

(b) dimension on the outside...

$$h[X] = n \ln(\det(\sqrt{2\pi}e(\Sigma^{1/2})^{1/n})) = n \ln(\sqrt{2\pi}e \det(\Sigma^{1/2})^{1/n}) \\ \approx n \ln(4 \det(\Sigma^{1/2})^{1/n})$$

(since $\ln(x) = \frac{n}{n} \ln(x) = n \ln(x^{1/n})$)

• Interpretation: $\det(\Sigma^{1/2}) = \prod_{j=1}^n \sigma_j$ so $\det(\Sigma^{1/2})^{1/n} = \left(\prod_{j=1}^n \sigma_j\right)^{1/n} = \bar{\sigma}$

where $\bar{\sigma}$ = geometric average of the principal s.d.'s

then: $h[X] = n \ln(\sqrt{2\pi}e \bar{\sigma}) \approx n \ln(4 \bar{\sigma})$

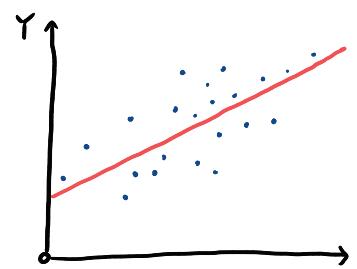
\downarrow scale
 \uparrow dimension
 \uparrow standard normal

- $\left[\begin{array}{l} \cdot X \sim N(\mu, \bar{\sigma}^2) \text{ is as uncertain as} \\ \quad (i) n \text{ i.i.d. draws from uniform on interval length } 4\bar{\sigma} \\ \quad (ii) n \text{ i.i.d. draws from normal w/ s.d. } \bar{\sigma} \end{array} \right]$

Computing Mutual Information - Example II: Correlated Gaussian Variables

• Aim: Compute $I[X; Y]$ for:

1. $X \& Y \sim N$ univariate
2. $X \& Y \sim N$ multivariate



1. Univariate: $X \sim N(\mu_x, \sigma_x^2)$, $Y \sim N(\mu, \sigma_y^2)$, $\text{Corr}[X, Y] = r_{xy}$

• What is $I[X; Y]$?

$$I[X; Y] = \begin{cases} h[X] - h[X|Y] & \leftarrow X|Y=y \text{ is } N \forall y \\ h[Y] - h[Y|X] & \leftarrow Y|X=x \text{ is } N \forall x \\ h[X] + h[Y] - h[X, Y] & \leftarrow \text{only need} \end{cases}$$

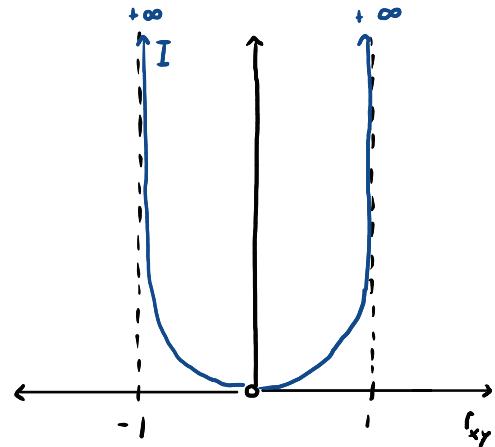
could compute
would need formula for
conditionals

formula for entropy
of N's

$$\begin{aligned}
I[X; Y] &= h[X] + h[Y] - h[X, Y] \\
&= \frac{1}{2} \ln(2\pi e \sigma_x^2) + \frac{1}{2} \ln(2\pi e \sigma_y^2) - \frac{1}{2} \ln((2\pi e)^2 \det(\underbrace{\begin{bmatrix} \sigma_x^2 & \sigma_x \sigma_y r_{xy} \\ \sigma_x \sigma_y r_{xy} & \sigma_y^2 \end{bmatrix}}_{\text{for } [X, Y]})) \\
&= \left(\frac{1}{2} \ln(2\pi e) + \frac{1}{2} \ln(2\pi e) - \frac{1}{2} \ln(2\pi e) \right) \\
&\quad + \frac{1}{2} \ln(\sigma_x^2) + \frac{1}{2} \ln(\sigma_y^2) - \frac{1}{2} \ln(\sigma_x^2 \sigma_y^2 - \sigma_x^2 \sigma_y^2 r_{xy}^2) \\
&= \ln(\sigma_x) + \ln(\sigma_y) - \ln(\sigma_x \sigma_y \sqrt{1 - r_{xy}^2}) \\
&= \ln\left(\frac{\sigma_x \sigma_y}{\sigma_x \sigma_y \sqrt{1 - r_{xy}^2}}\right) = -\frac{1}{2} \ln(1 - r_{xy}^2)
\end{aligned}$$

so: $I[X; Y] = -\frac{1}{2} \ln(1 - r_{xy}^2)$ if $X \sim N$, $Y \sim N$ (univariate)
 and $\text{Corr}[X, Y] = r_{xy}$

the more correlated
 the larger the mutual information



- notice:
 - (i) if $r_{xy} = 0$, $X \perp\!\!\!\perp Y$, $I[X; Y] = 0$
 - (ii) if $|r_{xy}| \rightarrow 1$, Y predicts X exactly, $I \rightarrow \infty$
 - (iii) I is independent of μ_X, μ_Y and σ_X, σ_Y
... scale invariant, depends only on correlation...
- } intrinsic? or just scale invariant?

2. Multivariate: $X \sim N$ n-dimensional, $Y \sim N$ m-dimensional

$$\begin{bmatrix} X \\ Y \end{bmatrix} \sim N\left(\begin{bmatrix} \mu_X \\ \mu_Y \end{bmatrix}, \begin{bmatrix} \Sigma_{XX} & \Sigma_{XY} \\ \Sigma_{YX} & \Sigma_{YY} \end{bmatrix}\right) \quad \text{where} \quad \Sigma = \begin{bmatrix} \Sigma_{XX} & \Sigma_{XY} \\ \Sigma_{YX} & \Sigma_{YY} \end{bmatrix}$$

\longleftrightarrow

$$\begin{aligned}
\Sigma_{XX} &= \text{Cov}[X], \quad \Sigma_{YY} = \text{Cov}[Y] \\
\Sigma_{XY} &= \text{Cov}[X, Y] = \Sigma_{YX}^T.
\end{aligned}$$

then: $I[X; Y] = h[X] + h[Y] - h[X, Y]$

$$= \frac{1}{2} (\ln(\det(\Sigma_{XX})) + \ln(\det(\Sigma_{YY})) - \ln(\det(\Sigma)))$$

can we reduce?

- let $V_{\text{var}}[X] = [V[X_1], V[X_2], \dots, V[X_n]]$
- $V_{\text{var}}[Y] = [V[Y_1], V[Y_2], \dots, V[Y_m]]$

↑
"variance of"

- let $D_v = \text{diag}(v_1, v_2, \dots)$

- then: $\Sigma_{xx} = D_{\text{var}[X]}^{\frac{1}{2}} R_{xx} D_{\text{var}[X]}^{\frac{1}{2}} = \begin{bmatrix} SD[X_1] & & & \\ & SD[X_2] & & \\ & & \ddots & \\ & & & SD[X_n] \end{bmatrix} \begin{bmatrix} r_{11} & r_{12} & \dots & r_{1n} \\ r_{21} & r_{22} & \dots & r_{2n} \\ \vdots & \vdots & \ddots & \vdots \\ r_{n1} & r_{n2} & \dots & r_{nn} \end{bmatrix} \begin{bmatrix} SD[X_1] & & & \\ & SD[X_2] & & \\ & & \ddots & \\ & & & SD[X_n] \end{bmatrix}$

$\Sigma_{yy} = D_{\text{var}[Y]}^{\frac{1}{2}} R_{yy} D_{\text{var}[Y]}^{\frac{1}{2}}$

$\underbrace{SD}_{SD} \quad \underbrace{\text{Corr}}_{\text{Corr}} \quad \underbrace{SD}_{SD}$

$r_{ij} = \text{Corr}[X_i, X_j]$

- and: $\Sigma = \begin{bmatrix} D_{\text{var}[X]}^{\frac{1}{2}} & & \\ & D_{\text{var}[Y]}^{\frac{1}{2}} & \\ & & \end{bmatrix} \begin{bmatrix} R_{xx} & R_{xy} \\ R_{yx} & R_{yy} \end{bmatrix} \begin{bmatrix} D_{\text{var}[X]}^{\frac{1}{2}} & \\ & D_{\text{var}[Y]}^{\frac{1}{2}} \end{bmatrix}$

$R_{xy_{ij}} = \text{Corr}[Y_i, X_j]$

- now: $\log(\det(\Sigma)) = \log(\det(D_v^{\frac{1}{2}}) \det(R)) = \log(\det(D_v)) + \log(\det(R))$

- so: $\log(\det(\Sigma_{xx})) + \log(\det(\Sigma_{yy})) - \log(\det(\Sigma))$

$$= [\log(\det(D_{\text{var}[X]})) + \log(\det(D_{\text{var}[Y]})) - \log(\det([D_{\text{var}[X]} \ D_{\text{var}[Y]}]))]$$

$$+ \log(\det(R_{xx})) + \log(\det(R_{yy})) - \log(\det([R_{xx} \ R_{xy} \ R_{yx} \ R_{yy}]))]$$

$\hookrightarrow = \log\left(\prod_{i=1}^n V[X_i] \prod_{j=1}^m V[Y_j]\right) - \log\left(\prod_{i=1}^n V[X_i] \prod_{j=1}^m V[Y_j]\right) = 0 \dots \text{all var. terms cancel out}$

$$= \log(\det(R_{xx}) \det(R_{yy})) - \log(\det([R_{xx} \ R_{xy} \ R_{yx} \ R_{yy}]))$$

- so: $I[X; Y] = \frac{1}{2} [\log(\det(R_{xx}) \det(R_{yy})) - \log(\det([R_{xx} \ R_{xy} \ R_{yx} \ R_{yy}]))]$

$$= \frac{1}{2} [\log(\det([R_{xx} \ R_{xy} \ 0 \ R_{yy}])) - \log(\det([R_{xx} \ R_{xy} \ R_{yx} \ R_{yy}]))]$$

$$= -\frac{1}{2} [\log(\det([R_{xx} \ R_{xy} \ R_{yy}])) - \log(\det([R_{xx} \ 0 \ R_{yy}]))]$$

$$= -\frac{1}{2} [\log(\det([R_{xx} \ R_{xy} \ R_{yy}])) + \log(\det([R_{xx}^{-1} \ 0 \ R_{yy}^{-1}]))]$$

$$\begin{aligned}
\text{so: } I[X; Y] &= \frac{1}{2} \left[\log(\det(R_{xx}) \det(R_{yy})) - \log(\det(\begin{bmatrix} R_{xx} & R_{xy} \\ R_{yx} & R_{yy} \end{bmatrix})) \right] \\
&= \frac{1}{2} \left[\log(\det(\begin{bmatrix} R_{xx} & 0 \\ 0 & R_{yy} \end{bmatrix})) - \log(\det(\begin{bmatrix} R_{xx} & R_{xy} \\ R_{yx} & R_{yy} \end{bmatrix})) \right] \\
&= -\frac{1}{2} \left[\log(\det(\begin{bmatrix} R_{xx} & R_{xy} \\ R_{yx} & R_{yy} \end{bmatrix})) - \log(\det(\begin{bmatrix} R_{xx} & 0 \\ 0 & R_{yy} \end{bmatrix})) \right] \\
&= -\frac{1}{2} \left[\log(\dots) + \log(\det(\begin{bmatrix} R_{xx}^{-1} & 0 \\ 0 & R_{yy}^{-1} \end{bmatrix})) \right] \\
&= -\frac{1}{2} \left[\log(\dots) + \log(\det(\begin{bmatrix} R_{xx}^{-1/2} & 0 \\ 0 & R_{yy}^{-1/2} \end{bmatrix})^2) \right] \\
&= -\frac{1}{2} \log(\det(\begin{bmatrix} R_{xx}^{-1/2} & 0 \\ 0 & R_{yy}^{-1/2} \end{bmatrix}) \det(\begin{bmatrix} R_{xx} & R_{xy} \\ R_{yx} & R_{yy} \end{bmatrix}) \det(\begin{bmatrix} R_{xx}^{-1/2} & 0 \\ 0 & R_{yy}^{-1/2} \end{bmatrix})) \\
&= -\frac{1}{2} \log(\det(\begin{bmatrix} R_{xx}^{-1/2} & 0 \\ 0 & R_{yy}^{-1/2} \end{bmatrix} \begin{bmatrix} R_{xx} & R_{xy} \\ R_{yx} & R_{yy} \end{bmatrix} \begin{bmatrix} R_{xx}^{-1/2} & 0 \\ 0 & R_{yy}^{-1/2} \end{bmatrix}))
\end{aligned}$$

$$I[X; Y] = -\frac{1}{2} \log(\det(\begin{bmatrix} I_{n \times n} & M \\ M^T & I_{m \times m} \end{bmatrix})) \quad \text{where } M = R_{xx}^{-1/2} R_{xy} R_{yy}^{-1/2}$$

• so $I[X; Y]$ is independent of all X & Y variances (scale independent), and depends "M"

to simplify, use the block determinant formula

$$\det(\begin{bmatrix} A & B \\ C & D \end{bmatrix}) = \det(A) \det(D - CA^{-1}B) \quad \text{if } A \text{ & } D \text{ are square...}$$

so:

$$\det(\begin{bmatrix} I & M \\ M^T & I \end{bmatrix}) = \det(I) \det(I - M^T I M) = \det(I - M^T M) = \prod_{j=1}^m (1 - \sigma_j(M)^2)$$

singular values of M

so: given $X, Y \sim N$ w/ Correlations R_{xx}, R_{xy}, R_{yy}

$$I[X; Y] = -\frac{1}{2} \log(\det(I - M^T M)) = -\frac{1}{2} \log\left(\prod_{j=1}^m (1 - \sigma_j(M)^2)\right) \quad \text{where } M = R_{xx}^{-1/2} R_{xy} R_{yy}^{-1/2}$$

• fully generalizes the 1D case

• in both cases I only depends on the correlations, is scale independent.

$$I[X; Y] = -\frac{1}{2} \log(1 - r_{xy}^2)$$