# Linear Methods for Classification

## Robert Schmidt

#### Abstract

The following are notes on the key results from the **Elements of Statistical Learning** text. They were primarily derived from course notes and readings in the Stanford STATS 315: *Modern Applied Statistics* series.

## Contents

1	Discriminant analysis rule derivations	2
	1.1 LDA rule derivation	2
	1.2 QDA rule derivation	
<b>2</b>	Discriminant analysis computations	4
	2.1 LDA computation: sphering the data	4
	2.2 Computations for reduced-rank LDA	
	2.3 Rayleigh quotient and canonical discriminant analysis	
3	LDA vs. least squares fit	į
	3.1 Least squares regression coefficient is identical to LDA coefficient, up to scalar multiple	ŀ
	3.2 Difference between LDA and OLS coefficient	
4	Logistic regression	6
	4.1 Derivation of logistic rule	6
	4.2 Two-class algorithm for logistic regression	
	4.3 Newton-Raphson IRLS algorithm	

### 1 Discriminant analysis rule derivations

We need an expression for  $\mathbb{P}(G \mid \mathbf{X})$  in order to perform classification. Adopting the notation in ESL, let:

$$\begin{cases} f_k(x) \to \text{class-conditional density of } \mathbf{X} \text{ in class } G = k \\ \pi_k \to \text{ prior probability of class } k \\ \sum_{k=1}^K \pi_k = 1 \end{cases}$$

Via Bayes rule, the desired probability is:

$$\mathbb{P}(G = k \mid \mathbf{X} = x) = \frac{f_k(x)\pi_k}{\sum_{\ell=1}^K f_\ell(x)\pi_\ell}$$

In the case that each class density is multivariate Gaussian, the densities are of the form:

$$f_k(x) = \frac{1}{(2\pi)^{p/2} |\mathbf{\Sigma}_k|^{1/2}} \exp\left(-\frac{1}{2}(x - \mu_k)^T \mathbf{\Sigma}_k^{-1} (x - \mu_k)\right)$$

Here, we consider the two-class case (K = 2). In this case, we examine the log ratio of the conditional probabilities to find the classification boundary:

$$L_{k\ell} = \log\left(\frac{\mathbb{P}(G=k \mid \mathbf{X} = x)}{\mathbb{P}(G=\ell \mid \mathbf{X} = x)}\right) = \log\frac{f_k(x)}{f_\ell(x)} + \log\frac{\pi_k}{\pi_\ell}$$

The degree to which this formula can be simplified depends upon the assumptions on the covariance matrix  $\Sigma_k$ . To this end, we will consider LDA and QDA.

#### 1.1 LDA rule derivation

For LDA, we assume the  $f_k$  are multivariate normal with common covariance  $\Sigma$  and separate  $\mu_k$ . In other words:

$$f_k = c \cdot \exp\left(-\frac{1}{2}(x - \mu_k)^T \mathbf{\Sigma}^{-1}(x - \mu_k)\right)$$
$$f_\ell = c \cdot \exp\left(-\frac{1}{2}(x - \mu_\ell)^T \mathbf{\Sigma}^{-1}(x - \mu_\ell)\right)$$

Here, the leading coefficient terms are the same c since the two classes share a common covariance matrix. This greatly simplifies the following calculations. First, let us examine the ratio of the class densities:

$$\frac{f_k}{f_\ell} = \exp\left(-\frac{1}{2}(x - \mu_k)^T \mathbf{\Sigma}^{-1}(x - \mu_k) + \frac{1}{2}(x - \mu_\ell)^T \mathbf{\Sigma}^{-1}(x - \mu_\ell)\right) 
\log \frac{f_k}{f_\ell} = -\frac{1}{2} \left[ (x - \mu_k)^T \mathbf{\Sigma}^{-1}(x - \mu_k) - \frac{1}{2}(x - \mu_\ell)^T \mathbf{\Sigma}^{-1}(x - \mu_\ell) \right] 
= -\frac{1}{2} \left[ x^T \mathbf{\Sigma}^{-1} x - 2\mu_k^T \mathbf{\Sigma}^{-1} x + \mu_k^T \mathbf{\Sigma}^{-1} \mu_k - x^T \mathbf{\Sigma}^{-1} x + 2\mu_\ell^T \mathbf{\Sigma}^{-1} x - \mu_\ell^T \mathbf{\Sigma}^{-1} \mu_\ell \right] 
= -\frac{1}{2} \left[ -2\mu_k^T \mathbf{\Sigma}^{-1} x + \mu_k^T \mathbf{\Sigma}^{-1} \mu_k + 2\mu_\ell^T \mathbf{\Sigma}^{-1} x - \mu_\ell^T \mathbf{\Sigma}^{-1} \mu_\ell \right] 
= x^T \mathbf{\Sigma}^{-1} (\mu_k - \mu_\ell) - \frac{1}{2} \left[ \mu_k^T \mathbf{\Sigma}^{-1} \mu_k - \mu_\ell^T \mathbf{\Sigma}^{-1} \mu_\ell \right]$$

For a moment, let us consider in more detail the  $-\frac{1}{2} \left[ \mu_k^T \mathbf{\Sigma}^{-1} \mu_k - \mu_\ell^T \mathbf{\Sigma}^{-1} \mu_\ell \right]$  term.

This should simplify to the  $-\frac{1}{2}(\mu_k + \mu_\ell)^T \mathbf{\Sigma}^{-1}(\mu_k - \mu_\ell)$  term found in ESL (4.9). Let us prove that the two are equivalent:

$$-\frac{1}{2}(\mu_k + \mu_\ell)^T \mathbf{\Sigma}^{-1}(\mu_k - \mu_\ell) = -\frac{1}{2} \left[ \mu_k^T \mathbf{\Sigma}^{-1} \mu_k - \mu_\ell^T \mathbf{\Sigma}^{-1} \mu_\ell + \mu_\ell^T \mathbf{\Sigma}^{-1} \mu_k - \mu_k^T \mathbf{\Sigma}^{-1} \mu_\ell \right]$$
$$= -\frac{1}{2} \left[ \mu_k^T \mathbf{\Sigma}^{-1} \mu_k - \mu_\ell^T \mathbf{\Sigma}^{-1} \mu_\ell \right]$$

Hence, we arrive at the final LDA expression from ESL equation (4.9):

$$L_{k\ell} = \log \frac{\pi_k}{\pi_\ell} - \frac{1}{2} (\mu_k + \mu_\ell)^T \mathbf{\Sigma}^{-1} (\mu_k - \mu_\ell) + x^T \mathbf{\Sigma}^{-1} (\mu_k - \mu_\ell)$$

### 1.2 QDA rule derivation

Another way to look at the LDA expression is in terms of the discriminant functions  $\delta_k$ , where  $L_{k\ell} = \delta_k - \delta_\ell$ . Here, each discriminant function has the form:

$$\delta_k = x^T \mathbf{\Sigma}^{-1} \mu_k - \frac{1}{2} \mu_k^T \mathbf{\Sigma}^{-1} \mu_k + \log \pi_k$$

In terms of classification, we classify to class k if  $\delta_k > \delta_\ell$ .

This decision boundary will be linear (hence "linear" discriminant analysis).

To derive the QDA result, we first start from the same log-ratio as LDA:

$$L_{k\ell} = \log\left(\frac{\mathbb{P}(G = k \mid \mathbf{X} = x)}{\mathbb{P}(G = \ell \mid \mathbf{X} = x)}\right) = \log\frac{f_k(x)}{f_{\ell}(x)} + \log\frac{\pi_k}{\pi_{\ell}}$$

The key difference between QDA and LDA is that, unlike LDA, the  $\Sigma_k$  are not assumed to be equal. First, consider the QDA densities:

$$f_k(x) = \frac{1}{(2\pi)^{p/2} |\mathbf{\Sigma}_k|^{1/2}} \exp\left(-\frac{1}{2} (x - \mu_k)^T \mathbf{\Sigma}_k^{-1} (x - \mu_k)\right)$$
$$f_{\ell}(x) = \frac{1}{(2\pi)^{p/2} |\mathbf{\Sigma}_{\ell}|^{1/2}} \exp\left(-\frac{1}{2} (x - \mu_{\ell})^T \mathbf{\Sigma}_{\ell}^{-1} (x - \mu_{\ell})\right)$$

As was the case in the LDA derivation, we need an expression for the log-ratio of the densities.

$$\frac{f_k}{f_\ell} = \frac{|\Sigma_k|^{1/2}}{|\Sigma_\ell|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu_k)^T \Sigma_k^{-1}(x-\mu_k) + \frac{1}{2}(x-\mu_\ell)^T \Sigma_\ell^{-1}(x-\mu_\ell)\right) 
\frac{f_k}{f_\ell} = \frac{|\Sigma_k|^{1/2}}{|\Sigma_\ell|^{1/2}} \exp\left(-\frac{1}{2}\left[(x-\mu_k)^T \Sigma_k^{-1}(x-\mu_k) - \frac{1}{2}(x-\mu_\ell)^T \Sigma_\ell^{-1}(x-\mu_\ell)\right]\right)$$

Here, the terms cannot be factored nicely. However, we can still develop a final expression for the discriminant functions in order to match equation (4.12) in ESL.

$$\delta_k = -\frac{1}{2}\log|\mathbf{\Sigma}_k| - \frac{1}{2}(x - \mu_k)^T \mathbf{\Sigma}_k^{-1}(x - \mu_k) + \log \pi_k$$

This expression is quadratic in x, which yields a quadratic decision boundary.

- 2 Discriminant analysis computations
- 2.1 LDA computation: sphering the data
- 2.2 Computations for reduced-rank LDA
- 2.3 Rayleigh quotient and canonical discriminant analysis

- 3 LDA vs. least squares fit
- 3.1 Least squares regression coefficient is identical to LDA coefficient, up to scalar multiple
- 3.2 Difference between LDA and OLS coefficient

- 4 Logistic regression
- 4.1 Derivation of logistic rule
- 4.2 Two-class algorithm for logistic regression
- ${\bf 4.3}\quad {\bf Newton\text{-}Raphson~IRLS~algorithm}$