# Assignment 4

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### Question 1

e)

```
a)
food_prices <- readr::read_delim("food_prices_kg2019.csv", delim = ",", col_types = readr::cols())</pre>
theta_est <- IQR(food_prices$Data_value) %>% round(3)
theta_est
## [1] 6.675
  b)
set.seed(1)
N <- 1e4
boot IQR <- 1:N %>%
  lapply(function(i) sample(food_prices$Data_value, replace = TRUE)) %>%
  sapply(IQR) %>%
  round(3)
## standard error of estimator
sd(boot_IQR) %>% round(3)
## [1] 1.197
## standard 95% bootstrap confidence interval
(theta_est + 1.96*c(-1, 1)*sd(boot_IQR)) %>% round(3)
## [1] 4.328 9.022
  c)
## Efron's interval
quantile(boot_IQR, probs = c(0.025, 0.975)) %>% round(3)
     2.5% 97.5%
   5.290 10.105
  d)
## Hall's interval
hall <- (2 * theta_est - quantile(boot_IQR, probs = <math>c(0.975, 0.025))) \%% round(3)
names(hall) <- c("2.5%", "97.5%")
hall
## 2.5% 97.5%
## 3.245 8.060
```

```
## bias
bias <- (mean(boot_IQR) - theta_est) %>% round(3)
bias
```

#### ## [1] 0.522

```
## size of bias in relation to the std error
bias_size <- (bias/sd(boot_IQR)) %>% round(3)
bias_size
```

### ## [1] 0.436

The bias is approximately 44% of the s.e.( $\hat{\theta}$ ). The size of this bias is considerable.

f)

```
## bias corrected Efron interval
(quantile(boot_IQR, probs = c(0.025, 0.975)) - bias) %>% round(3)
```

```
## 2.5% 97.5%
## 4.768 9.583
```

The lower bound of the confidence interval is above \$4. We reject the hypothesis that the test IQR could be below 4NZD at the 5% confidenc interval.

### Question 2

a)

- 1. Calculate the observed  $\hat{\beta}$  and  $\hat{\sigma}^2$  from the observed data
- 2. Draw a sample of the observations with replacement and calculate a new estimate  $\hat{\beta}_h^*$  from the sample
- 3. Repeat step 2 N times
- 4. Calculate  $s.e.(\hat{\beta}^*)$  as the standard error over the results of the bootstrapped samples
- 5. Calculate  $\hat{\beta} \pm 1.96 \times s.e.(\hat{\beta}^*)$  (for 95% confidence interval)

b)

```
set.seed(1)
bootstrap_velocity <- 1:N %>%
  lapply(function(i) {
    beta_est*distance + rnorm(n, sd = sqrt(sigma_est * distance))})
bootstrap_beta <- bootstrap_velocity %>%
  sapply(calc_beta, distance)
estimate <- mean(bootstrap beta)</pre>
ci \leftarrow beta_est + 1.96 * c(-1, 1) * sd(bootstrap_beta)
results <- c(estimate, ci) %>% round(3)
names(results) <- c("Estimate", "Lower", "Upper")</pre>
results ## boostrap results for beta
## Estimate
                Lower
                         Upper
     76.036
               59,630
                        92.351
bootstrap_sigma2 <- bootstrap_velocity %>%
  sapply(calc_sigma, distance)
bootstrap_sigma2_est <- mean(bootstrap_sigma2)</pre>
ci <- sigma_est + 1.96 * c(-1, 1) * sd(bootstrap_sigma2)
sigma2_results <- c(bootstrap_sigma2_est, ci)</pre>
names(sigma2_results) <- c("Estimate", "Lower", "Upper")</pre>
sigma2_results ## bootstrap results for sigma squared
## Estimate
                  Lower
                             Upper
## 8834.016 2164.639 16987.697
## Reserve results for part d)
bootstrap_results <- rbind(results, sigma2_results)</pre>
rownames(bootstrap_results) <- c("Beta", "Sigma2")</pre>
Using the boot package
boot_beta <- function(dataset, beta_est, sigma_est) {</pre>
  v <- beta_est*dataset$d + rnorm(n, sd = sqrt(sigma_est * dataset$d))
  calc beta(v, dataset$d)
}
```

```
names(boot_results) <- c("Estimate", "Lower", "Upper")</pre>
boot_results
## Estimate
                                 Lower
                                                    Upper
           76.035
                               59.630
                                                  92.350
     c)
y_i \sim N(\beta x_i, \sigma^2 x_i)
L(\beta, \sigma | \boldsymbol{y}) = \prod_{i=1}^{n} \left[ (2\pi\sigma^{2}x_{i})^{1/2} \exp(-\frac{(y_{i} - \beta x_{i})^{2}}{2\sigma^{2}x_{i}}) \right]
               \ell = -\frac{1}{n} \sum_{i=1}^{n} \log(2\pi\sigma^2) - \frac{1}{n} \sum_{i=1}^{n} \log(x_i) - \frac{1}{2\sigma^2} \sum_{i=1}^{n} x_i^{-1} (y_i - \beta x_i)^2
Solving for \hat{\beta} we only need the last term.
          \frac{\partial \ell}{\partial \beta} = \frac{1}{\sigma^2} \sum_{i=1}^{n} (y_i - \beta x_i) \quad \text{setting equal to } 0
\frac{1}{\sigma^2} \sum_{i=1}^{n} y_i = \beta \frac{1}{\sigma^2} \sum_{i=1}^{n} x_i
            \beta = \frac{\sum_{i=1}^{n} y_i}{\sum_{i=1}^{n} x_i}
Solving for \hat{\sigma}^2
\frac{\partial \ell}{\partial \sigma^2} = -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^4} \sum_{i=1}^n x_i^{-1} (y_i - \beta x_i)^2 \quad \text{setting equal to } 0
\frac{n}{2\sigma^2} = \frac{1}{2\sigma^4} \sum_{i=1}^{n} x_i^{-1} (y_i - \beta x_i)^2
\frac{2\sigma^4}{2\sigma^2} = \frac{1}{n} \sum_{i=1}^{n} x_i^{-1} (y_i - \beta x_i)^2
   \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^{n} \frac{1}{x_i} (y_i - \beta x_i)^2
loglikelihood <- function(params, x, y) {</pre>
     n <- length(x)
    beta_ <- params[1]</pre>
    sigma2_ <- params[2]
     -(n/2)*\log(2*pi*sigma2_) - 0.5*sum(\log(x)) - (1/(2*sigma2_)) * sum((y - beta_*x)^2 / x)
result <- optim(par = c(40, 250),
                                   fn = loglikelihood,
                                  x = distance, y = velocity,
lower = c(0, 250), upper = c(Inf, Inf),
method = "L-BFGS-B",
```

control = list(fnscale = -1),

```
hessian = TRUE)

opt_mle_se <- solve(-result$hessian) %>% diag() %>% sqrt() %>% round(3)
ci <- result$par[1] + c(-1, 1) * 1.96 * opt_mle_se[1]
likelihood_beta_results <- c(result$par[1], ci) %>% round(3)
names(likelihood_beta_results) <- c("Estimate", "Lower", "Upper")
#likelihood_beta_results

ci <- result$par[2] + c(-1, 1) * 1.96 * opt_mle_se[2]
likelihood_sigma_results <- c(result$par[2], ci) %>% round(3)
names(likelihood_sigma_results) <- c("Estimate", "Lower", "Upper")
#likelihood_sigma_results

output <- rbind(likelihood_beta_results, likelihood_sigma_results)
rownames(output) <- c("Beta", "Sigma2")
#output</pre>
```

knitr::kable(bootstrap\_results, caption = "Bootstrap Estimates")

Table 1: Bootstrap Estimates

	Estimate	Lower	Upper
Beta	76.036	59.630	92.351
Sigma2	8834.016	2164.639	16987.697

knitr::kable(output, caption = "optim Estimates")

Table 2: optim Estimates

	Estimate	Lower	Upper
Beta	75.990	59.548	92.433
Sigma2	9576.071	2223.131	16929.011

The estimates for  $\hat{\beta}$  are very similar. The 95% confidence interval widths are 32.721 and 32.885

The estimates for  $\hat{\sigma}^2$  are more divergent. The 95% confidence interval widths are 14823.1 and 14705.9. This is not surprising as the log-likelihood function is not symmetric around the test statistic.

### Question 3

a)