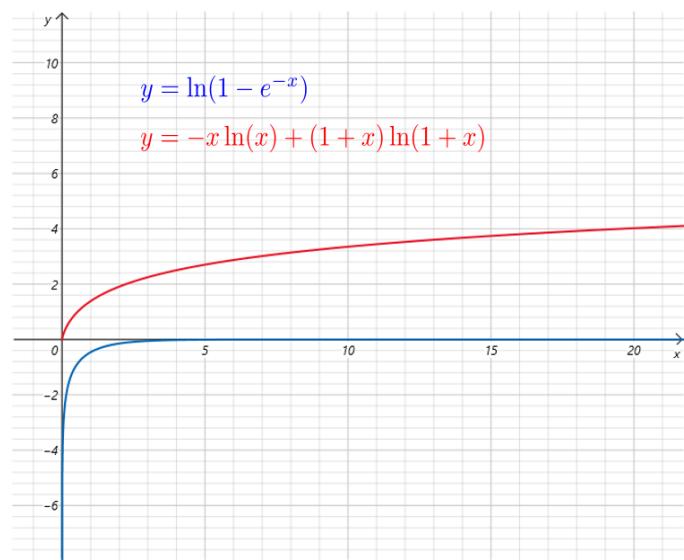


Figure 89.4: $f(x) = \ln(1 - e^{-x})$ and its dual.

Chapter 90

Variational Bayesian Approximation via D_{KL}

For more info and references about this topic, see Ref.[147].

A Variational Bayesian Approximation (VBA) is when we approximate a probability distribution by another probability distribution that depends on a continuous “variational parameter”. This parameter is adjusted within its range of possible values, to make the approximation as good as possible. There are many VBA methods. VBA methods are inspired by ancient methods used in Calculus of Variations applied to Physics and Engineering problems.

In this chapter, we do VBA via the Kullback-Leibler divergence D_{KL} . We approximate the probability distribution $P(h|\vec{x})$, where h are the hidden variables and \vec{x} is the data. More precisely, suppose $\underline{h} \in S_{\underline{h}}$ and $\underline{q} \in S_{\underline{h}}$. Suppose $\underline{\vec{x}} \in S_{\underline{x}}^{nsam}$ is a vector of $nsam$ samples and the samples $\underline{x}[\sigma] \in S_{\underline{x}}$ are i.i.d.. The VBA is simply an approximation $P_{\underline{q}|\underline{\vec{x}}}$ to $P_{\underline{h}|\underline{\vec{x}}}$:

$$P_{\underline{h}|\underline{\vec{x}}}(h|\vec{x}) \approx P_{\underline{q}|\underline{\vec{x}}}(h|\vec{x}) \quad (90.1)$$

obtained by minimizing the Kullback-Leibler divergence $D_{KL}(P_{\underline{q}|\underline{\vec{x}}} \parallel P_{\underline{h}|\underline{\vec{x}}})$ over all $P_{\underline{q}|\underline{\vec{x}}}$. The minimization is usually subject to some constraints on the admissible forms of $P_{\underline{q}|\underline{\vec{x}}}$.

$D_{KL}(Q \parallel P) \neq D_{KL}(P \parallel Q)$; i.e., D_{KL} is not symmetric. So why do we use $D_{KL}(P_{\underline{q}|\underline{\vec{x}}} \parallel P_{\underline{h}|\underline{\vec{x}}})$ instead of $D_{KL}(P_{\underline{h}|\underline{\vec{x}}} \parallel P_{\underline{q}|\underline{\vec{x}}})$? Because $D_{KL}(P_{\underline{h}|\underline{\vec{x}}} \parallel P_{\underline{q}|\underline{\vec{x}}})$ requires knowledge of $P_{\underline{h}|\underline{\vec{x}}}$, but calculating $P_{\underline{h}|\underline{\vec{x}}}$ is what we are trying to do in the first place.

See Fig.90.1 for some intuition on what minimizing $D_{KL}(P_{\underline{q}|\underline{\vec{x}}} \parallel P_{\underline{h}|\underline{\vec{x}}})$ means.

Suppose $\underline{h} = (\underline{h}_0, \underline{h}_1, \dots, \underline{h}_{nh-1})$ and $\underline{q} = (\underline{q}_0, \underline{q}_1, \dots, \underline{q}_{nh-1})$ where $\underline{h}_i \in S_{\underline{h}_i}$ and $\underline{q}_i \in S_{\underline{h}_i}$ for all i . We say \underline{q} and \underline{h} have nh mirroring components and those of \underline{q} are independent at fixed $\underline{\vec{x}}$ if

$$P_{\underline{q}|\underline{\vec{x}}}(h|\vec{x}) = \prod_i P_{\underline{q}_i|\underline{\vec{x}}}(h_i|\vec{x}) . \quad (90.2)$$

The bnet Fig.90.2 describes the scenario that we have in mind: The samples $\underline{x}[\sigma]$ are

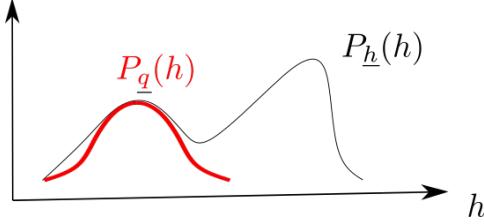


Figure 90.1: If $P_{\underline{q}}(h)$ is Gaussian shaped and $P_{\underline{h}}(h)$ has multiple bumps (modes) then $D_{KL}(P_{\underline{q}} \parallel P_{\underline{h}})$ is minimized when $P_{\underline{q}}$ fits one of the modes of $P_{\underline{h}}$. That is because $D_{KL}(P_{\underline{q}} \parallel P_{\underline{h}}) = \sum_h P_{\underline{q}}(h) \ln \frac{P_{\underline{q}}(h)}{P_{\underline{h}}(h)}$ is a weighted average with weights $P_{\underline{q}}$, so nothing going on outside the support of $P_{\underline{q}}$ influences much the final average.

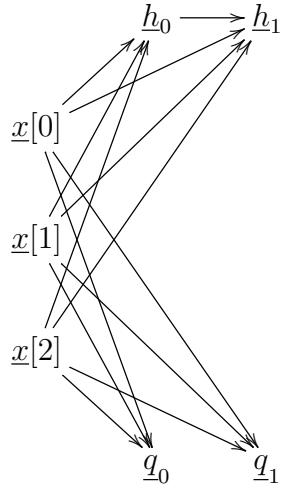


Figure 90.2: \underline{q} and \underline{h} have $nh = 2$ mirroring components and those of \underline{q} are independent at fixed \vec{x} .

i.i.d.. Each component \underline{h}_i of \underline{h} has a mirroring component \underline{q}_i in \underline{q} . The components of \underline{h} are correlated whereas those of \underline{q} are independent at fixed \vec{x} .

Claim 146 *If \underline{q} and \underline{h} have nh mirroring components and those of \underline{q} are independent at fixed \vec{x} and $\bar{D}_{KL}(P_{\underline{q}|\vec{x}} \parallel P_{\underline{h}|\vec{x}})$ is minimum over all $P_{\underline{q}|\vec{x}}$, then*

$$P_{\underline{q}_i|\vec{x}}(q_i|\vec{x}) = \mathcal{N}(!q_i) e^{E_{(q_j)_{j \neq i}} [\ln P_{\underline{h}|\vec{x}}(h=q|\vec{x})]} \quad (90.3)$$

$$= \mathcal{N}(!q_i) e^{E_{(q_j)_{j \neq i}} [\ln P_{\underline{h},\vec{x}}(h=q,\vec{x})]} \quad (90.4)$$

for all i .

proof:

Since all quantities in Eq.(90.3) are conditioned on \vec{x} , let us omit all mention of \vec{x} in this proof.

Let

$$\mathcal{L} = \mathcal{L}_0 + \mathcal{L}_1 \quad (90.5)$$

where

$$\mathcal{L}_0 = D_{KL}(P_{\underline{q}} \| P_{\underline{h}}) \quad (90.6)$$

$$= \sum_h P_{\underline{q}}(h) \ln \frac{P_{\underline{q}}(h)}{P_{\underline{h}}(h)} \quad (90.7)$$

$$= \sum_h P_{\underline{q}}(h) \ln P_{\underline{q}}(h) - \sum_h P_{\underline{q}}(h) \ln P_{\underline{h}}(h) \quad (90.8)$$

$$= \sum_i \sum_{h_i} P_{\underline{q}_i}(h_i) \ln P_{\underline{q}_i}(h_i) - \sum_h P_{\underline{q}}(h) \ln P_{\underline{h}}(h) \quad (90.9)$$

and

$$\mathcal{L}_1 = \sum_i \lambda_i \left[\sum_{h_i} P_{\underline{q}_i}(h_i) - 1 \right]. \quad (90.10)$$

Then

$$\delta \mathcal{L} = \sum_i \sum_{h_i} \delta P_{\underline{q}_i}(h_i) \left[\ln P_{\underline{q}_i}(h_i) + 1 + \lambda_i - \frac{1}{nh} \sum_{(h_j)_{j \neq i}} \prod_{(h_j)_{j \neq i}} \{P_{\underline{q}_j}(h_j)\} \ln P_{\underline{h}}(h) \right]. \quad (90.11)$$

Hence,

$$P_{\underline{q}_i}(h_i) = \mathcal{N}(!h_i) e^{\sum_{(h_j)_{j \neq i}} \left\{ \prod_{(h_j)_{j \neq i}} P_{\underline{q}_j}(h_j) \right\} \ln P_{\underline{h}}(h)}. \quad (90.12)$$

QED

Note that Eq.(90.3) yields a system of nh nonlinear equations in nh unknowns $(P_{\underline{q}_i|\vec{x}})_{i=0,1,\dots,nh-1}$. This system is usually solved recursively.

90.1 Free Energy $\mathcal{F}(\vec{x})$

To simplify the notation below, let us introduce the following abbreviations:

$$P(h|\vec{x}) = P_{\underline{h}|\vec{x}}(h|\vec{x}) \quad (90.13)$$

$$P(h, \vec{x}) = P_{\underline{h}, \vec{x}}(h, \vec{x}) \quad (90.14)$$

$$P(\vec{x}) = P_{\vec{x}}(\vec{x}) \quad (90.15)$$

Note that

$$D_{KL}(P_{\underline{q}|\vec{x}} \| P_{\underline{h}|\vec{x}}) = \sum_h P_{\underline{q}|\vec{x}}(h|\vec{x}) \ln \frac{P_{\underline{q}|\vec{x}}(h|\vec{x})}{P(h|\vec{x})} \quad (90.16)$$

$$= \sum_h P_{\underline{q}|\vec{x}}(h|\vec{x}) \ln \frac{P_{\underline{q}|\vec{x}}(h|\vec{x})}{P(h, \vec{x})} + \ln P(\vec{x}) \quad (90.17)$$

$$= \mathcal{F}(\vec{x}) + \ln P(\vec{x}) \quad (90.18)$$

Hence, the **Free energy** $\mathcal{F}(\vec{x})$ is defined as

$$\mathcal{F}(\vec{x}) = \sum_h P_{\underline{q}|\vec{x}}(h|\vec{x}) \ln \frac{P_{\underline{q}|\vec{x}}(h|\vec{x})}{P(h, \vec{x})} \quad (90.19)$$

$$= E_{\underline{q}|\vec{x}} \left[\ln \frac{P_{\underline{q}|\vec{x}}(q|\vec{x})}{P_{\underline{h}, \vec{x}}(q, \vec{x})} \right]. \quad (90.20)$$

The name free energy is justified because

$$\mathcal{F}(\vec{x}) = \underbrace{- \sum_h P_{\underline{q}|\vec{x}}(h|\vec{x}) \ln P_{\underline{h}, \vec{x}}(h, \vec{x})}_{U, \text{ Internal Energy}} + \underbrace{\sum_h P_{\underline{q}|\vec{x}}(h|\vec{x}) \ln P_{\underline{q}|\vec{x}}(h|\vec{x})}_{-S, \text{ minus Entropy}}. \quad (90.21)$$

It is also common to define a quantity called “ELBO” to be the negative of the free energy.

$$ELBO(\vec{x}) = -\mathcal{F}(\vec{x}) \quad (90.22)$$

ELBO stands for “Evidence Lower BOund”. That name is justified because

$$\underbrace{\ln P_{\vec{x}}(\vec{x})}_{\text{evidence} \leq 0} = \underbrace{D_{KL}(P_{\underline{q}|\vec{x}} \| P_{\underline{h}|\vec{x}})}_{\geq 0} - |ELBO(\vec{x})|. \quad (90.23)$$

Some properties of \mathcal{F} are:

- \mathcal{F} is non-negative.

$$\underbrace{D_{KL}(P_{\underline{q}|\vec{x}} \| P_{\underline{h}|\vec{x}})}_{\geq 0} + \underbrace{\ln \frac{1}{P_{\vec{x}}(\vec{x})}}_{\geq 0} = \mathcal{F}(\vec{x}) \quad (90.24)$$

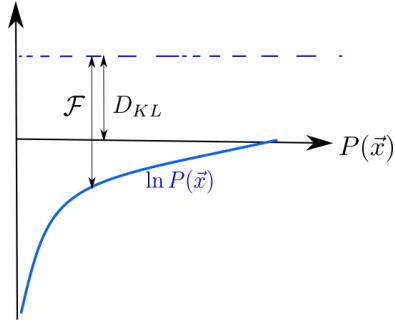


Figure 90.3: $D_{KL} + \ln \frac{1}{P(\vec{x})} = \mathcal{F}$.

- KL divergence is min iff \mathcal{F} is min at fixed $P(\vec{x})$.

During a variation δ that holds $P(\vec{x})$ fixed, the KL divergence and \mathcal{F} change by the same amount:

$$\delta D_{KL}(P_{\underline{g}|\vec{x}} \parallel P_{\underline{h}|\vec{x}}) = \delta \mathcal{F}(\vec{x}) \quad (90.25)$$

Chapter 91

XGBoost

This paper is based on the original XGBoost paper Ref.[6] and the excellent StatQuest videos (highly recommended) [59].

Extreme Gradient Boosting (XGBoost) (Chen, Guestrin, 2016) improves gradient boosting (Friedman, 1999) in a number of ways, such as by using a quadratic rather than linear approximation for the variational function.

In XGBoost, one calculates a sequence of functions where each function tries to correct the errors of the previous function. Then a series (i.e., linear combination) of those functions yields an estimate \hat{y}^σ of the target attribute y^σ . As will be shown in this chapter, XGBoost can be used in two cases: Regression (continuous target attribute) and Binary Classification (binary target attribute)¹. An implementation of the XGBoost algorithm is available as open source. It's written in C++, with Python and R interfaces.

Boosting (see this chapter on XGBoost and Chapter 1 on AdaBoost) and bagging (see Chapter 66 on Random Forest) are two methods of building a classifier function from an ensemble of classifier functions. These two methods are most commonly applied to dtrees: Boosting for an ensemble of small dtrees, and Bagging for a random forest (which is an ensemble of dtrees that are usually much more complicated than small dtrees).

91.1 Divergences

To set up a cost function for XGBoost, we begin by defining 2 types of “divergences” and calculating the first and second derivatives of those divergences:

- **Divergence for regression** (i.e., continuous classification, continuous target attribute). For $x, y \in \mathbb{R}$

$$D_{reg}(x, y) = \frac{1}{2}(x - y)^2 \quad (91.1)$$

¹XGBoost can also be used for classification into more than two classes, as long as a suitable Divergence function can be defined.

- **Divergence for binary classification** (i.e., binary target attribute). For $p, q \in [0, 1]$.

$$D_{bc}(p, q) = -\{p \ln q + (1-p) \ln(1-q)\} \quad (91.2)$$

$$= CE(\{p, 1-p\} \| \{q, 1-q\}) \quad (91.3)$$

The cross-entropy $CE()$ is defined in Chapter C.

Claim 147

$$\partial_{\hat{y}} D_{reg}(y, \hat{y}) = \hat{y} - y \quad (91.4)$$

$$\partial_{\hat{y}}^2 D_{reg}(y, \hat{y}) = 1 \quad (91.5)$$

$$D(y, \hat{y} + h) \approx D(y, \hat{y}) + (\hat{y} - y)h + \frac{1}{2}h^2 \quad (91.6)$$

proof: Obvious.

QED

Recall from Section C.20 in Chapter C that

$$\hat{y} = \text{smoid}(\text{lodds}(\hat{y})) \quad (91.7)$$

Let us abbreviate $s() = \text{smoid}()$ and $l() = \text{lodds}()$ so

$$\hat{y} = s(l(\hat{y})). \quad (91.8)$$

Claim 148 ²

$$\partial_l D(y, \hat{y}(l)) = \hat{y} - y \quad (91.9)$$

$$\partial_l^2 D(y, \hat{y}(l)) = \hat{y}(1 - \hat{y}) \quad (91.10)$$

$$D(y, \hat{y}(l + \Delta l)) \approx D(y, \hat{y}(l)) + (\hat{y} - y)\Delta l + \frac{1}{2}\hat{y}(1 - \hat{y})(\Delta l)^2 \quad (91.11)$$

proof:

²Note that $D_{bc}(y, \hat{y}) \neq D_{bc}(\hat{y}, y)$; i.e., D_{bc} is not symmetric in its two arguments. Normally $0 < \hat{y} < 1$ and $y \in \{0, 1\}$. Since the log is applied only to the second argument, to avoid logs of zero, it is better to have \hat{y} as the second argument.

$$\frac{\partial D(y, s)}{\partial s} = -\partial_s \{y \ln s + (1 - y) \ln(1 - s)\} \quad (91.12)$$

$$= -\frac{y}{s} + \frac{1-y}{1-s} \quad (91.13)$$

$$= \frac{-y(1-s) + (1-y)s}{s(1-s)} \quad (91.14)$$

$$= \frac{-y+s}{s(1-s)} \quad (91.15)$$

Recall that $\text{smoid}'(l) = \text{smoid}(l)[1 - \text{smoid}(l)]$ so

$$\frac{\partial s}{\partial l} = s(1-s). \quad (91.16)$$

$$\frac{\partial D(y, s)}{\partial l} = \frac{\partial s}{\partial l} \frac{\partial D(y, s)}{\partial s} = -y + s = -y + \hat{y} \quad (91.17)$$

$$\frac{\partial^2 D(y, s)}{\partial l^2} = \frac{\partial s}{\partial l} = s(1-s) \quad (91.18)$$

QED

91.2 Minimizing Cost function for single tree

Let

$\sigma \in \Sigma$ be an individual in a population Σ

$x^\sigma \in S_{\underline{x}}$ be a feature vector $x^\sigma = (x_i^\sigma)_{i=0,1,\dots,nf-1}$

$t \in \{0, 1, \dots, nt - 1\}$ be the tree index

\mathcal{L}_t be the set of leafs of tree t

$\ell \in \mathcal{L}_t$ be a leaf in tree t

$w_t^\ell \in \mathbb{R}$

$f_t : S_{\underline{x}} \rightarrow \mathbb{R}$

$\ell_t : \Sigma \rightarrow \mathcal{L}_t$, $\ell_t(\Sigma) = \mathcal{L}_t$.

$\ell_t(\sigma)$ be the leaf of individual σ in tree t

$\Sigma_t^\ell = \{\sigma \in \Sigma : \ell_t(\sigma) = \ell\}$

Define the function f_t by

$$f_t(x^\sigma) = \sum_{\ell \in \mathcal{L}_t} w_t^\ell \mathbb{1}(\sigma \in \Sigma_t^\ell) = w_t^{\ell_t(\sigma)}. \quad (91.19)$$

$f_t(x^\sigma)$ gives the output value for tree t and feature vector x^σ .

Define the **cost function** for tree t as follows

$$\mathcal{C}_t = \sum_{\sigma} D(\hat{y}_t^{\sigma}, y^{\sigma}) + \underbrace{\sum_{\ell \in \mathcal{L}_t} \left[\gamma + \frac{\lambda}{2} (w_t^{\ell})^2 \right]}_{\text{regulator}}, \quad (91.20)$$

where $\gamma > 0, \lambda > 0$ are regulator parameters.

The estimate \hat{y}_t^{σ} , using trees from 0 to t , of the target attribute y^{σ} , is defined by

$$\hat{y}_0^{\sigma} = f_0(x^{\sigma}) = \text{arbitrary constant, XGBoost default for this is 0.5} \quad (91.21)$$

$$\hat{y}_1^{\sigma} = f_0(x^{\sigma}) + f_1(x^{\sigma}) \quad (91.22)$$

$$\hat{y}_2^{\sigma} = f_0(x^{\sigma}) + f_1(x^{\sigma}) + f_2(x^{\sigma}) \quad (91.23)$$

$$\vdots \quad (91.24)$$

$$\hat{y}_t^{\sigma} = \sum_{t' \leq t} f_{t'}(x^{\sigma}) = \hat{y}_{t-1}^{\sigma} + f_t(x^{\sigma}) \quad (91.25)$$

In the cost function \mathcal{C}_t , we approximate the divergence $D()$ by a second order Taylor approximation:

$$D(y^{\sigma}, \hat{y}_t^{\sigma}) = D(y^{\sigma}, \hat{y}_{t-1}^{\sigma} + \underbrace{f_t(x^{\sigma})}_{\delta}) \quad (91.26)$$

$$\approx D(y^{\sigma}, \hat{y}_{t-1}^{\sigma}) + a_t^{\sigma} \delta + \frac{1}{2} b_t^{\sigma} \delta^2 \quad (91.27)$$

$$= D(y^{\sigma}, \hat{y}_{t-1}^{\sigma}) + a_t^{\sigma} w_t^{\ell_t(\sigma)} + \frac{1}{2} b_t^{\sigma} (w_t^{\ell_t(\sigma)})^2, \quad (91.28)$$

where

$$a_t^{\sigma} = [\partial_{\hat{y}} D(y^{\sigma}, \hat{y})]_{\hat{y}=\hat{y}_{t-1}^{\sigma}}, \quad (91.29)$$

$$b_t^{\sigma} = [\partial_{\hat{y}}^2 D(y^{\sigma}, \hat{y})]_{\hat{y}=\hat{y}_{t-1}^{\sigma}}. \quad (91.30)$$

a_t^{σ} is called g for gradient and b_t^{σ} is called h for Hessian in Ref.[6]. Table 91.1 gives the values for a_t^{σ} and b_t^{σ} for Regression (reg) and Binary classification (bc).

Define the **residual** for tree t and individual σ by

$$r_t^{\sigma} = y^{\sigma} - \hat{y}_{t-1}^{\sigma}. \quad (91.31)$$

Note that

$$\sum_{\sigma} = \sum_{\ell \in \mathcal{L}_t} \sum_{\sigma \in \Sigma_t^{\ell}}. \quad (91.32)$$

Define

	Regression ($y^\sigma \in \mathbb{R}$, $\hat{y}_t^\sigma \in \mathbb{R}$, $D = D_{reg}$)	Binary Classification ($y^\sigma \in \{0, 1\}$, $\hat{y}_t^\sigma \in [0, 1]$, $D = D_{bc}$)
a_t^σ	$\hat{y}_{t-1}^\sigma - y^\sigma = \text{neg.residual}$	$\hat{y}_{t-1}^\sigma - y^\sigma = \text{neg.residual}$
b_t^σ	1	$\hat{y}_{t-1}^\sigma(1 - \hat{y}_{t-1}^\sigma)$

Table 91.1: The first (a_t^σ) and second (b_t^σ) derivatives for Regression (reg) and Binary classification (bc).

$$A_t^\ell = \sum_{\sigma \in \Sigma_t^\ell} a_t^\sigma \quad (91.33)$$

and

$$B_t^\ell = \sum_{\sigma \in \Sigma_t^\ell} b_t^\sigma. \quad (91.34)$$

Now we can rewrite the cost function as

$$\mathcal{C}_t = \underbrace{\sum_{\sigma} D(\hat{y}_{t-1}^\sigma, y^\sigma)}_{\mathcal{K}} + \sum_{\ell \in \mathcal{L}_t} \left[A_t^\ell w_t^\ell + \frac{1}{2} B_t^\ell (w_t^\ell)^2 \right] + \sum_{\ell \in \mathcal{L}_t} \left[\gamma + \frac{\lambda}{2} (w_t^\ell)^2 \right] \quad (91.35)$$

$$= \sum_{\ell \in \mathcal{L}_t} \left[\gamma + A_t^\ell w_t^\ell + \frac{1}{2} (B_t^\ell + \lambda) (w_t^\ell)^2 \right] \quad (\text{absorbed } \mathcal{K} \text{ into } \gamma) \quad (91.36)$$

The cost function \mathcal{C}_t can be minimized over w_t^ℓ :

$$0 = \delta \mathcal{C}_t = \sum_{\ell \in \mathcal{L}_t} \delta w_t^\ell [A_t^\ell + (B_t^\ell + \lambda) w_t^\ell] \quad (91.37)$$

Therefore, the cost is minimized when

$$w_t^\ell = \frac{-A_t^\ell}{B_t^\ell + \lambda}. \quad (91.38)$$

The optimum cost is

$$\mathcal{C}_t = \sum_{\ell \in \mathcal{L}_t} \underbrace{\left[\gamma - \frac{(A_t^\ell)^2}{2(B_t^\ell + \lambda)} \right]}_{\mathcal{C}_t^\ell} \quad (91.39)$$

\mathcal{C}_t^ℓ is called the **similarity score** for leaf ℓ and tree t . Note that an increase in similarity decreases the cost.

91.3 Leaf Splitting

Suppose leaf ℓ_0 splits into leafs ℓ_L and ℓ_R . Then

$$\Sigma_t^{\ell_0} = \Sigma_t^{\ell_L} \cup \Sigma_t^{\ell_R}, \quad \Sigma_t^{\ell_L} \cap \Sigma_t^{\ell_R} = \emptyset \quad (91.40)$$

so

$$A_t^{\ell_0} = A_t^{\ell_L} + A_t^{\ell_R} \quad (91.41)$$

and

$$B_t^{\ell_0} = B_t^{\ell_L} + B_t^{\ell_R}. \quad (91.42)$$

Hence,

$$\mathcal{C}_t^{\ell_j} = \gamma - \frac{(A_t^{\ell_j})^2}{2(B_t^{\ell_j} + \lambda)} \quad \text{for } j = L, R \quad (91.43)$$

$$\mathcal{C}_t^{\ell_0} = \gamma - \frac{(A_t^{\ell_L} + A_t^{\ell_R})^2}{2(B_t^{\ell_L} + B_t^{\ell_R} + \lambda)} \quad (91.44)$$

Define the **XGBoost branch Gain** for a binary tree branch by

$$Gain_{XGB} = \mathcal{C}_t^{\ell_0} - \mathcal{C}_t^{\ell_L} - \mathcal{C}_t^{\ell_R} \quad (91.45)$$

$$= \underbrace{\left\{ SS_t^{\ell_L} + SS_t^{\ell_R} - SS_t^{\ell_0} \right\}}_{\Delta SS_t} - \gamma \quad (91.46)$$

By splitting a leaf, we create a new binary branch with 2 new leafs. We successively add new branches to the tree until some stopping criterion is satisfied. We continue adding branches to the tree as long as they have a positive gain, and as long as the number of levels is smaller than an upper bound input parameter (XGBoost's default maximum number of tree levels is six). XGBoost also rules out leafs that have a denominator quantity B_t^ℓ (called the **cover**) smaller than a lower bound input parameter. This branching process is illustrated in Figs.91.1 and 91.2.

91.4 Pruning

If we raise γ or raise λ , branches that previously had positive gain may acquire a negative gain. Get rid of branches from highest level that now have negative gain. This will generate new branches. Get rid of new branches from the highest level that have negative gain. Continue this process until all highest level branches have positive gain. This may reduce a tree to a single node or even rule out the entire tree.

λ measures **level of insensitivity** to observations, and γ measures **level of tree simplicity**.

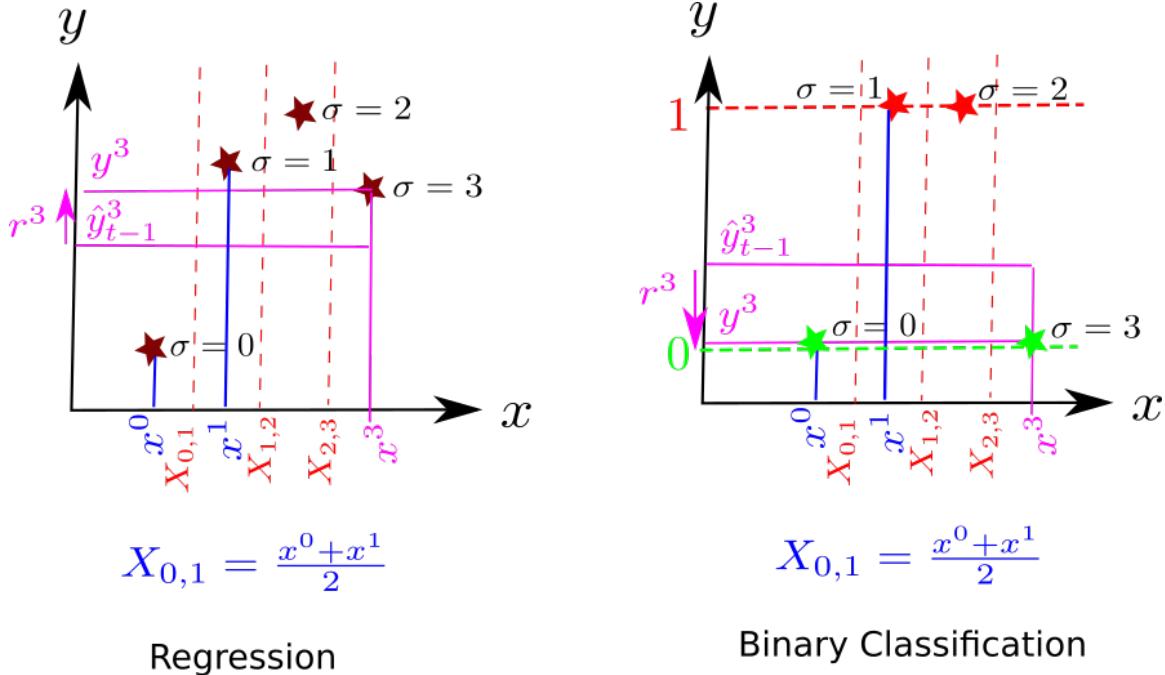
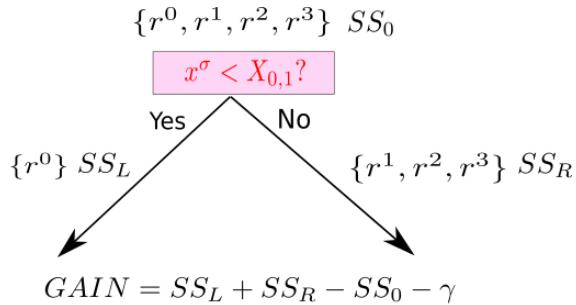


Figure 91.1: Plot of target attribute y^σ versus a feature vector x^σ with a single component. More generally, the feature vector $x^\sigma = (x_j^\sigma)_{j=0,1,\dots,nf-1}$ can have multiple components called features or attributes. The left plot refers to a population of 4 individuals and regression so $y^\sigma, \hat{y}^\sigma \in \mathbb{R}$. The right plot refers to a population of 4 individuals and binary classification so $y^\sigma \in \{0, 1\}, \hat{y}^\sigma \in [0, 1]$. $X_{j,j+1}$ for $j = 0, 1, 2$ is the average between 2 adjacent values of x^σ .



Repeat for $x^\sigma < X_{1,2}$? and $x^\sigma < X_{2,3}$?
 Choose inequality with highest GAIN

Figure 91.2: This figure refers to the situation of Fig.91.1. Out of all allowed tree branches, we choose the one with the highest gain.

91.5 Feature Binning

For really large population sizes $|\Sigma|$, it is convenient to bin the set $\{x_i^\sigma : \sigma \in \Sigma\}$ for each i .

A common bin type is quantiles. Quantiles are bins $[X_{j-1}, X_j]$ for $j = 0, 1, \dots, nbins - 1$ that all contain approximately the same number of points. For example, in Fig.91.2, with 1 point quantile bins, use $[X_{j-1}, X_j]$ for $j = 0, 1, 2, 3$. With 2 point quantile bins, use $[X_{j-1}, X_j]$ for $j = 0, 1$.

Use the right edge of each bin as the X_j in the question $x^\sigma < X_j?$, and choose the question which yields the highest gain.

91.6 Final estimate of target attribute

In this section, we will give a formula for the final estimate of the target attribute. Previously, we set the learning rate η to one. Here we restore η to an arbitrary value between 0 and 1.

Instead of using

$$f(x^\sigma) = \sum_{t=0}^{nt-1} f_t(x^\sigma), \quad (91.47)$$

we will use

$$f(x^\sigma) = \sum_{t=0}^{nt-1} (\eta)^t f_t(x^\sigma), \quad (91.48)$$

where $\eta \in [0, 1]$ is called the **learning rate**. This has the effect of compensating for the f_t 's with $t > nt - 1$ that would have been included had we used an infinite series. Also, think of Eq.(91.47) as an approximation (a truncated Taylor expansion in powers of Δx) of a function $f(x + \Delta x)$, and think of Eq.(91.48) as an analogous approximation of the function $f(x + \eta \Delta x)$.

For the case of Regression, we get:

$$f(x^\sigma) = \sum_t (\eta)^t \left(\frac{-A_t^{\ell_t(\sigma)}}{B_t^{\ell_t(\sigma)} + \lambda} \right). \quad (91.49)$$

For the case of Binary Classification, we get

$$f(x^\sigma) = \underbrace{\sum_{t=0}^{nt-1} (\eta)^t \text{smoid} \left(\frac{-A_t^{\ell_t(\sigma)}}{B_t^{\ell_t(\sigma)} + \lambda} \right)}_{\substack{\text{lodds(probability)} \\ \text{probability}}}. \quad (91.50)$$

91.7 Bnet for XGBoost

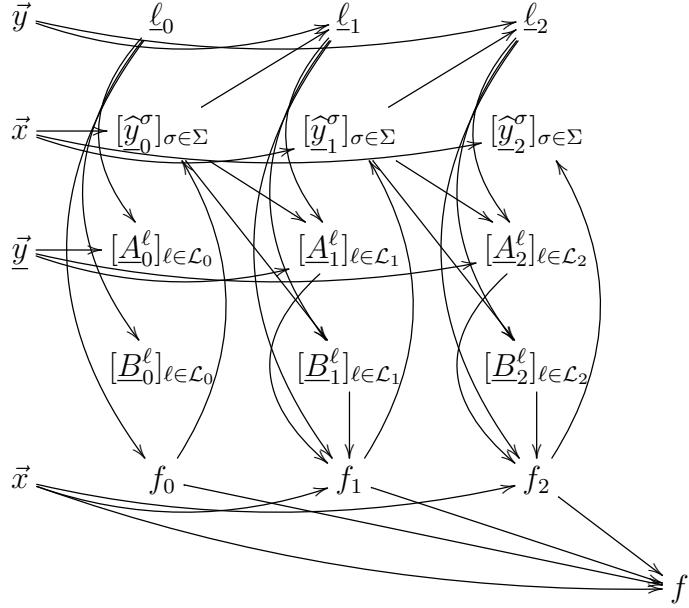


Figure 91.3: Bnet for XGBoost assuming $nt = 3$. Nodes that appear multiple times (namely \vec{x} and \vec{y}) should be considered the same node, drawn multiple times for clarity. The parameters λ, γ, η are taken to be global. Recall that $\ell_t : \Sigma \rightarrow \mathcal{L}_t$. From the function $\ell_t()$, we can derive its range $\mathcal{L}_t = \{\ell_t(\sigma) : \sigma \in \Sigma\}$, and $\Sigma_t^\ell = \{\sigma \in \Sigma : \ell_t(\sigma) = \ell\}$ for all $\ell \in \mathcal{L}_t$.

Fig.91.3 gives our bnet for the XGBoost algo assuming $nt = 3$. The TPMs, printed in blue, for this bnet, are as follows:

For $t = 0$, ℓ_0 describes a single node “tree” such that $\ell_0(x^\sigma) = f_0(x^\sigma) = 0.5$ for all $\sigma \in \Sigma$. For $t \geq 1$,

$$P(\ell_t | \vec{y}, [\hat{y}_{t-1}^\sigma]_\sigma) = \begin{array}{l} \text{build tree } \ell_t \text{ using } y^\sigma \text{ and } \hat{y}_{t-1}^\sigma \text{ for all } \sigma \in \Sigma. \\ \lambda \text{ and } \gamma \text{ are used here.} \end{array} \quad (91.51)$$

$$P([\hat{y}_t^\sigma]_\sigma | f_t, \vec{x}) = \prod_{\sigma \in \Sigma} \mathbb{1}(\hat{y}_t^\sigma = f_t(x^\sigma)) \quad (91.52)$$

$$P([A_t^\ell]_{\ell \in \mathcal{L}_t} | \vec{y}, [\hat{y}_{t-1}^\sigma]_\sigma, \ell_t) = \prod_{\ell \in \mathcal{L}_t} \mathbb{1}(A_t^\ell = - \sum_{\sigma \in \Sigma_t^\ell} \underbrace{(y^\sigma - \hat{y}_{t-1}^\sigma)}_{r_t^\sigma}) \quad (91.53)$$

$$P(B_t^\ell | \ell \in \mathcal{L}_t | [\hat{y}_{t-1}^\sigma]_\sigma, \ell_t) = \prod_{\ell \in \mathcal{L}_t} \mathbb{1}(B_t^\ell = \sum_{\sigma \in \Sigma_t^\ell} \left\{ \begin{array}{ll} 1 & \text{if reg.} \\ \hat{y}_{t-1}^\sigma (1 - \hat{y}_{t-1}^\sigma) & \text{if b.c.} \end{array} \right\}) \quad (91.54)$$

$$P(f_t | [A_t^\ell]_{\ell \in \mathcal{L}_t}, [B_t^\ell]_{\ell \in \mathcal{L}_t}, \ell_t, \vec{x}) = \prod_{\sigma \in \Sigma} \mathbb{1} \left(f_t(x^\sigma) = \left\{ \begin{array}{ll} 1 & \text{if reg.} \\ \text{smoid} & \text{if b.c.} \end{array} \right\} \frac{-A^{\ell_t(\sigma)}}{2(B^{\ell_t(\sigma)} + \lambda)} \right) \quad (91.55)$$

$$P(f | [f_t]_t, \vec{x}) = \prod_{\sigma \in \Sigma} \mathbb{1}(f(x^\sigma) = \sum_t (\eta)^t f_t(x^\sigma)) \quad (91.56)$$

Chapter 92

Zero Information Transmission (Graphoid Axioms)

This chapter assumes that you have read Chapter 21 on d-separation.

The following quantities play a very prominent role in the d-separation Theorem that we enunciated in Chapter 21.

- the mutual information (MI)
(a.k.a. information transmission) $H(\underline{a} : \underline{b})$
- the conditional mutual information (CMI)
(a.k.a. conditional information transmission) $H(\underline{a} : \underline{b} | \underline{c})$

MI can be viewed as the special case of CMI, when the set of variables being conditioned on is empty. Particularly prominent in d-separation discussions are probability distributions for which CMI vanishes. The goal of this chapter is to study such probability distributions.

Recall that CMI is non-negative and symmetric in its first two variables (i.e., $H(\underline{a} : \underline{b} | \underline{c}) = H(\underline{b} : \underline{a} | \underline{c})$). Another very useful property of CMI is its chain rule

Claim 149 (*Chain Rule for CMI*)

$$H(\underline{y} : \underline{x}^n) = \sum_i H(\underline{y} : \underline{x}_i | \underline{x}_{<i}), \quad (92.1)$$

where $\underline{x}^n = (\underline{x}_0, \underline{x}_1, \dots, \underline{x}_{n-1})$ and $\underline{x}_{<i} = (\underline{x}_0, \underline{x}_1, \dots, \underline{x}_{i-1})$.

proof:

$$\frac{P(y|x_{<i+1})}{P(y)} = \frac{P(y|x_i, x_{<i})}{P(y|x_{<i})} \cdot \frac{P(y|x_{<i})}{P(y)} \quad (92.2)$$

Therefore,

$$\ln \frac{P(y|x_{<i+1})}{P(y)} = \ln \frac{P(y|x_i, x_{<i})}{P(y|x_{<i})} + \ln \frac{P(y|x_{<i})}{P(y)} \quad (92.3)$$

If we now apply $\sum_{y,x^n} P(y, x^n)$ to both sides, we get

$$H(\underline{y} : \underline{x}_{<i+1}) = H(\underline{y} : \underline{x}_i | \underline{x}_{<i}) + H(\underline{y} : \underline{x}_{<i}) \quad (92.4)$$

QED

A trivial but very useful consequence of the chain rule for CMI is:

$$H(\underline{y} : \underline{x}^n) = 0 \iff H(\underline{y} : \underline{x}_i | \underline{x}_{<i}) = 0 \text{ for all } i. \quad (92.5)$$

92.1 Consequences of Eq.(92.5)

Table 92.1 gives a set of statements about CMI referred to as the Graphoid Axioms in chapter 1 of Ref.[46]. See Ref.[46] to learn the history of these axioms. The purpose of this section is to prove that the graphoid axioms are all a simple consequence of Eq.(92.5).

Symmetry	$\underline{a} \perp_P \underline{b} \implies \underline{b} \perp_P \underline{a}$ $H(\underline{a} : \underline{b}) = 0 \implies H(\underline{b} : \underline{a}) = 0$
Decomposition	$\underline{a} \perp_P \underline{b}, \underline{c} \implies \underline{a} \perp_P \underline{b} \text{ and } \underline{a} \perp_P \underline{c}$ $H(\underline{a} : \underline{b}, \underline{c}) = 0 \implies H(\underline{a} : \underline{b}) = 0 \text{ and } H(\underline{a} : \underline{c}) = 0$
Weak Union	$\underline{a} \perp_P \underline{b}, \underline{c} \implies \underline{a} \perp_P \underline{b} \underline{c} \text{ and } \underline{a} \perp_P \underline{c} \underline{b}$ $H(\underline{a} : \underline{b}, \underline{c}) = 0 \implies H(\underline{a} : \underline{b} \underline{c}) = 0 \text{ and } H(\underline{a} : \underline{c} \underline{b}) = 0$
Contraction	$\underline{a} \perp_P \underline{b} \underline{c} \text{ and } \underline{a} \perp_P \underline{c} \implies \underline{a} \perp_P \underline{b}, \underline{c}$ $H(\underline{a} : \underline{b} \underline{c}) = 0 \text{ and } H(\underline{a} : \underline{c}) = 0 \implies H(\underline{a} : \underline{b}, \underline{c}) = 0$
Intersection	$\underline{a} \perp_P \underline{b} \underline{c}, \underline{d} \text{ and } \underline{a} \perp_P \underline{d} \underline{c}, \underline{b} \implies \underline{a} \perp_P \underline{b}, \underline{d} \underline{c}$ $H(\underline{a} : \underline{b} \underline{c}, \underline{d}) = 0 \text{ and } H(\underline{a} : \underline{d} \underline{c}, \underline{b}) = 0 \implies H(\underline{a} : \underline{b}, \underline{d} \underline{c}) = 0$

Table 92.1: Graphoid Axioms

Claim 150 *Table 92.1 is true.*

proof:

- **Symmetry**

Follows trivially from $H(\underline{a} : \underline{b}) = H(\underline{b} : \underline{a})$.

- **Decomposition**

From the chain rule for CMI, we have

$$H(\underline{a} : \underline{b}, \underline{c}) = H(\underline{a} : \underline{b}|\underline{c}) + H(\underline{a} : \underline{c}), \quad (92.6)$$

and

$$H(\underline{a} : \underline{b}, \underline{c}) = H(\underline{a} : \underline{c} | \underline{b}) + H(\underline{a} : \underline{b}) . \quad (92.7)$$

Hence,

$$H(\underline{a} : \underline{b}, \underline{c}) = 0 \quad (92.8)$$

implies

$$H(\underline{a} : \underline{b} | \underline{c}) = H(\underline{a} : \underline{c}) = 0 , \quad (92.9)$$

and

$$H(\underline{a} : \underline{c} | \underline{b}) = H(\underline{a} : \underline{b}) = 0 . \quad (92.10)$$

- **Weak Union**

Already proven in proof of Decomposition.

- **Contraction**

From chain rule for CMI, we have

$$H(\underline{a} : \underline{b}, \underline{c}) = H(\underline{a} : \underline{b} | \underline{c}) + H(\underline{a} : \underline{c}) . \quad (92.11)$$

- **Intersection**

From the chain rule for CMI, we have

$$H(\underline{a} : \underline{b}, \underline{d} | \underline{c}) = H(\underline{a} : \underline{b} | \underline{d}, \underline{c}) + H(\underline{a} : \underline{d} | \underline{c}) , \quad (92.12)$$

and

$$H(\underline{a} : \underline{b}, \underline{d} | \underline{c}) = H(\underline{a} : \underline{d} | \underline{b}, \underline{c}) + H(\underline{a} : \underline{b} | \underline{c}) . \quad (92.13)$$

Thus,

$$H(\underline{a} : \underline{b}, \underline{d} | \underline{c}) = 0 \quad (92.14)$$

implies

$$H(\underline{a} : \underline{b} | \underline{d}, \underline{c}) = H(\underline{a} : \underline{d} | \underline{c}) = 0 , \quad (92.15)$$

and

$$H(\underline{a} : \underline{d} | \underline{b}, \underline{c}) = H(\underline{a} : \underline{b} | \underline{c}) = 0 . \quad (92.16)$$

QED

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