

Statistical Inference for Everybody and a Linguist

It wasn't me!

Textbooks in Language Sciences



Textbooks in Language Sciences

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Statistical Inference for Everybody and a Linguist

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Preface

1 The Need for A New Introduction to Statistics

2 Methods and the Progression of Topics

Readers should be aware that Chapter 2 progresses very slowly, is very repetitive, and contains a lot of detail and material for illustrative purposes. This changes in later chapters, and we speed up the narrative significantly. However, a slow start is clearly desirable given what we said in Section 1.

Acknowledgments

1 Scientific Inference and Error

2 Guessing and Counting

Overview

In this chapter, we introduce the notion of frequentist probability. It's the probability of a certain event under the assumption of pure randomness. Pure randomness is not a very precise technical term, but we hope it gives the right impression. For example, frequentist probability of rolling six pips with a fair die is $1 \div 6$ (or roughly 0.17) before you roll the die. This also implies that the proportion of rolls of a fair die showing six pips in a very long or endless sequence of rolls will be 0.17 (or 17%). Notice that after rolling the die, it either shows six pips or it doesn't show six pips, and there is no longer a probability attached to the result. It becomes a fact. We apply reasoning around frequentist probability to lotteries, parlour games with self-proclaimed psychics, and corpus studies. It is shown that unexpected outcomes under the assumption of pure randomness are those that have a low frequentist probability. In other words, those outcomes would be rare if we repeated the experiment over and over again, and if only pure randomness were in control of the outcomes (rather than some influencing factor, such as the manipulation of a die). To quantify these notions and make some careful and limited scientific inferences based on this quantification, we introduce several mathematical concepts. The [binominal coefficient](#) is a way of calculating the number of ways of choosing a specific number of elements from a set of elements regardless of their order and without replacement. The [p-value calculated in Fisher's Exact Test](#) is the frequentist probability of obtaining a certain result or a more extreme result in simple experiments involving two two-valued variables and counts of their instances. The [effect size](#) for Fisher's Exact Test is a measure of how strongly two variables interact. Finally, [probability distribution functions](#) are general functions to calculate (and plot) such probabilities that have specific mathematical properties.

Readers should be warned that this chapter is a tad narrative and slow in order to create insight into the basic idea of frequentist inference. We consider this vital in order to remedy problems with frequentist practice (not with the frequentist philosophy). Take your time and be assured that we will up the pace in later chapters.

Problem Statement: If there's Nothing Going On ...

Let's consider three rather simple questions: (i) You know that you aren't prescient, but you decide to play the lottery anyway. How surprised would you be if you won the big prize? (ii) You don't believe that your friend, who claims to be a psychic, actually has psychic abilities. Nevertheless, you give them a chance and invite them to a party where they have to guess the phone numbers of all other guests. How surprised would you be if they guessed the phone numbers of all your other guests correctly? (iii) Given that most grammatical theories (which have something to say about passives) claim that the verb *sleep* cannot be passivised at all or only under very marginal circumstances, how surprised would you be to find ten passives of *sleep* in a corpus of English? [Please think thoroughly about your answers to these questions before continuing on.](#)

2.1 Unexpected Outcomes

Did you think about the questions from the Problem Statement? Really? Good. Here's one possible discussion. Most importantly, the questions from the Problem Statement cannot be answered properly in their given form, simply because there are significant data missing. As for (i), the question doesn't specify what kind of lottery we're considering. Is it a simple urn at a funfair from which you get to draw one out of a thousand lots, and only one of the thousand lots is a win? Or is it the Eurojackpot, where you have to guess five numbers out of fifty (plus some additional single numbers) correctly to win the big prize? Most likely, you either decided that you can't answer the question, or you answered it with respect to some specific type of lottery by way of example. Maybe you even wondered whether the lottery is supposed to be fair or not. However, when presented with this specific example, most people typically don't worry too much about how the lottery was conducted and whether it was fair. At least with big national lotteries, they tend to put trust in there being sufficient oversight and the draw being—here it comes—properly random. Most importantly (and disappointingly), they see no way to rigging the lottery in their own favour. Considering the urn at the funfair, people likely assume that it's rigged anyway, but they don't care

(at least in the Free World).¹

The scenario in (ii) is similar, and there's also relevant information missing. You probably decided whether your degree of surprise would critically depend on the number of guests at the party and the number of digits phone numbers have. In my youth, smaller German villages (like Twiste, located in the Twistetal district) still had three-digit phone numbers, for example. If the local Twiste psychic only had to guess one such phone number, guessing that number correctly even without psychic powers would be much less awe-inspiring than guessing the ten-digit U.S. phone numbers of 28 guests at a party in NYC with the same accuracy, for example. Furthermore—and most likely because it involves a psychic—this scenario makes people much more suspicious of whether and how it was ensured that the psychic didn't cheat. Maybe they have a secret app that exploits a vulnerability in close-by mobile phones, and they simply read the numbers off of peoples' phones. Maybe the party was announced in a group chat on some messenger app, and they tracked all the guests' numbers down in the app before the party. Maybe the host or some other guest conspired with the psychic and gave them all the numbers, either as a practical joke or even because they want to get people to pay for the psychic's services (tracking down dead ancestors who lived as maids and servants at the court of Louis XIV of France).

Example (iii) is much more intricate and, in a way, boring, which is why it only intrigues linguists. Some linguists would smirk at you and claim that they don't care about corpus examples because it was determined once and for all by a cherubic figure (who never laughs) that examples from corpora don't count for anything. Some linguists, on the other hand, would take the ten sentences as conclusive evidence that whatever random modification to their theory they came up with is provably correct, or that somebody else's theory is provably incorrect.² What were your thoughts? We certainly hope that you don't belong to either of the aforementioned tribes of linguists and that you saw the parallels to the first two scenarios. Above all, quantitative considerations play a role, among others: How large is the corpus? How often does *sleep* occur in the corpus, regardless of its voice? How many active and passive verbs occur in the corpus? Also, the question of whether it was a fair draw are vastly more complicated than in the case of a lottery. For example, is it a corpus of language produced by native speakers, children, L2 learners of English, frontier large language models,

¹*It doesn't get more American than this, my friend. Fatty foods, ugly decadence, rigged games.* (Murray Bauman, Episode 7 of *Stranger Things* 3)

²*Whenever I find even one example that contradicts a claim, I consider that claim refuted.* (an unnamed linguist, p. c.)

or even some cute language bot from 1998? Does the composition matter considering your research question? What exactly is your research question? Finally, the underlying theory from which it allegedly follows that *sleep* cannot be passivised needs further inspection. Does it also exclude the figura etymologica for such unergative verbs? Maybe all ten sentences are instances of silliness such as (1). Would the result still count as unexpected, regardless of any quantitative evaluation?

- (1) The sleep of Evil has been slept by many a demon.

It's a muddle! Therefore, we'll use a simple non-linguistic example in Section 2.2 to introduce some important statistical concepts that concern the numerical side of this muddle. The example is about tea, and it's extremely famous, so anyone interested in statistical inference should be aware of it, even if they're not in Tea Studies. Then, we'll return to the questions from the problem statement.

2.2 Tea and Milk

2.2.1 An Introduction to Unexpected Outcomes

What unites the examples in the Problem Statement is that they describe a confrontation with chance. Given this confrontation, you're asked **what kind of a result would be unexpected under the assumption that there is nothing going on:** (i) you're not prescient, (ii) the psychic isn't actually a psychic, (iii) *sleep* cannot be passivised. In this section, we formalise the notion of **unexpected outcome** in relation to experiments.

First of all, an *unexpected outcome* cannot be one which is deemed totally impossible. If you saw no chance of winning the lottery, you wouldn't play it. If you absolutely knew for certain that your psychic friend couldn't guess phone numbers, you wouldn't ask them to guess numbers at your party, except maybe if there were others who didn't know for certain that the psychic didn't have the ability in question. Finally, if you were absolutely certain anyway that *sleep* cannot be passivised, you wouldn't bother to do a corpus search for passivised forms of that verb. In fact, that's what many self-described theoretical linguists do. Clearly, unexpected outcomes are not miracles where everything we know about the world is up for debate.

What we usually mean when we deem an outcome *unexpected* is that it had a very slim chance of occurring before we made it occur. Mathematically, the most

straightforward case is the one with the urn at the funfair. If there are a thousand lots in the urn, one of them is a win, all the others are duds, and you draw exactly one, most people know you have a chance of one in a thousand (or 1:1 000) to win. Usually, it is understood intuitively that this means that if you played this game over and over again, you would end up winning in one of a thousand rounds on average. (Playing the game over and over again, each time with a fresh urn of one thousand lots, not gradually emptying one and the same urn, of course.) That's why playing it once and winning is unexpected or surprising: winning is a rare event given the way the urn was set up (one winning lot and 999 duds). The maths is slightly more complex for the Eurojackpot because you have to choose five numbers out of fifty and not one lot out of a thousand, but it essentially follows the same logic. For the psychic guessing phone numbers, the idea is also the same once the number of phone numbers and the number of the digits per phone number has been determined. We will return to the third scenario (the corpus study) later, but we can apply a similar logic even to that example.

2.2.2 The Total Number of Possible Outcomes

In each of the scenarios, we need to know the number of potential outcomes in order to quantify how unexpected a single specific outcome is. The higher the number of overall possible outcomes, the more unexpected a specific outcome is. A seminal application of this idea to scientific reasoning is reported in Fisher1935a, and we'll introduce it here before applying the same reasoning to the scenarios from the Problem Statement. In that book, Ronald A. Fisher reports an event where Muriel Bristow, herself a scientist, claimed that she could taste whether the milk or the tea was poured into a cup first. While it is not impossible that some physical properties of the mixed liquids differ depending on their order of being poured into the cup, some doubt was in order. Therefore, Fisher devised an experiment to shed some light on the substance of Bristow's claim. She was presented with eight cups, four tea-first cups and four milk-first cups. Otherwise, the cups were identical. Her task in the experiment was to find the four tea-first cups merely by tasting. Very much like winning a lottery after buying just a single lot, some outcomes of this experiment might surprise us by being relatively unexpected if Bristow didn't have the ability which she claims to have. We still wouldn't consider it proven above all doubt that she does indeed have the ability if that happened. However, we'd at least not consider her claims of being a tea expert refuted if she guessed a surprising number of cups correctly. The question is: what's a surprising number? How many cups does she have to get right for us to call it an unexpected outcome?

2 *Guessing and Counting*

Statistics doesn't offer a final answer to this question. However, it provides the maths upon which we need to base our answer. Remember that Muriel Bristow has to choose four cups out of eight, and we first need to calculate how many distinct sets of four cups out of eight she could potentially choose, without even considering whether she chose the right ones. Let's do it. In Figure 2.1, we illustrate the eight cups. While they would all look exactly the same in the real experiment, we've made it easier to follow the argument by showing the tea-first cups with steam and the milk-first ones without steam. Furthermore, we've coloured the cups to make them identifiable individually. There is one gray, one red, one blue, and one green cup for each of the conditions (milk-first or tea-first). Again, in the real experiment, cups should have the exact same physical properties.

Choices: 8								
Chosen: 0								

Figure 2.1: Four tea-first cups (steaming) and four milk-first cups (not steaming) for Muriel Bristow to choose from; in the actual experiment, they'd all look exactly the same (without colours).

Before choosing her first cup, Ms Bristow obviously has 8 choices. She could pick the red steaming cup, the gray steaming cup, the gray non-steaming cup, etc. Figure 2.2 shows the situation after a first cup was chosen.

Choices: 7								
Chosen: 1								

Figure 2.2: That's the first cup chosen! Choices left: 7.

Before she continues on and picks the second cup, only 7 choices are still available. Notably, *for each of the 8 distinct choices she had in the beginning, she now has 7 distinct subsequent choices*. In the illustration, she chose the blue steaming cup and has the other 7 cups still available. Had she chosen the red steaming cup instead in the first step, she'd now be confronted with a different set of 7 options. That means that after picking another cup (Figure 2.3), she has already decided

on one specific choice from among $8 \cdot 7 = 56$ possible choices.³ Put differently, she has taken 1 out of 56 possible decision paths to choose 2 out of 8 cups.

Choices: 6								
Chosen: 2								

Figure 2.3: After another cup was chosen, there are now 6 choices left.

The story goes on in a similar vein. Let’s assume she chooses the red steaming cup as her third pick, as in Figure 2.4. Now, she has made 1 of $8 \cdot 7 \cdot 6 = 336$ possible choices, since for each of the $8 \cdot 7$ options left over after her previous decision, she had 6 distinct choices available.

Choices: 5								
Chosen: 3								

Figure 2.4: As Ms Bristow picks another cup, we’re down to 5 choices.

To make things more interesting, she now makes her first incorrect guess. She picks the green non-steaming one, which is a tea-first cup. She has now decided on 1 specific configuration from $8 \cdot 7 \cdot 6 \cdot 5 = 1,680$ possible configurations. Or has she? As we mentioned in Footnote 3, there’s a catch.

Not chosen: 4								
Chosen: 4								

Figure 2.5: Ms Bristow has picked her last cup.

First, she chose the blue steaming cup, then the gray steaming cup, then the red steaming cup, and finally the the green non-steaming cup. But is this really the only way to arrive at the same result? In the first step, she had 8 options, and

³This is not exactly true. There’s a catch to which we’ll return presently. Do you remember from grammar school maths what it is?

2 Guessing and Counting

she chose the blue steaming cup, leaving her with 7 choices, etc. She could have chosen the red steaming one and still arrived at the same end result via a different path. For example, she could have chosen the red steaming cup first, then the gray steaming cup, followed by the blue steaming one, and finally the green non-steaming one. Put in quantitative terms, there are groups within the set of 1,680 decision chains that yield identical results, at least if the order in which the cups were selected is irrelevant. And for the purpose of this experiment, the order is indeed irrelevant.

How do we know how many of the 1,680 decision paths lead to identical results? Well, how many different ways of ordering four cups are there? Imagine you had to put 4 cups on a table from left to right one by one. In the first step, you can choose from among the 4 cups. For each of these 4 distinct choices, there are 3 distinct subsequent choices because you'll have 3 cups left. Then there are 2 choices each, then just 1. Hence, the groups of identical outcomes should each have a size of $4 \cdot 3 \cdot 2 \cdot 1 = 24$. There are thus

$$\frac{8 \cdot 7 \cdot 6 \cdot 5}{4 \cdot 3 \cdot 2 \cdot 1} = \frac{1,680}{24} = 70$$

truly distinct sets of four cups to be chosen from among a set of eight cups.⁴

This just gives us the number of distinct sets of four cups we can choose from 8 cups. So, how unexpected is her performance (3 cups detected correctly) given that there are 70 different ways of choosing 4 cups out of 8? This would have been easier to quantify if she had guessed all four cups correctly. Clearly, there is only 1 such immaculate result. Had Ms Bristow chosen the green steaming cup instead of the green non-steaming cup, it would have been the immaculate guess. Obviously, by merely guessing (without any special sensory capability), Muriel Bristow would produce such a perfect result in 1 out of 70 runs if the experiment were repeated over and over again. Put differently, there is a 1:70 chance of guessing the four cups by mere luck. In technical terms, the **frequentist probability** of hitting the tea jackpot by uninformed guessing is $1 \div 70 \approx 0.014$. This probability is sometimes converted to a percentage, in this case 1.4%.⁵ Would this be a highly unexpected result? So unexpected maybe that you'd doubt that Muriel Bristow merely got lucky? Well, you tell me!

⁴And for those who are beginning to remember their elementary stochastics: we get different results if the order is not irrelevant and if a cup can be chosen more than once (replacement).

⁵From our perspective, this conversion to a percentage is not at all wrong inasmuch as 1.4% of an endless sequence of tries would result in an immaculate result, even if the taster is really just guessing. However, in scientific contexts, probabilities are expressed properly as real values between 0 and 1, not as percentages.

2.2.3 The Number of Some Less Than Perfect Outcomes

Before proceeding to such delicate matters of scientific inference, let's calculate the probability of guessing 3 cups correctly and getting 1 wrong, as in the example. To do that, we first introduce a convenient general notation for the maths of choosing k items out of n . First, notice that in the calculations above we often multiplied a natural number with the next smaller natural number, then the next smaller number, and so on. For example, we calculated $4 \cdot 3 \cdot 2 \cdot 1$. Such an operation, where we multiply a natural number repeatedly with its next smaller neighbour until we reach 1, is called a **factorial**, and it is expressed as $n!$ such that $4! = 4 \cdot 3 \cdot 2 \cdot 1$ if $n = 4$. To calculate the number of possible decision chains for 4 out of 8 cups, we calculated $8 \cdot 7 \cdot 6 \cdot 5$, but then we didn't go down all the way to 1. For two cups out of eight, we'd have calculated $8 \cdot 7$ (two choices, then stop), etc. In general and using n as the variable encoding the number of items and k as the variable encoding the number of items to choose, this can be expressed as

$$\frac{n!}{(n-k)!}$$

To illustrate, we insert the concrete numbers from our example above:

$$\frac{8!}{(8-4)!} = \frac{8!}{4!} = \frac{8 \cdot 7 \cdot 6 \cdot 5 \cdot 4 \cdot 3 \cdot 2 \cdot 1}{4 \cdot 3 \cdot 2 \cdot 1} = 8 \cdot 7 \cdot 6 \cdot 5$$

To account for the decision paths that led to identical results, we divided this further by $4!$ because $k = 4$ items can be arranged in $4 \cdot 3 \cdot 2 \cdot 1$ ways. Generally, we need to divide by $k!$. This gives us the **binomial coefficient** used to calculate the number of distinct sets of k items from n items without replacement and irrespective of their order:

$$\binom{n}{k} := \frac{n!}{(n-k)!k!}$$

The binomial coefficient is usually read n choose k , but it's okay to read it as k chosen from n etc. Since the factorial results in very large numbers even for relatively low input numbers, it is often practically infeasible to calculate the binomial coefficient using the above formula, and several alternative methods for calculating it are available. They can be found on Wikipedia or in any book teaching applied maths.

With this, we can now finally calculate how many ways there are of choosing 3 tea-first cups and 1 milk-first cup out of 8 cups in total where there are 4 tea-first and 4 milk-first cups. This is easy if we regard the 8 cups as two sets of 4

factorial

binomial
coefficient

2 Guessing and Counting

cups (milk-first and tea-first). Muriel Bristow thus chose 3 cups out of 4 correctly and 1 cup out of 4 incorrectly, hence:

$$\binom{4}{3} \cdot \binom{4}{1} = \frac{4!}{1!3!} \frac{4!}{3!1!} = \frac{24}{6} \cdot \frac{24}{6} = 4 \cdot 4 = 16 \quad (2.2)$$

There are thus 16 distinct sets of 3 correct and 1 incorrect cups. We already know that there are 70 ways of choosing any 4 cups out of 8, and hence $16 \div 70 \approx 0.23$ is the frequentist probability of achieving such a result by chance. Would this be a highly unexpected result? Most likely not, because in about a quarter (23%) of an endless sequence of runs of such an experiment, anyone would reach this level of accuracy just by guessing.

There is one final amendment that we should make. In order to evaluate how unexpected a result with three correctly chosen cups is, it is more informative to ask how often such a result [or an even better result](#) would be when someone's just guessing. Hence, we should add the number of possible configurations with four correct cups, which is 1:

$$\binom{4}{4} \binom{4}{0} + \binom{4}{3} \binom{4}{1} = 1 \cdot 1 + 4 \cdot 4 = 1 + 16 = 17$$

The probability of obtaining 3 or 4 correct cups by chance is thus $17 \div 70 \approx 0.24$. Again, in the light of this number, choosing 3 correct cups wouldn't be an unexpected result at all. It would provide no grounds whatsoever for rejecting the hypothesis that Muriel Bristow is just guessing. Incidentally, she allegedly chose all four cups correctly when the experiment was actually conducted. So, do you find this impressive?

2.2.4 Towards Making an Inference

What does this statistical anecdote show? We hope to have illustrated that there is a way of calculating how unexpected specific results of certain experiments are before the experiment is conducted and under the assumption that the experiment works essentially like a lottery: the person does not have the ability in question, the effect predicted by a theory or a specific hypothesis doesn't exist, the results are completely random. Then, if we then run the experiment and the outcome happens to be one of the highly unexpected ones, our surprise might lead us to believe that the experiment wasn't just a lottery, but that there is something going on: the person does indeed have the ability in question, the effect predicted by a theory or a specific hypothesis does indeed exist, etc. In general, we'd assume that there isn't nothing going on. We would be making a

statistical inference based on the outcome of the experiment. Since inference is a process fraught with peril, much more will have to be said about it, for example in Section 2.3.

However, a typical type of misunderstandings that comes with ill ways of phrasing the mathematical results of frequentist statistics must be mentioned immediately. Some people have a tendency to say things like “*Because guessing three or more cups correctly only has a probability of 0.24, the probability that Muriel Bristow has the ability to detect tea-first cups is $1-0.24$ or 0.76 .*” or “*We can be 76% certain that she has the capability in question.*” This is all completely wrong, and you should actually find such interpretations surprisingly unwarranted if you followed our argument. First, consider statements such as “*the probability that she has the ability in question is 0.76* ” and how nonsensical they are regardless of the specific number. In reality, it’s a fact that she either has the ability, or she doesn’t have it, and there’s no probability attached to a fact. Probability only plays a role when there is a confrontation with chance, and before the dice are rolled. The probability we calculated above **was** the probability of obtaining the result which we then actually obtained **before we conducted the experiment and under the assumption that there was no ability, effect, etc. but just pure randomness**. We frequently call this hypothetical absence of an effect the **assumption of nothing-going-on**. Second, we’d like to ask you what it really means to be so-and-so percent certain of something. Honestly, please provide any sound definition of this notion! We consider this notion (sometimes also called a *subjective probability* or similar) to be utterly undefined and above all irrelevant to science as science is about how things really are and not about our subjective certainty of how they are. However, we suggest you also give those people a chance who think the complete opposite. Google *subjective Bayesianism*, read a few blogs, and then read Senn2011. However, everybody (including—I think—Bayesians) agrees that the **frequentist probabilities calculated in this book have nothing to do with “being certain” about anything to a quantifiable degree**. They serve to **test scientific claims** in order to slowly establish them as unrefuted in a piecemeal fashion. Please keep the following Big Point in mind when reading the next section, which is about inference.

Big Point: Unexpected Outcomes

The outcome of an experiment is unexpected if it had a low probability

before the experiment was conducted under the assumption of nothing-going-on. After that, the outcome is a fact and doesn't have a meaningful frequentist probability attached to it. A low probability of a specific outcome means that it would be rare if the experiment were conducted very often.

2.3 Inference With Fisher's Exact Test

2.3.1 P-Values and Sig Levels

There are two important aspects with respect to making inferences from unexpected outcomes. Can you be *sure* that Muriel Bristow had the capability of discerning tea-first cups from milk-first cups just because in the real experiment, she pointed out 4 correct cups, not just 3 as in our mathematical illustration? The answer is clearly negative because, well, $p \approx 0.014$ and not $p = 0$. To illustrate, consider my grandfather on my father's side, Carl. Carl used to play the German national lottery with a bunch of friends, and in 1962, they won the big prize. (True story!) They guessed 6 numbers out of 49 correctly. Would you take their win as evidence that they were prescient and could foresee the number that would be drawn? Even better, would you take it as evidence that extra-sensory perception (ESP) exists? After all, the p-value is bloody low:

$$p = \frac{1}{\binom{49}{6}} = \frac{1}{13,983,816} \approx 0.0000000715$$

Clearly, most readers would not make such an inference, regardless of how low the p-value is. The interpretation of a statistical result needs to be well-informed about the mechanism that brought about the result, and it must be made with great care. First of all, no known physical and/or cognitive mechanism that we know of could account for ESP, which is why most people don't even bother to run experiments investigating ESP. Hypotheses about why ESP should be real simply don't exist. Also, notice that we did not, in fact, conduct an experiment about my grandfather's ESP capabilities, but I merely told a story about him winning the lottery. At some point I just started calling it an experiment, probably without most people noticing. I could have told a similar true story about any random person I knew (a friend's girlfriend's uncle or whomever) if it so happened that that person won the lottery at some arbitrary point in the past instead of my

grandfather. If you allow yourself to look anywhere for evidence of something, you're bound to find it somewhere, and a low p -value becomes utterly meaningless. Interestingly, Carl and his friends repeated the "experiment" over and over again, playing the national lottery every week for roughly ten years in total (over five hundred draws) without ever winning any considerable amount of money ever again. This really makes it look like they were just guessing numbers and got lucky exactly once. Obviously, repeating experiments (so-called **replication**) is a very good way of testing inferences made based on unexpected outcomes further.

replication

Astonishingly, researchers in soft sciences are often satisfied if $p < 0.05$ in order to proceed with a substantive inference from an experiment. Such thresholds are often called the α level, although we prefer **sig level**. Setting $\text{sig} := 0.05$ means that researchers are satisfied to make an inference if the outcome of an experiment would only be expected in 1 out of 20 experiments under the assumption that there is nothing going on, i. e., that there really isn't any effect. While this may be justified in some cases, automatically assuming such a sig level is ill-advised and outright insane. We'll come back to this point again and again, but to show you that a chance of 1 out of 20 usually wouldn't give you much confidence when there is any important matter at stake, think about the old game show *Let's Make a Deal*. In that show, contestants regularly had to choose one door out of three, and there was a big prize behind one of the doors and duds called *Zonk* or less impressive prizes behind the other two. Let's modify the rules slightly: In the soft-sciences version of *Let's Make a Deal* there are 20 doors. Behind 19 of them, there are prizes worth a significant fortune (grant money, which means money plus prestige), but one of the doors hides an automated gun turret which instantly kills the contestant if they choose that door. Would you expect any sane person to participate in such a game show? Of course you wouldn't! People who—given a choice—wouldn't take any substantial risk in real life with a 1 in 20 chance are happy to bet the future of linguistics or social psychology on such a chance by setting $\text{sig} := 0.05$ (and not doing replication). On the other hand, we have seen that even $p \approx 0.0000000715$ might be meaningless. Clearly, doing maths is easy, but making good inferences isn't. Therefore, two of the major themes in this book are (i) that you shouldn't take unnecessarily high risks in making scientific inferences from data and (ii) that there is no recipe-like procedure that leads to good inferences.⁶

sig level

In the next section, we will return to the corpus example from the Problem Statement and formalise the procedure described above in the form of the so-

⁶Anyone who thinks they're a Bayesian and just shouted "Yeah, except in Bayesian inference!" should stop reading and take a long, hard look in the mirror.

called Fisher Exact Test or just Fisher Test. You'll see that the logic of the Tea Tasting Experiment lies behind the omnipresent 2x2 tables that often pop up in the corpus literature, especially in research on collocations and collocations (StefanowitschGries2003, Evert2008).

2.3.2 Fisher's Exact Test and Null Hypotheses

Our examples from the Problem Statement and the Tea Tasting Experiment are all about comparing counts of events. How often was a correct cup of tea chosen relative to the maximal number of events of choosing a cup of tea in the experiment? The comparison of these numbers allowed us to calculate the probability of an outcome as extreme as or more extreme than the actual outcome under the assumption that the process generating the guesses (for example, Muriel Bristow) is completely random. Such counts are customarily summarised in **contingency tables**, see Table 2.1.

contingency
tables

Table 2.1: A typical 2x2 contingency table for the Tea Tasting Experiment with row sums, column sums, and a grand total

		Reality		
		Tea-First	Milk-First	
Bristow	Tea-First	3	1	4
	Milk-First	1	3	4
		4	4	8

A contingency table tabulates counts of events characterised by two variables, each having two or more discrete possible values. Here, one variable (called *Reality* in Table 2.1) characterises the event as **involving a real tea-first or a milk-first cup**, and the other variable (called *Bristow* in Table 2.1) characterises the event as **involving a cup designated by Ms Bristow as a tea-first or a milk-first cup**. One variable is shown in columns, the other one in rows, and the table shows us how often which values of the two variables co-occur. In row 1, we put the 4 events involving tea-first cups according to Muriel Bristow. In row 2, we put the 4 events involving milk-first cups as assigned by her. The row sums in the last column reflect the fact that there were indeed 4 events (and cups) for each condition. In the two columns, we count the event involving real tea-first and milk-first cups. As you can see, of the three real tea-first cups, Muriel Bristow classified 3 as tea-first (cell 1,1) and 1 as milk-first (cell 2,1).⁷ The opposite is true for the real milk-first cups.

⁷In matrices and tables, it is customary to index cells first by rows, then by columns. Hence, cell

To calculate the p-value for Fisher's Exact Test—which is exactly what we've been introducing in this chapter—we only need to consider either the rows or the columns and the corresponding sums. In Table 2.2, the cells shaded in blue and the cells shaded in green suffice to calculate the probability of drawing 3 from 4 and independently 1 from 4, which corresponds exactly to the binomial coefficient calculated in Equation 2.2.

Table 2.2: The same contingency table with the relevant components for Fisher's Exact Test highlighted

		Reality		
		Tea-First	Milk-First	
Bristow	Tea-First	3	1	4
	Milk-First	1	3	4
		4	4	8

As was shown above, we need to add the probability of obtaining a more extreme result, which is illustrated for completeness in Table 2.3.

Table 2.3: The contingency table for the *even more extreme result*

		Reality		
		Tea-First	Milk-First	
Bristow	Tea-First	4	0	4
	Milk-First	0	4	4
		4	4	8

Finally, we turn to the corpus example from the Problem Statement.⁸ As we pointed out in Section 2.1, the Problem Statement mentions 10 passives of the verb *sleep*, but we need more information. Let's introduce that information and the argument that comes with it, roughly following the logic behind collostructional analysis.⁹ Assume that we drew 90 active sentences containing *sleep* in addition to the 10 passives. Furthermore, assume that the corpus contains 1,100 sentences

1,1 is the upper-left cell. Cell 2,1 is the lower-left cell. The upper-right and lower-right cells are indexed 1,2 and 2,2, respectively.

⁸Be warned that people actually use the test like this, but that this is a slightly incorrect use. There will be ample discussion of this caveat throughout the book.

⁹Paradoxically, we consider collostructional analysis to be suboptimal as an illustration of Fisherian logic of inference. However, as it used to be the most prominent use case of Fisher's Exact Test in linguistics, we use it to introduce the test and critique the specific use of it in the process.

2 Guessing and Counting

in total, 890 of them active sentences, 210 passive sentences. The corresponding contingency table is shown in Table 2.4.

Table 2.4: A contingency table approximately as found in collostruc-tional analysis

		Voice		
		Active	Passive	
Verb	<i>sleep</i>	90	10	100
	Other	800	200	1,000
		890	210	1,100

Descriptively, it is the case that *sleep* occurs less frequently in the passive than all other verbs. Only 10% of all occurrences of *sleep* are passives, but 20% of all other verbs are passives. Using the maths introduced in this chapter, we can attempt to quantify how unexpected this result is under the assumption of nothing-going-on, and the story goes as follows. If we drew 100 sentences randomly from this corpus, what would be the frequentist probability of drawing exactly 90 active sentences and 10 passive sentences, given the overall distribution of voice in the corpus? Clearly, it would be:

$$\frac{\binom{890}{90}\binom{210}{10}}{\binom{1,100}{100}} \approx \frac{6.573 \cdot 10^{141}}{1.423 \cdot 10^{144}} \approx 0.005$$

This probability is useful because [we obtained the result by querying for all sentences containing *sleep*, not by drawing random sentences](#). Based on this, we can now inch our way towards making an inference about *sleep*. Our theory states that there should be no or at least very few passives of verbs like *sleep* (unaccusatives) compared to other verbs, many of which can be readily passivised (transitives, unergatives). If *sleep* behaved like the average verb, a sample of sentences containing it would be expected to resemble a random sample from the corpus with respect to the number of actives and passives. The more extreme the distribution of verbal voice in the *sleep* sample is, the more we are inclined to assume that *sleep* does not behave like the rest of the verbs with respect to passivisation. This is actually a similar logic as in the Tea Tasting Experiment. An unexpected result under the assumption of mere guessing on Muriel Bristow's part is conceptually very similar to an unexpected result regarding the distribution of verbal voice in a corpus sample under the assumption that the sample is not in some way different from the rest of the corpus. The Fisherian type of the frequentist logic of inference is based on this kind of argument: Since it is

very difficult to come up with quantitative evidence in favour of research hypotheses (see Chapter 1), a **Null Hypothesis** (or simply *Null*, symbolised H_0) is constructed. The Null states in some way that **the effect predicted by the theory is absent**. If the result obtained in the experiment has a very low frequentist probability under the assumption that the Null is true (i. e., it's an unexpected result), the experiment is assumed to lend some limited support for the theory (or rather for a minor hypothesis which is part of the theory). Unexpected results are not in any way taken as proof of the theory or a part of the theory, and it's obvious why. First, we never know whether a rare (unexpected) event has occurred by chance, regardless of how unexpected it was **before we conducted the experiment**. Our calculations are based on the realisation that it is not at all impossible to obtain unexpected results by chance. On the contrary, the p-value quantifies the expectedness of the actual result under a given Null (i. e., by chance) with perfect precision. Second, take the abovementioned "experiment" regarding my grandfather's ESP capabilities with $p = 0.0000000715$. It's practically irrelevant how low the p-value is, most people will not take it as evidence that my grandfather or one of his friends could foresee the numbers drawn in next week's national lottery. Thus, the strength of the evidence depends on many factors, among them the design of the experiment and the p-value.

Let's keep this in mind and complete the maths. The value of $p = 0.005$ calculated above is just the probability of obtaining exactly 10 passives and 90 actives under the informally stated H_0 : **The verb *sleep* is passivised as often as all other verbs**. Since any more extreme (i. e., even lower) number of passives would be at least as good evidence against the Null, we should include them. Hence:

$$\frac{\binom{890}{90}\binom{210}{10}}{\binom{1,100}{100}} + \frac{\binom{890}{91}\binom{210}{9}}{\binom{1,100}{100}} + \dots + \frac{\binom{890}{100}\binom{210}{0}}{\binom{1,100}{100}} \approx 0.008$$

This is indeed a possible p-value as calculated in Fisher's Exact Test. It's not the only possible p-value, as we will show immediately.

2.3.2.1 The Direction of the Null Hypothesis

Let's take stock. We've shown that the Tea Tasting Experiment and other confrontations with chance can be analysed statistically with a very simple logic. If Ms Bristow has no special tea-tasting capabilities, she would still guess 3 or 4 cups correctly in 24% of an infinite or at least very long sequence of trials. If *sleep* did indeed behave like the totality of all other verbs with respect to passivisation in the fictional corpus described above, 100 instances of *sleep*, randomly sampled

from a corpus would still lead to just 10 or less passives in 0.8% of an infinite or at least very long sequence of trials. If my grandfather and his friends weren't prescient, they'd still win the big prize in 0.00000715% of an infinite or at least very long sequence of draws.

In the case of the corpus study, we had a specific substantive hypothesis generated by many theories of syntax and the lexicon in mind, namely that unaccusative verbs cannot be passivised or are at least very rarely passivised. Hence, we calculated the probability of obtaining a sample with as few or fewer passives of *sleep* in order to weaken the Null and mildly support the family of hypotheses to which the substantive hypothesis belongs.¹⁰ This means, however, that the Null couldn't have been *The verb 'sleep' passivises as often as the totality of all other verbs*. If this were the Null, then **any strong deviation (of *sleep*) from the overall distribution of passives would be sufficient to weaken it**. We'd have to take a result with, for example, 90 passives and 10 actives of *sleep* (i. e., a strong deviation in the other direction) as equally good evidence against the Null as the result that was actually obtained. The test would be a **two-sided test** because it tests for deviations to both sides. As we'll see in Section 2.6, in this case we'd get $p \approx 0.016$.

two-sided
test

Based on our substantive hypothesis, however, we were only interested in deviations into one direction, namely a **lower relative number (proportion) of occurrences of *sleep* in the passive voice compared to the overall proportion of active and passive voice in the corpus**. Hence, the actual Null was something like: *The verb 'sleep' passivises as often or less often than the totality of all other verbs verb*. The test was a **single-sided test**, in this case a right-sided test (see Section 2.6). If $q_{\text{PASSIVE}(\textit{sleep})}$ denotes the proportion of passives of *sleep* and $q_{\text{PASSIVE}(\neg\textit{sleep})}$ denotes the proportion of passives of all the other verbs, then the mathematical formulation of the Null is:

single-sided
test

$$H_0 : q_{\text{PASSIVE}(\textit{sleep})} \geq q_{\text{PASSIVE}(\neg\textit{sleep})}$$

There might be verbs for which the same theory predicts that they passivise more often than the average verb. For example, *arrest* could be such a verb, as it is not only transitive but also often used in reports of arrests with phrases like “*A man was arrested by the police.*” In order to test a Null based on this substantive hypothesis, we have to reverse the relation between the proportions:

$$H_0 : q_{\text{PASSIVE}(\textit{arrest})} \leq q_{\text{PASSIVE}(\neg\textit{arrest})}$$

¹⁰Please review Chapter 1 if the exact wording of this sentence is less than crystal clear to you.

Table 2.5: The fictional result for the voice distribution of *arrest*

		Voice		
		Active	Passive	
Verb	<i>arrest</i>	69	31	100
	Other	800	200	1,000
		869	231	1,100

In this case, a result as convincing as the result for *sleep* is described numerically in Table 2.5, assuming there were also exactly 100 instances of *arrest* overall. The p-value for Fisher’s Exact Test can be calculated following the know well-established method as follows:

$$\frac{\binom{869}{69}\binom{231}{31}}{\binom{1,100}{100}} + \frac{\binom{869}{68}\binom{231}{32}}{\binom{1,100}{100}} + \dots \frac{\binom{869}{0}\binom{231}{100}}{\binom{1,100}{100}} \approx 0.009$$

Since we are limited by the corpus size of 1,100 and the fixed sample size of 100, we cannot reach exactly 0.008. However, 0.009 is close enough. Such a result could be interpreted as weakening the Null which states that *arrest* is passivised as often or more often than all other verbs. While we can’t argue that a precise substantive hypothesis was directly supported, at least a global negation of such a hypothesis failed the test. However, the directionality involved in formulating the Null is obviously informed by the substantive hypothesis. Hence, it wouldn’t be correct to think that Fisher’s testing philosophy completely ignores substantive hypotheses. In Chapter 8, we’ll turn to methods of directly supporting hypotheses within the Neyman-Pearson philosophy and Deborah Mayo’s Severity approach to frequentist inference.

2.4 Sample Size and Effect Size

When we obtain a low p-value, we can assume that [the Null is false or a rare event has occurred](#). The Null usually states that some effect is absent, that there is nothing going on beyond random chance. We have already illustrated (using the soft-sciences version of *Let’s Make a Deal* and my grandfather’s lottery win) that there cannot be a fixed threshold for what counts as rare. In this section, we examine how the size of the sample and the p-value are related. Then, we informally introduce the notion of the effect size, to which we will return repeatedly in later chapters, very crucially in Chapter 8.

2 Guessing and Counting

Assume that Fisher and Ms Bristow, after fiercely discussing the theoretical foundations of her potential tea-tasting abilities after the first experiment (as one should), decide to repeat the experiment but with a much larger sample. This time, she agrees to taste 80 cups of tea, 40 of which are tea-first. After the Tea-Tasting Marathon, Fisher counts the successes and failures and aggregates the numbers in Table 2.6.

Table 2.6: Results of a Tea-Tasting Marathon

		Reality		
		Tea-First	Milk-First	
Bristow	Tea-First	30	10	40
	Milk-First	10	30	40
		40	40	80

In this second experiment, the results are similar to those reported in Table 2.2 inasmuch as she guessed 75% of the tea-first cups correctly, and she wrongly detected 25% of the milk-first cups incorrectly as tea-first cups. However, if we calculate the frequentist probability of obtaining this result or an even better result without tea-tasting abilities by chance (the p-value), we get:¹¹

$$\frac{\binom{40}{30}\binom{40}{10}}{\binom{80}{40}} + \frac{\binom{40}{31}\binom{40}{9}}{\binom{80}{40}} + \dots + \frac{\binom{40}{40}\binom{40}{0}}{\binom{80}{40}} \approx 7.44 \cdot 10^{-6}$$

Do you recognise the exponential notation? With a negative exponent of -6 , it means that you have to move the point 6 digits to the left, hence $7.44 \cdot 10^{-6} = 0.00000744$. Although the rate of accuracy is the same, the p-value is much lower than the p-value of 0.24 corresponding to Table 2.2.

2.5 The Distribution of P-Values

2.6 Probability Distributions

2.7 Some Afterthoughts on Tests and Collostructions

¹¹Many of these calculations are virtually impossible to do by hand. We double-checked everything. Promised!

3 Describing Data

4 Sampling Accuracy and Confidence

5 Inferences About Means

5.1 Population Means and Sample Means

Problem Statement: Z-Tests

Let's assume you know the mean reaction time for a critical region when native speakers process a certain type of relative clause. This mean reaction time and the corresponding variance in measurements are extremely well established parameters. They were predicted by a robust theory of syntactic processing, and this prediction has been corroborated by a large number of diverse experiments. For an emergent subtype of this kind of relative clause, the theory predicts considerably higher processing effort and thus longer reaction times. You conduct an experiment and measure reaction times in the critical region. Which outcomes of the experiment would you interpret as indicating that reaction times are indeed longer for the emergent type of relative clause?

5.1.1 Introducing the Logic

The Problem Statement exemplifies a common question: Given a known mean value, do means under a specific condition diverge from this known mean? In this section, we show through simple frequentist reasoning how measurements from experiments can provide evidence to tackle such questions. The simplest test for such tasks is the **z-Test**. Notice that for the z-Test to be applicable, the given population mean (and the corresponding variance) must be truly known, which is why the Problem Statement stresses that the mean was predicted by a robust theory and that the prediction was tested in a long series of experiments. If these conditions are not met, other tests apply, and we're going to introduce such tests as we go along.

For the sake of illustration, let's assume that the population mean is $\mu = 120$ (for example milliseconds) and the population variance is $\sigma^2 = 16$, which corre-

z-Test

5 Inferences About Means

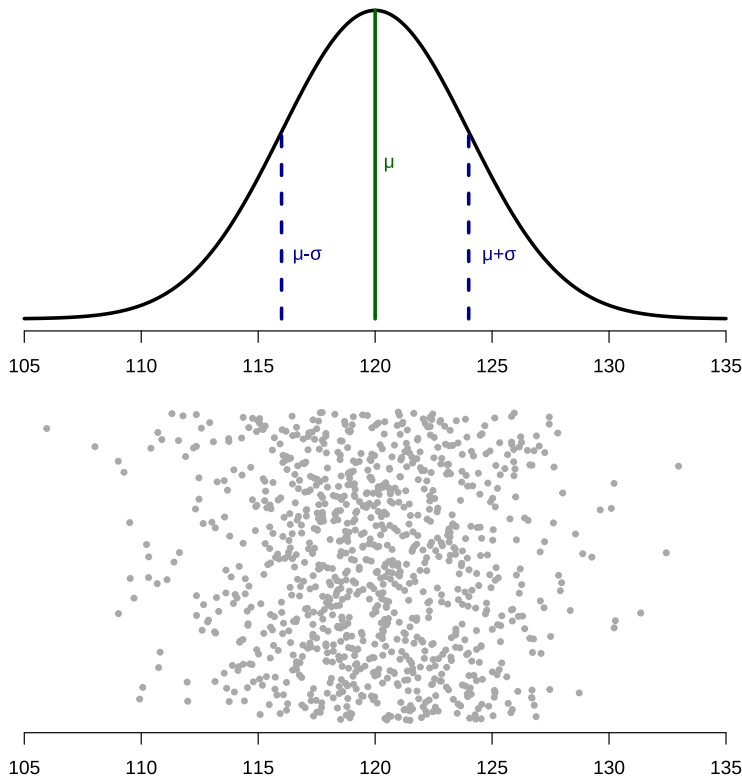


Figure 5.1: Theoretical population distribution for a normal distribution with $\mu=120$ and $\sigma=4$ and a simulated random sample of $n=1000$ measurements from the population

sponds to a standard deviation of $\sigma = 4$. If the population values are generated according to a normal distribution, values are distributed according to the bell curve in Figure 5.1.

To recapitulate, this curve plots the probability density for the known population (for example, of reaction times). In the simplest terms, it shows for each measurement (x-axis) the probability with which it occurs (y-axis).¹ Informally speaking, the curve shows that if we measure random values from this population, the probability of measuring a value close to μ is highest, and measurements deviate on average by the standard deviation σ from μ . The dashed lines show

¹Technically, the probability of each point measurement is 0, and non-zero probabilities are only defined as integrals of the density curve for intervals. This mathematical detail is mostly irrelevant for practical applications, but it should be kept in mind.

the standard deviation in each direction from the mean. As a result, a very much expected sample of $n = 1000$ measurements is shown in the form of the point cloud below the curve. The measurements are indeed centred around the mean, and they seem to follow the normal distribution.

If, however, we draw a sample from a different population where the true mean is higher (for example because we're measuring reaction times under a condition that is more difficult to process) we expect samples to turn out differently and have a higher sample mean compared to the known population mean. However, this expectation can be treacherous because individual samples are not in any way *guaranteed* to represent their population well, as we have shown in Chapter 4. Very similar to Ronald A. Fisher in his experiment with Muriel Bristow (see Chapter 2), we need to ask whether the actual sample warrants any inference regarding the underlying mechanism by being very much unexpected (albeit not impossible) under the assumption that the desired inference is not correct. In the case of the reaction times described in the Problem Statement, the desired inference is that reaction times are higher with the emergent subtype of relative clauses because of assumed processing penalties. However, especially if our sample is small, inferring anything from a specific result is tricky, as will be shown. Figure 5.2 shows a possible outcome with $n = 16$.

Let's call the sample plotted in Figure 5.2 x , a vector of 16 measurements x_1 through x_{16} . The mean of x is $\bar{x} = 122.1$. As we've shown, inferences in frequentist logic (see Chapter 2) are always made by taking into account what the outcome of an experiment could have been under one or several possibly correct hypotheses. In the case at hand, we're interested in the hypothesis that the true mean under the experimental condition—call it μ_1 —is larger than the known mean μ . In other words, we would like to **gather evidence in support of H, where H: $\mu_1 > \mu$** . For several reasons, we cannot gather evidence that supports this hypothesis directly. First, μ_1 is obviously not observable. It's a hypothesised mean in a population that exists as separate from the known population if H is correct. If that population is not substantially different from the known one, then we have $\mu_1 = \mu$. Given that μ_1 is a non-observable, all we've got are 16 data points from x and the sample mean \bar{x} calculated from them. While we certainly hope that \bar{x} is a good indicator of the true value μ_1 , we have no guarantees whatsoever that it actually is. Second, as we're still Fisherians on this page of this book, we have no formal method of gathering positive evidence. Only Neyman-Pearson philosophy and the Severity approach will give us this power (pun intended) later in Chapter 8, but with some important caveats. Therefore, we can only check **whether the data x are in accord with a Null Hypothesis (Null for**

Null
Hypothesis

5 Inferences About Means

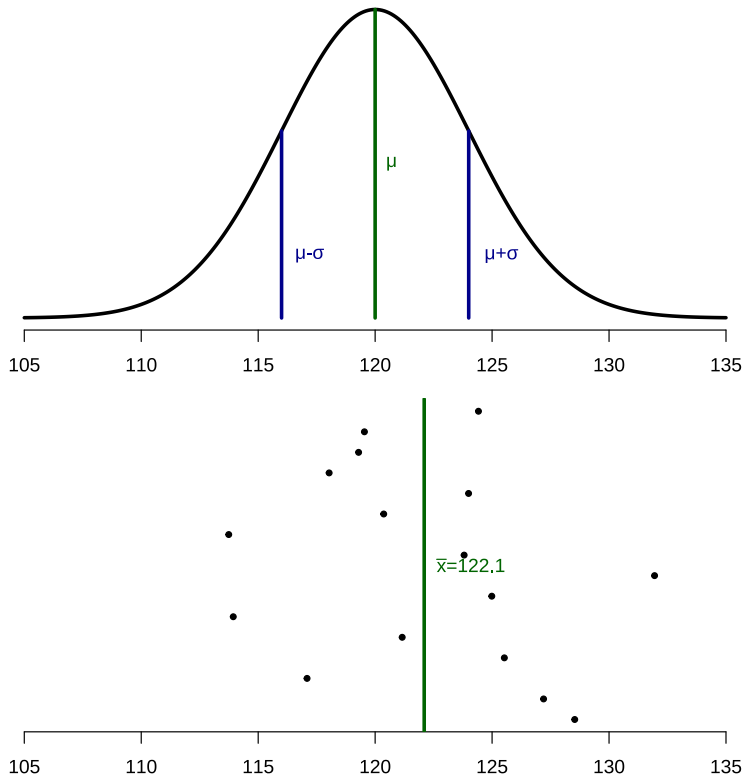


Figure 5.2: Theoretical population distribution for a normal distribution with $\mu=120$ and $\sigma=4$ and a simulated random sample of $n=16$ measurements from some population

short) $N: \mu_1 \not> \mu$.²

To test this hypothesis, the Fisherian framework assigns a certain well-defined kind of probability to (obtaining) the data x given that N is correct, formally $p(x|N)$ or *the probability of x given N* . This probability can be used to assess whether the data x are in accord with the Null N . Before moving on to the calculations, it is vital to consider the question of which inferences are warranted in case x is or isn't in accord with N . First, what do we infer from the data x (in other words, from our experiment) if they are compatible with N ? The answer

²The Null Hypothesis N is sometimes designated as H_0 , in which case H is often called H_1 . This nomenclature is—in our view—where the confusion between Fisher and Neyman-Pearson begins. Furthermore, as the Fisherian Null Hypothesis is not a proper hypothesis anyway but rather a non-hypothesis, we call it Null or N .

is: absolutely nothing! If the data are in accord with N , we haven't found any evidence that it is not the case that μ_1 is not greater than μ , and that's the end of it. If that sounds uselessly messed up and disappointing, that's because it is.³ If you infer anything from such a result, you're not only wrong, but also Baeyesians with Dutch names will come to haunt you (or at least unfollow you on the messaging platform of your choice)—and rightly so. Second, what do we infer from the data x if they are not compatible with N ? This is the much more interesting case, but it's difficult to define the admissible inferences without creating false ideas if it's done without the maths and without a deeper look at the way H and N were set up. Let's say rather informally that in such a case we've found some evidence in support of H because N and H partition the range of possible values of μ_1 : either it's greater than μ (H) or it is not greater than μ (N). Finding no accordance with N despite serious attempts to do so (see Chapter 8) provides at least some indication that H might be correct. If you're looking for proof of anything, we recommend that you stick to pure theory, logic, theoretical maths, or pseudoscience. There is no proof to be found in (non-trivial) experiments, and statistical inferences are weak and fragile.

5.1.2 Doing the Maths

We have argued that in Fisherian inference, we have to assess whether x is compatible or in accord with N . But how do we do this? The most naive but not at all wrong thing to do is calculate the difference between the known population mean μ and the mean of the obtained sample \bar{x} . In our case, this is $\bar{x} - \mu = 2.1$. Clearly, a minimal requirement for any further calculations is that this difference is positive. If it were negative, the sample could hardly be interpreted as evidence against N : $\mu_1 \not> \mu$.⁴

However, solid empirical inference requires us to evaluate how significant this difference actually is. This evaluation follows the same logic as in our introduction to Fisher's philosophy (Chapter 2). In the analysis of the Tea Tasting experiment, we asked how often someone who merely guesses would classify one, two, three, or four cups correctly by chance. In the case at hand, simply counting events is not informative as the events are occurrences of specific values, namely individual reaction times. Hence, the question becomes: how often would we expect to see a sample of 16 measurements with a sample mean of 122.1 or larger

³Whether you're a linguist or not, please consider that *finding no evidence that A is true* is not the same as *finding evidence that A is false*.

⁴This way of putting it is slightly sloppy and informal. We will return to this notion and make it more precise. However, in practice it is blatantly obvious that we would never take an experiment that showed lower reaction times as evidence for higher reaction times, etc.

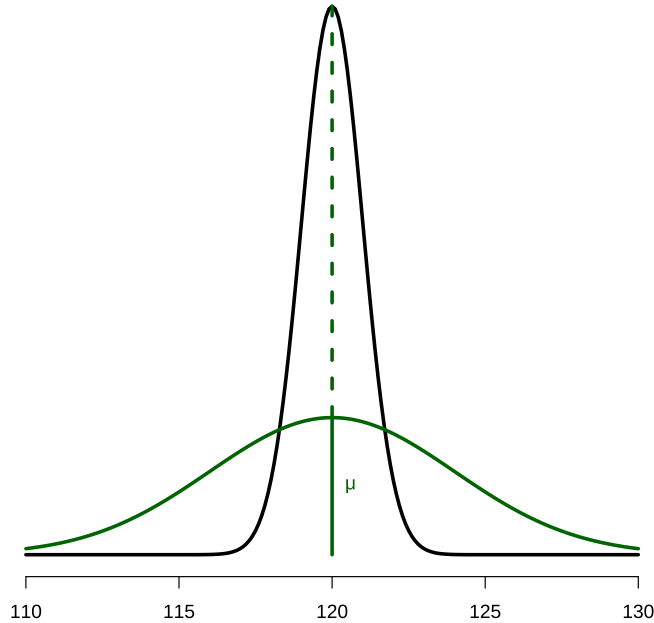


Figure 5.3: Theoretical population distribution for a normal distribution with $\mu=120$ and $\sigma=4$ (green) the distribution of sample means for samples from this distribution with $n=16$ (black)

standard
error

if the true mean is that specified by the Null, which is $\mu = 120$? Luckily, we have already introduced the tool that we need: the **standard error** of the mean. The standard error for $n = 16$ and the known variance $s = 4$ (see p. 27) tells us how much samples of this size differ from the true mean on average. The standard error of the mean is:

$$S(n, \sigma) = \frac{4}{\sqrt{16}} = 1$$

Remember what the standard error is all about (Chapter 4). If the mean in a population is μ and the standard deviation is σ , then the sample means \bar{x}_i of repeated samples of size n are themselves normally distributed with the standard error S being the standard deviation of that normal distribution. Furthermore,

keep in mind that we're talking about the distribution of the known population. Under the Null, it is also the distribution from which our small sample was drawn. Figure 5.3 contrasts the density of the distribution of individual data points (in our example: individual reaction times) with the much narrower distribution of sample means. Mathematically, it is narrower because the standard error is always smaller or equal to the standard deviation. Intuitively, it should be narrower because on average sample means from samples with $n > 1$ approximate the true mean better than single measurements.⁵

Since (i) the distribution of sample means is normal, (ii) we know its mean, (iii) we know its variance/standard deviation, we can calculate how many samples of infinitely many samples (or at least a lot of samples) drawn from the known population have a mean of 122.1 or larger. In other words, we can calculate how many sample means would deviate by 2.1 from the population mean anyway due to expected sampling error if we took a lot of samples of size $n = 16$. This is exactly parallel to the argument regarding the Tea Tasting experiment where we calculated how often we could expect certain outcomes anyway, even if the person performing the Tea Tasting task had no ability to detect which liquid was poured into the cup first. It gives us a very precise and well-defined measure of how surprising the obtained result would be were the Null true.

There are many ways of calculating the number of interest. Since the area under the normal curve sums up to 1 (as should be the case with probability density functions), we could integrate it over the interval $[122.1, \infty]$. Alternatively, we could use the so-called cumulative density function, to which we will return later. To make things much simpler in practice, the most widely used way in applied statistics is based on counting the distance between the distribution mean and the sample mean as multiples of the standard error, which gives us the **z-score**:

z-score

$$z = \frac{\bar{x} - \mu}{S(n, \sigma)} = \frac{122.1 - 120}{1} = \frac{2.1}{1} = 2.1$$

Figure 5.4 shows the distribution of sample means for the known population, the true mean μ , the obtained sample value \bar{x} , and the area under the normal curve which defines how many samples of size $n = 16$ (in the limit) have means of \bar{x} or greater. In addition, the red line measures the distance from μ to \bar{x} , which corresponds to the z-score. By dividing the the distance between the sample mean and the population mean with the standard error, the z-score normalises said

⁵Understanding this argument is crucial. If you're not following it, you should go back to Chapter 4 for an introduction to the distribution of sample means, especially the argument concerning samples of size $n = 1, 2, \dots$

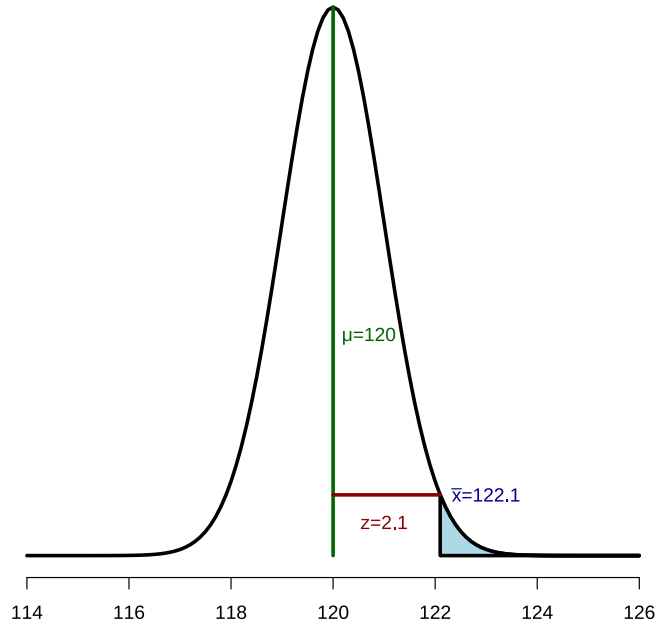


Figure 5.4: Distribution of sample means for mean μ and standard error S with obtained sample mean \bar{x} and corresponding z -value; area under the curve for results equal to or greater than \bar{x} is shaded

distance, which itself varies with the standard deviation in the population and subsequently with the standard error. Hence, regardless of the slope of the concrete distribution, $z = 2.1$ tells us how (im)probable the sample mean (or a more extreme sample mean) is under the Null. It gives us another infamous **p-value**. In the olden days (and in this book), tables were used to look up p-values corresponding to z -scores, and modern statistics software has functions to achieve the same. We will discuss those tables later in this chapter, but for the time being, let's take it for granted that

$$Pr(\bar{x}|N) = p_{\text{Norm}}(2.1) = 0.02$$

Let's go through this step by step. First, $Pr(\bar{x}|N)$ reads *the probability of the sample mean \bar{x} given the Null N* . Remember that the Null was specified as $N: \mu_1 \neq$

μ , which means *the mean under the condition of interest* (the mean reaction time in the emergent type of relative clause) *is not greater than the population mean* (the reaction time in other relative clauses). Second, this probability is equated to $p_{\text{Norm}}(2.1)$, which is the p-value corresponding to the z-value of 2.1 as calculated above, which is 0.02.

5.1.3 Interpreting the Results

At this point, a little exercise is in order. From the following statements, choose the one which is a correct interpretation of the calculations above given the scenario described in the Problem Statement. There is no pressure. Nobody can read your mind, nobody even cares whether you pick a wrong statement, and you can only gain by actually *thinking* about each statement thoroughly. Do not decide on one statement because it is intuitively correct, but because you know why it is correct.

1. The probability that hypothesis H (reaction times are longer in the emergent type of relative clause) is true is 0.02.
2. The probability that hypothesis H (reaction times are longer in the emergent type of relative clause) is true is 0.98.
3. The results prove that the emergent type of relative clause incurs longer reaction times than other relative clauses.
4. The p-value is very small, which indicates that mean reaction times in the emergent type of relative clause are substantially longer than in other relative clauses.
5. The probability of obtaining another sample with a mean of 122.1 or greater in an exact replication of the experiment is 0.02.
6. Based on this outcome, we can reject the possibility that reaction times under the condition of interest are actually *smaller* than in the population with a certainty of $1 - 0.02 = 0.98$ (or 98%).
7. The probability that we actually drew a sample with a mean of 122.1 is 0.02.
8. The experiment has shown that reaction times are normally distributed.
9. The experiment provides evidence in favour of the underlying theory of linguistic processing.

[TODO continue here]

Big Point: Interpretation of the P-Value in Z-Tests

The p-value in a z-test is the frequentist probability of drawing a sample x with a mean as extreme as or more extreme than the one that was actually obtained if the Null were true. The frequentist probability is the probability of an event occurring before it has actually occurred. After the sample has been drawn, the probability that it was drawn is 1, regardless of how extreme its mean is.

5.1.4 Differences Between the Fisher Test and the Z-Test

5.1.5 The Distribution of P Values

5.2 The Unknown Population

5.3 Means of the Unknowns

6 Differences in Means and Variances

7 Some Other Scenarios

8 Positive Inferences

9 Quantifying Correlation Between Variables

10 Modelling Linear Relationships

11 Modelling Arbitrary Outcomes

12 Modelling With Groups

Statistical Inference for Everybody and a Linguist

This book is good.