

# Econometrics I - Simple hypothesis testing

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$$H_0 : \beta_j = \beta_{j,0}$$

$$H_A : \beta_j \neq \beta_{j,0}$$

The decision to “reject” or “fail to reject”  $H_0$  may begin by the researcher *subjectively* deciding on a *significance level* and then doing one or more of the following:

- ▶ Calculating a ( $p$ -value) and comparing it to the significance level.
- ▶ Seeing whether or not  $\beta_{j,0}$  is contained in a confidence interval.
- ▶ Calculating a test statistic and seeing if it exceeds a critical value.

We need an *estimator* for the *standard error of the estimator* being used to assess the hypothesis.

For example, the t-test statistic for testing:

$$H_0 : \beta_j = 0$$

$$H_A : \beta_j \neq 0$$

is:

$$t = \frac{b_j}{s.\hat{e.}(b_j)}$$

We need this quantity  $s.\hat{e.}(b_j)$ , which is called the *estimated standard error*.

# Estimating $\sigma^2$

Know a lot about estimating  $\beta$ . Another parameter in the model:  $\sigma^2$  – the variance of each  $\epsilon_i$ . We need to estimate  $\sigma^2$  so that we can get an estimate for the covariance matrix of the LS estimator:

$$V(\mathbf{b}) = \sigma^2 (X'X)^{-1}.$$

Let's derive an estimator for  $\sigma^2$ . Begin by noting that

$$\sigma^2 = \text{var}(\epsilon_i) = E \left[ (\epsilon_i - E(\epsilon_i))^2 \right] = E(\epsilon_i^2),$$

due to assumption A.3.

The sample counterpart to this population parameter ( $\sigma^2$ ) is the sample average of the “residuals”:

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n e_i^2 = \frac{1}{n} \mathbf{e}' \mathbf{e},$$

which is the method of moments, and the maximum likelihood estimator. However, there is a distortion in this estimator of  $\sigma^2$ . Although the mean of the  $e_i$ 's is zero (if there is an intercept in the model), not all of  $e_i$ 's are independent of each other: only  $(n - k)$  of them are.

We should consider what properties  $\hat{\sigma}^2$  has as an estimator of  $\sigma^2$ , before we use it. Is this a *good* estimator? What properties of the LS estimator did we evaluate? We will write  $\mathbf{e}'\mathbf{e}$  in terms of only  $\boldsymbol{\epsilon}$  (which we have made assumptions about), and then derive its expected value:

$$\mathbf{e} = (\mathbf{y} - \hat{\mathbf{y}}) = (\mathbf{y} - \mathbf{X}\mathbf{b}) = \mathbf{y} - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y} = \mathbf{M}\mathbf{y}$$

where

$$\mathbf{M} = \mathbf{I}_n - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}' \quad ; \quad \text{idempotent, and } \mathbf{M}\mathbf{X} = \mathbf{0}$$

So,

$$\mathbf{e} = \mathbf{M}\mathbf{y} = \mathbf{M}(\mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}) = \mathbf{M}\boldsymbol{\epsilon}$$

and

$$\mathbf{e}'\mathbf{e} = (\mathbf{M}\boldsymbol{\epsilon})'(\mathbf{M}\boldsymbol{\epsilon}) = \boldsymbol{\epsilon}'\mathbf{M}\boldsymbol{\epsilon} \quad ; \quad \text{a scalar}$$

From this, it can be shown that:

$$\begin{aligned} E(\mathbf{e}'\mathbf{e}) &= E[\boldsymbol{\epsilon}'M\boldsymbol{\epsilon}] = E[\text{tr}(\boldsymbol{\epsilon}'M\boldsymbol{\epsilon})] = E[\text{tr}(M\boldsymbol{\epsilon}\boldsymbol{\epsilon}')] \\ &= \text{tr}[ME(\boldsymbol{\epsilon}\boldsymbol{\epsilon}')] = \text{tr}[M\sigma^2 I_n] = \sigma^2 \text{tr}(M) \\ &= \sigma^2(n - k) \end{aligned}$$

We won't cover the trace operator ( $\text{tr}$ ), we will not discuss why  $\text{tr}(M) = \sigma^2(n - k)$ . However, you need to be aware that an important step in considering whether an estimator should be used is to examine its *bias*, and in the case of  $\hat{\sigma}^2$ :

$$E[\hat{\sigma}^2] = E\left[\frac{1}{n}\mathbf{e}'\mathbf{e}\right] = \frac{1}{n}(n - k)\sigma^2 < \sigma^2$$

The method of moments and maximum likelihood estimator,  $\hat{\sigma}^2$ , is *biased*.

It is easy to convert this biased estimator to an *unbiased* one:

$$s^2 = \frac{1}{(n - k)} e' e$$

Some notes:

- ▶  $(n - k)$  is the “degrees of freedom” – number of independent sources of information in the  $n$  residuals (the  $e_i$ ’s).
- ▶ We can use  $s$  as an estimator of  $\sigma$ , but it is a biased estimator. Even though it is biased,  $s$  is typically used in practice as the bias is small.
- ▶  $s$  is called the “standard error of the regression”, or the “standard error of estimate”.
- ▶  $s^2$  is a statistic. It has its own sampling distribution, etc.



Let's see one immediate application of  $s^2$  and  $s$ . Recall the sampling distribution for the LS estimator,  $\mathbf{b}$ :

$$\mathbf{b} \sim N \left[ \boldsymbol{\beta}, \sigma^2 (X'X)^{-1} \right]$$

So, the variance of the  $i^{th}$  LS estimator is the  $i^{th}$  diagonal of the covariance matrix of  $\mathbf{b}$ :  $\text{var}(b_i) = \sigma^2 \left[ (X'X)^{-1} \right]_{ii}$ , but  $\sigma^2$  is *unobservable*. If we want to report the variability associated with  $b_i$  as an estimator of  $\beta_i$ , we need to use an estimator of  $\sigma^2$ . The estimated variance of the  $i^{th}$  LS estimator is then:

$$\widehat{\text{var}(b_i)} = s^2 \left[ (X'X)^{-1} \right]_{ii}$$

The square-root of the above is called the “standard error” of  $b_i$ . This quantity will be very important when it comes to constructing *interval estimates* of our regression coefficients, and when we construct *tests of hypotheses* about these coefficients.

# Standard errors in R.

Start by setting the random seed and sample size, and then generate some data:

```
1 set.seed(7010)
2 n <- 10
3 x <- rnorm(n)
4 y <- rnorm(n)
```

Estimate and summarize a model:

```
1 mod <- lm(y ~ x)
2 summary(mod)
```

```
1 Coefficients:
2             Estimate Std. Error t value Pr(>|t|)
3 (Intercept)   0.2472     0.3734   0.662   0.526
4 x            -0.2455     0.6348  -0.387   0.709
5
6 Residual standard error: 1.178 on 8 degrees of freedom
7 Multiple R-squared:  0.01835, Adjusted R-squared:  -0.1044
8 F-statistic: 0.1495 on 1 and 8 DF,  p-value: 0.7091
```

R is reporting the standard errors of  $b_1$  and  $b_2$  as 0.3734 and 0.6348 respectively. We will see one way how statistical packages can calculate these numbers. Start by getting the estimate  $s^2$ :

```
1 s2 <- sum(mod$residuals ^ 2) / (n - 2)
```

Note that  $k = 2$  above. If we take the square root, we get the “residual standard error” reported in the R output above:

```
1 sqrt(s2)
```

```
1 [1] 1.178469
```

Next we will calculate the  $V(\mathbf{b})$  matrix. Start by arranging the  $\mathbf{x}$  data into a matrix (and take a look at the  $\mathbf{X}$  matrix):

```
1 X <- matrix(c(rep(1, n), x), n, 2)
2 X
```

```
1      [,1]      [,2]
2 [1,]      1 -0.2214732
3 [2,]      1  0.6051370
4 [3,]      1  0.7208573
5 [4,]      1 -0.2230900
6 [5,]      1 -0.2662395
7 [6,]      1 -1.0890823
8 [7,]      1 -0.5655553
9 [8,]      1  0.5330395
10 [9,]      1  0.6225352
11 [10,]     1 -0.4755066
```

The  $s^2(X'X)^{-1}$  matrix is then:

```
1 s2 * solve(t(X) %*% X)
```

```
1      [,1]      [,2]  
2 [1,] 0.13939930 0.01448197  
3 [2,] 0.01448197 0.40297318
```

I have used the `solve()` function to find the inverse. Taking the square root of any of the diagonal elements gives the standard error reported in the `summary()` output above:

```
1 sqrt(s2 * solve(t(X) %*% X))[1, 1]
```

```
1 [1] 0.3733622
```