# Assignment 3

Rory Sarten 301005654 29 September, 2020

#### Question 1

a)

A Runs Test tests a set of binary variables  $X_1, ..., X_n$  to verify if the variables occur randomly.

 $H_0$ : variables occur randomly, i.e. knowing  $X_1, ..., X_n$  does not help predict  $X_{n+1}$ .

 $H_A$ : variables are not random, i.e. knowing some part of the sequence can help predict subsequent variables.

As the variables are binary, they will take the value 0 or 1. The number of 0s is  $n_0$  and the number of 1s is  $n_1$ , where:

$$n_0 = n - \sum_{i=1}^n X_i$$

$$n_1 = \sum_{i=1}^n X_i$$

To perform a Runs Test the observations are combined into one collection of  $n = n_0 + n_1$  observations and arranged in increasing order of magnitude or observation. They are labeled according to which set they originally came from. A run is a group of two or more sequential values of 0 or 1.

Let R denote the number of runs in the combined ordered sample of  $X \in \{0,1\}$ . Under  $H_0$ , R can be approximated as a normally distributed random variable, assuming both  $n_0$  and  $n_1$  are sufficiently large.

$$R = 1 + \sum_{i=2}^{n} I_{(X_i, X_{i-1})}$$
, where  $I_{(X_i, X_{i-1})} = 0$  if  $X_i = X_{i-1}$  and  $I_{(X_i, X_{i-1})} = 1$  if  $X_i \neq X_{i-1}$ 

$$\bar{R} = \frac{2n_0n_1}{n} + 1$$

$$Var(\bar{R}) = \frac{2n_0n_1(2n_0n_1 - n)}{n^2(n-1)}$$

With test statistic 
$$Z = \frac{R - \bar{R}}{\sqrt{Var(\bar{R})}}$$
 where  $Z \sim N(0,1)$ 

b)

A small number of runs (a small value for R) would indicate that  $X_i$  is more likely to be the same as  $X_{i-1}$ . A large number means that X is fluctuating regularly between values and X is less likely to be the same as  $X_{i-1}$ .

c)

```
R <- 1 + sum(X[2:n] != X[1:n-1])
R_est <- 1 + 2*n_0*n_1/n
R_var <- (2*n_0*n_1*(2*n_0*n_1 - n))/(n^2*(n - 1))
Z <- (R - R_est)/sqrt(R_var)
p <- pnorm(Z)
cbind(R, R_est, R_var, Z, p)</pre>
```

```
## R R_est R_var Z p
## [1,] 15 20.65957 7.974767 -2.004125 0.02252834
```

With p-value of 0.0225 we reject  $H_0$  at the 5% level. We conclude that values are not randomly ordered.

d)

```
set.seed(101)

calc_R <- function(input)
  1 + sum(input[2:length(input)] != input[1:length(input)-1])

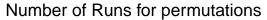
sample_R <- function(iter, input)
  input %>% sample() %>% calc_R()

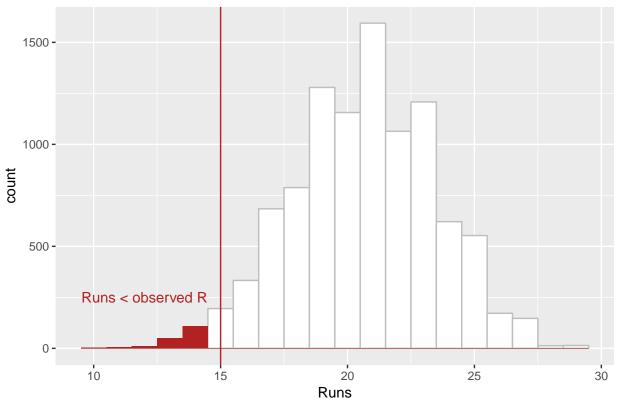
obs_R <- calc_R(X)

N <- 1e4
perm_R <- 1:N %>% sapply(sample_R, input = X)

p <- sum(perm_R < obs_R)/length(perm_R)
p</pre>
```

#### ## [1] 0.0179





Because the number of runs is not a continuous variable, there is some ambiguity around whether the p-value should be calculated comparing values < or <= or even using some half point. I have decided to use < as it gives the test the highest power.

Given this we find a p-value of 0.0179. We reject  $H_0$  at the 5% level. We conclude that values are not randomly ordered.

#### Question 2

a)

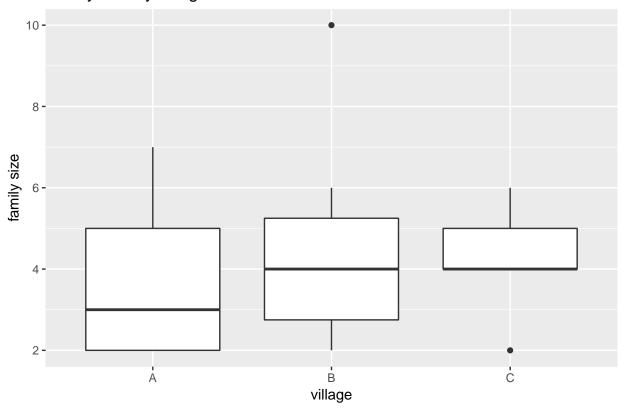
We could use a Poisson distribution

b)

```
dataset <- data.frame(
    size = c(2, 3, 2, 7, 5, 5, 3, 2, 6, 10, 3, 2, 2, 5, 6, 4, 4, 5, 4, 4, 6, 5, 4, 2),
    village = rep(c("A", "B", "C"), c(9, 8, 7))
)

ggplot(dataset, aes(x = village, y = size)) +
    geom_boxplot() +
    ggtitle("Family size by village") +
    ylab("family size")</pre>
```

## Family size by village



```
## village means
tapply(dataset$size, dataset$village, mean) %>% round(digits = 4)
```

```
## A B C
## 3.8889 4.5000 4.2857
```

```
calc_F <- function(input)
   input[["fstatistic"]]["value"]

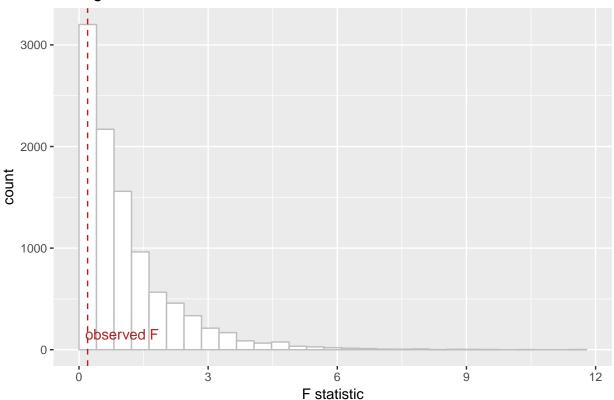
obs_F <- lm(size ~ village, data = dataset) %>%
   summary() %>% calc_F() %>% round(digits = 3)

N <- 1e4
set.seed(150)
perm_F <- 1:N %>%
   lapply(function(i) sample(dataset$village)) %>%
   lapply(function(groups) lm(dataset$size ~ groups)) %>%
   lapply(summary) %>%
   sapply(calc_F) %>%
   round(digits = 3)

p <- sum(perm_F >= obs_F)/N
p
```

## [1] 0.8319

### Histogram of the F Statistic



Given a p-value of 0.8319 we fail to reject  $H_0$  and find that there is not a difference in the mean family sizes.

```
set.seed(150)
perms <- 1:niter %>%
    lapply(function(i) sample(dataset$village)) %>%
    lapply(function(groups) lm(dataset$size ~ groups)) %>%
    lapply(summary) %>%
    sapply(test_func) %>%
    round(digits = 3)

draw_hist(perms, test_stat)
    sum(comparator(perms, test_stat))/niter
}
```

e)

A p-value is the probability of obtaining a test statistic as, or more, extreme than the test statistic observed under the null hypothesis. The p-value is a suitable test statistic to use in a permutation test as the relationship between the F statistic and the p-value means that the result of the ANOVA is preserved. Increasing the F statistic will always result in a decrease in the p-value so long as the degrees of freedom remain the same. Meaning that for all  $F_x > F_y$ , there will be  $p_x < p_y$ .

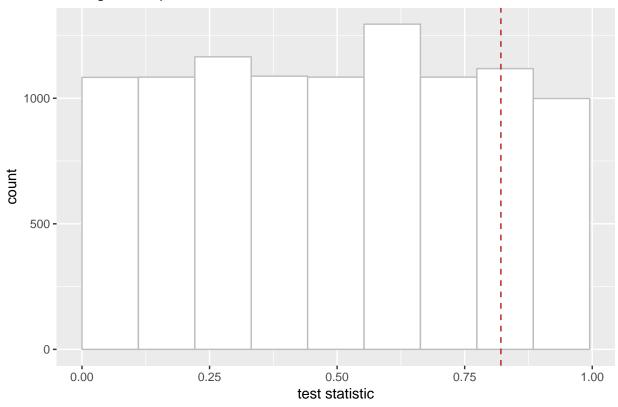
Using the same set.seed value will mean that the permutations of grouping values will be the same across tests even where the test statistic in question changes. Given the case above, the ratio of cases above/below the observed statistic will remain the same, giving the same final p-value.

The generated p-values appear to be drawn from a uniform distribution, and this is what we would expect under a null hypothesis.

```
calc_p <- function(input)
  pf(input$fstatistic, input$df[1]-1, input$df[2], lower.tail = FALSE)["value"]

p <- permutation_test(calc_p, dataset, 1e4, `<=`)</pre>
```

## Histogram of permutation values



p

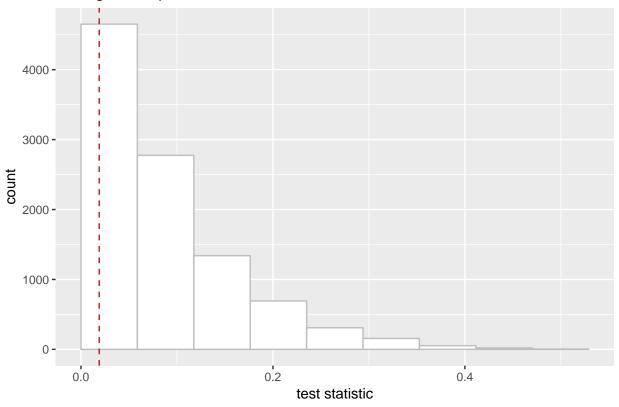
## [1] 0.8319 f)

A large  $R^2$  value means the linear model fits the data well. This would also indicate that the F-statistic would be large as at least one of the coefficients in the model would be non-zero. This would mean that for all  $F_x > F_y$  we have  $R_x^2 > R_y^2$ .

```
calc_r2 <- function(input)
  input[["r.squared"]]

r2 <- permutation_test(calc_r2, dataset, 1e4, `>=`)
```

## Histogram of permutation values



r2

## [1] 0.8319

## Question 3

a)

The chi-squared test statistic:

 $\chi^2 = \sum \frac{(O_i - E_i)}{E_i}$  where  $O_i$  is the observed value in cell i and  $E_i$  is the expected value in cell i.

```
living_standards <- data.frame(
    significant_hardship = c(34, 27),
    fairly_comfortable = c(18, 22),
    comfortable = c(23, 52),
    good = c(9, 20),
    row.names = c("with_dep_children", "no_dep_children"))

chisq.test(living_standards)</pre>
```

```
##
## Pearson's Chi-squared test
##
## data: living_standards
## X-squared = 10.245, df = 3, p-value = 0.0166
```

With p-value of 0.0166 we reject  $H_0$  at the 5% level. We conclude that there is a difference in living standards between families with dependent children and families without dependent children.

b)

The procedure for the permutation chi-square test is as follows:

- 1. Construct a data frame with a column indicating the child dependency status and another indicating the living standard. There will be a single row giving the observed combination for each family.
- 2. Calculate the observed  $\chi^2$  value  $(\chi^2_{obs})$ . Tabulate the observed table O then calculate the expected matrix E. The expected table is calculated as the outer product of the tabulated row and column sums of the observed counts.

$$E = \frac{rowsum(O) * colsum^T(O)}{n}$$

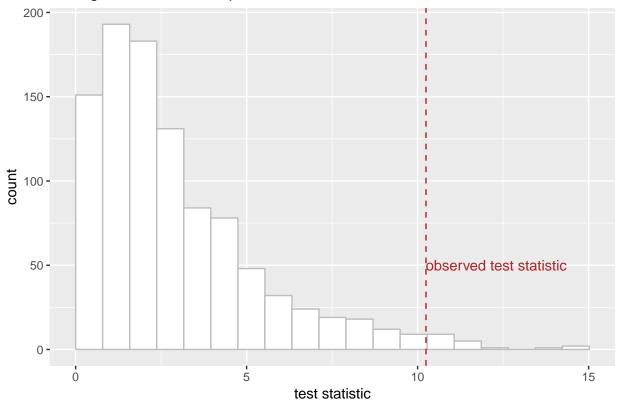
The chi-square test statistic can then be calculated using the formula in a).

- 3. For each permutation randomly reallocate the values for dependent children, while holding the living standard fixed. Under the null hypothesis each value for the dependent children is equally likely for the levels of living standards. The chi-square statistic for each permutation  $\chi^2_{perm}$  is calculated as above.
- 4. p-value is calculated as:

$$\text{p-value} = \frac{\text{No of } \chi^2_{perm} >= \chi^2_{obs}}{\text{No of permutations}}$$

```
families <- data.frame(</pre>
  dependent_children = c(rep("Yes", 84L), rep("No", 121L)),
  living_standard = c(rep("significant_hardship", 34L),
                       rep("fairly_comfortable", 18L),
                       rep("comfortable", 23L),
                       rep("good", 9L),
                       rep("significant_hardship", 27L),
                       rep("fairly_comfortable", 22L),
                       rep("comfortable", 52L),
                       rep("good", 20L)))
## calculates E, then formula for chi-squared
calc X2 <- function(observed) {</pre>
  expected <- outer(rowSums(observed), colSums(observed))/sum(observed)</pre>
  sum((observed - expected)^2/expected)
## calculate observed chi-square test statistic
obs_X2 <- calc_X2(table(families)) %>% round(digits = 4)
N <- 1000
set.seed(101)
perm_X2 <- 1:N %>%
  ## shuffle dependent child values
  lapply(function(i) sample(families$dependent_children)) %>%
  ## generate table against fixed living standard values
  lapply(table, families$living_standard) %>%
  ## calculate chi-square test statistic
  sapply(calc_X2) %>%
```

## Histogram of the chi-square statistic



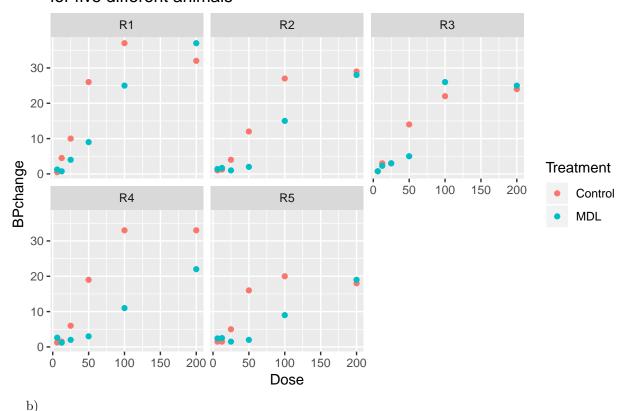
With p-value of 0.019 we reject  $H_0$  at the 5% level. We conclude that there is a difference in living standards between families with dependent children and families without dependent children.

The histogram above clearly demonstrates that  $\chi^2_{obs}$  lies well into the tail of the  $\chi^2_{perm}$  distribution.

#### Question 4

a)
rabbits <- readr::read\_csv("bp\_rabbits.csv", col\_types = readr::cols())
ggplot(data = rabbits, aes(x = Dose, y = BPchange, colour = Treatment)) +
 geom\_point() +
 facet\_wrap(~ Animal) +
 ggtitle("Change in blood pressure (BPchange) by Dose and Treatment\nfor five different animals")</pre>

# Change in blood pressure (BPchange) by Dose and Treatment for five different animals



model <- lm(BPchange ~ Dose + Treatment + Animal, data = rabbits)
summary(model)</pre>

```
##
  lm(formula = BPchange ~ Dose + Treatment + Animal, data = rabbits)
##
## Residuals:
       Min
                10
                    Median
                                 3Q
                                        Max
## -11.342 -3.702
                   -1.263
                                    14.554
                              2.027
##
##
  Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
                 8.74087
                            1.93350
                                       4.521
                                              3.5e-05
## (Intercept)
                 0.13992
                                     12.902
## Dose
                            0.01084
                                              < 2e-16 ***
## TreatmentMDL -4.68000
                            1.46786
                                     -3.188
                                              0.00240 **
## AnimalR2
                -5.30417
                            2.32089
                                      -2.285
                                              0.02632 *
## AnimalR3
                -4.85000
                            2.32089
                                      -2.090
                                              0.04146 *
## AnimalR4
                -4.28750
                            2.32089
                                      -1.847
                                              0.07028
## AnimalR5
                -7.38333
                            2.32089
                                      -3.181
                                              0.00245 **
##
                  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Signif. codes:
##
## Residual standard error: 5.685 on 53 degrees of freedom
## Multiple R-squared: 0.7796, Adjusted R-squared: 0.7547
## F-statistic: 31.25 on 6 and 53 DF, p-value: 9.529e-16
```

#### anova(model)

```
## Analysis of Variance Table
##
## Response: BPchange
##
            Df Sum Sq Mean Sq F value
                                          Pr(>F)
## Dose
             1 5380.3 5380.3 166.4726 < 2.2e-16 ***
## Treatment
             1
                328.5
                        328.5
                               10.1653 0.002402 **
## Animal
             4
               351.4
                         87.9
                                2.7185 0.039217 *
## Residuals 53 1712.9
                         32.3
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

Animal is included as a predictor to account for any of the effects or differences between individual rabbits. There are multiple observations of a single rabbit, each of which may have different reactions and baseline blood pressures when subjected to the drug, this needs to be accounted for.

c)

The procedure for the permutation test is as follows:

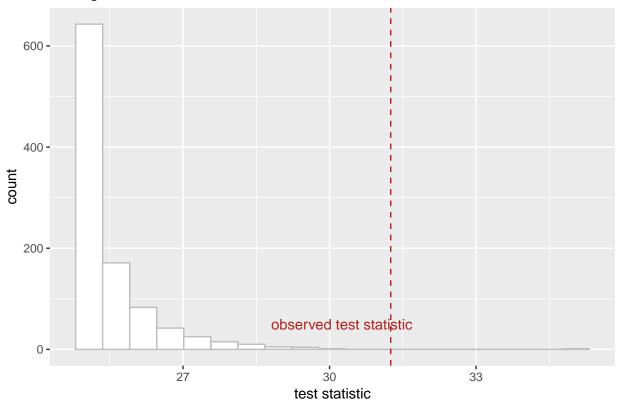
- 1. Calcualte the observed F statistic  $F_{obs}$
- 2. For each permutation randomly allocate the 60 values for Treatment while hold Dose and Animal in place.
- 3. Calculate the permuted F statistic  $F_{perm}$
- 4. The p-value is calculate as:

$$p
-value = \frac{\text{No of } F_{perm} > F_{obs}}{\text{No of permutations}}$$

```
calc_F <- function(treatment, dataset) {
  lm(BPchange ~ Dose + treatment + Animal, data = dataset) %>%
    summary() %>% .[["fstatistic"]] %>% .["value"] %>% unname()
}
obs_F <- calc_F(rabbits$Treatment, rabbits)
obs_F</pre>
```

#### ## [1] 31.252

## Histogram of the F statistic



With p-value of 0.001 we reject  $H_0$  at the 5% level. We conclude that Treatment has an effect on the change in blood pressure after allowing for Dose and Animal.