

### Conditional Gaussian Multivariate Distribution

The multivariate Gaussian distribution on  $\mathbb{R}^n$  is defined as,

$$\rho(\mathbf{x}; \boldsymbol{\mu}, \Sigma) = \frac{1}{(2\pi)^{n/2} |\det \Sigma|^{\frac{1}{2}}} \exp \left( -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right)$$

calculate the conditional distribution,

$$\begin{aligned} & \rho([\mathbf{x}_1^T, \mathbf{x}_2^T]^T; \boldsymbol{\mu}, \Sigma) / \rho(\mathbf{x}_2; \boldsymbol{\mu}_2, \Sigma_2) \\ & \Sigma = \int_{\mathbb{R}^n} (\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^T f(\mathbf{x}_1, \mathbf{x}_2) d\mathbf{x} \\ & = \left( \begin{array}{cc} \int_{\mathbb{R}^n} (\mathbf{x}_1 - \boldsymbol{\mu}_1)(\mathbf{x}_1 - \boldsymbol{\mu}_1)^T f(\mathbf{x}_1, \mathbf{x}_2) & \int_{\mathbb{R}^n} (\mathbf{x}_1 - \boldsymbol{\mu}_1)(\mathbf{x}_2 - \boldsymbol{\mu}_2)^T f(\mathbf{x}_1, \mathbf{x}_2) \\ \int_{\mathbb{R}^n} (\mathbf{x}_2 - \boldsymbol{\mu}_2)(\mathbf{x}_1 - \boldsymbol{\mu}_1)^T f(\mathbf{x}_1, \mathbf{x}_2) & \int_{\mathbb{R}^n} (\mathbf{x}_2 - \boldsymbol{\mu}_2)(\mathbf{x}_2 - \boldsymbol{\mu}_2)^T f(\mathbf{x}_1, \mathbf{x}_2) \end{array} \right) \\ & = \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix} \\ & \Sigma_{12}^T = \Sigma_{21} \quad \Sigma_{11}^T = \Sigma_{11} \quad \Sigma_{22}^T = \Sigma_{22} \end{aligned}$$

let's ignore the constant term in the front and denote it  $A'$ . Conditional Gaussian PDF,

$$f(\mathbf{x}_1 | \mathbf{x}_2) = A' \exp \left( -\frac{1}{2} [(\mathbf{x} - \boldsymbol{\mu})^T \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}) - (\mathbf{x}_2 - \boldsymbol{\mu}_2)^T \Sigma_{22}^{-1} (\mathbf{x}_2 - \boldsymbol{\mu}_2)] \right)$$

Recall,

$$\begin{bmatrix} A & B \\ B^T & D \end{bmatrix}^{-1} = \begin{bmatrix} I & 0 \\ -D^{-1}B^T & I \end{bmatrix} \begin{bmatrix} (A - BD^{-1}B^T)^{-1} & 0 \\ 0 & D^{-1} \end{bmatrix} \begin{bmatrix} I & -BD^{-1} \\ 0 & I \end{bmatrix}$$

let  $\Sigma_{11} = A$ ,  $\Sigma_{12} = B$ ,  $\Sigma_{21} = B^T$ ,  $\Sigma_{22} = D$ .

$$\begin{aligned} & = A' \exp \left( -\frac{1}{2} \{ (\mathbf{x}_1 - \boldsymbol{\mu}_1)^T \quad (\mathbf{x}_2 - \boldsymbol{\mu}_2)^T \} \begin{bmatrix} I & 0 \\ -D^{-1}B^T & I \end{bmatrix} \begin{bmatrix} (A - BD^{-1}B^T)^{-1} & 0 \\ 0 & D^{-1} \end{bmatrix} \begin{bmatrix} I & -BD^{-1} \\ 0 & I \end{bmatrix} \right. \\ & \quad \left. \begin{bmatrix} I & -BD^{-1} \\ 0 & I \end{bmatrix} \begin{pmatrix} \mathbf{x}_1 - \boldsymbol{\mu}_1 \\ \mathbf{x}_2 - \boldsymbol{\mu}_2 \end{pmatrix} - (\mathbf{x}_2 - \boldsymbol{\mu}_2)^T \Sigma_{22}^{-1} (\mathbf{x}_2 - \boldsymbol{\mu}_2) \} \right) \\ & = A' \exp \left( -\frac{1}{2} \{ (\mathbf{x}_1 - \boldsymbol{\mu}_1 - BD^{-1}(\mathbf{x}_2 - \boldsymbol{\mu}_2))^T (A - BD^{-1}B^T)^{-1} (\mathbf{x}_1 - \boldsymbol{\mu}_1 - BD^{-1}(\mathbf{x}_2 - \boldsymbol{\mu}_2)) + \right. \\ & \quad \left. (\mathbf{x}_2 - \boldsymbol{\mu}_2)^T \Sigma_{22}^{-1} (\mathbf{x}_2 - \boldsymbol{\mu}_2) - (\mathbf{x}_2 - \boldsymbol{\mu}_2)^T \Sigma_{22}^{-1} (\mathbf{x}_2 - \boldsymbol{\mu}_2) \} \right) \\ & = A' \exp \left( -\frac{1}{2} \{ (\mathbf{x}_1 - \boldsymbol{\mu}_1 - BD^{-1}(\mathbf{x}_2 - \boldsymbol{\mu}_2))^T (A - BD^{-1}B^T)^{-1} (\mathbf{x}_1 - \boldsymbol{\mu}_1 - BD^{-1}(\mathbf{x}_2 - \boldsymbol{\mu}_2)) \} \right) \end{aligned}$$

hence,

$$\boldsymbol{\mu}_{\mathbf{x}_1 | \mathbf{x}_2} = \boldsymbol{\mu}_1 + \Sigma_{12} \Sigma_{22}^{-1} (\mathbf{x}_2 - \boldsymbol{\mu}_2) \quad (1)$$

$$\Sigma_{\mathbf{x}_1 | \mathbf{x}_2} = \Sigma_{11} - \Sigma_{12} \Sigma_{22}^{-1} \Sigma_{21} \quad (2)$$

Furthermore, suppose  $\mathbf{x} \sim \mathcal{N}(\mu_x, \Sigma_x)$  and a variable  $\mathbf{y}$  resulting from affine transformation of  $\mathbf{x}$ , so that  $\mathbf{y} = A\mathbf{x} + \mathbf{b}$  where  $\mathbf{b} \sim \mathcal{N}(0, \Sigma_b)$ .

$$\begin{aligned}\Sigma_{\mathbf{x}} &= E[\tilde{\mathbf{x}}\tilde{\mathbf{x}}^T] \\ \Sigma_{\mathbf{xy}} &= E(\tilde{\mathbf{x}}\tilde{\mathbf{y}}^T) \\ &= E[(\mathbf{x} - \bar{\mathbf{x}})(A\mathbf{x} + \mathbf{b} - A\bar{\mathbf{x}})^T] \\ &= E[(\mathbf{x} - \bar{\mathbf{x}})((\mathbf{x} - \bar{\mathbf{x}})^T A^T + \mathbf{b}^T)] \\ &= E[(\mathbf{x} - \bar{\mathbf{x}})(\mathbf{x} - \bar{\mathbf{x}})^T] A^T = \Sigma_x A^T\end{aligned}$$

Similarly,

$$\Sigma_{\mathbf{yx}} = A\Sigma_{\mathbf{x}}$$

and,

$$\begin{aligned}\Sigma_{\mathbf{yy}} &= E[(A(\mathbf{x} - \bar{\mathbf{x}}) + \mathbf{b})(A(\mathbf{x} - \bar{\mathbf{x}}) + \mathbf{b})^T] \\ &= AE[(\mathbf{x} - \bar{\mathbf{x}})(\mathbf{x} - \bar{\mathbf{x}})^T] A^T + E[\mathbf{b}\mathbf{b}^T] \\ &= A\Sigma_{\mathbf{x}} A^T + \Sigma_b\end{aligned}$$

hence, comparing to (1) and (2)

$$\begin{aligned}\boldsymbol{\mu}_{\mathbf{x}|\mathbf{y}} &= \boldsymbol{\mu}_{\mathbf{x}} + \Sigma_{\mathbf{x}} A^T (A\Sigma_{\mathbf{x}} A^T + \Sigma_b)^{-1} (\mathbf{y} - A\bar{\mathbf{x}}) \\ \Sigma_{\mathbf{x}|\mathbf{y}} &= \Sigma_{\mathbf{x}} - \Sigma_{\mathbf{x}} A^T (A\Sigma_{\mathbf{x}} A^T + \Sigma_b)^{-1} A\Sigma_{\mathbf{x}}\end{aligned}$$

references:

1. [http://www.cs.columbia.edu/liulp/pdf/linear\\_normal\\_dist.pdf](http://www.cs.columbia.edu/liulp/pdf/linear_normal_dist.pdf)
2. Chirikjian, G S. 2009. Stochastic Models, Information Theory, and Lie Groups: Classical Results and Geometric Methods vol 1.
3. Timothy D. Barfoot. 2017. State Estimation for Robotics.