

Statistics ST2334 Topic 4,5 addendum: Joint Distributions

2011/2012 Semester 2

Joint Distributions: Some Theory

Recall

- ▶ that a random variable X is a function that maps outcomes $\omega \in \Omega$ to $X(\omega) \in \mathbb{R}$.
- ▶ that a probability function P , maps events $A \subseteq \Omega$ to a number, and obeys the axioms of probability.
- ▶ that we can then apply P to X by specifying $P(X \in M) := P(\{\omega : X(\omega) \in M\})$ for $M \subseteq \mathbb{R}$.
- ▶ we then call this entire description of the values and the corresponding probability of X its probability distribution.

We can similarly define probability distributions for multiple random variables.

Joint Distributions: Some Theory

- ▶ Let X, Y be random variables, then

$$P(X \in M, Y \in N) := P(\{\omega : X(\omega) \in M\} \cap \{v : Y(v) \in N\})$$

where $\omega, v \in \Omega$.

- ▶ Note that there is nothing new going on on the RHS. It is applying the probability function P on a subset of Ω .
- ▶ This description of possible values X, Y can take and their corresponding probabilities is called the *joint* probability distribution of X and Y .
- ▶ We can of course extend this definition to any n number of random variables.

Joint Distributions: Discrete Case

- ▶ For discrete random variables, we extend our probability mass function to a function of multiple variables.
- ▶ E.g. For two discrete variables X_1 and X_2 , their joint probability mass function is defined as

$$f(x_1, x_2) = P(X_1 = x_1, X_2 = x_2)$$

where x_1, x_2 are possible values of X_1, X_2 respectively.

- ▶ Similarly, For n discrete random variables X_1, \dots, X_n , their joint probability mass function is simply

$$f(x_1, \dots, x_n) = P(X_1 = x_1, \dots, X_n = x_n)$$

Marginal Distributions: Discrete Case

- ▶ When given the joint distribution of say X_1, X_2 , we call the distribution of X_1 (or X_2) alone the *marginal* distribution.
- ▶ Note that the marginal distributions are embedded in the joint distribution. We can recover the probability mass function f_1 of X_1 from the joint probability mass function f by

$$f_1(x_1) = P(X_1 = x_1) = \sum_{\text{all } x_2} P(X_1 = x_1, X_2 = x_2) = \sum_{\text{all } x_2} f(x_1, x_2)$$

- ▶ Similarly, if f is the joint probability mass function of X_1, \dots, X_n , then the marginal probability mass functions are

$$f_i(x_i) = \sum_{x_1} \cdots \sum_{x_{i-1}} \sum_{x_{i+1}} \cdots \sum_{x_n} f(x_1, \dots, x_n)$$

Conditional Probability Distributions: Discrete Case

- ▶ Consistent with the definition of conditional probability of *events*, we can define the conditional probability distribution of X_1 given $X_2 = x_2$ as

$$\begin{aligned} f_1(x_1|x_2) &= P(X_1 = x_1|X_2 = x_2) = \frac{P(X_1 = x_1, X_2 = x_2)}{P(X_2 = x_2)} \\ &= \frac{f(x_1, x_2)}{f_2(x_2)}, \text{ provided } f_2(x_2) \neq 0 \end{aligned}$$

- ▶ If $f_1(x_1|x_2) = f_1(x_1)$ for all x_1 and x_2 , then we say X_1 and X_2 are independent. This is equivalent to

$$f(x_1, x_2) = f_1(x_1)f_2(x_2)$$

for all x_1 and x_2 .

Example: Discrete Case

Let X_1 and X_2 have the joint probability distribution in the table below.

		$f(x_1, x_2)$		
		x_1		
		-1	0	1
x_2	-1	0.1	0.3	0.1
	1	0.15	0.25	0.1

- Find $P(X_1 + X_2 > 1)$.

$$P(X_1 + X_2 > 1) = P(X_1 = 1, X_2 = 1) = 0.1$$

- Find $P(X_1 X_2 = 0)$.

$$P(X_1 X_2 = 0) = P(X_1 = 0) = 0.3 + 0.25 = 0.55$$

Example: Discrete Case (continued)

- Find the marginal distributions of X_1 and X_2 .

	$f_1(x_1)$				$f_2(x_2)$	
x_1	-1	0	1	x_2	-1	1
$f_1(x_1)$	0.25	0.55	0.1	$f_2(x_2)$	0.5	0.5

- Find the conditional distribution of X_2 given $X_1 = -1$

$$P(X_2 = x_2 | X_1 = -1) = f_2(x_2 | -1) = \frac{f(-1, x_2)}{f_1(-1)}$$

	$f_2(x_2 -1)$	
x_2	-1	1
$f_2(x_2 -1)$	$0.1/0.25 = 0.4$	$0.15/0.25 = 0.6$

- Are X_1 and X_2 independent?

Joint Distributions: Continuous Case

- ▶ As we did before for single random variables, we introduce probability density functions to handle continuous random variables.
- ▶ Suppose X_1, X_2, \dots, X_n are n continuous random variables. Then the joint probability density function of these variables $f(x_1, x_2, \dots, x_n)$ is such that

$$\begin{aligned} P(a_1 \leq X_1 \leq b_1, a_2 \leq X_2 \leq b_2, \dots, a_n \leq X_n \leq b_n) \\ = \int_{a_1}^{b_1} \int_{a_2}^{b_2} \dots \int_{a_n}^{b_n} f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n \end{aligned}$$

- ▶ For the axioms of probability to hold, we need
 - ▶ $f(x_1, x_2, \dots, x_n) \geq 0$
 - ▶ $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n = 1$

Joint Distributions: Continuous Case

We define

- ▶ the joint cumulative distribution function

$$F(x_1, x_2, \dots, x_n) := \int_{-\infty}^{x_1} \int_{-\infty}^{x_2} \dots \int_{-\infty}^{x_n} f(u_1, u_2, \dots, u_n) du_1 du_2 \dots du_n$$

- ▶ the marginal probability density function for X_i

$$f_i(x_i) = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f(u_1, u_2, \dots, u_n) du_1 \dots du_{i-1} du_{i+1} \dots du_n$$

- ▶ the marginal cdf for X_i

$$F_i(x_i) = \int_{-\infty}^{x_i} f_i(u_i) du_i$$

Conditional Distribution: Continuous Case

- ▶ Just as we did for the probability mass functions, we define the conditional probability density function

$$f_1(x_1|x_2) = \frac{f(x_1, x_2)}{f_2(x_2)}$$

- ▶ And continuous random variables X_1, X_2 are independent if

$$f(x_1, x_2) = f_1(x_1)f_2(x_2)$$

for all values of x_1, x_2 .

- ▶ This is equivalent to

$$F(x_1, x_2) = F_1(x_1)F_2(x_2)$$

for all values of x_1, x_2 .

Example: Continuous Case

If two variables X_1 and X_2 have a joint density function

$$f(x_1, x_2) = \begin{cases} \frac{1}{18} e^{-x_1} x_2^2, & \text{for } x_1 > 0, -3 < x_2 < 3 \\ 0, & \text{otherwise} \end{cases}$$

find the probabilities that

- (a) the first random variable will take on a value between 0 and 1 and the second random variable a value between 1 and 3;

$$\begin{aligned} \int_0^1 \int_1^3 f(x_1, x_2) dx_1 dx_2 &= \int_1^3 \left\{ \int_0^1 \frac{1}{18} e^{-x_1} x_2^2 dx_1 \right\} dx_2 \\ &= \frac{1}{18} \int_1^3 [-e^{-x_1}]_0^1 x_2^2 dx_2 \\ &= \frac{1}{18} \cdot (1 - e^{-1}) \left[\frac{1}{3} x_2^3 \right]_1^3 \\ &= \dots = 0.304 \end{aligned}$$

Example: Continuous Case continued

(b) find the joint cumulative distribution function

$$F(x_1, x_2) = \int_{-\infty}^{x_1} \int_{-\infty}^{x_2} f(x_1, x_2) dx_1 dx_2$$

► we need to be careful about the limits: split into 3 regions

$$\begin{aligned} &= \begin{cases} 0, & \text{if } x_1 \leq 0, x_2 \leq -3 \\ \int_0^{x_1} \int_{-3}^{x_2} \frac{1}{18} e^{-x_1} x_2^2 du_1 du_2, & \text{if } x_1 > 0, -3 < x_2 < 3 \\ \int_0^{x_1} \int_{-3}^3 \frac{1}{18} e^{-x_1} x_2^2 du_1 du_2, & \text{if } x_1 > 0, x_2 \geq 3 \end{cases} \\ &= \begin{cases} 0, & \text{if } x_1 \leq 0, x_2 \leq -3 \\ \frac{1}{54}(1 - e^{-x_1})(x_2^3 + 27), & \text{if } x_1 > 0, -3 < x_2 < 3 \\ 1 - e^{-x_1}, & \text{if } x_1 > 0, x_2 \geq 3 \end{cases} \end{aligned}$$

Example: Continuous Case continued continued

(c) are X_1 and X_2 independent?

$$\begin{aligned}f_1(x_1) &= \int_{-\infty}^{\infty} f(x_1, u_2) du_2 \\&= \begin{cases} \int_{-3}^3 \frac{1}{18} e^{-x_1} u_2^2 du_2 = e^{-x_1}, & \text{if } x_1 > 0 \\ 0, & \text{otherwise} \end{cases}\end{aligned}$$

Similarly,

$$f_2(x_2) = \int_{-\infty}^{\infty} f(u_1, x_2) du_1 = \begin{cases} \frac{1}{18} x_2^2, & \text{if } -3 < x_2 < 3 \\ 0, & \text{otherwise} \end{cases}$$

Thus

$$f_1(x_1)f_2(x_2) = f(x_1, x_2)$$

They are independent.

Expectation of multiple random variables

- In the discrete case,

$$E[g(X_1, X_2, \dots, X_n)] = \sum_{x_1} \sum_{x_2} \dots \sum_{x_n} g(x_1, x_2, \dots, x_n) f(x_1, x_2, \dots, x_n)$$

- In the continuous case,

$$E[g(X_1, X_2, \dots, X_n)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} g(x_1, x_2, \dots, x_n) \times f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n$$

Expectation of multiple random variables continued

- ▶ For the previous example, taking $g(x_1, x_2) = x_1 + x_2$, we have

$$E(X_1 + X_2) = \int_0^{\infty} \int_{-3}^3 (x_1 + x_2) \left(\frac{1}{18} e^{-x_1} x_2^2 \right) dx_1 dx_2$$

- ▶ You can work it out directly, but recall for linear functions we can use the following theorem:

For **any** random variables X_1, \dots, X_n and constants a_0, a_1, \dots, a_n , we have

$$E(a_0 + a_1 X_1 + \dots + a_n X_n) = a_0 + a_1 E(X_1) + \dots + a_n E(X_n)$$

- ▶ Define the covariance of X_i and X_j

$$\text{Cov}(X_i, X_j) = E\{(X_i - EX_i)(X_j - EX_j)\}$$

- ▶ This can also be written as

$$\text{Cov}(X_i, X_j) = E(X_i X_j) - E(X_i)E(X_j)$$

- ▶ We have already seen that for discrete random variables: if X_i, X_j are independent, the above expression = 0 (Topic 4c slide 4).

Covariance continued

- ▶ Similarly, for continuous X_i, X_j , if they are independent, then

$$\begin{aligned} E(X_i X_j) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x_i x_j f(x_i, x_j) dx_i dx_j \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x_i x_j f_i(x_i) f_j(x_j) dx_i dx_j \\ &= \int_{-\infty}^{\infty} x_i f_i(x_i) dx_i \int_{-\infty}^{\infty} x_j f_j(x_j) dx_j \\ &= E(X_i) E(X_j) \end{aligned}$$

Hence, X_i, X_j independent $\Rightarrow \text{Cov}(X_i, X_j) = 0$

- ▶ Note that the converse is NOT true. i.e. two random variables with covariance 0 are not necessarily independent.

- ▶ For **any** random variables X_1, \dots, X_n , it's not hard to see

$$\text{Var}(X_1 + \dots + X_n) = \text{Var}(X_1) + \dots + \text{Var}(X_n) + 2 \sum_{j>i} \text{Cov}(X_i, X_j)$$

- ▶ You can see now that our earlier identity follows: If X_1, \dots, X_n are **independent**, then

$$\text{Var}(X_1 + \dots + X_n) = \text{Var}(X_1) + \dots + \text{Var}(X_n)$$